

## A HÁJEK-RÉNYI TYPE INEQUALITY FOR STOCHASTIC VECTORS WITH APPLICATIONS TO SIMULTANEOUS CONFIDENCE REGIONS<sup>1</sup>

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For a sequence of stochastic vectors forming either a forward or a reverse martingale, a Hájek-Rényi type inequality is derived, and its applications in some problems of simultaneous confidence regions are stressed.

**1. Statement of the result.** A variety of multivariate Chebyshev inequalities is available in the literature; we refer to Karlin and Studden (1966) and to Mudholkar and Rao (1967) which include earlier references. In the present note, for stochastic vectors, a simultaneous inequality comparable to Chow's (1960) semi-martingale extension of the Hájek-Rényi (1955) inequality is considered.

Let  $\{\mathbf{Z}_i, i \geq 1\}$  be a sequence of stochastic  $p$ -dimensional column vectors, where  $p \geq 1$ . Let  $\mathcal{B}_i = \mathcal{B}(\mathbf{Z}_1, \dots, \mathbf{Z}_i)$  and  $\mathcal{C}_i = \mathcal{C}(\mathbf{Z}_i, \mathbf{Z}_{i+1}, \dots)$  be the  $\sigma$ -fields generated by  $(\mathbf{Z}_1, \dots, \mathbf{Z}_i)$  and  $(\mathbf{Z}_i, \mathbf{Z}_{i+1}, \dots)$  respectively,  $i \geq 1$ ; clearly,  $\mathcal{B}_i$  is  $\uparrow$  while  $\mathcal{C}_i$  is  $\downarrow$  in  $i$ . Suppose that  $E\mathbf{Z}_i = \mathbf{0}$  and  $E\mathbf{Z}_i\mathbf{Z}_i'$  exists for all  $i \geq 1$ . Also, assume that either of the following two conditions holds:

- (1)  $E(\mathbf{Z}_n | \mathcal{B}_k) = \mathbf{Z}_k$  almost surely (a.s.) for all  $n \geq k \geq 1$ ,  
 (2)  $E(\mathbf{Z}_k | \mathcal{C}_n) = \mathbf{Z}_n$  a.s., for all  $n \geq k \geq 1$ .

Let  $\mathbf{A}$  be an arbitrary  $(p \times p)$  positive definite (p.d.) matrix, and let

$$(3) \quad \zeta_n = E(\mathbf{Z}_n' \mathbf{A}^{-1} \mathbf{Z}_n), \quad \zeta_{n+1}^* = E[(\mathbf{Z}_{n+1} - \mathbf{Z}_n)' \mathbf{A}^{-1} (\mathbf{Z}_{n+1} - \mathbf{Z}_n)], \quad n \geq 1,$$

where  $\mathbf{A}\mathbf{A}^{-1} = \mathbf{I}_p$ . Then, we have the following theorem.

**THEOREM 1.** For a non-increasing sequence  $\{c_i\}$  of positive constants, under (1), for every  $\varepsilon > 0$ ,  $n \geq 1$ , and  $N \geq 1$ ,

$$(4) \quad P[\max_{n \leq k \leq n+N} c_k \{\sup_{\lambda \neq 0} (\lambda' \mathbf{A} \lambda)^{-\frac{1}{2}} |\lambda' \mathbf{Z}_k|\} > \varepsilon] \leq \varepsilon^{-2} \{c_n^2 \zeta_n + \sum_{k=n+1}^{n+N} c_k^2 \zeta_k^*\};$$

for a non-decreasing sequence  $\{c_i\}$  of positive constants, under (2),

$$(5) \quad P[\max_{n \leq k \leq n+N} c_k \{\sup_{\lambda \neq 0} (\lambda' \mathbf{A} \lambda)^{-\frac{1}{2}} |\lambda' \mathbf{Z}_k|\} > \varepsilon] \leq \varepsilon^{-2} \{c_{n+N}^2 \zeta_{n+N} + \sum_{k=n+1}^{n+N} c_k^2 \zeta_k^*\}.$$

It may be noted that in (4) or (5), when  $N = 0$ , the second term on the right-hand side should be taken as equal to zero. If we let  $c_k = c_n$ ,  $n \leq k \leq n+N$ , we obtain the Kolmogorov-type inequality, while for  $N = 0$ , this reduces to the Chebyshev-type inequality. Some applications are considered in Section 3.

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**2. Proof of the theorem.** By the Schwarz-inequality

$$(6) \quad \sup_{\lambda \neq 0} (\lambda' \mathbf{A} \lambda)^{-\frac{1}{2}} |\lambda' \mathbf{Z}_k| = (\mathbf{Z}_k' \mathbf{A}^{-1} \mathbf{Z}_k)^{\frac{1}{2}} (\geq 0), \quad \text{for all } k \geq 1.$$

Hence, from (6), we have

$$(7) \quad P\{\max_{n \leq k \leq n+N} c_k [\sup_{\lambda \neq 0} (\lambda' \mathbf{A} \lambda)^{-\frac{1}{2}} |\lambda' \mathbf{Z}_k|] > \varepsilon\} = P\{\max_{n \leq k \leq n+N} c_k^2 Y_k > \varepsilon^2\},$$

where  $Y_k = (\mathbf{Z}_k' \mathbf{A}^{-1} \mathbf{Z}_k)$ ,  $k \geq 1$ . Now, under (1),

$$(8) \quad E[Y_n | \mathcal{B}_k] = Y_k + 2E\{(\mathbf{Z}_n - \mathbf{Z}_k)' | \mathcal{B}_k\} \mathbf{A}^{-1} \mathbf{Z}_k + E[(\mathbf{Z}_n - \mathbf{Z}_k)' \mathbf{A}^{-1} (\mathbf{Z}_n - \mathbf{Z}_k) | \mathcal{B}_k] \\ = Y_k + E[(\mathbf{Z}_n - \mathbf{Z}_k)' \mathbf{A}^{-1} (\mathbf{Z}_n - \mathbf{Z}_k)' | \mathcal{B}_k] \geq Y_k \text{ a.s., for all } n \geq k \geq 1.$$

Hence,  $\{Y_k, \mathcal{B}_k, k \geq 1\}$  forms a nonnegative semi-martingale sequence. Consequently, on using the second inequality in Theorem 1 of Chow (1960) [which provides the semi-martingale extension of the Hájek-Rényi (1955) inequality], the right-hand side of (4) directly follows from (3) and (7).

By reversing the ordering of the index set  $\{i\}$  in (2), the reverse martingale property of  $\{\mathbf{Z}_i, \mathcal{C}_i, i \geq 1\}$  can be converted into forward martingale property, and hence, the same proof as in (4) applies. This completes the proof of (5).  $\square$

**3. Some applications to simultaneous confidence regions.** We consider here the following three problems.

(I) Let  $\omega = \{\mathbf{X}_1, \mathbf{X}_2, \dots, \text{ad inf}\}$  be a sequence of independent stochastic  $p$  dimensional column vectors, where  $E\mathbf{X}_i = \boldsymbol{\mu}_i$  and  $\mathbf{V}(\mathbf{X}_i) = \boldsymbol{\Sigma}_i, i \geq 1$ . Let  $\mathbf{X}_n = n^{-1} \sum_{i=1}^n \mathbf{X}_i, \bar{\boldsymbol{\mu}}_n = n^{-1} \sum_{i=1}^n \boldsymbol{\mu}_i$ , and let

$$(9) \quad \mathbf{T}_n = n(\bar{\mathbf{X}}_n - \bar{\boldsymbol{\mu}}_n) = \sum_{i=1}^n (\mathbf{X}_i - \boldsymbol{\mu}_i), \quad n \geq 1.$$

It follows that  $\{\mathbf{T}_n, \mathcal{B}_n, n \geq 1\}$  forms a forward martingale sequence, i.e., (1) holds. If we let  $v_i^* = \text{Trace}(\boldsymbol{\Sigma}_i \mathbf{A}^{-1}), i \geq 1$ , we have from (3),  $\zeta_{n+1}^* = v_{n+1}^*$  and  $\zeta_n = v_n = v_1^* + \dots + v_n^*, n \geq 1$ . Hence, from (4), we obtain that

$$(10) \quad P[\max_{n \leq k \leq n+N} (kc_k) [\sup_{\lambda \neq 0} (\lambda' \mathbf{A} \lambda)^{-\frac{1}{2}} |\lambda' (\bar{\mathbf{X}}_k - \bar{\boldsymbol{\mu}}_k)|] > t] \leq t^{-2} \\ \cdot [c_n^2 v_n + \sum_{k=n+1}^{n+N} c_k^2 v_k^*].$$

In particular, if  $\boldsymbol{\Sigma}_i = \boldsymbol{\Sigma}, \forall i \geq 1$ , and we let  $c_k = k^{-1}, \mathbf{A} = \boldsymbol{\Sigma}$ , we obtain from (10) that

$$(11) \quad P[|\lambda' (\bar{\mathbf{X}}_k - \bar{\boldsymbol{\mu}}_k)| \leq t(\lambda' \boldsymbol{\Sigma} \lambda)^{\frac{1}{2}}, \forall \lambda \neq 0, n < k < n+N] \\ \geq 1 - pt^{-2} [n^{-1} + \sum_{k=n+1}^{n+N} k^{-2}] \geq 1 - p(2N+n)/[n(n+N)t^2].$$

For  $N = 0$ , (11) is analogous to the Scheffé-type (cf. [7] page 68) simultaneous confidence region for  $\lambda' \bar{\boldsymbol{\mu}}_n$  (or  $\lambda' \boldsymbol{\mu}$  when all the  $\boldsymbol{\mu}_i$  are equal) under the Chebyshev set up (i.e., under no assumption of normality, inherent in [7]). For  $N \geq 1$ , it is an extension along the lines of the Kolmogorov inequality.

(II) If the  $\mathbf{X}_i$  are identically distributed with mean  $\boldsymbol{\mu}$  and dispersion matrix  $\boldsymbol{\Sigma}$ ,  $\{(\bar{\mathbf{X}}_k - \bar{\mathbf{X}}_{n+N}), \mathcal{C}_k, n < k < n+N\}$  has the reverse martingale property, for all  $n \geq 1$ ; that is, (2) holds for  $\mathbf{Z}_k = \bar{\mathbf{X}}_k - \bar{\mathbf{X}}_{n+N}, n \leq k \leq n+N$ . Hence, by (5),

$$(12) \quad P[\max_{n \leq k \leq n+N} n^{\frac{1}{2}} |\boldsymbol{\lambda}'(\bar{\mathbf{X}}_k - \bar{\mathbf{X}}_{n+N})| < \varepsilon(\boldsymbol{\lambda}'\boldsymbol{\Sigma}\boldsymbol{\lambda})^{\frac{1}{2}}, \forall \boldsymbol{\lambda} \neq \mathbf{0}] \\ \geq 1 - p\varepsilon^{-2}N/(n+N) \geq 1 - \eta,$$

whenever  $N \leq \delta n$  and  $\delta/\varepsilon^2 \leq \eta (> 0)$ . (12) establishes the 'uniform continuity in probability' with respect to  $n^{-\frac{1}{2}}$  [in the sense of Anscombe (1952)] for all possible linear compounds  $\{\boldsymbol{\lambda}'\mathbf{X}_k, \boldsymbol{\lambda} \neq \mathbf{0}\}$ . This result is useful for the study of sequential (simultaneous) confidence regions for all possible linear compounds of  $\boldsymbol{\mu}$ .

(III) Consider now a  $p$ -variate separable semi-martingale  $\{\mathbf{Z}_t, t \geq 0\}$ , such that (i)  $E\|\mathbf{Z}_t\| < \infty, \forall t > 0$ , where  $\|\mathbf{x}\|$  stands for the Euclidean norm of a vector  $\mathbf{x}$ . Let  $f(t)$  be a non-decreasing positive function on  $[0, \tau]$  where  $\tau > 0$ , and let  $E\mathbf{Z}_t = \mathbf{0}, \forall t \geq 0$ ,

$$(13) \quad \zeta(t) = E[\mathbf{Z}_t' \mathbf{A}^{-1} \mathbf{Z}_t], \quad t \geq 0,$$

where  $\mathbf{A}$  is p.d., and assume that (i)  $\zeta(t)/f^2(t) \rightarrow a_0 (< \infty)$  as  $t \rightarrow 0$ , and (ii)  $\int_0^\tau [f(t)]^{-2} d\zeta(t)$  exists. Then, by virtue of (6), we obtain on proceeding as in Theorem 5.1 of Birnbaum and Marshall (1961) that

$$(14) \quad P\{\sup_{t \in [0, \tau]} [\sup_{\boldsymbol{\lambda} \neq \mathbf{0}} (\boldsymbol{\lambda}'\mathbf{A}\boldsymbol{\lambda})^{-\frac{1}{2}} |\boldsymbol{\lambda}'\mathbf{Z}_t|/f(t)] \geq 1\} \leq a_0 + \int_0^\tau [f(t)]^{-2} d\zeta(t).$$

The last inequality provides a multivariate extension of Theorem 5.1 of Birnbaum and Marshall and also an extension of [4] to separable semi-martingale processes.

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