

SLEEP DURATION/SLEEP QUALITY

Effects of Wind Turbine Noise on Self-Reported and Objective Measures of Sleep

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Study Objectives: To investigate the association between self-reported and objective measures of sleep and wind turbine noise (WTN) exposure.

Methods: The Community Noise and Health Study, a cross-sectional epidemiological study, included an in-house computer-assisted interview and sleep pattern monitoring over a 7 d period. Outdoor WTN levels were calculated following international standards for conditions that typically approximate the highest long-term average levels at each dwelling. Study data were collected between May and September 2013. One adult, aged 18–79 y (606 males, 632 females) was randomly selected from households (1 per) between 0.25 and 11.22 kilometers from operational wind turbines in two Canadian provinces. Self-reported sleep quality over the past 30 d was assessed using the Pittsburgh Sleep Quality Index. Additional questions assessed the prevalence of diagnosed sleep disorders and the magnitude of sleep disturbance over the previous year. Objective measures for sleep latency, sleep efficiency, total sleep time, rate of awakening bouts, and wake duration after sleep onset were recorded using the wrist worn Actiwatch2[®] from a subsample of 707 participants (320 males, 387 females) for a total of 4,371 sleep nights.

Results: Participant response rate for the interview was 78.9%. Outdoor WTN levels reached 46 dB(A) with an arithmetic mean of 35.6 and a standard deviation of 7.4. Self-reported and objectively measured sleep outcomes consistently revealed no apparent pattern or statistically significant relationship to WTN levels. However, sleep was significantly influenced by other factors, including, but not limited to, the use of sleep medication, other health conditions (including sleep disorders), and annoyance with blinking lights on wind turbines.

Conclusions: Study results do not support an association between exposure to outdoor WTN up to 46 dB(A) and an increase in the prevalence of disturbed sleep. Conclusions are based on WTN levels averaged over 1 y and, in some cases, may be strengthened with an analysis that examines sleep quality in relation to WTN levels calculated during the precise sleep period time.

Keywords: actigraphy, annoyance, multiple regression models, PSQI, sleep, wind turbine noise

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Significance

This study provides the most comprehensive assessment to date of the potential association between exposure to wind turbine noise (WTN) and sleep. As the only study to include both subjective and objective measures of sleep, the results provide a level of insight that was previously unavailable. The absence of an effect of WTN on sleep is based on an analysis of self-reported and objectively measured outcomes in relation to long term outdoor average sound levels. Knowledge in this area may be strengthened by future research to consider the potential transient changes in WTN levels throughout the night, which may influence subtle measures of sleep not assessed in the current study.

INTRODUCTION

Sleep loss has been implicated in a variety of negative health outcomes¹ including cardiovascular abnormalities,² immunological problems,³ psychological health concerns,⁴ and neurobehavioral impairment that can lead to accidents.⁵ Sleep loss may be related to total sleep time restriction and/or reduced sleep quality in the sleep time obtained. Sleep disorders such as insomnia and obstructive sleep apnea are associated with an increased incidence of hypertension, heart failure, and stroke.^{6,7}

Sleep can clearly be disrupted with noise.⁸ It has long been recognized that electroencephalography (EEG) arousals can be induced with external environmental stimuli, but are modulated by sleep state.⁹ The World Health Organization (WHO) Guidelines for Community Noise recommend that, for continuous noise, an indoor sound level of 30 dB(A) should not be exceeded during the sleep period time to avoid sleep disturbance.¹⁰ More recently, the WHO's Night Noise Guidelines for

Europe¹¹ suggest a nighttime annual average outdoor level of 40dB(A) to reduce negative health outcomes from sleep disturbance even among the most vulnerable groups.

Sleep can be measured by subjective and objective means¹² although due to the fundamental nature of unconsciousness in this state, people are unable to introspect on their sleep state. As such, an individual may surmise the quality of his or her sleep, with descriptions of what his or her presumed sleep was like, periods of awakening, and consequences of the state. However, sleep state misperception is a common clinical phenomenon, whereby patients with some degree of insomnia may report much worse quality of sleep than what actually occurred.¹³ Subjective interpretation of sleep state is thus subject to biased reporting from the individual and therefore subjective and objective measures of sleep are frequently discordant. Therefore, objective physiological measures of sleep can provide a more accurate reflection of what actually happened during an individual's sleep and form the

basis of an unprejudiced understanding of the actual biological effect of factors such as noise on sleep.

Although the current study is the first to include objective measures in the assessment of sleep quality in the context of wind turbine noise (WTN) exposure, the psychological experience of the individual must be considered, though this factor may be more prone to subjective interpretation. Numerous subjective scales of sleep have been devised. The Pittsburgh Sleep Quality Index (PSQI)¹⁴ is a measure of the subjective experience of sleep that has had detailed psychometric assessment,¹⁵ validation in numerous populations,^{16–18} and is one of the most common subjective methodologies used in sleep research.

The PSQI has been administered in a study to compare subjective sleep quality among 79 subjects living near two different wind farms wherein it was reported that sleep quality was worse among the group living closer to the wind turbines.¹⁹ Pedersen²⁰ found that self-reported sleep disturbance for any reason from any source was inconsistently related to the level of WTN. Bakker et al.²¹ showed that self-reported sleep disturbance was correlated to WTN level, but when noise annoyance from wind turbines was brought into a multiple regression, sleep disturbance appeared to be highly correlated to the annoyance, but not to WTN level and only annoyance was statistically correlated to WTN level. This is consistent with the study by van den Berg et al.²² wherein noise annoyance was reported as a better predictor of self-reported sleep disturbance than noise level for transportation, industrial, and neighbor noise.

Several studies have provided objectively measured assessments of transportation noise-induced sleep disturbance.^{23–26} Although it is clear that noise is among the many factors that contribute to sleep disturbance^{23,24,27,28} there has been no study to date that has provided an assessment of sleep disturbance in the context of WTN exposures using objective measures such as actigraphy.

The current study was designed to objectively measure sleep in relation to WTN exposure using actigraphy, which has emerged as a widely accepted tool for tracking sleep and wake behavior.^{29,30} The objective measures of sleep, when considered together with self-report, provide a more comprehensive evaluation of the potential effect that WTN may have on sleep.

This study was approved by the Health Canada and Public Health Agency of Canada Review Ethics Board (Protocol #2012-0065 and #2012-0072).

METHOD

Sample Design

Target population, sample size, and sampling frame strategy

Several factors influenced the determination of the final sample size, including having adequate statistical power to assess the study objectives, and adequate time allocation for collection of data, influenced by the length of the personal in dwelling interview and the time needed to collect the physical measures. Overall statistical power for the study was based on the study's primary objective to assess WTN-associated effects on sleep quality. Based on an initial sample size of 2,000 potential dwellings, it was estimated that there would be 1,120 completed survey responses. For 1,120 survey responses there should be

sufficient statistical power to detect at least a 7% difference in the prevalence of sleep disturbances with 80% power and a 5% false positive rate (Type I error). There was uncertainty in the power assessment because the current *Community Noise and Health Study*, was the first to implement objectively measured endpoints to study the possible effects of WTN on sleep. How these power calculations applied to actigraphy-measured sleep was also unknown. In the absence of comparative studies, a conservative baseline prevalence for reported sleep disturbance of 10% was used.^{31,32} Sample size calculation also incorporated the following assumptions: (1) approximately 20% to 25% of the targeted dwellings would not be valid dwellings (i.e., demolished, unoccupied seasonal, vacant for unknown reasons, under construction, institutions, etc.); and (2) of the remaining dwellings, there would be a 70% participation rate. These assumptions were validated (see response rates and sample characteristics related to sleep).

Study locations were drawn from areas in southwestern Ontario (ON) and Prince Edward Island (PEI) where there were a sufficient number of dwellings within the vicinity of wind turbine installations. The ON and PEI sampling regions included 315 and 84 wind turbines, respectively. The wind turbine electrical power outputs ranged between 660 kW to 3 MW (average 2.0 ± 0.4 MW). All turbines were modern monopole tower design with three pitch-controlled rotor blades (~80 m diameter) upwind of the tower and most had 80 m hub heights. All identified dwellings within approximately 600 m from a wind turbine and a random selection of dwellings between 600 m and 11.22 km were selected from which one person per household between the ages of 18 and 79 y was randomly selected to participate. The final sample size in ON and PEI was 1,011 and 227, respectively. Participants were not compensated in any way for their participation.

Wind turbine sound pressure levels at dwellings

Outdoor sound pressure levels were estimated at each dwelling using both ISO 9613-1³³ and ISO 9613-2³⁴ as incorporated in the commercial software CadnaA version 4.4.³⁵ The resulting calculations represent long-term (1 y) A-weighted equivalent continuous outdoor sound pressure levels (LAeq). Therefore, calculated sound pressure levels can only approximate with a certain degree of uncertainty the sound pressure level at the dwelling during the reference time periods that are captured by each measure of sleep. The time reference period ranges from 1–7 d (actigraphy), to 30 d for the PSQI and the previous year for the assessment of the percentage highly sleep disturbed. Van den Berg³⁶ has shown that, in the Dutch temperate climate, the long-term average WTN level for outdoor conditions is 1.7 ± 1.5 dB(A) below the sound pressure level at 8 m/sec wind speed. Accordingly, a best estimate for the average nighttime WTN level is approximately 2 dB(A) below the calculated levels reported in this study.

Calculations included all wind turbines within a radius of 10 km, and were based on manufacturers' octave band sound power spectra at a standardized wind speed of 8 m/sec and favorable sound propagation conditions. Favorable conditions assume the dwelling is located downwind of the noise source, a stable atmosphere, and a moderate ground-based temperature inversion. Although variations in wind speeds and temperature as a func-

tion of height could not be considered in the model calculations due to a lack of relevant data, 8 m/sec was considered a reasonable estimate of the highest noise exposure conditions. The manufacturers' data were verified for consistency using on-site measurements of wind turbine sound power. The standard deviation in sound levels was estimated to be 4 dB(A) up to 1 km, and at 10 km the uncertainty was estimated to be between 10 dB(A) and 26 dB(A). Although calculations based on predictions of WTN levels reduces the risk of misclassification compared to direct measurements, the risk remains to some extent. The calculated levels in the current study represent reasonable worst-case estimates expected to yield outdoor WTN levels that typically approximate the highest long-term average levels at each dwelling and thereby optimize the chances of detecting WTN-induced sleep disturbance. The few dwellings beyond 10 km were assigned the same calculated WTN value as dwellings at 10 km. Unless otherwise stated, all decibel references are A-weighted. A-weighting filters out low frequencies in a sound that the human auditory system is less sensitive to at low sound pressure levels.

In the current study, low-frequency noise was estimated by calculating C-weighted sound pressure levels. No additional benefit was observed in assessing low frequency noise because C- and A-weighted levels were so highly correlated. Depending on how dB(C) was calculated and what range of data was assessed, the correlation between dB(C) and dB(A) ranged from $r = 0.84$ to $r = 0.97$.³⁷

Background nighttime sound levels at dwellings

As a result of certain meteorological phenomena (atmospheric stability and wind gradient) coupled with a tendency for background sound levels to drop throughout the day in rural/semi-rural environments, WTN can be more perceptible at the dwelling during nighttime.^{38–41} In Canada, it is possible to estimate background nighttime sound pressure levels according to the provincial noise regulations for Alberta, Canada,⁴² which estimates ambient noise levels in rural and suburban environments. Estimates are based on dwelling density per quarter section, which represents an area with a 451 m radius and distance to heavily travelled roads or rail lines. When modeled in accordance with these regulations, estimated levels can range from 35 dB(A) to 51 dB(A). The possibility that exposure to high levels of road traffic noise may create a background sound pressure level higher than that estimated using the Alberta regulations was considered. In ON, road noise for the six-lane concrete Highway 401 was calculated using the United States Federal Highway Administration (FHWA) Traffic Noise Model⁴³ module in the CadnaA software.³⁵ This value was used when it exceeded the Alberta noise estimate, making it possible to have levels above 51 dB(A).

Data Collection

Questionnaire administration and refusal conversion strategies

The questionnaire instrument included modules on basic demographics, noise annoyance, health effects, quality of life, sleep quality, sleep disorders, perceived stress, lifestyle behaviors, and prevalence of chronic disease. To avoid bias, the true intent of the study, which was to assess the community response to

wind turbines, was masked. Throughout the data collection, the study's official title was: *Community Noise and Health Study*. This approach is commonly used to avoid a disproportionate contribution from any group that may have distinct views toward wind turbines. Data collection took place through in-person interviews between May and September 2013 in southwestern ON and PEI. After a roster of all adults aged 18 to 79 y living in the dwelling was compiled, a computerized method was used to randomly select one adult from each household. No substitution was permitted; therefore, if the targeted individual was not at home or unavailable, alternate arrangements were made to invite them to participate at a later time.

All 16 interviewers were instructed to make every reasonable attempt to obtain interviews, which included visiting the dwelling at various times of the day on multiple occasions and making contact by telephone when necessary. If the individual refused to participate, they were then contacted a second time by either the senior interviewer or another interviewer. If, after a second contact, respondents refused to participate, the case was coded as a final refusal.

Self-reported sleep assessment

Long-term self-reported sleep disturbance included an assessment of the magnitude of sleep disturbance experienced at home (of any type for any reason) over the past year. Participants were requested to describe their level of sleep disturbance at home over the past year using one of the following categories: “not at all,” “slightly,” “moderately,” “very” or “extremely,” where the top two categories were collapsed and considered to reflect “highly sleep disturbed.” For the purposes of this analysis the bottom three categories reflect “low sleep disturbance.” These categories and the classification of “highly sleep disturbed” is consistent with the approach adopted for annoyance⁴⁴ and facilitates comparisons to self-reported sleep disturbance functions developed for transportation noise sources.⁴⁵ Data were collected on prevalence of diagnosed sleep disorders. In addition, participants completed the PSQI, which provided an assessment of sleep quality over the previous 30 d. The seven components of the PSQI are scored on a scale from 0 (better) to 3 (worse); therefore the global PSQI is a score ranging between 0–21, where a value of greater than 5 is thought to represent poor sleep quality.^{14,16–18}

Objectively measured sleep

An Actiwatch2[®] (Philips Healthcare, Andover, MA, USA) sleep watch was given to all consenting and eligible participants aged 18 to 79 y who were expected to sleep at their current address for a minimum of 3 of the 7 nights following the interview. There were 450 devices at hand that were cycled throughout the study. In order to receive the device, respondents also needed to have full mobility in the arm on which the watch was to be worn. Respondents were asked to wear the device on their wrist during all hours of the day and night for the 7 d following their interview. The Actiwatch2[®] provides key information on sleep patterns (based on movement), including timing and duration of sleep as well as awakenings, and has been compared with polysomnography in some patient samples,⁴⁶ but does not replace polysomnography due to imperfect sensitivity and specificity for detecting wake periods. However, this tool can provide

reasonable estimates for assessing subjects objectively for more prolonged periods of time than conventional assessment tools, with minimal participant burden.⁴⁷ The devices were configured to continuously record a data point every 60 sec for the entire 7 d period. Data analysis was conducted using Actiware® Version 5.70⁴⁸ with the software set to default settings (i.e., sensitivity setting of medium and a minimum minor rest interval size of 40 min). With these settings an epoch of 40 counts (i.e., accelerometer activity above threshold) or less is considered sleep and epochs above 40 counts are considered wake. However, any given epoch is scored using a 5-epoch weighting scheme. This procedure weighs the 2 epochs adjacent to the epoch in question. The 5-epoch weighting is achieved by multiplying the number of counts in each respective epoch by the following: 1/25, 1/5, 1, 1/5, 1/25, whereby an average above 40 indicates “awake” for the central epoch. Within the rest interval the sleep start parameter was automatically calculated by the Actiware® software determined by the first 10 min period in which no more than one 60 sec epoch was scored as mobile. An epoch is scored as mobile if the number of activity counts recorded in the epoch is greater than or equal to the epoch length in 15 sec intervals (i.e., in a 60 sec epoch an activity value of 4 or higher). Endpoints of interest from wrist actigraphy included sleep efficiency (total sleep time divided by measured time in bed), sleep latency (how long it took to fall asleep), wake after sleep onset (WASO) (the total duration of awakenings), total sleep time, and the number of awakening bouts (WABT) (during a sleep period). The WABT data was analysed as the rate of awakening bouts per 60 min in bed. The raw sleep data extracted from the Actiwatch2® to the Actiware® software should be visually inspected for the purposes of manually assigning the “rest interval” for each sleep night, when necessary. In our experience, the default determination of the “rest interval” was unreliable when participants did not wear the sleep watch continuously. This screening measure is essential as the “rest interval” is used to determine all of the aforementioned sleep endpoints.

To help interpret the measured data, respondents were asked to complete a basic sleep log each night of the study. The log contained information about whether the respondent slept at home or not, presence of windows in the room where they slept, and whether or not the windows were open. After the 7 d collection period, respondents were asked to return the completed sleep log with the sleep watch in a prepaid package.

Statistical Methodology

The analysis follows the description in Michaud et al.,⁴⁹ which provides a summary of the study design and objectives, as well as a proposed data analysis. Briefly, the Cochran Mantel-Haenszel chi-square test was used to detect associations between self-reported magnitude or contributing sources of sleep disturbance and WTN exposure groups while controlling for province. Because a cut-off value of 5 for the global PSQI score provided a sensitive and specific measure distinguishing good and poor sleep, the PSQI score was dichotomized with the objective to model the proportion of individuals with poor sleep quality (i.e., PSQI > 5).¹⁴ As a first step to develop the best model to predict the dichotomized PSQI score, univariate logistic regression models only adjusting for WTN exposure groups and province were carried out. It should be emphasized that variables considered in

the univariate analysis have been previously demonstrated to be related to the modeled endpoint and/or considered by the authors to conceptually have a potential association with the modeled endpoint. The analysis of each variable only adjusts for WTN category and province; therefore, interpretation of any individual relationship must be made with caution.

The primary objective in the current analysis was to use multiple regression models to identify the best predictors for (1) reporting a PSQI score greater than 5; and (2) the actigraphy endpoints. All explanatory variables that were statistically significant at the 20% level in the univariate analysis for each respective endpoint were considered in the multiple regression models. To develop the best model to predict each endpoint of interest, the stepwise method, which guards against issues of multicollinearity, was used for multiple regression models.

The stepwise regression was carried out in three different ways wherein the base model included: (1) WTN exposure category and province; (2) WTN exposure category, province, and an adjustment for individuals who reported receiving personal benefit from having wind turbines in the area; and (3) WTN category and province, only for those who received no personal benefit.

For the analysis of PSQI, multiple logistic regression models were developed using the stepwise method with a 20% significance entry criterion and a 10% significance criterion to remain in the model. The WTN groups were treated as a continuous variable, giving an odds ratio (OR) for each unit increase in WTN level, where a unit reflects a 5 dB(A) WTN category. The Nagelkerke pseudo R² is reported for logistic regression models.

Repeated-measures data from all wrist actigraphy measurements were modeled using the generalized estimating equations (GEE) method, as available in SAS (Statistical Analysis System) version 9.3 PROC GENMOD.^{50–52} Univariate GEE regression models only adjusting for WTN exposure groups, province, day of the week, and the interaction between WTN groups and day of the week were carried out. The base model for the multiple GEE regression models included the same variables as in the univariate GEE models. The same stepwise methodology that was applied to build the PSQI models was used to develop multiple GEE regression models for each actigraphy endpoint. The within-subjects correlations were examined with different working correlation matrix structures (unstructured, compound symmetry, and autoregressive of first order). An unstructured variance-covariance structure between sleep nights was applied to all endpoints with the exception of sleep latency, where compound symmetry was used. The advantage of the GEE method is that it uses all available data to estimate individual subject variability (i.e., if 1 or more nights of data is missing for an individual, the individual is still included in the analysis).

The wrist actigraphy endpoints of sleep efficiency and rate of awakening bouts do not follow a normal distribution, because one is a proportion ranging between 0 and 1 (sleep efficiency) and the other is a count (awakening bouts). Therefore, to analyze awakening bouts a Poisson distribution was assumed. The number of awakening bouts was analyzed with respect to the total time spent in bed and is reported as a rate of awakening bouts per 60 min in bed. Sleep efficiency, sleep latency, and WASO were transformed in order to normalize the data and stabilize the variance.^{53–55} In the GEE models, statistical tests were based on transformed data in order to satisfy the normality and

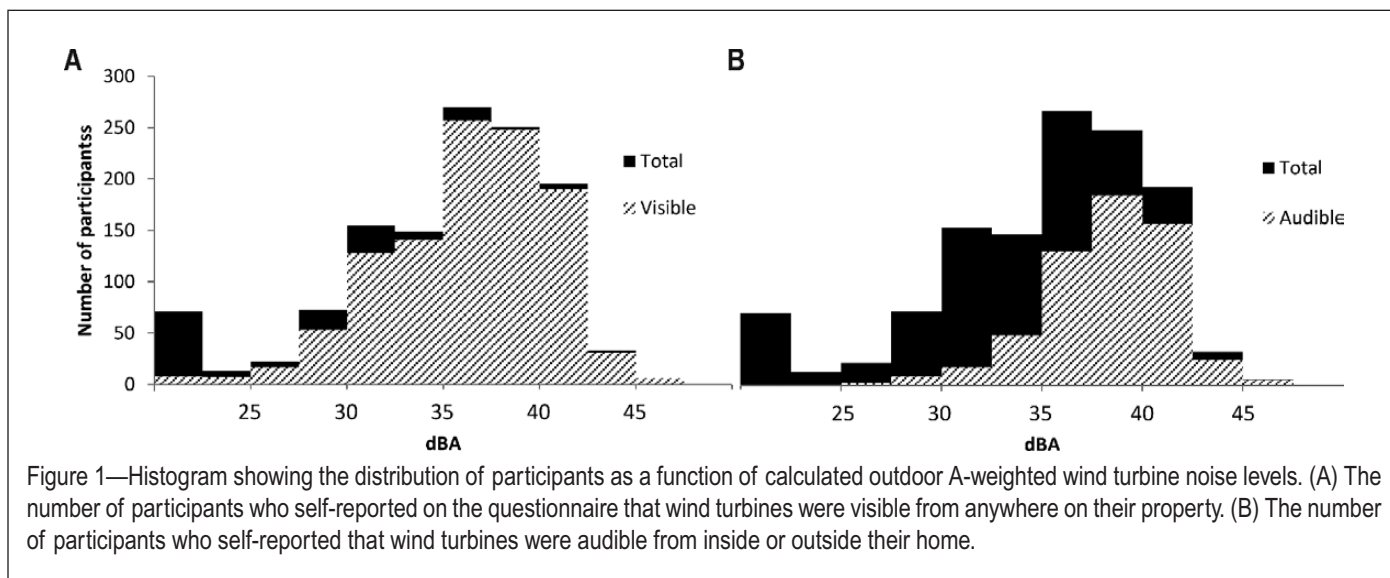


Figure 1—Histogram showing the distribution of participants as a function of calculated outdoor A-weighted wind turbine noise levels. (A) The number of participants who self-reported on the questionnaire that wind turbines were visible from anywhere on their property. (B) The number of participants who self-reported that wind turbines were audible from inside or outside their home.

constant variance assumptions. Because back-transformation was not possible for some endpoints, the arithmetic mean (least squares mean [LSM]) is presented for all endpoints.

All regression models for PSQI and actigraphy endpoints were adjusted for provincial differences. Province was initially assessed as an effect modifier. Because the interaction was not statistically significant for any of the multiple regression models, province was treated as a confounder in the models with associated adjustments, as required. Statistical analysis was performed using SAS version 9.3. A 5% statistical significance level was implemented throughout unless otherwise stated and Tukey corrections were applied to account for all pairwise comparisons to ensure that the overall Type I (false positive) error rate was less than 0.05.

Actigraphy Data Screening

The sleep actigraphy file consisted of 5,280 nights of actigraphy measured sleep (i.e., sleep nights) data from 821 participants. The following adjustments to the file were made: sleep nights with data from shift workers ($n = 732$), and data related to sleep nights away from home ($n = 147$) were removed. Removal of these data supported the objective to relate typical sleep behavior to noise exposure from wind turbines at the participants' dwelling. Sleep starting after 05:00 with awakening on the same day before 23:59 was considered day sleep and removed from the analysis ($n = 30$). The final sample size consisted of 4,371 sleep nights and 707 participants. Any sleep that started after midnight, but before 05:00 was re-coded and considered as sleep for the previous night to avoid having two sleep observations for the same night. For the remaining data, all available data was used whether the person wore the watch for 1 d or for the maximum 7 d.

RESULTS

Wind Turbine Sound Pressure Levels at Dwellings

Calculated outdoor sound pressure levels at the dwellings determined by ISO 9613-1³³ and ISO 9613-2³⁴ reached levels as high as 46 dB(A). Results are considered to have an uncertainty of ± 4 dB(A) within distances that would have the strongest effect on

sleep (i.e., ~ 600 m). Figure 1 illustrates the distribution of participants as a function of WTN levels and identifies the number of participants who reported wind turbines were visible from anywhere on their property (panel A) and audible (panel B) while they were either outside or inside their dwelling.

Background Nighttime Sound Pressure Levels

Modeled background nighttime sound (BNTS) levels ranged between 35 and 61 dB(A) in the sample. Average BNTS was highest in the WTN group 30–35 dB(A) and lowest in areas where modeled WTN levels were between 40–46 dB(A).³⁷ In the univariate analysis of global PSQI, the proportion of people with poor sleep (i.e., global scores above 5) was statistically similar among the BNTS levels ($P = 0.9727$) (see supplemental material).

Response Rates and Sample Characteristics Related to Sleep

A detailed breakdown of the response rates, along with personal and situational variables by WTN category, is presented by Michaud.³⁷ Of the 2,004 potential dwellings, 1,570 were valid and 1,238 agreed to participate in the survey (606 males, 632 females), resulting in a final overall response rate of 78.9%. Of the 1,238 participants, 1,208 completed the PSQI in its entirety (97.6%) and 821 participated in the sleep actigraphy portion of the study (66%). Sleep actigraphy participation rates were in line with projections based on an unpublished pilot study designed to assess different sleep watch devices and participant compliance. Participation rate was equally distributed across WTN categories.

The prevalence of reporting a diagnosed sleep disorder was unrelated to WTN levels ($P = 0.3102$).³⁷ In addition, the use of sleep medication at least once a week was significantly related to WTN levels ($P = 0.0083$). The prevalence was *higher* among the two lowest WTN categories (< 25 dB(A) and 25–30 dB(A)).³⁷ Factors that may affect sleep quality, such as self-reported health conditions, chronic illnesses, quality of life, and noise sensitivity were all found to be equally distributed across WTN categories.^{37,56} In response to the general question on magnitude of sleep disturbance for any reason over the past year while at home, a total of 757 participants (61.3%) reported at least a “slight” magnitude of sleep disturbance (includes ratings of “slightly,” “moderately,” “very” and “extremely”),

Table 1—Self-reported magnitude and contributing sources of sleep disturbance.

Variable	Wind Turbine Noise, dB(A)					Overall	CMH P value ^a
	< 25	25–30	30–35	35–40	40–46		
n	83	95	304	519	234	1,235	
Self-reported sleep disturbance n (%)							
Not at all	29 (34.9)	44 (46.3)	112 (36.8)	208 (40.1)	85 (36.3)	478 (38.7)	
At least slightly ^b	54 (65.1)	51 (53.7)	192 (63.2)	311 (59.9)	149 (63.7)	757 (61.3)	0.7535
Highly ^c	13 (15.7)	11 (11.6)	41 (13.5)	75 (14.5)	24 (10.3)	164 (13.3)	0.4300
Source of sleep disturbance (among participants at least slightly sleep disturbed) n (%)							
n ^d	53	51	186	298	138	726	
Wind turbine	0 (0.0)	2 (3.9)	4 (2.2)	45 (15.1)	31 (22.5)	82 (11.3)	< 0.0001
Children	9 (17.0)	12 (23.5)	21 (11.3)	36 (12.1)	20 (14.5)	98 (13.5)	0.2965
Pets	7 (13.2)	12 (23.5)	9 (4.8)	45 (15.1)	22 (15.9)	95 (13.1)	0.3582
Neighbors	6 (11.3)	5 (9.8)	9 (4.8)	13 (4.4)	5 (3.6)	38 (5.2)	0.0169
Other	41 (77.4)	35 (68.6)	162 (87.1)	232 (77.9)	87 (63.0)	557 (76.7)	0.0128
Stress/anxiety	6 (11.3)	2 (3.9)	21 (11.3)	33 (11.1)	11 (8.0)	73 (10.1)	0.8938
Physical pain	11 (20.8)	9 (17.6)	50 (26.9)	48 (16.1)	18 (13.0)	136 (18.7)	0.0289
Snoring	5 (9.4)	6 (11.8)	17 (9.1)	20 (6.7)	12 (8.7)	60 (8.3)	0.4126

Participants were asked to report their magnitude of sleep disturbance over the last year while at home by selecting one of the following five categories: not at all, slightly, moderately, very, or extremely. Participants that indicated at least a slight magnitude of sleep disturbance were asked to identify all sources perceived to be contributing to sleep disturbance. ^aThe Cochran Mantel-Haenszel chi-square test was used to adjust for provinces. ^bAt least slightly sleep disturbed includes participants indicating the slightly, moderately, very or extremely categories. ^cHighly sleep disturbed includes participants who reported the very or extremely categories. The prevalence of reported sleep disturbance was unrelated to wind turbine noise levels. ^dOf the 757 participants who reported at least a slight amount of sleep disturbance, 31 did not know what contributed to their sleep disturbance. Of the remaining 726, at least one source was identified. Columns may not add to sample size totals as some participants did not answer questions and/or identified more than one source as the cause of their sleep disturbance.

Table 2—Summary of Pittsburgh Sleep Quality Index scores.

	Wind Turbine Noise, dB(A)					Overall
	< 25	25–30	30–35	35–40	40–46	
Mean (95% CI)	6.22 (5.32, 7.11)	5.91 (5.05, 6.77)	6.00 (5.51, 6.50)	5.74 (5.33, 6.16)	6.09 (5.55, 6.64)	5.94 (5.72, 6.17)
n (%) score > 5 ^a	40 (49.4)	45 (48.9)	138 (46.5)	227 (44.4)	106 (46.7)	556 (46.0)

^aPittsburgh Sleep Quality Index score above 5 is considered to represent poor sleep. CI, confidence interval.

with a total of 164 (13.3%) classified as “highly” sleep disturbed (i.e., either “very” or “extremely”). The levels of WTN were not found to have a statistically significant effect on the prevalence of sleep disturbance whether the analysis was restricted to only participants highly sleep disturbed ($P = 0.4300$), or if it included all participants with even a slight disturbance ($P = 0.7535$) (Table 1). When assessing the sources reported to contribute to sleep disturbance among all participants with even slight disturbance, reporting wind turbines was significantly associated with WTN categories ($P < 0.0001$). The prevalence was $\geq 15.1\%$ among the participants living in areas where WTN levels were ≥ 35 dB(A) compared to $\leq 3.9\%$ in areas where WTN levels were below 35 dB(A). However, wind turbines were not the only, nor the most prevalent, contributing source at these sound levels (see Table 1).

PSQI Scores

For the 1,208 participants who completed the PSQI in its entirety, the average PSQI score across the entire sample was 5.94 with 95% confidence interval (CI) (5.72, 6.17). The Cronbach alpha

for the global PSQI was 0.76 (i.e., greater than the minimum value of 0.70 in order to validate the score). Table 2 presents the summary statistics for PSQI as both a continuous scale and a binary scale (the proportion of respondents with poor sleep; i.e., PSQI above 5) by WTN exposure categories. Analysis of variance was used to compare the average PSQI score across WTN exposure groups (after adjusting for provinces). There was no statistical difference observed in the mean PSQI scores between groups ($P = 0.7497$) as well as no significant difference between provinces ($P = 0.7871$) (data not shown). Similarly, when modeling the proportion of respondents with poor sleep (PSQI > 5) in the logistic regression model, no statistical differences between WTN exposure groups ($P = 0.4740$) or provinces ($P = 0.6997$) were observed (see supplemental material).

Effects of Personal and Situational Variables on PSQI Scores and Actigraphy

A univariate analysis of the personal and situational variables in relation to the PSQI scores (logistic regression) and actigraphy

endpoints (GEE) was conducted. The list of variables considered was extensive and included, but was not limited to, age, sex, income, education, body mass index, caffeine consumption, housing features, diagnosed sleep disorders, health conditions, annoyance, household complaints, and personal benefit (i.e., rent, payments or other indirect benefits through community improvements) from having wind turbines in the area. The analysis of these and several other variables in relation to the endpoints has been made available in the supplemental material.

Multiple Logistic Regression Models for PSQI

Table 3 provides a summary of the variables retained in the multiple regressions for the PSQI and actigraphy endpoints. A detailed description of the statistical results, including the direction of change and the pairwise comparisons made among the groups within each variable is available in the supplemental material.

Table 4 presents the results from stepwise multiple logistic regression modeling of the proportion of respondents with “poor sleep” (i.e., scores above 5 on the PSQI). The final models for the three approaches to stepwise regression as listed in the Statistical Methods section produced nearly identical results to one another. Therefore, results are only presented for the regression method where the variables WTN category, province, and personal benefit were forced into the model that fit the data well (Hosmer-Lemeshow test, $P > 0.05$). Using stepwise regression, the predictive strength of the final model was 37%. There was no observed relationship between the proportion of respondents with poor sleep and WTN levels ($P = 0.3165$).

Participants who had improved sleep quality after closing their bedroom window were found to have the same odds of poor sleep when compared to those who did not need to close their window ($P = 0.0565$). Participants who stated that closing their window did not improve sleep quality had higher odds of poor sleep in comparison with both those who had improved sleep quality after closing windows and those who did not need to close windows ($P \leq 0.0006$, in both cases). Unemployed individuals had higher odds of poor sleep compared with those who were employed (OR [95% CI]: 1.55 [1.12, 2.15]).

Long-term sleep disturbance (of any type by any source) was included in the study because dose-response relationships have been published for this measure in relation to other community noise sources⁴⁵ and this endpoint provides a longer time reference period than the previous 30 d assessed using the PSQI. Those who reported a very or extremely high level of sleep disturbance (i.e., percentage highly sleep disturbed) by any source while at home had 6 times higher odds of poor sleep assessed with the PSQI (OR [95%CI]: 6.28 [3.46, 11.40]) when compared to those with no, slight, or moderate reported sleep disturbance. Finally, participants suffering from migraines/headaches, asthma, arthritis and a diagnosed sleep disorder (e.g., sleep apnea or insomnia) had higher odds of poor sleep when compared to those not suffering from these health and chronic conditions.

Sleep Actigraphy

The majority of participants (66%) wore the watch for the full 7 nights (mean number of days 6.18, SD = 1.49). The frequency across the days of the week was equally distributed (data not

shown). Response rates for the sleep actigraphy were equally distributed across WTN exposure groups ($P = 0.4488$), although a higher proportion of participants were noted in PEI, in comparison to ON ($P = 0.0110$).

Table 5 presents the summary data for each sleep actigraphy endpoint analyzed. Although mean values appear stable between one sleep night to the next within an endpoint, the standard deviation is observed to fluctuate between sleep nights (data not shown). The observed correlations between the PSQI and the actigraphy endpoints are presented as supplemental material.

Multiple GEE Regression Models for Actigraphy

Multiple regression models for the five sleep actigraphy endpoints were developed. Variables that were associated with each endpoint (i.e., significant at the 10% level) are summarized in Table 3. Specific information on these variables, including the direction of change, P values, and pairwise comparisons has been made available in the supplemental material. Table 6 presents the LSM and the P values for the exposure of interest, the WTN exposure categories, obtained from the GEE regression models for the sleep actigraphy endpoints. Unadjusted results reflect the base model (including WTN, province, day of the week, and the interaction between WTN and day of the week) whereas adjusted results come from the multiple regression models obtained through the stepwise method and take into account factors beyond the base model. The level of exposure to WTN was not found to be related to sleep efficiency ($P = 0.0519$), sleep latency ($P = 0.1783$), total sleep time ($P = 0.7348$), WASO ($P = 0.3596$), or the number of awakening bouts ($P = 0.1673$). Because WTN was marginally associated with sleep efficiency ($P = 0.0519$) a pairwise comparison (PWC) was conducted, however no consistent pattern was revealed (data not shown).

DISCUSSION

The effects on health and well-being associated with accumulated sleep debt have been well documented.^{1-5,57} The sound pressure levels from wind turbines can exceed the WHO recommended annual average nighttime limit of 40 dB(A) for preventing health effects from noise-induced sleep disturbance.¹¹ The calculated outdoor A-weighted WTN levels in this study reached a maximum of 46 dB(A), with 18% of dwellings found to exceed 40 dB(A). Within an uncertainty of approximately 4 dB(A), the calculated A-weighted levels in the current study can be compared to the WHO outdoor nighttime annual average threshold of 40 dB(A).^{11,58} With the average façade attenuation with windows completely opened of 14 ± 2 dB(A),⁵⁸ the average bedroom level at the highest façade level, 46 dB(A), will be 32 ± 2 dB(A), which is close to the 30 dB(A) indoor threshold in the WHO's Guidelines for Community Noise.¹⁰ Considering the uncertainty in the calculation model and input data, only dwellings in the highest WTN category are expected to have indoor levels above 30 dB(A) and thus sensitivity to sleep disturbance. However, with windows closed, the indoor to outdoor level difference is approximately 26 dB, which should result in an indoor level around 20 dB(A) in the current study.

Factors including, but not limited to, medication use, other health effects (including sleep disorders), and annoyance with blinking lights on wind turbines were found to statistically influence reported and/or actigraphically measured sleep out-

Table 3—Variables retained in multiple generalized estimating equations and multiple logistic regression models.

	Sleep Efficiency (%)	Sleep Latency (min)	Total Sleep Time (min)	WASO (min)	Rate of Awakening Bouts (per 60 min)	PSQI (score > 5)
Base model						
WTN levels dB(A)	+					
Province						+
Day of the week			++			
Demographic variables						
Sex	++	++	++		++	
BMI group		++				
Age					++	
Employment		++	++			++
Alcohol consumption				++	++	
Education				++		
Income	+					
Situational variables						
Bedroom location	+		+			
Bedroom window type		+				
Sleep improved by closing window						++
Façade type		+				
BNTS level dB(A)				+		
Audible rail noise						++
Wind turbine related variables						
Personal benefits		+				++
Annoyance with blinking lights			++			
Annoyance with vibrations			++			
Years hearing WTN					+	
Personal and health related variables						
Self-reported sleep disturbance ^a						++
Sleep disturbed by other ^b						++
Annoyed by snoring						+
Sleep medication ^c			++		++	
Migraines			++			++
Dizziness			++			+
Chronic pain						+
Asthma				+		++
Arthritis						++
Diabetes	++			++		
Diagnosed sleep disorder						++

A summary of significant variables retained in multiple generalized estimating equations and multiple logistic regression models for objectively measured and self-reported sleep endpoints, respectively. The specific direction of change, level of statistical significance, pairwise comparisons between variable groups and full description of the variable names is provided in supplemental material. ^aEvaluates the magnitude of reported sleep disturbance at home from not at all to extremely, for any reason over the previous year. ^bThe source identified by participants as contributing to their sleep disturbance was not wind turbines, children, pets or neighbors. ^cUse of sleep medication was not considered in the multiple regression model for PSQI since it is one of the seven components that make up the global PSQI score. +, ++ denotes statistically significant, $P < 0.10$, $P < 0.05$, respectively. BMI, body mass index; BNTS, background nighttime sound level; PSQI, Pittsburgh Sleep Quality Index; WASO, wake after sleep onset; WTN, wind turbine noise.

Table 4—Multiple logistic regression model for Pittsburgh Sleep Quality Index.

Variable	Groups in Variable ^b	Model: WTN, Province, and Personal Benefit Forced in PSQI ^a	
		P value ^c (n = 933, R ² = 37%, H-L P = 0.9252) ^h	OR (CI) ^d
WTN, dB(A) ^e		0.3165	0.93 (0.80, 1.07)
Province	PEI/ON	0.0810	1.46 (0.95, 2.25)
Personal benefit	No/Yes	0.0499	1.82 (1.00, 3.30)
Sleep improved by closing window (overall P value < 0.0001)	Yes	0.0565	1.41 (0.99, 2.00)
	No Did not need to close windows	< 0.0001	8.48 (3.11, 23.14) Reference
Employment	No/Yes	0.0085	1.55 (1.12, 2.15)
Audible rail noise	No/Yes	0.0380	1.56 (1.03, 2.37)
Reported cause for sleep disturbance Other ^f	Yes/No	< 0.0001	2.55 (1.86, 3.48)
Self-reported sleep disturbance ^g	High/Low	< 0.0001	6.28 (3.46, 11.40)
Annoyed by snoring	High/Low	0.0693	2.16 (0.94, 4.94)
Migraines	Yes/No	0.0062	1.76 (1.17, 2.64)
Dizziness	Yes/No	0.0696	1.46 (0.97, 2.20)
Chronic pain	Yes/No	0.0754	1.47 (0.96, 2.25)
Asthma	Yes/No	0.0166	2.01 (1.14, 3.56)
Arthritis	Yes/No	0.0497	1.45 (1.00, 2.10)
Diagnosed sleep disorder	Yes/No	0.0001	2.99 (1.71, 5.23)

^aThe logistic regression is modeling the probability of having a PSQI score above 5. ^bWhere a reference group is not specified it is taken to be the last group. ^cP value significance is relative to the reference group. ^dOR (CI) odds ratio and 95% confidence interval based on logistic regression model. ^eThe exposure variable, WTN level, is treated as a continuous scale in the logistic regression model. ^fThe source identified by participants as the cause of closing bedroom windows to reduce noise levels was not road traffic, aircraft, rail or wind turbines. ^gEvaluates the magnitude of reported sleep disturbance at home from not at all to extremely for any reason over the previous year. ^hH-L P > 0.05 indicates a good fit. CI, confidence interval; H-L, Hosmer-Lemeshow test; ON, Ontario; OR, odds ratio; PEI, Prince Edward Island; PSQI, Pittsburgh Sleep Quality Index; WTN, wind turbine noise.

comes. However, there was no evidence for any form of sleep disturbance found in relation to WTN levels. Studies published to date have been inconsistent in terms of self-reported evidence that WTN disrupts sleep,^{59,60} and none of these studies assessed sleep using an objectively measured method. These inconsistent findings are not entirely surprising considering that sleep disturbance reported as a result of transportation noise exposure occurs at sound pressure levels that exceed WTN levels calculated in the current study.^{27, 28, 45} Study results concur with those of Bakker et al.,²¹ with outdoor WTN levels up to 54 dB(A), wherein it was concluded that there was no association between the levels of WTN and sleep disturbance when noise annoyance was taken into account.

The current study employed a wide range of self-reported and objectively measured endpoints related to sleep to provide a comprehensive assessment of the potential effects that WTN exposure may have on sleep. Self-reported diagnosed sleep disorders³⁷ and self-reported highly sleep disturbed for any reason

were factors found to be unrelated to WTN exposure. Furthermore, taking medication at least once per week was more commonly reported among participants living in areas where WTN levels were below 30 dB(A). Scores on the PSQI, either analyzed as a proportion above 5, or as a mean score, were also unrelated to WTN level. Actigraphy-measured sleep latency, sleep efficiency, the rate of awakening bouts, WASO and total sleep time were all found to be unrelated to WTN exposure. The results of the current study do not support conclusions that exposure to WTN up to 46 dB(A) has any statistically significant effect on self-reported or objectively measured sleep. However, annoyance with blinking lights on wind turbines (used as aircraft warning signals) may be related to reduced total sleep time. On the other hand, high annoyance toward vibrations/rattles was found to be related to *higher* total sleep time.

This study has some important limitations. Objective measures of sleep were assessed for up to 7 d, whereas the PSQI and the reported highly sleep disturbed outcomes represent time periods

Table 5—Summary of Actiwatch2® data.

		Wind Turbine Noise, dB(A)				
n (weekday, weekend)		< 25 (220, 90)	25–30 (228, 85)	30–35 (761, 297)	35–40 (1331, 497)	40–46 (632, 230)
Sleep Actigraphy Endpoint	Sleep Night	Mean (SD)	Mean (SD)	Mean (SD)	Mean (SD)	Mean (SD)
Sleep latency, min	Weekday	12.58 (19.27)	9.11 (11.96)	12.43 (21.05)	12.71 (19.55)	12.67 (19.57)
	Weekend	15.48 (20.84)	9.28 (16.29)	13.36 (25.19)	12.11 (19.34)	15.26 (22.04)
Sleep efficiency, %	Weekday	84.86 (7.20)	86.06 (7.51)	83.86 (9.45)	84.64 (8.04)	84.00 (7.69)
	Weekend	84.69 (7.34)	86.35 (7.37)	83.66 (9.40)	84.38 (8.14)	83.42 (8.07)
WASO, min	Weekday	54.88 (28.18)	49.78 (36.46)	52.49 (30.44)	50.55 (28.27)	53.32 (28.36)
	Weekend	53.31 (27.29)	49.31 (30.28)	55.26 (31.55)	53.00 (27.45)	54.47 (27.88)
Total sleep time, min	Weekday	415.08 (77.07)	408.30 (82.14)	398.75 (85.47)	401.98 (81.56)	395.99 (77.47)
	Weekend	420.22 (75.34)	416.04 (89.79)	410.87 (84.90)	411.66 (83.27)	404.80 (86.25)
Number of awakenings bouts, count	Weekday	23.36 (8.58)	22.18 (10.70)	24.30 (9.87)	23.27 (9.06)	24.16 (8.94)
	Weekend	23.31 (9.74)	22.71 (9.90)	25.44 (10.63)	24.25 (9.58)	24.90 (9.25)
Time in bed, min	Weekday	489.43 (83.04)	475.63 (89.21)	475.76 (89.10)	474.85 (85.59)	471.33 (82.03)
	Weekend	496.27 (77.68)	481.22 (91.46)	491.35 (88.68)	487.40 (86.22)	485.00 (92.45)
Rate of awakening bouts per 60 min in bed	Weekday	2.87 (0.96)	2.78 (1.16)	3.05 (1.09)	2.94 (1.01)	3.07 (0.96)
	Weekend	2.82 (1.12)	2.84 (1.15)	3.07 (1.05)	2.98 (1.06)	3.08 (1.01)

SD, standard deviation; WASO, wake after sleep onset.

Table 6—Generalized estimating equations regression models for sleep actigraphy endpoints.

	Sleep Efficiency, %	Sleep Latency, min	Total Sleep Time, min	WASO, min	Awakening Bouts, #
Full n (sleep nights)	707 (4371)	707 (4371)	707 (4371)	707 (4371)	707 (4371)
P value, unadjusted ^a	0.1327	0.3156	0.5198	0.5028	0.3738
Reduced n (sleep nights) ^b	550 (3384)	620 (3805)	620 (3825)	672 (4140)	575 (3551)
P value adjusted ^c	0.0519	0.1783	0.7348	0.3596	0.1673
Adjusted ^c WTN, dB(A)	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d
<25	83.95 (81.93, 85.97)	9.89 (6.24, 13.55)	447.90 (422.62, 473.18)	60.92 (54.24, 67.60)	23.08 (20.34, 25.83)
25–30	86.04 (84.07, 88.01)	4.44 (1.40, 7.48)	442.71 (412.79, 472.64)	58.58 (50.61, 66.56)	21.70 (18.89, 24.50)
30–35	82.78 (80.76, 84.81)	8.14 (5.30, 10.99)	438.49 (416.41, 460.57)	62.66 (57.08, 68.24)	25.23 (23.18, 27.28)
35–40	83.87 (82.15, 85.60)	8.51 (6.24, 10.79)	444.36 (423.08, 465.65)	60.81 (55.64, 65.98)	23.91 (22.07, 25.75)
40–46	83.50 (81.70, 85.30)	9.91 (7.39, 12.43)	438.46 (416.05, 460.88)	64.06 (57.81, 70.30)	25.06 (23.17, 26.95)
Unadjusted ^a WTN, dB(A)	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d	LSM (95% CI) ^d
<25	84.95 (83.53, 86.37)	12.74 (9.63, 15.86)	415.23 (399.66, 430.81)	52.61 (47.16, 58.05)	22.92 (20.82, 25.02)
25–30	86.28 (84.69, 87.87)	8.44 (6.33, 10.55)	406.59 (387.82, 425.36)	48.99 (41.52, 56.46)	21.97 (19.55, 24.38)
30–35	83.83 (82.54, 85.12)	12.20 (10.22, 14.18)	401.96 (392.10, 411.82)	53.58 (49.68, 57.48)	24.48 (23.19, 25.77)
35–40	84.67 (83.85, 85.49)	11.96 (10.50, 13.41)	404.32 (396.67, 411.96)	50.96 (48.18, 53.74)	23.29 (22.37, 24.22)
40–46	84.05 (83.03, 85.08)	12.83 (10.98, 14.67)	399.04 (389.08, 409.00)	52.89 (49.09, 56.70)	24.14 (22.95, 25.33)

^aThe base model for the multiple generalized estimating equations (GEE) regression models for all endpoints included wind turbine noise (WTN) exposure groups, province, day of the week, and the interaction between WTN groups and day of the week; ^bSample size for the adjusted GEE regression models. The sample size is reduced due to missing values of additional variables in the adjusted regression model; ^cA complete list of the other variables included in each multiple GEE regression model based on the stepwise methodology is presented in Table 3; ^dLSM, least squares means, for each group after adjusting for all other variables in the multiple GEE regression model and corresponding 95% confidence interval (CI). P values for both the adjusted and unadjusted models are based on the transformed variable in order to satisfy model assumptions of normality and constant variance.

of 30 d and 1 y, respectively. The concern is that 7 d of actigraphy may not represent long-term average sleep patterns. However, the selected time frame for actigraphy measures is typical,

supported in the literature and considered more than adequate for evaluating sleep in a nonclinical study sample.^{30,61} If there were situational factors (e.g. an ill child) that made sleep worse in the

actigraphy-assessed week, it would not be expected to bias against the effect of wind turbines on sleep, and in fact, would overstate the effect of recent situational events as compared to the long-term theoretical concern about WTN-induced sleep disturbance. As previously discussed, the analysis of actigraphy results was based on nightly average sleep patterns in relation to long-term WTN levels. Although WTN calculations would be expected to produce the highest sound pressure levels at the dwelling, they do not take into consideration the influence that night-to-night variations in outdoor WTN levels may have had on actigraphy results. Similarly, an analysis based on long-term average sound level does not fully account for transient deviations in WTN levels that could potentially interfere with sleep. An analysis based on a time-matched comparison between operational turbine data and actigraphy would permit a more refined assessment of the possible effect that night-to-night variations in WTN levels may have on sleep. These limitations extend to the fact that fluctuations in indoor sound levels during sleep remain unknown.

The possibility that wind turbine operators may have intentionally altered the output of their turbines in order to reduce potential WTN effects on sleep has been one of the concerns raised during the external peer review of this paper. When the *Community Noise and Health Study* was originally announced several months preceding data collection the study locations were unknown. Although awareness of the precise study locations would have become greater as data collection commenced, the deployment of the sleep watches took place over several months among a subsample of participants across the entire study sample. Furthermore, the reference period time for self-reported sleep disturbance was over the previous year and previous 30 d (PSQI). Finally, the subsets of sound power measurements were consistent with manufacturer-supplied data. In the authors' opinion, there is no evidence to suggest that wind turbine operators intentionally altered the output of their turbines to minimize potential effects on sleep at any point in the study.

CONCLUSIONS

The potential association between WTN levels and sleep quality was assessed over the previous 30 d using the PSQI, the previous year using percentage highly sleep disturbed, together with an assessment of diagnosed sleep disorders. These self-reported measures were considered in addition to several objective measures including total sleep time, sleep onset latency, awakenings, and sleep efficiency. In all cases, in the final analysis there was no consistent pattern observed between any of the self-reported or actigraphy-measured endpoints and WTN levels up to 46 dB(A). Given the lack of an association between WTN levels and sleep, it should be considered that the study design may not have been sensitive enough to reveal effects on sleep. However, in the current study it was demonstrated that the factors that influence sleep quality (e.g. age, body mass index, health conditions) were related to one or more self-reported and objective measures of sleep. This demonstrated sensitivity, together with the observation that there was consistency between multiple measures of self-reported sleep disturbance and among some of the self-reported and actigraphy measures, lends strength to the robustness of the conclusion that WTN levels up to 46 dB(A) had no statistically significant effect on any measure of sleep quality.

The WHO's¹¹ health-based limit for protecting against sleep disturbance is an annual average outdoor level of 40 dB(A). This level was exceeded in 18% of the cases, but by no more than 6 dB(A) and as such represents a limit to detecting a potential effect on sleep. It is therefore important to acknowledge that no inferences can be drawn from the current results to areas where WTN levels exceed 46 dB(A). Likewise, assuming a baseline prevalence of 10%, the study was designed so that the statistical power would be sufficient to detect at least a 7% difference in the prevalence of self-reported sleep disturbance. A larger sample size would be required to detect smaller differences. The statistical power of a study design is a limitation that applies to all epidemiological studies.

Although it may be tempting to generalize the current study findings to other areas, this would have required random selection of study locations from all communities living near wind turbines in Canada. Despite the fact that participants in the study were randomly selected, the locations were not and for this reason the level of confidence one has in generalizing the results to other areas can only be based on a certain level of scientific judgment regarding the level of exposure and the similarity between the current study sample and others. Despite limitations in generalizing the results of this analysis beyond the study sample, the current study is the largest and most comprehensive analysis of both self-reported and objectively measured sleep disturbance in relation to WTN levels published to date.

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