

1 **Dissociation between the neural correlates of conscious**
2 **face perception and visual attention**

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1 **Abstract**

2 Given the higher chance to recognize attended compared to non-attended stimuli, the
3 specific neural correlates of these two processes, attention and awareness, tend to be
4 intermingled in experimental designs. In this study, we dissociated the neural correlates of
5 conscious face perception from the effects of visual attention. To do this, we presented faces
6 at the threshold of awareness and manipulated attention through the use of exogenous pre-
7 stimulus cues. We show that the N170 component, a scalp EEG marker of face perception,
8 was modulated independently by attention and by awareness. An earlier P1 component was
9 not modulated by either of the two effects and a later P3 component was indicative of
10 awareness but not of attention. These claims are supported by converging evidence from: 1)
11 modulations observed in the average evoked potentials, 2) correlations between neural and
12 behavioral data at the single-subject level, and 3) single-trial analyses. Overall, our results
13 show a clear dissociation between the neural substrates of attention and awareness. Based
14 on these results, we argue that conscious face perception is triggered by a boost in face-
15 selective cortical ensembles that can be modulated by, but are still independent from, visual
16 attention.

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2 **Introduction**

3 When an image is flashed to an observer, retinal stimulation triggers a series of neural
4 processes, a subset of which lead to conscious perception (C. Koch, 2004). In order to
5 identify these brain processes, a standard approach is to compare the brain responses
6 elicited by “seen” and “unseen” visual stimuli. For example, we previously showed that the
7 amplitude of a face-selective occipito-temporal EEG component, peaking ~170 ms after
8 stimulus onset (i.e., the N170), was larger whenever subjects reported seeing a face,
9 although the stimuli eliciting the seen and unseen conditions, at the threshold of conscious
10 perception, was exactly the same (J. Navajas, Ahmadi, & Quian Quiroga, 2013). In fact, with
11 the peak amplitude of this signal it was possible to decode above chance the subjects’
12 conscious reports. One limitation of this study, which also applies to most related studies
13 (Aru et al., 2012; Fisch et al., 2009; Pitts, Martínez, & Hillyard, 2012; Pitts, Metzler, &
14 Hillyard, 2014; Sergent, Baillet, & Dehaene, 2005; Shafto & Pitts, 2015) is that the observed
15 modulations in the neural responses might just reflect trial-by-trial fluctuations of attention
16 rather than the perceptual state of the observer (J. Navajas, Rey, & Quian Quiroga, 2014). In
17 other words, high (low) levels of attention are more likely to elicit “seen” (“unseen”) trials and
18 the difference between the “seen” and “unseen” conditions may be due to a varying
19 allocation of attentional resources.

20 To address this issue, we developed a paradigm where we present faces at the
21 threshold of awareness and manipulate involuntary attention through the use of exogenous
22 cues (M. I. Posner, 1994; M. I. Posner, 1980). We exploited the fact that approximately one
23 second after the presentation of an exogenous flash, attentional resources are diminished in
24 the location where the flash was presented, an effect known as “inhibition of return” (Klein,
25 2000). Based on previous findings using faces as stimuli (Taylor & Therrien, 2005; Taylor &
26 Therrien, 2008), we predicted that “valid” trials (i.e., when the location of the flash matched

1 the one of the stimulus) should elicit larger reaction times, due to a lower level of attention,
2 than invalid trials. This experimental setting led to a 2×2 design that allowed a factorial
3 analysis of the neural correlates of conscious face perception and attention. Moreover, here
4 we rule out other possible confounds that are pervasive in the literature (J. Navajas et al.,
5 2014), such as differences in the physical stimuli used to elicit the different conditions.

6 Altogether we found that the pre-stimulus cues modulated attention, as measured by
7 the participants' reaction times, but had no impact on the probability to perceive the faces.
8 The neural data showed that the N170 component was modulated by attention, but still gave
9 a reliable index of conscious perception for some attentional conditions. An earlier positive
10 component (the P1) was uninformative of both attention and awareness. Finally, a later EEG
11 response with centro-parietal topography (the P3), which has been previously linked to
12 subjective awareness (S. Dehaene & Changeux, 2011), was indicative of conscious reports
13 but not of the attentional state of the subjects. Overall, these results provide a dissociation of
14 the neural mechanisms of attention and conscious face perception in scalp occipito-temporal
15 and centro-parietal EEG responses.

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17 **Methods**

18 **Stimuli**

19 Twenty pictures of faces in frontal view were selected from the Psychological Image
20 Collection at Stirling (<http://pics.psych.stir.ac.uk/>), and twenty front views of cars were
21 downloaded from the internet. Twelve different levels of zero-mean Gaussian noise (equally
22 spaced between 0 and 0.3) were added to the images using the Matlab function *imnoise*.
23 The brightness and spatial frequency of the entire dataset were normalized using the SHINE
24 toolbox (Willenbockel et al., 2010). Stimuli were presented on a CRT Monitor with a
25 resolution of 1024×768 pixels and a refresh rate of 100 Hz. The viewing distance was
26 approximately 50 cm.

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3 **Paradigm**

4 To dissociate attention and awareness, we combined an experimental design previously
5 used to study conscious face perception (Navajas et al., 2013) with a classic paradigm to
6 manipulate involuntary attention through the presentation of pre-stimulus exogenous cues
7 (M. I. Posner, 1980). Each trial (**Fig. 1a**) began with a central fixation cross on a gray
8 background and two empty squares centered 3.6 degrees off the vertical meridian (size: 2.6
9 degrees) that were shown for a randomly chosen time between 800 and 900 ms. After this
10 interval, a flash appeared for 100 ms on either of the two squares and after another 900-
11 1100 ms, a face with added noise was presented for 100 ms on one of the two squares.
12 Participants were instructed to report as soon and as accurately as possible whether or not
13 they saw a face regardless of its location. Responses were given by pushing one of two
14 buttons of a millisecond accurate Black Box Toolkit (BBTK) response box (Plant & Quinlan,
15 2013) using the right hand.

16 **Task and Procedure**

17 The experiment consisted in 1680 trials of two different types: 1568 “staircase” trials and 112
18 “catch” trials. The entire experiment lasted approximately 90 minutes. In the “staircase” trials,
19 we presented faces with a variable amount of added noise using a double staircase
20 procedure which kept running throughout the experiment (Navajas et al., 2014). The upper
21 (lower) staircase always started from a noise level of 0.3 (0) with a fixed step size of 0.025.
22 The stimulus location (left or right visual field) and the attention manipulation (valid or invalid
23 cue) were counterbalanced across the “staircase” trials. Thus, cues were not predictive of
24 the stimulus location. “Catch” trials were cars presented with the same level of noise as the
25 “staircase” (face) trials. This type of trials was used to assess the reliability of the

1 participants' reports by computing a false positive rate (i.e., the proportion of trials in which
2 participants reported to have seen a face in these trials).

3 To ensure that the visual input was constant across conditions, we selected for each subject
4 the level of noise that led to a face recognition rate closest to 50% (Navajas et al., 2013;
5 Navajas et al., 2014). The number of trials in this level of noise was different for each subject
6 (mean: 564 trials, s.d.: 114 trials). These trials were the ones used to compute ERPs across
7 the 8 conditions of our experiment (left / right × valid / invalid × seen / unseen). The contrast
8 between the ERPs elicited in the “unseen” vs. “seen” condition could be interpreted as the
9 neural signals that correlate with the switch from reporting “I did not see a face” to “I saw a
10 face” at constant retinal stimulation. Following previous studies (Navajas et al., 2013; Shafto
11 & Pitts, 2015), we refer to these signals as the “neural correlates of conscious face
12 perception”.

13 **Participants**

14 29 healthy adults volunteered to participate in this study. All of them were right handed, had
15 no history of neurological disorders and a normal or corrected-to-normal vision. All
16 participants were recruited through advertisement at the University of Leicester, gave written
17 informed consent, and were paid £15 for their participation. Data from 4 participants were
18 discarded due to an excessive number of false positives (more than 13%, while the range of
19 false positive rates was [0-3.5] % for the remaining participants). We discarded data from 2
20 other subjects because, after artefact rejection, they had less than 40 trials per condition to
21 compute the ERPs. Hence, we report results from 23 participants (15 female, mean age:
22 23.6 years, range: 18-36). For these participants, we computed ERPs based on 71 ± 15
23 (mean \pm s.d.) trials (range: 48-94 trials) per condition. We confirmed, using subsamples of
24 our data that this number of trials was sufficient to measure reliable ERPs (**Supplementary**
25 **Fig. 1**).

26 **EEG Recordings and ERP Analysis**

1 Signals were recorded with a sampling rate of 256 Hz using a Biosemi Active Two 64-
2 channel EEG system. The reference was set to the average across all electrodes. Epochs
3 including 1 sec before and 1 sec after stimulus onset were extracted for further analysis and
4 band-pass filtered with a second-order Butterworth filter between 1 and 70 Hz. This
5 reference and filter parameters were shown to be optimal for the P1 and N170 components
6 (Joyce & Rossion, 2005; Rousselet et al., 2007).

7 After examination of the evoked topographies (averaging all conditions together), we
8 distinguished three clear components: P1, N170, and P3 (see top panels of **Fig. 2**). For
9 stimuli presented in the left visual field, the P1 (N170) component had its peak in electrode
10 PO8 at 130 ms (190 ms). For stimuli presented in the right visual field, both the P1 and N170
11 had their peaks at the same time in electrode PO7. Therefore, we selected electrodes PO7/8
12 to further analyse these components, which is in line with previous studies (e.g., Rossion
13 and Jacques, 2008; Navajas et al., 2013).

14 Given the global nature of the P3 component, an average reference is not optimal in
15 this case and we therefore re-referenced samples to mastoid electrodes (Joyce & Rossion,
16 2005). Because the P3 wave is slower than the P1/N170, we used different filter parameters
17 to focus on this component (0.1-30 Hz). We found that the P3, averaged across all
18 conditions, had its peak in electrode CPz at 400 ms (for stimuli presented from both sides),
19 and thus selected this electrode for further analysis. ERP traces from other electrode sites
20 are shown in **Supplementary Fig. 3**.

21 Trials containing blinks, eye-movements, and other artifacts were rejected by a
22 semiautomatic procedure applied to each subject separately. First, we epoched and filtered
23 the data on a time window of ± 1 sec around the stimulus onset. Then, we set a threshold in
24 amplitude of ± 30 μV for a selection of channels along the midline (Fz, FCz, Cz, CPz, Pz, and
25 Oz) and pre-selected all trials that crossed this threshold. Finally, we visually inspected all of
26 these trials to only reject them if they were actual artifacts. Overall, this procedure led to a

1 rejection rate of 10.4% of the trials (range: 1.4%-21.1%) for the 23 subjects considered in
2 our study.

3 For statistical analysis, we computed for each channel and subject the mean average
4 ERP on specific time windows. For the P1 and N170, we selected time windows using the
5 same procedure in Navajas et al. (2013); we considered 40 ms windows centred in the peak
6 of each component (110-150 ms for the P1, 170-210 ms for the N170). For the P3, we used
7 a 200 ms window centred in its peak (300-500 ms), as in previous studies (Pitts et al., 2012;
8 Pitts et al., 2014; Shafto & Pitts, 2015).

9 Stimuli presented at different locations (ipsilateral and contralateral visual field) were
10 analyzed separately. Each component was submitted to a two-way repeated-measures
11 ANOVA (rm-ANOVA) with factors “attention” (two levels: “valid” or “invalid” trials) and
12 “conscious report” (two levels: “seen” and “unseen”). Statistical differences were assessed
13 separately for each component. For all pairwise comparisons, we performed t-tests and
14 report for each condition the mean \pm s.e.m., along with a standardized measure of effect size
15 (i.e., 95% confidence intervals for the Hedges’ g value, see Hentschke & Stuttgart, 2011).

16 **Assessment of correlations between behavioral and neural data**

17 To evaluate the correlation between the observed effects in the behavioral and the neural
18 data, we measured for each participant the mean reaction time (mRT) and the mean ERP
19 amplitude ($mERP$) for the P1, N170, and P3 across all conditions. Both for the behavioral
20 and the neural data, for each subject we computed the effect of cue validity (ECV) as (**Fig.**
21 **4a,b**):

[1]

$$22 \quad ECV(X) = \frac{X(v, s) + X(v, u) - (X(i, s) + X(i, u))}{\max(X)}$$

23 where X denotes either mRT or $mERP$, v and i denote the valid and invalid conditions, and s
24 and u refer to the seen and unseen conditions, respectively. This allowed us to test if the

1 behavioral effect of cue validity, as measured from the reaction times, was correlated to
2 amplitude changes on the ERP components on a subject-by-subject basis.

3 We also aimed at evaluating if the effects of attention and awareness on the ERPs
4 were correlated on a subject-by-subject basis (**Fig. 4c,d**). To this end, we computed the
5 effect of awareness on the ERP amplitudes as:

$$6 \quad EA(mERP) = \frac{mERP(v, s) + mERP(i, s) - (mERP(v, u) + mERP(i, u))}{\max(mERP)} \quad [2]$$

7

8 **Single-Trial Analysis and Decoding**

9 We assessed the information provided by the single-trial traces using a decoding analysis.
10 Single-trial ERPs were extracted by means of a denoising algorithm that implements a
11 wavelet decomposition of the single-trial ERPs and denoises them by reconstructing the
12 signals using only the wavelet coefficients related to the evoked responses, setting to zero
13 all the others (Ahmadi & Quian Quiroga, 2013; R. Quian Quiroga, 2000). The set of wavelet
14 coefficients used to extract the single-trial traces was kept constant across subjects and
15 conditions. The single-trial N170 responses were identified as the local minimum between
16 140 and 220 ms, and for the P3 we used the local maximum between 250 and 550 ms.

17 With these peaks, we implemented a linear discriminant analysis to decode the conscious
18 report (“seen” or “unseen”) and attentional condition (cue “valid” or “invalid”) using the Matlab
19 function *classify* and a leave-one-out cross validation. Chance levels were defined by
20 repeating the procedure 1,000 times per subject using a random permutation of the labels.
21 Statistical differences were assessed using paired t-tests between the 23 observed decoding
22 performances and the 23 mean performances obtained with random permutations.
23 Significance levels were corrected for multiple comparisons using Bonferroni’s method.

24 **Results**

1 **Behavior**

2 The double-staircase procedure converged to a noise level of 0.078 ± 0.013
3 (mean \pm SEM), with a recognition performance of 48.7 ± 7.4 %. To evaluate the effectiveness
4 of the attention manipulation, we compared the mean reaction times (**Fig. 1b**) obtained in
5 the different conditions using a two-way rm-ANOVA with “attention” and “conscious report”
6 as factors. We observed a significant effect of attention ($F(1,22)=10.05$, $p<0.01$) but not of
7 conscious report ($F(1,22)=2.09$, $p=0.16$), and no significant interaction ($F(1,22)=3.21$,
8 $p=0.11$). These results suggest that the experimental design succeeded at manipulating
9 attention without affecting the perceptual decisions of the subjects. To further validate this
10 observation, we measured the proportion of “seen” trials in “valid” and “invalid” conditions.
11 We reasoned that, if higher levels attention correlated with a higher probability of consciously
12 perceiving the faces, we should then observe less “seen” trials upon the “valid” condition,
13 where attention is diminished through inhibition of return. However, the proportion of “seen”
14 trials in the “invalid” condition was not significantly different from the one obtained in the
15 “valid” condition (valid trials: (47.8 ± 1.83) %, invalid trials: (48.1 ± 1.78) %, paired t-test, $t(22) =$
16 0.15 , $p=0.87$).

17 **Evoked Responses**

18 We identified three clear event-related potentials (ERPs) in response to faces (**Fig 2**). The
19 first one was a positive component in occipital electrodes peaking at ~ 130 ms (the P1). The
20 second one was a negative deflection at ~ 190 ms with an occipito-temporal location (the
21 N170). Lastly, we observed a later (~ 400 ms) centro-parietal wave, corresponding to the P3
22 component. To evaluate which of these responses were informative of the subjects' attention
23 and conscious perception, we computed the P1 and N170 in electrodes PO7/8 and the P3 in
24 electrode CPz (see Methods). The P1 failed to show any significant main effect or
25 interaction, both when the stimulus appeared in the left (conscious perception: $F(1,22)=2.53$,
26 $p=0.12$; attention: $F(1,22)=0.21$, $p=0.65$, interaction: $F(1,22)=1.67$, $p=0.21$) and in the right

1 visual field (conscious perception: $F(1,22)=0.03$, $p=0.86$; attention: $F(1,22)=0.13$, $p=0.72$,
2 interaction: $F(1,22)=0.40$, $p=0.53$).

3 The largest N170 responses were observed in the right hemisphere (i.e., electrode
4 PO8) upon contralateral stimulus presentation (**Fig. 3a**). In this case, we observed both a
5 significant effect of attention ($F(1,22)=6.27$, $p=0.02$) and conscious perception
6 ($F(1,22)=21.54$, $p<0.001$), and no significant interaction ($F(1,22)=0.92$, $p=0.35$). Post-hoc
7 tests showed that the difference between “seen” and “unseen” trials was significant both for
8 the valid (“unseen” condition: (-1.9 ± 0.4) μV ; “seen” condition: (-2.4 ± 0.4) μV ; Hedges’ g CI =
9 $[0.01 - 0.53]$; $t(22)=2.1$, $p<0.05$) and invalid (“unseen” condition: (-2.1 ± 0.4) μV ; “seen”
10 condition: (-2.9 ± 0.4) μV ; Hedges’ g CI = $[0.18 - 0.61]$; $t(22)=4.6$, $p<10^{-3}$) conditions. No
11 effects were observed for this electrode when the stimulus was presented in the ipsilateral
12 field (conscious perception: $F(1,22)=1.85$, $p=0.18$; attention: $F(1,22)=3.10$, $p=0.10$,
13 interaction: $F(1,22)=0.19$, $p=0.67$). Also, no effects were observed in the left hemisphere
14 (electrode PO7) both when the stimulus was presented in the contralateral (conscious
15 perception: $F(1,22)=3.25$, $p=0.09$; attention: $F(1,22)=0.69$, $p=0.41$, interaction: $F(1,22)=0.08$,
16 $p=0.77$) or in the ipsilateral field (conscious perception: $F(1,22)=1.07$, $p=0.13$; attention:
17 $F(1,22)=0.63$, $p=0.43$, interaction: $F(1,22)=1.04$, $p=0.32$). These results are consistent with
18 previous studies showing that the N170 is maximal in the right hemisphere (B. Rossion,
19 Joyce, Cottrell, & Tarr, 2003) and contralateral to the stimulus location (Feng, Martinez, Pitts,
20 Luo, & Hillyard, 2012) (see Discussion).

21 Centro-parietal P3 responses were significantly modulated by conscious perception
22 but not by attention (**Fig. 3b**). These effects were present both for visual stimuli presented in
23 the left (electrode CPz, conscious perception: $F(1,22)=18.62$, $p<0.001$; attention:
24 $F(1,22)=0.76$, $p=0.38$, interaction: $F(1,22)=0.08$, $p=0.76$) and right visual field (electrode
25 CPz, conscious perception: $F(1,22)=11.51$, $p<0.01$; attention: $F(1,22)=0.16$, $p=0.68$,
26 interaction: $F(1,22)<0.01$, $p=0.92$). Post-hoc tests showed that the effect of awareness on
27 the P3 was present for both valid (“unseen” condition: (0.3 ± 0.7) μV ; “seen” condition:

1 (1.7±0.7) μV ; Hedges' g CI = [0.10 – 0.75]; $t(22)=2.9$, $p<0.01$) and invalid (“unseen”
2 condition: (0.2±0.6) μV ; “seen” condition: (1.4±0.6) μV ; Hedges' g CI = [0.10 – 0.67];
3 $t(22)=3.2$, $p<0.01$) trials (**Fig. 3b**).

4 We also observed a negative drift in the baseline at electrode CPz which
5 corresponds to the contingent negative variation (CNV) component. This signal is a
6 signature of the anticipation of an upcoming event (Macar & Vidal, 2003), a new trial in our
7 case. Because the CNV has a slow temporal dynamics, we only observe it in **Fig. 3b**, where
8 signals were high-pass filtered at 0.1 Hz (as opposed to 1 Hz when in **Fig. 3a**, see **Methods**
9 for details). This CNV component was present in most electrodes, including PO8
10 (**Supplementary Fig. 3**), but we did not observe any significant difference in the pre-
11 stimulus CNV amplitude (conscious perception: $F(1,22)=0.94$, $p=0.34$; attention:
12 $F(1,22)=0.24$, $p=0.63$, interaction: $F(1,22)=1.45$, $p=0.24$).

13

14 **Correlation between behavioral and neural responses**

15 We further examined the effect of attention on the N170 (**Supplementary Fig. 2**). First, we
16 observed that, averaging “seen” and “unseen” trials, the N170 amplitudes were significantly
17 larger in the “invalid” compared to the “valid” condition (“valid” condition: (-2.2±0.4) μV ,
18 “invalid” condition: (-2.5±0.4) μV ; Hedges' g CI = [0.01 – 0.33]; $t(22)= 2.3$, $p=0.02$). This
19 effect was independent of the assumption of normality in our data (AUROC = 0.43, 95% CI =
20 [0.26-0.70]; Wilcoxon sign rank test, $z=2.3$, $p=0.02$). Post-hoc tests showed that that the
21 effect of attention was significant for seen trials (“valid” condition: (-2.4±0.4) μV , “invalid”
22 condition: (-2.9±0.4) μV ; Hedges' g CI = [0.01 – 0.29]; $t(22)=2.12$, $p<0.05$), but it was absent
23 in the unseen condition (“valid” condition: (-1.9±0.4) μV , “invalid” condition: (-2.1±0.4) μV ;
24 Hedges' g CI = [-0.02 – 0.48]; $t(22)=1.14$, $p=0.26$).

25 To gain more confidence that attention indeed modulated the N170 component, we
26 performed a new analysis. We evaluated if the varying modulations in the ERP components

1 among subjects correlated with their varying degrees of attention in the task, quantified by
2 their reaction times. To this end, we measured the effect of cue validity on the mean reaction
3 times (i.e. an index of how effective the attentional manipulation was at the behavioral level)
4 and correlated it with the one observed for the N170 and P3 amplitudes (see Methods,
5 Equation [1]). We observed a positive correlation between the modulation of reaction times
6 due to attention, and the changes in the N170 amplitude due to the same effect (**Fig. 4a**,
7 Pearson correlation coefficient $r=0.47$, $p=0.02$).

8 Consistent with the abovementioned finding that the P3 did not change with attention,
9 we did not observe such a correlation for this component (**Fig. 4b**, Pearson correlation
10 coefficient $r=-0.29$, $p=0.16$). We then studied if the different significant effects observed at
11 the average level (**Fig. 3**) were correlated with each other on a subject-by-subject basis. We
12 observed that the modulations in the N170 due to awareness were uncorrelated with the
13 effect of attention on the same component (**Fig. 4c**), and with the effect of awareness on the
14 P3 component (**Fig. 4d**). This supports the view that these two signals reflect independent
15 neural processes.

16 **Discarding residual effects of attention on the N170**

17 Valid and invalid trials led, on average, to conditions with different levels of spatial attention
18 (**Fig 1b**). However, it could still be possible that other attentional processes (not captured by
19 our manipulation), led to trial-to-trial fluctuations in attention *within* each condition. To test
20 this possibility, we performed a single-trial analysis of the EEG traces. We reasoned that, if
21 residual sources of attention modulated neural signals within each condition, then trial-to-trial
22 changes in reaction time (our proxy for attention) should correlate with single-trial
23 fluctuations in the evoked responses.

24 We focused on the N170 peak amplitude as this was the only evoked response that showed
25 significant modulation with attention at the average level (**Fig. 2, Fig. 3a, Fig. 4**). Single-trial
26 peaks were extracted using a wavelet-denoising algorithm (see Methods). The algorithm

1 consists in decomposing the EEG traces using a discrete wavelet transform and
2 reconstructing the signal using only the wavelet coefficients that are related to the average
3 evoked responses. This procedure was previously shown to improve the estimation of
4 single-trial peaks (Ahmadi & Quian Quiroga, 2013; R. Quian Quiroga, 2000), including N170
5 responses at the threshold of awareness (J. Navajas et al., 2013).

6 Fluctuations in single-trial N170 amplitudes were compared to trial-to-trial changes in
7 reaction time. Lumping all conditions together, we observed that trials with higher attention
8 (i.e. faster reaction times) had a larger N170 effect (i.e., a more negative N170). **Fig. 5a**
9 shows this effect for one representative participant. To quantify this observation at the
10 population level, we fit a linear regression to each participant and compared the distribution
11 of slopes with zero. We found that this distribution was significantly larger than zero (**Fig. 5b**,
12 two-tailed sign-test, $p=0.01$). This finding is consistent with our previous observation that the
13 N170 is larger for invalid trials.

14 We then examined each condition separately, and observed that the N170 amplitudes were
15 uncorrelated with fluctuations in reaction time (**Fig. 5c,d**, two-tailed sign-tests of distribution
16 of best-fitting slopes, $p>0.4$ for all conditions). This indicates that, if there were any residual
17 effect of attention within each condition, these effects did not modulate the N170. Therefore,
18 the effects reported in this study cannot simply be explained by trial-to-trial changes in
19 residual attentional processes.

20 **Decoding**

21 Next, we asked whether the modulations in the ERP responses were strong enough to
22 predict, at the single-trial level, both the attention (attended or not) and perceptual (seen or
23 not) conditions. From the single-trial denoised traces, we identified the N170 and P3 peaks
24 and implemented a decoding analysis. To assess the significance of the decoding algorithm,
25 we compared the empirically measured performances with the ones obtained using random
26 permutations (see Methods). We focused this analysis on trials in which the stimulus was

1 presented in the left visual field, as these were the trials that yielded significant results in the
2 evoked responses.

3 We did not observe any significant modulation in the single-trial latencies of the ERP
4 components, and consequently, we focused on the single-trial peak amplitudes. **Table 1**
5 summarizes the decoding results with the corresponding statistical significances. We first
6 aimed at decoding the four conditions of the experiment (attended and unattended; seen
7 and unseen) based on the amplitudes of these components. Decoding performance was
8 significantly above chance both for the N170 (measured performance: $(28.1 \pm 0.5)\%$, chance
9 level: $(25.0 \pm 0.2)\%$, paired t-test $t(22)=5.5$, $p < 10^{-4}$, Bonferroni corrected $\alpha=0.025$) and the P3
10 component (measured performance: $(27.4 \pm 0.8)\%$, chance level: $(24.9 \pm 0.1)\%$, paired t-test
11 $t(22)=3.1$, $p < 10^{-4}$, Bonferroni corrected $\alpha=0.025$).

12 Next, we asked whether we could decode conscious reports disregarding the
13 attentional condition. In this case, decoding performance was significantly above chance
14 both using the N170 peak (measured performance: $(54.0 \pm 0.6)\%$, chance level: $(50.0 \pm 0.3)\%$,
15 paired t-test $t(22)=5.9$, $p < 10^{-5}$, Bonferroni corrected $\alpha=0.025$) and the P3 peak (measured
16 performance: $(54.1 \pm 1.1)\%$, chance level: $(50.1 \pm 0.3)\%$, paired t-test $t(22)=3.5$, $p < 10^{-3}$,
17 Bonferroni corrected $\alpha=0.025$). Similarly, we studied if we could decode the attentional
18 condition regardless of the conscious report given by the participant. Consistently with the
19 effects observed in the evoked responses, decoding performance was significantly above
20 chance when using the N170 (measured performance: $(53.1 \pm 0.4)\%$, chance level:
21 $(49.9 \pm 0.3)\%$, paired t-test $t(22)=6.1$, $p < 10^{-5}$, Bonferroni corrected $\alpha=0.025$) but not when
22 using the P3 peak (measured performance: $(50.9 \pm 1.5)\%$, chance level: $(49.8 \pm 0.3)\%$, paired
23 t-test $t(22)=0.8$, $p=0.4$, Bonferroni corrected $\alpha=0.025$).

24 We then examined if conscious reports could be decoded at a constant attention
25 level. Using the N170 peak, decoding performance was above chance both for invalid
26 (measured performance: $(53.8 \pm 0.9)\%$, chance level: $(49.7 \pm 0.3)\%$, paired t-test $t(22)=4.5$,
27 $p < 10^{-3}$, Bonferroni corrected $\alpha=0.012$) and valid trials (measured performance: $(55.6 \pm 1.0)\%$,

1 chance level: (49.7±0.4)%, paired t-test $t(22)=5.2$, $p<10^{-4}$, Bonferroni corrected $\alpha=0.012$.
2 Similarly, conscious reports could also be decoded using the P3 component both with invalid
3 (measured performance: (53.4±1.2)%, chance level: (50.1±0.3)%, paired t-test $t(22)=3.0$,
4 $p<0.01$, Bonferroni corrected $\alpha=0.012$) and valid trials (measured performance: (55.4±1.1)%,
5 chance level: (50.0±0.3)%, paired t-test $t(22)=5.1$, $p<10^{-4}$, Bonferroni corrected $\alpha=0.012$).

6 Complementing these results, we studied if the attention condition could be decoded
7 from single-trial peaks both for the seen and unseen trials. Decoding of attention using the
8 N170 peak was above chance for the seen (measured performance: (54.4±0.8)%, chance
9 level: (49.7±0.3)%, paired t-test $t(22)=5.6$, $p<10^{-4}$, Bonferroni corrected $\alpha=0.012$) and unseen
10 trials (measured performance: (52.7±0.6)%, chance level: (50.3±0.3)%, paired t-test
11 $t(22)=3.4$, $p<0.01$, Bonferroni corrected $\alpha=0.012$). On the contrary, decoding performance
12 was not significantly different from chance using the P3 peak both for seen (measured
13 performance: (52.3±1.4)%, chance level: (50.1±0.3)%, paired t-test $t(22)=1.6$, $p=0.11$,
14 Bonferroni corrected $\alpha=0.012$) and unseen trials (measured performance: (48.2±2.6)%,
15 chance level: (50.3±0.3)%, paired t-test $t(22)=0.8$, $p=0.43$, Bonferroni corrected $\alpha=0.012$).
16 This pattern of results at the single-trial level is consistent with the analysis of the average
17 evoked responses: the amplitude of N170 is informative of both attention and conscious
18 perception whereas the P3 component is only predictive of conscious reports.

19 **Discussion**

20 In everyday life, attention and awareness are profoundly intertwined. We move our eyes
21 approximately three times per second (Yarbus, 1967), only becoming aware of what we
22 overtly attend to (Burr, Morrone, & Ross, 1994; De Pisapia, Kaunitz, & Melcher, 2010; J.
23 Navajas, Sigman, & Kamienkowski, 2014). We can even fail to perceive very salient stimuli
24 in plain sight if our attention is engaged in another task (Most et al., 2001). In order to
25 dissociate these two cognitive processes, several studies proposed different experimental
26 paradigms that allow a factorial analysis of attention and awareness (S. Dehaene,

1 Changeux, Naccache, Sackur, & Sergent, 2006; C. Koch & Tsuchiya, 2007; V. A. F. Lamme,
2 2003). With this approach, it has been shown that BOLD signals in the primary visual cortex
3 are modulated by attention but not by awareness (Watanabe et al., 2011), that attentional
4 load reduces visual discrimination (Lee, Itti, Koch, & Braun, 1999) and modulates brain
5 responses in the absence of awareness (Bahrami, Lavie, & Rees, 2007), and that the
6 frequency band of cortical signals carrying information about awareness is different than the
7 one related to attention (Wyart & Tallon-Baudry, 2008). Moreover, attention and awareness
8 were shown to have opposing effects on the duration of afterimages (van Boxtel, Tsuchiya,
9 & Koch, 2010).

10 However, these studies contrasted neural and behavioral responses upon conditions
11 using different stimuli (van Boxtel et al., 2010; Watanabe et al., 2011), different cognitive
12 load (Bahrami et al., 2007; Lee et al., 1999), or different expectations given by predictive
13 cues (Wyart & Tallon-Baudry, 2008). The main caveat of these approaches is that the
14 contrast between different conditions might conflate modulations in attention or awareness
15 with other brain processes (J. Navajas et al., 2014). To overcome all these issues, we
16 developed a paradigm that dissociates attention from awareness in a 2×2 factorial design
17 where all conditions have constant retinal stimulation, cognitive load, and expectancy.

18 To this end, we adapted a classic paradigm that uses pre-stimulus exogenous cues
19 and leads to inhibition of return (M. I. Posner, 1994; M. I. Posner, 1980), which impairs
20 attention at the cued location approximately one second after the cue onset (Klein, 2000).
21 Because cues and faces were separated by an interval of at least 900 ms, we can rule out
22 that the evoked responses given by the faces were mixed up with those ones given by the
23 cue. Shorter intervals (up to approximately 300 ms) would have led to enhanced attentional
24 resources at cued locations (Klein, 2000) but only at the cost of not being able to disentangle
25 the neural responses of the cue and the stimulus. Moreover, cues were not predictive of the
26 stimulus location, which allowed dismissing effects of expectancy. This would have been

1 impossible if we had used endogenous cues, as these need to be predictive of the stimulus
2 location in order to modulate attention (M. I. Posner, 1980).

3 In valid trials, when the stimuli matched the location of the pre-stimulus cue, we
4 observed that reaction times were slower compared to invalid trials (Taylor & Therrien,
5 2008). This effect was independent of the perceptual report given by the participants, which
6 allowed asserting that, behaviorally, attention and awareness were manipulated in an
7 orthogonal way. This finding was further supported by the fact that the pre-stimulus cue did
8 not change the probability to consciously perceive the faces in that location. In other words,
9 this result suggests that our paradigm was able to empirically dissociate attention and
10 awareness.

11 We recorded scalp EEG activity in response to faces and identified three clear
12 components: P1, N170, and P3 (**Fig. 2**). At the average ERP level, we observed three
13 significant effects (**Fig. 3**): 1) the N170 amplitude was modulated by attention for the same
14 perceptual state; 2) the N170 was also modulated by awareness at constant attentional
15 level; and 3) the P3 component was modulated by awareness but not by attention. These
16 results were further supported by consistent findings at the single-subject level (**Fig. 4a**), and
17 by a single-trial decoding analysis (**Table 1**). We did not find any significant interaction (**Fig.**
18 **3**) or correlation (**Fig. 4c,d**) between these effects. Moreover, the P1 component was not
19 modulated by either attention or awareness.

20 We observed that the N170 component was independently modulated by attention
21 and awareness in the right occipito-temporal cortex (electrode PO8) with contralateral
22 stimulation. Modulations with stimuli presented in the ipsilateral field and in the left occipito-
23 temporal cortex (PO7) were not statistically significant. In line with these findings, previous
24 studies have shown evidence supporting that the right-lateralization of face processing in the
25 brain leads to smaller N170 responses in the left hemisphere (B. Rossion & Jacques, 2011;
26 B. Rossion et al., 2003; B. Rossion et al., 1999; B. Rossion & Jacques, 2008). Also, as with
27 many other visually-evoked EEG potentials, the N170 was larger for stimuli presented in the

1 contralateral visual field (Feng et al., 2012). Based on these previous findings, we believe
2 that, in trials where the stimuli were presented in the right visual field, the signal-to-noise
3 ratio in both hemispheres was not sufficient to measure sizeable effects.

4 Previous studies have found that attention can sometimes amplify post-perceptual
5 signals such as the P3 component (Polich, 2007), but in our study we did not observe such
6 modulation. This is in consistent with theoretical accounts proposing that attention operates
7 on a variety of cognitive systems depending on task demands, selectively processing
8 relevant information (Hillyard, Vogel, & Luck, 1998; Lavie, 1995; S. Luck & Hillyard, 2000).
9 For example, if the task consisted in memorizing letters, or the identity of certain faces, we
10 would have expected to observe modulations by attention at a timing consistent with the P3
11 (S. J. Luck, Woodman, & Vogel, 2000). Thus, the lack of a modulation by attention in the P3
12 component in our study can be attributed to the nature of our paradigm, which matched
13 perceptual features, expectations, and cognitive loads upon all conditions.

14 A similar argument applies to explaining why we did not observe a modulation in the
15 P1 component, as it has been reported in the literature (Heinze, Luck, Mangun, & Hillyard,
16 1990; Hillyard et al., 1998; S. J. Luck, Heinze, Mangun, & Hillyard, 1990; Mangun & Hillyard,
17 1991). These studies were based on paradigms where subjects were instructed to detect
18 whether the stimulus was flashed on the left or on the right visual field. In contrast, the task
19 in our experiment was to report whether the stimulus was a face or not. We believe that the
20 absence of a P1 modulation in our study, and the presence of an effect of attention in the
21 later N170 signal, is because attention was focused to high-level visual features of the
22 stimuli – driven by the task instructions and the use of images with equal low-level visual
23 features – which are processed by face-selective neural populations sourcing the N170
24 signal (B. Rossion & Jacques, 2011). Therefore, neural ensembles contributing to the P1
25 component were task-irrelevant in our experiment (B. Rossion & Caharel, 2011). This is also
26 in line with previous findings showing that attention rapidly activates task-related cortical
27 modules (Schoenfeld, Hopf, Merkel, Heinze, & Hillyard, 2014). Alternatively, it might be

1 possible that paradigms based on inhibition of return simply do not modulate P1 responses
2 (for a review, see Martin-Arevalo et al., 2016).

3 In past years, there has been an extensive debate regarding the neural signals
4 involved in conscious perception and their spatio-temporal dynamics. A vast corpus of
5 evidence has provided support for a theoretical framework asserting that consciousness is
6 reflected by “late” (>300 ms) global activity spread throughout the cortex (S. Dehaene &
7 Naccache, 2001; S. Dehaene, Kerszberg, & Changeux, 1998; S. Dehaene, Sergent, &
8 Changeux, 2003; S. Dehaene et al., 2006; S. Dehaene & Changeux, 2011). Other studies
9 have found that localized cortical activity at a relatively earlier latency (100-200 ms after
10 stimulus onset) is already predictive of conscious perception (Fisch et al., 2009; Koivisto &
11 Revonsuo, 2010; V. A. F. Lamme, 2000; J. Navajas & Kaunitz, 2016; J. Navajas et al., 2013;
12 Pitts et al., 2012; Pitts et al., 2014; Sandberg et al., 2013; Shafto & Pitts, 2015). Our findings
13 suggest that the earliest neural correlate of conscious face perception indeed occurs at ~170
14 ms in the occipito-temporal cortex (J. Navajas et al., 2014; J. Navajas et al., 2013). This
15 information undergoes further processing, which leads to a modulation by awareness after
16 300 ms with a distributed topography (S. Dehaene & Changeux, 2011), potentially triggered
17 by hippocampal activations to consciously perceived stimuli (R. Quian Quiroga, Kraskov,
18 Mormann, Fried, & Koch, 2014; R. Quian Quiroga, Mukamel, Isham, Malach, & Fried, 2008).
19 This second correlate of consciousness was recently linked to post-perceptual processing of
20 sensory information (Shafto & Pitts, 2015), and is believed to guide task-relevant behavior
21 such as the production of conscious reports (Pitts et al., 2014; Tsuchiya, Wilke, Frässle, &
22 Lamme, 2015). In line, a recent study has found that this second stage may even occur after
23 a motor response, which strongly suggests that this is indeed a post-perceptual process
24 (Railo et al., 2015). Building up on this evidence, we have previously argued that such late
25 component reflects *contextual awareness*, following a first process of *perceptual awareness*
26 (J. Navajas et al., 2014). In other words, the process of recognizing a specific person or item
27 is then followed by putting it into context, e.g. by recalling related experiences. In line with

1 the distinction between an early and late conscious process, we showed that only the earlier
2 correlate of consciousness, indexed by the N170 amplitude, was also modulated by
3 attention. Altogether, our results provide a clear dissociation of the neural correlates of
4 conscious face perception from the effects of visual attention.

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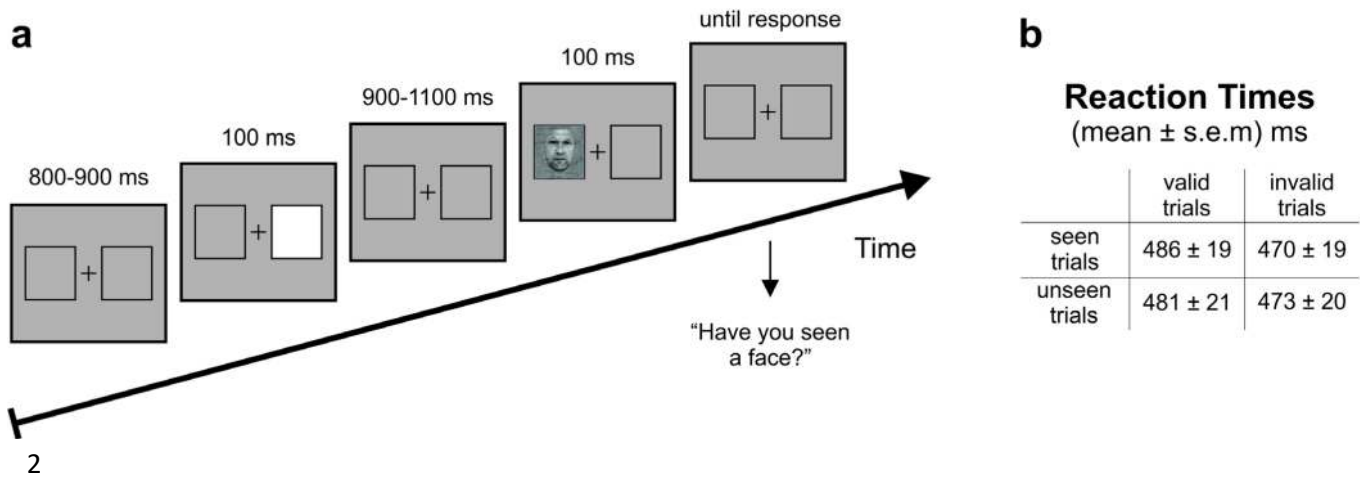
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1 Figure 1



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4 **Figure 1.** Paradigm and behavioral results. **a)** Temporal structure of the experiment. Trials

5 began with a central fixation cross and two lateral squares. After 800-900 ms, a bright flash

6 (100 ms) was presented in one of the two sides. After another interval of 900-1100 ms, a

7 face or a car was presented for 100 ms on a location that matched (valid trials) or did not

8 match (invalid trials) the one of the flash. We added zero-mean Gaussian noise to the stimuli

9 and manipulated the variance of the noise using a double-staircase procedure that kept

10 running throughout the experiment. Participants were instructed to report whether or not they

11 saw a face by pressing one of two buttons in a response box. **b)** Reaction times as a

12 function of experimental condition (rows: perceptual conditions; columns: attentional

13 conditions). Participants were faster in invalid trials compared to valid trials, irrespective of

14 their conscious report.

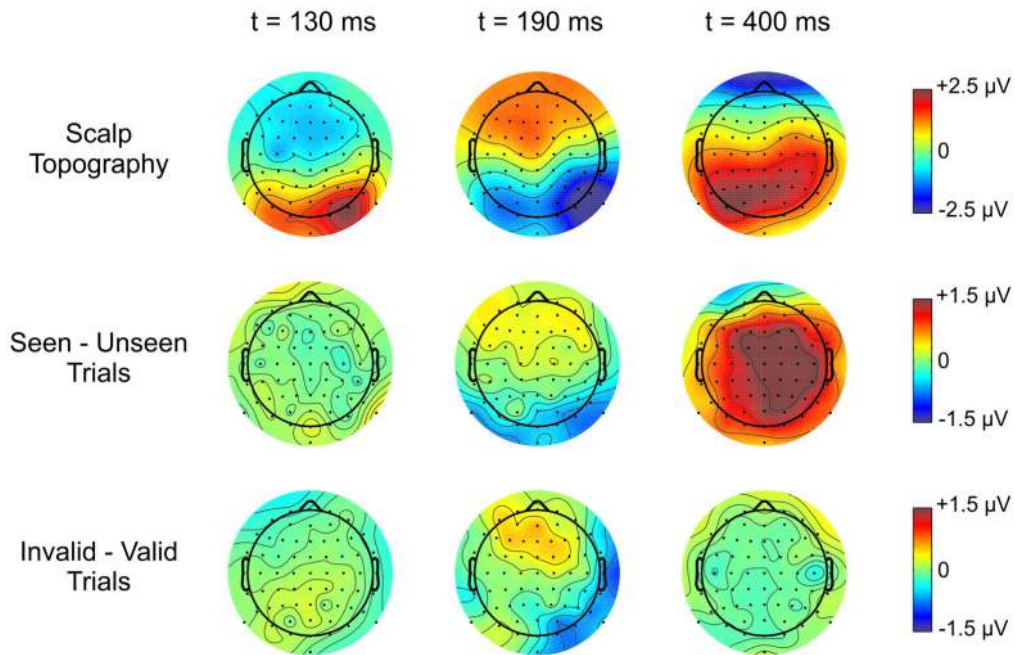
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2 **Figure 2**



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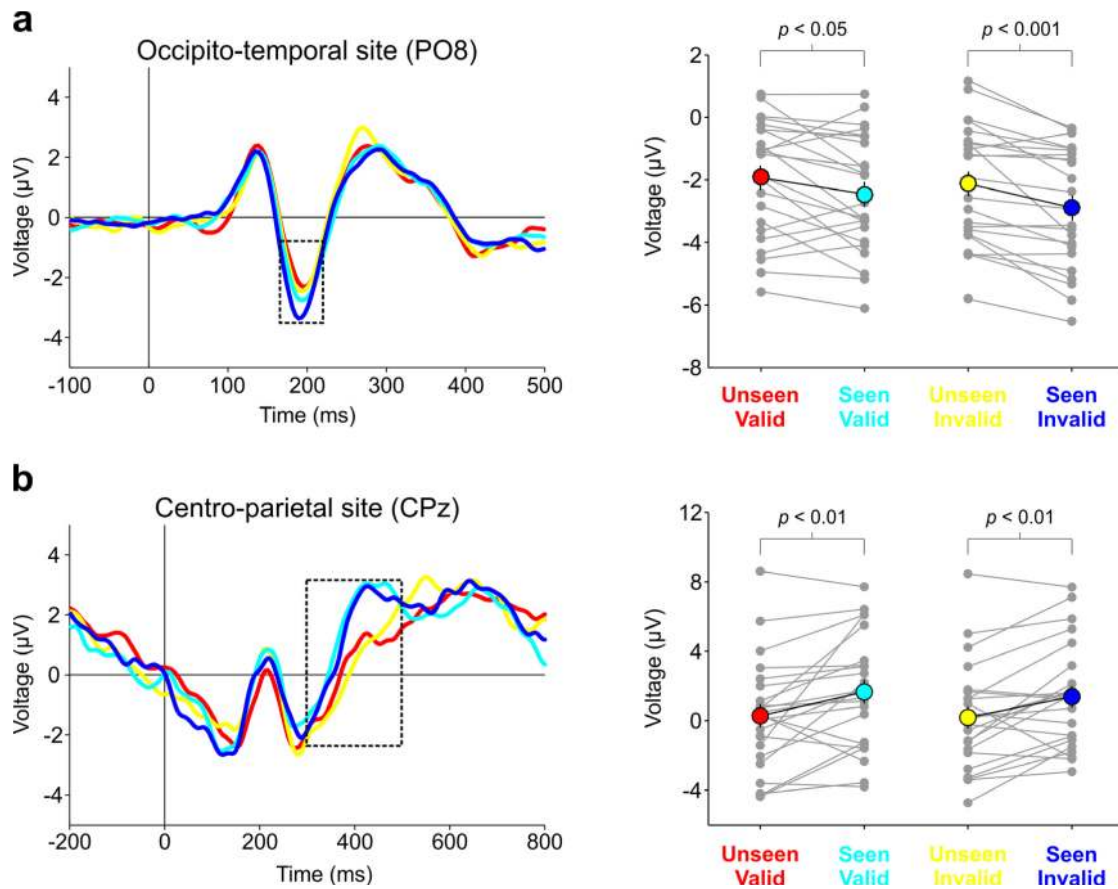
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5 **Figure 2.** Scalp topography and differences between conditions, for stimuli presented in the
6 left visual field, at the time of the three observed components. Columns: The left column
7 corresponds to the P1 component, which was maximal at 130 ms, the central column
8 corresponds to the N170, which peaked at 190 ms, and the right column corresponds to the
9 P3 component which was maximal at 400 ms. Rows: The upper row shows the scalp
10 topography of each component, the middle row displays the difference between seen and
11 unseen trials, and the lower row shows the difference between invalid and valid trials.

12

13

1 **Figure 3**

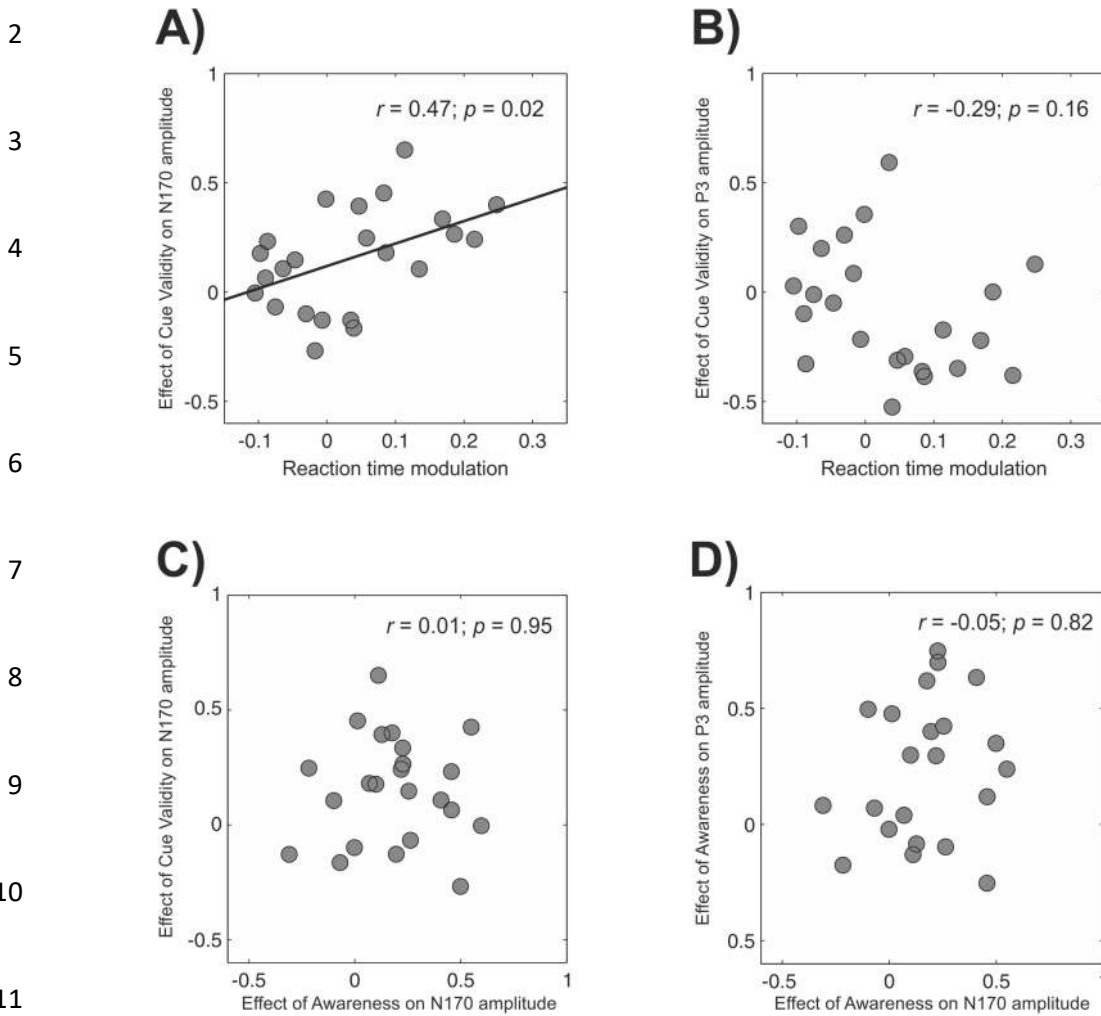


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3 **Figure 3.** Event-related potentials (ERPs) for the four conditions in this study, for stimuli
 4 presented in the left visual field. **a)** Left panel: ERPs at electrode PO8 located at an occipito-
 5 temporal site. The dotted lines mark the time window used to measure the N170 component.
 6 Right panel: Colored circles and vertical lines show the mean and s.e.m. of the N170
 7 amplitudes for each condition. Grey circles connected by lines show individual data. There
 8 was a significant effect of attention and also of conscious perception. **b)** Left panel: ERPs at
 9 electrode CPz located at a centro-parietal site. The dotted lines remark the time window
 10 used to measure the P3 component. Right panel: same as **a)** for the P3 responses. There
 11 was a significant effect of conscious perception but not of attention. Asterisks show
 12 statistically significant comparisons.

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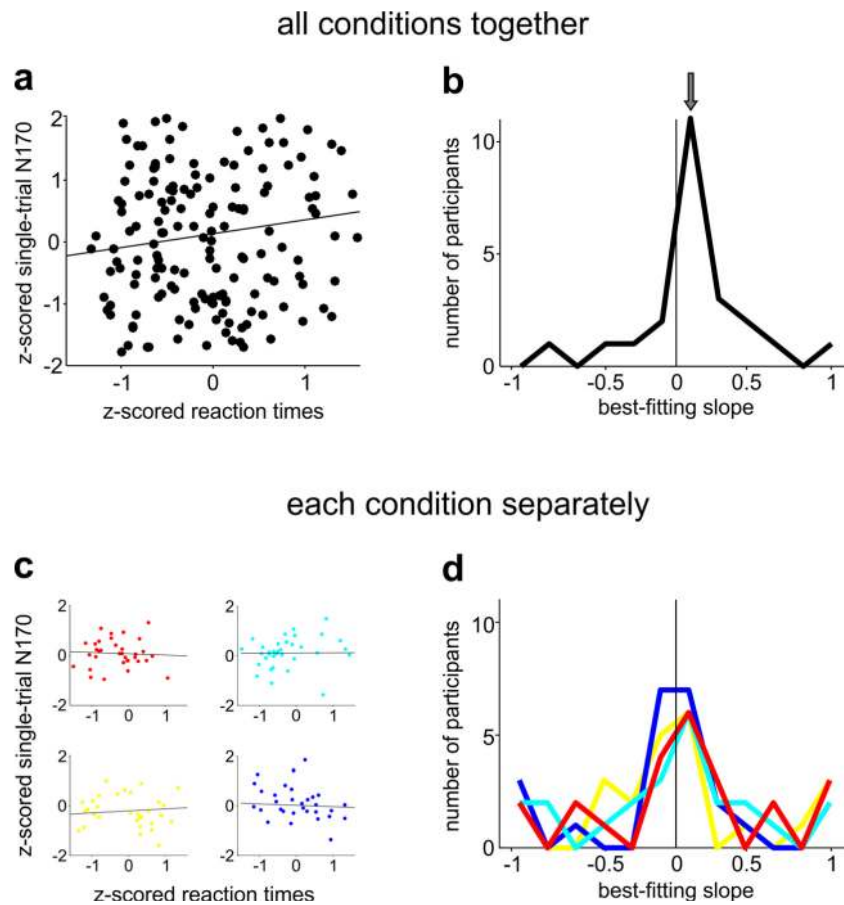
1 **Figure 4**



12 **Figure 4. a,b)** Subject-by-subject correlation of reaction time changes (valid vs. invalid trials)
13 and N170 (a) and P3 (b) changes. See Methods and Equation 1 for details. Each dot
14 represents results for a single subject. Pearson correlation coefficients (r) and their
15 associated p-values are shown for each plot. The line shows the linear fit in the only
16 significant comparison. c,d) Subject-by-subject correlation between the different significant
17 effects observed at the average level (see Fig. 3, and Equation 2 for details). c) The effect of
18 awareness on the N170 is uncorrelated with the effect of attention (cue validity) on the same
19 component. d) The effect of awareness on the N170 and P300 were not correlated with each
20 other.

21

1 **Figure 5**



2

3 **Figure 5.** Testing for residual effects of attention on the N170. **a,b)** Considering all
4 conditions together, the single-trial N170 amplitude was positively correlated with reaction
5 time. **a)** z-scored single-trial N170 versus z-scored reaction times for one representative
6 participant. **b)** Distribution of best-fitting slopes among participants. The arrow points to the
7 value of the participant selected in **(a)**, which is the median participant. **c,d)** Considering
8 each condition separately, the single-trial N170 amplitude was uncorrelated with reaction
9 time. **c)** Based on the data of the same participant selected in **(a)**, panels show z-scored
10 single-trial N170 amplitudes versus z-scored reaction times for each condition. Colors are
11 the same as in **Fig. 3**: red for “unseen valid”, cyan for “seen valid”, yellow for “unseen
12 invalid”, and blue for “seen invalid”. **d)** Distribution of best-fitting slopes among participants
13 for each condition.

1

2 **Table 1**

Decoded variable	N170	P3
Conscious reports & Attention	(28.1±0.5)%, p<10 ⁻⁴	(27.4±0.8), p<10 ⁻⁴
Conscious reports	(54.0±0.6)%, p<10 ⁻⁵	(54.1±1.1)%, p<0.001
Attention	(53.1±0.4)%, p<10 ⁻⁵	(50.9±1.5)%, p=0.42
Conscious reports for attended (invalid) trials	(53.8±0.9)%, p<0.001	(53.4±1.2)%, p<0.01
Conscious reports for unattended (valid) trials	(55.6±1.0)%, p<10 ⁻⁴	(55.4±1.1)%, p<10 ⁻⁴
Attention for seen trials	(54.4±0.8)%, p<10 ⁻⁴	(52.3±1.4)%, p=0.11
Attention for unseen trials	(52.7±0.6)%, p<0.01	(48.2±2.6)%, p=0.43

3

4 **Table 1. Decoding perception and awareness with different event-related potentials**

5 Decoding performance (mean±s.e.m.) for different conditions (rows) using different
6 components (columns). Statistical comparisons were performed by comparing the measured
7 performance with the one obtained with random permutations.

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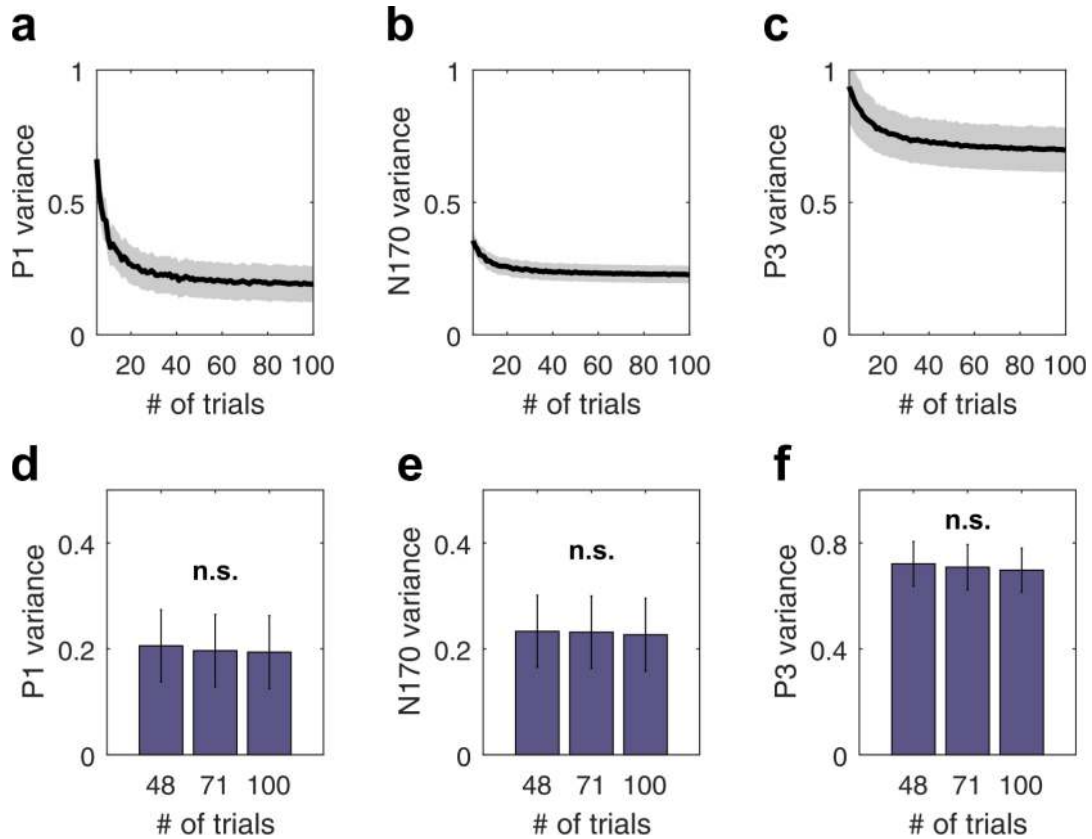
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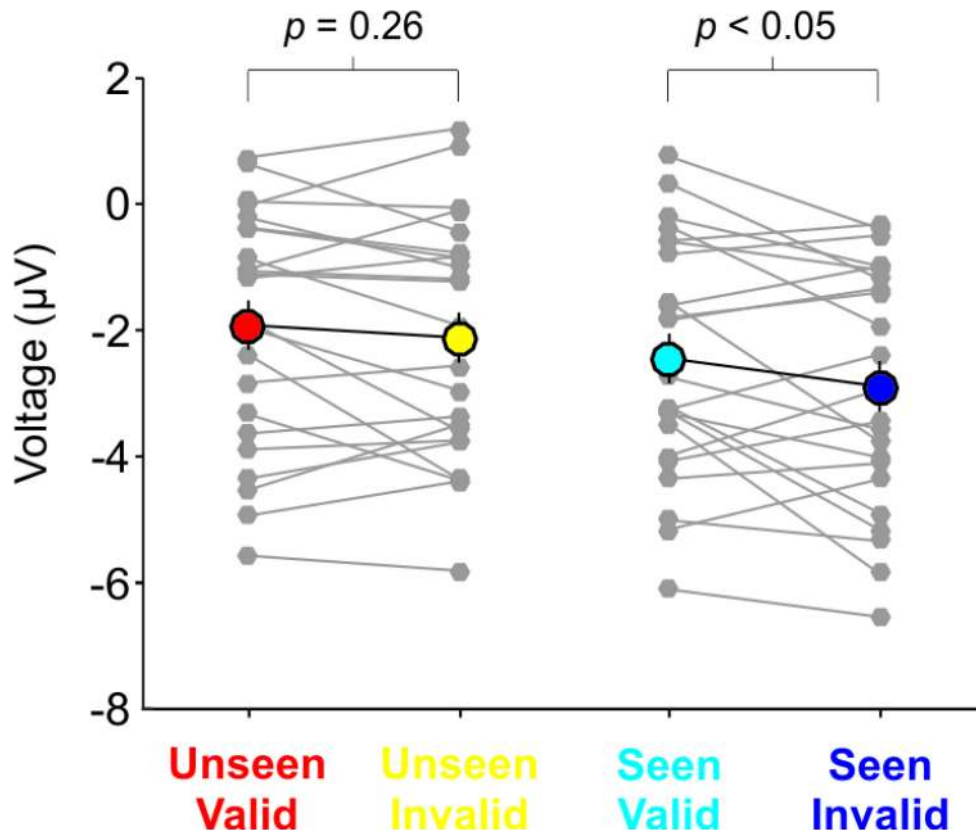


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4 **Supplementary Figure 1. Assessing the reliability of different ERP components.** We
5 mixed all conditions together and tested the reliability of our ERP components. **a-c)** For each
6 number of trials (x-axis), we randomly selected 200 combinations of trials and measured the
7 P1 (**a**), N170 (**b**), and P3 (**c**) for each iteration. The y-axis show the variance in the estimate
8 of each component as a function of increasing number of trials. The black line is the mean
9 across subjects and the grey shaded area is the s.e.m. We observed that variance
10 decreases for increasing number of trials, but that this effect reaches a plateau at around
11 ~40 trials. **d-f)** To formally test this, we compared the variance in the estimate of each
12 component for 48 trials (the minimum number of trials per condition that we have in our
13 study), 71 trials (the average number of trials per condition in our study), and 100 trials.

1 Sign-rank tests confirmed that the difference in reliability for this range was statistically
2 indistinguishable from chance ($z < 0.45$, $p > 0.6$ for all pairwise comparisons).

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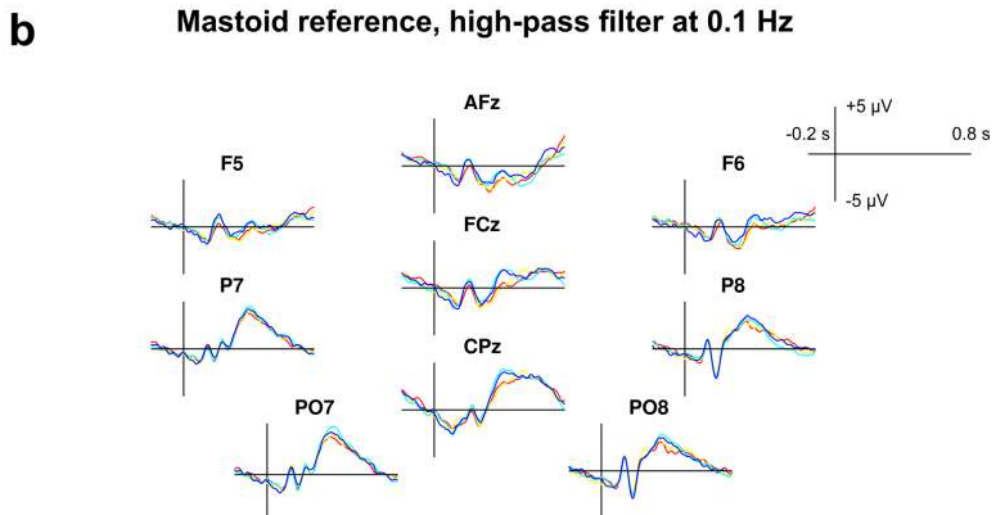
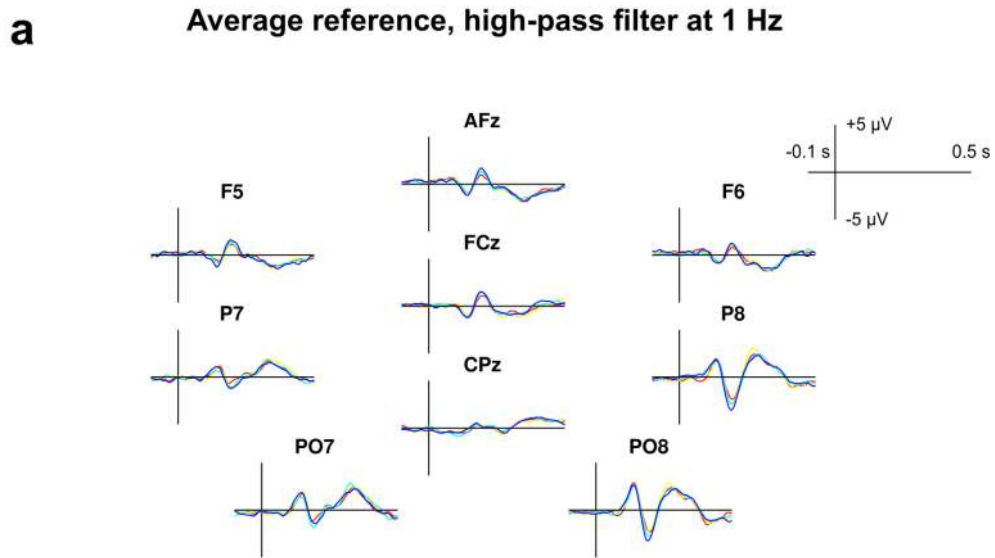
8 **Supplementary Figure 2. Post-hoc comparisons for the effect of cue validity on the**
9 **average N170 component.** Same as Fig. 3a but changing the order of the conditions (x-
10 axis) to allow pairwise comparisons across valid and invalid trials. Each grey dot is a
11 different participant and lines connect the same participants across conditions. Post-hoc
12 tests showed that the effect of cue validity was significant for seen trials (“valid” condition: (-
13 2.4 ± 0.4) μV , “invalid” condition: (-2.9 ± 0.4) μV ; Hedges’ g CI = [0.01 – 0.29]; $t(22)=2.12$,
14 $p < 0.05$), but it was absent in the unseen condition (“valid” condition: (-1.9 ± 0.4) μV , “invalid”
15 condition: (-2.1 ± 0.4) μV ; Hedges’ g CI = [-0.02 – 0.48]; $t(22)=1.14$, $p=0.26$).

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5 **Supplementary Figure 3. Event-related potentials in different electrode sites.** For the
6 two references and filter parameters used in this study, we show ERPs for stimuli presented
7 in the left visual field at 9 different electrode sites. The colors depict the 4 conditions of the
8 experiment as in **Fig 3**. Red: unseen valid; cyan: seen valid; yellow: unseen invalid; blue:
9 seen valid. **a)** We filtered the epochs using a high-pass filter at 1 Hz and used the average

1 across all electrodes as reference. This was done to focus on the P1/N170 components (as
2 in **Fig. 3a**). **b**) We filtered our data with a high-pass filter at 0.1 Hz and referenced all
3 channels to the average between two electrodes located at the mastoids. This analysis was
4 performed to focus on the P3 component (**Fig. 3b**).