

ASYMPTOTIC FORMULAS FOR SIGNIFICANCE LEVELS OF CERTAIN DISTRIBUTIONS

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1. Introduction. The purpose of this paper is to derive asymptotic formulas for the significance levels, or per cent points, of certain well-known statistical distributions.¹ Although we restrict ourselves here to two distributions, those of Chi-Square and of Student's t , it will be apparent that the methods used are applicable to many other distributions as well.

The following results are obtained. Let y_p be the p per cent point of the normal distribution, that is, the distribution defined by

$$(1.1) \quad \Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}v^2} dv,$$

so that

$$(1.2) \quad \Phi(y_p) = 1 - p.$$

If $\chi_{p,n}^2$ and $t_{p,n}$ denote the p per cent points of the Chi-Square and Student's t distributions with n degrees of freedom respectively, then

$$(1.3) \quad \chi_{p,n}^2 = n + y_p \sqrt{2n} + \frac{2}{3} (y_p^2 - 1) + \frac{y_p^3 - 7y_p}{9\sqrt{2n}} + o\left(\frac{1}{\sqrt{n}}\right),$$

and

$$(1.4) \quad t_{p,n} = y_p + \frac{y_p^3 + y_p}{4n} + o\left(\frac{1}{n}\right).$$

These formulas approximate the true values of $\chi_{p,n}^2$ and $t_{p,n}$ to a high degree of accuracy. Tables of comparative values for several values of p and n are given in Section 4.

We shall need the following theorem due to Cramér [3, p. 81; see also pp. 86-87].

THEOREM 1: *Let X_1, X_2, \dots be a sequence of independent, identically distributed random variables having an absolutely continuous distribution function with mean value zero, dispersion σ and finite fifth absolute moment. Let $H_n(x)$ be the distribution function of $(X_1 + \dots + X_n)/(\sigma\sqrt{n})$, and let $n^{-\frac{1}{2}(r-2)}\lambda_r$ denote the r -th semi-invariant of $H_n(x)$. Then*

$$(1.5) \quad \Phi(x) - H_n(x) = \frac{\lambda_3}{3! \sqrt{n}} \Phi^{(3)}(x) - \frac{\lambda_4}{4! n} \Phi^{(4)}(x) - \frac{10\lambda_3^2}{6! n} \Phi^{(6)}(x) + o(n^{-3/2}).$$

¹ This problem was proposed to the author by J. H. Curtiss.

2. The Chi-Square distribution. A random variable X is said to be distributed according to Chi-Square with n degrees of freedom ($X = \chi_n^2$) if its distribution function is

$$(2.1) \quad F_n(x) = \begin{cases} \int_0^x \frac{t^{n-1} e^{-t}}{2^{\frac{n}{2}} \Gamma(\frac{n}{2})} dt, & x \geq 0, \\ 0, & x < 0. \end{cases}$$

The variable $(\chi_n^2 - n)/\sqrt{2n}$ then has the distribution function

$$(2.2) \quad H_n(x) = F_n(n + x\sqrt{2n}).$$

If we write

$$(2.3) \quad \chi_{p,n}^2 = n + y_p\sqrt{2n} + a_n,$$

so that

$$(2.4) \quad F_n(\chi_{p,n}^2) = 1 - p,$$

and let $z_{pn} = y_p + a_n/\sqrt{2n}$, then $H_n(z_{pn}) = 1 - p$, and it follows from (1.1) and (1.2) that

$$(2.5) \quad \Phi(z_{pn}) - H_n(z_{pn}) = \Phi(z_{pn}) - \Phi(y_p) = \frac{1}{\sqrt{2\pi}} \frac{a_n}{\sqrt{2n}} e^{-\frac{1}{2}(y_p + \theta a_n/\sqrt{2n})^2},$$

where $0 < \theta < 1$. Then by a theorem of Liapounoff's [3, p. 77],

$$\frac{|a_n|}{\sqrt{2n}} e^{-\frac{1}{2}(y_p + \theta a_n/\sqrt{2n})^2} \leq \frac{K \log n}{\sqrt{n}},$$

where K denotes a constant independent of n . But if $\lim_{n \rightarrow \infty} |a_n|/\sqrt{2n} = \infty$, then $\lim_{n \rightarrow \infty} H_n(z_{pn})$ is either 0 or 1. Hence $a_n = o(\sqrt{n})$.

Fisher [1, p. 81] has suggested the use of

$$\chi_{p,n}^2 \doteq \frac{1}{2}[y_p + \sqrt{2n-1}]^2.$$

A closer approximation,

$$\chi_{p,n}^2 \doteq n \left[1 - \frac{2}{9n} + y_p \sqrt{\frac{2}{9n}} \right]^3,$$

has been obtained by Wilson and Hilferty [2]. It is interesting to note that, according to (1.3), this last approximation is correct to terms of the zero-th order in n .

We apply Theorem 1 to the variables $X_j = (\chi_j^2 - 1)/\sqrt{2}$, $j = 1, 2, \dots$. Then $\sigma = 1$, and, by the additive property of the Chi-Square distribution [3, p. 45], $H_n(x)$, the distribution function of the variable $(X_1 + \dots + X_n)/\sqrt{n}$,

is related to $F_n(x)$ by (2.2). Thus, $\lambda_3 = 2\sqrt{2}$. It follows from (1.5) and (2.3) that

$$(2.6) \quad \begin{aligned} \lim_{n \rightarrow \infty} \sqrt{2n} [\Phi(z_{pn}) - H_n(z_{pn})] &= \lim_{n \rightarrow \infty} \frac{2}{3\sqrt{2\pi}} (z_{pn}^2 - 1) e^{-\frac{1}{2}z_{pn}^2} \\ &= \frac{2}{3\sqrt{2\pi}} (y_p^2 - 1) e^{-\frac{1}{2}y_p^2}, \end{aligned}$$

since $a_n = o(\sqrt{n})$. Then by (2.5) and (2.6)

$$\lim_{n \rightarrow \infty} a_n = \frac{2}{3}(y_p^2 - 1).$$

According to (2.3) we may now write

$$\chi_{p,2n}^2 = 2n + 2y_p\sqrt{n} + 2r_p + 2b_n,$$

where

$$(2.7) \quad r_p = \frac{1}{3}(y_p^2 - 1),$$

and $b_n = o(1)$. A simple change of variables in (2.1) yields

$$(2.8) \quad F_{2n}(\chi_{p,2n}^2) = \int_{\frac{y_p + r_p}{\sqrt{n}}}^{\frac{y_p + b_n}{\sqrt{n}}} \frac{\sqrt{n} e^{-n} n^n}{\Gamma(n+1)} e^{-v\sqrt{n-r_p}} \left[1 + \frac{v}{\sqrt{n}} + \frac{r_p}{n}\right]^{n-1} dv.$$

If we let

$$(2.9) \quad J_n = \int_{y_p}^{y_p + \frac{b_n}{\sqrt{n}}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}v^2} dv,$$

then

$$(2.10) \quad nJ_n = \frac{b_n\sqrt{n}}{\sqrt{2\pi}} e^{-\frac{1}{2}(y_p + \delta_n)^2},$$

where $\delta_n = o(1)$. By (1.2) and (2.4),

$$(2.11) \quad J_n = \Phi\left(y_p + \frac{b_n}{\sqrt{n}}\right) - F_{2n}(\chi_{p,2n}^2).$$

Using Stirling's formula for $\Gamma(n+1)$ in (2.8), (2.11) becomes

$$\begin{aligned} J_n &= \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} \left[\exp\left\{-\frac{1}{12n} + o(n^{-3})\right\} \right. \\ &\quad \left. + (n-1) \log\left(1 + \frac{v}{\sqrt{n}} + \frac{r_p}{n}\right) + \frac{v^2}{2} - v\sqrt{n} - r_p\right] - 1 \Big] dv \\ &= \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} [e^{4n} - 1] dv, \end{aligned}$$

where $A_n = \frac{1}{n} \left(-\frac{1}{12} - r_p - \frac{r_p^2}{2} + \frac{v^2}{2} + v^2 r_p - \frac{v^4}{4} \right) + \frac{1}{\sqrt{n}} \left(\frac{v^3}{3} - v - v r_p \right) + f_p(v)$, $n f_n(v)$ being dominated by $P(|v|)$, where P is a polynomial in v independent of n , and $f_n(v) = 0(n^{-3/2})$. Then

$$(2.12) \quad J_n = \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{n\sqrt{2\pi}} \left[-\frac{1}{12} - r_p - \frac{r_p^2}{2} + v^2 \left(1 + 2r_p + \frac{r_p^2}{2} \right) - v^4 \left(\frac{7}{12} + \frac{r_p}{3} \right) + \frac{v^6}{18} \right] dv + \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi n}} \left[\frac{v^3}{3} - v - v r_p \right] dv + \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} g_n(v) dv + \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} \left(\sum_{j=3}^{\infty} \frac{A_n^j}{j!} \right) dv,$$

where $g_n(v)$ has the same properties given above for $f_n(v)$. If we call these last integrals K_1 , K_2 , K_3 and K_4 respectively, we see that

$$(2.13) \quad \lim_{n \rightarrow \infty} nK_3 = \int_{y_p}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} \lim_{n \rightarrow \infty} n g_n(v) dv = 0.$$

Also, since A_n^j , $j = 1, 2, \dots$ is dominated by $P_j(|v|)$, $P_j(v)$ being a polynomial in v independent of n , we see that

$$\sum_{j=3}^{\infty} \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} \frac{|A_n|^j}{j!} dv \leq \sum_{j=3}^{\infty} e^{-\frac{1}{2} \left(y_p + \frac{b_n}{\sqrt{n}} \right)^2} \frac{Q_j \left(y_p + \frac{b_n}{\sqrt{n}} \right)}{j!},$$

where Q_j is a polynomial. Since this last sum converges, we have

$$(2.14) \quad nK_4 = \sum_{j=3}^{\infty} \int_{y_p + \frac{b_n}{\sqrt{n}}}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{\sqrt{2\pi}} \frac{nA_n^j}{j!} dv,$$

and by the uniform convergence of (2.14),

$$(2.15) \quad \lim_{n \rightarrow \infty} nK_4 = \sum_{j=3}^{\infty} \int_{y_p}^{\infty} \frac{e^{-\frac{1}{2}v^2}}{j! \sqrt{2\pi}} \lim_{n \rightarrow \infty} nA_n^j dv = 0,$$

since $A_n^j = 0(n^{-j/2})$.

Integrating by parts we obtain

$$nK_2 = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \left(y_p + \frac{b_n}{\sqrt{n}} \right)^2} \left(\frac{2}{3} y_p b_n + \frac{b_n^2}{3\sqrt{n}} \right),$$

and since $b_n = o(1)$,

$$(2.16) \quad \lim_{n \rightarrow \infty} nK_2 = 0.$$

Further integration by parts and the use of (2.7) yields

$$(2.17) \quad \lim_{n \rightarrow \infty} nK_1 = \frac{e^{-4y_p^2}}{36\sqrt{2\pi}} (y_p^3 - 7y_p).$$

Then, by (2.10), (2.12), (2.13), (2.15), (2.16) and (2.17),

$$\lim_{n \rightarrow \infty} b_n \sqrt{n} = \frac{1}{36} (y_p^3 - 7y_p),$$

so that

$$\chi_{p,2n}^2 = 2n + 2y_p \sqrt{n} + \frac{2}{3} (y_p^2 - 1) + \frac{y_p^3 - 7y_p}{18\sqrt{n}} + o\left(\frac{1}{\sqrt{n}}\right).$$

Equation (1.3) now follows at once.

3. Student's t . If the random variable Y_n has the distribution function $\Phi(x/\sqrt{n})$, then $t_n = Y_n/\chi_n$ is distributed according to Student's distribution for n degrees of freedom and has the distribution function

$$G_n(x) = \int_{-\infty}^x \frac{1}{\sqrt{n\pi}} \frac{\Gamma[\frac{1}{2}(n+1)]}{\Gamma[\frac{1}{2}n]} \left(1 + \frac{t^2}{n}\right)^{-\frac{1}{2}(n+1)} dt.$$

If $\sigma = \sqrt{n/(n-2)}$, the variable t_n/σ then has the distribution function

$$(3.1) \quad H_n(x) = G_n(x\sigma).$$

If we write

$$(3.2) \quad t_{p,n} = y_p + a_n,$$

so that

$$(3.3) \quad G_n(t_{p,n}) = 1 - p,$$

and let $z_{pn} = t_{p,n}/\sigma$, then $H_n(z_{pn}) = 1 - p$, and it follows from (1.1) and (1.2) that

$$(3.4) \quad \begin{aligned} \Phi(z_{pn}) - H_n(z_{pn}) &= \Phi(z_{pn}) - \Phi(y_p) \\ &= \frac{1}{\sqrt{2\pi}} \left[y_p \left(\frac{1}{\sigma} - 1 \right) + \frac{a_n}{\sigma} \right] e^{-\frac{1}{2} \left[y_p + \theta \left(y_p \left(\frac{1}{\sigma} - 1 \right) + a_n \right) \right]^2}, \end{aligned}$$

where $0 < \theta < 1$. Then by Liapounoff's Theorem [3, p. 77],

$$\left| y_p \left(\frac{1}{\sigma} - 1 \right) + \frac{a_n}{\sigma} \right| e^{-\frac{1}{2} \left[y_p + \theta \left(y_p \left(\frac{1}{\sigma} - 1 \right) + a_n \right) \right]^2} \leq \frac{K \log n}{\sqrt{n}},$$

where K denotes a constant independent of n . But if $\lim_{n \rightarrow \infty} |a_n| = \infty$, then

$\lim_{n \rightarrow \infty} H_n(z_{pn})$ is either 0 or 1.

Hence $a_n = o(1)$.

We apply Theorem 1 to the variables $X_j = Y_n/\chi_n$, $j = 1, 2, \dots$. Then $\sigma = \sqrt{n/(n-2)}$, and by the additive property of the normal distribution, $H_n(x)$, the distribution function of $(X_1 + \dots + X_n)/(\sigma\sqrt{n})$, satisfies the relation (3.1). Thus $\lambda_3 = 0$ and $\lambda_4 = 6n/(n-4)$. It follows from (1.5) and (3.2) that

$$(3.5) \quad \begin{aligned} \lim_{n \rightarrow \infty} n[\Phi(z_{pn}) - H_n(z_{pn})] &= \lim_{n \rightarrow \infty} \frac{n}{4(n-4)\sqrt{2\pi}} (z_{pn}^3 - 3z_{pn})e^{-\frac{1}{2}z_{pn}^2} \\ &= \frac{1}{4\sqrt{2\pi}} (y_p^3 - 3y_p)e^{-\frac{1}{2}y_p^2}, \end{aligned}$$

since $a_n = o(1)$. By (3.4) and (3.5) we have

$$\lim_{n \rightarrow \infty} n \left[y_p \left(\frac{1}{\sigma} - 1 \right) + \frac{a_n}{\sigma} \right] = \frac{y_p^3 - 3y_p}{4}.$$

But $\lim_{n \rightarrow \infty} n(1 - \sigma)/\sigma = -1$, so that

$$\lim_{n \rightarrow \infty} na_n = \frac{y_p^3 + y_p}{4}.$$

Hence

$$a_n = \frac{y_p^3 + y_p}{4n} + o\left(\frac{1}{n}\right),$$

and equation (1.4) follows at once.

4. Tables. The following tables compare the true values of $\chi_{p,n}^2$ and $t_{p,n}$ with those obtained from (1.3) and (1.4). The true values [4], [5], (to three decimal places) are shown in *italics*.

TABLE 1
 $\chi_{p,n}^2$

$p \backslash n$.01	.05	.1	.5	.9
10	23.253 <i>23.209</i>	18.318 <i>18.307</i>	15.989 <i>15.987</i>	9.333 <i>9.342</i>	4.875 <i>4.865</i>
30	50.908 <i>50.892</i>	43.777 <i>43.773</i>	40.257 <i>40.256</i>	29.333 <i>29.336</i>	20.600 <i>20.599</i>
50	76.163 <i>76.154</i>	67.507 <i>67.505</i>	63.168 <i>63.167</i>	49.333 <i>49.335</i>	37.689 <i>37.689</i>
100	135.811 <i>135.807</i>	124.343 <i>124.342</i>	118.499 <i>118.498</i>	99.333 <i>99.334</i>	82.358 <i>82.358</i>

TABLE II

		$t_{p,n}$				
		.0125	.025	.05	.125	.25
10	p	2.579	2.197	1.797	1.212	0.700
	n	2.634	2.228	1.813	1.221	0.700
30	p	2.354	2.039	1.696	1.171	0.683
	n	2.360	2.042	1.697	1.173	0.683
60	p	2.298	2.000	1.670	1.161	0.679
	n	2.299	2.000	1.671	1.162	0.679
120	p	2.270	1.980	1.658	1.156	0.677
	n	2.270	1.980	1.658	1.156	0.677

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