

Critical subgraphs of a random graph

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Abstract

We prove that the threshold for a random graph to have a k -core is equal to the threshold for having a subgraph which meets a necessary condition of Gallai for being k -critical.

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1 Introduction

In this paper, we examine the random graph $G_{n,M}$ formed by taking n vertices and choosing M edges where each of the $\binom{n}{M}$ possible edgesets is equally likely to be chosen. In particular, we will be concerned with the chromatic number of such a graph when $M = O(n)$.

Equivalently, we often discuss the *random graph process* in which we start with the graph with n vertices and no edges, and repeatedly add an edge chosen uniformly at random from amongst all edges not currently present. Note that $G_{n,M}$ is equivalent to the state of the random graph process after exactly M iterations.

One of the most tantalizing open problems in random graph theory (see for example [11]) is that of determining

$$d_k = \sup\{d \mid \text{a.s.}^1 \chi(G_{n,M=dn}) \leq k\},$$

where $\chi(G)$ is the chromatic number of G . As is the trend in the study of k -chromatic graphs, the case $k = 2$ is well-understood - $d_2 = 0$ - while the case $k \geq 3$ seems much

¹We say that $G_{n,p}$ *almost surely* (a.s.) has a property P if $\lim_{n \rightarrow \infty} \{\mathbf{Pr}(G_{n,p} \text{ has } P)\} = 1$.

more difficult. In fact, it was only recently shown that for $k \geq 3$ the threshold for k -colourability is sharp (see [1] and [14]).

If G is not k -colourable, then it must have a $(k+1)$ -critical subgraph, i.e. a subgraph $H \subseteq G$ such that $\chi(H) = k+1$, but $\chi(H - e) = k$ for any edge $e \in E(G)$. The most well known necessary conditions for a graph to be $(k+1)$ -critical are (see [9]):

- (a) it must have minimum degree at most k ;
- (b) it must be 2-connected.

Often a property P will a.s. occur for the first time during the random graph process at the exact step in which a weaker property P' first occurs. For example, a.s. the first graph to have a perfect matching will be the first graph to have minimum degree one [13], and a.s. the first graph to be Hamiltonian will be the first graph to have minimum degree two [17] (see also [8]). Given the nature of these two examples, it is natural to wonder if a.s. the first graph to be $(k+1)$ -chromatic will be the first graph to have a subgraph with minimum degree k .

Bollobás was the first to consider this approach to the problem of determining d_k . He defined the k -core of a graph to be its unique maximal subgraph with minimum degree at least k , if such a subgraph exists. Until recently, all lower bounds on d_k have been achieved by bounding the relaxation:

$$c_k = \sup\{c \mid \text{a.s. } G_{n, M=cn} \text{ has no } k\text{-core}\}.$$

Bollobás [6] proved that a.s. the first k -core to appear is k -connected, thus showing that there would be little benefit to incorporating condition (b) in our search for d_k . It was here that he first asked the much repeated question of whether $c_k = d_k$ for all $k \geq 3$.

In [10], Chvátal showed that for $c < c^* = 1.442\dots$, the expected number of subgraphs of $G_{n, M=cn}$ with minimum degree 3 is subexponentially small, while for $c > c^*$, the expected number of such subgraphs is exponentially large. An immediate corollary is that $c_3 \geq c^*$.

In [19], Molloy and Reed showed that $1.67 \leq c_3 \leq 1.78$, and that $d_3 \leq 2.571$. Shortly thereafter, Pittel, Spencer and Wormald [22] successfully determined

$$\begin{aligned} c_k &= \min_{y>0} y/2\pi_k(y) \\ &= \frac{1}{2}k + \frac{1}{2}\sqrt{k \log k} + O(\log k), \end{aligned}$$

with $\pi_k(y) = \mathbf{Pr}\{\text{Poisson}(y) \geq k-1\}$. For example, $c_3 = 1.67\dots$. Furthermore, the appearance of a k -core has a very sharp threshold.

Molloy [20] observed that this result can be used to show $d_k > c_k$ for $k \geq 4$. More recently, Achlioptas and Molloy [2] showed that $d_3 > 1.94\dots > c_3$, and so there is a significant gap between these two thresholds. Among other things, this indicates that the quest to determine d_k may require consideration of other properties of critical graphs.

This paper marks the first time that a nontrivial criticality condition is incorporated in the study of the chromatic number of a random graph.

Given a graph H with minimum degree at least k , we define the *low graph* of H , $L(H)$, to be the subgraph induced by the vertices of degree k . In [16] Gallai characterized the set of graphs which are the low graphs of $(k+1)$ -critical graphs:

Theorem L is the low graph of some $(k+1)$ -critical graph H , $k \geq 2$, iff

- (a) L has maximum degree at most k ;
- (b) each block of L is either a clique or a chordless odd cycle.

We say that a graph is k -Gallai if it has minimum degree k and it contains no even cycles whose vertices are all of degree k and do not induce a clique. Thus a graph with minimum degree k is k -Gallai iff its low graph satisfies the conditions of this theorem, and so it implies that any $(k+1)$ -critical graph must be k -Gallai.

A natural question to ask, particularly in light of the now known gap between c_k and d_k , is whether c_k is also the threshold for the appearance of a k -Gallai subgraph. In this paper we answer this question in the affirmative:

Theorem 1.1 For $k \geq 3$ and any $\epsilon > 0$, a.s. $G_{n, M=(c_k+\epsilon)n}$ has a k -Gallai subgraph.

One implication of this theorem is that in order to compute d_k , we may need to study even further properties of k -critical graphs.

We define the *edge-density* of a graph to be the ratio of the number of edges to the number of vertices in the graph. Let $\Omega_{n, M}^k$ be the set of all simple graphs with n vertices and M edges and with minimum degree k , and define $G_{n, M}^k$ to be a uniformly random member of $\Omega_{n, M}^k$. It is well-known (see, for example, [19] or [22]) that every member of $\Omega_{n, M}^k$ is equally likely to occur as the k -core of the random graph $G_{n', M'}$, for any n', M' , and so upon exposing the number of vertices and edges of the k -core of $G_{n, M}$, we can then choose the k -core from the $G_{n, M}^k$ model.

Define t_k to be the solution to

$$\frac{t_k^{k-1}}{(k-2)!} = e^{t_k} - \sum_{i=0}^{k-2} \frac{t_k^i}{i!} \quad (1)$$

and set

$$\ell_k = \frac{\sum_{i \geq k} \frac{it_k^i}{i!}}{2 \sum_{i \geq k} \frac{t_k^i}{i!}}.$$

For example, $\ell_3 = 1.7932\dots$, $\ell_4 = 2.5377\dots$, $\ell_5 = 3.2541\dots$

The main step towards proving Theorem 1.1 is the following:

Theorem 1.2

- (a) For $\ell < \ell_k$, $G_{n,M=\lceil \ell n \rceil}^k$ is k -Gallai with probability at least $1 - z^n$ for some constant $z < 1$.
- (b) For $\ell > \ell_k$, the probability that $G_{n,M=\lceil \ell n \rceil}^k$ is k -Gallai tends to a constant $0 < p_k(\ell) < 1$.

It is straightforward, using the results of [22] to determine the edge-density of the k -core when it first appears. Perhaps surprisingly, it is asymptotic to precisely ℓ_k :

Fact 1.3 For $k \geq 3$, a.s. the first k -core to appear during the random graph process has edge-density $\ell_k + o(1)$.

Theorem 1.1 will follow from Fact 1.3 and Theorem 1.2(b), along with a little more work. We will present this proof in the final section of the paper.

The crux of the proof of Theorem 1.2 lies in showing that ℓ_k is a threshold for $G_{n,M}^k$ in the sense of the celebrated Double Jump Threshold discovered by Erdős and Rényi. In [12] they show that the component structure of $G_{n,M=cn}$ undergoes a dramatic change at the point $c = \frac{1}{2}$. For $c < \frac{1}{2}$ a.s. all components of $G_{n,M=cn}$ are quite small, very few of them are cyclic, and none have more than one cycle. For $c > \frac{1}{2}$, a.s. $G_{n,M=cn}$ has a *giant component* on $\Theta(n)$ vertices and containing at least $\Theta(n)$ cycles.

In [21], Molloy and Reed showed that a similar phenomenon occurs for random graphs with a given degree sequence. This work was originally developed specifically to be applied to the results in this paper, and plays an important role here. Thus, before proceeding further, it is necessary to summarize the main theorem.

Definition 1.4 An asymptotic degree sequence is a sequence of integer-valued functions $\mathcal{D} = d_0(n), d_1(n), \dots$ such that

1. $d_i(n) = 0$ for $i \geq n$;
2. $\sum_{i \geq 0} d_i(n) = n$.

Given an asymptotic degree sequence \mathcal{D} , we set \mathcal{D}_n to be the degree sequence $\{c_1, c_2, \dots, c_n\}$, where $c_j \geq c_{j+1}$ and $|\{j : c_j = i\}| = d_i(n)$ for each $i \geq 0$. We say that \mathcal{D}_n is an *incident* of \mathcal{D} . We define $\Omega_{\mathcal{D}_n}$ to be the set of all graphs with vertex set $[n]$ with degree sequence \mathcal{D}_n . A random graph on n vertices with degree sequence \mathcal{D} is a uniformly random member of $\Omega_{\mathcal{D}_n}$.

Definition 1.5 An asymptotic degree sequence \mathcal{D} is **feasible** if $\Omega_{\mathcal{D}_n} \neq \emptyset$ for all $n \geq 1$.

Definition 1.6 An asymptotic degree sequence \mathcal{D} is **smooth** if there exist constants λ_i such that $\lim_{n \rightarrow \infty} d_i(n)/n = \lambda_i$.

Definition 1.7 Given a smooth asymptotic degree sequence, \mathcal{D} , $Q(\mathcal{D}) = \sum_{i \geq 1} i(i-2)\lambda_i$.

Definition 1.8 An asymptotic degree sequence \mathcal{D} is **sparse** if $\sum_{i \geq 0} id_i(n)/n = K + o(1)$ for some constant K .

We omit the definition of *well-behaved*, but suffice it to say that if for all n , $d_i(n) = 0$ whenever $i > \Delta$ for some fixed Δ then \mathcal{D} is well-behaved. This will be the case for any degree sequence considered in this paper (although the definition of *well-behaved* allows for more general situations and so the statement of this theorem is more general than is needed for this paper).

Theorem 1.9 Let $\mathcal{D} = d_0(n), d_1(n), \dots$ be a well-behaved sparse asymptotic degree sequence for which there exists $\epsilon > 0$ such that for all n and $i > n^{\frac{1}{4}-\epsilon}$, $d_i(n) = 0$. Let G be a graph with n vertices, $d_i(n)$ of which have degree i , chosen uniformly at random from amongst all such graphs. Then:

- (a) If $Q(\mathcal{D}) > 0$ then there exist constants $\zeta_1, \zeta_2 > 0$ dependent on \mathcal{D} such that G a.s. has a component with at least $\zeta_1 n$ vertices and $\zeta_2 n$ cycles. Furthermore, if $Q(\mathcal{D})$ is finite then G a.s. has exactly one component of size greater than $\gamma \log n$ for some constant γ dependent on \mathcal{D} .
- (b) If $Q(\mathcal{D}) < 0$ and for some function $0 \leq \omega(n) \leq n^{\frac{1}{8}-\epsilon}$, $d_i(n) = 0$ for all $i \geq \omega(n)$, then for some constant R dependent on $Q(\mathcal{D})$, G a.s. has no component with at least $R\omega(n)^2 \log n$ vertices, and a.s. has fewer than $2R\omega(n)^2 \log n$ cycles. Also, a.s. no component of G has more than one cycle.

Consistent with the model $G_{n,M}$, we call the component referred to in Theorem 1.9(a) the *giant component*.

As the reader has no doubt guessed, it is a simple matter to use Theorem 1.9 to show that for $\ell < \ell_k$, a.s. $L(G_{n,M=\lceil \ell n \rceil}^k)$ has a giant component, while for $\ell > \ell_k$, a.s. all components of $L(G_{n,M=\lceil \ell n \rceil}^k)$ are small. In the next section we do this and show that if $L(G_{n,M=\lceil \ell n \rceil}^k)$ has a giant component then with subexponentially high probability $L(G_{n,M=\lceil \ell n \rceil}^k)$ has an even cycle whose vertices do not induce a clique, and if all components of $L(G_{n,M=\lceil \ell n \rceil}^k)$ are small then the probability that no such cycle exists tends to $p_k(\ell)$, thus proving Theorem 1.2. The reader who is willing to accept these facts on faith may wish to skip Section 2. The reader whose faith is shaken because it seems counterintuitive that $L(G_{n,M=\lceil \ell n \rceil}^k)$ a.s. has a giant component when ℓ is *below* the threshold should note that when ℓ is small more vertices of $G_{n,M=\lceil \ell n \rceil}^k$ will have degree k , and so the low graph will tend to have higher edge-density.

In the final section, we prove Fact 1.3 and complete the proof of Theorem 1.1.

Throughout this paper, all asymptotics are taken as $n \rightarrow \infty$ and we only claim statements to hold for sufficiently large n . By $A \sim B$, we mean that $\lim_{n \rightarrow \infty} A/B = 1$.

2 A Gallai Threshold

In this section we will prove Theorem 1.2. As is common in the study of random graphs with restricted degree sequences, we will work with the configuration model, introduced in this form by Bollobás[5] and motivated in part by the work of Bender and Canfield[4]. This model arose in a somewhat different form in the work of Bekessy, Bekessy and Komlos[3] and Wormald[23, 24]. Given the degree sequence of a graph, we construct a random configuration with the same degree sequence by taking $\deg(v)$ copies of each vertex v , and then choosing a uniformly random pairing of the vertex-copies. Note that every configuration has an underlying multigraph with the desired degree sequence. Note further that every simple graph with that degree sequence occurs as the underlying multigraph with the same probability. As we shall see in the proof of Theorem 1.2(b), if the degree sequence has bounded maximum degree then the underlying multigraph of a random configuration is simple with probability asymptotic to $e^{-\mu_1-\mu_2}$ where μ_1, μ_2 are constants dependent on the degree sequence. An immediate and very useful consequence is the following:

Lemma 2.1 *If a random configuration, F , on a particular degree sequence with constant maximum degree a.s. has a property P , then a random (simple) graph, G , on the same degree sequence a.s. has P . Moreover, the probability that G does not have P is at most a constant multiple of the probability that F does not have P .*

We will use this lemma to allow us to work with the configuration model a number of times in this section.

Note that choosing a random configuration amounts to nothing more than selecting a uniformly random pairing of its vertices, and that we are free to choose this pairing any way we like so long as the distribution remains uniform. It is frequently useful to repeatedly take any unpaired vertex-copy we please and pair it with another randomly selected unexposed vertex-copy. In the proof of Theorem 1.9 in [21], we did this, each time being precise about the vertex-copy that we chose to pair. Essentially, we would expose the components of the underlying multigraph one at a time. If any vertices already known to be in the component had some unpaired copies, then we would choose one such copy (arbitrarily) to be the next one paired. At each step, we refer to such vertices as *partially exposed*, and we refer to the subgraph induced by the exposed edges of the component currently being exposed as a *partial component*. If there were no partially exposed vertices, then we would arbitrarily choose a vertex-copy, effectively starting a new component. We mention this here, because during the proof of Theorem 1(a) we will refer to a lemma from [21] regarding the evolution of the configuration during this exposure.

Now we begin our proof of Theorem 1.2. The first step is to show that λ_k is the threshold for $L(G_{n,M=\lambda n}^k)$ to a.s. have a giant component.

Lemma 2.2 For $\ell < \ell_k$, a.s. $L(G_{n,M=\lceil \ell n \rceil}^k)$ has a giant component on $\theta(n)$ vertices, while for $\ell > \ell_k$, a.s. all components of $L(G_{n,M=\lceil \ell n \rceil}^k)$ have size at most $O(\log n)$.

Proof First we expose the degree sequence of $G_{n,M=\lceil \ell n \rceil}^k$, letting d_i denote the number of vertices of degree i for each $i \geq k$. Using the results of [18] one can show (see [15] or [22]) that a.s. this degree sequence has an asymptotic truncated Poisson distribution with sharp concentration. More precisely, with $\nu_i = \frac{t^i}{i!}e^{-t}$ where t is chosen so that

$$\frac{\sum_{i \geq k} i \nu_i}{\sum_{i \geq k} \nu_i} = 2\ell,$$

a.s. $d_i = (\nu_i / \sum_{i \geq k} \nu_i)n + o(n)$, for each $i \geq k$. Furthermore, for any $\epsilon_1 > 0$ there exists $\zeta_1 < 1$ such that with probability at least $1 - \zeta_1^n$, $|d_k - \mu_k n| < \epsilon_1 n$.

Consider a random configuration, F , on this degree sequence. First we expose the degree sequence of $L(F)$, the low graph of the underlying multigraph of F . Each vertex-copy is paired to a copy of a vertex of degree k with probability $kd_k/2\ell n = \alpha$. Therefore a.s. a vertex of degree k in F has degree i in $L(F)$ with probability approximately $\beta_i = \binom{k}{i} \alpha^i (1 - \alpha)^{k-i}$. In fact it is straightforward to show that for each $0 \leq i \leq k$, the number of vertices of degree i in $L(F)$, b_i , is sharply concentrated around $\beta_i n$ in the sense that d_i is sharply concentrated, i.e. for any $\epsilon_2 > 0$ there exists $\zeta_2 < 1$ such that with probability at least $1 - \zeta_2^n$, $|b_i - \beta_i n| < \epsilon_2 n$. Therefore, if \mathcal{D}_n is the degree sequence of $L(F)$, then a.s. \mathcal{D}_n is an incident of an asymptotic degree sequence \mathcal{D} with

$$\begin{aligned} Q(\mathcal{D}) &= \sum_{i=1}^k i(i-2) \binom{k}{i} \alpha^i (1-\alpha)^{k-i} \\ &= k\alpha((k-1)\alpha - 1) \end{aligned}$$

Solving $\alpha = 1/(k-1)$ for ℓ yields $\ell = \ell_k$, showing that for $\ell < \ell_k$ ($\ell > \ell_k$), there exist $\epsilon = \epsilon(\ell) > 0, \zeta_3 = \zeta_3(\ell) < 1$ such that $Q(\mathcal{D}) > \epsilon$ ($Q(\mathcal{D}) < -\epsilon$) with probability at least $1 - \zeta_3^n$, and apply Theorem 1.9. ■

And now we prove that the threshold for $L(G_{n,M=\lceil \ell n \rceil}^k)$ to a.s. have a giant component is the threshold for $G_{n,M=\lceil \ell n \rceil}^k$ to a.s. not be Gallai, as described in the statement of Theorem 1.2.

Proof of Theorem 1.2(a): First we consider the case $\ell < \ell_k$. We will first expose the degree sequence $\mathcal{D} = \mathcal{D}_n$ of $L(G_{n,M=\lceil \ell n \rceil}^k)$. As shown in the preceding proof, there is some $\epsilon > 0, \zeta_3 < 1$ such that \mathcal{D}_n is an incident of an asymptotic degree sequence \mathcal{D} , where $Q(\mathcal{D}) > \epsilon$, with probability at least $1 - \zeta_3^n$. Let E_1 be the event that this happens. Note that, conditioning on E_1 , $L(G_{n,M=\lceil \ell n \rceil}^k)$ a.s. has a giant component and so, intuitively, it is not surprising that it is a.s. not k -Gallai.

Given the degree sequence of $L(G_{n,M=\lceil \ell n \rceil}^k)$, we expose a random configuration, F , on that degree sequence, in the manner used in [21] as described above. It is straightforward

to check that every simple graph with degree sequence \mathcal{D} is equally likely to occur as $L(G_{n,M=\lceil \ell n \rceil}^k)$, and so we are justified in using the configuration model here. After i pairs have been exposed, we let X_i denote the number of unexposed copies of partially exposed vertices. If E_1 holds, then by Lemma 9 of [21] there exists $I, \gamma > 0, \zeta_4 < 1$ such that with probability at least $1 - \zeta_4^n$, $X_I \geq \gamma n$. Let E_2 be the event that this happens.

If E_2 holds, consider any spanning tree T of the partial component currently being exposed. Because the maximum degree in T is at most k , it is easy to show that T has a vertex v such that, considering T to be rooted at v , v has a child v_1 such that the subtree rooted at v_1 has at least $\frac{\gamma n}{2^k}$ but not more than $\frac{2\gamma n}{k}$ open vertex-copies. Furthermore, we can repeatedly choose a child v_{i+1} of v_i such that the subtree rooted at v_{i+1} has at least $\frac{\gamma n}{2^{k^i}}$ open vertex-copies. Setting $u = v_{k+1}$, and letting P be the unique uv -path in T , we have that P has length $k + 1$, and upon deleting the edges of P from T , the components containing u and v , T_u, T_v , each have at least $(\gamma/2k^{k+1})n$ open vertex-copies. Similarly, T_u contains a path P_u of length $k + 1$ such that the components of $T_u - P_u$ containing the endpoints of P_u, T_u^1, T_u^2 each have at least $(\gamma/4k^{2k+2})n$ open vertex-copies.

Therefore, there exists $\zeta_5 < 1$ such that at least one pair of the configuration will contain a point from each of T_u^1, T_v , and at least one pair will contain a point from each of T_u^2, T_v , with probability at least $1 - \zeta_5^n$. The two cycles induced in T by these edges intersect in P , and so their union must contain an even cycle of length at least $k + 2$. Since F has maximum degree k , it contains no $(k + 2)$ -clique, and so it cannot be k -Gallai.

Therefore, applying Lemma 2.1 the probability that $L(G_{n,M=\lceil \ell n \rceil}^k)$ is k -Gallai is at most

$$\Pr\{L(G_{n,M=\lceil \ell n \rceil}^k) \text{ is } k\text{-Gallai} \mid E_1 \wedge E_2\} + \Pr\{\bar{E}_2 \mid E_1\} + \Pr\{\bar{E}_1\} < \zeta^n,$$

for some $\zeta > 0$. ■

To prove Theorem 1.2(b), we will essentially show that the number of even cycles in $L(G_{n,M=\lceil \ell n \rceil}^k)$ is asymptotically equivalent to a Poisson variable with mean μ . The result then follows by setting $p_k(\ell) = e^{-\mu}$.

Proof of Theorem 1.2(b): First we expose the degree sequence of $L(G_{n,M=\lceil \ell n \rceil}^k)$, and then we choose $L(G_{n,M=\lceil \ell n \rceil}^k)$ by taking a random graph on that degree sequence. Again, we work with the configuration model, taking a random configuration F on the same degree sequence.

As in Lemma 2.2, a.s. the number of vertices of degree i is $\lambda_i n + o(n)$ for some $\lambda_0, \dots, \lambda_k > 0$ such that $\sum_{i=1}^k i(i-2)\lambda_i < \epsilon = \epsilon(\ell)$, and we let E_3 be the event that this holds.

For constant $r \geq 1$, let C_r denote the number of cycles of length r in F . We will show that C_1, C_2, \dots are asymptotic to independent Poisson variables with means μ_1, μ_2, \dots

The first step is to compute μ_r . Define $K = \sum_{i \geq 1} i\lambda_i$.

$$\begin{aligned} \mu_r &\approx \frac{r!}{(Kn)^r} \sum_{a_2+\dots+a_k=r} \prod_{i=2}^k \binom{\lambda_i n}{a_i} (i(i-1))^{a_i} \\ &\approx \frac{r!}{K^r} \sum_{a_2+\dots+a_k=r} \prod_{i=2}^k \frac{(i(i-1)\lambda_i)^{a_i}}{a_i!} \\ &= \frac{r!}{K^r} [x^r] \prod_{i=2}^k e^{i(i-1)\lambda_i x} \\ &= \left(\frac{\sum_{i=2}^k i(i-1)\lambda_i}{K} \right)^r. \end{aligned}$$

Note that if E_3 holds, then μ_r tends to zero as r grows, since $\sum_{i \geq 1} i(i-2)\lambda_i < 0$ implies that $\rho = \sum_{i=2}^k i(i-1)\lambda_i < K$. Note further that for any $r = r(n)$, the expected number of cycles of length r in F is at most ρ^r .

The next step is to compute the second moment of C_r . To do this we will compute Δ_{r_1, r_2} , the sum over all pairs of intersecting cycles in K_n of length r_1, r_2 of the probability that they both appear in $L(G_{n, M=\lceil \ell n \rceil}^k)$.

To compute Δ , we sum first over all cycles, H_1 , of length r_1 , and then over all intersecting cycles H_2 of length r_2 . Noting that $H_2 - E(H_1)$ is a collection of a paths for some $a \geq 1$, we sum over the number of choices for these paths, setting l_1, \dots, l_a to be their lengths.

$$\begin{aligned} \Delta_{r_1, r_2} &\leq \sum_{a_2+\dots+a_k=r_2} \prod_{i=2}^k \binom{\lambda_i n}{a_i} (i(i-1))^{a_i} \\ &\quad \times \sum_{a=1}^{r_2} \sum_{\substack{l_1, \dots, l_a \geq 1 \\ l_1+\dots+l_k < r_2}} \prod_{j=1}^a \sum_{b_2+b_3+\dots=b_j-1} \binom{r_1}{2} k^2 (k-1)^2 \frac{(l_j-1)!}{(Kn)^{l_j}} \\ &\quad \times \prod_{i=2}^k \binom{\lambda_i n}{b_i} (i(i-1))^{b_i} \\ &\leq \rho^{r_1} \sum_{a=1}^{r_2} \binom{r_1}{2} k^2 (k-1)^2 n^{-1} \left(\sum_{l=1}^{r_2-1} \rho^l \right)^a \\ &= O(n^{-1}) \end{aligned}$$

Therefore, $\mathbf{Exp}(C_r^2)$ is asymptotic to the expected number of pairs of non-intersecting cycles of length r in F which is easily calculated to be $\frac{\mu_r^2}{2} + o(1)$. Similarly, for $r_1 \neq r_2$, $\mathbf{Exp}(C_{r_1} C_{r_2}) \approx \mu_{r_1} \mu_{r_2}$.

Similar calculations show that the expected number of t -tuples of non-edge-disjoint cycles of lengths r_1, \dots, r_t is $o(1)$ for any $t \geq 2$, and so for any x_1, \dots, x_s ,

$\mathbf{Exp}(C_{r_1}^{x_1}, \dots, C_{r_s}^{x_s}) \approx \prod_{i=1}^s \frac{\mu_{r_i}^{x_i}}{x_i!}$, and so C_0, C_1, \dots are asymptotically independent Poisson variables as claimed.

Note that for any $\delta > 0$, there exists $R = R(\delta)$ such that the expected number of cycles of length at least R is less than δ and so, by Markov's Inequality, the probability that F contains a cycle of length at least R is less than δ . Therefore, the probability that F contains no even cycle of length at least 4 is asymptotic to

$$p_k(\ell) = e^{-\sum_{r \geq 2} \rho^{2r}}.$$

Furthermore, this event is asymptotically independent of the event that F contains no cycles of length 1 or 2, and so the probability that a random graph on the same degree sequence has no even cycle is asymptotic to $p_k(\ell)$.

Define E_4 to be the event that $G_{n, M=\lceil \ell n \rceil}^k$ has no 4-clique. It is straightforward to compute that $\mathbf{Pr}(E_4) > 1 - n^{-1}$. Therefore, again applying Lemma 2.1,

$$\begin{aligned} |\mathbf{Pr}\{G_{n, M=\lceil \ell n \rceil}^k \text{ is } k\text{-Gallai}\} - p_k(\ell)| &\leq \mathbf{Pr}(\bar{E}_3) + \mathbf{Pr}(\bar{E}_4) \\ &= o(1). \end{aligned}$$

■

3 The first k -core to appear

Here, we see that a.s. the first k -core to appear has edge density precisely ℓ_k .

Proof of Fact 1.3: It is implicit in [22] that a.s. the edge-density of the first k -core to appear is:

$$\frac{\sum_{i \geq k} \frac{i \tau_k^i}{i!}}{2 \sum_{i \geq k} \frac{\tau_k^i}{i!}} + o(1),$$

where τ_k minimizes

$$\frac{\tau}{1 - e^{-\tau} \sum_{i=0}^{k-2} \frac{\tau^i}{i!}}.$$

Thus, it is our goal to show that $\tau_k = t_k$. To solve for τ_k , we set $f(\tau) = (1 - e^{-\tau} \sum_{i=0}^{k-2} \frac{\tau^i}{i!})/\tau$, yielding

$$\begin{aligned} f'(\tau) &= -\frac{1}{\tau^2} - \left(-e^{-\tau} \left(\sum_{i=0}^{k-2} \frac{\tau^{i-1}}{i!} \right) + e^{-\tau} \left(-\frac{1}{\tau^2} + \sum_{i=0}^{k-4} \frac{\tau^i}{(i+2)i!} \right) \right) \\ &= -\frac{e^{-\tau}}{\tau^2} \left(e^\tau - \left(1 + \tau + \sum_{i=2}^{k-2} \tau^i \left(\frac{1}{(i-1)!} - \frac{1}{i(i-2)!} \right) + \frac{\tau^{k-1}}{(k-2)!} \right) \right) \\ &= -\frac{e^{-\tau}}{\tau^2} \left(e^\tau - \sum_{i=0}^{k-2} \frac{\tau^i}{i!} - \frac{\tau^{k-1}}{(k-2)!} \right), \end{aligned}$$

and so $\tau_k = t_k$. ■

And this yields our main theorem:

Proof of Theorem 1.1: Consider any $\epsilon > 0$ and set $M = (c_k + \epsilon)n$. By the main result of [22], $G_{n,M}$ a.s. has a k -core, and by Fact 1.3, the edge-density of the k -core is at least $t_k + \delta$ for some $\delta > 0$. Let T denote this k -core.

Once we have exposed n' , the number of vertices, and M' , the number of edges in T , it is a uniformly random member of $\Omega_{n',M'}^k$ as each member of $\Omega_{n',M'}^k$ is the k -core of the same number of graphs with n vertices and M edges. Therefore Theorem 1.2(b) can be applied to T .

Define $T_1 \subset T$ to be the subgraph of T induced by the components of $L(T)$ which contain even cycles. Recall that in the proof of Theorem 1.2(b) we show that the number of even cycles in a random configuration with the same degree sequence as $L(T)$ has an asymptotically Poisson distribution. In the same manner, it is straightforward to show that (a) a.s. $|T_1| < |T|$, (b) a.s. no vertex $v \in T$ of degree at least $k + 1$ has more than 1 neighbour in T_1 , and finally (c) a.s. the subgraph induced by $(T \cup N(T_1)) \setminus T_1$ contains no even cycles. This follows from a straightforward first-moment analysis; we omit the details. Intuition towards the truth of this statement comes from the fact that the expected size of T_1 is $O(1)$.

Therefore, a.s. $T \setminus T_1$ is k -Gallai. ■

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