



A Central Limit Theorem for the Volumes of High Excursions of Stationary Associated Random Fields

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Abstract We prove that under certain conditions the excursion sets volumes of stationary positively associated random fields converge after rescaling to the normal distribution as the excursion level and the size of the observation window grow. In addition, we provide a number of examples.

Keywords Association, Excursion Set, Central Limit Theorem, Gaussian Random Field, Shot-Noise

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1. Introduction

The study of excursion set properties plays an increasingly important role within modern theory of random fields. Such objects arise in connection to a wide range of stochastic models (see, e.g., [1, 2]). Recently a number of works appeared ([3, 4, 5]) focused on asymptotic properties of the excursion sets volumes corresponding to fixed excursion levels and sequences of growing (in a certain sense) observation windows. Spodarev gives in [6] an overview of the recent asymptotic results concerning the geometry of excursion sets of stationary associated random fields.

In this paper, we show that under certain conditions the excursion sets volumes of stationary positively associated random fields converge after rescaling to the normal distribution as both the excursion level and the size of the observation window grow. Note that a similar model was studied in the monograph of Ivanov and Leonenko [7] for the case of a Gaussian random field. For associated random fields on lattices the asymptotic behaviour of excursion sets cardinalities corresponding to a growing excursion level was examined in [8].

The present paper is organized as follows. In section 2 we provide the necessary definitions. Section 3 contains the main theorem. In section 4 we show how its conditions can be verified for Gaussian random fields, fields with regularly varying tails and shot-noise random fields.

2. Preliminaries

We assume that all random objects are defined on a complete probability space $(\Omega, \mathcal{F}, \mathbb{P})$.

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Consider a family of random variables $X = \{X_t, t \in T\}$. For $I \subset T$ set $X_I = \{X_t, t \in I\}$. Let $M(d)$ denote the class of real-valued bounded coordinate-wise nondecreasing Borel functions on $\mathbb{R}^d, d \in \mathbb{N}$. According to [9], X is called *positively associated* ($X \in PA$) or *weakly associated*, if for arbitrary disjoint finite sets $I, J \subset T$ and any functions $f \in M(\text{card}(I)), g \in M(\text{card}(J))$ one has

$$\text{Cov}(f(X_I), g(X_J)) \geq 0.$$

Here $\text{card}(I)$ denotes the cardinality of I . It is known that quite a wide range of random fields possesses the property of positive association (see [10]). For example, any Gaussian random field with nonnegative covariance function is positively associated.

Let $A_u(X, T) = \{t \in T : X_t \geq u\}$ be the *excursion set* of the random field $X = \{X_t, t \in \mathbb{R}^d\}$ at the level $u \in \mathbb{R}$ corresponding to the observation window $T \subset \mathbb{R}^d$ and let $v_d(B)$ be the volume of measurable $B \subset \mathbb{R}^d$. By $\mathbb{I}(C)$ we denote the indicator function of the set C . Obviously,

$$V_u(X, T) = v_d(A_u(X, T)) = \int_T \mathbb{I}(X_t \geq u) dt.$$

Let $\|\cdot\|$ and $\|\cdot\|_\infty$ be, respectively, the Euclidean norm and the maximum norm in \mathbb{R}^d and $\mathcal{O}(\cdot)$ — the Landau notation.

3. Central Limit Theorem

Let X be a measurable strictly stationary positively associated random field on \mathbb{R}^d . We also assume that X is square integrable and its covariance function $r(t) = \text{Cov}(X_0, X_t), t \in \mathbb{R}^d$, is continuous. Note that due to the latter condition X is also associated (see [10] for the definition of association and related dependence types).

For any $u \in \mathbb{R}$ and bounded measurable $B, C \subset \mathbb{R}^d$ Fubini's theorem implies

$$\mathbb{E} V_u(X, B) = v_d(B) \mathbb{P}(X_0 \geq u),$$

$$\text{Cov}(V_u(X, B), V_u(X, C)) = \int_B \int_C \text{Cov}(\mathbb{I}(X_s \geq u), \mathbb{I}(X_t \geq u)) ds dt \geq 0. \tag{1}$$

Consider an increasing sequence of excursion levels $\{u_n\}_{n \in \mathbb{N}}, 0 < u_n \rightarrow \infty, n \rightarrow \infty$. Define $S_n = V_{u_n}(X, [0, n]^d), n \in \mathbb{N}$, and let $\sigma_n^2 = \text{Var } S_n$.

Suppose that the random variable X_0 has a bounded density $f(\cdot)$ and set $\gamma(x) = \sup_{t \geq x} f(t), x \in \mathbb{R}$.

Theorem 1

Assume that the random field X satisfies the following conditions:

- (A1) $r(t) = \mathcal{O}(\|t\|^{-\mu})$ for some $\mu > 3d$;
- (A2) $\sigma_n \rightarrow \infty, n \rightarrow \infty$;
- (A3) $\delta_n = n^d \gamma^{2/3}(u_n) \sigma_n^{-(1+\mu/3d)} \rightarrow 0, n \rightarrow \infty$.

Then

$$\frac{S_n - \mathbb{E} S_n}{\sqrt{\text{Var } S_n}} \xrightarrow{d} \mathcal{N}(0, 1),$$

where \xrightarrow{d} denotes convergence in distribution.

The following lemma can be proved in the same way as Lemma 7.3.4 from [10].

Lemma 1

Under the conditions of Theorem 1

$$\text{Cov}(\mathbb{I}(X_s \geq u), \mathbb{I}(X_t \geq u)) \leq 3 \cdot 2^{2/3} \gamma^{2/3}(u) r^{1/3}(s-t), \quad u \in \mathbb{R}, s, t \in \mathbb{R}^d. \tag{2}$$

Proof of Theorem 1. One can find a sequence $m_n \rightarrow \infty, n \rightarrow \infty$, such that $m_n^d/\sigma_n \rightarrow 0$ and

$$\frac{n^d}{\sigma_n^2} \gamma^{2/3}(u_n) m_n^{d-\mu/3} \rightarrow 0, \quad n \rightarrow \infty. \tag{3}$$

To see this, consider $m_n = \max \left\{ \sigma_n^{1/(2d)}, \delta_n^{1/(\mu-3d)} \sigma_n^{1/d} \right\}$.

Set $q_n = \sigma_n^{1/(2d)} m_n^{1/2}$ and $r_n = \left\lfloor \frac{n}{q_n} \right\rfloor$. Note that

$$\frac{m_n}{q_n} \rightarrow 0, \quad \frac{q_n^d}{\sigma_n} \rightarrow 0, \quad n \rightarrow \infty. \tag{4}$$

Using (1), (A1) and (2), we also get $\sigma_n^2 = \mathcal{O}(n^d), n \rightarrow \infty$. Hence, $r_n \rightarrow \infty$ as $n \rightarrow \infty$. Clearly, there exists $N_0 \in \mathbb{N}$ such that $0 < m_n^d \leq q_n^d \leq \sigma_n$ for all $n > N_0$.

Define

$$\tilde{S}_n = V_{u_n}(\mathbf{X}, [0, r_n q_n]^d) = \sum_{k \in \mathbb{Z}^d \cap [0, r_n]^d} \zeta_{n,k}, \quad Z_n = \sum_{k \in \mathbb{Z}^d \cap [0, r_n]^d} \xi_{n,k},$$

where $\zeta_{n,k} = V_{u_n}(\mathbf{X}, [0, q_n]^d \oplus kq_n)$, $k \in \mathbb{Z}^d \cap [0, r_n]^d$, and $\{\xi_{n,k}, k \in \mathbb{Z}^d \cap [0, r_n]^d\}$ are independent random variables distributed in the same way as $V_{u_n}(\mathbf{X}, [0, q_n]^d)$, $n > N_0$. Approximating integrals with finite sums in L_1 (see [4]) it is not difficult to show that, since the random field \mathbf{X} is positively associated and its covariance function is continuous, $\{\zeta_{n,k}, k \in \mathbb{Z}^d \cap [0, r_n]^d\} \in \text{PA}$.

Before we continue with the proof of the theorem, we establish the following lemma.

Lemma 2

Under the conditions of Theorem 1 it holds that

$$\frac{\text{Var } Z_n}{\text{Var } S_n} \rightarrow 1, \quad n \rightarrow \infty.$$

Proof. Due to (1), the variance of Z_n is not greater than the variance of S_n . Thus, it suffices to prove that

$$\limsup_{n \rightarrow \infty} \frac{\text{Var } S_n}{\text{Var } Z_n} \leq 1. \tag{5}$$

Observe that for $n > N_0$

$$\begin{aligned} \text{Var } S_n &= \text{Var } Z_n + \sum_{k \in \mathbb{Z}^d \cap [0, r_n]^d} \text{Cov} (V_{u_n}(\mathbf{X}, [0, q_n]^d \oplus kq_n), V_{u_n}(\mathbf{X}, [0, n]^d \setminus ([0, q_n]^d \oplus kq_n))) \\ &\quad + \text{Cov} (V_{u_n}(\mathbf{X}, [0, n]^d \setminus [0, r_n q_n]^d), V_{u_n}(\mathbf{X}, [0, n]^d)) = \text{Var } Z_n + \Sigma_1 + \Sigma_2. \end{aligned}$$

Below we show that Σ_1 and Σ_2 admit the following estimates

$$\Sigma_1 \leq C_1 n^d \gamma^{2/3}(u_n) m_n^{d-\mu/3} + C_2 \frac{m_n}{q_n} \text{Var } Z_n, \tag{6}$$

$$\Sigma_2 \leq C_3 r_n^{-1} \text{Var } S_n, \quad n > N_0. \tag{7}$$

Here, and in the following, C_1, C_2, C_3, \dots are some positive real numbers which may only depend on d, r and μ . Applying (6) and (7), we have

$$\frac{\text{Var } S_n}{\text{Var } Z_n} \left(1 - C_3 r_n^{-1} - C_1 \frac{n^d}{\sigma_n^2} \gamma^{2/3}(u_n) m_n^{d-\mu/3} \right) \leq 1 + C_2 \frac{m_n}{q_n}, \quad n > N_0.$$

Due to (3) and (4), the latter inequality implies (5).

Now let us obtain (6). It can be easily seen that (1) yields

$$\begin{aligned} & \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d \oplus kq_n), V_{u_n}(\mathbf{X}, [0, n]^d \setminus ([0, q_n]^d \oplus kq_n))) \\ & \leq \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d), V_{u_n}(\mathbf{X}, [-n, 2n]^d \setminus [0, q_n]^d)), k \in \mathbb{Z}^d \cap [0, r_n]^d. \end{aligned}$$

Thus, $\Sigma_1 \leq R_1 + R_2$, where

$$\begin{aligned} R_1 &= r_n^d \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d), V_{u_n}(\mathbf{X}, [-n, 2n]^d \setminus [-m_n, q_n + m_n]^d)), \\ R_2 &= r_n^d \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d), V_{u_n}(\mathbf{X}, [-m_n, q_n + m_n]^d \setminus [0, q_n]^d)). \end{aligned}$$

Using (2) and (A1), we get

$$\begin{aligned} R_1 &\leq C_4 \gamma^{2/3}(u_n) r_n^d q_n^d \int_{\|x\|_\infty > m_n} r^{1/3}(x) dx \\ &\leq C_5 \gamma^{2/3}(u_n) n^d \int_{\|x\|_\infty > m_n} \|x\|^{-\mu/3} dx \leq C_1 n^d \gamma^{2/3}(u_n) m_n^{d-\mu/3}. \end{aligned}$$

Now we show that $R_2 \leq C_2 (m_n/q_n) \text{Var } Z_n$. Set $\tau = \lfloor \frac{2q_n}{m_n} \rfloor$. Clearly,

$$\begin{aligned} \frac{m_n}{q_n} \text{Var } Z_n &= \frac{m_n}{q_n} r_n^d \text{Var } V_{u_n}(\mathbf{X}, [0, q_n]^d) \geq C_6 \frac{m_n}{q_n} r_n^d \text{Var } V_{u_n}(\mathbf{X}, [-2q_n, 3q_n]^d) \\ &\geq C_6 \frac{m_n}{q_n} r_n^d \text{Cov}(V_{u_n}(\mathbf{X}, [-2q_n, 3q_n]^d), V_{u_n}(\mathbf{X}, [-2q_n, 3q_n]^d \setminus [0, q_n]^d)) \\ &\geq C_6 \frac{m_n}{q_n} r_n^d \text{Cov}(V_{u_n}(\mathbf{X}, [-2q_n, 3q_n]^d), \\ &\quad \sum_{i=1}^{\tau} V_{u_n}(\mathbf{X}, [-im_n, q_n + im_n]^d \setminus [-(i-1)m_n, q_n + (i-1)m_n]^d)) \\ &\geq C_6 \frac{m_n}{q_n} r_n^d \sum_{i=1}^{\tau} \text{Cov}(V_{u_n}(\mathbf{X}, [-(i-1)m_n, q_n + (i-1)m_n]^d), \\ &\quad V_{u_n}(\mathbf{X}, [-im_n, q_n + im_n]^d \setminus [-(i-1)m_n, q_n + (i-1)m_n]^d)) \\ &\geq C_6 r_n^d \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d), V_{u_n}(\mathbf{X}, [-m_n, q_n + m_n]^d \setminus [0, q_n]^d)). \end{aligned}$$

Hence, the desired inequality holds for $C_2 = 1/C_6$. To conclude the proof of the lemma, it remains to note that the method we used to estimate R_2 also yields (7). Lemma 2 is proved. \square

Due to (7) and Lemma 2 the terms $\frac{S_n - \mathbb{E} S_n}{\sqrt{\text{Var } S_n}}$ and $\frac{\tilde{S}_n - \mathbb{E} \tilde{S}_n}{\sqrt{\text{Var } \tilde{S}_n}}$ have the same limiting distribution as $n \rightarrow \infty$. Thus, it suffices to show that $\frac{\tilde{S}_n - \mathbb{E} \tilde{S}_n}{\sqrt{\text{Var } \tilde{S}_n}}$ has the same limiting distribution as $\frac{Z_n - \mathbb{E} Z_n}{\sqrt{\text{Var } Z_n}}$ and that $\frac{Z_n - \mathbb{E} Z_n}{\sqrt{\text{Var } Z_n}}$ converges in distribution to the standard Gaussian law as $n \rightarrow \infty$.

Applying Newman's inequality (cf. [10, Corollary 1.5.5]), we get

$$\begin{aligned} & \left| \mathbb{E} \exp \left(it \frac{\tilde{S}_n - \mathbb{E} \tilde{S}_n}{\sqrt{\text{Var } \tilde{S}_n}} \right) - \mathbb{E} \exp \left(it \frac{Z_n - \mathbb{E} Z_n}{\sqrt{\text{Var } Z_n}} \right) \right| \\ & \leq \sum_{\substack{k, l \in \mathbb{Z}^d \cap [0, r_n]^d \\ k \neq l}} \frac{2t^2}{\text{Var } Z_n} \text{Cov}(V_{u_n}(\mathbf{X}, [0, q_n]^d \oplus kq_n), V_{u_n}(\mathbf{X}, [0, q_n]^d \oplus lq_n)) \\ & \leq \frac{2t^2}{\text{Var } Z_n} \Sigma_1, \end{aligned} \tag{8}$$

where Σ_1 appears in the proof of Lemma 2. Now it is implied by (A2), (6), (3), Lemma 2 and (4) that the right-hand side of (8) tends to zero as $n \rightarrow \infty$.

As for the convergence of $\frac{Z_n - \mathbb{E} Z_n}{\sqrt{\text{Var } Z_n}}$ to $\mathcal{N}(0, 1)$, it follows from the central limit theorem in the form of Lindeberg. Indeed, the Lindeberg function

$$L_n(\varepsilon) = \frac{1}{\text{Var } Z_n} \sum_{k \in \mathbb{Z}^d \cap [0, r_n]^d} \mathbb{E}(\xi_{n,k} - \mathbb{E} \xi_{n,k})^2 \mathbb{I} \left(|\xi_{n,k} - \mathbb{E} \xi_{n,k}| > \varepsilon \sqrt{\text{Var } Z_n} \right)$$

is equal to zero for any $\varepsilon > 0$ and $n > n_0(\varepsilon)$, since for all k

$$\frac{|\xi_{n,k} - \mathbb{E} \xi_{n,k}|}{\sqrt{\text{Var } Z_n}} \leq \frac{q_n^d}{\sqrt{\text{Var } Z_n}} \rightarrow 0, \quad n \rightarrow \infty. \tag{9}$$

In order to obtain (9), we employed (4) and Lemma 2. The other conditions of the Lindeberg theorem are also easily checked. Theorem 1 is proved. \square

4.1. Gaussian Random Fields

Let $X = \{X_t, t \in \mathbb{R}^d\}$ be a measurable stationary positively associated Gaussian random field with $\mathbb{E} X_0 = 0$ and $\text{Var } X_0 = 1$. Suppose that the covariance function $r(\cdot)$ is continuous and fulfills (A1). Consider

$$u_n \leq \sqrt{c \log n}, \quad c = \frac{d(\mu - 3d)}{(1 + \nu)(\mu + d)}, \quad n \in \mathbb{N},$$

where ν is some positive number. For any $s, t \in \mathbb{R}^d$ it holds (see [4]) that

$$\text{Cov}(\mathbb{I}(X_s \geq u_n), \mathbb{I}(X_t \geq u_n)) = \frac{1}{2\pi} \int_0^{r(t-s)} \frac{1}{\sqrt{1-x^2}} \exp\left(-\frac{u_n^2}{1+x}\right) dx, \quad n \in \mathbb{N}.$$

Consequently,

$$\text{Var } S_n \geq \frac{1}{2\pi} \exp(-u_n^2) \int_{[0,n]^d} \int_{[0,n]^d} r(t-s) dt ds.$$

Since $r(\cdot)$ is continuous and $r(0) = 1$, we have

$$\liminf_{n \rightarrow \infty} \frac{\text{Var } S_n}{n^d \exp(-u_n^2)} > 0. \tag{10}$$

It is easy to show that (10) implies (A2). Applying (10) and the fact that

$$\gamma(u_n) = \frac{1}{\sqrt{2\pi}} \exp(-u_n^2/2),$$

we also obtain

$$\delta_n = \mathcal{O} \left(n^{-(\mu-3d)/6} \exp\left(\frac{1}{6} (1 + \mu/d) u_n^2\right) \right), \quad n \rightarrow \infty.$$

It is easily seen that for our choice of c

$$\frac{n^{c(1+\mu/d)}}{n^{\mu-3d}} \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Therefore, $\delta_n \rightarrow 0$ as $n \rightarrow \infty$.

The case of a Gaussian random field with non-integrable covariance function was considered by Ivanov and Leonenko in [7]. It is interesting to note that the central limit theorem they obtained features the sequence $\{u_n, n \in \mathbb{N}\}$ growing rapidly enough for $\text{Var } S_n$ to converge to 0 as $n \rightarrow \infty$.

4.2. Random Fields with Regularly Varying Tails

Let $X = \{X_t, t \in \mathbb{R}^d\}$ be a measurable stationary square integrable positively associated random field. Assume that its covariance function $r(\cdot)$ belongs to the $C^2(\mathbb{R}^d)$ class and fulfills (A1). We also assume that X_0 has a bounded density. Let F be the distribution function of X_0 . Suppose that it admits, for some $\alpha > 0$, the following representation

$$1 - F(x) = x^{-\alpha} L(x), \quad x > 0,$$

where $L(\cdot)$ is a slowly varying function on $(0, \infty)$, i.e. $1 - F(\cdot)$ is regularly varying in the sense of Karamata (see [11]). We will assume $\alpha > 2$. Note that for every $\varepsilon > 0$

$$x^\varepsilon L(x) \rightarrow \infty, \quad \frac{L(x)}{x^\varepsilon} \rightarrow 0, \quad x \rightarrow \infty. \tag{11}$$

Since $r(\cdot)$ attains its maximum at zero, $\nabla r(0) = 0$. Therefore, using Taylor's expansion, one can find $b > 0$ and $\delta \in (0, 1/2)$ such that

$$r(0) - r(t) \leq b \|t\|^2 \text{ for all } t \in [0, \delta]^d. \tag{12}$$

We show that under certain restrictions on the growth rate of $u_n, n \rightarrow \infty$, X satisfies the conditions of Theorem 1. We need to estimate the following covariance from below

$$\text{Cov}(\mathbb{I}(X_0 \geq u_n), \mathbb{I}(X_t \geq u_n)), \quad t \in [0, \delta]^d, \quad n \in \mathbb{N}.$$

Clearly,

$$\begin{aligned} \text{Cov}(\mathbb{I}(X_0 \geq u_n), \mathbb{I}(X_t \geq u_n)) &= \mathbb{P}(X_0 \geq u_n, X_t \geq u_n) - \mathbb{P}^2(X_0 \geq u_n) \\ &\geq \mathbb{P}(X_0 \geq 2u_n, X_0 - X_t \leq u_n) - \mathbb{P}^2(X_0 \geq u_n) \\ &\geq \mathbb{P}(X_0 \geq 2u_n) - \mathbb{P}(X_0 - X_t \geq u_n) - \mathbb{P}^2(X_0 \geq u_n), \quad n \in \mathbb{N}. \end{aligned}$$

Using Markov's inequality, (12) and the fact that $2(r(0) - r(t)) = \mathbb{E}(X_0 - X_t)^2$, we obtain

$$\begin{aligned} &\mathbb{P}(X_0 \geq 2u_n) - \mathbb{P}(X_0 - X_t \geq u_n) - \mathbb{P}^2(X_0 \geq u_n) \\ &\geq 2^{-\alpha} u_n^{-\alpha} L(2u_n) - 2b \|t\|^2 u_n^{-2} - u_n^{-2\alpha} L^2(u_n) \\ &= \Delta_1(n) - \Delta_2(n) - \Delta_3(n), \quad t \in [0, \delta]^d, \quad n \in \mathbb{N}. \end{aligned}$$

Applying (11), it is easy to show that for arbitrary $\varepsilon > 0$

$$\frac{2b u_n^{-\alpha-2\varepsilon} + \Delta_3(n)}{\Delta_1(n)} \rightarrow 0, \quad n \rightarrow \infty.$$

Hence, for sufficiently large n and $\|t\| \leq u_n^{1-\alpha/2-\varepsilon}$ we have

$$\Delta_1(n) - \Delta_2(n) - \Delta_3(n) \geq \frac{1}{2} \Delta_1(n) \geq A u_n^{-\alpha-\nu}$$

for any $\nu > 0$ and some $A = A(L, \alpha, \nu) > 0$.

Therefore, for any $\varepsilon, \nu > 0$ and large enough n , it holds that

$$\int_{[0, \delta]^d} \text{Cov}(\mathbb{I}(X_0 \geq u_n), \mathbb{I}(X_t \geq u_n)) dt \geq \int_{[0, u_n^{1-\alpha/2-\varepsilon}]^d} \text{Cov}(\mathbb{I}(X_0 \geq u_n), \mathbb{I}(X_t \geq u_n)) dt \geq A u_n^{d-(d+2)\alpha/2-\varepsilon d-\nu}.$$

Consequently,

$$\sigma_n^2 \geq (n/2)^d \int_{[0, \delta]^d} \text{Cov}(\mathbb{I}(X_0 \geq u_n), \mathbb{I}(X_t \geq u_n)) dt \geq A 2^{-d} n^d u_n^{d-(d+2)\alpha/2-\varepsilon d-\nu}.$$

Thus, $\sigma_n \rightarrow \infty$ if $u_n = \mathcal{O}(n^a)$, $n \rightarrow \infty$, where $a < d/((d + 2)\alpha/2 - d)$.

Since the density of X_0 is bounded, we can always find $c \geq 0$ such that

$$\gamma(x) = \mathcal{O}(x^{-c}), \quad x \rightarrow \infty.$$

We have

$$\begin{aligned} \delta_n &= n^d \gamma^{2/3}(u_n) \sigma_n^{-(1+\mu/(3d))} \\ &\leq A_2 u_n^{-2c/3} n^{d/2-\mu/6} u_n^{(1+\mu/(3d))((d+2)\alpha/2-d+\varepsilon d+\nu)/2} \\ &\leq A_2 n^{-1/6(\mu-3d)} n^{\max\{0, -2ac/3+a(\mu+3d)((d+2)\alpha/2-d+\varepsilon d+\nu)/(6d)\}}, \end{aligned}$$

where $A_2 > 0$ does not depend on n . It is not difficult to show that $\delta_n \rightarrow 0$, $n \rightarrow \infty$, if

$$a < \frac{2d(\mu - 3d)}{\max\{0, (\mu + 3d)((d + 2)\alpha - 2d) - 8cd\}},$$

where $s/0 = \infty$, $s \in \mathbb{R}$. Note that if

$$f(x) = x^{-\alpha-1} L_1(x), \quad x > 0,$$

where f is the density of X_0 and L_1 is a slowly varying function on $(0, \infty)$, then c can be taken to be any nonnegative number less than $\alpha + 1$. Consequently, $\delta_n \rightarrow 0$, $n \rightarrow \infty$, provided

$$a < \frac{2d(\mu - 3d)}{\max\{0, (\mu + 3d)((d + 2)\alpha - 2d) - 8d(\alpha + 1)\}}.$$

4.3. Shot-Noise Random Fields

In this section, the requirements of Theorem 1 are checked for certain shot-noise random fields and for the excursion level $u_n = \mathcal{O}((\log n)^{1-\beta})$, $n \rightarrow \infty$, where $\beta \in (0, 1)$. The following definition is taken from [12].

Let ϕ be a stationary Poisson counting measure with intensity $\lambda > 0$. Also, let $g : \mathbb{R} \rightarrow \mathbb{R}$ be a deterministic function with $\int_{\mathbb{R}} g(x) dx < \infty$ and $\int_{\mathbb{R}} g^2(x) dx < \infty$. The random field $X = \{X_t, t \in \mathbb{R}\}$ defined by

$$X_t = \int_{\mathbb{R}} g(t - x) \phi(dx), \quad t \in \mathbb{R}, \tag{13}$$

is called a shot-noise random field. The function g is called a response function.

Here we consider shot-noise random fields with the intensity $\lambda > 1/2$ and the response function

$$g(t) = \exp(-|t|), \quad t \in \mathbb{R}.$$

In this case, X is square integrable and for any $\mu > 3$ (A1) is fulfilled.

The characteristic function $\varphi_{X_0}(\cdot)$ of X_0 is (see [10, Lemma 1.3.7]) given by

$$\varphi_{X_0}(s) = \exp\left(\lambda \int_{\mathbb{R}} \left(e^{ise^{-|t|}} - 1\right) dt\right), \quad s \in \mathbb{R}.$$

It is not difficult to show that for $\lambda > 1/2$ $\varphi_{X_0}(\cdot)$ is integrable and, consequently, X_0 has a bounded density.

In addition, one can show that the distribution of X_0 is infinitely divisible (see, e.g., [13]) with Lévy measure

$$\psi(dx) = 2 \lambda/x \mathbb{I}(x \in (0, 1)) dx.$$

In order to find an approximation for the distribution of X_0 , we consider upper records. Let $\{Z_n : n \in \mathbb{N}\}$ be a sequence of independent and identically distributed observations. The observation Z_j , $j \in \mathbb{N}$, is called an upper record, if its value exceeds all previous observations, that is $Z_j > \max\{Z_1, Z_2, \dots, Z_{j-1}\}$. Since the distribution

of X_0 is infinitely divisible and the Lévy measure ψ has a density l on $[0, \infty)$ with $\int_0^\infty l(y) dy = \infty$, it follows from [14, Theorem 2] that

$$(n + 1) - T_n \longrightarrow X_0 \quad \text{a.s.},$$

where $T_n, n \in \mathbb{N}$, is the sum of the first $(n + 1)$ upper records from a certain distribution H . Moreover, the density of H is known and, in our case, is given by

$$h(x) = 2\lambda (1 - x)^{2\lambda - 1} \mathbb{I}(x \in (0, 1)), \quad x \in \mathbb{R}.$$

Thus, H is the distribution function of a Beta-distributed random variable with parameters 2λ and 1. Such Beta records are considered in [15]. There it is shown that

$$\lim_{n \rightarrow \infty} (n + 1) - T_n \stackrel{d}{=} \sum_{i=0}^\infty \prod_{j=0}^i U_j^{\frac{1}{2\lambda}}, \tag{14}$$

where $U_j, j \in \mathbb{Z}_+$, are independent random variables distributed uniformly on $(0, 1)$.

Vervaat showed in [16, Theorem 4.7.7 and Lemma 4.7.9] that the density $f(x)$ of the random variable in the right-hand side of (14) (which is also the density of X_0) is nonincreasing, for sufficiently large x , and that

$$f(x) = \exp(-(1 + o(1)) x \log x), \quad x \rightarrow \infty. \tag{15}$$

To check the requirements of Theorem 1, the following result from [17] is used.

Let $\mathbf{N} = \mathbf{N}(\mathbb{R})$ be the space of integer-valued σ -finite measures κ on \mathbb{R} and let \mathcal{N} be the smallest σ -algebra making the mappings $\kappa \rightarrow \kappa(B), \kappa \in \mathbf{N}$, measurable for all $B \in \mathcal{B}(\mathbb{R})$, where $\mathcal{B}(\mathbb{R}^d)$ is the Borel σ -algebra of $\mathbb{R}^d, d \in \mathbb{N}$. For a measurable $\psi : \mathbf{N} \rightarrow \mathbb{R}$ define

$$D_y \psi(\kappa) = \psi(\kappa + \delta_y) - \psi(\kappa), \quad \kappa \in \mathbf{N},$$

where δ_y is the Dirac measure at $y \in \mathbb{R}$. For $k \in \mathbb{N}$ and $(y_1, \dots, y_k) \in \mathbb{R}^k$ set

$$D_{y_1, \dots, y_k}^k \psi = D_{y_1}^1 (D_{y_2, \dots, y_k}^{k-1} \psi),$$

where $D^1 = D$ and $D^0 \psi = \psi$. Define

$$R_k \psi(y_1, \dots, y_k) = \mathbb{E} D_{y_1, \dots, y_k}^k \psi(\phi),$$

if this expectation exists, where ϕ appears in (13). In [17, Theorem 1.1] it is shown that for any $\psi, \eta \in L_2(\mathbf{N}, \mathbb{P}_\phi)$, where \mathbb{P}_ϕ is the distribution of ϕ in $(\mathbf{N}, \mathcal{N})$, the following relation holds

$$\mathbb{E} \psi(\phi) \eta(\phi) = \mathbb{E} \psi(\phi) \mathbb{E} \eta(\phi) + \sum_{k=1}^\infty \frac{1}{k!} \langle R_k \psi, R_k \eta \rangle_{L_2(\mathbb{R}^k, \lambda^k v_k)}.$$

Here, v_k is the Lebesgue measure on $(\mathbb{R}^k, \mathcal{B}(\mathbb{R}^k)), k \in \mathbb{N}$.

Define $\psi_n(\phi) = \int_{[0, n)} \mathbb{I}(X_t \geq u_n) dt$. In order to establish (A2), it is sufficient to show that

$$\langle R_1 \psi_n, R_1 \psi_n \rangle_{L_2(\mathbb{R}, \lambda v_1)} = \lambda \int_{\mathbb{R}} \left(\int_0^n \mathbb{P}(u_n - e^{-|t-x|} \leq X_t < u_n) dt \right)^2 dx \rightarrow \infty, \quad n \rightarrow \infty. \tag{16}$$

Applying Vervaat's result, it holds that, for sufficiently large n ,

$$\int_{\mathbb{R}} \left(\int_0^n \mathbb{P}(u_n - e^{-|t-x|} \leq X_t < u_n) dt \right)^2 dx \geq f^2(u_n) \int_{\mathbb{R}} \left(\int_0^n e^{-|t-x|} dt \right)^2 dx.$$

Therefore,

$$\langle R_1 \psi_n, R_1 \psi_n \rangle_{L_2(\mathbb{R}, \lambda_{v_1})} \geq \lambda f^2(u_n) (n - 2). \quad (17)$$

Since $u_n = \mathcal{O}((\log n)^{1-\beta})$, $n \rightarrow \infty$, $\beta \in (0, 1)$, (15) and (17) imply (16).

It remains to show that $\delta_n \rightarrow 0$, $n \rightarrow \infty$. For n large enough, the inequality (17) yields

$$\delta_n = n \gamma^{\frac{2}{3}}(u_n) \sigma_n^{-(1+\mu/3)} \leq A \frac{n (f(u_n))^{\frac{2}{3}}}{(f^2(u_n) n)^{\frac{1}{6}(\mu+3)}} = A n^{-\frac{1}{6}(\mu-3)} (f(u_n))^{-\frac{1}{3}(\mu+1)}, \quad (18)$$

where $A = A(\lambda) > 0$. It is easy to see that the right-hand side of (18) converges to zero as $n \rightarrow \infty$.

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