# A Distributed Algorithm for Managing Multi-Target Identities in Wireless Ad-hoc Sensor Networks 

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#### Abstract

This paper presents a scalable distributed algorithm for computing and maintaining multi-target identity information. The algorithm builds on a novel representational framework, Identity-Mass Flow, to overcome the problem of exponential computational complexity in managing multi-target identity explicitly. The algorithm uses local information to efficiently update the global multi-target identity information represented as a doubly stochastic matrix, and can be efficiently mapped to nodes in a wireless ad hoc sensor network. The paper describes a distributed implementation of the algorithm in sensor networks. Simulation results have validated the Identity-Mass Flow framework and demonstrated the feasibility of the algorithm.


## 1 Introduction

A wireless ad-hoc sensor network (WASN) is a network of sensor nodes with limited on-node sensing, processing, and communication capabilities. At the heart of many WASN applications such as object tracking and environmental monitoring is the problem of estimating non-local parameters or states of the physical phenomenon being observed using only local information available to each node. This problem poses unique challenges that are not addressed in a centralized setting or fixed network:

- Scalable distributed information fusion: The global parameters or states of interest must be estimated, updated, and maintained using only local information.
- Sensor tasking: A sensor node may be tasked to compute or store state information based on its relevance and utility to the current task, as well as cost constraints.

In this paper, we study the problem of distributed multi-target identity management in WASN. The goal is to maintain information about who is who over time given targets' position estimates. In addition to its central importance in many of the monitoring and surveillance applications, the problem highlights the need for distributed information fusion and sensor tasking. The study of this problem is an important step towards establishing a general methodology for the design of scalable WASN algorithms.

Multi-target identity management is closely related to the multi-target tracking problem. The main difficulty in both problems is the exponential complexity in associating target position estimates with target identities. In the past three decades, a number of approaches have been developed for the multi-target tracking problem, mostly for centralized computational platforms. MHT ([2]) explicitly maintains associations, or hypotheses, over time and prunes the associations using a rank function. JPDA ([4]) computes an association matrix at each time and updates it with a combination of all new measurements weighted by their marginal probabilities. While widely used in practice, both MHT and JPDA algorithms still suffer from their computational complexity, in the worst case exponential in the number of targets or time steps. Moreover, for WASN applications, a significant challenge lies in distributing the information and computation to each node.

This paper develops an efficient distributed approach to computing and udpating multi-target identity information. The main contribution of this work is twofold. First, it introduces a new distributed representation called identity belief matrix, a doubly stochastic matrix, that describes how identity information of each target is distributedly represented (a row in the matrix). The key advantage of this representation is that when the matrix is mapped to a set of nodes in a WASN, a node could efficiently maintain possible identifies of a target it is tracking (a column in the matrix), using its local evidence only. Second, the paper develops a distributed algorithm for computing and updating the cross-node identity belief matrix in a WASN. The algorithm exploits doublystochasticity of the matrix to renormalize identity probability masses stored on different WASN nodes. As soon as a piece of local/marginal evidence is available to a node, the local belief of target identity is updated, and the information is propagated through the network to other nodes. The computational complexity of our algorithm is $O\left(N^{2}\right)$, where $N$ is the number of targets, a significant advantage over the exponential complexity of MHT and JPDA.

The rest of the paper is organized as follows. Section 2 introduces identitymass flow (IMF) framework to represent multi-target identity information. To ease the introduction of mathematical materials and focus on key representational issues, the identify representation is first developed in a centralized setting. Sections 3 and 4 develop algorithms to distribute the computation and map the algorithm to WASN nodes. Section 3 describes an algorithm for updating global identity information using local identity information. Section 4 maps the identity representation and algorithm to a set of WASN nodes. Finally, Section 5 describes an implementation of the algorithm, and presents simulation results
that validate the correctness of the representational framework and the basic structure of the algorithm.

## 2 Multiple Identity Management: A Mathematical Framework

Consider the following problem. There are $N$ targets moving in the region of interest and their position-estimates are reported periodically. Assuming the initial identities are known, the goal of the multiple identity management is to maintain and update the identity information of all the targets as they are moving.

Apparently, the dynamics of the target - how they are moving - seems to be the only information available to do the job. Unlike the radar system, however, the subset of the nodes in the sensor network are very close to the targets and are able to sense more than position-estimate - the target signature information. This signature information is clear and dependable in the sparse target configuration, but is not informative when the targets are close to one another. Figure 1 illustrates the two target configurations, which are named as configuration of high uncertainty $(\mathrm{COHU})$ and configuration of low uncertainty (COLU), respectively. The actions/decisions based on the information in COHU could be false and it would be nice to have a representational framework that could fix the poor actions/decisions in COHU using the better information in COLU. This is the main motivation behind our mathematical formulation in the upcoming subsections.


Fig. 1. Sparse(COLU) and crowded(COHU) target configurations

### 2.1 Formulation

In this section, we formulate the problem of multiple identity management as identity-mass distribution problem. The identity set $I=\{1, \cdots, N\}$ is a set of
identities of all the targets and the position-estimates of $N$ targets at time $k$ is $X(k)=\left\{x_{i}(k) \in \mathbf{R}^{2} \mid i=1, \cdots, N\right\} .{ }^{3}$ The identity management algorithm is supposed to compute the correct permutation of $I$ given $X(k)$. The natural approach to this is to maintain all the possible permutations ${ }^{4}$ at each time, although the number of possible permutations increases exponentially. Even with good rank functions and pruning heuristics, the number of possible permutations can easily go unmanageably large in a very short period of time.

To overcome the above combinatorial complexity and maintain the computational complexity as constant over time, we propose the idea of Identity-Mass Flow(IMF) to approximately represent all the possible permutations. Figure 2 (a) shows the basic idea behind our approach; Initially, a unit mass is assigned to each identity. Whenever the new position-estimate $X(k)$ is available, the whole or partial masses from $X(k-1)$ flow into $X(k)$ and the identities are mixed in this way. There, however, need to be constraints regarding how masses flow to make the resulting mixed identities physically meaningful. Figure 2 (b) and (c) explain the two constraints. (b) says no mass can be created or destroyed during the Identity-Mass Flow and (c) says the sum of all the masses arriving at $x_{i}(k)$ is one.


Fig. 2. Identity Mass Flow

To formulate the above idea, we define the identity belief matrix $B(k)$ and the mixing matrix $M(k)$.

[^0]Definition 1. The identity belief matrix $B(k)$ is a $N \times N$ doubly-stochastic matrix ${ }^{5}$, whose entry $b_{i j}(k)$ represents the amount of identity mass from $i \in I$ that arrives at $x_{j}(k)$. The $j_{t h}$ column $\mathbf{b}_{\mathbf{j}}(k)$ of $B(k)$ is called the identity belief vector of $x_{j}(k)$.

$$
B(k)=\left[\mathbf{b}_{\mathbf{1}}(k) \mathbf{b}_{\mathbf{2}}(k) \cdots \mathbf{b}_{\mathbf{N}}(k)\right] \in[0,1]^{N \times N}
$$

where

$$
\mathbf{b}_{\mathbf{i}}(k)=\left[\begin{array}{c}
p\left(x_{i}(k) \text { 's } I D\right. \text { is } \\
p\left(x_{i}(k) \text { 's } I D\right. \text { is } \\
\vdots \\
p\left(x_{i}(k) \text { 's } I D\right. \text { is } \\
\hline
\end{array}\right] \in[0,1]^{N \times 1}
$$

Definition 2. The mixing matrix $M(k)$ is a $N \times N$ doubly-stochastic matrices, whose entry $m_{i j}(k)$ represents the probability of $x_{j}(k)$ being originated from $x_{i}(k-1)$, and is statistically independent with $M(l)$ for all $l \neq k$.

Given the definition of $M(k)$ and $B(k)$, the following theorem presents the basic equation relating the two quantities.

Theorem 1. Let $B(k)$ and $M(k)$ be the identity belief matrix and the mixing matrix at time $k$ as defined above, then the following is true.

$$
B(k+1)=B(k) M(k+1)
$$

Proof. From the above definitions, the identity belief of $x_{j}(k+1)$ is computed as follows

$$
\mathbf{b}_{\mathbf{j}}(k+1)=\sum_{l=1}^{N} m_{l j}(k+1) \mathbf{b}_{\mathbf{l}}(k)=B(k) \mathbf{m}_{\mathbf{j}}(k+1)
$$

where $\mathbf{m}_{\mathbf{j}}(k+1)$ is the $j_{t h}$ column of $M(k+1)$. Therefore,

$$
B(k+1)=B(k) M(k+1)
$$

and this concludes the proof.
The above theorem shows that we can recursively compute the identity information $B(k)$ by computing $M(k)$ from $X(k)$ and $X(k-1)$ and is illustrated in Figure 3. The details on how to compute $M(k)$ is investigated in the next section.

The following lemma explains how the uncertainty in the system changes over time in this formulation.

## Lemma 1.

$$
H(B(k)) \geq H(B(k-1))^{6}
$$

where $H(\cdot)$ is the statistical entropy of a probability distribution.

[^1]

Fig. 3. Updating $B(k)$ using $M(k)$

Proof. The mixing matrix $M(k)$ can be represented as a convex sum of permutation matrices as follows

$$
M(k)=\sum_{i=1}^{N!} \alpha_{i} \Phi_{i}
$$

where $\sum_{i} \alpha_{i}=1$ and $\Phi_{i}$ is the $i_{t h} N \times N$ permutation matrix. Then,

$$
\begin{aligned}
H(B(k)) & =H(B(k-1) M(k)) \\
& =H\left(\sum_{i=1}^{N!} \alpha_{i} B(k-1) \Phi_{i}\right) \\
& \geq \sum_{i=1}^{N!} \alpha_{i} H\left(B(k-1) \Phi_{i}\right) \\
& =\sum_{i=1}^{N!} \alpha_{i} H(B(k-1)) \\
& =H(B(k-1))
\end{aligned}
$$

where the inequality comes from the strict concavity of the entropy function. This concludes the proof.

The statistical entropy is a quantitative measure of uncertainty in the probability distribution and the above lemma shows the uncertainty on the possible identity associations does not decrease over time. Therefore, the uncertainty will grow until every identity association becomes equally like without any additional information available. As we have mentioned in section 2, the target identity information is more likely to be available in COLU and the proper use of this local information could make the uncertainty decrease in IMF formulation. The details on how to exploit the local identity information is investigated in section 3

### 2.2 Computing Mixing Matrix $M(k)$

The mixing matrix $M(k)$ is a collection of marginal association probabilities and can be computed from the joint association probability ${ }^{7}$ in theory. Computing the joint association probabilities is very expensive and should be avoided in practice. In this paper, we propose a simple heuristic with $O\left(N^{2}\right)$ empirical complexity that only requires the information on how fast the targets are moving. Let's assume that the speed information is given as a probability density function $p(s(k))$, where $s(k)$ is $|x(k)-x(k-1)| / \Delta T .^{8}$ Then, we compute a non-negative matrix $L(k) \in \mathbf{R}^{N \times N}$, whose $(i, j)$ entry is $l_{i j}(k)=p\left(s(k)=\mid x_{j}(k)-x_{i}(k-\right.$ $1) \mid / \Delta T)$. In general, $L(k)$ is not doubly-stochastic and an optimization need to be done to transform this into a doubly-stochastic matrix. We use the Iterative Scaling algorithm to transform $L(k)$ into a doubly-stochastic matrix. The details on the Iterative Scaling algorithm are explained in section 3.2.

## 3 Multi-target Identity Update Using Local Information

In WASN, the ability to use local information efficiently is critical for distributed algorithms since non-local information only comes at the cost of communication. The IMF approach in multi-target identity management does provide a natural setting for exploiting local evidences. Figure 4 illustrate this in a simple two targets crossover scenario. Two targets are moving cross each other and their identity masses are mixed at the intersection of the two tracks. After the mixing, the identity belief matrix $B(k)$ becomes un-informative - each association is almost equally likely. When the two targets are well-separated, i.e., in COLU, one of the nodes near the bottom target observes a local evidence ${ }^{9}$ that the bottom target is more likely to be yellow ${ }^{10}$. This observation increases the yellow-mass of the bottom target $b_{\text {yellow, } \operatorname{bottom}}(k)$ and the rest of the elements in $B(k)$ can be updated from the doubly-stochasticity of $B(k)$. Therefore, the local information about the bottom target directly leads to the information about the top target in a unique way.

From the above example, only a local evidence seems to be enough to update the whole identity belief $B(k)$ uniquely. For the general $N$ target case, however, the doubly-stochasticity of $B(k)$ is not enough to guarantee a unique solution since there are more unknowns ( $N^{2}$ ) than the number of constraints $(N)$. Therefore, we need more constraints or optimizations to compute a unique $B(k)$ given a local evidence. The upcoming sub-sections deal with this problem of computing $B(k)$ given a local evidence in a centralized setting. Section 4 discusses the distributed implementation of the algorithm for WASN.

[^2]

Fig. 4. Example of using local evidence: The top left ID is colored as blue and the bottom left ID is yellow. The blue and yellow look dark and lighter grey, respectively, in a black and white printout.

### 3.1 Bayesian Normalization

Before we introduce a practical solution to compute unique $B(k)$ given a local evidence, we study the desirable properties of the perfect solution by computing $B(k)$ as a Bayesian posterior belief distribution given a local evidence, assuming the whole history of the mixing events are known. In this case, we assume that the joint probability distribution $\pi(k) \in\left\{Z \in \mathbf{R}^{N!\times 1} \mid \sum_{i} z_{i}=1, z_{i} \geq 0\right\}$ over all the possible $N$ ! associations at time $k$, from which $M(k)$ can be derived, is available for all $k$. Note that, for some $k$ 's, $\pi(k)$ 's are deterministic and their associated $M(k)$ are permutation matrices. These $\pi(k)$ do not contribute in computing the $B(k)$ and its posterior. To consider only those random mixing events, we define $K \subset\{0,1, \cdots, N\}$ be the set of time indices associated with $\left|\pi\left(k_{i}\right)\right|>2$, where $k_{i}$ is the $i_{t h}$ element in $K$ and $|\cdot|$ represents the cardinality of a set. We also introduce a sequence of random variables $R_{i}$ associated with $\pi\left(k_{i}\right)$, which takes values $j \in\left\{1, \cdots,\left|\pi\left(k_{i}\right)\right|\right\}$ with probability of $\pi_{j}\left(k_{i}\right)$, i.e., $j_{t h}$ value in $\pi\left(k_{i}\right)$.

Figure 5 illustrates the above formulation, where each box can be considered as a probabilistic switch governed by $R_{i}$, i.e., a specific permutation is chosen according to the value of $R_{i}$ with some probability. Then, a single identity association at time $k$ is a point $\varepsilon$ in the joint event space $S=\left\{\left(R_{1}, R_{2}, \cdots, R_{|K|}\right) \mid R_{i} \in\right.$ [1 $\left.\left.\left|\pi\left(u_{i}\right)\right|\right]\right\}$ with $|S|=\prod_{i}\left|\pi\left(u_{i}\right)\right|$ and the probability of this event can be easily computed due to the statistical independence assumption in section 2.1.

$$
\begin{aligned}
\left.p\left(\left(R_{1}, \cdots, R_{|K|}\right)=\varepsilon\right)\right) & \left.=p\left(R_{1}=\varepsilon_{1}\right) \cdots p\left(R_{|K|}=\varepsilon_{|K|}\right)\right) \\
& =\pi_{\varepsilon_{1}}\left(u_{1}\right) \cdots \pi_{\varepsilon_{|K|}}\left(u_{|K|}\right)
\end{aligned}
$$

Using the above equation, we can compute the posterior $B(k)$ given a local evidence $L$, which is a set of events in $S$ satisfying the local observation, say, $I D\left(x_{i}(k)\right)=j$, using the following theorem.


Fig. 5. Bayesian Normalization

Theorem 2. Let $E_{i j}(k)$ be the subset of $S$ satisfying $I D\left(x_{j}(k)\right)=i$ and $L$ be the subset of $S$ satisfying the local observation, then

$$
p\left(b_{i j}(k) \mid L\right)=p\left(E_{i j} \mid L\right)=\frac{\sum_{\varepsilon_{l} \in E_{i j} \cup L} p\left(\varepsilon_{l}\right)}{\sum_{\varepsilon_{l} \in L} p\left(\varepsilon_{l}\right)}
$$

Proof. The proof is trivial using the Bayes' rule.
Now that we know how to update $B(k)$ given a local evidence $L$, but the effect of the evidence to the other columns is not obvious from the equation. The following lemma describes how the local evidence effect the uncertainty of the other columns.

Lemma 2. The local observation $L$ does not increase the entropies of the columns, i.e.

$$
H\left(\mathbf{b}_{\mathbf{i}}(k) \mid L\right) \leq H\left(\mathbf{b}_{\mathbf{i}}(k)\right)
$$

Proof. See [6] for the proof.
The lemma says that the local evidence does not increase the entropy of the other columns on the average.

The ideal solution obtained in the above theorem exhibits one very interesting characteristic, in which there are some elements in $B(k)$ - in addition to the zero elements in $B(k)$ - that are not affected by the local evidence. This can potentially save huge amount of communication energy in practice. The following theorem formally presents the property of the Bayesian solution.

Theorem 3. Let $b_{p q}(k)$ be the entry that becomes 1 from the local evidence $L$, then the columns with zero at $p_{\text {th }}$ entry and the rows with zero at $q_{\text {th }}$ entry do not change.

Proof. Let's first prove that the columns with zero at $p_{t h}$ entry do not change. If $\mathbf{b}_{\mathbf{i}}(k)$ is such a column, then there does not exist any event in $S$ that guarantees a path from $x_{p}(0)$ to $x_{i}(k)$, i.e., no path originated from $x_{p}(0)$ reaches $x_{i}(k)$. The local evidence $L$ defines a subset of $S$, which guarantees the existence of at least one path between $x_{p}(0)$ and $x_{q}(k)$. None of these paths between $x_{p}(0)$ and $x_{q}(k)$, however, affect the paths arriving at $x_{i}(k)$. Due to the statistical independence assumption on the mixing events, $\mathbf{b}_{\mathbf{i}}(k)$ does not change given the local evidence $L$. The row case can be proved in the same way. This concludes the proof.

The above theorem reduces the number of variables to be updated in $B(k)$ given a local evidence. In addition to that, this can help the number of communications required to update $B(k)$ given a local evidence assuming that each column $\mathbf{b}_{\mathbf{i}}(k)$ is maintained by a node in the sensor network. The details on the distributed computation is discussed in section 4.

In addition to these rows and columns, the zero elements in $B(k)$ do not change given a local evidence $L$.

### 3.2 Iterative Scaling

In practice, the Bayesian formulation in the previous section is infeasible due to its exponential complexity. This section presents the practical alternative called the Iterative scaling. First, we present a version of the Iterative Scaling algorithm to achieve a doubly-stochastic matrix $A$ given a $N \times N$ non-negative matrix $B$.

```
B := A;
B_old := A;
for k = 1 to maximum_number_of_iteration
    for i = 1 to number_of_row
        row_sum := 0;
        for k = 1 to number_of_column
            row_sum := row_sum + B(i,k);
        end
        for j = 1 to number_of_column
            B(i,j) := B(i,j)/row_sum;
        end
    end
    for i = 1 to number_of_column
        column_sum := 0;
        for k = 1 to number_of_row
            column_sum := column_sum + B(i,k);
        end
        for j = 1 to number_of_row
            B(j,i) := B(j,i)/column_sum;
        end
    end
    if |B - B_old| < error
```

```
        terminate;
    end
    B_old := B;
end
```

Basically, the algorithm divides each element in $i_{t h}$ row(column) by the sum of the $i_{t h}$ row(column) and repeats the normalization until the error margin is small. The following observations are made from the numerical simulations and we list them here without proofs.

- The algorithm converges to a unique doubly-stochastic matrix given an initial matrix.
- The ordering of row/column normalization does not affect the convergence.
- The total number of iteration is not affected by the size of the matrix.


Fig. 6. Typical convergence of iterative scaling: The flat part of the plots are due to the Matlab precision limit.

From these observations, the Iterative Scaling algorithm scales as $O\left(N^{2}\right)$. The proof of the complexity result remains as an immediate task for future research. Figure 6 shows an example of the convergence behavior of the algorithm. Three different sizes of matrices $(10 \times 10,100 \times 100,1000 \times 1000)$ are generated randomly using Matlab $\operatorname{rand}(\cdot)$ function, in which each entry is generated according to the uniform probability density over [ $\left.\begin{array}{lll}0 & 1\end{array}\right]$. The plot shows that the Iterative Scaling method has fast convergence and the size of a matrix does not affect the convergence ratio. What seems to affect the convergence rate is how different
the (scaled) initial matrix from being the doubly-stochastic matrix, although we do not have a proper quantitative measure for this at this point. This is why the larger matrices in the above figure converges a little faster than the smaller ones, since all the row/column sums of the larger matrices are close to 0.5 N due to the Law of Large Numbers and effectively close to a doubly-stochastic matrix after scaling.

To benefit from the theorem 3, the Iterative scaling algorithm should be able to deal with the non-square matrix with the fixed row/column sums. Let $\mathbf{r}$ and $\mathbf{c}$ be $N \times 1$ vector of row and column sum satisfying $\sum_{i} c_{i}=\sum_{j} r_{j}$, then we can modify the above pseudo code as follows.

```
B(i,j) = B(i,j)/row_sum*r(i);
B(j,i) = B(j,i)/column_sum*c(i);
```


## 4 Distributed Implementation in WASN

The basic quantity to be distributed is $B(k)$, the belief matrix. One way of distributing $B(k)$ is to let each node in the sensor network maintain its own version of $B(k)$. This method is very robust and fault-tolerant due to the information redundancy in the system. However, this idealistic distribution is infeasible and non-scalable for the following the reason. To update the information, each node need to compute its version of $M(k)$, which requires information from at least one of the other nodes. This is exactly the scenario in the landmark paper [14], where per-node throughput goes to zero as the number of nodes goes to infinity even under optimal routing and power control scheme. Therefore, the idealistic distribution of each node maintaining $B(k)$ is impossible and this argument is true for all the algorithms of sensor networks that estimates the global quantity.

To overcome this problem, we adopt and extend the approach in [16] and [17], where a leader-based single target tracking in WASN is introduced. The basic idea is that, only a small number of nodes called leaders are active and responsible for maintaining/updating the information of interest. When the leaders are no longer the good subset of nodes for the information, they handoff the information to the other nodes based on a utility function and the nodes receiving the information become the new leaders. In this approach, whereabout of the information is easily maintained at the risk of the reduced robustness and fault-tolerance.

Applying the leader-based approach to our algorithm, each column $\mathbf{b}_{\mathbf{i}}(k)$ of $B(k)$ and its position estimate $x_{i}(k)$ is maintained by each leader. When the mixing happens, the local version of $M(k)^{11}$ can be easily computed by the leader. When a leader node observes a local evidence about the ID of its target, say $I D\left(x_{i}(k)\right)=j$, then the leader needs to talk to some of the other leaders, who also think what they are tracking can be of the same ID. This type of multi-cast

[^3]

Fig. 7. Flowchart of the distributed algorithm for the multi-target identity management
communication in network is usually dealt by a group management protocol, which maintains and updates the different groups in the network according to a predefined membership. In our case, the $i_{t h}$ group is the set of the leader that have non-zero probability at $i_{t h}$ entry of their $\mathbf{b}(k)$ and we assume there exists a good (or optimal) group management protocol for our purpose. Figure 7 shows the main procedures and their relations of the distributed algorithm that each node is running.

## 5 Experimental Results

We make the following assumptions for the simulation.

- Each node can sense the positions of the targets within its sensing range.
- Each node can talk to all the nodes within its communication range.
- Initially, the number of the targets are known to the initial leaders.
- Each node has a signal processing module for the signal classification and the module performs better in COLU.
- Each node knows the relative positions of all the nodes in the communication range.

The initial leaders are manually selected for the simulation, although it's possible to detect them as long as their initial positions are well separated. Each
leader updates its belief vector $\mathbf{b}_{\mathbf{i}}(\mathbf{k})$ using the local version of $M(k)$ and handoffs the information to the next leader, which is selected based on its geographical location. The signal classification module of the leaders will keep collecting the information and initiates the identity normalization process by talking to the other leaders that are in the same group when the identity information from the signal classification is better than $\mathbf{b}_{\mathbf{i}}(k)$ and some threshold.


Fig. 8. Simulation Example

Figure 8 shows the three screen shots from the simulation of the algorithm. Four targets - one tank and three dots - are moving along the straight lines for 10 seconds. The tank signature is much different from those of the other three, so it can be identified with high probability in COLU. This local identity information about the tank is the only available information and used to normalize the belief $\mathbf{b}_{\mathbf{i}}(k)$ of the other leaders. The four leaders are colored differently and their corresponding beliefs are displayed with the same color. Figure 8 (a) is the initial configuration of the targets and their associated leaders. (b) shows that the belief $\mathbf{b}_{\mathbf{i}}(k)$ of each leader gets uncertain after some number of mixing events at $t=7$ and the leaders are no longer sure of the identities of the targets. (c), however, shows the beliefs get much better after the normalization using Iterative Scaling algorithm given local identity information.

In figure 9 , how the identity uncertainty of each target evolves during the simulation is depicted using the entropy of each identity belief $\mathbf{b}_{\mathbf{i}}$. The increases in the uncertainty are due to the mixing events and the decreases are by the local evidences. The two pieces of the local evidence on target 1 have reduced the uncertainties of all the other targets in this example, since the identity mass from the target 1 is mixed with all the other identities during the mixing events.


Fig. 9. An example of how uncertainty of each belief changes. The target numbers correspond to the track numbers in the previous figure.

## 6 Summary and Future Work

We have developed a scalable distributed algorithm for the multi-target identity management problem, first presented as a mathematical framework for a centralized setting, and then mapped to a distributed WASN. Simulation results have demonstrated the effectiveness of the framework and the efficiency of the algorithm.

Since the target identity computation is at the heart of many WASN tracking and classification applications, the work presented here is an important step towards building a comprehensive system for distributed inference in sensor networks. As our future work, we plan to relax some of the assumptions and models used in our framework. For example, we may exploit a signal processing model for target source separation, localization, and signature classification to obtain additional target identity information when targets are in close proximity of each other, and incorporate the ability to handle dynamic addition and deletion of targets. A more theoretical task is the convergence proof for the normalization step given multiple local evidences regardless of their chronological order.

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[^0]:    ${ }^{3} x_{i}(k)$ and $x_{i}(k+1)$ are not related. $i$ 's are just random labels that come with the position estimate.
    ${ }^{4}$ In multi-target tracking community, this is also called an association.

[^1]:    ${ }^{5}$ The doubly-stochastic matrix is a $N \times N$ non-negative matrix, whose rows and columns sum to one.
    ${ }^{6}$ To be more precise, $H(B(k))$ here means the entropy of any joint identity association that can be derived from $B(k)$.

[^2]:    ${ }^{7}$ Probabilities of permutations/associations, i.e., how likely each permutation is.
    ${ }^{8} p(s(k))$ is usually stationary, i.e., does not depend on $k$.
    ${ }^{9}$ A local evidence is the information enough to determine the whole or partial entries in $b_{i}(k)$. The simple example is " $x_{i}(k)$ is of ID $j$ ".
    ${ }^{10}$ Lighter color in a black and white printout

[^3]:    ${ }^{11}$ In the two target mixing, only non-zero entries in the $i_{t h}$ and $j_{t h}$ columns of $M(k)$ are required to update $b_{i}$ and $b_{j}$ and they are locally computable.

