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Antipsychotic Use in Pregnancy and the Risk for Congenital Malformations

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IMPORTANCE The frequency of antipsychotic (AP) use during pregnancy has approximately doubled during the last decade. However, little is known about their safety for the developing fetus, and concerns have been raised about a potential association with congenital malformations.

OBJECTIVE To examine the risk for congenital malformations overall and cardiac malformations associated with first-trimester exposure to APs.

DESIGN, SETTING, AND PARTICIPANTS This nationwide sample of 1360 101 pregnant women enrolled in Medicaid with a live-born infant constituted the pregnancy cohort nested in the Medicaid Analytic Extract database, which included data from January 1, 2000, to December 31, 2010. Participants were enrolled in Medicaid from 3 months before their last menstrual period through at least 1 month after delivery. Relative risks (RRs) were estimated using generalized linear models with fine stratification on the propensity score to control for the underlying psychiatric disorders and other potential confounders. Data were analyzed during 2015.

EXPOSURES Use of APs during the first trimester, the etiologically relevant period for organogenesis.

MAIN OUTCOMES AND MEASURES Major congenital malformations overall and cardiac malformations identified during the first 90 days after delivery.

RESULTS Of the 1341715 pregnancies that met inclusion criteria (mean [SD] age of women, 24.02 [5.77] years), 9258 (0.69%) filled at least 1 prescription for an atypical AP and 733 (0.05%) filled at least 1 prescription for a typical AP during the first trimester. Overall, 32.7 (95% CI, 32.4-33.0) per 1000 births not exposed to APs were diagnosed with congenital malformations compared with 44.5 (95% CI, 40.5-48.9) per 1000 births exposed to atypical and 38.2 (95% CI, 26.6-54.7) per 1000 births exposed to typical APs. Unadjusted analyses suggested an increased risk for malformations overall for atypical APs (RR, 1.36; 95% CI, 1.24-1.50) but not for typical APs (RR, 1.17; 95% CI, 0.81-1.68). After confounding adjustment, the RR was reduced to 1.05 (95% CI, 0.96-1.16) for atypical APs and 0.90 (95% CI, 0.62-1.31) for typical APs. The findings for cardiac malformations were similar. For the individual agents examined, a small increased risk in overall malformations (RR, 1.26; 95% CI, 1.02-1.56) and cardiac malformations (RR, 1.26; 95% CI, 0.88-1.81) was found for risperidone that was independent of measured confounders.

CONCLUSIONS AND RELEVANCE Evidence from this large study suggests that use of APs early in pregnancy generally does not meaningfully increase the risk for congenital malformations overall or cardiac malformations in particular. The small increase in the risk for malformations observed with risperidone requires additional study.

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Corresponding Author: Krista F. Huybrechts, MS, PhD, Division of Pharmacoepidemiology and Pharmacoeconomics, Department of Medicine, Brigham and Women's Hospital and Harvard Medical School, 1620 Tremont St, Ste 3030, Boston, MA 02120 (khuybrechts@bwh .harvard.edu). xposure to antipsychotics (APs) during pregnancy is increasingly common. Most newer atypical drugs are less likely to affect fertility than the older typical APs. This unimpaired fecundity combined with deinstitutionalization of patients with psychiatric illness and more widespread offlabel use of these drugs have resulted in a doubling in the use of APs during pregnancy in the last decade.^{1,2}

Since the first systematic review on this topic,³ clinicians continue to have very little information regarding the safety of these drugs for the developing fetus despite the growing use of atypical APs. Registry studies have been small (<250 exposed women) and have yielded mixed results; one study⁴ reported higher than expected rates of malformations (particularly cardiac), whereas another study⁵ did not find this association. Results from epidemiologic studies have also yielded inconsistent evidence. The 2 largest epidemiologic studies available to datewhich included 570 and 561 exposed women^{6,7}-reported a 1.5to 2-fold increase in the risk for congenital malformations, in particular atrial and ventricular septal defects. Another study based on 151 exposed women¹ suggested that atypical APs are not associated with an increased risk for major malformations. However, this study was underpowered to detect a less than 5-fold increase in risk. Given the need for additional information to enable psychiatrists and their patients to weigh the risks of using these medications during early pregnancy, our objective was to examine the risk for congenital malformations, and in particular cardiac malformations, associated with firsttrimester exposure to different APs in a large cohort.

Methods

Data Source and Study Cohort

We used the pregnancy cohort nested in the nationwide Medicaid Analytic Extract database, which included data from January 1, 2000, to December 31, 2010. The Medicaid Analytic Extract data include demographic and insurance enrollment information, medical visits and hospitalizations, inpatient and outpatient diagnoses and procedures, and prescriptions filled on an outpatient basis. The development of this cohort has been described in detail elsewhere.8 The cohort consisted of all pregnancies that resulted in live births for which Medicaid covered the health care expenses. Women and girls aged 12 to 55 years (hereinafter referred to as women) were required to have coverage through Medicaid from 3 months before the date of their last menstrual period (start of pregnancy) to 1 month after delivery. Infants were required to have coverage through Medicaid for the first 3 months of life unless they died sooner. Pregnancies with exposure to a known teratogenic medication during the first trimester (n = 8246) and pregnancies with a chromosomal abnormality (n = 2550) were excluded. This study was approved by the institutional review board of Brigham and Women's Hospital, which waived the need for informed consent.

Antipsychotics

Exposure to AP was defined based on filling at least 1 prescription during the first 90 days of pregnancy (first trimester), **Key Points**

Question Does the use of antipsychotics during pregnancy increase the risk for congenital malformations?

Findings In this cohort study of 1.3 million pregnant women, after accounting for psychiatric conditions and other potential confounding variables, no increased risk for congenital malformations was found for typical or atypical antipsychotics, with a possible exception for risperidone.

Meaning These findings suggest that antipsychotics have no important teratogenic effects; the initial safety signal for risperidone requires further study.

which is considered the etiologically relevant period for organogenesis. We evaluated typical and atypical APs (eTable 1 in the Supplement) as 2 separate classes, as well as the most frequently used individual medications, including aripiprazole, olanzapine, quetiapine fumarate, risperidone, and ziprasidone. Women were considered unexposed if they did not fill an AP prescription during the 3 months before the start of pregnancy or during the first trimester.

Congenital Malformations

The presence of congenital malformations was defined on the basis of inpatient or outpatient diagnoses and procedure codes from the International Classification of Diseases, Ninth Revision (ICD-9) in the maternal (first month after delivery) or infant (first 3 months after date of birth) record. The maternal record was considered because Medicaid claims are sometimes recorded under the mother before the infant's eligibility has been processed.⁹ We defined 13 specific malformation groups (central nervous system, ear, eye, cardiovascular, other vascular, respiratory, oral cleft, gastrointestinal tract, genital, urinary, musculoskeletal, limb, or other) and considered a malformation to be present if (1) an ICD-9 diagnosis was recorded for the specific malformation on more than 1 date, (2) a diagnosis was recorded on 1 date as well as a relevant surgery or procedure code, or (3) a diagnosis was recorded on 1 date and the infant died before 90 days. If the malformation was identified using codes from the maternal record only and these codes were also present during the first 105 days of pregnancy, the outcome was excluded under the assumption that it reflects a preexisting malformation in the mother. An infant was considered to have a major congenital malformation if any of the 13 specific malformations were present. Because of earlier findings of a possible increased risk,^{4,6,7} we evaluated the risk for cardiac malformations separately (eTable 2 in the Supplement).

Covariates

We considered a broad range of potential confounders or proxies for potential confounders, including calendar year, age, race, smoking, multiple gestation, indications for APs, other maternal morbidity, concomitant medication use, and general markers of the burden of illness. Maternal morbidity and concomitant medication use were measured from 3 months before the start of pregnancy to the end of the first trimester.

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General markers of the burden of illness were assessed during the 3 months before pregnancy to avoid these measures of intensity of health care use being affected by early awareness of possible pregnancy complications (eFigure 1 in the Supplement). Indications for AP use included schizophrenia, bipolar disorder, psychosis, depression, anxiety, attention-deficit/ hyperactivity disorder, and other psychiatric disorders. Other maternal morbidities included other psychiatric and neurologic conditions (personality disorder, adjustment disorder, delirium, sleep disorders, chronic fatigue syndrome, alcohol and other drug abuse or dependence, epilepsy or convulsions, migraine, and other headaches) and chronic maternal illness (diabetes, hypertension, renal disease, and obesity). Concomitant medications assessed included other psychotropic medications (anticonvulsants, antidepressants, anxiolytics, benzodiazepines, other hypnotics, barbiturates, and stimulants), antidiabetic medications, antihypertensives, and suspected teratogens, as well as methadone hydrochloride, naloxone hydrochloride, naltrexone hydrochloride, and opioid use as possible proxies for drug abuse or dependence. Finally, general markers of comorbid illness included the Obstetric Morbidity Index¹⁰ and numbers of distinct prescriptions for medications other than APs, distinct diagnoses, outpatient visits, hospitalizations, and emergency department visits.11

Statistical Analysis

Data were analyzed in 2015. Baseline characteristics were assessed in women treated and not treated with the various APs, and balance was assessed using the standardized mean difference. We calculated absolute risks for any congenital malformation and for cardiac malformations and unadjusted relative risks (RRs) with their 95% CIs. Use of the robust variance estimator to account for correlations within women with multiple pregnancies did not appreciably change the CIs, so correlation structures were omitted from all analyses.

In the first adjusted analysis, we accounted for psychiatric and neurologic conditions and the use of other psychotropic medications to adjust for the possible confounding effect of the underlying disease and its associated factors. In the second adjusted analysis, we accounted for all potential confounding variables described earlier. Given the large number of covariates considered, propensity scores (PS)-estimated using logistic regression-were used as a data reduction technique. We trimmed the cohort by excluding observations from the nonoverlapping regions of the PS distributions and created 50 equal-sized PS strata based on the distribution among the treated women.¹² In the outcome models, the untreated observations were weighted using the distribution of the treated women among PS strata. Adjusted RRs were estimated using generalized linear models (PROC GENMOD [SAS Institute Inc] with a weight statement and loglink function).

Sensitivity Analyses

Sensitivity analyses were conducted to test the robustness of the findings. To account for potential residual confounding, we conducted high-dimensional PS analyses that included 200 empirically defined covariates, in addition to the investigatordefined covariates, to account for additional proxies of unmeasured confounders.¹³ To ensure that the results do not change when restricted to women with the main US Food and Drug Administration-approved indications for APs, we restricted the treated and untreated populations to women with schizophrenia, bipolar disorder, or psychosis and adjusted for other psychiatric disorders. To evaluate the potential impact of exposure and outcome misclassification, we redefined exposure as having filled at least 2 prescriptions for an AP during the first trimester and having at least a 1-day supply that overlapped with the first trimester, and we defined the outcomes based on infant claims only. We evaluated the presence of a dose-response association using strata based on dose tertiles for the first prescription filled during the first trimester. Because the cohort included live births only, we examined the potential impact of differences in the proportion of terminations among women treated with APs vs those untreated within levels of covariates used in the adjustment.¹⁴

Exploratory Analyses

Exploratory analyses were conducted to further test potential safety signals emerging from the study. We adjusted for high-dimensional PS in the analyses in which exposure was defined based on filling at least 2 prescriptions. We compared the risk for malformations between women who were treated with APs during the 3 months before the start of pregnancy and continued treatment during the first trimester (continuers) and women who discontinued treatment before the first trimester (discontinuers). The rationale for this analysis is that discontinuers might be more comparable to continuers than women who were never treated with APs. We increased the length of the gap between the last prescription fill and the start of pregnancy for the discontinuers (0, 4, 6, and 8 weeks) to reduce the likelihood of misclassifying as unexposed any women who still had medication from their last fill available to consume early in pregnancy. All analyses were conducted using the SAS System for Unix (version 9.3; SAS Institute Inc).

Results

The study cohort included 1 341 715 pregnancies (mean [SD] age of women, 24.02 [5.77] years) who met the inclusion criteria (eFigure 2 in the **Supplement**). Among those, 9258 women (0.69%) filled a prescription for an atypical AP during the first trimester, and 733 women (0.05%) filled a prescription for a typical AP. The most frequently used atypical AP was quetiapine (n = 4221), followed by aripiprazole (n = 1756), risperidone (n = 1566), olanzapine (n = 1394), and ziprasidone (n = 697). A total of 1 331 910 women had no AP dispensed from 3 months before the start of pregnancy through the end of the first trimester.

Women taking APs during the first trimester were older, more likely to be white, and more likely to deliver prematurely. They had a much higher burden of comorbid illness than untreated women as judged by the differences in baseline characteristics. They had more psychiatric and neurologic conditions and other comorbid conditions, used more psychotropic medications and suspected teratogens, and were generally in poorer health. These patterns were observed for atypical and typical APs and for each of the individual atypical agents (**Table 1** and **Table 2** and eTables 3-9 in the Supplement).

The absolute risks for congenital malformations overall per 1000 live-born infants were higher among treated women, with 38.2 (95% CI, 26.6-54.7) for those treated with typical APs and 44.5 (95% CI, 40.5-48.9) for those treated with atypical APs vs 32.7 (95% CI, 32.4-33.0) for untreated women. Higher risks were also observed for women treated with all individual atypical medications, ranging between 37.3 and 51.1 per 1000. Likewise, increased risks for cardiac malformations were observed for all AP exposures (Table 3). Unadjusted analyses suggested a significantly increased risk for malformations overall for atypical APs (RR, 1.36; 95% CI, 1.24-1.50) and for the individual agents within that class with the exception of ziprasidone (Figure 1A and eTable 10 in the Supplement). The findings for cardiac malformations were similar, but the effects were less precisely estimated (Figure 2A and eTable 10 in the Supplement). A small, nonsignificant increase in the unadjusted risk was seen for typical APs (RR for overall malformations, 1.17 [95% CI, 0.81-1.68]; RR for cardiac malformations, 1.18 [95% CI, 0.64-2.18]).

These associations were attenuated when we adjusted for psychiatric indications, but the risk for malformations overall, although small, remained significantly increased for atypical APs (RR, 1.12; 95% CI, 1.02-1.23), with the largest increase seen for risperidone (RR, 1.31; 95% CI, 1.05-1.63) (Figure 1B). The corresponding RRs for cardiac malformations were 1.15 (95% CI, 0.98-1.35) and 1.39 (95% CI, 0.96-2.01) for atypical APs and risperidone, respectively (Figure 2B and eTable 10 in the Supplement).

The AP-treated and untreated populations were well balanced in terms of baseline characteristics once we accounted for all potential confounding variables and their proxies, as evidenced by the absolute value of the standardized differences less than 0.1 (Tables 1 and 2). In this fully adjusted analysis, the RR further shifted to the null for atypical APs for malformations overall (RR, 1.05; 95% CI, 0.96-1.16) and for cardiac malformations (RR, 1.06; 95% CI, 0.90-1.24). In contrast, the risk remained elevated for risperidone for overall malformations (RR, 1.26; 95% CI, 1.02-1.56) and cardiac malformations (RR, 1.26; 95% CI, 0.88-1.81). None of the other medications were associated with an increased risk (Figures 1C and 2C and eTable 10 in the Supplement).

Sensitivity Analyses

Use of the high-dimensional PS and restriction to women with schizophrenia, bipolar disorder, or psychosis did not affect the findings, but the associations were less precisely estimated in the restricted analyses (eTable 11 in the Supplement). Redefining exposure as having filled at least 2 prescriptions for an AP or having at least a 1-day supply that overlapped with the first trimester also did not meaningfully change the results, although the association appeared to strengthen somewhat for risperidone in analyses based on filling at least 2 prescriptions (RR for any malformation, 1.46 [95% CI, 1.01-2.10]; RR for cardiac malformations, 1.87 [95% CI, 1.09-3.19]). Restricting the outcome definition to malformation codes in the in-

fant's record only did not alter the results (eTable 12 in the Supplement). We found no evidence of a dose-response relationship for any of the individual APs except for risperidone. Dosages of at least 2 mg/d were associated with an increased risk for cardiac malformation (RR, 2.08; 95% CI, 1.32-3.28) (eTable 13 in the Supplement). The RR for any congenital malformation associated with atypical AP use increased from 1.05 to 1.09 and 1.14 after accounting for potential differences in the probability of termination of malformed fetuses among treated and untreated women, respectively, using the best estimates based on the literature (eResults in the Supplement).

Exploratory Analyses

Adjusting for the high-dimensional PS in analyses in which exposure was defined based on filling at least 2 prescriptions for risperidone did not attenuate the associations (eTable 14 in the Supplement). Finally, comparing risperidone continuers with discontinuers, the strength of the association between risperidone continuation and congenital malformations was null without a gap between the last prescription fill and the start of pregnancy for the discontinuers but increased in a monotonic manner when defining discontinuation with increasing levels of strictness (eTable 15 in the Supplement).

Discussion

We examined the association between AP use during the first trimester of pregnancy and the risk for congenital malformations overall and cardiac malformations in a cohort of 1341715 Medicaid-insured women. We did not observe a significant increased risk for either outcome for typical or atypical APs after controlling for potential confounding by mental and physical comorbid conditions and their associated behaviors, with the exception of risperidone.

The small increase in absolute risk and RR for malformations observed with risperidone should be interpreted with caution because no apparent biological mechanism can readily explain this outcome, and the possibility of a chance finding cannot be ruled out. This finding should therefore be interpreted as a potential safety signal that will require follow-up in other studies, which also should attend carefully to confounding and have adequate statistical power. Ennis and Damkier¹⁵ reported an unadjusted RR of 1.5 (95% CI, 0.9-2.2) for risperidone for major congenital malformations based on a systematic review of the evidence available through May 2014 (432 exposed pregnancies). Although this estimate is similar to our unadjusted estimate (RR, 1.56), it is based on a population reference value with no confounding control. The extent to which this association can be attributed to confounding is therefore not known.

Our study has several strengths, including its large size. Our cohort included 9258 women exposed to an atypical AP during the first trimester compared with 570 in the largest study available to date.⁶ This number enabled us to assess the risks associated with individual agents. The richness of the source data allowed for careful control of potential confounders through the use of PS methods, including more than

	Unadjusted			Adjusted		
Characteristic	Atypical AP (n = 9258) ^a	Untreated (n = 1 331 910)	Standardized Difference	Atypical AP (n = 9237)	Untreated (n = 1 289 826) ^b	Standardized Difference ^b
Age, mean (SD), y	25.39 (6.42)	24.01 (5.77)	0.23	25.39 (6.43)	25.48 (6.33)	-0.01
Race, No. (%) ^c						
White	5856 (63.25)	532 864 (40.01)	0.48	5842 (63.25)	901761 (67.78)	-0.10
Black	2001 (21.61)	446 430 (33.52)	-0.27	1996 (21.61)	245 285 (18.44)	0.08
Hispanic	472 (5.10)	196 589 (14.76)	-0.33	471 (5.10)	59334 (4.46)	0.03
Other or unknown	929 (10.03)	156 027 (11.71)	-0.05	928 (10.05)	124 119 (9.33)	0.02
Multiple gestation, No. (%)	384 (4.15)	45 188 (3.39)	0.04	384 (4.16)	55 927 (4.20)	0.00
AP indications, No. (%)						
Bipolar disorder	3165 (34.19)	9548 (0.72)	0.98	3144 (34.04)	397 813 (29.90)	0.09
Psychosis	537 (5.80)	1655 (0.12)	0.34	521 (5.64)	62 138 (4.67)	0.04
Depression	2985 (32.24)	62 931 (4.72)	0.76	2972 (32.17)	480 077 (36.08)	-0.08
Schizophrenia	761 (8.22)	873 (0.07)	0.42	740 (8.01)	67 722 (5.09)	0.12
Anxiety	2373 (25.63)	41 813 (3.14)	0.68	2362 (25.57)	374 207 (28.13)	-0.06
ADHD	768 (8.30)	9965 (0.75)	0.37	763 (8.26)	115 961 (8.72)	-0.02
Other	964 (10.41)	14 930 (1.12)	0.41	946 (10.24)	130 588 (9.81)	0.01
Other mental and neurologic con						
Adjustment disorder	387 (4.18)	13 153 (0.99)	0.20	385 (4.17)	61206 (4.60)	-0.02
Personality disorder	363 (3.92)	2130 (0.16)	0.27	353 (3.82)	46 201 (3.47)	0.02
Alcohol abuse or dependence	393 (4.24)	6669 (0.50)	0.25	385 (4.17)	53 274 (4.00)	0.01
Other drug abuse or dependence	916 (9.89)	16 256 (1.22)	0.39	908 (9.83)	127 411 (9.58)	0.01
Chronic maternal illness, No. (%)						
Diabetes	311 (3.36)	23 780 (1.79)	0.10	307 (3.32)	46 169 (3.47)	-0.01
Hypertension	392 (4.23)	25 028 (1.88)	0.14	391 (4.23)	58 987 (4.43)	-0.01
Smoking	882 (9.53)	39 418 (2.96)	0.27	877 (9.49)	130 899 (9.84)	-0.01
Other psychotropic medications,	No. (%)					
Anticonvulsants	2804 (30.29)	21 055 (1.58)	0.85	2786 (30.16)	378 039 (28.41)	0.04
Antidepressants	6629 (71.60)	107 131 (8.04)	1.71	6608 (71.54)	1 070 319 (80.44)	-0.21
Anxiolytics	398 (4.30)	4305 (0.32)	0.27	394 (4.27)	60 021 (4.51)	-0.01
Barbituates	288 (3.11)	12 992 (0.98)	0.15	287 (3.11)	46 987 (3.53)	-0.02
Benzodiazepines	2733 (29.52)	35 889 (2.69)	0.78	2725 (29.50)	417 032 (31.34)	-0.04
Other hypnotics	1838 (19.85)	43 700 (3.28)	0.54	1829 (19.80)	271 134 (20.38)	-0.01
Stimulants	937 (10.12)	7789 (0.58)	0.43	935 (10.12)	138 814 (10.43)	-0.01
Other medications, No. (%)	(/			(/	()	
Antidiabetics	151 (1.63)	8973 (0.67)	0.09	149 (1.61)	21942 (1.65)	0.00
Antihypertensives	723 (7.81)	28 278 (2.12)	0.26	722 (7.82)	108 126 (8.13)	-0.01
Opioids	3616 (39.06)	254 845 (19.13)	0.45	3609 (39.07)	558 292 (41.96)	-0.06
Suspected teratogens, No. (%) ^d	1596 (17.24)	110 453 (8.29)	0.27	1589 (17.20)	246 150 (18.50)	-0.03
No. of distinct prescriptions,	4.75 (3.98)	1.63 (2.31)	0.96	4.74 (3.98)	5.04 (3.95)	-0.08
mean (SD) Hospitalization, No. (%)	779 (8.41)	48 687 (3.66)	0.20	763 (8.26)	98 525 (7.41)	0.03
Proxies for severity	() (0.71)	10 007 (0.00)	0.20	,03 (0.20)	55525 (7.41)	0.05
Obstetric Comorbidity Index, mean (SD) ^e	1.60 (1.84)	0.89 (1.39)	0.44	1.59 (1.84)	1.62 (1.86)	-0.01
No. of diagnoses, mean (SD)	5.56 (4.66)	2.56 (3.10)	0.76	5.53 (4.61)	5.90 (4.71)	-0.08
(JD)						

Abbreviations: ADHD, attention-deficit/hyperactivity disorder;

AP, antipsychotic; PS, propensity scores.

^a All APs were administered orally except for risperidone in 22 users, who received injections.

^c Determined on the basis of information submitted to the Centers for Medicare & Medicaid Services by individual states, which was based on information that had been collected and coded from Medicaid applications.
^d Included fluconazole, methimazole, danazol, propylthiouracil, progestins, and

^b All characteristics included in PS. To account for PS, the untreated observations were weighted using the distribution of the treated observations among 50 PS strata. Observations from the nonoverlapping regions of the PS distributions were trimmed.

^d Included fluconazole, methimazole, danazol, propylthiouracil, progestins, and corticosteroids. Women exposed to known teratogens (warfarin, angiotensin-converting enzyme inhibitors, antineoplastic agents, lithium,

isotretinoin, misoprostol, and thalidomide) were excluded from the cohort.

^e Scores range from 0 to 45, with higher scores indicating greater comorbidity.

	Unadjusted			Adjusted		
Characteristic	Typical AP (n = 733)	Untreated (n = 1 331 910)	Standardized Difference	Typical AP (n = 727)	Untreated (n = 1 297 638) ^a	Standardized Difference ^a
Age, mean (SD), y	26.96 (6.12)	24.01 (5.77)	0.50	26.97 (6.14)	27.08 (6.20)	-0.03
Race, No. (%) ^b						
White	356 (48.57)	532 864 (40.01)	0.17	354 (48.69)	685 149 (51.53)	-0.06
Black	253 (34.52)	446 430 (33.52)	0.02	250 (34.39)	433 204 (32.58)	0.04
Hispanic	49 (6.68)	196 589 (14.77)	-0.26	48 (6.60)	79 658 (5.99)	0.03
Other or unknown	75 (10.23)	156 027 (11.71)	-0.05	75 (10.32)	131 697 (9.90)	0.01
Multiple gestation, No. (%)	27 (3.68)	45 188 (3.39)	0.02	27 (3.71)	49 216 (3.70)	0.00
AP indications, No. (%)						
Bipolar disorder	201 (27.42)	9548 (0.72)	0.83	195 (26.82)	390 813 (29.39)	-0.06
Psychosis	95 (12.96)	1655 (0.12)	0.54	90 (12.38)	137 893 (10.37)	0.06
Depression	192 (26.19)	62 931 (4.72)	0.62	189 (26.00)	385 850 (29.02)	-0.07
Schizophrenia	160 (21.83)	873 (0.07)	0.74	154 (21.18)	225 370 (16.95)	0.11
Anxiety	134 (18.28)	41 813 (3.14)	0.50	131 (18.02)	278 458 (20.94)	-0.07
ADHD	24 (3.27)	9965 (0.75)	0.18	24 (3.30)	48 059 (3.61)	-0.02
Other	73 (9.96)	14 930 (1.12)	0.39	69 (9.49)	130 848 (9.84)	-0.01
Other mental and neurologic condition	s, No. (%)					
Adjustment disorder	20 (2.73)	13 153 (0.99)	0.13	20 (2.75)	41 997 (3.16)	-0.02
Personality disorder	37 (5.05)	2130 (0.16)	0.31	36 (4.95)	67 751 (5.10)	-0.01
Alcohol abuse or dependence	38 (5.18)	6669 (0.50)	0.28	34 (4.68)	61 418 (4.62)	0.00
Other drug abuse or dependence	59 (8.05)	16 256 (1.22)	0.33	57 (7.84)	117 023 (8.80)	-0.03
Chronic maternal illness, No. (%)						
Diabetes	41 (5.59)	23 780 (1.79)	0.20	39 (5.36)	75 392 (5.67)	-0.01
Hypertension	36 (4.91)	25 028 (1.88)	0.17	36 (4.95)	66 835 (5.03)	0.00
Smoking	68 (9.28)	39 418 (2.96)	0.27	65 (8.94)	126 589 (9.52)	-0.02
ے 2) Other psychotropic medications, No.						
Anticonvulsants	184 (25.10)	21 055 (1.58)	0.74	179 (24.62)	363 812 (27.36)	-0.06
Antidepressants	389 (53.07)	107 131 (8.04)	1.12	383 (52.68)	801 052 (60.24)	-0.15
Anxiolytics	25 (3.41)	4305 (0.32)	0.23	24 (3.30)	48 088 (3.62)	-0.02
Barbituates	33 (4.50)	12 992 (0.98)	0.22	33 (4.54)	71 660 (5.39)	-0.04
Benzodiazepines	173 (23.60)	35 889 (2.69)	0.65	169 (23.25)	359 406 (27.03)	-0.09
Other hypnotics	175 (23.87)	43 700 (3.28)	0.63	172 (23.66)	351 977 (26.47)	-0.06
Stimulants	29 (3.96)	7789 (0.58)	0.23	154 (21.18)	225 370 (16.95)	0.11
Other medications, No. (%)						
Antidiabetics	22 (3.00)	8973 (0.67)	0.17	21 (2.89)	42 893 (3.23)	-0.02
Antihypertensives	52 (7.09)	28 278 (2.12)	0.24	52 (7.15)	101 745 (7.65)	-0.02
Opioids	255 (34.79)	254 845 (19.13)	0.36	254 (34.94)	510 679 (38.41)	-0.07
Suspected teratogens, No. (%) ^c	99 (13.51)	110 453 (8.29)	0.17	98 (13.48)	199 383 (14.99)	-0.04
No. of distinct prescriptions, mean	4.19 (3.92)	1.63 (2.31)	0.79	4.14 (3.96)	4.71 (4.28)	-0.14
SD) Hospitalization, No. (%)	74 (10.10)	48 687 (3.66)	0.26	68 (9.35)	116 279 (8.74)	0.02
Proxies for severity	,>/			- (
Obstetric Comorbidity Index, mean (SD) ^d	1.59 (1.91)	0.89 (1.39)	0.42	1.57 (1.92)	1.62 (1.95)	-0.04
No. of diagnoses, mean (SD)	5.48 (5.22)	2.56 (3.10)	0.68	5.37 (5.10)	5.88 (5.16)	-0.11
No. of outpatient physician visits, mean (SD)	8.87 (11.31)	2.80 (3.95)	0.72	8.53 (10.79)	8.09 (9.20)	0.04

^a All characteristics included in PS. To account for PS, the untreated

observations were weighted using the distribution of the treated observations among 50 PS strata. Observations from the nonoverlapping regions of the PS distributions were trimmed.

^c Included fluconazole, methimazole, danazol, propylthiouracil, progestins, and corticosteroids. Women exposed to known teratogens (warfarin, $angiotens in-converting enzyme \ inhibitors, \ antineoplastic \ agents, \ lithium,$ isotretinoin, misoprostol, and thalidomide) were excluded from the cohort. ^d Scores range from 0 to 45, with higher scores indicating greater comorbidity.

^b Determined on the basis of information submitted to the Centers for Medicare

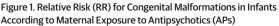
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Table 3. Absolute Risk for Congenital Malformations in Women With and Without AP Exposure

		Any Malformat	ion	Cardiac Malformation		
Exposure Group	Total No.	No. of Events	Risk per 1000 Population (95% CI)	No. of Events	Risk per 1000 Population (95% CI)	
Unexposed	1 331 910	43 494	32.7 (32.4-33.0)	15 405	11.6 (11.4-11.7)	
Typical AP	733	28	38.2 (26.6-54.7)	а	13.6 (7.4-24.9)	
Atypical AP	9258	412	44.5 (40.5-48.9)	150	16.2 (13.8-19.0)	
Aripiprazole	1756	75	42.7 (34.2-53.2)	27	15.4 (10.6-22.3)	
Olanzapine	1394	59	42.3 (33.0-54.2)	20	14.3 (9.3-22.1)	
Quetiapine	4221	182	43.1 (37.4-49.7)	70	16.6 (13.1-20.9)	
Risperidone	1566	80	51.1 (41.2-63.1)	29	18.5 (12.9-26.5)	
Ziprasidone	697	26	37.3 (25.6-54.1)	а	12.9 (6.8-24.4)	

Abbreviation: AP, antipsychotic.

^a Cell sizes of 10 or less have been suppressed in accordance with the Centers for Medicare & Medicaid Services cell size suppression policy.





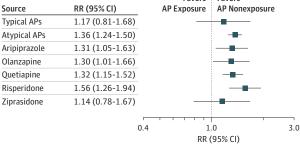
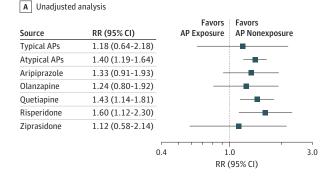
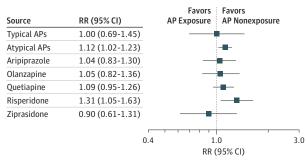


Figure 2. Relative Risk (RR) for Cardiac Malformations in Infants According to Maternal Exposure to Antipsychotics (APs)



B Adjusted for psychiatric indications

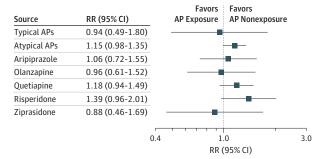




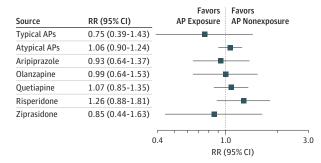
C Fully adjusted

Source	RR (95% CI)	Favors AP Exposure	Favors AP Nonexposure	
Typical APs	0.90 (0.62-1.31)			
Atypical APs	1.05 (0.96-1.16)	-	-	
Aripiprazole	0.95 (0.76-1.19)		—	
Olanzapine	1.09 (0.85-1.41)			
Quetiapine	1.01 (0.88-1.17)	-		
Risperidone	1.26 (1.02-1.56)		—	
Ziprasidone	0.88 (0.60-1.28)			
			· · ·	
		0.4 1	.0	3.0
		RR	(95% CI)	

The graph shows the results for the unadjusted analyses, analyses adjusted for psychiatric indications, and analyses adjusted for all potential confounders by exposure groups using Medicaid data from 2000 to 2010.







The graph shows the results for the unadjusted analyses, analyses adjusted for psychiatric indications, and analyses adjusted for all potential confounders by exposure groups using Medicaid data from 2000 to 2010.

50 variables in conventional analyses and more than 250 variables in high-dimensional PS analyses.

The study limitations are those characteristic of studies that use large health care utilization databases or nationwide registries and center around the potential for misclassification and selection bias. To guard against exposure misclassification (ie, false-positives), we favored specificity over sensitivity in our primary exposure definition, which required women to have filled a prescription during the first trimester. Requiring women to have filled at least 2 prescriptions under the assumption that filling multiple prescriptions increases the likelihood that the medication is being taken as prescribed did not affect the conclusions. However, neither of these definitions guarantees that the medication was actually-although likely-consumed during the first trimester. No risk for recall bias exists given the data source used, and no risk for false-negatives given that APs are not available over the counter. Similarly, we used a highly specific outcome definition because this will result in an unbiased estimate of the RR as long as the sensitivity is nondifferential. Using this definition, the risks among the unexposed were in line with expectations given the known prevalence of congenital malformations, and we previously reproduced the known associations between diabetes and overall and organ-specific malformations.¹⁶ Moreover, the definition for cardiac malformations has previously been validated.¹⁷ The absence of a validation study for the noncardiac malformations is a limitation, however. Potential confounding by variables incompletely or not captured in the data source (eg, smoking, obesity, alcohol or other drug abuse or dependance, or body

mass index) is not a major concern given our null findings. Whereas residual confounding could theoretically be an explanation for the increased risk observed with risperidone, it is unlikely to affect the findings for risperidone but not for other agents. Finally, a potential for selection bias exists because the cohort was restricted to live births. Spontaneous abortions, still-births, or planned terminations owing to congenital malformations diagnosed early in pregnancy would therefore be missed. Previous studies^{1,7} have documented that the rate of therapeutic abortions may be 10% to 20% higher among AP users. The quantitative bias analysis that assessed the potential impact of such missed terminations suggested that selection bias is not a major explanatory factor for our findings.

Whenever possible, given incomplete knowledge regarding the reproductive safety profiles of many pharmacologic agents, medications should be avoided during pregnancy. However, frequently avoidance is not possible, as is the case for women with schizophrenia, bipolar disorder, or major depressive disorder, in which few alternative treatment options are available.

Conclusions

Our findings suggest that use of APs early in pregnancy does not meaningfully increase the risk for congenital malformation or cardiac malformation, with the possible exception of risperidone. The findings for risperidone should be viewed as an initial safety signal that will require confirmation in other studies.

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Dr Hernández-Díaz reports consulting for AstraZeneca and UCB for unrelated topics. Dr L. Cohen reports receiving grant support from AstraZeneca Pharmaceuticals, Alkermes, Bristol-Myers Squibb/Otsuka, Sunovion Pharmaceuticals, Inc, Bayer HealthCare Pharmaceuticals, Ortho-MCNeil Janssen Pharmaceuticals, Ortho-MCNeil Janssen Pharmaceuticals, Inc, Pfizer, Inc, Forest Laboratories, Inc, Cephalon, Inc, GlaxoSmithKline, Takeda/Lundbeck, National Institute on Aging, National Institutes of Health, and National Institute of Mental Health (NIMH) and personal fees for consultancy from JDS Therapeutics LLC, Noven Pharmaceuticals, and PamLab LLC. No other disclosures were reported.

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REFERENCES

1. McKenna K, Koren G, Tetelbaum M, et al. Pregnancy outcome of women using atypical antipsychotic drugs: a prospective comparative study. J Clin Psychiatry. 2005;66(4):444-449.

2. Mitchell AA, Gilboa SM, Werler MM, Kelley KE, Louik C, Hernández-Díaz S; National Birth Defects Prevention Study. Medication use during pregnancy, with particular focus on prescription drugs: 1976-2008. *Am J Obstet Gynecol*. 2011;205 (1):51.e1-51.e8.

3. Gentile S. Clinical utilization of atypical antipsychotics in pregnancy and lactation. *Ann Pharmacother*. 2004;38(7-8):1265-1271.

4. Kulkarni J, Worsley R, Gilbert H, et al. A prospective cohort study of antipsychotic medications in pregnancy: the first 147 pregnancies and 100 one year old babies. *PLoS One*. 2014;9(5): e94788.

5. Cohen LS, Viguera AC, McInerney KA, et al. Reproductive safety of second-generation antipsychotics: current data from the Massachusetts General Hospital National Pregnancy Registry for Atypical Antipsychotics. *Am J Psychiatry*. 2016;173(3):263-270.

6. Reis M, Källén B. Maternal use of antipsychotics in early pregnancy and delivery outcome. *J Clin Psychopharmacol*. 2008;28(3):279-288. 7. Habermann F, Fritzsche J, Fuhlbrück F, et al. Atypical antipsychotic drugs and pregnancy outcome: a prospective, cohort study. *J Clin Psychopharmacol*. 2013;33(4):453-462.

8. Palmsten K, Huybrechts KF, Mogun H, et al. Harnessing the Medicaid Analytic Extract (MAX) to evaluate medications in pregnancy: design considerations. *PLoS One*. 2013;8(6):e67405.

9. Centers for Medicare & Medicaid Services. Medicaid Analytic Extract (MAX) general information: MAX 1999-2005 state claims anomalies from the "2005 Files" zipped file within the MAX Data 2005 to 2008 General Information, Data Dictionaries, Data Element Lists, Data Anomalies, Validation Table Measures and SAS Loads zipped file. https://www.cms.gov/Research -Statistics-Data-and-Systems/Computer-Data-and -Systems/MedicaidDataSourcesGenInfo /MAXGeneralInformation.html. Accessed December 22, 2014. **10**. Bateman BT, Mhyre JM, Hernández-Díaz S, et al. Development of a comorbidity index for use in obstetric patients. *Obstet Gynecol*. 2013;122(5): 957-965.

11. Schneeweiss S, Seeger JD, Maclure M, Wang PS, Avorn J, Glynn RJ. Performance of comorbidity scores to control for confounding in epidemiologic studies using claims data. *Am J Epidemiol*. 2001;154 (9):854-864.

12. Desai R, Rothman K, Bateman B, Hernández-Díaz S, Huybrechts K. Confounding control using propensity scores when the exposure is infrequent: making the case for a fine stratification approach. *Pharmacoepidemiol Drug Saf*. 2015;24(24)(suppl 1):217.

13. Schneeweiss S, Rassen JA, Glynn RJ, Avorn J, Mogun H, Brookhart MA. High-dimensional propensity score adjustment in studies of treatment effects using health care claims data. *Epidemiology*. 2009;20(4):512-522. **14**. Huybrechts KF, Palmsten K, Avorn J, et al. Antidepressant use in pregnancy and the risk of cardiac defects. *N Engl J Med*. 2014;370(25):2397-2407.

15. Ennis ZN, Damkier P. Pregnancy exposure to olanzapine, quetiapine, risperidone, aripiprazole and risk of congenital malformations: a systematic review. *Basic Clin Pharmacol Toxicol.* 2015;116(4): 315-320.

16. Bateman BT, Hernández-Díaz S, Fischer MA, et al. Statins and congenital malformations: cohort study. *BMJ*. 2015;350:h1035.

17. Palmsten K, Huybrechts KF, Kowal MK, Mogun H, Hernández-Díaz S. Validity of maternal and infant outcomes within nationwide Medicaid data. *Pharmacoepidemiol Drug Saf.* 2014;23(6):646-655.