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For numbered affiliations see end of article. **Correspondence to:** Dr Whanhee Lee, School of the Environment, Yale University, New Haven, CT 06511, USA; whanhee.lee@yale.edu

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Received: 1 September 2021 Accepted: 21 March 2022 Associations between long term air pollution exposure and first hospital admission for kidney and total urinary system diseases in the US Medicare population: nationwide longitudinal cohort study

Whanhee Lee ⁽¹⁾, ^{1,2,3}, Xiao Wu,⁴ Seulkee Heo,¹ Kelvin C Fong,¹ Ji-Young Son,¹ M Benjamin Sabath,⁵ Danielle Braun,^{6,7} Jae Yoon Park,^{8,9} Yong Chul Kim,¹⁰ Jung Pyo Lee ⁽¹⁾, ¹¹ Joel Schwartz,¹² Ho Kim,^{13,14} Francesca Dominici,⁴ Michelle Bell¹

ABSTRACT

OBJECTIVE To estimate the associations between long term exposure to air pollution and the first hospital admission related to kidney and total urinary system diseases.

DESIGN Nationwide longitudinal cohort study. **SETTING** Data were collected from the Medicare fee-for-service for beneficiaries living in 34 849 zip codes across the continental United States from 2000 to 2016. Exposure variables were annual averages of traffic related pollutants (fine particles ($PM_{2.5}$) and nitrogen dioxide (NO_2)) that were assigned according to the zip code of residence of each beneficiary with the use of validated and published hybrid ensemble prediction models. **PARTICIPANTS** All beneficiaries aged 65 years or older who were enrolled in Medicare part A fee-forservice (n=61097767).

PRIMARY AND SECONDARY OUTCOME

MEASURES First hospital admission with diagnosis codes for total kidney and urinary system disease or chronic kidney disease (CKD), analyzed separately. **RESULTS** The average annual concentrations of air pollution were $9.8 \,\mu\text{g/m}^3$ for PM_{2.5} and 18.9 ppb for NO₂. The total number of first admissions related to total kidney and urinary system disease and CKD were around 19.0 million and 5.9 million, respectively (2000-16). For total kidney and urinary

WHAT IS ALREADY KNOWN ON THIS TOPIC

- ⇒ A few studies have reported that long term exposure to air pollution is associated with an increased risk of incident chronic kidney disease (CKD)
- ⇒ The generalizability of these results is limited because of the nonrepresentative nature of the study populations

WHAT THIS STUDY ADDS

- ⇒ The findings of this large nationwide study suggest that long term exposures to fine particles (PM_{2.5}) and nitrogen dioxide (NO₂) were associated with an increased risk of first hospital admission related to diseases of the kidney and urinary system or CKD
- ⇒ These associations were found to persist at levels below the current annual National Ambient Air Quality Standards for PM₂, and NO₂

HOW THIS STUDY MIGHT AFFECT RESEARCH, PRACTICE, OR POLICY

⇒ The findings of this study suggest an adverse effect of air pollution on diseases of the kidney and urinary system, which could help establish public health interventions system disease, hazard ratios were 1.076 (95% confidence interval 1.071 to 1.081) for a 5 μ g/m³ increase in PM_{2.5} and 1.040 (1.036 to 1.043) for a 10 ppb increase in NO₂. For CKD, hazard ratios were 1.106 (1.097 to 1.115) for a 5 μ g/m³ increase in PM_{2.5} and 1.013 (1.008 to 1.019) for a 10 ppb increase in NO₂. These positive associations between PM_{2.5} and kidney outcomes persisted at concentrations below national health based air quality standards. **CONCLUSIONS** The findings suggest that higher annual air pollution levels were associated with increased risk of first hospital admission related to diseases of the kidney and urinary system or CKD in the Medicare population.

Introduction

Kidney disease has been recognized as a worldwide public health concern owing to the increased risk of early death, cardiovascular disease, and poor quality of life.¹⁻³ Globally, over 750 million people have kidney diseases,^{4 5} annual mortality attributable to kidney disease might be as high as 5 million,⁶ and its prevalence has increased during recent decades.⁷⁸ Among the types of kidney disease, chronic kidney disease (CKD) is an important public health priority worldwide.⁴⁷⁹ In 2016, CKD was the 13th highest cause of death on a global scale and projections suggest that it will become the world's fifth leading cause of death by 2040.¹⁰ In the United States, about 15% of adults (37 million) were estimated to have CKD in 2021 and the disease is more common in people aged 65 years or older (38.1%).¹¹ Kidney disease results in high medical costs and financial burden for Medicare. According to the 2020 US Renal Data System annual data report, the total Medicare fee-for-service (FFS) spending for beneficiaries with CKD who did not have end stage renal disease exceeded \$81bn in 2018 (£61.7bn, €72.7bn; 22.3% of Medicare FFS spending), and annual expenditure for beneficiaries aged 66 years or older was estimated at \$23691 per person.¹² The total Medicare FFS expenditures for beneficiaries with end stage renal disease was \$36.6bn in 2018 (7.2% of Medicare spending).

Recently, air pollution has been suggested as a potential risk factor for low glomerular filtration

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rate and kidney disease.¹³⁻¹⁶ Inhaled air pollution can cause oxidative stress and DNA damage to kidney tissue,¹⁷ and cardiovascular and endocrine toxicities related to air pollution, such as disturbances in the autonomic nervous and circulatory systems and pulmonary inflammatory disease, have also been suggested as biological mechanisms linking air pollution exposure and kidney disease.^{14 18 19} In spite of these plausible biological mechanisms, associations between air pollution and kidney disease have been studied less than those for cardiovascular and respiratory diseases.¹³ Several recent cohort studies investigated the association between air pollution exposure and the incidence of kidney disease.¹³ ¹⁴ ²⁰⁻²² Although these studies consistently reported the positive association between exposure to air pollution and the risk of kidney disease, these studies had limitations related to the type of participants (eg. US veterans who were mostly white men),¹³¹⁴ locations,²² and sample size (less than 200000)^{20 21}; therefore, the results were limited in providing generalizable estimates for the association between air pollution and kidney disease. Additionally, air pollution concentrations in the US have been reduced in the past few decades; therefore, it is crucial to evaluate whether the relation between air pollution and kidney disease persists at lower concentrations.²³

To resolve these gaps in knowledge, we performed a nationwide longitudinal cohort study of Medicare Part A FFS beneficiaries from 2000 to 2016, a population of over 61 million. with 410 million person years of follow-up (that is, an average 6.7 follow-up years per person). We investigated the association between long term exposure to air pollution and the first hospital admission related to total kidney and urinary system disease or CKD. A nationwide air pollution dataset integrating validated zip code level fine particulate matter (PM_{2,5}) and nitrogen dioxide (NO₂) prediction models was used. These pollutants are key components of pollution from traffic, although they have other sources such as energy production and industry. Based on the large sample size and geographical scope, the study aimed to provide generalizable and statistically reliable results on the association between traffic related air pollution and the first hospital admission related to total kidney and urinary system disease or CKD with state-of-the-science air pollution prediction models.

Methods

The findings are reported in accordance with the strengthening the reporting of observational studies in epidemiology (STROBE) guidelines (see online supplemental material A).

Study design and participants

Our longitudinal study cohort consisted of beneficiaries who were enrolled in Medicare part A FFS (age 65 years or older) in the US from 1 January 2000 to 31 December 2016. We used the Medicare inpatient hospital claims from the Medicare Provider and Analysis Review files that contain one record per hospital admission. People become eligible to enter Medicare at age 65 years. Follow-up for each beneficiary started on 1 January 2000 or 1 January of the year after entry into the cohort and continued until the first admission with diagnosis codes related to each kidney outcome (total kidney and urinary system disease or CKD, analyzed separately), death, or end of the study period (31 December 2016), whichever came first. For each beneficiary, we extracted the date of the first admission, age, sex, self-reported race, zip code of residence, and Medicaid eligibility (as a proxy for low socioeconomic status²⁴ in each follow-up vear). Additionally, if the residential address changed during the follow-up period, exposure assessment incorporated any address changes in the analysis. For each zip code, the mean population was around 7500 and the median land area is about $92 \, \text{km}^2$.

We used international classification of diseases (ICD) codes to identify diseases of the kidney and the urinary system (ICD-9: 580-599; ICD-10: N00-N39) and CKD (ICD-9: 585; ICD-10: N18) as the primary or secondary diagnosis during the study period. Analysis is based on each participant's first hospital admission related to a given cause. The kidney and bladder are linked by ureters and are closely related. For example, neurogenic bladder and bacterial infections in the bladder and urinary tract are major causes of kidney diseases.^{25 26} Additionally, kidney stones are an important risk factor for urinary diseases, such as hematuria and urinary tract infections.²⁷ This study did not subdivide kidney and urinary disease, and previous studies also applied the combined ICD codes for total kidney and urinary system disease (N00-N39) as total renal disease.²⁸⁻³⁴ Because kidney disease has causal associations with hypertension, diabetes, and cardiac diseases, and is also closely associated with various complications including heart disease, stroke, and other metabolic diseases,³⁵ we considered the primary and secondary diagnosis ICD codes in the main analysis. Therefore, the outcomes of this study should be interpreted as the first hospital admission related to total kidney and urinary system disease or CKD.

Air pollution data

We obtained daily concentration predictions of ambient levels of $PM_{2.5}$ and NO_2 at $1 \times 1 \text{ km}^2$ spatial resolution across the contiguous US from validated, published models.^{36 37} We used these data to calculate the daily mean concentration of each air pollutant for each zip code by averaging the predictions at the grid cells with centroid points inside the boundary of that zip code.^{23 24 38} For each calendar

		Participants in low exposure	Participants not in
	Participants in full cohort (n=61 097 767)	cohortt (n=7 022 642)	low exposure cohort‡ (n=54 075 125)
Age at entry, years			
65-74	46886857 (76.7)	5 399 807 (76.9)	41 487 050 (76.7)
75-84	10 4 3 4 1 5 0 (17.1)	1 188 202 (16.9)	9 2 4 5 9 4 8 (17.1)
<u>></u> 85	3776760 (6.2)	434633 (6.2)	3 3 4 2 1 2 7 (6.2)
Sex			
Men	27 413 250 (44.9)	3 258 875 (46.2)	24 154 385 (44.7)
Women	33684507 (55.1)	3763767 (53.8)	29920740 (55.3)
Race			
White	51 577 679 (84.4)	6610683 (94.1)	44966996 (83.2)
Black	5 372 559 (8.8)	122114 (1.7)	5 2 5 0 4 4 5 (9.7)
Other*	3 4 8 0 7 6 3 (5.7)	223421 (3.2)	3 2 5 7 3 4 2 (6.0)
Medicaid eligibility			
Not eligible	53 561 222 (87.7)	6 355 624 (90.5)	47 205 598 (87.3)
Eligible	7 536 545 (12.3)	667018 (9.5)	6869527 (12.7)
Air pollution concentration			
PM _{2.5} (μg/m ³)	9.8 (3.1)	7.2 (2.2)	10.2 (3.0)
NO ₂ (ppb)	19.0 (10.2)	9.9 (3.6)	20.3 (10.2)
Potential confounder (community level characteristics)			
Below poverty level (%)	10.5 (8.1)	9.3 (7.0)	10.7 (8.2)
Population density (people per km²)	3300.5 (8590.2)	266.3 (702.8)	3744.6 (9109.1)
Median home value (\$1000)	203.8 (167.6)	161.8 (127.2)	209.3 (171.9)
Black (%)	11.9 (18.4)	2.8 (7.1)	13.2 (19.1)
Hispanic (%)	7.2 (8.7)	4.6 (5.9)	7.5 (9.0)
Median annual household income (\$1000)	52.6 (23.1)	47.7 (16.0)	53.3 (23.9)
Owner occupied housing (%)	67.4 (17.8)	75.2 (11.6)	66.3 (18.2)
Below high school education (%)	27.0 (16.1)	23.5 (14.5)	27.6 (16.3)
Ever smoked (%)	46.0 (7.2)	48.8 (8.0)	45.6 (6.9)
Body mass index	27.5 (1.0)	27.6 (1.0)	27.5 (1.0)

Table 1 | Descriptive cohort characteristics for US Medicare beneficiaries from 2000 to 2016. Values are number (%) or mean (SD)

\$1.00=£0.76, €0.90. SD=standard deviation.

*Other included Asian, Hispanic, American Indian or Alaskan Native, and unknown. Data on race were obtained from Medicare beneficiary files.

†Fine particles (PM₂₅)<12 µg/m³ and nitrogen dioxide (NO₂)<20 ppb.

 $PM_{2} \ge 12 \mu g/m^3$ and NO₂ $\ge 20 \text{ ppb}$.

year, we assigned the annual average concentrations for each zip code to Medicare enrollees according to the zip code of residence and calendar year to calculate long term exposures to air pollutants.^{23 24 38} The coefficients of determination (cross validated R²) for annual PM_{2.5} and NO₂ were 0.89 and 0.84, respectively, across the study area.³⁶⁻³⁸ Detailed information on the air pollution data is given in online supplemental material B.

Statistical analysis

We applied a two pollutant Cox equivalent reparameterized Poisson approach to manage the computational challenges of the conventional Cox proportional hazard model (eg, insufficient memory size and lengthy computation time).²³ This Poisson model was mathematically identical to a time varying Cox proportional hazard model under an Anderson-Gill representation.^{23 39} Specifically, for each of two kidney outcomes (first hospital admission related to total kidney and urinary system disease or CKD), a stratified Poisson model with a log link function was fit to estimate the association between time varying annual air pollutant and the first hospital admissions related to total kidney and urinary system disease or CKD. The dependent variable was the number of first hospital admissions in each follow-up year, calendar year, and zip code within stratums specified by individual variables: age at study entry in two year categories (65-66, 67-68, ..., ≥85), sex, race (white, black, and other-Asian, Hispanic, American Indian, Alaskan Native, and unknown), and Medicaid eligibility. Data

Table 2 Admissions related to total kidney and urinary
system disease and chronic kidney disease (CKD), and
association between air pollution and first hospital
admission related to kidney diseases

	PM _{2.5}	NO ₂			
Total kidney and urinary system disease					
Total cohort					
No of admis- sions	19 195 342				
Total person years	410 303 502				
Median follow- up year	5				
Hazard ratio (95% Cl)	1.076 (1.071 to 1.081)	1.040 (1.036 to 1.043)			
Low pollution coho	ort				
No of admis- sions	5233054				
Total person years	113564232				
Median follow- up year	6				
Hazard ratio (95% Cl)	1.133 (1.121 to 1.145)	1.021 (1.009 to 1.033)			
Chronic kidney di	sease				
Total cohort					
No of admis- sions	5902129				
Total person years	454 497 761				
Median follow- up year	6				
Hazard ratio (95% Cl)	1.106 (1.097 to 1.115)	1.013 (1.008 to 1.019)			
Low pollution cohort					
No of admis- sions	1638976				
Total person years	126 479 286				
Median follow- up year	6				
Hazard ratio (95% Cl)	1.202 (1.165 to 1.240)	0.980 (0.920 to 1.040)			

The low pollution cohort includes subset of cohort with annual exposures $r_{12} \mu g/m^3$ for fine particles (PM₂) and 20 ppb for nitrogen dioxide (NO₂) over the study period. Hazard ratio: for each 5 $\mu g/m^3$ increase in PM₂₅ and 10 ppb increase in NO₂.

on race were obtained from Medicare beneficiary files and further divisions of race were not possible owing to the structure of the data. The corresponding total person time of beneficiaries within each stratum was used as the offset. Confidence intervals were calculated using m out of n bootstrap.⁴⁰ This study did not consider missing data imputation because the data were complete for the considered variables (eg, age, residential zip code, ICD code). To measure the associations between air pollution and kidney outcomes, we estimated a hazard ratio for each 5 µg/m³ increase in PM_{2,5} and 10 ppb increase in NO₂.

For potential confounding, we also adjusted for neighborhood level socioeconomic status indicators.^{23 24} These indicators were available at zip code or county level and included population density, proportion of the population below the federally

defined poverty level, median household income, racial composition, education level, smoking rate, average body mass index, and average air temperature. Additionally, to adjust potential residual confounding by spatial and temporal trends, we included indicator variables for region (Northeast, Southeast, Midwest, Southwest, and West) and calendar years. Detailed information about the indicators is given in online supplemental material C. Finally, to investigate the potential non-linear association between air pollutants and kidney outcomes, we replaced the linear term of air pollution in the main model with a B spine function with three equally distributed internal knots (at the 25th, 50th, and 75th percentiles of the air pollution concentrations) for each pollutant.⁴¹

Low pollution cohort

To estimate the associations between air pollution and kidney outcomes at low concentrations, we repeated the main analysis but only included a subset of the main cohort with annual exposures lower than $12 \mu g/m^3$ and 20 ppb for PM_{2.5} and NO₂, respectively, over the study period. The current health based National Ambient Air Quality Standards (NAAQs) are $12 \mu g/m^3$ for PM_{2.5} and 53 ppb for NO₂.

Subgroup analysis

To identify subpopulations with higher vulnerability, we evaluated potential effect modification by race, age group (65-74 and \geq 75 years), sex, and Medicaid eligibility (as a proxy for low socioeconomic status) by including additional interaction terms between these variables and air pollutants in the main analysis. We also conducted region stratified analyses to assess potential heterogeneities among regions, which could relate to differences in chemical compositions, environmental and climatic conditions, and populations.

Sensitivity analysis

We performed sensitivity analyses to examine whether our main results are robust to the selection of confounders, inclusion of participants with prevalent disease, exposure time window, and potential outcome misclassification. To exclude those with potentially prevalent disease, we repeated the analyses using data excluding beneficiaries who had their first hospital admission for the outcomes in the first two years of follow-up. We also applied one year lag period exposures as an alternative exposure window. Finally, we performed a sensitivity analysis restricting kidney outcomes to those with primary diagnosis codes only.

Patient and public involvement

Patients and the public were not involved in the design, or conduct, or reporting, or dissemination plans of this research. This study did not include



Figure 1 | Nationwide first hospital admissions for total kidney and urinary system disease (upper panel) and chronic kidney disease (lower panel) per 100 000 Medicare fee-for-service beneficiaries across the contiguous United States from 2000 to 2016

plans to recruit participants and only used preexisting Medicare part A datasets. All data used in this study were prerecorded and deidentified.

Results

The full cohort dataset included 61097767 beneficiaries living in 34849 zip codes. Table 1 presents descriptive information about the beneficiaries. There were 410.3 million person years of follow-up for total kidney and urinary system disease and 454.5 million for CKD (table 2). The median follow-up was five years for total kidney and urinary system disease and six years for CKD. For hospital admissions, 27.6% for total kidney and urinary system disease (5 291 896) and 0.98% for CKD (57875) were identified as the primary discharge diagnosis code. The total number of first admissions was around 19.0 million for total kidney and urinary system disease and 5.9 million for CKD. Figure 1 shows the geographical distribution of first hospital admissions for total kidney and urinary system disease, and CKD. The low pollution cohort included 113.5 million person years of follow-up for total kidney and urinary system disease and 126.5 million for CKD (table 2).

The average annual concentrations of air pollution over the study period were $9.8 \,\mu\text{g/m}^3$ for PM_{2.5} and 18.9 ppb for NO₂. The highest PM_{2.5} concentrations were in California and the eastern and southeastern US. The highest NO₂ concentrations were observed in metropolitan areas such as New York, Los Angeles, Denver, and Chicago (figure 2).

In the full cohort, long term exposures to air pollution were positively associated with total kidney and urinary system disease for both pollutants, with hazard ratios of 1.076 (95% confidence interval 1.071 to 1.081) for a 5 µg/m³ increase in PM_{2.5} and 1.040 (1.036 to 1.043) for a 10 ppb increase in NO₂ (table 2). In the low pollution cohort, the associations remained significant: 1.133 (1.121 to 1.145) for PM_{2.5} and 1.021 (1.009 to 1.033) for NO₂. CKD also showed positive associations with PM_{2.5} (1.106, 1.097 to 1.115) and NO₂ (1.013, 1.008 to 1.019) in the full cohort. In the low pollution cohort, the risk related to PM_{2.5} was higher in the low pollution cohort than in the full cohort with a hazard ratio of 1.202





(1.165 to 1.240) for the low pollution cohort compared with the full cohort, although there was no evidence of a statistically significant relation with NO_2 .

Figure 3 shows the concentration-response associations for each kidney outcome and each pollutant. Table 2 presents results from the linear model. For $PM_{2.5}$, the non-linear association between pollutant level and hazard ratio showed a consistent rise, with an increased risk at higher levels of $PM_{2.5}$ for both kidney outcomes in general, and the slopes for both curves were steeper at concentrations lower than $8\mu g/m^3$. NO₂ also showed a consistently increased risk with higher exposure for total kidney and urinary system disease, whereas the association with CKD showed a decreasing pattern with larger confidence intervals for NO₂ concentrations higher than around 20 ppb (approximately 60th percentile of the NO₃ distribution).

Table 3 shows the results from subgroup analyses. For both kidney outcomes, white participants had higher associations with $PM_{2.5}$ compared with the black population, and beneficiaries who were eligible for Medicaid generally had lower estimated associations with both pollutants than those who were not eligible for Medicaid. Women showed

lower air pollution related risks on kidney outcomes than men. Additionally, positive associations with air pollution were generally observed in all regions of the US for both kidney outcomes (online supplemental table S1), with the highest $PM_{2.5}$ and NO_2 association estimates observed in the Southwest and Midwest, respectively.

The findings from our sensitivity analysis were generally consistent with the main results. Our findings were robust to adjustment for other confounders, removal of people with prevalent disease, and use of a different exposure window (online supplemental table S2). For total kidney and urinary system disease, exclusion of those identified by secondary diagnostic codes did not change the main results, although all estimates slightly increased; however, CKD showed negative associations with all air pollutants when those identified by secondary diagnostic codes were excluded.

Discussion

Principal findings

This large nationwide prospective cohort study of over 61 million Medicare beneficiaries found that



Figure 3 | Concentration-response curves showing association between long term air pollution exposure and kidney diseases. Total kidney and urinary system disease (upper panels) and chronic kidney disease (lower panels). Dashed vertical lines represent 10th and 90th percentiles of each air pollution concentration and shaded areas indicate 95% confidence intervals. Reference exposure points: $0 \mu g/m^3$ for fine particles (PM_{2.5}) and 0 ppb for nitrogen dioxide (NO₂)

long term exposures to $PM_{2.5}$ and NO_2 were associated with an increased risk of first hospital admission related to diseases of the kidney and the urinary

Table 3 Association between air pollution and first hospital admission for total kidney and urinary system disease or chronic kidney disease for specific subpopulations					
	PM _{2.5}	NO ₂			
Total kidney and	urinary system disease				
Age at entry, years					
<u>≥</u> 75 <i>v</i> 65-74	0.983 (0.979 to 0.987)	1.012 (1.012 to 1.017			
Sex					
Women v men	0.985 (0.983 to 0.988)	0.987 (0.985 to 0.989)			
Race					
Black v white	0.976 (0.969 to 0.982)	1.014 (1.009 to 1.019			
Other v white	1.003 (0.991 to 1.015)	0.950 (0.942 to 0.958)			
Medicaid eligibility					
Eligible <i>v</i> not eligible	0.980 (0.974 to 0.986)	0.962 (0.957 to 0.966)			
Chronic kidney di	sease				
Age at entry, years					
<u>≥</u> 75 <i>v</i> 65-74	0.932 (0.927 to 0.938)	1.007 (1.004 to 1.010)			
Sex					
Women v men	0.935 (0.931 to 0.939)	0.955 (0.952 to 0.958)			
Race					
Black <i>v</i> white	0.988 (0.978 to 0.997)	1.018 (1.013 to 1.024)			
Other v white	0.993 (0.976 to 1.010)	1.012 (1.006 to 1.019			
Medicaid eligibility					
Eligible <i>v</i> not eligible	1.018 (1.011 to 1.026)	0.978 (0.974 to 0.982)			
Data are hazard ratio (95% confidence interval) for each 5 μ g/m ³ increase in fine particles (PM ₂) and 10 ppb increase in nitrogen dioxide (NO ₂).					

in fine particles (PM_{2,y}) and 10 ppb increase in nitrogen dioxide (NO₂). *Other included Asian, Hispanic, American Indian or Alaskan Native, and unknown. Data on race were obtained from Medicare beneficiary files. system or CKD. These associations persisted even at levels below the current annual NAAQS for $PM_{2.5}$ and NO_2 . Furthermore, the study found that white people and those not eligible to receive Medicaid are more vulnerable than the general population. These findings suggest that improving air quality can contribute to public health benefits overall and also provide evidence on the benefit of a more stringent air pollution standard for people who potentially have kidney disease.

Comparision with other studies

The findings of this study are consistent with previous research studying the association between long term exposure to air pollution and kidney diseases. Two recent US Veteran cohort studies (about 2-2.5 million cohort participants who were mostly white men) showed that exposures to higher concentrations of particulate matter no larger than $10 \mu g/m^3$ (PM₁₀), PM, , carbon monoxide, and NO, are associated with increased risk of incident CKD and end stage renal disease.^{13 14} Another cohort study conducted in four US counties (10997 participants) reported that higher annual average PM25 was associated with increased urinary albumin-creatinine ratio and higher risk of incident CKD.²² Cohort studies in China (71151 participants) and Taiwan (around 100000-160000 participants) found that exposure to long term air pollution was associated with an increased risk of CKD, end stage renal disease,^{20 21} and membranous nephropathy.⁴² Nevertheless, owing to the limited cohort participants and different study regions, the findings from these previous studies have limited generalizability. This study provides more generalizable estimates for the risks of air pollution relating to total kidney and urinary system disease and CKD for the US Medicare part A FFS population by using a cohort of almost 61 million beneficiaries across the contiguous US with high spatial resolution air pollution data.

Strengths and limitations

We observed a significant association between $PM_{2.5}$ and the first hospital admission related to diseases of the kidney and the urinary system at low concentrations of air pollution, suggesting that the health benefit per unit decrease in the concentration of $PM_{2.5}$ might be consistent across concentrations below the current NAAQS. This potential saturation effect (lower air pollution risk at high ranges of annual air pollution) can be explained by smaller risk changes in annual mean air pollution concentration in regions with higher baseline air pollution levels.⁴¹ Furthermore, from the non-linear concentration-response curves, we found no evidence of a threshold value (a concentration below which air pollution exposure does not affect kidney outcomes) for harmful pollution.

The large sample from the Medicare part A FFS cohort enabled us to estimate air pollution risks relating to kidney diseases among subpopulations.²⁴ However, the results should be carefully interpreted because of small sample sizes for some subpopulations, such as racial minority groups and people eligible for Medicaid, compared with others. Further research is needed that investigates the association between air pollution and kidney disease in potentially vulnerable subpopulations, including potential differences in diagnosis and access to healthcare, especially given the high number of people with undiagnosed disease. The estimated associations between air pollution and kidney outcomes were generally higher in white participants and those not eligible for Medicaid. This finding conflicts with results from previous US studies reporting higher air pollution mortality risk in non-white populations and those eligible for Medicaid,^{24 43} although it is consistent with some other study results. A few plausible explanations exist; one possibility is competing risks. A previous study in Taiwan reported that people with comorbidities showed a lower risk related to PM2, and CKD development.²⁰ Another study in Korea reported that patients with CKD and healthy lifestyles (normal weight, non-smokers, and non-drinkers) showed a higher air pollution mortality risk than patients with CKD and non-healthy lifestyles.⁴⁴ These findings suggest the possibility of competing risk between air pollution and biological or behavioral factors that affects the incidence of kidney disease; the results by race and Medicaid eligibility could be affected by the potential competing risks of underlying health conditions and other risk factors. Another possibility is that the longer life expectancy of high income and white populations could affect the results of this study. The life expectancy of those with high income is longer than that of low income populations,⁴⁵ and white populations also have a longer life expectancy than black populations.⁴⁶ Because the older population is more vulnerable to the incidence of kidney disease,¹¹⁴⁷ the relatively longer life expectancy of Medicaid non-eligible and white populations could be related to the higher air pollution related risks on kidney outcomes in this study. Finally, because low income and racial or ethnic minority populations generally have less access to medical facilities,⁴⁸ underdiagnosis could potentially affect the results. Subgroup results should be examined in future studies with more detailed clinical and socioeconomic indicators.

This study has several limitations. Firstly, the outcome of first hospital admission with diagnosis codes for kidney diseases has limitations when interpreted as the onset of kidney disease. According to the Centers for Disease Control and Prevention, most adults with CKD do not recognize they have CKD,¹¹ and hospital admission can occur at more advanced stages of the disease or for treating complications attributable to kidney disease. Moreover, kidney disease is generally diagnosed through laboratory tests (eg, the estimated glomerular filtration rate, creatinine, and the accumulation of end products of nitrogen metabolism),^{35 49–51} and previous studies have reported that the incidence of kidney disease was substantially underestimated when ICD codes were used to define kidney disease.^{52 53} Therefore, the hospital admission records do not fully represent disease incidence, and our cohort probably undercounted the total kidney disease incidence. Additionally, because the use of ICD codes limits the classification of individual disease and the small sample size of specific disease, we did not investigate specific kidney diseases except for CKD and instead used the total kidney and urinary system diseases based on earlier published studies.^{28–34} Therefore, we were not able to investigate kidney and urinary disease separately, although an association exists between some kidney and urinary diseases.^{30–33} Therefore, our study did not investigate the potential heterogeneity of biological mechanisms among specific kidney and urinary diseases related to air pollution.

Secondly, the potential competing events (eg, hypertension, diabetes, and cardiovascular diseases; not death) that might affect the first hospital admission related to kidney disease^{35 49 54} were not considered in this study because of limited data availability. Thirdly, patients with CKD defined in this study represent a subset of those with total kidney disease; these patients were not independent of the total number of patients with kidney and urinary system disease. Therefore, the apparent air pollution related risks for total kidney and urinary system disease could be due to confounding with CKD,⁵⁵ and bias could exist in the results of total kidney and urinary system disease resulting from estimating the effects on CKD. Furthermore, additional research is needed on the association between air pollution and CKD in relation to diagnosis, especially given the large number of participants with undetected disease. In this study CKD was identified as the primary diagnosis in less than 1%; most participants with CKD were identified through hospital admission for comorbidities.⁵⁴ Other areas of interest are the impacts of air pollution on recurrent events. This study investigated first hospital admissions only, but subsequent hospital admissions might also be of interest. Importantly, the association between air pollution and CKD was present when primary and secondary ICD codes were considered, but not when the primary code only was considered, which warrants further investigation.

Fourthly, our results only represent the Medicare part A FFS population that does not include all Medicare beneficiaries (the Medicare FFS population covers up to around 75% of the Medicare population).²³ We had no information on Medicare Health Maintenance Organization claims, therefore, we could not cover the entire Medicare population. Fifthly, although we adjusted for several neighborhood level indicators, Medicare claims do not include extensive individual level data on behavioral and socioeconomic risk factors, which could be crucial confounders. Additionally, owing to limitations of the study data, we were not able to consider the potential effect modifications by comorbidities and CKD stages. These unmeasured individual level confounders might help to elucidate the biological mechanisms underlying the observed association between air pollution and kidney disease, and they should be looked at in future studies. Finally, future research could also investigate complex air pollution mixtures. We focused on traffic related air pollutants of NO₂ and PM₂₅, although each of these pollutants also has non-traffic sources, and actual air pollution mixtures are complexes with other pollutants and different chemical structures of PM, ... We also used state-of-the-science air pollution prediction data, which are only available until 2016; therefore, future studies should investigate the association between air pollution and kidney disease with more recent data.

Conclusion

We found that long term exposures to higher levels of air pollution were associated with increased risk of a first hospital admission related to total kidney and urinary system disease or CKD, even at low pollution concentrations. Our findings suggest that improvement of air quality might have some benefits in reducing kidney and urinary system disease.

AUTHOR AFFILIATIONS

¹School of the Environment, Yale University, New Haven, CT, USA ²Department of Environmental Medicine, College of Medicine, Ewha Womans University, Seoul, Republic of Korea

³Institute of Ewha-SCL for Environmental Health (IESEH), Seoul, Republic of Korea

⁴Department of Biostatistics, Harvard University T H Chan School of Public Health, Boston, MA, USA $^{\rm 5}\mbox{Harvard}$ University Faculty of Arts and Sciences, Cambridge, MA, USA

⁶Harvard University T H Chan School of Public Health, Boston, MA, USA

⁷Department of Data Science, Dana-Farber Cancer Institute, Boston, MA, USA

⁸Department of Internal Medicine, Dongguk University Ilsan

Hospital, Goyang, Gyeonggi-do, Republic of Korea ⁹Department of Internal Medicine, Dongguk University College of

Medicine, Goyang, Gyeonggi-do, Republic of Korea ¹⁰Department of Internal Medicine, Seoul National University

Hospital, Jongno-gu, Seoul, Republic of Korea

¹¹Department of Internal Medicine, Seoul National University

Boramae Medical Center, Seoul, Republic of Korea ¹²Environmental Health, Harvard University T H Chan School of

Public Health, Boston, MA, USA

¹³Graduate School of Public Health, Seoul National University, Gwanak-gu, Seoul, Republic of Korea

¹⁴Institute for Sustainable Development, Graduate School of Public Health, Seoul, Republic of Korea

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ORCID iDs

Whanhee Lee http://orcid.org/0000-0001-5723-9061 Jung Pyo Lee http://orcid.org/0000-0002-4714-1260

REFERENCES

- Gansevoort RT, Correa-Rotter R, Hemmelgarn BR, et al. Chronic kidney disease and cardiovascular risk: epidemiology, mechanisms, and prevention. Lancet 2013;382:339–52. doi:10.1016/S0140-6736(13)60595-4
- 2 Luyckx VA, Al-Aly Z, Bello AK. Sustainable development goals relevant to kidney health: an update on progress. *Nature Reviews Nephrology*2020:1–18. doi:10.1038/s41581-021-00473-9
- 3 Perlman RL, Finkelstein FO, Liu L, et al. Quality of life in chronic kidney disease (CKD): a cross-sectional analysis in the renal research Institute-CKD study. Am J Kidney Dis 2005;45:658–66. doi:10.1053/j.ajkd.2004.12.021
- 4 Jager KJ, Kovesdy C, Langham R, et al. A single number for advocacy and communication—worldwide more than 850 million individuals have kidney diseases. Nephrology Dialysis Transplantation 2019;34:1803–5. doi:10.1093/ndt/gfz174
- 5 GBD 2015 DALYS and HALE Collaborators. Global, regional, and national disability-adjusted life-years (DALYS) for 315 diseases and injuries and healthy life expectancy (HALE), 1990-2015: a systematic analysis for the global burden of disease study 2015. Lancet 2016;388:1603-58. doi:10.1016/S0140-6736(16)31460-X
- 6 Luyckx VA, Tonelli M, Stanifer JW. The global burden of kidney disease and the sustainable development goals. *Bull World Health Organ* 2018;96:414–22. doi:10.2471/BLT.17.206441
- 7 Bikbov B, Purcell CA, Levey AS, et al. Global, regional, and national burden of chronic kidney disease, 1990–2017: a systematic analysis for the global burden of disease study 2017. The Lancet 2020;395:709–33. doi:10.1016/S0140-6736(20)30045-3
- 8 Lameire NH, Bagga A, Cruz D, et al. Acute kidney injury: an increasing global concern. Lancet 2013;382:170–9. doi:10.1016/ S0140-6736(13)60647-9
- 9 Levin A, Tonelli M, Bonventre J, et al. Global kidney health 2017 and beyond: a roadmap for closing gaps in care, research, and policy. Lancet 2017;390:1888–917. doi:10.1016/S0140-6736(17)30788-2
- 10 Foreman KJ, Marquez N, Dolgert A, et al. Forecasting life expectancy, years of life lost, and all-cause and cause-specific mortality for 250 causes of death: reference and alternative scenarios for 2016–40 for 195 countries and territories. *The Lancet* 2018;392:2052–90. doi:10.1016/S0140-6736(18)31694-5
- 11 Centers for Disease Control and Prevention. Chronic kidney disease in the United States, 2021. Atlanta, GA: US Department of Health and Human Services, Centers for Disease Control and Prevention, 2021.
- 12 United States Renal Data System. 2020 USRDS annual data report: epidemiology of kidney disease in the United States. Bethesda, MD: National Institutes of Health, National Institute of Diabetes and Digestive and Kidney Diseases.
- 13 Bowe B, Xie Y, Li T, et al. Associations of ambient coarse particulate matter, nitrogen dioxide, and carbon monoxide with the risk of kidney disease: a cohort study. Lancet Planet Health 2017;1:e267–76. doi:10.1016/S2542-5196(17)30117-1
- 14 Bowe B, Xie Y, Li T, *et al.* Particulate matter air pollution and the risk of incident CKD and progression to ESRD. *J Am Soc Nephrol* 2018;29:218–30. doi:10.1681/ASN.2017030253
- 15 Ye J-J, Wang S-S, Fang Y, et al. Ambient air pollution exposure and risk of chronic kidney disease: a systematic review of the literature and meta-analysis. Environ Res 2021;195:110867. doi:10.1016/j. envres.2021.110867
- 16 Xu X, Nie S, Ding H, et al. Environmental pollution and kidney diseases. Nat Rev Nephrol 2018;14:313–24. doi:10.1038/ nrneph.2018.11
- 17 Mehta AJ, Zanobetti A, Bind M-AC, *et al.* Long-Term exposure to ambient fine particulate matter and renal function in older men: the Veterans administration normative aging study. *Environ Health Perspect* 2016;124:1353–60. doi:10.1289/ehp.1510269
- 18 Chin MT. Basic mechanisms for adverse cardiovascular events associated with air pollution. *Heart* 2015;101:253–6. doi:10.1136/ heartjnl-2014-306379
- 19 Rajagopalan S, Brook RD. Air pollution and type 2 diabetes: mechanistic insights. *Diabetes* 2012;61:3037–45. doi:10.2337/db12-0190

- 20 Chan T-C, Zhang Z, Lin B-C, et al. Long-Term exposure to ambient fine particulate matter and chronic kidney disease: a cohort study. Environ Health Perspect 2018;126:107002. doi:10.1289/EHP3304
- 21 Lin S-Y, Ju S-W, Lin CL, et al. Air pollutants and subsequent risk of chronic kidney disease and end-stage renal disease: a populationbased cohort study. Environ Pollut 2020;261:114154. doi:10.1016/j. envpol.2020.114154
- 22 Blum MF, Surapaneni A, Stewart JD, *et al.* Particulate matter and albuminuria, glomerular filtration rate, and incident CKD. *Clin J Am Soc Nephrol* 2020;15:311–9. doi:10.2215/CJN.08350719
- 23 Shi L, Wu X, Danesh Yazdi M, Yazdi MD, et al. Long-term effects of PM₂₅ on neurological disorders in the American Medicare population: a longitudinal cohort study. *Lancet Planet Health* 2020;4:e557–65. doi:10.1016/S2542-5196(20)30227-8
- 24 Di Q, Wang Y, Zanobetti A, *et al.* Air pollution and mortality in the Medicare population. *N Engl J Med Overseas Ed* 2017;376:2513–22. doi:10.1056/NEJM0a1702747
- 25 Kemec Z, Tüzün C, Gürel A. Neurogenic bladder and acute kidney injury in leukodystrophy. *Cureus* 2020;12:e8707. doi:10.7759/ cureus.8707
- 26 Salo J, Ikáneimo R, Tapiainen T, *et al.* Childhood urinary tract infections as a cause of chronic kidney disease. *Pediatrics* 2011;128:840–7. doi:10.1542/peds.2010-3520
- 27 Alelign T, Petros B. Kidney stone disease: an update on current concepts. Adv Urol 2018;2018:1–12. doi:10.1155/2018/3068365
- 28 Hansen AL, Bi P, Ryan P, et al. The effect of heat waves on hospital admissions for renal disease in a temperate city of Australia. Int J Epidemiol 2008;37:1359–65. doi:10.1093/ije/dyn165
- 29 Kim E, Kim H, Kim YC, et al. Association between extreme temperature and kidney disease in South Korea, 2003–2013: stratified by sex and age groups. Sci Total Environ 2018;642:800–8. doi:10.1016/j.scitotenv.2018.06.055
- 30 Hansen A, Bi P, Ryan P. The effect of Heatwaves on hospital admissions for renal disease in Adelaide, South Australia. *Epidemiology* 2008;19:S105–6.
- 31 Borg M, Bi P, Nitschke M, *et al.* The impact of daily temperature on renal disease incidence: an ecological study. *Environmental Health* 2017;16:1–30. doi:10.1186/S12940-017-0331-4
- 32 Liu Y, Hoppe BO, Convertino M. Threshold evaluation of emergency risk communication for health risks related to hazardous ambient temperature. *Risk Anal* 2018;38:2208–21. doi:10.1111/risa.12998
- 33 Kovats RS, Hajat S, Wilkinson P. Contrasting patterns of mortality and hospital admissions during hot weather and heat waves in greater London, UK. Occup Environ Med 2004;61:893–8. doi:10.1136/ 0em.2003.012047
- 34 Ballotari P, Ranieri SC, Luberto F, *et al.* Sex differences in cardiovascular mortality in diabetics and nondiabetic subjects: a population-based study (Italy). *Int J Endocrinol* 2015;2015:1–10. doi:10.1155/2015/914057
- 35 Webster AC, Nagler EV, Morton RL, et al. Chronic kidney disease. The Lancet 2017;389:1238–52. doi:10.1016/S0140-6736(16)32064-5
- 36 Di Q, Kloog J, Koutrakis P, et al. Assessing PM2.5 exposures with high spatiotemporal resolution across the continental United States. Environ Sci Technol 2016;50:4712–21. doi:10.1021/acs.est.5b06121
- 37 Di Q, Amini H, Shi L, et al. Assessing NO concentration and model uncertainty with high spatiotemporal resolution across the contiguous United States using ensemble model averaging. Environ Sci Technol 2020;54:1372–84. doi:10.1021/acs.est.9b03358
- 38 Wei Y, Yazdi MD, Di Q, et al. Emulating causal dose-response relations between air pollutants and mortality in the Medicare population. Environmental Health 2021;20:1–10. doi:10.1186/ S12940-021-00742-X
- 39 Andersen PK, Gill RD. Cox's regression model for counting processes: a large sample study. Ann Stat 1982:1100–20. doi:10.1214/ aos/1176345976
- 40 Bickel PJ, Götze F, van Zwet WR. Resampling fewer than N observations: gains, losses, and remedies for losses. Selected works of Willem van Zwet: Springer, 2012: 267–97.
- 41 Liu C, Chen R, Sera F, *et al.* Ambient particulate air pollution and daily mortality in 652 cities. *N Engl J Med* 2019;381:705–15. doi:10.1056/NEJM0a1817364
- 42 Xu X, Wang G, Chen N, et al. Long-Term exposure to air pollution and increased risk of membranous nephropathy in China. J Am Soc Nephrol 2016;27:3739–46. doi:10.1681/ASN.2016010093
- 43 Pope CA, Lefler JS, Ezzati M, et al. Mortality risk and fine particulate air pollution in a large, representative cohort of U.S. adults. Environ Health Perspect 2019;127:077007. doi:10.1289/EHP4438
- 44 Jung J, Park JY, Kim YC, *et al.* Effects of air pollution on mortality of patients with chronic kidney disease: a large observational cohort study. *Sci Total Environ* 2021;786:147471. doi:10.1016/j. scitotenv.2021.147471
- 45 National Academies of Sciences E, Medicine Committee on Population. *The growing gap in life expectancy by income: implications for federal programs and policy responses*. National Academies Press, 2015.

6

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- 46 Arias E, JQ X. United States life tables, 2018. National vital statistics reports. *National Center for Health Statistics* 2020;69.
- 47 Ravani P, Quinn R, Fiocco M, *et al.* Association of age with risk of kidney failure in adults with stage IV chronic kidney disease in Canada. *JAMA Netw Open* 2020;3:e2017150–e50. doi:10.1001/jamanetworkopen.2020.17150
- 48 Dubay LC, Lebrun LA. Health, behavior, and health care disparities: disentangling the effects of income and race in the United States. Int J Health Serv 2012;42:607–25. doi:10.2190/HS.42.4.c
- Bellomo R, Kellum JA, Ronco C. Acute kidney injury. *The Lancet* 2012;380:756–66. doi:10.1016/S0140-6736(11)61454-2
 Levey AS, Coresh J. Chronic kidney disease. *The Lancet*
- 2012;379:165–80. doi:10.1016/S0140-6736(11)60178-5 51 Gaitonde DY, Cook DL, Rivera IM. Chronic kidney disease: detection
- and evaluation. Am Fam Physician 2017;96:776–83.
 Logan R, Davey P, De Souza N, et al. Assessing the accuracy of ICD-
- 10 coding for measuring rates of and mortality from acute kidney

injury and the impact of electronic alerts: an observational cohort study. *Clin Kidney J* 2020;13:1083-1090. doi:10.1093/ckj/sfz117

- 53 Vlasschaert MEO, Bejaimal SAD, Hackam DG, *et al.* Validity of administrative database coding for kidney disease: a systematic review. *Am J Kidney Dis* 2011;57:29–43. doi:10.1053/j. ajkd.2010.08.031
- 54 Chen TK, Knicely DH, Grams ME. Chronic kidney disease diagnosis and management. JAMA 2019;322:1294. doi:10.1001/ jama.2019.14745
- 55 Sun X, Briel M, Walter SD, *et al.* Is a subgroup effect believable? updating criteria to evaluate the credibility of subgroup analyses. *BMJ* 2010;340:C117. doi:10.1136/bmj.C117

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