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# Breaking Substitution Ciphers Using a Relaxation Algorithm 

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Substitution ciphers are codes in which each letter of the alphabet has one fixed substitute, and the word divisions do not change. In this paper the problem of breaking substitution ciphers is represented as a probabilistic labeling problem. Every code letter is assigned probabilities of representing plaintext letters. These probabilities are updated in parallel for all code letters, using joint letter probabilities. Iterating the updating scheme results in improved estimates that finally lead to breaking the cipher. The method is applied successfully to two examples.

Key Words and Phrases: cryptography, substitution ciphers, probabilistic classification, relaxation

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## 1. Introduction

Let $\sum$ be an alphabet consisting of letters and a space symbol. A substitution cipher is a permutation in which every letter of the alphabet in the message $M=m_{1} \ldots m_{l}$ except for the space symbol is replaced consistently by another letter to give the coded message $C=c_{1} \ldots c_{l}$. A key for a substitution cipher is a transformation $K: \sum \rightarrow$ $\sum$ such that $M=K\left(c_{1}\right) \ldots K\left(c_{l}\right)$. Manual methods for breaking such codes are well-known [1, 7], and are based primarily on the relative frequencies of single letters, and on certain combinations of letters in certain positions.

[^0]In this paper, a completely automatic method for breaking substitution ciphers is presented, based on relaxation methods. Relaxation algorithms have recently been introduced in image processing [4, 6]. They are iterative parallel classification algorithms, where every element in a graph structure tries to estimate its class membership probabilities based on those of its neighbors. The process is iterated until a satisfactory classification is achieved. A new formulation of relaxation [4,5], based on probability theory, paves the way for more general applications of relaxation. The use of relaxation for domains other than image classification is demonstrated in this paper.

Section 2 describes the relaxation approach to probabilistic graph labeling; Section 4 discusses the application of this approach to substitution ciphers; and Section 5 summarizes the results obtained.

## 2. Probability Updating

The problem of probabilistic graph labeling is as follows:

Let $G=(V, E)$ be a graph with $V=\left\{v_{1}, \ldots, v_{n}\right\}$ the set of nodes and $E$ the set of arcs, and let $\Lambda=$ $\left(\lambda_{1}, \ldots, \lambda_{L}\right)$ be a set of labels. With every node $v_{i} \in V$ a random variable $l_{i}$ is associated, specifying the probabilities of the possible labels for that node. Initially, based on some measurements, a probability distribution $P_{i}^{(0)}$ : $\Lambda \rightarrow[0,1]$ is estimated for every random variable $l_{i}$. In this section the updating of these distributions is discussed, based on the distributions at neighboring nodes, and on statistical relations among the random variables $l_{i}$.

Given a graph $G=(V, E)$, a set of labels $\Lambda$, and an estimate for a discrete probability distribution $P_{i}: \Lambda \rightarrow$ $[0,1]$ for each random variable $l_{i}$, new estimated probability distributions for $l_{i}$ are to be computed.

We can regard the probability estimate vectors $P_{i}$ as events, i.e., we can think of them as being chosen from. a space of possible vectors. We shall now consider various prior and conditional probabilities involving these $P_{i}$ events and the outcomes of the random variables $l_{i}$. In particular, we shall consider probabilities of the form
(1) $\operatorname{Prob}\left(l_{i}=\alpha \mid P_{i}\right)$; this is just the probability that node $v_{i}$ has label $\alpha$, given that its estimated probability distribution of labels is $P_{i}$. We denote this probability by $P_{i}\left(l_{i}=\alpha\right)$.
(2) $\operatorname{Prob}\left(P_{i} \mid l_{i}=\alpha\right)$; this is the probability that the distribution estimate for node $v_{i}$ is $P_{i}$, given that the true label of $v_{i}$ is $\alpha$.

Evidently we have $P_{i}\left(l_{i}=\alpha\right)=\operatorname{Prob}\left(P_{i} \mid l_{i}=\alpha\right) \cdot \operatorname{Prob}$ $\left(l_{i}=\alpha\right) / \operatorname{Prob}\left(P_{i}\right)$. (The problem of actually calculating $\operatorname{Prob}\left(P_{i}\right)$ will not be considered yet.)

Let $v_{i}, v_{j}$, and $v_{k}$ be three nodes such that $v_{j}$ and $v_{k}$ are neighbors of $v_{i}$. We can consider the joint events ( $l_{i}=$ $\alpha, l_{j}=\beta, l_{k}=\gamma$ ) and ( $P_{i}, P_{j}, P_{k}$ ), and write

$$
\begin{align*}
& \operatorname{Prob}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma \mid P_{i}, P_{j}, P_{k}\right) \\
& \equiv P_{i j k}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right) \\
& \quad \operatorname{Prob}\left(P_{i}, P_{j}, P_{k} \mid l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)  \tag{1}\\
& =\frac{\operatorname{Prob}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)}{\operatorname{Prob}\left(P_{i}, P_{j}, P_{k}\right)}
\end{align*}
$$

We now assume that

$$
\begin{aligned}
& \operatorname{Prob}\left(P_{i}, P_{j}, P_{k} \mid l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right) \\
& =\operatorname{Prob}\left(P_{i} \mid l_{i}=\alpha\right) \operatorname{Prob}\left(P_{j} \mid l_{j}=\beta\right) \operatorname{Prob}\left(P_{k} \mid l_{k}=\gamma\right) \\
& =\frac{\operatorname{Prob}\left(l_{i}=\alpha \mid P_{i}\right) \operatorname{Prob}\left(P_{i}\right)}{\operatorname{Prob}\left(l_{i}=\alpha\right)} \\
& \cdot \frac{\operatorname{Prob}\left(l_{j}=\beta \mid P_{j}\right) \operatorname{Prob}\left(P_{j}\right)}{\operatorname{Prob}\left(l_{j}=\beta\right)} \\
& \frac{\operatorname{Prob}\left(l_{k}=\gamma \mid P_{k}\right) \operatorname{Prob}\left(P_{k}\right)}{\operatorname{Prob}\left(l_{k}=\gamma\right)} \\
& \quad P_{i}\left(l_{i}=\alpha\right) P_{j}\left(l_{j}=\beta\right) P_{k}\left(l_{k}=\gamma\right) \\
& \operatorname{Prob}\left(P_{i}\right) \operatorname{Prob}\left(P_{j}\right) \operatorname{Prob}\left(P_{k}\right) \\
& \operatorname{Prob}\left(l_{i}=\alpha\right) \operatorname{Prob}\left(l_{j}=\beta\right) \operatorname{Prob}\left(l_{k}=\gamma\right)
\end{aligned}
$$

The meaning of this assumption is that the probability estimate $P_{i}$ is directly dependent only on $l_{i}$, and once $l_{i}$ is given, the probability of the estimate being $P_{i}$ is independent of $P_{j}, j \neq i$, and of $l_{j}, j \neq i$. Under this assumption, (1) becomes

$$
\begin{align*}
& P_{i j k}\left(l_{i}=\right.\left.\alpha, l_{j}=\beta, l_{k}=\gamma\right) \\
&=P_{i}\left(l_{i}=\alpha\right) P_{j}\left(l_{j}=\beta\right) P_{k}\left(l_{k}=\gamma\right) \\
& \cdot \frac{\operatorname{Prob}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)}{\operatorname{Prob}\left(l_{i}=\alpha\right) \operatorname{Prob}\left(l_{j}=\beta\right) \operatorname{Prob}\left(l_{k}=\gamma\right)} \\
& \cdot \frac{\operatorname{Prob}\left(P_{i}\right) \operatorname{Prob}\left(P_{j}\right) \operatorname{Prob}\left(P_{k}\right)}{\operatorname{Prob}\left(P_{i}, P_{j}, P_{k}\right)} \tag{2}
\end{align*}
$$

From now on, we will denote

$$
\frac{\operatorname{Prob}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)}{\operatorname{Prob}\left(l_{i}=\alpha\right) \operatorname{Prob}\left(l_{j}=\beta\right) \operatorname{Prob}\left(l_{k}=\gamma\right)}
$$

by $r_{i j k}(\alpha, \beta, \gamma)$. The quantities $r_{i j k}(\alpha, \beta, \gamma)$ are independent of the estimated distributions $P_{i}$. These quantities are computed in advance by using some model to find the required probabilities. See [3] for an approach to computing these coefficients in the absence of such models. Now

$$
P_{i j k}\left(l_{i}=\alpha\right)=\sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{i j k}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right),
$$

and

$$
\sum_{\alpha \in \Lambda} \sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{i j k}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)=1 .
$$

We thus have

$$
\begin{align*}
& P_{i j k}\left(l_{i}=\alpha\right) \\
& \sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{i j k}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right) \\
& \sum_{\lambda \in \Lambda} \sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{i j k}\left(l_{i}=\lambda, l_{j}=\beta, l_{k}=\gamma\right)  \tag{3}\\
& P_{i}\left(l_{i}=\alpha\right) \cdot \sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{j}\left(l_{j}=\beta\right) \\
&= \frac{P_{k}\left(l_{k}=\gamma\right) r_{i j k}(\alpha, \beta, \gamma)}{\sum_{\lambda \in \Lambda} P_{i}\left(l_{i}=\lambda\right) \cdot \sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{j}\left(l_{j}=\beta\right)} \\
& P_{k}\left(l_{k}=\gamma\right) r_{i j k}(\lambda, \beta, \gamma)
\end{align*}
$$

since the factor
$\frac{\operatorname{Prob}\left(P_{i}\right) \operatorname{Prob}\left(P_{j}\right) \operatorname{Prob}\left(P_{k}\right)}{\operatorname{Prob}\left(P_{i}, P_{j}, P_{k}\right)}$
cancels out.
Note that $P_{i j k}\left(l_{i}=\alpha\right)$ is dependent on the nodes $v_{i}, v_{j}$, and $v_{k}$. In a similar approach, rules for any number of nodes can be derived. A rule using two nodes, rather than three as is done here, can be found in [4].

Given all the possible pairs of neighbors, all the estimates suggested by all the pairs are to be combined into one estimate. Two methods can be used.

The first method takes the average of all the estimates given by all pairs of neighbors. In this case, the iterative updating expression is
$P_{i}^{(n+1)}\left(l_{i}=\alpha\right)=$ Average $\left\{P_{i j k}^{(n+1)}\left(l_{i}=\alpha\right)\right\}$
where $P_{i j k}^{(n+1)}\left(l_{i}=\alpha\right)$ is computed from the $P_{i}^{(n)}$ 's by expression (3), and the average is taken over all possible pairs of neighbors.

The second method, developed in [2], considers the updating as done pair after pair. Define $Q_{i j k}^{(n)}(\alpha)$ to be
$Q_{i j k}^{(n)}(\alpha)=\sum_{\beta \in \Lambda} \sum_{\gamma \in \Lambda} P_{j}^{n}\left(l_{j}=\beta\right) P_{k}^{n}\left(l_{k}=\gamma\right) r_{i j k}(\alpha, \beta, \gamma)$.
Using $Q_{i j k}^{(n)}(\alpha)$ changes (3) into
$P_{i j k}^{(n+1)}\left(l_{i}=\alpha\right)=\frac{P_{i}^{(n)}\left(l_{i}=\alpha\right) \cdot Q_{i j k}^{(n)}(\alpha)}{\sum_{\beta \in \Lambda} P_{i}^{(n)}\left(l_{i}=\beta\right) Q_{i j k}^{(n)}(\beta)}$
As was shown in [2], the estimate for the total effect of all pairs of labels is
$P_{i}^{(n+1)}\left(l_{i}=\alpha\right)=\frac{P_{i}^{(n)}\left(l_{i}=\alpha\right) \cdot \prod_{j, k} Q_{i j}^{(n)}(\alpha)}{\sum_{\beta \in \Lambda} P_{i}^{(n)}\left(l_{i}=\beta\right) \prod_{j, k} Q_{i j k}^{(n)}(\beta)}$
where $j, k$ varies over all possible pairs of neighbors for node $v_{i}$.

## 3. The Coefficients

This paper handles problems from the domain of English text, and a node in the graph will represent a letter of the coded message. The relations that will be used are the probabilities of certain combinations of letters appearing in English plaintext. Only trigrams will be used; the event ( $l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma$ ) means that nodes $v_{i}, v_{j}$, and $v_{k}$ represent the sequence $\alpha \beta \gamma$. The a priori probability $\operatorname{Prob}\left(l_{i}=\alpha, l_{j}=\beta, l_{k}=\gamma\right)$ is equal to the probability that a randomly chosen trigram from an English text is $\alpha \beta \gamma$, and $\operatorname{Prob}\left(l_{i}=\alpha\right)$ is equal to the probability that a randomly chosen letter from English text will be $\alpha$. The above probabilities can be estimated from a long sample of English text. They were computed from the novel Wuthering Heights by Emily Bronte, which contains approximately one million letters. The coefficients $r_{i j k}(\alpha, \beta, \gamma)$ used in the updating rule are then

[^1]$r_{i j k}(\alpha, \beta, \gamma)=\frac{\operatorname{Prob}(\alpha \beta \gamma)}{\operatorname{Prob}(\alpha) \operatorname{Prob}(\beta) \operatorname{Prob}(\gamma)}$
where $v_{i}$ represents a letter preceding $v_{j}$, $v_{k}$ represents a letter following $v_{j}, \operatorname{Prob}(\alpha \beta \gamma)$ is the a priori probability of a randomly chosen trigram being $\alpha \beta \gamma$, and $\operatorname{Prob}(\alpha)$ is the probability of a randomly chosen letter being $\alpha$.

## 4. Decoding

Let $\sum$ be the English alphabet together with a space symbol. A key $K$ for a coded message $C=c_{1} \ldots c_{l}$ is the transformation $K: \sum \rightarrow \sum$ such that $M=K\left(c_{1}\right) \ldots K\left(c_{l}\right)$ is the original message. In this section the key will be obtained from the coded message in two steps. Every node $v_{\alpha}$ in the graph will represent a letter $\alpha$ in the coded message. At the first step initial probabilities are assigned to every letter for every node. $P_{\alpha}^{(0)}(\beta)$ is the initial probability that $K(\alpha)$ will be $\beta$. The second step involves iterative application of relaxation to the probabilistically labeled graph to obtain a less ambiguous labeling. For every iteration, a key $K^{(n)}$ is constructed from $P^{(n)}$ such that $K^{(n)}(\alpha)=\beta$ if $P_{\alpha}^{(n)}(\beta)$ is the maximal element in $P_{\alpha}^{(n)}$. It will be seen that the number of elements in $K^{(n)}$ which agree with $K$ increases with the number of iterations performed.

### 4.1 Initial Probabilities

First order statistics are used to obtain the initial probabilistic labeling. Let $f_{e}(\beta)$ be the relative frequency of the letter $\beta$ in English. An initial probabilistic labeling can be found by using a multinomial model for the appearance of letters in a text [2]. In such a model, every letter in a text has the probability $f_{e}(\beta)$ of being the English letter $\beta$. For every such $\beta$, we can consider the binomial event of a letter being $\beta$ or being something else, using $f_{e}(\beta)$ as the parameter of that binomial distribution. Given a coded message of length $n$, let the code letter $\alpha$ appear $k$ times with relative frequency $f_{c}(\alpha)=$ $k / n$. Using the binomial model, we can estimate the probability of the code letter $\alpha$ being the code for the English letter $\beta$ (or the probability of $K(\alpha)=\beta$ ).

Using Bayes' rule, we have

$$
\begin{align*}
& \operatorname{Prob}\left(K(\alpha)=\beta \mid f_{\mathrm{c}}(\alpha)=k / n\right) \\
& =\frac{\operatorname{Prob}\left(f_{c}(\alpha)=k / n \mid K(\alpha)=\beta\right) \cdot \operatorname{Prob}(K(\alpha)=\beta)}{\operatorname{Prob}\left(f_{c}(\alpha)=k / n\right)} \tag{6}
\end{align*}
$$

The following three expressions are used to compute (6):
$\operatorname{Prob}(K(\alpha)=\beta)=\frac{1}{|\Sigma|}$
This means that all codes are equally likely.

$$
\begin{align*}
& \operatorname{Prob}\left(f_{c}(\alpha)=k / n\right) \\
& =\sum_{\beta \in \Sigma} \operatorname{Prob}\left(f_{c}(\alpha)=k / n \mid K(\alpha)=\beta\right)  \tag{8}\\
& \cdot \operatorname{Prob}(K(\alpha)=\beta)
\end{align*}
$$

This expression is an identity in probability theory.
$\operatorname{Prob}\left(f_{c}(\alpha)=k / n \mid K(\alpha)=\beta\right)$
$=\binom{n}{k}\left[f_{e}(\beta)\right]^{k}\left[1-f_{e}(\beta)\right]^{n-k}$
This arises from the assumption of a binomial distribution with probability $f_{e}(\beta)$ for a letter being $\beta$.

By substituting (7), (8), and (9) in (6), we can derive a simplified expression

$$
\begin{align*}
& \operatorname{Prob}\left(K(\alpha)=\beta \mid f_{c}(\alpha)=k / n\right) \\
& =\frac{\left[f_{e}(\beta)\right]^{k}\left[1-f_{e}(\beta)\right]^{n-k}}{\sum_{\lambda \in \Sigma}\left[f_{e}(\lambda)\right]^{k}\left[1-f_{e}(\lambda)\right]^{n-k}} \tag{10}
\end{align*}
$$

The values computed in (10) are used as initial probability estimates, and

$$
\begin{equation*}
P_{\alpha}^{(0)}\left(l_{\alpha}=\beta\right)=\operatorname{Prob}\left(K(\alpha)=\beta \mid f_{c}(\alpha)=k / n\right) . \tag{11}
\end{equation*}
$$

It was found in the experiments that the particular way in which initial probabilities are assigned does not significantly change the results of the relaxation process. Even when an arbitrary expression like
$P_{\alpha}^{(0)}\left(l_{\alpha}=\beta\right) \cong\left(1-\frac{\left|f_{c}(\alpha)-f_{e}(\beta)\right|}{f_{c}(\alpha)+f_{e}(\beta)}\right)^{4}$
(which assigns a probability of $\alpha$ being $\beta$ in accordance with how close $\beta$ 's frequency in English is to $\alpha$ 's frequency in the message) was used to compute the initial estimates, the results did not change significantly.

Note that no initial estimate is computed for the space symbol, since it is assumed that the space symbol is not changed by the code.

Table I. Relative frequencies of letters in the reference text, the technical report paragraph, and the Gettysburg Address.

| Letters | English | Report | Gettysburg |
| :---: | :---: | :---: | :---: |
| A | 0.078 | 0.072 | 0.089 |
| B | 0.014 | 0.008 | 0.012 |
| C | 0.024 | 0.036 | 0.027 |
| D | 0.048 | 0.046 | 0.050 |
| E | 0.129 | 0.117 | 0.144 |
| F | 0.022 | 0.025 | 0.023 |
| G | 0.021 | 0.035 | 0.024 |
| H | 0.066 | 0.031 | 0.070 |
| I | 0.072 | 0.086 | 0.059 |
| J | 0.001 | 0.001 | 0 |
| K | 0.008 | 0.007 | 0.003 |
| L | 0.040 | 0.032 | 0.037 |
| M | 0.025 | 0.027 | 0.011 |
| N | 0.072 | 0.089 | 0.067 |
| O | 0.075 | 0.071 | 0.081 |
| P | 0.016 | 0.029 | 0.013 |
| Q | 0.001 | 0.001 | 0.001 |
| R | 0.058 | 0.060 | 0.069 |
| S | 0.060 | 0.071 | 0.038 |
| T | 0.086 | 0.088 | 0.110 |
| U | 0.030 | 0.031 | 0.018 |
| V | 0.009 | 0.015 | 0.021 |
| W | 0.021 | 0.006 | 0.024 |
| X | 0.002 | 0.003 | 0 |
| Y | 0.022 | 0.008 | 0.009 |
| Z | 0.000 | 0.001 | 0 |

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Table IV. Same as Table II, But Using the Product Updating Rule (13).

| IIERATION: | 0 | 1 | 2 | 3 | 4 | 4 | 4 | 7 | 8 | 9 | 10 | 11 | 12 | 1.3 | 14 | 15 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| A | T/ 5®゙ | A/ 100 | A/100 | A/100 | A/100 | A/100 | A/100 | A/100 | A/100 | A/100 | A/100 | A. 100 | A/100 | A/ 100 | A/100 | A/100 |
| 日 | B/ 33 | B/100 | B/100 | B/100 | B/100 | B/100 | B/100 | B/100 | B/100 | B/ 100 | B/100 | B/100 | 13/100 | B/100 | 13/100 | B/100 |
| c | M/ 21 | C/ 98 | C/ 100 | C/100 | C/ 100 | C/ 100 | c/100 | $\mathrm{C} / 100$ | C/ 100 | $\mathrm{C} / 100$ | c/100 | C/ 100 | C/ 100 | c/100 | c/100 | c/100 |
| D | D/ 42 | S/ 80 | D/100 | D/100 | D/100 | D/ 100 | D/ 100 | D/100 | D/100 | D/100 | D/100 | D/100 | D/100 | D/100 | D/100 | D/100 |
| E | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 | E/100 |
| F | C/ 16 | M/ 91 | F/100 | F/100 | F/100 | F/100 | F/100 | F/100 | F/100 | F/100 | F/100 | $F / 100$ | F/100 | F/100 | F/100 | F/100 |
| G | c/ 17 | C/ 68 | 6/100 | G/100 | G/100 | G/100 | G/100 | G/100 | G/100 | G/100 | 6/100 | G/100 | G/100 | G/100 | G/100 | G/100 |
| H | I/ 19 | R/ 96 | H/100 | H/100 | H/100 | H/100 | H/100 | H/100 | H/100 | H/100 | Hi100 | H/100 | H/100 | H/100 | H/100 | H/100 |
| 1 | 5/ 29 | 1/98 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 | 1/100 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | x/ 44 | K/100 | K/100 | K/100 | K/100 | K/100 | K/100 | K/100 | k/100 | K/100 | K/100 | K/100 | K/100 | K/100 | K/100 | $k / 100$ |
|  | L/ 52 | L/100 | L/100 | L/100 | L/ 100 | L/100 | L/100 | 1/100 | L/100 | L/100 | 1./100 | L/100 | 1.1100 | L/100 | L/100 | L/100 |
|  | B/ 29 | P/ 72 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 | M/100 |
|  | H/ 21 | $N / 100$ | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 | N/100 |
|  | A/ 26 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 | 0/100 |
|  | B/ 35 | B/ 98 | B/100 | B/100 | 4/100 | -1/100 | B/100 | - /100 | B/100 | B/100 | - / 100 | B/100 | B/100 | B/100 | B/100 | B/100 |
|  | Q/ 28 | x/ 98 | x/ 56 | Q/ 97 | Q/100 | Q/100 | Q/100 | - /100 | 0/100 | Q/100 | Q/100 | Q/100 | Q/100 | Q/100 | Q/100 | Q/100 |
|  | H/ 19 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 | R/100 |
|  | L/ 61 | 5/ 81 | 5/100 | 5/100 | S/100 | 5/100 | S/100 | S/100 | S/100 | S/100 | 5/100 | S/100 | S/100 | \$/100 | S/100 | 5/100 |
|  | E/ 86 | T/100 | T/100 | 1/100 | T/100 | 1/100 | T/100 | T/100 | T/100 | T/100 | T/100 | T/100 | T/100 | T/100 | 1/100 | T/100 |
|  | G/ 16 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 | U/100 |
|  | G/ 16 | C/ 42 | $v / 100$ | $v / 100$ | V/100 | $v / 100$ | V/100 | V/100 | $v / 100$ | V/100 | V/100 | V/100 | V/100 | V/100 | V/100 | $V / 100$ |
|  | C/ 17 | W/ 95 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 | W/100 |
|  | V/ 42 | Y/ 87 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 | Y/100 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| CORRECT. | 5 | 15 | 21 | 22 | 22 | 22 | 22 | 22 | 22 | 22 | 22 | 22 | 2? | 22 | 2e | 22 |



### 4.2 Relaxation

The graph $G=(V, E)$ used consists of 27 nodes, $V=\left\{v_{\#}, v_{a}, \ldots, v_{z}\right\}$, where $v_{\#}$ represents the space symbol and $\nu_{\alpha}$ represents the code letter $\alpha$ for all letters that can appear in $C$. The "arcs" in the graph are triples of nodes, $E \subseteq V^{3}$. Every arc represents an occurrence of a trigram in $C$. The arcs are
$E=\left\{\left(v_{\alpha}, v_{\beta}, v_{\gamma}\right) \mid\right.$ There is one occurrence of $\alpha \beta \gamma$ in $\left.C\right\}$.
An $\operatorname{arc}\left(v_{\alpha}, v_{\beta}, v_{\gamma}\right)$ will appear in $E$ as many times as the sequence $\alpha \beta \gamma$ appears in $C$.

Let $E_{\lambda}$ be the set of all arcs passing through $v_{\lambda}$ :
$E_{\lambda}=\left\{\left(v_{\alpha}, v_{\lambda}, v_{\beta}\right) \mid\left(v_{\alpha}, v_{\lambda}, v_{\beta}\right) \in E\right.$.
Given the probability vectors $P^{(n)}$ at the $n$th iteration, the probabilities at the $(n+1)$ st iteration using (4) are

$$
\begin{equation*}
P_{\lambda}^{(n+1)}\left(l_{\lambda}=\omega\right)=\frac{1}{\left|E_{\lambda}\right|} \sum_{\left(v_{\alpha}, \nu_{\lambda}, \nu_{\beta}\right) \in E_{\lambda}} P_{\lambda \alpha \beta}^{(n+1)}\left(l_{\lambda}=\omega\right) \tag{12}
\end{equation*}
$$

where the $P_{\lambda \alpha \beta}^{(n+1)}$ are computed from the $P^{(n)}$,s using (3). Since $P_{\lambda \alpha \beta}^{(n+1)}$ is the probability vector for $\nu_{\lambda}$ as suggested by the sequence $\alpha \lambda \beta, P_{\lambda}^{(n+1)}$ is the average (vector) of all probability vectors suggested by all the occurrences of $\lambda$ in $C$, together with the preceding and the succeeding letters.

We can alternatively use (5) to compute the $P^{(n+1)}$ vectors from the $P^{(n)}$ vectors. Let $Q_{\lambda \alpha \beta}^{(n)}(\omega)$ be

$$
Q_{\lambda \alpha \beta}^{(n)}(\omega)=\sum_{\delta \in \Sigma} \sum_{\epsilon \in \Sigma} P_{\alpha}\left(l_{\alpha}=\delta\right) P_{\beta}\left(l_{\beta}=\epsilon\right) r_{\alpha \lambda \beta}(\delta, \omega, \epsilon)
$$

where $\left(v_{\alpha}, v_{\lambda}, v_{\beta}\right) \in E_{\lambda} \cdot Q_{\lambda \alpha \beta}^{(n)}(\omega)$ can be interpreted as the support that interpretation $\omega$ in node $v_{\lambda}$ gets from the probabilistic labeling at a predecessor $v_{\alpha}$ of $v_{\lambda}$ and a successor $v_{\beta}$ of $v_{\lambda}$. From the $Q$ 's a new estimate is computed analogous to (5):

$$
\begin{align*}
& P_{\lambda}^{(n+1)}\left(l_{\lambda}=\omega\right) \\
& \quad=\frac{P_{\lambda}^{(n)}\left(l_{\lambda}=\omega\right) \cdot \prod_{(\alpha, \lambda, \beta) \in E_{\lambda}} Q_{\lambda \alpha \beta}^{(n)}(\omega)}{\sum_{\delta} P_{\lambda}^{(n)}\left(l_{\lambda}=\delta\right) \cdot \prod_{(\alpha, \lambda, \beta) \in E_{\lambda}} Q_{\lambda \alpha \beta}^{(n)}(\delta)} \tag{13}
\end{align*}
$$

## 5. Examples

As examples we use two short passages. One is a paragraph taken from a recent technical report (996 characters), and the other is Lincoln's Gettysburg Address ( 1149 characters). The frequencies of letters in our reference text and the two messages are given in Table I. The messages were coded using the identity substitution, so that the coded message is identical to the original message. The key in this case is, of course, the identity transformation where $K(\alpha)=\alpha$ for every $\alpha \in \sum$. Using such an encoding makes no sense for a human cryptoanalyst, since every person familiar with the language can
immediately recognize the message, and there is no need for decoding. But for the computer system, which does not know English, finding the identity key is just as difficult as finding any other key.

Tables II and III summarize the maximum-probability substitution $K^{(n)}$ for each iteration, together with the corresponding probabilities, for the two examples using the average rule (12). The usage of one product rule (13) for the two examples is shown in Tables IV and V.

It can be observed from the examples that the multiplicative updating rule (13) converges faster than the average updating rule (12). But for the technical passage this speed caused a wrong classification of one letter, that was correctly classified by the "slow" averaging scheme.

## 6. Concluding Remarks

This paper has demonstrated the application of relaxation methods to the solution of substitution ciphers. We used ciphers in which blanks were left intact, but the method should work well even if another character were substituted for blank, since blank has significantly higher frequency than any other letter. Further experiments with messages where blanks were eliminated also resulted in a mostly correct key, but took more iterations to achieve.

The results illustrate how relaxation methods can be useful in solving a variety of probabilistic graph labeling problems. The application of relaxation to the problem of ambiguous segmentation of handwritten words with uncertain interpretation can be found in [5].

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