A STUDY OF THE DISTRIBUTION OF MEANS ES-TIMATED FROM SMALL SAMPLES BY THE METHOD OF MAXIMUM LIKELIHOOD FOR PEARSON'S TYPE II CURVE

By JOHN L. CARLSON

The object of this paper is to study the distribution of estimates of the parameter of location for Pearson's Type II Curve. estimated by the method of maximum likelihood from small samples.

R. A. Fisher has assumed, and Professor Hotelling has proved that in large categories of cases the distribution of an optimum statistic approaches normality as the sample size increases. This normality has been assumed to hold for optimum statistics in general whether calculated from large samples or small ones, and it has also been assumed that optimum statistics have minimum variance and always give better fits than do statistics calculated by the method of moments. That this is the case whenever the sample is large and the distribution of optimum statistics normal is made plausible by the reasoning of R. A. Fisher. In case the sample is small, however, there may be reason to doubt that the normality of distribution of optimum statistics holds, and that the other conclusions hold. It is with this phase of the subject that we shall be concerned in what follows.

Before entering into the topic under discussion it will be convenient to review some of the more elementary facts regarding the curve with which we are to be concerned. We shall take first the general equation for the curve in the form,

(1)
$$y = y_0 \left(1 - \frac{(x-m)^2}{Q^2}\right)^P$$

¹On the Mathematical Foundations of Theoretical Statistics, R. A. Fisher, Phil. Trans. Series A, Vol. 222, 1922. Pp. 309-368.

²The Consistency and Ultimate Distribution of Optimum Statistics. Harold Hotelling, Trans. Amer. Math. Soc. Vol. 32, No. 4. Pp. 847-859.

⁸Ibid, Pp. 328-368.

and determine the effect of variation of the constants.

When p=0 the equation reduces to the straight line $y=y_0$ When p=+1 the equation is that of a parabola with $y=y_0$ at the point x=m and the x intercepts at the points $x=\pm a$. It of course meets the x axis at an angle.

When p = +2 the equation y = 0 is of the fourth degree in x with double roots at the points $x = \pm a$.

In general when p=n the equation y=0 is of degree 2n in x and has two sets of n-fold multiple roots $x=\pm a$.

Since our curve is to be a probability curve it will of necessity have unit area, and this fact makes it possible for us to evaluate y_0 in terms of the parameters a and p. In order to do this we shall perform the integration below.

Area =
$$1 = \int_{x-a}^{x+a} y \, dx = 2y_0 \int_{0}^{a} \left(\left(-\frac{x^2}{az} \right)^p \right) dx$$

whence

$$1 = 2 a y_o \int_0^{\frac{\pi}{2}} \cos^{2p+1} \theta d\theta$$

but now since

(2)
$$\int_{a}^{\frac{\pi}{2}} \cos^{n}\theta d\theta = \frac{\sqrt{\pi}}{2} \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{n+2}{2})}$$

we have
$$1 = 2ay_0 \sqrt{\frac{\pi}{2}} \frac{\Gamma(\rho+1)}{\Gamma(\rho+\frac{3}{2})}$$

and so

$$y_0 = \frac{\Gamma(\rho + \frac{3}{2})}{\alpha \sqrt{\pi} \Gamma(\rho + 1)}$$

Therefore

(3)
$$y = \frac{\Gamma(\rho + \frac{3}{2})}{a\sqrt{\pi}\Gamma(\rho + 1)} \left[1 - \frac{(x - m)^2}{a^2}\right]^{\rho}$$

Formula (2) we shall find to be of value a number of times and formula (3) is the form in which equation (1) will be used throughout the remainder of the paper.

It will be worth while now to consider the likelihood function L together with its first and second partial derivatives with respect to m. (We shall hereafter refer to m as the parameter of location, a as the parameter of scaling, and p as the parameter of shape.) We are to use m to denote the estimate of m obtained by the method of maximum likelihood, in accordance with the convention introduced by Fisher, and it will be with this parameter that we shall concern ourselves in this investigation.

We have from (3) on the preceding page

(4)
$$L = n \log \frac{\Gamma(\rho + \frac{3}{2})}{a f \overline{h} \Gamma(\rho + 1)} + \rho \sum_{i=1}^{n} \log \left[1 - \frac{(x_i - m)^2}{a^2}\right]$$

and so

(5)
$$\frac{\partial L}{\partial m} = 2\rho \sum_{i=1}^{n} \frac{x_i - m}{a^2 - (x_i - m)^2}$$

and

(6)
$$\frac{\partial^2 L}{\partial m^2} = -2p \sum_{i=1}^n \frac{\alpha^2 + (x_i - m)^2}{[\alpha^2 - (x_i - m)^2]^2}$$

¹Thid Pp. 309-368.

At this point we shall stop to consider the effect of variation of the parameters a and p upon our estimate of \widehat{m} . Let us first consider p. Since the method of maximum likelihood is here merely the method of the differential calculus it follows from a consideration of equation (5) that our estimate of \widehat{m} will be independent of p, for any particular sample. Such is not the case when we consider a, however, for any change in a allows a change in the variance of \widehat{m} for the particular sample.

We shall find it advantageous to cover as much of the theoretical work as possible before embarking upon our experimental check and its great amount of numeral calculations, for it is only by means of a check between theory and experiment that we are able in the present state of knowledge, to judge of the applicability of the method of maximum likelihood to small samples. Of course it will be necessary to consider the distribution of our estimates of \hat{m} , in order to make this statistic of practical use, and it is desirable to know the theoretical variances of \bar{z} the arithmetic mean, \hat{m} , and the experimentally obtained variance of the distribution of our estimates of \hat{m} . The first of these we can obtain from theory, the second from an approximation valid only in the limit, and the last by means of calculation based on actual sampling.

We shall be concerned first with the theoretical variance of \bar{z} . It is well known that the variance of \bar{z} is equal to the variance of the distribution divided by \bar{z} . We must first, therefore, find the variance of the distribution.

From (3) page 88 it follows that

$$\sigma^{2} = \frac{2\Gamma(\rho + \frac{3}{2})}{a\sqrt{\pi}\Gamma(\rho+1)} \int_{0}^{a} x^{2} \left(1 - \frac{x^{2}}{a^{2}}\right)^{\rho} dx$$

$$\sigma^{2} = \frac{2\Gamma(\rho + \frac{3}{2})}{a\sqrt{\pi}\Gamma(\rho+1)} \int_{0}^{\frac{\pi}{2}} \left[\cos^{2\rho+1}\theta - \cos^{2\rho+3}\theta\right] d\theta$$

remembering now (2) page 87

$$\sigma^2 = \frac{a^2}{2p+3}$$

whence it follows that

(7)
$$\sigma_{\bar{z}}^2 = \frac{\sigma^2}{n} = \frac{\alpha^2}{n(2\rho+3)}.$$

We shall now calculate the limiting form of the variance of \widehat{m} . Fisher has proved that if the distribution of optimum statistics is normal the variance of an optimum statistic is equal to the negative reciprocal of the mathematical expectation of the second partial derivative of the logarithm of the likelihood with respect to the parameter in question.

We may write, therefore

$$-\frac{1}{\sigma_{m}^{2}} = \frac{-4 \pi \rho \Gamma(\rho + \frac{3}{2})}{a \sqrt{\pi} \Gamma(\rho + 1)} \int_{a}^{a} \frac{(a^{2} + x^{2})}{(a^{2} - x^{2})} \left[1 - \frac{x^{2}}{a^{2}}\right]^{\rho} dx$$

 $-\frac{1}{\sigma_{m}^{2}} = \frac{-4 \, n\rho \, \Gamma\left(\rho + \frac{3}{2}\right)}{a \sqrt{\pi} \, \Gamma\left(\rho + 1\right)} \int_{0}^{\frac{\pi}{2}} (2 \cos^{2\rho - 3} - \cos^{2\rho - 1} \theta) d\theta$

and so again referring to (2) page 87 we have

$$-\frac{1}{\sigma_{\hat{m}}^2} = -\frac{2\pi\rho\left(\rho + \frac{1}{2}\right)}{a^2(\rho - 1)}$$

hence we have

(8)
$$\sigma_{\widehat{m}}^2 = \frac{a^2(p-1)}{np(2p+1)}$$

The efficiency of the mean is then

(9)
$$E = \frac{(p-1)(2p+3)}{p(2p+1)} = \frac{\sigma_{2}^{2}}{\sigma_{z}^{2}}$$

The next problem with which we must concern ourselves is that of experimental verification of the assumptions under discus-

¹Ibid Pp. 327-328

sion. The problem is briefly that of choosing a number of small samples from a population which obeys our law of frequency, estimating m for these samples, and then calculating the variance for the distribution. Also we shall draw an histogram of the distribution and observe the general type of the distribution, so nearly as that is possible from our samples.

The problem of choosing our samples is not the least of our difficulties, for we can not take all types of samples. They must be of a very special nature: they must be from a population of the Type II. In order to accomplish this it will be necessary to have a table of areas corresponding to given values of x for the Type II Curve. There is no such table available to the knowledge of the writer, and it is therefore necessary to construct the table before we can procede with the choosing of the samples. After the table has been built, we can with the aid of Tippett's Tables,1 choose our samples with ease. The manner of choosing is as fol-Take the numbers from Tippett's Tables as areas under the Type II Curve and look up in the table of areas the values of x corresponding to the smallest area containing the area found from the Random Numbers. This will give the value of x to be taken. Since we will take four digits let the fifth digit determine the sign. If it is odd take the sign - , if it is even take the sign + .

There are two ways in which a table of this nature can be prepared, and the method employed must in any case be determined by the degree of accuracy attainable and the amount of labor involved. One of these methods is that of the calculus of finite differences, determining the zero order differences by means of algebra, and from these by the process of addition building up the table. This method is best used when a dependable listing adding machine is at hand. The other method is that of direct integration, and it is found that with the aid of two calculating machines this is by far the quicker. It was this method that was applied in the building of the table on page 92 and 93.

¹Tracts For Computers, No. XV.

Table of Areas Under Pearson's Type II Curve, Correct to 9 Places of Decimals. The Areas are included between ordinates located $\pm \varkappa$ units from the parameter of location.

Constants. $y_0 = \frac{15}{16}a$, m = 0, p = 2, and a = 1

x	Area		x	Area
0.00	0.000 000 000		30	529 661 250
01	018 <i>7</i> 48 <i>7</i> 50		31	545 084 843
02	037 490 001		32	560 298 291
03	056 212 259		33	575 295 077
04	074 920 038		34	590 073 828
05	093 593 867		35	604 625 820
06	112 230 292		36	618 947 482
07	130 821 880	1	37	633 034 148
08	149 361 229	1	38	646 881 319
09	167 840 964		.39	660 484 657
10	186 253 750	L	40	673 840 000
11	204 592 289	İ	41	686 943 358
12	222 849 331		42	699 790 921
13	241 017 673		43	712 379 067
14	259 090 168	l	44	724 704 358
15	<i>277</i> 059 <i>727</i>	1	45	736 763 555
16	294 919 322	l	46	748 553 612
17	312 661 995		47	<i>7</i> 60-071 688
18	330 280 859		48	773 151 488
19	347 769 104		49	782 281 572
<i>2</i> 0	3 65 120 040	L	50	792 968 750
21	382 326 904		. 51	803 374 696
22	399 383 261		52	813 497 651
23	416 282 613		53	823 336 081
24	433 018 598		54	832 888 688
25	449 584 961		55	842 154 414
26	465 975 552	1	56	851 132 442
27	482 184 334		57	860 009 702
28	498 205 389		58	868 223 379
29	514 032 918		59	876 335 911
30	529 661 250		60	884 160 000

61	892 546 290	81	985 203 165
62	898 944 981	82	987 317 441
63	905 907 620	83	989 230 274
64	912 585 318	84	990 949 478
65	919 120 273	85	992 483 242
66	925 092 472	86	993 840 132
67	930 925 941	87	995 029 095
68	936 482 509	88	996 059 469
69	941 764 926	89	996 940 979
70	946 776 250	90	997 683 750
71	951 519 851	91	998 298 304
72	955 999 411	92	998 795 571
73	960 294 310	93	999 186 888
74	964 182 748	94	999 484 008
75	967 895 508	95	999 699 102
76	971 362 202	96	999 844 762
77	974 588 156	97	999 934 010
78	977 579 039	98	999 980 299
79	980 340 865	99	999 997 519
80	982 880 000	100	1.000 000 000

The table on the preceding page was built by direct integration of (3) page 88, with p=2 and a=1. These values for the parameters were chosen so as to save as much labor in calculation as possible, and at the same time maintain the desired shape of the curve. There has of course been no less of generality in setting a=1 but we have limited ourselves quite definitely in using the value p=2. The accuracy of the table to 9 places of decimals has been assured by calculating all values to 13 places and then determining the maximum error over the whole range which was 625 in the 13 place due to the use of the decimal equivalent of $\frac{2}{3}$ in the third degree term.

Our table of areas being complete and the problem of sampling thus solved, we must next consider the task of estimating \hat{m} . Since we have already taken a=1 and p=2 in building our tables we shall continue to use these values throughout the work.

The next question to be settled before beginning our work is the size of the samples with which we are to deal. The case $\pi = 1$ is of course of no interest for the best estimate of a single observation is the observation itself. The case $\pi = 2$ is likewise of no interest for in this case the arithmetic mean coincides with the solution by the method of maximum likelihood. Therefore it is the case $\pi = 3$ with which we shall be concerned. Our results are not then trivial, and at the same time we have the case which is the easiest to deal with, since the number of numerical calculations for each sample is reduced to the minimum.

Before going ahead with any attempt at solving an equation such as (5) page 83, it is well to have in mind a picture of what we are actually trying to accomplish. With such a picture in mind we are better able to realize the difficulties of the situation and so are better able to cope with them. To this end we have included a graph of the problem involved which will make clear at a glance just what must be done to find the true value of 277.

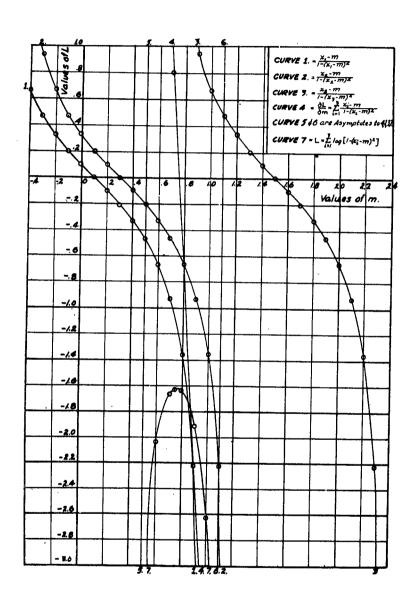
We have drawn, page 96, from plotted points the curves representing L, $\frac{\partial L}{\partial m}$ and the three terms of $\frac{\partial L}{\partial m}$. Also we have drawn in the asymptotes to the curve which are of significance. If we are able to find the point at which $\frac{\partial L}{\partial m} = 0$ we have the solution to our problem. This involves solving a fifth degree equation. We shall use Newton's method of successive approximation. Fisher states that in some cases at least, we may start with an inefficient statistic and by a single approximation obtain an efficient one. Whether or not this is the case for small samples we shall see when our calculations have been analyzed.

It will be seen upon examining the graph that as is allowed to vary each of the therms of $\frac{\partial L}{\partial m}$ varies from $-\infty$ to $+\infty$,

¹Theory of Statistical Estimation, R. A. Fisher, Proc. Cam. Phil. Soc. Vol. XXII part 5. Pp. 708-709.

and so their sum also varies between these limits. It will be seen that the asymptote corresponding to the largest allowable value of m can be found by adding the value of a to the smallest observation, and that the asymptote corresponding to the smallest allowable value of m may be found by subtracting the value of a from the greatest observation. It is thus evident that as dispersion in the sample increases the variance of m must surely decrease. That this fact is of fundamental importance in choosing our first estimate of m will be seen from consideration of the following case, since it is well known that in the case of a curve with a real finite pole Newton's method may lead us to erroneous results. Let us consider the sample consisting of the three observations $x_2 = + .99$, $x_3 = + .99$, for which the arithmetic mean is $\bar{z} = + .33$. If now we take as our first approximation $m_i = \bar{z}$ we will immediately lead ourselves to an erroneous value of m_2 , our estimate of \hat{m} . That this is the case will be easily seen by means of the check given on page 94 for using a=1 we locate the asymptotes at $z = \pm$.01. The true value of m located between these asymptotes, is not therefore even to be approached should we take $m_1 = \bar{z}$. This is an extreme case and fortunately not to be anticipated very often, or the arithmetic mean would be deprived of any value whatsoever. We should always be sure that the difference between any observation and the value of m that we are using is not greater than the value of a when dealing with the Type II Curve. This holds no matter what our manner of attack may be.

We shall now be concerned with the calculations for our 100 samples of 3. All of the data are tabulated in such a way as to be self-explanatory, and so we shall not bother to give sample calculations. The next ten pages cover these calculations. The discussion is continued on page 104.



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Sa	am-		$m_i = \tilde{\chi}$		<u> </u>	/+(x;-m)2	$\frac{\partial^2 L}{\partial m^2}$ $\frac{3}{4} \frac{1+(x_i-m)^4}{2}$	m_2
N	1.2	$ x_i $	13 xi	1-(x;-m)	$\frac{3}{1-1}\frac{x_1-m}{1-(x_1-m)}$	$\left[\left(-\frac{(x_i-m)^2}{2}\right)^2\right]$		$\frac{\partial^2 L}{\partial m^2} \frac{\partial L}{\partial m}$
1	1 2 3	+.17 23 15	0700	+ .254667 164200 080520	+ .009840	- 1.120600 - 1.051500 - 1.01,2800	- 3.184900	066879
-2	1 2 3	03 15 +.34	+.0533	083882 212064 + .312676	+.016730	- 1.021059 - 1.133055 - 1.284716	- 3.438830	+.049737
3	1 2 3	+.67 11 +.73	+.4300	+ .254668. 762281 + .329670	177943	- 1.190833 - 2.573777 - 1.316266	- 5.080876	+.465022
4	1 2 3	54 33 +.06	2700	291230 060216 + .370328	÷.018882	- 1,248262 -1.010865 -1.396495	- 3.655622	264835
5	1 2 3	41 +.11 +.32	+.0067	481688 + .104445 + .347440	029803	-1.718538 -1.032609 -1.350292	- 4.101439	+.013936
6	1 2 3	26 +.12 +.42	+.0967	408701 +.023312 +.361036	024353	-1.479856 -1.001630 -1.377417	- 3.858903	+.103011
7	1 2 3	+.54 50 69	2167	41.770451 308021 609932	+.853498	-4.424677 -4.277016 -1.277016	- 9.978709	131268
8	1 2 3	+.19 64 66	3700	+.815850 291230 316628	+.207992	-2.218101 -1.248262 -1.292329	- 4.758692	326292
9	1 2 3	12 48 73	4433	+ .361036 036749 312376	+ .011911	- 1.377433 - 1.004049 - 1.284716	- 3.666198	440051
10	1 2 3	- 49 77 70	6533	+ .167774 118311 046802	+ .002661	- 1.083693 - 1.041802 - 1.006566	- 3.132061	652450
11	1 2 3	07 +.04 38	1367	+ .066896 + .182285 258727	009546	- 1.013446 - 1.098764 - 1.196675	- 3.308885	139555
12	1 2 3	40 72 46	5267	+ .128735 200836 + .066967	005134	- 1.049476 - 1.119217 - 1.013434	- 3.182127	528283
13	1 2 3	30 65 +.16	2633	036749 454723 + .104414	387058	1.004049 -1.589322 -1.750185	-4.343556	352411
14	1 2 3	14 50 16	2667	+ .128767 246729 + .107928		-1.049476 -1.179312 -1.034813	- 3.263601	2 6977 5
			-	-				

	·			DISTRI	BUTTON (OF MEANS		
e.	am-		$m_i = \bar{x}$		<u>dL</u>		$\frac{\partial^2 L}{\partial m^2}$	m ₂
pl N	٩.	$ x_i $	$\frac{13}{3}\sum_{i=1}^{3}x_{i}$	$\frac{x_i - m}{1 - (x_i - m)^2}$	$\sum_{i=1}^{3} \frac{x_i - m}{1 - (x_i - m)}$	$\frac{1+(x_i-m)^2}{[1-(x_i-m)^2]^2}$	$\sum_{i=1}^{3} \frac{1+(x_i-m)^2}{[1-(x_i-m)^2]}$	$\frac{\partial^2 L}{\partial m^2} \frac{\partial L}{\partial m} = \frac{\partial L}{\partial m}$
15	1 2 3	09 20 43	2367	+ .149926 + .036749 183962	+ .002713	- 1.066950 - 1.004049 - 1.106718	- 3.177717	235846
16	1 2 3	29 +.12 +.27	+.0333	360623 + .087663 + .251104	021856	- 1.376579 - 1.022996 - 1.185618	- 3.585193	+.039096
17	1 2 3	10 39 23	2400	+ .142798 153452 + .010001	000653	- 1.060774 -1.070113 - 1.000300	- 3.131187	240209
18	1 2 3	+.26 42 17	1100	+ .428687 342958 060216	+ .025513	- 1.526159 - 1.341557 - 1.010865	- 3.878581	103422
19	1 2 3	+ .23 21 43	1367	+ .423670 073695 320905	+ .029076	- 1.514353 - 1.016264 - 1.300082	- 3 .830699	12 9111
20	1 2 3	+ .37 09 + .67	+.3167	+ .053552 483174 + .403835	025787	- 1.008563 - 1.673112 - 1.468552	- 4.150227	+ .322813
21	1 2 3	+ .21 76 + .66	+.0367	+ .178665 -2.181131 +1.019301	983165	- 1.094805 -12.252377 - 3.713282	-17.060464	094328
22	1 2 3	20 + .23 22	0633	139303 + .320905 160644	+ .020958	- 1.057853 - 1.300082 - 1.076786	- 3.434721	057198
23	1 2 3	44 62 +.16	3000	142798 356506 + .583460	+ .084156	- 1.060744 - 1.368375 - 1.949243	- 4.378292	280779
24	1 2 3	28 + .20 + .45	+.1233	481639 + .077153 + .365736	038750	-1.658197 -1.017823 -1.387011	- 4.063031	132837
25	1 2 3	13 56 02	2367	+ .107928 361036 + .408551	+ .155433	-1.034813 -1.377417 -1.152673	- 3.564903	193096
26		+.10 +.14 27	0100	+ .111347 + .153452 278850	014051	-1.024632 -1.070113 -1.228016	- 3.322761	01422 ⁹
27	1 2 3	79 09 08	3200	603260 + .242846 + .254668	105746	-2.011377 -1.142660 -1.190833	- 4.344870	344338
28		08 +.43 +.18	+.1760	273953 + .271517 + .004000	+ .001564	-1.028078 -1.216408 -1.000048	- 3.244534	175518

			777 ₁ = Z		<u>dL</u>		$\frac{\partial^2 L}{\partial m^2}$	m ₂
Sar ple No.	n- i	x _i	$\frac{\frac{1}{3}\sum_{i=1}^{9}x_{i}}{3\sum_{i=1}^{9}x_{i}}$	$\frac{\chi_i - m}{1 - (\chi_i - m)^2}$	\(\frac{\chi_{i=1}}{\infty} \frac{\chi_{i} - m}{(\chi_{i} - m)^2} \)	+(x;-m)2 [/-(x;-m)2]2	$\sum_{i=1}^{3} \frac{1 + (x_i - m)^2}{[1 - (x_i - m)^2]}$	021 m, 01 0mm m, 0m 021 /0m2
29	1 2 3	‡.50 59 14	0767	+ .864077 696923 063554	+.103600	- 2.991572 - 2.329134 - 1.012101	- 6.332807	060341
30	1 2 3	+.02 +.17 +.46	+.2160	- ,203830 - ,046097 + ,259446	+ .009519	- 1.123044 - 1.006379 - 1.197929	- 3.327343	4.213139
31	1 2 3	35 18 35	2930	057185 + .114461 057185	+.000091	- 1.009800 - 1.039137 - 1.009800	- 3.058737	292970
32	1 2 3	+.29 16 16	0100	+ .329670 153452 153452	+ .022766	- 1.316266 - 1.070113 - 1.070113	- 3.456492	003414
33	1 2 3	45 60 35	4667	+ .017004 135395 + .118623	000222	- 1.000867 - 1.054671 - 1.042022	- 3.097560	463667
34	1 2 3	23 +.21 20	0733	160644 + .308021 128767	+.018610	- 1.076786 - 1.057853 - 1.049476	- 3.184115	067455
35	1 2 3	43 05 +.33	0500	444132 + .000000 + .444132	+.000000.+	± 1.563278 = 0.000000 = 1.563278	- 2.832959	050000
36	1 2 3	62 59 +.80	1367	630593 570533 +7.640723	+6.439597	- 2.100061 - 1.909640 - 124.918362	-1 2 8.9280 63	086753
37	1 2 3	44 +.50 12	0200	509956 + .712719 101010	+.101753	- 1.734292 - 2.386551 - 1.030507	- 5.151350	000247
38	1 2 3	14 +.49 +.09	+.1467	312376 + .387806 056882	+.018548	- 1.284716 - 1.436499 - 1.009696	- 3,730911	+130837
39	1 2 3	32 +.74 +.15	+.1900	689282 + .788530 040064	+ .059184	- 2.301754 - 2.677252 - 1.004812	- 5.983818	179792
40	1 2 3	17 +.08 46	1833	+ .013002 + .285611 300020	001407	- 1.000507 - 1.233970 - 1.263129	- 3.497606	-183402
41	1 2 3	30 +.07 +.22	0033	325339 + .073695 + .235018	016626	- 1.308220 - 1.016264 - 1.162947	- 3.487431	008067
42	1 2 3	+.01 10 37	1533	+ .167774 + .053451 227577	006152	- 1.909102 - 1.008563 - 1.153673	- 4.070338	154811

Saı			m,=x		<u>dL</u>		$\frac{\partial^2 L}{\partial m^2}$	m ₂
ple No	Li	x_i	$\int_{\frac{3}{3}}^{\frac{3}{2}} \sum_{i=1}^{\infty} x_{i}$	$\frac{\chi_i - m}{1 - (\chi_i - m)^2}$	$\sum_{i=1}^{3} \frac{x_i - m}{1 - (x_i - m)^2}$	$\frac{1+(x_{i}-m)^{2}}{[1-(x_{i}-m)^{2}]^{2}}$	$\sum_{i=1}^{3} \frac{1+(x_i-m)^2}{(1-(x_i-m)^2)^2}$	34 m de
43	1 2 3	+.27 +.03 28	+ .0067	+ .283036 + .023412 295486	+ .010962	- 1.234771 - 1.001644 - 1.313099	- 3.549514	+.003512
44	1 2 3	28 13 22	2100	070344 + .080515 010001	+ .000170	- 1.014820 - 1.019406 - 1.000300	- 3.034526	-209944
45	1 2 3	09 +.19 +.06	₹.053 3	146304 + .139303 + .006700	+ .000301	- 1.063775 - 1.057853 - 1.00013#	- 3.121762	4 053396
46	1 2 3	08 28 19	1833	+ .104414 097612 006700	+ .000102 ·	- 1.032590 - 1.028495 - 1.000134	- 3.061219	183267
47	1 2 3	10 70 11	3033	+ .212064 470788 + .200802	057922	- 1.133055 - 1.630045 - 1.119458	- 3.882558	-318219
48	1 2 3	+.67 +.07 33	∔.1367	+ .745508 066896 596458	+ .082154	- 2.509221 - 1.013405 - 1.989832	- 5.512458	÷.121697
49	1 2 3	- 28 - 15 - 18	2033	077153 +.053466 +.023312	000381	- 1.017823 - 1.008563 - 1.001630	- 3.028016	203426
50	1 2 3	33 +.45 23	0367	320905 + .637773 200802	+ .116066	- 1.300082 - 2.123915 - 1.119458	- 4.5434 5 5	-0111 54
51	1 2 3	+.44 79 08	1433	+.883184 -1.112848 +.063251	166413 ·	- 3.074927 - 4.196875 - 2.019971	~ 9.291723	-160910
52	1 2 3	15 +.36 63	1400	010001 +.666667 644821	+ .011845	-1.000300 -2.222222 -2.147552	- 5.270074	-137752
53	1 2 3	+.09 +.31 +.47	+.2900	208333 +.020008 +.186027	014143	- 1.128472 - 1.001200 - 1.102697	- 3,232369	±290044
54	1 2 3	+.21 29 +.37	+.0967	+.114773 454693 +.295361	- .0445 59	- 1.039349 - 1.589382 - 1.255198	- 3.883929	÷108173
55	1 2 3	+,39 -,21 +,45	+.2100	+.186027 509956 +.254668	069261	-1.102697 -2.600554 -1.190833	- 4.894084	t224152
56	1 2 3	30 22 67	3967	+.097612 +.182394 295361	015355	-1.028495 -1.098764 -1.255198	- 3.382457	. 401 240

				JOH	N L. CAF	CLSON		101
			777, = Z		<u>dL</u> dm		$\frac{\partial^2 L}{\partial m^2}$	me
Sa ple No	i	x _i	$\int_{0}^{1} \sum_{i=1}^{3} x_{i}$	$\frac{\chi_i - m}{1 - (\chi_i - m)^2}$	$\sum_{i=1}^{3} \frac{x_i - m}{1 - (x_i - m)^2}$	$\frac{1+(x_i-m)^2}{[1-(x_i-m)^2]^2}$	$\sum_{i=1}^{3} \frac{1 + (x_i - m)^2}{1 - (x_i - m)^2}$	$\frac{\partial^2 L}{\partial m^2} \frac{\partial L}{\partial m^2}$
57	1 2 3	41 71 +.78	1133	325339 926626 +4.418971	810074	- 1.324459 - 3.371716 -44.057350	- 48.783525	129916
58	1 2 3	+ .04 69 + .04	20 33	+ .258608 637773 + .258608	120557	- 1.196675 - 2.123915 - 1.196675	- 4.517265	229988
59	1 2 3	+.43 48 01	0200	+ .564263 583460 + .010001	009196	- 1.890704 - 1.949243 - 1.000300	- 4.840247	021900
60	1 2 3	+.69 17 09	+ .1433	+ .779753 347399 246729	+ .185625	- 2.643417 - 1.350213 - 1.179312	- 5.172942	+.107416
61	1 2 3	+.12 +.09 33	0400 ,	+ .164203 + .132234 316628	020191	- 1.080198 - 1.052162 - 1.183644	- 3.316004	046089
62	1 2 3	+.63 +.36 +.05	+ .3467	+ .308021 + .013302 325339	004016	- 1.277016 - 1.000530 - 1.308220	- 3.585766	+.347820
63	1 2 3	00 40 +.69	+ .0967	097509 659155 + .915915	+ .163267	- 1.028435 - 1.654673 - 2.087008	- 4.770116	+.062373
64	1 2 3	+.13 69 +.73	+ .0567	+ .073695 -1.057383 +1.231645	+ .247951	- 1.016264 - 7.938848 - 4.863167	- 13.818279	+.038756
65	1 2 3	39 13 44	3200	07.0344 + .197115` 121753	+ ,005018	- 1.014820 - 1.115161 - 1.044258	- 3.174239	318419
66	1 2 3	12 +.51 +.29	+.2267	394067 + .308021 + .063554	022492	- 1.447201 - 1.255198 - 1.012101	- 3.714500	+.232755
67	1 2 3	16 +.22 +.02	+.0267	193442 + .200802 006700	+.000660	- 1.110956 - 1.119458 - 1.000134	- 3.230548	+.026496
68	1 2 3	51 -:44 20	3833	129081 057185 +.189340	+.09 3074	- 1.049717 - 1.009800 1.106349	- 3.165866	-382029
69	1 2 3	+.35 +.43 +.16	+3133	+ .036749 + .118311 156989	001929	- 1.004049 - 1.041802 - 1.073357	- 3.119208	+ 321050
70	1 2 3	+.59 +.03 +.14	+.2533	+ .379751 235018 114773	÷.029960	- 1.416284 - 1.162947 - 1.039349	- 3.618580	+245021

Sam			$m_i = \bar{\chi}$		ðL		ð2/L	
			•		δm		<u>∂²L</u> ∂m²	m_2
ple .	- 1		$\frac{1}{3}\sum_{i=1}^{3}\chi_{i}$	xi-m	3 ×i-m	1+(x;-11)2	$\frac{3}{5} \frac{1+(x_1-m)^2}{1+(x_1-m)^2}$	Om m. Om
No	i	x_i	اءًا ا	l-(x;-m)2	=1 /-(xi-m)2	$[I-(x_i-m)^2]^2$	[=1[1-(x;-m)2]	021/0m2
71 2		19 38 12	2300	+ .040064 153452 + .111347	002041	- 1.004812 - 1.070113 - 1.037044	- 3.111969	230656
72 2		45 16 +.11	1667	307638 + .007000 + .300020	000618	- 1.276344 - 1.000147 - 1.263129	- 3.539620	167175
73 2		+.71 48 07	∔ .053 3	+1.155256 745580 125203	+1284473	- 4.424677 - 2.508265 - 1.046789	- 7.979731	4.017651
74	1 2 3	33 +.15 17	1167	223467 + .287122 053451	+ .010204	- 1.147540 - 1.251454 - 1.008563	- 3.407557	113705
75	1 2 3	19 39 +.23	1167	073695 295361 + .394067	+ .025011	- 1.016264 - 1.255198 - 1.447201	- 2.718663	109974
76	1 2 3	01 73 27	3367	+ .365736 465270 + .066998	057547	- 1.387011 - 1.615943 - 1.013446	- 4.016400	351028
77	1 2 3	63 09 +.53	0683	820607 021710 + .931877	+ .089560	- 2.707683 - 1.001413 - 3.294334	- 7.103430	055692
78	1 2 3	04 +.47 61	0600	+ .020008 + .737032 815850	058810	- 1.001200 - 2.477060 - 2.788101	- 6.266361	069385
79	1 2 3	+.39 +.38 +.10	+.1900	+ .208333 + .197115 090734	+ .314714	- 1.128472 - 1.115161 - 1.024631	- 3.268264	+.093706
	1 2 3	77 16 +.15	2600	689282 + .101010 + .492847	095425	- 2.301754 - 1.030507 - 1.687865	- 5.020126	279008
	1 2 3	03 01 64	2267	+ .204504 + .227262 498612	066846	- 1.123849 - 1.152521 - 1.703355	- 3.979725	243397
82	1 2 3	+.21 +.71 +.49	+ .4700	278850 + .254668 + .020008	004174	- 1.228016 - 1.190833 - 1.001200	- 3.420049	+.364611
83	1 2 3	26 28 08	2067	053451 073695 + .128767	+ .001621	- 1.008563 - 1.016264 - 1.049476	- 3.074303	206173
	1 2 3	16 +.23 35	≟ .0933	066998 + .361036 274808	+ .019230	- 1.013446 - 1.057853 - 1.221582	- 3.292881	087460
	1 2 3	31 +.06 +.52	+.0900	476190 030027 + .527542	+ .021325	-1.643990 -1.002704 -1.783445	- 4,430139	+.085186

	JOHN L, CARLSON 103								
			777, =X		<u>dL</u> dm		<u>31</u> 0m²	7772	
Sai ple No	i	x_i	<u> </u>	$\frac{x_i-m}{1-(x_i-m)^2}$	3 <u>zi-m</u> i=1 1-(zi-m)*	$[-(x_i-m)^2]^2$		<u>34</u> 3m2 ^m , 3m 32L/3m2	
86	1 2 3	+.36 +.36 35	+.1233	+ .250748 + .250748 609932	108436	- 1.185102 - 1.185102 - 1.162837	- 3.533041	+.153992	
87	1 2 3	+.56 60 43	1567	+1.473657 551721 295361	+ .626575	- 6.399500 - 1.853371 - 1.255198	- 9.508069	090801	
88	1 2 3	45 29 +.19	1833	287122 107928 + .433743	+ .038693	- 1.241454 - 1.034813 - 1.538183	- 3.814450	173156	
89	1 2 3	09 39 +.24	0800	010001 341851 + .357620	+ .005768	- 1.000300 - 1.341559 - 1.368275	- 3.710134	078445	
90	1 2 3	09 44 22	2500	+ .164203 197115 + .030027	002885	- 1.080198 - 1.115161 - 1.002704	- 3.198063	-,250902	
91	1 2 3	+.24 +.06 06	+.0800	+ .164203 020008 142798	+ .001397	- 1.080198 - 1.001200 - 1.060774	- 3.142172	±079555	
92	1 2 3	+.05 +.12 43	0867	+ .139303 + .215139 389164	034122	- 1.057853 - 1.137879 - 1.436499	- 3.632231	-,096094	
93	1 2 3	10 +.03 +.32	+.0833	189672 053451 + .250748	+ .007625	- 1.106718 - 1.008563 - 1.185102	- 3.300383	+.080990	
94	1 2 3	08 +.26 +.19	+.1233	212064 + .139303 + .066998	005763	- 1.133055 - 1.057853 - 1.013446	- 3.204354	+.125098	
95	1 2 3	40 +.30 57	2233	183494 + .720642 394067	+ .143083	- 1.098764 - 2.415765 - 1.447201	- 4.961730	-,194463	
96	1 2 3	+.01 13 09	0700	+ .080515 060216 020008	+ .000291	- 1.019406 - 1.010865 - 1.001200	- 3.031471	069904	
97	1 2 3	04 +.28 11	+.0433	083882 + .250748 156989	+ .009877	- 1.021059 - 1.185102 - 1.073357	- 3.279518	+.040288	
98	1 2 3	-,21 -,18 -,53	3067	+ .097612 + .128767 235018	018516	- 1.028495 - 1.049476 - 1.162947	- 3.240918	312413	
99	1 2 3	02 12 07	0700	+ .050125 050125 + .000000	+.000000	- 1.007531 - 1.007531 - 0.000000	- 2.015062	070000	
100	1 2 3	09 06 03	0600	030027 +.000000 +.030027	+ .000000	- 1.032758 - 0.000000 -1.032758	- 2.065516	060000	

In order to analyze the results of the calculations on the preceding ten pages it is necessary that we find the theoretical variance for \mathbb{Z} and \widehat{m} from the formulae derived for this purpose lier in this paper.

From (7) page 90, we get after setting a=1, $\rho=2$, and n=3,

(10)
$$\sigma_{\overline{z}}^2 = \frac{1}{2I} = .047619$$

and from (8) page 90 after similar substitutions

(11)
$$\sigma_{\widehat{m}}^2 = \frac{1}{30} = .033333$$

From (9) page 90 we get for the efficiency of the mean when

(12)
$$\rho = 2 E = .70$$

Now by actual calculation from the ungrouped data the mean square deviation from zero which we shall designate as

$$\frac{\Sigma \bar{z}^2}{N} = .048611$$

and

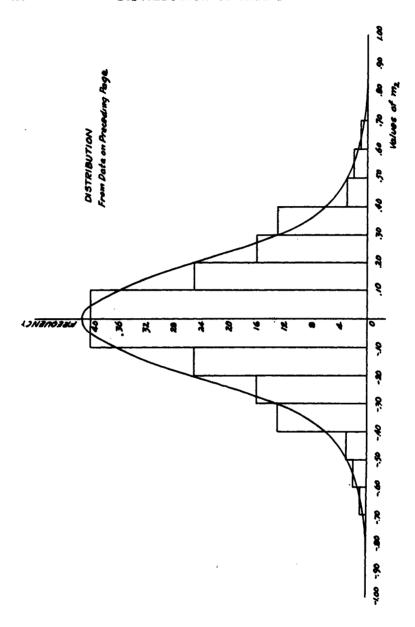
(14)
$$\frac{\sum m_2^2}{N} = .047612$$

Comparing (10) and (13) it is evident that such a difference can be said to be well within the limits of random sampling. The difference between (11) and (14) is of such magnitude that we can not say that it might be expected in the course of random sampling. Now since m_2 is our estimate of \hat{m} it is evident that either a single approximation by Newton's method is not adequate to give the best results or the approximation,

$$\sigma_{\widehat{m}}^2 = \frac{\alpha^2(p-1)}{n\rho(2p+1)} ,$$

to the variance is not valid in the case of small samples. That the latter seems to be the case the writer firmly believes. The reason for this belief lies in the fact that in a subsequent case a sample of 3 was examined by Newton's method, and starting with m, -.07 the values $m_2 = -.066879$ and $m_3 = -.066791$ were obtained. There is not sufficient improvement here to cause one to suppose that by taking a third approximation we would obtain a variance in keeping with the one derived from theory. Also considering the mean square deviation from zero for \overline{z} and m_z it seems that the gain in accuracy to be expected from the use of the method of maximum likelihood solution instead of the arithmetic mean is not sufficiently great, in the case of samples of three, to warrant the additional labor involved in calculation. We must be sure, however, that in using the arithmetic mean we are using an approximation to m which complies with the qualifications given on pages 95 and 96. A graph of the distribution as found from the calculations is given on page 106. The histogram represents the grouped data while the smoothed curve is a rough approximation to the actual form of the distribution.

Histogram Data										
	m ₂ + - Totals									
.01 .11 .21 .31 .41 .51 .61 .71	0.00 10 .20 .30 .40 .50 .60 .70 .80 .90	2 16 8 4 4 1 1 0 0 0	2 20 17 12 9 2 1 1 0 0	0 40 25 16 13 3 2 1 0						
		36	64	100						



The totals have been used in the histogram which has been forced to be symmetrical so as to give the effect of a sample of twice the size.

The tabulation of the histogram data draws attention to the great excess of samples having negative values for \overline{z} and m_2 . This has caused the writer no little concern. In examining the signs of the observations we note that there are 183 – and only 117 + values. Assuming + and – values to be equally likely this gives a deviation from 150 of 33 or 3.81 times its standard error which is incredible. It seems therefore that Tippett's numbers are not random in this respect, and that it perhaps would have been better to toss a coin to determine the signs.

As a final check and in an effort to place the type of the distribution the value of β_z has been calculated, and found to be $\beta_z=3.056$

 \mathcal{S} , was not calculated as the excess of negative signs would lead to an erroneous value. There is every reason to believe that \mathcal{S} , should be zero. These facts suggest that the curve is very near to the normal curve, but perhaps slightly more leptokurtic. But, why, if this is the case, there is not better agreement between (11) and (14) page 104 the writer is unable at the present to state.

Johnsbarlson