## A COEFFICIENT OF LINEAR CORRELATION BASED ON THE METHOD OF LEAST SQUARES AND THE LINE OF BEST FIT.

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Given N points in a plane, corresponding to N pairs of values for two variables, X and Y, we find the line of best fit and the line of worst fit, by the method of least squares\*. Then we derive a cofficient of correlation based on the sum of the squares of the distances of the points from these two lines.

The line of best fit is in the line such that the sum of the squares of the distances of the points from it is a minimum. The line of worst fit is the one from which the sum of the squares of the distances of the points is a maximum. We shall refer to them, respectively, as the minimum and maximum lines.

For convenience we take the origin at the centre of gravity of the points, letting  $\varkappa$  and y denote deviations of X and Y, respectively, from their arithmetic means.

1. The two lines pass thru the arithmetic means of the X's, and of the Y's.

y = mx + b may represent any line of the plane. The distance,  $d_i$ , of a point,  $(x_i, y_i)$ , from the line is

$$\frac{y_i - mz_i - b}{\sqrt{1 + m^2}}$$
. The sum of the squares of the distances of

the N points from the line will be

<sup>\*</sup>For a general discussion of this method of fitting when *Q* variables are involved, see, Pearson, Karl, "On Lines and Planes of Closest Fit to Systems of Points in Space", Phil. Mag., 6th series, vol. ii, 1901, P. 559.

$$(1) \qquad \sum \underline{d}^2 = \frac{\sum y^2 + m^2 \sum x^2 + Nb^2 - 2m\sum xy - 2b\sum y + 2mb\sum x}{1 + m^2}$$

Using  $\Sigma x = \Sigma y = 0$ , in the condition for maximum or minimum values in (1), the condition reduces to b = 0, and the theorem follows.

2. To find the slopes of the two lines.

Equation (1) now becomes

The condition under which (2) will have a maximum or minimum values, is that

$$m^2 \Sigma xy + m (\Sigma x^2 - \Sigma y^2) - \Sigma xy = 0.$$

This condition is satisfied by two values of m, namely;

(3) 
$$m_{j} = \frac{\sum y^{2} - \sum x^{2} + \sqrt{(\sum x^{2} - \sum y^{2})^{2} + 4(\sum xy)^{2}}}{2\sum xy}.$$

(4) 
$$m_{2^{2}} \frac{\Sigma y^{2} \cdot \Sigma x^{2} - \sqrt{(\Sigma x^{2} \cdot \Sigma y^{2})^{2} + 4(\Sigma xy)^{2}}}{2\Sigma xy}.$$

It is found by considering the second derivative that (4) is the condition under which  $\Sigma \underline{\mathcal{L}}^2$  will have a maximum value, and (3) is the condition for a minimum value.

The equation of the minimum line is  $y = \gamma \gamma_1, x$ , and that of the maximum line is  $y = \gamma \gamma_2 x$ . The value of  $\gamma \gamma_1$ , and  $\gamma \gamma_2$  are those given in (3) and (4).

3. The maximum and minimum lines are perpendicular to each other.

That  $m_1 = -1/m_2$  is easily shown from (3) and (4).

Further  $m_1$  has the same sign as  $\Sigma \times y$ , and  $m_2$ , the opposite sign, since their numerators are, respectively, positive and negative.

4. The minimum line lies between the two lines of regression, or coincides with them.

If 
$$\Sigma xy > 0$$
,

$$m_1 \leq \frac{\sum y^2 - \sum x^2 + \sqrt{(\sum x^2 - \sum y^2)^2 + 4\sum x^2 \sum y^2}}{2\sum xy}$$

since  $\sum x^2 \sum y^2 \ge (\sum xy)^2$ .

Hence  $m_1 \le \sum y^2 / \sum xy = m_{xy}$ , the slope of the line of regression of x on y.

Rationalizing the numerator of (3), and noting that  $\mathbb{Z} z^2 \mathbb{Z} y^2 \stackrel{>}{=} (\mathbb{Z} zy)^2$ , we obtain

$$m_1 \ge \frac{2\Sigma xy}{\sum x^2 - \sum y^2 + \sqrt{(\sum x^2 - \sum y^2)^2 + 4\Sigma x^2 \sum y^2}} = \frac{\sum xy}{\sum x^2} = m_{yx},$$

the slope of the line of regression of y on x.

In the same way it may be shown that if  $\sum xy < 0$ , then  $m_{xy} \le m_1 \le m_{yx}$ .

The condition that  $m_I$  be equal to the slope of one line of regression is the same that it be equal to the other, so that the minimum line coincides with both lines of regression, or else lies between the two.

5. To find the sum of the squares of the distances of the points from the *minimum* line; also from the *maximum* line.

Let  $\alpha$  be the distance of a point from the line  $y=m_1x$ 

$$\sum d^2 = \sum \left(\frac{y - m_i x}{\sqrt{1 + m^2}}\right)^2 = \frac{\sum y^2 - 2 m_i \sum xy + m_i^2 \sum x^2}{1 + m_i^2}$$

Substituting for m from (3) and reducing, we obtain

Similarly, if  $\mathcal{D}$  represents the distance of a point from the line  $y = m_2 x$ , we obtain

(6) 
$$\sum D^2 = \left[ \sum_{x}^2 + \sum_{y}^2 + \sqrt{(\sum_{x}^2 - \sum_{y}^2)^2 + 4(\sum_{xy})^2} \right] / 2$$
.

6. To find a coefficient of linear correlation.

Let  $q = \sqrt{\Sigma d^2/\Sigma D^2}$  Substituting from (5) and (6), and reducing, we obtain

(7a) 
$$Q = \frac{\sum x^2 + \sum y^2 - \sqrt{\sum x^2 - \sum y^2} + 4(\sum xy)^2}{2\sqrt{\sum x^2 \sum y^2 - (\sum xy)^2}}, \text{ or }$$

(7b) 
$$Q = \frac{2\sqrt{\sum x^2 \sum y^2 - (\sum xy)^2}}{\sum x^2 + \sum y^2 + \sqrt{(\sum x^2 - \sum y^2)^2 + 4(\sum xy)^2}}$$

q represents the ratio of the root-mean-squares of the distances of the point from the *minimum* and *maximum* lines. This ratio

is a measure of the closeness of fit of the points to a line, and should furnish a measure of linear correlation. The value of q may vary from  $\mathbf{0}$  to  $\mathbf{1}$ . q = 0 indicates that the points are all on a straight line, hence that the correlation is perfect. It is of interest to note that when q = 0, (7b) gives  $\sum xy/No_x o_y$   $\mathbf{0} = \mathbf{1} = \pm 1$ . When q is 1 the mean squares of the distances of the points from two lines at right angles is the same, and linear correlation is lacking. Hence 1-q would furnish a coefficient conforming to the customs that it have the value 1 for perfect correlation, and  $\mathbf{0}$  for lack of correlation.

Values of Q found from (7a) or (7b) would involve the units in which X and Y are given. Hence these forms would be objectionable, in that Q could be made to assume different values for the same data, by changing the units in which x and y are expressed. However, this objection may be removed by taking  $\sigma_x$  and  $\sigma_y$  as units in which to express X and Y (7b) then reduces to

$$q = \frac{\sqrt{N^2 - (\sum xy)^2/\sigma_x^2 \sigma_y^2}}{N + |\sum xy/\sigma_x \sigma_y|}$$

The coefficient 1-q may now be expressed as

(8) 
$$r_c = 1 - q = 1 - \frac{N - |\Sigma_{xy}/\sigma_x \sigma_y|}{N + |\Sigma_{xy}/\sigma_x \sigma_y|} .$$

The sign of the coefficient should agree with that of the slope of the line to which the points are fitted. Hence, when the value for l-q has been found it should be given the same sign as the slope of the *minimum* line. But the slope of the *minimum* line is determined by that of  $\Sigma \times y$ , so that the sign given to l-q should be that of  $\Sigma \times y$ .

The coefficient 1-q may be expressed immediately in terms of the Pearson coefficient, r, which is equal to  $\sum xy/N\sigma_x\sigma_y$ . Making this substitution in (8) we have

$$r_c = 1 - q = 1 - \sqrt{\frac{1 - |r|}{1 + |r|}}$$
.

In the table are shown values of l-q corresponding to some given values of r. The values for  $l-q^2$  have also been listed corresponding to the same set of values for r. The maximum difference occurs between l-q and r when r=839 and l-q=704, a difference of 135 by which l-q is smaller.

r	$r_c=1-q$	1-q2
1 .99 .95 .9 .839 .8 .7 .6	1 .929 .840 .771 .704 .667 .580 .500 .423	1 .995 .974 .947 .912 .889 .824 .750 .667
.4 .3 .2 .1	.345 .266 .183 .095	.371 .462 .333 .182

