Original Contribution

Chronic Particulate Exposure, Mortality, and Coronary Heart Disease in the Nurses' Health Study

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Adverse health effects of exposures to acute air pollution have been well studied. Fewer studies have examined effects of chronic exposure. Previous studies used exposure estimates for narrow time periods and were limited by the geographic distribution of pollution monitors. This study examined the association of chronic particulate exposures with all-cause mortality, incident nonfatal myocardial infarction, and fatal coronary heart disease (CHD) in a prospective cohort of 66,250 women from the Nurses' Health Study in northeastern US metropolitan areas. Nonfatal outcomes were assessed through self-report and medical record review and fatalities through death certificates and medical record review. During follow-up (1992–2002), 3,785 deaths and 1,348 incident fatal CHD and nonfatal myocardial infarctions occurred. In age- and calendar-time-adjusted models, 10-μg/m^3 increases in 12-month average exposures to particulate matter <10 μm in diameter were associated with increased all-cause mortality (16%, 95% confidence interval: 5, 28) and fatal CHD (43%, 95% confidence interval: 10, 86). Adjustment for body mass index and physical activity weakened these associations. Body mass index and smoking modified the association between exposure to particulate matter <10 μm in diameter and fatal CHD. In this population, increases in such exposures were associated with increases in all-cause and CHD mortality. Never smokers with higher body mass indexes were at greatest risk of fatal CHD.

air pollution; cohort studies; coronary disease; environmental exposure; incidence; particulate matter; risk factors

Adverse effects associated with acute air pollution exposures have been well studied; however, evidence of chronic exposure effects has been more limited. The Harvard Six Cities Study (1) and the American Cancer Society Study (2), prominent early studies of long-term exposure to air pollution, both found an increased risk of mortality associated with long-term exposure to particulates. More recent key studies have shown an association between adverse health outcomes and chronic particulate exposure, even as particulate levels have decreased over time (3–6). Despite these findings, questions remain regarding the persistence of effects because particulate exposures accumulate over longer time periods. Additionally, gaps in current knowledge regarding susceptible subgroups of the population add to the importance of investigating the impact of risk factors that may change over time. Answers to these questions would inform targeted public health prevention strategies, as well as environmental policy.

The current study, using data from the Nurses' Health Study, an ongoing prospective cohort, examined the relation of chronic particulate exposures with all-cause mortality and incident and fatal coronary heart disease (CHD), addressing these questions by combining improved exposure, covariate, and outcome assessment. We estimated particulate matter less than 10 μm in diameter (PM$_{10}$) at every residential address for each nurse (updated biennially)

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during each month of the study period by using a statistical model combining spatial smoothing and land use regression. The model used PM_{10} monitoring data and variables that affect pollution generation and dispersion, such as population density, distance to roadways, elevation, urban land use, point-source PM_{10} emissions, precipitation, and wind speed. Repeated information on lifestyle and behavioral risk factors enabled us to examine potential confounding and effect modification by a variety of time-varying, individual-level risk factors.

MATERIALS AND METHODS

The Nurses' Health Study was initiated in 1976 when 121,700 US registered female nurses, aged 30–55 years, completed a mailed questionnaire. At study inception, they resided in 11 states scattered over the United States (California, Connecticut, Florida, Massachusetts, Maryland, Michigan, New Jersey, New York, Ohio, Pennsylvania, Texas). Since that time, participants have moved into all 50 states. The study population for this analysis was limited to nurses living in Metropolitan Statistical Areas in 13 states in the northeastern region of the United States (Maine, New Hampshire, Vermont, Massachusetts, Connecticut, Rhode Island, New York, New Jersey, Pennsylvania, Delaware, Ohio, Maryland, Michigan). This contiguous region contains 63% of the Nurses' Health Study population. The density of PM_{10} monitoring locations was inadequate for detailed exposure assessment prior to 1988. Therefore, we began follow-up in 1992 to allow for as long as 4 years of exposure estimation prior to the outcomes of interest. Since the baseline and subsequent follow-up Nurses' Health Study questionnaires are mailed to participants in June of even years, cohort follow-up began on June 30, 1992.

From 1976 until the present, study participants have completed biennial questionnaires on behavioral and lifestyle exposures and health outcomes. Only 6% of nurses available for follow-up no longer respond to questionnaires. Questions have addressed such risk factors as physical activity, body mass index (weight (kg)/height (m)^2), smoking status, and medical history. Our outcomes of interest were fatal CHD, nonfatal myocardial infarction, incident CHD (defined as first incident nonfatal myocardial infarction or fatal CHD), and all-cause mortality. These outcomes were selected on the basis of prior Nurses' Health Study research and were verified (7–10). Women who self-reported nonfatal outcomes were asked for permission to release their medical records, which were reviewed by a physician blinded to exposure status.

Confirmed cases of nonfatal myocardial infarction met World Health Organization criteria: symptoms and diagnostic electrocardiographic changes or elevations in cardiac enzymes (11). Cases of nonfatal myocardial infarctions were designated as probable if an interview or letter confirming hospitalization for the infarction was obtained and medical records were unavailable. We included probable and confirmed cases in the analysis. Information on deaths was obtained through next of kin, death certificates, postal

authority reports, or the National Death Index. Fatal CHD was confirmed by hospital records, autopsy, or death certificate if CHD was the underlying and most plausible cause. We also included sudden deaths if CHD was the plausible cause listed on the death certificate. For all-cause mortality analyses, we excluded 7,884 women with a history of cancer (other than nonmelanoma skin cancer) before 1992, the beginning of the follow-up period. For analyses of incident nonfatal myocardial infarction, we excluded women with a history of nonfatal myocardial infarction. Analyses of all-cause mortality excluded accidental deaths. For our analysis of fatal CHD, women with nonfatal myocardial infarctions prior to baseline were excluded, and women experiencing nonfatal myocardial infarctions during follow-up were censored at the time of the event.

Exposure assessment

PM_{10} exposures were estimated for each participant by using a geographic information system–based spatial smoothing model that predicted residence-specific monthly outdoor PM_{10} concentrations in the northeastern region of the United States (12). This model, a generalized additive mixed model, used PM_{10} data from monitoring sites in the US Environmental Protection Agency's Air Quality System, the IMPROVE network (Interagency Monitoring of PROtected Visual Environments, a network of monitors located in national parks and wilderness areas), and Harvard University research studies to estimate monthly smooth spatial terms and smooth regression terms of geographic information system–derived and meteorologic covariates.

The model was fit in 2 stages. In the first-stage equation, $y_{i,t} = u_i + f_1(Z_{i,t,1}) + \ldots + f_P(Z_{i,t,P}) + (a_t + g_t(s_i)) + e_{i,t}; e_{i,t} \sim N(0, \sigma_{e,t}^2)$, $y_{i,t}$ is the natural-log-transformed PM_{10} monthly site average, t is the number of months, u_i is an estimated site-specific intercept that represents the long-term adjusted mean, $g_t(s_i)$ is the time-period-specific residual spatial variability, s_i is the projected spatial coordinate pair for the ith location, and $Z_{i,t,1}$ through $Z_{i,t,P}$ are time-varying covariates derived from the geographic information system. This model stage estimates site-specific terms ($\hat{u}_i$), adjusting for time-varying covariates and time-varying residual spatial variability. The second-stage equation, $\hat{u}_i = a + d_1(X_{i,1}) + \ldots + d_Q(X_{i,Q}) + g(s_i) + b_i; b_i \sim N(0, \sigma_b^2)$, where b_i is a site-specific random effect, models the site-specific $\hat{u}_i$ terms by using smooth functions of site-specific, time-invariant geographic information system–derived covariates ($X_{i,1}$ through $X_{i,Q}$) and residual time-invariant spatial variability ($g(s_i)$). This model structure allows for highly spatially and temporally resolved predictions of chronic PM_{10} levels, even for individuals living in areas with no nearby monitors (albeit with greater uncertainty for locations with distant monitors).

Geographic information system–derived model covariates included block group and county-level population density, distance to the nearest road by Census Feature Class Code road class for A1–A3 roads, elevation from the US Geological Survey National Elevation Dataset, urban land use from the US Geological Survey National Land Cover Dataset, point- and area-source primary PM_{10} emissions

information from the US Environmental Protection Agency National Emissions Inventory, and meteorologic information on wind speed and precipitation from the National Climatic Data Center (12). Cross-validation results showed that this geographic information system–based spatial smoothing model (cross-validation $R^2 = 0.62$) performed substantially better than other approaches using inverse distance weighting (cross-validation $R^2 = 0.29$) or nearest-neighbor spatial interpolators (cross-validation $R^2 = 0.22$).

Statistical analysis

Person-time was calculated from baseline (June 30, 1992) until the end of follow-up (June 30, 2002), death, or disease outcome, whichever occurred first. Those who died, reported cancer, or reported an outcome of interest (e.g., nonfatal myocardial infarction) prior to June 30, 1992, were excluded from the analysis. Nurses who moved outside a Metropolitan Statistical Area or the northeastern region of the United States were excluded from the period of follow-up during which these conditions were true but were allowed to reenter the risk set if the situation reversed. Incidence rates were determined as the number of new cases divided by person-months of follow-up. Time-varying Cox proportional hazards models on a monthly time scale were used to calculate hazard ratios and 95% confidence intervals evaluating the association between each outcome of interest and average PM_{10} exposure. In separate models, PM_{10} exposure was modeled as the month prior to the outcome; a 3-month moving average (calculated as the 3 months of exposure previous to the month in which an outcome occurred); and 12-, 24-, 36-, and 48-month moving averages. We stratified by age in months, as determined at the beginning of each month, and adjusted for state of residence (indicator variables), year (linear term), and season (indicator variables). State-level indicator variables were included to control for large-scale spatial patterns in mortality that might be caused by factors other than pollution, thereby estimating the effect of pollution based on within-state variation (6, 13). All statistical analyses used SAS version 9.1 software (SAS Institute, Inc., Cary, North Carolina).

Evaluation of confounders and effect modifiers

Although data for outcomes and pollutant exposures were available on a monthly basis, information on potential confounders and effect modifiers were available on a biennial basis only. Therefore, each woman was assigned the same covariate values for all of the months following a questionnaire return until updated values were available from the next questionnaire. Hypertension (yes, no), physician-diagnosed diabetes (yes, no), high cholesterol level (yes, no), physical activity (<3, $3-<9$, $9-<18$, $18-<27$, or ≥ 27 metabolic-equivalent hours/week), body mass index (continuous), and smoking status (never, former, or current) and pack-years of smoking were updated every 2 years. Family history of myocardial infarction (yes, no) was assessed at enrollment in 1976 and again in 1984. Previous studies (14, 15) have found an influence of socioeconomic status on the relation between air pollution exposure and health

Table 1. Characteristics of the US Study Population in Selected Categories From the Nurses' Health Study During the Follow-up Period 1992–2002[a]

Characteristic	Nation	Northeast	Northeast MSAs
No. of women	104,645	75,809	66,250
Mean (SD) age, years	62.2 (7.7)	62.5 (7.6)	62.4 (7.6)
Body mass index in kg/m², %			
<25.0	43.7	41.9	42.3
$25.0-<30.0$	33.9	34.4	34.2
≥ 30.0	22.4	23.7	23.5
Smoking status, %			
Never	45.0	44.1	43.6
Current	12.7	13.4	13.5
Former	42.3	42.5	42.9
Mean (SD) pack-years of smoking	25.0 (21.3)	24.8 (21.0)	24.8 (21.0)
High cholesterol level, %	49.7	49.9	49.5
Diabetes, %	7.4	7.5	7.4
Hypertension, %	40.5	40.6	40.4
Family history of myocardial infarction, %	33.5	34.2	34.2
Physical activity in MET-hours/ week, %			
<3	21.9	22.1	22.1
$3-<9$	23.2	23.7	23.8
$9-<18$	20.8	21.0	21.0
$18-<27$	12.8	13.0	12.8
≥ 27	21.2	20.5	20.3
Mean (SD) predicted PM_{10} exposure, µg/m³		21.2 (4.5)	21.6 (4.3)
Mean (SD) median family income in US $, thousands[b]		64.2 (23.8)	67.0 (24.1)
Mean (SD) median house value in US $, 10 thousands[b]		15.7 (10.1)	16.6 (10.4)

Abbreviations: MET, metabolic equivalent; MSA, Metropolitan Statistical Area; PM_{10}, particulate matter <10 µm in diameter; SD, standard deviation.

[a] Percentages are based on complete information for participants.

[b] Estimated for census tract of residence using 2000 US Census data.

outcomes. Given that adequate individual-level socioeconomic status information was unavailable from the Nurses' Health Study questionnaires, we developed 2 area-level socioeconomic status variables that were evaluated as potential confounders. Median household value and median family income were determined from 2000 US Census data for the census tract of residence. Effect modification by body mass index, physical activity, diabetes, smoking,

Table 2. Hazard Ratios and 95% CIs for All-Cause and Cause-Specific Mortality Associated With a 10-$\mu g/m^3$ Change in Predicted PM_{10} Exposure[a] for the US Nurses' Health Study During the Follow-up Period 1992–2002

	No. of Cases	No. of Person-months	The Month Before		3-Month Moving Average		12-Month Moving Average		48-Month Moving Average	
			Hazard Ratio	95% CI	Hazard Ratio	95% CI	Hazard Ratio	95% CI	Hazard Ratio	95% CI
All-cause mortality	3,785	600,752	1.04	0.98, 1.11	1.14	1.05, 1.23	1.16	1.05, 1.28	1.15	1.04, 1.28
Myocardial infarction										
First CHD event	1,348	597,456	1.08	0.98, 1.19	0.96	0.84, 1.09	1.10	0.94, 1.29	1.09	0.92, 1.29
Fatal CHD[b]	494	597,456	1.16	0.98, 1.36	1.21	0.98, 1.48	1.43	1.10, 1.86	1.43	1.09, 1.88
Nonfatal myocardial infarction	854	597,458	1.03	0.91, 1.18	0.83	0.71, 0.98	0.94	0.77, 1.15	0.93	0.75, 1.15

Abbreviations: CHD, coronary heart disease; CI, confidence interval; PM_{10}, particulate matter <10 μm in diameter.
[a] Modeled stratifying by age in months, adjusting for state of residence, year, and season.
[b] Including sudden deaths, excluding prior nonfatal myocardial infarction.

hypercholesterolemia, and hypertension was evaluated through stratification.

RESULTS

In 1992, the Nurses' Health Study consisted of 104,645 living nurses who had no reported history of cancer (excluding nonmelanoma skin cancer). Of these nurses, 75,809 were residing in our 13 selected states (referred to as the northeast region), with 66,250 living in Metropolitan Statistical Areas. Most women lived in Pennsylvania, New York, Ohio, and Massachusetts (4 of the original states at inception of the cohort), while fewer than 1% resided in Vermont, Rhode Island, New Hampshire, Delaware, or Maine. Table 1 shows little difference in the characteristics of nurses in the main cohort, in the northeastern region, and in northeastern region Metropolitan Statistical Areas. Mean age during follow-up was approximately 62 years. Forty-four percent of women living in the northeastern region reported never smoking, and 58% met the World Health Organization's defined category of being overweight (body mass index >25) (16–18). During follow-up, the mean 12-month moving average of predicted PM_{10} exposure in the northeastern region was 21.6 $\mu g/m^3$, with a standard deviation of 4.3 and an interquartile range of 6.0. Mean values of 12-month-moving-average PM_{10} were compared across categories of characteristics, and differences among categories were slight (data not shown). For example, the mean values ranged from 21.4 $\mu g/m^3$ to 21.6 $\mu g/m^3$ across categories of physical activity and 21.0 $\mu g/m^3$ to 22.8 $\mu g/m^3$ across median census tract family income.

During the follow-up period from 1992 to 2002, 3,785 nonaccidental deaths occurred. When we stratified by age and adjusted for state of residence, year, and season, each increase of 10 $\mu g/m^3$ in PM_{10} in the 12 months prior to death was associated with a hazard ratio of 1.16 (95% confidence interval (CI): 1.05, 1.28) (Table 2). Risk was not significantly elevated for exposure in the month prior to death; however, risk appeared to increase as the window of exposure increased to 24 months prior to the event (data not shown).

From 1992 until 2002, 854 nonfatal myocardial infarctions and 494 CHD deaths occurred, excluding prior nonfatal myocardial infarctions (Table 2). The hazard ratios for first incident nonfatal myocardial infarction or fatal CHD events associated with an increase of 10 $\mu g/m^3$ in PM_{10} were elevated, but not statistically significant, for the majority of the exposure time periods considered. Each 10-$\mu g/m^3$ increase in PM_{10} in the 12 months prior to CHD death was associated with a hazard ratio of 1.43 (95% CI: 1.10, 1.86). Risk of CHD death was also elevated significantly for PM_{10} exposure periods longer than 12 months. The risk associated with PM_{10} in the month prior to a nonfatal myocardial infarction was elevated slightly; however, none of the other time periods of exposure showed an association.

To assess the impact of potential confounders or modifying risk factors, univariate and multivariate models were adjusted for each potential confounder (Table 3). Physical activity and smoking each attenuated the risk estimates for all-cause mortality, whereas median census tract house value increased the risk estimate. In the full model, after adjustment for smoking status and pack-years, body mass index, diabetes, family history of myocardial infarction, hypercholesterolemia, hypertension, physical activity, and median family income and median house value in the census tract of residence, the hazard ratio was attenuated to 1.07 (95% CI: 0.97, 1.18). If physical activity was not included in the full model, then the hazard ratio was 1.11 (95% CI: 1.01, 1.23). The hazard ratio for CHD death remained statistically significant regardless of which potential confounder was included in the model (Table 3). However, physical activity and smoking attenuated and census tract median house value increased the strength of the association. In the fully adjusted multivariate model, the hazard ratio was 1.30 (95% CI: 1.00, 1.71); in the fully adjusted model, excluding physical activity, was 1.35 (95% CI: 1.03, 1.77).

The effect of PM_{10} exposure on all-cause mortality was not modified by body mass index, physical activity, hypercholesterolemia, hypertension, or diabetes. However, the association between CHD death and PM_{10} exposure was modified by body mass index. For CHD death, stratified results revealed a stronger risk of fatal CHD for women with a body mass index of 30 or higher (Table 4). Stratified

Table 3. Hazard Ratios and 95% CIs for Adjusted Associations of All-Cause Mortality and Fatal CHD[a] With a 10-μg/m^3 Change in Average Predicted PM$_{10}$ in the 12 Months Prior to Death for the US Nurses' Health Study During the Follow-up Period 1992–2002

Model	All-Cause Mortality		Fatal CHD	
	Hazard Ratio	95% CI	Hazard Ratio	95% CI
Basic[b]	1.16	1.05, 1.28	1.43	1.10, 1.86
Smoking	1.12	1.02, 1.23	1.39	1.07, 1.81
Body mass index	1.15	1.05, 1.27	1.40	1.08, 1.82
Diabetes	1.15	1.04, 1.27	1.41	1.08, 1.83
Family history of myocardial infarction	1.16	1.05, 1.28	1.44	1.11, 1.87
Hypercholesterolemia	1.16	1.05, 1.28	1.43	1.10, 1.86
Hypertension	1.15	1.05, 1.27	1.42	1.09, 1.84
Physical activity	1.08	0.98, 1.19	1.34	1.03, 1.74
Median family income	1.16	1.05, 1.28	1.42	1.09, 1.84
Median house value	1.19	1.08, 1.31	1.47	1.13, 1.93
Full[c]	1.07	0.97, 1.18	1.30	1.00, 1.71
Full excluding physical activity	1.11	1.01, 1.23	1.35	1.03, 1.77

Abbreviations: CHD, coronary heart disease; CI, confidence interval; PM$_{10}$, particulate matter <10 μm in diameter.
[a] Including sudden deaths.
[b] Modeled stratifying by age in months, adjusting for state of residence, year, and season.
[c] Modeled stratifying by age in months, adjusting for state of residence, year, season, smoking status, pack-years of smoking, family history of myocardial infarction, high cholesterol level, diabetes, hypertension, median family income in the census tract of residence, physical activity, and median house value in the census tract of residence.

Table 4. Hazard Ratios and 95% CIs for the Association Between Fatal CHD[a] and a 10-μg/m^3 Change in Average Predicted PM$_{10}$ in the 12 Months Prior to Death Stratified by Body Mass Index[b] for the US Nurses' Health Study During the Follow-up Period 1992–2002

	Body Mass Index <30		Body Mass Index ≥30	
No. of cases[c]	310		149	
No. of person-months	432,557		132,250	
	Hazard Ratio	95% CI	Hazard Ratio	95% CI
Base model[d]	1.20	0.86, 1.68	2.16	1.35, 3.45
Adjusted model[e]	1.08	0.76, 1.52	1.99	1.23, 3.22

Abbreviations: CHD, coronary heart disease; CI, confidence interval; PM$_{10}$, particulate matter <10 μm in diameter.
[a] Including sudden deaths.
[b] Weight (kg)/height (m)^2.
[c] Risks are based on complete information for participants.
[d] Modeled stratifying by age in months, adjusting for state of residence, year, and season.
[e] Modeled stratifying by age in months, adjusting for state of residence, year, season, smoking status, family history of myocardial infarction, high cholesterol level, diabetes, hypertension, median family income in the census tract of residence, physical activity, and median house value in the census tract of residence.

results also indicated that never smokers were at higher risk of fatal CHD associated with PM$_{10}$ exposure than former or current smokers (Table 5). Furthermore, among women with a body mass index of less than 30, the hazard ratio for those who had never smoked was 1.41 (95% CI: 0.82, 2.42) for a 10-μg/m^3 change in PM$_{10}$, whereas hazard ratios for former (hazard ratio = 0.98, 95% CI: 0.58, 1.64) and current (hazard ratio = 0.85, 95% CI: 0.47, 1.53) smokers were not elevated. Women with body mass indexes of 30 or higher showed a similar, but stronger pattern of risk, with a hazard ratio of 2.82 (95% CI: 1.40, 5.71) for never smokers.

DISCUSSION

In a population of women residing in Metropolitan Statistical Areas in the northeastern region of the United States, all-cause mortality was statistically significantly associated with average PM$_{10}$ exposures in the time period 3–48 months prior to death. The association was strongest with average PM$_{10}$ exposure in the 24 months prior to death (hazard ratio = 1.16, 95% CI: 1.05, 1.28) and weakest with exposure in the month prior to death (hazard ratio = 1.04, 95% CI: 0.98, 1.11). Risk of fatal CHD was significantly associated with chronic exposure to PM$_{10}$. The association

of greatest magnitude occurred with mean exposure in the 24 months prior to death (hazard ratio = 1.42, 95% CI: 1.11, 1.81). We did not find an association between nonfatal myocardial infarctions and PM$_{10}$ exposure in this population.

We examined the effects of potential confounders on these health outcomes, with physical activity and smoking attenuating the strength of the relation. Risk of all-cause mortality and fatal CHD associated with PM$_{10}$ in the 12 months before death was significantly elevated after adjustment for smoking status and pack-years, body mass index, diabetes, family history of myocardial infarction, hypercholesterolemia, hypertension, and median census tract family income and house value. Controlling for physical activity in addition to these factors resulted in elevated, but nonsignificant risks of all-cause mortality and fatal CHD. Few previous studies have controlled for physical activity, and mechanisms underlying this attenuation are unclear.

Examination of possible effect modification showed that the association between fatal CHD and long-term PM$_{10}$ exposure was influenced by body mass index and smoking. Women with a higher body mass index were at increased risk of fatal CHD associated with PM$_{10}$. Women who had never smoked showed the strongest risk, regardless of body mass index. These results suggest that the health effects of smoking may mask the impact of air pollution. Reductions in PM$_{10}$ levels may produce greater benefits in healthier populations.

Our observations are consistent with a growing body of literature on chronic particulate matter exposure, cardiovascular events, and mortality. The size fraction in these studies varies (PM$_{10}$ or particulate matter <2.5 μm in diameter (PM$_{2.5}$)), precluding direct comparisons in some cases. In the extended Harvard Six Cities Study (4) and the American

Table 5. Hazard Ratios and 95% CIs for the Association Between Fatal CHD[a] and a 10-µg/m³ Change in Average Predicted PM$_{10}$ in the 12 Months Prior to Death Stratified by Body Mass Index[b] and Smoking Status[c] for the US Nurses' Health Study During the Follow-up Period 1992–2002

	Never Smoker		Former Smoker		Current Smoker	
All						
No. of cases[a]	160		190		125	
No. of person-months	251,153		244,466		77,153	
	Hazard Ratio	**95% CI**	**Hazard Ratio**	**95% CI**	**Hazard Ratio**	**95% CI**
Basic model[d]	1.87	1.24, 2.81	1.32	0.89, 1.96	0.90	0.54, 1.51
Adjusted model[e]	1.83	1.20, 2.79	1.22	0.82, 1.83	0.88	0.52, 1.48
Body mass index <30						
No. of cases[a]	97		120		100	
No. of person-months	181,451		180,120		64,912	
	Hazard Ratio	**95% CI**	**Hazard Ratio**	**95% CI**	**Hazard Ratio**	**95% CI**
Basic model[d]	1.42	0.83, 2.41	1.01	0.60, 1.68	0.88	0.50, 1.58
Adjusted model[e]	1.41	0.82, 2.42	0.98	0.58, 1.64	0.85	0.47, 1.53
Body mass index ≥30						
No. of cases[a]	59		67		23	
No. of person-months	59,425		59,897		11,614	
	Hazard Ratio	**95% CI**	**Hazard Ratio**	**95% CI**	**Hazard Ratio**	**95% CI**
Basic model[d]	2.85	1.44, 5.65	1.89	0.99, 3.58	0.98	0.31, 3.14
Adjusted model[e]	2.82	1.40, 5.71	1.64	0.86, 3.13	1.03	0.31, 3.42

Abbreviations: CHD, coronary heart disease; CI, confidence interval; PM$_{10}$, particulate matter <10 µm in diameter.
[a] Including sudden deaths.
[b] Weight (kg)/height (m)².
[c] Risks are based on complete information for participants.
[d] Modeled stratifying by age in months, adjusting for state of residence, year, and season.
[e] Modeled stratifying by age in months, adjusting for state of residence, year, season, smoking status, family history of MI, high cholesterol level, diabetes, hypertension, median family income in census tract of residence, physical activity, and median house value in census tract of residence.

Cancer Society Study (6), the adjusted hazard ratios for a 10-µg/m³ change in PM$_{2.5}$ exposure were 1.14 (95% CI: 1.06, 1.22) and 1.06 (95% CI: 1.02, 1.11), respectively. Eftim et al. (3) evaluated the association between all-cause mortality and fine particulates by using Medicare data in the same counties. They observed a 10.9% (95% CI: 9.0, 12.8) increase in all-cause mortality for a 10-µg/m³ increase in PM$_{2.5}$ in the American Cancer Society Study counties and a 20.8% (95% CI: 14.8, 27.1) increase in the Harvard Six Cities Study counties.

In a case-control study of Swedish men and women using 30-year averaged, traffic-related PM$_{10}$, Rosenlund et al. (19) did not find an association with nonfatal myocardial infarctions (odds ratio = 0.92, 95% CI: 0.71, 1.19 for an approximately 5-µg/m³ change in PM$_{10}$). For fatal myocardial infarction, there was a nonsignificant elevated risk (odds ratio = 1.39, 95% CI: 0.94, 2.07) (19). Results from our study were stronger than those in this study, which included both genders. The Adventist Health Study on Health Effects of Smog showed associations of fatal CHD with PM$_{10}$, PM$_{2.5}$, and PM$_{10-2.5}$ for women but not for men (20).

The relation between chronic PM$_{2.5}$ exposure and cardiovascular events was examined in the Women's Health Initiative observational study (5). Although results are not directly comparable, a significant increase in deaths from CHD and a nonsignificant increased risk of first myocardial infarction were also found. Similar to our study, the authors of this study observed effect modification by body mass index. The overall associations were higher than those reported here, likely because of differences in the particle size studied.

The time pattern of the association we found is interesting. Distributed lag-time series analyses of PM$_{10}$ exposure out to a month prior to death (21, 22) reported larger associations than those seen a few days preceding death, very similar to our findings regarding monthly exposure variables. Recent reanalyses of the Harvard Six Cities Study (23, 24) examined whether previous findings are due to decades of exposure or shorter-term changes. The reported association was with exposure in the previous 2 years. These results indicate a developing coherence of the air pollution mortality literature, and the mortality risk benefits from reducing

air pollution would be expected within a few years of intervention.

Most previous cohort studies have not examined the effect of physical activity on the association of particulate exposure with all-cause mortality and fatal CHD (1, 2). It was assessed in the Women's Health Initiative study, which did not find evidence of confounding (5). Whether physical activity is a confounder, is a risk modifier, or is in the causal pathway warrants further exploration.

PM_{10} can deposit in the thoracic area of the respiratory system. The biologic mechanisms through which PM_{10} contributes to mortality and CHD are not clearly understood. Hypothesized pathways include oxidative stress (25) and inflammation leading to accelerated atherosclerosis (26, 27) or endothelial dysfunction (28), and ischemic responses in the myocardium (28).

Limitations of our study include potential bias from obtaining initial information through self-report. Use of medical records, interviews, autopsy reports, and death certificate information reduces the possibility, and it is unlikely that underreporting would be associated with exposure. Findings may differ for other geographic areas. Selecting the northeastern region of the United States enabled us to focus on contiguous states with a denser and more evenly distributed study population than in other regions of the country. The similarity of sources and model complexity increased accuracy but necessitated initial assessment and validation in a smaller area. We did not adjust for additional ambient pollutants. Particulate exposures were assigned by geocoded residential locations; however, work locations were unavailable. In addition, we were unable to determine time spent outdoors, housing characteristics, or whether nurses were living at reported addresses year-round. Some addresses were unable to be geocoded because of abbreviations and the like. High correlations between time windows of exposure ($\rho > 0.95$ for correlations of all measures between 12 and 48 months) and potential residual confounding by incomplete adjustment for season limited our ability to determine the most relevant time period. Although estimates of pollution exposure were available on a monthly basis, covariates were assessed only every 2 years.

Through a combination of updated covariates, improved outcome assessment, and more accurate long-term exposure assessment, the current study provides valuable information on the associations of chronic PM_{10} exposure with all-cause mortality and CHD. We identified women nonsmokers with a higher body mass index as a potentially susceptible population. Most importantly, with residential addresses that were updated every 2 years and monthly predictions from our geographic information system–based exposure model, we had the unique opportunity to assess exposures on a finer spatial and temporal scale than in previous long-term studies. Finally, our findings add to a growing coherence of the literature across multiple time scales indicating that the public health benefits of reducing particle concentrations will be realized within years, not decades, of the reduction. This study also suggests that measures taken to limit particulate air pollution should benefit population health over extended periods of time.

ACKNOWLEDGMENTS

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