

## CHANGES OF LAW, MARTINGALES AND THE CONDITIONED SQUARE FUNCTION

NORIIHIKO KAZAMAKI

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Let  $(\Omega, F, P)$  be a complete probability space, given an increasing sequence  $(F_n)$  of sub  $\sigma$ -fields of  $F$  such that  $F = \bigvee_{n \geq 0} F_n$ . If  $f = (f_n, F_n)$  is a martingale with difference sequence  $d = (d_n)_{n \geq 1}$ , we shall set  $f^* = \sup_{n \geq 0} |f_n|$ ,  $S(f) = (\sum_{n=1}^{\infty} d_n^2)^{1/2}$  and  $s(f) = (\sum_{n=1}^{\infty} E[d_n^2 | F_{n-1}])^{1/2}$ . Let us assume that  $f_0 = 0$ . The operator  $s(f)$ , which is not of matrix type, is called the conditioned square function. It was studied by Burkholder and Gundy [3]. Let  $s_n(f) = (\sum_{k=1}^n E[d_k^2 | F_{k-1}])^{1/2}$ . Clearly,  $s_n(f)$  is  $F_{n-1}$ -measurable. Throughout the paper, we fix a BMO-martingale  $M_n = \sum_{k=1}^n m_k$ ,  $M_0 = 0$  such that  $-1 + \delta < m_k$ , ( $k \geq 1$ ) for some constant  $\delta$  with  $0 < \delta \leq 1$ , and consider the process  $Z$  given by the formula  $Z_n = \prod_{k=1}^n (1 + m_k)$ ,  $Z_0 = 1$ .  $Z$  is a positive uniformly integrable martingale which satisfies the condition

$$(A_p) \quad Z_n E[Z_{\infty}^{-1/(p-1)} | F_n]^{p-1} \leq C_p, \quad n \geq 0$$

for some  $p > 1$ ; see [6]. As  $Z_{\infty} > 0$  a.s., the weighted probability measure  $d\hat{P} = Z_{\infty} dP$  is equivalent to  $dP$ . Note that for every  $\hat{P}$ -integrable random variable  $Y$

$$\hat{E}[Y | F_n] = E[Z_{\infty} Y | F_n] / Z_n \quad \text{a.s., under } dP \text{ and } d\hat{P},$$

where  $\hat{E}$  denotes the expectation over  $\Omega$  with respect to  $d\hat{P}$ .

Our aim is to prove the following:

**THEOREM.** *Let  $0 < p \leq 2$ . Then the inequality*

$$(1) \quad \hat{E}[(f^*)^p] \leq c_p \hat{E}[s(f)^p]$$

*is valid for all martingales  $f = (f_n)$ .*

*Furthermore, if  $2 \leq p < \infty$  and  $Z$  satisfies the  $(A_p)$  condition, then we have*

$$(2) \quad \hat{E}[s(f)^p] \leq C_p \sup_{n \geq 0} \hat{E}[|f_n|^p].$$

*Here, the choice of  $c_p$  and  $C_p$  depends only on  $p$ .*

This result is well-known for the case where  $Z \equiv 1$ ; see Theorem 5.3 of [3]. To prove the theorem, we need several lemmas, which will

be stated without proof in the following. The letter  $C_p$  denotes a positive constant, not necessarily the same number from line to line.

LEMMA 1. *If  $\{a_n\}$  is a sequence of non-negative random variables, then for  $p \geq 1$*

$$E\left[\left(\sum_{n=1}^{\infty} E[a_n | F_n]\right)^p\right] \leq p^p E\left[\left(\sum_{n=1}^{\infty} a_n\right)^p\right].$$

See Theorem 3.2 of [2].

LEMMA 2. *Let  $1 < p < \infty$ . If  $Z$  satisfies  $(A_p)$ , then the inequality*

$$\widehat{E}[(f^*)^p] \leq C_p \sup_{n \geq 0} \widehat{E}[|f_n|^p]$$

*is valid for all martingales  $f = (f_n)$ .*

In our case,  $(A_p)$  implies  $(A_{p-\varepsilon})$  for some  $\varepsilon > 0$ , and so this inequality follows from Theorem 2 of [5]. It is proved in [5] that the converse to Lemma 2 is true.

LEMMA 3. *Let  $1 \leq p < \infty$ . If  $f = (f_n)$  is a martingale, then*

$$c_p \widehat{E}[S(f)^p] \leq \widehat{E}[(f^*)^p] \leq C_p \widehat{E}[S(f)^p].$$

For the proof, see [4]. The inequality corresponding to the continuous parameter case was obtained by Bonami and Lépingle [1] and Sekiguchi [8] independently.

LEMMA 4. *Let  $0 < \varepsilon \leq 1$ . Then we have*

$$Z_n^\varepsilon \leq C_\varepsilon E[Z_\infty^\varepsilon | F_n], \quad n \geq 1.$$

This is Lemma 1 of [7]. It is easy to see that for a martingale  $f = (f_n)$  with difference sequence  $d = (d_n)$ , the process  $\widehat{f} = (\widehat{f}_n)$  defined by  $\widehat{f}_n = \sum_{k=1}^n d_k / (1 + m_k)$  is a martingale with respect to  $d\widehat{P}$ . The following lemma is proved in [7].

LEMMA 5. *Let  $1 \leq p < \infty$ , and set  $\widehat{d}_k = d_k / (1 + m_k)$ . Then we have*

$$\widehat{E}\left[\left(\sum_{k=1}^{\infty} \widehat{d}_k^2\right)^{p/2}\right] \leq C_p E\left[\left(\sum_{k=1}^{\infty} Z_{k-1}^{2/p} d_k^2\right)^{p/2}\right].$$

PROOF OF THEOREM. Let  $s(\widehat{f})$  denote the conditioned square function  $(\sum_{k=1}^{\infty} \widehat{E}[\widehat{d}_k^2 | F_{k-1}])^{1/2}$  relative to  $d\widehat{P}$ . Since  $\delta < Z_k / Z_{k-1} = 1 + m_k \leq \|M\|_{\text{BMO}}$  and  $\widehat{E}[\widehat{d}_k^2 | F_{k-1}] = E[d_k^2 / (1 + m_k) | F_{k-1}]$ , we have  $(1 + \|M\|_{\text{BMO}})^{-1/2} s(f) \leq s(\widehat{f}) \leq \delta^{-1/2} s(f)$  and  $\widehat{E}[d_k^2 | F_{k-1}] \leq (1 + \|M\|_{\text{BMO}}) E[d_k^2 | F_{k-1}]$ .

We start with the case  $0 < p \leq 2$ . From Lemma 3 it follows that

$$\begin{aligned} \widehat{E}[(f^*)^2] &\leq C\widehat{E}[S(f)^2] = C\widehat{E}\left[\sum_{n=1}^{\infty} \widehat{E}[d_n^2 | F_{n-1}]\right] \\ &\leq C\widehat{E}\left[\sum_{n=1}^{\infty} E[d_n^2 | F_{n-1}]\right] = C\widehat{E}[s(f)^2]. \end{aligned}$$

Thus (1) is proved for  $p = 2$ . Now, let us consider the case  $0 < p < 2$ . Following the idea of Garsia, we define a martingale transform  $g$  by the formula  $g_n = \sum_{k=1}^n s_k(f)^{(p-2)/2} d_k$ . Then by the definition of  $s(g)$

$$s(g)^2 = \sum_{n=1}^{\infty} s_n(f)^{p-2} E[d_n^2 | F_{n-1}] = \sum_{n=1}^{\infty} s_n(f)^{p-2} \{s_n(f)^2 - s_{n-1}(f)^2\}.$$

But, if  $0 < a \leq b$ , then  $b^{p-2}(b^2 - a^2) \leq 2(b^p - a^p)/p$ . This gives  $s(g)^2 \leq 2s(f)^p/p$ . Therefore,  $\widehat{E}[(g^*)^2] \leq C\widehat{E}[s(g)^2] \leq C_p \widehat{E}[s(f)^p]$ . On the other hand,

$$f_n = \sum_{k=1}^n s_k(f)^{1-p/2} (g_k - g_{k-1}) = g_n s_n(f)^{1-p/2} - \sum_{k=1}^n g_k \{s_k(f)^{1-p/2} - s_{k-1}(f)^{1-p/2}\}$$

and so  $f^* \leq 2g^* s(f)^{1-p/2}$ . Then we apply Hölder's inequality with exponents  $2/p$  and  $2/(2 - p)$ :

$$\begin{aligned} \widehat{E}[(f^*)^p] &\leq 2^p \widehat{E}[(g^*)^p s(f)^{p(1-p/2)}] \leq 2^p \widehat{E}[(g^*)^2]^{p/2} \widehat{E}[s(f)^p]^{1-p/2} \\ &\leq C_p \widehat{E}[s(f)^p]^{p/2} \widehat{E}[s(f)^p]^{1-p/2} \leq C_p \widehat{E}[s(f)^p]. \end{aligned}$$

Thus the desired inequality (1) is obtained.

Next we deal with the case  $2 \leq p < \infty$ . Let us assume that  $Z$  satisfies  $(A_p)$ ; namely, the weighted norm inequality stated in Lemma 2 holds. As  $s(f) \leq cs(\widehat{f})$ , we have  $\widehat{E}[s(f)^p] \leq C_p \widehat{E}[s(\widehat{f})^p] \leq C_p \widehat{E}[(\widehat{f}^*)^p]$ ; the right hand side inequality is well-known. See Theorem 5.3 (i) of [3]. By Lemmas 4 and 5 we have

$$\begin{aligned} \widehat{E}[(\widehat{f}^*)^p] &\leq C_p E\left[\left(\sum_{n=1}^{\infty} Z_n^{2/p} d_n^2\right)^{p/2}\right] \leq C_p E\left[\left(\sum_{n=1}^{\infty} Z_n^{2/p} d_n^2\right)^{p/2}\right] \\ &\leq C_p E\left[\left(\sum_{n=1}^{\infty} E[Z_n^{2/p} d_n^2 | F_n]\right)^{p/2}\right] \end{aligned}$$

and by Lemma 1 the expectation on the right hand side is smaller than  $E[(\sum_{n=1}^{\infty} Z_n^{2/p} d_n^2)^{p/2}] = \widehat{E}[(\sum_{n=1}^{\infty} d_n^2)^{p/2}]$ . Then this combined with Lemmas 2 and 3 yields (2) as desired. Thus the theorem is established.

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DEPARTMENT OF MATHEMATICS  
TOYAMA UNIVERSITY  
TOYAMA, 930 JAPAN