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Authors

Chung, Bowen Ong, Michael Ettner, Susan L <u>et al.</u>

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SUPPLEMENT

12-Month Outcomes of Community Engagement Versus Technical Assistance to Implement Depression Collaborative Care

A Partnered, Cluster, Randomized, Comparative Effectiveness Trial

Bowen Chung, MD, MSHS*; Michael Ong, MD, PhD; Susan L. Ettner, PhD; Felica Jones, AA; James Gilmore, MBA; Michael McCreary, MPP; Cathy Sherbourne, PhD; Victoria Ngo, PhD; Paul Koegel, PhD; Lingqi Tang, PhD; Elizabeth Dixon, PhD; Jeanne Miranda, PhD; Thomas R. Belin, PhD; and Kenneth B. Wells, MD, MPH*

Background: Depression collaborative care implementation using community engagement and planning (CEP) across programs improves 6-month client outcomes in minority communities, compared with technical assistance to individual programs (resources for services [RS]). However, 12-month outcomes are unknown.

Objective: To compare effects of CEP and RS on mental healthrelated quality of life (MHRQL) and use of services among depressed clients at 12 months.

Design: Matched health and community programs (n = 93) in 2 communities randomly assigned to receive CEP or RS. (ClinicalTrials.gov: NCT01699789).

Measurements: Self-reported MHRQL and services use at baseline, 6 months, and 12 months.

Setting: Los Angeles, California.

Patients: 1018 adults with depressive symptoms (8-item Patient Health Questionnaire score \geq 10), 88% of whom were an ethnic minority.

Intervention: CEP and RS to implement depression collaborative care.

Measurements: The primary outcome was poor MHRQL (12-item mental health composite score \leq 40) at baseline, 6 months, and 12 months; the secondary outcome was use of services at 12 months.

Results: At 6 months, the finding that CEP outperformed RS to reduce poor MHRQL was significant but sensitive to underlying statistical assumptions. At 12 months, some analyses suggested that CEP was advantageous to MHRQL, whereas others did not confirm a significant difference favoring CEP. The finding that CEP reduced behavioral health hospitalizations at 6 months was less evident at 12 months and was sensitive to underlying statistical assumptions. Other services use did not significantly differ between interventions at 12 months.

Limitation: Data are self-reported, and findings are sensitive to modeling assumptions.

Conclusion: In contrast to 6-month results, no consistent effects of CEP on reducing the likelihood of poor MHRQL and behavioral health hospitalizations were found at 12 months. Still, given the needs of underresourced communities, the favorable profile of CEP, and the lack of evidence-based alternatives, CEP remains a viable strategy for policymakers and communities to consider.

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* Former Robert Wood Johnson Foundation Clinical Scholar.

Depression and depressive symptoms are main causes of disability in the United States (1, 2), where racial disparities persist in access to and quality and outcomes of care (3–9). Depression collaborative care provided in primary care settings can improve quality and outcomes of care for depressed adults while reducing outcome disparities by race (10–18), but safety-net primary care settings generally have limited capacity for full implementation of collaborative care (19–21). Encouraging safety-net clinics to collaborate with other key agencies (for example, social services or faith-based organizations) using community engagement (22–26) may support successful implementation of depression collaborative care across underresourced communities.

Community Partners in Care (CPIC) was designed to compare the effects of 2 depression collaborative care implementation approaches: 1) community engagement and planning (CEP), which supports collaborative planning and implementation across myriad community programs, and 2) more traditional resources for services (RS) models, which rely on time-limited expert technical assistance for collaborative care to individual programs (27–29). Earlier

See also:

Web-Only Data Supplements

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Table 1. Baseline Characteristics of Depressed Clients, by Intervention*

Characteristic	Overall ($n = 1018$)	RS Group ($n = 504$)	CEP Group ($n = 514$)
Service sector, n (%)			
Primary care or public health	290 (28.5)	134 (26.6)	156 (30.4)
Mental health services	195 (19.2)	110 (21.8)	85 (16.5)
Substance abuse	230 (22.6)	111 (22.0)	119 (23.2)
Homeless services	162 (15.9)	92 (18.3)	70 (13.6)
Community-based	141 (13.9)	57 (11.3)	84 (16.3)
Mean age (SD), y	44.8 (12.7)	44.2 (12.3)	45.3 (13.0)
Female, n (%)	595 (58.4)	286 (56.7)	309 (60.1)
Ethnicity, n (%)†			
Latino	409 (40.2)	194 (38.5)	215 (41.8)
African American	488 (47.9)	239 (47.4)	249 (48.4)
Non-Hispanic white	86 (8.4)	45 (8.9)	41 (8.0)
Other	35 (3.4)	26 (5.2)	9 (1.8)
Married or living with partner, n (%)	231 (22.7)	116 (23.0)	115 (22.5)
Less than high school education, n (%)	446 (43.8)	221 (43.9)	224 (43.7)
\geq 3 chronic medical conditions, <i>n</i> (%)	548 (53.8)	270 (53.6)	278 (54.0)
Family income from work in the past 12 mo \leq \$10 000, <i>n</i> (%)	755 (74.1)	374 (74.2)	381 (74.0)
Family income under federal poverty level, n (%)	750 (73.7)	373 (74.0)	377 (73.3)
No health insurance, n (%)	545 (53.5)	286 (56.7)	259 (50.4)
Working for pay, n (%)	205 (20.1)	105 (20.9)	100 (19.4)
12-mo depressive disorder, n (%)	629 (61.8)	311 (61.8)	318 (61.8)
Probable depression (PHQ-8 score \geq 10), <i>n</i> (%)	992 (97.7)	490 (97.4)	502 (98.0)
Mean PHQ-8 score (SD)	14.9 (4.1)	15.0 (4.2)	14.8 (4.1)
Alcohol abuse or use of illicit drugs in the past 12 mo, n (%)	398 (39.1)	180 (35.8)	218 (42.4)
Poor mental health-related quality of life, n (%)‡	546 (53.6)	271 (53.7)	275 (53.5)

CEP = community engagement and planning; PHQ-8 = 8-item Patient Health Questionnaire; RS = resources for services.

* Data were multiply imputed. The chi-square test was used to compare the groups, taking into account the design effect of the cluster randomization. + P > 0.30 for all comparisons except for ethnicity, for which P = 0.030.

 \ddagger 12-item Mental Composite Score \leq 40 (1 SD below the population mean).

studies concluded that at 6 months, compared with RS, CEP reduced the probability of poor mental health-related quality of life (MHRQL) among depressed clients, increased their physical activity, and reduced risk factors for homelessness (28-30). Moreover, CEP reduced behavioral health hospitalizations and specialty medication visits among visitors to mental health specialists while increasing use of primary care, faith-based, and park-based services for depression among such clients. To our knowledge, CPIC is the first randomized U.S. study of the added value of CEP beyond more traditional expert assistance to individual programs and the first depression collaborative care study to span the health care and social community sectors.

We examined the effects of CEP over RS on poor MHRQL and services use at 6 and 12 months, as well as changes in outcomes from baseline to 6 months and baseline to 12 months. We hypothesized that compared with RS, CEP would decrease the proportion of clients with poor MHRQL at 12 months.

METHODS

Study Design

Community Partners in Care is a group-level randomized comparative effectiveness trial comparing CEP with RS. Both interventions were designed to provide extensive depression collaborative care training to mental health, medical, and community-based agencies. The RS model provided preset, time-limited training to individual agencies, whereas CEP encouraged these diverse agencies to develop a strategy and training plan to jointly provide care for depression (Table 1 in Data Supplement 4, available at www.annals.org). The interventions and study methods are described elsewhere (28-31).

The study and CEP intervention were guided by community-partnered participatory research principles (32-35), a community-based participatory research variant (36, 37) promoting equal authority among community and academic partners (Data Supplement 1, available at www.annals.org). The study council, co-led by the University of California, Los Angeles (UCLA); RAND Corporation; Healthy African American Families II; Behavioral Health Services; and QueensCare Health and Faith Partnership supported workgroups and community forums for study input (27-31, 38, 39).

Setting

The study took place in 2 Los Angeles County communities: South Los Angeles and Hollywood-Metro. These communities have high rates of poverty and avoidable hospitalizations and low rates of insurance (40-42). We hosted community meetings to identify community-based settings that support vulnerable depressed populations, with the aim of oversampling from these settings, which included mental health, primary care, public health, substance abuse, and social services; faith-based programs; parks; senior centers; hair salons; and exercise clubs. South Los Angeles partners emphasized inclusion of large samples of substance abuse clients and African American persons, whereas Hollywood-Metro emphasized homeless clients and seniors.

Participants and Randomization Programs

We began by identifying a pool of relevant agencies and organizations through county program lists and nominations from community partners. We then contacted each to assess interest, eligibility, and enrollment. This process resulted in a pool of 60 potentially eligible agencies with 194 programs. Programs were eligible if they served 15 or more clients per week, had 1 or more staff members, and did not focus exclusively on psychotic disorders or home services. A total of 133 of these 194 programs were potentially eligible.

Within each community, programs or clusters of smaller programs were paired on the basis of location, service sector, size, population served, services provided, and funding streams; 2 larger agencies were their own single stratum. Within pairs, one program or cluster was randomly assigned to CEP and the other to RS (43). A statistician uninvolved with recruitment supported council members in producing seed numbers for randomization (44).

The 133 programs were randomly assigned to the interventions (65 to RS and 68 to CEP) (Figure). After randomization, RAND staff who were blinded to assignment conducted site visits to finalize enrollment; 20 programs were determined ineligible, 18 declined participation, and 95 programs from 50 consenting agencies were enrolled (46 in RS and 49 in CEP) (Figure).

Administrators were informed of intervention status by letter. Participating and nonparticipating agencies were similar in terms of clients' age, sex, and race; population density; and client income by ZIP code (each P > 0.10), as determined by analysis of census tract data.

Clients

To achieve a 6-month follow-up sample of 780 depressed clients, we planned to enroll 557 to 600 clients per intervention (assuming 65% to 70% retention). We powered the study to identify a detectable effect size ranging from 0.20 to 0.22 and a percentage point difference between groups ranging from 9.98 to 10.91, assuming 80% power with a 2-sided α value of 0.05 and an intraclass correlation coefficient assumed to range from 0.00 to 0.02 (43, 45, 46).

Within programs, clients were screened in waiting rooms (approached consecutively) or at events (approached randomly) from March to November 2010. Allocating 2 to 3 days per program, RAND staff who were blinded to intervention status approached 4649 adults (aged ≥ 18 years) allocating 2 to 3 days per program; of these, 4440 (95.5%) agreed to screening. Eligibility was limited to clients who provided contact information and had at least mild depressive symptoms, as indicated by a score of 10 or greater on the 8-item Patient Health Questionnaire (47). Of 4440 persons screened, 1322 (29.8%) were eligible; 1246 (94.3%) of eligible persons enrolled, a high rate for such studies (11, 14, 16).

Between April 2010 and January 2011, 981 clients (79% of those enrolled) completed a baseline telephone survey conducted by RAND survey staff who were blinded to intervention. Of 1093 participants approached for 6-month telephone follow-up surveys between November 2010 and August 2011, 759 (70%) participated. Of 974 participants approached for 12-month telephone follow-up surveys between May 2011 and March 2012, 733 (75%) participated. **Data Supplement 2** (available at www.annals .org) shows the survey questions. We did not attempt to contact 272 participants because their survey response at baseline (n = 153) or 6 months (n = 119) was final refusal, ill, incarcerated, unable to contact, or dead.

Our analytic sample comprises 1018 individuals (77% of eligible; 82% of enrolled) who completed 1 or more surveys at baseline, 6 months, or 12 months (Figure). Characteristics of persons who completed the 12-month survey differed from those did not complete it by intervention: The RS group had significantly higher nonresponse rates among men, clients recruited from substance abuse programs, and those with no health insurance. In the CEP group, responders were more likely to have lower family income and to be African American (Tables B2 to B4 in Data Supplement 3, available at www.annals.org).

Interventions

The compared interventions, CEP and RS, were designed to expose a range of health care and social community agencies to the same depression collaborative care toolkits. Between December 2009 and July 2011, CEP supported program administrators to work as councils: one in Hollywood-Metro and another in South Los Angeles. Each council met biweekly over 5 months to adapt depression care toolkits and trainings to each community. In addition, each council developed plans for a coordinated services network across health care and social community programs to support depressed adults. Planning was co-led by community and academic council members following community-partnered participatory research principles (for example, shared authority and 2-way knowledge exchange) (39) (Table 1 in Data Supplement 4, available at www.annals.org).

In RS, technical assistance was offered to assigned programs for the depression care toolkits by using a "train-thetrainer" model. Between December 2009 and July 2010, training was conducted through 10 webinars plus site visits to primary care for each community (39). Trainers included a nurse care manager, a cognitive-behavioral therapist who was a licensed psychologist, 3 board-certified psychiatrists for medication management, and community

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Five programs (2 in the RS group and 3 in the CEP group) had no clients with data for outcome analysis. CEP = community engagement and planning; RS = resources for services.

service administrators to support participation and cultural competence.

The CPIC Council modified depression collaborative care toolkits (48) that supported clinician assessment, medication management, case management (screening, care coordination, and patient education), patient education, and cognitive-behavioral therapy (14, 16, 17, 48, 49), adding a lay health worker manual and team support tools (50, 51). Toolkits introduced to programs before randomization at 1-day kick-off conferences in each community were avail-

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able online, on flash drives, and on hard copy (27, 38, 39), for participating programs in both interventions (Table 1 in Data Supplement 4).

After randomization and enrollment, within each intervention, training invitations were offered by phone, e-mail, and postcards to staff who attended prior CPIC study meetings; circulation to all eligible staff was encouraged. Providers and clients in enrolled programs could use intervention resources for free, even if they were not individually enrolled as participants. Incentives to participate in training included continuing education credits and food during training sessions. Enrolled client lists were provided to CEP but not to RS administrators, except at 1 agency that had a shared waiting room where both were given lists.

The institutional review boards of RAND and other participating agencies approved the study procedures before initiation. The National Institutes of Health did not consider the study a clinical trial when it was funded in 2007, and no data safety monitoring board was required. After data collection, the study was registered with ClinicalTrials.gov (NCT01699789). No major design changes were made after recruitment began.

Outcomes and Follow-up

All outcomes were based on client self-report during telephone surveys and were assessed by RAND staff at baseline, 6 months, and 12 months. Baseline measures include program intervention assignment and sector, and client data from the screening and baseline survey on demographic characteristics (age and sex), presence of 3 or more chronic conditions (among 18 conditions), education level, race/ethnicity, physical health composite score and mental health composite score (MCS-12) from the 12-item Short-Form Health Survey (52, 53), and meeting federal criteria for family poverty (54). Using the Mini-International Neuropsychiatric Interview (55), we assessed for the following conditions on the basis of Diagnostic and Statistical Manual of Mental Disorders, fourth edition, criteria: 12-month major depressive or dysthymic disorder, current manic episode, recent anxiety disorder (panic or post traumatic stress in the past month, or generalized anxiety disorder within the past 6 months), and alcohol abuse or use of illicit drugs in the past 12 months.

The primary study outcome was percentage of clients with poor MHRQL, as indicated by an MCS-12 of 40 or less (1 SD below the U.S. population mean) at 12 months (52, 53). A sensitivity analysis was conducted using the MCS-12 as a continuous measure. Secondary outcomes were services use indicators (for example, primary care visits and nights spent in a hospital because of behavioral health problems).

We report 6- and 12-month outcomes, and the change in outcomes from baseline to 6 and 12 months. For all outcomes, we also report results of sensitivity analyses, with survey follow-up time as a class variable and varying imputed data assumptions.

We assessed services use in the past 6 months for behavioral health (mental health, alcohol abuse, and substance abuse). In particular, we asked clients about the number of nights spent in a hospital; use of overnight substance abuse rehabilitation facilities; emergency department visits; outpatient mental health or self/family group visits; hotline calls; and use of outpatient primary care or public health clinics, substance abuse or social services programs, parks and community centers, and faith-based and other community programs. Services for which the client reported receiving information, referral, counseling, or medication management for depression or emotional problems were classified as depression-related visits. We developed indicators for any service use and being above the baseline median number of visits and counts of contacts. Because a single overnight stay could reflect emergency department use, we performed a sensitivity analysis that included 4 or more hospital nights. To account for potential bias in self-report, we asked participants to provide names and addresses for up to 4 providers per sector; for high utilizers and "other" locations, we confirmed sector and count feasibility through searching the Internet and calling programs.

Statistical Analysis

We conducted intention-to-treat analyses of repeated measures that included all participants with available data at baseline, 6 months, or 12 months by using SAS software, version 9.2 (SAS Institute). Initial explorations of 3-level, random-effects logistic models using SAS proc glimmix for binary outcomes yielded unstable estimates for program-specific random effects.

We analyzed dichotomous and count outcomes by using a generalized estimating equation framework. Specifically, we fitted logistic regression models for binary outcomes and Poisson models for count data using SAS proc genmod, specifying exchangeable correlation at the program level, with regression adjustment for baseline covariates (age, sex, \geq 3 chronic conditions, education, race/ ethnicity, family poverty, alcohol abuse or use of illicit drugs in the past 12 months, depressive disorder in the past 12 months, and community). We then developed a contrast involving a linear combination of coefficients to test intervention effects at each end point (baseline, 6 months, and 12 months) and tested differences between intervention groups in change from baseline to 6 and 12 months.

The results of analyses of binary outcomes are presented as odds ratios, and the results of Poisson regression analyses of count data are presented as rate ratios. We summarized effect sizes by presenting unadjusted means and proportions by intervention group (16, 56) (Data Supplement 3). We treated time as a continuous variable and examined the fixed effects for time and intervention, and their interactions. We included quadratic terms (squared effect of time and its interaction with the intervention),

Quality of Life Measure	Clients, n	CEP Group	RS Group	Odds Ratio (95% CI)	
				Primary Analysist	
				CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline
MCS-12 ≤40, <i>n/N</i> (%)					
Baseline	962	257/479 (53.7)	259/483 (53.6)	1.00 (0.79 to 1.28)	-
6 mo	755	166/376 (44.1)	198/379 (52.2)	0.71 (0.55 to 0.91)**	0.71 (0.51 to 0.98)**
12 mo	717	160/361 (44.3)	181/356 (50.8)	0.77 (0.61 to 0.97)**	0.77 (0.55 to 1.07)
Mean MCS-12 (SD)				Group Difference	Group Difference in Change From Baseline
Baseline	962	39.2 (7.3)	39.2 (7.5)	0.12 (-0.82 to 1.06)	_
6 mo	755	40.3 (7.0)	39.7 (7.4)	0.60 (-0.36 to 1.57)	0.48 (–0.61 to 1.58)
12 mo	717	40.7 (7.0)	40.4 (7.1)	0.35 (-0.73 to 1.44)	0.23 (-1.21 to 1.67)

Table 2. Proportion of Clients With Poor Mental Health-Related Quality of Life and Mean MCS-12 at 6 and 12 Months*

CEP = community engagement and planning; MAR = missing at random; MCS-12 = 12-item mental health composite score; NMAR = nonignorable missing at random; RS = resources for services.

* Adjusted analyses used multiply imputed data based on 1018 clients. A generalized estimating equation logistic regression model was used for the binary variable MCS-12 \leq 40 (presented as odds ratios). A 3-level mixed-effect regression model was used for the continuous variable MCS-12 score (presented as mean difference), adjusted for age, sex, \geq 3 chronic conditions, education, race/ethnicity, income, family income below the federal poverty level, 12-mo alcohol abuse or use of illicit drugs, 12-mo depressive disorder, and community.

+ Analysis with time as a continuous variable and multiple imputation procedures assuming that the missing data are MAR.

‡ Analysis with time as a continuous variable and NMAR by multiplying the ignorable model's imputed data by 1.1.

§ Analysis with time as a continuous variable and NMAR by multiplying the ignorable model's imputed data by 0.9.

Analysis with time as a categorical variable with 2 indicators for 6- and 12-mo time points and MAR.

** P < 0.05.

which allowed insight into whether changes are greater from baseline to 6 months or subsequent months.

In analyzing continuously scaled MCS-12 as the dependent variable, we used a 3-level, mixed-effect regression model by using SAS proc mixed. We accounted for the multilevel data structure with clients nested within programs and repeated measurements nested within clients. To account for the intraclass correlation expected in the data, we specified random effects at the program level and an autoregressive (1) covariance structure within clients to account for within-client correlation over time.

We used item-level imputation for missing data and wave-level imputation for missing surveys to adjust findings to the observed analytic sample (n = 1018). In our prior outcome study (30), we used weights to account for nonenrollment and nonresponse. In the current study, we used a model-based approach with unweighted data (56). As a result, the current study's baseline and 6-month estimates differ slightly from those in prior reports (28, 29).

We conducted sensitivity analyses for alternative representations of time as a continuous or class variable and for alternative weighting approaches. To investigate possible nonignorable effects, we used 2 methods. For continuous measures (such as MCS-12 and number of service visits), we multiplied ignorable-model imputations alternatively by 1.1 and 0.9 to reveal sensitivity to 10% departures from ignorable-model predictions with dichotomized versions of continuous measures (MCS-12 ≤40) based on the imputed continuous value. For categorical imputations where reference cells were based on an underlying continuous measure (that is, predicted response propensity) in-

cluding an indicator for any utilization and adjusted Bayesian bootstrap imputations reflecting unit nonresponse at a particular time point, nonignorable imputations for cases in nonboundary reference cells were generated by borrowing values from the reference cell with either the next higher or next lower value of the underlying continuous measure (57).

Role of the Funding Source

The National Institute of Mental Health, UCLA Clinical and Translational Science Institute, Robert Wood Johnson Foundation, National Library of Medicine, and California Community Foundation supported the study. The National Institute of Mental Health project officer served as an advisor to the CPIC Council, but otherwise, funders had no role in design, conduct, or analysis of the study; interpretation of the data; manuscript preparation; or the decision to submit the manuscript for publication.

RESULTS

Baseline Participant Characteristics

Of 1018 depressed clients in 12-month outcome analyses, 58.4% were female, 88.1% were Latino or African American, 43.8% had less than a high school education, 73.7% had an income below the federal poverty level, 53.5% had no health insurance, and 20.1% were employed. The percentage with 12-month depressive disorder was 61.8%, whereas 39.1% had substance or alcohol abuse and 53.8% had 3 or more chronic medical conditions (Table 1). The CEP group had more Latino and African American participants than the RS group (85.9% vs

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Odds Ratio (95% CI)							
Higher NMAR Value†‡		Lower N/	NAR Value†§	Value†§ Categorical Time Value†			
CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline		
0.97 (0.76 to 1.25)	-	1.03 (0.81 to 1.31)	-	1.00 (0.78 to 1.27)	-		
0.73 (0.58 to 0.93)**	0.75 (0.55 to 1.04)	0.79 (0.61 to 1.02)	0.77 (0.56 to 1.07)	0.72 (0.55 to 0.95)**	0.73 (0.51 to 1.04)		
0.83 (0.66 to 1.06)	0.86 (0.61 to 1.21)	0.86 (0.67 to 1.09)	0.83 (0.59 to 1.18)	0.79 (0.59 to 1.05)	0.79 (0.56 to 1.13)		
Group Difference	Group Difference in Change From Baseline	Group Difference	Group Difference in Change From Baseline	Group Difference	Group Difference in Change From Baseline		
0.27 (-0.73 to 1.26)	C	-0.02 (-0.96 to 0.91)	C C	0.11 (-0.84 to 1.05)	Ū		
0.68 (-0.35 to 1.7)	0.41 (-0.74 to 1.56)	0.53 (-0.41 to 1.47)	0.55 (-0.53 to 1.63)	0.61 (-0.51 to 1.73)	0.50 (-0.77 to 1.77)		
0.44 (–0.71 to 1.58)	0.17 (-1.37 to 1.71)	0.27 (-0.79 to 1.33)	0.29 (-1.10 to 1.69)	0.28 (-1.03 to 1.60)	0.18 (–1.43 to 1.78)		

90.2%; P = 0.030). There were no other significant differences by intervention in baseline characteristics. **Outcomes**

In planned analyses comparing study end points, CEP compared with RS significantly decreased the odds of reduced MHRQL at 6 months (P = 0.009) and 12 months (P = 0.028) (Table 2; Figure 1 in Data Supplement 4). In an analysis of change from baseline in likelihood of poor MHRQL, CEP also showed a significant advantage at 6 months (P = 0.038), but not at 12 months.

A modest degree of nonignorability in imputations for missing data or changing the representation of time in statistical models from a continuous to a categorical variable affects interpretations, with most findings becoming either borderline significant or nonsignificant, but with a direction favoring CEP. In addition, sensitivity analyses reflecting MCS-12 on a continuous scale did not reveal any significant differences between interventions at 6 or 12 months (**Table 2**; **Figures 2** and **3** in **Data Supplement 4**).

Service Utilization

Analyses comparing percentages of any behavioral health hospitalizations in the prior 6 months confirmed a significant reduction in the CEP group at 6 months (P = 0.042), but no significant difference at 12 months (**Table 3**). When analyzed as change from baseline, CEP showed significant reductions in likelihood of behavioral health hospitalizations at 6 months (P = 0.002) and 12 months (P = 0.002).

At 6 months, qualitatively similar findings were observed for 4 or more behavioral health hospital nights, but at 12 months, the change from baseline was only borderline significant. No observed significant differences between CEP and RS were observed for other services use measures. For certain sectors (such as parks), there were too few users to develop reliable estimates of mean depression visits at 12 months.

Sensitivity analyses with time as a class variable and varying imputed data assumptions confirmed favorable effects of CEP at 6 months on any behavioral health hospitalizations and 4 or more behavioral health hospital nights, but all 12-month results on behavioral health hospitalizations were sensitive to analysis choices.

DISCUSSION

Although the significance of study findings was sensitive to underlying statistical assumptions and CEP effects were not significant in terms of a continuous MCS-12, CEP was found to have advantages over RS in that it reduced the likelihood of poor MHRQL (MCS-12 \leq 40), the primary outcome for depressed clients from health care and social community programs in underresourced, communities of color in Los Angeles. Evidence of persistence of CEP intervention effects at 12 months is less clear, with greater sensitivity of findings to underlying statistical assumptions.

Our analyses confirm the effect of CEP on reducing behavioral hospitalizations at 6 months (30), but the significance of a similar effect from baseline to 12 months is more speculative, owing to sensitivity to statistical methods. We found no significant differences by intervention status on utilization variables, including health care-based depression treatments (medication or counseling). For some sectors (such as parks), there were too few users of depression services to estimate differences in mean visits. Overall, the shift of outpatient visits toward alternative sectors reported at 6 months was not apparent at 12 months (30). In addition, the baseline findings reported here differ slightly from those in prior publications, owing to differences in weighting and statistical analysis procedures (29, 30).

The effects of CEP at 12 months may have been due to decreased intervention support after the first 6 months, or the variable level of CEP implementation resulting in clients with positive outcomes being outweighed by clients with no evidence of positive outcomes. Future research should examine whether additional implementation sup-

Table 3.	Services	Utilization	at 6	and	12	Months*
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Measure	Clients, n	CEP Group	RS Group	Odds or Rate Ratio (95% CI)		
				Primary Analysist		
				CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	
Any behavioral h	ealth hospitalizations in	the past 6 mo, <i>n</i> (%)				
Baseline	970	77/482 (16.0)	58/488 (11.9)	1.38 (0.91–2.11)	-	
6 mo	759	22/379 (5.8)	35/380 (9.2)	0.60 (0.37–0.98)**	0.43 (0.26–0.74)**	
12 mo	731	18/367 (4.9)	17/364 (4.7)	0.70 (0.4–1.22)	0.51 (0.27–0.96)**	
≥4 behavioral he	ealth hospital nights, <i>n</i> (·%)				
Baseline	970	44/482 (9.1)	33/488 (6.8)	1.33 (0.85–2.08)	-	
6 mo	759	8/379 (2.1)	19/380 (5.0)	0.47 (0.23–0.97)**	0.35 (0.17–0.75)**	
12 mo	730	10/367 (2.7)	9/363 (2.5)	0.70 (0.33–1.48)	0.52 (0.25–1.11)	
≥2 emergency de	epartment visits, n (%)					
Baseline	970	167/482 (34.6)	177/488 (36.3)	0.88 (0.62–1.23)	-	
6 mo	759	91/379 (24.0)	107/380 (28.2)	0.79 (0.57–1.11)	0.91 (0.6–1.38)	
12 mo	730	75/367 (20.4)	88/363 (24.2)	0.75 (0.55–1.03)	0.86 (0.58–1.27)	
Any mental healt	h outpatient visit, n (%)				
Baseline	970	277/482 (57.5)	286/488 (58.6)	1.13 (0.73–1.77)	-	
6 mo	758	206/378 (54.5)	207/380 (54.5)	1.19 (0.78–1.81)	1.04 (0.69–1.57)	
12 mo	728	165/366 (45.1)	163/362 (45.0)	1.05 (0.66–1.66)	0.92 (0.56–1.52)	
Any primary care	visit. <i>n</i> (%)					
Baseline	969	335/482 (69.5)	333/487 (68.4)	0.98 (0.68-1.43)	_	
6 mo	759	262/379 (69.1)	262/380 (68.9)	1.06 (0.76–1.47)	1.07 (0.78–1.48)	
12 mo	729	263/366 (71.9)	231/363 (63.6)	1.31 (0.98–1.76)	1.33 (0.90–1.99)	
Any primary care	visit with depression se	ervice, n (%)				
Baseline	952	195/472 (41.3)	196/480 (40.8)	0.93 (0.66–1.31)	-	
6 mo	756	117/377 (31.0)	111/379 (29.3)	1.05 (0.77–1.45)	1.13 (0.82–1.56)	
12 mo	728	106/366 (29.0)	89/362 (24.6)	1.03 (0.74–1.42)	1.10 (0.74–1.65)	
Mean (SD) couns	eling visits from mental	health specialty or primary of	are. n			
Baseline	948	7.3 (13.7)	6.4 (3.8)	1.11 (0.81–1.52)	_	
6 mo	755	7.1 (12.5)	8.8 (22.9)	0.73 (0.45–1.19)	0.66 (0.41–1.06)	
12 mo	724	4.7 (10.3)	5.1 (11.2)	0.87 (0.58–1.30)	0.78 (0.51–1.21)	
Any faith-based	program participation, <i>n</i>	(%)				
Baseline	967	280/481 (58.2)	299/486 (61.5)	0.84 (0.62–1.14)	-	
6 mo	759	217/379 (57.3)	229/380 (60.3)	0.84 (0.60–1.17)	1.00 (0.72–1.40)	
12 mo	729	195/366 (53.3)	214/363 (59.0)	0.79 (0.60–1.05)	0.95 (0.67–1.34)	
Any use of parks	or community centers,	n (%)				
Baseline	967	225/481 (46.8)	239/486 (49.2)	0.88 (0.67–1.16)	-	
6 mo	759	150/379 (39.6)	161/380 (42.4)	1.00 (0.72–1.39)	1.14 (0.83–1.55)	
12 mo	730	132/366 (36.1)	133/364 (36.5)	0.97 (0.72–1.32)	1.10 (0.79–1.53)	
Took antidoproce	ant >2 ma in the next (- ma n (9/)				
Baseline	ant = 2 mo in the past 6 945	151/471 (22 1)	145/474 (30.6)	1 30 (0 86-1 96)	_	
6 mo	757	125/277 (22.1)	1/9/380 (39.2)	0.91 (0.55-1.50)	-	
12 mo	730	108/366 (29.5)	123/364 (33.8)	0.87 (0.55–1.39)	0.67 (0.43–1.04)	
				•	· · ·	
Mean (SD) total	outpatient contacts for c	lepression, n				
Baseline	929	28.2 (54.5)	30.4 (51.0)	0.93 (0.68–1.26)	-	
6 mo	759	21.6 (43.9)	21.0 (46.8)	0.96 (0.60–1.53)	1.03 (0.65–1.64)	
12 mo	719	18.0 (40.4)	19.4 (43.6)	0.91 (0.65–1.29)	0.99 (0.69–1.41)	

CEP = community engagement and planning; MAR = missing at random; MCS-12 = 12-item mental health composite score; NMAR = nonignorable missing at random; RS = resources for services.

* Adjusted analyses used multiply imputed data based on 1018 clients. A generalized estimating equation logistic regression model was used for a binary variable (presented as odds ratios). A generalized estimating equation Poisson regression model was used for a count variable (presented as rate ratios), adjusted for age, sex, \geq 3 chronic conditions, education, race/ethnicity, income, family income below the federal poverty level, 12-mo alcohol abuse or use of illicit drugs, 12-mo depressive disorder, and community.

⁴ Analysis with time as a continuous variable and multiple imputation procedures assuming that the missing data are MAR.
[‡] Analysis with time as a continuous variable and NMAR by multiplying the ignorable model's imputed data by 1.1.

§ Analysis with time as a continuous variable and NMAR by multiplying the ignorable model's imputed data by 1.1. § Analysis with time as a continuous variable and NMAR by multiplying the ignorable model's imputed data by 0.9. || Analysis with time as a categorical variable with 2 indicators for 6- and 12-mo time points and MAR. ** P < 0.05.

Table 3—Continued

Odds or Rate Ratio (95% CI)						
Higher NMAR Value†‡		Lower N/	NMAR Value†§ Categorical Time Va		l Time Value†∥	
CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	CEP vs. RS at Specific Time	CEP vs. RS in Change From Baseline	
1 34 (0 89–2 03)	_	1 35 (0 88–2 05)	_	1 41 (0 93–2 16)	_	
0.71 (0.42–1.18)	0.53 (0.31_0.91)**	0.55 (0.24–1.23)	0.41 (0.19_0.89)**	0.53 (0.30-0.95)**	0 38 (0 20-0 72)**	
0.79 (0.40–1.55)	0.59 (0.29–1.18)	0.65 (0.30–1.41)	0.48 (0.21–1.09)	0.92 (0.50–1.66)	0.65 (0.34–1.24)	
1.31 (0.82–2.09)	-	1.28 (0.82-2.00)	_	1.39 (0.87–2.20)	-	
0.55 (0.27–1.10)	0.42 (0.21-0.82)**	0.38 (0.16-0.91)**	0.29 (0.12-0.69)**	0.36 (0.16-0.83)**	0.26 (0.10-0.65)**	
0.79 (0.31–2.00)	0.60 (0.24–1.50)	0.57 (0.26–1.26)	0.45 (0.20-1.00)**	1.01 (0.43–2.34)	0.73 (0.33–1.58)	
0.87 (0.62–1.22)	-	0.88 (0.63–1.23)	-	0.89 (0.63–1.26)	-	
0.81 (0.59–1.12)	0.93 (0.64–1.37)	0.87 (0.61–1.25)	0.99 (0.67–1.46)	0.76 (0.52–1.12)	0.86 (0.53–1.37)	
0.77 (0.53–1.11)	0.88 (0.57–1.37)	0.77 (0.52–1.14)	0.87 (0.55–1.38)	0.77 (0.54–1.10)	0.87 (0.58–1.30)	
1.12 (0.72–1.75)	-	1.12 (0.72–1.75)	-	1.15 (0.74–1.79)	-	
1.29 (0.80–2.08)	1.15 (0.72–1.84)	1.23 (0.81–1.86)	1.09 (0.72–1.66)	1.14 (0.76–1.71)	0.99 (0.67–1.45)	
1.20 (0.74–1.93)	1.06 (0.66–1.71)	1.04 (0.69–1.55)	0.92 (0.58–1.47)	1.08 (0.68–1.73)	0.94 (0.57–1.54)	
0.98 (0.67–1.42)	-	1.00 (0.69–1.46)	-	1.00 (0.69–1.45)	-	
1.11 (0.78–1.57)	1.13 (0.79–1.62)	1.08 (0.70–1.68)	1.08 (0.71–1.64)	0.98 (0.66–1.46)	0.99 (0.67–1.46)	
1.38 (1.03–1.86)**	1.41 (0.97–2.05)	1.33 (0.90–1.97)	1.33 (0.83–2.13)	1.44 (1.05–1.99)**	1.45 (0.94–2.23)	
0.95 (0.67_1.34)	_	0.95 (0.68_1.33)	_	0.94 (0.66 - 1.34)	_	
1 14 (0 82–1 58)	1 20 (0 86–1 67)	1 17 (0 83–1 66)	1 24 (0 89–1 72)	0.96 (0.66–1.38)	1 02 (0 68–1 52)	
1.14 (0.82 1.50)	1 19 (0 80–1 76)	1 10 (0 78–1 54)	1 16 (0 78–1 73)	1 12 (0 79–1 58)	1.19 (0.79–1.80)	
1.12 (0.81-1.53)	-	1.10 (0.80–1.51)	-	1.13 (0.83–1.54)	-	
0.73 (0.44–1.21)	0.65 (0.40-1.07)	0.74 (0.46–1.17)	0.67 (0.43–1.05)	0.70 (0.43–1.15)	0.62 (0.38–1.01)	
0.88 (0.58–1.32)	0.79 (0.50–1.23)	0.86 (0.58–1.28)	0.78 (0.51–1.19)	0.96 (0.64–1.43)	0.85 (0.53-1.35)	
0.83 (0.60–1.13)	_	0.81 (0.58–1.13)	_	0.83 (0.61–1.14)	-	
0.81 (0.56-1.16)	0.98 (0.68–1.42)	0.82 (0.59–1.13)	1.01 (0.70–1.46)	0.83 (0.62–1.12)	1.00 (0.71-1.40)	
0.73 (0.54–0.98)**	0.88 (0.62–1.25)	0.81 (0.60–1.10)	1.01 (0.70–1.44)	0.80 (0.58–1.11)	0.96 (0.66–1.39)	
0.00 (0.69, 1.10)		0.91 (0.69, 1.20)				
1.00 (0.73-1.38)	- 1 11 (0 82–1 50)	1.04 (0.76-1.20)	- 1 16 (0 80–1 67)	-	-	
0.99(0.75 - 1.38)	1.10 (0.82–1.50)	1.04 (0.73-1.43)	1.10 (0.80–1.67)	1.05 (0.76-1.45)	1 15 (0 81–1 65)	
0.99 (0.70-1.41)	1.10 (0.80-1.21)	1.00 (0.75-1.56)	1.12 (0.78-1.00)	1.05 (0.76-1.45)	1.13 (0.01-1.05)	
1.27 (0.83–1 92)	_	1.28 (0.85–1 92)	_	1.30 (0.87–1.96)	_	
1.02 (0.62–1.67)	0.81 (0.55–1.18)	0.93 (0.62–1.40)	0.73 (0.51-1.03)	0.86 (0.50–1.47)	0.66 (0.40-1.08)	
0.90 (0.59–1.39)	0.71 (0.48–1.06)	0.87 (0.57–1.34)	0.68 (0.42–1.11)	0.93 (0.60–1.44)	0.72 (0.47–1.10)	
0.02 (0.00 1.07)				0.02 (0.62, 1.27)		
0.93 (0.69–1.27)		0.92 (0.68-1.26)		0.93 (0.68–1.27)	-	
0.92 (0.58 - 1.56)	1.02 (0.63 - 1.67)	0.96(0.62-1.51)	1.04 (0.67 - 1.62)	0.95(0.57 - 1.56)	1.02 (0.62-1.69)	
0.92 (0.65-1.30)	0.99 (0.69-1.41)	0.91 (0.64–1.29)	0.99 (0.69-1.42)	0.93 (0.64–1.36)	1.01 (0.69–1.47)	

port would offer more consistent evidence of sustained CEP effects beyond 6 months.

Our study has important limitations. We did not have a usual care group, but rather compared 2 active interventions that are each likely to be effective relative to usual care. We did not have data on hospitalization and medication use for general health conditions other than behavioral health. Because our sample includes only 1018 clients, precision was low for definitive services use estimates. The study was conducted in 2 Los Angeles communities where study leaders have a long history of applying communitypartnered participatory research to depression (58–63). It

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is unknown whether applying this approach in communities without this history would yield similar effects.

In addition, response rates were moderate for agencies and high for programs. Although initial client enrollment rates were high, retention was lower relative to other studies of quality improvement in depression care, but similar to that in studies of clients in safety-net settings (64, 65). Client outcomes relied on self-reported data, and clinical process data linking programs to clients were unavailable. We did not adjust significance for multiple comparisons because, as noted in our protocol, we focused on 1 primary outcome: poor MHRQL.

Finally, the significance of CEP effects was sensitive to underlying statistical assumptions of representation of time in models (class or continuous variable); to possible departures from nonignorable model predictions for imputed values; and to whether we used a generalized estimating equation longitudinal analysis with an exchangeable working correlation assumption or a design-based analysis using SUDAAN (RTI International) to incorporate sampling and nonresponse weights for 12-month outcomes (**Table B8** in **Data Supplement 3**).

In conclusion, our results confirm the short-term effect of CEP on reducing the percentage of depressed clients with poor MHRQL and behavioral health hospitalizations at 6 months, with less evident effects at 12 months. Shortterm change in avoiding poor quality of life and behavioral health hospitalizations, and possibly longer term, are clinically important owing to consistent mental health disparities (3-5), depression-related costs (66, 67), and the recurrent chronicity of depression (68, 69). Given the unmet needs of underresourced communities, the absence of evidence-based alternatives, and the modestly favorable profile and limited risk of CEP, community engagement remains a viable strategy that policymakers and communities could consider for collaborative care implementation (70) to improve population-based health outcomes of depression among vulnerable individuals.

From Harbor-UCLA Medical Center/Los Angeles Biomedical Research Institute, Torrance; Center for Health Services and Society, Semel Institute for Neuroscience and Human Behavior, David Geffen School of Medicine at UCLA, Greater Los Angeles Veterans Affiars Healthcare System, UCLA Fielding School of Public Health, Healthy African American Families II, and School of Nursing at UCLA, Los Angeles; RAND Corporation, Santa Monica; and Behavioral Health Services, Gardena, California.

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Requests for Single Reprints: Bowen Chung, MD, MSHS, Department of Psychiatry, Harbor-UCLA Medical Center, 1000 West Carson Street, Box 498, Torrance, CA 90509; e-mail, bchung@mednet.ucla .edu.

Current Author Addresses: Dr. Chung: Department of Psychiatry, Harbor-UCLA Medical Center, 1000 West Carson Street, Box 498, Torrance, CA 90509.

Dr. Ong: Division of General Internal Medicine and Health Services Research, David Geffen School of Medicine at UCLA, Box 957394, 10940 Wilshire Boulevard, Suite 700, Los Angeles, CA 90095-7394.

Dr. Ettner: Division of General Internal Medicine and Health Services Research, David Geffen School of Medicine at UCLA, Box 951736, 911 Broxton Plaza, Los Angeles, CA 90095-1736.

Ms. Jones: Healthy African American Families II, 4305 Degnan Boulevard, Suite 105, Los Angeles, CA 90008.

Mr. Gilmore: Behavioral Health Services, 15519 Crenshaw Boulevard, Gardena, CA 90249.

Mr. McCreary and Drs. Tang, Miranda, and Wells: Center for Health Services and Society, Semel Institute for Neuroscience, Box 957082, 10920 Wilshire Boulevard, Suite 300, Los Angeles, CA 90095-7082.

Drs. Sherbourne, Ngo, and Koegel: RAND Corporation, 1776 Main Street, PO Box 2138, Santa Monica, CA 90407-2138.

Dr. Dixon: UCLA School of Nursing, Box 956919, 5-660 Factor, Los Angeles, CA 90095-6919.

Dr. Belin: UCLA School of Public Health, Department of Biostatistics, Box 951772, 51-267 CHS, Los Angeles, CA 90095-1772.

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Author Contributions: Conception and design: B. Chung, M.K. Ong, S.L. Ettner, J. Gilmore, C.D. Sherbourne, V. Ngo, L. Tang, T.R. Belin, K.B. Wells.

Analysis and interpretation of the data: B. Chung, M.K. Ong, S.L. Ettner, J. Gilmore, M. McCreary, C.D. Sherbourne, L. Tang, J. Miranda, T.R. Belin, K.B. Wells.

Drafting of the article: B. Chung, M.K. Ong, M. McCreary, C.D. Sherbourne, P. Koegel, L. Tang, J. Miranda.

Critical revision of the article for important intellectual content: B.

Chung, M.K. Ong, S.L. Ettner, F. Hones, C.D. Sherbourne, V. Ngo, P. Koegel, L. Tang, J. Miranda, T.R. Belin, K.B. Wells.

Final approval of the article: B. Chung, M.K. Ong, S.L. Ettner, F. Hones, J. Gilmore, C.D. Sherbourne, V. Ngo, P. Koegel, L. Tang, E. Dixon, J. Miranda, T.R. Belin, K.B. Wells.

Provision of study materials or patients: J. Gilmore, M. McCreary, V. Ngo, E. Dixon, K.B. Wells.

Statistical expertise: B. Chung, M.K. Ong, S.L. Ettner, L. Tang, J. Miranda, T.R. Belin.

Obtaining of funding: M.K. Ong, C.D. Sherbourne, K.B. Wells, B. Chung.

Administrative, technical, or logistic support: B. Chung, M.K. Ong, S.L. Ettner, J. Gilmore, M. McCreary, V. Ngo, P. Koegel, L. Tang, E. Dixon, K.B. Wells.

Collection and assembly of data: B. Chung, J. Gilmore, M. McCreary, C.D. Sherbourne, L. Tang, E. Dixon, K.B. Wells.

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