

STI/HIV testing and treatment in all US states and the District of Columbia as of October 2021. Most states also allowed minors to consent independently to STI/HIV prevention, including human papillomavirus vaccination and HIV preexposure and postexposure prophylaxis. However, some states required clinicians to apply specific criteria before minors can consent. Most states neglected or only cursorily addressed confidentiality obligations for clinicians who care for independently consenting minors. In states that permit or require that STI/HIV services be kept confidential from minors' guardians, clinicians will need to identify and implement practices to avoid inadvertent disclosure via insurance billing or electronic health records. Clinicians may need to consult additional state or federal regulations, such as the 21st Century Cures Act, to develop these procedures. This study did not assess municipal or federal law or changes after October 2021.

Minor consent laws are structured to protect clinicians who rely on minors' independent consent when providing STI/HIV services. These statutes therefore benefit both minors and clinicians, allowing minors to obtain STI/HIV services without involving their guardians, and enabling clinicians to provide these services to minors without risking legal sanctions. Due to low levels of knowledge about these laws and a dearth of institutional policies and procedures to support their use, minors often do not receive the services they need.⁶ Trainings, policies, and procedures that support and routinize the application of these statutes may empower clinicians to rely on them more confidently in practice. Ensuring that clinicians, researchers, and minors understand and trust these minor consent laws may expand access to STI/HIV services for youth.

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1. Leichter JS, Copen C, Dittus PJ. Confidentiality issues and use of sexually transmitted disease services among sexually experienced persons aged 15-25 years: United States, 2013-2015. *MMWR Morb Mortal Wkly Rep*. 2017;66(9):237-241. doi:10.15585/mmwr.mm6609a1
2. Culp L, Caucci L. State adolescent consent laws and implications for HIV pre-exposure prophylaxis. *Am J Prev Med*. 2013;44(1)(suppl 2):S119-S124. doi:10.1016/j.amepre.2012.09.044
3. English A, Bass L, Boyle AD, Eshragh F. *State Minor Consent Laws: A Summary*. 3rd ed. Center for Adolescent Health & the Law; 2010.
4. Guttmacher Institute. Minors' access to STI services. Accessed March 23, 2022. <https://www.guttmacher.org/state-policy/explore/minors-access-sti-services>
5. Anderson ED, Tremper C, Thomas S, Wagenaar AC. Measuring statutory law and regulations for empirical research. In: Wagenaar AC, Burris SC, eds. *Public Health Law Research: Theory and Methods*. Wiley; 2013:237-260.
6. Pampati S, Lidson N, Dittus PJ, Adkins SH, Steiner RJ. Confidentiality matters but how do we improve implementation in adolescent sexual and reproductive health care? *J Adolesc Health*. 2019;65(3):315-322. doi:10.1016/j.jadohealth.2019.03.021

Association Between BNT162b2 Vaccination and Long COVID After Infections Not Requiring Hospitalization in Health Care Workers

Survivors of COVID-19 may present with long-lasting symptoms.¹ Some factors have been associated with the development of post-COVID conditions (also referred to as "long COVID"),² including hospitalization.³ A study of older US veterans showed 15% reduction of long COVID after vaccination; however, study limitations included the low number of women and suboptimal vaccination schedules.⁴

Methods | The study was approved by the Humanitas Research Hospital institutional review board. Each participant provided written informed consent.

We conducted an observational cohort study from March 2020 to April 2022 in individuals working in 9 Italian health care facilities.^{5,6} Polymerase chain reaction (PCR) tests for SARS-CoV-2 were conducted every week (in COVID wards) or 2 weeks (in other wards) for hospital personnel, or if they developed symptoms or were exposed to cases. All health care workers were required to receive 3 doses of vaccine (BNT162b2), with the first and second doses administered in January-February 2021 and the booster dose in November-December 2021.



Supplemental content

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women and suboptimal vaccination schedules.⁴

Table 1. Characteristics of the Nonhospitalized Study Population of Routinely Tested Health Care Personnel With COVID-19 (N = 739)

| | Had long COVID | | Did not have long COVID | | P value |
|--|----------------|-------------------------|-------------------------|-------------------------|--------------------|
| | No. | % (95% CI) ^a | No. | % (95% CI) ^a | |
| No. | 229 | 31.0 (27.7-34.5) | 510 | 69.0 (65.5-72.3) | .11 ^b |
| Women | 180 | 32.7 (28.8-36.8) | 371 | 67.3 (63.2-71.2) | |
| Men | 49 | 26.1 (19.9-33.0) | 139 | 73.9 (67.0-80.1) | |
| Age, mean (SD), y | 44.3 (10.7) | | 41.2 (11.4) | | <.001 ^c |
| BMI, mean (SD) | 24.3 (4.3) | | 23.5 (3.7) | | .01 ^c |
| COVID-19 wave ^d | | | | | <.001 ^b |
| 1 | 74 | 48.1 (39.9-56.2) | 80 | 51.9 (43.8-60.1) | |
| 2 | 108 | 35.9 (30.5-41.6) | 193 | 64.1 (58.4-69.5) | |
| 3 | 47 | 16.5 (12.4-21.4) | 237 | 83.5 (78.6-87.6) | |
| Vaccine doses before SARS-CoV-2 infection ^e | | | | | <.001 ^b |
| 0 | 176 | 41.8 (37.0-46.7) | 245 | 58.2 (53.3-63.0) | |
| 1 | 3 | 30.0 (6.7-65.2) | 7 | 70.0 (34.8-93.3) | |
| 2 | 8 | 17.4 (7.8-31.4) | 38 | 82.6 (68.6-92.2) | |
| 3 | 42 | 16.0 (11.8-21.0) | 220 | 84.0 (79.0-88.2) | |
| Comorbidities | | | | | |
| Allergies | 104 | 36.5 (30.9-42.4) | 181 | 63.5 (57.6-69.1) | .01 ^b |
| Heart and cardiovascular diseases | 34 | 40.0 (29.5-51.2) | 51 | 60.0 (48.8-70.5) | .07 ^b |
| Obstructive lung disease (asthma/COPD/bronchiectasis) | 28 | 46.7 (33.7-60.0) | 32 | 53.3 (40.0-66.3) | .009 ^b |
| Autoimmune and rheumatic diseases | 21 | 43.8 (29.5-58.8) | 27 | 56.2 (41.2-70.5) | .07 ^b |
| Metabolic disease | 18 | 34.0 (21.5-48.3) | 35 | 66.0 (51.7-78.5) | .74 ^b |
| Cancer | 5 | 21.7 (7.5-43.7) | 18 | 78.3 (56.3-92.5) | .46 ^b |
| Pregnancy or breastfeeding | 5 | 33.3 (11.8-61.6) | 10 | 66.7 (38.4-88.2) | .79 ^b |
| Anemia/hemoglobinopathies/coagulation disorders | 3 | 23.1 (5.0-53.8) | 10 | 76.9 (46.2-95.0) | .76 ^b |
| Mental health conditions | 3 | 60.0 (14.7-94.7) | 2 | 40.0 (5.3-85.3) | .18 ^f |
| IBD | 2 | 40.0 (5.3-85.3) | 3 | 60.0 (14.7-94.7) | .65 ^f |
| GERD | 2 | 100.0 (15.8-100) | 0 | 0.0 (0-84.2) | .09 ^f |

Abbreviations: BMI, body mass index (calculated as weight in kilograms divided by height in meters squared); COPD, chronic obstructive pulmonary disease; GERD, gastroesophageal reflux disease; IBD, inflammatory bowel disease.

^a The 95% CIs for the prevalence data were calculated using the Clopper-Pearson method.

^b χ^2 test.

^c Mann-Whitney U test.

^d Wave 1: February-September 2020 (wild-type variant), wave 2, October 2020-July 2021 (Alpha variant), and wave 3, August 2021-March 2022 (Delta and Omicron variants).

^e The average periods of the vaccine administration were January 2021 (first dose), February 2021 (second dose), and November 2021 (third dose).

^f Fisher exact test.

Between February and April 2022, each participant completed a survey including demographics, comorbidities, a list of SARS-CoV-2-related symptoms at the time of infection and their duration (survey in the [Supplement](#)), and vaccination status. We defined long COVID as reporting at least 1 SARS-CoV-2-related symptom with a duration of more than 4 weeks. Hospitalized individuals were excluded to avoid bias related to severe disease, as were individuals with a date of infection less than 28 days before the survey. We included asymptomatic infections in the acute infection group (they could not have long COVID by definition) to avoid overestimating the prevalence of long COVID. The analysis was restricted to health care workers who were tested every 1 or 2 weeks with complete demographic data and a documented positive result for SARS-CoV-2 between March 2020 and March 2022.

By the date of infection, we divided the patients into 3 groups corresponding to the peaks in our data and circulation of variants of concern in Italy (wave 1, February-September 2020 [wild-type variant]; wave 2, October 2020-July 2021 [Alpha]; and wave 3, August 2021-March 2022 [Delta and Omicron]) (eFigure in the [Supplement](#)). A multivariable logistic regression model was used to assess the relationship between long COVID and characteristics, including participant sex, age, SARS-CoV-2 infection, wave, and vaccination status

14 days prior to infection. Time since second vaccination was assessed among vaccinated individuals.

The Clopper-Pearson method was used to calculate 95% CIs and the Mann-Whitney U test or the *t* test for continuous variables and the χ^2 -test for categorical variables to calculate *P* values. The significance threshold was defined as *P* < .05 (2-sided). Analyses were done in Python, version 3.8.3.

Results | Of 2560 participants, 739 individuals (29%) had COVID-19 (89 asymptomatic), of whom 229 (31.0%; 95% CI, 27.7%-34.5%) had long COVID (**Table 1**). The prevalence of long COVID varied across the pandemic waves, from 48.1% (95% CI, 39.9%-56.2%) in wave 1 to 35.9% (95% CI, 30.5%-41.6%) in wave 2 to 16.5% (95% CI, 12.4%-21.4%) in wave 3. The number of vaccine doses was associated with lower long COVID prevalence: 41.8% (95% CI, 37.0%-46.7%) in unvaccinated patients, 30.0% (95% CI, 6.7%-65.2%) with 1 dose, 17.4% (95% CI, 7.8%-31.4%) with 2 doses, and 16.0% (95% CI, 11.8%-21.0%) with 3 doses. Older age, higher body mass index, allergies, and obstructive lung disease were associated with long COVID.

With a reference group of unvaccinated females in wave 1 with no allergies or comorbidities (**Table 2**), male sex (odds ratio [OR], 0.65; 95% CI, 0.44-0.98, *P* = .04), 2 vaccine doses (OR, 0.25; 95% CI, 0.07-0.87, *P* = .03), and 3 vaccine doses (OR, 0.16;

Table 2. Multivariable Logistic Regression Analysis of the Association of Long COVID (N = 229) With Patient Characteristics^a

| | OR (95% CI) | P value |
|-----------------------------------|------------------|---------|
| Male sex | 0.65 (0.44-0.98) | .04 |
| Age ^b | 1.23 (1.01-1.49) | .04 |
| BMI ^b | 1.10 (0.92-1.31) | .30 |
| Allergies | 1.50 (1.06-2.11) | .02 |
| No. of comorbidities ^c | 1.32 (1.04-1.68) | .03 |
| COVID-19 wave | | |
| 2 | 0.72 (0.48-1.08) | .11 |
| 3 | 1.34 (0.26-7.01) | .73 |
| Vaccine dose ^d | | |
| 1 | 0.86 (0.21-3.49) | .83 |
| 2 | 0.25 (0.07-0.87) | .03 |
| 3 | 0.16 (0.03-0.84) | .03 |

Abbreviations: BMI, body mass index; OR, odds ratio.

^a Reference model: women in COVID-19 wave 1 with 0 doses of vaccine, with no allergies and no comorbidities.

^b Age and BMI have been standardized (mean = 0; SD = 1). Age SD = 11.3 years; BMI SD = 3.9.

^c Number of comorbidities is a discrete variable ranging from 0 to 4, where 4 represents 4 or more different comorbidities.

^d At least 14 days prior to infection.

95% CI, 0.03-0.84, $P = .03$) were associated with a lower probability of long COVID. Older age (OR, 1.23; 95% CI, 1.01-1.49, $P = .04$), allergies (OR, 1.50; 95% CI, 1.06-2.11, $P = .02$), and an increasing number of comorbidities (OR, 1.32; 95% CI, 1.04-1.68, $P = .03$) were associated with a higher probability. No statistically significant association with infection wave was found. Among vaccinated individuals ($n = 265$), time between the second vaccination dose and infection was not associated with long COVID (OR, 0.66; 95% CI, 0.34-1.29).

Discussion | In this longitudinal observational study conducted among health care workers with SARS-CoV-2 infections not requiring hospitalization, 2 or 3 doses of vaccine, compared with no vaccination, were associated with lower long COVID prevalence. Study limitations include that symptoms and duration were self-reported, and causality cannot be inferred.

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1. Groff D, Sun A, Ssentongo AE, et al. Short-term and long-term rates of postacute sequelae of SARS-CoV-2 infection: a systematic review. *JAMA Netw Open*. 2021;4(10):e2128568. doi:10.1001/jamanetworkopen.2021.28568
2. Su Y, Yuan D, Chen DG, et al; ISB-Swedish COVID-19 Biobanking Unit. Multiple early factors anticipate post-acute COVID-19 sequelae. *Cell*. 2022;185(5):881-895.e20. doi:10.1016/j.cell.2022.01.014
3. Chen C, Hauptert SR, Zimmermann L, Shi X, Fritsche LG, Mukherjee B. Global prevalence of post COVID-19 condition or long COVID: a meta-analysis and systematic review. *J Infect Dis*. Published online April 16, 2022. doi:10.1093/infdis/jiac136
4. Al-Aly Z, Bowe B, Xie Y. Long COVID after breakthrough SARS-CoV-2 infection. *Nat Med*. Published online May 25, 2022. doi:10.1038/s41591-022-01840-0
5. Levi R, Azzolini E, Pozzi C, et al. One dose of SARS-CoV-2 vaccine exponentially increases antibodies in individuals who have recovered from symptomatic COVID-19. *J Clin Invest*. 2021;131(12):149154. doi:10.1172/JCI149154
6. Darwich A, Pozzi C, Fornasa G, et al; ICH COVID-19 Task-force. BNT162b2 vaccine induces antibody release in saliva: a possible role for mucosal viral protection? *EMBO Mol Med*. 2022;14(5):e15326. doi:10.15252/emmm.202115326

COMMENT & RESPONSE

Medical Need and Transplant Accessibility

To the Editor A recent Viewpoint¹ discussed inequities in solid organ transplants in the US and presented policy proposals to ameliorate these disparities. However, an additional barrier to transplant accessibility that should be considered is the stigma and ableism faced by individuals with intellectual and developmental disabilities (IDD).