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# ASYMPTOTIC PROPERTIES OF THE MAXIMUM LIKELIHOOD ESTIMATOR FOR STOCHASTIC PARABOLIC EQUATIONS WITH ADDITIVE FRACTIONAL BROWNIAN MOTION

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ABSTRACT. A parameter estimation problem is considered for a diagonalizable stochastic evolution equation using a finite number of the Fourier coefficients of the solution. The equation is driven by additive noise that is white in space and fractional in time with the Hurst parameter  $H \geq 1/2$ . The objective is to study asymptotic properties of the maximum likelihood estimator as the number of the Fourier coefficients increases. A necessary and sufficient condition for consistency and asymptotic normality is presented in terms of the eigenvalues of the operators in the equation.

## 1. INTRODUCTION

In the classical statistical estimation problem, the starting point is a family  $\mathbf{P}^\theta$  of probability measures depending on the parameter  $\theta$  in some subset  $\Theta$  of a finite-dimensional Euclidean space. Each  $\mathbf{P}^\theta$  is the distribution of a random element. It is assumed that a realization of one random element corresponding to one value  $\theta = \theta_0$  of the parameter is observed, and the objective is to estimate the values of this parameter from the observations.

The intuition is to select the value  $\theta$  corresponding to the random element that is *most likely* to produce the observations. A rigorous mathematical implementation of this idea leads to the notion of the regular statistical model [4]: the statistical model (or estimation problem)  $\mathbf{P}^\theta$ ,  $\theta \in \Theta$ , is called regular, if the following two conditions are satisfied:

- there exists a probability measure  $\mathbf{Q}$  such that all measures  $\mathbf{P}^\theta$  are absolutely continuous with respect to  $\mathbf{Q}$ ;
- the density  $d\mathbf{P}^\theta/d\mathbf{Q}$ , called the likelihood ratio, has a special property, called local asymptotic normality.

If at least one of the above conditions is violated, the problem is called singular.

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In regular models, the estimator  $\hat{\theta}$  of the unknown parameter is constructed by maximizing the likelihood ratio and is called the maximum likelihood estimator (MLE). Since, as a rule,  $\hat{\theta} \neq \theta_0$ , the consistency of the estimator is studied, that is, the convergence of  $\hat{\theta}$  to  $\theta_0$  as more and more information becomes available. In all known regular statistical problems, the amount of information can be increased in one of two ways: (a) increasing the sample size, for example, the observation time interval (large sample asymptotic); (b) reducing the amplitude of noise (small noise asymptotic).

In finite-dimensional models, the only way to increase the sample size is to increase the observation time. In infinite-dimensional models, in particular, those provided by stochastic partial differential equations (SPDEs), another possibility is to increase the dimension of the spatial projection of the observations. Thus, a consistent estimator can be possible on a finite time interval with fixed noise intensity. This possibility was first suggested by Huebner et al. [2] for parabolic equations driven by additive space-time white noise, and was further investigated by Huebner and Rozovskii [3], where a necessary and sufficient condition for the existence of a consistent estimator was stated in terms of the orders of the operators in the equation.

The objective of the current paper is to extend the model from [3] to parabolic equations in which the time component of the noise is fractional with the Hurst parameter  $H \geq 1/2$ . More specifically, we consider an abstract evolution equation

$$u(t) + \int_0^t (\mathcal{A}_0 + \theta \mathcal{A}_1)u(s)ds = W^H(t), \quad (1.1)$$

where  $\mathcal{A}_0, \mathcal{A}_1$  are known linear operators and  $\theta \in \Theta \subseteq \mathbb{R}$  is the unknown parameter; the zero initial condition is taken to simplify the presentation. The noise  $W^H(t)$  is a formal series

$$W^H(t) = \sum_{j=1}^{\infty} w_j^H(t)h_j, \quad (1.2)$$

where  $\{w_j^H, j \geq 1\}$  are independent fractional Brownian motions with the same Hurst parameter  $H \geq 1/2$  and  $\{h_j, j \geq 1\}$  is an orthonormal basis in a Hilbert space  $\mathbf{H}$ ;  $H = 1/2$  corresponds to the usual space-time white noise. Existence and uniqueness of the solution for such equations are well-known for all  $H \in (0, 1)$  (see, for example, Tindel et al. [14, Theorem 1]).

The main additional assumption about (1.1), both in [3] and in the current paper, is that the equation is *diagonalizable*:  $\{h_j, j \geq 1\}$  from (1.2) is a common system of eigenfunction of the operators  $\mathcal{A}_0$  and  $\mathcal{A}_1$ :

$$\mathcal{A}_0 h_j = \rho_j h_j, \quad \mathcal{A}_1 h_j = \nu_j h_j. \quad (1.3)$$

Under certain conditions on the numbers  $\rho_j, \nu_j$ , the solution of (1.1) is a convergent Fourier series  $u(t) = \sum_{j \geq 1} u_j(t)h_j$ , and each  $u_j(t)$  is a fractional Ornstein-Uhlenbeck (OU) process. An  $N$ -dimensional projection of the solution is then an  $N$ -dimensional fractional OU process with independent components. A Girsanov-type formula (for example, from Kleptsyna et al. [7, Theorem 3]) leads to a maximum likelihood estimator  $\hat{\theta}_N$  of  $\theta$  based on the first  $N$  Fourier coefficients  $u_1, \dots, u_N$  of the solution

of (1.1). An explicit expression for this estimator exists but requires a number of additional notations; see formula (3.8) on page 8 below.

The following is the main results of the paper.

**Theorem 1.1.** *Define  $\mu_j = \theta\nu_j + \rho_j$  and assume that the series  $\sum_j(1 + |\mu_j|)^{-\gamma}$  converges for some  $\gamma > 0$ . Then the maximum likelihood estimator  $\hat{\theta}_N$  of  $\theta$  is strongly consistent and asymptotically normal, as  $N \rightarrow \infty$ , if and only if the series  $\sum_j \nu_j^2 \mu_j^{-1}$  diverges; the rate of convergence of the estimator is given by the square root of the partial sums of this series: as  $N \rightarrow \infty$ , the sequence  $\left(\sum_{j \leq N} \nu_j^2 \mu_j^{-1}\right)^{1/2} (\hat{\theta}_N - \theta)$  converges in distribution to a standard Gaussian random variable.*

If the operators  $\mathcal{A}_0$  and  $\mathcal{A}_1$  are elliptic of orders  $m_0$  and  $m_1$  on  $L_2(M)$ , where  $M$  is a  $d$ -dimensional manifold, and  $2m = \max(m_0, m_1)$ , then the condition of the theorem becomes  $m_1 \geq m - (d/2)$ ; in the case  $H = 1/2$  this is known from [3]. Thus, beside extending the results of [3] to fractional-in-time noise, we also generalize the necessary and sufficient condition for consistency of the estimator.

While parameter estimation for the finite-dimensional fractional OU and similar processes has been recently investigated by Tudor and Viens [15] for all  $H \in (0, 1)$ , our analysis in infinite dimensions requires more delicate results: an explicit expression for the Laplace transform of a certain functional of the fractional OU process, as obtained by Kleptsyna and Le Breton [6], and for now this expression exists only for  $H \geq 1/2$ .

## 2. STOCHASTIC PARABOLIC EQUATIONS WITH ADDITIVE FBM

In this section we introduce a diagonalizable stochastic parabolic equation depending on a parameter and study the main properties of the solution.

Let  $\mathbf{H}$  be a separable Hilbert space with the inner product  $(\cdot, \cdot)_0$  and the corresponding norm  $\|\cdot\|_0$ . Let  $\Lambda$  be a densely-defined linear operator on  $\mathbf{H}$  with the following property: there exists a positive number  $c$  such that  $\|\Lambda u\|_0 \geq c\|u\|_0$  for every  $u$  from the domain of  $\Lambda$ . Then the operator powers  $\Lambda^\gamma$ ,  $\gamma \in \mathbb{R}$ , are well defined and generate the spaces  $\mathbf{H}^\gamma$ : for  $\gamma > 0$ ,  $\mathbf{H}^\gamma$  is the domain of  $\Lambda^\gamma$ ;  $\mathbf{H}^0 = \mathbf{H}$ ; for  $\gamma < 0$ ,  $\mathbf{H}^\gamma$  is the completion of  $\mathbf{H}$  with respect to the norm  $\|\cdot\|_\gamma := \|\Lambda^\gamma \cdot\|_0$  (see for instance Krein et al. [8]). By construction, the collection of spaces  $\{\mathbf{H}^\gamma, \gamma \in \mathbb{R}\}$  has the following properties:

- $\Lambda^\gamma(\mathbf{H}^r) = \mathbf{H}^{r-\gamma}$  for every  $\gamma, r \in \mathbb{R}$ ;
- For  $\gamma_1 < \gamma_2$  the space  $\mathbf{H}^{\gamma_2}$  is densely and continuously embedded into  $\mathbf{H}^{\gamma_1}$ :  $\mathbf{H}^{\gamma_2} \subset \mathbf{H}^{\gamma_1}$  and there exists a positive number  $c_{12}$  such that  $\|u\|_{\gamma_1} \leq c_{12}\|u\|_{\gamma_2}$  for all  $u \in \mathbf{H}^{\gamma_2}$ ;
- For every  $\gamma \in \mathbb{R}$  and  $m > 0$ , the space  $\mathbf{H}^{\gamma-m}$  is the dual of  $\mathbf{H}^{\gamma+m}$  relative to the inner product in  $\mathbf{H}^\gamma$ , with duality  $\langle \cdot, \cdot \rangle_{\gamma, m}$  given by

$$\langle u_1, u_2 \rangle_{\gamma, m} = (\Lambda^{\gamma-m} u_1, \Lambda^{\gamma+m} u_2)_0, \text{ where } u_1 \in \mathbf{H}^{\gamma-m}, u_2 \in \mathbf{H}^{\gamma+m}.$$

In the above construction, the operator  $\Lambda$  can be bounded, and then the norms in all the spaces  $\mathbf{H}^\gamma$  will be equivalent. A more interesting situation is therefore when  $\Lambda$  is unbounded and plays the role of the first-order operator.

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space and let  $\{w_j^H, j \geq 1\}$  be a collection of independent fractional Brownian motions on this space with the same Hurst parameter  $H \in (0, 1)$ :

$$\mathbb{E}w_j^H(t) = 0, \quad \mathbb{E}(w_j^H(t)w_j^H(s)) = \frac{1}{2}(t^{2H} + s^{2H} - |t - s|^{2H}).$$

Consider the following equation:

$$\begin{cases} du(t) + (\mathcal{A}_0 + \theta\mathcal{A}_1)u(t)dt = \sum_{j \geq 1} g_j(t)dw_j^H(t), & 0 < t \leq T, \\ u(0) = u_0 \end{cases} \quad (2.1)$$

where  $\mathcal{A}_0, \mathcal{A}_1$  are linear operators,  $g_j$  are non-random, and  $\theta$  is a scalar parameter belonging to an open set  $\Theta \subset \mathbb{R}$ .

**Definition 2.1.**

- (1) Equation (2.1) is called *diagonalizable* if the operators  $\mathcal{A}_0, \mathcal{A}_1$ , have a common system of eigenfunctions  $\{h_j, j \geq 1\}$  such that  $\{h_j, j \geq 1\}$  is an orthonormal basis in  $\mathbf{H}$  and each  $h_j$  belongs to  $\bigcap_{\gamma \in \mathbb{R}} \mathbf{H}^\gamma$ .
- (2) Equation (2.1) is called  *$(m, \gamma)$ -parabolic* for some numbers  $m \geq 0$  and  $\gamma \in \mathbb{R}$  if
  - the operator  $\mathcal{A}_0 + \theta\mathcal{A}_1$  is uniformly bounded from  $\mathbf{H}^{\gamma+m}$  to  $\mathbf{H}^{\gamma-m}$  for  $\theta \in \Theta$ : there exists a positive real number  $C_1$  such that

$$\|(\mathcal{A}_0 + \theta\mathcal{A}_1)v\|_{\gamma-m} \leq C_1 \|v\|_{\gamma+m} \quad (2.2)$$

for all  $\theta \in \Theta, v \in \mathbf{H}^{\gamma+m}$ ;

- there exists a positive number  $\delta$  and a real number  $C$  such that, for every  $v \in \mathbf{H}^{\gamma+m}, \theta \in \Theta$ ,

$$-2\langle (\mathcal{A}_0 + \theta\mathcal{A}_1)v, v \rangle_{\gamma,m} + \delta \|v\|_{\gamma+m}^2 \leq C \|v\|_{\gamma}^2. \quad (2.3)$$

**Remark 2.2.** If equation (2.1) is  $(m, \gamma)$ -parabolic, then condition (2.3) implies that

$$\langle (2\mathcal{A}_0 + 2\theta\mathcal{A}_1 + CI)v, v \rangle_{\gamma,m} \geq \delta \|v\|_{\gamma+m}^2,$$

where  $I$  is the identity operator. The Cauchy-Schwartz inequality and the continuous embedding of  $\mathbf{H}^{\gamma+m}$  into  $\mathbf{H}^\gamma$  then imply

$$\|(2\mathcal{A}_0 + 2\theta\mathcal{A}_1 + CI)v\|_{\gamma} \geq \delta_1 \|v\|_{\gamma}$$

for some  $\delta_1 > 0$  uniformly in  $\theta \in \Theta$ . As a result, we can take  $\Lambda = (2\mathcal{A}_0 + 2\theta^*\mathcal{A}_1 + CI)^{1/(2m)}$  for some fixed  $\theta^* \in \Theta$ . If the operator  $\mathcal{A}_0 + \theta\mathcal{A}_1$  is unbounded, it is natural to say that  $\mathcal{A}_0 + \theta\mathcal{A}_1$  has order  $2m$  and  $\Lambda$  has order 1.

*From now on, if equation (2.1) is  $(m, \gamma)$ -parabolic and diagonalizable, we will assume that the operator  $\Lambda$  has the same eigenfunctions as the operators  $\mathcal{A}_0, \mathcal{A}_1$ ; by Remark 2.2, this leads to no loss of generality.*

For a diagonalizable equation, condition (2.3) can be expressed in terms of the eigenvalues of the operators in the equation.

**Theorem 2.3.** *Assume that equation (2.1) is diagonalizable and*

$$\mathcal{A}_0 h_j = \rho_j h_j, \quad \mathcal{A}_1 h_j = \nu_j h_j.$$

*With no loss of generality (see Remark 2.2), we also assume that*

$$\Lambda h_j = \lambda_j h_j.$$

*Then equation (2.1) is  $(m, \gamma)$ -parabolic if and only if there exist positive real numbers  $\delta, C_1$  and a real number  $C_2$  such that, for all  $j \geq 1$  and  $\theta \in \Theta$ ,*

$$\lambda_j^{-2m} |\rho_j + \theta \nu_j| \leq C_1; \tag{2.4}$$

$$-2(\rho_j + \theta \nu_j) + \delta \lambda_j^{2m} \leq C_2. \tag{2.5}$$

*Proof.* We show that, for a diagonalizable equation, (2.4) is equivalent to (2.2) and (2.5) is equivalent to (2.3). Indeed, note that for every  $\gamma, r \in \mathbb{R}$ ,

$$\|h_j\|_{\gamma+r} = \|\Lambda^r h_j\|_\gamma = \lambda_j^r \|h_j\|_\gamma.$$

Then (2.4) is (2.2) and (2.5) is (2.3), with  $v = h_j$ . Since both (2.4) and (2.5) are uniform in  $j$  and the collection  $\{h_j, j \geq 1\}$  is dense in every  $\mathbf{H}^\gamma$ , the proof of the theorem is complete.  $\square$

**Remark 2.4.** (a) As conditions (2.4), (2.5) do not involve  $\gamma$ , we conclude that a diagonalizable equation is  $(m, \gamma)$ -parabolic *for some*  $\gamma$  if and only if it is  $(m, \gamma)$ -parabolic *for every*  $\gamma$ . As a result, in the future we will simply say that the equation is  $m$ -parabolic.

(b) If the operators  $\mathcal{A}_0 + \theta \mathcal{A}_1$  and  $\Lambda$  are unbounded, then (2.5) implies that  $\mu_j(\theta) = \rho_j + \theta \nu_j$  is positive for all sufficiently large  $j$ .

From now on we will assume that equation (2.1) is diagonalizable and fix the basis  $\{h_j, j \geq 1\}$  in  $\mathbf{H}$ . Since each  $h_j$  belongs to every  $\mathbf{H}^\gamma$  and, by construction,  $\bigcap_\gamma \mathbf{H}^\gamma$  is dense in  $\bigcup_\gamma \mathbf{H}^\gamma$ , every element  $f$  of  $\bigcup_\gamma \mathbf{H}^\gamma$  has a unique expansion  $\sum_{j \geq 1} f_j h_j$ , where  $f_j = \langle f, h_j \rangle_{0,m}$  for a suitable  $m$ .

**Definition 2.5.** *The space-time fractional Brownian motion  $W^H$  is an element of  $\bigcup_{\gamma \in \mathbb{R}} \mathbf{H}^\gamma$  with the expansion*

$$W^H(t) = \sum_{j \geq 1} w_j^H(t) h_j. \tag{2.6}$$

**Definition 2.6.** *Let  $W^H$  be a space-time fractional Brownian motion. The solution of the diagonalizable equation*

$$\begin{cases} du(t) + (\mathcal{A}_0 + \theta \mathcal{A}_1)u(t)dt = dW^H(t), & 0 < t \leq T, \\ u(0) = u_0 \end{cases} \tag{2.7}$$

$u_0 \in \mathbf{H}$ , *is a random process with values in  $\bigcup_\gamma \mathbf{H}^\gamma$  and an expansion*

$$u(t) = \sum_{j \geq 1} u_j(t) h_j, \tag{2.8}$$

where

$$u_j(t) = (u_0, h_j)_0 e^{-(\theta\nu_j + \rho_j)t} + \int_0^t e^{-(\theta\nu_j + \rho_j)(t-s)} dw_j^H(s). \quad (2.9)$$

Notice that, due to the special structure of the equation, Definition 2.6 implies both existence and uniqueness of the solution.

To simplify further notations we write

$$\mu_j(\theta) = \theta\nu_j + \rho_j. \quad (2.10)$$

By (2.5), if equation (2.1) is  $m$ -parabolic and diagonalizable, then, for every  $\theta \in \Theta$ , there exists a positive integer  $J$  such that

$$\mu_j(\theta) > 0 \quad \text{for all } j \geq J.$$

**Theorem 2.7.** *Assume that*

- (1)  $H \geq 1/2$ ;
- (2) equation (2.1) is  $m$ -parabolic and diagonalizable;
- (3) There exists a positive real number  $\gamma$  such that

$$\sum_{j \geq 1} (1 + |\mu_j(\theta)|)^{-\gamma} < \infty. \quad (2.11)$$

Then, for every  $t > 0$ ,

- (1)  $W^H(t) \in L_2(\Omega; \mathbf{H}^{-m\gamma})$ ;
- (2)  $u(t) \in L_2(\Omega; \mathbf{H}^{-m\gamma + 2mH})$ .

*Proof.* Condition (2.11) implies that  $\lim_{j \rightarrow \infty} |\mu_j| = \infty$ , and consequently the operators  $\mathcal{A}_0 + \theta \mathcal{A}_1$  and  $\Lambda$  are unbounded. The parabolicity assumption and Theorem 2.3 then imply that, for all sufficiently large  $j$ ,

$$1 + |\mu_j(\theta)| \leq C_2 \lambda_j^{2m},$$

uniformly in  $\theta \in \Theta$ .

$$\mathbb{E} \|W^H(t)\|_{-m\gamma}^2 = t^{2H} \sum_{j \geq 1} \lambda_j^{-2m\gamma} \leq C_2 t^{2H} \sum_{j \geq 1} (1 + |\mu_j(\theta)|)^{-\gamma} < \infty.$$

Next, the properties of the fractional Brownian motion imply

$$\mathbb{E} u_j^2(t) = H(2H - 1) e^{-2\mu_j(\theta)t} \int_0^t \int_0^t e^{\mu_j(\theta)(s_1 + s_2)} |s_1 - s_2|^{2H-2} ds_1 ds_2;$$

see, for example, Pipiras and Taqqu [11, formulas (4.1), (4.2)]. By direct computation,

$$\lim_{j \rightarrow \infty} |\mu_j(\theta)|^{2H} \mathbb{E} u_j^2(t) = H(2H - 1) \int_0^\infty x^{2H-2} e^{-x} dx = H(2H - 1) \Gamma(2H - 1). \quad (2.12)$$

Consequently,

$$\sum_{j=1}^{\infty} (1 + |\mu_j(\theta)|)^{-\gamma + 2H} \mathbb{E} |u_j(t)|^2 < \infty, \quad (2.13)$$

and the second conclusion of the theorem follows.  $\square$

**Example 2.8.** (a) For  $0 < t \leq T$  and  $x \in (0, 1)$ , consider the equation

$$du(t, x) - \theta u_{xx}(t, x)dt = dW^H(t, x) \quad (2.14)$$

with periodic boundary conditions, where  $u_{xx} = \partial^2 u / \partial x^2$ . Then  $\mathbf{H}^\gamma$  is the Sobolev space on the unit circle (see, for example, Shubin [13, Section I.7]) and  $\Lambda = \sqrt{I - \Delta}$ , where  $\Delta$  is the Laplace operator on  $(0, 1)$  with periodic boundary conditions. Direct computations show that equation (2.14) is diagonalizable; it is 1-parabolic if and only if  $\theta > 0$ . Also,  $\mu_j = -\theta\pi^2 j^2$ , so that, by Theorem 2.7 the solution  $u(t)$  of (2.14) is an element of  $L_2(\Omega; \mathbf{H}^{-\gamma+2H})$  for every  $t > 0$ ,  $\gamma > 1/2$ , and  $\theta > 0$ .

(b) Let  $G$  be a smooth bounded domain in  $\mathbb{R}^d$ . Let  $\Delta$  be the Laplace operator on  $G$  with zero boundary conditions. It is known (for example, from Shubin [13]), that

- (1) the eigenfunctions  $\{h_j, j \geq 1\}$  of  $\Delta$  are smooth in  $G$  and form an orthonormal basis in  $L_2(G)$ ;
- (2) the corresponding eigenvalues  $\sigma_j, j \geq 1$ , can be arranged so that  $0 < -\sigma_1 \leq -\sigma_2 \leq \dots$ , and there exists a number  $c > 0$  such that  $|\sigma_j| \sim cj^{2/d}$ , that is,

$$\lim_{j \rightarrow \infty} |\sigma_j| j^{-2/d} = c. \quad (2.15)$$

We take  $\mathbf{H} = L_2(G)$ ,  $\Lambda = \sqrt{I - \Delta}$ , where  $I$  is the identity operator. Then  $\|\Lambda u\|_0 \geq \sqrt{1 - \sigma_1} \|u\|_0$  and the operator  $\Lambda$  generates the Hilbert spaces  $\mathbf{H}^\gamma$ , and, for every  $\gamma \in \mathbb{R}$ , the space  $\mathbf{H}^\gamma$  is the closure of the set of smooth compactly supported function on  $G$  with respect to the norm

$$\left( \sum_{j \geq 1} (1 + j^2)^\gamma |\varphi_j|^2 \right)^{1/2}, \quad \text{where } \varphi_j = \int_G \varphi(x) h_j(x) dx,$$

which is an equivalent norm in  $\mathbf{H}^\gamma$ . Then, for every  $\theta \in \mathbb{R}$ , the stochastic equation

$$du - (\Delta u + \theta u)dt = dW^H(t, x) \quad (2.16)$$

is diagonalizable and 1-parabolic. Indeed, we have  $\mathcal{A}_1 = I$ ,  $\mathcal{A}_0 = -\Delta$ , and

$$-2\langle \mathcal{A}_0 v, v \rangle_{\gamma,1} = -2\|v\|_{\gamma+1}^2 + 2\|u\|_\gamma^2,$$

so that (2.3) holds with  $\delta = 2$  and  $C = 2 - \theta$ . Finally, by (2.15) we see that (2.11) holds for every  $\gamma > d/2$ . As a result, by Theorem 2.7, the solution  $u(t)$  of (2.16) is an element of  $L_2(\Omega; \mathbf{H}^{-\gamma+2H})$  for every  $t > 0$ ,  $\gamma > d/2$ , and  $\theta \in \mathbb{R}$ .

### 3. THE MAXIMUM LIKELIHOOD ESTIMATOR AND ITS PROPERTIES

Consider the diagonalizable equation

$$du(t) + (\mathcal{A}_0 + \theta \mathcal{A}_1)u(t)dt = dW^H(t) \quad (3.1)$$

with solution  $u(t) = \sum_{j \geq 1} u_j(t) h_j$  given by (2.9); for simplicity, we assume that  $u(0) = 0$ . Suppose that the processes  $u_1(t), \dots, u_N(t)$  can be observed for all  $t \in [0, T]$ . The problem is to estimate the parameter  $\theta$  using these observations.

Recall the notation  $\mu_j(\theta) = \rho_j + \nu_j\theta$ , where  $\rho_j$  and  $\nu_j$  are the eigenvalues of  $\mathcal{A}_0$  and  $\mathcal{A}_1$ , respectively. Then each  $u_j$  is a fractional Ornstein-Uhlenbeck process satisfying

$$du_j(t) = -\mu_j(\theta)u_j(t)dt + dw_j^H(t), \quad u_j(0) = 0, \quad (3.2)$$

and, because of the independence of  $w_j^H$  for different  $j$ , the processes  $u_1, \dots, u_N$  are (statistically) independent.

Let  $\Gamma$  denote the Gamma-function (see (2.12)). Following Kleptsyna and Le Breton [6], we introduce the notations

$$\kappa_H = 2H\Gamma\left(\frac{3}{2} - H\right)\Gamma\left(H + \frac{1}{2}\right), \quad k_H(t, s) = \kappa_H^{-1}s^{\frac{1}{2}-H}(t-s)^{\frac{1}{2}-H}; \quad (3.3)$$

$$\lambda_H = \frac{2H\Gamma(3-2H)\Gamma\left(H + \frac{1}{2}\right)}{\Gamma\left(\frac{3}{2} - H\right)}, \quad \mathbf{w}_H(t) = \lambda_H^{-1}t^{2-2H}; \quad (3.4)$$

$$M_j^H(t) = \int_0^t k_H(t, s)dw_j^H(s), \quad Q_j(t) = \frac{d}{d\mathbf{w}_H(t)} \int_0^t k_H(t, s)u_j(s)ds; \quad (3.5)$$

$$Z_j(t) = \int_0^t k_H(t, s)du_j(s). \quad (3.6)$$

By a Girsanov-type formula (see, for example, Kleptsyna et al. [7, Theorem 3]), the measure in the space of continuous,  $\mathbb{R}^N$ -valued functions, generated by the process  $(u_1, \dots, u_N)$  is absolutely continuous with respect to the measure generated by the process  $(w_1^H, \dots, w_N^H)$ , and the density is

$$\exp\left(-\sum_{j=1}^N \mu_j(\theta) \int_0^T Q_j(s)dZ_j(s) - \sum_{j=1}^N \frac{|\mu_j(\theta)|^2}{2} \int_0^T Q_j^2(s)d\mathbf{w}_H(s)\right). \quad (3.7)$$

Maximizing this density with respect to  $\theta$  gives the Maximum Likelihood Estimator (MLE):

$$\widehat{\theta}_N = -\frac{\sum_{j=1}^N \int_0^T \nu_j Q_j(s)(dZ_j(s) + \rho_j Q_j(s)d\mathbf{w}_H(s))}{\sum_{j=1}^N \int_0^T \nu_j^2 Q_j^2(s)d\mathbf{w}_H(s)}. \quad (3.8)$$

An important feature of (3.8) is that the process  $Z_j$  is a semi-martingale ([6, Lemma 2.1]), and so there is no stochastic integration with respect to fractional Brownian motion:  $\int_0^T \nu_j Q_j(s)dZ_j(s)$  is an Itô integral. Notice that, when  $H = 1/2$ , we have  $k_H = 1$ ,  $\mathbf{w}_H(s) = s$ ,  $Q_j(s) = Z_j(s) = u_j(s)$ , and (3.8) becomes

$$\widehat{\theta}_N = -\frac{\sum_{j=1}^N \int_0^T \nu_j u_j(s)(du_j(s) + \rho_j u_j(s)ds)}{\sum_{j=1}^N \int_0^T \nu_j^2 u_j^2(s)du_j(s)}, \quad (3.9)$$

which is the MLE from [3].



Let us also emphasize that an implementation of (3.8) is impossible without the knowledge of  $H$ .

The following is the main result of the paper.

**Theorem 3.1.** *Under the assumptions of Theorem 2.7, the following conditions are equivalent:*

$$(1) \quad \sum_{j=J}^{\infty} \frac{\nu_j^2}{\mu_j(\theta)} = +\infty; \quad (3.10)$$

$$(2) \quad \lim_{N \rightarrow \infty} \hat{\theta}_N = \theta \text{ with probability one,} \quad (3.11)$$

where  $J = \min\{j : \mu_i(\theta) > 0 \text{ for all } i \geq j\}$ .

*Proof.* Following Kleptsyna and Le Breton [6, Equation (4.1)], we conclude that

$$\hat{\theta}_N - \theta = -\frac{\sum_{j=1}^N \int_0^T \nu_j Q_j(s) dM_j^H(s)}{\sum_{j=1}^N \int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)}. \quad (3.12)$$

Both the top and the bottom on the right-hand side of (3.12) are sums of independent random variables; moreover, it is known from [6, page 242] that

$$\mathbb{E} \left( \int_0^T Q_j(s) dM_j^H(s) \right)^2 = \mathbb{E} \int_0^T Q_j^2(s) d\mathbf{w}_H(s) ds. \quad (3.13)$$

From the expression for the Laplace transform of  $\int_0^T Q_j^2(s) d\mathbf{w}_H(s) ds$  (see [6, Equation (4.2)]) direct computations show that

$$\lim_{j \rightarrow \infty} \mu_j(\theta) \mathbb{E} \int_0^T Q_j^2(s) d\mathbf{w}_H(s) ds = \frac{T}{2} > 0 \quad (3.14)$$

and, with  $\text{Var}(\xi)$  denoting the variance of the random variable  $\xi$ ,

$$\lim_{j \rightarrow \infty} \mu_j^3(\theta) \text{Var} \left( \int_0^T Q_j^2(s) d\mathbf{w}_H(s) ds \right) = \frac{T}{2} > 0; \quad (3.15)$$

a detailed derivation of (3.14) and (3.15) is given in the appendix, Lemmas A.1 and A.2 respectively.

We now see that if (3.10) does not hold, then, by (3.14), the series

$$\sum_{j \geq 1} \int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)$$

converges with probability one, which, by (3.12), means that (3.11) cannot hold.

On the other hand, if (3.10) holds, then

$$\sum_{n \geq J} \frac{\nu_n^2 \mu_n^{-1}}{\left( \sum_{j=1}^n \nu_j^2 \mu_j^{-1} \right)^2} < \infty. \quad (3.16)$$

Indeed, setting  $a_n = \nu_n^2 \mu_n^{-1}$  and  $A_n = \sum_{j=1}^n a_j$ , we notice that

$$\sum_{n \geq J} \frac{a_n}{A_n^2} \leq \sum_{n \geq J+1} \left( \frac{1}{A_n} - \frac{1}{A_{n-1}} \right) = \frac{1}{A_J}.$$

Then the strong law of large numbers, together with the observation

$$\mathbb{E} \int_0^T Q_j(s) dM_j^H(s) = 0, \quad j \geq 1,$$

implies

$$\lim_{N \rightarrow \infty} \frac{\sum_{j=1}^N \int_0^T \nu_j Q_j(s) dM_j(s)}{\sum_{j=1}^N \mathbb{E} \int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)} = 0 \quad \text{with probability one.}$$

Next, it follows from (3.16) and (2.11) that

$$\sum_{n \geq J} \frac{\nu_n^4 \mu_n^{-3}}{\left( \sum_{j=J}^n \nu_j^2 \mu_j^{-1} \right)^2} < \infty. \quad (3.17)$$

Then another application of the strong law of large numbers implies that

$$\lim_{N \rightarrow \infty} \frac{\sum_{j=1}^N \int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)}{\sum_{j=1}^N \mathbb{E} \int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)} = 1 \quad (3.18)$$

with probability one, and (3.11) follows.  $\square$

**Corollary 3.2.** *Under assumptions of Theorem 2.7, if (3.10) holds, then*

$$\lim_{N \rightarrow \infty} \sqrt{\sum_{j=J}^N \frac{\nu_j^2}{\mu_j(\theta)}} \left( \hat{\theta}_N - \theta \right) = \zeta \quad (3.19)$$

in distribution, where  $\zeta$  is a Gaussian random variable with mean zero.

*Proof.* This follows from (3.12), (3.18), and the central limit theorem for the sum of independent random variables.  $\square$

Let us now consider a more general equation

$$du = (\mathcal{A}_0 + \theta \mathcal{A}_1) u dt + \mathcal{B} dW^H(t),$$

where  $\mathcal{B}$  is a linear operator. If  $\mathcal{B}^{-1}$  exists, the equation reduced to (3.1) by considering  $v = \mathcal{B}^{-1}u$ . If  $\mathcal{B}^{-1}$  does not exist, we have two possibilities:

- (1)  $(u_0, h_i)_0 = 0$  for every  $i$  such that  $\mathcal{B}h_i = 0$ . In this case,  $u_i(t) = 0$  for all  $t > 0$ , so that we can factor out the kernel of  $\mathcal{B}$  and reduce the problem to invertible  $\mathcal{B}$ .
- (2)  $(u_0, h_i)_0 \neq 0$  for some  $i$  such that  $\mathcal{B}h_i = 0$ . In this case,  $u_i(t) = u_i(0)e^{-\rho_i t - \nu_i \theta t}$  and  $\theta$  is determined exactly from the observations of  $u_i(t)$ :

$$\theta = \frac{1}{\nu_i(t-s)} \ln \frac{u_i(s)}{u_i(t)} - \frac{\rho_i}{\nu_i}, \quad t \neq s.$$

Let  $\mathcal{A}_0, \mathcal{A}_1$  be differential or pseudo-differential operators, either on a smooth bounded domain in  $\mathbb{R}^d$  or on a smooth compact  $d$ -dimensional manifold, and let  $m_0, m_1$ , be the orders of  $\mathcal{A}_0, \mathcal{A}_1$  respectively, so that  $2m = \max(m_0, m_1)$ . Then, under rather general conditions we have

$$\lim_{j \rightarrow \infty} |\nu_j| j^{m_1/d} = c_1, \quad \lim_{j \rightarrow \infty} \mu_j(\theta) j^{2m/d} = c(\theta) \quad (3.20)$$

for some positive numbers  $c_1, c(\theta)$ ; see, for example, Il'in [5] or Safarov and Vassiliev [12]. In particular, this is the case for the operators in equations (2.14) and (2.16).

If (3.20) holds, then condition (3.10) becomes

$$m_1 \geq m - (d/2), \quad (3.21)$$

which, in the case  $H = 1/2$ , was established by Huebner and Rozovskii [3]. In particular, (3.21) holds for equation (2.14) (where  $2m = m_1 = 2$ ), and for equation (2.16) if  $d \geq 2$  (where  $2m = 2, m_1 = 0$ ).

Note that, at least as long as  $H \geq 1/2$ , conditions (3.10) and (3.21) do not involve  $H$ .

The maximum likelihood estimator (3.8) has three features that are clearly attractive: consistency, asymptotic normality, and absence of stochastic integration with respect to fractional Brownian motion. On the other hand, actual implementation of (3.8) is problematic: when  $H > 1/2$ , computing the processes  $Q_j$  and  $Z_j$  is certainly nontrivial. Estimator (3.9) is defined for all  $H \geq 1/2$  and contains only the processes  $u_j$ , but, when  $H > 1/2$ , is not an MLE and is even harder to implement because of the stochastic integral with respect to  $u_j$ .

With or without condition (3.10), a consistent estimator of  $\theta$  is possible in the large time asymptotic: for every  $j \geq 1$ ,

$$\lim_{T \rightarrow \infty} \frac{\int_0^T \nu_j Q_j(s) (dZ_j(s) + \rho_j Q_j(s) d\mathbf{w}_H(s))}{\int_0^T \nu_j^2 Q_j^2(s) d\mathbf{w}_H(s)} = -\theta \quad (3.22)$$

with probability one ([6, Proposition 2.2]). For  $H > 1/2$ , implementation of this estimator is essentially equivalent to the implementation of (3.8).

An alternative to (3.22) was suggested by Maslowski and Pospíšil [10] using the ergodic properties of the OU process. Let us first illustrate the idea on a simple example.

If  $a > 0$  and  $w = w(t)$  is a standard one-dimensional Brownian motion, then the OU process  $dX = -aX(t)dt + dw(t)$  is ergodic and its unique invariant distribution is normal with zero mean and variance  $(2a)^{-1}$ . In particular,

$$\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T X^2(t) dt = \frac{1}{2a} \quad (3.23)$$

with probability one, and so

$$\tilde{a}(T) = \frac{T}{2 \int_0^T X^2(t) dt} \quad (3.24)$$

is a consistent estimator of  $a$  in the long-time asymptotic. Note that the maximum likelihood estimator in this case is

$$\hat{a}(T) = -\frac{\int_0^T X(t)dX(t)}{\int_0^T X^2(s)ds} \quad (3.25)$$

and is strongly consistent for every  $a \in \mathbb{R}$  [9, Theorem 17.4].

Similarly, if  $a > 0$ , then the fractional OU process

$$dX(t) = -aX(t)dt + dw^H(t), \quad X(0) = 0 \quad (3.26)$$

is Gaussian, and, by (2.12) on page 6, converges in distribution, as  $t \rightarrow \infty$ , to the Gaussian random variable with zero mean and variance  $c(H)a^{-2H}$ , where

$$c(H) = H(2H - 1)\Gamma(2H - 1); \quad (3.27)$$

notice that, in the limit  $H \searrow 1/2$ , we recover the result for the usual OU process. Further investigation shows that, similar to (3.23),

$$\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T X^2(s)ds = \frac{c(H)}{a^{2H}}$$

(see [10]). As a result, for every  $j$  such that  $\theta\nu_j + \rho_j > 0$ , we have

$$\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T u_j^2(t)dt = \frac{c(H)}{(\theta\nu_j + \rho_j)^{2H}} \quad (3.28)$$

with probability one. Under an additional assumption that  $\nu_j \neq 0$ , we get an estimator of  $\theta$

$$\tilde{\theta}^{(j)}(T) = \frac{1}{\nu_j} \left( \frac{c(H)T}{\int_0^T u_j^2(t)dt} \right)^{\frac{1}{2H}} - \frac{\rho_j}{\nu_j}. \quad (3.29)$$

This estimator is strongly consistent in the long time asymptotic:  $\lim_{T \rightarrow \infty} |\tilde{\theta}^{(j)}(T) - \theta| = 0$  with probability one ([10, Theorem 5.2]). While not a maximum likelihood estimator, (3.29) is easier to implement computationally than (3.8). If, in Theorem 2.7 on page 6, we have  $\mathcal{A}_0 = 0$ ,  $\nu_j > 0$ , and  $\gamma < 2H$ , then a version of (3.30) exists using all the Fourier coefficients  $u_j$ ,  $j \geq 1$ :

$$\tilde{\theta}(T) = \left( \frac{c(H)T \sum_{j=1}^{\infty} \nu_j^{-2H}}{\sum_{j=1}^{\infty} \int_0^T u_j^2(t)dt} \right)^{\frac{1}{2H}}; \quad (3.30)$$

see [10, Theorem 5.2].

An interesting open question related to both (3.8) and (3.29), (3.30) is how to combine estimation of  $\theta$  with estimation of  $H$ .

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## APPENDIX

Below, we prove equalities (3.14) and (3.15).

**Lemma A.1.** *For every  $\theta \in \Theta$  and  $H \in [1/2, 1)$ ,*

$$\lim_{j \rightarrow \infty} \mu_j(\theta) \mathbb{E} \int_0^T Q_j^2(s) d\mathbf{w}_H(s) = \frac{T}{2}.$$

*Proof.* Denote by  $\Psi_T^H(a, \mu_j)$  the Laplace transform of  $\int_0^T Q_j^2(s) d\mathbf{w}_H(s)$ , namely

$$\Psi_T^H(a, \mu_j(\theta)) = \mathbb{E} \exp \left\{ -a \int_0^T Q_j^2(s) d\mathbf{w}_H(s) \right\}, \quad a > 0. \quad (\text{A.1})$$

We will use the expression for  $\Psi_T^H$  from [6, page 242], and write it as follows

$$\Psi_T^H(a, \mu_j) = \alpha e^{\frac{(\mu_j - \alpha)T}{2}} [\Delta_T^H(\mu_j, \alpha)]^{-\frac{1}{2}}$$

where  $\mu_j = \mu_j(\theta)$ ,  $\alpha := \sqrt{\mu_j^2 + 2a}$ ,

$$\begin{aligned} \Delta_T^H(\mu_j, \alpha) &= \frac{\pi\alpha T e^{-\alpha T} (\alpha^2 - \mu_j^2)}{4 \sin(\pi H)} I_{-H} \left( \frac{\alpha T}{2} \right) I_{H-1} \left( \frac{\alpha T}{2} \right) \\ &\quad + e^{-\alpha T} \left[ \alpha \sinh \left( \frac{\alpha T}{2} \right) + \mu_j \cosh \left( \frac{\alpha T}{2} \right) \right]^2, \end{aligned}$$

and  $I_p$  is the modified Bessel function of the first kind and order  $p$ .

Note that

$$\mathbb{E} \int_0^T Q_j^2(s) d\mathbf{w}_H(s) = - \frac{\partial \Psi_T^H(a, \mu_j)}{\partial a} \Big|_{a=0}.$$

Direct evaluations (for example, using Mathematica computer algebra system) give

$$\frac{\partial \Psi_T^H(a, \mu_j)}{\partial a} \Big|_{a=0} = \frac{2 + 2e^{\mu_j T} (1 - \mu_j T) - \mu_j \pi T I_{H-1} \left( \frac{\mu_j T}{2} \right) I_{-H} \left( \frac{\mu_j T}{2} \right) \csc(H\pi)}{4\mu_j^2 e^{\mu_j T}},$$

where  $\csc(x) = 1/\sin(x)$ . By combining formulas (6.106), (6.155), and (6.162) in [1], we conclude that, for all  $p \in (-1, 1)$ ,  $p \neq 0$ , we have  $I_p(x) \sim e^x / \sqrt{2\pi x}$ ,  $x \rightarrow \infty$ , that is,

$$\lim_{x \rightarrow +\infty} \sqrt{2\pi x} e^{-x} I_p(x) = 1. \quad (\text{A.2})$$

Therefore

$$\frac{\partial \Psi_T^H(a, \mu_j)}{\partial a} \Big|_{a=0} \sim \frac{2 + 2e^{\mu_j T} (1 - \mu_j T) - e^{\mu_j T} \csc(H\pi)}{4\mu_j^2 e^{\mu_j T}} \sim -\frac{T}{2\mu_j}, \quad j \rightarrow \infty,$$

$$\lim_{j \rightarrow \infty} \mu_j \frac{\partial \Psi_T^H(a, \mu_j)}{\partial a} \Big|_{a=0} = -\frac{T}{2},$$

and the lemma is proved.  $\square$

**Lemma A.2.** *For every  $\theta \in \Theta$  and  $H \in [1/2, 1)$*

$$\lim_{j \rightarrow \infty} \mu_j^3(\theta) \text{Var} \left( \int_0^T Q_j^2(s) d\mathbf{w}_H(s) \right) = \frac{T}{2}.$$

*Proof.* Note that

$$\mathbf{V} := \text{Var} \left( \int_0^T Q_j^2(s) d\mathbf{w}_H(s) \right) = \left[ \frac{\partial^2 \Psi_T^H(a, \mu_j)}{\partial a^2} - \left( \frac{\partial \Psi_T^H(a, \mu_j)}{\partial a} \right)^2 \right]_{a=0}, \quad (\text{A.3})$$

with  $\Psi_T^H$  from (A.1). Direct evaluation of the right hand side of (A.3) (for example, using Mathematica computer algebra system) gives

$$\begin{aligned} \mathbf{V} = & \frac{1}{8\mu_j^4 e^{2T\mu_j}} \left( 2 - 8e^{\mu_j T} (1 + \mu_j T) + 2e^{2\mu_j T} (-5 + 2\mu_j T) \right. \\ & + \pi\mu_j T \csc(\pi H) \left[ -2e^{\mu_j T} \mu_j T I_{1-H} \left( \frac{\mu_j T}{2} \right) I_{H-1} \left( \frac{\mu_j T}{2} \right) \right. \\ & + I_{-H} \left( \frac{\mu_j T}{2} \right) \left\{ 4(-1 + e^{\mu_j T} (1 + \mu_j T)) I_{H-1} \left( \frac{\mu_j T}{2} \right) \right. \\ & \left. \left. - 2e^{\mu_j T} \mu_j T I_H \left( \frac{\mu_j T}{2} \right) + \pi\mu_j T I_{H-1}^2 \left( \frac{\mu_j T}{2} \right) I_{-H} \left( \frac{\mu_j T}{2} \right) \csc(H\pi) \right\} \right], \end{aligned}$$

where  $\csc(x) = 1/\sin(x)$  and  $I_p$  is the modified Bessel function of the first kind and order  $p$ .

Using (A.2), we conclude that

$$\begin{aligned} \lim_{j \rightarrow \infty} \mu_j^3(\theta) \mathbf{V} &= \lim_{j \rightarrow \infty} \mu_j^3 \left( \frac{-10 + 4 \csc(H\pi) + \csc^2(H\pi)}{8\mu_j^4} + \frac{1}{4\mu_j^4 e^{2\mu_j T}} \right. \\ &\quad \left. - \frac{\csc(H\pi) + 2 + 2\mu_j T}{2\mu_j^4 e^{\mu_j T}} + \frac{T}{2\mu_j^3} \right) \\ &= \frac{T}{2} \end{aligned}$$

and complete the proof of the lemma.  $\square$

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