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Abstract

In this paper we study the effect of optimistic income expectations on life satisfaction amongst the Chinese population. Using a large scale household survey conducted in 2002 we find that the level of optimism about the future is particularly strong in the countryside and amongst rural-to-urban migrants. The importance of these expectations for life satisfaction is particularly pronounced in the urban areas, though also highly significant for the rural area. If expectations were to reverse from positive to negative, we calculate that this would have doubled the proportion of unhappy people and reduced proportion of very happy people by 48%. We perform several robustness checks to see if the results are driven by variations in precautionary savings or reverse causality.

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1 Introduction

Over the last 20 years, China has experienced significant economic and social changes. The degree, sheer size, and the speed of these changes are unprecedented in human history. Although the nation's economic changes are closely followed around the world, China's social changes are less apparent. These social changes have undoubtedly impacted upon individual day-to-day lives, including individual feelings and perceptions of the current and future states of the world. In turn, these same changes pose a significant impact on social and political stability.

Many authors have commented on the relationship between economic expectations, and social and political stability in China during periods of transition.¹ For example, in 1992, Richard Baum alleged that economic growth was the main reason behind the ability of the Chinese political system to avoid the collapse of communism experienced in the ex-Soviet regions. Similarly, Zeng (2003) contends that the legitimacy of the ruling party derives almost entirely from positive expectations, arguing that only optimistic expectations prevent emerging social problems like inequality and the uncertainty that followed the various employment and social welfare reforms. A strong indication of such perceptions within China comes from the self-identification of the regime with economic growth. Chinese leaders have actively promoted the idea that the political status quo shields economic growth. This pursuit became clear with the current Chinese President, Hu Jintao, explicitly highlighting political stability a key factor for ensuring high economic growth (AFX News, 2005).²

In this paper, we intend to quantify the role of optimistic economic expectations in keeping the Chinese happy, amidst all the social and economic

¹Examples of this idea from public think-tanks and political scientists include Holbig (2006), Giessmann (2007), and Zeng (2003).

²Younis et al. (2008), looking at the differential growth experience across South Asian countries in the past few decades, indeed also suggest that political stability and economic growth go hand in hand, though they cannot ascertain the main direction of causality.

changes.³ The main research questions we pursue are: who has optimistic economic expectations? To what extent do these optimistic economic expectations make the Chinese happy? And, by how much does life satisfaction fall, if expectations change from good to bad?

We try to answer these questions using a large-scale household survey conducted for the year 2002. The survey contains information on both future income expectations and happiness, in addition to extensive information about socioeconomic characteristics and personality traits. We also relate observed expectations to official records of the number of labour disputes (at the provincial level), in order to ascertain whether expectations matter for observed aggregate behaviour.

The following section reviews the institutional background and relevant literature. Section 3 describes the data. Section 4 analyses expectations and their determinants. In Section 5, we investigate how expectations relate to happiness. Section 6 tests the robustness of our results, given a variety of possible critiques. Conclusions are provided in Section 7.

2 Background and literature

2.1 The Chinese situation

The economic reforms which began in 1978 and led China from a planned economy to a market economy have generated unprecedented income growth and dramatic social changes. Since the late 1980s, GDP growth has been around 10% per year, a historical record in terms of sustained growth. Accompanying this extraordinary growth record is an equally noteworthy increase in income inequality. According to the China Statistical Yearbook, the

³Though economists have so far not addressed the question of happiness and social stability, psychologists and sociologists have reported positive correlations between the happiness of regions and their levels of social stability (eg. Diener and Suh 2000). Quite generally, low levels of happiness are related to instability in personal lives and group lives.

urban-rural income ratio increased from 190% in 1986 to 330% in 2006 (NBS, various years). Similarly, the Gini coefficient increased from 0.15 in 1988 to 0.32 in 2002 for the cities (Gusstafson, Li, and Sicular, 2008), and from 0.30 in the mid-1980s to 0.45 in the mid-2000s for rural areas (Benjamin, Brandt, and Giles, 2007).

In addition to the rapid income growth and change in relative income positions, the urban Chinese have also been subject to two important sources of social change. The first being the erosion of the social welfare system. Urban residents used to enjoy a ‘cradle to grave’ social welfare system during the pre-reform era. Since the mid 1990s, a new system has taken shape which places significant emphasis on individual responsibilities. Housing reform has led to the removal of subsidized housing, forcing urban households to purchase and/or rent housing from the market. The health care system transformed from a full state-covered medical service to an one-third coverage rate for state employees and null cover for private sector employees. Similarly, full pension coverage has changed to an individual retirement savings account, covering less than 50 per cent of all employees. In addition, high tuition fees and compulsory donations are charged at the primary, secondary, as well as tertiary education levels. The second source of social change for the urban population is that lifetime employment has been abolished and, as a result, some 15 million state sector employees were made redundant between 1995 and 1999 (Meng, 2000; MOLSS, 2003; Cai and Meng, 2003; Fan, 2000; Garnaut, Song, Wang, and Yao, 2001).

For rural people, the most important social changes are generated by large scale rural-urban migration. In the pre-reform era, individuals born in the countryside were not permitted to move. There was complete segregation of rural and urban economies. Rural-urban migration did not take place until the mid-1980s, although at very restricted levels. Since the early 1990s, however, large scale migration has accelerated. There are currently somewhere between 120 and 130 million migrants working in Chinese cities, with an es-

estimated further 150 million to be realised within the course of the next few decades. Compared to most developed countries, where similar population movements occurred over a hundred years, China is experiencing the phenomenon on a much larger and faster scale. Although current rural-urban migration is less restrictive than before, migrants in cities have access only to jobs which the city dwellers are unwilling to take. In addition, they have no equal right to access urban social benefits. The unprecedented scale and pace of the migration movements (with the discriminatory nature) has also altered the life course for millions of Chinese farmers.

Against this background, our study attempts to ascertain to what degree optimistic economic expectations are keeping individuals happy, hence devoting stability to the Chinese society in times of transition.

2.2 The theoretical background

Income expectations can affect current utility in two different ways. Expectations affect utility indirectly via choices or, alternatively, as a direct consumption good.

In the standard economic model where expectations affect utility only via choices, individuals maximise $E\{\sum_{t=0}^T \delta^t U(X_{it})\}$ where future utility is discounted by δ^t , and the consumption bundle, X_{it} , is dependent on the choices made before time period t . The choices are made, as such, to maximise the expected discounted stream of utility where, under the Von Neumann-Morgenstern assumptions, individual expectations are presumed to be rational, hence equal to the mathematical expectation. Note that expectations themselves have no place in the utility function.

In this framework, where no direct link between expectation and utility is present, both a positive or a negative correlation between current utility and expectations may be observed. A positive correlation can arise if the higher future income expectations are somewhat exogenous to current costly investments. For example, when arising due to circumstances completely beyond

individual control (say, the weather or an inheritance). In this situation, a rational individual would reduce precautionary savings today, increasing present consumption, and consequently giving rise to a positive correlation between income expectations and current utility. A negative correlation can arise under the exact opposite scenario, i.e. when high future expectations result from costly investments made today. Consider, for instance, a two-period model where individuals only differ with respect to their time discounting, keeping utility functions, $U(X_{it})$, homogenous. Given an increase in incomes, individuals who care more about the future will save more today, in order to enjoy higher consumption levels tomorrow. In this case, higher income expectations today are due to higher sacrifices made today for the sake of higher income in the future, implying a lower utility level today for individuals with higher income expectations.

This standard theoretical perspective will be important within the Chinese context, and, hence, later in our empirical strategy. After presenting our main argument for the importance of expectations to China, we explicitly examine whether savings and consumption can explain the found positive correlation. Additionally, we will seek evidence of a spurious negative relation between expectations and utility via unobserved variables such as discount rates.

An alternative theory linking income expectations to utility defines expectations as consumption goods. Support for this hypothesis derives from the literature within psychology and neuroscience. Findings point to distinct neural pathways by which individuals obtain psychic rewards from expectations of the future (eg. Berns et al. 2006). In its simplest form, this means the utility function, $U(\cdot)$, is not merely a function of current consumption (captured by a vector X_{it} , that includes income), but also contains a large role for subjective expectations, $E_{it}[X_{it+1}]$, of future goods. Hence, here $U(\cdot)$ is expressed as a function $U(X_{it}, E_{it}[X_{it+1}])$ rather than the standard $U(X_{it})$.

We regard this second possibility as the ‘true’ effect of expectations on

happiness as it relates to direct consumption benefits not captured via other variables.

2.3 Previous economic literature on expectations

The literature on subjective expectations in macro-economics is vast relative to that in micro-economics. The majority of the macro-economics studies deal with the usefulness of subjective expectations in predicting macro-economic variables such as inflation and growth (see, for example, Mankiw et al. 2003 and Souteles 2004). The few micro-economics studies on subjective expectations have so far mainly focussed on whether expectations of income and happiness conform to the rational expectations hypothesis (Hamermesh, 2004; Das and Van Soest 1999; Stutzer 2004; Hagerty 2003, and Frijters et al. 2008).

Whilst our paper uses explicit information on the expectations of individuals about their future income changes, the focus of the paper is on the importance of these expectations for other outcomes, rather than if expectations in themselves are rational. In a political sense, it does not matter whether or not these expectations are perfectly rational. However, what remains important is how expectations contribute to economic stability via effects on overall happiness.

To date, analysis of the effect of expectations on individual happiness remains absent from the empirical happiness literature, despite some theories hypothesising the importance of income expectations for happiness (e.g., the tunnel effect hypothesis by Senik (2005); and the theory of erroneous income expectations (Easterlin, 2001)). To our knowledge, the only available paper examining the causal effect of expectations on happiness is a recent study by Senik (2008) who implements the Russian Longitudinal Monitoring Survey (1994-2004), finding a strong effect of expectations on life satisfaction. Senik also finds that expectations improve self-rated health, reinforcing the notion that there is an actual benefit of expectations for current utility. The

main differences between our study and that of Senik (2008) are that our contexts differ and that we have access to many variables which are lacking in the Russian Longitudinal Monitoring Survey data, including self-reported relative income position, individual personality traits, mode of the day, and village/city level characteristics. Our rich data set allows us to better reveal the causal relationship between expectations and life satisfaction.

One of the main reasons why economists studying happiness avoid the role of expectations is due to traditional economic theory simply rejecting the notion of any direct effects from expectations on utility. Yet, the psychology literature has for a long time argued that expectations are consumption goods and, hence, have observable physical effects on individual well-being. For instance, according to Cannon (1914), negative expectations lead to fear. Fear is physically observable and unpleasant, making it a negative consumption good irrelevant of whether the event occurs. The idea that expectations themselves have a consumption value is a relatively new concept within economics, with only a handful of authors, such as Brunnermeier and Parker (2004), being notable exceptions. Their work addressed the issue of optimal savings when expectations of future consumption contain consumption value.

3 Data

We use data from the 2002 China Income Project Survey (CHIPs). The survey was conducted by the Institute of Economics at the Chinese Academy of Social Sciences (in early 2003) and comprises three sub-samples: urban households, rural households, and rural-urban migrant households. The rural survey was implemented in 22 out of the 30 provinces in China, while the urban survey was conducted in 12 provinces. Questionnaires for the three sub-samples are largely consistent, however, slight discrepancies are present. The total rural sample comprises of 9,200 households and 37,969 individuals. The urban sample includes 6,835 households with 20,548 individuals. And,

the migrant sample covers 2,000 households with 5,318 individuals. Most of the questions were asked of all individuals who were living in the household, with only the subjective questions inquired to one person in each household (household heads or spouses). Thus, our final sample includes only individuals of whom the income expectation questions were asked.

The survey questions one individual in each household on how they think their household income would change in the next five years. The possible answers are:

1. A reduction
2. No change
3. A slight improvement
4. A significant improvement.

In the analysis, we will refer to ‘1. A reduction’ as ‘pessimistic’, ‘2. No change’ and ‘3. A slight improvement’ as ‘neutral’, and ‘4. A significant improvement’ as ‘optimistic’.

In addition, individuals are asked to rate their general happiness (life satisfaction) on a scale ranging from 1 (not happy at all) to 5 (very happy). Figure 1 presents the distributions of these variables for the various subsamples.

For the total sample, around 10.6 and 24.6 per cent of individuals expect their incomes to ‘reduce’ or remain ‘unchanged’ in the next five years, respectively; while the remaining 74.8 per cent believe that their income will ‘increase’. Comparing this proportion to those found in Das and Soest (1999) for the Netherlands, Chinese households seem to have much higher income growth expectations. Das and Soest (1999) find that during a recession around 11 per cent of individuals believe that their income will increase in the next 12 months, while during an economic boom this figure increases to 33 per cent.

With regard to life satisfaction, slightly more than 10 per cent of the sample either regard themselves as being ‘not satisfied at all’ or ‘not very satisfied’, 32 per cent are considered to be ‘fair’, while 58 per cent of individuals are either ‘satisfied’ or ‘very satisfied’. We compare these findings to ones from the US, where around 73 per cent of respondents said (in Gallup polls) that they were ‘satisfied’ or ‘very satisfied’ with their lives. This number being 84% in 2007.⁴ Yet, the number of Americans who are dissatisfied is in the 10 to 15 per cent of the whole of the 1980-2007 period. The numbers for other OECD countries fall in a similar range (see Clark et al. 2008), suggesting that the Chinese are less happier than the average OECD respondent, with simultaneously there not being a relatively large group who is dissatisfied. Rather, there are fewer satisfied Chinese and more in the neutral range.

Figure 1 indicates that urban residents (on average) have the worst income expectations, with migrants second, and rural residents being the most optimistic. With regard to happiness, though, the pattern is not as clear. On average, the proportion of individuals who are unhappy is slightly higher for urban residents, while the proportion that regards themselves as being fairly happy is highest for migrants. Once again, rural residents are the happiest group.

Figure 2 presents the relationship of expectations and happiness with income levels for each of the three sub-samples. Within each group, individuals with higher life satisfaction and high income expectations have higher income levels. However, if we examine the income levels across groups, this relationship does not seem to be clear. The happiest rural Chinese have much lower income levels relative to the least happy urban Chinese. This seems to suggest that income itself does not matter to a great extent, and that it is perchance mainly the relative income position which defines whether individuals are happy. This finding is consistent with the literature (see,

⁴<http://www.gallup.com/poll/103483/Most-Americans-Very-Satisfied-Their-Personal-Lives.aspx>

for example, Clark et al. 2008) and was already analysed quite extensively for this data by the works of Knight and Gunatilaka (2008) and Song and Appleby (2008).⁵

Table 1 gives summary statistics of all the other variables used in the analyses, disaggregated by sub-sample. On average, urban households have the highest per capita household income and expenditure, followed by migrants and rural households. A striking aspect is that although the income level of migrants is only three quarters of that of urban households, their savings rate is around 3 percentage points higher than both urban and rural residents. This is a reflection of the temporary nature of their current status and a reflection that migration is seen as an investment decision with inter-temporal costs and benefits. Not surprisingly, hence, migrants have the lowest level of net assets.

In the rural and urban household surveys, households were asked to report their incomes in the preceding five years. This was not asked of the migrants. Using this information, we are able to calculate changes in household income for these two sub-samples. The summary statistics show that, in the early years, the degree of income changes for urban and rural household are quite similar, while in later years a much higher income growth is observed for urban households than for their rural counterparts.

4 The determinants of income expectations

In this section, we examine what determines individual income expectations. The literature on individual level income expectations is quite thin (Das and Van Soest 1999; and Ramos, 2006). The prime focus in this literature has been to question whether individual level expectations conform to the rational expectations hypothesis. Typical findings report that individuals

⁵These two papers give extensive additional background information to the Institutions in China in this period and analyse the income-happiness relation for this sample.

make predictable but smallish mistakes. This contradicts the relatively large literature within psychology which argues that observed expectations differ systematically from outcomes (for a survey, see Rabin 1998).

4.1 Methodology

We model the income expectation, IE_i , of individual i as the result of a transformation of a latent variable measured on a 4-point scale:

$$\begin{aligned} IE_i^* &= x_i\beta + \epsilon_i \\ IE_i &= k \Leftrightarrow \lambda_{k-1} \leq IE_i^* < \lambda_k \\ \lambda_0 &= 0, \lambda_4 = +\infty, \epsilon_i|x_i \sim N(0, 1), \end{aligned}$$

where x is a vector of observed individual variables, IE_{it}^* is the latent income expectation, λ_k - the thresholds increasing in k , and ϵ_i is a normally distributed error-term. This assumption makes the model a standard ordered Probit model. In line with much of the literature on limited dependent variables, we also run simple OLS models of income expectations of which the coefficients are more intuitive as they have a direct size interpretation.⁶ The results, using both estimation methodologies, are largely consistent. This is a usual finding in the literature on limited dependent variables, notably for happiness (see Ferrer et al. 2004). For simplicity of interpretation, only OLS results are reported in Table 2 and discussed below, with the Ordered Probit model results reported in Appendix A1.

As regressors, we include a set of variables which are common to all samples, such as individual and household characteristics, a log per capita income, and a self-assessed relative income position in the city/village where the respondents reside. For the urban and rural samples, we then estimate

⁶Implicitly, when running an OLS on income expectations, we assume that $IE_i = IE_i^*$ and that $E[\epsilon_i|x_i] = 0$.

additional specifications using variables only available for these very samples (Model 2). The additional specifications being, notably, retrospective annual income changes and self-assessed “good mood”, where the inclusion of mood is meant to overcome the well-known dependence of satisfaction answers on transient emotions.⁷

4.2 Interpretations of results

Starting from the total sample (first column of Table 2), we find that log per capita income and its squared term are not statistically significant for income expectations, while individual self-assessed relative income position in the city/village has a large positive effect.⁸ Individuals who regard themselves as being positioned at the top-end of the income distribution (within their home city/village) have higher income expectations than their counterparts who place themselves at the lower-end of the income distributions. This is noteworthy as it suggests a time of widening income distributions, with those already enjoying a positional advantage being better able to take the opportunities that come along. A change from the lowest to highest income position increases income expectation scores by 0.47 points.

Another interesting finding is that urban individuals, despite having much higher incomes (conditional on all the other variables), have much lower income growth expectations relative to their migrant and rural counterparts.

⁷The question addressing mood differs slightly between the rural and urban questionnaires. In the former, the question was asked about individuals’ mood on that day, while the question posed to the latter was framed in terms of “recently”. Further, the question on “relative economic position” has a different scale for the urban/migrant sample and the rural sample. In the urban/migrant surveys, answers are on a four point scale (the lowest 25%, low middle 25%, high middle 25%, and top 25%), while in the rural survey the answer is a five point scale (lowest 20%, low middle 20%, middle 20%, high middle 20%, and top 20%). To construct a consistent measure we transformed the different scale into a consistent continuous variable which is bounded between zero and 1 and denotes the mid-points of the cumulative distribution of the answer categories.

⁸When log per capita income entered as a linear term into the income expectation equation, the effect is positive and statistically significant at the 10 per cent level.

At the same time, rural people seem to have the highest income expectations among the three groups. Such findings conform with intuition since individuals with the most ground to make up believe that they stand to gain most in the near future. Together with the finding on positional effects, overall it would seem that it is the poor who expect to gain the most. On the other hand, within each group, it is those already best-placed who expect to perform better.

The effects of household composition and individual characteristics on income expectations all seem reasonable: individuals from couple only households have lower income expectations than their counterparts from other types of households. Additional children and adults increase the income expectation scores by 0.02 and 0.05, respectively. Males seem to have higher income expectations than females. Age has a U-shape relationship with income expectations, while Households with higher level of average years of schooling have more optimistic expectations. Healthy individuals and those whose spouses are healthy have high expectations, whereas neither own party membership nor spouse party membership affect income expectations. Being unemployed, or having a spouse who is unemployed, lowers income expectations. A similar result holds for individuals who work longer hours, though the effect is minute. Overall, the findings suggest that individuals with the greatest amount of human capital and those with the ability to spread efforts (i.e. members of large households) possess the most optimistic expectations.

Turning to the estimated results of Model 1 for the three separate samples (columns 2, 4, and 5 of Table 1), we find large variations in the determinants of income expectations. Income has a large inverse U-shape relationship with respect to income expectations for the urban sample, while no statistically significant effects are found for either the migrant or rural sample. These relationships are presented in Figure 3. The figure illustrates for the urban sample that the expectation score for the lowest income is around 2.8, increasing to 3.2 when log income is increased (to 8.5), and then declines with

an increase in income. At the highest income level, the expectation score actually reduces to below the level for the lowest income group, perhaps indicating that at the very top-end individuals expect to have reached their peak. For migrants, the relationship is positive, almost linear, in line with the notion that migrants were, in 2002, only starting to take the opportunities available to them. There seems to be no effect of log income on income expectations for the rural sample.

The negative effect of weekly hours worked on income expectations is only found for the migrant sample, which seems plausible given the extreme number of hours that migrants work on aggregate (71 hours a week, vs. 44 for other urban residents).

The results for Model 2 are reported in columns 3 and 6 for the urban and rural samples, respectively, with very plausible coefficients for the added variables: income expectations increase with past income increases and current mood. Rural households with migrated members have significantly higher income expectations than households without migrants, reflecting the positive effect of migration on rural household income. The other remaining relations appear to be hardly effected by the additional variables.

5 The relationship between income expectations and happiness

Our next question is how income expectations are associated with individual subjective wellbeing.

5.1 Methodology

We model the happiness level H_i of individual i as the result of a transformation of a latent variable measured on a 4-point scale:

$$\begin{aligned}
H_i^* &= x_i\gamma + \sum_j (IE_i = j)\delta_j + u_i \\
H_i &= k \Leftrightarrow \mu_{k-1} \leq H_i^* < \mu_k \\
\mu_0 &= 0, \mu_5 = +\infty, u_i|x_i \sim N(0, 1),
\end{aligned}$$

with x a set of observed individual variables, H_i^* denoting latent happiness, μ_k the thresholds increasing in k , u_i a normally distributed error-term, and $(IE_i = j)$ a set of dummy variables with $j = 1, \dots, 4$.

We first estimate a standard microeconomic happiness equation. The specification for Model 1 is the same as the expectation function except that we include the income expectation variable as an explanatory variable in the happiness equation. In Model 2, we exclude the variable indicating the previous income change (changes between 1999-2001). We should mention, here, that the inclusion or exclusion of past income changes makes little difference to the effect of income expectations. The most important coefficients are reported in Table 3.⁹

5.2 Interpretations

We first investigate the relationship between income and happiness. As normally found in this literature, income brings happiness to individuals. The relationship, however, differs among different samples. For the urban population the relationship is positive and non-linear, whereas for the migrant and rural samples a linear positive relationship is observed.¹⁰ The income gradient is much higher for the urban sample, while rural people seem to achieve the same level of happiness with much less income due to their higher baseline happiness level (see Figure 4). Among the three sample groups, at each

⁹The full results are available upon request from the authors, and the Ordered Probit model results are presented in Appendix C.

¹⁰When log per capita household income entered as a linear term it is statistically significant at the 1 per cent level for both rural and migrant samples.

particular income level (apart from log per capita income levels below 7), migrants have the lowest happiness level. Relative to income levels, the association between the relative income position and happiness is much stronger. The coefficients on the self-assessed relative income position indicate that a change from the lowest level of relative income position to the average level (where the relative income score equals 0.5) increases urban, migrant, and rural individual happiness scores by 0.41, 0.36, and 0.45 points (half of the observed coefficients), respectively. The equivalent increase in log income needed to achieve the same increase in happiness would be an increase of 1, 6, and 100 folds for the three samples, respectively. Hence, relative income dominates absolute income in terms of importance for happiness.

As discussed before, both rural and urban residents are significantly happier than the migrants. The difference being 0.36 and 0.12 scores, respectively. A salient aspect is that rural residents are the poorest group in the sample, and all of the migrants were once rural residents whose migration increased their absolute income dramatically. One of explanations for this paradox is that the migration process leads the migrants to start comparing themselves to a group richer than themselves, i.e. the urban residents. This idea is consistent with the fact that it would take a 0.33 increase in perceived relative position for the migrants to be as happy as the urban residents.¹¹ Of course, relative income is not the only reason for the unhappiness of the migrants. Other things, such as unfair treatment they receive in cities, the hard work they perform, and the fact that they are away from their families should also deplete happiness.

All the other variables seem to be consistent with the literature on the individual correlates of happiness (eg.Frey and Stutzer 2002). For example, age has a U-shape relationship with happiness, females on average are happier than males, and married people are happier than singles. Health brings

¹¹When comparing the answers of the migrant group with the urban group it is found that on the urban scale, migrants are about 0.2 lower than the other urban residents which is thus about 2/3 of the happiness difference between the migrants and the urban residents.

happiness to people, while unemployment reduces happiness. It is interesting to find that party members seem to be happier for the urban sample but not the other samples. Perhaps the benefits of party membership are higher in the cities. As expected, mood is positively associated with one's happiness and this effect is stronger for the rural sample than for the urban sample. Further, rural households with member(s) who have migrated are less happier than their counterparts, despite the fact that the same variable gives them higher income expectations (see Table 2). This suggests that migration is probably best viewed as an investment for both the migrating member and the remaining rural household.

The most important finding for this paper is that income expectations are positively associated with individual happiness. When treated as a linear variable, the coefficients range between 0.14 to 0.18. If we simply enter each response possibility as a separate dummy variable, we find that relative to individuals with pessimistic expectations, those who expect their future income to be unchanged report around 0.13 to 0.38 points higher happiness levels. If we compare individuals with pessimistic expectations to others carrying optimistic income expectations, the happiness difference increases by 0.38 to 0.64 points for the three samples. This is a 8 to 13 per cent increase in happiness levels, making expectations even more important than relative income. Unlike relative income, high expectations are not a zero-sum game.

As a confirmatory mind experiment, we can ask how important expectations are relative to log income. If we compare the coefficients and ask how much increase in log-income would be equivalent to a change in expectations from neutral to significant improvement, we find the answer is 0.6, 1.65, and 2.7 for the urban, migrant, and rural samples, respectively. This translates to an income increase of 85%, 420% and 1400% respectively. Even at the current economic growth rates experienced in China, this is not a realistic income increase for any individual to expect, even if spread out over a long period of time. Hence, the effect of expectations is far greater in terms of the

effect on life satisfaction than the possible effect of higher income could be. This suggests that the importance of expectations does not run via material consumption alone. Expectations matter beyond their material component. Note that, it does not mean that individuals expect their immaterial welfare to improve. It may simply be the feeling of material progress that gives respondents satisfaction over and beyond actual consumption. A good analogy of this feeling is the feeling one gets from the prospect of achieving one of life's main aims, such as having children. It is not only the actual achievement that gives satisfaction, but also the mere prospect that this is going to happen.

5.3 Micro-simulations

To further show the importance of income expectations in determining happiness, we perform micro-simulations. First, we use the estimated results (from Model 1) to predict for each individual the (predicted) probability of being at each happiness level. We compare the results to the actual average proportion of each sample sitting at each of the happiness levels. The motivation behind this is to show the general ability of the Ordered Probit models in predicting the sample proportions. Second, we anticipate individual happiness levels given pessimistic expectations. Finally, we repeat the second step, however this time under optimistic expectations. The results are reported in Table 4.

The results presented in Panel A are actual happiness distributions for the total sample and for the three separate samples. Panel B reports the three predicted happiness distributions from our estimated model. Comparing results from the two panels, it is clear that our model mimics the actual distributions closely, indicating a good fit for the Ordered Probit model.

In Panel C, we show the predicted happiness distributions assuming that everyone had pessimistic income expectations. For the total sample, we find that had individuals all expected falling income, their happiness level would

have decreased significantly. Comparing results in Panel C with those in Panel A, we observe that the proportion of individuals documented as being not happy at all would have increased two fold (from 1.5 per cent to 3.0 per cent). The proportion reported as being very happy would have fallen from 11.4 per cent to 5.4 per cent (a drop of 48% of the original level). Similar patterns are observed for all the three separate samples. The most dramatic changes are observed for the migrant sample: the proportion who would be very happy would have fallen to only 36 per cent of the actual proportion.

Panel D assigns every individual optimistic income expectations. Here we observe a significant increase in happiness. For the total sample, the extremely unhappy group would have declined to 37 per cent of the actual observed level. Conversely, the extremely happy group would have increased by 51 per cent. These effects are found to be most profound for the urban sample. Assuming everyone possessed optimistic income expectations would have reduced the extremely unhappy group to 19 per cent of the actual level and increased the extremely happy group by 136 per cent. This reflects mainly the fact that actual expectations of the urban group are the lowest given the three sub-samples. Hence, a switch to universal optimistic expectations would have the greatest effect on this group.

6 Alternative hypotheses

Having made our central argument, we now attempt to dislodge our findings by presenting alternative hypotheses.

6.1 Is the effect of expectations all about consumption and savings?

Within textbook economic theory, expectations themselves have no direct causal effect on utility. Consequently, expectations should have no direct

effect on life satisfaction, if life satisfaction is to be interpreted as an empirical proxy for utility. Rather, the effect of expectations on observed ‘utility’ runs via the effect of expectations on current choices that affect current consumption. Mainly, positive expectations about future wealth translate themselves into a reduced motivation for precautionary savings, thereby increasing present consumption. If this is true, then, by adding indicators of current consumption we would expect those with higher expectations to engage less in savings and for the life satisfaction effect to disappear.

Our data allows for such predictions from standard theory. In Table 5, we show, for each of the 3 samples, the savings rates for individuals with pessimistic, neutral, and optimistic expectations. The information revealed from Table 5 does not indicate any clear pattern. Nonetheless, the migrant sample is an exception. Migrants who expected income reductions saved much more than the rest of the group, clearly demonstrating precautionary saving behaviour. For the other two groups, there is no prima facie evidence of precautionary savings and, hence, no clear link between expectations and current consumption.

Additionally, we re-estimate both the income expectation and life satisfaction regressions whilst adding indicators of current savings and net assets to the original list of regressors. In the income expectations regression, the effect of saving is only significant for the migrant sample. This is consistent with Table 5. When these variables are added to the happiness equation, the effect of assets is positive for all the samples (as one would expect). The effect of savings on happiness is negative and significant for the migrant and rural samples, again complying with intuition: higher savings, conditional on income, imply lower current consumption.¹²

Table 6 compares the coefficient of income expectations on happiness from regressions with and without savings and net assets variables. As evident, their inclusion hardly changes the found effects of expectations, implying that

¹²The full results of the models with saving and net assets variables are in Appendix D.

the effect of expectations on consumption is not responsible for the strong effect on happiness.

6.2 Is there reverse causality due to personality?

So far, we have implicitly taken the error terms of income expectations and happiness (ϵ_i and u_i) to be orthogonal. Yet, we know that reverse causality plays a large role in the coefficients of many regressors on life satisfaction (see Ferrer and Frijters 2004). Unmeasured personality traits, in particular, can affect both a regressor and life satisfaction, leading to a spurious correlation between the two variables. In the context of expectations, optimistic personality traits lead to both high expectations and high life satisfaction, without there being a causal relation between expectations and life satisfaction *per se*.

One reply to such criticism is that our list of current variables already includes an indicator of current mood which is highly significant and was solely included to trace personality traits. Nevertheless, given the low explanatory power of the model, it is hard to completely dismiss the possibility that the found effects of expectations are not due to reverse causality.

As a first check, on whether personality traits are the missing variables leading to a spurious correlation between life satisfaction and expectations, we can include all available personality factors (regressors) that arguably reflect personality traits. There are no true psychologically recognised personality factors in our dataset, however there are questions in the urban and rural surveys which may reveal individual personality traits. For example, whether people follow the news and feel they are trusted by their boss.¹³ We

¹³For the urban survey, we include two indices which are generated using factor analysis from a group of questions regarding how often people read newspapers, books, and magazines; whether they listen to the radio and whether they follow topics on the economy, policy and politics. In addition, we include the following questions on individuals' description of their own personality: 1. I do my best to keep myself fit; 2. I always feel happy when I am at home; 3. In the next 10 years my health will go down the hill; 4. Maybe my lifestyle is not good, but I could not be bothered to change it; 5. My boss really trust

include these variables in model 2. The results are reported in the last panel of Table 6. Including these personality traits increased adjusted R^2 for the urban sample from 0.26 to 0.28, and from 0.30 to 0.31 for the rural sample, even though most of these personality indicators are statistically significant in the happiness equations. The results in Table 6 show that the expectation variables remain highly significant and only drop by 9 to 13% for the urban sample and from 5 to 8% for the rural sample. Hence, to the degree that our sample is suited for this question, personality traits do not appear to be able to capture the strong effect of expectations on happiness.

6.3 Is there reverse causality due to other unobservables?

A final check on the issue of reverse causality is to identify random variation. For this, we need instruments that affect expectations but do not directly affect happiness. Whilst we do not have laboratory-type instruments in our data, a few candidates seem to be reasonable for the rural sample.

Our main candidate instruments derive from indicators of the financial prospects of the villages in which the rural respondents live. These are unlikely to be related to individual personality or even individual choices. To this end, the rural survey comprises a village module which was answered by heads of villages from where the sample was drawn. In that module, the village heads provide information on village characteristics, income changes, and financial situations. We select variables from the village module which may affect income expectations of the sample households but should have

me; 6. After a day's work, I feel exhausted; 7. I often work after hours; 8. Many people come to talk to me about news and current affairs; and 9. I am a very trendy person.

For the rural sample, we include one variable regarding how often the individual helps his/her neighbours and relatives; two indices about the level of importance of family/friends/nice life/health/leisure and work/religion/politics; a dummy regarding whether the local elections are important to the individual, and one about whether income is important to the individual.

no direct effect on their happiness. These are; per capita village level debt owed by others and the proportion of the sample households in each village of which some members migrated.

Per capita village level debt owed by others indicates the village financial management ability. Villages which are better at financial management should be able to recover most of their debt owed by others. This should affect a village member's income expectations, while having no direct current effect on their happiness since we control for current income. Similarly, migrant workers send and/or bring home remittances and generate higher incomes for the whole village. Thus, the proportion of the sample households with migrated members, controlling for whether the household itself has a member migrated, should impact a household's income expectations but should have no effect on subjective wellbeing.

An additional instrument derives from the possibility of habit formation in expectations. If individuals use the past to predict the future (a learning habit often hypothesised to hold in macro-economics), then these prior experiences will affect expectations. Whilst, the most recent income change experience, current consumption and wealth measures should prevent these previous experiences from having any direct effect on life satisfaction. We use household income changes between 1999 and 2001 (log per capita income in 2001 minus log per capita income in 1999) to capture past income changes.

6.3.1 Instrumental variable methodology

Instrumental variable estimation in a bivariate Ordered Probit model is not a standard option in existing software packages. We thus explicitly model the endogeneity between income expectations and happiness as arising through the correlation of the error terms:

$$\begin{aligned}
H_i^* &= x_i\gamma + \sum_j (IE_i = j)\delta_j + u_i \\
IE_i^* &= x_i\beta + Z_i\eta + \epsilon_i \\
u_i, \epsilon_i | x_i, Z_i &\sim N\left(0, \begin{bmatrix} 1 & \rho^2 \\ \rho^2 & 1 \end{bmatrix}\right)
\end{aligned}$$

Although this model would be weakly identified from the normality assumption, the main source of identification of the effect δ_j comes from the existence of the instruments Z_i . The likelihood function for an individual observation is:

$$\begin{aligned}
L(H_i = H, IE_j = J | \beta, \gamma, \delta, \eta, \rho, \lambda, \mu) = \\
\int_{\lambda_{J-1} - x_i\beta - Z_i}^{\lambda_J - x_i\beta - Z_i} [\Phi(\mu_H - x_i\gamma - \delta_J | \epsilon_i) - \Phi(\mu_{H-1} - x_i\gamma - \delta_J | \epsilon_i)] f(\epsilon_i) d\epsilon_i
\end{aligned}$$

This model was programmed into Gauss, with standard tests implemented for the internal consistency of the instruments. We present these results in Table 6. The model includes all the personality traits and saving and net asset variables.

6.3.2 Interpretations

The results in Table 7 show that past income increases have a significant positive impact on individual income expectations. Per capita village debt owed by others reduces income expectations. The intuition for this is simple: if a village mismanages their finances and has higher levels of outstanding debt owed by others, village members may have low income expectations over and above their own household income earning potentials. As expected, the proportion of sample households with migrated members also has a strong positive impact on income expectations. The significance of each of the three

instruments is high, a necessary requirement for instruments.

The IV estimate of high expectations versus the default (negative expectations) is 0.56. It is basically the same as the coefficient of the direct ordered Probit estimation (which is 0.57). An interesting aspect of our findings is that the estimated ρ is very close to zero which suggest no significant endogeneity problems. Our preferred interpretation of this is that the omitted traits that would lead to a spurious positive relation (like unobserved consumption) cancel out the omitted traits that would lead to a spurious negative relation (such as work effort).

The test for the internal validity are shown at the bottom of Table 7. They are based on additional specifications in which the instruments were allowed to have non-zero coefficients in the happiness equation, interpreting the increase in the likelihood as a ratio-test of their instrument validity. It shows that we cannot reject the internal validity of the instruments with a 1% significant level though a 5% significant level means a rejection of the joint validity of the first two instruments. The important thing to report here, though, is that the coefficients on income expectations and ρ are virtually identical (no more than 1% difference) across these auxilliary specifications. Hence, any choice of two of these 3 instruments as ‘valid’ begets the same result on the effect of expectations on happiness.

6.4 Do expectations matter for behaviour?

As a final critique to our main analyses we can pose the question whether expectations truly matter for any observed behaviour, let alone political stability. Perhaps self-reported expectations are just cheap-talk. In this line, the relation between expectations and happiness is one between unobserved subjective traits common to both and not indicative at all of any great significance of expectations.

When looking for observed behaviour, we face the unfortunate circumstance that hardly any social and political stability measure exists in China

that would allow us to test directly the effect of income expectations on stability. Nevertheless, we are able to find some provincial level data on urban annual labour disputes. In Figure 5, we plot our sample average income expectations at the provincial level (12 data points) for our urban sample against the log of the number of labour disputes in 2004. The number of observations is obviously too small to conduct any meaningful analyses. However, the graph does indicate that the higher the income expectations, the lower the number of labour dispute cases. Hence, in the one dimension of political stability, the correlation with expectations is as hypothesised.

7 Conclusions

In this paper, we examined the importance of optimistic income expectations for the happiness levels of the Chinese. We found that optimistic expectations were amongst the most important explanatory variables, roughly equal in size to the importance of perceived relative income. We found that the difference between optimistic expectations and pessimistic expectations was worth about 9 to 15% on a happiness scale. The effect is particularly strong and important for the rural-to-urban migrants, which is predicted to grow to around half the total population in the next 20 years. Given that the Chinese are roughly as happy as a middle-income country like Nigeria or Croatia, a shift towards negative expectations would bring them to the happiness levels of relatively poor and unstable countries, such as India or Bangladesh.¹⁴ This means that continued optimistic expectations could indeed well be a large part of the explanation for the relative stability of China during the immense transition that this country is going through.

Our treatment of expectations on a proxy of utility is a deviation from the standard economic assumption, where expectations themselves are not

¹⁴Here, we have taken the comparison levels of the mentioned other countries from the World Value Survey (Veenhoven, 2004).

considered consumption goods but rather matter indirectly via choices. We found the effect of expectations to be very large and robust to the inclusion of incomes and savings, which are the choices theoretically associated with income expectations. As Senik (2008) points out, the direct importance of expectations for utility opens up a whole new set of questions. In this paper, we have taken the line that expectations are important for keeping the unhappiness associated with large societal transitions to a minimum. The role of expectations and expectation manipulation, in normal times, is a field still wide open to economists, both experimentally and empirically.

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Table 1: Summary Statistics

	<u>Total sample</u>		<u>Urban sample</u>		<u>Migrant sample</u>		<u>Rural sample</u>	
	Freq	%	Freq	%	Freq	%	Freq	%
Income expectations:								
Reduction	1,821	10.57	1,260	19.23	177	9.56	384	4.36
No change	4,235	24.59	2,109	32.18	531	28.67	1,595	18.09
Slight improvement	10,043	58.31	3,039	46.38	1013	54.70	5,991	67.95
Significant improvement	1,123	6.52	145	2.21	131	7.07	847	9.61
Life satisfaction (happiness)								
Not satisfied (happy) at all	262	1.53	144	2.21	32	1.73	86	0.98
Not very satisfied (happy)	1,534	8.97	663	10.19	178	9.63	693	7.93
Fair	5,415	31.67	2045	31.44	827	44.75	2,543	29.09
Satisfied (happy)	7,940	46.44	3,188	49.01	693	37.5	4,059	46.43
Very satisfied (happy)	1,945	11.38	465	7.15	118	6.39	1,362	15.58
	Mean	Std. Dev.	Mean	Std. Dev.	Mean	Std. Dev.	Mean	Std. Dev.
Per capita income in 2002			8508	5611	6896	6770	2772	2333
Per capita expenditure in 2002			6284	4481	4746	3819	2208	2192
Saving rate 2002			0.23	0.27	0.26	0.28	0.23	0.34
Net total assets			132401	166754	23625	87313	37330	41301
Income change 2001-2002	0.10	0.43	0.16	0.35			0.05	0.48
Income change 2000-2001	0.06	0.36	0.08	0.22			0.04	0.44
Income change 1999-2000	0.09	0.30	0.08	0.22			0.10	0.35
Income change 1998-1999	0.06	0.26	0.06	0.29			0.06	0.29
Number of children living in the household			0.57	0.56	0.66	0.73	1.16	0.98
Living in couple only household			0.19		0.31		0.07	
Proportion married			0.95		0.91		0.96	
Proportion of male household head			0.46		0.62		0.75	
Age of the household head			47.08	10.94	35.33	8.96	45.84	10.15
Years of schooling of household head			10.79	3.18	7.92	2.79	7.10	2.56
Years of schooling of spouse			10.23	3.49	7.47	2.75	6.01	2.88
Proportion of household head is party member			0.34		0.03		0.17	
Proportion of spouse is party member			0.27		0.02		0.05	
Proportion of household head unemployed			0.08		0.01		0.01	
Proportion of spouse unemployed			0.08		0.02		0.01	
Proportion of household head healthy			0.60		0.90		0.80	
Proportion of spouse healthy			0.56		0.72		0.75	
No. of observations				6569		1858		8831

Table 2: Determinants of income expectations, OLS

	<u>Total sample</u>	<u>Urban sample</u>		<u>Migrant sample</u>	<u>Rural sample</u>	
	Model 1	Model 1	Model 2	Model 1	Model 1	Model 2
Log per capita income	-0.064 [0.085]	0.932 [0.329]***	0.930 [0.328]***	-0.296 [0.296]	-0.032 [0.121]	-0.100 [0.122]
(Log per capita income) ²	0.005 [0.005]	-0.054 [0.019]***	-0.054 [0.018]***	0.020 [0.017]	0.003 [0.008]	0.006 [0.008]
Income change (2001-2002)			-0.013 [0.031]			0.040 [0.018]**
Income change (1999-2001)			0.171 [0.034]***			0.066 [0.017]***
Income position in the city/village	0.467 [0.022]***	0.576 [0.042]***	0.536 [0.043]***	0.435 [0.072]***	0.395 [0.027]***	0.366 [0.028]***
Good mood			0.139 [0.021]***			0.087 [0.015]***
Dummy indicating hh with members migrated						0.079 [0.017]***
Couple only households	-0.039 [0.020]**	0.045 [0.036]	0.051 [0.036]	0.042 [0.055]	-0.294 [0.032]***	-0.284 [0.032]***
No. of children age 0-18 at home	0.019 [0.008]**	-0.006 [0.025]	-0.009 [0.025]	-0.024 [0.036]	0.015 [0.009]*	0.010 [0.009]
No. of adults age >18	0.053 [0.007]***	0.054 [0.018]***	0.056 [0.018]***	0.020 [0.043]	0.032 [0.008]***	0.027 [0.008]***
Dummy for married	-0.111 [0.028]***	-0.123 [0.050]**	-0.124 [0.049]**	-0.211 [0.080]***	-0.015 [0.038]	-0.009 [0.038]
Own age	-0.026 [0.004]***	-0.057 [0.007]***	-0.055 [0.007]***	-0.018 [0.013]	0.014 [0.005]***	0.013 [0.005]**
(Own age) ² /10	0.002 [0.000]***	0.005 [0.001]***	0.005 [0.001]***	0.001 [0.002]	-0.002 [0.001]***	-0.002 [0.001]***
HH mean schooling year aged>20	0.010 [0.002]***	0.012 [0.004]***	0.012 [0.004]***	0.015 [0.007]**	0.005 [0.003]	0.006 [0.003]*
Own gender (male==1)	0.064 [0.013]***	0.090 [0.022]***	0.086 [0.022]***	0.039 [0.037]	0.006 [0.017]	0.008 [0.017]
Own unemployment	-0.093 [0.033]***	-0.044 [0.044]	-0.039 [0.044]	-0.041 [0.157]	-0.090 [0.074]	-0.083 [0.073]
Spouse being unemployed	-0.062 [0.029]**	-0.060 [0.037]	-0.050 [0.037]	0.235 [0.132]*	-0.021 [0.068]	-0.008 [0.067]
Own weekly working hours /10	-0.006 [0.003]*	-0.005 [0.010]	-0.006 [0.010]	-0.026 [0.009]***	0.001 [0.004]	0.002 [0.004]
Dummy for own healthy	0.093 [0.015]***	0.077 [0.024]***	0.058 [0.024]**	0.222 [0.064]***	0.077 [0.020]***	0.072 [0.019]***
Dummy for spouse being healthy	0.054 [0.015]***	0.089 [0.025]***	0.079 [0.025]***	0.017 [0.057]	0.025 [0.019]	0.021 [0.019]
Dummy for own party membership	0.012 [0.018]	0.007 [0.022]*	0.006 [0.022]	0.075 [0.094]	0.024 [0.019]	0.024 [0.019]
Dummy for spouse party member	0.011 [0.014]	0.043 [0.024]	0.034 [0.024]	0.049 [0.144]	-0.022 [0.034]	-0.019 [0.034]
rural	0.253 [0.026]***					
urban	-0.232 [0.024]***					
Region	Yes	Yes	Yes	Yes	Yes	Yes
Observations	17033	6489	6474	1841	8703	8579
R-squared	0.16	0.08	0.09	0.11	0.09	0.10

Standard errors in brackets

* significant at 10%; ** significant at 5%; *** significant at 1%

Table 3: Determinants of happiness (OLS)

	Total sample		Urban sample		Migrant sample	Rural sample	
	Model 1	Model 1	Model 2	Model 1	Model 1	Model 2	
Expectation (=no change)	0.263 [0.022]***	0.296 [0.027]***	0.278 [0.027]***	0.378 [0.066]***	0.146 [0.044]***	0.131 [0.042]***	
Expectation (=slight improvement)	0.368 [0.021]***	0.388 [0.026]***	0.358 [0.026]***	0.383 [0.063]***	0.303 [0.041]***	0.263 [0.040]***	
Expectation (=sig. improvement)	0.521 [0.030]***	0.644 [0.067]***	0.610 [0.066]***	0.608 [0.090]***	0.434 [0.049]***	0.375 [0.047]***	
Log per capita income	0.104 [0.092]	1.348 [0.320]***	1.313 [0.319]***	0.417 [0.308]	0.105 [0.147]	0.164 [0.145]	
(Log per capita income) ²	0.002 [0.006]	-0.064 [0.018]***	-0.060 [0.018]***	-0.016 [0.018]	0.000 [0.010]	-0.003 [0.009]	
Income change (2001-2002)			-0.163 [0.029]***			-0.067 [0.018]***	
Income position in the city/village	0.884 [0.025]***	0.836 [0.041]***	0.751 [0.042]***	0.726 [0.075]***	0.913 [0.034]***	0.751 [0.033]***	
Good mood			0.260 [0.020]***			0.469 [0.017]***	
Dummy indicating hh with members migrated						-0.038 [0.020]*	
Couple only households	0.004 [0.021]	-0.023 [0.035]	-0.025 [0.034]	0.021 [0.057]	0.079 [0.039]**	0.034 [0.038]	
No. of children age 0-18 at home	0.004 [0.009]	0.005 [0.024]	0.009 [0.024]	0.072 [0.037]*	-0.002 [0.011]	0.001 [0.010]	
No. of adults age >18	-0.003 [0.008]	0.007 [0.018]	0.013 [0.017]	0.007 [0.044]	0.002 [0.010]	0.007 [0.010]	
Dummy for married	0.222 [0.030]***	0.223 [0.048]***	0.204 [0.048]***	0.107 [0.083]	0.213 [0.047]***	0.215 [0.045]***	
Own age	-0.017 [0.004]***	-0.029 [0.007]***	-0.026 [0.007]***	-0.006 [0.013]	-0.015 [0.006]**	-0.012 [0.006]*	
(Own age) ² /10	0.002 [0.000]***	0.003 [0.001]***	0.003 [0.001]***	0.001 [0.002]	0.002 [0.001]***	0.002 [0.001]**	
HH mean schooling year aged>20	0.000 [0.003]	-0.003 [0.004]	-0.005 [0.004]	-0.007 [0.007]	0.001 [0.004]	0.000 [0.004]	
Own gender (male==1)	-0.043 [0.014]***	-0.051 [0.021]**	-0.061 [0.021]***	0.032 [0.039]	-0.049 [0.021]**	-0.046 [0.020]**	
Own unemployment	-0.116 [0.036]***	-0.121 [0.043]***	-0.121 [0.042]***	0.139 [0.163]	-0.132 [0.090]	-0.057 [0.086]	
Spouse being unemployed	-0.081 [0.032]**	-0.052 [0.036]	-0.047 [0.035]	-0.276 [0.137]**	-0.116 [0.082]	-0.112 [0.079]	
Own weekly working hours /10	-0.007 [0.004]**	0.008 [0.010]	0.010 [0.010]	-0.003 [0.009]	-0.012 [0.005]**	-0.012 [0.005]**	
Dummy for own healthy	0.150 [0.016]***	0.133 [0.024]***	0.098 [0.024]***	0.025 [0.067]	0.171 [0.024]***	0.141 [0.023]***	
Dummy for spouse being healthy	0.114 [0.016]***	0.102 [0.024]***	0.080 [0.024]***	0.147 [0.059]**	0.125 [0.023]***	0.096 [0.022]***	
Dummy for own party membership	0.048 [0.015]***	0.071 [0.021]**	0.037 [0.023]***	0.101 [0.149]	-0.031 [0.041]	-0.032 [0.022]	
Dummy for spouse party member	0.056 [0.020]***	0.050 [0.023]***	0.063 [0.021]*	-0.005 [0.097]	0.044 [0.023]*	0.028 [0.040]	
rural	0.357 [0.028]***						
urban	0.118 [0.026]***						
Regions	Yes	Yes	Yes	Yes	Yes	Yes	
Observations	17033	6489	6474	1841	8703	8579	
R-squared	0.23	0.23	0.26	0.16	0.24	0.30	

Standard errors in brackets

* significant at 10%; ** significant at 5%; *** significant at 1%

Table 4: Actual and predicted happiness distribution with different income expectations

Panel A: Actual	Total sample	Urban sample	Migrant sample	Rural sample
Not happy at all	0.015	0.022	0.017	0.010
Not happy	0.090	0.102	0.097	0.079
Fair	0.317	0.314	0.446	0.291
Happy	0.465	0.491	0.376	0.464
Very happy	0.114	0.072	0.064	0.156
Average score	3.570	3.490	3.372	3.677
Panel B: Predicted				
Not happy at all	0.014	0.021	0.017	0.009
Not happy	0.090	0.103	0.097	0.079
Fair	0.318	0.318	0.449	0.293
Happy	0.463	0.486	0.373	0.463
Very happy	0.115	0.073	0.065	0.156
Panel C: Predicted (assuming all expected income fall)				
Not happy at all	0.030	0.038	0.044	0.020
Not happy	0.149	0.159	0.178	0.129
Fair	0.389	0.382	0.513	0.361
Happy	0.377	0.389	0.243	0.404
Very happy	0.054	0.032	0.022	0.088
Panel D: Predicted (assuming all expect income improve significantly)				
Not happy at all	0.006	0.004	0.006	0.005
Not happy	0.052	0.037	0.052	0.052
Fair	0.257	0.204	0.372	0.245
Happy	0.513	0.586	0.457	0.489
Very happy	0.172	0.169	0.113	0.210
No. of observations	17033	6489	1814	8703

Table 5: Saving by income expectations

	<u>Urban Sample</u>		<u>Migrant Sample</u>		<u>Rural Sample</u>	
	yuan	%	yuan	%	yuan	%
Reduced	5982	18.13	5370	19.90	2651	-5.19
Not change	5563	19.28	4161	8.96	2812	8.30
Improved	6836	23.79	5081	12.03	2850	10.73
Significantly improved	6153	17.44	7445	0.67	3918	-4.11
Total	6249	21.12	5015	11.10	2938	8.16

Table 6: Comparison of estimated expectation effects with and without saving and assets variables

	<u>Total sample</u>		<u>Urban sample</u>		<u>Rural sample</u>	
	Model 1	Model 1	Model 2	Model 1	Model 1	Model 2
Model without saving and assets						
Expectation (=no change)	0.263 [0.022]***	0.296 [0.027]***	0.277 [0.027]***	0.379 [0.066]***	0.146 [0.044]***	0.136 [0.042]***
Expectation (=slight improvement)	0.368 [0.021]***	0.388 [0.026]***	0.357 [0.026]***	0.384 [0.063]***	0.303 [0.041]***	0.264 [0.040]***
Expectation (=sig. improvement)	0.521 [0.030]***	0.644 [0.067]***	0.615 [0.066]***	0.604 [0.090]***	0.434 [0.049]***	0.376 [0.047]***
Model with saving and assets						
Expectation (=no change)	0.263 [0.022]***	0.297 [0.027]***	0.278 [0.027]***	0.372 [0.066]***	0.149 [0.044]***	0.138 [0.042]***
Expectation (=slight improvement)	0.369 [0.021]***	0.39 [0.026]***	0.359 [0.026]***	0.377 [0.063]***	0.309 [0.041]***	0.271 [0.040]***
Expectation (=sig. improvement)	0.52 [0.030]***	0.637 [0.067]***	0.608 [0.066]***	0.595 [0.090]***	0.439 [0.049]***	0.381 [0.047]***
Model with personality traits						
Expectation (=no change)			0.243 [0.027]***			0.129 [0.042]***
Expectation (=slight improvement)			0.311 [0.026]***			0.248 [0.040]***
Expectation (=sig. improvement)			0.561 [0.065]***			0.354 [0.047]***

Table 7: IV estimates of happiness regression

	ML Ordered Probit	
	Happiness IV	Expectations
Income expectations = no change	0.205 [0.086]***	
Income expectations = slight improvement	0.370 [0.146]***	
Income expectations = significant improvement	0.556 [0.234]***	
Log per capita income	0.290 [0.201]	-0.304 [0.235]
(Log per capita income) ²	-0.088 [0.136]	0.213 [0.155]
Income change (2001-2002)	-0.098 [0.029]***	0.049 [0.033]*
Good mood	0.665 [0.031]***	0.153 [0.029]***
Dummy indicating hh with members migrated	-0.044 [0.031]	0.092 [0.037]***
Income change (1999-2001)		0.109 [0.027]***
Per capita village negative debt/10		-0.011 [0.004]***
Proportion of sample hh in the vill has migrants		0.256 [0.074]***
Chi-square test for over-identification of the first instrument		4.90 [0.03]
Chi-square test for over-identification of the second instrument		4.70 [0.04]
Chi-square test for over-identification of the third instrument		0.05 [0.82]
rho	-0.025	
Observations	8417	
Mean Log-lik	-1.95	

Standard errors in brackets, except that the figure in the bracket below the Chi-square test is the p-value.

* significant at 10%; ** significant at 5%; *** significant at 1%

The Chi-square-test is based on the likelihood ratio of an additional estimation where the first instrument was allowed to have a non-zero effect on happiness.

Figure 1: Expectations and happiness

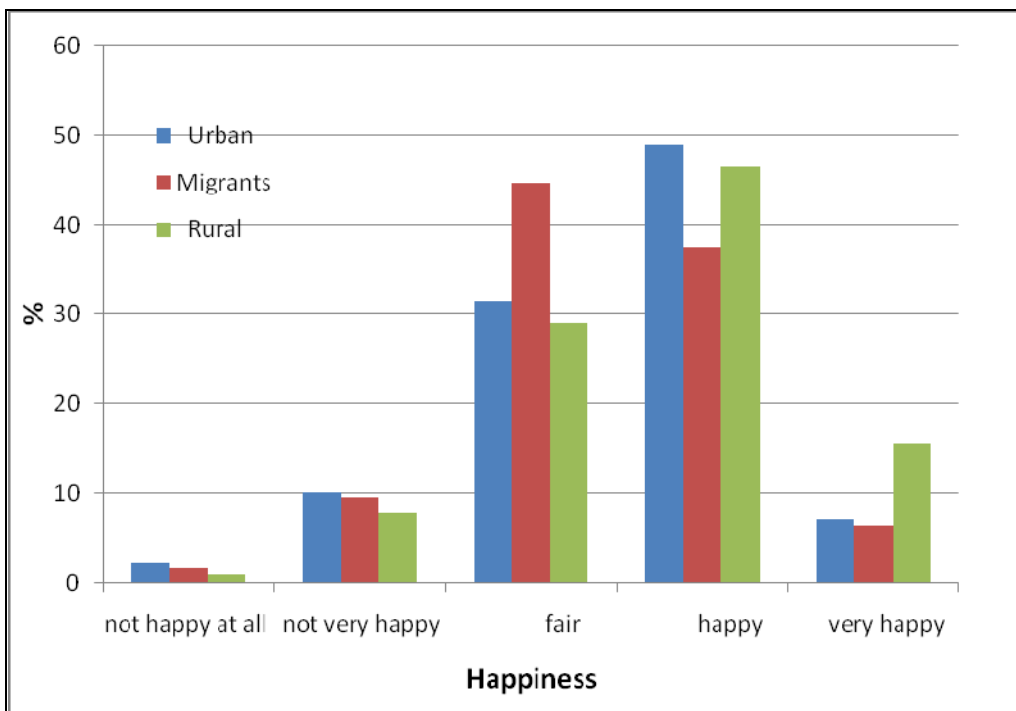
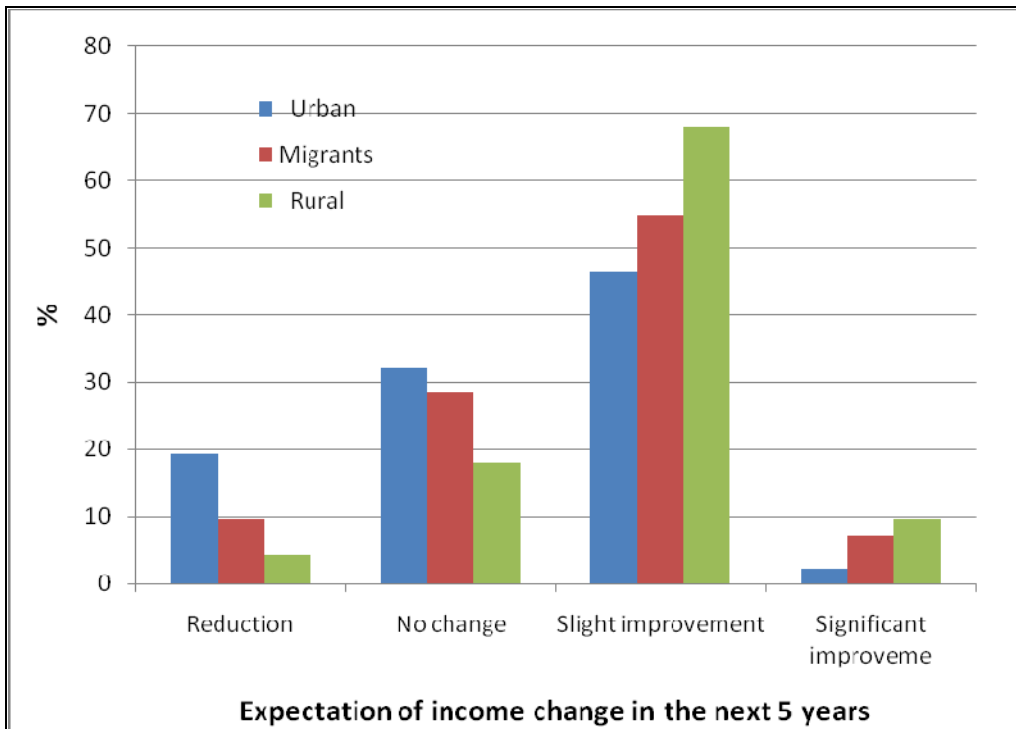


Figure 2: Income and expectation and happiness

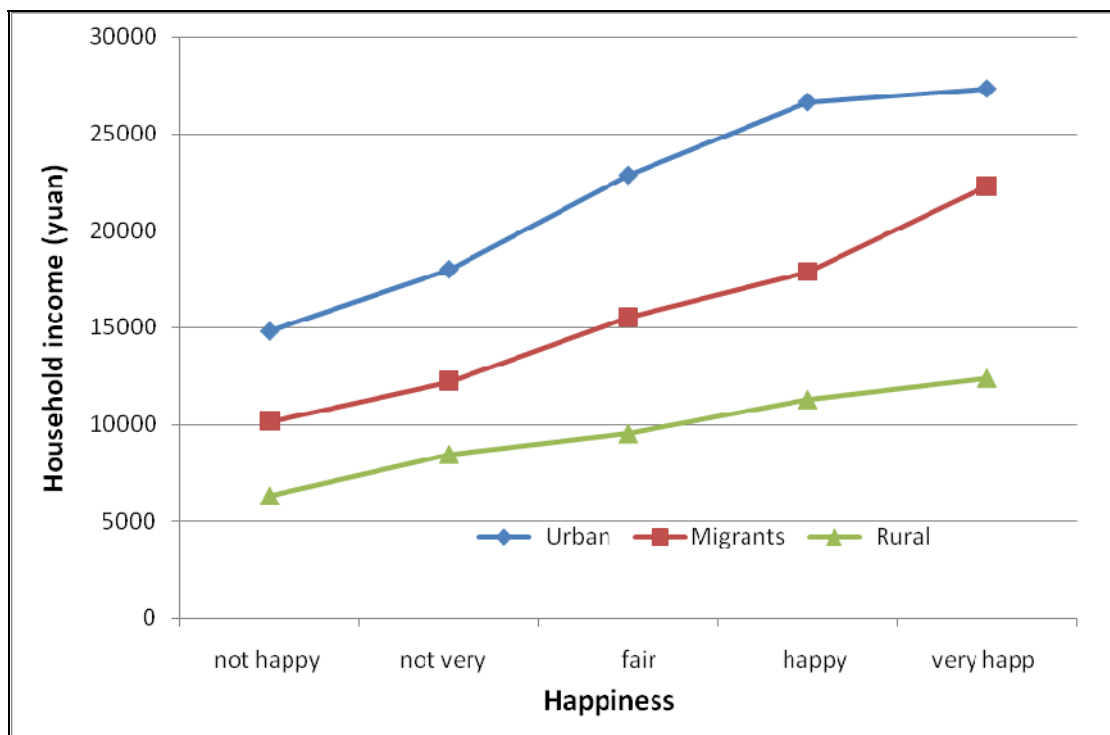
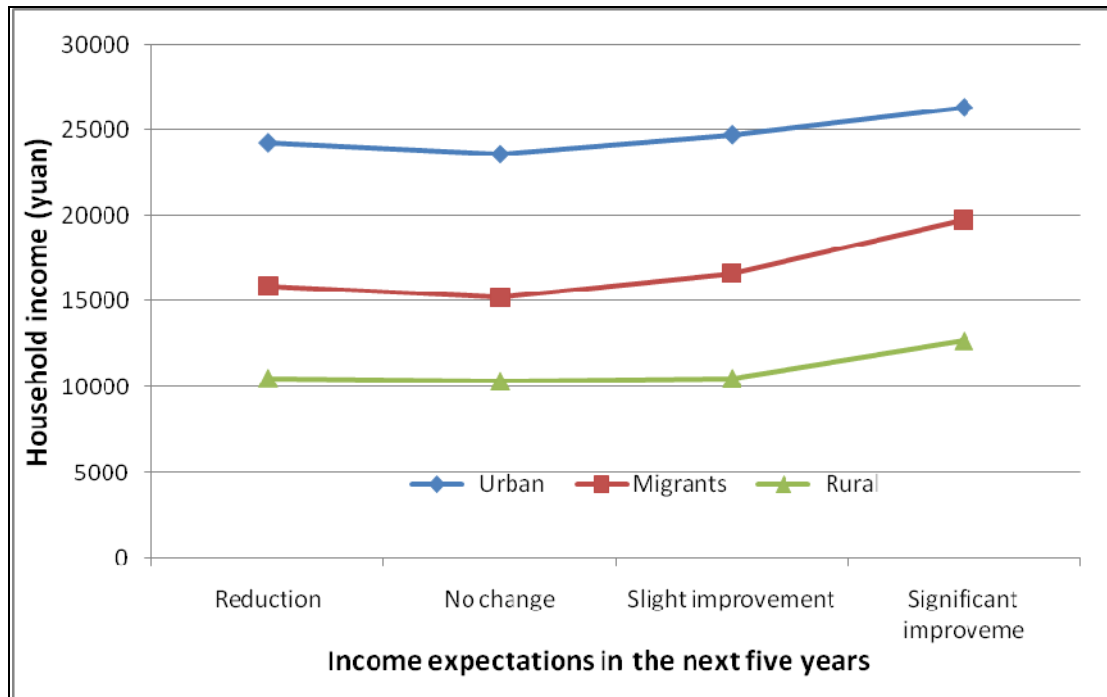


Figure 3: Predicted relationship between log per capita income and income expectations

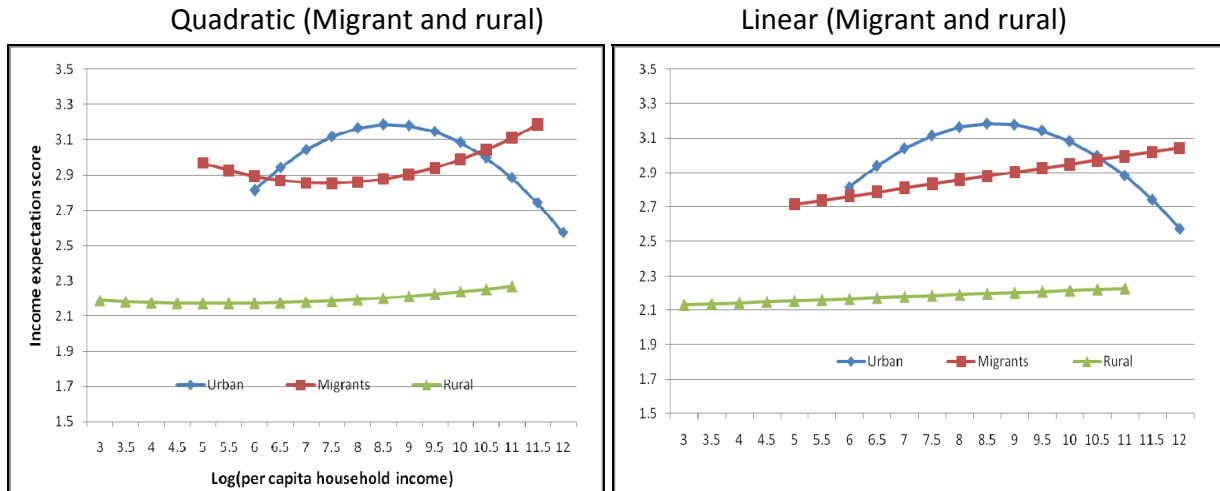


Figure 4: Predicted relationship between log per capita income and happiness

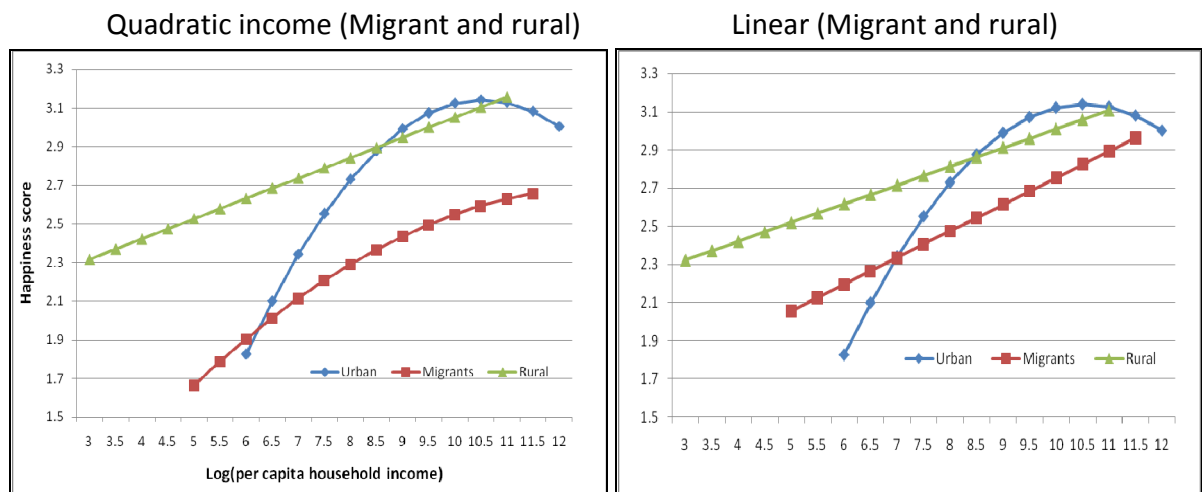
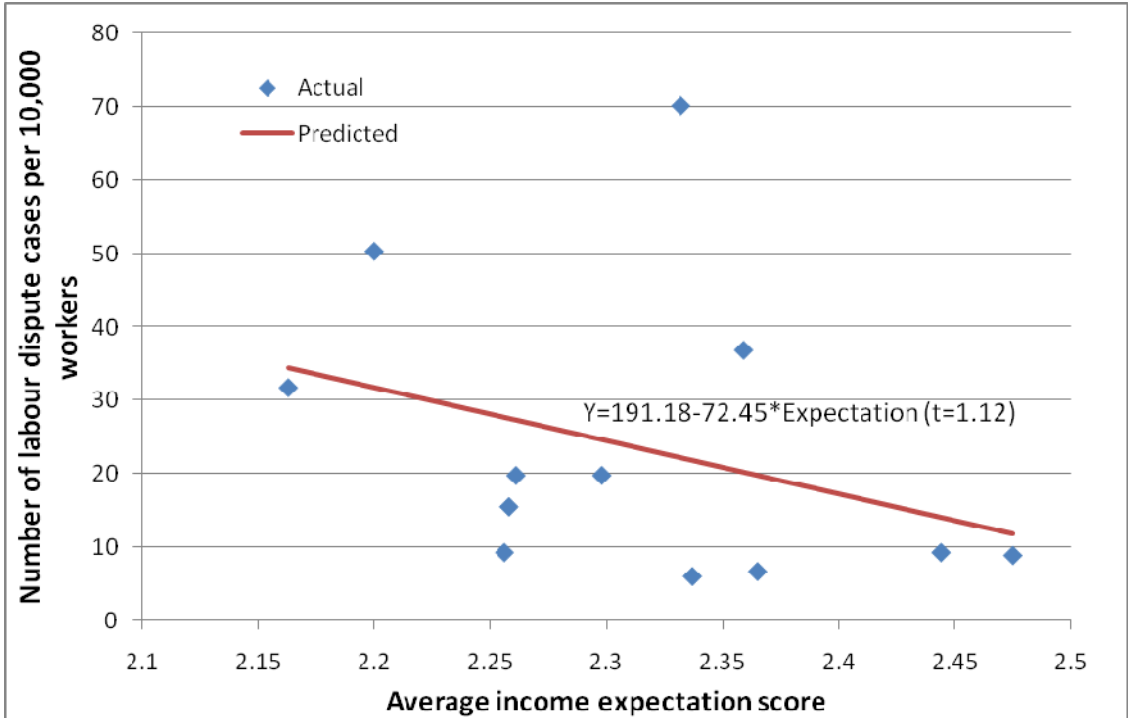


Figure 5: Relationship between average household income expectations and actual number of labour dispute cases



Appendix A: Determinants of income expectations, Ordered Probit

	Total sample		Urban sample		Migrant sample	Rural sample	
	Model 1	Model 1	Model 2	Model 1	Model 1	Model 2	
Log per capita income	-0.110 [0.139]	1.247 [0.475]***	1.265 [0.477]***	-0.575 [0.452]	-0.130 [0.224]	-0.254 [0.228]	
(Log per capita income) ²	0.009 [0.009]	-0.072 [0.027]***	-0.074 [0.027]***	0.039 [0.027]	0.010 [0.015]	0.016 [0.015]	
Income change (2001-2002)			-0.023 [0.045]			0.069 [0.033]**	
Income change (1999-2001)			0.242 [0.050]***			0.119 [0.032]***	
Income position in the city/village	0.765 [0.037]***	0.834 [0.061]***	0.783 [0.063]***	0.676 [0.111]***	0.746 [0.052]***	0.695 [0.053]***	
Good mood			0.199 [0.030]***			0.162 [0.027]***	
Dummy indicating hh with members migrated						0.151 [0.031]***	
Couple only households	-0.065 [0.032]**	0.069 [0.052]	0.076 [0.052]	0.064 [0.084]	-0.489 [0.059]***	-0.475 [0.059]***	
No. of children age 0-18 at home	0.029 [0.014]**	-0.004 [0.036]	-0.008 [0.036]	-0.034 [0.055]	0.027 [0.016]*	0.020 [0.016]	
No. of adults age >18	0.090 [0.012]***	0.084 [0.026]***	0.086 [0.026]***	0.035 [0.066]	0.061 [0.015]***	0.051 [0.015]***	
Dummy for married	-0.173 [0.045]***	-0.175 [0.072]**	-0.180 [0.072]**	-0.326 [0.124]***	-0.016 [0.071]	-0.006 [0.072]	
Own age	-0.039 [0.006]***	-0.081 [0.010]***	-0.079 [0.010]***	-0.032 [0.019]	0.021 [0.010]**	0.019 [0.010]*	
(Own age) ² /10	0.003 [0.001]***	0.007 [0.001]***	0.007 [0.001]***	0.002 [0.002]	-0.003 [0.001]***	-0.003 [0.001]***	
HH mean schooling year aged>20	0.015 [0.004]***	0.017 [0.006]***	0.018 [0.006]***	0.023 [0.011]**	0.008 [0.006]	0.010 [0.006]	
Own gender (male==1)	0.100 [0.021]***	0.131 [0.032]***	0.127 [0.032]***	0.068 [0.058]	0.008 [0.032]	0.012 [0.032]	
Own unemployment	-0.138 [0.053]***	-0.064 [0.063]	-0.056 [0.064]	-0.100 [0.241]	-0.159 [0.135]	-0.150 [0.135]	
Spouse being unemployed	-0.086 [0.047]*	-0.091 [0.053]*	-0.077 [0.053]	0.351 [0.205]*	-0.040 [0.125]	-0.015 [0.125]	
Own weekly working hours /10	-0.010 [0.006]*	-0.009 [0.015]	-0.010 [0.015]	-0.039 [0.013]***	0.002 [0.007]	0.004 [0.007]	
Dummy for own healthy	0.145 [0.024]***	0.112 [0.035]***	0.085 [0.036]**	0.333 [0.098]***	0.138 [0.036]***	0.129 [0.036]***	
Dummy for spouse being healthy	0.078 [0.024]***	0.126 [0.036]***	0.112 [0.036]***	0.010 [0.088]	0.046 [0.035]	0.038 [0.035]	
Dummy for own party membership	0.018 [0.023]	0.009 [0.032]*	0.050 [0.032]	0.128 [0.224]	-0.032 [0.035]	0.051 [0.064]	
Dummy for spouse party member	0.017 [0.030]	0.062 [0.035]	0.007 [0.035]	0.082 [0.145]	0.050 [0.064]	-0.027 [0.035]	
rural	0.407 [0.041]***						
urban	-0.362 [0.039]***						
Region							
Observations	17033	6489	6489	1841	8703	8703	

Standard errors in brackets

* significant at 10%; ** significant at 5%; *** significant at 1%

Appendix B: Determinants of happiness (Ordered Probit)

	Total sample		Urban sample		Migrant sample	Rural sample	
	Model 1	Model 1	Model 2	Model 1	Model 1	Model 2	
Expectation (=no change)	0.341 [0.031]***	0.391 [0.040]***	0.371 [0.040]***	0.528 [0.096]***	0.195 [0.063]***	0.183 [0.063]***	
Expectation (=slight improvement)	0.492 [0.030]***	0.533 [0.039]***	0.497 [0.039]***	0.533 [0.092]***	0.417 [0.059]***	0.380 [0.060]***	
Expectation (=sig. improvement)	0.744 [0.044]***	1.001 [0.101]***	0.967 [0.102]***	0.876 [0.131]***	0.624 [0.070]***	0.565 [0.071]***	
Log per capita income	0.137 [0.133]	1.499 [0.472]***	1.492 [0.479]***	0.585 [0.448]	0.053 [0.214]	0.135 [0.221]	
(Log per capita income) ²	0.003 [0.008]	-0.068 [0.027]**	-0.065 [0.027]**	-0.022 [0.026]	0.006 [0.014]	0.003 [0.014]	
Income change (2001-2002)			-0.230 [0.043]***			-0.100 [0.028]***	
Income position in the city/village	1.254 [0.036]***	1.234 [0.062]***	1.131 [0.064]***	1.059 [0.111]***	1.286 [0.050]***	1.102 [0.051]***	
Good mood			0.411 [0.031]***			0.698 [0.027]***	
Dummy indicating hh with members migrated						-0.058 [0.030]**	
Couple only households	0.006 [0.031]	-0.022 [0.051]	-0.026 [0.052]	0.031 [0.083]	0.108 [0.056]*	0.045 [0.057]	
No. of children age 0-18 at home	0.007 [0.013]	0.014 [0.036]	0.018 [0.036]	0.106 [0.054]**	-0.001 [0.015]	0.004 [0.016]	
No. of adults age >18	-0.007 [0.011]	0.010 [0.026]	0.020 [0.026]	0.009 [0.064]	0.002 [0.014]	0.009 [0.014]	
Dummy for married	0.293 [0.044]***	0.297 [0.071]***	0.274 [0.071]***	0.160 [0.121]	0.270 [0.067]***	0.287 [0.068]***	
Own age	-0.024 [0.006]***	-0.043 [0.010]***	-0.039 [0.010]***	-0.010 [0.019]	-0.021 [0.009]**	-0.017 [0.009]*	
(Own age) ² /10	0.003 [0.001]***	0.005 [0.001]***	0.004 [0.001]***	0.002 [0.002]	0.003 [0.001]***	0.003 [0.001]**	
HH mean schooling year aged>20	-0.001 [0.004]	-0.005 [0.006]	-0.007 [0.006]	-0.011 [0.010]	0.002 [0.006]	0.000 [0.006]	
Own gender (male==1)	-0.064 [0.020]***	-0.075 [0.031]**	-0.092 [0.032]***	0.047 [0.056]	-0.075 [0.030]**	-0.076 [0.030]**	
Own unemployment	-0.148 [0.051]***	-0.155 [0.062]**	-0.157 [0.063]**	0.211 [0.236]	-0.205 [0.129]	-0.103 [0.130]	
Spouse being unemployed	-0.105 [0.045]**	-0.065 [0.052]	-0.060 [0.052]	-0.406 [0.199]**	-0.179 [0.118]	-0.182 [0.119]	
Own weekly working hours /10	-0.010 [0.005]*	0.013 [0.014]	0.015 [0.015]	-0.006 [0.013]	-0.017 [0.007]**	-0.017 [0.007]**	
Dummy for own healthy	0.213 [0.023]***	0.197 [0.035]***	0.146 [0.035]***	0.036 [0.097]	0.239 [0.034]***	0.206 [0.035]***	
Dummy for spouse being healthy	0.165 [0.023]***	0.154 [0.036]***	0.122 [0.036]***	0.219 [0.086]**	0.180 [0.033]***	0.147 [0.034]***	
Dummy for own party membership	0.072 [0.022]***	0.075 [0.031]**	0.057 [0.035]***	-0.015 [0.141]	0.070 [0.060]	-0.046 [0.034]	
Dummy for spouse party member	0.080 [0.029]***	0.112 [0.035]***	0.102 [0.032]*	0.159 [0.219]	-0.043 [0.033]**	0.049 [0.061]	
rural	0.525 [0.040]***						
urban	0.173 [0.038]***						
Regions	Yes	Yes	Yes	Yes	Yes	Yes	
Observations	17033	6489	6474	1841	8703	8579	

Standard errors in brackets

* significant at 10%; ** significant at 5%; *** significant at 1%