

Association between autoimmune diseases and COVID-19 as assessed in both a test-negative casecontrol and population case-control design

Rossella Murtas

Agency for Health protection of Milan https://orcid.org/0000-0002-8989-7605

Anita Andreano

Agency for Health protection of Milan

Federico Gervasi

Agency for Health protection of Milan

Davide Guido

Agency for Health protection of Milan

David Consolazio

Agency for Health protection of Milan

Sara Tunesi

Agency for Health protection of Milan

Laura Andreoni

Agency for Health protection of Milan

Maria Teresa Greco

Agency for Health protection of Milan

Maria Elena Gattoni

Agency for Health protection of Milan

Monica Sandrini

Agency for Health protection of Milan

Antonio Riussi

Agency for Health protection of Milan

Antonio Giampiero Russo (Zagrusso@ats-milano.it)

Agency for Health protection of Milan https://orcid.org/0000-0002-5681-5861

Research Article

Keywords: Autoimmune diseases, covid-19, test-negative design

Posted Date: August 18th, 2020

DOI: https://doi.org/10.21203/rs.3.rs-58948/v1

Version of Record: A version of this preprint was published at Autoimmunity Highlights on October 6th, 2020. See the published version at https://doi.org/10.1186/s13317-020-00141-1.

Abstract

Background: COVID-19 epidemic has paralleled with the so called infodemic, where countless pieces of information have been disseminated on putative risk factors for COVID-19. Among those, emerged the notion that people suffering from autoimmune diseases (AIDs) have a higher risk of SARS-CoV-2 infection.

Methods: The cohort included all COVID-19 cases residents in the Agency for Health Protection (AHP) of Milan that, from the beginning of the outbreak, developed a web-based platform that traced positive and negative cases as well as related contacts. AIDs subjects were defined ad having one the following autoimmune disease: rheumatoid arthritis, systemic lupus erythematosus, systemic sclerosis, Sjogren disease, ankylosing spondylitis, myasthenia gravis, Hashimoto's disease, acquired autoimmune hemolytic anemia, and psoriatic arthritis. To investigate whether AID subjects are at increased risk of SARS-CoV-2 infection, and whether they have worse prognosis than AIDs-free subjects once infected, we performed a combined analysis of a test-negative design case-control study, a case-control with test-positive as cases, and one with test-negative as cases (CC-NEG).

Results: During the outbreak, the Milan AHP endured, up to April 27th 2020, 20,364 test-positive and 34,697 test-negative subjects. We found no association between AIDs and being positive to COVID-19, but a statistically significant association between AIDs and being negative to COVID-19 in the CC-NEG. If, as likely, test-negative subjects underwent testing because of respiratory infection symptoms, these results imply that autoimmune diseases may be a risk factor for respiratory infections in general (including COVID-19), but they are not a specific risk factor for COVID-19. Furthermore, when infected by SARS-CoV-2, AIDs subjects did not have a worse prognosis compared to non-AIDs subjects. Results highlighted a potential unbalance in the testing campaign, which may be correlated to the characteristics of the tested person, leading specific frail population to be particularly tested.

Conclusions: Lack of availability of sound scientific knowledge inevitably lead unreliable news to spread over the population, preventing people to disentangle them form reliable information. Even if additional studies are needed to replicate and strengthen our results, these findings represent initial evidence to derive recommendations based on actual data for subjects with autoimmune diseases.

Background

Since its first appearance in China in December 2019, and even more with its worldwide subsequent diffusion, COVID-19 epidemic has paralleled with the so called infodemic. Countless pieces of information (often lacking scientific validity) have been disseminated on putative risk factors for COVID-19 trough traditional and social media channels [1]. Among this information, the notion that people suffering from autoimmune diseases (AIDs) have a higher risk of SARS-CoV-2 infection emerged. In fact, one can speculate that subjects with AIDs might be at greater risk of infections for the AID itself, but also because of immunomodulatory treatment and secondary chronic conditions. In literature, subjects with

rheumatoid arthritis (RA) have been found to have higher risk of death from infections [2–4] and higher risk of nonfatal infections [5, 6] compared to the general population. Tektonidou et al. [7] found an increased risk of hospitalization for serious infections in subjects with systemic lupus erythematosus (SLE) and Bosch et al [8] concluded that subjects with SLE have an increased overall risk for infections (including pneumonia). Also, an increased risk of pneumonia from different coronavirus infection has been reported in immunocompromised subjects [9].

However, limited scientific evidence is currently available on the association between AIDs and COVID-19. Articles on the subject, published in peer-reviewed medical journals, are mostly general recommendations or systemic reviews providing an overview on viral infectious risk in AIDs subjects [10]. In a literature review, Favalli et al. [11], hypothesized a two-way association between rheumatoid arthritis (RA) and COVID-19: microorganisms can indeed produce acute and chronic arthritis through direct colonisation of the joints or inducing an autoimmune response to the infection. At the same time, an increase in risk of infection in RA subjects compared to the general population due to the impairment of the immune system typical of autoimmune disorders is well documented. Askanase and colleagues [12] pointed out a lack of knowledge in the COVID-related respiratory complications in subjects with autoimmune diseases, in particular for SLE subjects that may be susceptible to the more severe manifestations of COVID-19, such as pneumonia. They even suggested that high type I interferon levels, found clustered in SLE families [13], may exert, on the contrary, a protective effect on COVID-19. Few studies have been focusing on quantifying the relationship between AIDs and susceptibility to SARS-COV-2 infection or to severe COVID-19 disease. D'Silva et al [14] investigated differences in manifestations and outcomes of coronavirus disease 2019 infection between subjects with rheumatic disease (RD) and subjects without RDs. They found similar characteristic between RD and non-RD subjects on hospitalizations, and significantly different prognosis in RD subjects requiring more often intensive care admission and mechanical ventilation. Liu et al [15], through a meta-analysis, showed that AID subjects had a 21% increased risk of severe COVID-19 disease and a 31% increased risk of mortality in subjects with COVID-19. However, no details on which diseases were considered was available, and none of the found increases in risk was statistically significant overall or by country. In a second study, Emmi et al [16], in a sample of subjects with AID residing in Tuscany, found a prevalence of COVID-19 comparable to that observed in the general population of Tuscany.

Most of the available studies attributed the potential association between AID and susceptibility to COVID-19, infection or severe disease, to immunosuppressive or immunomodulatory therapies used to treat AIDs [17–19]. Subjects treated with high-dose corticosteroids are overall considered at significant risk of serious infection [20, 21]. On the other hand, conventional disease-modifying anti-rheumatic drugs are not considered a risk factor for COVID-19 [21, 22], and some immunosuppressive medications (such as tocilizumab), have been found effective in alleviating symptoms and even recommended for severe COVID-19 management in addition to standard therapy [23].

Methods

Data sources

The cohort included all COVID-19 cases in the study area, covered by the Agency for Health Protection (AHP) of Milan, corresponding to 193 municipalities in the northern Italian region of Lombardy, with a total population of 3,48 million inhabitants. The study area includes the municipality of Codogno that was at the origin of the first Italian epidemic outbreak. From the beginning of the outbreak, all tracing activities were included in a web-based platform, developed by the Epidemiologic Unit of the AHP, called Milano COV, including cases and related contacts (details on the information system are described in the Online-Only Methods). A confirmed-case is defined as a person with a real-time polymerase chain reaction (RT-PCR) positive result of SARS-COV-2 infection, irrespective of clinical signs and symptoms. Contacts are defined as all individuals who are associated with a case's sphere of activity, thus potentially exposed to the same source of contagion. Cases and close contacts underwent epidemiological investigation to provide description of the clinical presentation of COVID-19 and its clinical course. Furthermore, data were collected to estimate the serial interval of SARS-COV-2 infection, the symptomatic proportion of COVID-19 cases, and to identify possible routes of transmission.

From the beginning of the outbreak, all tracing activities were included in a web based platform, developed by the Epidemiologic Unit of the AHP, called Milano COV, including cases and related contacts. During the outbreak, it has been necessary to communicate the nominative list of identified cases and close contacts of each municipality to the Mayor's office, to verify social support needs and to assess possible guarantine violations. The prefectures of the province of Milan and Lodi needed the same information to enforce restrictive measures provided for by temporary laws issued during the lockdown phase. Moreover, in order to allow the clinical management of identified cases and close contacts, information related to their registered patients was released to general practitioners (GPs). In order to manage the dynamic lists of cases and contacts, updated with information relating to death, hospitalization or discharge home, an additional web portal called ATS-Milano COR was set up. The portal allowed the selective visualization of subjects by town of residence (for Mayor's offices and other stakeholders), province of residence (for prefectures) or according to their GP. This dynamic information structure is capable also to track the clinical evolution of each case from symptom onset, or swab date, up to a negative PCR result, and of each close contact to the end of quarantine. For cases and contacts reported by general practitioners, in addition to the active surveillance and contact tracing described above, a massive SMS system was activated both to reinforce the indications for segregation and to provide links to information material on how to maintain the guarantine at home.

In order to expand the outbreak reporting system, general practitioners could add symptomatic cases that did not undergo a nasopharyngeal swab, and their close contacts. In the AHP of Milan 95% of the residents are registered with a GP affiliated with the Lombardy Regional Health System (RHS). From the beginning of the epidemic, the Lombardy Region daily sent the list of hospitalized COVID cases to each AHP, including the date of entry and the name of the hospital where the patient was admitted, and each variation (transfer, home discharge, death) was communicated in the following daily dataflow. Additional information relating to hospitalizations of patients in the ATS Milano COR was derived from the regular

administrative data discharge flow, which is consolidated for hospital admissions up to the end of April 2020. We integrated between them, and verified with the demographic information in the Health Service Register of the Lombardy Region (age, gender, place of residence), the described data sources in the Integrated Datawarehouse for COVID Analysis in Milan, anonymized with a random unique patient id. The same id was assigned to those subjects in all other administrative databases of the AHP, anonymized prior to analysis. Individual level comorbidities data were derived using the chronic disease administrative database of the AHP of Milan, according to the algorithms specified in the Regional Act X/6164 [24] and X/7655 [25] of 2017, and summarized in English in E-Table 1. Vital status was derived from the early notification system of the AHP of Milan, set-up from the beginning of the epidemic, in which deaths are communicated from the Civil Registry of each Municipality to the AHP and manually introduced in the Health Service Register, or directly from the GP and Mayor's offices for the subjects already in the Milano COV database through the web-based information system. We determined vital status at 30-day from diagnosis, which was defined for confirmed-cases as the first date between registered symptom onset and the swab positivity result. The date of symptom onset in the database was derived from the epidemiological interview or from the date of first access to an emergency department or first thorax CT scan, in this order of priority. If none of these dates was available and the patient had been hospitalized, the date of hospital admission was used. For a minority of patients, infected in the early phase of the epidemic and for whom no onset dates were available, we uniformly random imputed the date of symptoms onset between February 10th and 17th. For symptomatic cases, date of diagnosis was the date of symptoms onset reported by the GP or, if missing, the date in which the subject was introduced in the web-system by the GP. For this analysis, we considered as alive patients with a date of death more than 30 days after the date of diagnosis. The vital status was assessed on May 23th 2020.

Study Population

From the COVID-19 database of the Milan AHP we extracted, on April 27, 2020, all subjects with nasopharyngeal positive and negative swabs confirmed SARS-COV-2. Through the database, we collected demographic information on age, gender, municipality of residence, ASST (geographical and administrative partition of the territory of the Milan AHP). We defined as exposed all subjects with the following autoimmune diseases: AR, SLE, systemic sclerosis, Sjogren disease, ankylosing spondylitis, myasthenia gravis, Hashimoto's disease, acquired autoimmune hemolytic anemia, and psoriatic arthritis. Presence of any autoimmune disease was identified in the chronic disease administrative database of the AHP of Milan, where information on the presence of 64 chronic conditions is recorded, for every resident registered with the RHS, using outpatient exams and visits, hospital discharge sheets, pharmaceutical, and exemption from co-payment databases according to the algorithms defined in the Regional Act X/6164 and X/7655 of 2017 [24, 25]. Number of comorbidities was derived from the same database. For test-positive subjects, death and hospitalization status were updated to June 11, 2020.

Study Design

To evaluate the association between autoimmune status and occurrence of COVID-19 *we performed a combined analysis of a test-negative design (TND) case-control study, a case-control with test-positive as cases (CC-POS design), and one with test-negative as cases (CC-NEG design). As proposed by Vandenbroucke <i>et al* [26], *the combination of these studies* will serve to evaluate if autoimmune diseases are specific risk factors for COVID-19 or generally for respiratory diseases with similar symptoms. We also performed a conventional matched case-control design for comparison.

TNDs evaluate the association between an exposure and an outcome by comparing test-positive (cases) with test-negative (controls) subjects. Cases were defined as all subjects with a positive swab collected from the ATS-Milano COR system, no exclusions were performed. Controls as all subjects with a negative swab included in the same database. The idea is that test-negative controls underwent testing because they presented symptoms attributable to COVID-19 but resulted negative, thus having a different infection which may lead to similar symptoms, for example another respiratory infection. In fact, being susceptible to the same selection mechanisms as tested-positives will protect from common case-control biases. The TND design is potentially capable of identifying the effect of autoimmunity on COVID-19 if it has a different magnitude, or even direction, compared to the effect that it has in other respiratory infections [26]. To control for different spatial correlation among subjects, and to control for time trends in the administration of swabs, we matched TND's cases and controls by ASST and date of swab, date of positive swab for cases and date of negative swab for controls within 7 days of the case index date, only matched cases and controls will be considered.

At the same time, comparing test-positive subjects with general population controls will allow to estimate the effect of autoimmunity on COVID-19 infection compared to a control without respiratory symptoms. On the other hand, comparing test-negative subjects to the general population will allow to assess if autoimmunity is a risk factor for respiratory infections in general. Consequently, we designed two additional case-controls: in the first one cases were those subjects who tested positive (CC-POS design), while in the second one cases were those who tested negative (CC-NEG). Test-negative subjects were the same subjects selected as controls in the TND design [26]. In order to compare CC-POS's (and CC-NEG's) results with TND, a number of controls equal to 4 times the number of test-positive plus test-negative subjects was randomly sampled from the general population of the Milan AHP, thus resulting in a single control group for both CC-POS and CC-NEG designs.

For comparison, we performed also a classical population case-control design (hereafter name casecontrol design 2), where cases (test-positive subjects) and controls are matched by age (5 years), gender and municipality of residence. Controls were randomly sampled from the general population of the Milan AHP, also with a ratio of 1:4, only cases matched with 4 controls will be considered.

To evaluate the association between autoimmune status and a proxy of disease severity, defined as *non-hospitalized and alive, hospitalized and alive, and deceased,* we performed a cohort study using all COVID-19 test-positive subjects.

Statistical Analysis

To measure the association between autoimmune diseases and COVID-19 occurrence we used logistic regression models and conditional logistic regression models in matched designs, presenting results as the ORs of having an autoimmune disease in cases compared to controls and their 95% CIs.

Models for the TND design were adjusted for age (categorized as <17, [18-40), [40-70), \geq 70 years), gender and number of non-AIDs chronic conditions (categorized as no conditions, 1-3, and \geq 4). Models for CC-POS and CC-NEG were adjusted for gender, age (categorized as <17, [18-40), [40-70), \geq 70 years), number of non-AIDs chronic conditions (categorized as no conditions, 1-3, and \geq 4) and municipality of residence. Models for the classical Population Case-Controls Design were adjusted for number of non-AIDs chronic conditions (categorized as no conditions, 1-3, and \geq 4).

To measure the association between autoimmune diseases and severity of COVID-19 disease we used ordinal (cumulative) and multinomial logistic models. Ordinal logistic regression, in its cumulative formulation, requires the assumption of proportional odds [27]. When any covariate did not satisfy this assumption, a partial proportional odds model was fitted allowing non-proportionality for the selected variables [28]. Ordinal and multinomial logistic models were adjusted for gender, age (categorized as <17, [18-40), [40-70), \geq 70 years), number of non-AIDs chronic conditions (categorized as no conditions, 1-3, and \geq 4) and ASST (n=6). We decided to adjust for ASST, and not for municipality, of residence given that there are 193 municipality in the territory of the AHP, which may have led to very few cases in each stratum. Results were displayed as ORs with 95% CIs. ORs for the ordinal logistic model will be interpreted in their cumulative formulation that is the odds of deceased versus the combined categories hospitalized and alive, and non-hospitalized and alive. ORs for the multinomial logistic regression will be interpreted as usual ORs with non-hospitalized and alive as the reference category. The analyses were performed using SAS Software version 9.4 (SAS Institute Inc., Cary, NC, USA).

Results

Description of cases and controls

During the outbreak, the Milan AHP endured, up to April 27th 2020, 20,364 test-positive and 34,697 testnegative subjects (overall demographic and clinical characteristics are reported in Table 1). Demographic and clinical characteristics of test-negative matched to test-positive subjects by ASST and date of swab, and of population controls for CC-POS and CC-NEG designs are reported in Additional file 1.

The proportion of males among tested-positives was 47.5%, similar to population controls (47.5% for case-control design 2 and 48.4% for the CC-POS and CC-NEG designs) and higher than test-negative subjects (39.4% overall, and 40.7% after matching). The majority of tested-positives had more than 70 years (47.5%), as well as population controls for the case-control design 2 (47.4%). On the contrary, the

proportion of people older than 70 years was lower in test-negative subjects (26.6% overall, 26.1% after matching) and controls for the CC-POS and CC-NEG designs (18.1%). However, mean age in test-negative subjects was 54.8 years (s.d. 20.8, Table 1) in the overall test-negative group and 54.8 (s.d. 20.6) in the matched test-negative group (Additional file 1), higher than in the random sample of the general population controls used for CC-POS and CC-NEG where the mean age was 45.4 years (s.d. 23.7, Additional file 1).

Most of cases presented at least one non-AIDs comorbidity (60%), and 15.5% of tested-positives had at least 4 non-AIDs comorbidities. Among test-negative subjects, the number of non-AIDs comorbidities was quite different, with the majority of people having no comorbidities (54.4%) and approximately 45% having at least one non-AIDs comorbidity. Only 32% of the general population controls used for CC-POS and CC-NEG had at least one non-AIDs comorbidity while the majority of controls for the case-control design 2 had at least one comorbidity (55%).

Among tested-positives, 665 (3.2%) had an AID disease with the majority having Hashimoto's thyroiditis (48.6%) followed by rheumatoid arthritis (25.1%). The figures were slightly higher in test-negative subjects, with 1,297 (3.9%) having an autoimmune disease, 56.7% having Hashimoto's thyroiditis and 17.2% having rheumatoid arthritis. The proportions of AIDs subjects among population controls (case-control design 2) were considerably different compared to test-positive and test-negative subjects, with 2,169 persons with an autoimmune disease (2.7%), most of them again having Hashimoto's thyroiditis (52%) followed by rheumatoid arthritis (22.1%).

Almost 41% of test-positive subjects were non-hospitalized and alive, 39% were hospitalized and alive, and 20% were deceased.

Autoimmune disease and COVID-19 occurrence

The adjusted OR *of having an autoimmune disease in COVID-19 test-positive compared to test-negative subjects was* 0.86 (95% CI 0.76-0.96) in the TND design analysis (Table 2). Comparing test-positive subjects to a random sample of population controls (CC-POS), the unadjusted OR *of having an autoimmune disease* was 1.45 (95% CI 1.32-1.58) which shrunken to no-association when adjusted by covariates (OR=0.98, 95% CI 0.90-1.08). On the other hand, the adjusted OR *of having an autoimmune disease* was 1.19 (1.09-1.29) for test-negative subjects compared to a random sample of population controls (CC-NEG).

In the case-control design 2, matching for age, gender and municipality of residence (among 20,364 testpositive subjects, 20,327 where matched with 4 controls), we found a positive association between being diagnosed with COVID-19 and autoimmune disease, with an adjusted OR *of having an autoimmune disease in tested-positives compared to controls* of 1.16 (95% CI 1.06-1.26).

Autoimmune disease and COVID-19 severity

Among autoimmune subjects, 44% were non-hospitalized and alive, and 17.1% died (Table 1). Treating the proxy variable of disease severity as ordinal, we found no association between having a more severe outcome and having an autoimmune disease, with an adjusted OR of 0.96 (95% CI 0.83-1.12), which *is the odds of deceased versus the combined categories* hospitalized and alive, and non-hospitalized and alive, and of *the combined categories deceased, and* hospitalized and alive versus non-hospitalized and alive (Table 3). Similar results were obtained using multinomial logistic regression, the adjusted OR of having an autoimmune disease was 1.05 (95% CI 0.88-1.25) for the group being hospitalized and alive compared to non-hospitalized and alive.

Discussion

In this work, we evaluated the association between autoimmune disease and risk of COVID-19 in the population of the AHP of Milan, an area populated by 3.48 million of inhabitants, particularly damaged by the virus, which includes the municipality of Codogno (where the Italian outbreak started). Concerning autoimmunity and COVID-19, we found a negative association between autoimmune status and risk of COVID-19 for tested-positives compared to tested-negatives controls. When comparing test-positive and test-negative subjects with a random sample of the population (by CC-POS and CC-NEG designs), we found no association between the exposure and a positive swab result, but a statistically significant association between the exposure and a negative swab result. If we are willing to assume that testnegative subjects underwent testing because of different disease but presented symptoms attributable to COVID-19, such as another respiratory infection, these results seem to imply that autoimmune diseases may be a risk factor for respiratory infections in general (including COVID-19), but they might not be a specific risk factor for COVID-19. On the other hand, the higher proportion of autoimmune diseases in the tested population compared to non-tested population (3.2% in test-positives, 3.9% in test-negatives, and 2.7% or 2.2% in the general population) highlights a potential unbalance of the exposure in the testing campaign, that is AIDs subjects were more likely to get tested than general population. However, comparing test-positives with the general population (by case-control design 2), we found that autoimmune disease is a risk factor for COVID-19. Vandenbroucke et al [26] suggested that, when testing campaigns are correlated to the characteristics of the tested person, or general practitioners may serve specific frail populations, usual case-controls designs could produce biased results given that tested subjects might be self-selected contrary to the underlying population. In fact, the results had shown that test-negative subjects were older and had higher proportions of comorbidities compared to general population. In addition, the results suggested that AIDs subjects, when infected by SARS-CoV-2, do not have a worse prognosis compared to non-AIDs subjects.

Being one of the first study aimed at investigating the association between autoimmune diseases and COVID-19, the present study is scarcely comparable to previously published results. The prevalence of RA subjects found in a cohort of 2154 SARS-CoV-2 positive subjects in a large healthcare system in Massachusetts was similar [14], while smaller proportions of autoimmune conditions (with no specification) were found in literature compared to our study [29–32].

Strengths and limitations

A limitation of the present study is the selection bias due to biased selection of cases and controls, plausible in our situation given that, to date, only symptomatic cases were tested. On the other hand, comparing test-positives with test-negatives we compared subjects with similar characteristics and potentially similar probability of being tested.

One of the strength of this work is the use of the TND, CC-POS and CC-NEG designs, which combined, helped to evaluate potential risk factors for COVID-19 that differ from those for other respiratory infections. Furthermore, the numerousness of the cohorts considered which ensure generalizability of our results.

Conclusions

Lack of availability of sound scientific knowledge inevitably lead unreliable news to spread over the population, preventing people to disentangle them form reliable information. The rapid circulation of information disseminated in television and social media channels led the World Health Organization to acknowledge that: "We are not just fighting an epidemic, we're fighting an infodemic" making fake news spread more easily than the virus. Even if additional studies are needed to replicate and strengthen our results, these findings represent initial evidence to derive recommendations based on actual data for subjects with autoimmune diseases.

Abbreviations

COVID-19 = Coronavirus Disease 2019

AID = autoimmune diseases

- AHP = Agency for Health Protection
- RA = rheumatoid arthritis
- SLE = systemic lupus erythematosus
- GP = general practitioner
- RHS = Regional Health System
- TND = test-negative design
- CC-POS = case-control with test-positive as cases

OR = Odd Ratio

Declarations

Ethics approval and consent to participate: Ethics approval and consent to participate were not required, as this is an observational study based on data routinely collected by the Agency for Health Protection (ATS) of Milan, a public body of the Regional Health Service – Lombardy Region. The ATS has among its institutional functions, established by the Lombardy Region legislation (R.L. 23/2015), the government of the care pathway at the individual level in the regional social and healthcare system, the evaluation of the services provided to, and the outcomes of, patients residing in the covered area. This study is also ethically compliant with the National Law (D.Lgs. 101/2018) and the "General Authorisation to Process Personal Data for Scientific Research Purposes" (n.8 and 9/2016, referred to in the Data Protection Authority action of 13/12/2018). Data were anonymized with a unique identifier in the different datasets before being used for the analyses.

Consent for publication: Not applicable

Availability of data and materials: Data are not publicly available due to property restrictions.

Competing interests: none.

Funding: none.

Authors' contributions: All authors have made substantial contributions to the conception and design of the work; RM and AGR conceptualized the study and defined the methodology. RM analysed the dataset and wrote the original version of the manuscript. AGR supervised and administered the project. AA, FG, DG, DC, ST, MTG, and MEG contributed to the interpretation of the data, definition of methodology and revising the paper. AR and MS collected the data and contributed to the revision of the manuscript.

Acknowledgements: Not applicable

References

- 1. Orso D, Federici N, Copetti R, et al (2020) Infodemic and the spread of fake news in the COVID-19-era. Eur J Emerg Med. https://doi.org/10.1097/MEJ.0000000000000713
- 2. Mitchell DM, Spitz PW, Young DY, et al (1986) Survival, prognosis, and causes of death in rheumatoid arthritis. Arthritis Rheum 29:706–714. https://doi.org/10.1002/art.1780290602
- 3. Wolfe F, Mitchell DM, Sibley JT, et al (1994) The mortality of rheumatoid arthritis. Arthritis Rheum 37:481–494. https://doi.org/10.1002/art.1780370408

- 4. Symmons DP, Jones MA, Scott DL, Prior P (1998) Longterm mortality outcome in patients with rheumatoid arthritis: early presenters continue to do well. J Rheumatol 25:1072–1077
- 5. Smitten AL, Choi HK, Hochberg MC, et al (2008) The risk of hospitalized infection in patients with rheumatoid arthritis. J Rheumatol 35:387–393
- 6. Doran MF, Crowson CS, Pond GR, et al (2002) Frequency of infection in patients with rheumatoid arthritis compared with controls: a population-based study. Arthritis Rheum 46:2287–2293. https://doi.org/10.1002/art.10524
- Tektonidou MG, Wang Z, Dasgupta A, Ward MM (2015) Burden of Serious Infections in Adults With Systemic Lupus Erythematosus: A National Population-Based Study, 1996-2011. Arthritis Care Res (Hoboken) 67:1078–1085. https://doi.org/10.1002/acr.22575
- Bosch X, Guilabert A, Pallarés L, et al (2006) Infections in systemic lupus erythematosus: a prospective and controlled study of 110 patients. Lupus 15:584–589. https://doi.org/10.1177/0961203306071919
- 9. Pene F, Merlat A, Vabret A, et al (2003) Coronavirus 229E-related pneumonia in immunocompromised patients. Clin Infect Dis 37:929–932. https://doi.org/10.1086/377612
- 10. Horisberger A, Moi L, Ribi C, Comte D (2020) [Autoimmune diseases in the context of pandemic COVID-19]. Rev Med Suisse 16:827–830
- 11. Favalli EG, Ingegnoli F, De Lucia O, et al (2020) COVID-19 infection and rheumatoid arthritis: Faraway, so close! Autoimmun Rev 19:102523. https://doi.org/10.1016/j.autrev.2020.102523
- 12. Askanase AD, Khalili L, Buyon JP (2020) Thoughts on COVID-19 and autoimmune diseases. Lupus Sci Med 7:e000396. https://doi.org/10.1136/lupus-2020-000396
- 13. Niewold TB (2011) Interferon alpha as a primary pathogenic factor in human lupus. J Interferon Cytokine Res 31:887–892. https://doi.org/10.1089/jir.2011.0071
- 14. D'Silva KM, Serling-Boyd N, Wallwork R, et al (2020) Clinical characteristics and outcomes of patients with coronavirus disease 2019 (COVID-19) and rheumatic disease: a comparative cohort study from a US 'hot spot.' Annals of the Rheumatic Diseases
- 15. Liu M, Gao Y, Zhang Y, et al (2020) The association between severe or death COVID-19 and autoimmune disease: a systematic review and meta-analysis. Journal of Infection. https://doi.org/10.1016/j.jinf.2020.05.065
- 16. Emmi G, Bettiol A, Mattioli I, et al (2020) SARS-CoV-2 infection among patients with systemic autoimmune diseases. Autoimmun Rev 19:102575. https://doi.org/10.1016/j.autrev.2020.102575
- 17. International MG/COVID-19 Working Group, Jacob S, Muppidi S, et al (2020) Guidance for the management of myasthenia gravis (MG) and Lambert-Eaton myasthenic syndrome (LEMS) during the COVID-19 pandemic. J Neurol Sci 412:116803. https://doi.org/10.1016/j.jns.2020.116803
- Price E, MacPhie E, Kay L, et al (2020) Identifying rheumatic disease patients at high risk and requiring shielding during the COVID-19 pandemic. Clin Med (Lond). https://doi.org/10.7861/clinmed.2020-0160

- Matucci-Cerinic M, Bruni C, Allanore Y, et al (2020) Systemic sclerosis and the COVID-19 pandemic: World Scleroderma Foundation preliminary advice for patient management. Annals of the Rheumatic Diseases 79:724–726. https://doi.org/10.1136/annrheumdis-2020-217407
- 20. Pope JE (2020) What Does the COVID-19 Pandemic Mean for Rheumatology Patients? Curr Treatm Opt Rheumatol 1–4. https://doi.org/10.1007/s40674-020-00145-y
- 21. Price E, MacPhie E, Kay L, et al (2020) Identifying rheumatic disease patients at high risk and requiring shielding during the COVID-19 pandemic. Clin Med (Lond). https://doi.org/10.7861/clinmed.2020-0160
- 22. Monti S, Balduzzi S, Delvino P, et al (2020) Clinical course of COVID-19 in a series of patients with chronic arthritis treated with immunosuppressive targeted therapies. Annals of the Rheumatic Diseases 79:667–668. https://doi.org/10.1136/annrheumdis-2020-217424
- 23. Nicastri E, Petrosillo N, Bartoli TA, et al (2020) National Institute for the Infectious Diseases "L. Spallanzani" IRCCS. Recommendations for COVID-19 Clinical Management. Infectious Disease Reports 12:. https://doi.org/10.4081/idr.2020.8543
- 24. Attivazione della presa in carico dei pazienti cronici e fragili: DGR n. X/6164 del 30.01.17 "Governo della domanda."

https://www.regione.lombardia.it/wps/portal/istituzionale/HP/DettaglioRedazionale/servizi-einformazioni/Enti-e-Operatori/sistema-welfare/attuazione-della-riforma-sociosanitarialombarda/avvio-presa-carico-pazienti-cronici-fragili/dgr2017-6164-avvio-presa-carico-pazienticronici-fragili. Accessed 14 Jul 2020

- 25. Avvio del percorso di presa in carico dei pazienti cronici e fragili: DGR n. X/7655 del 28.12.17. https://www.regione.lombardia.it/wps/portal/istituzionale/HP/DettaglioRedazionale/servizi-einformazioni/Enti-e-Operatori/sistema-welfare/attuazione-della-riforma-sociosanitarialombarda/dgr2017-7655-avvio-presa-carico-cronici/dgr2017-7655-avvio-presa-carico-cronici. Accessed 14 Jul 2020
- 26. Vandenbroucke JP, Brickley EB, Vandenbroucke-Grauls CMJE, Pearce N (2020) The test-negative design with additional population controls: a practical approach to rapidly obtain information on the causes of the SARS-CoV-2 epidemic. arXiv:200406033 [q-bio, stat]
- 27. Kelly S Fitting a Cumulative Logistic Regression Model. 5
- 28. 22954 The PROC LOGISTIC proportional odds test and fitting a partial proportional odds model. https://support.sas.com/kb/22/954.html. Accessed 17 Jun 2020
- 29. Wang L, He W, Yu X, et al (2020) Coronavirus disease 2019 in elderly patients: Characteristics and prognostic factors based on 4-week follow-up. J Infect 80:639–645. https://doi.org/10.1016/j.jinf.2020.03.019
- 30. Du R-H, Liu L-M, Yin W, et al (2020) Hospitalization and Critical Care of 109 Decedents with COVID-19 Pneumonia in Wuhan, China. Ann Am Thorac Soc. https://doi.org/10.1513/AnnalsATS.202003-2250C

- 31. Wei Y-Y, Wang R-R, Zhang D-W, et al (2020) Risk factors for severe COVID-19: Evidence from 167 hospitalized patients in Anhui, China. Journal of Infection. https://doi.org/10.1016/j.jinf.2020.04.010
- 32. Chen T, Wu D, Chen H, et al (2020) Clinical characteristics of 113 deceased patients with coronavirus disease 2019: retrospective study. BMJ 368:. https://doi.org/10.1136/bmj.m1091

Tables

Table 1. Overall demographic and clinical characteristics of test-positive and test-negative subjects, and population controls (for the case-control design 2) by autoimmune status.

Table 2. Un-adjusted and adjusted Odds Ratios for COVID-19 according to autoimmune disease from case-control designs.

Table 3. Un-adjusted and adjusted Odds Ratios for health status among test-positive subjects according to autoimmune disease from ordinal and multinomial logistic models. Health status was defined as non-hospitalized and alive (n=8451, 41.5%), hospitalized and alive (n=8073, 39.6%), and deceased (n=3840, 18.9%).

Supplementary Files

This is a list of supplementary files associated with this preprint. Click to download.

- eTable1.pdf
- Additionalfile13.pdf

	Test-positive subjects			Test-negative subjects			Controls		
							(for ca	se-control de	sign 2)
		Autoim	imune		Autoim	imune		Autoim	mune
		sta	tus		stat	tus		sta	tus
	Overall	No	Yes	Overall	No	Yes	Overall	No	Yes
	n=20364	n=19699	n=665	n=34697	n=33400	n=1297	n=81414	n=79245	n=2169
		96.7%	3.3%		96.3%	3.7%		97.3%	2.7%
Gender = Male (%)	9666	9496	170	13658	13438	220	38659	38127	532
	(47.5)	(48.2)	(25.6)	(39.4)	(40.2)	(17)	(47.5)	(48.1)	(24.5)
Age mean (sd)	65.5	65.6	64.5	54.8	54.8	56.9	65.5	65.5	66.2
	(19.4)	(19.5)	(16.1)	(20.8)	(20.95)	(16.4)	(19.4)	(19.5)	(16.1)
Age class (years) (%)									
<17	171	170	1	905	902	3	684	682	2
	(0.8)	(0.9)	(0.2)	(2.6)	(2.7)	(0.2)	(0.8)	(0.9)	(0.1)
[18-40)	1907	1868	39	7521	7353	168	7629	7512	117
	(9.4)	(9.5)	(5.9)	(21.7)	(22)	(13)	(9.4)	(9.5)	(5.4)
[40-70)	8622	8264	358	17047	16241	806	34488	33452	1036
	(42.3)	(42)	(53.8)	(49.1)	(48.6)	(62.1)	(42.4)	(42.2)	(47.8)
≥70	9664	9397	267	9224	8904	320	38613	37599	1014
2,0	(47.5)	(47.7)	(40.2)	(26.6)	(26.7)	(24.7)	(47.4)	(47.4)	(46.7)
Geographic location (%)	(1710)	(1)))	(10.2)	(2010)	(2017)	(21.7)	(1/11)	(1)11)	(1017)
Lodi province (start of the outburst)	2939	2836	103	5185	5001	184	11675	11356	319
	(14.4)	(14.4)	(15.5)	(14.9)	(15)	(14.2)	(14.3)	(14.3)	(14.7)
City of Milan									
City of Milan	7235	7025	210	13188	12761	427	29185	28437	748
Miles annia a	(35.5)	(35.7)	(31.6)	(38)	(38.2)	(32.9)	(35.8)	(35.9)	(34.5)
Milan province	10190	9838	352	16324	15638	686	40554	39452	1102
	(50)	(49.9)	(52.9)	(47)	(46.8)	(52.9)	(49.8)	(49.8)	(50.8)
Setting (%)									
Home	6122	5861	261	-	-	-	-	-	-
	(30.1)	(29.8)	(39.2)						
Residential	3381	3319	62	-	-	-	-	-	-
	(16.6)	(16.8)	(9.3)						
Residential followed by hospitalization	556	548	8	-	-	-	-	-	-
	(2.7)	(2.8)	(1.2)						
Hospitalized	10305	9971	334	-	-	-	-	-	-
	(50.6)	(50.6)	(50.2)						
Health status									
Non-hospitalized and alive	8292	7998	294	-	-	-	-	-	-
	(40.7)	(40.6)	(44.2)						
Hospitalized and alive	7949	7692	257	-	-	-	-	-	-
	(39.0)	(39.0)	(38.6)						
Deceased	4123	4009	114	-	-	-	-	-	-
	(20.2)	(20.4)	(17.1)						
Number of comorbidities									
(Excluded autoimmune diseases)									
None	8265	8105	160	18870	18457	413	36658	36131	527
	(40.6)	(41.1)	(24.1)	(54.4)	(55.3)	(31.8)	(45)	(45.6)	(24.3)
1-3	8951	8581	370	11959	11286	673	36156	34862	1294
	(44)	(43.6)	(55.6)	(34.5)	(33.8)	(51.9)	(44.4)	(44)	(59.7)
≥4	3148	3013	135	3868	3657	211	8600	8252	348
	(15.5)	(15.3)	(20.3)	(11.1)	(10.9)	(16.3)	(10.6)	(10.4)	(16)
Comorbidition $(9/)$	(10.0)	(10.0)	(20.0)	(****)	(10.0)	(10.0)	(10.0)	(10,1)	(10)

Comorbidities (%)

Transplanted any time = Yes (%)	77	75	2	227	221	6	159	155	4
	(0.4)	(0.4)	(0.3)	(0.7)	(0.7)	(0.5)	(0.2)	(0.2)	(0.2)
	30	27	3	64	57	7	85	80	5
Blood and Hematopoietic organs = Yes (%)	(0.1)	(0.1)	(0.5)	(0.2)	(0.2)	(0.5)	(0.1)	(0.1)	(0.2)
	72	72	0	224	220	4	261	259	2
HIV infection or AIDS = Yes (%)	(0.4)	(0.4)	(0)	(0.6)	(0.7)	(0.3)	(0.3)	(0.3)	(0.1)
	1024	923	101	2027	1869	158	3303	3023	280
Tumor in first line treatment = Yes (%)	(5)	(4.7)	(15.2)	(5.8)	(5.6)	(12.2)	(4.1)	(3.8)	(12.9)
	788	759	29	1128	1071	57	3153	3050	103
Tumor in follow-up, 1-5years = Yes (%)	(3.9)	(3.9)	(4.4)	(3.3)	(3.2)	(4.4)	(3.9)	(3.8)	(4.7)
	1087	1038	49	1337	1269	68	4238	4088	150
Tumor in remission after 5 years = Yes (%)	(5.3)	(5.3)	(7.4)	(3.9)	(3.8)	(5.2)	(5.2)	(5.2)	(6.9)
	28	28	0	79	76	3	58	54	4
Type 1 Diabetes = Yes (%)	(0.1)	(0.1)	(0)	(0.2)	(0.2)	(0.2)	(0.1)	(0.1)	(0.2)
	2426	2357	69	2537	2434	103	7660	7433	227
Type 2 Diabetes = Yes (%)	(11.9)	(12)	(10.4)	(7.3)	(7.3)	(7.9)	(9.4)	(9.4)	(10.5)
	423	408	15	527	503	24	750	722	28
Complicated DM Type 1 and $2 = $ Yes (%)	(2.1)	(2.1)	(2.3)	(1.5)	(1.5)	(1.9)	(0.9)	(0.9)	(1.3)
	2301	2216	85	2614	2495	119	9607	9306	301
Hypercholesterolaemia = Yes (%)	(11.3)	(11.2)	(12.8)	(7.5)	(7.5)	(9.2)	(11.8)	(11.7)	(13.9)
	8187	7896	291	9454	9016	438	32202	31267	935
Arterial hypertension = Yes (%)	(40.2)	(40.1)	(43.8)	(27.2)	(27)	(33.8)	(39.6)	(39.5)	(43.1)
	2364	2297	67	2744	2617	127	7497	7273	224
Ischemic heart disease = Yes (%)	(11.6)	(11.7)	(10.1)	(7.9)	(7.8)	(9.8)	(9.2)	(9.2)	(10.3)
	475	458	17	678	647	31	1328	1287	41
Valvular heart disease = Yes (%)	(2.3)	(2.3)	(2.6)	(2)	(1.9)	(2.4)	(1.6)	(1.6)	(1.9)
	2310	2241	69	2614	2503	111	6850	6634	216
Cardiomyopathy with arrhythmia = Yes (%)	(11.3)	(11.4)	(10.4)	(7.5)	(7.5)	(8.6)	(8.4)	(8.4)	(10)
Cardiomyopathy without arrhythmia = Yes	1699	1641	58	2025	1941	84	5278	5129	149
(%)	(8.3)	(8.3)	(8.7)	(5.8)	(5.8)	(6.5)	(6.5)	(6.5)	(6.9)
	1401	1346	55	1764	1680	84	3567	3463	104
Chronic hearth failure = Yes (%)	(6.9)	(6.8)	(8.3)	(5.1)	(5)	(6.5)	(4.4)	(4.4)	(4.8)
	595	573	22	820	784	36	1334	1290	44
Peripheral artery Disease = Yes (%)	(2.9)	(2.9)	(3.3)	(2.4)	(2.3)	(2.8)	(1.6)	(1.6)	(2)
	201	192	9	281	259	22	580	562	18
Venous diseases = Yes (%)	(1)	(1)	(1.4)	(0.8)	(0.8)	(1.7)	(0.7)	(0.7)	(0.8)
	822	799	23	932	884	48	1653	1607	46
Cerebrovascular disease = Yes (%)	(4)	(4.1)	(3.5)	(2.7)	(2.6)	(3.7)	(2)	(2)	(2.1)
	1137	783	354	2149	1357	792	4106	2872	1234
Thyroid diseases = Yes (%)	(5.6)	(4)	(53.2)	(6.2)	(4.1)	(61.1)	(5)	(3.6)	(56.9)
	78	70	8	173	150	23	225	211	14
Other endocrine diseases = Yes (%)	(0.4)	(0.4)	(1.2)	(0.5)	(0.4)	(1.8)	(0.3)	(0.3)	(0.6)
	359	22	337	601	47	554	1095	63	1032
Other autoimmune diseases= Yes (%)	(1.8)	(0.1)	(50.7)	(1.7)	(0.1)	(42.7)	(1.3)	(0.1)	(47.6)
	271	257	14	393	380	13	598	577	21
Epilepsy = Yes (%)	(1.3)	(1.3)	(2.1)	(1.1)	(1.1)	(1)	(0.7)	(0.7)	(1)
	669	655	14	579	562	17	989	968	21
Alzheimer and dementias $=$ Yes (%)	(3.3)	(3.3)	(2.1)	(1.7)	(1.7)	(1.3)	(1.2)	(1.2)	(1)
	310	305	5	287	279	8	616	596	20
Parkinson and parkinsonisms = Yes (%)	(1.5)	(1.5)	(0.8)	(0.8)	(0.8)	(0.6)	(0.8)	(0.8)	(0.9)
	99	94	5	232	218	14	263	257	6
Nervous system diseases, others = Yes (%)	(0.5)	(0.5)	(0.8)	(0.7)	(0.7)	(1.1)	(0.3)	(0.3)	(0.3)
· · · · · · · · · · · · · · · · ·	(0.0)	(0.0)	(0.0)	(0.7)	(0.7)	(1.1)	(0.0)	(0.0)	(0.0)

Chronic hepatitis and cirrhosis = Yes (%)	433	413	20	705	666	39	1293	1248	45
	(2.1)	(2.1)	(3)	(2)	(2)	(3)	(1.6)	(1.6)	(2.1)
Digestive system diseases, others = Yes	254	243	11	459	434	25	763	710	53
(%)	(1.2)	(1.2)	(1.7)	(1.3)	(1.3)	(1.9)	(0.9)	(0.9)	(2.4)
	897	864	33	1089	1039	50	2433	2345	88
COPD = Yes (%)	(4.4)	(4.4)	(5)	(3.1)	(3.1)	(3.9)	(3)	(3)	(4.1)
	97	91	6	151	142	9	221	207	14
RF or oxygen therapy = Yes (%)	(0.5)	(0.5)	(0.9)	(0.4)	(0.4)	(0.7)	(0.3)	(0.3)	(0.6)
	443	422	21	833	776	57	1369	1303	66
Asthma = Yes (%)	(2.2)	(2.1)	(3.2)	(2.4)	(2.3)	(4.4)	(1.7)	(1.6)	(3)
	590	557	33	716	681	35	1464	1409	55
CKD = Yes (%)	(2.9)	(2.8)	(5)	(2.1)	(2)	(2.7)	(1.8)	(1.8)	(2.5)
	151	147	4	443	427	16	122	121	1
Dialysis dependent CKD = Yes (%)	(0.7)	(0.7)	(0.6)	(1.3)	(1.3)	(1.2)	(0.1)	(0.2)	(0)
Autoimmune diseases									
Acquired autoimmune hemolytic anemia	-	-	10	-	-	11	-	-	19
			(1.5)			(0.8)			(0.9)
Systemic sclerosis	-	-	21	-	-	47	-	-	51
			(3.2)			(3.6)			(2.4)
Ankylosing spondylitis	-	-	13	-	-	22	-	-	42
			(2)			(1.7)			(1.9)
Myasthenia gravis	-	-	20	-	-	21	-	-	39
			(3)			(1.6)			(1.8)
Rheumatoid arthritis	-	-	167	-	-	223	-	-	479
			(25.1)			(17.2)			(22.1)
Psoriatic arthritis	-	-	69	-	-	139	-	-	284
			(10.4)			(10.7)			(13.1)
Systemic lupus erythematosus	-	-	16	-	-	38	-	-	57
			(2.4)			(2.9)			(2.6)
Sjögren disease	-	-	26	-	-	60	-	-	70
			(3.9)			(4.6)			(3.2)
	_	-	323	_	-	736	-	-	1128
Hashimoto's disease									

			OR (95% CI)		
Designs	Exposure		Un-adjusted	Adjusted	
Test-negative case-control*	Autoimmune status	No	Referent	Referent	
		Yes	0.86 (0.77- 0.96)	0.90 (0.80- 1.01)	
Positive case-control (CC-POS)*	Autoimmune status	No	Referent	Referent	
		Yes	1.51 (1.38-	1.00 (0.92-	
			1.64)	1.10)	
Negative case-control (CC- NEG)*	Autoimmune status	No	Referent	Referent	
		Yes	1.75 (1.62- 1.90)	1.15 (1.05- 1.25)	
Population case-control (case-control design 2)**	Autoimmune status	No	Referent	Referent	
		Yes	1.24 (1.13- 1.35)	1.16 (1.06- 1.26)	

*Odds ratios and corresponding 95% confidence intervals were calculated from a multivariable logistic model adjusted for gender, age (categorized as <17, [18-40), [40-70), \geq 70 years), number of non-autoimmune chronic conditions (categorized as no conditions, 1-3, \geq 4), and municipality of residence (for CC-POS and CC-NEG). Testpositive subjects were matched to test-negatives by ASST and date, date of positive swab for cases and date of negative swab for controls within 7 days of the case index date.

**Odds ratios and corresponding 95% confidence intervals were calculated from a conditional multivariable logistic model adjusted for number of non-autoimmune chronic conditions (categorized as no conditions, 1-3, ≥ 4); cases and controls matched by gender, age and municipality of residence.

Model	Health Status	Exposure		Un-	Adjusted OR
				adjusted OR	(95% CI)
				(95% CI)	
Ordinal Logistic	-	Autoimmune	No	Referent	Referent
Model* [¥]		status			
			Yes	0.85 (0.73-	0.96 (0.83-
				0.98)	1.12)
Multinomial	Being hospitalized and alive vs non	Autoimmune	No	Referent	Referent
Logistic Model*	hospitalized and alive	status			
			Yes	0.91 (0.77-	1.05 (0.88-
				1.08)	1.25)
	Being deceased vs non hospitalized and	Autoimmune	No	Referent	Referent
	alive (95% CI)	status			
			Yes	0.77 (0.62-	0.95 (0.74-
				0.96)	1.22)

*Odds ratios and corresponding 95% confidence intervals were calculated from logistic models adjusted for gender, age (categorized as <17, [18-40), [40-70), \geq 70 years), number of non-autoimmune chronic conditions (categorized as no conditions, [1-4), \geq 4), and ASST.

[¥]Age, gender and ASST did not satisfy the proportional odds assumption thus they have been inserted as having unequal slopes in the ordinal logistic regression model.