

Applying polygenic risk scores to postpartum depression

Enda M. Byrne · Tania Carrillo-Roa · Brenda W. J. H. Penninx · Hannah M. Sallis · Alexander Viktorin · Brett Chapman · Anjali K. Henders · Psychiatric Genomic Consortium Major Depressive Disorder Working Group · Michele L. Pergadia · Andrew C. Heath · Pamela A. F. Madden · Patrick F. Sullivan · Lynn Boschloo · Gerard van Grootheest · George McMahon · Debbie A. Lawlor · Mikael Landén · Paul Lichtenstein · Patrik K. E. Magnusson · David M. Evans · Grant W. Montgomery · Dorret I. Boomsma · Nicholas G. Martin · Samantha Meltzer-Brody · Naomi R. Wray

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Abstract The etiology of major depressive disorder (MDD) is likely to be heterogeneous, but postpartum depression (PPD) is hypothesized to represent a more homogenous subset of MDD. We use genome-wide SNP data to explore this hypothesis. We assembled a total cohort of 1,420 self-report cases of PPD and 9,473 controls with genome-wide genotypes from Australia, The Netherlands, Sweden and the UK. We estimated the total variance attributable to genotyped variants.

We used association results from the Psychiatric Genomics Consortia (PGC) of bipolar disorder (BPD) and MDD to create polygenic scores in PPD and related MDD data sets to estimate the genetic overlap between the disorders. We estimated that the percentage of variance on the liability scale explained by common genetic variants to be 0.22 with a standard error of 0.12, $p=0.02$. The proportion of variance (R^2) from a logistic regression of PPD case/control status in all

Psychiatric Genomic Consortium Major Depressive Disorder Working Group is listed in the [Supplementary Material](#)

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E. M. Byrne (✉) · T. Carrillo-Roa · N. R. Wray
Queensland Brain Institute, The University of Queensland, Upland Road, St. Lucia, Brisbane, QLD 4072, Australia
e-mail: enda.byrne@uq.edu.au

B. W. J. H. Penninx · L. Boschloo · G. van Grootheest
Department of Psychiatry, VU University Medical Center, Amsterdam, The Netherlands

H. M. Sallis · G. McMahon · D. A. Lawlor · D. M. Evans
MRC Integrative Epidemiology Unit, University of Bristol, Bristol, UK

A. Viktorin · M. Landén · P. Lichtenstein · P. K. E. Magnusson
Department of Medical Epidemiology and Biostatistics, Karolinska Institute, Stockholm, Sweden

E. M. Byrne · T. Carrillo-Roa · B. Chapman · A. K. Henders · G. W. Montgomery · N. G. Martin
Queensland Institute of Medical Research, Brisbane, Australia

M. L. Pergadia · A. C. Heath · P. A. F. Madden
Department of Psychiatry, Washington University, St. Louis, MO, USA

P. F. Sullivan
Department of Genetics and Psychiatry, University of North Carolina at Chapel Hill, CB# 7264, 5097 Genomic Medicine, Chapel Hill, NC 27599-27264, USA

D. I. Boomsma
Department of Biological Psychology, Vrije Universiteit, Amsterdam, The Netherlands

S. Meltzer-Brody
Department of Psychiatry, University of North Carolina at Chapel Hill, Chapel Hill, NC 27599-7570, USA

H. M. Sallis · G. McMahon · D. A. Lawlor · D. M. Evans
School of Social and Community Medicine, University of Bristol, Bristol, UK

D. M. Evans
University of Queensland Diamantina Institute, Translational Research Institute, Brisbane, Australia

four cohorts on a SNP profile score weighted by PGC-BPD association results was small (0.1 %) but significant ($p=0.004$) indicating a genetic overlap between BPD and PPD. The results were highly significant in the Australian and Dutch cohorts ($R^2>1.1\%$, $p<0.008$), where the majority of cases met criteria for MDD. The genetic overlap between BPD and MDD was not significant in larger Australian and Dutch MDD case/control cohorts after excluding PPD cases ($R^2=0.06\%$, $p=0.08$), despite the larger MDD group affording more power. Our results suggest an empirical genetic evidence for a more important shared genetic etiology between BPD and PPD than between BPD and MDD.

Keywords Depression · Postpartum · Bipolar

Introduction

Postpartum depression (PPD) is defined as a subtype of major depression occurring within the first 3 months postpartum, and it can have far-reaching consequences for the woman, her children and family (Gaynes et al. 2005; Marmorstein et al. 2004; Flynn et al. 2004). PPD is associated with poorer maternal-infant attachment (Stein et al. 1991) and parenting behaviour (Gavin et al. 2005; Britton 2007). Treatment options include antidepressants and cognitive behavioural therapy, and many women experience improvement in symptoms (Miller 2002).

Differences in assessment criteria and the length of time that subjects have been followed have given rise to some inconsistency regarding the prevalence of PPD, with estimates ranging from 10 to 20 % (O'Hara and Swain 1996). PPD has partial genetic etiology with heritability of postpartum depressive symptoms estimated to be 0.38 (Treloar et al. 1999) (estimated in a sample that partly overlaps with one used in the present study), implying both genetic and environmental risk factors but with evidence of a genetic component partially distinct from major depressive disorder (MDD). A number of studies have demonstrated that women with a prior history of bipolar disorder (BPD) or MDD are at elevated risk of postpartum mood episodes. Similarly, women with siblings with BPD or MDD are also at increased risk of postpartum episodes, indicating that there are likely shared genetic risk factors between mood disorders and postpartum depression.

The advent of reasonably cheap genotyping chips that can survey a large proportion of the common genetic variants in the genome has led to a number of new insights into the genetic underpinnings of psychiatric disorders. Genome-wide association studies (GWAS) that test each genetic variant for association with the disorder of interest have been successful in identifying individual variants associated with psychiatric disorders (Sullivan et al. 2012), most notably for schizophrenia (SCZ) (Ripke et al. 2013) and BPD (Sklar et al.

2011). Specifically for BPD, three distinct regions have been reliably identified as associated with the disorder, implicating the *ANK3*, *CACNA1C* and *ODZ4* genes. However, each identified common variant has had an odds ratio (OR) of 1.2 or less, and thus very large samples are required to detect them. Studies that have examined many SNPs in aggregate rather than just one at a time have shown that many common (minor allele frequency (MAF) >0.01) SNPs of small effect account for a large proportion of the overall heritability of psychiatric disorders (Lee et al. 2012b; Sullivan et al. 2012). Moreover, when analysing large numbers of SNPs, it has been shown that many common variants that increase risk to disease are shared between disorders (Purcell et al. 2009). Genetic risk profiles constructed using results from schizophrenia GWAS studies were shown to be significantly associated with case/control status in an independent bipolar dataset, indicating that they share common genetic risk alleles.

Genome-wide association studies of MDD have not proven to be as successful in identifying genetic risk variants (Sullivan et al. 2009; Shi et al. 2011; Muglia et al. 2010; Shyn et al. 2011; Lewis et al. 2010; Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium 2013). This may reflect, in part, the phenotypic, genetic and environmental heterogeneity that characterize this disorder. PPD may represent a more homogenous subset of MDD that is more amenable to genetic analysis (females only, age-banded, all exposed to the same bio-psychosocial event). The existing GWAS studies of MDD have included PPD cases among the MDD cases, aiming to identify SNPs associated with the broad diagnostic class of MDD. Yet, evidence from a twin study suggests that a proportion of the genetic risk to PPD is distinct from MDD (Treloar et al. 1999). Furthermore, there is an increased rate of conversion to bipolar disorder in PPD cases relative to women with MDD with onset that was not after childbirth, implying that there are clinical differences between the two disorders.

To investigate the genetic architecture of PPD, we assembled a total cohort of 1,420 self-report cases of PPD and 9,473 controls with genome-wide genotypes from Australia, The Netherlands, Sweden and the UK. We sought to investigate whether there is evidence for differences in genetic risk factors between PPD and MDD without postpartum onset and specifically if there is evidence that BPD shares more genetic risk factors with PPD than with MDD. We used SNP association results from the Psychiatric Genomics Consortia (PGC) of BPD and MDD to create polygenic scores in PPD and related MDD data sets and estimate the proportion of variance (R^2) explained in case/control status by the scores. We also conducted a GWAS, but the study was underpowered to detect common risk variants with effect sizes typical of those found for other psychiatric disorders. Our study was adequately powered for performing multi-SNP profile scoring analyses.

Methods

The Australian QIMR sample

Phenotypic information was obtained from seven studies undertaken at the Queensland Institute of Medical Research (QIMR). Participants were drawn from the Australian Twin Registry and also included relatives of the twin pairs. Studies were carried out between 1980 and 2001 and consisted of mailed health questionnaires and follow-up telephone interviews. In some studies, participants were asked ‘Have you ever had a period of at least 2 weeks when you were feeling depressed or down most of the day nearly every day?’ (‘yes/no’) and if ‘yes’, female participants were asked ‘Did this depression occur around the time of childbirth?’ Later studies included a comprehensive psychiatric interview designed to assess MDD and other psychiatric disorders according to DSM-III-R and DSM-IV criteria. In other studies, participants were asked ‘Did you feel depressed after the birth of any of your children?’ (‘yes/no’), and if ‘yes’, ‘How many weeks did this go on for?’ PPD cases were defined as those endorsing depressed feelings around the time of childbirth for a period of two or more weeks and who had at least one child ($n=1,856$). Participants with no recorded history of MDD, who had at least one child, did not qualify for diagnoses of PPD, and those who did not have a sister who met the PPD criteria were selected as controls ($n=2,621$). A summary of cases drawn from each study is shown in Supplementary Table 1. A more detailed description of the samples and data collection is given in Treloar et al. (Treloar et al. 1999). Across a 10-year period, the test-retest reliability of reporting depressive symptoms after live birth was high ($r=0.75$, standard error (SE)=0.06).

After removing those who had not been genotyped, a total of 564 cases and 1,571 controls remained of which 486 cases and 1,056 controls were unrelated and used in genetic analyses (Supplementary Table 1). Depending upon the genotyping study in which they were included, participants were

genotyped either on the Illumina 317, 370, or 610 K platform (Supplementary Methods).

The QIMR PDD case/control sample partly overlaps with the QIMR samples included in the PGC-MDD study (Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium 2013) (Table 1). The PGC-MDD-QIMR cases and controls had all completed full diagnostic psychiatric interviews. Some PPD cases not in the PGC-MDD-QIMR sample had a relative in the PGC-MDD sample (either case or control). The PPD controls are a partially overlapping subset of the PGC-MDD-QIMR controls. The PGC-MDD-QIMR cases that did not report PPD were used in additional profile scoring analyses to compare how well the PGC association results can predict MDD without PPD versus MDD with PPD.

The Dutch NESDA/NTR sample

The NESDA/NTR sample is a subset of the GAIN MDD cohort (Boomsma et al. 2008; Sullivan et al. 2009; Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium 2013), which draws participants from The Netherlands Twin Registry (NTR) (Boomsma et al. 2006) and The Netherlands Study of Depression (NESDA) (Penninx et al. 2008). For this analysis, PPD cases came from the NESDA study; controls came from both NESDA and NTR studies. Details of genotyping and quality control procedures in the sample have been described in detail elsewhere (Boomsma et al. 2008; Sullivan et al. 2009), and the imputation procedure is described in the PGC-MDD study (Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium 2013). Lifetime MDD was assessed using the Composite International Diagnostic Interview (CIDI, version 2.1) (Wittchen et al. 1991). Cases of PPD were selected using a modified retrospective version of the Edinburgh Postnatal Depression Scale (EPDS) (Cox et al. 1987) that was only administered to the NESDA cohort. The EPDS is commonly used to assess current symptoms of depression and

Table 1 Profile scoring target sets from the QIMR and GAIN samples and the number of cases and controls in each set

Analysis Group	Analysis Name	Description	Sex	QIMR	NESDA/NTR	STR	ALSPAC
Group 1	PPD	PPD cases	F	484	208	104	616
		PPD-screened controls	F	1,024	761	1,351	6,311
Group 2	MDD	MDD cases	M+F	1,450	1,699		
		MDD-screened controls	M+F	1,703	1,765		
Group 3	MDD ex PPD	MDD cases ex. PPD	M+F	1,103	1,491		
		MDD-screened controls	M+F	1,703	1,765		
Group 4	PPD_all controls	PPD cases	F	484	208		
		MDD-screened controls	M+F	1,703	1,765		
Group 5	MDD_female	MDD female cases	F	932	1,180		
		MDD-screened female controls	F	988	1,095		

anxiety in the postpartum period. A score of >11 is considered as a cutoff for identifying those most likely to meet the criteria for a depression diagnosis (Cox et al. 1987; Wisner et al. 2002). The EPDS was expanded to include two initial screening questions: (1) At any point in your childbearing, did you experience symptoms of depression or anxiety that began during pregnancy or postpartum? (2) Were you ever diagnosed or treated for PND? Those women who answered yes to either question were asked to complete the EPDS based on the symptoms they experienced in the worst episode. Women scoring more than 11 were considered to be cases. Women from NTR with no history of mood disorders, a low factor score based on indices of depression, neuroticism and anxiety (Boomsma et al. 2008) and who reported having at least one child were selected as controls. A total of 214 cases and 755 controls were included in the analysis. All of these participants were included in the PGC-MDD GWAS analysis. NESDA/NTR MDD cases from the PGC-MDD study that did not have PPD were also included in separate profile scoring analyses.

The Swedish STR sample

The Swedish sample consisted of participants drawn from the Swedish Twin Registry (STR) (Magnusson et al. 2013). The phenotypic information was obtained from paper-questionnaire self-reports from the SALTY study initiated in 2007. Target population were twins born in Sweden 1943–1958. The first requests for participation in the SALTY study were sent out in early 2009, and the data collection was completed in the summer of 2010 when a total of 24,916 twins had been contacted. The survey was answered by 11,372 respondents that gave informed consent (46 %), and 54.3 % were females (Magnusson et al. 2013). PPD cases were those scoring >11 on the retrospective EPDS. A total of 104 PPD cases and 1,351 controls had available genome wide genotype data (Illumina OmniExpress platform) of which 100 cases and 1,209 controls were unrelated. Controls were not screened for psychiatric disorders such as MDD but did not report depression after childbirth.

The ALSPAC sample

The Avon Longitudinal Study of Parents and Children (ALSPAC) is a prospective birth cohort in which all pregnant women in the former county of Avon, in the South West of England, with an expected date of delivery between 1 April 1991 and 31 December 1992 were eligible to take part. The cohort has been described in detail elsewhere (Fraser et al. 2013). Phenotype information was obtained from the EPDS, which was administered to the mothers at several time points during pregnancy and postpartum to monitor depression. Scores recorded at 8 weeks postpartum were used to establish presence of PPD, with women scoring >12 classified as cases. The data reported in this paper include 6,927 women (616

cases and 6,311 controls) with both genome-wide SNP data and EPDS scores at 8 weeks postpartum. Genotyping was performed on the Illumina Human660W-quad array, and imputation to HapMap 2 was performed.

Further information on genotyping and QC in each cohort is provided in the [Supplementary Methods](#).

The study was approved by the ethics board of each of the participating institutions.

Association analyses

Genome-wide association analyses were conducted in each cohort, and a fixed-effect meta-analysis was performed in PLINK (Purcell et al. 2007). Our study was underpowered, and no genome-wide significant associations were found. Details are presented in the supplementary material ([Supplementary Methods](#), [Supplementary Tables 1–5](#), [Supplementary Figures 1–3](#)), and full results are available from the authors to allow future meta-analyses.

GREML analysis

We estimated the proportion of variance explained by the common SNPs together in the Australian, Dutch and Swedish samples using the genomic relationship matrix restricted maximum likelihood method (GREML) implemented in GCTA (Yang et al. 2011). The combined GREML analysis requires access to the raw genotypes of each cohort, and this access was unavailable for the ALSPAC sample. ALSPAC was therefore not included in the GREML analysis. We removed at random one of any pair of individuals with genetic relatedness >0.025 ($n=312$). A total of 739 cases and 2,739 controls were included in the analysis. Study cohort and ancestry principal components were included as covariates.

Polygenic profile score analyses

Despite being underpowered for association analysis, our sample is well powered as a target sample in a polygenic profile score analysis, in which the sample size of the discovery sample (in which associated SNPs are identified and their effect sizes estimated) is more critical. Assuming a significance threshold of 0.05, we have 100 % power to detect whether the polygenic scores explain 0.1 % or more of the variance in PPD case/control status in our sample (Dudbridge 2013).

Analysis groups for polygenic profile scoring

The aim of the profile scoring was to test for overlap in common genetic risk factors between BPD and PPD and MDD and PPD. We then compared the overlap in genetic risk factors for BPD and PPD and for BPD and MDD, respectively, to test if BPD shares more genetic risk with PPD than with

MDD. We therefore conducted five different profile scoring analyses to compare the genetic overlap of BPD and PPD and BPD and MDD without PPD. The groups for profile scoring analyses are as follows (Table 1):

Group 1 (PPD) This group includes all PPD cases and controls. Information on PPD case/control status was available in all of the cohorts. Profile scoring was conducted in each of the cohorts separately, and the results were combined to give an estimate of how well the BPD genetic risk score can predict PPD case/control status.

The remaining groups included only QIMR and NTR/NESDA samples as they provided information on MDD case/control status. No diagnostic information on MDD was available in the STR and ALSPAC studies.

Group 2 (MDD) All MDD cases and controls from the QIMR and NESDA/NTR studies that were included in the PGC-MDD study. This group allowed for estimation of the genetic overlap between BPD and MDD in these two cohorts regardless of whether an episode of MDD occurred postpartum or not. The NESDA/NTR PPD cases and controls are a direct subset of the NESDA/NTR cohort included in the PGC-MDD study, and so they are included in this analysis, along with MDD cases without postpartum onset. The QIMR PPD case/control set is not strictly a subset of the QIMR cohort in the PGC-MDD cohort, although there is a partial overlap between the two datasets (Table 1).

Group 3 (MDD ex PPD) PGC-MDD cohorts from the QIMR and NESDA/NTR with the PPD cases removed. In the QIMR cohort, cases in the PGC-MDD with a relative included as a PPD case in the present study were removed. This analysis allowed for testing of how well the bipolar polygene score can predict MDD case/control status in those who did not experience MDD postpartum (both males and females).

Group 4 (PPD_all controls) PPD cases compared to the controls from the PGC-MDD study. This set allows for comparing the predictive power of the bipolar polygene score in a group of PPD cases and a larger control group that has not been screened for a postpartum MDD episode. Not all of the

controls are female and have not been screened for having children. This particular analysis was included because we wanted to compare to the results of using PPD cases and only women who have had children as controls. Naively, including more controls should improve the power and therefore increase the accuracy of the profile scores.

Group 5 (MDD_female) Female cases and controls from the QIMR and NESDA/NTR PGC-MDD studies. This analysis allowed for testing of whether results from predicting PPD case/control status using bipolar polygene scores are due to a sex-specific effect in females that is not attributable to PPD.

We used the PGC-BPD (Sklar et al. 2011) (downloaded from <http://www.broadinstitute.org/mpg/ricopili/>) and PGC-MDD (Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium 2013) samples as the discovery cohorts clumped based on LD with $r^2 \leq 0.25$. As both the QIMR and GAIN samples were parts of the PGC-MDD study, we re-analysed the PGC-MDD data with the QIMR samples and NESDA/NTR samples removed to obtain the MDD polygenic profile score. A total of 6,324 cases and 6,678 controls remained in the PGC-MDD discovery sample.

We used the profile score method (Purcell et al. 2009), constructing a score for each case and control in the target samples as the sum of the log odds ratios of the risk alleles weighted by the number of risk alleles. We used different MDD and PPD cohorts as target samples (Table 1). The MDD cases and controls from the Australian and Dutch samples that were used as target samples were all included in the PGC-MDD study. Information on MDD case status was unavailable in the STR and ALSPAC samples, so comparison between MDD and PPD cases was not possible in those samples. Different SNP sets were used in the predictor based on the association p values in the discovery sample. We report the Nagelkerke's R^2 attributable to the polygenic score after fitting covariates.

To combine results across cohorts, each individual's profile score was transformed into a z-score within each cohort. The cohorts were then combined together, and a logistic regression of case/control status on the profile z-scores was performed.

Results

GREML analyses

We estimated that the percentage of variance on the liability scale explained by SNPs in the combined QIMR, GAIN and

Swedish PPD case/control samples, assuming a prevalence of 0.13, was 0.22 with a standard error of 0.12, $p=0.02$ (null hypothesis: percentage of variance on the liability scale explained by SNPs=0).

Polygenic profile scoring analyses

Profile scores based on the PGC-MDD (excluding QIMR and NESDA/NTR samples) association results were not significant in any of the four cohorts (see Supplementary Table 6). When all four samples were analysed together, profile scores based on the PGC-BPD GWAS results explained a small but significant proportion of the variance across all samples. The most significant prediction of PPD case/control status came when using PGC-BPD SNPs with $p<0.1$. The variance explained was 0.1 % ($p=0.004$, Supplementary Table 7). From analysing the results by cohort, it is clear that the predictive signal of PPD case/control status is driven primarily by the QIMR and NESDA/NTR cohorts (Fig. 1). The PGC-BPD profile scores significantly predicted PPD case/control status in both the QIMR and NESDA/NTR, and the direction of effect was such that carrying increasing numbers of BPD risk alleles increased the chance of being a PPD case. The PGC-BPD profile scores were not significantly associated with PPD case/control status in the Swedish or UK samples, and in both cohorts, the direction of the estimate of effect was in the opposite direction to that in QIMR and NESDA/NTR.

Results for all SNP sets are provided in the Supplementary Table 7, and results based on all SNPs ($n=108,824$ SNPs) are shown in Fig. 1. The diagnostic interviews conducted in the QIMR and NESDA/NTR samples allowed for diagnosis of MDD, and hence comparison of profile scoring results when applied to MDD cases and controls and PPD cases and controls. BPD profile scores were applied to different QIMR and

NESDA/NTR MDD or PPD cohorts (Table 1, Fig. 1). The PGC-BPD profile score significantly predicts MDD but explains more variance and is more significant when applied to the PPD cases and PPD controls (i.e. women who have had at least one child). The pattern of results was consistent in both cohorts.

We confirmed the significance levels via permutation analysis in the NESDA/NTR sample (Supplementary Material). The results indicated that the increased prediction of case/control status of PPD cases and controls when compared to MDD cases was unlikely to have occurred by chance ($p=0.02$).

When analysing the MDD datasets with the PPD cases removed, there was no significant prediction using PGC-BPD profile scores in the NESDA/NTR sample, implying that the prediction is mostly coming from the PPD cases. In the QIMR sample, removing the PPD cases from the overall MDD sample and trying to predict using PGC-BPD only gives significant results when using SNPs with $p<0.1$ ($R^2=0.4\%$, $p=0.02$). By contrast, the R^2 was 1.64 %, $p=3.04\times 10^{-5}$ for the QIMR PPD case/control sample for the same SNP set, despite a reduced sample size (Table 1).

Discussion

Despite the adverse impact of PPD on women and their newborns, little is understood about the genetic and environmental components affecting this disease (Mahon et al. 2009). Our estimate of variance explained by all SNPs (0.22 with a standard error of 0.12) provides direct evidence for a polygenic architecture for PPD, although large sample sizes are needed to increase the accuracy of this estimate and to identify individual associated loci. This estimate is approximately half

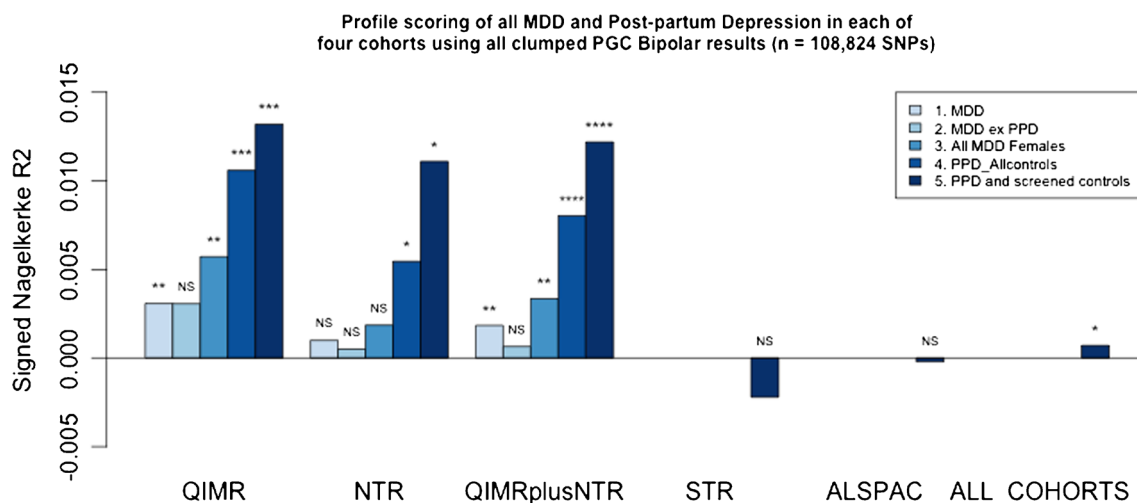


Fig. 1 Profile scoring results with polygenic scores constructed from PGC-BPD association results. **** $p<0.00005$; *** $p<0.0005$; ** $p<0.005$; * $p<0.05$; NS= $p>0.05$. Negative R^2 values represent

negative effect sizes i.e. that higher profile scores increased likelihood of being a control rather than a case

of the total heritability of PPD estimated in a twin study. The method used in this paper is not dependent upon the same assumptions as twin studies and therefore supports the results from the twin analysis showing that PPD is heritable. The observation that approximately half of the heritability of liability to PPD is tagged by common variants is in line with the results from the same analyses of schizophrenia (Lee et al. 2012a) and BPD (Lee et al. 2013) (approximately one third of the heritability explained). Two previous studies have estimated the SNP heritability of MDD, with estimates of 21 % (Cross-Disorder Group of the Psychiatric Genomics et al. 2013) and estimated as 30 % (Lubke et al. 2012) for the GAIN sample. The estimated SNP heritability of PPD lies between those estimates, although it should be noted that some of the samples included in the analysis here were also included in those analyses. In terms of the contribution of common SNPs, PPD demonstrates similar genetic architecture to other psychiatric disorders. This implies that vastly increased sample sizes for GWAS will identify common variants that increase PPD risk in the population.

The profile scoring results show evidence that the genetic risk factors for PPD overlap with those for BPD and suggest a stronger genetic relationship (at least of common variants) between BPD and PPD than between BPD and MDD. These results replicate in the QIMR and NESDA/NTR cohorts. In both cohorts, the predictive power of the PGC-BPD profile scores is reduced when non-PPD MDD cases are included in the target sample. Predicting non-PPD MDD case status is also reduced when compared to predicting PPD case status, in spite of the larger sample size in the non-PPD MDD target sample. A permutation analysis where the same number of cases and controls from the NESDA/NTR set as were in the PPD analysis were randomly selected 1,000 times, but allowing cases to have either PPD or MDD, demonstrated that the stronger prediction when analysing PPD cases and controls was unlikely to have occurred by chance.

While the overall amount of variance in PPD risk explained by the BPD scores across all cohorts is low (0.1 %), this does not imply that the actual amount of genetic overlap between the disorders is small. Similarly, the negative results when using MDD profile scores from the PGC to predict PPD case/control status do not imply that there is no genetic correlation between them. The results of the profile scoring analysis depend on several factors; the primary one being that the estimates of the SNP effects in the discovery sample should be as accurate as possible. This accuracy depends upon the power. The power of the PGC-BPD study is likely greater than that of the PGC-MDD study. Since MDD is a more prevalent disorder, sample sizes three- to fivefold greater are needed for MDD compared to BPD to afford the same power (Wray et al. 2012). Greater heterogeneity in MDD may also contribute to lower accuracy of the polygenic predictor. The sample size in the target sample also affects the power. In general, the method

of estimating the SNP effects one at a time, then summing their effects to generate a predictor, is not optimal owing to the errors in the estimation of the SNP effects (for a review of profile scoring see (Wray et al. 2013)). As an example of this, profile scoring was first used to demonstrate that a predictor based on a schizophrenia GWAS in a discovery sample could explain 3 % of the variance in schizophrenia case/control status in an independent dataset (Purcell et al. 2009). The true genetic correlation between schizophrenia cases in one sample and another is of course much greater than this. In the same study, the schizophrenia profile score could explain approximately 1 % of the variance in an independent BPD dataset. The genetic correlation between SCZ and BPD estimated from population data is 0.6 (Lichtenstein et al. 2009), while the estimate of genetic correlation based on common SNPs is 0.64 (Cross-Disorder Group of the Psychiatric Genomics et al. 2013). So, while the estimates of variance explained using profile scoring are often small except in the case where very large sample sizes are used, they reflect a genetic overlap that is far more substantial. An overall estimate of 0.1 % variance explained by PGC-BPD scores in the independent PPD samples is therefore not trivial. The estimates of the variance explained in the QIMR and NESDA/NTR PPD case/control groups are greater than 1 % similar to what was found when SCZ profile scores were used to predict BPD.

In both QIMR and NESDA/NTR, comparing PPD cases to all controls (both PPD and MDD controls) reduces the predictive power of the bipolar polygenic score, despite the fact that, naively, the larger numbers of controls should afford more power. One potential explanation for this is that the additional controls include men and women who carry BPD risk alleles, but having not experienced the environmental trigger of being pregnant or giving birth, they are at reduced risk of developing a mood disorder. The observation that the predictive ability of the polygenic score is reduced when using all females compared to the PPD only sample (i.e. female controls with children) indicates that the PPD results cannot be explained as a sex-specific effect.

The results in the QIMR and NESDA/NTR cohorts support trends from previous epidemiological studies of postpartum mood disorders. A recent study using Danish population registries that followed up women who presented to a psychiatrist for the first time and who were not given a BPD diagnosis found that almost 14 % of women who presented shortly after childbirth went on to be given a BPD diagnosis in the next 15 years. This was a threefold increase over women who first presented to a psychiatrist and were not given a BPD diagnosis at a time other than after childbirth. The risk decreased with increasing number of days postpartum the patient presented (Munk-Olsen et al. 2012). Specifically, among those women given a diagnosis of unipolar depression upon first presentation, women who presented in the postpartum period had a relative risk of 2.88 (95 % CI 1.51–5.92) of a

subsequent BPD diagnosis, compared to women presenting at any other time. Another study that compared depressive symptoms in BPD patients to unipolar patients found a greatly increased reporting of a postpartum episode in BPD patients compared to unipolar (OR=7.9, 95 % CI 0.8–378.1) (Ghaemi et al. 2004). However, the small sample size of the study meant that the null hypothesis of equality of postpartum episodes could not be rejected.

A limitation of our study was that PPD was defined by the use of self-reported data obtained from questionnaires (retrospective in three of the studies), which may be less homogeneous and not representative of a clinically derived sample. However heritability for postnatal depressive symptoms screened in the QIMR sample has previously been reported (12), and the measures showed good test-retest reliability. The modified EPDS shows strong internal consistency (Cronbach's $\alpha=0.82$ (Meltzer-Brody et al. 2013)) and may be used in the future as a screen for lifetime postpartum depression in a clinical setting. Our study did not assess postnatal mania, which is a distinct diagnosis from postnatal depression, but is much less common with a rate of approximately 1 in 1,000 births (Andrews-Fike 1999; Kendell et al. 1987). Questionnaires administered to PPD cases in subsequent studies also did not assess criteria for BPD, so it was not possible to estimate how many women went on to a subsequent diagnosis of BPD. Future studies should systematically assess postpartum mania and should include a follow-up of symptoms in the months and years after the postpartum mood episode to get a clearer picture of the diagnosis.

Another source of heterogeneity in our study is the definition of cases being different in the Australian sample when compared to the Dutch, Swedish and UK samples. Out of the 486 cases in the Australian discovery sample, 354 had been qualified as MDD by DSM-IV lifetime criteria. Of the remaining 132 Australian PPD cases, only 53 had completed a diagnostic interview that allowed DSM classification. All of the cases in the NESDA/NTR sample met the DSM-IV criteria for MDD as well as having an EPDS score above 11, and controls were screened for psychiatric disorders. The STR and ALSPAC samples used the EPDS to ascertain cases and controls, but no further information on psychiatric disorders such as MDD was available, and hence the controls were not screened. This heterogeneity in ascertainment may explain some of the differences in the profile scoring results seen across cohorts.

Further studies of the genetic relationship between BPD and PPD are warranted. Specifically, such studies should include cohorts where the controls have been fully screened for MDD and other psychiatric disorders. Studies investigating the prevalence of postpartum mood episodes in women with BPD showed that 67 % experienced an episode within 1 month of delivery, and these were almost exclusively depressive episodes with no psychotic features (Freeman et al. 2002). Allied to this, another study showed that >50 % of

women given a diagnosis of PPD were misdiagnosed and were subsequently given a lifetime diagnosis of BPD (Sharma et al. 2008). Our results support the hypothesis that postpartum depression is more closely related to BPD and highlight the need for proper screening for BPD in patients presenting with PPD. Increasing the accuracy of diagnosis could aid in selecting the best treatment and help to reduce the risk of adverse events for mother and child.

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