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Association Between Tracheal Intubation During Pediatric In-Hospital Cardiac Arrest and Survival

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Abstract

IMPORTANCE—Tracheal intubation is common during pediatric in-hospital cardiac arrest, although the relationship between intubation during cardiac arrest and outcomes is unknown.

OBJECTIVE—To determine if intubation during pediatric in-hospital cardiac arrest is associated with improved outcomes.

DESIGN, SETTING, AND PARTICIPANTS—Observational study of data from United States hospitals in the Get With The Guidelines-Resuscitation registry. Pediatric patients (<18 years) with index in-hospital cardiac arrest between January 2000 and December 2014 were included. Patients who were receiving assisted ventilation, had an invasive airway in place, or both at the time chest compressions were initiated were excluded.

EXPOSURES—Tracheal intubation during cardiac arrest.

MAIN OUTCOMES AND MEASURES—The primary outcome was survival to hospital discharge. Secondary outcomes included return of spontaneous circulation and neurologic outcome. A favorable neurologic outcome was defined as a score of 1 to 2 on the pediatric cerebral performance category score. Patients being intubated at any given minute were matched with patients at risk of being intubated within the same minute (ie, still receiving resuscitation) based on a time-dependent propensity score calculated from multiple patient, event, and hospital characteristics.

RESULTS—The study included 2294 patients; 1308 (57%) were male, and all age groups were represented (median age, 7 months [25th-75th percentiles, 21 days, 4years]). Of the 2294 included patients, 1555 (68%) were intubated during the cardiac arrest. In the propensity score-matched cohort (n = 2270), survival was lower in those intubated compared with those not intubated (411/1135 [36%] vs 460/1135 [41%]; risk ratio [RR], 0.89 [95% CI, 0.81–0.99]; P= .03). There was no significant difference in return of spontaneous circulation (770/1135 [68%] vs 771/1135 [68%]; RR, 1.00 [95% CI, 0.95–1.06]; P= .96) or favorable neurologic outcome (185/987 [19%] vs 211/983 [21%]; RR, 0.87 [95% CI, 0.75–1.02]; P= .08) between those intubated and not intubated. The association between intubation and decreased survival was observed in the majority of the sensitivity and subgroup analyses, including when accounting for missing data and in a subgroup of patients with a pulse at the beginning of the event.

CONCLUSIONS AND RELEVANCE—Among pediatric patients with in-hospital cardiac arrest, tracheal intubation during cardiac arrest compared with no intubation was associated with decreased survival to hospital discharge. Although the study design does not eliminate the potential for confounding, these findings do not support the current emphasis on early tracheal intubation for pediatric in-hospital cardiac arrest.

Survival rates from pediatric in-hospital cardiac arrest are improving, yet mortality remains high.¹ Most pediatric in-hospital cardiac arrests are precipitated by acute respiratory failure (hypoxia), and relatively few are attributable to primary arrhythmias (eg, ventricular fibrillation or ventricular tachycardia).^{2,3} Therefore, the standard paradigm for pediatric in-

hospital cardiopulmonary resuscitation (CPR) focuses on early effective respiratory support to establish adequate oxygenation and ventilation and on high-quality chest compressions to provide adequate myocardial and cerebral blood flow.⁴ Although there is increasing evidence supporting the importance of high-quality chest compressions for survival from pediatric in-hospital cardiac arrests,⁵ less is known about the importance of invasive airway management during CPR.

Invasive airway management for children is often a challenging and potentially dangerous task. Therefore, it is not surprising that tracheal intubation in the prehospital setting did not improve outcomes compared with bag-valve-mask ventilation during pediatric emergencies, including cardiac arrest.⁶ Consequently, recommendations from the International Liaison Committee On Resuscitation Pediatric Task Force and the American Heart Association support the use of bag-valve-mask ventilation as opposed to tracheal intubation in patients with out-of-hospital pediatric cardiac arrest and with a short transport time.^{7,8}

In contrast to the prehospital setting, a large proportion of pediatric in-hospital cardiac arrests occur in an intensive care unit, emergency department, or interventional setting (eg, catheterization laboratory or operating room),⁹ where the hospital-based personnel are presumably highly trained for invasive airway management and ventilatory support.

The goal of the current investigation was to examine, among pediatric patients with inhospital cardiac arrest, whether tracheal intubation during cardiac arrest is associated with survival to hospital discharge. Because acute respiratory failure is a common cause of pediatric in-hospital cardiac arrest, it was hypothesized that tracheal intubation during cardiac arrest would be associated with improved survival.

Methods

Data Source

Data from the Get With The Guidelines-Resuscitation (GWTG-R) registry were used. This is an American Heart Association-sponsored prospective quality improvement registry of inhospital cardiac arrests in the United States. The data collection process and reliability have been described in full detail.^{10,11} In the registry, cardiac arrest is defined as pulselessness, or a pulse with inadequate perfusion, requiring chest compressions, defibrillation, or both and with a hospital-wide or unit-based emergency response by acute care facility personnel. Hospital-level data (teaching status and type of hospital) were obtained from the American Hospital Association's Annual Survey from 2013.¹²

All participating hospitals in the GWTG-R registry are required to comply with local regulatory guidelines. Because data are used primarily at the local site for quality improvement, sites are granted a waiver of informed consent under the common rule.

Study Population

The cohort included data submitted to the GWTG-R registry between January 2000 and December 2014. All patients younger than 18 years who had an index cardiac arrest with 1 minute or more of chest compressions were included. Patients were excluded if they were

receiving assisted ventilation, had an invasive airway (including a tracheostomy tube) in place at the time chest compressions were initiated, or both. Furthermore, hospital visitors and cardiac arrests in the delivery room or neonatal intensive care unit were excluded. For the primary analysis and majority of analyses, patients with missing or inconsistent data on tracheal intubation, timing of intubation, timing of the end of resuscitation, timing of pulselessness, survival, or relevant covariates were excluded (537 patients [19%], see Figure 1 for details). These patients were included in a preplanned sensitivity analysis accounting for missing data (see statistical analysis below).

Intubation and Study Outcomes

Intubation was defined as insertion of a tracheal tube during the cardiac arrest event. Unsuccessful intubation attempts were not counted as intubations. The time to intubation was defined as the time interval in whole minutes from the start of chest compressions until the tracheal tube was inserted. All times in the GWTG-R registry are collected in whole minutes. As such, a time to intubation of 0 minutes indicates that the intubation was performed within the same whole minute as chest compressions were started; a time of 1 minute indicates that intubation was performed within the next whole minute, etc.

The primary outcome was designated a priori as survival to hospital discharge. Secondary outcomes were return of spontaneous circulation (ROSC) and favorable neurologic outcome at hospital discharge. ROSC was defined as no further need for chest compressions (including initiation of cardiopulmonary bypass) that was sustained for longer than 20 minutes. Neurologic outcome was assessed as recommended by the Utstein guidelines¹³ using the pediatric cerebral performance category (PCPC) score,¹⁴ where a PCPC score of 1 indicates no neurologic deficit; 2, mild cerebral disability; 3, moderate cerebral disability; 4, severe cerebral disability; 5, coma or vegetative state; and 6, brain death. A PCPC score of 1 and 2 was considered a favorable neurologic outcome. Prospectively designated sensitivity analyses were performed with different definitions of favorable neurologic outcome: (1) a PCPC score of 1 or 2, or no increase from baseline; (2) a PCPC score of 1,2, or 3; and (3) a PCPC score of 1,2, or 3, or no increase from baseline.¹⁵

Data abstractors were not blinded to outcomes but were unaware of the hypothesis of the current study.

Statistical Analyses

The study population was characterized using descriptive statistics. Categorical variables are presented as counts and frequencies and continuous variables as medians with 25th and 75th percentiles. The χ^2 test was used to compare categorical data between independent groups and the Wilcoxon rank sum test to compare continuous variables. The Cochran-Armitage test was used to test for linear trends in tracheal intubation over time.

To assess the adjusted association between tracheal intubation during CPR and survival to hospital discharge, time-dependent propensity score matching was used.¹⁶ A similar methodology has been successfully applied to the adult cardiac arrest population and accounts for the time-dependent nature of interventions during cardiac arrest.¹⁷ If the timing

The propensity score was calculated using a multivariable Cox proportional hazards model as explained in additional details in the eMethods in the Supplement. All variables presented in Table 1 were included in the propensity score model (see eTable 1 in the Supplement for definitions of preexisting conditions). All variables were chosen a priori based on prior work or clinical reasoning derived from prior publications related to pediatric in-hospital cardiac arrest.^{1,11,19,20}

Matching (1:1) on the propensity score was performed using a nearest neighbor-matching algorithm with a maximum caliper of 0.01 of the propensity score. Patients being intubated at any given minute (from minute 0 to minute 15) were separately and sequentially propensity score matched with a patient who was "at risk" of being intubated within the same minute. Patients were matched up until 15 minutes after the cardiac arrest, because intubation at any given minute after 15 minutes was uncommon. At-risk patients included those still undergoing resuscitation (ie, excluding patients with ROSC or resuscitation terminated without ROSC) and who did not receive intubation before or within the same minute. At-risk patients also included patients intubated at a later point, because the matching should not be dependent on future events.^{16,17,21,22} For example, a patient intubated at minute 3 was matched with a patient who was at risk of intubation at minute 3 (ie, either a patient intubated after minute 3 or a patient never intubated). Replacement of controls was used to reduce the number of unmatched exposed patients.

To assess the performance of the matching, baseline categorical variables were compared between the matched groups using the Cochran-Mantel-Haenszel test, and standardized differences were calculated for which a difference between -0.1 and 0.1 is generally considered negligible.²³ Using the matched cohort, modified Poisson regression was performed to assess the association between intubation during CPR and survival to hospital discharge obtaining risk ratios (RRs) with robust variance estimates.^{24,25} To account for the matching, the correlation between resampled controls, and potential clustering within hospitals, generalized estimating equations were used as described by Miglioretti and Heagerty, accounting for the nonnested nature of these clusters.²⁶ The results from the regression models are reported as RRs with 95% confidence intervals.

Whether the association between intubation during CPR and survival varied with the duration of the event was examined by adding an interaction term between the intubation variable (yes/no) and time to intubation/matching (ie, time from start of chest compressions to the minute the patient was matched) to the modified Poisson regression model. Time was first treated as a linear continuous variable and next as a categorical variable (0–5, 6–10, and 11–15 minutes).

Three predefined sensitivity analyses were performed. First, multiple imputations were performed to account for missing data. Categorical variables were imputed using the fully conditional specification method, and continuous time variables were imputed using Poisson

distributions as described in more detail in the eMethods in the Supplement. Second, the primary analysis was performed after excluding those patients who received cardiopulmonary bypass (extracorporeal membrane oxygenation) during the event. Third, the main analysis was performed including only those patients who had 2 minutes or more of chest compressions and not including patients intubated 2 minutes or less after the start of chest compressions. A post hoc sensitivity analysis assessing the potential influence of an unmeasured confounder was performed using the method as described by Lin et al.²⁷ Additional details are provided in the eMethods in the Supplement. As a second post hoc sensitivity analysis, we conducted a traditional propensity score-matched analysis (ie, not accounting for the timing of intubation). Additional details are provided in the eMethods in the Supplement.

As a predefined subgroup analysis, the cohort was restricted to those patients who lost pulse at any time during the event. For this analysis, the time to intubation was calculated as the time from loss of pulse to intubation, and patients who were intubated before pulselessness were excluded. Then an approach similar to that described above was used to assess whether intubation during CPR was associated with outcomes in this subgroup of patients. As a post hoc analysis, a subgroup analysis was also conducted in only those who received CPR for bradycardia and inadequate perfusion (ie, pulse present at initiation of CPR).

Exploratory analyses tested whether the association between intubation during CPR and survival to hospital discharge differed according to multiple subgroups. The subgroups included age group, illness category, whether the patients had preceding respiratory insufficiency, and the location of the event. This was tested by adding interactions between the intubation variable and the subgroup variable of interest to the modified Poisson regression model in the propensity-matched cohort. Given the exploratory nature of these analyses, a more stringent 2-sided P value (<.01) was considered significant.

All hypothesis tests were 2-sided, with a significance level of P < .05 apart from the exploratory analyses of subgroup associations. All secondary analyses should be considered exploratory, as no adjustments except as noted were made for multiple comparisons. No power calculation was performed before the study. All statistical analyses were conducted using SAS version 9.4 (SAS Institute Inc).

Results

Characteristics of the Patient Population

The study included 2294 patients (Figure 1); 1308 (57%) were male, and all age groups were represented (median age, 7 months [25th-75th percentiles, 21 days, 4 years]). Of these, 1555 (68%) were intubated during the cardiac arrest. For these patients, the median documented time to successful tracheal intubation from start of compressions was 5 minutes (25th-75th percentiles, 2–11), see Figure 2. There was a decrease in the proportion of patients being intubated over time (79% in 2000 to 62% in 2014, P= .001 for linear trend). Only 13 of those intubated (0.8%) received a tracheostomy tube. Baseline characteristics according to intubation status are presented in Table 1.

Overall Outcomes and Unadjusted Analyses

Overall, 1162 patients (51%) survived to hospital discharge. In unadjusted analysis, tracheal intubation during CPR was associated with decreased survival to hospital discharge (667/1555 [43%] vs 495/739 [67%]; RR, 0.64 [95% CI, 0.59-0.69]; P < .001). ROSC was achieved in 1766 patients (77%). In unadjusted analysis, tracheal intubation during CPR was associated with decreased ROSC (1130/1555 [73%] vs 636/739 [86%]; RR, 0.84 [95% CI, 0.81–0.88]; P < .001). Data on neurologic outcomes at hospital discharge was missing for 414 patients (18%), all of whom survived, since nonsurvivors were classified as poor neurologic outcome. In patients with data on neurologic outcome, 557 of 1880 (30%) had a favorable neurologic outcome, which corresponds to 557 of 748 survivors (74%). In unadjusted analysis, tracheal intubation during CPR was associated with decreased favorable neurologic outcome (313/1318 [24%] vs 244/562 [43%]; RR, 0.55 [95% CI, 0.48–0.63]; P < .001).

Time-Dependent Propensity-Matched Multivariable Analyses

Of the 2294 patients, 2270 were matched based on the propensity score. The baseline characteristics of the matched cohort are presented in Table 2. The groups were well matched on all included variables (all *P* values >.10 and standardized differences between -0.1 and 0.1). In the propensity score-matched cohort, survival to hospital discharge was lower in those intubated compared with those not intubated during CPR (411/1135 [36%] vs 460/1135 [41%]; RR, 0.89 [95% CI, 0.81–0.99]; *P*= .03). There was no significant difference in ROSC (770/1135 [68%] vs 771/1135 [68%]; RR, 1.00 [95% CI, 0.95–1.06]; *P* = .96) or favorable neurologic outcome (185/987 [19%] vs 211/983 [21%]; RR, 0.87 [95% CI, 0.75–1.02]; *P*= .08) between those intubated and not intubated during CPR. The results are summarized in Figure 3. The results were similar in magnitude with each of the different sensitivity definitions of favorable neurologic outcome (eTable 2 in the Supplement). None of the interactions as described in the methods sections were significant, indicating no significant subgroup difference in this study for age group, illness category, preceding respiratory insufficiency, or the location of the event.

Timing of Intubation

There was no significant interaction between tracheal intubation during CPR and the timing of the matching for any of the outcomes for either the linear (P = .50) or the categorical (P = .86) version of time. Thus, the association between tracheal intubation during CPR and survival did not change with the duration of the cardiac arrest.

Sensitivity Analyses

The sensitivity analysis accounting for missing data included 2831 patients. Between 2736 and 2768 patients were matched in the 10 imputed data sets. The combined estimate from the 10 data sets showed that tracheal intubation during CPR was associated with decreased survival (RR, 0.88 [95% CI, 0.79–0.99]; P= .03) and decreased favorable neurologic outcome (RR, 0.85 [95% CI, 0.73–0.98]; P= .03). There was no association with ROSC (RR, 0.99 [95% CI, 0.93–1.05]; P= .79).

One hundred eleven patients (5%) received cardiopulmonary bypass (ie, extracorporeal CPR) during the event, 2 patients had missing data on this variable, and 2181 patients were included in the analysis excluding those patients; 2116 patients were matched based on the propensity score. In this cohort, tracheal intubation during CPR was associated with decreased survival (RR, 0.90 [95% CI, 0.81–0.99]; P= .04). Tracheal intubation during CPR was not associated with ROSC (RR, 1.01 [95% CI, 0.95–1.08]; P= .73) and not associated with favorable neurologic outcome (RR, 0.95 [95% CI, 0.82–1.09]; P= .44). Survival to hospital discharge was 41% (46/111) in those who received cardiopulmonary bypass, and 87% (97/111) were intubated during CPR.

In the sensitivity analysis only including patients receiving at least 2 minutes of chest compression, 1762 were propensity-score matched. In this group, tracheal intubation during CPR was associated with decreased survival to hospital discharge (RR, 0.90 [95% CI, 0.80– 1.00]; P = .04). Intubation was not associated with ROSC (RR, 1.00 [95% CI, 0.94–1.07]; P = .91) or favorable neurologic outcomes (RR, 0.89 [95% CI, 0.73–1.10]; P = .29).

The results for the post hoc sensitivity analysis exploring the potential influence of an unmeasured confounder are presented in the eFigure in the Supplement. Based on the findings of this analysis, it is unlikely that an unmeasured confounder is masking a positive association (ie, that intubation is associated with improved survival). For example, for the point estimate of the adjusted RR to be 1.20 (ie, a clinically meaningful effect), the unmeasured binary confounder would have to be prevalent and strongly associated with intubation (eg, a prevalence of 20% in those not intubated and 40% in those intubated) and be very strongly associated with increased mortality (an RR between 3.5 and 4.0), independent of measured confounders. None of the measured confounders (Table 1) comes close to fulfilling these criteria.

For the post hoc analysis using traditional propensity score matching, 1388 patients were matched. In this analysis, intubation compared with no intubation was associated with decreased survival (368/694 [53%] vs 458/694 [66%]; RR, 0.80 [95% CI, 0.71–0.91]; P < . 001), decreased ROSC (555/694 [80%] vs 591/694 [85%]; RR, 0.94 [95% CI, 0.88–1.00]; P = .05), and decreased good neurologic outcome (193/573 [34%] vs 231/536 [43%]; RR, 0.78 [95% CI, 0.66–0.93]; P = .005).

Subgroup Analyses

In the subgroup of patients documented as pulseless (ie, excluding those with severe bradycardia requiring CPR), 1494 patients were included. Of these patients, 1019 (68%) were intubated during CPR, and overall survival to hospital discharge was 40%. The median time from pulselessness to successful intubation in these patients was 5 minutes (25th-75th percentiles, 2–10). For the subgroup analyses, 1706 patients were matched; the groups were well matched on all included variables after propensity score matching (all *P* values >.10 and standardized differences between -0.1 and 0.1; see eTable 3 in the Supplement). In this subgroup, tracheal intubation during CPR was not associated with survival to hospital discharge (RR, 0.88 [95% CI, 0.77–1.01]; *P*=.07). Intubation also was not associated with ROSC (RR, 0.98 [95% CI, 0.91–1.05]; *P*=.53) or favorable neurologic outcomes (RR, 0.83 [95% CI, 0.68–1.02]; *P*=.07).

The post hoc analysis of the subgroup of patients who started the event with a pulse (ie, severe bradycardia requiring CPR) included 935 patients, with an overall survival to hospital discharge of 66%. Of these, 223 (24%) lost their pulse during the event. Five hundred seventy-three patients (61%) were intubated, with a median time from start of chest compressions to intubation of 6 minutes (25th-75th percentiles, 2–11). Six hundred fifty patients were matched. The groups were reasonably well matched on all included variables after propensity score matching (all *P* values >.01 and standardized differences between -0.18 and 0.18; see eTable 4 in the Supplement). In this subgroup, tracheal intubation during CPR was associated with decreased survival to hospital discharge (RR, 0.86 [95% CI, 0.75– 0.98]; *P* = .03) and decreased favorable neurologic outcome (RR, 0.72 [95% CI, 0.58–0.90]; *P* = .003). Intubation was not associated with ROSC (RR, 1.00 [95% CI, 0.95–1.06]; *P* = .90).

Discussion

Although tracheal intubation to achieve effective oxygenation and ventilation is a basic tenet of advanced CPR for pediatric in-hospital cardiac arrests, this study was unable to demonstrate any association of improvement in survival with tracheal intubation during CPR when using time-dependent propensity score matching in this large multicenter study of pediatric in-hospital cardiac arrest. Importantly, this study found that tracheal intubation during CPR was associated with decreased survival to hospital discharge in the primary propensity-matched analysis. This association was consistently demonstrable with several sensitivity and subgroup analyses.

Tracheal intubation is the current standard of care in the setting of in-hospital cardiac arrest and is commonly prioritized and accomplished during in-hospital pediatric cardiac arrest. The data confirmed that this is the most common approach, as 68% of children who were not already intubated or receiving assisted ventilation at the time of CPR were intubated during CPR. Although recent guidelines have deemphasized the focus on tracheal intubation during cardiac arrest, especially in the adult population, the optimal approach to airway management during cardiac arrest in general and particularly pediatric in-hospital cardiac arrest remains unknown.^{7,28} Underlying pulmonary or airway disease and a presumed respiratory etiology are common in pediatric in-hospital cardiac arrest,^{2,3,29} providing the rationale for invasive airway management. However, tracheal intubation during CPR also carries potential complications and risk of adverse events.³⁰ Intubation may result in interruptions of chest compressions³¹ or in delay of other interventions such as defibrillation, achieving vascular access, or epinephrine administration, which could influence outcomes.^{15,32} In addition, the resuscitation leader may also be the most skilled operator, so the choice to intubate during CPR may prevent the resuscitation leader from attending to non-airway-related resuscitation efforts. Another potential detrimental effect of tracheal intubation is hyperventilation and increased intrathoracic pressure, which has been found to be common in cardiac arrest and can result in worse outcomes in animal models. ^{33–35} Importantly, a malpositioned tracheal tube could result in poor oxygenation and ventilation, particularly if not recognized rapidly.

Existing literature on tracheal intubation during pediatric in-hospital cardiac arrest is scarce. ³⁶ Gupta et al³⁶ conducted a retrospective review of pediatric intensive care unit cardiac arrests at a single center. Of 391 patients, 194 (50%) were not intubated or receiving invasive mechanical ventilation prior to the arrest. In this subgroup, the authors found no association between the timing of invasive airway placement (tracheal intubation) and survival or favorable neurologic outcome, but the power of their study to find a difference was limited. Their findings are consistent with those reported for adult in-hospital cardiac arrests³⁷ and consistent with those reported here. However, the analytic methodologies, multicenter data, and size of the database differ substantially. In the current analysis, the association between the timing of intubation and the timing of intubation and matching to the model. This interaction was nonsignificant, indicating that the association between intubation during CPR and outcomes with the arrest.

In the out-of-hospital setting, a quasi-randomized trial including 591 children in cardiac arrest found no difference in outcomes between patients in cardiac arrest receiving intubation compared with bag-valve-mask ventilation (24/301 vs 24/290 survivors).⁶ In a large observational study of out-of-hospital cardiac arrest in adults, Hasegawa et al³⁸ found that tracheal intubation as compared with bag-valve-mask ventilation was associated with decreased odds of survival and decreased favorable neurologic outcome. Given the substantial differences between in-hospital and out-of-hospital cardiac arrest³ and between pediatric and adult cardiac arrest,¹¹ any direct comparison to the current study is challenging.

In this study of in-hospital pediatric cardiac arrest, time-dependent propensity score matching was used, which allows for adjustment for covariates, including time-dependent covariates, and adjustment for the time-dependent risk of being intubated during CPRsomething not possible using traditional regression.^{16,21} As described in detail in the Methods section, this approach is superior to traditional regression or propensity score matching in situations in which a time-dependent intervention is assessed. Furthermore, the practical relevance of using this approach has previously been shown in relation to 2 highimpact observational studies regarding epinephrine administration in out-of-hospital cardiac arrest. The 2 studies used the same out-of-hospital database, but one used time-dependent propensity score matching,¹⁷ whereas the other used traditional propensity score matching, ³⁹ ultimately yielding differing results. One potential reason for these inconsistent results is that an intracardiac arrest intervention (eg, epinephrine or intubation) can be a function of the duration of the cardiac arrest (ie, patients are more likely to receive the intervention if the cardiac arrest is prolonged), thereby confounding the results. Moreover, the intervention might influence the duration of the event (ie, a patient being intubated might have ROSC because of the intubation), and adjusting for duration of the arrest is therefore difficult when using traditional propensity score matching or regression. The use of time-dependent propensity matching addresses these issues and allows for slightly different interpretation of the results. Specifically, the effect estimates (risk ratios) should be interpreted as the risk of the outcome in those being intubated during CPR at a given minute (from 0 to 15) relative to the risk of the outcome in those not being intubated during CPR at that given minute, which include those being intubated at later time points. This interpretation is more clinically

relevant, because a clinician might consider intubating a patient at any given minute with future events being unknown.

The result of the analysis using traditional propensity score matching indicated a stronger association between intubation and poor outcome. However, as discussed above, this analysis should be interpreted with caution, as it did not take into account the timing of intubation.

Although an association between intubation and decreased survival to hospital discharge was found, no significant association between intubation and neurologic outcome was seen in the main time-dependent propensity-matched multivariable analysis. There are a number of possible explanations for this finding. First, the current data set had a relatively high frequency of missing data on neurologic outcome (18%, corresponding to 36% of survivors), which limits the power and interpretability of the findings. Second, neurologic outcome was based on a nongranular scoring system,⁴⁰ for which there might be significant variability in reporting, biasing the results toward the null hypothesis. These potential explanations are supported by the findings of similar point estimates for the risk ratios for survival and favorable neurologic outcome but wider confidence intervals for the latter, resulting in nonsignificant findings. These concerns prompted this study, a priori, to use survival–an objective and almost uniformly reported outcome–as the primary outcome measurement.

The results of the current study should be interpreted in the context of the study design and several limitations. First, this was an observational study, and unmeasured confounders could have influenced the results. The registry does not have data on the specific indication for intubation, data on clinician experience or background, quantitative data on the quality of cardiopulmonary resuscitation, or data on the effectiveness of noninvasive oxygenation and ventilation prior to intubation during CPR. However, the sensitivity analysis assessing the potential influence of an unmeasured confounder indicated that any unmeasured confounder would be unlikely to mask a positive association between intubation and survival. The lack of data on clinician experience and background also limits the ability to assess whether intubation could be beneficial if provided by more experienced clinicians and the ability to adjust for clustering by clinician. Second, intubation attempts are not recorded in the registry, and unsuccessful intubation attempts are classified as "not intubated" and were therefore included in that group in the analyses. This is in contrast to the intent-to-treat principle most often used in randomized clinical trials and could have biased the results, especially if failed intubation attempts during CPR led to poor outcomes. However, this likely would have biased the results toward the null. Third, the GWTG-R registry has very limited information on post-ROSC care, and we were therefore unable to describe this aspect. Fourth, the study may have been underpowered to detect some clinically relevant associations, especially for neurologic outcomes and interactions.

Conclusions

Among pediatric patients with in-hospital cardiac arrest, tracheal intubation during cardiac arrest compared with no intubation was associated with decreased survival to hospital discharge. Although the study design does not eliminate the potential for confounding, these

findings do not support the current emphasis on early tracheal intubation for pediatric inhospital cardiac arrest.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Key Points

Question

Is tracheal intubation during pediatric in-hospital cardiac arrest associated with a difference in survival?

Findings

In this observational study including 2270 matched patients, tracheal intubation during pediatric cardiac arrest was significantly associated with decreased survival to hospital discharge (36% for those intubated compared with 41% for those not intubated).

Meaning

Although the study design does not eliminate the potential for confounding, these findings do not support the current emphasis on early tracheal intubation for pediatric inhospital cardiac arrest.

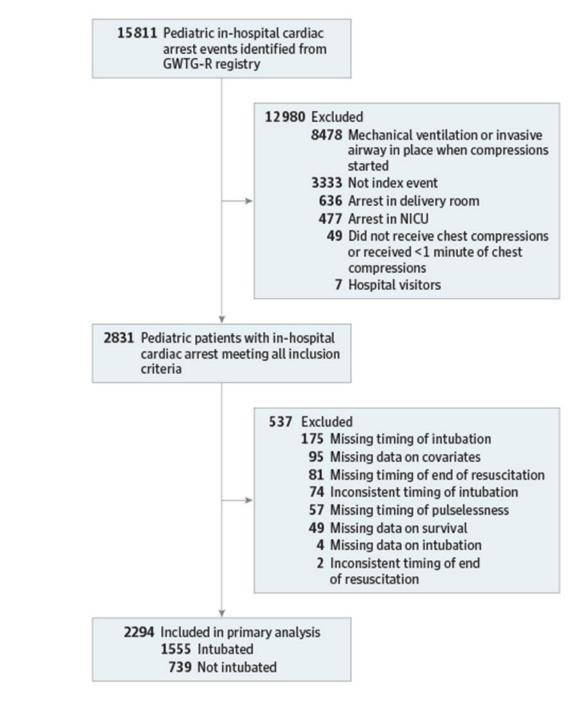


Figure 1.

Patient Selection for Study of Tracheal Intubation During Pediatric Cardiac Arrest GWTG-R, Get With The Guidelines-Resuscitation; IHCA indicates in-hospital cardiac arrest; NICU, neonatal intensive care unit.

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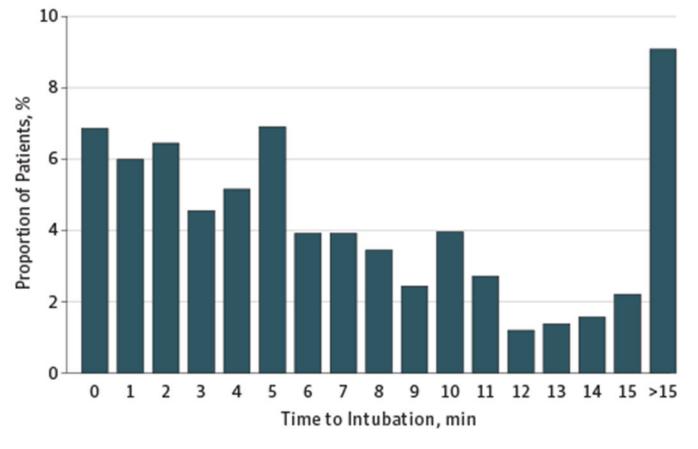


Figure 2.

Distribution of Time to Intubation

Distribution of time from start of chest compression to documented successful intubation during cardiopulmonary resuscitation (CPR) in those intubated. Of the 1555 patients (68%) intubated during the CPR event, the median time to intubation from start of chest compressions was 5 minutes (25th-75th percentiles, 2–11). Data on time to intubation were recorded in whole minutes. Time to intubation of 0 minutes indicates that intubation was performed in the same whole minute as the start of chest compressions.

2.0

	Not Intubated	Intubated			
	Patients With Events, No./Total (%)	Patients With Events, No./Total (%)	Risk Ratio (95% CI)	Favors No Favor Intubation Intub	
Survival					
Unadjusted analysis	495/739 (67)	667/1555 (43)	0.64 (0.59-0.69)	⊢ ∎-	
Main adjusted analysis	460/1135 (41)	411/1135 (36)	0.89 (0.81-0.99)	⊢ ∎(
Sensitivity analyses					
Imputation of missing data	580/1376 (42)	513/1376 (37)	0.88 (0.79-0.99)	⊢ ∎{	
Excluding patients who received CPB	426/1058 (40)	382/1058 (36)	0.90 (0.81-0.99)	H	
Excluding patients who received <2 min of chest compressions	343/881 (39)	307/881 (35)	0.90 (0.80-1.00)	⊢ ∎-1	
Subgroup analyses					
Patients who were pulseless at any time	288/853 (34)	254/853 (30)	0.88 (0.77-1.01)	⊢	
Patients with pulse present at initiation of CPR	193/325 (59)	166/325 (51)	0.86 (0.75-0.98)	├∎	
ROSC					
Unadjusted analysis	636/739 (86)	1130/1555 (73)	0.84 (0.81-0.88)	i= -i	
Main adjusted analysis	771/1135 (68)	770/1135 (68)	1.00 (0.95-1.06)	F∎H	
Sensitivity analyses					
Imputation of missing data	952/1376 (69)	944/1376 (69)	0.99 (0.93-1.05)	⊢	
Excluding patients who received CPB	690/1058 (65)	698/1058 (66)	1.01 (0.95-1.08)	H e -I	
Excluding patients who received <2 min of chest compressions	590/881 (67)	592/881 (67)	1.00 (0.94-1.07)	⊦∎⊣	
Subgroup analyses					
Patients who were pulseless at any time	530/853 (62)	518/853 (61)	0.98 (0.91-1.05)	⊢■⊣	
Patients with pulse present at initiation of CPR	281/325 (86)	282/325 (87)	1.00 (0.95-1.06)	F∎H	
Favorable neurologic outcome					
Unadjusted analysis	244/563 (43)	313/1318 (24)	0.55 (0.48-0.63)	⊢ ∎	
Main adjusted analysis	211/983 (21)	185/987 (19)	0.87 (0.75-1.02)	⊢_ ∎]	
Sensitivity analyses					
Imputation of missing data	351/1376 (26)	298/1376 (22)	0.85 (0.73-0.98)	⊢_ ∎	
Excluding patients who received CPB	178/899 (20)	173/923 (19)	0.95 (0.82-1.09)	⊢	
Excluding patients who received <2 min of chest compressions	154/762 (20)	139/769 (18)	0.89 (0.73-1.10)	⊢_ ∎1	
Subgroup analyses					
Patients who were pulseless at any time	132/758 (17)	111/768 (14)	0.83 (0.68-1.02)	⊢ ∎ <u></u>	
Patients with pulse present at initiation of CPR	107/254 (42)	79/260 (30)	0.72 (0.58-0.90)	⊢	
				0.4 1.0 Risk Ratio (95% CI)	

Figure 3.

Main and Sensitivity or Secondary Analyses According to Outcome

Unadjusted and adjusted analyses as well as sensitivity analyses and the subgroups of pulseless patients and patients with a pulse at the beginning of the event for the primary outcome survival to hospital discharge and the secondary outcomes return of spontaneous circulation (ROSC) and favorable neurologic outcome at hospital discharge. ROSC was defined as no further need for chest compressions that was sustained for greater than 20 minutes. Neurologic outcome was determined using the pediatric cerebral performance category (PCPC) score, for which a PCPC score of 1 indicates no neurologic deficit; 2, mild cerebral disability; 3, moderate cerebral disability; 4, severe cerebral disability; 5, coma or vegetative state; and 6, brain death. A PCPC score of 1 and 2 was considered a favorable neurologic outcome. The "main adjusted" analysis refers to the multivariable time-dependent propensity

score-matched analysis. CPB indicates cardiopulmonary bypass; CPR, cardiopulmonary resuscitation.

Table 1.

Characteristics of the Study Population

	No. (%)				
Characteristic	No Intubation (n = 739)	Intubation (n = 1555)	P Value		
Demographics					
Sex					
Female	328 (44)	658 (42)	25		
Male	411 (56)	897 (58)	.35		
Age group					
Neonate (<1 mo)	194 (26)	414 (27)	_		
Infant (1 mo to <1 y)	226 (31)	434 (28)			
Child (1 y to 12 y)	226 (31)	485 (31)	.50		
Adolescent (>12 y)	93 (13)	222 (14)	•		
Illness category					
Medical			_		
Cardiac	155 (21)	301 (19)	•		
Noncardiac	326 (44)	690 (44)	•		
Surgical			• <.001		
Cardiac	98 (13)	258 (17)	•		
Noncardiac	111 (15)	157 (10)	•		
Newborn ^a	49 (7)	149 (10)	•		
Preexisting conditions ^b					
Heart failure prior to this admission	53 (7)	138 (9)	.17		
Heart failure this admission	59 (8)	151 (10)	.18		
Hypotension	123 (17)	246 (16)	.62		
Respiratory insufficiency	326 (44)	685 (44)	.98		
Hepatic insufficiency	23 (3)	45 (3)	.77		
Renal insufficiency	45 (6)	125 (8)	.10		
Metabolic or electrolyte abnormality	70 (9)	165 (11)	.40		
Acute nonstroke CNS event	36 (5)	70 (5)	.69		
Baseline depression in CNS function	102 (14)	215 (14)	.99		
Metastatic or hematologic malignancy	27 (4)	75 (5)	.20		
Pneumonia	48 (7)	107 (7)	.73		
Septicemia	63 (9)	158 (10)	.21		
Location and Time of Arrest					
Location					
Emergency department	99 (13)	328 (21)			
Intensive care unit	363 (49)	657 (42)	• <.001		
Floor			•		

	No. (%)			
Characteristic	No Intubation (n = 739)	Intubation (n = 1555)	P Valu	
Without telemetry	122 (17)	261 (17)	_	
With telemetry/step-down unit	39 (5)	79 (5)	-	
Other ^C	116 (16)	230 (15)		
Time of week				
Weekend ^d	206 (28)	469 (30)	26	
Weekday	533 (72)	1086 (70)	.26	
Time of day				
Nighttime ^e	226 (31)	436 (28)	21	
Daytime	513 (69)	1119 (72)	.21	
Year of arrest			·	
2000–2002	46 (6)	150 (10)	_	
2003–2004	72 (10)	203 (13)		
2005–2006	129 (17)	229 (15)		
2007–2008	110 (15)	247 (16)	.005	
2009–2010	122 (17)	257 (17)	.005	
2011–2012	135 (18)	253 (16)		
2013–2014	125 (17)	216 (14)	•	
Arrest Characteristics				
Witnessed				
Yes	664 (90)	1349 (87)	02	
No	75 (10)	206 (13)	.03	
Monitored				
Yes	606 (82)	1128 (73)	.001	
No	133 (18)	427 (27)	• <.001	
Pulseless				
At the beginning of the event	377 (51)	982 (63)	_	
During the event	41 (6)	182 (12)	<.001	
Never	321 (43)	391 (25)		
Initial pulseless rhythm ^f				
Nonshockable	278 (67)	876 (75)		
Shockable	61 (15)	106 (9)	<.001	
Unknown	79 (19)	182 (16)		
Hospital Characteristics				
Type of hospital				
Primarily children	317 (43)	607 (39)	00	
Primarily adult	422 (57)	948 (61)	.08	
Teaching status				

	No. (%)	No. (%)		
Characteristic	No Intubation (n = 739)	Intubation (n = 1555)	<i>P</i> Value	
Major	518 (70)	933 (60)		
Minor	157 (21)	424 (27)	<.001	
Nonteaching	64 (9)	198 (13)	-	

Abbreviation: CNS, central nervous system.

^aDefined as being born on the current admission.

^bSee eTable 1 in the Supplement for definitions.

^CIncluding ambulatory or outpatient clinics, diagnostic or interventional areas, operating room, postanesthesia recovery room, rehabilitation units, same-day surgical areas, and "other."

^d Friday 11 PM to Monday 7 AM.

^e11:00 PM to 6:59 AM.

f Only including those patients who lost pulse at any time (n = 418 not intubated and n = 1164 intubated).

Table 2.

Characteristics of the Study Population in the Matched Cohort

	No. (%)	P Value		
Characteristic	No Intubation (n = 1135) Intubation (n = 1135)		Standardized Differenc	
Demographics				
Sex				
Female	498 (44)	481 (42)	1.5	-0.030
Male	637 (56)	654 (58)	.46	0.030
Age group				
Neonate (<1 mo)	308 (27)	290 (26)		-0.03
Infant (1 mo to <1 y)	294 (26)	302 (27)		0.01
Child (1 y to 12 y)	355 (31)	364 (32)	.53	0.01
Adolescent (>12 y)	178 (16)	179 (16)	•	0.002
Illness category				
Medical				
Cardiac	228 (20)	230 (20)	•	0.00
Noncardiac	505 (44)	495 (44)	•	-0.01
Surgical			.66	
Cardiac	176 (16)	198 (17)	•	0.05
Noncardiac	114 (10)	113 (10)		-0.00
Newborn ^a	112 (10)	99 (9)	•	-0.04
Preexisting conditions ^b				
Heart failure prior to this admission	101 (9)	111 (10)	.47	0.02
Heart failure this admission	108 (10)	118 (10)	.47	0.03
Hypotension	188 (17)	187 (16)	.95	-0.00
Respiratory insufficiency	507 (45)	476 (42)	.18	-0.05
Hepatic insufficiency	36 (3)	34 (3)	.81	-0.01
Renal insufficiency	73 (6)	97 (9)	.06	-0.02
Metabolic or electrolyte abnormality	132 (12)	125 (11)	.65	-0.02
Acute nonstroke CNS event	44 (4)	48 (4)	.68	0.01
Baseline depression in CNS function	176 (16)	165 (15)	.51	-0.02
Metastatic or hematologic malignancy	65 (6)	62 (5)	.78	-0.01
Pneumonia	75 (7)	78 (7)	.80	0.01
Septicemia	119 (10)	118 (10)	.94	-0.00
Location and Time of Arrest				
Location				
Emergency department	265 (23)	256 (23)		-0.01
Intensive care unit	468 (41)	469 (41)	.37	0.00
Floor			-	

	No. (%)			
Characteristic	No Intubation (n = 1135)	Intubation (n = 1135)	P Value	Standardized Difference
Without telemetry	187 (16)	187 (16)		0.00
With telemetry/step-down unit	65 (6)	57 (5)		-0.03
Other ^C	150 (13)	166 (15)		0.04
Time of week				
Weekend ^d	351 (31)	332 (29)	20	-0.03
Weekday	784 (69)	803 (71)	38	0.03
Time of day				
Nighttime ^e	315 (28)	310 (27)		-0.01
Daytime	820 (72)	825 (73)	80	0.01
Year of arrest				
2000–2002	97 (9)	123 (11)		0.07
2003–2004	161 (14)	163 (14)		0.00
2005–2006	174 (15)	179 (16)	•	0.01
2007–2008	178 (16)	177 (16)	.20	-0.00
2009–2010	207 (18)	177 (16)	•	-0.07
2011–2012	180 (16)	168 (15)	•	-0.02
2013–2014	138 (12)	148 (13)	•	0.02
Arrest Characteristics				
Witnessed				
Yes	972 (86)	975 (86)	0.5	0.00
No	163 (14)	160 (14)	.85	-0.00
Monitored				
Yes	805 (71)	807 (71)		0.00
No	330 (290)	328 (29)	.92	-0.00
Pulseless				
At the beginning of the event	788 (69)	774 (68)		-0.02
During the event	135 (12)	134 (12)	.64	-0.00
Never	212 (19)	227 (20)		0.03
Initial pulseless rhythm ^f				
Nonshockable	699 (76)	693 (76)		0.01
Shockable	84 (9)	88 (10)	.89	0.02
Unknown	140 (15)	127 (14)	•	-0.03
Hospital Characteristics				
Type of hospital				
Primarily children	447 (39)	420 (37)	20	-0.04
Primarily adult	688 (61)	715 (63)	.20	0.04

	No. (%)	No. (%)		
Characteristic	No Intubation (n = 1135)	Intubation (n = 1135)	P Value	Standardized Difference
Major	679 (60)	658 (58)		-0.038
Minor	302 (27)	322 (28)	.52	0.040
Nonteaching	154 (14)	155 (14)		0.003

Abbreviation: CNS, central nervous system.

^{*a*}Defined as being born on the current admission.

^bSee eTable 1 for definitions.

^CIncluding ambulatory or outpatient clinics, diagnostic or interventional areas, operating room, postanesthesia recovery room, rehabilitation units, same-day surgical areas, and "other."

^d Friday 11 PM to Monday 7 AM

^e11:00 PM to 6:59 AM

f Only including those patients who lost pulse at any time (n = 923 not intubated and n = 908 intubated).