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Bank insolvency risk and time-varying Z-score measures

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Abstract

We compare the different existing approaches to the construction of timevarying Z-score measures, plus an additional alternative one, using a panel of banks for the G20 group of countries covering the period 1992–2009. We examine which ways of estimating the moments used in these different approaches best fit the data, using a simple root mean squared error criterion. Our results are supportive of our alternative time-varying Z-score measure: it uses mean and standard deviation estimates of the return on assets calculated over full samples combined with current values of the capital-asset ratio, and is thus straightforward to implement.

JEL classification: G21

Key words: insolvency risk, Z-score, time-varying, mean squared error

January 8, 2013

1. Introduction

The Z-score is a risk measure commonly used in the empirical banking literature to reflect a bank's probability of insolvency. It is generally attributed to Boyd and Graham (1986), Hannan and Hanweck (1988) and Boyd et al. (1993), and plays an important role in the assessment of both individual bank risk as well as overall financial stability. Its use in cross-sectional studies has become widespread due to its simplicity and the fact that it can be constructed using only accounting information; in contrast to market-based risk measures, it is also applicable to the substantial number of unlisted financial institutions. Starting with work by Boyd et al. (2006), Hesse and Čihák (2007) and Yeyati and Micco (2007), it is now also increasingly being implemented as a time-varying measure in panel studies. Despite this growing popularity, there appears so far to be a certain lack of consensus on what the best way of constructing such time-varying Z-score measures might be.

In this paper, we discuss the time-varying Z-score measures in use so far and propose a related alternative one; we then compare these measures using data on commercial, cooperative and savings banks for the G20 group of countries covering the period 1992–2009. We further examine which of the various ways of estimating the moments used in the different approaches to the computation of these time-varying Z-score measures best fit the data, using a simple root mean squared error criterion. Our results overall support the use of the alternative time-varying Z-score measure we propose, which uses mean and standard deviation estimates of the return on assets that are calculated over the full sample and combines these with current values of the capital-asset ratio. This approach to the construction of time-varying Z-score measures is furthermore straightforward to implement and does not drop initial observations, an inherent problem with rolling moments approaches.

Section 2 discusses the different approaches to the construction of timevarying Z-score measures considered, Section 3 evaluates them for a panel of commercial, cooperative and savings banks for the G20 group of countries, and Section 4 concludes the paper.

2. Different approaches to time-varying Z-score measures

To discuss the different approaches to the construction of time-varying Z-score measures currently in use in the literature, let us first recapitulate the established rationale for the use of Z-score measures more generally. As is common in the literature, we define bank insolvency as a state where $(car + roa) \leq 0$, with *car* being the bank's capital-asset ratio and *roa* its return on assets. Then, if *roa* is a random variable with mean μ_{roa} and finite variance σ_{roa}^2 , Hannan and Hanweck (1988) and Boyd et al. (1993) point out that the Bienaymé-Chebyshev inequality allows us to state an upper bound of the probability of insolvency as

$$p(roa \le -car) \le Z^{-2} \tag{1}$$

where
$$Z \equiv \frac{car + \mu_{roa}}{\sigma_{roa}} > 0$$
 (2)

The Z-score Z defined in Equation (2) is thus inversely related to an upper bound of the probability of insolvency $p(roa \leq -car)$, even for the weakest of distributional assumptions.¹

¹Strobel (2011) derives a related upper bound of the probability of insolvency for the special case where the bank's distribution of returns is unimodal.

The implementation of Z-score measures for cross-sectional analysis is largely uncontroversial; by comparison, the construction of time-varying Zscore measures to be used in panel analysis may appear less straightforward. However, it is sufficient to reiterate that the probabilistic interpretation of Z-score measures implies that μ_{roa} and σ_{roa} are moments of the distribution of roa that are possibly time-varying and need to be estimated as $\mu_{roa,t}$ and $\sigma_{roa,t}$ for time periods t. On the other hand, (-car) is simply a "safety first" level of returns (in the spirit of Roy, 1952) delimiting the insolvency case; it could therefore plausibly be made directly time-varying using current period t values of car_t , or otherwise be represented by its mean $\mu_{car,t}$ which then would also need to be estimated for time periods t. Taking those two elements together would thus suggest time-varying Z-scores of either of the two following forms

$$Z_t = \frac{car_t + \mu_{roa,t}}{\sigma_{roa,t}} \quad \text{or} \quad Z_t = \frac{\mu_{car,t} + \mu_{roa,t}}{\sigma_{roa,t}} \tag{3}$$

The various approaches to the construction of time-varying Z-score measures currently in use in the literature, together with their first exponents, can then be classified as follows:

- Approach Z1: Boyd et al. (2006, section III.A) use moving mean and standard deviation estimates $\mu_{car,t}(n)$, $\mu_{roa,t}(n)$ and $\sigma_{roa,t}(n)$ (with window width n = 3) that are calculated for each period $t \in \{1 \dots T\}$.
- Approach Z2: Yeyati and Micco (2007) use moving mean and standard deviation estimates $\mu_{roa,t}(n)$ and $\sigma_{roa,t}(n)$ (with window width n = 3) that are calculated for each period $t \in \{1 \dots T\}$, and combine these with current period t values of car_t .

- Approach Z3: Hesse and Čihák (2007) use standard deviation estimates σ_{roa} that are calculated over the full sample $[1 \dots T]$, and combine these with current period t values of car_t and roa_t .
- Approach Z4: Boyd et al. (2006, section III.B) use what might be called "instantaneous" standard deviation estimates $\sigma_{roa,t}^{inst} = |roa_t - \mu_{roa}|$, where the mean estimate μ_{roa} is calculated over the full sample $[1 \dots T]$, and combine these with current period t values of car_t and roa_t .

Clearly, all these approaches are consistent with the probabilistic interpretation of Z-score measures discussed above. A further approach that is also consistent with this interpretation, but does not seem to have been used in the literature so far, would be

• Approach Z5: One could use mean and standard deviation estimates μ_{roa} and σ_{roa} that are calculated over the full sample $[1 \dots T]$, and combine these with current period t values of car_t .

Which of these five approaches to constructing time-varying Z-score measures is then most appropriate in a given context is an inherently empirical question and will depend on the data under consideration; we will examine this issue further in the following section.

3. Evaluating different time-varying Z-score measures for the G20 countries

We now examine how the different ways of computing time-varying Zscore measures discussed in Section 2 compare when taken to the data. To this end, we examine a dataset of commercial, cooperative and savings banks for the G20 group of countries extracted from BvD Bankscope, covering the period 1992–2009. We manually clean for obvious outliers/erroneous data, and retain for each bank j the longest contiguous run of observations T_j while imposing a minimum of five observations; we end up with data for 14658 banks with an average of 10.4 years of observations. Table 1 gives a breakdown of the number of banks in our sample by country and bank type.

In Table 2 we give descriptive statistics of the five different time-varying Z-score measures. We observe three distinct clusters amongst these measures, consistently for all banks, for all three types of bank and for all G20 countries. Z1 and Z2 are very similar, with average means and standard deviations of these measures, as calculated per bank, in a medium range with an average coefficient of variation of 1.1. Z3 and Z5 are also very close to each other, with average means and standard deviations in a generally lower range and an also rather low average coefficient of variation of 0.1. Z4, on the other hand, gives results that are very different from the other measures, with average means and standard deviations in a much higher range, even running to four digits in the case of the US, and a much larger average coefficient of variation of 2.2. These three clusters are confirmed when examining the average correlation coefficients of these different measures, as presented in Table 3. In order to better understand the markedly different behavior of Z4, we examine some descriptive statistics of the components of these timevarying Z-score measures in Table 4. We note that the "instantaneous" standard deviation estimates $\sigma_{roa,t}^{inst}$ used in the calculation of Z4 can obtain extremely small values compared to the other measures, particularly for the case of the US, leading consequently to potentially very large values of Zscores and quite substantial volatility in these measures, as observed in Table 2.

[Insert Tables 1–4]

We then take our investigation further by examining which of the various mean and standard deviation estimates that are used to compute the timevarying Z-score measures Z1–Z5 best fit the data, in the sense of producing minimum one-period-ahead forecast errors. Given the rather short time dimension of our panel, we opt for a simple root mean squared error (RMSE) criterion to evaluate which of the proposed approaches to calculate the means $\mu_{car,t}$ and $\mu_{roa,t}$ best fit the data. For this we explore which of the different estimates² $\mu_{car,t}^{est} \in {\mu_{car,t}(2), \mu_{car,t}(3), \mu_{car,t}(4), \mu_{car,t}(5), \mu_{car}, car_t}$ and $\mu_{roa,t}^{est} \in {\mu_{roa,t}(2), \mu_{roa,t}(3), \mu_{roa,t}(4), \mu_{roa,t}(5), \mu_{roa}, roa_t}$, respectively, minimize the (weighted) average RMSE of the N banks j given by

$$\sum_{j=1}^{N} \frac{T_j}{\sum_{j=1}^{N} T_j} \sqrt{\frac{1}{T_j} \sum_{t=1}^{T_j} \left(car_{j,t} - \mu_{car,j,t-1}^{est} \right)^2}$$
(4)

$$\sum_{j=1}^{N} \frac{T_j}{\sum_{j=1}^{N} T_j} \sqrt{\frac{1}{T_j} \sum_{t=1}^{T_j} \left(roa_{j,t} - \mu_{roa,j,t-1}^{est} \right)^2}$$
(5)

Note that we also examine rolling windows of size two, four and five in addition to the one of three most used in the literature previously.

²As defined in Section 2, moving mean estimates $\mu_{car,t}(n)$, $\mu_{roa,t}(n)$ (with window width n) are calculated for each period $t \in \{1 \dots T\}$, the mean estimates μ_{car} , μ_{roa} are calculated over the full sample $[1 \dots T]$, and car_t , roa_t are current period t values.

Finding a realized volatility measure is unfortunately much less straightforward when relying on (mostly annual) accounting data than in other contexts, where higher frequency data can normally be used to calculate these. We opt to rely on the "instantaneous" standard deviation estimates $\sigma_{roa,t}^{inst} = |roa_t - \mu_{roa}|$ introduced previously to represent realized volatility, and thus investigate which of the different standard deviation estimates³ $\sigma_{roa,t}^{est} \in \{\sigma_{roa,t}(2), \sigma_{roa,t}(3), \sigma_{roa,t}(4), \sigma_{roa,t}(5), \sigma_{roa}, \sigma_{roa,t}^{inst}\}$ minimizes the (weighted) average RMSE of the N banks j given by

$$\sum_{j=1}^{N} \frac{T_j}{\sum_{j=1}^{N} T_j} \sqrt{\frac{1}{T_j} \sum_{t=1}^{T_j} \left(\sigma_{roa,j,t}^{inst} - \sigma_{roa,j,t-1}^{est}\right)^2}$$
(6)

Results are presented in Tables 5–7; all these measures are calculated for the full sample of all G20 countries, and then further broken down by bank type and individual country. We also test whether the average RMSE are significantly different from the minimum ones using two-sided paired t-tests.

We find that current values of the capital-asset ratio car_t provide the lowest average RMSE for the full sample of G20 countries, for commercial, cooperative and savings banks, and all individual countries in the sample (Table 5). We further observe that the mean of the return on assets as calculated over the full sample (μ_{roa}) provides the lowest average RMSE for the full sample of G20 countries, for commercial and cooperative banks, and all countries in the sample except China, Indonesia, Turkey and the USA

³Analogously to footnote 2, moving standard deviation estimates $\sigma_{roa,t}(n)$ (with window width n) are calculated for each period $t \in \{1 \dots T\}$, whereas the standard deviation estimate σ_{roa} is calculated over the full sample $[1 \dots T]$.

(Table 6). Current values of the return on assets roa_t , on the other hand, give, or are indistinguishable from, the lowest average RMSE for these particular four countries, and savings banks overall. Lastly, the standard deviation of the return on assets as calculated over the full sample (σ_{roa}) provides the lowest average RMSE for the full sample of G20 countries, for all three bank types, and all individual countries except Turkey and Indonesia, the latter of which however gives results that are indistinguishable across all measures (Table 7).

Taken together, these results overall support the use of the time-varying Z-score measure Z5, which uses mean and standard deviation estimates of the return on assets that are calculated over the full sample and combines these with current values of the capital-asset ratio. For specific subsamples, such as savings banks, or banks in China, Indonesia or the USA, the time-varying Z-score measure Z3, which uses standard deviation estimates of the return on assets calculated over the full sample together with current values of the capital-asset ratio and the return on assets, might be more appropriate given our results. This is consistent with our previous observation that these two measures do in fact behave very similarly for the sample under consideration (see Tables 2 and 3). Both measures have the added advantage of allowing the construction of time-varying Z-scores that are available over the full sample; this compares favorably with the rolling moment approach used in measures Z1 and Z2, which requires some initial observations to be dropped. Lastly, they also represent very straightforward approaches to implement, making them practical yet well-founded ways of constructing time-varying Z-score measures for a wide range of empirical issues in the banking and financial stability related literature.

[Insert Tables 5–7]

4. Conclusion

We discussed and compared different approaches to the construction of time-varying Z-score measures in use in the literature, proposing a related alternative one; for this we used a panel of commercial, cooperative and savings banks for the G20 group of countries covering the period 1992–2009. We also explored which ways of estimating the moments used in the different approaches to computing these time-varying Z-score measures best fit the data, using a simple root mean squared error criterion. Our results were overall supportive of the use of the alternative time-varying Z-score measure we proposed: this measure uses mean and standard deviation estimates of the return on assets that are calculated over the full sample and combines these with current values of the capital-asset ratio, making it a very straightforward measure to implement in the assessment of individual bank insolvency risk and financial stability more generally. This measure furthermore displays a fairly low level of intertemporal volatility on a per bank level, consistently for all three types of bank and for all G20 countries, stressing the importance of avoiding the introduction of potentially "spurious" volatility in the construction of such time-varying bank insolvency risk measures more generally.

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	Commercial banks	Cooperative banks	Savings banks	Total number of banks	Average obs. per bank
Argentina	3	0	0	3	14.3
Australia	42	1	0	43	10.5
Brazil	144	2	0	146	9.5
Canada	59	20	3	82	10.2
China	13	1	0	14	9.5
France	267	100	30	397	11.2
Germany	248	1319	595	2162	11.5
India	66	6	0	72	11.7
Indonesia	16	0	0	16	9.9
Italy	182	512	64	758	11.5
Japan	152	521	1	674	11.1
Korea	28	2	2	32	9.9
Mexico	34	0	0	34	10.0
Russia	236	0	0	236	6.3
Saudi Arabia	11	0	0	11	14.2
South Africa	24	0	0	24	10.5
Turkey	39	2	0	41	7.4
UK	190	0	5	195	11.2
USA	8666	14	1038	9718	10.2
Total	10420	2500	1738	14658	10.4

Table 1. Number of banks by country and type for G20 countries and period 1992–2009 $\,$

 $\mathbf{Z1}$ $\mathbf{Z2}$ Z3 $\mathbf{Z4}$ Z5Mean St.dev. Mean Mean St.dev. Mean St.dev. Mean St.dev. St.dev. \mathbf{per} \mathbf{per} \mathbf{per} per per \mathbf{per} \mathbf{per} \mathbf{per} \mathbf{per} \mathbf{per} bank G20 banks 174192175195454.918454126454.5Mean 389981 St.dev. 342634349646 7717174118 7717Min 0.810.0760.290.0570 0.370 0.0750 1 13180 23926 13121 24796 4789 17542.10E + 074.70E + 074789 1753Max Commercial banks 123120 120 37 37 4.2 4.5381879 Mean 123St.dev. 171316170306 422013215 39519 4220Min 1.1 0.11.60.380.0570 0.370 0.0750 Max 5921 122585939 125372017 17541338834 4015505 2017 1753Cooperative banks 326 402 332 412 73 5.8540 1095 74 5.4Mean St.dev. 601 1060 61510881558.94144 112601559 Min 1.50.151.20.290.250.0440.510.60.90.1213180 23926 13121 463650 4789 Max 24796 4789249142438 249Savings banks Mean 257321261328 515.612486 27917 515.4St.dev. 483 10345001065566.1504253 1127557 566.2Min 0.810.076 1 0.30.430.0590.650.220.360.0628961 19434 9789 19654 1361 109 2.10E + 074.70E + 071360 110Max Argentina 36 Mean 57 565248 20 5.244 20530 66 294.7St.dev. 81 797363 4.452Min 7.14.67.64.92.1254.52.11.2Max 1501481371205410120975410Australia 73 74 74 76 21 5.293 126 204.8 Mean 8585 154.8124189154.7St.dev. 1511514.74.10.0570.130.550.380.140.29Min 1.41.5Max 529960 53296366245297116623Brazil 38 3538 37 14 4.2 104 183 14 3.8Mean 38 5339 57145.15011075 145.1St.dev. Min 2.50.642.70.850.110.491 1 0.180.115609 Max 21035122536287 561217588 55Canada Mean 95919592315151234314.7St.dev. 9419393 196296.8329618296.80.690.610.0003 Min 6.74.38.7 4.90.570.70.47Max 691 1662694 16901745828524787 17357China Mean 15512814812541111692594111113St.dev. 10298105249 184316249.3322236121.422121.9Min 1918370 354379 35895670 980 95Max 3434France 1531613939Mean 1431477.52484637 St.dev. 165317173336 458.210052054458.10.731.61.1 0.460.481.90.380.520.17Min 1.1Max 12263547 1274 3763 411 7316750 2877341072Germany 8.6 4435681209 8.3 Mean 45358389 607 89 690 St.dev. 1335713137917644091195917644440.280.850.0910.12 Min 2.51.40.371.50.530.3910104 23926 24796 4789 1754142438 463650 4789 Max 11051 1753India 3.8Mean 94101102120284.21111792897181147327 41176391413.8St.dev. 3.8100.822.20.25Min 5.211 6 2.54.53.4Max 5011178 1108 2652303 251141 2866 30425

Table 2. Descriptive statistics of different time-varying Z-score measures (calculated per bank), for G20 countries and period 1992—2009

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$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	Turkey										
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		33	27	32	28	13	3	60	81	13	2.4
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$											
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$\begin{array}{c c c c c c c c c c c c c c c c c c c $					-10				10	-	2.0
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$			103	90	104	28	6.1	101	414	28	57
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$											
Max 1959 2903 2005 2949 210 29 6379 23688 210 30 USA Mean 127 120 128 121 39 4 2574 5810 39 3.7 St.dev. 147 281 147 281 34 4.3 213705 478636 34 4.4 Min 0.81 0.076 1 0.3 0.43 0 0.65 0 0.36 0											
USA Mean 127 120 128 121 39 4 2574 5810 39 3.7 St.dev. 147 281 147 281 34 4.3 213705 478636 34 4.4 Min 0.81 0.076 1 0.3 0.43 0 0.65 0 0.36 0											
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $		1998	2909	2000	2949	210	29	0379	20000	210	J U
St.dev. 147 281 147 281 34 4.3 213705 478636 34 4.4 Min 0.81 0.076 1 0.3 0.43 0 0.65 0 0.36 0		107	100	100	101	90	4	0574	F010	90	07
Min 0.81 0.076 1 0.3 0.43 0 0.65 0 0.36 0											
Max 5921 12258 5939 12537 576 149 2.10E+07 4.70E+07 576 150											
	Max	5921	12258	5939	12537	576	149	2.10E+07	4.70E + 07	576	150

The time-varying Z-score measure Z1 uses three-period moving mean and standard deviation estimates $\mu_{car,t}(3)$, $\mu_{roa,t}(3)$ and $\sigma_{roa,t}(3)$; Z2 uses moving mean and standard deviation estimates $\mu_{roa,t}(3)$ and $\sigma_{roa,t}(3)$, combined with period t values of car_t ; Z3 uses standard deviation estimates σ_{roa} calculated over the full sample, combined with period t values of car_t and roa_t ; Z4 uses "instantaneous" standard deviation estimates $\sigma_{roa,t} = |roa_t - \mu_{roa}|$, where the mean estimate μ_{roa} is calculated over the full sample, combined with period t values of car_t and roa_t ; Z5 uses mean and standard deviation estimates σ_{roa} is calculated over the full sample, combined with period t values of car_t and roa_t ; Z5 uses mean and standard deviation estimates μ_{roa} and σ_{roa} calculated over the full sample, combined with period t values of car_t .

Table 3. Average correlation coefficients of different time-varying Z-score measures (calculated per bank), for G20 countries and period 1992-2009

	Z1	Z2	Z3	Z4
G20 banks			20	
Z2	0.98			
Z3	0.12	0.17		
Z4	0.12	0.13	0.11	
Z5	0.068	0.11	0.91	0.092
Commercial banks	0.008	0.11	0.91	0.092
Z2	0.08			
Z2 Z3	0.98 0.12	0.17		
Z3 Z4	0.12	0.17	0.11	
Z4 Z5		$0.13 \\ 0.11$	0.11 0.9	0.090
	0.054	0.11	0.9	0.089
Cooperative banks	0.00	T	T	
Z2	0.99	0.15		
Z3	0.12	0.15	0.14	
Z4	0.15	0.15	0.14	0.11
Z5	0.11	0.14	0.93	0.11
Savings banks				1
Z2	0.99			
Z3	0.13	0.17		
Z4	0.1	0.098	0.095	
Z5	0.093	0.13	0.93	0.081
Argentina				
Z2	0.93			
Z3	-0.11	0.093		
Z4	0.14	0.27	0.52	
Z5	-0.2	0.0017	0.95	0.41
Australia				
Z2	0.96			
Z3	0.18	0.26		
Z4	0.2	0.22	0.16	
Z5	0.13	0.22	0.9	0.14
Brazil				1
Z2	0.92			
Z3	0.068	0.21		
Z4	0.18	0.2	0.14	
Z5	0.041	0.2	0.9	0.13
Canada				
Z2	0.98			
Z3	0.057	0.13		
Z4	0.15	0.14	0.029	
Z5	0.0078	0.079	0.91	0.018
China	0.0010	0.010	0.01	0.010
Z2	0.97			
Z3	0.13	0.22		
Z3 Z4	0.13	0.22	0.058	
Z4 Z5	0.16	0.1	0.038	0.042
France	0.10	0.20	0.91	0.042
Z2	0.97			
Z2 Z3	0.97	0.18		
Z3 Z4	0.11 0.11	0.18	0.12	
				0.19
Z5	0.09	0.17	0.96	0.12
Germany	0.00	I	I	
Z2	0.99			
Z3	0.1	0.12	0.11	
Z4	0.13	0.13	0.11	
Z5	0.096	0.12	0.93	0.091

India				
Z2	0.96			
Z3	0.22	0.22		
			0.11	
Z4	0.094	0.1	0.11	0.1
Z5	0.19	0.2	0.94	0.1
Indonesia				
Z2	0.99			
Z3	0.23	0.29		
Z4	0.02	0.026	0.32	
Z5	0.21	0.27	0.98	0.3
Italy				
Z2	0.98			
Z3	0.0069	0.053		
Z4	0.096	0.1	0.13	
Z5	-0.0086	0.037	0.13	0.11
	-0.0080	0.037	0.94	0.11
Japan	0.00	[]		
Z2	0.98			
Z3	0.34	0.37		
Z4	0.17	0.18	0.27	
Z5	0.31	0.35	0.9	0.19
Korea				-
Z2	0.99			
Z3	0.47	0.5		
Z4	0.36	0.35	0.23	
Z4 Z5	0.30	$0.35 \\ 0.49$	0.25	0.26
	0.40	0.49	0.89	0.20
Mexico		[]		
Z2	0.94			
Z3	0.3	0.39		
Z4	0.31	0.35	0.3	
Z5	0.26	0.36	0.96	0.28
Russia				
Z2	0.91			
Z3	0.11	0.28		
Z4	0.27	0.32	0.16	
Z5	0.047	0.23	0.92	0.15
Saudi Arabia	0.047	0.25	0.32	0.15
	1			
Z2	1	0.000		
Z3	-0.076	-0.062		
Z4	-0.0052	-0.0099	-0.088	
Z5	-0.073	-0.051	0.89	-0.097
South Africa				
Z2	0.97			
Z3	-0.086	-0.022		
Z4	0.24	0.24	0.064	
Z5	-0.11	-0.046	0.95	0.057
Turkey	V.11	0.010	0.00	0.001
Z2	0.93			
		0.01		
Z3	0.092	0.21	0.82	
Z4	0.019	0.076	0.36	0.00
Z5	0.045	0.17	0.9	0.33
United Kingdom				
Z2	0.93			
Z3	0.019	0.13		
Z4	0.063	0.082	0.098	
Z5	0	0.11	0.94	0.11
USA	~	··**		~.**
Z2	0.99			
		0.16		
72				
Z3	0.12		0.000	
Z3 Z4 Z5	$0.12 \\ 0.12 \\ 0.05$	$0.10 \\ 0.12 \\ 0.093$	$\begin{array}{c} 0.096 \\ 0.9 \end{array}$	0.078

Definitions of time-varying Z-scores Z1-Z5: see Table 2.

Table 4. Descriptive statistics of components of time-varying Z-scores, for G20 countries and period $1992\mathaccurrencement 2009$

	car_t	roa_t	σ_{roa}	$\sigma_{roa,t}^{inst}$	$\sigma_{roa,t}(3)$
C20 hardes		$r \circ \mathfrak{A}_t$	° roa	o _{roa,t}	
G20 banks Mean	10.84	0.78	0.72	0.54	0.36
St. dev. Min	$\begin{array}{c}10.11\\0\end{array}$	$2.76 \\ -165.83$	$1.97 \\ 0.002$	2.03 1.2e-07	$1.38 \\ 0.000049$
Max	100	193.57	72.47	176.21	91.97
Commercial b		0.05	0.00	0.05	0.49
Mean	12.06	0.95	0.86	0.65	0.43
St. dev.	11.04	3.03	2.12	2.21	1.49
Min	0	-165.83	0.0038	1.1e-06	0.0003
Max	100	193.57	66.91	176.21	91.36
Cooperative b		0.00			0.10
Mean	7.32	0.36	0.27	0.2	0.16
St. dev.	4.32	0.58	0.34	0.39	0.3
Min	0	-17.02	0.002	3.0e-06	0.000077
Max	100	17.54	8.55	16.48	10.33
Savings banks				,	
Mean	9.25	0.44	0.58	0.44	0.27
St. dev.	9.18	2.95	2.27	2.3	1.61
Min	0	-104.43	0.0044	1.2e-07	0.000049
Max	100	153.16	72.47	135.09	91.97
Argentina	-				
Mean	36.27	-0.73	3.39	2.71	2.09
St. dev.	31.58	3.98	1.77	2.71	2.08
Min	3.77	-14.43	1.29	0.25	0.08
Max	95.61	6.4	5.56	12.35	7.68
Australia					
Mean	9.23	0.47	1.84	1.19	0.62
St. dev.	11.84	7.33	6.65	6.8	3.74
Min	0.62	-151.2	0.072	0.0013	0.0027
Max	100	15.69	45.43	136.24	70.25
Brazil				1	
Mean	20.17	2.12	3.23	2.36	2.02
St. dev.	18.39	5.81	4.03	4.59	3.98
Min	0.3	-97.18	0.07	0.00024	0.013
Max	99.52	58.15	36.23	92.25	57.28
Canada	00101	00110	00120	02120	01.20
Mean	11.38	1.05	2.18	1.43	0.97
St. dev.	16.03	8.06	7.28	7.47	5.7
Min	0.45	-14.26	0.04	0.00051	0.0008
Max	100	193.57	55.72	176.21	91.36
China	100	100.01	00.12	110.21	01.00
Mean	20.51	0.71	0.55	0.44	0.25
St. dev.	25.92	0.93	0.49	0.6	0.23 0.37
Min	1.47	-4.19	0.45	0.0014	0.0073
Max	99.46	4.28	1.6	4.36	2.59
France	00.10	1.20	1.0	1.00	2.00
Mean	10.39	0.64	0.94	0.7	0.47
St. dev.	12.82	3.65	2.11	2.21	1.48
Min	0	-79.92	0.03	0.000079	0.0004
Max	99.75	66.18	31.77	58.18	33.91
Germany	33.10	00.10	01.11	50.10	33.31
Germany Mean	6.11	0.28	0.26	0.19	0.14
St. dev.	6.5	1.47	1.19	$0.19 \\ 1.27$	$0.14 \\ 0.96$
	0.5	-111.32	0.002	1.27 3.0e-06	0.96 0.000049
Min Max	100	-111.32 100.98	35.23	3.0e-06 96.67	0.000049 55.08
	100	100.98	əə.2ə	90.07	55.08
India	7 59	0.84	0.69	0.47	0.99
Mean	7.53	0.84	0.62	0.47	0.33
St. dev.	7.9	1.03	0.63	0.75	0.53
Min	0.06	-6.65	0.06	0.0011	0.00073
Max	98.55	9.38	3.59	6.93	4

Indonesia					
Mean	16.72	1.88	1.07	0.82	0.54
St. dev.	11.3	2.03	1.13	1.32	0.7
Min	3.87	-5.16	0.15	0.0012	0.04
Max	69.28	16.98	4.67	13.02	5.73
Italv	11			11	
Mean	11.94	0.73	0.52	0.39	0.26
St. dev.	6.75	1.35	1.01	1.07	0.63
Min	0.6	-42.34	0.05	0.000052	0.00037
Max	100	54.27	19.77	44.21	24.88
Japan					
Mean	5.72	0.03	0.5	0.37	0.3
St. dev.	4.34	1.17	0.87	0.93	0.6
Min	0	-59.84	0.03	0.000018	0.00014
Max	100	21.23	19.76	53.27	25.26
Korea	200	21120	10110	00.21	20120
Mean	5.09	0.14	1.09	0.77	0.6
St. dev.	2	1.48	0.82	1.12	0.86
Min	0.31	-9.2	0.11	0.00046	0.0036
Max	17.36	5.15	3.39	8.03	4.06
Mexico	1	0.10	0.00	0.00	2.00
Mean	18.46	0.86	2.35	1.89	1.41
St. dev.	17.43	4.41	2.3	2.7	1.89
Min	0.44	-24.59	0.23	0.01	0.01
Max	99.24	26.24	8.7	19.21	11.9
Russia					
Mean	20.92	1.59	1.43	1.11	0.91
St. dev.	14.49	2.58	1.65	1.89	1.36
Min	0.81	-31.12	0.06	0.000046	0.0013
Max	97.87	29.83	12.98	27.88	16.15
Saudi Arabia	11			II	
Mean	11.26	1.92	1.09	0.76	0.49
St. dev.	2.99	1.44	0.85	1.16	0.84
Min	2.87	-6.28	0.13	0.01	0.01
Max	27.05	12.55	3.18	10.64	4.99
South Africa	•				
Mean	16.71	1.64	1.82	1.31	1.19
St. dev.	15.92	3.67	2.39	2.71	2.26
Min	0.85	-26.82	0.11	0.002	0.0022
Max	81.05	25.61	8.64	26.01	14.61
Turkey				· ·	
Mean	18.56	1.79	3.23	2.37	2.01
St. dev.	17.93	5.62	3.75	4.35	3.72
Min	1.87	-55.95	0.22	0.0011	0.03
Max	91.64	23.45	20.6	50.22	29.03
United Kingdo	om				
Mean	15.76	1.19	1.68	1.2	0.81
St. dev.	18.34	4.96	3.42	3.62	2.74
Min	0.14	-67.19	0.01	0.00015	0.00055
Max	100	80.99	29.24	74.58	37.38
USA					
Mean	11.88	0.92	0.75	0.57	0.36
St. dev.	9.93	2.84	1.94	2	1.28
Min	0	-165.83	0.01	1.2e-07	0.0003
Max	100	188.44	72.47	147.28	91.97

 car_t is banks' capital-asset ratio and roa_t their return on assets; σ_{roa} is the standard deviation estimate of roa over full samples; $\sigma_{roa,t}^{inst} = |roa_t - \mu_{roa}|$ is the "instantaneous" standard deviation estimate, where the mean estimate μ_{roa} is calculated over the full sample; $\sigma_{roa,t}(3)$ is the three-period moving standard deviation estimate of roa.

	Number	mber Average Root Mean Squared Error (RMSE)						
	of banks	$\mu_{car,t}(2)$	$\mu_{car,t}(3)$	$\mu_{car,t}(4)$	$\mu_{car,t}(5)$	μ_{car}	car_t	
G20 banks	14658	1.547	1.680	1.810	1.992	1.803	1.424	
Commercial banks	10420	1.874	2.022	2.168	2.388	2.131	1.738	
Cooperative banks	2500	0.654	0.738	0.816	0.895	0.873	0.570	
Savings banks	1738	1.131	1.253	1.371	1.511	1.427	1.018	
Argentina	3	12.796	13.125	13.018	13.039	12.335	11.105	
Australia	43	4.547	4.855	5.078	5.290	4.994	4.050	
Brazil	146	7.045	7.569	8.175	8.553	7.517	6.636	
Canada	82	4.642	5.120	5.533	5.831	5.448	4.088	
China	14	5.743	6.803	7.541	7.980	7.168	4.557	
France	397	2.740	2.958	3.141	3.326	3.087	2.515	
Germany	2162	0.763	0.849	0.929	1.010	1.019	0.689	
India	72	1.409	1.577	1.856	2.121	1.901	1.253	
Indonesia	16	3.593	3.878	4.347	4.806	4.371	3.339	
Italy	758	1.240	1.393	1.529	1.682	1.546	1.067	
Japan	674	0.726	0.791	0.865	0.931	0.835	0.671	
Korea	32	1.172	1.294	1.399	1.499	1.366	1.075	
Mexico	34	8.367	9.101	9.841	10.353	9.294	7.566	
Russia	236	$5.468^{ m NSD}$	6.021	6.672	7.196	6.092	5.175	
Saudi Arabia	11	1.744	1.821	1.843	1.898	1.821	1.577	
South Africa	24	5.004	5.542	5.810	6.202	5.529	4.332	
Turkey	41	$5.007^{ m NSD}$	5.547	6.082	6.331	$5.309^{ m NSD}$	4.927	
United Kingdom	195	4.505	5.004	5.471	5.931	5.339	4.284	
USA	9718	1.530	1.648	1.765	1.971	1.756	1.416	

Table 5. Average Root Mean Squared Error for various estimators of time-varying means of the capital-asset ratio, for G20 subsamples and period 1992--2009

Minimum average RMSE are highlighted in grey; the superscript NSD marks values not significantly different at the 5% level from the minimum ones using a two-sided paired t-test. μ_{car} is the mean of the capital-asset ratio *car* calculated over full samples; *car_t* is the period *t* value of *car*; $\mu_{car,t}(2)$ is the two-period moving average of *car*, analogously for $\mu_{car,t}(3)$ etc.

	Number	mber Average Root Mean Squared Error (RMSE)						
	of banks	$\mu_{roa,t}(2)$	$\mu_{roa,t}(3)$	$\mu_{roa,t}(4)$	$\mu_{roa,t}(5)$	μ_{roa}	roa_t	
G20 banks	14658	0.709	0.732	0.753	0.775	0.676	0.700	
Commercial banks	10420	0.860	0.888	0.913	0.940	0.820	0.846	
Cooperative banks	2500	0.293	0.289	0.290	0.293	0.253	0.307	
Savings banks	1738	0.525	0.553	0.581	0.602	0.534	0.507	
Argentina	3	3.897	3.774	3.822	3.968	3.400	4.060	
Australia	43	2.293	2.310	2.322	2.338	2.108	2.326	
Brazil	146	3.654	3.544	3.486	3.486	3.034	4.011	
Canada	82	2.986	2.890	2.887	2.890	2.556	3.399	
China	14	0.617	0.661	0.687	0.714	0.611	0.555	
France	397	0.909	0.909	0.935	0.971	0.850	0.926	
Germany	2162	0.255	0.258	0.260	0.267	0.230	0.269	
India	72	0.603	0.634	0.643	0.654	0.555	0.591	
Indonesia	16	$0.773^{ m NSD}$	0.766	0.819	0.950	0.869	$0.788^{ m NSD}$	
Italy	758	0.473	0.480	0.492	0.514	0.452	0.471	
Japan	674	0.532	0.518	0.518	0.516	0.438	0.563	
Korea	32	1.221	1.234	1.196	1.162	0.928	1.272	
Mexico	34	2.434^{NSD}	2.454	2.579	2.698	2.333	2.508	
Russia	236	1.704	1.721	1.755	1.836	1.552	1.746	
Saudi Arabia	11	1.365	1.412	1.405	1.388	1.279	$1.297^{ m NSD}$	
South Africa	24	2.580	2.560	2.500	2.471	2.167	2.847	
Turkey	41	1.443	1.928	2.286	2.407	1.919	1.512^{NSD}	
United Kingdom	195	1.660	1.609	1.622	1.644	1.460	1.830	
USA	9718	0.747	0.784	0.813	0.840	0.734	0.711	

Table 6. Average Root Mean Squared Error for various estimators of time-varying means of the return on assets, for G20 subsamples and period 1992--2009

Minimum average RMSE are highlighted in grey; the superscript NSD marks values not significantly different at the 5% level from the minimum ones using a two-sided paired t-test. μ_{roa} is the mean of the return on assets *roa* calculated over full samples; roa_t is the period t value of roa; $\mu_{roa,t}(2)$ is the two-period moving average of *roa*, analogously for $\mu_{roa,t}(3)$ etc.

	Number	Average Root Mean Squared Error (RMSE)						
	of banks	$\sigma_{roa,t}(2)$	$\sigma_{roa,t}(3)$	$\sigma_{roa,t}(4)$	$\sigma_{roa,t}(5)$	$\sigma_{\scriptscriptstyle roa}$	$\sigma_{roa,t}^{inst}$	
G20 banks	14658	0.583	0.553	0.537	0.538	0.476	0.541	
Commercial banks	10420	0.704	0.668	0.650	0.652	0.579	0.656	
Cooperative banks	2500	0.227	0.220	0.217	0.216	0.191	0.225	
Savings banks	1738	0.457	0.428	0.411	0.403	0.351	0.398	
Argentina	3	2.795	2.255	2.209	2.103	1.861	2.870	
Australia	43	2.026	1.999	1.997	2.004	1.659	1.907	
Brazil	146	2.744	2.705	2.709	2.695	2.233	2.928	
Canada	82	2.564	2.543	2.544	2.537	2.000	2.675	
China	14	0.490	0.457	0.440	0.425	0.358	0.384^{NSD}	
France	397	0.751	0.741	0.711	0.706	0.636	0.735	
Germany	2162	0.212	0.206	0.204	0.205	0.180	0.207	
India	72	0.486	0.450	0.444	0.455	0.406	0.447	
Indonesia	16	$0.708^{ m NSD}$	$0.682^{\rm NSD}$	$0.672^{\rm NSD}$	0.699^{NSD}	$0.663^{ m NSD}$	0.643	
Italy	758	0.384	0.361	0.348	0.354	0.325	0.353	
Japan	674	0.396	0.385	0.388	0.389	0.335	0.391	
Korea	32	0.809	0.848	0.868	0.895	0.768	1.085	
Mexico	34	1.861	1.757	1.692	1.661	1.451	1.716	
Russia	236	1.365	1.265	1.268	1.281	1.128	1.321	
Saudi Arabia	11	1.150	1.152	1.127	1.116	0.974	1.201	
South Africa	24	1.964	1.867	1.899	1.874	1.575	2.126	
Turkey	41	1.740	1.780	2.028	2.231	2.047	1.195	
United Kingdom	195	1.387	1.361	1.356	1.392	1.237	1.403	
USA	9718	0.618	0.580	0.558	0.557	0.497	0.557	

Table 7. Average Root Mean Squared Error for various estimators of time-varying standard deviations of the return on assets, for G20 subsamples and period 1992--2009

Minimum average RMSE are highlighted in grey; the superscript NSD marks values not significantly different at the 5% level from the minimum ones using a two-sided paired t-test. σ_{roa} is the standard deviation of the return on assets *roa* calculated over full samples; $\sigma_{roa,t}^{inst} = |roa_t - \mu_{roa}|$ is the "instantaneous" standard deviation, where the mean μ_{roa} is calculated over full samples of *roa*; $\sigma_{roa,t}(2)$ is the two-period moving standard deviation of *roa*, analogously for $\sigma_{roa,t}(3)$ etc.