

1 A DSM-5 AMPD and ICD-11 compatible measure for an early identification of
2 personality disorders in adolescence – LoPF-Q 12-18 latent structure and short
3 form

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Abstract

The LoPF-Q 12-18 (Levels of Personality Functioning Questionnaire) was designed for clinical use and to promote early detection of personality disorder (PD). It is a self-report measure with 97 items to assess personality functioning in adolescents from 12 years up. It operationalizes the dimensional concept of personality disorder (PD) severity used in the Alternative DSM-5 Model for Personality Disorders and the ICD-11. In this study, we investigated the factorial structure of the LoPF-Q 12-18. Additionally, a short version was developed to meet the need of efficient screening for PD in clinical and research applications.

To investigate the factorial structure, several confirmatory factor analysis models were compared. A bifactor model with a strong general factor and four specific factors showed the best nominal fit (CFI = .91, RMSEA = .04, SRMR = .07).

The short version was derived using the ant colony optimization algorithm. This procedure resulted in a 20-item version with excellent fit for a hierarchical model with four first order factors to represent the domains and a secondary higher order factor to represent personality functioning (CFI = .98, RMSEA = .05, SRMR = .04). Clinical validity (effect size $d = 3.1$ between PD patients and controls) and clinical utility (cutoff ≥ 36 providing 87.5% specificity and 80.2% sensitivity) for detecting patients with PD were high for the short version. Both, the long and short LoPF-Q 12-18 version are ready to be used for research and diagnostic purposes.

Keywords: Personality functioning, self-report, ant optimization algorithm, adolescents, youth, early detection, personality disorder, bifactor model

46 DSM-5 AMPD and ICD-11 compatible measure for an early identification of personality
47 disorders in adolescence – LoPF-Q 12-18 latent structure and short form

48 **Introduction**

49 **Recent changes in the conceptualisation of personality disorders (PD)**

50 The conceptualisation of personality disorder (PD) and, subsequently, the diagnostic system for
51 PDs is currently transitioning from a categorical system (e.g. narcissistic or avoidant PD) which is still the
52 official system in the DSM-5 [1] to a dimensional approach. The dimensional approach is used both in the
53 ICD-11 [2] as well as in the Alternative DSM-5 Model for Personality Disorders (AMPD). The AMPD is
54 described in the section “emerging measures and models” of the DSM-5. The reason for the fundamental
55 change of the guidelines for diagnosing PD are well documented shortcomings of the categorical system:
56 For instance, individual differences in the expression of PD characteristics are not dichotomous but
57 appear continuously distributed, thresholds (i.e. number of present symptoms required to assign a
58 diagnosis) for categorical PD diagnoses have been criticised as largely arbitrary, and the empirical
59 covariation of the individual criteria does not fully correspond with their assignment to the ten distinct PD
60 categories in the diagnostic manual [3]. As a consequence, the categorical approach is no longer regarded
61 as the only valid taxonomy and has been criticised as a hindrance to research and practice [4]. A growing
62 number of publications in the field now argue in favour of the dimensional approach.

63 The AMPD and the ICD-11 dimensional PD models are conceptually similar. They each contain
64 two assessment modules to characterise PDs: A first diagnostic module is the evaluation of the self- and
65 the interpersonal functioning of the patients to represent general features and the severity of the PD. This
66 is referred to as ‘criterion A’ in the AMPD. Criterion A is constructed from four domains: identity, self-
67 direction (those account for self-related functioning), empathy and intimacy (accounting for interpersonal
68 functioning). The second module is used to evaluate maladaptive personality traits to represent stylistic
69 differences in the expression of PD [5]. This assessment of ‘trait domain qualifiers’ in the ICD-11 is not

70 mandatory for a diagnosis but helps to refine the assessment. In contrast, in the AMPD, the usage of this
71 second module, ‘Criterion B’, is required.

72 As a further major shift in the PD diagnosis paradigm, experts now broadly agree on the
73 importance of early detection and treatment of PD [6]. Strong evidence has been delivered showing that
74 PD is a valid diagnosis in youth [7]. PDs can have their origin during childhood and can emerge in early
75 adolescence [8,9]. Early detection and treatment are important as adolescence is a critical and formative
76 period which lays the foundation in terms of psychosocial functioning for the wellbeing and productivity
77 of the adult [10]. Additionally, from a neurocognitive perspective, adolescence also represents a window
78 of opportunity to effectively and efficiently treat mental disorders [8]. By providing early interventions in
79 adolescents, clinicians are trying to avert the harmful psychosocial consequences of a developing disorder
80 and prevent chronification. This is important as PDs can have a heavily incapacitating impact on the
81 patients and their environment, including the somatic health and life expectancy of those affected [11].
82 Additionally, societal costs of untreated PD are high (e.g. direct healthcare costs and loss of productivity)
83 [12,13]. To account for a perspective of PD across lifetime, both the ICD-11 and the AMPD have
84 abolished the age limit for PD diagnoses. Multiple manualised psychotherapies are available for young
85 PD patients [14–17]. However, to allow for early treatment, age-adequate assessment procedures to
86 detect PDs in adolescents according to the dimensional PD concept are required.

87 **Need for psychometric instruments**

88 These two-fold changes (dimensional approach and earlier diagnosis) in the diagnostic
89 systems of PDs pose a challenge for mental health care services on a global level. The World
90 Health Organization emphasises the ICD-11 system for PDs needs to be useful and usable also
91 for health care workers in lower-resource settings who are not highly trained specialists [18]. To
92 overcome this challenge, evidence-based assessments are of critical importance (19).
93 Zimmermann et al. [5] provide a brief but comprehensive review on research regarding the

94 dimensional PD models summarizing currently available measures. Birkhölzer et al. [19] provide
95 an updated review for instruments to measure criterion A.

96 The ICD-11 model for PD is relatively new. Tools specifically targeting the ICD-11
97 operationalisation of PD diagnosing are currently being developed and validated. Based on
98 strong similarities of ICD-11 and the AMPD regarding personality functioning, Bach & First
99 [20] propose that assessment tools developed for the DSM-5 AMPD model can also be used to
100 support an ICD-11 dimensional PD diagnosis. To comply with the new ICD-11 lifetime
101 perspective on mental disorders in general, all psychometric instruments will in principle have to
102 be adapted for younger ages.

103 To date, the Levels of Personality Functioning Questionnaire for Adolescents (LoPF-Q
104 12-18) [21] is the only available self-report questionnaire to assess personality functioning
105 according to the AMPD that was developed specifically for adolescents from 12 years up. The
106 items of the LoPF-Q 12-18 have been carefully designed to take into account the developmental
107 stage and life situation of adolescents [22]. It has been optimised for use in clinical practice,
108 providing several descriptive subscales matching classical psychological concepts in addition to
109 the total score and the four domain scores. This is supposed to inform differentiated diagnoses
110 and therapy planning and to facilitate the upcoming fundamental changes in diagnostic
111 guidelines for PD. First developed in German language, it has been translated and culturally
112 adapted by expert teams for English [23], Spanish [24], Turkish [25,26] and Lithuanian [27].
113 Adaptations for Slovenian, Russian, French, Danish, Swahili (Tanzania), Hebrew, Chinese and
114 Romanian are currently under development, showing that this instrument is supported by an
115 international clinical and research community including low- and middle-income countries. The
116 LoPF-Q 12-18 shows excellent scale reliability and accurately detects patients with personality

117 disorders [21]. It can be requested for free for research purposes and is also available in
118 electronic format at the project website (academic-tests.com).

119 **Test construction and psychometric properties of the LoPF-Q 12-18**

120 The LoPF-Q 12-18 is a 97-item self-report measure for adolescents between 12 and 18
121 years (+/- 2 years) to assess the dimensions of personality functioning: Identity, Self-direction,
122 Empathy, and Intimacy. It is designed to enable a dimensional differentiation between healthy
123 and impaired personality functioning to promote early detection of PD (criterion A). The
124 construction was inspired by the AMPD [28] and the ICD-11 beta draft capturing the full scope
125 of self- and interpersonal functioning. To operationalize the LoPF-Q 12-18, all descriptors of the
126 four AMPD domains were carefully analyzed and enriched with available concepts from child
127 and adolescent psychology with focus on clinical validity. This led to a detailed structure for
128 operationalizing the domains of functioning (see S1 Table), building the basis for a deductive
129 item formulation. The derived item pool was then revised in an empirically informed iterative
130 process to make them appropriate for a self-rating instrument for adolescents. Accordingly, the
131 four resulting primary scales identity, self-direction, empathy, and intimacy are composed of two
132 subscales per scale. These subscales are reported in addition to the total score and scale scores to
133 support detailed clinical decision making. Because they represent less abstract and more
134 commonly shared concepts (like e.g. Purposefulness or Prosociality), they may be helpful to
135 better understand a patient's situation or to trace developments over time.

136 The process of test construction as well as psychometric properties have been described
137 in detail in [22]. The main psychometric targets were clinical validity, good applicability for
138 older and younger adolescents, and good scale reliabilities. The LoPF-Q 12-18 shows good scale
139 reliability (Cronbach's alpha of .96 for the total scale, .92, .94, .87, and .92 for the primary scales

140 and between .76 and .96 for the subscales), good construct validity and substantial clinical
141 validity. The LoPF-Q 12-18 total score distinguished between adolescents from the general
142 population and $n = 96$ SCID-II diagnosed PD patients at a highly significant level and with a
143 large effect size of 2.1 standard deviations [21].

144 As all four dimensions of personality functioning were designed to build upon the joint
145 construct of PD severity, and since the AMPD defines a current PD as the presence of
146 impairments in two or more of them, scales were expected and found to be highly intercorrelated
147 (Pearson correlation coefficients ranged between .41 and .83). Exploratory factor analysis on
148 item level supported a one-factor solution (i.e., strong first factor and a ratio of first to second
149 factors' eigenvalue of 5.1) speaking for a common factor of "personality pathology". This is
150 intended and in line with the goal of creating an assessment of the generalised severity of
151 personality pathology. However, all four domains of functioning had been operationalized
152 independently and in careful contrast to each other to make sure that each domain only covers
153 one of the described aspects of PD-related impairments with minimum overlap. Each item had to
154 show: sufficient item-total correlation as part of the assigned a) subscale, b) primary scale, and c)
155 total scale, respecting an internal consistent structure on all scale levels, and a reasonable effect
156 size for discriminating the school population and the PD patient sample as a sign of clinical
157 validity. Factor analytic approaches were not used to empirically select the final item set.
158 However, in an exploratory factor analysis on item level, a model with four factors accounted for
159 39.9% of the variance, and 72.2% of the items showed a loading $> .30$ on the factor that
160 corresponded to the theoretically assumed domain. This was interpreted as preliminary evidence
161 for the appropriateness of using the four domain scores [22]. However, with a Turkish translation

162 of the LoPF-Q 12-18, a four-factor model did not show adequate fit in a CFA [26]. Therefore,
163 the factor-analytical basis of the four domain scores has not yet been fully clarified.

164 **The Current Study**

165 The first goal of this study was an in-depth investigation of the factorial structure of the
166 LoPF-Q 12-18 items. Based on the preliminary analyses reported above, we expected that the
167 LoPF-Q 12-18 is essentially unidimensional, in the sense that most of the reliable variance of the
168 total score is due to a general factor. This is in line with research showing that different measures
169 of PD severity capture a strong common factor and can therefore be scaled along a single latent
170 continuum [29]. Nevertheless, previous research also indicates that specific factors might still
171 play a role even when a strong first factor is present [30–33]. Hence, Goth et al., (23, p. 687)
172 hypothesized that a bifactor structure might be suitable for the LoPF-Q 12-18, taking into
173 account a strong general factor as well as four empirically distinguishable domains.

174 Our second goal for the current study was to achieve a considerably shorter version
175 which maintains the structure of the questionnaire in terms of the four domains and the high
176 clinical validity of the original long version. With 97 5-point likert scale items, the LOPF-Q 12-
177 18 can be considered a somewhat long measure, at least for many research and clinical
178 applications with a focus on fast and efficient screening. Length can, therefore, be considered a
179 barrier for its usage. For instance, individuals with mental health problems often present in non-
180 specialised settings like primary care, school psychologist offices, or emergency departments
181 [34]. A shorter version would allow for administration of the instrument in a resource saving
182 manner. This is important, as the LoPF-Q 12-18 might not be the only instrument that needs to
183 be administered at a certain time. A short version can reduce burden for the patients and,

184 additionally, it reduces resources required for the scoring of the questionnaire. Taken together,
185 we expect that a short version will have a high impact on the practicability of the instrument.

186 **Materials and methods**

187 **Participants and procedures**

188 The current analyses were conducted on the same samples previously described in Goth
189 et al. 2018 [22]. In short, a school sample of $n = 351$ students was assessed at three public
190 schools. The BPFSC-11 (Borderline Personality Features Scale for Children, 11 Item Version;
191 [35]) was used to screen for the PD related health status, $n = 337$ were below the Cut-Off ≥ 34
192 and was taken as healthy control group. The study was reviewed and approved by the ethics
193 committee "Ethikkommission Beider Basel" which is now "Ethikkommission Nordwest- und
194 Zentralschweiz". Written informed consent has been obtained from all participants. A clinical
195 sample of $n = 415$ patients was recruited at inpatient and outpatient units of six child and
196 adolescent psychiatric hospitals in Basel, Innsbruck, Berlin, Mainz, Idar-Oberstein, and
197 Heidelberg. Inclusion criteria were age of 12 to 20 years, sufficient language and cognitive skills,
198 no autistic disorder, and no current psychotic episodes. Diagnoses were based on the results of
199 the clinical interviews Structured Clinical Interview for DSM-IV Axis II (SCID-II; [36]), the
200 Children's Diagnostic Interview for Psychiatric Diseases (K-DIPS; [37]) and a classification
201 conference. Patients with a PD diagnosis were assigned to the PD group independently from
202 Axis I diagnoses. Of the total clinical sample, $n = 96$ patients (23.1%) met the DSM-IV criteria
203 of one or more PDs (44.8% BPD). The total sample of $n = 766$ adolescents consists of 44.4%
204 boys and 55.6% girls, the age range was 12-20 years ($M = 15.5$, $SD = 1.9$). For details, please
205 see the description of the full study [22].

206 **Measures**

207 The LoPF-Q 12-18 [21] has been described above. It contains 97 items to be answered on
208 a 5-point scale ranging from 0 (no), 1 (more no), 2 (part/part), 3 (more yes) to 4 (yes). The
209 resulting four scales Identity, Self-Direction, Empathy, and Intimacy are coded towards
210 pathology and add up to a total score Personality Functioning, ranging from no impairment to
211 severe impairment. For descriptive reasons, two subscales per scale are included, matching
212 classical psychological concepts to facilitate interpretation. The test is available on the self-
213 publishing project website (academic-tests.com).

214 **Investigation of the latent structure**

215 For the investigation of the latent structure of the LoPF-Q 12-18, confirmatory factor
216 analyses on item level were used. Scale reliabilities were evaluated using McDonald's Omega.
217 The analyses were conducted with the software 'R' [38] and the package 'lavaan' [39]. Fig 1
218 illustrates models representing different factorial assumptions that were tested in order to
219 compare their fit.

220

221 *--- Fig 1: Different configural assumptions tested for the long version ---*

222

223 The following fit indices are reported: Comparative Fit Index (CFI), Root Mean Squared Error of
224 Approximation (RMSEA) and Standardized Root Mean Square Residual (SRMR) [40]. The
225 following combination of indices was used as the cut-off to determine acceptable models: CFI > .
226 90 acceptable and CFI > .95 good, RMSEA < .05, SRMR < .08 [41,42]. The model fit of the short
227 version created by ACO method (see below) was investigated using the same criteria. Scale
228 reliabilities were estimated using the package 'semTools' [43]. We report an ordinal version of
229 coefficient alpha according to Zumbo et. al [44], as well as omega hierarchical according to

230 McDonald [45]. Note that in bifactor models, omega hierarchical corresponds to OmegaH for the
231 total score and to OmegaHS for the subscale scores [46].

232

233 **Ant Colony Optimization to create a short version**

234 For small item pools it can be an option to iterate through all possible item combinations
235 or to apply simple iterative methods, e.g. a Stepwise Confirmatory Factor Analysis Approach
236 (SCOFA) in order to find a well-suited combination of items that can be used as a short version
237 of a test [47]. Considering that the LoPF-Q 12-18 consists of 97 items, the number of possible
238 item combinations reaches a level where this is no longer possible. To illustrate that such an
239 attempt is impossible, we calculated the number of possible combinations based on “n over k”,
240 as suggested previously [48]. The S2 Table illustrates the estimated computation time, memory,
241 and amount of energy that would be required to execute these iterations. Results show that the
242 algorithm would need to run for billions of years while using multiple times the global estimated
243 yearly energy consumption. Consequently, it is inevitable to run an optimization algorithm which
244 approximates a close-to-optimal short version of the questionnaire in a shorter period of time.

245 The Ant Colony Optimization (ACO) meta-heuristic [49] was used to select a set of items
246 for the short version. For our use case, this algorithm ran for less than 5 hours CPU time (see S2
247 Table). The ACO consists of virtual ‘ants’ who explore the selection of sets of items and
248 attribute a ‘pheromone’ level to the items of the selection according to a statistical criterion
249 (defined below). Items with a higher pheromone level have a higher probability to be re-selected
250 in future item sets to be explored. The pheromone level fades over time (“evaporation”). The
251 ACO was started with 30 ants, and the algorithm stopped after 20 runs without improvement.

252 This method has already proven useful in the construction of the 34-item “Personality Inventory
253 for DSM-5, Brief Form Plus” (PID5BF+) [50].

254 Our goal was to generate a short version of the LoPF-Q 12-18 with a total of 20 items,
255 including 5 items for each domain (Identity, Self-direction, Intimacy, Empathy). For this purpose,
256 the ACO method was set up to select a subset of 5 items from each domain of the original
257 version. The criterion to calculate the pheromone used by the ants was a combination of model
258 fit, reliabilities of the domain scales and clinical validity i.e. the capacity to discriminate between
259 patients with personality disorder (n = 96) and students without signs of personality disorder
260 according to the BPFSC-11 (n = 337). Model fit was based on a confirmatory factor analysis
261 with 4 first order factors to represent the domains and a secondary higher order factor to
262 represent generalised severity. The loadings of the domains on the higher order factor were
263 constrained to be equal, thereby ensuring a balanced interpretation of the general severity
264 continuum. Pheromones were calculated based on logistic transformations (ϕ) of fit measures
265 (CFI, RMSEA), measures of reliability (McDonald’s Omega and minimum factor loading) and
266 criterion validity (adjusted R²). The ultimately optimised (i.e. maximised) pheromone was based
267 on the sum of all three ϕ -values. Please refer to S3 Materials for the formulas used for
268 calculating pheromone levels.

269 Finally, to show the advantage of the applied ACO-algorithm over the iterative approach,
270 we compared reliability, CFA model fit as well as criterion validity with 100,000 random
271 combinations of items. Due to the high computational load, the calculation of these 100,000
272 models was performed at the sciCORE (<http://scicore.unibas.ch/>), the scientific computing centre
273 of the University of Basel. All analyses have been conducted using R (> version 4.0.1), as well
274 as the R packages ‘lavaan’ [39] and ‘semTools’ [43] for Confirmatory Factor Analyses. Receiver

275 operating characteristic (ROC) analysis was used to determine the clinical utility of the LoPF-Q
276 short version and to derive empirical cut-off scores for defining clinically relevant thresholds.

277 **Results**

278 **Investigation of the factorial structure of the original (97-item) version**

279 Table 1 shows the parameters of different confirmatory factor analyses (CFA) testing
280 different factorial assumptions.

281
282 Table 1: *Confirmatory Factor Analyses (CFA) testing different factorial assumptions (long*
283 *version)*

Model (id)	Factors	par	χ^2	CFI	RMSEA	SRMR
1-dim (1)	1	485	12341.2	0.851	0.057	0.090
2-dim (2)	2	486	11584.6	0.865	0.054	0.087
4-dim (3)	4	491	11019.2	0.876	0.052	0.084
2+bifactor (4)	2	582	9286.3	0.907	0.046	0.070*
4+bifactor (5)	4	582	9099.5*	0.911*	0.045*	0.072
2-dim hierarchical (6a) ⁱ	(2)	(487)	(21882.8)	(0.667)	(0.085)	(0.087)
2-dim hierarchical (6b)	2	486	11584.6	0.865	0.054	0.087
4-dim hierarchical (7a)	4	489	11089.0	0.875	0.052	0.085
4-dim hierarchical (7b)	4	486	11397.4	0.869	0.054	0.086

284 Note. *npar*: number of estimated parameters; *df*: degrees of freedom; *CFI*: Comparative Fit
285 *Index*; *RMSEA*: Root Mean Squared Error of Approximation; *SRMR*: Standard Root Mean
286 *Residual*. Best fit indices are highlighted with asterisks. ⁱ For model 6a standard errors could not
287 be computed and the information matrix could not be inverted. This may be a symptom of a non-
288 identifiable model. Therefore, the parameters of this model are shown in parenthesis.

289

290
 291 All factorial assumptions are related to the basic personality functioning concept, highlighting
 292 either the joint construct of PD severity, the two areas of Self-related and Interpersonal
 293 functioning, or the four domains Identity, Self-Direction, Empathy, and Intimacy according to
 294 the AMPD. Overall, the bifactor models performed slightly better than all other correlated or
 295 hierarchical factor models. A four-dimensional bifactor was the only model to show acceptable
 296 fit based on all three fit measures, RMSEA (<.05) and SRMR (<.08) and CFI (>.90). The two
 297 best fitting bifactor models (“two-dimensional bifactor” and “four-dimensional bifactor”)
 298 showed very similar fit indices with only a subtle difference on RMSEA. Table 2 summarizes the
 299 model-based scale reliabilities. Based on the best fitting model (model 5), ordinal alpha was
 300 excellent for the general factor as well as all four domains (>.90). However, while OmegaH
 301 showed excellent reliability of the total score (.94), OmegaHS was substantially lower for the
 302 four domains scales (.07, .11, .50, .20) (see S4 Tables – sheet 1 for additional details). Factor
 303 loadings of model 5 indicated that several items from the domain of empathy did not
 304 substantially (> .30) load on the general factor, and several items from the domain of identity had
 305 even negative loadings on the respective specific factor (see S4 Tables – sheet 2). S4 Tables –
 306 sheet 3 shows the factor intercorrelations for the bi-factor models.

307

308 Table 2: *Factor reliabilities for the long and short version*

	Long version (models)									Short Version ⁱ
	1	2	3	4	5 ⁱⁱ	6a	6b	7a	7b	
<i>Ordinal Alpha</i>										
Total score	0.98	0.98	0.98	0.97	0.98	0.98	0.97	0.97	0.97	.91
Identity		0.97	0.94	0.97	0.94	0.97	0.97	0.94	0.94	.80
Self-direction			0.96		0.96			0.96	0.96	.84

	Long version (models)								Short	
Empathy	0.95	0.90	0.95	0.90	0.95	0.95	0.90	0.90	.71	
Intimacy		0.94		0.94			0.94	0.94	.78	
<i>Hierarchical Omega</i>										
Total score	1.00	1.00	1.00	0.99	0.94	1.00	1.00	1.00	1.00	.93
Identity		0.99	0.96	0.17	0.07	0.99	0.99	0.96	0.99	.78
Self-direction			0.97		0.11			0.97	0.99	.84
Empathy		0.92	0.77	0.22	0.50	0.92	0.92	0.77	0.61	.68
Intimacy			0.95		0.20			0.95	0.93	.75

309 *Note. Alpha = ordinal alpha; Omega = hierarchical Omega*

310 ⁱ *the optimized short version corresponds to the factorial assumptions of model 7b.*

311 ⁱⁱ *best fitting model in Confirmatory Factor Analysis (see Table 1)*

312

313 Taken together, this suggests that although model 5 has the best fit, it is not a very satisfying
 314 representation of the structure of the LoPF-Q 12-18 [51]. In consequence, we have chosen a
 315 hierarchical model with four lower-order factors (i.e., model 7b) for developing the short version.

316

317 **Creating a short version**

318 As intended, a 20-item version with 5 items for each of the domain scales was obtained.
 319 Fit indices presented in Table 3 show that the optimised short version had a very good fit on all
 320 fit indices (CFI = .980, RMSEA = .046, SRMR = .038).

321

322 *Table 3: Model fit indices and external validity for the Ant Colony Optimised Short version*
 323 *compared to 100,000 random combinations of items*

324

Model (id)	Factors	par	χ^2	CFI	RMSEA	SRMR	Adj. R ²
ACO short version ^a	101	169	252.3*	0.980*	0.046*	0.038*	0.425*
100,000 random combinations ^a	101	169	716.7	0.918	0.080	0.067	0.389
				(0.02)	(0.01)	(0.01)	(0.02)

325 Note. *npar*: number of estimated parameters; *df*: degrees of freedom; *CFI*: Comparative Fit
326 Index; *RMSEA*: Root Mean Squared Error of Approximation; *SRMR*: Standard Root Mean
327 Residual. Best fit indices are highlighted with asterisks. *Adj. R²* = external validity (variance
328 explained). ^a modelling corresponds to the factorial assumptions of model 7b.

329

330 Due to the smaller number of items, ordinal alpha was slightly lower in comparison to the long
331 version (.91 total scale, .71 - .84 domains). Omega hierarchical of the total score was similar
332 (.93). Factor loadings of the optimised short version are depicted in Fig 2.

333

334 --- Fig 2: Factor loadings of short version ---

335 Fig 2 - legend: Optimized model using ant colony optimization to develop a short version to
336 identify personality disorders in adolescence. The configuration corresponds to model 7b in Fig

337 
338 1

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340

341 The final short version was optimized in a way that the lower-order factor loadings were kept
342 constant across all four domains. This procedure yielded very high and equidistant factor
343 loadings (.95) on the general factor.

344 For comparison, Table 2 additionally shows the average model fits (and standard

deviations) for 100,000 randomly selected item combinations testing the same hierarchical factor

345 model. Fig 3 visually compares the 20-item solution that was generated with the ACO with
346 100'000 random combinations of items regarding external validity and model fit. Compared to
347 the random combinations, the combination of model fit and external validity (the ability to
348 differentiate between healthy controls and PD patients) of the short ACO version are excellent
349 with an adjusted R square of .425 (i.e. 42,5% explained variance).

350

351 *--- Fig 3: Fit and external validity of short version ---*

352 *Fig 3 - legend: Model fit and external validity of the optimized short version in comparison to*
353 *100,000 random item combinations.*

354

355

356 Expressed with the more traditional effect size Cohens d, the LoPF-Q Short total score differs
357 between the PD patients and the healthy controls with $d = 3.1$ standard deviations. ROC analysis
358 showed an area under the curve (AUC) of .92 ($p < .001$; 95 % confidence interval .89-.95). A
359 preliminary cut-off score for the LoPF-Q Short total score was defined to be ≥ 36 using Youden
360 index, corresponding to a T-score of 74. Specificity for detecting patients with personality
361 disorder compared to healthy controls was 87.5% and sensitivity was 80.2%. Reliability
362 coefficients for the short version can be seen in Table 2.

363

364

Discussion

365 The current study had two aims: First, to investigate the latent structure of the LoPF-Q
366 12-18 original 97-item version. We assumed that a bifactor model with a general factor –
367 representing a joint construct of PD severity – with four specific factors matching the four

368 domains of functioning according the AMPD (DSM-5) would perform well. The second goal
369 was the construction of an optimised short version to meet the needs for an efficient screening
370 instrument for PD in adolescents.

371 **Factorial structure of LoPF-Q 12-18 long version**

372 As hypothesised, the nominally best fitting models were bifactor models when compared
373 to correlated factor or hierarchical factor models (see Fig 1). The model fit of both bifactor
374 models (including either two broad dimensions or four narrow domains) was acceptable when
375 considering all evaluated criteria. However, other aspects besides model fit should be considered
376 when interpreting bifactor models [51] due i.a. to their less restrictive nature which results in a
377 higher overall chance of good fit, even when using random data. In addition to an acceptable fit,
378 the items should show substantial loadings on the general factor, and the specific scales should
379 have sufficient reliability after controlling for variance of the general factor (i.e., Omega HS). In
380 both these respects the estimates of the bi-factor models were lacking. Conclusively, bi-factor
381 models did not satisfactorily represent the structure of the questionnaire despite the acceptable fit.
382 Nevertheless, the following important conclusions can be drawn from the performed analyses.
383 First, as expected, the item level data collected with the LoPF-Q 12-18 contain a very strong
384 general factor. This can be seen, for example, in the fact that model fit was only moderately
385 improved by extracting more than one factor, that the four domains in model 3 were very highly
386 correlated (S4 Tables – sheet 3), and that the reliable variance in the total score in model 5 (i.e.,
387 Omega H) was almost entirely attributable to the general factor (S4 Tables – sheet 1). This
388 support for a general factor of personality functioning is very much in line with the usage of the
389 LoPF-Q 12-18 in the framework of diagnostic procedures of both the AMPD (DSM-5) and the
390 ICD-11. In both diagnostic models, personality functioning is seen as an overarching construct

391 important to establish a PD diagnosis and to judge its severity. Importantly, according to [52],
392 the general factor of personality pathology has been primarily described in adult populations.
393 The current study might be one of the first to describe this general factor in a sample of younger
394 patients.

395 Second, the four domain subscales, with the possible exception of empathy,
396 contain hardly any reliable variance beyond general severity. In other words: Although the four
397 domain scores were reliable in their own right (i.e., ordinal alpha > .90), their very high
398 correlation in the underlying sample makes it seem unlikely that distinctive and clinically
399 interpretable profiles will emerge in individual cases. This contrasts with PD criteria from DSM-
400 IV [32,53] or items of the Inventory of Personality Organization [31], which tend to warrant
401 scoring of subscales in adult samples. At least on a group level, [22] found first evidence of
402 distinctive profiles, for example, the empathy scale was severely impaired only in patients with
403 narcissistic and antisocial PD, whereas the identity scale was particularly impaired in patients
404 with Borderline or anxious-avoidant PD. The specific clinical variation of the empathy scale may
405 be an explanation for why only this one showed an independent variance beyond the general
406 factor. In sum, whether the use of each of the four domain scores is clinically meaningful needs
407 to be investigated in clinical trials with different types of PD patients and optimally with
408 different therapeutic approaches in a longitudinal design.

409 The debate on the meaningfulness of the domain scales is important as mental health care
410 workers tend to find the primary scales and subscales of the LoPF-Q 12-18 useful for the
411 interpretation of the assessments regarding clinical decision making and therapy planning. This
412 is comprehensible as the less abstract denomination of the subscales appear to be closer to
413 commonly shared concepts and can be used to find a shared language with the patients and their

414 families. According to the authors [22] the LoPF-Q 12 -18 has been primarily developed to meet
415 the needs of clinical practitioners and to cover a wide range of symptoms related to the four
416 domains, because often specific aspects of functioning (identity pathology, problems with self-
417 regulation or problems with social interaction etc.) are the primary target for psychotherapy. This
418 discrepancy between the authors' experiences and intentions and our current findings cannot be
419 conclusively clarified. The currently investigated sample consisted mainly of subjects without
420 signs of PD (351 from schools and 319 patients without PD vs 96 patients with PD). The general
421 factor might turn out being less pronounced and the domain scales more independent from each
422 other when investigating clinical samples of PD patients [54]. Similarly, Watts et al. found that
423 the inclusion of undiagnosed individuals causes more positive correlations in psychopathological
424 data, leading to a stronger p -factor [55]. For a further optimization of the structure in a short
425 form of the LoPF-Q 12-18, it seemed reasonable to keep the four domains in terms of content
426 validity, but to put the focus of the optimization on the general factor.

427 **LoPF-Q Screener (20-item version)**

428 The short version was optimised for clinical validity and internal consistency accounting
429 for a structure with four first order factors to represent the personality functioning domains and a
430 secondary higher order factor to represent the general personality functioning denoting PD
431 severity. Thanks to the optimisation, the short version performed excellently regarding both
432 external validity and internal consistency. The optimisation was done with the ACO heuristic
433 which had already proven useful in previous studies for creating short personality assessments
434 [31,50] and performed very well in the current study (see Fig 3). The derived short version
435 "LoPF-Q Screener" contains 20 items and preserves the four scales Identity, Self-direction,
436 Empathy and Intimacy as well as the total scale Personality Functioning. It showed an excellent

437 model fit concerning all parameters and good scale reliabilities. Most importantly, it showed
438 excellent clinical validity, with the total scale differentiating significantly and with an effect size
439 of 3.1 standard deviations between PD patients and healthy adolescents.

440 The LoPF-Q Screener can be used in contexts where employing the longer version is not
441 feasible or inconvenient. This flexibility cannot be overestimated in the presence of a general
442 global mental health gap [56] in adolescents and a specific gap regarding personality disorders in
443 youth [6,7]. Tools that can help address these gaps are required, and while diagnostic tools
444 cannot solve this issue alone, they are one of the cornerstones to advance research and
445 interventions. The results on psychometric properties of this short version are still preliminary
446 and need to be verified with test data that were not used for its construction. The data ideally
447 needs to be collected with this short version in order to validate it, since using a subsample of
448 items of data collected with the long version might potentially introduce bias (e.g. memory
449 effects, effects of the sequential order of items, attention span of the subject etc.). Finally, the
450 question arises whether an even shorter version wouldn't be better in terms of practicality of the
451 assessment and, thus, versatility in clinical contexts. However, an even shorter version may come
452 at the expense of inferior measurement precision and diagnostic validity, both of which are
453 highly relevant for clinical usage [57]. The 20-item version of the LoPF Q 12-18 is likely to
454 present a solid compromise between psychometric precision and practicality.

455 **Research recommendations**

456 Research on the usefulness of the levels of personality functioning model for clinical
457 decision making such as selection of appropriate treatment and treatment customisation is needed.
458 The long and short version need to be compared in future studies regarding their usability and
459 user experience of the different stakeholders. For instance, do users benefit from the more

460 comprehensive data collection of the long version or are they looking for more efficient tools? A
461 further question is the preparation of a pathway towards shorter versions for different cultural
462 settings. The authors of the LoPF-Q 12-18 pursue a strategy in which they emphasise the
463 importance of the same set of items for all cultural settings and actively support the development
464 of cultural adaptations and networking among interested colleagues. A shared set of items across
465 culturally adapted versions is necessary because it facilitates scientific exchange and
466 management of the different versions and enables joint data analyses in cross-cultural settings.
467 This possibility is particularly important because the development of PD in early adolescence is
468 an under-researched area and data pooling is key. In addition, LoPF-Q versions for informant
469 report and for even younger age groups (from 6 years up) are under development, and the
470 seamless and clear transferability of the assessed scales in all cultural adaptations is crucial,
471 especially for longitudinal studies. Future research will show whether the optimised short version
472 LoPF-Q Screener will provide measurement invariance across different cultural settings and
473 translations.

474 The current study highlights the usefulness of a more detailed and more time-efficient
475 assessment of personality functioning in adolescence. Whereas there is no doubt about a
476 common core, i.e. a general latent construct, there is somewhat mixed evidence regarding the
477 usefulness of the lower-order domains (identity, self-direction, empathy, intimacy). Earlier
478 research on alcohol use disorders has shown that determining the factor structure in a sample
479 including individuals with no clinical symptoms may have a debilitating impact on the
480 discrimination of sub-factors [54]. Future research on the LoPF-Q 12-18 and the introduced
481 LoPF-Q Screener short version might provide more comprehensive insights by comparing the
482 factor structure between clinical and non-clinical samples.

484

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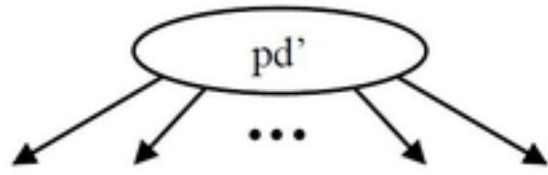
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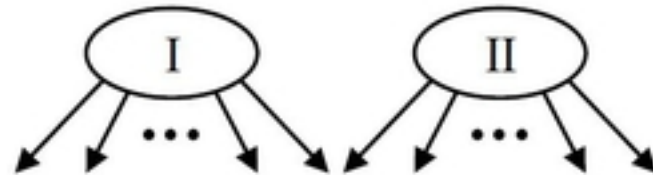
Model 1

1-dimensional



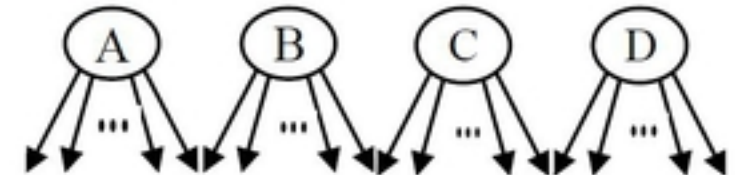
Model 2

2-dimensional
(factors allowed to correlate)



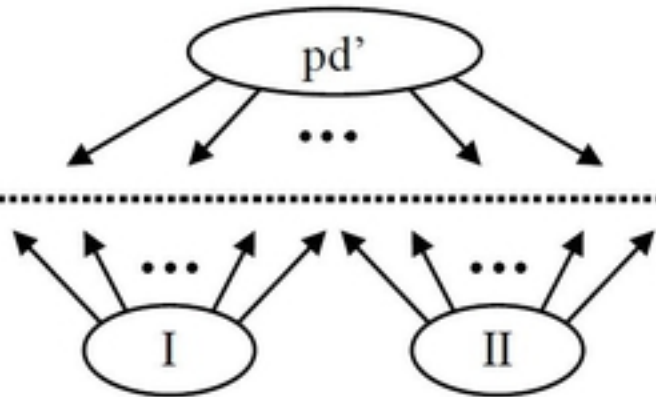
Model 3

4-dimensional
(factors allowed to correlate)



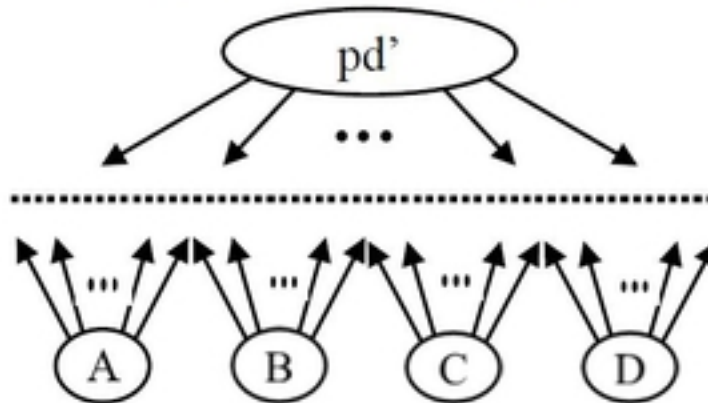
Model 4

2-dimensional + bifactor
(all factors orthogonal)



Model 5

4-dimensional + bifactor
(all factors orthogonal)



Factors:

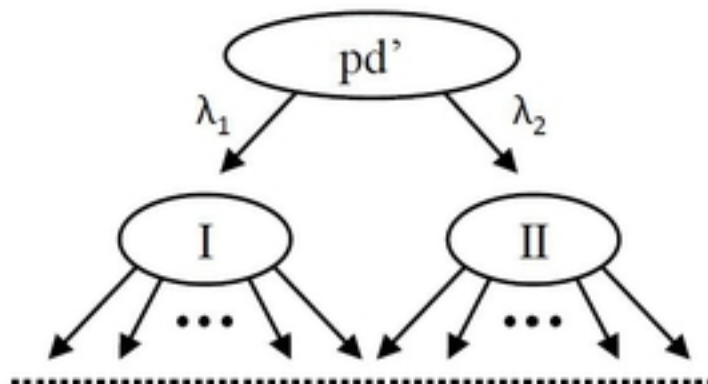
pd': General factor of personality functioning

I: Interpersonal
II: Intrapersonal

A: Identity
B: Self-direction
C: Intimacy
D: Empathy

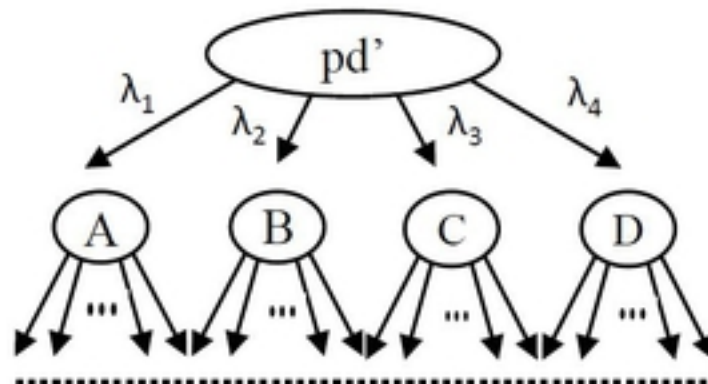
Models 6a, 6b

2-dimensional
(hierarchical)



Models 7a, 7b

4-dimensional
(hierarchical)



Model restrictions:

Model 6b: $\lambda_1 = \lambda_2$

Model 7b: $\lambda_1 = \lambda_2 = \lambda_3 = \lambda_4$

Indicators:

97 items

Figure 1

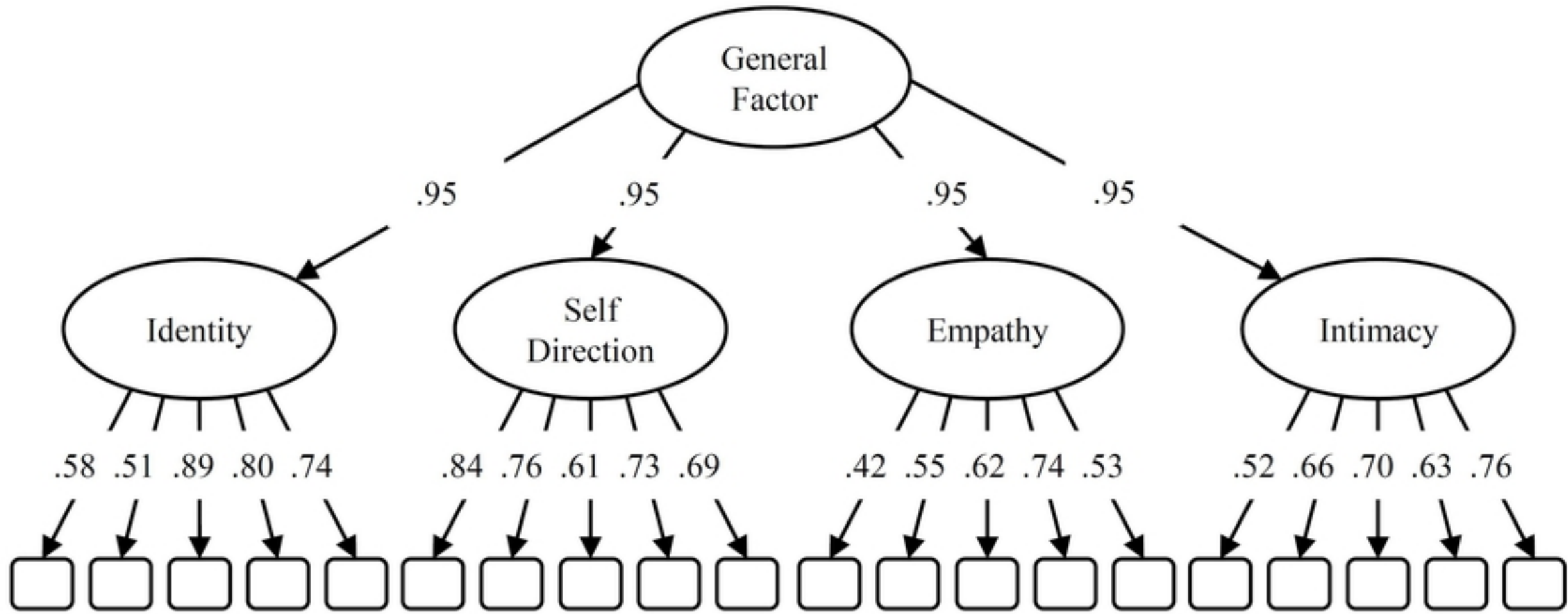


Figure 2

● 100k random combinations ● Short version

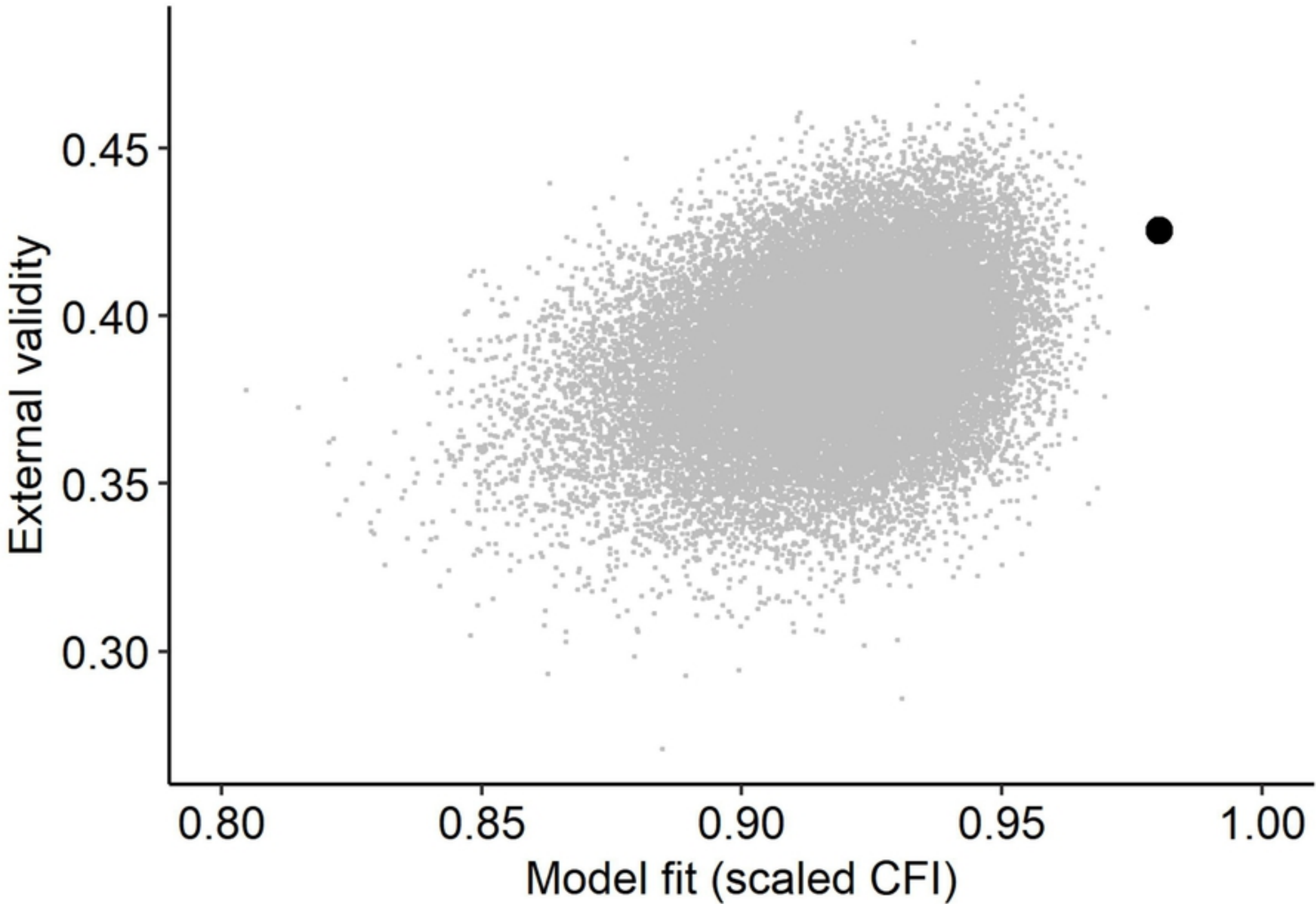


Figure 3