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The association of genetic instrumental variables for lung function on coronary artery disease risk: A two-sample Mendelian randomization study

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Abstract

Background: Lung function, assessed by forced expiratory volume in 1 second (FEV₁) and forced vital capacity (FVC), is inversely associated with coronary artery disease (CAD) but these associations could be due to confounding or reversed causality. We conducted a two-sample Mendelian randomization study, using publicly available data from relevant genome wide association studies (GWAS), to examine the role of FEV₁ or FVC on CAD.

Methods: We used the most recent GWAS on lung function to extract genetic instruments related to FEV₁ and FVC (n=92,749). Data on the association between genetic instruments and CAD were obtained from CARDIoGRAMplusC4D 1000 Genomes based GWAS (60,801 CAD cases and 123,504 controls). We used inverse variance weighting with a multiplicative random effect to estimate the genetic instrumented association of FEV₁ and FVC on CAD. Sensitivity analyses included weighted median and MR-Egger methods.

Results: Each standard deviation (SD) greater FEV₁ was associated with a lower risk of CAD (Odds ratio (OR): 0.78 per SD; 95% CI: 0.62 to 0.98) with a similar magnitude for FVC on CAD risk (OR: 0.82 per SD; 95% CI: 0.64 to 1.06). Estimates for FEV₁ were similar when using MR-Egger method (OR: 0.80 per SD; 95% CI: 0.33 to 1.94) although the magnitude was smaller for weighted median method (OR: 0.93 per SD; 95% CI: 0.75 to 1.17). Estimates for FVC in the sensitivity analyses were attenuated (median) or changed direction (MR-Egger).

Conclusion: Our study suggested an inverse relation between FEV₁ and CAD but for FVC, evidence is less clear.

Keywords: Coronary artery disease; lung function; Mendelian randomization

1 **Introduction**

2 Reduced lung function is consistently associated with higher risk of cardiovascular diseases
3 (CVD), including coronary artery disease (CAD), in observational studies.¹⁻⁵ Possible
4 mechanisms linking reduced lung function and cardiovascular diseases include increased
5 inflammation as people with higher lung function tend to have lower levels of inflammatory
6 markers such as C-reactive protein and fibrinogen,⁶ inefficient cardio-toxic substance
7 removal, or a mismatch of ventilation/perfusion ratio.⁶ However, these inverse associations
8 could be a result of residual confounding or reverse causality. Smoking and height, which are
9 both strongly related to lung function and cardiovascular disease in directions that would
10 result in the direction of association between these two are particularly important sources of
11 confounding.^{2, 7} Furthermore, early CAD may result in reduced perfusion and reduced lung
12 function, i.e. reverse causation.⁸ If improving lung function can reduce CAD risk this is a
13 potentially modifiable risk factor that could have important public health benefit beyond that
14 of improving lung function.

15
16 Mendelian randomization studies may provide more credible evidence concerning the causal
17 role of lung function in CAD than that possible with conventional observational approaches.
18 For example, previous studies have shown that Mendelian randomization studies generate
19 comparable results with randomized controlled trials.⁹ A Mendelian randomization study to
20 test the effect of lung function on CAD would use genetic variants that are robustly
21 associated with lung function as instrumental variables for its causal effect. As these variants
22 are randomly allocated from parents to their offspring at conception they are less susceptible
23 to reverse causation than observational studies.¹⁰ Furthermore, the random allocation also
24 implies that the genetic variants were less likely to be affected by the many factors that
25 commonly confound observational studies, which is supported by empirical evidence.^{10, 11} To

1 our knowledge, no previous Mendelian randomization studies have explored the effect of
2 lung function on CAD. Therefore, we conducted a two-sample Mendelian randomization
3 study to clarify the relation of lung function (forced expiratory volume in 1 second (FEV₁)
4 and forced vital capacity (FVC)) and CAD risk using summary statistics from large genome
5 wide association studies (GWAS).^{12, 13} We hypothesized higher lung function (higher FEV₁
6 and FVC) would be associated with lower risk of CAD. Further, as FVC is particularly
7 related to height (more so than FEV₁), we hypothesized that inverse association between
8 FVC and CAD would be weaker than those between FEV₁ and CAD in our Mendelian
9 randomization study, if height is a key confounder of FVC with CAD.

10

11 **Method**

12 The data and methods are publicly available.

13

14 *Summary of Study design*

15 We used two-sample Mendelian randomization with summary data from publicly available
16 GWAS databases.^{12, 14} Firstly, we identified independent ($r^2 < 0.05$) genetic instruments
17 (single nucleotide polymorphisms (SNP)) strongly associated with FEV₁ or FVC ($p < 5 \times 10^{-8}$)
18 from GWAS of lung function and obtained the estimates between each of these SNPs and
19 lung function parameters (i.e. difference in mean FEV₁ or in FVC and their corresponding
20 standard errors) – sample one.¹⁵⁻¹⁷ Secondly, we obtained the corresponding genetic
21 associations for each of the lung function SNPs with CAD risk from
22 CARDIoGRAMplusC4D 1000 Genomes-based GWAS – sample two.¹⁸ We then used these
23 summary results to calculate the Mendelian randomization estimates of FEV₁ and FVC on
24 CAD risk, using various statistical methods as described in the statistical analysis section.^{19, 20}

25

1 *GWAS of lung function (FEV₁ and FVC – Sample 1)*

2 For this study we used data from the most recent GWAS (which included both FEV₁ and
3 FVC as outcomes),¹⁵ which built on previous GWAS,^{16, 17} and had a combined discovery and
4 replication sample size of 92,749 (discovery: 38,199; replication: 54,550). It was undertaken
5 in European origin participants, and was imputed based on the 1000 Genomes Project Phase
6 1 reference panel.¹⁵ Variants with imputation information <0.3, assessed using r2.hat (MACH
7 and minimac) or .info (IMPUTE2), were excluded. The GWAS used linear regression
8 adjusted for age, age², sex, height, and principal component for population structure,
9 separately for never and ever smokers. The residuals from FEV₁ and FVC were converted to
10 ranks and then transformed to normally distributed z scores, which were then used in an
11 additive genetic model, separately for never and ever smokers, with the results from these
12 two (never and ever smokers) strata then meta analyzed. In the GWAS, only SNPs previously
13 not identified were genotyped and analyzed in both discovery and replication stages whereas
14 previously known SNPs were only genotyped and analyzed in the discovery stage. As such,
15 we extracted information for 17 SNPs related to FEV₁ (10 from both stages and 7 from stage
16 1) and 11 SNPs related to FVC (5 from both stages and 6 from stage 1).¹⁵

17

18 *GWAS of CAD: CARDIoGRAMplusC4D 1000 Genomes (Outcome – Sample 2)*

19 We obtained summary data on the estimates of each lung function GWAS SNP with CAD
20 from CARDIoGRAMPLUSC4D (downloading these data from
21 www.CARDIOGRAMPLUSC4D.ORG).¹⁸ CARDIoGRAMplusC4D 1000 Genomes-based
22 GWAS is a meta-analysis of GWAS of CAD case-control studies of people of mainly
23 European (~91%), South Asian, and East Asian descent imputed using the 1000 Genomes
24 phase 1 v3 training set with 38 million variants. The study interrogated 9.4 million variants
25 and included 60,801 CAD cases and 123,504 controls.¹⁸ Case status was defined as having

1 CAD if they had received a diagnosis of myocardial infarction, acute coronary syndrome,
2 chronic stable angina, or coronary stenosis >50%. Diagnoses were ascertained in various
3 ways in different studies; extracted from medical records, clinical diagnosis at the time of
4 study, procedures (coronary angiography results or by-pass surgery), medications or
5 symptoms that indicate angina, or self-report of a doctor diagnosis, as described elsewhere.¹⁸
6 Covariate adjustment in the GWAS included study-specific covariates (e.g. age and sex) and
7 genomic control. The GWAS reported the corresponding log odds and standard error on
8 CAD for each SNP effect allele using an additive model.¹⁸

9

10 *Harmonization of the two samples*

11 To ensure correct allele harmonization, SNPs were orientated so that the effect allele was the
12 same for both the exposure and the outcome datasets. Palindromic SNPs were checked and
13 corrected using the effect allele frequency (EAF) reported in the exposure and outcome
14 GWAS to identify corresponding strands between the two GWAS. If EAF is >0.42 and ≤ 0.50 ,
15 a proxy non-palindromic SNP ($R^2=1$) would be used instead to avoid ambiguity in strand
16 direction since alleles of palindromic SNPs are the same on both strands.²¹ Linkage
17 disequilibrium of the SNPs were checked via 1000 Genomes project to avoid inclusion of
18 correlated SNPs. If two SNPs had an R^2 of >0.05, the SNP with the smaller p value was
19 retained and the other SNP was discarded.

20

21 *Statistical analysis*

22 *Main Mendelian randomization analysis*

23 Inverse variance weighting (IVW) with multiplicative random effects was used to give our
24 main Mendelian randomization effect estimate, which is a weighted regression of SNP-
25 outcome association (log odds of CAD per effect allele) on SNP-exposure association (mean

1 difference in FEV₁ or FVC per effect allele), with the intercept constrained to zero.²² One key
2 assumption for this method to produce valid estimate is that there is no other way SNPs could
3 affect the outcome (CAD here) than through the risk factor that the SNPs are instrumenting
4 for (FEV₁ or FVC). Specifically with MR it assumes no horizontal pleiotropy.¹² We also
5 calculated Wald ratio for each SNP ((i.e. SNP-outcome association divided by SNP-exposure
6 association), with the standard error calculated by dividing the standard error of the SNP-
7 outcome association by the effect size of the SNP-exposure association). We then meta-
8 analysed these ratios using a fixed effect method and compared the results to the
9 multiplicative random effects IVW. Both methods are supposed to estimate the same overall
10 effect, but standard errors will be larger when using multiplicative random effects IVW in the
11 presence of heterogeneity of effect across SNPs. We also estimated the I² statistics to explore
12 whether there was heterogeneity in effects between SNPs, which could be suggestive that one
13 or more SNPs are invalid instruments.

14

15 *Sensitivity analyses*

16 We undertook a number of sensitivity analyses to test the robustness of our main analyses.
17 Since the IVW multiplicative random effect models can give biased estimates if one or more
18 instruments are invalid due to horizontal pleiotropy,²⁰ we particularly focus on sensitivity
19 analyses that allow the assumption of no horizontal pleiotropy to be relaxed. These analyses
20 have different assumptions to the IVW (main analyses) method, and also to each other. In
21 comparison to the IVW method they have less statistical efficiency; this is particularly the
22 case with MR-Egger. Therefore, the aim of these sensitivity analyses was mostly to examine
23 if the magnitude and direction of effect estimates were consistent across methods.

24

25 *i. Leave one out analysis*

1 The IVW estimates were recalculated removing one SNP at a time aiming to explore whether
2 the overall estimate was primarily driven by any (outlying) SNP.

3

4 *ii. MR-Egger method*

5 Similar to the IVW, the MR-Egger method is a weighted regression of the SNP-outcome on
6 the SNP-exposure association, but unlike the IVW method, the intercept is not constrained to
7 be zero.¹⁹ This means that if there is unbalanced pleiotropy the Mendelian randomization
8 (slope of the regression line) effect estimate will differ from the IVW estimate, and the
9 intercept of the regression line will be non-zero. The MR-Egger method slope will give
10 unbiased estimate even if all instruments are invalid (i.e. due to the presence of horizontal
11 pleiotropy).¹⁹ However, it has an additional assumption known as the InSIDE (instrument
12 strength independent of direct effect), which requires no correlation between the strength of
13 the direct effect of instrument on outcome and the instrument strength (strength of association
14 between genetic variants for lung function and lung function). One possible scenario through
15 which the InSIDE assumption could be violated is if the pleiotropic effects of genetic
16 instruments are mediated via the same exposure-outcome confounder.²⁰ The MR-Egger
17 method is statistically inefficient and expected to have considerably wider confidence
18 intervals than our main analyses or other sensitivity analyses. As such we focus on whether
19 the odds ratio is consistent in terms of magnitude and direction with our main analysis results
20 and also whether the intercept value suggests evidence of pleiotropy.

21

22 *iii. Weighted median method*

23 The weighted median method is a further method that allows some relaxation of the
24 horizontal pleiotropy assumption. It weights each of the Wald ratio effects of each SNP by
25 the corresponding precision and then takes the median weighted ratio as the causal effect.²⁰ It

1 will provide an unbiased estimate so long as more than 50% of the weights were derived
2 from valid instruments.²⁰ This means that the estimate will be biased if one single SNP is
3 pleiotropic and contributes more than 50% of the weight for the overall effect, or if a group
4 of SNPs are pleiotropic and together they contribute more than 50% of the weight.

5

6 *iv. Difference in the imputation reference panel for lung function GWAS*

7 To examine whether the results were robust to the difference in the imputation reference
8 panel, we also extracted genome wide significant SNPs for lung function from the most
9 recent GWAS which used HapMap reference populations (CEU) and repeated the analyses.^{16,}
10 ¹⁷ For FEV₁, the SNPs were extracted from a GWAS with a sample size of 94,612 (discovery:
11 48,201; replication: 46,411) of mainly European origin.¹⁷ For FVC, the SNPs were extracted
12 from a GWAS with a sample size of 85,170 (discovery: 52,253; replication: 32,917) of
13 mainly European origins.¹⁶ Similar adjustments, stratification by smoking and analytical
14 methods, and imputation quality control threshold were used in these GWAS as in the one
15 used in our main analyses with imputation using the 1000 Genome Project Phase 1 reference
16 panel although the FVC GWAS used absolute FVC (mL) as the outcome.¹⁶ To improve
17 comparability with the other analyses, we rescaled the Mendelian randomization estimates of
18 FVC (HapMap) into per standard deviation (SD). The SD was derived by pooling the SDs
19 reported in the European cohorts in the FVC GWAS.

20

21 *v. Exclusion of potential pleiotropic SNPs*

22 Considering that height and smoking are key potential confounders, we checked whether any
23 of our selected FEV₁ or FVC genetic instruments were related to height or smoking (ever
24 smoking status and cigarettes smoked per day) using summary statistics from the relevant
25 GWAS consortia (GIANT and TAG, respectively).^{23, 24} Similarly, we explored whether any

1 of our FEV₁ or FVC genetic instruments were related to established cardiovascular risk
2 factors by looking for their associations using data from relevant GWAS: triglycerides and
3 LDL cholesterol (from GLGC);^{25, 26} body mass index (GIANT);²⁷ and systolic blood pressure
4 (UKB).^{28, 29} We considered a SNP as potentially pleiotropic if there was evidence of it being
5 associated with height, smoking or cardiovascular risk factors after applying a Bonferroni
6 correction to account for multiple testing ($p < 0.0004$), calculated by 0.05 (nominal p value)
7 /126 (number of comparisons)). We repeated the main analysis (IVW estimate) after
8 excluding the SNPs classified as pleiotropic.

9

10 *vi. Multivariable Mendelian randomization analysis*

11 Since some of the SNPs predicted both FEV₁ and FVC (i.e. with p value $< 5 \times 10^{-8}$ for both
12 parameters) based on the most recent GWAS,¹⁵ we also conducted a multivariable Mendelian
13 randomization analysis to adjust effects of FEV₁ by FVC, and vice-versa, in order to
14 disentangle the effects of each exposure on CAD risk.³⁰

15

16

17 *Exploring instrument strength*

18 Instrument strength was assessed by presenting the first stage (regression of FEV₁ and FVC
19 on genetic instruments) F-statistics and R² values. For both of these statistics higher values
20 related to stronger instruments.

21

22 We calculated the first stage F-statistics for the main analyses using the equation 1 in the
23 supplementary material. The GWAS only reported R² of the instruments (known and novel)
24 for the discovery stage.¹⁵ We directly extracted the R² for known loci as reported in the
25 GWAS discovery stage. However, the R² for novel loci in the discovery stage could be

1 susceptible to winner's curse.³¹ As such, we calculated the R^2 for novel loci using equation 2
2 in the supplementary material using parameters from each novel loci based on the combined
3 analyses (discovery + replication). Then, we summed up the R^2 and compute the F statistics
4 using the sample size as stated in the GWAS.

5

6 All analyses were performed using R Version 3.3.2 (R Development Core Team, Vienna,
7 Austria) with the R package (TwosampleMR). This study only used publicly available data
8 and hence no ethical approval was required to conduct this study.

9

10

11 **Results**

12 Amongst 17 SNPs for FEV₁ (1000 Genomes imputation), rs1985524 was palindromic and
13 was replaced by rs1989154 ($R^2=1$) to avoid ambiguity in the strand direction, rs11383346
14 was removed as this was an indel and was not available in CARDIoGRAMplusC4D, hence
15 16 SNPs were used. Amongst 11 SNPs for FEV₁ (HapMap imputation), rs17331332 and
16 rs3995090 were in linkage disequilibrium with other SNPs and were removed, and hence 9
17 SNPs remained in these analyses. Amongst 11 SNPs for FVC (1000 Genomes imputation),
18 rs1430193 was palindromic and was replaced by rs1430194 to avoid ambiguity in the strand
19 direction; rs11383346 was removed as this was an indel and was not available in
20 CARDIoGRAMplusC4D, thus 10 SNPs remained in these analyses. Amongst 6 SNPs for
21 FVC (HapMap imputation), rs1430193 was palindromic and was replaced by rs1430194.

22

23 Supplemental Tables 1-2 summarize the genetic associations as reported in the relevant
24 GWAS after allele harmonization. The first-stage F statistics and variance explained by the
25 genetic instruments and were 49.5 and 0.8% for FEV₁ and 38.4 and 0.4% for FVC.

1
2 Supplemental Table 3 shows the associations of all SNPs that were used as instrumental
3 variables with FEV₁ and FVC with smoking, height, BMI, LDLc, Triglycerides, and systolic
4 blood pressure. After correcting for multiple testing, 5 of the SNPs we were using as genetic
5 instruments were associated with height (rs1032296; rs134041; rs1430194; rs6923462; and
6 rs7155279), one SNP (rs2036527) associated with number of cigarettes smoked per day, one
7 with BMI (rs4237643) and 4 with systolic blood pressure (rs1989154; rs2571445; rs2863171;
8 rs7068966); none of our genetic instruments were associated with ever smoking, LDLc or
9 triglycerides. Based on the lung function GWAS, rs6681426, rs6441207, rs2274116,
10 rs10850377 were associated with both FEV₁ and FVC.¹⁵

11
12 FEV₁ was inversely associated with CAD risk in the inverse variance weighting with
13 multiplicative random effect model (Odds ratio: 0.78 per standard deviation (SD) increase,
14 95% Confidence Interval (CI): 0.62 to 0.98, I²=62%) (Table 1). Findings for the association
15 of FEV₁ with CAD risk were consistent across the range of sensitivity analyses. The leave
16 one out plot suggested the overall estimate was not driven by any single outlying SNP
17 (Figure 1). The estimate from MR-Egger, which is more robust to invalid variants was
18 consistent with the main estimate (OR: 0.80, 95%CI: 0.33 to 1.94), although had wider
19 confidence as expected because of the lower statistical power of this approach. In addition,
20 there was no evidence for directional pleiotropy based on the MR-Egger intercept value
21 (1000 Genomes imputation: -0.0009; p=0.96; HapMap imputation: -5.4 x 10⁻⁵; p=0.998)
22 (Table 1). The weighted median estimate was somewhat attenuated in comparison to the
23 main inverse variance weighted analyses, but consistent with those and the MR-Egger results
24 (OR: 0.93 per SD, 95% CI 0.75 to 1.17). Similar findings were observed when we used
25 estimates from the GWAS imputed using HapMap reference panel (Table 1).

1
2 There was also an inverse association of genetic instruments for FVC with CAD risk in the
3 inverse variance weighting with multiplicative random effect model using 1000 genome
4 imputation (Odds ratio: 0.82 per SD increase, 95% CI: 0.64 to 1.06, $I^2=35\%$) (Table 1). This
5 point estimate is similar to that for the association of FEV₁ with CAD, but with wider
6 confidence intervals. The estimate was weaker (closer to the null) using the weighted median
7 method (OR: 0.92, 95%CI: 0.69 to 1.23) and the MR-Egger method yielded a directionally
8 different estimate with wide confidence intervals (OR: 1.14, 95%CI: 0.36 to 3.61). There was
9 no evidence for directional pleiotropy based on the MR-Egger intercept value (1000
10 Genomes imputation: -0.011, $p=0.59$; HapMap imputation: -0.022, $p=0.48$) (Table 1). The
11 leave one out plot suggested the overall estimate was not driven by any single SNP, although
12 it was somewhat attenuated when rs4237643 or rs6681426 were removed (Figure 2). When
13 we used estimates from the GWAS with the HapMap imputation the main inverse variance
14 weighed estimates were similar to the 1000 genome imputation (OR: 0.85, 95%CI: 0.49 to
15 1.48), as were the results for weighted median analyses (attenuating towards the null) and
16 MR-Egger (showing a positive, rather than inverse association) (Table 1).

17
18 Repeating the analyses with the exclusion of SNPs related to height, smoking phenotypes,
19 body mass index, and systolic blood pressure did not notably change the results
20 (Supplemental Tables 4-5). The fixed effect Wald ratio meta-analysis gave similar results
21 (Supplemental Table 6). The multivariable Mendelian randomization analysis produced
22 directionally similar results compared to the MR-Egger regression for FEV₁ (OR: 0.82 per
23 SD increase, 95% CI: 0.53 to 1.28) and FVC (OR: 1.84 per SD increase. 95% CI: 0.61 to
24 5.50).

25

1 **Discussion**

2 To our knowledge, this is the first Mendelian randomization study examining the role of lung
3 function on CAD. Consistent with previous multivariable regression analyses, we found an
4 inverse relation between FEV₁ and CAD.^{1,2} In our main analyses the association of FVC on
5 CAD was similar to that of FEV₁. However, whilst the association of FEV₁ with CAD was
6 robust to numerous sensitivity analyses, those for FVC were less so, with MR-Egger and the
7 multivariable regression analyses (in which we adjusted for effect of SNPs on FEV₁)
8 suggesting a possible positive association of FVC on CAD, albeit with wide confidence
9 intervals including the null. Thus, our study does not corroborate previous multivariable
10 regression studies of FVC, which suggested an inverse association of FVC with CAD risk.⁶
11 Overall, our findings suggest that higher FEV₁ may reduce CAD risk, but that previous
12 inverse associations of FVC with CAD from multivariable regression analyses are possibly
13 explained by residual confounding, in particular due to height.

14

15 The inverse association of FEV₁ with CVD risk has been consistently seen in observational
16 studies.¹⁻⁴ However, smoking is a strong determinant of FEV₁ and CAD, and it is difficult to
17 fully adjust for it without very detailed information on life course frequency, amount and
18 intensity of smoking. Therefore, residual confounding by smoking might have explained
19 those previous results.^{2,7} Our Mendelian randomization study, which is less susceptible to
20 such confounding,¹⁰ suggests that greater FEV₁ may be related to reduced CAD risk. This
21 conclusion is particularly supported by the consistency of results across several different
22 Mendelian randomization approaches, including a meta-analysis of Wald ratio estimates for
23 each SNP, MR-Egger and weighted median methods, as well as when results from FEV₁
24 GWAS with a different imputation reference panel used to identify genetic instruments, and
25 when we removed SNPs with known associations with smoking, height and BMI (for which

1 we found some possible evidence of pleiotropy), and multivariable Mendelian randomization
2 analysis. These different approaches each have differing underlying assumptions, and hence
3 key sources of bias, and thus the triangulation of findings across them increases confidence
4 that FEV₁ is causally related to CAD risk.³² We did not find robust and consistent (across our
5 sensitivity analyses) evidence of an association of FVC on CAD. It is possible that the null
6 findings are due to lack of statistical power although some of these methods are less
7 statistically efficient and with FEV₁ confidence intervals were wide and included the null in
8 some sensitivity analyses. The difference between the sensitivity analyses of FEV₁ and CAD
9 and of FVC and CAD is that the former were all directionally consistent and had broadly
10 similar magnitudes, whereas with FVC two of the sensitivity analyses suggested positive,
11 rather than inverse associations, which suggests less robust Mendelian randomization
12 evidence for an inverse association.

13
14 That our study suggests possible differences in the association of FEV₁ and FVC on CAD is
15 perhaps not surprising given they reflect very different aspects of lung function, as reflected
16 by the differing genetic architecture observed in genome wide association studies.^{16, 17} FEV₁
17 represents severity of airway obstruction,³³ whereas FVC is a measure overall lung capacity.
18 These differences result in somewhat different confounding structures between the two
19 measurements, with smoking being more strongly associated with FEV₁ and height with FVC.
20 Thus, our results suggest that inability to fully account for confounding by height in previous
21 multivariable regression analyses might have resulted in spurious inverse associations of it
22 with CAD. Our results also suggest that airway obstruction, rather than lung size or capacity,
23 may increase CAD risk.

24

25 *Strengths and Limitations*

1 Key strengths of this study are the large sample size, with over 60,000 CAD cases, and the
2 use of Mendelian randomization to minimize the impact of confounding and reverse
3 causation. The limitations of our study in inferring causal relation relate to the extent to
4 which the key Mendelian randomization assumptions were satisfied – that the genetic
5 instruments are (i) robustly associated with risk factor of interest (here FEV₁ and FVC), (ii)
6 not associated with confounders of the risk factor-outcome association, and (iii) not related to
7 the outcome via any path other than through the risk factor (i.e. that there is no horizontal
8 pleiotropy). In two-sample Mendelian randomization as used here, weak instrument bias is
9 expected to bias findings towards the null (rather than towards the confounded result as
10 occurs in one-sample Mendelian randomization). However, this will depend on the extent of
11 overlap between the two samples, with substantial overlap resulting in similar bias to that
12 seen for one-sample Mendelian randomization.¹² Only one cohort (n=837, 0.5%) was used in
13 both the lung function and CAD GWAS samples, and the high F-statistic and R² for the
14 combined genetic instrumental variables, suggest that weak instrument bias is unlikely to
15 have impacted our results. Horizontal pleiotropy is unlikely given the similarity of results
16 across the different methods that we have used, and the intercept values from the MR-Egger
17 being close to zero. One limitation of using summary GWAS data is that we are unable to
18 explore the association of the genetic instruments with a broad range of observed potential
19 confounders, but we did find similar results when we excluded SNPs that are known to be
20 related to smoking and height, the two confounders we were mostly concerned about, one
21 SNP associated with BMI, and 4 SNPs associated with systolic blood pressure. We did not
22 find evidence for any of our genetic instrumental variables being robustly statistically
23 associated with LDLc and triglycerides, which are important risk factors for CAD. In these
24 analyses, the fact that the GWAS of both FEV₁ and FVC adjusted for height and smoking
25 could introduce spurious associations through a phenomenon known as collider bias.

1 However, results were consistent after removing SNPs that had a residual association with
2 height and smoking (Supplemental Table 4), suggesting that collider bias is unlikely to fully
3 account for the inverse association between FEV₁ and CAD. Nevertheless, research using UK
4 Biobank may help verify the findings since some of these biases can only be explored with
5 individual data.³⁴ Lastly, we were unable to explore potential reverse causation using a bi-
6 directional Mendelian randomization design because summary data across the whole genome
7 for the lung function GWAS are currently not available. Whilst there may be a bidirectional
8 association (i.e. a propensity to CAD causing reduced FEV₁ as well as vice versa), this would
9 not negate the association of FEV₁ in CAD that we have identified here.

10

11 *Conclusion*

12 In this first Mendelian randomization study of the relation of FEV₁ and FVC with CAD risk
13 that we are aware of, we find evidence for an inverse association of FEV₁ with CAD risk, but
14 less robust evidence for an association of FVC with CAD risk. Whilst it has been suggested
15 that increased systemic inflammation in those with poor lung function, inefficient cardio-
16 toxic substance removal, or a mismatch of ventilation/perfusion ratio, are potential ways in
17 which lower FEV₁ is related to increased CAD risk the exact mechanisms are unclear.⁶ As
18 our findings now provide more concrete evidence that there may be a causal effect that is not
19 explained by confounding due to smoking or height, further research to explore potential
20 mechanisms could identify novel targets for the prevention of CAD.

21

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17

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22

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3

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- 7

1 Table 1: Mendelian randomization estimates for lung function on coronary artery disease risk

*Lung function	Imputation	Inverse variance weighting with multiplicative random effects		MR-Egger			Weighted median	
		Odds ratio Per 1SD	95% CI	Odds ratio Per 1SD	95% CI	Intercept (p value)	Odds ratio Per 1SD	95% CI
FEV ₁ (16 SNPs)	1000 Genomes	0.78	0.62 to 0.98	0.80	0.33 to 1.94	-0.0009 (0.96)	0.93	0.75 to 1.17
FEV ₁ (9 SNPs)	HapMap	0.77	0.57 to 1.04	0.77	0.28 to 2.09	-5.4 x 10 ⁻⁵ (0.998)	0.80	0.62 to 1.04
FVC (10 SNPs)	1000 Genomes	0.82	0.64 to 1.06	1.14	0.36 to 3.61	-0.011 (0.59)	0.92	0.69 to 1.23
FVC (6 SNPs)	HapMap	0.85	0.49 to 1.48	2.37	0.17 to 33.5	-0.022 (0.48)	0.92	0.53 to 1.57

2 *FEV₁: Forced expiratory volume in 1 second; FVC: Forced vital capacity.

3

1 **Figure legends**

2

3 Figure 1: Leave one out plot for the association of forced expiratory volume in one second

4 (FEV₁), per standard deviation, and coronary artery disease risk using Mendelian

5 randomization (Single nucleotide polymorphisms with 1000 Genomes imputation)

6

7 Figure 2: Leave one out plot for the association of forced vital capacity (FVC), per standard

8 deviation, and coronary artery disease risk using Mendelian randomization (Single nucleotide

9 polymorphisms with 1000 Genomes imputation)