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ENTREPRENEURSHIP, HAPPINESS
AND SUPERNORMAL RETURNS:
EVIDENCE FROM BRITAIN AND THE US

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ABSTRACT

Do entrepreneurs earn supernormal returns, or does competitive pressure ensure that entrepreneurs receive the same utility level as workers? If those who run their own businesses get supernormal returns (or 'rents') they should be happier than those who work as employees. The paper tests this hypothesis. It uses survey data from Britain and the USA to show that, in comparison with those in regular forms of employment, the self-employed report significantly higher levels of utility as proxied by overall satisfaction data.

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Entrepreneurship, Happiness and Supernormal Returns: Evidence from Britain and the US

David G. Blanchflower and Andrew J. Oswald

1. Introduction

This paper studies the returns to entrepreneurship. It postulates, and tries to discriminate empirically between, two different hypotheses about the supply of entrepreneurs⁽¹⁾. The first, which is denoted *model 1*, sees the entrepreneurial labour market as akin to the market for any other kind of labour. This is the type of theory proposed by authors such as Kanbur (1982), Khilstrom and Laffont (1979) and Grossman (1984). In this form of framework there is free entry by potential entrepreneurs and, on average, they compete away any excess return to running a business. Hence the typical entrepreneur enjoys the same utility as the typical employee. By contrast, the second approach, *model 2*, views the market for entrepreneurs as working in an intrinsically different way from conventional labour markets. Its assumption is that not all workers have the vision or start-up capital to become entrepreneurs and that this allows the average entrepreneur to earn a supernormal return or 'rent'.

The paper may also be relevant to another, more complex, issue. An argument often made by politicians is that in a free-market economy the supply of entrepreneurs tends to be too low. This forms the basis for the common view that it is desirable to provide government encouragement and subsidies to those who set up in business on their own. However, there is little or no systematic evidence to suggest that a market economy can be trusted to supply approximately the correct number of secretaries and professional footballers but not to do the same for the supply of small businesses.

Some recent empirical research suggests one reason why there might be distortions to the supply of entrepreneurs. Evans and Leighton (1989) and Evans and Jovanovic (1989) for the US, and Blanchflower and Oswald (1990a, 1992) for Great Britain, argue that imperfect credit markets may constrain entrepreneurs. The US studies show that, all

else equal, people with greater family assets are more likely to switch to self-employment from employment. Evans and his collaborators draw the conclusion that capital and liquidity constraints bind, and that this holds back the effective supply of entrepreneurs. The British evidence reveals that (i) when interviewed the self-employed say they are constrained by shortage of capital and (ii) those who inherit or are given money or goods are much more likely later to run their own businesses. Blanchflower and Oswald's evidence is consistent with the view that capital-constrained individuals are prevented from leaving regular employment to form their own enterprises.

An entrepreneurial shortage could take two different forms. First, if entrepreneurial individuals create beneficial externalities, free entry into entrepreneurship would not be sufficient to generate the socially-optimal number of entrepreneurs. Without subsidies, there would be an under-supply of such people. Second, a different kind of entrepreneurial shortage could be said to exist if, in free-market equilibrium, there is some impediment to entry into the small-business sector. Economic and social efficiency requires that the expected utility of the marginal entrant to entrepreneurship be equal to the expected utility of the marginal employee. Given sufficiently large numbers of individuals, and smoothly distributed tastes within the population, the infra-marginal entrepreneur should also earn the same utility as the infra-marginal employee. This implies that average utilities could be the same in each sector.

There is a little evidence to suggest that many individuals would prefer to be self-employed. The 1989 British Social Attitude Survey asked a random sample of 780 British adults the question

"Suppose you were working and could choose between different kinds of jobs. Which of the following would you choose?"

I would choose ...

(i) Being an employee

(ii) Being self-employed

(iii) Can't choose."

A remarkable 46.9% of individuals said they would like to be self-employed (see Jowell et al 1990). This compares with an actual proportion of self-employed people of just over 10%.

The paper attempts to test whether entrepreneurs earn supernormal returns. The empirical analysis uses survey data -- from two nationally representative samples -- on the overall utility from work as measured by reported levels of job satisfaction. Section 2 of the paper contrasts two theoretical views of entrepreneurship. Section 3 discusses British and American data on the satisfaction levels of the employed and self-employed (such data have apparently not previously been used by economists). The econometric findings from these data sets are given in Section 4. Section 5 states the paper's conclusions.

2. Modelling the Supply of Entrepreneurs

Consider the following two approaches to the economics of entrepreneurship. In model 1, entrepreneurs and employees get the same utility and there are no rents. In model 2, those who run their own businesses earn rents.

Model 1: Free entry and no entrepreneurial rents

This model is just a version of the familiar competitive theory of the labour market, and applications to the entrepreneurial labour market are contained in sources such as Kanbur (1982), so no formal version of it is presented here. According to the theory, individuals are able to enter whichever occupational labour market they wish, and running an enterprise comprises one of those occupational markets. Individual i weighs up his or her talents and interests and chooses a career path. If entrepreneurship gives higher utility than being a regular worker, individual i will be more prone to choose to enter that line of work. Thus he or she will be more likely to set up in business. Under plausible assumptions, entry into the entrepreneurial labour market drives down the marginal product of labour there, and all individuals receive an income equal to the marginal entrant's marginal product of labour. Given free entry, long-run equilibrium has the characteristic

that the utility of entrepreneurs has been pushed down until it equals the utility of employees.

Model 2: Supply constraints and entrepreneurial rents

This model captures a rather different way of viewing the supply of entrepreneurs. It is one in which, because some individuals have neither entrepreneurial vision nor the necessary start-up capital, there can be a shortage of small businesses that allows entrepreneurs to earn supernormal returns.

Assume, following Knight (1921) and others, that entrepreneurial opportunities cannot be assigned probabilities. Lack of knowledge, rather than risk, therefore plays a role in the analysis. Entrepreneurs are also assumed to be constrained by the amount of capital they own. This can be justified by assuming that when bankers cannot make probabilistic judgements about new business ventures they optimize by offering loans backed by collateral. The potential entrepreneur's assets then constrain his or her ability to set up in business.

This approach makes genuine uncertainty a central feature of the analysis. By contrast, the recent work by Kanbur (1979), Khilstrom and Laffont (1981) and Grossman (1984) breaks with the principal tenets of earlier thought on entrepreneurial activity. Kanbur *et al* develop an essentially standard neoclassical approach in which

- (i) productive business opportunities are ex ante feasible for, and visible to, all individuals (most simply choose not to exploit them),
- (ii) there is an objective probability distribution governing business risk, and everyone knows that distribution,
- (iii) entrepreneurs receive the same expected utility as their workers,
- (iv) the entrepreneur is likely to be someone with unusually low risk-aversion (see especially Khilstrom and Laffont, 1981).

These are different from the main assumptions and arguments of classic sources such as Schumpeter (1939), Knight (1921) and Kirzner (1973). In contrast to modern

theory, based upon the four points above, the classic writings about the nature of the entrepreneur stressed the following:

- (i) most individuals are not sufficiently alert or innovative to perceive business opportunities,
- (ii) there is no objective probability distribution governing business risks,
- (iii) an innovative entrepreneur may receive higher expected utility than he or she would as a regular worker,
- (iv) attitude to risk is not the central characteristic which determines who becomes an entrepreneur.

The model of this paper draws upon the older, but recently neglected, current of thought. Six assumptions are made.

A.1 Assume that proportion β of the population have entrepreneurial vision. This group of individuals can see business opportunities where proportion $1-\beta$ see none.

A.2 There is, in the economy, an array of viable entrepreneurial projects. These projects are indexed by e on the unit interval. The profit from project e is $\pi(e)$. This function describes the returns from the different entrepreneurial ventures in the economy. It is assumed to be strictly decreasing, so that $\pi(0)$ is the most profitable project and $\pi(1)$ the least profitable.

A.3 Each project requires capital k . Only Z individuals in the economy have k or more capital.

A.4 Individuals receive utility:

$$u = \pi + i \quad \text{in self-employment}$$

$$u = w \quad \text{in conventional employment,}$$

where w is the wage paid for non-entrepreneurial work, π is entrepreneurial profit, and i is the utility from being independent and one's own boss.

A.5 Anyone can find work at wage w in the non-entrepreneurial part of the economy. Output there is produced according to a concave production function $f(\cdot)$ defined on

employment. It is assumed that w equals the marginal product of labour.

A.6 The working population (divided between entrepreneurs and regular workers) is fixed at P . In equilibrium the number of entrepreneurs is e^* and the number of workers is $P - e^*$.

These assumptions lead to a model where, to make the key points as simply as possible, probabilistic business risk has been assumed away. The model is one in which, by their nature, some potential entrepreneurs may be capital-constrained.

People enter entrepreneurship until capital shortages hold them back or until the rewards from being an entrepreneur are too low. Therefore, in equilibrium, either (i) capital or vision constraints are binding in aggregate or (ii) the utility from running a business has been driven down to equal that from wage-work. In the former case

$$w < \pi(e^*) + i$$

and in the latter,

$$w = \pi(e^*) + i.$$

Here e^* is the marginal entrepreneurial venture. At this point all business projects with higher profitability are already being undertaken.

It can now be shown that entrepreneurs are better off than regular workers, and that the mean gap in utility between the two kinds of work is higher if there are fewer numbers of people with capital.

Proposition 1 *Entrepreneurs get higher utility than regular workers.*

Proof

In equilibrium in this economy it cannot be the case that

$$\pi(e^*) + i < w,$$

because otherwise entrepreneurs would leave for the wage-sector paying w . Thus either marginal entrepreneurial utility $\pi(e^*) + i$ is equal to w or, because people are held back by

capital constraints, it exceeds it. As $\pi(e)$ is a decreasing function, all other entrepreneurs earn higher profit than the one operating the marginal project. Hence, all but the marginal entrepreneur receives strictly more utility than regular workers, and the marginal entrepreneur gets weakly greater utility than regular workers.

If capital constraints bind, so that people are prevented from setting up the marginal business even though they would get higher utility than working in the wage-sector, the following result holds.

Proposition 2. *When capital constraints bind, the larger is Z , the number of people in the economy who have capital, the smaller is the utility gap between entrepreneurs and workers.*

Proof The sum of entrepreneurs' utilities is given by

$$\int_0^{e^*} (\pi(e) + i) de$$

and average entrepreneurial utility by

$$\frac{\int_0^{e^*} (\pi(e) + i) de}{\beta Z}$$

Each worker in the wage sector gets utility equal to the wage w . There are $P - \beta Z$ individuals working in that sector. This is because the supply of entrepreneurs is constrained to be the product of β (those with entrepreneurial vision) and Z (those with capital).

Assuming that the equilibrium is one where there is an aggregate shortage of individuals with capital, free-entry does not eliminate the difference in returns to the marginal entrepreneur between the wage sector and the entrepreneurial sector. Let the

average utility gap between the entrepreneurial sector and the wage sector be denoted v . It is given algebraically by

$$v = \frac{\int_0^{\beta Z} [\pi(e) + i] de}{\beta Z} - f(P - \beta Z)$$

where the latter term is the marginal product of labour in the wage sector. A rise in Z , the total number of individuals with sufficient capital to run their own business, increases the numbers setting up enterprises. This drives down the marginal entrepreneurial return and, by inducing workers to leave the wage-sector, raises the marginal product of labour there. Hence the utility difference, v , changes by:

$$\frac{\partial v}{\partial Z} = \frac{1}{Z} [\pi(e^*) + i] - \frac{1}{\beta Z^2} \int_0^{\beta Z} [\pi(e) + i] de + \beta f'(P - \beta Z)$$

The third of these three terms is unambiguously negative, by the concavity of the production function, so to establish the Proposition it is sufficient to show that the first two terms sum to a negative number. Informally this can be seen from the fact that the sum of these two terms equals one over Z multiplied by the difference between the marginal entrepreneur's return and the average entrepreneur's return.

More formally, by the First Mean Value Theorem, there exists a point μ , in the interval from zero to βZ , such that

$$\beta Z [\pi(\mu) + i] = \int_0^{\beta Z} [\pi(e) + i] de$$

which (noting that profits at μ exceed profits at e^*) when substituted into the previous expression shows that the utility gap, v , is a decreasing function of the number of

individuals with capital, Z . This establishes the proposition.

These two models form the basis for the empirical testing carried out in Section 4. The analysis focuses on self-employed people as the empirical equivalent of those in the models who run their own enterprises. It is necessary for the test to compare utility levels between work and entrepreneurial activity. The reported earnings of self-employed individuals are known to be unreliable, and it is likely that such individuals get a non-pecuniary benefit from being their own boss, so some proxy for utility is required. This paper follows the psychology literature in using survey data on overall job satisfaction. It is established there (see, for example, Argyle 1988 and Warr 1985) that these reported satisfaction numbers are highly correlated with observable measures of individual well-being such as quitting behaviour and physiological symptoms. The small economics literature includes Hamermesh (1977), Borjas (1979), Freeman (1978) and Clark and Oswald (1992), but has not studied the self-employed.

3. NCDS and GSS Data

The *National Child Development Study* is a longitudinal birth cohort which takes as its subjects all those living in Great Britain who were born between the 3rd and 9th March 1958. The survey has been sponsored by five UK Government Departments - the Departments of Health and Social Security (DHSS), Education and Science (DES), Employment (DE), Environment (DOE) and the Manpower Services Commission (MSC). Since the original Peri-natal Mortality Study was undertaken in 1958, major surveys were also carried out in 1965 (NCDS1), 1969 (NCDS2), 1974 (NCDS3), 1981 (NCDS4) and most recently in 1991 (NCDS5) (2). For the purposes of the first three surveys, the birth cohort was augmented by including those new immigrants born in the relevant week, and information was obtained separately from parents, teachers, and doctors, as well as member of the NCDS cohort. The 1981 survey differs in that no attempt was made to include new immigrants since 1974 and information was obtained only from the subject.

In this paper we make use only of data from NCDS4, which was conducted when all the respondents were aged twenty-three.

The 1981 survey contained a total of 12,537 interviews, namely, approximately 76 percent of the original target sample and 93 percent of those traced and contacted by interviewers. The interview survey was carried out between August 1981 and March 1982. For further details of the surveys, see Elias and Blanchflower (1988, 1989).

In 1981, when the respondents were twenty-three years of age, 9.3% of the sample were unemployed, 14.2% were out-of-the-labor-force and 2.4% were in full-time education. Out of 9178 individuals who were employed at the time of interview, 8657 worked as employees while a further 521 were self-employed on a full-time basis⁽³⁾. Hence, one in fifteen young people who were working at the time of interview had a job which they had created themselves. However, in 138 of these cases the job was in a family firm. The self-employed are predominantly men (79% of the self-employed are male compared to 55% for employees).

The empirical analysis that follows is restricted to individuals who were either employed or self-employed in their main activity at the time of interview in 1981. This gives an initial sample size of 9176 observations, of whom 519 individuals (5.7%) were self-employed in their main job.

The *General Social Surveys* are administered by the National Opinion Research Center at the University of Chicago. Surveys have been conducted in the US during February, March and April of 1972-8, 1980 and 1982-1990⁽⁴⁾. There are a total of 26,265 completed interviews. The median length of interview was about one and a half hours. Each survey is an independently drawn sample of English speaking persons 18 years of age or over, living in non-institutional arrangements within the United States. Block quota sampling was used in the 1972-4 surveys and for half of the 1975 and 1976 surveys. Full probability sampling was used in the second halves of the 1975 and 1976 surveys as well as in the 1977, 1978, 1980 and 1982-1990 surveys. In 1982 and 1987 an

oversampling of blacks was conducted (354 and 353 extra cases respectively). For further details of the sample design, see General Social Surveys, 1972-1990: Cumulative Codebook, Appendix A.

In the United States over the years 1972-1990, an average of 13.3% of workers in the GSS sample were self-employed at the time of interview. This is a considerably higher rate than we reported above in our sample of young people in Great Britain. In large part, however, this is because the GSS sample includes the whole age range. As can be seen below the incidence of self-employment rises strongly with age. The self-employment rate (the proportion of workers that are self-employed) among those under 24 in the US is roughly in line with that reported in the NCDS (4.8% and 5.7% respectively).

Age group	% self-employed	No. of observations
16-19	4.6%	246
20-24	4.8%	1690
25-34	10.6%	4527
35-44	13.8%	3685
45-54	16.2%	2706
55-64	17.3%	1924
65-74	30.0%	498
≥75	42.7%	93

The self-employed in the US are also predominantly male -- 65.9% of the self-employed compared with 52.9% of employees

The empirical analysis that follows is restricted to individuals who were either employed or self-employed at the time of interview. This gives a sample size of 15271 observations, of whom 2025 individuals (13.3%) reported that they were self-employed as their main activity⁽⁵⁾.

Respondents in both our data files were asked how satisfied they were with their current jobs. In Great Britain the following question was asked:

"Taking everything into consideration, how satisfied or dissatisfied are you with your job as a whole" (Q19j, p.9: NCDS4 questionnaire)

The responses to this question were coded into 5 categories and the results are reported in part A of Table 1.

In the US a slightly different question was asked:

"On the whole how satisfied are you with the work you do - would you say you are very satisfied, moderately satisfied, a little dissatisfied or very dissatisfied" (Q180, p.233, GSS codebook)

Responses to this question are reported in part B of Table 1. In both countries the vast majority of people reported that they were satisfied, very satisfied or moderately satisfied with their present job (80% in Great Britain and 87% in the US).

4. Statistical Results

It is useful to begin by examining the patterns in the raw data. Table 1 provides cross-tabulations which bear on the hypothesis that the entrepreneurs get higher utility than workers. Part A gives the British means and part B the American means.

Consistent with Proposition 1 outlined in Section 2 above, the self-employed in both countries report that they are more satisfied than is the case for regular workers. The differences between the two groups are large. For example, 46.2% of the British self-employed say that they are 'very satisfied' (the top category of a five-point satisfaction scale), whereas 29.1% of the employed give this response. In the US, 63.3% of those who are self-employed answer this on a four point scale, whereas only 47.3% of employees do so.

The next step is to see whether this effect continues to hold after a set of personal control variables are introduced. This section uses the information on job satisfaction reported in Table 1 to estimate ordered probits for both countries. Ordered probits are the appropriate statistical procedure where, as in this case, respondents express their preferences in the form of an ordinal ranking⁽⁶⁾. The ordered probit is based on the following specification:

$$\begin{aligned} z &= \beta'x + \varepsilon \\ \varepsilon &\sim N[0, 1] \\ y &= 0 \text{ if } z \leq \mu_0 \\ &1 \text{ if } \mu_0 < z \leq \mu_1 \end{aligned}$$

$$\begin{aligned}
 & 2 \text{ if } \mu_1 < z \leq \mu_2 \\
 & \dots \\
 & J \text{ if } z > \mu_{J-1}
 \end{aligned}$$

where z is a latent (unobserved) indicator of job satisfaction, y is the reported categorical indicator, and e is a random disturbance. Since the scale of z is not observed, we employ the usual probit normalization and set the variance of z to unity. The μ 's are unobserved thresholds to be estimated; different values of y are realized as the latent indicator z crosses these thresholds. The dependent variable is coded 1, 2, ..., J .

There is little published econometric work based upon satisfaction data. Hamermesh (1977) and the small literature cited earlier find some influence from conventional microeconomic control variables, such as gender and union status, and these kinds of variables are included in the later equations. Previous work has not looked at the satisfaction levels of the self-employed.

The central issue is whether, *ceteris paribus*, the self-employed report higher levels of satisfaction than do employees. Because work satisfaction will be influenced by wage income, some stance must be taken on whether or not a wage measure is to be included in the probit equations. The theory of Section 2 makes clear that the appropriate test is to omit income variables. This is because the focus of interest is the total utility of individuals -- to allow a comparison of the utility 'package' associated with each kind of work -- and not just the satisfaction level with income held constant. A pragmatic further reason is that the earnings levels reported by the self-employed are unreliable.

British results for NCDS data are given in Tables 2 and 3. US results for GSS are provided in Table 4. Included as controls in the British equations are dummy or continuous variables for self-employment, union membership, marital status, gender, disabled status, region, highest educational qualification, part-time, ever unemployed in the previous 5 years, a dummy for problems with arithmetic, months of experience, and job tenure. There is no age variable because all NCDS individuals are the same age (23 in this year of the sample). Month-of-interview dummies are included. A set of industry

dummies are also included in Table 2, and for illustrative purposes are omitted in Table 3. Results are also given, in Tables 2 and 3, for the sub-samples of people who did and did not inherit money. These form the second and third columns of the Tables.

The US specifications of Table 4 were chosen to be as close as possible to the British ones. Included as controls in these American equations are dummy or continuous variables for self-employment, union membership, ethnic status, age, marital status, gender, disabled status, region, years of schooling, ever unemployed in the previous 5 years, months of experience, and job tenure. Year dummies are also included in the specification. A set of industry dummies are entered in the specification in the left hand column of Table 4.

It is apparent that the self-employment dummy variable is strongly significant and positive. This is true for both Britain and the USA. The most comparable general equations for the two countries are the first column of Table 2 and the first column of Table 4. The t-statistics are, respectively, 4.9 and 9.0. Because these equations are ordered probits, the two countries' coefficients' sizes cannot be compared without adjusting for the different distributions, but it is clear from the Tables that the coefficient on the self-employment dummy is much larger than those on personal variables such as gender. Consistent with the cross-tabulations presented in Table 1, self-employment has a major effect on reported satisfaction levels. As the equations exclude income measures, the self-employment variables are not capturing merely the non-monetary return to being one's own boss.

Given the paucity of work with data like these, the other controls are also of interest. Women are more satisfied; blacks (in the US) are less satisfied; married people are more satisfied. Those who have been unemployed are less satisfied. Union members are also less content: this replicates the main finding of the earlier literature of Borjas (1979) and Freeman (1978). Older people (in the US) say they are more content, and years of schooling and being part-time are also positively associated with satisfaction. There are

large differences in satisfaction across different industries, but not across different regions.

As an experiment into the effects of access to capital, and guided by the finding in Blanchflower and Oswald (1990a, 1992) that an inheritance increases the likelihood of self-employment, the British data were split into two sub-samples. The second columns of Tables 2 and 3 are estimated with data on the 6887 people who reported themselves as having received no inheritance or gift of money or goods exceeding £500. The third column of each of Tables 2 and 3 gives estimates for the sub-sample of 987 people who had received this kind of inheritance or gift. There is some evidence that the self-employment dummy variable has a smaller effect in the group who inherited: in Table 2 the dummy even goes negative. Such evidence might be taken to be consistent with the idea that those with capital -- through an inheritance -- are more able to enter the self-employment sector and drive down the rents available there. This argument can only be suggestive, but indicates an area where further research may be fruitful.

Although there is some evidence here that entrepreneurs get higher utility than conventional employees, one particular caveat should be borne in mind in interpreting the paper's findings. The use of satisfaction data to proxy utility levels is unconventional in economics research. It may be that reported satisfaction levels are subject to important biases, and economists have much to learn about how correctly to interpret these kinds of survey responses. Directly reported measures of satisfaction have a variety of possible applications in economics. This paper should be seen as an exploratory attempt to exploit the potential of such data.

5. Conclusions

This paper examines the hypothesis that entrepreneurs receive supernormal returns (or 'rents'). It uses a cross-section sample of approximately eight thousand Britons in 1981 and a series of cross-section samples covering approximately fifteen thousand Americans over the years 1972 to 1990.

The paper considers two alternative theories of entrepreneurial activity. According to model 1, the market for entrepreneurs behaves like the economist's traditional view of a competitive labour market. Free entry eliminates the difference in utility between those who run businesses and those who work in regular employment. According to model 2, by contrast, there exist shortages of entrepreneurial vision and of capital, and entrepreneurs get rents. To test for this utility premium the paper uses survey data on employed and self-employed individuals' reported levels of overall satisfaction with their jobs. Such data are rarely used by economists but are the foundation for a large psychology literature and have been shown in that literature to be strongly correlated with observable measures of physiological and mental well-being. The analysis treats these as proxy utility data.

The paper's main conclusion is that, controlling for other variables, the self-employed report significantly higher levels of well-being than employees. This is true both for the US and Great Britain, and the size of the effect appears to be substantial. The data therefore accord with the second of the models discussed in Section 2, which predicts that an entrepreneur will enjoy higher utility than a worker. One of the paper's indirect, and necessarily more tentative, aims has been to consider the hypothesis that a market economy fails to produce a sufficient supply of entrepreneurs. This is a complex issue and a proper test of the hypothesis would require data on the marginal individual: economic efficiency and social optimality demand that that person should be indifferent between going into business and being a worker. The paper's empirical analysis can look only at the average individual, but suggests that those who run their own businesses feel happier than employees. According to these results the typical entrepreneur does earn rents.

Endnotes

1. The present paper contributes to the quickly-expanding economics literature on entrepreneurship and self-employment. Fuchs (1982) and Rees and Shah (1986) were early econometric contributions. The new literature includes Evans and Leighton (1989), Evans and Jovanovic (1989), Holmes and Schmitz (1990, 1991), Borjas and Bronars (1989), Blanchflower and Oswald (1990a, 1990b, 1992), Blanchflower and Meyer (1991), Casson (1990), Lentz and Laband (1990), Meyer (1990), Reid and Jacobsen (1988), and Reid (1990).
2. At the time of writing the NCDS 5 data seem likely to be available in 1993.
3. In addition, 94 individuals worked as employees in their first job but were self-employed part-time. These are included here in the employees' sample.
4. There were no surveys in either 1979 or 1981.
5. This rate is higher than the official rate of 8.7% reported in the Monthly Labor Review, which uses a more restrictive definition of self-employment.
6. The ordered probit model is discussed in Zavoina and McElvey (1975).

Table 1. Job Satisfaction in Great Britain and the US (%).**A) Great Britain**

	Self-employed	Employees	All workers
Very dissatisfied	1.7	2.8	2.8
Dissatisfied	2.9	9.6	9.2
Neither	6.7	8.2	8.1
Satisfied	42.4	50.2	49.8
Very satisfied	46.2	29.1	30.1
N	519	8657	9176

Base: individuals in employment at the time of interview
 Source: National Child Development Study, 1981.

b) United States

	Self-employed	Employees	All workers
Very dissatisfied	2.0	3.9	3.7
A little dissatisfied	5.5	10.3	9.6
Moderately satisfied	29.2	38.5	37.3
Very satisfied	63.3	47.3	49.4
No. of observations	2025	13246	15271

Base: individuals in employment at the time of interview
 Source: General Social Surveys: 1972-1990

Table 2. Job Satisfaction in Britain 1981: Ordered Probit (with industry controls)

Variable	All		No Inheritance		Inheritance	
	Coefficient	T-ratio	Coefficient	T-ratio	Coefficient	T-ratio
Self-employed	0.4235	4.930	0.4911	5.278	-0.0266	0.106
Female	0.1156	3.958	0.1311	4.189	0.0417	0.470
Disabled	-0.1062	0.061	-0.0034	0.019	-0.2521	0.165
Number problems	-0.1442	2.483	-0.1449	2.380	-0.1663	0.795
Married	0.0694	2.580	0.0786	2.711	-0.0074	0.092
Divorced	-0.0468	0.367	-0.0831	0.635	0.6753	0.942
Separated	0.0778	0.906	0.1085	1.178	-0.4742	1.808
South East	-0.0229	0.513	-0.0230	0.471	-0.0293	0.247
South West	0.0449	0.784	0.0544	0.878	-0.0557	0.355
Wales	0.0294	0.444	0.0385	0.543	0.0160	0.078
West Midlands	0.0538	1.023	0.0429	0.768	0.0965	0.544
East Midlands	0.0890	1.553	0.0847	1.360	0.1084	0.665
East Anglia	-0.0437	0.595	-0.0456	0.571	-0.0394	0.197
Yorshire & Humberside	0.0407	0.746	0.0371	0.630	0.0880	0.548
North West	0.0205	0.404	-0.0398	0.073	0.2261	1.544
North	0.0243	0.384	0.0503	0.742	-0.1632	0.892
Scotland	0.0513	0.899	0.0465	0.758	0.0630	0.383
Higher degree	-0.2766	0.884	-0.4175	1.177	0.4933	0.491
Degree	-0.1277	1.695	-0.0944	1.124	-0.2173	1.123
Teacher	0.3149	1.941	0.3220	1.855	0.2968	0.606
HNC/HND	-0.0370	0.605	-0.0390	0.594	-0.0597	0.325
Nurse	0.2474	2.824	0.2211	2.331	0.4370	1.825
2+ A's	-0.1020	1.480	-0.1003	1.322	-0.0755	0.419
A or ONC or TEC	-0.0046	0.094	-0.0094	0.178	0.0880	0.598
5+ O's or Craft	-0.0343	0.768	-0.0345	0.730	-0.0286	0.193
O's +	0.0760	1.008	0.0159	0.197	0.5443	2.188
O's Only	-0.0081	0.196	-0.0174	0.399	0.0609	0.440
No O's - something	0.0108	0.114	-0.0442	0.453	0.5992	1.360
Other qualifications	0.1557	1.596	0.1863	1.833	-0.2432	0.646
Part-time	0.1290	2.095	0.0948	1.445	0.4536	2.382
Union member	-0.0484	1.744	-0.0498	1.681	0.0177	0.204
Ever unemployed	-0.1938	7.076	-0.1777	6.073	-0.2959	3.563
Energy & Water Supply	-0.2269	1.801	-0.2245	1.704	-0.3218	0.677
Other Mineral & Ore Extraction	-0.4276	3.450	-0.4347	3.330	-0.4232	0.985
Metal Goods, Eng. & Vehicles	-0.4783	4.239	-0.4885	4.123	-0.4291	1.072
Other Manufacturing	-0.4119	3.641	-0.4197	3.532	-0.4087	1.023
Construction	-0.2232	1.896	-0.2531	2.045	-0.0162	0.039
Distribution, Hotels & Catering	-0.3496	3.142	-0.3707	3.170	-0.1767	0.444
Transport & Communication	-0.1236	1.043	-0.1543	1.235	0.1234	0.298
Banking, Finance & Insurance	-0.2095	1.830	-0.2272	1.878	-0.1418	0.356
Other Services	-0.0850	0.765	-0.0531	0.484	-0.2855	0.723
Month of Interview - August	0.0898	0.879	-0.0536	0.484	0.4749	1.699
Month of Interview - October	-0.0573	2.008	-0.0666	2.176	0.0418	0.494
Month of Interview - November	-0.0545	1.424	-0.0636	1.553	0.0180	0.157
Month of Interview - December	-0.1407	2.045	-0.1586	2.127	-0.0409	0.208
Completed Apprenticeship	-0.0236	0.556	-0.0143	0.316	-0.0904	0.673

Experience (months)	0.0008	0.776	0.0008	1.501	-0.0032	1.207
Tenure in current job (months)	-0.0001	0.016	-0.0001	0.262	-0.0007	0.425
Constant	2.1924	15.007	2.1152	13.541	2.5307	5.281
Threshold (1)	0.7748	28.343	0.7709	26.437	0.8324	9.272
Threshold (2)	1.1207	38.541	1.1160	35.929	1.1971	12.656
Threshold (3)	2.5344	78.574	2.5393	73.538	2.5892	25.303
Log Likelihood	-9536.3		-8323.8		-1184.9	
Restricted Log L.	-9717.9		-8497.0		-1219.9	
Chi Squared (49)	363.27		346.35		70.048	
N	7874		6887		987	

Excluded - Greater London, Agriculture, September

Table 3. Job Satisfaction in Britain 1981: Ordered Probit (no industry controls)

Variable	All		No Inheritance		Inheritance	
	Coefficient	T-ratio	Coefficient	T-ratio	Coefficient	T-ratio
Self-employed	0.4885	9.311	0.5219	9.026	0.2915	2.168
Female	0.1353	5.105	0.1482	4.189	0.0681	0.876
Disabled	-0.0300	0.179	-0.0232	0.138	-0.2786	0.176
Number problems	-0.1247	2.256	-0.1308	2.239	-0.0548	0.300
Married	0.0712	2.844	0.0760	2.801	0.0211	0.290
Divorced	-0.0108	0.089	-0.0823	0.645	0.9522	1.819
Separated	0.0168	0.209	0.0389	0.456	-0.2843	1.148
South East	-0.0331	0.800	-0.0349	0.774	-0.0337	0.306
South West	0.0743	1.408	0.0986	1.727	-0.1008	0.703
Wales	0.0602	0.976	0.0547	0.825	0.1327	0.720
West Midlands	0.0067	0.137	-0.0003	0.005	0.0518	0.328
East Midlands	0.0502	0.929	0.0532	0.904	0.0475	0.324
East Anglia	-0.0183	0.265	-0.0047	0.059	-0.1021	0.545
Yorkshire & Humberside	0.0214	0.421	0.0147	0.268	0.0940	0.638
North West	0.0019	0.040	-0.0247	0.484	0.1964	1.420
North	-0.0083	0.137	0.0203	0.314	-0.1911	1.133
Scotland	0.0284	0.535	0.0262	0.460	0.0034	0.022
Higher degree	-0.0261	0.096	-0.2875	0.823	0.9570	1.278
Degree	-0.0163	0.234	0.0069	0.089	-0.0687	0.373
Teacher	0.4569	2.955	0.4679	2.838	0.4298	0.917
HNC/HND	0.0503	0.882	0.0390	0.636	0.1207	0.722
Nurse	0.3918	4.847	0.3904	4.491	0.4190	1.827
2+ A's	-0.0282	0.443	-0.0209	0.301	-0.0332	0.194
A or ONC or TEC	0.0446	0.994	0.0443	0.917	0.1010	0.744
5+ O's or Craft	0.0236	0.571	0.0130	0.299	0.0941	0.691
O's +	0.1253	1.800	0.0935	1.249	0.3789	1.822
O's Only	0.0284	0.748	0.0258	0.645	0.0782	0.620
No O's - something	0.0381	0.428	0.0150	0.165	0.3302	0.807
Other qualifications	0.1240	1.359	0.1818	1.871	-0.3721	1.232
Part-time	0.1266	2.346	0.1226	2.098	0.2163	1.372
Union member	-0.0268	1.080	-0.0160	0.604	-0.0882	1.158
Ever unemployed	-0.1960	7.700	-0.1752	6.445	-0.3351	4.417
Month of Interview - August	0.0460	0.459	0.0055	0.050	0.3933	1.352
Month of Interview - October	-0.0446	1.671	-0.0491	1.714	0.0147	0.189
Month of Interview - November	-0.0389	1.099	-0.0330	0.870	0.0584	0.562
Month of Interview - December	-0.070	1.116	-0.0637	0.938	0.1146	0.633
Completed Apprenticeship	-0.0684	1.765	-0.0587	1.421	-0.1294	1.088
Experience (months)	0.0006	0.609	0.0011	1.091	-0.0009	0.383
Tenure in current job (months)	-0.0000	0.058	0.0002	0.332	-0.0011	0.705
Constant	1.8694	20.218	1.801	17.908	2.1247	8.125
Threshold (1)	0.7517	29.762	0.7508	27.820	0.7723	9.868
Threshold (2)	1.0968	40.781	1.0942	38.042	1.1365	13.664
Threshold (3)	2.4857	83.402	2.4931	78.025	2.4866	27.683
Log Likelihood	-10909		-9519.8		-1368.4	
Restricted Log L.	-11055		-9648.8		-1403.2	
Chi Squared (49)	291.07		258.17		69.654	
N	8973		7837		1136	

Table 4. Job Satisfaction in the US 1972-1990: Ordered Probit

Variable	Coefficient	T-ratio	Coefficient	T-ratio
Self-employed	0.2798	9.013	0.2754	9.392
Female	0.0509	2.401	0.0565	2.901
Union member	-0.1261	4.370	-0.1151	4.114
Black	-0.2334	7.993	-0.2158	7.467
Other non-white	-0.1259	1.896	-0.1252	1.894
Age	0.0146	17.289	0.0153	18.386
Married	0.1362	5.184	0.1398	5.363
Widowed	0.0510	0.868	0.0465	0.795
Divorced	0.0637	1.740	0.0738	2.029
Separated	0.2268	4.055	0.2308	4.142
Ever unemployed last 5 years	-0.1792	5.964	-0.1860	6.061
1975 year dummy	0.1576	3.320	0.1508	3.202
1976 year dummy	0.0885	1.869	0.0771	1.633
1977 year dummy	0.0305	0.647	0.0305	0.647
1978 year dummy	0.1529	3.292	0.1376	2.975
1980 year dummy	-0.0211	0.454	-0.0362	0.780
1982 year dummy	-0.0439	1.024	-0.0414	0.968
1983 year dummy	0.1503	3.266	0.1436	3.154
1984 year dummy	-0.0655	1.486	-0.0751	1.711
1985 year dummy	-0.0256	0.569	-0.0274	0.611
1986 year dummy	0.1054	2.206	0.1022	2.141
1987 year dummy	-0.0151	0.353	-0.0224	0.528
1988 year dummy	0.0325	0.703	0.0204	0.444
1989 year dummy	-0.0072	0.163	-0.0058	0.130
1990 year dummy	-0.0049	0.107	-0.0090	0.197
Middle Atlantic	-0.0710	1.506	-0.0743	1.587
East North Central	-0.0062	0.136	-0.0180	0.394
West North Central	0.0169	0.316	0.0123	0.232
South Atlantic	0.0767	1.633	0.0770	1.653
East South Central	0.1161	2.053	0.0982	1.750
West South Central	0.0645	1.237	0.0673	1.298
Mountain	-0.0574	0.987	-0.0522	0.903
Pacific	-0.0542	1.111	-0.0488	1.006
Construction	0.0861	1.718		
Lumber & Wood	-0.0937	0.720		
Furniture & Fixtures	-0.2136	1.637		
Stone, Clay & Glass	-0.1317	1.050		
Primary Metals	-0.2603	2.954		
Fabricated Metals	-0.1660	1.903		
Not Specified Metal Industries	-0.1190	1.818		
Machinery Excluding Electricals	-0.1627	1.989		
Electrical Machinery & Equipment	0.0001	0.001		
Professional & Photographic Equipmt.	0.0563	0.552		
Food and Kindred Products	-0.2342	3.008		
Miscellaneous Manufacture	-0.0381	0.358		
Textiles	-0.1442	1.284		
Apparel	-0.4121	4.924		
Paper & Allied Products	-0.0647	0.521		
Printing, Publishing etc.	0.0002	0.003		

Chemicals and Petroleum	-0.2048	2.728		
Rubber	-0.2845	2.494		
Leather	-0.0258	0.146		
Transportation	0.0035	0.057		
Communications	0.0996	1.646		
Wholesale Trade	-0.0449	0.730		
Retail Trade	-0.2105	5.388		
Finance, Insurance & Banking	0.0018	0.037		
Other Professional Services	0.0978	2.672		
Business Services, Repairs etc.	-0.0643	1.513		
Agriculture	-0.0356	0.530		
Years of Schooling	0.0144	4.017	0.0246	7.529
Constant	1.0797	12.836	0.8728	111.560
Threshold (1)	0.7121	40.407	0.7065	40.687
Threshold (2)	1.8972	91.927	1.8834	92.430
Log-Likelihood	-15599.	15681.		
Restricted (Slopes=0) Log-L.	-16185.	16185.		
Chi-Squared (62)	1171.0	1007.5		
N	15190	15190		

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