## NBER WORKING PAPER SERIES

# ARE WHITES STILL "FLEEING"? <br> RACIAL PATTERNS AND ENROLLMENT <br> SHIFTS IN URBAN PUBLIC <br> SCHOOLS, 1987-1996 

Charles T. Clotfelter

Working Paper 7290
http://www.nber.org/papers/w7290

NATIONAL BUREAU OF ECONOMIC RESEARCH<br>1050 Massachusetts Avenue<br>Cambridge, MA 02138<br>August 1999

The views expressed herein are those of the authors and not necessarily those of the National Bureau of Economic Research.
© 1999 by Charles T. Clotfelter. All rights reserved. Short sections of text, not to exceed two paragraphs, may be quoted without explicit permission provided that full credit, including $\odot$ notice, is given to the source.

Are Whites Still "Fleeing"? Racial Patterns and Enrollment Shifts in Urban Public Schools, 1987-1996
Charles T. Clotfelter
NBER Working Paper No. 7290
August 1999
JEL No. R2, I2


#### Abstract

The effect of interracial contact in public schools on the enrollment of whites has been an important concern in assessments of desegregation since the 1970s. It has been feared that "white flight" -- meaning exit from or avoidance of racially mixed public schools -- could undermine the racial contact that desegregation policy seeks to enhance. This study examines this question using recent data. It also expands coverage from large urban districts to entire metropolitan areas, paying attention to the spatial context within which enrollment decisions are made. To do so, it examines data for 1987 and 1996 on racial composition and enrollment in all schools and school districts in 238 metropolitan areas. The study finds that white losses appear to be spurred both by interracial contact in districts where their children attend school and by the opportunities available in metropolitan areas for reducing that contact. Implications for metropolitan segregation are examined.


Charles T. Clotfelter
Box 90245
Duke University
Durham, NC 27708-0245
and NBER
Charles.Clotfelter@Duke.edu

# Are Whites Still "Fleeing"? Racial Patterns and Enrollment Shifts in Urban Public Schools, 1987-1996* 

Charles T. Clotfelter

In sharp contrast to the firestorm of controversy it generated into the 1970 s, school desegregation has receded from prominence as a national issue in the last two decades. One reason for its virtual disappearance has been the Supreme Court's clear resolve to place limits on aggressive efforts to achieve and maintain racial balance in public schools. ${ }^{1}$ Yet school desegregation survives as national policy, and most of the issues that surrounded its implementation in the 1960s remain important today. Among the most significant issues is what effect desegregation itself has on the tendency of white and middle class families to leave - or avoid - racially mixed, largely urban school districts. Such departures are important, of course, because they tend to undercut desegregation, to the extent that they result in greater racial disparities among districts.

The aim of this paper is to examine recent changes in the racial composition and enrollment levels in urban public schools and, in particular, the factors associated with white losses. These white losses, sometimes dubbed "white flight," arise not only when white families move from one district to another or enroll in private schools, but also when they simply avoid

[^0]moving into districts with high interracial contact. Utilizing data covering all public schools and districts in a sample of metropolitan areas, the study assesses trends over the period 1987 to 1996 for all major districts in these metropolitan areas, not just central city districts. It examines changes in white public school enrollments as well as changes in racial compositions, levels of interracial contact, and segregation. The data, which are taken from the Department of Education's Common Core of Data, include the enrollment by race and ethnic group of virtually every public school in 238 metropolitan areas in both 1987 and 1996. ${ }^{2}$ In contrast, the comparable data used in most previous research were collected only for selected districts, making it necessary to restrict attention to large urban, mostly central city districts. ${ }^{3}$ The data used in the current study cover all districts, thus making it feasible to obtain a complete picture of each metropolitan area's public schools and to examine patterns over time in those metropolitan areas. Thus it is possible to see if the kinds of patterns observed in the 1960s and 1970s are still evident, with the maturation of school desegregation as a policy. It is particularly interesting to trace the effects of recent demographic trends, including significant ethnic shifts arising from immigration, as well as public policies aimed at fostering integration. It will be important as well to reflect insofar as possible the tremendous heterogeneity among metropolitan areas - by size, region, age, number of central cities, and number of school districts. One dimension of heterogeneity is the physical distance separating districts. Fortunately, the data set's complete coverage of districts allows one to take into account the geographical proximity of districts to which whites might consider moving in order to escape high concentrations of nonwhites.

The study addresses two sets of questions. First, it will be useful to describe recent trends in enrollment and racial composition. Is the movement of whites out of central cities slowing or
accelerating? Are interracial contact and segregation increasing or decreasing? And, is there evidence that the South, whose previous patterns were deeply influenced by the existence of legal segregation in schools, is evincing patterns increasingly similar to the North? Second, it is important to raise again the perennial question of whether and how desegregation contributes to white enrollment losses. Is there evidence of tipping points in these losses, as previous research has suggested? ${ }^{4}$ Do white reactions to interracial contact differ by nonwhite group?

What makes these questions important to ask are the ramifications of racial patterns in schools for socially significant outcomes. Racial contact in schools may affect such things as the level and distribution of academic achievement in the population, racial attitudes, subsequent social and economic outcomes of students, and patterns of residential integration. ${ }^{5}$ These important issues, however, are beyond the scope of the present paper.

Section I provides background for the analysis, including a brief review of related statistical studies. Section II describes the data used and the variables defined in the current study. Section III presents and discusses regression analysis for a sample of districts, and section IV examines the implications for changes at the metropolitan level. There is a brief concluding section.

## I. Background

Although the landmark Brown v. Board of Education was handed down in 1954, it was not until the late 1960s that many Southern school districts had undertaken substantial efforts to desegregate their schools. But when desegregation did take place, it typically produced dramatic changes in interracial contact. Between 1968 and 1972 the percentage of black students going to
schools with 90 to 100 percent minority enrollments fell from 78 to 25 percent in the South, by far the biggest change in any region (Orfield 1983, p. 4). Concomitantly, the exposure of whites to blacks increased, largely but not exclusively in the South. ${ }^{6}$ In urban areas this desegregation often took the form of cross-town busing, a policy that was endorsed by the Supreme Court in its 1971 decision Swann v. Charlotte-Mecklenberg Board of Education.

## Research on Desegregation and "White Flight"

One major concern that developed in the wake of this and other federal court decisions was whether these actions would cause white families to leave desegregating districts, thereby undercutting the potential for racial integration. To examine the question of whether school desegregation itself caused "white flight,"7 James Coleman and two colleagues (Coleman et al.1975) examined enrollments in a sample of districts data over time, concluding that larger white enrollment declines were associated with declines in measured segregation and with higher proportions of black students. Although this conclusion met with initial opposition, the bulk of subsequent empirical research has supported the main thrust of that study.

With a few exceptions, the empirical studies of the effect of desegregation on white enrollment have utilized school-level data on enrollment by race, collected since 1967 by the Office for Civil Rights (OCR), data which allow the calculation of indices of interracial contact and segregation as well as changes in total enrollment. The unit of observation has generally been the school district, with most studies focusing on urban districts. In general, the studies have attempted to distinguish changes in white enrollment that are attributable to desegregation policies from those that would have occurred in their absence. That is, since whites have been
declining as a percentage of all public school students in the country, and many central cities have witnessed the decline in white population for years, it would be incorrect to attribute all of those declines to desegregation policies. Two studies that sought to assess the Coleman et al. results were Clotfelter (1979) and Farley, Richards, and Wurdock (1980), the former reanalyzing the original data set and the latter using a sample of 104 urban school districts. Both studies supported the basic Coleman et al. findings.

Subsequent studies likewise suggested that white enrollments are sensitive to the presence of nonwhites in the public schools. Wilson (1985) contended that racial contact, irrespective of the form of the policy, is the central cause of such losses. But a number of subsequent studies concluded that the form of the policy was important, with white losses being greater under mandatory desegregation plans than under voluntary plans. Welch and Light (1987), Rossell (1990, 1994), and Rossell and Armor (1996) all compared the effects of different desegregation techniques, and concluded that mandatory plans such as pairing or clustering lead to bigger white losses. Some recent evidence offered by Armor (1995, p.179) further supports the hypothesis that whites are sensitive to changes in racial composition in schools attended by their children. Focusing on formerly minority schools to which whites were assigned as part of mandatory desegregation plans, Armor shows abnormally high proportions of "no-shows" among the whites. Similar sensitivity to racial composition is apparent in Lankford and Wyckoff's (1997) study using the school choices of individual families in metropolitan areas in New York State. All of this empirical work supports the working assumption stated by James (1989, p. 966): "white parents make decisions based on the actual or potential exposure of their children to blacks, not how equally students are assigned to schools by race."

## Recent Patterns and Trends

Before analyzing recent data on school racial patterns and white enrollment changes, it is useful to consider broader patterns and trends in metropolitan areas. A basic fact is that public schools at the metropolitan level tend to be quite segregated, in that the observed patterns of enrollment by school depart markedly from racial balance. This divergence is greatest in the largest metropolitan areas. By region, metropolitan-level segregation tends to be most severe in the Northeast and Midwest, least in the West and South. The bulk of this observed segregation can be attributed to disparities in racial composition among the various school districts in metropolitan areas, as opposed to segregation within districts. ${ }^{8}$ Thus a considerable portion of existing school segregation in metropolitan areas is associated with segregated housing patterns. Combined with the Supreme Court's decision in the 1974 Milliken v. Bradley case, this residential segregation virtually guarantees public school segregation in urban America for the foreseeable future. Although Cutler, Glaeser and Vigdor (1999) find that neighborhood segregation declined after 1970 and Farley and Frey (1994) report some lessening of residential segregation in metropolitan areas during the 1980s (especially in younger, smaller areas), those residential patterns remain highly segregated. Moreover, Massey and Hujnal (1995) argue that segregation among jurisdictions has been increasing. Movements inside metropolitan areas continue to be dominated by suburbanization by the middle class, both white and nonwhite, leaving concentrations of poor people in parts of central cities.

Another set of trends with important implications are demographic. The racial and ethnic composition of the school-age population is changing. Owing to immigration and differences in birth rates, the black and Hispanic school-age population is growing faster than that of the white
population. As shown in Table 1, between 1986 and 1996, while the number of whites in public elementary and secondary schools was increasing by 3 percent, the number of black students rose by 14 percent, and the number of Hispanic students rose by 45 percent. According to Frey (1995), the recent, ethnically diverse immigration that underlies some of these changes has profoundly influenced patterns of internal migration, so that immigrants have come to supplant natives in some metropolitan areas. Combining these demographic trends with ongoing suburbanization, it should come as no surprise that public schools are becoming, on average, increasingly nonwhite. ${ }^{9}$ Beyond that, however, the implications for school segregation and white enrollments are by no means obvious.

Another set of changes, related to education policy rather than demography, arise from consolidations of school districts. Although the findings in the present study suggest that any movement in this regard is relatively minor, the temporal comparisons examined in this paper require consistency in the definition of districts over time. Thus it will be necessary to pay attention to consolidations and other changes in district boundaries.

A final trend with obvious relevance has to do with changes in attitude and race relations. Although changes in this area are nothing if not complex, it appears that, over the last two decades, interracial contact has slowly increased and white racial prejudice has declined steadily. ${ }^{10}$

## II. Data and Methodology

The principal source of data used in the present paper is the National Center for Education Statistics' Common Core of Data, which includes information on the racial
composition of individual schools. ${ }^{1}$ The research strategy for the current study was to collect information for similarly-defined metropolitan areas at two points in time in order to assess changes over that period. Data were not available for all metropolitan areas, however. The data are supplied voluntarily by states, and, while the number of participating states has increased steadily, a significant number did not participate during the 1980s. Weighing the advantages of a longer period against the disadvantages of a smaller sample size, the decision was made to use 1987 as the beginning year. The Common Core for the fall of that year lacked data for 17 states, compared to only two states that were missing in 1996. ${ }^{12}$

Calculations were made for metropolitan areas in both 1987 and 1996, using the component counties (or, in New England, towns and cities) for each defined by the Census in 1990. Thus, while the definitions of metropolitan areas are periodically updated, the present paper utilizes the definitions as of 1990 in order to achieve comparability over time. Given these definitions and the states for which data were available, sufficient data were available for 238 metropolitan areas to make comparable calculations for both 1987 and 1996. In 1996 these 238 areas contained almost 4,000 districts. As in the case of the metropolitan areas, it was desirable to have consistent definitions of the school districts over the period. However, this aim was frustrated by changes in district boundaries over time, most commonly consolidations of two or more 1987 districts into one 1996 district. Through careful accounting, comparable districts were formed, typically by combining the components of the consolidated district into a "virtual" district in 1987. These virtual districts constituted only about 2 percent of the districts in the sample. ${ }^{13}$

## Measures and Variables

The basic measure of white behavior is changes in white public school enrollment. This measure is necessarily a net measure, reflecting the difference between departures and new enrollments. As such, the measure cannot measure the absolute number of departures, nor can it identify whether departures take the form of moving to another district, enrolling in a private school, or simply graduation. Nor can it reflect the choice not to move into a district in the first place, a decision having much the same impact as a departure. Where $\mathrm{W}_{0}$ and $\mathrm{W}_{\mathrm{t}}$ are white enrollments in years 0 and t , the change in white enrollment is expressed as the exponential growth rate $g$ in the equation

$$
\mathrm{W}_{\mathrm{t}}=\mathrm{W}_{0} \mathrm{e}^{\mathrm{gt}} .
$$

Expressed as a percentage, the growth rate between 1987 and 1996 is
$100 \mathrm{~g}=100 \ln \left(\mathrm{~W}_{96} / \mathrm{W}_{87}\right) / 9$.
Measures of composition and segregation are based entirely on racial and ethnic categories used in the survey. No measure of economic status is included. ${ }^{14}$ The basic measure of interracial contact is the exposure rate of whites to nonwhites, which is the percent nonwhite in the average white student's school. It is defined as:

$$
\begin{equation*}
E=(1 / W) \sum_{i} W_{i}\left[N_{i} /\left(W_{i}+N_{i}\right)\right], \tag{1}
\end{equation*}
$$

where $W_{i}$ and $N_{i}$ are the number of whites and nonwhites, respectively, in school i and W and N are their totals for the district. If schools were racially balanced, each white child would attend a school whose racial composition was $\mathrm{PCN}=\mathrm{N} /(\mathrm{W}+\mathrm{N})$, the overall proportion of students who are nonwhite. The gap between this theoretical maximum and the actual rate of racial contact,
expressed as a proportion of the area's racial composition, represents one measure of the extent of segregation. ${ }^{15}$ This gap-based index,

$$
S=\left(P C N-E_{W N}\right) / P C N,
$$

ranges from zero, signifying perfect racial balance among schools, to one, signifying total segregation. Since the measures of exposure and segregation are based on school-level enrollment data, they do not measure racial contact in classrooms or in school groups. Nor do these measures differentiate between voluntary or mandatory desegregation plans.

To account for differences in overall growth among metropolitan areas, the population growth rate, expressed as a percentage, is included as an explanatory variable:

$$
\mathrm{PG}=100 \ln \left(\mathrm{P}_{90} / \mathrm{P}_{80}\right) / 10
$$

Finally, to account for the differing legal history of school segregation and other, otherwise unmeasured regional differences, metropolitan areas were assigned to one of five regions: South, Border, Northeast, Midwest, and West. ${ }^{16}$

Patterns and Trends in the Data
Before discussing the last set of variables, it is useful to examine the sample of 238 metropolitan areas for patterns and trends. It is helpful to begin by examining aggregate figures. Table 1 compares total public school enrollment for the nation with the enrollment in the 238 metropolitan areas analyzed in the current study. Not only does it give a sense of overall demographic changes over time, the table also allows an assessment of how representative the data used in the present analysis are. As indicated by the total enrollment figures, the present sample -- although it omits all schools outside of metropolitan areas as well as metropolitan areas
for which complete data were not available -- still covers more than half of all public school students. As expected, the racial composition of these metropolitan areas differs from the nation at large, featuring higher percentages of blacks, Hispanics, and other nonwhites than the U.S. as a whole. It is clear also that the racial composition of the nation and metropolitan areas is changing over time, with nonwhites assuming a continually increasing share. The reason behind this changing composition is illustrated in the table's last column, which lists average annual growth rates for each racial and ethnic group. Enrollment of whites grew at 0.5 percent a year, slower than the rate for blacks ( 2.0 percent) , and considerably slower than the 5.4 percent rate for Hispanics and the 4.4 percent for other nonwhite students.

To illustrate the sort of data examined, Table 2 presents several measures for 16 metropolitan areas of various sizes and from different regions. The table makes plain the great differences in size of metropolitan areas: public school enrollment in these 16 areas in 1987 ranged from 1.3 million in the 83 districts in the Los Angeles-Long Beach PMSA to 32,000 in the two districts covered by the Tallahassee MSA. In keeping with the findings of Clotfelter (forthcoming), larger metropolitan areas, and those in the Northeast and Midwest, tended to have more districts than smaller areas and those in the South and West. The areas also differed markedly in racial and ethnic composition. The percentage of students who were nonwhite (which includes Hispanic whites) ranged from 3 percent in Johnstown to 71 percent in Los Angeles-Long Beach, with other nonwhites outnumbering blacks in the latter. Blacks were relatively most numerous in the Southern areas. Columns F and G of the table give the average annual growth rate in white and nonwhite enrollment for metropolitan areas. Reflecting the national trends shown in Table 1, nonwhite enrollments were increasing in all these metropolitan
areas, while whites increased in only about two thirds of them.
Racial segregation at the metropolitan level differed quite a bit among the 16 areas shown, but the calculated indices for individual metropolitan areas did not change greatly over the period. Clotfelter (forthcoming) shows that segregation in the public schools is most pronounced in the largest metropolitan areas, and the calculations for the areas shown in this table bear out that generalization; Detroit, in fact, was found to have the most segregated schools among all metropolitan areas studied in 1994. However, some of the areas, in particular the seven with indices less than 0.20 , exhibited quite low levels of segregation, suggesting near racial balance throughout the public schools in each metropolitan area.

Because much of the interest in white withdrawals from public schools has focused on individual school districts, especially central city districts, it is useful to examine data for some illustrative districts as well. Table 3 presents information for the two largest districts in each of the 16 metropolitan areas shown in Table 2. For a majority of the metropolitan areas, the largest district also shares the name of the area's primary central city. In 1987 these districts ranged in size from 582,871 in Los Angeles to 3,104 in Somerset, PA. As a percentage of their respective metropolitan areas, they ranged from 74.1 percent in Hillsborough County, Florida to 2.9 percent in Lynn, Massachusetts. In most cases, the largest district in each metropolitan area had both a higher nonwhite percentage and a slower growth rate of white students than the metropolitan area as a whole. (Exceptions were in Raleigh-Durham and Tallahassee.) Their growth in nonwhite enrollments was less predictable, however, with 19 of the 32 districts having faster nonwhite growth than their metropolitan areas. Revealing their generally low levels of segregation, exposure rates were quite close to their overall nonwhite proportions.

The last three columns relate to the opportunities open to white families to reduce their children's exposure to nonwhites in the public schools by moving to other districts in the metropolitan area. As one simple measure of such opportunities, column $H$ shows, for each district, the number of districts in its metropolitan area that satisfied these conditions in 1987: enrollment 5,000 or more, no more than 10 miles away, and an exposure rate to nonwhites at least 10 percentage points less than the listed district. Using this criterion, one discovers that "whiter" alternative districts did exist for almost half of these districts, with as many as 21 such districts, in the case of Boston. The last two columns present more refined measures of access that are discussed in the following section.

To get a more representative picture of racial composition and white enrollment growth in districts, Table 4 presents weighted mean values by size and region of metropolitan area. The table shows clearly that the nonwhite percentage tended to rise with metropolitan area size: while the smallest metropolitan areas had enrollments that averaged 25 percent nonwhite, those in the largest metropolitan areas averaged 54 percent. By region, the components of the nonwhite percentage differed markedly, with blacks being most numerous in Midwest districts and Hispanics and other nonwhites most numerous in the West. Overall, white enrollments were declining over the nine-year period. They declined most rapidly in districts contained in the largest metropolitan areas, and they grew slightly in districts within the smallest metropolitan areas.

The Spatial Context of Enrollment Choices
An important aspect of enrollment shifts in metropolitan areas lies in their spacial
context. White families which, for whatever reason, desire to enroll their children in public schools with a smaller proportion of nonwhites usually can bring this about by moving to another district. If they are moving into a metropolitan area, they can simply avoid districts with high nonwhite enrollments. But locating in predominantly white districts may entail longer commutes to work. Thus the desirability of moving from or avoiding a given urban district will likely depend on the existence of alternative public school districts with lower nonwhite percentages, the extent of the difference in racial compositions, and the distance from that given district to the alternative districts.

To illustrate the variegated jurisdictional landscape in metropolitan areas, Table 5 presents data for the largest districts in three metropolitan areas. The largest of these three, Detroit, features a dominant central city district surrounded by a multitude of other districts, many of which would accurately be described as "suburban." True to the stereotype of big U.S. urban areas, the Detroit district is overwhelmingly nonwhite in composition, while most of the surrounding districts are predominantly white. Within the Detroit district, the racial makeup of individual schools deviated from racial balance, as indicated by the segregation index of 26.6. During the nine-year sample period, the number of whites in the Detroit district fell by a rate of more than 5 percent a year. White families in Detroit who wanted to move from or avoid moving into the Detroit district could certainly find alternative districts with lower exposure rates to nonwhites, but locating in one of the biggest of these would require being more than 10 miles away from the center of the Detroit district. Many virtually all-white enclaves did exist, however, among the more than 100 smaller districts not listed in the table.

In contrast to the dominating size of one central city district, the Anaheim metropolitan
area features much more balance among the top eight districts. The racial composition also differs markedly, with Hispanics and other nonwhites being much more numerous than blacks. Distances to the largest district, Santa Ana, tended to be less, but so too were the differences in potential exposure rates for whites. As in the Detroit area, the largest district was characterized by an exposure rate over 80 percent and a declining white enrollment. A third pattern is illustrated by the smaller Raleigh-Durham metropolitan area, which contains only five districts. The largest of these, the county containing Raleigh, features growing rather than declining white enrollment. It is the second-largest district, Durham, that is characterized by the high exposure rate and white decline usually associated with central city districts.

As these three examples amply illustrate, districts can differ greatly in how easy it is for white families to lower their exposure to nonwhites by avoiding or moving from some districts. As illustrated by the largest two districts in each of the 16 illustrative metropolitan areas shown above, comparatively white alternatives are not uncommon, but in some metropolitan areas they simply do not exist. As an illustration of how common such alternatives were in 1987, a tabulation showed that over a third of white students in the 975 largest districts in the current sample could have lowered their exposure rate by at least 0.10 by moving to a district within 10 miles. ${ }^{17}$ If one wishes to model white enrollment shifts, it is important to take such alternatives into account. While a few studies of school segregation and white flight have included variables that are designed to reflect the opportunities for avoidance on a very aggregated level within metro areas, ${ }^{18}$ none to my knowledge has incorporated data on distance among districts in a metropolitan area. In the present paper I propose two measures to characterize the options facing white families in metropolitan areas. Each of these measures is designed to account for three
spatially-related factors that would impinge on any white family's decision to move to another district, to avoid moving into a district in the first place, or to enroll in a private school:
a)How many other options for enrollment are available in the local housing market? One way of measuring this is by enrollment: how large are the whiter alternative districts compared to the enrollment of the family's own school system?
b)How close in distance are those other districts? If one must move a long way away to attend a different district, that might mean changing jobs or commuting farther, given that the current location is otherwise logical, if not optimal. The farther away, the less good the option.
c) How much can the exposure to nonwhites be reduced (the postulated objective for changing schools) by moving? If other districts offer similar rates of exposure, there would be little gain from moving no matter how far away they are.

Two variables of accessibility to whiter districts, which take these three factors into account, are defined below. District i is the "origin" district, from whose perspective the calculations are made; districts j are possible alternative districts. $\mathrm{D}_{\mathrm{ij}}$ is the distance between districts $i$ and $j . E_{i}$ is the average exposure rate of whites to nonwhites in district $i$ and $T_{i}$ is the enrollment in district $\mathrm{i} . \mathrm{E}_{\mathrm{j}}$ and $\mathrm{T}_{\mathrm{j}}$ are defined similarly. The bigger, the more numerous, and the closer the alternative districts featuring lower exposure rates to nonwhites, the greater are the options for reducing nonwhite exposure through relocation. The first measure selects as possible alternatives those districts that are within 10 miles distance and feature an exposure rate at least 10 percentage points $(0.10)$ less than the origin district. In order to account for the relative size of these possible destination districts, this measure of accessibility is based on the ratio of the enrollment of such districts to that of the origin district:

$$
A 1=\left[\sum_{j}\left(T_{j} / T_{i}\right)\right]^{-r} \text { for all } j \text { s.t. } E_{i}-E_{j} \geq 0.10 \text { and } D_{i j} \leq 10 \text { miles, }
$$

where $r$ is a factor reflecting the importance of alternative enrollments. If $r=1$, the accessibility proxy rises in proportion to the enrollment of neighboring districts offering the 10 percentagepoint potential reduction in exposure; if $r=0.5$, it rises with the square root. The threshold values of 10 miles for distance and 10 percent for exposure are quite arbitrary, of course. In order to account for distance and exposure on a continuous basis, a second measure adds these two factors multiplicatively:

$$
A 2=\left[\sum_{j}\left(T_{j} / T_{i}\right)\left(1 / D_{i j}\right)^{-a}\left(E_{i}-E_{j}\right)\right]^{-s}, \text { for }\left(E_{i}-E_{j}\right)>0
$$

where a is a constant indicating the effect of distance on the attractiveness of a particular alternative district and s is another constant.

Both A1 and A2 are proxy variables having no exact interpretation. Their values are arbitrary, depending on the parameters $\mathrm{a}, \mathrm{r}$, and s . All that can be said about these proxy variables is that a given district will have large values for them the more a district is surrounded by close, relatively large, and predominantly white districts. In the analysis described below, the values of a and s are both assigned 0.5 , while $\mathrm{r}=1$. In order to approximate distances between pairs of districts, I measure the distance between the centroids of the zip codes in which the district offices are located. ${ }^{19}$

To give a sense of the values for A 1 and A2, Table 3 presents calculations of these indices for the two largest districts in each of the 16 metropolitan areas listed in Table 2. Both indices suggest considerable divergence in accessibility to whiter enclaves among this group of 32 districts. The indices are highly correlated: the districts in which white families have the best
opportunities to lessen their exposure to nonwhites tend to have higher values of both indices, although the differences are less stark with A2, owing to the continuity built into it. ${ }^{20}$ Districts where whites are deemed to have more opportunities for avoidance are Lynn, MA, Aldine, TX, Boston, and Providence, all of which feature high exposure rates and proximity to districts with much lower ones. At the other end of the opportunity spectrum are large districts with relatively low exposure rates: Leon County, FL, Wake County, NC, and Warwick, RI. Despite the arbitrary nature of these two accessibility measures, it seems clear that they both reflect the same basic phenomenon. Not surprisingly, these two accessibility measures yield similar results in tabulations and estimation. In the estimates presented below, A1 is used because it is the simpler of the two measures, but the qualitative results are not affected by which measure is used.

## III. Analysis of White Enrollment Changes by District

A perusal of Table 3 suggests that the districts having the fastest white enrollment declines tended to be those with high proportions of nonwhites and more opportunities for avoidance of nonwhites. To see whether this relationship exists more generally, Table 6 shows the average growth rate of whites in 374 larger districts (all with school enrollments of 5,000 or more and at least 10 percent of its metropolitan area's enrollment), broken down by accessibility (A1) and whites' exposure rate to nonwhties. In line with the tendencies noted in the previous table, the growth rate of white enrollment tends to decline as both variables increase, suggesting that white enrollment trends over this period were influenced both by the "push" of interracial contact and the "pull" of nearby whiter school districts.

In order to examine more fully the forces that influence changes in white enrollment, I estimated equations explaining the growth rate in white enrollment between 1987 and 1996. Previous analysis of "white flight" from urban school districts focused almost entirely on large, central city districts, most of which were in large metropolitan areas. Indeed, Coleman (1975, p. 11) stated: "The flight from integration appears to be principally a large-city phenomenon." Because of the continuing significance of their enrollment patterns, these big urban districts should obviously remain a major focus of research. However, it seems possible if not likely that urban districts differ in the enrollment patterns they exhibit. For example, small metropolitan areas, owing to the comparative ease of traversing them, may evince different residential and enrollment patterns than large metropolitan areas. Another possibility worth examining is that enrollment patterns for districts which are small relative to their metropolitan areas might be easier to avoid or be otherwise different from dominant districts. Given the large number of metropolitan areas and districts encompassed by the current sample, it is possible to analyze separately districts according to metropolitan area size. It is also possible to distinguish between districts that account for a significant share of their metropolitan area's total enrollment (defined here as a 10 percent share) and those that are small in relation to the whole.

The 3,933 districts in the sample were divided into four samples. The first three samples were restricted to districts at least 5,000 . Not only are these the most important districts to study, measured enrollment growth for them is less subject to error or undetected changes in district definitions. The first sample includes those sorts of districts which have traditionally received the bulk of attention in studies of white enrollment losses; these are districts in metropolitan areas with public school enrollments of 50,000 or more which account for at least 10 percent of
the total enrollment of their respective metropolitan area. The second sample includes districts in metropolitan areas with enrollments smaller than 50,000, again with at least 10 percent of the metropolitan total. The third sample includes districts with less than 10 percent of their metropolitan enrollment; given the 5,000 minimum size, these districts were thus all in metropolitan areas of 50,000 or more. The remaining fourth sample, composed of districts smaller than 5,000 , is by far the most numerous, containing over three quarters of all the districts studied.

Previous econometric studies of white enrollment losses have sought to explain white losses, usually measured by percentage change, as a function of variables describing racial composition, segregation, and other factors thought to be important in the decisions of white families. One early and influential study, by Coleman et al. (1975), noted above, used panel data on a set of large urban districts to estimate the equation:

$$
\% \Delta W=a+b_{1} \Delta S_{b}+b_{2} P_{b}+b_{3} \ln T+\lambda Z+u
$$

where $\% \Delta \mathrm{~W}$ is the percentage change in white enrollment, $\Delta \mathrm{S}_{\mathrm{b}}$ is the change in the segregation index (using blacks in place of nonwhites), $\mathrm{P}_{\mathrm{b}}$ is the proportion of blacks in the district, T is district enrollment, Z is a vector of other variables, a , the $\mathrm{b}_{\mathrm{i}}$ 's, and the vector $\lambda$ are coefficients, and $u$ is an error term. As noted above, Coleman et al. found that greater white losses were associated with higher black proportions and decreases in segregation, leading to their conclusion that school desegregation was causing white flight, and was therefore contributing to the resegregation of schools. ${ }^{21}$

The present study employs a similar model, but with a wider variety of school districts and with a single cross-section of one-year changes in place of the pooled annual changes analyzed by Coleman et al. Change in white enrollment is measured by the exponential annual growth rate. ${ }^{22}$ In place of the district percentage black, I used the exposure rate of whites to nonwhites, because exposure, not overall district racial composition, affects the experiences of whites, and to reflect the large and growing significance of Latinos and Asian-Americans in urban school systems. Accessibility to whiter alternative districts is measured by A1, as described above. The growth rate of metropolitan area population from 1980 to 1990 is included to account for overall metropolitan growth. It seems reasonable to treat this growth as an exogenous variable. Finally, a set of dummy variables is included to reflect regional differences in the growth of white enrollments.

Table 7 presents for each of the samples of districts the mean values for the basic variables used in estimation. Among the four samples, the growth rate of white enrollment, as well as interracial exposure and accessibility, differs noticeably. The two samples of districts in the largest metropolitan areas, I and III, show declines in white enrollments on average, whereas white enrollments in the average district in sample II remained steady over the period and those in the smallest districts grew. The sample I districts had much higher rates of exposure of whites to nonwhites than those in II and III, fitting the common image of large central city districts, while the smallest districts displayed the lowest exposure rates. Accessibility to districts with lower exposure rates also differed, with the smaller districts in III and IV having the highest indices. It is also worth noting that samples III and IV have relatively fewer districts in the South, owing to the large average size of districts in that region.

Table 8 presents the basic model explaining growth in white enrollments estimated for each of the four samples of school districts. Since the measured index of segregation depends in large part on the district's desegregation policy, and that policy could be influenced by white enrollment trends, it is not altogether clear that the change in segregation should be treated as exogenous as Coleman et al. did. Accordingly, the basic model is estimated using instrumental variables as an alternative to ordinary least squares.

The similarity of the estimates across the four samples is quite striking. Three variables are consistently important in explaining white growth rates in this basic model, regardless of the sample or estimation technique employed. First, the exposure of whites to nonwhites in 1987 has large and statistically significant estimated coefficients. Equation (1), for example, implies that in the most important districts in the large metropolitan areas an increase in exposure of $0.10-$ say from 0.15 to 0.25 - in 1987 would have been associated with an acceleration of white losses, in the form of a decrease in the growth rate of white enrollment of -0.7 percent a year. The estimated effect is somewhat smaller in sample II, but it is considerably larger in samples III and IV, implying a responsiveness half again as large in the smaller and less "significant" districts. The strong effect of exposure in all four samples is very much in line with Coleman et al.'s results for the district's proportion of blacks and with James' (1989) stated assumption that whites respond to interracial contact, not segregation per se.

A second consistent finding, indicated by the negative estimated coefficients for A1, is that white losses were greater where there were more opportunities in the metropolitan area for whites to find districts with lower rates of exposure - the "pull" factor. This finding is, of course, quite complementary to the first. ${ }^{23}$ The third consistent finding is that white enrollment
trends were - not surprisingly - influenced by overall metropolitan growth. Where the metropolitan population was growing, white enrollments in large districts tended to grow rather than shrink, but where the metropolis was stagnant there was less impetus to maintain white enrollments. This correspondence is much higher in the smaller metropolitan areas of sample II and the small districts in sample IV.

The major policy variable in the equations is the change in the segregation index. Holding constant interracial exposure in 1987, an increase in segregation would be expected to hold whites and thus increase the white enrollment growth rate (or decrease the rate of loss). This is what Coleman et al. (1975) found in their study of larger urban districts. And it is also what is implied in equation (1) in Table 8, the OLS equation covering the sample most similar to the ones they analyzed. The estimated coefficient in that equation implies that an increase in the segregation index $S$ of 0.10 would have decreased the rate of white enrollment decline by 0.83 , say from its mean of -1.55 to -0.72 percent a year. The corresponding coefficients in the other equations are not statistically significant, however. The estimated effect of this variable loses its statistical significance when it is treated as endogenous in the instrumental variables estimation in equation (5), as it also does in equations (6) and (7). Equation (8) actually yields a significant negative coefficient, but comparison to the OLS version suggests that that coefficient is quite unstable. Unfortunately, the first stage equations in all these equations are poor predictors of the change in segregation, resulting in estimated coefficients that are generally small relative to their standard errors, though with little change in the other coefficients of interest.

For the most part, the equations reveal little regional variation in white enrollment growth apart from that which is explained by the other included regressors. Only in samples III and IV is
there a statistically significant regional effect: in those relatively small districts, whites were less likely to leave or avoid districts in the South.

Central to these estimates is the reaction of white parents to the presence of nonwhites in the schools their children attend. It is especially germane to ask at least two questions related to that reaction. First, are whites more sensitive to one group of nonwhites than to others? To examine this question the first four equations in Table 9 split up the exposure rate of whites to nonwhites into three components, corresponding to blacks, Hispanics, and other nonwhites. For all but sample II, an F-test rejects at the 99 percent level the hypothesis that all the coefficients are the same. In each of those cases, the estimates suggest that whites respond most sharply to exposure to blacks ${ }^{24}$

The other exposure-related question addressed in Table 9 is whether the reaction of whites to nonwhite exposure in the schools is nonlinear, specifically whether there is a "tipping point," or a threshold exposure rate beyond which white departures and avoidance accelerate. To allow for such nonlinearities, both cubic and spline functions were estimated, with much the same result. Plots of the former are shown in the table and illustrated in Figure 1, which gives the predicted rate of growth of whites, calculated at mean values, as a function of exposure to nonwhites in 1987. They are remarkably similar, implying growth in white enrollments where exposure rates are below about 0.25 and losses beyond that point. Over most of the range of exposure rates the rate of loss is approximately linear. There is no evidence of a threshold beyond which losses accelerate. In fact, losses are reversed somewhat at very high exposure rates in samples I and III.

The conclusion that arises from these regressions is that the phenomenon of "white
flight," the loss of whites from school districts featuring significant interracial exposure, first identified and studied in the 1970 s, was still at work in the 1990 s. To be sure, white enrollment losses are made more likely by a slowing growth rate among the white population nationally and in metropolitan areas that are not growing, but they continue to be stimulated by exposure to nonwhites in the public schools, especially where those rates become large. Such white losses are moderated by configurations of school districts that minimize the opportunities for avoidance, such as the large county-wide districts common in parts of the South and West. But in the absence of a dramatic reversal of the Supreme Court's ban on cross-district desegregation plans, as well as the continuing influx of nonwhite immigrants, we are likely to see a continuation of these trends.

Before turning to the implications of these findings for school segregation, it is instructive to compare them to the earlier study by Coleman et al. (1975). Like Coleman et al., I find that districts with higher rates of white enrollment losses also had higher exposure rates to nonwhites. Changes in segregation are less clearly implicated, although the OLS estimates for sample I are consistent with Coleman et al.'s earlier findings. In order to compare present patterns of white withdrawal with those observed by Coleman et al. two decades earlier, a similar specification was estimated for sample I, the sample most similar to the group of large districts analyzed in the earlier study. The racial composition and segregation variables were redefined using blacks instead of nonwhites, to conform to Coleman et al. Table 10 compares the Coleman et al. regressions with this reestimated one using sample I. For both change in segregation and exposure, the responses implied using the recent data in sample I are considerably smaller than those implied by the sample of largest districts in 1968-73. They are, however, much closer in
magnitude to those based on the sample of 46 smaller central city districts. In another specification, not shown here, the estimated effect of between-district segregation based on the current sample I is quite similar to that reported by Coleman et al. ${ }^{25}$ Like the regressions presented above, this comparison suggests that the factors contributing to white losses from urban public school districts have not changed substantially in the 20 years since this topic first arose as a serious policy concern. While the magnitude of white response is not as large as that which Coleman et al. found for the very largest central city districts, it remains substantial.

## IV. Implications for Segregation in Schools

The estimates presented above make clear that white losses from urban public schools are not evenly distributed, but rather are systematically related to interracial contact and the ease of avoiding that contact. This kind of systematic avoidance was documented in research done in the 1970s. The present paper shows that it remained an important phenomenon in the 1990 s. Since a principal concern about white enrollment losses has been that they would lead to resegregation, it is useful to conclude by considering the impact of these white enrollment trends on measured segregation.

Overall, public schools in metropolitan areas became more segregated between 1987 and 1996. Figure 2 shows the distribution of metropolitan areas by segregation index S. It reveals a perceptible shift to the right, with a decline in the number of metro areas in the lowest two categories and increases in most of the higher categories.

Table 11 presents a more detailed summary of changes in metropolitan-level school segregation. The entries give the weighted average of the segregation index S for the 238
metropolitan areas in the present sample, broken down by size and region. In addition, the segregation indices are decomposed into two components: the portion that is attributable to segregation between districts and that which is attributable to segregation within districts. ${ }^{26}$ The table's top row shows that, for the entire sample of 238 metropolitan areas, school segregation increased over the nine-year period; the average value of $S$ increased from 0.302 to 0.317 . As there was actually a decline in within-district segregation, this overall increase was entirely attributable to an increase in between-district segregation. In other words, the racial compositions of school districts tended to diverge over this period, a change that would have been aided by systematic white losses from racially mixed school districts. Table 11 makes plain the tendency noted in Clotfelter (forthcoming) for segregation to rise with the size of the metropolitan area; this relationship is clear for both 1987 and 1996. The last set of columns shows, however, that segregation grew in all but the largest metropolitan areas over this period. When the data are classified by region, the effect of the balkanized metropolitan areas of the Northeast and Midwest are evident: not only do those regions feature the highest rates of segregation, they also have the biggest increases in segregation - again attributable to growing disparities between districts. The South had the highest level of within-district segregation in both years. Overall, the picture that arises is one of entrenched segregation, caused mainly by racial disparities between districts rather than segregation within school districts, and steady increases in that segregation. This picture is very much in line with Coleman et al.'s (1975, p. 80) statement: "The emerging problem of school segregation in large cities is a problem of metropolitan area residential segregation, black central cities and white suburbs, brought about by a loss of whites from the central cities." Barring any change in the legal status of
metropolitan desegregation, the only prospect for a reduction in school segregation is a lessening of residential segregation, a possibility raised by the recent work of Farley and Frey (1994).

## V. Conclusion

This paper uses recent data to examine an old question: what factors are associated with white losses from urban public schools? Using data covering all public schools in 238 metropolitan areas in 1987 and 1996, the present analysis suggests that much the same set of forces were at work in the 1990s as in the 1970s. The rate of white loss is affected by the "push" or exposure to nonwhites as well as the attraction of more predominantly white districts elsewhere in the same metropolitan area. Since segregation within districts by 1996 was rather mild in most districts, the key element in predicting whether whites would rapidly abandon central city districts is the size and homogeneity of all the districts in a metropolitan area. In particular, where the dominant districts are large, the prospects for avoiding large white losses are good. Furthermore, these forces appear to work similarly both inside and outside the South. To be sure, the world of urban public schools did change over the two decades. The proportion of nonwhites grew, in significant part due to immigration. In addition, the relative affluence of those at the upper end of the income distribution rose at the same time that Catholic parochial schools were in decline, probably increasing the socioeconomic gap between public and private schools.

Not surprisingly, the current paper leaves some important questions unanswered. For example, the models used do not distinguish between residential location and private school enrollment as alternative avenues for white avoidance of racial exposure. The relative cost of
these options surely affects their use; this relative cost is only crudely proxied by the measures used here to reflect accessibility to white enclaves. Nor does the paper examine the effects of contact across socioeconomic groups, as opposed to the racial and ethnic groupings used here and elsewhere. In addition, the data used in the paper contain no information on racial contact within schools, which is affected by the extent of academic tracking. Nor do the measures used here distinguish between mandatory desegregation plans and the various alternative policies that have been used to desegregate schools. Such factors as these are likely to be important considerations for parents - both white and nonwhite - deciding where to send their children to school. Given the implications of these decisions for the racial composition and segregation of the public schools, research on this topic remains as important today as it was two decades ago.

## References

Armor, David J. 1995. Forced Justice: School Desegregation and the Law. New York: Oxford University Press.

Braddock, Jomills Henry II, Robert L. Crain, and James M. McPartland. 1984. "A Long-Term View of School Desegregation: Some Recent Studies of Graduates as Adults," Phi Delta Kappan 66 (December), 259-264.

Clotfelter, Charles T. 1978. "Alternative Measures of School Desegregation: A Methodological Note." Land Economics 54 (Aug.):373-380.

Clotfelter, Charles T. 1976. "The Detroit Decision and 'White Flight,'" Journal of Legal Studies 5 (January), 99-112.

Clotfelter, Charles T. Forthcoming. "Public School Segregation in Metropolitan Areas," Land Economics.

Clotfelter, Charles T. 1976. "School Desegregation, 'Tipping,' and Private School Enrollment," Journal of Human Resources 22 (Winter 1976), 29-50.

Clotfelter, Charles T. 1979. "Urban School Desegregation and Declines in White Enrollment: A Reexamination," Journal of Urban Economics 6, 352-370.

Coleman, James S. 1975. "Recent Trends in School Integration," Educational Researcher 4 (July-August), 3-12.

Coleman, James S. et al. 1975. Trends in School Segregation, 1968-73. Urban Institute Paper No. 722-03-01 (Aug.).

Cutler, David M., Edward L. Glaeser, and Jacob L. Vigdor. 1999. "The Rise and Decline of the

American Ghetto," Journal of Political Economy 107, 455.
DuBois, David L., and Barton J. Hirsch. 1990. "School and Neighborhood Friendship Patterns of Blacks and Whites in Early Adolescence." Child Development 61:524-536.

Farley, Reynolds, Toni Richards, and Clarence Wurdock. 1980. "School Desegregation and White Flight: An Investigation of Competing Models and their Discrepant Findings," Sociology of Education 53 (July), 123-139.

Farley, Reynolds, and William H. Frey. 1994. "Changes in the Segregation of Whites from Blacks During the 1980s: Small Steps Toward a More Integrated Society." American Sociological Review 59 (Feb.):23-45.

Frey, William H. 1995. "Immigration and Internal Migration 'Flight' from US Metropolitan Areas: Toward a New Demographic Balkanisation," Urban Studies 32, 733-757.

Fitzpatrick, Gary L. and Marilyn J. Modlin. 1986. Direct-Line Distances. Metuchen, N.J.: Scarecrow Press.

Hallinan, Maureen T. 1982. "Classroom Racial Composition and Children's Friendships." Social Forces 61 (Sept.):56-72.

James, David R. 1989. "City Limits on Racial Equality: The Effects of City-Suburb Boundaries on Public-School Desegregation, 1968-1976," American Sociological Review 54 (December), 963-985.

Jencks, Christopher and Meredith Phillips (eds.). 1998. The Black-White Test Score Gap (Washington: Brookings).

Lankford, Hamilton, and James Wyckoff. 1997. "The Effect of School Choice and Residential Location on the Racial Segregation of K-12 Students." Unpublished Paper, University of

Albany (June).
Massey, Douglas S. and Zoltan L. Hajnal. 1995. "The Changing Geographic Structure of BlackWhite Segregation in the United States," Social Science Quarterly 76 (September), 527542.

Orfield, Gary. 1983. Public School Desegregation in the United States, 1968-1980.

Washington, DC: Joint Center for Political Studies.
Orfield, Gary et al. 1997. "Deepening Segregation in American Public Schools." Harvard

Project on School Desegregation, Unpublished Paper, Harvard University (April 5).
Orfield, Gary, and Susan E. Eaton. 1996. Dismantling Desegregation: The Quiet Reversal of Brown v. Board of Education. New York: The New Press.

Orfield, Gary, and Frank Monfort. 1992. Status of School Desegregation: The Next Generation. Metropolitan Opportunity Project, Harvard University (March).

Rivkin, Steven G. 1994. "Residential Segregation and School Integration." Sociology of Education 67 (Oct.):279-292.

Rossell, Christine H. 1990. "The carrot or the Stick for School Desegregation Policy?" Urban Affairs Quarterly 25 (March), 474-499.

Rossell, Christine H. 1994. "The Progeny of Brown: From the Old Freedom of Choice to the New Freedom of Choice in Four Decades.", Urban Geography 15:435-453.

Rossell, Christine H., and David Armor. 1996. "The Effectiveness of School Desegregation Plans, 1968-1991." American Politics Quarterly 24 (July):267-302.

Smock, Pamela J. and Franklin D. Wilson. 1991. "Desegregation and the Stability of White Enrollments: A School-Level Analysis, 1968-84," Sociology of Education 64 (October),

Tuch, Steven A., Lee Sigelman, and Jason A. MacDonald. 1999. "Race Relations and American Youth, 1976-1995," Public Opinion Quarterly 63, 109-148.
U.S. Office for Civil Rights. 1974. Directory of Public Elementary and Secondary Schools in Selected Districts, Fall 1972. Washington: Government Printing Office, OCR74-5.

Welch, Finis. 1987. "A Reconsideration of the Impact of School Desegregation Programs on Public School Enrollment of White Students, 1968-76," Sociology of Education 60 (October), 215-221.

Welch, Finis, and Audrey Light. 1987. New Evidence on School Desegregation. Washington: United States Commission on Civil Rights (June).

Wilson, Franklin D. 1985. "The Impact of School Desegregation Programs on White PublicSchool Enrollment, 1968-1976," Sociology of Education 58 (July), 137-153.

## Endnotes

1. In particular, the Court in Board of Education of Oklahoma City v. Dowell (1991) and Freeman v. Pitts (1992) laid down conditions whereby districts that had been under court order could end affirmative efforts to desegregate their schools. For discussions of the legal issues, see Armor (1995, pp. 3-8, 17-58) and Orfield et al. (1997, pp. 6-7).
2.The sample excludes special, vocational, or alternative schools, and districts operated by the state or federal government. In the terms defined in the Common Core of Data, the present sample includes type 1 (regular) schools and districts of types 1-4. In addition, two districts for which no enrollment data were reported for 1996, East Cleveland City( Ohio) and Hillsboro UHS (Oregon) were excluded.
2. The sample of districts employed in the Office for Civil Rights surveys changed over time and was based on several different criteria, including whether districts were under court order, the coverage of minority enrollments, and the ability to project sample findings to national totals. For a description of the sampling criteria, see, for example, U.S. Office for Civil Rights (1974). Studies that have employed these data include Coleman et al (1975), Orfield (1983), and Welch and Light (1987).
3. Jencks and Phillips (1998, p. 45) state, for example: "once black enrollment in a neighborhood school expands past something like $20 \%$, most white parents become reluctant to move into the neighborhood." For references to earlier studies of this phenomenon, see Clotfelter (1976).
4. For a discussion of these effects, see Braddock, Crane and McPartland (1984) or Clotfelter (forthcoming).
5. For example, between 1968 and 1972, the percentage black in the typical white child's school rose from 4 to 12 percent in Dallas, 9 to 14 percent in Little Rock, and 9 to 48 percent in Norfolk.
Increases over the same period in cities Increases over the same period in cities outside the South were: 7 to 14 percent in Dayton, 6 to 14 percent in Denver, and 15 to 44 percent in Pasadena (Smock and Wilson 1991, Table 3).
6. As used by some researchers, this otherwise pejorative term does have a precise meaning. It is the loss of white students over and above that which would have been predicted simply on the basis of demographic factors alone.
7. For a fuller description of current patterns of school segregation, see Clotfelter (forthcoming).
8. Following the common usage of the terms, in this paper "white" refers to non-Hispanic whites; thus "nonwhite" refers to all others.
9. These trends are illustrated by annual surveys of high school seniors done as part of the Monitoring the Future project. The percentage of white seniors who reported that they "do things (conversation, eating together, playing sports) with people of other races" increased from an average of 50 percent in 1976-78 to 65 percent in 1993-95. (The percentage for blacks stayed about the same.) The percentage of white seniors who felt it would be desirable for their
"(future) children go to schools where some of the children are of other races" increased over the period from an average of 28 percent to 31 percent. Interestingly, the comparable average percentage for blacks declined, from 37 to 31 percent over the period. (Tuch, Sigelman, and MacDonald 1999, pp. 126-126, 143-144.
10. Specifically, the data were taken from Public Education Agency Universe and Public School Universe of the Common Core of Data (http:/nces.ed.gov/ccd/ and http./nces.ed.gov/surveys/SDDB/introd.html). Enrollment by racial group (American Indian, Asian or Pacific Islander, Hispanic, Black, non-Hispanic, White, non-Hispanic) was available for all schools for the academic year 1996-97. Virtually all districts reported consistent and clean data. A few did not; the sums for schools did not match the totals reported for districts. For those whose school-level data gave different sums, I based all calculations on the school-level data. Some districts reported no school-level data whatsoever; these had to be dropped.

A handful of (33) state-wide schools (such as the N.C. School of Science and Math) are listed as districts. Since they should not be considered to be part of the metropolitan areas where they are located, they were dropped from the sample.
12. Because one of the missing states in 1996, New Jersey, was not missing for 1987 , there were a total of 18 omitted states for the matched sample.
13. Data from the Common Core data sets for 1987-88 and 1996-97 were matched by district, and organized by county and metropolitan area. In an effort to see if there were errors in the data, printouts of schools by district were examined in detail. Growth in total enrollments by county were compared to growth in population. Where they deviated significantly or where district enrollments or numbers of schools differed greatly or where districts disappeared or appeared from the first to the second year, school names were used in matching to determine how the districts compared. In a few cases, schools in two districts were reorganized into two new districts, creating for the present sample two virtual districts, defined according to the schools each contained in 1996. Of the 875 districts with enrollments of 5,000 or more, there were 18 virtual districts created due to consolidation. Those that underwent consolidation were, on average, smaller and had lower nonwhite percentages, but there was no significant difference in the change in their segregation indices.
14. Counts of students receiving free and reduced price lunches are available in the data set, but this information was not used because of variations across states in eligibility criteria and the likelihood that existing criteria are unevenly applied.
15. This measure, denoted $R$, is used by Coleman et al. (1975). For a discussion of this measure and its relationship to measures of exposure, see Clotfelter (1978).

It is useful to note that the value of $S$ is invariant with respect to which of two groups is used as the basis for calculating the exposure rate. That is, S can be calculated using the exposure of nonwhites to whites, where W , the overall percentage of students who are white, is the maximum for this exposure rate: $\mathrm{S}=\left(\mathrm{W}-\mathrm{E}_{\mathrm{NW}}\right) / \mathrm{W}$.
16. Following Orfield and Monfort (1992, p. 2), I defined regions as follows: South: Alabama, Arkansas, Florida, Georgia, Louisiana, Mississippi, North Carolina, South Carolina, Tennessee,

Texas, Virginia; Border: Delaware, District of Columbia, Kentucky, Maryland, Missouri, Oklahoma, West Virginia; Northeast: Connecticut, Maine, Massachusetts, New Hampshire, New Jersey, New York, Pennsylvania, Rhode Island, Vermont; Midwest: Illinois, Indiana, Iowa, Kansas, Michigan, Minnesota, Nebraska, North Dakota, South Dakota, Ohio, Wisconsin; West: Arizona, California, Colorado, Idaho, Montana, Nevada, New Mexico, Oregon, Utah, Washington, Wyoming. In the 13 cases where a metropolitan area (MSA or PMSA) had components from two regions, I classified it in the region containing the largest enrollment.
17. Tabulation based on the 975 largest districts in the current sample. 77 percent of whites could have moved to a district with an exposure rate to nonwhites of 0.20 or less.
18. Coleman et al. (1975) include a measure of inter-district segregation to reflect racial disparities, and Clotfelter (1979) included the percentage of the metropolitan area in the central city district.
19. Where L 1 and L 2 are the latitudes of the centroids of the ZIP codes corresponding to districts 1 and 2 and DL is the difference in longitude between those centroids, the distance between district 1 and 2, measured in degrees of arc distance, is D in the equation:

$$
\operatorname{Cos} D=(\operatorname{Sin} L 1)(\operatorname{Sin} L 2)+(\operatorname{Cos} L 1)(\operatorname{Cos} L 2)(\operatorname{Cos} D L)
$$

(Fitzpatrick and Modlin 1986, p.XI). The distance in miles is 69.16 D.
The use of centroids for ZIP codes of the district offices generally yields locations that are quite central to the population center of each district, but not always. An example where this approach does not work as well is Chapel Hill, N.C., for which the centroid of the district office's ZIP code (27516) lies altogether outside the district boundaries.
20. The correlation between A 1 and A 2 in the three samples discussed below are $0.86,0.85$, and 0.86 .
21. The estimates of Coleman et al. (1975) are discussed below.
22. For comparison, Appendix Table A4 presents equations employing a slight modification of the Coleman et al. specification, wherein the percentage change in whites is replaced by its growth rate. Like the estimates obtained in the 1975 study, both change in segregation and district black proportion show up as statistically significant in all three equations. White losses are spurred by the black proportion and declines in measured segregation.
23. Regressions using A2 rather than A1 produce quite similar estimates.
24. The calculated test statistics based on the sum of squared residuals were $5.2,0.3,6.3$, and 45.1 in the four equations. The corresponding critical value for two restrictions at the 99 percent level of confidence ranges from 4.60 to 4.75 .
25. The variable RSMSA, which is the between-district segregation index for the metropolitan area based on blacks and whites, operates like A1 to indicate the availability of predominantly
white districts with the metropolitan area. Whereas the coefficient on this variable reported by Coleman et al is -21.0 and -10.2 in their two samples, it was -12.6 in the comparable equation estimated for sample I.
26. As shown in Clotfelter (forthcoming), the segregation measure S can be decomposed in the following way. Consider the hypothetical exposure rate for the metropolitan area that would occur if each district were to racially balance its schools. Just as any district's racial composition (measured by the percent nonwhite, N ) represents the maximum attainable exposure rate of whites to nonwhites, the maximum exposure rate for the metropolitan area that could be achieved within the constraints imposed by the existing racial compositions of school districts this hypothetical rate. Where this hypothetical exposure rate is $\mathrm{E}^{*}$, the gap that is due to inter-district disparities between districts is $S_{1}=\left(N-E^{*}\right) / N$. The gap due to segregation within districts is $S_{2}=S-S_{1}=\left(E^{*}-E\right) / N$, that is, the difference between the exposure rate if all districts were racially balanced and the actual exposure rate, as a proportion of the overall nonwhite proportion.

# Are Whites Still "Fleeing"? Racial Patterns and Enrollment 

Shifts in Urban Public Schools, 1987-1996

Tables and Figures


Table 2
Illustrative Data for 16 Metropolitan Areas

|  | A | B | C | D | E | F | G | H | 1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Metropolitan area | Number of districts | Metro enrollment 1987 | $\frac{198}{\text { Black }}$ | 7 percent Other nonwhite | e: $\qquad$ <br> Tota nwhite | 1987-199 rate in e White | 96 growth nrollment Nonwhite |  | $\frac{1}{1996}$ |
| Los Angeles--Long Beach, CA PMS | 83 | 1,301,780 | 0.14 | 0.57 | 0.71 | -2.1 | 2.8 | 0.35 | 0.34 |
| Detroit, MI PMSA | 109 | 712,284 | 0.28 | 0.03 | 0.31 | 0.1 | 1.0 | 0.73 | 0.72 |
| Houston, TX PMSA | 41 | 591,404 | 0.24 | 0.28 | 0.52 | 0.2 | 4.1 | 0.39 | 0.41 |
| Boston, MA PMSA | 113 | 353,727 | 0.11 | 0.09 | 0.21 | 0.2 | 3.7 | 0.46 | 0.47 |
| Tampa-St. Petersburg | 4 | 244,906 | 0.17 | 0.07 | 0.24 | 1.4 | 5.5 | 0.13 | 0.19 |
| Milwaukee, WI PMSA | 51 | 210,975 | 0.24 | 0.07 | 0.31 | 0.0 | 3.7 | 0.45 | 0.55 |
| Portland, OR PMSA | 63 | 187,371 | 0.05 | 0.08 | 0.13 | 1.2 | 5.7 | 0.16 | 0.13 |
| Cleveland, OH PMSA | 54 | 249,729 | 0.29 | 0.03 | 0.32 | 0.8 | 1.5 | 0.59 | 0.64 |
| Fresno, CA MSA | 46 | 126,694 | 0.06 | 0.51 | 0.57 | -0.3 | 5.1 | 0.26 | 0.27 |
| Raleigh--Durham, NC MSA | 7 | 102,132 | 0.34 | 0.02 | 0.36 | 2.7 | 3.9 | 0.17 | 0.14 |
| Providence, RI PMSA | 23 | 86,231 | 0.07 | 0.09 | 0.16 | -0.4 | 5.4 | 0.44 | 0.55 |
| Davenport--Rock Island--Moline, | 23 | 62,343 | 0.09 | 0.05 | 0.14 | -0.8 | 2.6 | 0.17 | 0.17 |
| Rockford, IL MSA | 13 | 45,165 | 0.13 | 0.06 | 0.19 | 0.1 | 4.4 | 0.19 | 0.19 |
| Eugene--Springfield, OR MSA | 16 | 44,121 | 0.01 | 0.05 | 0.06 | 0.3 | 5.7 | 0.02 | 0.02 |
| Johnstown, PA MSA | 23 | 36,807 | 0.02 | 0.00 | 0.03 | -0.9 | 1.3 | 0.14 | 0.11 |
| Tallahassee, FL MSA | 2 | 32,366 | 0.48 | 0.02 | 0.50 | 1.4 | 2.4 | 0.32 | 0.36 |

Source: U.S. Department of Education, Common Core of Data, 1987 and 1996; author's calculations.

123lx\metrolm30x
6/28/99

(a) Exposure rate of whites to nonwhites. See text.
(b) Measures of accessibility A1 and A2 are defined in the text. Number refers to the number of districts less than 10 miles away and with an exposure rate of whites to nonwhites at least 0.10 less than the listed district.
Source: Common Core of Data; author's calculations.

Table 4

## Growth Rate in White Enrollment at Larger Urban Districts


All
875
0.18
0.17
0.05
0.40
$-1.06$

By size of metropolitan area enrollment

| $5,000-50,000$ | 187 | 0.14 | 0.09 | 0.03 | 0.25 | 0.12 |
| :--- | :--- | :--- | :--- | :--- | :--- | ---: |
| $50,001-150,000$ | 240 | 0.19 | 0.15 | 0.03 | 0.37 | -0.45 |
| $150,001-350,000$ | 241 | 0.16 | 0.14 | 0.06 | 0.36 | -0.82 |
| $>350,000$ | 207 | 0.22 | 0.26 | 0.07 | 0.54 | -2.54 |

By region

| Border | 24 | 0.19 | 0.02 | 0.04 | 0.24 | -0.60 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- |
| Midwest | 255 | 0.27 | 0.05 | 0.03 | 0.35 | -1.26 |
| Northeast | 78 | 0.19 | 0.11 | 0.03 | 0.33 | -1.53 |
| South | 200 | 0.22 | 0.19 | 0.02 | 0.43 | -0.29 |
| West | 318 | 0.09 | 0.24 | 0.10 | 0.42 | -1.58 |

Source: U.S. Department of Education, Common Core of Data, 1987 and 1996; author's calculations tablem36.sas 5/12/99 13:03

M36
6/28/99


| District exposure rate of whites to nonwhites, 1987 | Accessibility to districts with lower exposure rates (A1) |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | NONE <br> 0 <br> (No such districts) | MEDIUM <br> 0.8 or less | HIGH Greater than 0.8 | ALL |
| LOW |  |  |  |  |
| 10\% or less | $\begin{array}{r} 1.10 \\ 99 \end{array}$ | (b) | (b) | $\begin{array}{r} 1.10 \\ 99 \end{array}$ |
| MEDIUM |  |  |  |  |
| 10\%, less than |  |  |  |  |
| or equal to $30 \%$ | $\begin{array}{r} 1.31 \\ 69 \end{array}$ | $\begin{array}{r} -0.99 \\ 39 \end{array}$ | $\begin{array}{r} -1.45 \\ 22 \end{array}$ | $\begin{array}{r} 0.46 \\ 130 \end{array}$ |
| HIGH |  |  |  |  |
| Greater than |  |  |  |  |
| 30\% | -1.53 | -2.92 | -3.20 | -2.68 |
|  | 35 | 45 | 65 | 145 |
| ALL | 0.50 | -2.52 | -2.84 | -1.15 |
|  | 203 | 84 | 87 | 374 |

Note: The sample consists of districts in the 238 metropolitan areas which had enrollments in 1987 of 5,000 and which had a least 10 percent of the metropolitan area's growth rate of white enrollment, weighted by district enrollment. The top number in each cell is the weighted average annual growth rate of white enrollment, weighted by district enrollment. The number of districts in each cell is below the mean in italics.
(a) For each district, A1 is positive only if there existed other districts in its metropolitan area not more than 10 miles away with a white-nonwhite exposusre rate at least 0.10 less than the district in question. A1 is the square root of the ratio of the enrollments of all such districts to the enrollment of the district in question. See text.
(b) No districts.

5/26/99 13:12
M39
6/28/99

Table 7
Sample Means of Variables for Districts, Weighted by Enrollment

| Sample | 1 | II | III | IV |
| :---: | :---: | :---: | :---: | :---: |
| Metropolitan enrollment | $\begin{gathered} 50,000 \text { or } \\ \text { more } \end{gathered}$ | $\begin{aligned} & \text { Under } \\ & 50,000 \end{aligned}$ | $\begin{aligned} & 50,000 \text { or } \\ & \text { more } \end{aligned}$ | All |
| District enrollment | 5,000 or more | 5,000 or more 5,000 or more Less than 5,000 |  |  |
| District share of metro enrollment | 10\% or more | 10\% or more | Under 10\% | - |
| N | 187 | 187 | 501 | 3,058 |
| Growth rate of white enrollment, 1987-1996 | -1.54 | 0.12 | -0.90 | 0.56 |
| 1987 exposure rate of whites to: |  |  |  |  |
| Nonwhites | 0.41 | 0.23 | 0.28 | 0.14 |
| Blacks | 0.19 | 0.12 | 0.07 | 0.05 |
| Hispanics | 0.17 | 0.08 | 0.14 | 0.07 |
| Other nonwhites | 0.05 | 0.03 | 0.06 | 0.02 |
| Accessibility to other districts with lower exposure rates |  |  |  |  |
| A1 | 0.40 | 0.24 | 0.72 | 0.55 |
| A2 | 0.28 | 0.14 | 0.58 | 0.57 |
| Change in segregation ( S ) | 0.001 | 0.005 | 0.013 | 0.001 |
| Metropolitan area growth rate, 1980-1990 | 1.59 | 1.27 | 1.74 | 0.79 |
| Region |  |  |  |  |
| South | 0.38 | 0.41 | 0.20 | 0.12 |
| West | 0.33 | 0.18 | 0.49 | 0.20 |
| Midwest | 0.21 | 0.31 | 0.24 | 0.43 |
| Border | 0.03 | 0.01 | 0.02 | 0.02 |
| Northeast (excluded category) | 0.05 | 0.09 | 0.05 | 0.23 |
| 5/26/99 10:49; 6/10/99 19:40 |  |  |  |  |
| M35 6/28/99 |  |  |  |  |

Table 8
Estimated Equations Explaining Growth Rate in White
Enrollment, 1987-1996

| Sample | I | II | III | IV | I | II | III | IV |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Equation | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ | $(7)$ | $(8)$ |
| Estimation | OLS | OLS | OLS | OLS | I.V. | I.V. | I.V. | I.V. |

Variable

| Intercept | 1.59 | 0.91 | 1.15 | 1.25 | 1.35 | 0.86 | 1.55 | 1.22 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.55 | 0.31 | 0.45 | 0.11 | 0.61 | 0.33 | 0.56 | 0.13 |
| Percent nonwhite, 1987 | -6.99 | -5.60 | -9.67 | -9.59 | -7.46 | -5.51 | -11.32 | -10.02 |
|  | 0.62 | 0.65 | 0.79 | 0.37 | 0.67 | 0.68 | 1.27 | 0.45 |
| Accessibility (A1) | -0.94 | -0.95 | -0.58 | -0.12 | -0.99 | -0.93 | -0.56 | -0.13 |
|  | 0.29 | 0.26 | 0.13 | 0.04 | 0.31 | 0.27 | 0.15 | 0.04 |
| Change in segregation (S) | 8.31 | 3.13 | -2.55 | -1.54 | -1.97 | 7.46 | -34.48 | -62.71 |
|  | 2.04 | 1.83 | 2.46 | 2.00 | 7.56 | 9.07 | 17.26 | 25.19 |
| Metropolitan area growth rate,1980-1990 | 0.38 | 0.79 | 0.30 | 0.81 | 0.43 | 0.76 | 0.29 | 0.88 |
|  | 0.14 | 0.08 | 0.12 | 0.06 | 0.15 | 0.10 | 0.14 | 0.08 |
| RegionSouth |  |  |  |  |  |  |  |  |
|  | 0.10 | -0.28 | 1.72 | 1.17 | 0.22 | -0.21 | 2.51 | 1.09 |
|  | 0.60 | 0.37 | 0.57 | 0.22 | 0.64 | 0.41 | 0.78 | 0.25 |
| West | -1.03 | -0.52 | 0.60 | 0.31 | -0.62 | -0.50 | 1.18 | 0.60 |
|  | 0.59 | 0.37 | 0.55 | 0.22 | 0.69 | 0.38 | 0.71 | 0.27 |
| Midwest | -0.66 | -0.29 | 0.05 | -0.05 | -0.16 | -0.28 | -0.01 | 0.03 |
|  | 0.54 | 0.31 | 0.49 | 0.14 | 0.68 | 0.32 | 0.57 | 0.16 |
| Border | -0.94 | -0.31 | 0.30 | 0.09 | -0.71 | -0.28 | 0.57 | -0.49 |
|  | 0.77 | 0.75 | 0.82 | 0.32 | 0.84 | 0.77 | 0.97 | 0.44 |
| Adjusted R-square | 0.63 | 0.64 | 0.55 | 0.31 | 0.59 | 0.63 | 0.47 | 0.25 |

Note: Numbers below coefficients are standard errors. Bold signifies coefficients that are significantly different from zero at the 95 percent level. Regressions are weighted by the square root of 1987 district enrollment. For definition of samples, see Table 7.

5/19/99 19:29; 6/10/99 19:40; 6/11/99 17:11
M29
7/27/99

Table 9

## Estimated Coefficients for Selected Variables

| Equation | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sample | I | II | III | IV | ( | II | III | IV |
| Variable |  |  |  |  |  |  |  |  |
| 1987 exposure rate of whites to: |  |  |  |  |  |  |  |  |
| Blacks | $\begin{array}{r} -8.34 \\ 1.00 \end{array}$ | $\begin{array}{r} -5.92 \\ 0.93 \end{array}$ | $\begin{array}{r} -10.39 \\ 1.31 \end{array}$ | $\begin{array}{r} -14.19 \\ 0.58 \end{array}$ |  |  |  |  |
| Hispanics | $\begin{array}{r} -6.48 \\ 0.70 \end{array}$ | $\begin{array}{r} -5.35 \\ 0.72 \end{array}$ | $\begin{array}{r} -9.43 \\ 0.85 \end{array}$ | $\begin{array}{r} -8.20 \\ 0.40 \end{array}$ |  |  |  |  |
| Other nonwhites | $\begin{array}{r} -6.98 \\ 2.32 \end{array}$ | $\begin{array}{r} -9.38 \\ 3.55 \end{array}$ | $\begin{array}{r} -10.19 \\ 2.12 \end{array}$ | $\begin{array}{r} -3.58 \\ 1.20 \end{array}$ |  |  |  |  |
| Exposure rate to nonwhites |  |  |  |  | $\begin{array}{r} -2.94 \\ 4.83 \end{array}$ | $\begin{array}{r} -4.95 \\ 3.77 \end{array}$ | $\begin{array}{r} -3.68 \\ 3.95 \end{array}$ | $\begin{array}{r} -3.07 \\ 1.75 \end{array}$ |
| Exposure rate squared |  |  |  |  | $\begin{array}{r} -17.17 \\ 11.73 \end{array}$ | $\begin{array}{r} 0.29 \\ 10.44 \end{array}$ | $\begin{array}{r} -23.22 \\ 10.28 \end{array}$ | $\begin{array}{r} -19.25 \\ 5.10 \end{array}$ |
| Exposure rate cubed |  |  |  |  | $\begin{array}{r} 16.12 \\ 8.28 \end{array}$ | $\begin{array}{r} -1.43 \\ 7.82 \end{array}$ | $\begin{array}{r} 20.55 \\ 7.52 \end{array}$ | $\begin{array}{r} 13.86 \\ 3.90 \end{array}$ |
| Adjusted R-square | 0.64 | 0.63 | 0.55 | 0.33 | 0.65 | 0.63 | 0.56 | 0.31 |

Note: Coefficients taken from regressions explaining growth rate in white enrollment. Other explanatory variables included were: the intercept, metropolitan growth rate, accessibility (A1), change in segregation, and regional dummy variables.
Numbers below coefficients are standard errors. Bold type signifies coefficients that are significantly different from zero at the 95 percent level. Regressions are weighted by the square root of 1987 district enrollment. For definition of samples, see Table 7.

5/19/99 19:29; 5/26/99 10:49
M29A
7/27/99

## Table 10

Comparison of Estimates with Coleman et. al (1975)

| Equation | (1) | (2) | (3) |
| :---: | :---: | :---: | :---: |
| Period | Annual changes 1968-73 | Annual changes 1968-73 | 1987-1996 |
| Sample | 21 largest central city districts, pooled | Next 46 central city districts, pooled | Metro districts at least 5,000 and 10\% share (Sample I) |
| Observations | 105 | 226 | 187 |
| Intercept | 1.3 | 45.2 | $\begin{array}{r} -0.68 \\ 1.79 \end{array}$ |
| Change in segregation | 27.9 | 5.6 | 6.17 |
|  | 6.2 | 2.6 | 1.32 |
| Proportion black in district | -13.3 | -9.0 | -6.21 |
|  | 2.8 | 1.4 | 0.85 |
| In (district enrollment) | 0.0 | -4.2 | 0.08 |
|  | 0.8 | 1.0 | 0.17 |
| R -square | 0.29 | 0.26 | 0.28 |

Note: coefficients of Coleman et al. are multiplied by 100 to reflect a dependent variable measured in percentages rather than proportions, for comparability to equation (3).

Source: equations (1) and (2): Coleman et. al (1975), "Insert," Revised
Table 14; equation (3): Common Core of Data, author's calculations.
6/14/99 19:47.
M45
6/28/99
Table 11

| N | ----1987 segregation ------- |  |  | -----1996 segregation ------- |  |  | --Change in segregation-- |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | Between districts | Within districts | Total | Between districts | Within districts | Total | Between districts | Within districts |
| 238 | 0.302 | 0.222 | 0.080 | 0.317 | 0.240 | 0.077 | 0.014 | 0.017 | -0.003 |


| 139 | 0.149 | 0.083 | 0.065 | 0.173 | 0.107 | 0.066 | 0.024 | 0.024 | 0.001 |
| ---: | ---: | ---: | :--- | :--- | :--- | :--- | :--- | :--- | ---: |
| 66 | 0.283 | 0.179 | 0.104 | 0.293 | 0.200 | 0.093 | 0.010 | 0.022 | -0.012 |
| 25 | 0.311 | 0.243 | 0.068 | 0.335 | 0.261 | 0.074 | 0.024 | 0.019 | 0.006 |
| 8 | 0.427 | 0.350 | 0.077 | 0.428 | 0.357 | 0.071 | 0.000 | 0.007 | -0.006 |
|  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |
| 7 | 0.187 | 0.113 | 0.073 | 0.184 | 0.121 | 0.063 | -0.003 | 0.007 | -0.010 |
| 76 | 0.403 | 0.340 | 0.062 | 0.426 | 0.367 | 0.059 | 0.023 | 0.026 | -0.003 |
| 36 | 0.349 | 0.303 | 0.046 | 0.375 | 0.341 | 0.034 | 0.026 | 0.038 | -0.011 |
| 70 | 0.259 | 0.135 | 0.124 | 0.260 | 0.144 | 0.116 | 0.001 | 0.009 | -0.008 |
| 49 | 0.242 | 0.172 | 0.070 | 0.257 | 0.183 | 0.075 | 0.015 | 0.011 | 0.005 |

Note: segregation is measured by S . See text.
Source: U.S. Department of Education, Common Core of Data, 1987 and 1996; author's calculations. 6/1/99 13:46.
M41
6/28/99

Table A1
Metropolitan Areas in the Sample

| Metropolitan area | Region | PMSA | ----------1987---------- |  | ---1987-96---- |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | code | Enrollment | Percent | Growth rate |
|  |  |  |  | nonwhite |  |
|  |  |  |  |  |  |
| 1 Abilene, TX MSA | S | 40 | 21,692 | 29.4 | 1.21 |
| 2 Akron, OH PMSA | M | 80 | 104,579 | 16.0 | 0.41 |
| 3 Altoona, PA MSA | N | 280 | 21,198 | 1.7 | -0.16 |
| 4 Amarillo, TX MSA | S | 320 | 35,331 | 25.9 | 1.09 |
| 5 Anaheim--Santa Ana, CA PMSA | W | 360 | 339,314 | 38.4 | 2.56 |
| 6 Anderson, IN MSA | M | 400 | 23,085 | 11.8 | -1.35 |
| 7 Anderson, SC MSA | S | 405 | 25,145 | 23.7 | 0.34 |
| 8 Ann Arbor, MI PMSA | M | 440 | 37,402 | 21.5 | 1.21 |
| 9 Appleton--Oshkosh--Neenah, WI MS | M | 460 | 46,609 | 4.5 | 2.18 |
| 10 Asheville, NC MSA | S | 480 | 26,406 | 12.5 | 0.80 |
| 11 Aurora--Elgin, IL PMSA | M | 620 | 76,309 | 22.3 | 2.60 |
| 12 Austin, TX MSA | S | 640 | 124,237 | 40.7 | 3.50 |
| 13 Bakersfield, CA MSA | W | 680 | 102,653 | 41.8 | 3.07 |
| 14 Battle Creek, MI MSA | M | 780 | 25,177 | 17.5 | -0.10 |
| 15 Beaumont--Port Arthur, TX MSA | S | 840 | 71,516 | 36.3 | 0.17 |
| 16 Beaver County, PA PMSA | N | 845 | 29,150 | 9.4 | -0.01 |
| 17 Bellingham, WA MSA | W | 860 | 18,939 | 11.0 | 2.64 |
| 18 Benton Harbor, MI MSA | M | 870 | 29,969 | 27.8 | -0.42 |
| 19 Bismarck, ND MSA | M | 1010 | 15,161 | 3.0 | 0.81 |
| 20 Bloomington, IN MSA | M | 1020 | 12,441 | 5.5 | 0.40 |
| 21 Bloomington--Normal, IL MSA | M | 1040 | 18,854 | 8.0 | 1.71 |
| 22 Boston, MA PMSA | N | 1120 | 353,727 | 20.6 | 1.01 |
| 23 Boulder--Longmont, CO PMSA | W | 1125 | 35,721 | 13.8 | 1.99 |
| 24 Bradenton, FL MSA | S | 1140 | 23,574 | 21.9 | 3.35 |
| 25 Brazoria, TX PMSA | S | 1145 | 38,975 | 29.0 | 1.89 |
| 26 Bremerton, WA MSA | W | 1150 | 32,173 | 13.8 | 2.78 |
| 27 Bridgeport--Milford, CT PMSA | N | 1160 | 59,026 | 32.2 | 1.61 |
| 28 Bristol, CT PMSA | N | 1170 | 11,889 | 5.7 | 0.84 |
| 29 Brockton, MA PMSA | N | 1200 | 29,045 | 14.0 | 0.30 |
| 30 Brownsville--Harlingen, TX MSA | S | 1240 | 68,427 | 91.8 | 1.65 |
| 31 Bryan--College Station, TX MSA | S | 1260 | 15,335 | 39.5 | 2.30 |
| 32 Burlington, NC MSA | S | 1300 | 16,871 | 26.2 | 1.28 |
| 33 Burlington, VT MSA | N | 1305 | 19,794 | 2.0 | 1.40 |
| 34 Canton, OH MSA | M | 1320 | 69,689 | 10.4 | 0.04 |
| 35 Cedar Rapids, IA MSA | M | 1360 | 28,941 | 5.4 | 0.78 |
| 36 Champaign--Urbana--Rantoul, IL M | M | 1400 | 23,253 | 21.4 | 0.13 |
| 37 Charleston, SC MSA | S | 1440 | 81,952 | 44.1 | 0.49 |
| 38 Charlotte--Gastonia--Rock Hill, | S | 1520 | 183,793 | 29.6 | 1.84 |
| 39 Chicago, IL PMSA | M | 1600 | 854,879 | 52.1 | 1.18 |
| 40 Chico, CA MSA | W | 1620 | 25,743 | 14.8 | 3.02 |
| 41 Cincinnati, OH--KY--IN PMSA | M | 1640 | 218,638 | 19.9 | 0.80 |
| 42 Clarksville--Hopkinsville, TN--K | S | 1660 | 24,410 | 27.6 | 2.53 |
| 43 Cleveland, OH PMSA | M | 1680 | 249,729 | 32.4 | 0.99 |
| 44 Colorado Springs, CO MSA | W | 1720 | 69,725 | 21.1 | 2.26 |


| 45 Columbia, SC MSA | S | 1760 | 74,384 | 39.6 | 1.31 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 46 Columbus, OH MSA | M | 1840 | 204,921 | 15.3 | 1.90 |
| 47 Corpus Christi, TX MSA | S | 1880 | 76,308 | 66.5 | 0.32 |
| 48 Dallas, TX PMSA | S | 1920 | 425,508 | 39.9 | 2.67 |
| 49 Danbury, CT PMSA | N | 1930 | 28,030 | 10.8 | 0.99 |
| 50 Davenport--Rock Island--Moline, | M | 1960 | 62,343 | 13.8 | -0.23 |
| 51 Dayton--Springfield, OH MSA | M | 2000 | 155,578 | 19.4 | -0.12 |
| 52 Daytona Beach, FL MSA | S | 2020 | 41,400 | 20.6 | 3.58 |
| 53 Decatur, IL MSA | M | 2040 | 21,149 | 20.5 | -1.07 |
| 54 Denver, CO PMSA | W | 2080 | 262,862 | 26.6 | 2.08 |
| 55 Des Moines, IA MSA | M | 2120 | 64,701 | 9.8 | 1.41 |
| 56 Detroit, MI PMSA | M | 2160 | 712,284 | 30.7 | 0.34 |
| 57 Dubuque, IA MSA | M | 2200 | 12,738 | 1.5 | -0.30 |
| 58 Duluth, MN--WI MSA | M | 2240 | 41,349 | 5.2 | -1.11 |
| 59 Eau Claire, WI MSA | M | 2290 | 21,611 | 3.3 | 1.12 |
| 60 El Paso, TX MSA | S | 2320 | 130,947 | 80.5 | 1.42 |
| 61 Elkhart--Goshen, IN MSA | M | 2330 | 26,996 | 9.3 | 1.48 |
| 62 Enid, OK MSA | B | 2340 | 10,497 | 10.2 | -0.27 |
| 63 Erie, PA MSA | N | 2360 | 40,669 | 10.1 | 0.67 |
| 64 Eugene--Springfield, OR MSA | W | 2400 | 44,121 | 6.2 | 0.73 |
| 65 Evansville, IN--KY MSA | M | 2440 | 44,307 | 9.4 | -0.09 |
| 66 Fall River, MA--RI PMSA | N | 2480 | 21,972 | 2.8 | -0.14 |
| 67 Fargo--Moorhead, ND--MN MSA | M | 2520 | 24,184 | 3.7 | 1.54 |
| 68 Fayetteville, NC MSA | S | 2560 | 44,039 | 45.3 | 1.60 |
| 69 Fayetteville--Springdale, AR MSA | S | 2580 | 19,377 | 3.6 | 2.54 |
| 70 Fitchburg--Leominster, MA MSA | N | 2600 | 12,268 | 14.1 | 2.38 |
| 71 Flint, MI MSA | M | 2640 | 87,099 | 31.1 | -0.57 |
| 72 Florence, SC MSA | S | 2655 | 22,471 | 51.5 | 0.37 |
| 73 Fort Collins--Loveland, CO MSA | W | 2670 | 29,921 | 10.1 | 2.31 |
| 74 Fort Lauderdale--Hollywood--Pomp | S | 2680 | 136,139 | 36.7 | 4.82! |
| 75 Fort Myers--Cape CoraL, FI MSA | S | 2700 | 37,202 | 21.3 | 3.45 |
| 76 Fort Pierce, FL MSA | S | 2710 | 28,733 | 30.6 | 4.11 |
| 77 Fort Smith, AR--OK MSA | S | 2720 | 33,151 | 14.0 | 1.27 |
| 78 Fort Walton Beach, FL MSA | S | 2750 | 24,467 | 16.7 | 2.08 |
| 79 Fort Wayne, IN MSA | M | 2760 | 59,994 | 16.4 | 0.34 |
| 80 Fort Worth--Arlington, TX PMSA | S | 2800 | 219,478 | 29.3 | 2.40 |
| 81 Fresno, CA MSA | W | 2840 | 126,694 | 56.9 | 3.09 |
| 82 Gainesville, FL MSA | S | 2900 | 27,483 | 34.2 | 1.96 |
| 83 Galveston--Texas City, TX PMSA | S | 2920 | 54,427 | 34.5 | 1.87 |
| 84 Gary--Hammond, IN PMSA | M | 2960 | 113,634 | 37.1 | -0.31 |
| 85 Grand Forks, ND MSA | M | 2985 | 11,384 | 8.2 | 0.64 |
| 86 Grand Rapids, MI MSA | M | 3000 | 103,545 | 16.0 | 2.44 |
| 87 Greeley, CO MSA | W | 3060 | 22,240 | 29.4 | 1.98 |
| 88 Green Bay, WI MSA | M | 3080 | 30,552 | 6.2 | 1.98 |
| 89 Greensboro--Winston-Salem--High | S | 3120 | 144,252 | 27.1 | 1.15 |
| 90 Greenville--Spartanburg, SC MSA | S | 3160 | 101,783 | 24.8 | 0.85 |
| 91 Hamilton--Middletown, OH PMSA | M | 3200 | 47,525 | 7.3 | 1.36 |
| 92 Harrisburg--Lebanon--Carlisle, P | N | 3240 | 91,204 | 13.2 | 0.92 |
| 93 Hartford, CT PMSA | N | 3280 | 112,595 | 27.5 | 1.06 |
| 94 Hickory--Morganton, NC MSA | S | 3290 | 36,791 | 12.2 | 0.92 |
| 95 Houston, TX PMSA | S | 3360 | 591,404 | 52.0 | 2.39 |
| 96 Indianapolis, IN MSA | M | 3480 | 204,616 | 19.4 | 0.96 |
| 97 lowa City, IA MSA | M | 3500 | 10,581 | 8.2 | 2.07 |


| 98 Jackson, MI MSA | M | 3520 | 23,759 | 10.9 | 0.27 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 99 Jackson, TN MSA | S | 3580 | 13,587 | 44.3 | 0.22 |
| 100 Jacksonville, FL MSA | S | 3600 | 141,815 | 32.5 | 2.43 |
| 101 Jacksonville, NC MSA | S | 3605 | 17,201 | 27.0 | 2.15 |
| 102 Janesville--Beloit, WI MSA | M | 3620 | 24,556 | 10.0 | 1.21 |
| 103 Johnstown, PA MSA | N | 3680 | 36,807 | 2.6 | -0.82 |
| 104 Joliet, IL PMSA | M | 3690 | 66,551 | 21.7 | 1.70 |
| 105 Kalamazoo, MI MSA | M | 3720 | 32,194 | 19.1 | 0.60 |
| 106 Kankakee, IL MSA | M | 3740 | 17,635 | 27.8 | 0.08 |
| 107 Kenosha, WI PMSA | M | 3800 | 19,979 | 15.1 | 2.71 |
| 108 Killeen--Temple, TX MSA | S | 3810 | 46,712 | 37.4 | 2.57 |
| 109 Knoxville, TN MSA | S | 3840 | 95,561 | 9.0 | 1.12 |
| 110 Kokomo, IN MSA | M | 3850 | 18,698 | 7.6 | -1.01 |
| 111 La Crosse, WI MSA | M | 3870 | 13,466 | 7.2 | 1.56 |
| 112 Lafayette--West Lafayette, IN MS | M | 3920 | 16,859 | 4.8 | 0.87 |
| 113 Lake County, IL PMSA | M | 3965 | 85,011 | 19.6 | 2.78 |
| 114 Lakeland--Winter Haven, FL MSA | S | 3980 | 59,331 | 25.8 | 1.95 |
| 115 Lancaster, PA MSA | N | 4000 | 56,470 | 11.8 | 1.99 |
| 116 Lansing--East Lansing, MI MSA | M | 4040 | 74,596 | 16.8 | -0.07 |
| 117 Laredo, TX MSA | S | 4080 | 31,642 | 95.5 | 3.86 |
| 118 Las Vegas, NV MSA | W | 4120 | 96,346 | 26.8 | 6.60 |
| 119 Lawrence, KS MSA | M | 4150 | 9,792 | 14.4 | 2.62 |
| 120 Lawton, OK MSA | B | 4200 | 22,558 | 34.0 | 0.27 |
| 121 Lexington-Fayette, KY MSA | B | 4280 | 53,337 | 16.1 | 0.23 |
| 122 Lima, OH MSA | M | 4320 | 29,320 | 12.4 | -0.26 |
| 123 Lincoln, NE MSA | M | 4360 | 29,774 | 6.3 | 1.82 |
| 124 Little Rock--North Little Rock, | S | 4400 | 87,065 | 33.0 | 0.39 |
| 125 Longview--Marshall, TX MSA | S | 4420 | 34,564 | 30.9 | 0.52 |
| 126 Lorain--Elyria, OH PMSA | M | 4440 | 48,707 | 19.4 | -0.60 |
| 127 Los Angeles--Long Beach, CA PMSA | W | 4480 | 1,301,780 | 71.1 | 1.56 |
| 128 Louisville, KY--IN MSA | B | 4520 | 146,569 | 21.9 | -0.05 |
| 129 Lubbock, TX MSA | S | 4600 | 40,666 | 46.0 | 0.17 |
| 130 Madison, WI MSA | M | 4720 | 49,076 | 8.9 | 2.31 |
| 131 Mansfield, OH MSA | M | 4800 | 23,236 | 11.1 | -0.41 |
| 132 Mcallen--Edinburg--Mission, TX M | S | 4880 | 101,340 | 94.1 | 2.73 |
| 133 Medford, OR MSA | W | 4890 | 24,061 | 6.0 | 1.82 |
| 134 Melbourne--Titusville--Palm Bay, | S | 4900 | 49,288 | 16.8 | 3.26 |
| 135 Merced, CA MSA | W | 4940 | 36,140 | 52.2 | 2.87 |
| 136 Miami--Hialeah, FL PMSA | S | 5000 | 251,740 | 77.3 | 3.01 |
| 137 Middletown, CT PMSA | N | 5020 | 11,927 | 13.6 | 1.70 |
| 138 Midland, TX MSA | S | 5040 | 20,758 | 37.7 | 1.70 |
| 139 Milwaukee, WI PMSA | M | 5080 | 210,975 | 31.1 | 1.52 |
| 140 Minneapolis--St. Paul, MN--WI MS | M | 5120 | 362,338 | 11.6 | 1.97 |
| 141 Modesto, CA MSA | W | 5170 | 66,324 | 31.9 | 3.05 |
| 142 Muncie, IN MSA | M | 5280 | 17,973 | 10.2 | -0.58 |
| 143 Muskegon, MI MSA | M | 5320 | 30,037 | 22.6 | 0.97 |
| 144 Naples, FL MSA | S | 5345 | 17,503 | 29.8 | 5.16 |
| 145 Nashville, TN MSA | S | 5360 | 151,621 | 22.2 | 1.82 |
| 146 New Bedford, MA MSA | N | 5400 | 26,377 | 13.5 | -0.82 |
| 147 New Britain, CT PMSA | N | 5440 | 16,931 | 25.0 | 2.71 |
| 148 New Haven--Meriden, CT MSA | N | 5480 | 68,638 | 27.2 | 1.49 |
| 149 New London--Norwich, CT--RI MSA | N | 5520 | 35,881 | 11.0 | 0.81 |
| 150 Norwalk, CT PMSA | N | 5760 | 16,551 | 25.3 | 1.53 |


| 151 Oakland, CA PMSA | W | 5775 | 293,630 | 44.8 | 1.69 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 152 Ocala, FL MSA | S | 5790 | 26,261 | 24.2 | 3.39 |
| 153 Odessa, TX MSA | S | 5800 | 25,089 | 44.9 | 0.72 |
| 154 Oklahoma City, OK MSA | B | 5880 | 166,292 | 24.6 | 1.07 |
| 155 Olympia, WA MSA | W | 5910 | 28,656 | 11.1 | 2.66 |
| 156 Omaha, NE--IA MSA | M | 5920 | 103,410 | 16.0 | 1.14 |
| 157 Orlando, FL MSA | S | 5960 | 145,743 | 27.9 | 3.96 |
| 158 Owensboro, KY MSA | B | 5990 | 13,547 | 7.1 | 0.18 |
| 159 Oxnard--Ventura, CA PMSA | W | 6000 | 109,219 | 38.5 | 1.45 |
| 160 Panama City, FL MSA | S | 6015 | 21,261 | 18.8 | 1.92 |
| 161 Pawtucket--Woonsocket--Attleboro | N | 6060 | 44,148 | 9.0 | 1.59 |
| 162 Pensacola, FL MSA | S | 6080 | 52,999 | 26.0 | 2.25 |
| 163 Peoria, IL MSA | M | 6120 | 57,131 | 14.5 | -0.26 |
| 164 Phoenix, AZ MSA | W | 6200 | 338,587 | 27.2 | 3.11 |
| 165 Pine Bluff, AR MSA | S | 6240 | 17,475 | 54.2 | -0.64 |
| 166 Pittsburgh, PA PMSA | N | 6280 | 272,863 | 14.3 | 0.36 |
| 167 Pittsfield, MA MSA | N | 6320 | 13,055 | 4.1 | 0.13 |
| 168 Portland, OR PMSA | W | 6440 | 187,371 | 12.8 | 1.85 |
| 169 Providence, RI PMSA | N | 6480 | 86,231 | 15.7 | 0.70 |
| 170 Provo--Orem, UT MSA | W | 6520 | 63,453 | 4.0 | 1.85 |
| 171 Pueblo, CO MSA | W | 6560 | 22,877 | 46.1 | 0.17 |
| 172 Racine, WI PMSA | M | 6600 | 28,809 | 25.0 | 0.19 |
| 173 Raleigh--Durham, NC MSA | S | 6640 | 102,132 | 36.2 | 3.14 |
| 174 Reading, PA MSA | N | 6680 | 49,243 | 11.7 | 2.45 |
| 175 Redding, CA MSA | W | 6690 | 24,721 | 10.9 | 1.84 |
| 176 Reno, NV MSA | W | 6720 | 32,929 | 18.0 | 4.49 |
| 177 Richland--Kennewick--Pasco, WA M | W | 6740 | 29,917 | 19.1 | 2.63 |
| 178 Riverside--San Bernardino, CA PM | W | 6780 | 403,022 | 39.9 | 4.45 |
| 179 Rochester, MN MSA | M | 6820 | 17,055 | 6.0 | 1.85 |
| 180 Rockford, IL MSA | M | 6880 | 45,165 | 18.6 | 1.02 |
| 181 Sacramento, CA MSA | W | 6920 | 225,993 | 30.3 | 2.78 |
| 182 Saginaw--Bay City--Midland, MI M | M | 6960 | 72,023 | 23.4 | -0.43 |
| 183 St. Cloud, MN MSA | M | 6980 | 33,164 | 1.7 | 1.78 |
| 184 Salem, OR MSA | W | 7080 | 43,787 | 10.8 | 2.17 |
| 185 Salem--Gloucester, MA PMSA | N | 7090 | 35,435 | 4.3 | 1.00 |
| 186 Salinas--Seaside--Monterey, CA M | W | 7120 | 55,066 | 57.3 | 1.68 |
| 187 Salt Lake City--Ogden, UT MSA | W | 7160 | 247,299 | 8.1 | 1.24 |
| 188 San Angelo, TX MSA | S | 7200 | 18,029 | 39.7 | 1.01 |
| 189 San Antonio, TX MSA | S | 7240 | 245,688 | 64.5 | 1.26 |
| 190 San Diego, CA MSA | W | 7320 | 351,980 | 42.6 | 2.41 |
| 191 San Francisco, CA PMSA | W | 7360 | 163,566 | 57.6 | 0.86 |
| 192 San Jose, CA PMSA | W | 7400 | 219,304 | 49.4 | 1.13 |
| 193 Santa Barbara--Santa Maria--Lomp | W | 7480 | 47,825 | 42.1 | 2.64 |
| 194 Santa Cruz, CA PMSA | W | 7485 | 33,084 | 34.4 | 1.39 |
| 195 Santa Rosa--Petaluma, CA PMSA | W | 7500 | 56,162 | 16.1 | 2.06 |
| 196 Sarasota, FL MSA | S | 7510 | 25,730 | 14.1 | 1.66 |
| 197 Scranton--Wilkes-Barre, PA MSA | N | 7560 | 96,628 | 2.3 | 1.38 |
| 198 Seattle, WA PMSA | W | 7600 | 264,408 | 18.0 | 2.52 |
| 199 Sharon, PA MSA | N | 7610 | 19,230 | 8.3 | 0.19 |
| 200 Sheboygan, WI MSA | M | 7620 | 17,346 | 6.3 | 1.45 |
| 201 Sherman--Denison, TX MSA | S | 7640 | 16,879 | 11.6 | 1.12 |
| 202 Sioux City, IA--NE MSA | M | 7720 | 20,515 | 8.8 | 0.85 |
| 203 South Bend--Mishawaka, IN MSA | M | 7800 | 37,541 | 19.7 | 0.26 |


| 204 Spokane, WA MSA | W | 7840 | 60,290 | 7.6 | 1.93 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 205 Springfield, IL MSA | M | 7880 | 28,407 | 14.4 | 0.87 |
| 206 Springfield, MA MSA | N | 8000 | 74,385 | 25.4 | 1.00 |
| 207 i Stamford, CT PMSA | N | 8040 | 23,391 | 27.6 | 2.04 |
| 208 State College, PA MSA | N | 8050 | 12,626 | 3.9 | 1.31 |
| 209 Stockton, CA MSA | W | 8120 | 85,946 | 48.2 | 2.17 |
| 210 Tacoma, WA PMSA | W | 8200 | 95,768 | 19.5 | 2.36 |
| 211 Tallahassee, FL MSA | S | 8240 | 32,366 | 50.1 | 1.88 |
| 212 Tampa--St. Petersburg--Clearwate | S | 8280 | 244,906 | 23.8 | 2.53 |
| 213 Terre Haute, IN MSA | M | 8320 | 21,820 | 6.2 | -0.30 |
| 214 Texarkana, TX--Texarkana, AR MSA | S | 8360 | 23,913 | 30.6 | -0.20 |
| 215 Toledo, OH MSA | M | 8400 | 97,212 | 21.8 | 0.02 |
| 216 Topeka, KS MSA | M | 8440 | 25,694 | 18.7 | 0.47 |
| 217 Tucson, AZ MSA | W | 8520 | 103,148 | 39.7 | 2.08 |
| 218 Tulsa, OK MSA | B | 8560 | 123,726 | 23.5 | 1.03 |
| 219 Tyler, TX MSA | S | 8640 | 28,197 | 33.7 | 0.57 |
| 220 Vallejo--Fairfield--Napa, CA PMS | W | 8720 | 67,876 | 36.1 | 2.43 |
| 221 Vancouver, WA PMSA | W | 8725 | 42,947 | 8.4 | 3.39 |
| 222 Victoria, TX MSA | S | 8750 | 15,019 | 52.7 | 0.41 |
| 223 Visalia--Tulare--Porterville, CA | W | 8780 | 64,071 | 52.7 | 2.58 |
| 224 Waco, TX MSA | S | 8800 | 32,315 | 39.1 | 1.77 |
| 225 Waterbury, CT MSA | N | 8880 | 29,709 | 23.6 | 1.90 |
| 226 Waterloo--Cedar Falls, IA MSA | M | 8920 | 24,495 | 11.9 | -0.49 |
| 227 Wausau, WI MSA | M | 8940 | 17,601 | 4.0 | 1.38 |
| 228 West Palm Beach--Boca Raton--Del | S | 8960 | 89,458 | 36.8 | 4.66 |
| 229 Wichita, KS MSA | M | 9040 | 79,865 | 18.9 | 1.46 |
| 230 Wichita Falls, TX MSA | S | 9080 | 21,188 | 26.0 | 0.58 |
| 231 Williamsport, PA MSA | N | 9140 | 19,850 | 2.7 | 0.27 |
| 232 Wilmington, NC MSA | S | 9200 | 19,192 | 30.9 | 1.31 |
| 233 Worcester, MA MSA | N | 9240 | 62,348 | 11.2 | 1.85 |
| 234 Yakima, WA MSA | W | 9260 | 36,279 | 35.4 | 2.64 |
| 235 York, PA MSA | N | 9280 | 58,544 | 7.5 | 1.79 |
| 236 Youngstown--Warren, OH MSA | M | 9320 | 82,795 | 18.1 | -0.74 |
| 237 Yuba City, CA MSA | W | 9340 | 22,348 | 30.7 | 2.19 |
| 238 Yuma, AZ MSA | W | 9360 | 20,713 | 57.0 | 3.25 |
|  |  |  |  |  |  |
| c: 1123 lx ${ }^{\text {m }}$ 26 |  |  |  |  |  |
| 4/15/99 |  |  |  |  |  |

N
高
E
N


Source: Table 9
Figure 2

Note: Data ranges exclude lower bound, i.e., .10-. 15 is greater than 10 and less than . 15 . Source: Common Core of Data; author's calculations.


[^0]:    * I am grateful to Jens Ludwig for helpful comments and to Thomas Anderson, Cathleen McHugh, and Randy Walsh for research assistance. c:lwp $|x| m e t r o l c h d r a f t ~$

