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1	Cryptic diversity and discordance in single-locus species delimitation methods within horned
2	lizards (Phrynosomatidae: Phrynosoma)
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23 Abstract

24 Biodiversity reduction and loss continues to progress at an alarming rate, and thus there is 25 widespread interest in utilizing rapid and efficient methods for quantifying and delimiting 26 taxonomic diversity. Single-locus species-delimitation methods have become popular, in part due 27 to the adoption of the DNA barcoding paradigm. These techniques can be broadly classified into 28 tree-based and distance-based methods depending on whether species are delimited based on a 29 constructed genealogy. Although the relative performance of these methods has been tested 30 repeatedly with simulations, additional studies are needed to assess congruence with empirical 31 data. We compiled a large data set of mitochondrial ND4 sequences from horned lizards 32 (Phrynosoma) to elucidate congruence using four tree-based (single-threshold GMYC, multiple-33 threshold GMYC, bPTP, mPTP) and one distance-based (ABGD) species delimitation models. 34 We were particularly interested in cases with highly uneven sampling and/or large differences in 35 intraspecific diversity. Results showed a high degree of discordance among methods, with 36 multiple-threshold GMYC and bPTP suggesting an unrealistically high number of species (29 37 and 26 species within the P. douglasii complex alone). The single-threshold GMYC model was 38 the most conservative, likely a result of difficulty in locating the inflection point in the 39 genealogies. mPTP and ABGD appeared to be the most stable across sampling regimes and 40 suggested the presence of additional cryptic species that warrant further investigation. These 41 results suggest that the mPTP model may be preferable in empirical data sets with highly uneven 42 sampling or large differences in effective population sizes of species.

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44 Key Words: ABGD, GMYC, ND4, Phrynosoma, PTP, speciation

45 Introduction

46

47 Proper identification and delimitation of species is of utmost importance for most fields of 48 biology (de Queiroz 2007). Many research programs address fundamental questions through a 49 comparative framework, necessitating the use of both a robust phylogeny and accurate species 50 assignments for hypothesis testing. Molecular-based species-delimitation methods can generally 51 be classified into single- or multilocus, and discovery and validation-based techniques (Carstens 52 et al. 2013). DNA barcoding threshold methods (Hebert et al. 2003, 2004; Hebert & Gregory 53 2005; Edgar 2010; Puillandre et al. 2012a) comprise one common example of a single-locus 54 technique, where threshold or cut-off values are used to differentiate inter- from intraspecific 55 divergences. The Refined Single Linkage (RESL) method, for example, is a popular clustering 56 algorithm implemented within the Barcode of Life Data Systems (Ratnasingham & Hebert 2013) 57 to delineate operational taxonomic units (OTUs) based on animal COI barcode data. Although 58 these threshold-type methods continue to be a quick and effective way to document and describe 59 diversity, they do not take into account tree structure and often rely on arbitrarily defined 60 thresholds (e.g., 2–3% pairwise sequence divergence) to delimit species (Blaxter 2004; Hebert & 61 Gregory 2005; Hamilton et al. 2011).

More recently, coalescent-based methods of species delimitation have become common, in part due to the continual ease in which researchers can generate vast quantities of molecular data (Leaché & Fujita 2010; Fujita *et al.* 2012). Bayesian multilocus coalescent methods, for example, can explicitly account for gene tree/species tree incongruence when delimiting species and estimating a species tree (Yang & Rannala 2014; Jones 2014). Although an attractive alternative to threshold-type methods, the utility of many Bayesian multilocus coalescent

68	methods for large data sets remains uncertain due to the relatively large computational demand
69	of the algorithms (Yang & Rannala 2010, 2014; Satler et al. 2013; Leaché et al. 2014).
70	Single-locus, coalescent-based methods like the General Mixed Yule Coalescent model
71	(GMYC; Pons et al. 2006; Fujisawa & Barraclough 2013) have become a popular tree-based
72	species-delimitation technique often applied to barcoding data (e.g., animal mitochondrial DNA).
73	The GMYC model uses maximum likelihood and an ultrametric gene tree to model the transition
74	between inter- and intraspecific branching patterns. Branching patterns older than the inferred
75	threshold represent speciation events (Yule process), whereas younger branching indicates
76	neutral coalescence within species. GMYC has been used in numerous empirical studies (e.g.,
77	Monaghan et al. 2009; e.g. Barraclough et al. 2009; Hamilton et al. 2011; Gebiola et al. 2012;
78	Esselstyn et al. 2012; Blair et al. 2015), and recent simulations and empirical data suggest that
79	the method is fairly robust to different assumptions (Reid & Carstens 2012; Esselstyn et al. 2012;
80	Fujisawa & Barraclough 2013; Talavera et al. 2013; Tang et al. 2014). The Poisson Tree
81	Processes (PTP/bPTP) model is similar in that it seeks to model the transition in branch lengths
82	between versus within species (Zhang et al. 2013). However, PTP estimates branching processes
83	using the expected number of substitutions (versus time in GMYC), and thus utilizes a non-
84	ultrametric phylogenetic tree as input. One limitation of the original PTP model is that it assumes
85	only two independent distributions to model branch lengths (one exponential distribution for
86	speciation and one exponential distribution for coalescence). This generally ignores the
87	stochastic variation among species due to different population sizes and demographic histories.
88	Conversely, the recently-developed multi-rate Poisson Tree Processes model (mPTP) fits
89	multiple independent exponential distributions to each delimited species to explicitly account for
90	differences in sampling intensity and/or population history (Kapli et al. 2016). Although the

91 mPTP model may potentially lead to more accurate delimitations versus other single-locus
92 methods, testing and comparison using empirical data characterized by highly heterogeneous
93 sampling intensity and/or large differences in genetic diversity among species is lacking.

94 Horned lizards (*Phrynosoma*) are a genus of phrynosomatid lizards consisting of 17–21 95 species distributed from Canada to Guatemala (Leaché & Linkem 2015; Montanucci 2015). The 96 unique morphology and behavior of these lizards, including ocular blood squirting (Sherbrooke 97 2003), has made them the target of numerous systematic investigations. Early studies based on 98 mitochondrial DNA (mtDNA) sequences yielded conflicting phylogenetic relationships with 99 both nuclear and morphological data (e.g. Hodges and Zamudio, 2004; Leaché and McGuire, 100 2006; Reeder and Montanucci, 2001), presumably because of mtDNA introgression (Leaché & 101 McGuire 2006; Mulcahy et al. 2006). More recently, next-generation sequencing has been used 102 to estimate a robust phylogeny for the genus, with results suggesting that cladogenesis initiated 103 in the Miocene around 22 million years ago (Leaché & Linkem 2015). Phylogeographic studies 104 have also been conducted on several species including P. douglasii (Zamudio et al. 1997), P. 105 mcallii (Mulcahy et al. 2006), P. platyrhinos (Jezkova et al. 2016), P. coronatum (Leaché et al. 106 2009), and P. orbiculare (Bryson et al. 2012). Many of these studies have indicated the presence 107 of undocumented cryptic diversity, but whether any of this diversity may warrant species status 108 has not been evaluated in detail for most groups (but see Leaché et al. 2009). These studies have 109 also revealed substantially different levels of intraspecific diversity, ranging from deep lineages 110 within P. orbiculare to low levels of genetic diversity in P. mcallii. More recent taxonomic work 111 on the genus has suggested that levels of diversity may be underestimated at a clade-wide level 112 (Nieto-Montes de Oca et al. 2014; Montanucci 2015).

113 Given the need for additional empirical studies to compare and contrast single-locus 114 "discovery-based" species delimitation methods, particularly in cases with large differences in 115 sampling intensity and levels of intraspecific diversity, in this study we assess congruence among 116 four tree-based and one distance-based methods of delimiting horned lizard species. We 117 particularly focus on the utility of one method (mPTP) with the ability to accommodate highly 118 heterogeneous data sets comprised of species with dramatically different levels of molecular 119 diversity. Based on our results, we provide further quantitative evidence for undescribed species 120 in the genus. 121 122 **Materials and Methods** 123 124 Data collection 125 We obtained from GenBank a total of 368 orthologs of the mitochondrial ND4 gene from 126 multiple representatives of *Phrynosoma*. Although many molecular phylogenetic studies have 127 been performed on the genus, there have been relatively few phylogeographic investigations of 128 species using the same marker. We extracted from GenBank multiple sequences from P. mcallii 129 (Mulcahy et al. 2006), P. platyrhinos (Jezkova et al. 2016), P. orbiculare (Bryson et al. 2012), 130 and the *P.douglasii* complex (Zamudio et al. 1997). We included singletons of several other 131 species, including P. cornutum, P. coronatum, P. solare, P. asio, and P. taurus. Although the 132 COI gene is the generally accepted standard for most animal barcoding studies (Hebert *et al.* 133 2003; Ratnasingham & Hebert 2013), ND4 likely serves as a good proxy due to the linked 134 inheritance of mtDNA. Full sampling information can be found in Supplementary Table S1. All 135 gene trees and alignments used in this study can be found on Dryad (doi:10.5061/dryad.r7989).

136

137 Phylogenetic analysis

138 The primary goal of our study was to test multiple methods of single-locus species delimitation 139 of horned lizards based on ND4 GenBank data. We were particularly interested in testing for 140 similarities and differences in those methods requiring an ultrametric tree (GMYC-type methods) 141 versus those that do not rely on temporal calibration (PTP-type methods). We were also 142 interested in the performance of both types of methods under scenarios with divergent rates of 143 coalescence across species. We tested a total of four tree-based methods on the data, including 144 the single-threshold GMYC (sGMYC; Pons et al. 2006; Fujisawa & Barraclough 2013), 145 multiple-threshold GMYC (mGMYC; Monaghan et al. 2009), bPTP (Zhang et al. 2013), and 146 mPTP (Kapli et al. 2016). All data were aligned using MUSCLE v.3.8.31 (Edgar 2004) 147 implemented in AliView v.1.17.1 (Larsson 2014). The total alignment consisted of 871 bp, 148 although sequence lengths for the *P. douglasii* complex were shorter (alignments available on 149 Dryad). Prior to species-delimitation analyses, we used RAxML v.8.0.0 (Stamatakis 2014) to 150 remove duplicate haplotypes from a matrix of 368 sequences. This left a total of 220 haplotypes 151 for species delimitation. Although identical sequences should generally be removed prior to tree-152 based methods of species delimitation (J. Zhang, pers. comm.), we also performed a series of 153 duplicate analyses using all 368 sequences for comparison. We used MEGA7 (Kumar et al. 154 2016) to calculate average within-species genetic distances using the Tamura-Nei model with 155 gamma distributed rate heterogeneity to account for multiple substitutions. In addition, we used 156 the R-packages APE (Paradis et al. 2004) and PEGAS (Paradis 2010) to calculate Watterson's 157 estimator of theta ($\theta = 4N_e\mu$) as an indicator of effective population sizes.

158	We used BEAST v.2.4.3 (Bouckaert et al. 2014) to generate ultrametric gene trees under
159	a strict clock and constant-size coalescent tree prior following the relative performance of clock
160	models in previous studies (Monaghan et al. 2009; Satler et al. 2013; Talavera et al. 2013). A
161	GTR+I+ Γ model of substitution was used as estimated using BIC in jModelTest2 (Darriba <i>et al.</i>
162	2012). We calibrated the rate of mtDNA substitution by specifying a normal prior with a mean of
163	0.00805 substitutions/site/million years and sigma of 0.001 (Bryson et al. 2012), a rate initially
164	estimated using vicariant scenarios in geckos (Macey et al. 1999). Similar substitution rates have
165	been applied in numerous studies to estimate divergence times in several vertebrate groups (see
166	Macey et al. 1999; Bryson et al. 2012 for examples), and our specified prior distribution
167	accommodated uncertainty in the estimate. Analyses were run for 20 million generations,
168	sampling every 2,000. Convergence and mixing were monitored in Tracer v.1.6 (Rambaut et al.
169	2014), and ESS values >200 indicated adequate sampling of the posterior. TreeAnnotator v.2.4.3
170	(Bouckaert et al. 2014) was used to create a maximum clade credibility (MCC) tree using mean
171	heights for node annotation. Maximum-likelihood (ML) phylogenetic analyses were
172	implemented in RAxML under a GTRGAMMA model by first implementing a rapid bootstrap
173	search (Stamatakis et al. 2008) with autoMRE bootstopping followed by a full ML search (-f a
174	option). Trees were rooted using P. asio (Leaché & Linkem 2015).
175	

176 Species delimitation analyses

177 Among the multiple methods of single-locus species delimitation currently available, the most

- 178 popular is the GMYC model (Pons et al. 2006; Fujisawa & Barraclough 2013). Previous studies
- 179 indicate that the GMYC model is fairly robust (Fujisawa & Barraclough 2013), especially if
- applied to an ultrametric tree constructed using BEAST (Tang *et al.* 2014), and that the choice of

181 clock model and tree prior has a relatively low impact on the results (Talavera *et al.* 2013). We 182 used the R package SPLITS (Ezard et al. 2009) to fit both the single- and multiple-threshold 183 models to the data. We initially included haplotypes only for species containing multiple 184 sequences (P. mcallii, P. platyrhinos, P. orbiculare, P. hernandesi, P. douglasii complex). 185 However, sGMYC results were not significantly different from the null model of coalescence. 186 Thus, we added singletons of P. cornutum, P. coronatum, P. solare, P. asio, and P. taurus to 187 increase the Yule portion of the tree and better fit the model to the data (Talavera *et al.* 2013). 188 bPTP analyses were performed using the online server (http://species.h-its.org/) and the 189 ML trees from RAxML. We ran the analyses for 500,000 generations with a thinning of 500 and 190 burnin of 0.1. Convergence was assessed by visualizing plots of MCMC iteration versus log-191 likelihood. We ran analyses both with and without the outgroup taxon (P. asio). As results were 192 qualitatively similar (not shown), all subsequent comparisons were made with the outgroup to 193 negate taxonomic discrepancy among analyses. We compared the results from bPTP to the 194 recently developed mPTP model that accommodates different rates of coalescence within clades 195 (Kapli *et al.* 2016). Discordant coalescent patterns could be due to uneven sampling intensity 196 among species or varying degrees of genetic structure arising from differences in evolutionary 197 processes and effective population sizes (N_e). We performed both ML and MCMC analyses on 198 RAxML ML trees using the standalone mPTP software (v.0.1.1). MCMC analyses were run for 199 100 million generations, sampling every 10,000. The first 2 million generations were discarded 200 as burnin and analyses started from the ML species delimitation estimate (identical results were 201 obtained when starting from both random and null delimitations). Convergence was again 202 assessed by monitoring the plot of generation versus log-likelihood. Both ML and MCMC 203 analyses utilized the --multi option to incorporate differences in rates of coalescence among

species and used a minimum branch length of 0.0001. We compared results among multiple
MCMC runs (10) to assess congruence.

206 Because our taxonomic sampling for the four target taxa was highly uneven (i.e., P. 207 *platyrhinos* was represented by ~3x the number of haplotypes), we re-ran all of the above 208 analyses after pruning haplotypes from P. platyrhinos to determine the potential influence of 209 highly heterogeneous sampling intensity on species delimitation. Haplotypes were selected (40 210 of 111 total) based on the previous phylogenetic analyses to maximize diversity within the 211 species. A new model of substitution was then calculated (HKY+I+ Γ) and all phylogenetic and 212 species delimitation analyses were repeated as described above. However, only unique 213 haplotypes (149) were included in this set of analyses.

214 To compare the results from the tree-based methods above, we ran Automatic Barcode 215 Gap Discovery (ABGD; Puillandre *et al.* 2012a) on both the full and pruned data sets. ABGD is 216 a computationally efficient distance-based method of species delimitation that has been shown to 217 perform well when compared to tree-based coalescent methods (Puillandre et al. 2012b; 218 Kekkonen & Hebert 2014; Kapli et al. 2016) and other threshold techniques (Ratnasingham & 219 Hebert 2013). The method seeks to quantify the location of the barcode gap that separates intra-220 from interspecific distances. As the presence of singletons may bias the analysis (Puillandre et al. 221 2012a), ABGD analyses were restricted to P. mcallii, P. orbiculare, P. platyrhinos, and the P. 222 douglasii complex. Default settings were used for the prior range for maximum intraspecific 223 divergence (0.001, 0.1). Results were compared using both JC69 and K80 corrected distances 224 and minimum slope increase (X) of 1.5 (default) and 1.0. 225 For all analyses, we reported the number of delimited species inferred by each method

along with the corresponding confidence intervals. In addition, we used current horned lizard

taxonomy (Leaché & Linkem 2015; Montanucci 2015), to compare the proportion of delimited
species matching taxonomic species, the proportion of taxonomic species lumped into a
delimited species, and the number of taxonomic species splits. We note that large values for
species lumps and splits may not necessarily indicate poor performance of methods, but they do
provide evidence that horned lizard taxonomy might be in need of revision. Next, we reported
the match ratio (following Ahrens *et al.* 2016) using the following formula:

233
$$match ratio = 2 * \frac{N_{match}}{(N_{delimited} + N_{morph})}$$

234 where N_{match} is the number of delimited species exactly matching taxonomic species, $N_{\text{delimited}}$ is 235 the total number of delimited species, and N_{morph} is the number of taxonomic, morphologically 236 defined species. Finally, we quantified performance of methods using the recently developed 237 Relative Taxonomic Resolving Power Index (R_{tax}) and the Taxonomic Index of Congruence 238 (C_{tax}) following Miralles & Vences (2013). The R_{tax} index quantifies the relative power of a 239 method to infer all estimated speciation events present in a data set (large R_{tax} = small Type-II 240 error), but does not necessarily imply correct delimitations (i.e. can lead to oversplitting). R_{tax} 241 metrics were calculated as follows:

242
$$Rtax(A) = \frac{nA}{n(A \cup B \cup C \cup D \cup E)}$$

where A,B,C,D,E represent the five species delimitation methods tested, the numerator (nA) represents the number of speciation events inferred by method A, and the denominator represents the cumulative number of speciation events inferred by all methods. An R_{tax} value of 1 would indicate that the method recovered all speciation events present across methods. The C_{tax} index is a measure of congruence in species assignments among two methods, with a

248 value of 1 indicating complete congruence. C_{tax} metrics were calculated as follows:

249
$$Ctax(AB) = \frac{n(A \cap B)}{n(A \cup B)}$$

250	where $A \cap B$ represents the number of speciation events shared by methods A and B, and $A \cup B$
251	represents the total number of speciation events inferred by method A and/or B . We refer the
252	reader to the original publication for additional descriptions of these metrics (Miralles & Vences
253	2013).
254	
255	Results and Discussion
256	
257	Uneven sampling
258	Inferred genealogies were congruent with previous studies based on mtDNA (Reeder &
259	Montanucci 2001; Hodges & Zamudio 2004; Leaché & McGuire 2006) and discordant from
260	nuclear phylogenies (Leaché & Linkem 2015), presumably due to historical introgression
261	(Leaché & McGuire 2006; Leaché & Linkem 2015). Our inferred divergence times based on an
262	assumed mtDNA substitution rate of 0.00805 substitutions per site per million years were
263	congruent with times previously inferred using secondary calibration information for the crown
264	age of phrynosomatids (Leaché & Linkem 2015). Average pairwise sequence divergence varied
265	considerably among species, from relatively low levels in <i>P. mcallii</i> (0.94%) to high divergence
266	within the <i>P. douglasii</i> complex (7.7%; Table 1). Similarly, effective population sizes varied by
267	an order of magnitude.
268	All MCMC species delimitation analyses indicated adequate convergence based on
269	visualizing plots of generation versus likelihood score. Because no differences were detected

270 when using a single run versus 10 independent MCMC runs, we present results from single runs

271	only. The number of horned lizard species inferred by each delimitation method was
272	substantially different between the full (220 haplotypes) and reduced (149 haplotypes) data sets.
273	For the full data set, sGMYC was the most conservative method, inferring a total of 10
274	species with wide confidence intervals (Table 1; Fig. 1). The comparatively small number of
275	species inferred with sGMYC was likely a result of the method having difficulty in locating the
276	threshold point in the data (Supplementary Fig. S1A). sGMYC is expected to work well when
277	there is a clear demarcation in branching rates between versus within species (Pons et al. 2006;
278	Reid & Carstens 2012; Esselstyn et al. 2012; Fujisawa & Barraclough 2013), which was not the
279	case in the horned lizard data as the threshold was placed relatively deep in the genealogy.
280	Conversely, mGMYC suggested an unrealistically large number of species (81), a high
281	proportion of splits (0.95), and low match ratio (0.13), and appeared to have an equally difficult
282	time placing transition points between inter- and intraspecific branching processes
283	(Supplementary Fig. S1B). mGMYC suggested up to 31 species within <i>P. platyrhinos</i> alone (Fig.
284	1), which is unlikely given the low levels of divergence within the species (Table 1). bPTP also
285	suggested an unrealistically high number of horned lizard species (52) with wide confidence
286	intervals from MCMC analyses. However, unlike mGMYC, bPTP inferred only four species
287	within P. platyrhinos (Fig. 1). Both mGMYC and bPTP inferred a large number of species
288	within the <i>P. douglasii</i> complex (29 and 26 species, respectively). mPTP analyses suggested an
289	intermediate number of total species (18) that was the most congruent with the relative levels of
290	structure in the inferred genealogies (Table 1; Fig. 1). These results were consistent with
291	previous findings that the original PTP model tends to oversplit, whereas mPTP is more
292	conservative and likely to represent true species clusters (Kapli et al. 2016).

293 ABGD analysis suggested a total of 12 species based on initial partitioning over a range 294 of prior values for maximum intraspecific divergence (Table 1; Supplementary Fig. S2). Results 295 based on JC69 and K80 corrected distances were identical. The number of species decreased to 296 five with a maximum intraspecific divergence prior value (P) of 0.021554, and to one species 297 with a value of 0.035938. Although there is still a lack of consensus of how to interpret 298 discordant ABGD results (Kekkonen & Hebert 2014), previous studies advocate using a P-value 299 of ~0.01 (Puillandre *et al.* 2012a), which in our data would result in the recognition of 12 species 300 of horned lizards (excluding the singletons). The relatively low value for the proportion of 301 species matches (0.17) and match ratio (0.20) was likely a result of both excluding singletons 302 from the analysis and the likely presence of multiple undescribed species in the data. Applying 303 the recursive algorithm resulted in a maximum of 20 species when X = 1.5 (Supplementary Fig. 304 S2A). This value increased to a maximum of 36 species when X = 1.0 (Supplementary Fig. S2B). 305 To be conservative, we focus on the results from the initial partitioning as this scheme has been 306 shown to be more stable across parameter settings and congruent with other species delimitation 307 methods (Puillandre et al. 2012a; Puillandre et al. 2012b; Kekkonen & Hebert 2014), including 308 those examined in this study.

309 R_{tax} values for the full 220 unique haplotype data set ranged from 0.05 for sGMYC to3100.96 for mGMYC (Table 2), further illustrating the tendency of mGMYC to delimit a large311number of (likely erroneous) species. Congruence among methods, based on C_{tax} values, was312highest between mPTP and ABGD (0.64) and lowest between sGMYC and mGMYC (0.05).313mPTP and ABGD also showed the largest mean C_{tax} among methods (Table 2).314Results of the full 368 sequence data sets (including identical sequences) revealed315varying degrees of sensitivity of methods to the presence of duplicates, with most algorithms

316 suggesting additional species (Supplementary Table S2). mGMYC was by far the most sensitive 317 of the methods compared, inferring a total of 164 species versus 81 species in the unique 220 318 haplotype data set. In general, the performance of all methods (except bPTP) was reduced based 319 on match ratios. Interestingly, mPTP failed to distinguish between P. cornutum and P. solare 320 even though these species are distantly related (Leaché & Linkem 2015). The performance of 321 ABGD also seemed to be impacted by the inclusion of identical sequences, concordant with 322 other recent findings (Ahrens et al. 2016). Thus, although previous studies suggest that the 323 GMYC model may be robust to the inclusion of identical sequences (Talavera et al. 2013), our 324 results suggest that additional bias may be introduced when duplicate haplotypes are not 325 removed.

326

327 Even sampling

328 Results from our analyses using a relatively even sampling scheme (after pruning the number of 329 P. platyrhinos haplotypes to 40) revealed varying levels of sensitivity among the species 330 delimitation methods to uneven sampling among species. sGMYC results were virtually identical 331 to the full analysis due to the same position of the inflection point (Supplementary Fig. S3A), but 332 the confidence interval was substantially reduced. In contrast, mGMYC was extremely sensitive 333 to sampling regime (Figs. 1,2; Table 3), and placed the four thresholds at different times in the 334 evenly sampled genealogy (Supplementary Fig. S3B). The proportion of species splits inferred 335 by mGMYC was substantially reduced in the pruned data set (0.71 vs. 0.95) and the match ratio 336 was increased from 0.13 to 0.39. For example, only a single species within *P. platyrhinos* was 337 inferred by mGMYC for the reduced data set (Fig. 2), compared to 31 species for the full data 338 (Fig. 1). These results were surprising since we pruned *P. platyrhinos* haplotypes evenly

339 throughout the original genealogy. mGMYC also inferred different numbers of species within P. 340 orbiculare and the P. douglasii complex. Thus, in contrast to sGMYC, mGMYC appears quite 341 sensitive to sampling conditions, which may further limit the utility of the method (Esselstyn et 342 al. 2012; Fujisawa & Barraclough 2013; Talavera et al. 2013). Conversely, bPTP and mPTP 343 appear less sensitive to sampling issues as the number of inferred species, proportion of matches, 344 lumps and splits and match ratios were similar between both sets of analyses (Tables 1,3). mPTP 345 was the most consistent, with 18 species inferred from the full data set and 17 in the reduced data 346 set (Figs. 1,2). The discrepancy in the single species arose from a slightly different gene tree for 347 the reduced data set within the *P. platyrhinos* clade. Zhang et al. (2013) also tested for the 348 influence of sampling evenness on species delimitation results and found that PTP outperformed 349 GMYC with even sampling, whereas GMYC was slightly more accurate with uneven sampling. 350 Due to the extreme sensitivity of mGMYC to sampling conditions, the largest R_{tax} value was 351 obtained from bPTP (1.00 vs. 0.59 for mGMYC) in the pruned data set (Table 4). Similar to the 352 full 220 haplotype analysis, pairwise C_{tax} was highest between mPTP and ABGD (0.69) and in 353 this case lowest between sGMYC and bPTP (0.11; Table 4).

Collectively, these results suggest mGMYC and bPTP were more sensitive to sampling regime, and that the large difference in the inferred number of species between bPTP and mPTP is likely due to the latter fitting multiple exponential branch length distributions to species to account for different rates of coalescence in heterogeneous data sets containing species with different N_e and demographic histories. Thus, our results provide further empirical evidence that mPTP may be a good choice for single-locus species delimitation based on accuracy, consistency, and speed (Kapli *et al.* 2016).

ABGD analyses on the evenly sampled (pruned) data set also indicated 12 species for most values of *P* using the initial partition (Table 3; Supplementary Fig. S4A,B). Once again, the recursive partition suggested a substantially higher number of species, particularly when P < 0.0129. There were also slight differences in the recursive partition between K80 and JC69 corrected distances (Supplementary Fig. S4A,B). However, K80 distances indicated the same number of groups (12) among the initial and recursive partition when P = 0.0129.

367

368 *Performance of methods*

369 Many studies have examined the GMYC model in detail using both simulated and empirical data 370 (Reid & Carstens 2012; Esselstyn et al. 2012; Fujisawa & Barraclough 2013; Talavera et al. 371 2013; Tang et al. 2014; Ahrens et al. 2016). Tang et al. (2014) quantified the influence of gene 372 tree reconstruction method and rate smoothing technique on the performance of both GMYC and 373 PTP and found that the former was generally more sensitive to the selected model. They found 374 that most of the sensitivity was likely due to errors during the smoothing step and subsequently 375 advocated the use of BEAST to generate ultrametric gene trees. Talavera et al. (2013) used a 376 large butterfly data set to test the performance of GMYC and suggested that the model is highly 377 stable under a variety of conditions including tree reconstruction method, the number of 378 singletons included, the number of species included in the gene tree, and sampling coverage. 379 They provided a summary table and chart with recommendations for running GMYC on 380 empirical data sets. Interestingly, their analysis suggested that sGMYC often overestimates the 381 number of species, in contrast to our analysis where the model may be underestimating true 382 diversity. However, our results were concordant in the sense that species delimitations did not 383 change significantly with different sample coverage, although it was slightly impacted by the

384 inclusion of identical sequences. Studies have also indicated that GMYC may be sensitive to 385 general phylogenetic history, sampling intensity, DNA sequence length, speciation rate, $N_{\rm e}$, and 386 differences in N_e among species (Reid & Carstens 2012; Esselstyn *et al.* 2012), and may 387 sometimes underestimate (sGMYC) or overestimate (mGMYC) the true number of species 388 (Esselstyn *et al.* 2012). This is a likely explanation for our sGMYC results that estimated only 10 389 species of horned lizards, as there was no abrupt change in branching rates between versus 390 within species. More recently, Ahrens et al. (2016) used both simulated and empirical data to 391 better understand potential biases in GMYC due to sampling and population genetic artifacts. 392 Their results suggested that the majority of bias is introduced by variation in $N_{\rm e}$ among species, 393 which can be exacerbated by uneven species abundance/sampling. In these cases, sGMYC tends 394 to lump species and return wide confidence intervals, which is consistent with our results for 395 horned lizards. To help overcome these issues, they suggest increasing the number of clades 396 examined to balance out the large skew in $N_{\rm e}$ among species. Although this solution may 397 alleviate some of the issues with sGMYC, mPTP may be more reliable in such cases as the 398 method can explicitly account for differences in $N_{\rm e}$ and rates of coalescence among species. 399 To our knowledge, few studies have examined the potential influence of methods for 400 summarizing node height information in BEAST analyses for subsequent species delimitation 401 using GMYC. To provide preliminary data on this issue, we performed a suite of additional 402 sGMYC and mGMYC analyses on both the full (368 sequences) and unique haplotype (220 403 sequences) data sets. Three methods of summarizing node heights were compared: mean heights, 404 median heights, common ancestor heights. Different results were detected depending on whether 405 the single- or multiple-threshold model was used (Table 5). Using common ancestor node 406 heights with sGMYC dramatically increased the number of delimited species, due to the

407 threshold point be pushed closer to the present (Supplementary Fig. S5). This effect was
408 negligible with mGMYC, which appeared to be more affected by the inclusion of identical
409 sequences (Supplementary Fig. S6) resulting in a doubling of the number of inferred species
410 (Table 5). Thus, there appears to be additional nuances of GMYC that should be considered
411 when utilizing these methods on empirical data. Based on relative concordance with our mPTP
412 and ABGD analyses, using sGMYC with mean or median node heights may be a good approach.
413 Additional simulation studies will be needed to test this prediction further.

414 Given the high levels of discordance observed among the methods tested, how should 415 researchers use these algorithms to discover and delimit diversity? As detailed above, mPTP has 416 numerous advantages over other methods. The consistency of mPTP to delimit putative species 417 in our study despite varied sampling depths and effective population sizes provides additional 418 evidence suggesting that the model may be appropriate for a wide variety of empirical data sets 419 (Kapli *et al.* 2016). Further, mPTP may be a good choice in data sets such as ours where sGMYC 420 has difficulty in placing the threshold point due to a more gradual slope in branching times, 421 possibly as a result of sampling a low species-to-individual ratio or due to large differences in $N_{\rm e}$ 422 among species (Ahrens et al. 2016). However, we agree with previous authors in that taxonomic 423 changes should not be made solely on the results of these methods (Lohse 2009; Puillandre et al. 424 2012a; Esselstyn et al. 2012; Zhang et al. 2013; Talavera et al. 2013), although concordance 425 using multiple analyses does tend to increase reliability (Puillandre et al. 2012b; Carstens et al. 426 2013; Satler et al. 2013). Rather, robust single-locus approaches should be used to form primary 427 taxonomic hypotheses that are subsequently tested with other types of data as part of an 428 integrative taxonomic framework (Fujita et al. 2012). Although ABGD has the potential to offer 429 a rapid and robust framework for assessing concordance (Puillandre et al. 2012a; Puillandre et

430	al. 2012b; Ahrens et al. 2016), additional work is needed to determine optimal parameter settings
431	and whether the recursive partition tends to oversplit. Moreover, additional studies are needed to
432	compare the performance of RESL through BOLD against some of the newer tree-based
433	methods (e.g., bPTP and mPTP).
434	The adoption of single-locus species delimitation methods to biodiversity research seems
435	particularly relevant to large metabarcoding studies (including microbial 16S rRNA sequencing)
436	as a rapid and cost-effective means to target groups for additional investigation. High-
437	throughput, multiplex amplicon sequencing using next-generation sequencing platforms allows
438	for the rapid generation of single-locus data from a large number of samples for primary species
439	delineation. Due to its speed and accuracy, mPTP seems to be an ideal analytical tool for these
440	large heterogeneous data sets consisting of species with different coalescent histories. We
441	anticipate that empiricists will continue to explore the utility of single-locus methods as
442	sequencing technologies improve and new analytical tools are developed.
443	
444	Taxonomy of horned lizards
445	Based on the performance and consistency of single-locus species delimitation methods tested in

this study, species-level diversity within *Phrynosoma* might be underestimated. Analyses of the
mitochondrial gene ND4 provide evidence to suggest at least 11 species could be present: *P. mcallii* (1 spp.), *P. platyrhinos* (2 spp.), *P. orbiculare* (5 spp.) and *P. douglasii* complex (3 spp.).
We note that this interpretation is conservative and based on congruence between mPTP and
ABGD analyses. Within *P. orbiculare*, both methods suggest the presence of one species in the
northern Sierra Madre Occidental of Chihuahua, one species in the southern Sierra Madre
Occidental, one in the northern Sierra Madre Oriental and Central Mexican Plateau, one in the

453 southern Sierra Madre Oriental and adjacent Trans-Mexican Volcanic Belt (Veracruz and 454 Puebla), and one in the central Trans-Mexican Volcanic Belt (Estado de México and Distrito 455 Federal). mPTP also suggested an additional species in the southern Sierra Madre Oriental in 456 Hidalgo (SMOr-H; Fig. 3). 457 Species delimitation scenarios for the P. douglasii complex varied among the methods. In 458 our RAxML gene trees, a strongly supported P. douglasii clade (bootstrap=93%) was nested 459 deep within *P. hernandesi* sensu lato (s.l.), albeit with weak support (<50% bootstrap; Fig. 3). 460 However, this relationship resulted in non-monophyly of *P. hernandesi* in RAxML gene trees. In 461 contrast, BEAST analyses reconstructed two major clades, one consisting of haplotypes from P. 462 hernandesi s.l. and one containing haplotypes from P. douglasii. This is one reason why the tree-463 based species delimitation methods failed to differentiate the two species. In all phylogenetic 464 analyses, however, a single divergent haplotype was found that may represent a third species 465 within this clade (PDU71587 from Arizona), a finding that warrants further evaluation in future 466 studies. ABGD further split *P. hernandesi* into two putative species that correspond to the "SER" 467 and "GB/CP" lineages of Zamudio et al. (1997). There was weak evidence to support the recently proposed taxonomy of Montanucci (2015), who resurrected the names P. brevirostris 468 469 and P. ornatissimum and described two new taxa (P. bauri and P. diminutum) based on 470 morphological comparisons. Phrynosoma bauri and P. ornatissimum were represented in our 471 data by one haplotype each, which were both strongly nested within *P. hernandesi* sensu stricto 472 (s.s.) in all phylogenetic analyses. In addition, nearly all species delimitation analyses (including 473 tree-based and distance-based) suggested that these taxa are not distinct from *P. hernandesi*. 474 Only bPTP and mGMYC analyses suggested that these species may be valid, but these analyses 475 generally tended to split *P. hernandesi* s.l. into a large number of singletons that are not

476 consistent with the recovered genealogies, and thus we view these results as unlikely.

477 *Phrynosoma brevirostris* was recovered as a clade deeply nested within *P. hernandesi* s.s. in all 478 the genealogies. Once again, only bPTP and mGMYC suggested that *P. brevirostris* is a distinct 479 species. Interestingly, mGMYC on the full 220 haplotype data set split *P. brevirostris* into 480 multiple species, whereas mGMYC on the evenly sampled data set lumped all haplotypes into a 481 single entity. Much more taxonomic work remains for the *P. douglasii* complex, and it is likely 482 that nuclear data will be required to resolve the discrepancies between morphology and mtDNA 483 data.

484 Within *P. platyrhinos*, our results suggested between two (ABGD) and four (mPTP) 485 species, with strong support from mPTP. Concordance among results suggests two samples from 486 Yuma Proving Ground in La Paz County, Arizona (LVT9951 and LVT9952) may represent a 487 new species, consistent with previous suggestions (Mulcahy et al. 2006; Jezkova et al. 2016). 488 mPTP results also suggested that sample LVT818 was a separate taxon as was a clade containing 489 samples DGM478 and DGM481, although the latter clade was weakly placed in alternative 490 positions in all of our ML and Bayesian analyses and unlikely to represent a distinct species. 491 Finally, the majority of analyses suggested that P. mcallii consists of a single species, concordant 492 with the low level of genetic diversity in the species (Mulcahy et al. 2006). However, mGMYC 493 on the full 368 sequence data set suggested a highly unrealistic estimate of 18 species within P. 494 *mcallii* (Supplementary Table S2), further illustrating the propensity of this method to oversplit. 495 We note that although many of the tree-based methods we tested are robust under a 496 variety of scenarios, poorly supported nodes and/or non-monophyly in gene trees may render 497 results unreliable until further data are collected—a broad limitation of the single-locus approach 498 (Esselstyn et al. 2012; Fujisawa & Barraclough 2013; Kapli et al. 2016). Thus, when delimiting

499 putative species based on single-locus data researchers should consider using both tree- and 500 threshold-based methods like ABGD to account for the shortcomings of each type of method 501 (Hamilton et al. 2011; Puillandre et al. 2012b). Prior to formal taxonomic changes, results should 502 be subsequently tested with additional types of data and analyses, such as morphological and/or 503 ecological data and multilocus coalescent-based methods (Puillandre et al. 2012b; Satler et al. 504 2013). This multiple-lines-of-evidence approach can therefore account for gene tree/species tree 505 discordance before formal taxonomic changes are implemented (Fujita et al. 2012; Carstens et al. 2013). Nevertheless, single-locus methods such as mPTP have the potential to quickly and 506 507 accurately target lineages that warrant additional investigation.

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656 Author Contributions

- 657 Designed study (CB), performed statistical analyses (CB), contributed to manuscript writing and
- 658 figure preparation (CB, RWB).
- 659

660 Data Accessibility

- 661 All multiple sequence alignments and gene trees can be found on the Dryad Digital Repository
- 662 (doi:10.5061/dryad.r7989).

Table 1. Number of horned lizard species (*Phrynosoma*) inferred by each single-locus species delimitation method tested on the full data set of unique haplotypes. Results from GMYC and PTP are from genealogies containing 220 ND4 haplotypes (n) with highly heterogeneous sampling intensity of target taxa. All bPTP and mPTP results are from Bayesian MCMC analyses. Confidence intervals for totals are in parentheses. ABGD results are based on the initial partitioning scheme with a maximum intraspecific diversity value of 0.012915 (K80 distances). Singletons were pruned prior to ABGD analysis. Taxonomy for *P. douglasii* complex follows Montanucci (2015) and includes *P. douglasii*, *P. hernandesi*, *P. bauri*, *P. brevirostris*, and *P. ornatissimum*. Also shown are average corrected pairwise distances (substitutions/site) for each species containing multiple haplotypes and Watterson's theta ($\theta = 4N_e\mu$). NA= Not applicable. 'Species matched' refers to the proportion of delimited species; 'Species splits' represents the proportion of taxonomic species splits by each delimitation method.

Taxon	n	Mean Tamura-Nei distance	Watterson's theta	GMYC single	GMYC multiple	bPTP	mPTP	ABGD
P. douglasii complex	41	0.0774	115.6936	2	29	26	2	4
P. orbiculare	34	0.0633	56.985	1	14	11	6	5
P. mcallii	29	0.0094	12.4772	1	2	1	1	1
P. platyrhinos*	111	0.0211	35.2124	1	31	4	4	2
P. asio	1	NA	NA	1	1	1	1	-
P. cornutum	1	NA	NA	1	1	1	1	-
P. coronatum	1	NA	NA	1	1	1	1	-
P. solare	1	NA	NA	1	1	1	1	-
P. taurus	1	NA	NA	1	1	1	1	-
Total	220	0.1258¶	121.1346†	10 (7-98) ^a	81 (69-103) ^b	52.40 (37-84) ^c	18 (16-21) ^d	12
Species matched (percentage of total	.)			0.80	0.07	0.15	0.33	0.17
Species lumped (percentage of total))			0.50	0.02	0.04	0.28	0.42
Species splits (percentage of total)				0.20	0.95	0.74	0.67	0.83
Match ratio				0.70	0.13	0.24	0.39	0.20

a = likelihood ratio versus null model 10.90506, p = 0.0043; b = likelihood ratio versus null model 23.85364, p < 0.001; c = mean number of species; d =

664 *Central Credible Interval;* * = -3x number of haplotypes versus P. douglasii complex, P. orbiculare, P. mcallii; $^{\dagger} = theta$ for all 220 sequences; $^{\$} = mean$ Tamura-665 *Nei distance for all 220 sequences.*

		_	C_{tax}	_		Mean C _{tax}	$R_{\rm tax}$	# species
	sGMYC	mGMYC	bPTP	mPTP	ABGD			
sGMYC	-	-	-	-	-	0.22	0.05	5
mGMYC	0.05	-	-	-	-	0.22	0.96	76
bPTP	0.12	0.50	-	-	-	0.3	0.53	42
mPTP	0.33	0.16	0.29	-	-	0.36	0.15	13
ABGD	0.36	0.15	0.27	0.64	-	0.36	0.14	12

Table 2. Calculation of the Taxonomic Index of Congruence (C_{tax}) and Relative Taxonomic Resolving Power Index (R_{tax}) for all species delimitation methods compared based on the full 220 unique haplotype data set.

666

Total number of speciation events across all methods (excluding singletons) = 78; Mean C_{tax} = average value for method across all pairwise comparisons. Singletons were excluded from calculations to allow fair comparisons (all singletons were pruned prior to ABGD analyses). See text for additional details.

Table 3. Number of horned lizard species (*Phrynosoma*) inferred by each single-locus species delimitation method tested on the pruned data set. Results from GMYC and PTP are from genealogies containing 149 ND4 haplotypes (n) with approximately even sampling intensity of target taxa. All bPTP and mPTP results are from Bayesian MCMC analyses. Confidence intervals for totals are in parentheses. ABGD results are based on the initial partitioning scheme with a maximum intraspecific diversity value of 0.012915 (K80 distances). Singletons were pruned prior to ABGD analysis. Taxonomy for *P. douglasii* complex follows Montanucci (2015) and includes *P. douglasii*, *P. hernandesi*, *P. bauri*, *P. brevirostris*, and *P. ornatissimum*. Also shown are average corrected pairwise distances (substitutions/site) for each species containing multiple haplotypes and Watterson's theta ($\theta = 4N_e\mu$). NA= Not applicable. 'Species matched' refers to the proportion of delimited species; 'Species splits' represents the proportion of taxonomic species splits by each delimitation method.

Tavon	n	Mean Tamura-Nei distance	Watterson's	GMYC	GMYC multiple	bPTP	mPTP	ABGD
14701		Tamura-Iver distance	tileta	single	multiple			
P. douglasii complex	41	0.0774	115.6936	2	11	22	2	4
P. orbiculare	34	0.0633	56.985	1	10	12	6	5
P. mcallii	29	0.0094	12.4772	1	1	1	1	1
P. platyrhinos	40	0.0256	34.7945	1	1	3	3	2
P. asio	1	NA	NA	1	1	1	1	-
P. cornutum	1	NA	NA	1	1	1	1	-
P. coronatum	1	NA	NA	1	1	1	1	-
P. solare	1	NA	NA	1	1	1	1	-
P. taurus	1	NA	NA	1	1	1	1	-
Total	149	0.1468¶	129.0831†	10 (3-19) ^a	28 (9-66) ^b	46.74 (34-68) ^c	17 (11-19) ^d	12
Species matched (percentage of total)				0.80	0.29	0.19	0.35	0.17
Species lumped (percentage of total)				0.50	0.14	0.00	0.29	0.42
Species splits (percentage of total)				0.20	0.71	0.73	0.65	0.83
Match ratio				0.70	0.39	0.30	0.40	0.20

a = likelihood ratio versus null model 10.02172, p = 0.0067; b = likelihood ratio versus null model 12.4399, p = 0.002; c = mean number of species; d = Central Credible Interval; [†] = theta for all 149 sequences; [¶] = mean Tamura-Nei distance for all 149 sequences.

-		_	C_{tax}	_		Mean C _{tax}	$R_{\rm tax}$	# species
	sGMYC	mGMYC	bPTP	mPTP	ABGD			
sGMYC	-	-	-	-	-	0.25	0.11	5
mGMYC	0.18	-	-	-	-	0.40	0.59	23
bPTP	0.11	0.59	-	-	-	0.33	1.00	38
mPTP	0.36	0.38	0.30	-	-	0.43	0.30	12
ABGD	0.36	0.43	0.30	0.69	-	0.45	0.30	12

Table 4. Calculation of the Taxonomic Index of Congruence (C_{tax}) and Relative Taxonomic Resolving Power Index (R_{tax}) for all species delimitation methods compared based on the pruned 149 unique haplotype data set.

Total number of speciation events across all methods (excluding singletons) = 37; Mean C_{tax} = average value for method across all pairwise comparisons. Singletons were excluded from calculations to allow fair comparisons (all singletons were pruned prior to ABGD analyses). See text for additional details.

Table 5. Comparison of the number of delimited horned lizard (*Phrynosoma*) species by the single-(sGMYC) and multiple-threshold (mGMYC) GMYC models based on different methods of annotating node height and sampling regimes. 'Unique' = only unique haplotypes (220); 'All' = all sequences (368). Three ways to summarize node heights on BEAST maximum clade credibility trees were evaluated ('Mean heights', 'Common ancestor heights', 'Median heights'). Values represent the number of ML entities with confidence intervals.

Data set	Node heights	sGMYC	mGMYC
Unique	Mean heights	10 (7-98)**	81 (69-103)***
Unique	Common ancestor heights	64 (58-72)***	66 (61-66)***
Unique	Median heights	9 (7-69)**	76 (68-103)***
All	Mean heights	13 (8-138)*	164 (157-164)***
All	Common ancestor heights	79 (76-87)***	100 (83-100)***
All	Median heights	9 (7-16)*	191 (13-299)***

 $^{\ast}=P<0.05;\ ^{\ast\ast}=P<0.005;\ ^{\ast\ast\ast}=P<0.001$

669

671 Figure Legends

672

673	Fig. 1. Comparison of species delimitation results of horned lizards (<i>Phrynosoma</i>) based on
674	analysis of 220 unique ND4 haplotypes and highly heterogeneous taxonomic sampling. Each
675	coloured bar represents a species delimited by each method tested. Gene tree is from a BEAST
676	analysis under a strict clock and constant size coalescent tree prior. Node height was determined
677	using mean heights across the posterior distribution. Node values represent Bayesian posterior
678	probabilities (>0.95) for major clades. sGMYC = single threshold GMYC; mGMYC = multiple
679	threshold GMYC; bPTP = single-rate Poisson Tree Processes model (MCMC analysis of a
680	RAxML gene tree) fitting a single branch length distribution to coalescent events; mPTP =
681	multi-rate Poisson Tree Processes model (MCMC analysis of a RAxML gene tree) fitting
682	multiple branch length distributions to coalescent events across distinct species. ABGD =
683	Automatic Barcode Gap Discovery. Note that singletons (i.e. P. asio, P. taurus, P. cornutum, P.
684	solare, P. coronatum) were removed prior to ABGD analysis. * = OTUs (PDU71587,
685	PDU71589) clustered together as a single species based on mPTP analysis of a RAxML gene
686	tree, which yielded a slightly different topology than the BEAST genealogy shown. See Fig. 3
687	for additional details. Refer to online version for a full colour representation of figure.
688	
689	Fig. 2. Comparison of species delimitation results of horned lizards (<i>Phrynosoma</i>) based on

analysis of 149 unique ND4 haplotypes with relatively even intraspecific sampling. Each
coloured bar represents a species delimited by each method tested. Gene tree is from a BEAST
analysis under a strict clock and constant size coalescent tree prior. Node height was determined
using mean heights across the posterior distribution. Node values represent Bayesian posterior

694	probabilities (>0.95) for major clades. sGMYC = single threshold GMYC; mGMYC = multiple
695	threshold GMYC; bPTP = single-rate Poisson Tree Processes model (MCMC analysis of a
696	RAxML gene tree) fitting a single branch length distribution to coalescent events; mPTP =
697	multi-rate Poisson Tree Processes model (MCMC analysis of a RAxML gene tree) fitting
698	multiple branch length distributions to coalescent events across distinct species. ABGD =
699	Automatic Barcode Gap Discovery. Note that singletons (i.e. P. asio, P. taurus, P. cornutum, P.
700	solare, P. coronatum) were removed prior to ABGD analysis. * = OTUs (PDU71587,
701	PDU71589) clustered together as a single species based on mPTP analysis of a RAxML gene
702	tree, which yielded a slightly different topology than the BEAST genealogy shown. See Fig. 3
703	for additional details. Refer to online version for a full colour representation of figure.
704	
705	Fig. 3. Species delimitation of horned lizards (<i>Phrynosoma</i>) based on MCMC mPTP analysis of
706	a RAxML gene tree constructed with 220 ND4 sequences under a GTRGAMMA model.
707	Branches are colour-coded to represent speciation (black) or coalescence (red) events. Values at
708	nodes indicate probability of a speciation event based on mPTP MCMC analysis (first value) and
709	maximum likelihood bootstrap proportions using the autoMRE bootstopping criterion in RAxML
710	(second value). YPG = samples LVT9951 and LVT9952 from Yuma Proving Ground, Arizona.
711	SMOcN = northern Sierra Madre Occidental; SMOcS = southern Sierra Madre Occidental;
712	SMOrS = southern Sierra Madre Oriental; TMVB = Trans-Mexican Volcanic Belt; TMVBc =
713	central Trans-Mexican Volcanic Belt; SMOr-H = Sierra Madre Oriental-Hidalgo; SMOrN =
714	
	northern Sierra Madre Oriental; CMP = Central Mexican Plateau. <i>PSS = Phrynosoma</i> singleton

- 716 bauri, P. brevirostris, P. ornatissimum. Horned lizard shown is a P. orbiculare from Coahuila,
- 717 Mexico. Refer to online version for a full colour representation of figure.