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Common variants at 2q37.3, 8q24.21, 15q21.3, and 16q24.1 influence chronic lymphocytic leukemia risk

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URLs

Online Inheritance in Man: <http://www.ncbi.nlm.nih.gov/sites/entrez>

The R suite can be found at <http://www.r-project.org/>

Detailed information on the tag SNP panel can be found at <http://www.illumina.com/>

dbSNP: <http://www.ncbi.nlm.nih.gov/entrez/query.fcgi?CMD=search&DB=snp>

HAPMAP: <http://www.hapmap.org/>

KBioscience: <http://kbioscience.co.uk/>

WGAVIEWER: <http://www.genome.duke.edu/centers/pg2/downloads/wgaviewer.php>

International immunogenetics information system: <http://imgt.cines.fr>

IWCLL: http://www.icr.ac.uk/research/research_sections/cancer_genetics/cancer_genetics_teams/molecular_and_population_genetics/fcl/index.shtml

The European Genome-phenome Archive (Wellcome Trust Case-Control Consortium [WTCCC]): <http://www.ebi.ac.uk/ega/page.php>

POLYPHEN: <http://www.bork.embl-heidelberg.de/PolyPhen/>

SIFT: <http://blocks.fhcr.org/sift/SIFT.html>

SNAP (SNP Annotation and Proxy Search): <http://www.broadinstitute.org/mpg/snap/>

Author Contributions

RSH designed, obtained funding, directed the study, and oversaw analyses. RSH drafted the manuscript, with help from DC-S, PB, MCDB and SED. MCDB performed statistical analyses. DC-S, MCDB, SED performed bioinformatics analyses. RSH and DC established the parent study. RSH, DC and DC-S developed patient recruitment, sample acquisition and performed sample collection of cases. For the GWA and UK-replication series 1, DC-S and PB supervised laboratory management and oversaw genotyping of cases. DC-S conducted sequencing. JMA and DJA conceived of the Newcastle-based CLL study. JMA established the study, supervised laboratory management and oversaw genotyping of cases and controls. NJS performed sample management of cases and controls. AGH developed the Newcastle Haematology Biobank, incorporating the Newcastle-based CLL study. TM-F, GHJ, GS, RJH, ARP, DO, DJA, JRB, GP, CP, and CF developed patient recruitment, sample acquisition and performed sample collection of cases. Coordinated of the Spain replication series was conducted by EC and AE. EC and AE provided CLL samples and AC controls and compiled detailed phenotypic information from cases and controls respectively. Genotyping was performed by MT and AC and MT supervised laboratory management and quality control. For the Swedish case-control study MM performed sample collection and prepared DNA. RR performed sample collection for all cases, while JJ, GJ and KES performed sample collection of cases in the SCALE study. PB and DC-S performed genotyping of cases and controls.

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Competing Interests Statement

The authors declare no competing financial interests.

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Abstract

To identify novel risk variants for chronic lymphocytic leukemia (CLL) we conducted a genome-wide association study of 299,983 tagging SNPs, with validation in four additional series totaling 2,503 cases and 5,789 controls. We identified four risk loci for CLL at 2q37.3 (rs757978, *FARP2*; odds ratio [OR] = 1.39; $P = 2.11 \times 10^{-9}$), 8q24.21 (rs2456449; OR = 1.26; $P = 7.84 \times 10^{-10}$), 15q21.3 (rs7169431; OR = 1.36; $P = 4.74 \times 10^{-7}$) and 16q24.1 (rs305061; OR = 1.22; $P = 3.60 \times 10^{-7}$). There was also evidence for risk loci at 15q25.2 (rs783540, *CPEBI*; OR = 1.18; $P = 3.67 \times 10^{-6}$) and 18q21.1 (rs1036935; OR = 1.22; $P = 2.28 \times 10^{-6}$). These data provide further evidence for genetic susceptibility to this B-cell hematological malignancy.

While B-cell chronic lymphocytic leukemia (CLL) shows a strong familial risk¹, the genetic basis of predisposition is largely unknown. We have previously reported the results of a genome-wide association (GWA) study of CLL based on the analysis of 299,983 tagging SNPs in 505 cases and 1,438 controls (henceforth referred to as Stage 1)² and through fast tracking analysis of the smallest *P*-values identified risk loci at 2q13, 2q37.1, 6p25.3, 11q24.1, 15q23, and 19q13.322. We have conducted further follow-up analyses making use of existing GWA data and report four novel susceptibility loci.

In Stage 2, we genotyped 180 SNPs in 540 UK cases (UK-replication series 1). The SNPs were chosen by a hypothesis-free strategy on the basis of *P*-values, excluding those correlated with an $r^2 > 0.8$ with previously identified association signals. After quality control 162 SNP genotypes were recovered for 519 of the cases. We used publicly accessible data on 2,695 UK-blood donor controls to compare genotype frequencies. In Stage 3, we genotyped the 19 SNPs that showed the strongest association from combined analysis of Stages 1 and 2 in two case-control series: UK-replication series 2 (660 cases, 809 controls), Spanish-replication series (424 cases, 450 controls). In Stage 4 we genotyped the 10 SNPs with the strongest association from a combined analysis of Stages 1-3 in the Swedish-replication series (395 cases, 397 controls) (Supplementary Figure 1).

The combined analysis of these data provided conclusive evidence of an association between seven SNPs and CLL (i.e. $P < 5.0 \times 10^{-7}$; Supplementary Table 1) mapping to five independent genomic regions. rs11668878 maps to 19q13.32, a region we have previously reported to be a risk locus for CLL (defined by rs11083846)². Linkage disequilibrium (LD) exists between rs11668878 and rs11083846 ($r^2 = 0.27, D' = 1.0$) and conditional analysis did not provide evidence for a second disease locus ($P > 0.05$). Mouthwash DNA was available from 89 of the cases in Stage 1 and Stage 2; SNP genotypes were 99% concordant with genotypes obtained from typing blood DNA. Together with the fact that these associations identified do not map to any of the regions of the genome commonly associated with copy number variation in CLL³⁻⁶ mitigates against bias from differential genotyping as a consequence of allelic imbalance influencing findings.

Under a fixed-effects model the strongest evidence for a novel association was attained with rs2456449 which maps to 8q24.21 at 128,262,163bp (OR=1.26; 95% CI: 1.17-1.35; $P = 7.84 \times 10^{-10}$; $P_{\text{het}} = 0.95, I^2 = 0\%$; Figure 1, Supplementary Table 1). rs2466024, localizing to 128,257,201bp, also provided significant evidence for the 8q24.21 association ($P = 4.61 \times 10^{-7}$; Supplementary Table 1). rs2456449 and rs2466024 are in strong LD ($r^2 = 0.75, D' = 1.0$) and map to a 40kb region of LD (128,241,868-128,282,415bp; Figure 2; Supplementary Figure 2). Conditional analysis provided no evidence for an independent role of rs2466024.

Genome-wide association studies have shown that the 128-130Mb genomic interval at 8q24.21 harbors multiple independent loci with different tumor specificities: prostate (rs16901979;128,194,098bp)⁷, breast (rs13281615;128,424,800bp)⁸, colorectal-prostate (rs6983267;128,482,487bp)^{9,10}, prostate (rs1447295;128,554,220bp)¹¹ and bladder (rs9642880;128,787,250bp)¹² cancer. The LD blocks defining these loci are distinct from the 8q24.21 CLL association signal. This is reflected in the LD metrics between rs2456449,

and rs6983267 ($r^2=0.00, D'=0.13$), rs13281615 ($r^2=0.00, D'=0.01$), rs16901979 ($r^2=0.01, D'=1.0$), rs1447295 ($r^2=0.04, D'=1.0$), rs9642880 ($r^2=0.02, D'=0.19$). The 8q24.21 region to which the cancer associations map is bereft of genes and predicted transcripts. rs6983267 has recently been shown to affect TCF4 binding to an enhancer for the *MYC* promoter, providing a mechanistic basis for this 8q24.21 association^{13,14}. It is possible that the effect of the other 8q24.21 cancer risk loci is via *MYC* through similar long range cis-acting mechanisms. We have shown that variation in *IRF4* influences CLL risk². If the 8q24.21 locus influences risk through differential *MYC* expression then the association is intriguing as *MYC* is a direct target of *IRF4* in activated B-cells¹⁵. While no relationship between rs2456449 genotype and *MYC* expression was shown in EBV-transformed lymphocytes (Supplementary Figure 3) this does not preclude a subtle effect on expression.

The second strongest statistical evidence for an association signal was provided by rs757978, mapping to exon 9 of *FARP2* at 2q37.3 (242,019,774bp; OR=1.39, 95%CI:1.25–1.56; $P=2.11 \times 10^{-9}$; $P_{\text{het}}=0.13$, $I^2=43\%$; Figure 1, Supplementary Figure 2, Supplementary Table 1). rs11681497, which maps to intron 4 of *FARP2* (241,993,006bp) and is in LD with rs757978 ($r^2=1, D'=1$), also provided strong evidence for the 2q37.3 association ($P=4.53 \times 10^{-9}$; Figure 2), however conditional analysis provided no evidence for an independent role. *FARP2* is involved in signaling downstream of G-protein-coupled receptors¹⁶. We examined if there was a relationship between genotype and *FARP2* expression in EBV-transformed lymphocytes. Data was unavailable for rs757978 but no association between rs11681497 genotype and mRNA expression level was shown (Supplementary Figure 3). rs757978 leads to the substitution of threonine for isoleucine at amino acid 260 (T260I) in the expressed protein. As *in silico* analysis predicts T260I to be functionally deleterious it is possible that the association signal at 2q37.3 is a consequence of T260I.

The third strongest association was provided by rs305061 (OR=1.22; 95% CI:1.12-1.32; $P=3.60 \times 10^{-7}$; $P_{\text{het}}=0.38$, $I^2=5\%$; Figure 1, Supplementary Table 1), which maps within a 30kb region of LD at 16q24.1 (84,533,160bp; Figure 2). rs305061 localizes 19kb telomeric to *IRF8*. Variation in *IRF8* represents a strong candidate for the association as *IRF8* regulates α and β interferon response. Moreover, *IRF8* is involved in B-cell lineage specification, immunoglobulin rearrangement, and regulation of germinal center reaction¹⁷. Variation at 16q24.1, defined by rs17445836 which maps 61kb telomeric to *IRF8* (84,575,164bp), has recently been shown to be a risk locus for multiple sclerosis (MS)¹⁸. MS has been reported to be associated with CLL risk^{19,20}. While rs17445836 and rs305061 are not strongly correlated ($r^2=0.13, D'=0.82$), it is possible that a common genetic basis to both diseases is mediated through the same causal variant at 16q24.1.

The fourth strongest association was identified with rs7169431 which maps to 15q21.3 (54,128,188bp; OR=1.36; 95%CI:1.21-1.53; $P=4.74 \times 10^{-7}$; $P_{\text{het}}=0.51$, $I^2=0\%$; Figure 1, Supplementary Table 1). rs7169431 localizes to a 25kb region of LD flanked by *NEDD4* and *RFX7* (Figure 2, Supplementary Figure 2). While there is no evidence for a direct role of *NEDD4* in CLL, it represents a credible candidate gene because of its role in regulating viral latency and pathogenesis of EBV. Specifically, *NEDD4* regulates EBV-LMP2A which mimics signaling induced by the B-cell receptor altering B-cell development²¹.

In addition to these four loci there was evidence for two disease loci at 15q25.2 (rs783540) and 18q21.1 (rs1036935), although associations did not attain genome-wide significance (OR=1.18; 95%CI:1.10-1.27; $P=3.67 \times 10^{-6}$; $P_{het}=0.07$, $I^2=53\%$ and OR=1.22; 95%CI: 1.12-1.32; $P=2.28 \times 10^{-6}$, $P_{het}=0.39$, $I^2=3\%$, respectively; Supplementary Table 1, Supplementary Figure 2). rs783540 localizes to intron 2 of *CPEB1* which has an established role in the regulation of cyclin B1 during embryonic cell division-differentiation. Two genes mapping centromeric to rs1036935, *CXXCI* and *MBDI* are involved in gene regulation. *MBDI* expression in EBV-transformed lymphocytes correlated with risk genotype (Supplementary Figure 3). Although *MBDI* has no documented role in CLL, it has potential to affect CLL development through translational control of *MYC* via MDBP binding²².

CLL shows male predominance and can be classified on the basis of the presence or absence of somatic hypermutations of immunoglobulin heavy-chain variable (IGVH) genes^{23,24}, with mutated-CLL having a better prognosis. 17p-deletion is also associated with inferior survival²⁵. We assessed the relationship between age, gender, family history of B-cell malignancy, IGVH-mutation and 17p-deletion status, and SNP-genotypes by case-only logistic regression (Supplementary Table 2). Furthermore, we examined the influence of genotype on overall survival (OS; Stages 1 and 4) and progression free survival (PFS; Stage 1; Supplementary Table 3). None of the SNPs displayed evidence of a relationship with age or gender (based on cases from all Stages). Having a family history of B-cell malignancy (Stages 1, 2) was associated with a higher frequency of *FARP2* rs11681497 and rs757978 risk genotypes ($P=0.031$, 0.037 respectively; Supplementary Table 2), compatible with familial cases being enriched for genetic susceptibility. No relationship between 17p deletion status and genotype was shown (Stage 1). Although there was evidence rs305061 risk genotype was associated with worse OS (Supplementary Table 3), IGVH mutational status (Stages 1,2,4) was highly correlated with rs305061, with risk genotype correlating with unmutated-CLL ($P=0.0002$; Supplementary Table 2). Since rs305061 maps to *IRF8* this relationship is compatible with dysfunctional B-cell activation signaling being associated with possession of risk genotype.

To gain insight into the allelic architecture of predisposition to CLL we examined for interactive effects between 2q37.3, 8q24.21, 15q21.3, and 16q24.1 variants and the previously identified loci at 2q13(rs17483466), 2q37.1(rs13397985), 6p25.3(rs872071), 11q24.1(rs735665), 15q23 (rs7176508) and 19q13.32(rs11083846). The only evidence for an interaction between loci was provided by rs305061 and rs7176508 ($P=0.045$), albeit non-significant after adjusting for multiple testing (Supplementary Table 4). The proportion of case and control subjects grouped according to the number of risk alleles that they carry is detailed in Figure 3. The distribution of risk alleles follows a normal distribution in both cases and controls, but with a shift toward a higher number of risk alleles in the cases.

The carrier frequencies of risk alleles are high in the European population and hence the 10 loci collectively make a major contribution to the development of CLL with the population attributable fraction ascribable to these loci being ~87%. Moreover, the risk of CLL increases with increasing numbers of variant alleles for the 10 loci (OR_{per-allele}=1.39, 95%CI:1.35-1.44; $P=2.73 \times 10^{-88}$; Figure 3, Supplementary Table 5). Individuals with 13+ risk alleles have a >7-fold increase in CLL risk compared to those with a median number of

risk alleles. We estimate that the 10 loci we have identified account for ~10% of the excess familial risk of CLL, assuming a polygenic model. While the estimated effect of each of the 10 risk variants we have identified may be upwardly biased due to the phenomenon referred to as the “winners curse”, ORs may be underestimates because the additive model on which analyses are based assumes equal weighting across the SNPs. The effect on susceptibility attributable to variation at the loci and the effect of the actual causal variant responsible for the association is however, likely to be greater. Furthermore, many of the loci may carry additional risk variants, potentially including low-frequency variants with larger influence on CLL risk. Genome-wide linkage scans have provided evidence, albeit non-significant, for moderate-high penetrance susceptibility loci for CLL26–28. As none of the variants we have identified map to these regions of linkage it is unlikely that any other allele at the loci we have identified is responsible for these linkage signals.

The power of our study to detect the major common loci conferring risks of 1.4 (e.g. rs872071) was high. Hence, we consider that there are unlikely to be many additional SNPs with similar effects for alleles with frequencies >0.4 in populations of European ancestry. In contrast, we had low power to detect alleles with smaller effects and/or MAFs<0.1 (e.g. rs7169431). By implication, variants with such profiles are likely to represent a much larger class of susceptibility loci for CLL, whether because of truly small effect sizes or sub-maximal LD with tagging SNPs. Furthermore, GWA-based strategies are not optimally configured to identify low frequency variants with potentially stronger effects. In addition, these arrays are not ideally formatted to capture copy number variants which may also impact on CLL risk. Thus it is likely that a large number of low penetrance variants remain to be discovered. This assertion is supported by the continued excess of associations observed over those expected, in addition to the regions studied herein. Further efforts to expand the scale of GWA meta-analyses, in terms of both sample size and SNP coverage, and to increase the number of SNPs taken forward to large-scale replication may identify additional risk variants.

These results provide evidence for low risk variants predisposing to CLL and insight into the development of this hematological malignancy. Ethnic differences in the risk of CLL are recognized and it is notable that the MAF of 8q24.21 and 15q21.3 SNPs provided evidence of differences in the Spanish and UK/Swedish populations (Supplementary Table 1). Hence it will be interesting to explore how our findings translate to non-Western countries with substantially low incidence rates²⁹. Furthermore, the reciprocal familial risks between CLL and other B-cell malignancies¹ raises the possibility that these variants may also influence the risk of related B-cell tumors.

Methods

Initial genome-wide scan

This study describes the follow-up of a previously reported GWA of CLL². Briefly 505 cases were genotyped using HumanCNV370-Duo BeadChips (Illumina, San Diego, USA). For controls we made use of publicly accessible HumanHap550K BeadChip genotype data generated on 1,438 individuals from the WTCCC British 1958 Birth Cohort. Quality control metrics included removal of samples with call rates <90% and SNP assays with call rates

<95%. Subjects with evidence of cryptic relatedness and non-European background were excluded from the analysis. We considered only the 299,983 autosomal SNPs with MAF>1% in cases and controls, and with no extreme evidence of departure from Hardy-Weinberg equilibrium (HWE; $P>10^{-5}$) in cases or controls. Comparison of observed and expected distributions showed little evidence for an inflation of test statistics (inflation factor $\lambda=1.053$).

Replication samples

In Stage 2 (**UK-replication series 1**) we genotyped SNPs in 540 CLL cases (353 male; mean age at diagnosis 60.3years; $SD\pm 12.2$) ascertained through the Royal Marsden NHS Hospitals Trust, 1998-2006. All cases were UK residents and self-reported to be of European ancestry. Genotyping was conducted using the Sequenom MassARRAY system (<http://www.sequenom.com/>). For controls we made use of publicly accessible genotype data generated on 2,736 UK-blood donor controls. We excluded SNPs that deviated from HWE at $P<1.0 \times 10^{-5}$ in cases or controls. We also removed SNPs with MAF<0.05. To identify and exclude individuals with non-Western European ancestry, case and control data was merged with individuals of different ethnicities from HapMap, genome-wide IBS distances for markers shared between HapMap and our SNP panel were determined, and dissimilarity measures used to perform principal component analysis. After imposing these stringent quality control measures 162 SNP genotypes were available on 519 cases and 2,695 controls and this dataset formed the basis of our analysis.

In Stage 3 we genotyped 19 SNPs showing the strongest evidence of an association from joint analysis of Stages 1 and 2 in two independent case-control studies. **UK-replication series 2:** Cases were ascertained through the Newcastle CLL Consortium - 660 UK Caucasian patients with CLL diagnosed between 1970 and 2006 who attended hematology units in the UK (Newcastle Royal Victoria Infirmary, Gateshead Queen Elizabeth Hospital, Birmingham Heartlands Hospital, Hull Royal Infirmary, Liverpool Royal University Hospital, Royal Bournemouth Hospital, University of Wales College of Medicine Hospital). Peripheral blood for DNA extraction was taken between 1998 and 2006. Controls comprised 809 UK Caucasian healthy individuals (509 male) aged 16-69 (mean age 48.2 years, $SD \pm 14.3$) recruited to a study of acute leukemia conducted 1991-1996, previously described³⁰. **Spanish-replication series:** 445 cases and 450 controls. Cases were Spanish residents and were of self-reported European ancestry. Controls were obtained from the MedXen Control population cohort (Xunta de Galicia) and were matched to cases by geographic origin, age, and sex.

In Stage 4 we genotyped 10 SNPs, showing the strongest evidence of an association from joint analysis of Stages 1-3 in the **Swedish-replication series:** 304 cases samples from SCALE (Scandinavian Lymphoma Etiology)³¹ and 91 samples from the biobank at Uppsala University Hospital, Uppsala, Sweden. For the SCALE samples, bloods were collected from 275 cases during 1999-2001, within a median of 4 months from diagnosis (0-29 months) while in 29 cases, follow-up samples ascertained in 2007-2008 were used. The samples received from the biobank at Uppsala University Hospital were collected between 1982 and 2005. The Swedish replication series totaled 148 females and 247 males (mean age at

diagnosis 62.6). Swedish controls were collected at Karolinska Institutet, Stockholm, Sweden and included 397 healthy individuals (251 male) aged 33-71 years (mean 61.3, SD \pm 6.8).

Data on both progression-free survival (PFS) and overall survival (OS) was available on 356 Stage 1 cases³³ who participated in the UK CLL-4 trial, a randomized phase III trial established to compare the efficacy of fludarabine, chlorambucil, and the combination of fludarabine plus cyclophosphamide as treatment for Binet stages B, C and A-progressive CLL³⁴. Progression free survival was defined as the survival time from the date of randomisation to relapse needing further therapy, progression or death from any cause. Data on OS was also available for 395 of the Swedish replication series (Stage 4).

In all studies the diagnosis of CLL has been confirmed in accordance with the current WHO classification guidelines³⁵.

Ethics

Collection of blood samples and clinico-pathological information from patients and controls was undertaken with informed consent and relevant institutional ethical review board approval in accordance with the tenets of the Declaration of Helsinki.

Replication series genotyping

DNA was extracted from EDTA venous blood samples using standard methodologies and Picogreen quantified (Invitrogen, Paisley, United Kingdom). To ensure quality control of genotyping, a series of duplicate samples were genotyped in the same batches. For all SNP assays >99% concordant results were obtained. Genotyping in UK-replication 1 was performed by Sequenom MassARRAY system (<http://www.sequenom.com/>). We attempted to type 180 SNPs, although it was not possible to design functional assays for 18 SNPs with this technology. Details of the methodology are available on request. To exclude technical artifacts of genotyping, we included a random series of 24 samples previously genotyped using Illumina 550K HapMap arrays (Illumina, San Diego, USA). We successfully genotyped 519 unique subjects passing quality control metrics, excluding validation samples, and study duplicates. Genotyping of UK-replication 2 and the Swedish-replication series were conducted by competitive allele-specific PCR KASPar chemistry (KBiosciences Ltd, Hertfordshire, UK; <http://www.kbioscience.co.uk/>); primer sequences and conditions available on request. Excluding validation samples and study duplicates, passing quality control metrics we genotyped 1,469 unique subjects in UK-replication series 2 and 792 unique subjects in the Swedish replication series. Genotyping of Spanish samples was conducted using the Sequenom MassARRAY system (<http://www.sequenom.com/>) in two i-Plex assays. Excluding validation samples and study duplicates, passing quality control metrics we genotyped 874 unique subjects.

IGVH mutational and 17p deletion status

IGVH mutational status was determined according to BIOMED-2 protocols as described previously^{36,37} or commercial reagents (*In Vivo*Scribe Technologies, San Diego, USA). Clonality was assessed by size discrimination of PCR products using semi-automated

ABI3730xl/ABI3700 sequencers in conjunction with Genescan software (Applied Biosystems, Foster City, USA). Sequences obtained were submitted to online database IMGT/V-QUEST38. In accordance with published criteria, sequences with a germline identity of 98% were classified as unmutated, and those displaying identity <98% as mutated. Deletion of 17p was defined as >10% loss in cells.

Statistical analysis

Statistical analyses were undertaken using R (v2.3.1) and STATA (v10.0) Software. Deviation of the genotype frequencies in the controls from those expected under HWE was assessed by the χ^2 -test (1d.f.). The association between each SNP and risk of CLL was assessed by the Cochran-Armitage trend test. Odds ratios (ORs) and associated 95% confidence intervals (CIs) were calculated by unconditional logistic regression. Relationships between multiple SNPs showing association with CLL risk in the same region were investigated using logistic regression analysis and the impact of additional SNPs from the same region was assessed by a likelihood-ratio test.

Associations by gender, age and mutation status were examined by logistic regression in case-only analyses. The combined effect of pairs of loci identified with risk was investigated by logistic regression modeling; evidence for interactive effects between SNPs assessed by a likelihood ratio test. The OR and trend test for increasing numbers of deleterious alleles was estimated by counting two for a homozygote and one for a heterozygote assuming equal weights.

Meta-analysis was conducted using standard methods³⁹. Cochran's Q-statistic to test for heterogeneity³⁹ and the I^2 statistic⁴⁰ to quantify the proportion of the total variation due to heterogeneity were calculated. Large heterogeneity is typically defined as $I^2 > 75\%$. The population attributable fraction was estimated from $1 - \prod_i 1 - (x_i - 1)/x_i$ where $x_i = (1-p)^2 + 2p(1-p)OR_1 + p^2OR_2$, p is the population allele frequency, and OR_1 and OR_2 are the ORs associated with hetero- and homozygosity respectively. The sibling relative risk attributable to a given SNP was calculated using the formula⁴¹:

$$\lambda^* = \frac{p(pr_2 + qr_1)^2 + q(pr_1 + q)^2}{[p^2r_2 + 2pqr_1 + q^2]^2}$$

where p is the population frequency of the minor allele, $q=1-p$, and r_1 and r_2 are the relative risks (estimated as OR) for heterozygotes and rare homozygotes, relative to common homozygotes. Assuming a multiplicative interaction the proportion of the familial risk attributable to a SNP was calculated as $\log(\lambda^*)/\log(\lambda_0)$, where λ_0 is the overall familial relative risk estimated from epidemiological studies, assumed to be 7.251.

Bioinformatics

We used Haploview software (v3.2) to infer the LD structure of the genome in the regions containing loci associated with disease risk. We applied two *in silico* algorithms, PolyPhen and SIFT, to predict the impact of non-synonymous SNPs on protein function.

Relationship between SNP genotype and expression levels

To examine for a relationship between SNP genotype and expression levels of *MYC*, *FARP2*, *NEDD4*, *MBD1*, and *CPEB1* in lymphocytes we made use of publicly available expression data generated from analysis of 90 Caucasian derived Epstein-Barr virus–transformed lymphoblastoid cell lines using Sentrix Human-6 Expression BeadChips (Illumina, San Diego, USA)^{42,43}. Online recovery of data was performed using WGAViewer Version 1.25 Software. Differences in the distribution of levels of mRNA expression between SNP genotypes were compared using a Wilcoxon-type test for trend⁴⁴.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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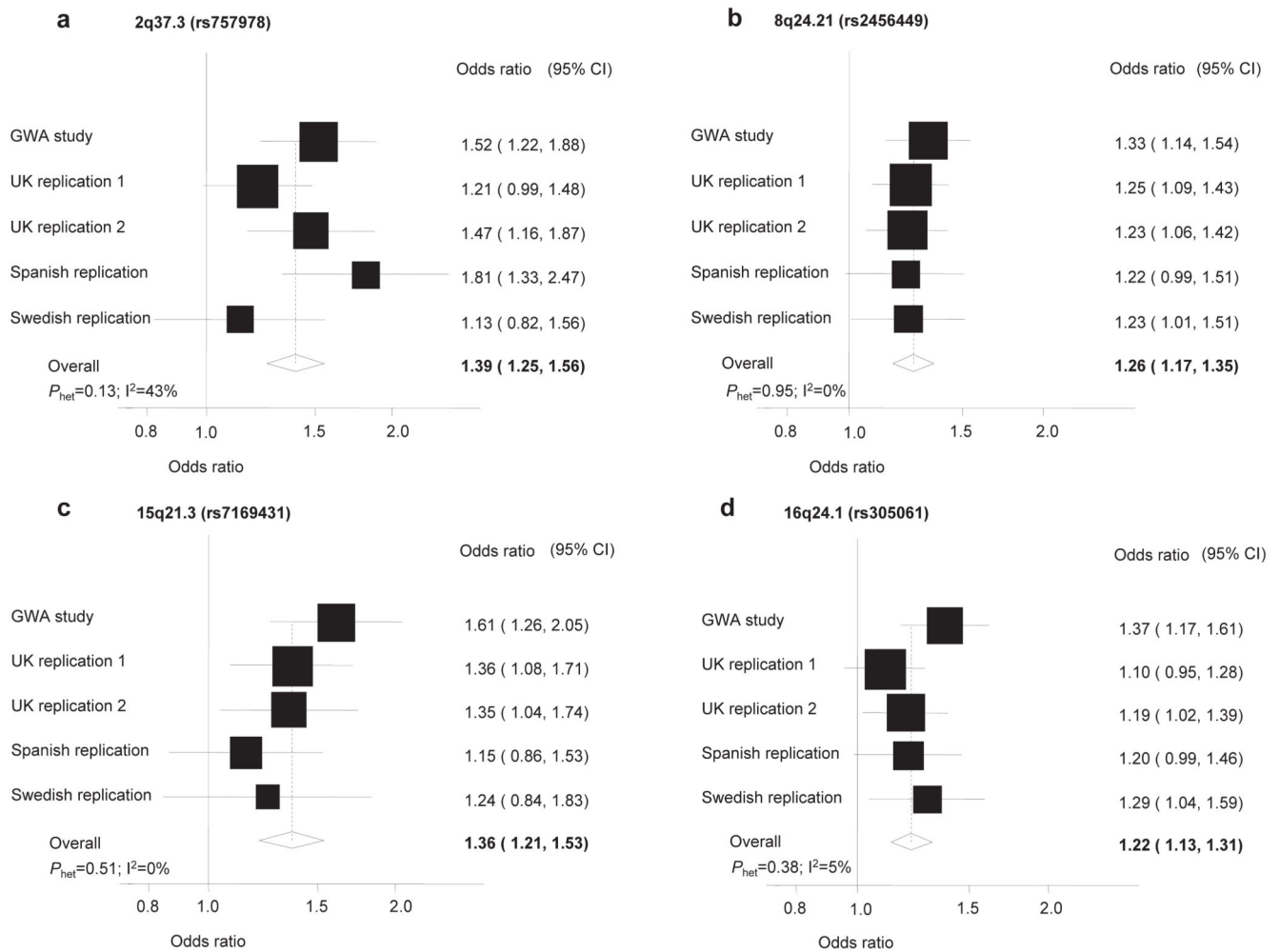


Figure 1. Forest plots of effect size and direction for the SNPs associated with CLL risk. (a) 2q37.3 (rs757978), (b) 8q24.21 (rs2456449), (c) 15q21.3 (rs7169431), (d) 16q24.1 (rs305061). Boxes denote OR point estimates, their areas being proportional to the inverse variance weight of the estimate. Horizontal lines represent 95% confidence intervals. The diamond (and broken line) represents the summary OR computed under a fixed effects model, with 95% confidence interval given by the width of the diamond. The unbroken vertical line is at the null value (OR = 1.0).

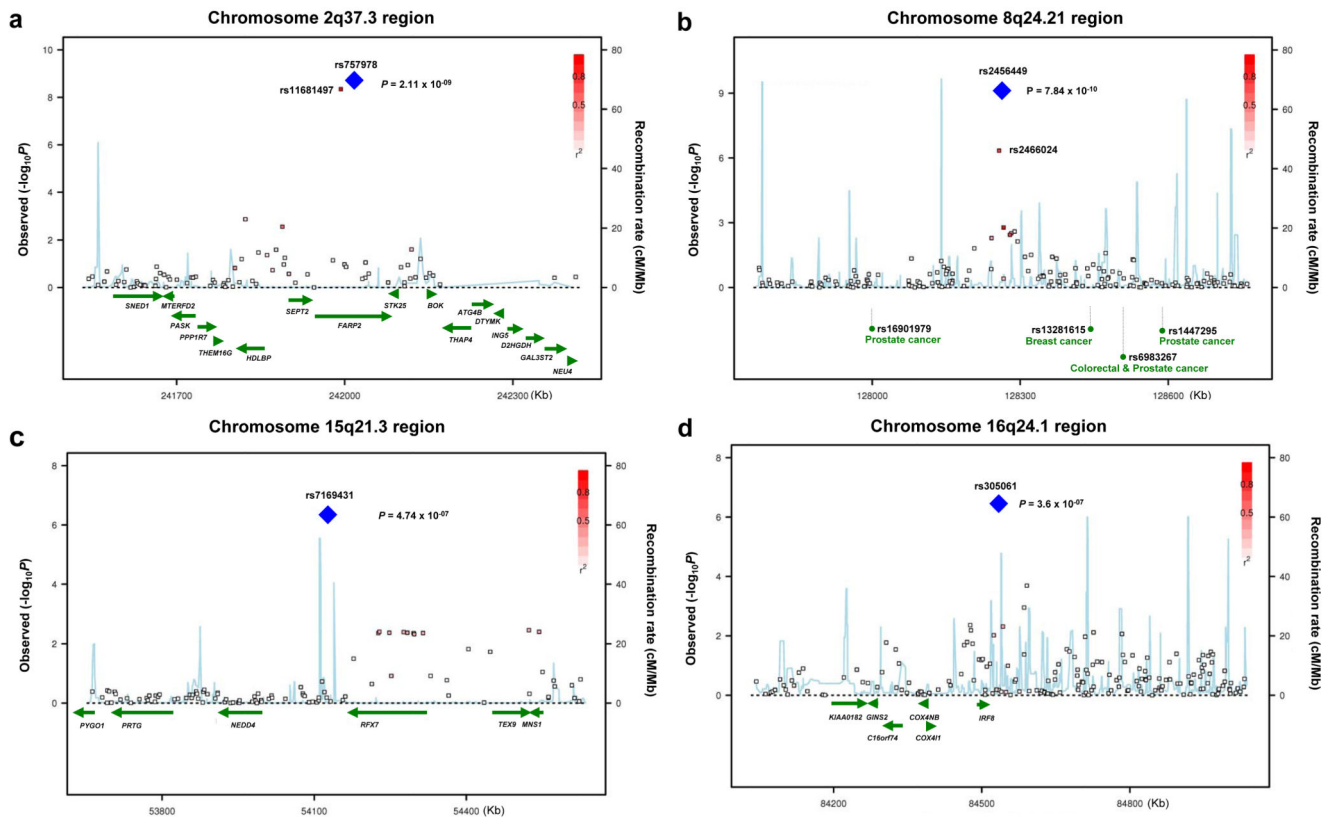


Figure 2. Four previously unidentified loci, 2q37.3, 8q24.21, 15q21.3, and 16q24.1 showing genome-wide level of evidence of association to CLL.

(a) Illustration of the **2q37.3** locus, with the local recombination rate plotted in light blue over this 600-kb chromosomal segment centered on **rs757978**. Each square represents a SNP found in this locus and the most associated SNP in the combined analysis, **rs757978**, is marked by a blue diamond. The color intensity of each square reflects the extent of LD with **rs757978** - red ($r^2 > 0.8$) through to white ($r^2 < 0.3$). Physical positions are based on build 36 of the human genome. **rs757978** is located in exon 9 of *FARP2*.

(b) Illustration of the **8q24.21** locus, with the most associated SNP in this locus, **rs2456449**, highlighted by a blue diamond. Here, we also present all SNPs found within a 600-kb window centered on **rs2456449** and define SNP colors based on LD with **rs2456449**.

(c) Illustration of the **15q21.3** locus, with the most associated SNP in this locus, **rs7169431**, highlighted by a blue diamond. Here, we also present all SNPs found within a 600-kb window centered on **rs7169431** and define SNP colors based on LD with **rs7169431**. *NEDD4* and *RFXDC2* map centromeric and telomeric to **rs7169431**.

(d) Illustration of the **16q24.1** locus, with the most associated SNP in this locus, **rs305061**, highlighted by a blue diamond. Here, we also present all SNPs found within a 600-kb window centered on **rs305061** and define SNP colors based on LD with **rs305061**. In this case, *IRF8* is the only gene found in the vicinity to the association signal.

Linkage disequilibrium maps are presented for all four loci in Supplementary Figure 1a–d online.

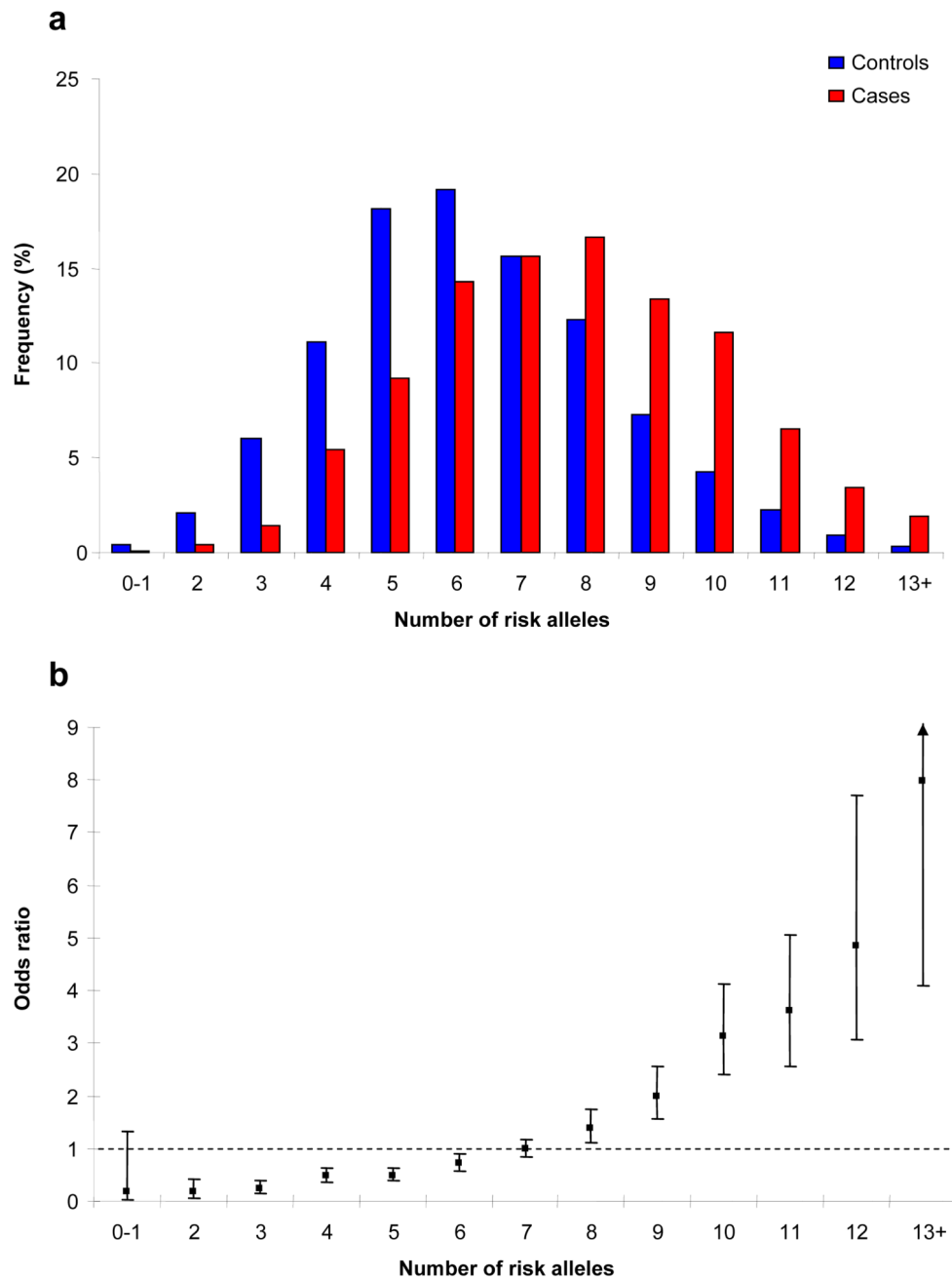


Figure 3. Cumulative impact of the 10 variants on CLL risk.

(a) Distribution of risk alleles in controls (blue bars) and CLL cases (red bars) for the 10 loci (rs757978, rs2456449, rs7169431 and rs305061 and the six previously identified loci2 - rs17483466, rs13397985, rs872071, rs735665, rs7176508, and rs11083846);

(b) Plot of the increasing ORs for CLL with increasing number of risk alleles. The ORs are relative to the median number of 7 risk alleles; Vertical bars correspond to 95% confidence intervals. The distribution of risk alleles follows a normal distribution in both case and controls, with a shift towards a higher number of risk alleles in cases. Analysis is

based on data from Stages 1, 2 and UK-replication series 2. Horizontal line denotes the null value (OR=1.0).