- 1 Development of imaging-based risk scores for prediction of intracranial haemorrhage
- and ischaemic stroke in patients taking antithrombotic therapy after ischaemic stroke

or TIA: a pooled analysis of individual patient data from cohort studies

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Jonathan G Best MD, Gareth Ambler PhD, Duncan Wilson PhD, Keon-Joo Lee MD, Jae-Sung 5 Lim MD, Masayuki Shiozawa MD, Masatoshi Koga MD, Linxin Li DPhil, Caroline Lovelock 6 7 FRACP, Prof Hugues Chabriat MD, Prof Michael Hennerici MD, Yuen Kwun Wong MSc, Henry Ka Fung Mak MD, Luis Prats-Sanchez PhD, Alejandro Martínez-Domeño MD, Shigeru 8 9 Inamura MD, Kazuhisa Yoshifuji PhD, Ethem Murat Arsava MD, Solveig Horstmann MD, Jan Purrucker MD, Bonnie Yin Ka Lam PhD, Adrian Wong PhD, Young Dae Kim MD, Tae-Jin 10 Song MD, Robin Lemmens PhD, Sebastian Eppinger MD, Thomas Gattringer MD, Ender 11 Uysal MD, Zeynep Tanriverdi MD, Prof Natan M Bornstein MD, Einor Ben Assayag PhD, 12 Hen Hallevi MD, Jeremy Molad MD, Masashi Nishihara MD, Jun Tanaka MD, Prof Shelagh 13 B. Coutts MD, Alexandros Polymeris MD, Benjamin Wagner MD, David J. Seiffge MD, Prof 14 Philippe Lyrer MD, Prof Ale Algra MD, Prof L. Jaap Kappelle MD, Prof Rustam Al-Shahi 15 Salman PhD, Prof Hans R Jäger FRCR, Prof Gregory Y.H. Lip MD, Prof Urs Fischer MD, 16 Prof Marwan El-Koussy MD, Prof Jean-Louis Mas MD, Laurence Legrand MD, Christopher 17 Karayiannis MD, Prof Thanh Phan MD, Sarah Gunkel MD, Nicolas Christ MD, Jill Abrigo 18 MD, Prof Thomas Leung MD, Prof Winnie Chu MD, Francesca Chappell PhD, Stephen Makin 19 20 PhD, Derek Hayden MD, Prof David J Williams PhD, Prof Werner H. Mess MD, Paul J. Nederkoorn MD, Carmen Barbato MD, Simone Browning BSc, Kim Wiegertjes MD, Anil M. 21 Tuladhar MD, Noortje Maaijwee PhD, Anne Cristine Guevarra MD, Chathuri Yatawara PhD, 22 Anne-Marie Mendyk RN, Christine Delmaire MD, Sebastian Köhler PhD, Prof Robert van 23 Oostenbrugge MD, Ying Zhou PhD, Chao Xu MD, Saima Hilal PhD, Bibek Gyanwali MD, 24

Christopher Chen FRCP, Prof Min Lou PhD, Julie Staals MD, Prof Régis Bordet MD,

Nagaendran Kandiah FRCP, Prof Frank-Erik de Leeuw MD, Robert Simister PhD, Prof Jeroen 26 Hendrikse MD, Prof Peter J Kelly MD, Prof Joanna Wardlaw MD, Yannie Soo MD, Felix Fluri 27 28 MD, Prof Velandai Srikanth PhD, Prof David Calvet MD, Prof Simon Jung MD, Vincent I.H. Kwa MD, Prof Stefan T Engelter MD, Prof Nils Peters MD, Prof Eric E Smith MD, Prof Hideo 29 Hara PhD, Prof Yusuke Yakushiji PhD, Prof Dilek Necioglu Orken MD, Prof Franz Fazekas 30 MD, Prof Vincent Thijs MD, Prof Ji Hoe Heo MD, Prof Vincent Mok MD, Prof Roland 31 32 Veltkamp MD, Hakan Ay MD, Toshio Imaizumi MD, Beatriz Gomez-Anson FRCR, Kui Kai Lau DPhil, Prof Eric Jouvent MD, Prof Peter M. Rothwell FMedSci, Prof Kazunori Toyoda 33 34 MD, Prof Hee-Joon Bae PhD, Prof Joan Marti-Fabregas PhD, and Prof David J Werring PhD, on behalf of the Microbleeds International Collaborative Network 35 36 Stroke Research Centre, Department of Brain Repair and Rehabilitation, UCL Queen Square 37 Institute of Neurology, London, UK (J G Best MD, D Wilson PhD, D J Seiffge MD, C Barbato 38 MD, Simone Browning BSc, R Simister PhD, D J Werring PhD); Department of Statistical 39 Science, University College London, Gower Street, London, UK (G Ambler PhD); New 40 Zealand Brain Research Institute, Christchurch, New Zealand (D Wilson PhD); Department of 41 Neurology, Seoul National University Bundang Hospital, Seoul National University College 42 of Medicine, Seongnam, Republic of Korea (K-J Lee MD, H-J Bae PhD); Department of 43 Neurology, Hallym University Sacred Heart Hospital, Hallym Neurological Institute, Hallym 44 University College of Medicine, Anyang, Republic of Korea (J-S Lim MD); Department of 45 Cerebrovascular Medicine, National Cerebral and Cardiovascular Centre, 6-1 Kishibe-46 shimmachi, Suita, Osaka 564-8565, Japan (M Shiozawa MD, M Koga MD, K Toyoda MD); 47 Wolfson Centre for Prevention of Stroke and Dementia, Nuffield Department of Clinical 48 Neurosciences, University of Oxford, UK (L Li DPhil, C Lovelock FRACP, P M Rothwell 49 FMedSci); APHP, Lariboisière Hospital, Department of Neurology, F-75475 Paris, France (H 50

Chabriat MD, E Jouvent MD); FHU NeuroVasc, Université de Paris and INSERM U1141, 51 Paris, France (H Chabriat MD, E Jouvent MD); Department of Neurology, Universitätsmedizin 52 53 Mannheim, University of Heidelberg, Mannheim, Germany (M Hennerici MD); Division of Neurology, Department of Medicine (Y K Wong MSc, K K Lau DPhil) and Department of 54 Diagnostic Radiology (H K F Mak MD), The University of Hong Kong, Hong Kong; 55 Department of Neurology, Hospital de la Santa Creu i Sant Pau, Biomedical Research Institute, 56 57 Barcelona, Spain (L Prats-Sanchez PhD, A Martínez-Domeño MD, J Marti-Fabregas PhD); Department of Neurosurgery, Kushiro City General Hospital, Kushiro, Japan (S Inamura MD, 58 59 K Yoshifuji PhD, T Imaizumi MD); Departments of Neurology and Radiology, Massachusetts General Hospital, Harvard Medical School, Boston MA, USA (EM Arsava MD); Department 60 of Neurology, Heidelberg University Hospital, Heidelberg, Germany (S Horstmann MD, J 61 Purrucker MD); Therese Pei Fong Chow Research Centre for Prevention of Dementia, Gerald 62 Choa Neuroscience Centre, Lui Che Woo Institute of Innovative Medicine, Department of 63 Medicine and Therapeutics, The Chinese University of Hong Kong, Hong Kong (B Y K Lam 64 PhD, A Wong PhD, V Mok MD); Department of Neurology, Yonsei University College of 65 Medicine, Seoul, South Korea (Y D Kim MD, J H Heo MD); Department of Neurology, Seoul 66 Hospital, Ewha Womans University College of Medicine, Seoul, South Korea (T-J Song MD); 67 Experimental Neurology, Department of Neurosciences, KU Leuven – University of Leuven; 68 VIB Center for Brain & Disease Research; Department of Neurology, University Hospitals 69 70 Leuven, Leuven, Belgium (R Lemmens PhD); Department of Neurology, Medical University of Graz, Auenbruggerplatz 22, 8036 Graz, Austria (S Eppinger MD, T Gattringer MD, F 71 Fazekas MD); Saglık Bilimleri University Sisli Etfal Education and Research Hospital 72 Department of Radiology, Istanbul, Turkey (E Uysal MD); İzmir Katip Çelebi University 73 Atatürk Education and Research Hospital Department of Neurology, İzmir Turkey (Z 74 Tanriverdi MD); Department of Neurology, Tel-Aviv Sourasky Medical Center, Tel-Aviv, 75

Israel; Sackler Faculty of Medicine, Tel-Aviv University, Tel-Aviv, Israel (N M Bornstein MD, 76 E Ben Assayag PhD, H Hallevi MD, J Molad MD); Department of Radiology (M Nishihara 77 78 MD) and Division of Neurology, Department of Internal Medicine (J Tanaka MD, H Hara PhD, Y Yakushiji PhD), Saga University Faculty of Medicine, 5-1-1, Nabeshima, Saga, Japan; 79 Calgary Stroke Program, Department of Clinical Neurosciences, Radiology and Community 80 Health Sciences, Hotchkiss Brain Institute, University of Calgary (S B Coutts MD, E E Smith 81 82 MD); Department of Neurology and Stroke Centre, University Hospital Basel and University of Basel, Switzerland (A Polymeris MD, B Wagner MD, D J Seiffge MD, P Lyrer MD, S T 83 Engelter MD, N Peters MD); Julius Centre for Health Sciences and Primary Care (A Algra 84 MD) and Department of Neurology and Neurosurgery (A Algra MD, L J Kappelle MD), 85 University Medical Centre Utrecht and Utrecht University, Utrecht, The Netherlands; Centre 86 for Clinical Brain Sciences, School of Clinical Sciences, University of Edinburgh, Edinburgh, 87 UK (R Al-Shahi Salman PhD); Lysholm Department of Neuroradiology and the 88 Neuroradiological Academic Unit, Department of Brain Repair and Rehabilitation, UCL 89 Institute of Neurology and the National Hospital for Neurology and Neurosurgery, London, 90 UK (H R Jäger FRCR); Liverpool Centre for Cardiovascular Science, University of Liverpool 91 and Liverpool Heart & Chest Hospital, Liverpool, United Kingdom (GYH Lip MD); Aalborg 92 Thrombosis Research Unit, Department of Clinical Medicine, Aalborg University, Aalborg, 93 Denmark (G Y H Lip MD); Department of Neurology (D J Seiffge MD, U Fischer MD, S Jung 94 95 MD) and Department of Diagnostic and Interventional Neuroradiology (M El-Koussy MD), University Hospital Inselspital Bern, University of Bern, Bern, Switzerland; Department of 96 Neurology (J-L Mas MD, D Calvet MD) and Department of Neuroradiology (L Legrand MD), 97 Sainte-Anne Hospital, Institut de Psychiatrie et Neurosciences de Paris (IPNP), UMR S1266, 98 INSERM, Université de Paris, F-75014, Paris, France; Peninsula Clinical School, Peninsula 99 Health (C Karayiannis MD, V Srikanth PhD) and Stroke and Ageing Research Group, School 100

of Clinical Sciences at Monash Health (T Phan MD), Monash University, Melbourne, 101 Australia; Department of Neurology, University Hospital of Würzburg, Josef-Schneider 102 Strasse 11, 97080, Würzburg, Germany (S Gunkel MD, N Christ MD, F Fluri MD); Department 103 of Imaging and Interventional Radiology (J Abrigo MD, W Chu MD) and Department of 104 Medicine and Therapeutics (T Leung MD, Y Soo MD), Prince of Wales Hospital, The Chinese 105 University of Hong Kong, Hong Kong; Centre for Clinical Brain Sciences, Edinburgh Imaging; 106 107 and UK Dementia Institute at the University of Edinburgh, Edinburgh, UK (F Chappell PhD, J Wardlaw MD); Institute of Applied Health Sciences, University of Aberdeen, Aberdeen, UK. 108 109 (S Makin PhD); The Neurovascular Research Unit and Health Research Board, Stroke Clinical Trials Network Ireland, University College Dublin, Dublin (D Hayden MD, P J Kelly MD); 110 Department of Medical Gerontology, Trinity College Dublin, College Green, Dublin 2, Ireland 111 (D Hayden MD); Department of Geriatric and Stroke Medicine, RCSI University of Medicine 112 and Health Sciences Dublin, Ireland and Beaumont Hospital Dublin, Ireland (D J Williams 113 PhD); Department of Clinical Neurophysiology, Maastricht University Medical Centre, 114 Maastricht, the Netherlands (W H Mess MD); Department of Neurology, Amsterdam 115 University Medical Centres, Location AMC, The Netherlands (P J Nederkoorn MD); 116 Comprehensive Stroke Service, University College London Hospitals NHS Trust, London, UK 117 (C Barbato MD, S Browning BSc, R Simister PhD); Department of Neurology, Donders 118 Institute for Brain, Cognition and Behaviour, Donders Centre for Medical Neuroscience, 119 120 Radboud University Medical Center, Nijmegen, The Netherlands (K Wiegertjes MD, A M Tuladhar MD, F-E de Leeuw MD); Lucerne State Hospital, Switzerland; Neurocenter, 121 Department for Neurology and Neurorehabilitation, Lucerne, Switzerland (N Maaijwee PhD); 122 Department of Neurology, National Neuroscience Institute, Singapore, Singapore (A C 123 Guevarra MD, C Yatawara PhD, N Kandiah FRCP); University of Lille, Inserm, CHU de Lille. 124 'Degenerative and vascular cognitive disorders' U1171. F-59000 Lille (A-M Mendyk RN, C 125

Delmaire MD, R Bordet MD); Department of Radiology. Fondation A de Rothschild. F-75019 126 Paris (C Delmaire MD); Department of Psychiatry and Neuropsychology, School for Mental 127 128 Health and Neuroscience (S Köhler PhD) and Department of Neurology, Cardiovascular Research Institute Maastricht (CARIM) (R van Oostenbrugge MD, J Staals MD), Maastricht 129 University Medical Centre, The Netherlands; Department of Neurology, The Second Affiliated 130 Hospital of Zhejiang University, School of Medicine (Y Zhou PhD, C Xu MD, M Lou PhD); 131 132 Memory Aging & Cognition Centre, Yong Loo Lin School of Medicine, National University of Singapore, Singapore (S Hilal PhD, B Gyanwali MD, C Chen FRCP); University Medical 133 134 Centre Utrecht, Utrecht University, The Netherlands (J Hendrikse MD); Department of Neurology, Onze Lieve Vrouwe Gasthuis, Amsterdam, The Netherlands (V I H Kwa MD); 135 Neurology and Neurorehabilitation, University Department of Geriatric Medicine FELIX 136 PLATTER; University of Basel, Switzerland (S T Engelter MD, N Peters MD); Department of 137 Neurology, Kansai Medical University, 2-5-1 Shinmachi, Hirakata, Osaka 573-1010, Japan (Y 138 Yakushiji PhD); Memorial Sisli Hospital, Department of Neurology, Istanbul, Turkey (D N 139 Orken MD); Stroke Division, Florey Institute of Neuroscience and Mental Health, University 140 of Melbourne, Heidelberg (V Thijs MD); Department of Neurology, Austin Health, 141 Heidelberg, Australia (V Thijs MD); Department of Brain Sciences, Imperial College London, 142 London, UK (R Veltkamp MD); Department of Neurology, Heidelberg University Hospital, 143 Heidelberg, Germany (R Veltkamp MD); A.A. Martinos Center for Biomedial Imaging, 144 Departments of Neurology and Radiology, Massachusetts General Hospital, Harvard Medical 145 School, Boston, MA, USA (H Ay MD); Takeda Pharmaceutical Company Limited, 146 Cambridge, MA, USA (H Ay MD); Unit of Neuroradiology, Hospital Santa Creu i Sant Pau, 147 Universitat Autonoma, Barcelona (B Gomez-Anson FRCR) 148

150	Correspondence to: Professor David J Werring, UCL Stroke Research Centre, Department of
151	Brain Repair and Rehabilitation, UCL Queen Square Institute of Neurology, Russell Square
152	House, 10 – 12 Russell Square, London WC1B 5EH, UK; d.werring@ucl.ac.uk
153	
154	Word count:
155	Manuscript: 3,448
156	Abstract: 298

Research in context panel: 445

Abstract

Background Balancing the risks of recurrent ischaemic stroke (IS) and intracranial haemorrhage (ICH) is important for patients treated with antithrombotic therapy after ischaemic stroke or transient ischaemic attack. However, existing predictive models offer limited performance, particularly for ICH. We aimed to develop new risk scores incorporating clinical variables and cerebral microbleeds (CMBs), an MRI biomarker of ICH and IS risk.

Methods We did a pooled analysis of individual-patient data from the Microbleeds International Collaborative Network, which comprises 38 hospital-based prospective cohort studies from 18 countries. All studies recruited participants with previous IS or TIA, acquired baseline MRI allowing quantification of CMBs, and followed up participants for IS and ICH. We excluded participants not taking antithrombotic drugs. We developed Cox regression models to predict the five-year risks of ICH and IS, selecting candidate predictors on biological relevance and simplifying models using backward elimination. We derived integer risk scores for clinical use. We assessed model performance in internal validation, adjusted for optimism using bootstrapping. We registered the study with the PROSPERO register of systematic reviews (registration: CRD42016036602).

Findings The included studies recruited participants between 28th August 2001 and 4th February 2018. 15,766 participants had follow-up for ICH, and 15,784 for IS. Over a median follow-up of two years, 184 ICH and 1,048 IS occurred. The risk models we developed included CMB burden and simple clinical variables. Optimism-adjusted c-indices were 0.73 (95% CI 0.69-0.77) for ICH and 0.63 for IS (95% CI 0.62-0.65); calibration slopes were 0.94 (95% CI 0.81-1.06) and 0.97 (95% CI 0.87-1.07) respectively, indicating good calibration.

Interpretation The MICON risk scores, incorporating clinical variables and CMBs, offer predictive value for the long-term risks of ICH and ischaemic stroke in patients prescribed antithrombotic therapy for secondary stroke prevention. External validation is warranted.

Funding British Heart Foundation and Stroke Association

Research in context

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Evidence before this study

We searched Medline from 1st January 1996 to 1st February 2020 using the following search strategy: (stroke[tiab] OR bleeding[tiab] OR haemorrhage[tiab] OR hemorrhage[tiab]) AND (prediction[tiab] OR risk stratification[tiab] OR risk score[tiab]). We identified studies in English which described or validated risk scores for ischaemic stroke or major bleeding, in patients taking antiplatelets or anticoagulants, with or without atrial fibrillation. Very few studies of bleeding risk scores reported their performance for intracranial haemorrhage specifically. A large cohort study (n=40,450) of patients with atrial fibrillation anticoagulated for stroke prevention found poor performance in predicting ICH for all bleeding risk scores assessed, including HEMORR2HAGES, HAS-BLED, ATRIA and ORBIT. The highest cindex obtained was 0.53, for HASBLED. A nationwide registry-based cohort study (n=182,678) assessing HASBLED and HEMORRH2HAGES in patients with atrial fibrillation also found limited performance, with c-indices between 0.58 and 0.62 in participants prescribed antithrombotics. Models developed for predicting ICH in patients taking antiplatelets specifically (including Intracranial-B2LEED3S and S2TOP-BLEED) also showed only moderate performance, with the highest reported c-index being 0.65, for S2TOP-BLEED. Risk scores for ischaemic stroke (including CHADS₂, CHAD₂S₂VASc and ATRIA) performed moderately, with c-indices typically between 0.60 and 0.70.

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Added value of this study

We present new clinical-radiological risk scores using cerebral microbleeds, an MRI marker of small vessel fragility, to predict ICH and ischaemic stroke in patients taking antithrombotic drugs for secondary prevention after ischaemic stroke or transient ischaemic attack, derived

from studies in the Microbleeds International Network (MICON), a large international collaboration of prospective cohort studies. The performance of our MICON-ICH score suggests it can usefully stratify patients by risk of antithrombotic-associated ICH in clinical practice. Our results also suggest that cerebral microbleeds add considerable value for predicting ICH, but not ischaemic stroke, clarifying the relative predictive importance of cerebral microbleeds for these outcomes. Our scores did not identify many patients with similar or greater predicted risk of ICH than ischaemic stroke, even in those with high cerebral microbleed burden and other risk factors. Our MICON scores are simple and widely applicable.

Implications of all the available evidence

Risk scores including cerebral microbleeds offer increased discrimination over clinical variables alone for the prediction of antithrombotic-associated ICH in a large, multicentre, international population. Although external validation is needed, this finding provides new evidence of how neuroimaging biomarkers can contribute to clinical prediction models. Identifying people at highest risk of ICH may facilitate timely and accurate prognostication to allow mitigation of reversible risk factors for bleeding (e.g. intensive blood pressure control), and selection of participants for clinical trials. While more complex combinations of clinical, biochemical, and radiological markers might further improve stroke risk prediction, balancing accuracy with simplicity will remain important.

Introduction

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Antithrombotic therapy is a key component of secondary prevention after ischaemic stroke or transient ischaemic attack. In patients without atrial fibrillation (AF), antiplatelet treatment reduces overall stroke risk by one-quarter, while oral anticoagulation in patients with AF reduces this risk by two-thirds.^{2,3} Although antithrombotic treatment increases the risk of intracranial haemorrhage (ICH) (by around one-quarter for antiplatelets, one-half for direct oral anticoagulants (DOACs), and two-fold for vitamin K antagonists (VKAs)), 1-3 the substantiallylower incidence of ICH overall means that antithrombotic treatment is recommended for most patients. However, deciding on appropriate antithrombotic therapy for a given patient can be challenging, especially in those with additional risk factors for bleeding. Ideally, this decision would be based on an individualised assessment of the risks of ischaemic stroke and ICH. To this end, risk scores for ischaemic stroke and major bleeding have been developed, mainly in patients with AF. Although these scores show reasonable discrimination for ischaemic stroke^{4,5} and all-cause major bleeding, 5,6 studies validating existing bleeding risk scores in predicting ICH have shown more limited performance, with c-indices between 0.50 and 0.62 in anticoagulated patients, 7,8 and 0.58 - 0.65 in patients taking antiplatelet drugs. 8,9 Most risk scores for ischaemic stroke and ICH only include clinical variables. More recently, scores using serum biomarkers have been developed, which may offer improved performance. 10-12 However, the role of magnetic resonance imaging (MRI) biomarkers for cerebrovascular disease (increasingly obtained as part of standard stroke care) in improving risk prediction remains uncertain. Cerebral microbleeds (CMBs) are an MRI biomarker of vascular fragility, associated with hypertensive microangiopathy (also known as arteriolosclerosis or deep perforator arteriopathy) and cerebral amyloid angiopathy, the two cerebral small vessel diseases that cause most spontaneous intracerebral haemorrhage.¹³

Accordingly, the potential of CMBs in predicting ICH has attracted particular interest. In a prospective observational study, the addition of CMB presence improved the c-index for ICH of the HASBLED bleeding risk score from 0·41 to 0·66, 14 while a recent large individual patient data meta-analysis confirmed a strong association between CMBs and ICH in patients with previous ischaemic stroke or TIA. 15 This study also found that CMBs are associated with IS risk, with a higher absolute risk of ischaemic stroke than ICH across all levels of CMB burden investigated.

Given these findings, we aimed to establish the added predictive value of CMBs for ICH and ischaemic stroke, by using the same large international dataset to develop risk models based on CMB burden and simple clinical variables, and to compare these to models using clinical variables alone. We aimed to derive from our models simple risk scores which could be easily used for risk stratification in clinical practice. We investigated whether the resulting scores identified a group of patients at similar or higher predicted risk of ICH than ischaemic stroke, and whether they performed better than existing risk scores.

Methods

Study design and participants

We used pooled individual patient data from the Microbleeds International Collaborative Network (MICON) of prospective observational studies, for which the full methodology and composition has been published. Briefly, MICON comprises 38 cohorts from 18 countries in North America, Europe, the Middle East, Asia, and Australasia, collectively including 20,322 participants with previous ischaemic stroke or TIA, baseline MRI including blood-sensitive paramagnetic sequences to detect CMBs, and at least three months' follow-up for ischaemic stroke, ICH, or a composite of both. We identified eligible cohorts through a systematic search of Medline and Embase from 01/01/1996 to 01/12/2018, clinical trial databases, scientific

abstracts, and the international METACOHORTS consortium of studies in cerebral small vessel disease. Published and unpublished studies were eligible. We assessed all studies identified for quality and risk of bias, including selection bias, using the Cochrane Collaboration tool. All included studies adjudicated events blinded to CMB burden. In the current prediction model development study, we included all MICON participants who were taking antithrombotic therapy and were followed up separately for ischaemic stroke or ICH. The study was approved by the UK Health Research Authority (reference: 8/HRA/0188). Included cohorts obtained ethical and regulatory approvals according to local requirements. Only fully-anonymised data was shared, so that individual consent was not required for this individual patient data pooled analysis. We registered the study protocol with the PROSPERO register of systematic reviews on April 5, 2016 (registration number: CRD42016036602, https://www.crd.york.ac.uk/prospero/display_record.php?RecordID=36602).

Outcomes

Our outcomes for prediction were the five-year risks of symptomatic ICH (including intracerebral, subdural, subarachnoid, and extradural haemorrhage) and ischaemic stroke (excluding TIA).

Prediction model development

We developed separate prediction models for ICH and ischaemic stroke using Cox regression, with robust standard errors calculated using the Huber-White sandwich estimator to allow for clustering within cohorts.¹⁸ We prespecified our candidate predictors, based on biological relevance and availability in the majority of our cohort, as: age; sex; presentation with transient ischaemic attack or ischaemic stroke; clinical history of hypertension; clinical history of type 1 or type 2 diabetes mellitus; previous ischaemic stroke before index stroke or TIA; previous

ICH; known AF; antithrombotic treatment after index event; CMB burden.; and type of MRI sequence used to detect CMBs (2D T2*-weighted gradient-recall echo (GRE) or susceptibilityweighted imaging (SWI, also including SWAN, SWIp and VenoBOLD sequences), in view of strong external evidence that CMB counts are systematically higher on these sequences than on GRE (appendix, p 3). We accounted for missing data using multiple imputation with chained equations (five imputations). We included a cluster-level variable indicating East Asian centres (Japan, Korea, China and South-East Asia), given the higher incidence of intracerebral haemorrhage and intracranial atherosclerosis in this region. 19 We categorised antithrombotic treatment as antiplatelet therapy only, anticoagulation with a VKA, or anticoagulation with a DOAC. The antiplatelet category included patients taking dual antiplatelets, and anticoagulant categories included participants taking a concomitant antiplatelet. We categorised CMB burden as none, one, two to four, five to ten, 11-19, and 20 or more, and assessed whether an interaction term between MRI sequence type and CMB burden was required. We investigated whether separate models were required for patients taking anticoagulants or antiplatelets using interaction terms and Wald tests. We simplified our models through backwards elimination at the 20% level (p=0·20). We scaled and rounded regression coefficients to produce integer scores for ease-of-use in clinical practice.

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Statistical analyses

We internally validated our models using bootstrapping.²⁰ As an additional test of model performance, we did internal-external cross validation,^{21,22} using five folds consisting of whole cohorts, repeated 20 times to reduce variance. We quantified discrimination using Harrell's cindex, and calibration through the calibration slope. We further assessed calibration by calculating predicted five-year risk for each outcome on the basis of the integer risk score,

dividing participants into lower, intermediate and highest-risk groups of roughly equal sizes, 339 and comparing predicted to observed risk using Kaplan-Meier plots. 340 341 To test the contribution of CMB burden to ICH and ischaemic stroke prediction, we developed purely clinical models in the same way as our main models, but excluding CMB burden and 342 MRI sequence type. We compared their discrimination to our main models, and tested if adding 343 CMB burden and MRI sequence type improved their fit. Next, we compared the performance 344 345 of our CMB-based ICH risk score (the form of our model that could most easily be used in clinical practice) to existing bleeding risk scores (ATRIA, ORBIT and HASBLED). Each 346 347 comparison used all participants for whom the additional variables required for calculation of the existing bleeding risk score were available. To apply HASBLED to patients not taking 348 vitamin K antagonists, we scored the 'labile INR' component as 0. As we made these 349 comparisons in a subset of the model development data, we adjusted for optimism using 350 bootstrapping. 351 We performed two sensitivity analyses. Firstly, we assessed the added predictive value of 352 additional variables that we considered potentially clinically relevant by adding each variable 353 individually to our final model for each outcome and testing if it improved model fit using a 354 Wald test²³, before comparing the discrimination of the base and augmented models if it did. 355 The additional variables were: clinical history of hypercholesterolaemia; current smoking 356 status; CMB distribution (strictly deep, strictly lobar, and mixed); and burden of white matter 357 hyperintensities on MRI assessed using the highest recorded Fazekas score from periventricular 358 and deep white matter regions. Secondly, we tested the performance of our ICH model for 359 intracerebral haemorrhage specifically. 360 Finally, we determined the number of participants with a predicted risk of ICH greater than 361 that of ischaemic stroke, and investigated their baseline characteristics. 362

Our statistical analyses used Stata version 16, and are reported following the TRIPOD guideline.²⁴

Role of the funding source

The funders of the study had no role in its design, the collection, analysis and interpretation of data, the writing of the report, or the decision to submit it for publication. All authors had full access to all the data in the study and final responsibility for the decision to submit for publication.

Results

Figure 1 describes the identification of studies in the MICON collaboration. From all 38 studies and 20,322 participants in the collaboration, we excluded one study comprising 3,335 participants that collected follow-up for a composite 'any stroke' outcome only. From the remaining 37 cohorts, we excluded 979 participants not taking antithrombotic medication, and a further 204 participants lacking follow-up for both ICH and ischaemic stroke, leaving a final study population of 15,784 participants, recruited between 28th August 2001 and 4th February 2018. Their characteristics are summarised in Table 1, and described by cohort in appendix pp4-6. All 15,784 participants had follow-up for ischaemic stroke, and 15,766 had follow-up for ICH. We imputed 2,747/15,784 (17.4%) observations for previous ICH, 2,002/15,784 (12.7%) for diabetes, and 1,097/15,784 (6.6%) for ischaemic stroke before index ischaemic stroke or TIA. We imputed fewer than 1% of observations for all other candidate predictors. During a total follow-up of 32,001 person-years for ICH (median 1·99yrs, IQR 0·61-2·87) and 31,468 person-years for ischaemic stroke (median 1·98yrs, IQR 0·56-2·80), 184 ICH

(including 146 intracerebral haemorrhages) and 1,048 ischaemic strokes occurred. The 387 annualised incidences were 0.57% for ICH, and 3.33% for ischaemic stroke. 388 389 Table 2 shows the hazard ratios from our final models for ICH and IS, and the resulting integer risk scores. Both models included age, CMB burden, MRI sequence type used to assess CMB 390 burden, history of ischaemic stroke prior to the index ischaemic stroke or TIA, and East Asian 391 392 centre location. Our ICH model also included previous ICH and antithrombotic treatment type. 393 We chose to retain antithrombotic treatment in this model on clinical grounds. Our ischaemic stroke model also included presentation with ischaemic stroke and history of diabetes mellitus, 394 395 and we found strong evidence of an interaction between antiplatelet treatment and AF (p = 0.0040), consistent with the known superior efficacy of anticoagulants for stroke prevention in 396 AF. We represented this in our model by combining AF, antithrombotic treatment type, and 397 their interaction into a single four-level variable, as the hazard ratios for DOAC and VKA 398 treatment were very similar. Appendix p7 shows the results of our other tests for interactions. 399 Apart from an interaction for ICH risk between antiplatelet use and previous ICH (p = 0.011), 400 which we attributed to treatment bias and chose to exclude, we found no compelling evidence 401 that other interaction terms were required. 402 The optimism-adjusted c-index for our final ICH model was 0.73 (95% CI 0.69–0.77), and the 403 calibration slope 0.94 (95% CI 0.81-1.06), indicating moderate discrimination and excellent 404 calibration. For our final ischaemic stroke model, the c-index was 0.63 (95% CI 0.62-0.65) 405 and the calibration slope 0.97 (95% CI 0.87-1.07), indicating reasonable discrimination and 406 excellent calibration. 407 In internal-external cross-validation, mean discrimination for ICH was 0.71, with a slightly 408 reduced mean calibration slope (0.85), partly explained by the reduced sample for model 409 development. Mean discrimination for IS was 0.60 and the mean calibration slope 0.76. For 410 each outcome, after combining participants into three groups on the basis of their total risk 411

score, we observed excellent agreement between predicted and observed risk (Figure 2, 412 appendix p 10). Figure 3 and appendix p11 show detailed calibration results for each outcome 413 across ten similarly-sized groups. Absolute ICH risk was moderately over-predicted in the 414 highest-risk decile. As 98.2% of participants received the same prediction across all five 415 imputations, we show calibration plots for the first imputation only. 416 The clinical-only models generated for comparison with our main, MRI-based models, 417 418 included the same variables as the main models apart from CMB burden and MRI sequence type. The clinical-only model for ICH showed reduced model fit and substantially lower 419 420 discrimination (difference in c-index 0.05, 95% CI 0.02 - 0.09, p < 0.0001). The clinical-only model for ischaemic stroke showed worse model fit (p = 0.00020) but similar discrimination 421 (c = 0.63 (95% CI 0.61-0.64)).422 Table 3 shows the results of comparisons between our new ICH risk score and the HASBLED, 423 ORBIT and ATRIA risk scores. Eleven cohorts from eight countries contributed to the 424 comparison for HASBLED, and eight cohorts from six countries to the comparison for ATRIA 425 and ORBIT. All comparisons included East Asian and European centres. For each comparison, 426 the estimate for the c-index of the new ICH risk score was higher, both in participants taking 427 any antithrombotics and when restricted to participants taking OAC. The optimism-adjusted 428 difference in c-index was substantial (range: 0.04 - 0.27) in all comparisons (*Table 3*), though 429 estimates were imprecise and the 95% confidence interval for comparisons with ATRIA and 430 ORBIT did not exclude 0. 431 In our planned sensitivity analyses, we found no evidence that any of the additional variables 432 tested improved model fit for ICH or ischaemic stroke (appendix p 8). The optimism-adjusted 433 c-index of our ICH model in predicting intracerebral haemorrhage specifically (rather than 434 intracranial haemorrhage in general) was 0.77 (95% CI 0.73-0.81), with calibration slope 0.95 435 (0.83-1.07). Having found evidence that using information on CMB burden from MRI 436

improves ICH prediction, we performed an additional sensitivity analysis testing the performance of our ICH prediction model according to MRI sequence type used. Performance was acceptable in both groups (*appendix p12*).

Of 11,953 participants for whom both risk scores could be calculated without imputed data, only 104 (0·87%) were in the 'highest risk' tertile for ICH and the 'lower risk' tertile for ischaemic stroke, in which the predicted five-year risks of ICH and ischaemic stroke were similar (6.7% and 7.2% respectively). Their baseline characteristics are described in *appendix p9*. An additional 999/11,953 participants (8·4%) were allocated to the 'highest risk' group for ICH and the 'intermediate risk' group for ischaemic stroke (predicted five-year risks 6.7% and 11·6% respectively). *Appendix p13* shows the full distribution of risk score predictions.

Discussion

Our most important result is the description of a novel risk score (MICON-ICH), including clinical variables and MRI-detected cerebral microbleeds, to predict ICH in patients taking antithrombotic therapy after ischaemic stroke or transient ischaemic attack. The addition of CMBs to a score based on clinical variables alone substantially improved performance, while a direct comparison with three existing bleeding risk scores also suggested superior discrimination of the new ICH risk score. Our risk score for ischaemic stroke showed modest discrimination, and CMBs appeared less important for predicting IS than ICH; nevertheless, this score can be used alongside our ICH score for straightforward and simultaneous estimation of ICH and ischaemic stroke risk. Both our scores showed excellent calibration in bootstrap validation, providing accurate estimates of absolute risk across low, medium, and high-risk groups. Discrimination was similar and calibration remained acceptable in internal-external

validation. A sensitivity analysis suggested that our ICH score might show higher

discrimination for the prediction of intracerebral haemorrhage specifically, the most serious form of non-aneurysmal ICH and the form most closely associated with cerebral microbleeds. Overall, the performance of our scores suggests they may be useful to estimate stroke risk and inform prognostication in clinical practice. Our scores have several features to ensure their ease-of-use in the clinical setting. Most importantly, they are simple: the clinical variables used are a standard part of the medical history for any stroke patient, and CMBs are familiar in stroke clinical practice (for example, in the diagnosis of cerebral amyloid angiopathy). CMBs are discrete lesions, which can be counted with very good inter-rater reliability,25 and the blood-sensitive GRE and SWI sequences required to image them (accounted for in our scores) are quick to acquire, widely available, and part of routine stroke imaging protocols in many centres. This offers an advantage over the use of serum biomarkers not usually measured clinically, as in the ABC bleeding score. Our scores include relatively few variables, allowing diagrammatic representation for quick reference (appendix pp14-15) and easy conversion to an online calculator or app. Finally, our scores are applicable to nearly all ischaemic stroke or TIA patients, whether taking antiplatelets or anticoagulants, with or without AF. Our scores are intended for use in patients in whom antithrombotic treatment is planned after ischaemic stroke or TIA. They are not applicable to patients in whom antithrombotic treatment is contraindicated, or for patients taking antithrombotics for primary prevention. They are not designed to help select the type of antithrombotic therapy to use (i.e. antiplatelet or anticoagulant), as this would require randomised data, rather observational data in which the relationship between antithrombotic type and outcomes is attenuated by selection bias. Rather, the MICON risk scores should be used to assess prognosis to inform clinical discussions and other aspects of care once the intended antithrombotic treatment has been chosen. The finding of a high predicted ICH risk might lead to more aggressive treatment of modifiable bleeding

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risk factors, such as hypertension and alcohol intake, review of concurrent medication, and consideration of non-pharmacological stroke prevention strategies if applicable, such as left atrial appendage occlusion in patients with AF. Our scores might also have applications in the selection of patients at high ICH risk for future clinical trials and mechanistic studies of ICH. The principal methodological strength of our study is the use of a large, multi-centre and truly international study population, increasing generalisability and allowing us to consider regional differences in stroke risk. We screened the prospective studies included for quality and risk of bias. These offered standardised baseline assessment and ascertainment of outcome events within each cohort, an advantage over registry-based studies, while we accounted statistically for within-cohort clustering. We performed both internal validation using bootstrapping and internal-external cross-validation, in accordance with TRIPOD guidelines and expert recommendations.^{22, 24} While we omitted some potentially clinically relevant variables from our model due to missing data, additional analyses suggested this did not reduce model performance. We acknowledge the limitations of our study. In particular, to maximise precision we used all available data to develop our scores. External validation of our scores in new data should be undertaken. While we compared our new ICH score to three existing bleeding risk scores, further comparison in a large, truly independent cohort would clarify the relative performance of these scores. Our model is applicable to antiplatelet and anticoagulant-treated patients, but we lacked data to make direct comparison with antiplatelet-specific scores such as Intracranial-B2LEED3S and S2TOP-BLEED, 9, 26-28 which should also be undertaken. Although large, our study cohort contained relatively few patients with very high CMB counts, reducing the precision of our estimates for ICH and ischaemic stroke risk in very high-risk categories. We lacked data on MRI field strength, which can influence CMB count, and on some additional risk factors which might have improved identification of high risk patients, notably cortical

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superficial siderosis, alcohol abuse, renal insufficiency and labile INR in VKA-treated patients. Hypertension, diabetes, and hyperlipidaemia were diagnosed according to local criteria for each cohort; we lacked data on their treatment, and on antithrombotic medication adherence. These factors may have reduced the association between these predictors and outcomes – for example, the unexpected absence of an association between hypertension and ICH. We did not have central formal adjudication of outcome events. Though we present data on the relative predicted risks of ICH and ischaemic stroke in our study sample, conclusions about the appropriateness of antithrombotic treatment are limited by the observational nature of our data. We also lacked data on functional outcomes, and it should be borne in mind that the morbidity and mortality of ICH is around twice that of ischaemic stroke.²⁹ Finally, our risk estimates are obtained from organised care systems with access to MRI, and may not be applicable to less developed settings. In summary, the MICON-ICH and MICON-IS scores we present here provide a new means by which to assess the long-term risk of ICH and ischaemic stroke. Although the MICON-ICH score appears promising and clinically useful, external validation is still required. Our results also clarify the relative predictive importance of CMBs for ICH and ischaemic stroke, and may facilitate the design of future randomised controlled trials of alternative stroke prevention strategies (e.g. of novel antithrombotic agents with potentially lower ICH risk) in patients at high predicted risk of ICH.

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Contributors

DJWe, DW, GA, and JM-F drafted the initial protocol, which was reviewed with critical revisions and approval by all authors. JGB and GA did the statistical analysis. JGB, GA and DJW accessed and verified the data, and wrote the first draft of the manuscript. All authors contributed to data acquisition, management, and brain imaging analyses. All authors contributed to critical revision of the manuscript and approved the final manuscript for submission.

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Declaration of Interests

MK reports grants from Ministry of Health, Labour and Welfare, Japan, and National Cerebral and Cardiovascular Center, and personal fees from Nippon Boehringer Ingelheim, Bayer Yakuhin, Daiichi-Sankyo Company, and Bristol-Myers Squibb (BMS)/Pfizer, outside the submitted work. HC reports personal fees from HOVID, outside the submitted work. EMA reports personal fees from Daiichi Sankyo, Pfizer, Bayer Healthcare, Nutricia, Abbott, Fresenius Kabi, and Sanofi, and grants from TUBITAK, outside the submitted work. JP reports personal fees from Abbott, Akcea, Boehringer Ingelheim, Daiichi Sankyo, and Pfizer, outside the submitted work. NB reports personal fees from Pfizer Israel, Ever Neuro Pharma, Shire Israel, and Boehringer Ingelheim Israel, outside the submitted work. DJS reports other funding from Bayer and Pfizer, outside the submitted work. PL reports other funding from Daiichi-Sankyo, Bayer, and Boehringer Ingelheim, outside the submitted work. GYHL reports consultancy and speaker fees from Bayer, Bayer/Janssen, BMS/Pfizer, Medtronic, Boehringer Ingelheim, Microlife, Roche, and Daiichi-Sankyo outside the submitted work. No fees received personally. UF reports grants from Medtronic, other funding from Medtronic, Stryker, and CSL Behring, and grants from Swiss National Science and Swiss Heart Foundation, outside the submitted work. DH reports grants from Bayer UCD Newman Fellowship, during the conduct of the study. JS reports grants from Adriana van Rinsum Ponsen Stichting, during the conduct of the study. NK reports personal fees from from Eisai Pjarmaceuticals, grants and personal fees from Novartis Pharmaceuticals and Schwabe Pharmaceuticals, and grants from Temasek Foundation, outside the submitted work. PJK reports grants from Health Research Board Ireland, during the conduct of the study. JMW reports grants from Wellcome Trust, Chest Heart Stroke Scotland, and Row Fogo Charitable Trust, during the conduct of the study; grants from Fondation Leducq, British Heart Foundation, UK DRI Ltd funded by Medical Research Council, Alzheimer's Society and Alzheimer's Research UK, and Stroke Association, outside the submitted work. VIHK reports grants from Netherlands Heart Foundation (grant 2001B071) during the conduct of the study. STE reports grants from Daiichi-Sankyo, and other funding from Bayer and BMS, outside the submitted work. NP reports grants from Swiss Heart Foundation, during the conduct of the study; other funding from Daiichi-Sankyo, Bayer, and Boehringer Ingelheim, and grants from Swiss National Science Foundation, outside the submitted work. EES reports personal fees from Alnylam Pharmaceuticals, Bayer, and Portola, outside the submitted work. VT reports personal fees from Boehringer Ingelheim, Pfizer, BMS, Bayer, and Medtronic, outside the submitted work. RV reports grants and personal fees from Bayer and BMS, grants from Boehringer, Daiichi-Sankyo, Medtronic and Biogen; and personal fees from Javelin, outside the submitted work. KKL reports grants, personal fees and nonfinancial support from Boehringer Ingelheim; grants and non-financial support from Pfizer; grants and personal fees from Amgen; grants from Eisai; grants and personal fees from Sanofi; and non-financial support from Daiichi Sankyo; outside the submitted work. PMR reports personal fees from Bayer, Abbott, and BMS, outside the submitted work. KT reports personal fees from Daiichi-Sankyo, Bayer Yakuhin, Bristol Myers Squibb, and Nippon Boehringer Ingelheim, outside the submitted work. BHJ reports grants from BMS Korea, Shinpoong Pharm. Co. Ltd., Bayer, Boehringer Ingelheim, and Daiichi-Sankyo, grants and personal fees

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from Esai, grants from AstraZeneca Korea, Servier Korea, Yuhan Corporation, Jeil Pharmaceutical Co, and Korean Drug Co., Ltd, grants and personal fees from Shire Korea, grants from JLK inspection, Chong Gun Dang Pharmaceutical Corp., and Dong-A Pharmaceutical, and personal fees from Amgen Korea, and Otsuka Korea, outside the submitted work. JMF reports grants from Instituto de Salud Carlos III, Fondo de Investigaciones Sanitarias, and Instituto de Salud Carlos III, RETICS INVICTUS PLUS RD16/0010/0019, during the conduct of the study. DJWe reports personal fees from Bayer, Alnylam and Portola outside the submitted work. All other authors declare no competing interests.

Acknowledgements

Funding for the included cohort studies was provided by the British Heart Foundation, Stroke Association, UCLH National Institute of Health Research (NIHR) Biomedical Research Centre, Wellcome Trust, Health Research Board Ireland, NIHR Biomedical Research Centre (Oxford, UK), Canadian Institutes of Health Research, Pfizer Cardiovascular Research award, Basel Stroke Funds, Science Funds Rehabilitation Felix-Platter-Hospital, Neurology Research Pool University Hospital Basel, Bayer AG, Fondo de Investigaciones Sanitarias Instituto de Salud Carlos III (FI12/00296; RETICS INVICTUS PLUS RD16/0019/0010; FEDER), Imperial College London NIHR Biomedical Research Centre, Dutch Heart Foundation, Servier, Association de Recherche en Neurologie Vasculaire and RHU TRT_cSVD (ANR-16-RHUS-004), Vidi innovational grant from The Netherlands ZonMw, Chest Heart Stroke Scotland, Medical Research Council, Fondation Leducq, The Row Fogo Charitable Trust, National Institute of Health (USA), Adriana van Rinsum-Ponsen Stichting, Japan Agency for Medical Research and Development (AMED), Ministry of Health, Labour and Welfare (Japan), and National Cerebral and Cardiovascular Center, Health and Medical Research Grant,

Singapore National Medical Research Council, and Dutch Heart Foundation. RS is part funded by the UCLH/UCL Biomedical Research Centre. RV is an investigator of Imperial BRC and partially funded by the European Union's Horizon 2020 research and innovation programme under grant agreement No. 754517 (PRESTIGE-AF).

Data Sharing Statement

Requests for access to anonymised study data for legitimate academic purposes may be directed to the corresponding author. Approval by the study steering committee and the principal investigator of each cohort in the study will be required before data can be shared.

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Figure Titles

Figure 1: Study flowchart

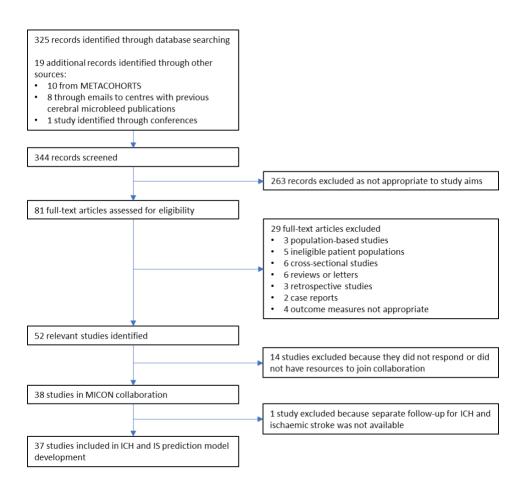


Figure 2: Kaplan-Meier plot and risk table for symptomatic ICH

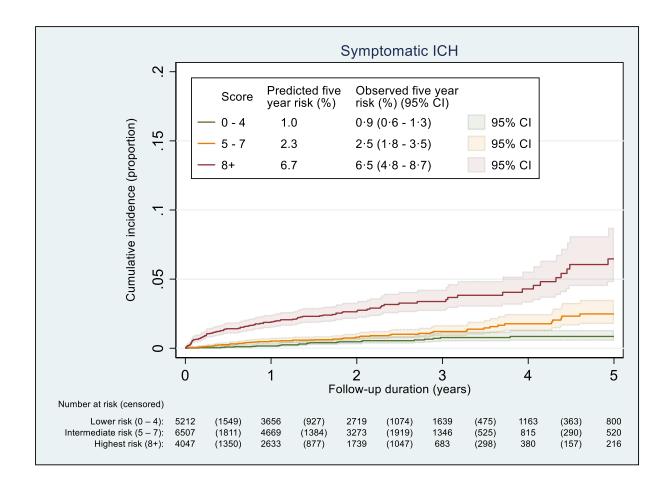
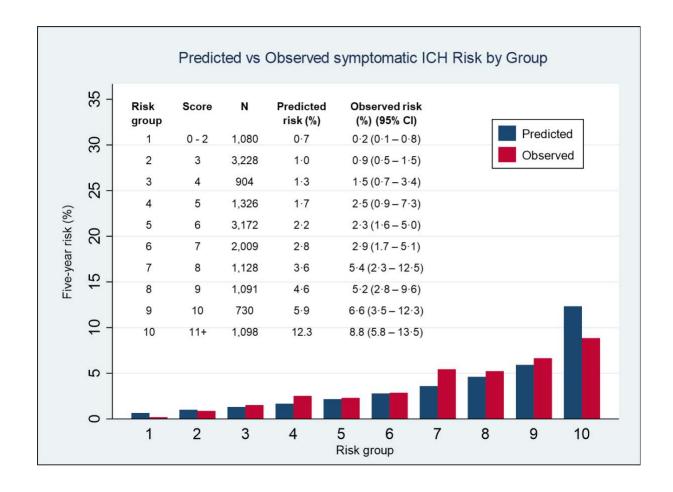


Figure 3: Model calibration – ICH



Tables

Table 1: Baseline characteristics

Values show prevalence for categorical variables, and mean (SD) for continuous variables.

Variable		Antiplatelet (n	Anticoagulant (n	Overall (n =
		= 8,736)	= 7,048)	15,784)
Age		67.4 (12.4)	74.7 (10.8)	70.7 (12.2)
Female sex		3,444/8,736	3,253/7,048	6,697/15,784
		(39.4%)	(46.2%)	(42.4%)
Male sex		5,292/8,736	3,795/7,048	9,087/15,784
		(60.6%)	(53.8%)	(57.6%)
East Asian populati	on	2,405/8,736	2,185/7,048	4,590/15,784
		(27.5%)	(31.0%)	(29·1%)
Hypertension		5,931/8,726	5,291/7,024	11,222/15,750
		(68.0%)	(75.33%)	(71.3%)
Atrial fibrillation		527/8,687	5,906/7,039	6,433/15,726
		(6.1%)	(83.9%)	(40.9%)
Diabetes mellitus (t	ype 1 or 2)	1,720/7,013	1,490/6,769	3,208/13,782
		(24.5%)	(22.0%)	(23.3%)
Ischaemic stroke be	efore	1,001/7,781	1,299/6,906	2,300/14,687
presenting stroke/T	IA	(12.9%)	(18.8%)	(16.5%)
Previous ICH		80/6,549	85/6,403 (1.31%)	165/12,872
		(1.22%)		(1.27%)
Presentation with ischaemic		6,632/8,735	6,172/7,039	12,804/15,774
stroke (vs TIA)		(75.9%)	(87.7%)	(81.2%)
CMB burden	0	6,418/8,733	5,202/6,970	11,620/15,703
		(73.5%)	(74.6%)	(74.0%)
	1	942/8,733	812/6,970	1,754/15,703
		(10.8%)	(11.7%)	(11.2%)
	2-4	785/8,733	671/6,970 (9.6%)	1,456/15,703
		(9.0%)		(9.3%)
	I .	1	1	1

	5-10	316/8,733	162/6,970 (2.3%)	478/15,703
		(3.6%)		(3.0%)
	11-19	157/8,733	59/6,970 (0.85%)	216/15,703
		(1.8%)		(1.4%)
	20 +	115/8,733	64/6,970 (0.92%)	179/15,703
		(1.3%)		(1.1%)
SWI sequence used	(vs T2*	2,422/8,734	2,335/7,025	4,757/15,759
GRE)		(27.7%)	(33.2%)	(30.2%)
Antithrombotic	AP only	8,736/8,736	NA	8,733/15,773
treatment		(100%)		(55.4%)
	Warfarin or	NA	4,752/7,037*	4,752/15,773
	VKA		(67.4%)	(30.1%)
	DOAC	NA	2,288/7,037	2,288/15,773
			(32.5%)	(14.5%)
Concomitant antiple	atelet with	NA	1,360/7,048	1,360/15,784
anticoagulant			(19.3%)	(8.6%)

^{*}Type of anticoagulant unknown for 11 participants

AP: antiplatelet; CMB: cerebral microbleed; DOAC: direct oral anticoagulant; VKA: vitamin

K antagonist; SWI: susceptibility-weighted imaging; GRE – gradient-recall echo

Table 2: Final models and risk scores for symptomatic ICH (MICON-ICH) and ischaemic stroke (MICON-IS)

			ICH		IS		ICH score	IS score
Predictor		Category	HR (95% CI)	P-	HR (95% CI)	P-	(/24)	(/34)
				value		value		
Number of	CMBS	0	1	<0.001	1	<0.001	0	0
		1	1.96 (1.38 - 2.80)		1.07 (0.86 - 1.34)		3	1
		2-4	2.18 (1.43 – 3.33)		1.29 (1.08 - 1.53)		3	2
		5-10	3.27 (1.71 - 6.24)		1.66 (1.21 - 2.27)		5	4
		11-19	4.93 (2.93 – 8.29)		*		6	4
		20+	9.26 (4.11 – 20.82)		1.91 (1.36 - 2.69)		9	5
T2*GRE se	equence used?	Yes	1.72 (0.80 - 3.70)	0.16	1.54 (0.82 - 2.89)	0.18	2	3
Age in year	·s	< 50	1	<0.001	1	<0.001	0	0
		50 - 59	1.05 (0.48 - 2.33)		1.03 (0.68 - 1.55)		0	0
		60 - 69	*		1.10 (0.77 - 1.57)		0	1
		70 -79	2.12 (0.95 - 4.75)		1.60 (1.11 - 2.29)		3	4
		80 +	2.66 (1.19 - 5.96)		1.72 (1.15 - 2.56)		4	4
East Asian	population	Yes	1.85 (0.82 - 4.15)	0.14	1.62 (0.78 - 3.37)	0.19	2	4
IS before p	resenting stroke/TIA	Yes	1.36 (1.00 - 1.87)	0.053	1.85 (1.48 - 2.31)	<0.001	1	5
ICH score	Previous ICH	Yes	3.91 (2.40 - 6.36)	<0.001	-	-	5	-
only	Antithrombotic	AP only	1.23 (0.69 - 2.18)	0.51	-	-	1	-

	treatment	Warfarin/VKA	1.30 (0.82 - 2.05)		-	-	1	-
		DOAC	1		-	-	0	-
IS score only	Presentation with ischaemic stroke	Yes	-	-	1.34 (0.91 - 1.98)	0.14	-	2
	Diabetes mellitus	Yes	-	-	1.32 (1.09 - 1.58)	0.004	-	2
	Antithrombotic	AP, has AF	-	-	3.14 (1.84 - 5.35)	<0.001	-	9
	treatment	AP, no AF	-	-	1.70 (1.16 - 2.51)		-	4
		OAC, other reason	-	-	1.36 (0.81 - 2.27)		-	2
		OAC, for AF	-	-	1		-	0

Baseline five-year survival for full ICH model: 99.53%; for full IS model: 97.15%

AF: atrial fibrillation; AP: antiplatelet; CMB: cerebral microbleed; DOAC: direct oral anticoagulant; GRE – gradient-recall echo; OAC (including vitamin K antagonists and direct oral anticoagulants); VKA: vitamin K antagonist

^{*} Category merged with preceding category to prevent inconsistent (non-monotonic) scoring

Table 3: Comparison of MICON-ICH score with existing bleeding risk scores

Comparator	Antithrombotics	N	C-index	C-index	Optimism-
			(Comparator)	(MICON)	adjusted
					difference
					(95% CI)
HASBLED*	All	5,510	0.47	0.75	0.27
					(0.18 - 0.37)
	OAC only	4,017	0.47	0.67	0.20
					(0.06 - 0.34)
ATRIA#	All	3,340	0.63	0.71	0.06
					(-0.06 - 0.18)
	OAC only	2,677	0.61	0.67	0.04
					(-0.08 - 0.17)
ORBIT#	All	3,340	0.60	0.71	0.09
					(-0.01 - 0.18)
	OAC only	2,677	0.58	0.67	0.08
					(-0.03 - 0.19)

^{*} Cohorts used for comparison: CROMIS-2, Graz, HERO, Kushiro City, NOACISP, IPAAC-Warfarin, SAMURAI-NVAF, TABASCO, UCLH, Wurzburg, Soo

^{*} Cohorts used for comparison: CROMIS-2, Graz, NOACISP, IPAAC-Warfarin, SAMURAI-NVAF, TABASCO, Soo

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List of MICON collaborators:

CROMIS-2: Kirsty Harkness, Louise Shaw, Jane Sword, Azlisham Mohd Nor, Pankaj Sharma, Deborah Kelly, Frances Harrington, Marc Randall, Matthew Smith, Karim Mahawish, Abduelbaset Elmarim, Bernard Esisi, Claire Cullen, Arumug Nallasivam, Christopher Price, Adrian Barry, Christine Roffe, John Coyle, Ahamad Hassan, Jonathan Birns, David Cohen, L Sekaran, Adrian Parry-Jones, Anthea Parry, David Hargroves, Harald Proschel, Prabel Datta, Khaled Darawil, Aravindakshan Manoj, Mathew Burn, Chris Patterson, Elio Giallombardo, Nigel Smyth, Syed Mansoor, Ijaz Anwar, Rachel Marsh, Sissi Ispoglou, Dinesh Chadha, Mathuri Prabhakaran, Sanjeevikumar Meenakishundaram, Janice O'Connell, Jon Scott, Vinodh Krishnamurthy, Prasanna Aghoram, Michael McCormick, Nikola Sprigg, Paul O'Mahony, Martin Cooper, Lillian Choy, Peter Wilkinson, Simon Leach, Sarah Caine, Ilse Burger, Gunaratam Gunathilagan, Paul Guyler, Hedley Emsley, Michelle Davis, Dulka Manawadu, Kath Pasco, Maam Mamun, Robert Luder, Mahmud Sajid, Jiaz Anwar, James Okwera, Elizabeth Warburton, Kari Saastamoinen, Timothy England, Janet Putterill, Enrico Flossman, Michael Power, Krishna Dani, David Mangion, Appu Suman, John Corrigan, Enas Lawrence, Djamil Vahidassr, Clare Shakeshaft, Martin M Brown, Andreas Charidimou, Hannah Cohen, Gargi Banerjee, Henry Houlden, Mark J White, Tarek A Yousry, Kirsty Harkness, Enrico Flossmann, Nigel Smyth, Louise J Shaw, Elizabeth Warburton, Keith W Muir. Bern: Leonidas Panos, Pascal Gratz, Heinrich Mattle. TABASCO: Amos D Korczyn, Efrat Kliper. PERFORM-MRI: Philippe Maeder, Achim Gass, Chahin Pachai, Luc Bracoub, Marie-Yvonne Douste-Blazy, Marie Dominique Fratacci, Eric Vicaut. SAMURAI NVAF: Shoichiro Sato, Kaori Miwa, Kyohei Fujita, Toshihiro Ide. Monash Stroke: Henry Ma, John Ly, Shaloo Singhal, Ronil Chandra, Lee-anne Slater, Cathy Soufan, Christopher Moran. Basel TIA and NOACISP: Christopher Traenka, Sebastian Thilemann, Joachim Fladt, Henrik Gensicke, Leo H Bonati. SNUBH Stroke cohort: Beom Joon Kim, Moon-Ku Han, Jihoon Kang, Eunbin Ko, Mi Hwa Yang, Myung Suk Jang. BIOSTROKE/TIA-Dublin: Sean Murphy, Fiona Carty, Layan Akijian, John Thornton, Mark Schembri. CASPER: Elles Douven. HERO: Raquel Delgado-Mederos, Rebeca Marín, Pol Camps-Renom, Daniel Guisado-Alonso, Fidel Nuñez, Santiago Medrano-Martorell, Elisa Merino. HAGAKURE: Kotaro Iida, Syuhei Ikeda, Hiroyuki Irie. Orken: Derya Selcuk Demirelli. CATCH: Jayesh Modi Medanta, Charlotte Zerna. MSS2: Maria C Valdés Hernández, Paul Armitage, Anna K Heye, Susana Muñoz-Maniega, Eleni Sakka, Michael J Thrippleton, Martin Dennis. Sainte-Anne: Ysoline Beigneux, Mauro Silva. Singapore: Narayanaswamy Venketasubramanian. HKU: Shu Leung Ho, Raymond Tak Fai Cheung, Koon Ho Chan, Kay Cheong Teo, Edward S Hui, Joseph Shiu Kwong Kwan, Richard S.K. Chang, Man Yu Tse, Chu Peng Hoi, Chung Yan Chan, Oi Ling Chan, Ryan Hoi Kit Cheung, Edmund Ka Ming Wong, IPAAC: Kam Tat Leung, Suk Fung Tsang, Hing Lung Ip, Sze Ho Ma, Karen Ma, Wing Chi Fong, Siu Hung Li, Richard Li, Ping Wing Ng, Kwok Kui Wong, Wenyan Liu, Lawrence Wong. MICRO: Lino Ramos, Els De Schryver, Joost Jöbsis, Jaap van der Sande, Paul Brouwers, Yvo Roos, Jan Stam, Stef Bakker, Henk Verbiest, Wouter Schoonewille, Cisca Linn, Leopold Hertzberger, Maarten van Gemert, Paul Berntsen. PARISK: Dianne van Dam-Nolen, M Eline Kooi, Aad van der Lugt, Peter J. Koudstaal. SIGNaL: Alexander Leff, Nicholas Ward, Parashkev Nachev, Richard J Perry, Hatice Ozkan, John Mitchell.

Supplementary Table 1: MRI sequence type and cerebral microbleed detection

Summarises studies comparing SWI and SWAN sequences to 2D GRE in the same patients

Study*	Population	N	Sequence	Prevalence (%) (SWI/SWAN)	Prevalence (%) (GRE)	Summary statistics# (SWI/SWAN)	Summary statistics# (GRE)
Vernooij 2008 ¹	General older population	200	SWI	71/200 (35.5)	42/200 (21.0)	Median 2.5 IQR: 1 – 9.5	Median 1 IQR: 1 - 4
Mori 2008 ²	Moya-Moya disease	50	SWI	21/50 (42.0)	16/50 (32.0)	-	-
Nandigam 2009 ³	Cerebral amyloid angiopathy	3	SWI	3/3 (100.0	3/3 (100.0)	Mean: 103.3	GRE: 34.3
Goos 2011 ⁴	Memory clinic patients	141	SWI	56/141 (39.7)	32/141 (22.7)	Median: 2 Range: 1 - 129	Median: 1 Range: 1 - 144
Cheng 2013 ⁵	Cerebral amyloid angiopathy	9	SWI	-	-	Median: 111 IQR: 48 – 192	Median: 57 IQR: 45 - 187
	Healthy controls	21	SWI	4/21 (19.0)	3/21 (14.3)	Median: 2	Median: 1
Guo 2013 ⁶	Hypertensive older population	273	SWI SWAN	SWI: 83/273 (30.4) SWAN: 88/273 (32.2)	54/273 (19.8)	SWI: Median: 8 Range: 1 – 15 SWAN: Median 8 Range: 1 - 17	GRE: Median 3 Range: 1 - 11
Shams 2015 ⁷	Memory clinic patients	246	SWI	50/246 (20.3)	43/246 (17.5)	Mean: 2.15	Mean: 1.48
Shao 2017 ⁸	Lacunar ischaemic stroke	60	SWI	26/60 (43.3)	15/60 (25.0)	-	-
	Healthy controls	60	SWI	8/60 (13.3)	4/60 (6.7)	-	-

[#]For patients with microbleeds detected

Supplementary Table 2: Baseline characteristics by cohort

Cohort	Location	N	Age (y)	Femal e Sex	AF	HTN	DM	Previo us IS	Previou s ICH	Index Event IS	SWI Used	CMB Presence	AP Only	VKA	DOAC	Follow Up (y)	ICH events (%)	IS events (%)
CROMIS-	UK	143	75.9	605/14	1435/1	897/141	241/14	137/14	8/1416	1199/143	0/1435	300/1435	36/143	874/14	525/14	2.34	14	56
29		5	(10.4)	35	435	3	34	12	(0.6)	5	(0.0)	(20.9)	5	35	35	(1.00)	(0.98)	(3.9)
				(42.2)	(100.0)	(63.5)	(16.8)	(9.7)		(83.6)			(2.5)	(60.9)	(36.6)			
HBS	US	504	67.7	209/50	120/50	373/504	140/50	116/50	-	454/504	2/504	71/504	394/50	109/50	1/504	0.23	0 (0)	0
			(15.4)	(41.5)	(23.8)	(74.0)	(27.8)	(23.0)		(90.1)	(0.4)	(14.1)	(78.2)	(21.6)	(0.2)	(0.05)	(0)	(0)
Bern ¹⁰	Switzerlan	245	66.6	106/24	79/202	144/245	32/245	33/245	_	245/245	245/24	49/245	171/24	66/245	8/245	0.30	0	5
Бети	d	243	(13.8)	5	(39.1)	(58.8)	(13.1)	(13.5)		(100.0)	5	(20.0)	5	(26.9)	(3.3)	(0.09)	(0)	(2)
			()	(43.3)	()	()	()	()		()	(100.0)	(,	(69.8)	()	()	(*****)	(-)	. ,
CU-	Hong	516	67.5	217/51	32/516	362/516	173/51	67/516	9/516	437/516	231/51	117/516	492/51	24/516	0/516	1.31	2	14
STRIDE ¹¹	Kong		(11.1)	6	(6.2)	(70.2)	6	(13.0)	(1.7)	(84.7)	6	(22.7)	6	(4.7)	(0.0)	(0.37)	(0.39)	(2.7)
				(42.1)			(33.5)				(44.8)		(95.3)					
TABASCO	Israel	378	67.4	171/37	29/374	221/374	89/374	0/378	0/378	275/378	0/378	59/378	345/37	33/378	0/378	4.06	0	54
12			(9.8)	8 (45.2)	(7.8)	(59.1)	(23.8)	(0.0)	(0.0)	(72.8)	(0.0)	(15.6)	8 (91.3)	(8.7)	(0.0)	(1.39)	(0)	(14)
Graz	Austria	385	65.9	142/38	91/385	299/385	84/385	77/385	5/385	342/385	0/385	75/385	315/38	58/385	12/385	1.75	13	52
Graz	Austria	363	(12.4)	5	(23.6)	(77.7)	(21.8)	(20.0)	(1.3)	(88.8)	(0.0)	(19.5)	515/56	(15.1)	(3.1)	(1.85)	(3.4)	(14)
			(12.1)	(36.9)	(23.0)	(,,,,,	(21.0)	(20.0)	(1.5)	(66.6)	(0.0)	(15.5)	(81.8)	(13.1)	(3.1)	(1.05)	(3.1)	(11)
PERFORM	France	105	67.7	370/10	16/105	887/105	324/10	120/10	3/1056	929/1056	0/1056	381/1056	1056/1	0/1056	0/1056	2.32	10	94
-MRI ¹³		6	(8.0)	56	6	6	56	56	(0.3)	(88.0)	(0.0)	(36.1)	056	(0.0)	(0.0)	(0.68)	(0.95)	(8.9)
14				(35.0)	(1.5)	(84.0)	(30.7)	(11.4)					(100.0)					
PARISK ¹⁴	Netherland	220	70.9	65/220	0/220	151/220	50/220	64/220	3/220	98/220	0/218	59/220	220/22	0/220	0/220	2.09	0	10
	S		(9.1)	(29.5)	(0.0)	(68.6)	(22.7)	(29.1)	(1.4)	(44.5)	(0.0)	(26.8)	0 (100.0)	(0.0)	(0.0)	(0.46)	(0)	(4.5)
SAMURAI	Japan	105	77.2	445/10	1051/1	977/105	208/10	225/10	19/1051	1007/105	771/10	250/1051	12/105	598/10	441/10	1.63	10	72
-NVAF ¹⁵		1	(9.8)	51	051	1	51	51	(1.8)	1	51	(23.8)	1	51	51	(0.72)	(0.95)	(6.9)
DI DI DI IGI	37.1.1.1	170	64.7	(42.3)	(100.0)	(93.0)	(19.8)	(21.4)	1/170	(95.8)	(73.4)	25/170	(1.1)	(56.9)	(42.0)	1.76		22
RUNDMC ¹	Netherland	178	64.7	63/178	19/178	144/178	36/178	46/178	1/178	90/178	0/178	35/178	159/17	19/178	0/178	4.76	2	23
	S		(8.7)	(35.4)	(10.7)	(80.9)	(20.2)	(25.8)	(0.6)	(50.6)	(0.0)	(19.7)	8 (89.3)	(10.7)	(0.0)	(0.75)	(1.1)	(13)
Wurzburg	Germany	343	70.7	154/34	99/343	276/343	73/343	77/343	12/343	270/343	151/34	75/343	219/34	40/341	82/341	0.34	1	19
Warzoung	Cermany	0.0	(13.3)	3	(28.9)	(80.5)	(21.3)	(22.4)	(3.5)	(78.7)	3	(21.9)	1	(11.7)	(24.0)	(0.22)	(0.29)	(5.5)
			()	(44.9)	()	()	(",		()	()	(44.0)		(64.2)		(,	()	()	()
Monash	Australia	356	75.0	172/35	356/35	283/356	92/355	97/356	6/356	305/356	336/35	153/356	0/356	319/35	37/356	1.74	7	9
Stroke ¹⁷			(10.7)	6	6	(79.5)	(25.9)	(27.2)	(1.7)	(85.7)	6	(43.0)	(0.0)	6	(10.4)	(1.24)	(2)	(2.5)
				(48.3)	(100.0)			121101		21121	(94.4)			(89.6)	21121			
Basel TIA ¹⁸	Switzerlan	181	69.3	67/181 (37.0)	24/181	134/181	31/181	13/181	-	0/181	0/181	20/181	148/18	33/181	0/181	0.25	0	24
	d		(12.3)	(37.0)	(13.3)	(74.0)	(17.1)	(7.2)		(0.0)	(0.0)	(11.0)	(81.8)	(18.2)	(0.0)	(0.00)	(0)	(13)
Yonsei ¹⁹	South	488	70.3	278/48	488/48	381/488	117/48	87/488	13/488	460/488	0/488	146/488	1/488	487/48	0/488	2.63	7	46
	Korea		(10.5)	8	8	(78.1)	8	(17.8)	(2.7)	(94.3)	(0.0)	(29.9)	(0.2)	8	(0.0)	(1.58)	(1.4)	(9.4)
				(57.0)	(100.0)		(24.0)							(99.8)				

BIO-	Ireland	240	67.9	91/240	73/236	141/238	38/237	19/236	-	89/240	0/240	24/240	167/24	73/240	0/240	0.47	0	13
STROKE/T IA ²⁰			(13.3)	(37.9)	(30.9)	(59.2)	(16.0)	(8.1)		(37.1)	(0.0)	(10.0)	0 (69.6)	(30.4)	(0.0)	(0.35)	(0)	(5.4)
Kushiro	Japan	631	71.5	257/63	86/631	407/631	182/63	115/63	17/631	631/631	0/631	268/631	568/63	63/631	0/631	0.15	20	99
City ²¹			(11.1)	(40.7)	(13.6)	(64.5)	(28.8)	(18.2)	(2.7)	(100.0)	(0.0)	(42.5)	(90.0)	(10.0)	(0.0)	(0.21)	(3.2)	(16)
IPAAC-	Hong	81	71.3	40/81	81/81	56/81	27/81	25/81	1/81	65/81	71/81	24/81	0/81	81/81	0/81	2.10	3	5
Warfarin ²²	Kong		(9.1)	(49.4)	(100.0)	(69.1)	(33.3)	(30.9)	(1.2)	(80.2)	(87.7)	(29.6)	(0.0)	(100.0)	(0.0)	(1.03)	(3.7)	(6.2)
CASPER ²³	Netherland	133	65.8	38/133	16/133	94/133	18/133	10/133	0/133	133/133	133/13	79/133	115/13	10/133	8/133	1.21	0	3
	S		(10.6)	(28.6)	(12.0)	(70.7)	(13.5)	(7.5)	(0.0)	(100.0)	(100.0)	(59.4)	(86.5)	(7.5)	(6.0)	(0.17)	(0)	(2.3)
HERO ²⁴	Spain	935	77.6	487/93	856/93	693/933	212/93	246/93	8/933	803/925	0/935	247/935	1/934	623/93	310/93	1.92	18	32
	- F		(6.6)	5	3	(74.3)	2	3	(0.9)	(86.8)	(0.0)	(26.4)	(0.1)	4	4	(0.58)	(1.9)	(3.4)
				(52.1)	(91.7)		(22.7)	(26.4)						(66.7)	(33.2)			
HAGAKU	Japan	350	73.1	141/35	102/35	263/347	116/35	50/350	10/349	317/350	28/350	127/350	197/35	93/350	60/350	2.15	9	23
RE			(13.0)	0	0	(75.8)	0	(14.3)	(2.9)	(90.6)	(8.0)	(36.3)	0	(26.6)	(17.1)	(1.08)	(2.6)	(6.6)
Leuven ²⁵	Belgium	487	72.2	(40.3) 192/48	(29.1) 103/48	313/487	(33.1) 92/487	61/487	_	354/487	0/487	129/487	(56.3) 354/48	133/48	0/487	2.12	4	32
Leuven	Beigiuiii	407	(9.4)	7	7	(64.3)	(18.9)	(12.5)	-	(72.7)	(0.0)	(26.5)	7	7	(0.0)	(0.72)	(0.82)	(6.6)
			(>)	(39.4)	(21.1)	(01.0)	(10.5)	(12.0)		(, 2, ,)	(0.0)	(20.0)	(72.7)	(27.3)	(0.0)	(01,2)	(0.02)	(0.0)
NOACISP	Switzerlan	290	78.3	132/29	290/29	226/290	55/289	49/289	12/289	262/290	284/29	79/290	10/290	67/290	213/29	1.84	9	19
	d		(9.1)	0	0	(77.9)	(19.0)	(17.0)	(4.2)	(90.3)	0	(27.2)	(3.4)	(23.1)	0	(0.74)	(3.1)	(6.6)
NG: T 26	CI.	106	64.4	(45.5)	(100.0)	00/10/		10/10/	7/10/	106/106	(97.9)	26/106	02/106	7/10/	(73.4)	0.20		
Min Lou ²⁶	China	106	64.4 (12.0)	34/106 (32.1)	16/106 (15.1)	80/106 (75.5)	-	18/106 (17.0)	7/106 (6.6)	106/106 (100.0)	106/10 6	36/106 (34.0)	92/106 (86.8)	7/106 (6.6)	7/106 (6.6)	0.39 (0.27)	0 (0)	(1.9)
			(12.0)	(32.1)	(13.1)	(13.3)		(17.0)	(0.0)	(100.0)	(100.0)	(34.0)	(80.8)	(0.0)	(0.0)	(0.27)	(0)	(1.9)
MICRO ²⁷	Netherland	397	65.3	165/39	30/396	218/397	54/397	35/397	0/397	35/397	0/397	72/397	357/39	40/397	0/397	3.25	11	21
	S		(12.2)	7	(7.6)	(54.9)	(13.6)	(8.8)	(0.0)	(8.8)	(0.0)	(18.1)	7	(10.1)	(0.0)	(1.63)	(2.8)	(5.3)
20				(41.6)									(89.9)					_
Orken ²⁸	Turkey	452	71.9 (12.1)	233/45	353/45 2	356/452 (78.8)	150/45 2	123/45 2	0/452 (0.0)	432/452 (95.6)	250/45 2	132/452 (29.2)	0/452 (0.0)	321/45 2	131/45 2	2.59 (2.07)	(0.66)	8 (1.8)
			(12.1)	(51.5)	(78.1)	(70.0)	(33.2)	(27.2)	(0.0)	(93.0)	(55.3)	(29.2)	(0.0)	(71.0)	(29.0)	(2.07)	(0.00)	(1.6)
CATCH ²⁹	Canada	392	67.6	154/39	26/392	218/392	54/392	0/392	0/392	236/392	0/392	62/392	325/39	67/392	0/392	0.26	1	13
			(13.9)	2	(6.6)	(55.6)	(13.8)	(0.0)	(0.0)	(60.2)	(0.0)	(15.8)	2	(17.1)	(0.0)	(0.09)	(0.26)	(3.3)
				(39.3)									(82.9)					
MSS2 ³⁰	UK	209	66.4	82/209	21/209	157/209	-	29/209	0/209	209/209	199/20	34/209	188/20	21/209	0/209	1.08	0	31
			(11.4)	(39.2)	(10.0)	(75.1)		(13.9)	(0.0)	(100.0)	9 (95.2)	(16.3)	9 (90.0)	(10.0)	(0.0)	(0.30)	(0)	(15)
Sainte-	France	302	78.6	154/30	302/30	215/302	56/302	39/302	6/302	302/302	0/279	80/302	0/302	122/30	180/30	1.53	5	20
Anne, Paris	Trance	302	(10.9)	2	2	(71.2)	(18.5)	(12.9)	(2.0)	(100.0)	(0.0)	(26.5)	(0.0)	2	2	(0.81)	(1.7)	(6.6)
			` /	(51.0)	(100.0)	, ,	, , ,	` /	. ,	, i	, ,	` '	, í	(40.4)	(59.6)		. ,	. ,
STROKDE	France	178	64.0	68/178	12/178	100/178	23/178	20/178	1/178	178/178	0/178	23/178	130/17	40/178	8/178	3.32	0	16
M			(12.7)	(38.2)	(6.7)	(56.2)	(12.9)	(11.2)	(0.6)	(100.0)	(0.0)	(12.9)	8	(22.5)	(4.5)	(1.61)	(0)	(9)
NUS	Cinconos	41	66.6	12/41	10/41	32/41	11//1	2/41	0/41	41/41	41/41	22/41	(73.0) 26/41	15/41	0/41	3.01	0	5
(Chen)	Singapore	41	66.6 (10.2)	12/41 (29.3)	(24.4)	(78.0)	11/41 (26.8)	(4.9)	(0.0)	(100.0)	(100.0)	(53.7)	(63.4)	(36.6)	(0.0)	(1.32)	(0)	(12)
FUTURE	Netherland	18	44.5	9/18	0/18	7/18	0/18	0/18	0/18	12/18	18/18	1/18	18/18	0/18	0/18	0.67	0	4
	S		(5.3)	(50.0)	(0.0)	(38.9)	(0.0)	(0.0)	(0.0)	(66.7)	(100.0)	(5.6)	(100.0)	(0.0)	(0.0)	(0.72)	(0)	(22)

Heidelberg ³	Germany	607	64.3	225/60	110/60	465/607	-	92/607	1/607	501/607	607/60	138/607	488/60	109/60	10/607	4.00	3	28
1			(14.0)	7	7	(76.6)		(15.2)	(0.2)	(82.5)	7	(22.7)	7	7	(1.6)	(1.27)	(0.49)	(4.6)
				(37.1)	(18.1)						(100.0)		(80.4)	(18.0)				
NNI	Singapore	182	57.7	56/182	28/181	142/182	59/182	26/182	0/182	182/182	0/182	49/182	150/18	23/182	9/182	0.80	0	0
			(11.5)	(30.8)	(15.5)	(78.0)	(32.4)	(14.3)	(0.0)	(100.0)	(0.0)	(26.9)	2	(12.6)	(4.9)	(0.63)	(0)	(0)
													(82.4)					
OXVASC ³²	UK	106	68.3	508/10	164/10	581/106	-	-	-	502/1067	0/1067	157/1067	949/10	112/10	6/1067	3.41	11	78
		7	(14.0)	67	66	6				(47.0)	(0.0)	(14.7)	67	67	(0.6)	(1.53)	(1)	(7.3)
				(47.6)	(15.4)	(54.5)							(88.9)	(10.5)				
HKU ³²	Hong	966	68.9	388/96	124/96	628/966	272/96	93/966	12/966	966/966	966/96	433/966	862/96	63/966	41/966	2.90	19	89
	Kong		(12.2)	6	6	(65.0)	6	(9.6)	(1.2)	(100.0)	6	(44.8)	6	(6.5)	(4.2)	(1.49)	(2)	(9.2)
				(40.2)	(12.8)		(28.2)				(100.0)		(89.2)					
Soo ³³	Hong	178	73.4	82/178	175/17	155/178	50/178	34/178	3/178	152/178	178/17	66/178	5/178	7/178	166/17	1.85	1	5
	Kong		(9.6)	(46.1)	8	(87.1)	(28.1)	(19.1)	(1.7)	(85.4)	8	(37.1)	(2.8)	(3.9)	8	(1.44)	(0.56)	(2.8)
					(98.3)						(100.0)				(93.3)			
SIGNaL	UK	206	72.4	85/206	65/206	146/206	49/206	55/206	8/206	185/206	140/20	92/206	163/20	9/206	34/206	0.60	1	24
			(14.0)	(41.3)	(31.6)	(70.9)	(23.8)	(26.7)	(3.9)	(89.8)	6	(44.7)	6	(4.4)	(16.5)	(0.20)	(0.49)	(12)
											(68.0)		(79.1)					
Total		157	70.7	6697/1	6882/1	11222/1	3208/1	2300/1	165/130	12804/15	4757/1	4164/157	8733/1	4759/1	2289/1	2.03	184	1048
		84	(12.2)	5784	5728	5750	3782	4687	37	774	5759	84	5781	5781	5781	(1.53)	(1.2)	(6.6)
				(42.4)	(43.8)	(71.3)	(23.3)	(15.7)	(1.3)	(81.2)	(30.2)	(26.4)	(55.3)	(30.2)	(14.5)			

Values shown are prevalence (%) or mean (SD). "ICH event (%)" and "IS event (%)" refer to the number and percentage of each cohort who experienced an event during follow-up. Studies without references are unpublished. FUTURE: Follow-Up of Transient ischemic attack and stroke patients and Unelucidated Risk factor Evaluation study. HAGAKURE: Hypertension, Amyloid, and aGe Associated Kaleidoscopic brain lesions on CT/MRI Undertaken with stroke REgistry. HBS: Heart Brain Interactions Study. NNI: National Neuroscience Institute, Singapore. NOACISP: Novel Oral Anticoagulants in Stroke Patients, Basel; NCT02353585. SIGNaL: Stroke Investigation in North and Central London. STROKDEM: Study of Factors Influencing Post-stroke Dementia.

Supplementary Table 3: Interaction terms

Each interaction was tested individually as an addition to a model comprising all candidate predictors for each outcome. The association of each variable tested is shown at each level of the interacting variable, including the interaction but not the main effect of the interacting variable. When testing interactions with antiplatelet vs anticoagulant treatment, we omitted the three-level antithrombotic treatment to avoid collinearity. CMB recoded describes CMB burden following recategorisation as a four-level variable to reduce sparseness.

A: Interactions with antithrombotic treatment

Variable	Anticoagulant (HR, 95% CI)	Antiplatelet (HR, 95% CI)	P-value for interaction
ICH			
CMB 0	1	1	0.36
CMB 1	2.11 (1.28 – 3.48)	1.72 (0.94 – 3.19)	
CMB 2 - 4	2.01 (0.98 – 4.11)	2.33 (1.37 – 3.96)	
CMB 5 - 10	1.25 (0.32 – 4.90)	4.66 (2.47 – 8.80)	
CMB 11 - 19	5.67 (2.17 – 14.8)	4.53 (2.74 – 7.49)	
CMB 20+	3.15 (0.42 – 23.45)	15.01 (7.06 – 31.92)	
Age (years)	1.04 (1.02 – 1.06)	1.03 (1.02 – 1.05)	0.66
Female sex	1.16 (0.81 – 1.66)	0.87 (0.54 – 1.40)	0.29
Presentation with ischaemic stroke	1.19 (0.49 – 2.89)	0.91 (0.34 – 2.48)	0.71
SWI MRI sequence used	0.73 (0.43 – 1.23)	0.42 (0.14 – 1.26)	0.28
Atrial fibrillation present	0.70 (0.29 – 1.73)	1.68 (0.96 – 2.91)	0.09
Hypertension present	0.85(0.55-1.30)	1.17 (0.62 – 2.23)	0.41
Diabetes present	1.63 (0.97 – 2.73)	0.83 (0.45 - 1.50)	0.10
Ischaemic stroke before index event	1.30 (0.92 – 1.84)	1.41 (0.84 – 2.39)	0.81
Previous intracranial haemorrhage	1.98 (0.83 – 4.74)	7.39 (4.11 – 13.29)	0.011
East Asian population	1.17 (0.70 – 1.99)	3.21 (0.96 – 10.7)	0.09
Ischaemic stroke			
CMB 0	1	1	054
CMB 1	0.92 (0.66 – 1.28)	1.16 (0.87 – 1.53)	
CMB 2 - 4	1.16 (0.91 – 1.48)	1.32 (1.06 – 1.66)	
CMB 5 - 10	0.95(0.47-1.91)	1.98 (1.32 – 2.96)	
CMB 11 - 19	1.43 (0.69 – 2.93)	1.44 (0.68 – 3.06)	
CMB 20+	1.85 (0.78 – 4.39)	1.94 (1.29 – 2.92)	
Age (years)	1.02 (1.00 – 1.03)	1.02 (1.01 – 1.03)	0.69
Female sex	1.10 (0.88 – 1.38)	0.88 (0.71 – 1.09)	0.15
Presentation with ischaemic stroke	1.14 (0.71 – 1.82)	1.42 (0.93 – 2.15)	0.39
GRE MRI sequence used	0.94 (0.60 – 1.45)	0.50 (0.23 – 1.10)	0.04
Atrial fibrillation present	0.72 (0.43 – 1.21)	1.81 (1.30 – 2.50)	0.0040
Hypertension present	1.07 (0.79 – 1.45)	1.07 (0.74 – 1.55)	0.98
Diabetes present	1.37 (1.08 – 1.73)	1.25 (1.03 – 1.53)	0.53
Ischaemic stroke before index event	1.82 (1.22 – 2.71)	1.86 (1.48 – 2.34)	0.92
Previous intracranial haemorrhage	1.04 (0.49 – 2.22)	1.63 (0.79 – 3.35)	0.40
East Asian population	1.77 (1.12 – 2.81)	1.45 (0.49 – 4.28)	0.67

B: Interactions with MRI sequence type

Variable	GRE	SWI	P-value for interaction
ICH			
CMB 0	1	1	0.50
CMB 1	2.28 (1.59 – 3.26)	1.09 (0.35 – 3.40)	
CMB 2 - 4	2.50 (1.60 – 3.89)	1.32 (0.51 – 3.37)	
CMB 5 - 10	2.91 (1.27 – 6.66)	3.40 (1.21 – 9.48)	
CMB 11 - 19	5.09 (2.82 – 9.19)	4.20 (1.87 – 9.45)	1
CMB 20+	9.14 (3.22 – 25.93)	8.35 (2.17 – 32.13)	
Ischaemic stroke			
CMB 0	1	1	0.0065
CMB 1	1.09 (0.80 – 1.49)	0.99 (0.69 – 1.41)	
CMB 2 - 4	1.30 (1.12 – 1.50)	1.17 (0.75 – 1.81)	
CMB 5 - 10	1.91 (1.24 – 2.93)	1.28 (0.77 – 2.14)	
CMB 11 - 19	2.19 (1.14 – 4.21)	0.44 (0.21 - 0.95)	
CMB 20+	1.92 (1.29 – 2.84)	1.83 (0.95 – 3.50)	
CMB recoded 0	1	1	0.18
CMB recoded 1	1.09 (0.80 – 1.49)	0.99 (0.69 – 1.41)	1
CMB recoded 2 – 4	1.30 (1.12 – 1.50)	1.17 (0.75 – 1.81)	
CMB recoded 5+	1.97 (1.38 – 2.82)	1.17 (0.75 – 1.83)	

Supplementary Table 4: Additional variables

Predictor	Prevalence or Median	HR (95% CI)*	P-value
ICH			
Hyperlipidaemia	5,880/13,128 (44.8%)	1.02 (0.66-1.57)	0.94
Current smoker	1,708/10,357 (16.5%)	1.01 (0.67-1.53)	0.94
Fazekas score (continuous)	1 (IQR 1 – 2)	1.06 (0.71-1.59)	0.78
Fazekas score 2+	3,777/9,366 (40.2%)	1.47 (0.41-1.70)	0.21
Strictly deep CMBs	1005/11,877 (8.5%)	1.55 (0.73-3.31)	0.26
Strictly lobar CMBs	1,146/11,874 (9.7%)	0.69 (0.31-1.51)	0.36
Mixed CMBs	938/11,878 (8.2%)	0.87 (0.45-1.66)	0.67
IS			
Hyperlipidaemia	5,889/13,146 (44.8%)	0.93 (0.72-1.20)	0.57
Current smoker	1,709/10,375 (16.5%)	1.10 (0.86-1.41)	0.43
Fazekas score (continuous)	1 (IQR 1 – 2)	1.02 (0.80-1.31)	0.85
Fazekas score 2+	3,786/9,414 (40.2%)	1.12 (0.82-1.53)	0.47
Strictly deep CMBs	1,007/11,895 (8.5%)	1.20 (0.95-1.52)	0.080
Strictly lobar CMBs	1,146/11,892 (9.7%)	0.96 (0.71-1.30)	0.80
Mixed CMBs	971/11,896 (8.2%)	0.79 (0.54-1.15)	0.23

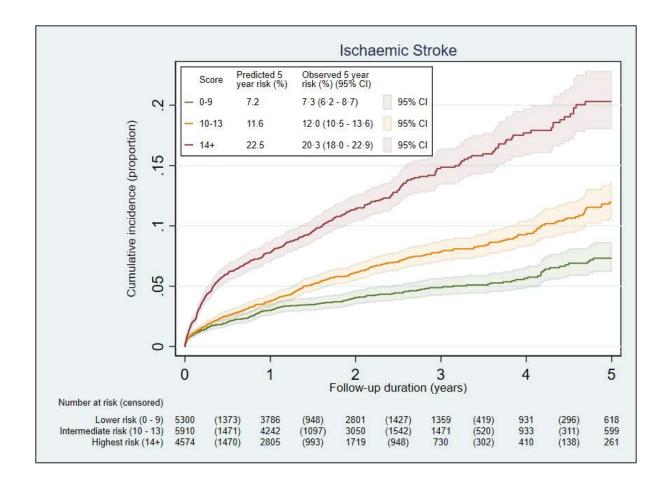
*Adjusted for other components of main model CMB: cerebral microbleed; ICH: intracranial haemorrhage; IS: ischaemic stroke

Supplementary Table 5: Characteristics of participants in highest-risk group for ICH and lower-risk group for IS

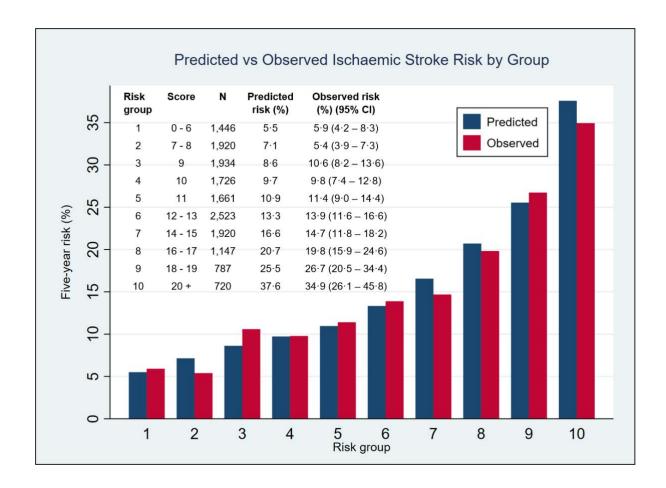
Values show prevalence for categorical variables, and mean (SD) or median (IQR) for continuous variables.

Variable		Group (n = 104)	Remainder (n = 11,849)
Age		80.9 (8.11)	71.4 (11.7)
Female sex		57/104 (54·8%)	5,055/11,849 (42.7%)
East Asian population		4/104 (3.85%)	4,423/711,849 (37.3%)
Hypertension		81/104 (77.9%)	8,613/11,849 (72.8%)
Atrial fibrillation		102/104 (98·1%)	5,898/11,849 (49.8%)
Previous IS		1/104 (0.96%)	1,878/11,849 (15.6%)
Previous ICH		17/104 (16·4%)	131/11,849 (1·11%)
Diabetes mellitus		9/104 (8.7%)	2,825/11,849 (23.8%)
Hyperlipidaemia		42/104 (41.6%)	5,236/11,849 (44.7%)
CMB burden	0	13/104 (12·5%)	8,531/11,849 (72.0%)
	1	60/104 (57.7%)	1,404/11,849 (11.9%)
	2 - 4	28/104(26.9%)	1,207/11,849 (10.2%)
	5 - 10	1/104 (1.0%)	382/11,849 (3.2%)
	11 - 19	1/104 (1.0%)	178/11,849 (1.5%)
	20 +	1/104 (1.0%)	147/11,849 (1.2%)
Antithrombotic treatment	AP only	2/104 (1.9%)	8,670/11,849 (48.6%)
	DOAC	19/104(18.3·3%)	2,215/11,849 (33·1%)
	Warfarin/VKA	83/104 (79 8.8%)	4.612/11.849 (18.3%)

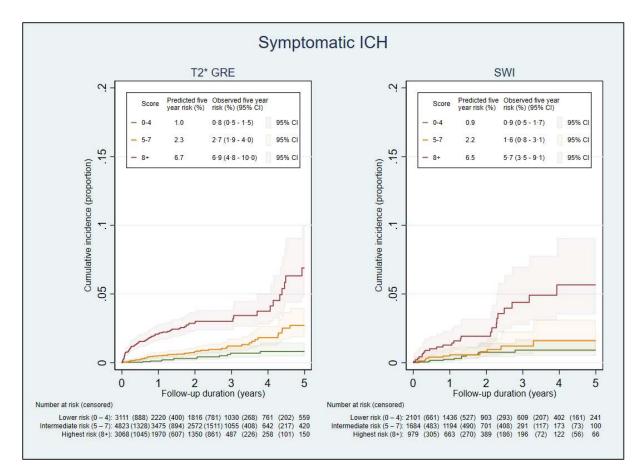
Supplementary Figure 1: Kaplan-Meier plot and risk table for ischaemic stroke model



Supplementary Figure 2: Model calibration – ischaemic stroke

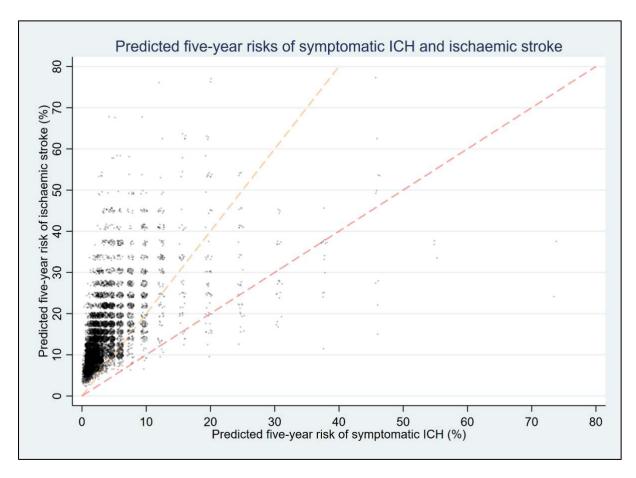


Supplementary Figure 3: ICH model performance by MRI sequence type



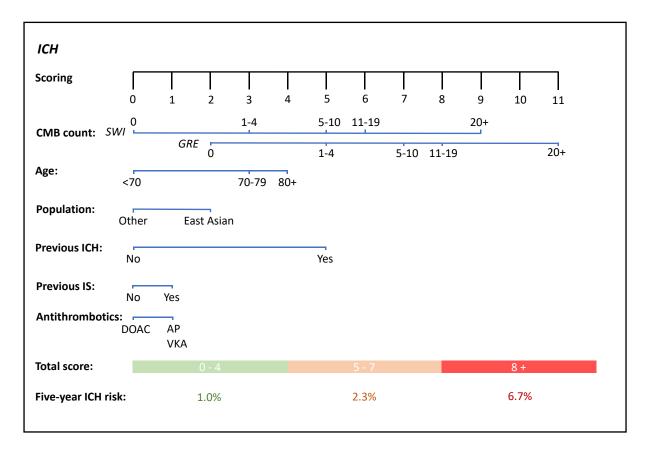
Performance measure	T2* GRE	SWI
C-index (optimism-adjusted)	0.75 (0.70 - 0.79)	0.70 (0.62 - 0.79)
Calibration slope	0.94 (0.76 – 1.12)	0.94 (0.79 – 1.09)

Supplementary Figure 4: Comparative risks of symptomatic ICH and ischaemic stroke



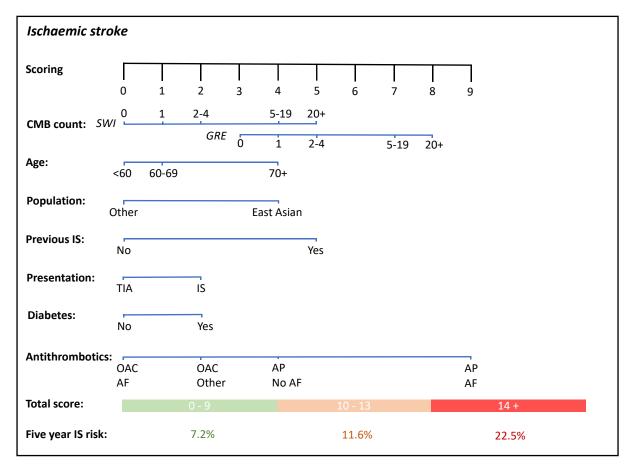
Predicted five-year risks from ICH and ischaemic risk scores for all 11,953 participants with all variables available without imputation. The red line indicates equality between predicted risks of ICH and IS; the orange line indicates predicted IS risk twice that of ICH. For presentation, markers are translucent and jittered.

Supplementary Figure 5: Nomogram for symptomatic ICH risk



For each variable, select the appropriate category, then read the score for that variable from the scoring bar. After summing the total score, select the corresponding 'total score' category and read the five-year ICH risk below.

Supplementary Figure 6: Nomogram for ischaemic stroke risk



For each variable, select the appropriate category, then read the score for that variable from the scoring bar. After summing the total score, select the corresponding 'total score' category and read the five-year IS risk below.

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