

Estimates for distribution of suprema of solutions to higher-order partial differential equations with random initial conditions

Yuriy Kozachenko^a, Enzo Orsingher^b, Lyudmyla Sakhno^a, Olga Vasylyk^{a,*}

^a*Department of Probability Theory, Statistics and Actuarial Mathematics,
Taras Shevchenko National University of Kyiv, 64 Volodymyrska str.,
01601, Kyiv, Ukraine*

^b*Department of Statistical Sciences, Sapienza University of Rome,
P.le Aldo Moro, 5, 00185, Rome, Italy*

ykoz@ukr.net (Yu. Kozachenko), enzo.orsingher@uniroma1.it (E. Orsingher),
lms@univ.kiev.ua (L. Sakhno), ovasylyk@univ.kiev.ua (O. Vasylyk)

Received: 25 July 2019, Revised: 10 October 2019, Accepted: 16 November 2019,
Published online: 17 December 2019

Abstract In the paper we consider higher-order partial differential equations from the class of linear dispersive equations. We investigate solutions to these equations subject to random initial conditions given by harmonizable φ -sub-Gaussian processes. The main results are the bounds for the distributions of the suprema for solutions. We present the examples of processes for which the assumptions of the general result are verified and bounds are written in the explicit form. The main result is also specified for the case of Gaussian initial condition.

Keywords Higher-order dispersive equations, random initial conditions, harmonizable processes, sub-Gaussian processes, distribution of supremum of solution, entropy methods

2010 MSC 35G10, 35R60, 60G20, 60G60

*Corresponding author.

1 Introduction

Numerous recent studies are concerned with evolution equations of the form

$$\frac{\partial u}{\partial t} + \sum_{j=1}^l \frac{\partial^{2j+1} u}{\partial x^{2j+1}} + u^k \frac{\partial u}{\partial x} = 0, \quad l, k \in \mathbb{N},$$

which are dispersive equations of order $2l + 1$ with a convective term $u^k \frac{\partial u}{\partial x}$; equations with coefficients of more general form and different kinds of nonlinearity are also the subject of active research.

The most celebrated equation of this class is the Korteweg–De Vries (KdV) equation

$$\frac{\partial u}{\partial t} = -\frac{\partial^3 u}{\partial x^3} + u \frac{\partial u}{\partial x}, \quad (1.1)$$

which describes the evolution of small amplitude long waves in fluids and other media.

The Kawahara equation with dispersive terms of the third and fifth orders

$$\frac{\partial u}{\partial t} + u \frac{\partial u}{\partial x} + \alpha \frac{\partial^3 u}{\partial x^3} + \beta \frac{\partial^5 u}{\partial x^5} = 0 \quad (1.2)$$

is used to model various dispersive phenomena such as plasma waves, capilarity-gravity water waves, etc. in situations when the cubic dispersive term is weak or not sufficient. In the most recent research, generalisations of Equations (1.1), (1.2) have been suggested and treated.

In the physical and mathematical literature the existence, uniqueness and analytic properties of solutions to the initial value problem have been intensively investigated for various linear and nonlinear dispersive equations. Boundary value problems for such equations were also considered. We refer, for example, to the comprehensive study undertaken in the book by Tao [16], among many other books and papers on the topic.

One should note the importance of the study of constant coefficient linear dispersive equations for its own sake and also because this provides prerequisites for the theory of nonlinear dispersive equations, since the latter are often obtained by perturbation of the linear theory ([16]). Developing the theory of linear equations is also essential for describing those evolution phenomena where the linear effects compensate or dominate nonlinear ones. In such situations, one can expect that the nonlinear solutions display almost the same behavior as the linear ones.

In the probabilistic literature significant attention has been paid to the equations of the form

$$\frac{\partial u}{\partial t} = k_m \frac{\partial^m u}{\partial x^m}, \quad x \in R, \quad t > 0, \quad m \geq 2. \quad (1.3)$$

The investigation of fundamental solutions to the equation (1.3) can be traced back to works by Bernštein and Lévy. Such solutions are sign-varying and, based on them, the so-called pseudoprocesses have been introduced and extensively investigated in the literature. Note, that Equations (1.3) of even and odd order possess

solutions of different structure and behaviour (for example, see [13]). We refer to [7], where a review of the recent results on this topic and additional literature are presented.

We note that in the probabilistic literature equations of the form (1.3) and their generalizations are often called higher-order heat-type equations.

Equations of the form (1.3) subject to random initial conditions were studied in [1], namely, the asymptotic behavior was analysed for the rescaled solution to the Airy equation with random initial conditions given by weakly dependent stationary processes.

More general odd-order equations of the form

$$\frac{\partial u}{\partial t} = \sum_{k=1}^N a_k \frac{\partial^{2k+1} u}{\partial x^{2k+1}}, \quad N = 1, 2, \dots, \quad (1.4)$$

subject to the random initial conditions represented by a strictly φ -sub-Gaussian harmonizable processes were considered in [2, 7]. Rigorous conditions were stated therein for the existence of solutions and some distributional properties of solutions were investigated.

The present paper continues the line of research initiated in the papers [2, 7]. Note that in the mathematical literature the initial value problems for partial differential equations have been studied within the framework of different functional spaces, including the most abstract ones. Here we take into consideration Equation (1.4) in the framework of special Banach spaces of random variables, which constitute a subclass of Orlicz spaces of exponential type, more precisely, we deal with the spaces of strictly φ -sub-Gaussian random processes. These spaces play an important role in extensions of properties of Gaussian and sub-Gaussian processes. Basic results on φ -sub-Gaussian processes and fields can be found, for example, in [3, 4, 8, 17].

The general methods and techniques developed for φ -sub-Gaussian processes, applied to the problems under consideration in the present paper, permit us to obtain bounds for the distributions of suprema of the solutions to the initial value problem for Equation (1.4). The bounds are presented in a form different than those obtained in the paper [7], and can be more useful in particular situations. In such a way, the results of the present paper complement and extend the results presented in [7].

To make the paper self-contained, in Sections 2 we present all important definitions and facts on harmonizable φ -sub-Gaussian processes, which will be used in the derivation of the main results. We also formulate the result on the conditions of existence of solutions to (1.4) with the φ -sub-Gaussian initial condition (see [2, 7] for its derivation). The main result on the bounds for the distributions of supremum of the solutions is stated in Section 3. We present examples of processes for which the assumptions of the general result are verified and bounds are written explicitly. The main result is also specified for the case of Gaussian initial condition.

2 Preliminaries

Since in the paper we consider a partial differential equation with random initial condition given by a real-valued harmonizable φ -sub-Gaussian process, in this section we present the necessary definitions and facts concerning such processes.

2.1 Harmonizable processes

Harmonizable processes are a natural extension of stationary processes to second-order nonstationary ones. Such class of processes allows us to retain advantages of the Fourier analysis. Harmonizable processes were introduced by Loève [12]. Recent developments on this theory are due to Rao [14] and Swift [15] among others.

Definition 2.1 ([12]). The second-order random function $X = \{X(t), t \in \mathbb{R}\}$, $\mathbf{E}X(t) = 0$, is called harmonizable if there exists a second-order random function $y = y(t), t \in \mathbb{R}, \mathbf{E}y(t) = 0$ such that the covariance $\Gamma_y(t, s) = \mathbf{E}y(t)\overline{y(s)}$ has finite variation and $X(t) = \int_{\mathbb{R}} e^{itu} dy(u)$, where the integral is defined in the mean-square sense.

Theorem 2.1 ([12] Loève theorem). *The second-order random function $X = \{X(t), t \in \mathbb{R}\}$, $\mathbf{E}X(t) = 0$, is harmonizable if and only if there exists a covariance function $\Gamma_y(u, v)$ with finite variation such that*

$$\Gamma_x(t, s) = \mathbf{E}X(t)\overline{X(s)} = \int_{\mathbb{R}} \int_{\mathbb{R}} e^{i(tu-sv)} d\Gamma_y(u, v). \quad (2.1)$$

Remark 2.1. In the theorem above, the covariance function Γ_y is of bounded Vitali variation (see [14, 15]). This fact guarantees that integral in (2.1) is in the Lebesgue sense. The function Γ_y is also called the spectral function or bi-measure of the process X .

Remark 2.2. In what follows, an integral of the type $\int \int_A f(t, s) dg(t, s)$ is understood as a common Lebesgue–Stieltjes integral, that is, the limit of the sum $\sum \sum f(t, s) \Delta_i \Delta_j g(t, s)$, and an integral of the type $\int \int_A f(t, s) |dg(t, s)|$ is understood as the limit of the sum $\sum \sum f(t, s) |\Delta_i \Delta_j g(t, s)|$.

Below we shall focus on real-valued harmonizable processes.

Definition 2.2. Real-valued second order random function $X = \{X(t), t \in \mathbb{R}\}$ is called harmonizable, if there exists a real-valued second order function $y(u), \mathbf{E}y(u) = 0, u \in \mathbb{R}$, such that $X(t) = \int_{-\infty}^{\infty} \sin tu dy(u)$ or $X(t) = \int_{-\infty}^{\infty} \cos tu dy(u)$ and the covariance function $\Gamma_y(t, s) = \mathbf{E}y(t)y(s)$ has finite variation. The integral is defined in the mean-square sense.

Theorem 2.2. *Real-valued second order function $X = \{X(t), t \in \mathbb{R}\}$, $\mathbf{E}X(t) = 0$, is harmonizable if and only if there exists the covariance function $\Gamma_y(u, v)$ with finite variation such that*

$$\Gamma_x(t, s) = \mathbf{E}X(t)X(s) = \int_{\mathbb{R}} \int_{\mathbb{R}} \kappa(tu)\kappa(sv) d\Gamma_y(u, v),$$

where $\kappa(v) = \cos v$ or $\kappa(v) = \sin v$.

The theorem above follows from Theorem 2.1.

2.2 φ -sub-Gaussian random variables and processes

Here we present some basic facts from the theory of φ -sub-Gaussian random variables and processes, as well as some necessary results.

A continuous even convex function φ is called an *Orlicz N -function*, if $\varphi(0) = 0, \varphi(x) > 0$ as $x \neq 0$, and the following conditions hold: $\lim_{x \rightarrow 0} \frac{\varphi(x)}{x} = 0$,

$\lim_{x \rightarrow \infty} \frac{\varphi(x)}{x} = \infty$. The function φ^* defined by $\varphi^*(x) = \sup_{y \in \mathbb{R}} (xy - \varphi(y))$ is called the *Young–Fenchel transform (or convex conjugate)* of the function φ [3, 11].

We say that φ satisfies the *Condition Q*, if φ is such an N -function that $\liminf_{x \rightarrow 0} \frac{\varphi(x)}{x^2} = c > 0$, where the case $c = \infty$ is possible [4, 8].

In what follows we will always deal with N -functions for which condition Q holds.

Examples of N -functions, for which the condition Q is satisfied:

$$\begin{aligned} \varphi(x) &= \frac{|x|^\alpha}{\alpha}, \quad 1 < \alpha \leq 2, \\ \varphi(x) &= \begin{cases} \frac{|x|^\alpha}{\alpha}, & |x| \geq 1, \alpha > 2; \\ \frac{|x|^2}{\alpha}, & |x| \leq 1, \alpha > 2. \end{cases} \end{aligned} \tag{2.2}$$

Definition 2.3 ([4, 8]). Let $\{\Omega, L, \mathbf{P}\}$ be a standard probability space. The random variable ζ is said to be φ -sub-Gaussian (or belongs to the space $\text{Sub}_\varphi(\Omega)$), if $\mathbf{E}\zeta = 0$, $\mathbf{E} \exp\{\lambda\zeta\}$ exists for all $\lambda \in \mathbb{R}$ and there exists a constant $a > 0$ such that the following inequality holds for all $\lambda \in \mathbb{R}$: $\mathbf{E} \exp\{\lambda\zeta\} \leq \exp\{\varphi(\lambda a)\}$.

The space $\text{Sub}_\varphi(\Omega)$ is a Banach space with respect to the norm [4, 8]

$$\tau_\varphi(\zeta) = \inf\{a > 0 : \mathbf{E} \exp\{\lambda\zeta\} \leq \exp\{\varphi(a\lambda)\}\},$$

which is called the φ -sub-Gaussian standard of the random variable ζ .

Centered Gaussian random variables $\zeta \sim N(0, \sigma^2)$ are φ -sub-Gaussian with $\varphi(x) = \frac{x^2}{2}$ and $\tau_\varphi^2(\zeta) = \mathbf{E}\zeta^2 = \sigma^2$. In the case $\varphi(x) = \frac{x^2}{2}$, φ -sub-Gaussian random variables are simply sub-Gaussian.

Definition 2.4. A family Δ of random variables $\zeta \in \text{Sub}_\varphi(\Omega)$ is called strictly φ -sub-Gaussian (see [5]), if there exists a constant C_Δ such that for all countable sets I of random variables $\zeta_i \in \Delta, i \in I$, the following inequality holds:

$$\tau_\varphi \left(\sum_{i \in I} \lambda_i \zeta_i \right) \leq C_\Delta \left(\mathbf{E} \left(\sum_{i \in I} \lambda_i \zeta_i \right)^2 \right)^{1/2}. \tag{2.3}$$

The constant C_Δ is called the *determining constant* of the family Δ .

The linear closure of a strictly φ -sub-Gaussian family Δ in the space $L_2(\Omega)$ is the strictly φ -sub-Gaussian with the same determining constant [5].

Definition 2.5. The random process $\zeta = \{\zeta(t), t \in \mathbf{T}\}$ is called (strictly) φ -sub-Gaussian if the family of random variables $\{\zeta(t), t \in \mathbf{T}\}$ is (strictly) φ -sub-Gaussian [5].

The following example of strictly φ -sub-Gaussian random process is important for our study. The solutions of partial differential equations considered in the next sections are of the same form as in this example.

Example 2.1 ([5]). Let K be a deterministic kernel and suppose that the process $X = \{X(t), t \in \mathbf{T}\}$ can be represented in the form

$$X(t) = \int_T K(t, s) d\xi(s),$$

where $\xi(t)$, $t \in \mathbf{T}$, is a strictly φ -sub-Gaussian random process and the integral above is defined in the mean-square sense. Then the process $X(t)$, $t \in \mathbf{T}$, is a strictly φ -sub-Gaussian random process with the same determining constant.

The notion of admissible function for the space $\text{Sub}_\varphi(\Omega)$ will be used to state the conditions of the existence of solutions of partial differential equations considered in the paper and to write down the bounds for suprema of these solutions.

Lemma 2.1 ([9, 10], p. 146). *Let $Z(u)$, $u \geq 0$, be a continuous, increasing function such that $Z(u) > 0$ and the function $\frac{u}{Z(u)}$ is non-decreasing for $u > u_0$, where $u_0 \geq 0$ is a constant. Then for all $u, v \neq 0$*

$$\left| \sin \frac{u}{v} \right| \leq \frac{Z(|u| + u_0)}{Z(|v| + u_0)} \quad (2.4)$$

Definition 2.6. Function $Z(u)$, $u \geq 0$, is called *admissible* for the space $\text{Sub}_\varphi(\Omega)$, if for $Z(u)$ conditions of Lemma 2.1 hold and for some $\varepsilon > 0$ the integral

$$\int_0^\varepsilon \Psi \left(\ln \left(Z^{(-1)} \left(\frac{1}{s} \right) - u_0 \right) \right) ds$$

converges, where $\Psi(v) = \frac{v}{\varphi^{(-1)}(v)}$, $v > 0$.

For example, the function $Z(u) = u^\rho$, $0 < \rho < 1$, is an admissible function for the space $\text{Sub}_\varphi(\Omega)$ if φ is defined in (2.2).

Characteristic feature of φ -sub-Gaussian random variables is the exponential bounds for their tail probabilities. For φ -sub-Gaussian processes estimates for their suprema are available in different forms, see, for example, the book [3].

To derive our main results, we shall use the following theorem on the distribution of supremum of a φ -sub-Gaussian random process, proved in the paper [6] (see also [3]).

Theorem 2.3. *Let $X = \{X(t), t \in \mathbf{T}\}$ be a φ -sub-Gaussian process and ρ_X be the pseudometrics generated by X , that is, $\rho_X(t, s) = \tau_\varphi(X(t) - X(s))$, $t, s \in \mathbf{T}$. Assume that the pseudometric space (\mathbf{T}, ρ_X) is separable, the process X is separable on (\mathbf{T}, ρ_X) and $\varepsilon_0 := \sup_{t \in \mathbf{T}} \tau_\varphi(X(t)) < \infty$.*

Let $r(x)$, $x \geq 1$, be a non-negative, monotone increasing function such that the function $r(e^x)$, $x \geq 0$, is convex, and for $0 < \varepsilon \leq \varepsilon_0$

$$I_r(\varepsilon) := \int_0^\varepsilon r(N(v)) dv < \infty, \quad (2.5)$$

where $\{N(v), v > 0\}$ is the massiveness of the pseudometric space (\mathbf{T}, ρ_X) , that is, $N(v)$ denotes the smallest number of elements in a v -covering of T , and the covering is formed by closed balls of radius of at most v .

Then for all $\lambda > 0$, $0 < \theta < 1$ and $u > 0$ it holds

$$\mathbf{E} \exp \left\{ \lambda \sup_{t \in \mathbf{T}} |X(t)| \right\} \leq Q(\lambda, \theta) \quad (2.6)$$

and

$$P\{\sup_{t \in \mathbf{T}} |X(t)| \geq u\} \leq 2A(\theta, u), \tag{2.7}$$

where

$$Q(\lambda, \theta) := \exp \left\{ \varphi \left(\frac{\lambda \varepsilon_0}{1 - \theta} \right) \right\} r^{(-1)} \left(\frac{I_r(\theta \varepsilon_0)}{\theta \varepsilon_0} \right),$$

$$A(\theta, u) = \exp \left\{ -\varphi^* \left(\frac{u(1 - \theta)}{\varepsilon_0} \right) \right\} r^{(-1)} \left(\frac{I_r(\theta \varepsilon_0)}{\theta \varepsilon_0} \right).$$

Remark 2.3. The integrals of the form

$$I(\varepsilon) := \int_0^\varepsilon g(N(v)) dv, \quad \varepsilon > 0, \tag{2.8}$$

with $g(v)$, $v \geq 1$, being a nonnegative nondecreasing function, are called entropy integrals. Entropy characteristics of the parametric set T with respect to the pseudometrics $\rho_X(t, s) = \tau_\varphi(X(t) - X(s))$, $t, s \in \mathbf{T}$, generated by the process $X = \{X(t), t \in \mathbf{T}\}$, and the rate of growth of metric massiveness $N(v)$, or metric entropy $H(v) := \ln(N(v))$, are closely related to the properties of the process X (see [3] for details).

The integrals (2.8) play an important role in the study of such properties as boundedness and continuity of sample paths of a process, these integrals appear in estimates for moduli of continuity and distribution of supremum.

General results of this kind for φ -sub-Gaussian processes are related to the convergence of the integrals (2.8), where for $g(v)$ one takes $\Psi(\ln(v))$ with $\Psi(v) = \frac{v}{\varphi^{(-1)}(v)}$, $v > 0$. Theorem 2.3 is more suitable for the case of “moderate” growth of the metric entropy and can lead to improved inequalities for upper bound for the distribution of supremum of the process, in comparison with more general inequalities involving the integrals based on the above function Ψ (see [3]).

Entropy methods are also used in the modern approximation theory. Theorem 2.3 was applied, for example, in [6], for developing uniform approximation schemes for φ -sub-Gaussian processes.

2.3 Solutions of linear odd-order heat-type equations with random initial conditions

Let us consider the linear equation

$$\sum_{k=1}^N a_k \frac{\partial^{2k+1} U(t, x)}{\partial x^{2k+1}} = \frac{\partial U(t, x)}{\partial t}, \quad t > 0, x \in \mathbb{R}, \tag{2.9}$$

subject to the random initial condition

$$U(0, x) = \eta(x), \quad x \in \mathbb{R}, \tag{2.10}$$

and $\{a_k\}_{k=1}^N$ are some constants.

The next theorem gives the conditions of the existence of the solutions of the equation above with a φ -sub-Gaussian initial condition $\eta(x)$ (see [2, 7]).

Theorem 2.4. *Let $\eta = \{\eta(x), x \in \mathbb{R}\}$ be a real harmonizable (see Definition 2.2) and strictly φ -sub-Gaussian random process. Also let $Z = \{Z(u), u \geq 0\}$ be a function*

admissible for the space $\text{Sub}_\varphi(\Omega)$. Assume that the following integral converges

$$\int_{\mathbb{R}} \int_{\mathbb{R}} |\lambda|^{2N+1} |\mu|^{2N+1} Z(u_0 + |\lambda|^{2N+1}) Z(u_0 + |\mu|^{2N+1}) d|\Gamma_y(\lambda, \mu)| < \infty. \tag{2.11}$$

Then

$$U(t, x) = \int_{-\infty}^{\infty} I(t, x, \lambda) dy(\lambda) \tag{2.12}$$

is the classical solution to the problem (2.9)–(2.10), that is, $U(t, x)$ satisfies Equation (2.9) with probability one and $U(0, x) = \eta(x)$. Here

$$I(t, x, \lambda) = \kappa \left(\lambda x + t \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k \right), \tag{2.13}$$

and $\kappa(v) = \cos v$ or $\kappa(v) = \sin v$ for the cases when $\eta(x) = \int_{\mathbb{R}} \cos(xu) dy(u)$ or $\eta(x) = \int_{\mathbb{R}} \sin(xu) dy(u)$ respectively.

Remark 2.4. Note that under the condition (2.11) all the integrals

$$\int_{\mathbb{R}} \lambda^s I(t, x, \lambda) dy(\lambda), \quad s = 0, 1, 2, \dots, 2N + 1, \tag{2.14}$$

converge uniformly in probability for $|x| \leq A$ and $0 \leq t \leq T$ for all $A > 0, T > 0$. This guarantees that the derivatives of orders $s = 1, 2, \dots, 2N + 1$ of solution $U(t, x)$ given by (2.12) exist with probability one. In this sense we can treat $U(t, x)$ as a classical solution. We refer for more details to [2].

Remark 2.5. Similar result can be stated for the case $\eta(x) = \int_{-\infty}^{\infty} (a \sin xu + b \cos xu) dy(u)$, where a and b are some real constants.

Remark 2.6. Let $\varphi(x) = \frac{|x|^p}{p}, p > 1$, for sufficiently large x . Then the statement of Theorem 2.4 holds if the following integral converges

$$\int_{\mathbb{R}} \int_{\mathbb{R}} |\lambda \mu|^{2N+1} (\ln(1 + \lambda) \ln(1 + \mu))^\alpha d|\Gamma_y(\lambda, \mu)|, \tag{2.15}$$

where α is a constant such that $\alpha > 1 - \frac{1}{p}$ (see [2]).

3 Main results

3.1 Some auxiliary estimates

Let us consider a separable metric space (\mathbf{T}, d) , where $\mathbf{T} = \{a_i \leq t_i \leq b_i, i = 1, 2\}$ and $d(t, s) = \max_{i=1,2} |t_i - s_i|, t = (t_1, t_2), s = (s_1, s_2)$.

Theorem 3.1. Let $X = \{X(t), t \in \mathbf{T}\}$ be a separable φ -sub-Gaussian random process such that $\varepsilon_0 = \sup_{t \in \mathbf{T}} \tau_\varphi(X(t)) < \infty$ and

$$\sup_{\substack{d(t,s) \leq h, \\ t,s \in \mathbf{T}}} \tau_\varphi(X(t) - X(s)) \leq \sigma(h), \tag{3.1}$$

where $\{\sigma(h), 0 < h \leq \max_{i=1,2} |b_i - a_i|\}$ is a monotonically increasing continuous function such that $\sigma(h) \rightarrow 0$ as $h \rightarrow 0$, and for $0 < \varepsilon \leq \gamma_0$

$$\tilde{I}_r(\varepsilon) := \int_0^\varepsilon r \left(\left(\frac{b_1 - a_1}{2\sigma^{(-1)}(v)} + 1 \right) \left(\frac{b_2 - a_2}{2\sigma^{(-1)}(v)} + 1 \right) \right) dv < \infty, \quad (3.2)$$

where $\gamma_0 = \sigma(\max_{i=1,2} |b_i - a_i|)$ and $r(x), x \geq 1$, is defined in Theorem 2.3. Then for all $0 < \theta < 1$ and $u > 0$ it holds

$$P\{\sup_{t \in \mathbf{T}} |X(t)| \geq u\} \leq 2\tilde{A}(\theta, u), \quad (3.3)$$

where

$$\tilde{A}(\theta, u) = \exp \left\{ -\varphi^* \left(\frac{u(1-\theta)}{\varepsilon_0} \right) \right\} r^{(-1)} \left(\frac{\tilde{I}_r(\min(\theta\varepsilon_0, \gamma_0))}{\theta\varepsilon_0} \right).$$

Proof. To prove this theorem, we apply Theorem 2.3 in the case of the separable metric space (\mathbf{T}, d) with $T = \{a_i \leq t_i \leq b_i, i = 1, 2\}$ and $d(t, s) = \max_{i=1,2} |t_i - s_i|$, $t = (t_1, t_2), s = (s_1, s_2)$.

In particular, condition (3.1) means that

$$\sup_{\substack{d(t,s) \leq h, \\ t,s \in \mathbf{T}}} \tau_\varphi(X(t) - X(s)) = \sup_{\substack{d(t,s) \leq h, \\ t,s \in \mathbf{T}}} \rho_X(t, s) \leq \sigma(h), \quad 0 < h \leq \max_{i=1,2} |b_i - a_i|.$$

From the fact that the process X is separable on (\mathbf{T}, ρ_X) and the function $\{\sigma(h), 0 < h \leq \max_{i=1,2} |b_i - a_i|\}$ is a monotonically increasing continuous function, we get that for $\varepsilon \leq \gamma_0$ the smallest number of elements in an ε -covering of the pseudometric space (\mathbf{T}, ρ_X) can be estimated as the smallest number of elements in a $(\sigma^{(-1)}(\varepsilon))$ -covering of the metric space (\mathbf{T}, d) as follows:

$$N(\varepsilon) \leq \left[\left(\frac{b_1 - a_1}{2\sigma^{(-1)}(\varepsilon)} + 1 \right) \left(\frac{b_2 - a_2}{2\sigma^{(-1)}(\varepsilon)} + 1 \right) \right] \quad \text{and} \quad N(\varepsilon) = 1 \quad \text{if } \varepsilon > \gamma_0.$$

Hence, from condition (3.2) we get that

$$I_r(\varepsilon) = \int_0^\varepsilon r(N(v)) dv \leq \tilde{I}_r(\varepsilon) < \infty, \quad \varepsilon \leq \gamma_0,$$

that is, conditions of Theorem 2.3 are satisfied for the process X . Finally, taking into account the estimates above and the properties of the function r we derive the estimate for distribution of supremum of the process X for all $0 < \theta < 1$ and $u > 0$:

$$\begin{aligned} P\{\sup_{t \in \mathbf{T}} |X(t)| \geq u\} &\leq 2 \exp \left\{ -\varphi^* \left(\frac{u(1-\theta)}{\varepsilon_0} \right) \right\} r^{(-1)} \left(\frac{I_r(\min(\theta\varepsilon_0, \gamma_0))}{\theta\varepsilon_0} \right) \\ &\leq 2 \exp \left\{ -\varphi^* \left(\frac{u(1-\theta)}{\varepsilon_0} \right) \right\} r^{(-1)} \left(\frac{\tilde{I}_r(\min(\theta\varepsilon_0, \gamma_0))}{\theta\varepsilon_0} \right) = 2\tilde{A}(\theta, u). \quad \square \end{aligned}$$

As can be seen from Theorem 3.1, it is crucial to guarantee some kind of continuity of X on T in the form (3.1), that is, with respect to the norm τ_φ induced by the process X itself. Fulfilment of condition (3.1) enables us to write down the upper bound (3.3) for the distribution of supremum of a φ -sub-Gaussian process $X = \{X(t), t \in \mathbf{T}\}$ defined on a separable metric space (\mathbf{T}, d) .

Below we present as a separate theorem the very useful result giving the conditions for the estimate (3.1) to hold for the field $\{U(t, x), a \leq t \leq b, c \leq x \leq d\}$ representing the solution to (2.9)–(2.10).

Theorem 3.2. *Let $y = \{y(u), u \in \mathbb{R}\}$ be a strictly φ -sub-Gaussian random process with a determining constant C_y and $U(t, x) = \int_{-\infty}^{\infty} I(t, x, \lambda) dy(\lambda)$, where $I(t, x, \lambda)$ is given in Theorem 2.4, $a \leq t \leq b, c \leq x \leq d$. Assume that $U(t, x)$ exists and is continuous with probability one (this condition holds if Theorem 2.4 holds). Let $\mathbf{E}y(t)y(s) = \Gamma_y(s, t)$. Assume that $\{Z(u), u \geq 0\}$ is an admissible function for the space $\text{Sub}_\varphi(\Omega)$. If the integral*

$$C_Z^2 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(Z\left(\frac{|\lambda|}{2} + u_0\right) + Z\left(\frac{1}{2} \left| \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k \right| + u_0\right) \right) \\ \times \left(Z\left(\frac{|\mu|}{2} + u_0\right) + Z\left(\frac{1}{2} \left| \sum_{k=1}^N a_k \mu^{2k+1} (-1)^k \right| + u_0\right) \right) d|\Gamma_y(\lambda, \mu)| \quad (3.4)$$

converges, then there exists the function

$$\sigma(h) = 2C_y C_Z \left(Z\left(\frac{1}{h} + u_0\right) \right)^{-1}, \quad 0 < h < \max(b - a, d - c), \quad (3.5)$$

such that

$$\sup_{\substack{t, t_1 \in [a, b]: |t - t_1| \leq h \\ x, x_1 \in [c, d]: |x - x_1| \leq h}} \tau_\varphi(U(t, x) - U(t_1, x_1)) \leq \sigma(h). \quad (3.6)$$

This result was obtained in the paper [7] as an intermediate statement in the course of the proof of Theorem 5.1. To make the present paper self-contained, we present in Appendix the main steps of the proof.

3.2 On the distribution of supremum of solution to the problem (2.9)–(2.10)

Now we have all the necessary tools to derive the estimate for the distribution of supremum of the field $U(t, x)$ representing the solution to (2.9)–(2.10).

Theorem 3.3. *Let $y = \{y(u), u \in \mathbb{R}\}$ be a strictly φ -sub-Gaussian random process with a determining constant C_y and $U(t, x) = \int_{-\infty}^{\infty} I(t, x, \lambda) dy(\lambda)$, where $I(t, x, \lambda)$ is given in Theorem 2.4, $a \leq t \leq b, c \leq x \leq d$. Assume that for $U(t, x)$ the conditions of Theorem 3.2 hold. Let $r(x), x \geq 1$, be a non-negative, monotone increasing function such that the function $r(e^x), x \geq 0$, is convex, and condition (3.2) is satisfied for $\sigma(h)$ given by (3.5).*

Then for all $0 < \theta < 1$ and $u > 0$ the following inequality holds true

$$P\left\{ \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} |U(t, x)| > u \right\} \leq 2\hat{A}(\theta, u), \quad (3.7)$$

where

$$\hat{A}(\theta, u) = \exp \left\{ -\varphi^* \left(\frac{u(1-\theta)}{\varepsilon_0} \right) \right\}_{r^{(-1)}} \left(\frac{\hat{I}_r(\min(\theta\varepsilon_0, \gamma_0))}{\theta\varepsilon_0} \right), \quad (3.8)$$

$$\begin{aligned} \hat{I}_r(\delta) = & \int_0^\delta r \left(\left[\left(\frac{b-a}{2} \left(Z^{(-1)} \left(\frac{2C_Z C_Y}{s} \right) - u_0 \right) + 1 \right) \right. \right. \\ & \left. \left. \times \left(\frac{d-c}{2} \left(Z^{(-1)} \left(\frac{2C_Z C_Y}{s} \right) - u_0 \right) + 1 \right) \right] \right) ds, \end{aligned} \quad (3.9)$$

$$\varepsilon_0 = \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} \tau_\varphi(U(t, x)),$$

$$\gamma_0 = \frac{2C_Y C_Z}{Z\left(\frac{1}{\varkappa} + u_0\right)}, \quad \varkappa = \max(b-a, d-c).$$

Proof. The assertion of this theorem follows from Theorems 3.1 and 3.2. Since the conditions of Theorem 3.2 are satisfied, then there exists the function $\sigma(h) = 2C_Y C_Z \left(Z\left(\frac{1}{h} + u_0\right) \right)^{-1}$, $0 < h < \max(b-a, d-c)$, such that

$$\sup_{\substack{t, t_1 \in [a, b]: |t-t_1| \leq h \\ x, x_1 \in [c, d]: |x-x_1| \leq h}} \tau_\varphi(U(t, x) - U(t_1, x_1)) \leq \sigma(h).$$

In this case,

$$\sigma^{(-1)}(v) = \left(Z^{(-1)} \left(\frac{2C_Y C_Z}{v} \right) - u_0 \right)^{-1}, \quad 0 < v < \frac{2C_Y C_Z}{Z\left(\frac{1}{\varkappa} + u_0\right)} = \gamma_0,$$

and for ε_0 the upper bound (A.1) holds (see Appendix A).

Since the conditions (3.1) and (3.2) of Theorem 3.1 also hold true, the final estimate directly follows. \square

Remark 3.1. The derivation of our main result is based on Theorem 2.3, and due to this we present the bounds for the distribution of the supremum of the process $U(t, x)$ in the form different than those obtained in the paper [7]. This form of bounds can be more useful in particular situations giving the possibility to calculate the explicit expressions for the bounds.

Now we will specify the statement of Theorem 3.3 for particular choices of the admissible function Z and the function φ .

Example 3.1. Consider $\varphi(x) = \frac{|x|^\alpha}{\alpha}$, $1 < \alpha \leq 2$. Then $\varphi^*(x) = \frac{|x|^\gamma}{\gamma}$, where $\gamma \geq 2$, and $\frac{1}{\alpha} + \frac{1}{\gamma} = 1$. Hence,

$$\hat{A}(\theta, u) = \exp \left\{ -\frac{u^\gamma (1-\theta)^\gamma}{\gamma \varepsilon_0^\gamma} \right\}_{r^{(-1)}} \left(\frac{\hat{I}_r(\min(\theta\varepsilon_0, \gamma_0))}{\theta\varepsilon_0} \right).$$

Now it is necessary to make the following steps:

1. to check the fulfilment of (3.4) with a particular function Z , admissible for a given φ ;
2. to calculate $\sigma(h)$ in (3.5);

3. to choose the function r satisfying (3.2);

4. to derive an estimate for $\hat{A}(\theta, u)$.

So, let us choose the admissible function $Z(u) = |u|^\rho$, $0 < \rho \leq 1$.

In this case $u_0 = 0$, $Z^{(-1)}(u) = u^{\frac{1}{\rho}}$, $u > 0$, and

$$C_Z^2 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(\left(\frac{|\lambda|}{2} \right)^\rho + \left| \frac{1}{2} \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k \right|^\rho \right) \quad (3.10)$$

$$\begin{aligned} & \times \left(\left(\frac{|\mu|}{2} \right)^\rho + \left| \frac{1}{2} \sum_{k=1}^N a_k \mu^{2k+1} (-1)^k \right|^\rho \right) d|\Gamma_y(\lambda, \mu)| \\ & \leq \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{1}{2^{2\rho}} \left(|\lambda|^\rho + \left(\sum_{k=1}^N |a_k| |\lambda|^{2k+1} \right)^\rho \right) \\ & \times \left(|\mu|^\rho + \left(\sum_{k=1}^N |a_k| |\mu|^{2k+1} \right)^\rho \right) d|\Gamma_y(\lambda, \mu)|. \end{aligned} \quad (3.11)$$

This integral converges if the following integral converges

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |\lambda \mu|^{(2N+1)\rho} d|\Gamma_y(\lambda, \mu)| < \infty. \quad (3.12)$$

Note that for the existence of the solution $U(t, x)$ we have to impose the condition (2.11), which for the admissible function $Z(u) = |u|^\rho$ takes the form

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |\lambda|^{(2N+1)(\rho+1)} |\mu|^{(2N+1)(\rho+1)} d|\Gamma_y(\lambda, \mu)| < \infty, \quad (3.13)$$

and implies the fulfilment of (3.12). Therefore, C_Z^2 is well defined.

If (3.13) holds true, then we can define $\sigma(h)$ by means of the formula (3.5), and, for our choice of Z , we have that

$$\sigma(h) = 2C_y C_Z h^\rho = Ch^\rho, \quad 0 < \rho \leq 1, \quad (3.14)$$

where we have denoted $C = 2C_y C_Z$.

Therefore, in view of Theorem 3.2, condition (3.6) holds with $\sigma(h)$ of the form (3.14), that is, we have the Hölder continuity of sample paths of the solution $U(t, x)$.

Let $r(v) = v^\beta - 1$, $0 < \beta < \rho/2$, then $r^{(-1)}(v) = (v + 1)^{1/\beta}$. For such $r(v)$ and the above choice of Z we have

$$\begin{aligned} \hat{I}_r(\delta) &= \int_0^\delta r \left(\left[\left(\frac{b-a}{2} \left(Z^{(-1)} \left(\frac{2C_Z C_y}{s} \right) - u_0 \right) + 1 \right) \right. \right. \\ & \quad \left. \left. \times \left(\frac{d-c}{2} \left(Z^{(-1)} \left(\frac{2C_Z C_y}{s} \right) - u_0 \right) + 1 \right) \right] \right) ds \\ & \leq \int_0^\delta \left(\left(\frac{\varkappa (2C_Z C_y)^{1/\rho}}{2 s^{1/\rho}} + 1 \right)^{2\beta} - 1 \right) ds. \end{aligned} \quad (3.15)$$

Consider $\delta \in (0, \theta \varepsilon_0]$ and let us choose θ such that $\frac{\varkappa}{2} \left(\frac{2C_Z C_Y}{\theta \varepsilon_0} \right)^{1/\rho} > 1$, that is, $\theta \in (0, \frac{2C_Z C_Y}{\varepsilon_0} \left(\frac{\varkappa}{2} \right)^\rho)$. Then we can write the following estimate:

$$\begin{aligned} \hat{I}_r(\delta) &\leq \int_0^\delta \left(\left(\varkappa \left(\frac{2C_Z C_Y}{s} \right)^{1/\rho} \right)^{2\beta} - 1 \right) ds \\ &= (2C_Z C_Y)^{2\beta/\rho} \varkappa^{2\beta} \left(1 - \frac{2\beta}{\rho} \right)^{-1} \delta^{1-2\beta/\rho} - \delta. \end{aligned}$$

Suppose that $\theta \varepsilon_0 < \gamma_0$, then

$$r^{(-1)} \left(\frac{\hat{I}_r(\min(\theta \varepsilon_0, \gamma_0))}{\theta \varepsilon_0} \right) \leq (2C_Z C_Y)^{2/\rho} \varkappa^2 \left(1 - \frac{2\beta}{\rho} \right)^{-1/\beta} (\theta \varepsilon_0)^{-2/\rho},$$

and

$$\hat{A}(\theta, u) \leq \exp \left\{ - \frac{u^\gamma (1 - \theta)^\gamma}{\gamma \varepsilon_0^\gamma} \right\} (2C_Z C_Y)^{2/\rho} \varkappa^2 \left(1 - \frac{2\beta}{\rho} \right)^{-1/\beta} (\theta \varepsilon_0)^{-2/\rho}.$$

If $\beta \rightarrow 0$, then $(1 - \frac{2\beta}{\rho})^{-1/\beta} \rightarrow e^{2/\rho}$, and we obtain

$$\hat{A}(\theta, u) \leq \exp \left\{ - \frac{u^\gamma (1 - \theta)^\gamma}{\gamma \varepsilon_0^\gamma} \right\} (2e C_Z C_Y)^{2/\rho} \varkappa^2 (\theta \varepsilon_0)^{-2/\rho},$$

for $\theta \in (0, \frac{2C_Z C_Y}{\varepsilon_0} \left(\frac{\varkappa}{2} \right)^\rho)$, and, therefore, for such θ we obtain

$$P \left\{ \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} |U(t, x)| > u \right\} \leq 2 \exp \left\{ - \frac{u^\gamma (1 - \theta)^\gamma}{\gamma \varepsilon_0^\gamma} \right\} (2e C_Z C_Y)^{2/\rho} \varkappa^2 (\theta \varepsilon_0)^{-2/\rho}. \tag{3.16}$$

Let $y = \{y(u), u \in \mathbb{R}\}$ be a centered Gaussian random process. Then $C_y = 1$, $\varphi(x) = \frac{x^2}{2}$, $\varphi^*(x) = \frac{x^2}{2}$.

Then the bound (3.16) takes place with $\gamma = 2$, provided that (3.13) holds.

Note that the constant $\rho \in (0, 1]$ should guarantee the convergence of the integral (3.13), and, therefore, the choice of ρ in (3.16) depends on the integrability properties of Γ_γ .

The choice of the function φ in Example 3.1 is reasoned by the fact that the corresponding class of random process is the natural generalization of Gaussian processes. This example is rather simple and, at the same time, it is very illustrative and instructive. Therefore, the derivations above are worth to be summarized as a separate statement.

Corollary 3.1. *Let $y = \{y(u), u \in \mathbb{R}\}$ be a real strictly φ -sub-Gaussian random process with $\varphi(x) = \frac{|x|^\alpha}{\alpha}$, $1 < \alpha \leq 2$, determining constant C_y and $\mathbb{E}y(t)y(s) = \Gamma_\gamma(s, t)$. Let $U(t, x) = \int_{-\infty}^\infty I(t, x, \lambda) dy(\lambda)$, where $I(t, x, \lambda)$ is given by (2.13),*

$a \leq t \leq b$, $c \leq x \leq d$, and $\varepsilon_0 = \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} \tau_\varphi(U(t, x))$. Further, let the constant $\rho \in (0, 1]$ be such that (3.13) holds. Then

(i) $U(t, x)$ exists, is continuous with probability one and for its sample paths the Hölder continuity holds in the form

$$\sup_{\substack{t, t_1 \in [a, b]; |t - t_1| \leq h \\ x, x_1 \in [c, d]; |x - x_1| \leq h}} \tau_\varphi(U(t, x) - U(t_1, x_1)) \leq 2C_Z C_Y h^\rho,$$

where C_Z is defined in (3.10);

(ii) for all $0 < \theta < 1$ such that $\theta \varepsilon_0 < 2C_Z C_Y (\varkappa/2)^\rho$ with $\varkappa = \max(b - a, d - c)$, γ such that $\frac{1}{\alpha} + \frac{1}{\gamma} = 1$, and $u > 0$ the following inequality holds true

$$P \left\{ \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} |U(t, x)| > u \right\} \leq 2 \exp \left\{ - \frac{u^\gamma (1 - \theta)^\gamma}{\gamma \varepsilon_0^\gamma} \right\} (2e C_Z C_Y)^{2/\rho} \varkappa^2 (\theta \varepsilon_0)^{-2/\rho}. \quad (3.17)$$

In particular, if the process $y = \{y(u), u \in \mathbb{R}\}$ is Gaussian, then

$$P \left\{ \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} |U(t, x)| > u \right\} \leq 2 \exp \left\{ - \frac{u^2 (1 - \theta)^2}{2\varepsilon_0^2} \right\} (2e C_Z)^{2/\rho} \varkappa^2 (\theta \varepsilon_0)^{-2/\rho} \quad (3.18)$$

for all $0 < \theta < 1$ such that $\theta \Gamma < 2C_Z (\varkappa/2)^\rho$ and $u > 0$.

Example 3.2. Consider $\varphi(x) = \exp\{|x|\} - |x| - 1$, $x \in \mathbb{R}$. Then $\varphi^*(x) = (|x| + 1) \ln(|x| + 1) - |x|$, $x \in \mathbb{R}$. Hence,

$$\begin{aligned} \hat{A}(\theta, u) &= \exp \left\{ - \left(\frac{u(1 - \theta)}{\varepsilon_0} + 1 \right) \ln \left(\frac{u(1 - \theta)}{\varepsilon_0} + 1 \right) + \frac{u(1 - \theta)}{\varepsilon_0} \right\} \\ &\times r^{(-1)} \left(\frac{\hat{I}_r(\min(\theta \varepsilon_0, \gamma_0))}{\theta \varepsilon_0} \right). \end{aligned}$$

Let us take the function $Z(u) = \ln^\alpha(u + 1)$, $u \geq 0$, $\alpha > 1$, as an admissible function for the space $\text{Sub}_\varphi(\Omega)$. In this case,

$$u_0 = e^\alpha - 1, \quad Z^{(-1)}(v) = \exp \left\{ v^{\frac{1}{\alpha}} \right\} - 1, \quad Z(v + u_0) = \ln^\alpha(v + e^\alpha),$$

$$\begin{aligned} C_Z^2 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(\ln^\alpha \left(\frac{|\lambda|}{2} + e^\alpha \right) + \ln^\alpha \left(\frac{1}{2} \left| \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k \right| + e^\alpha \right) \right) \\ &\times \left(\ln^\alpha \left(\frac{|\mu|}{2} + e^\alpha \right) + \ln^\alpha \left(\frac{1}{2} \left| \sum_{k=1}^N a_k \mu^{2k+1} (-1)^k \right| + e^\alpha \right) \right) d|\Gamma_y(\lambda, \mu)|. \end{aligned}$$

The above integral converges if the following integral converges

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \ln^\alpha(|\lambda| + e^\alpha) \ln^\alpha(|\mu| + e^\alpha) d|\Gamma_y(\lambda, \mu)|. \quad (3.19)$$

That is, if condition (3.19) holds true, then Theorem 3.3 holds with

$$\begin{aligned} \hat{I}_r(\delta) &= \int_0^\delta r \left(\left[\left(\frac{b-a}{2} \left(\exp \left\{ \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} - 1 - (e^\alpha - 1) \right) + 1 \right) \right. \right. \\ &\quad \left. \left. \times \left(\frac{d-c}{2} \left(\exp \left\{ \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} - 1 - (e^\alpha - 1) \right) + 1 \right) \right] \right) ds \\ &= \int_0^\delta r \left(\left[\left(\frac{b-a}{2} \left(\exp \left\{ \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} - e^\alpha \right) + 1 \right) \right. \right. \\ &\quad \left. \left. \times \left(\frac{d-c}{2} \left(\exp \left\{ \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} - e^\alpha \right) + 1 \right) \right] \right) ds. \end{aligned} \quad (3.20)$$

Let $\frac{d-c}{2} e^\alpha > 1$ and $\frac{b-a}{2} e^\alpha > 1$.

That is, we choose some $\alpha > \max \{1, \ln(\frac{2}{b-a}), \ln(\frac{2}{d-c})\}$. Then

$$\begin{aligned} \hat{I}_r(\delta) &\leq \int_0^\delta r \left(\frac{(b-a)(d-c)}{4} \exp \left\{ 2 \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} \right) ds \\ &\leq \int_0^\delta r \left(\frac{\varkappa^2}{4} \exp \left\{ 2 \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} \right) ds. \end{aligned} \quad (3.21)$$

In our case, the function $r(v) = \ln v, v \geq 1$, satisfies the conditions defined in Theorem 2.3 and is convenient for the estimation of $\hat{I}_r(\delta)$. Substituting it in the expression above, we get

$$\begin{aligned} \hat{I}_r(\delta) &\leq \int_0^\delta \ln \left(\frac{\varkappa^2}{4} \exp \left\{ 2 \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} \right\} \right) ds \\ &= \delta \ln \left(\frac{\varkappa^2}{4} \right) + \int_0^\delta 2 \left(\frac{2C_Z C_Y}{s} \right)^{\frac{1}{\alpha}} ds \\ &= \delta \ln \left(\frac{\varkappa^2}{4} \right) + \frac{2(2C_Z C_Y)^{\frac{1}{\alpha}} \delta^{1-\frac{1}{\alpha}}}{1-\frac{1}{\alpha}}. \end{aligned} \quad (3.22)$$

Since $r^{(-1)}(v) = e^v, v \geq 0$, for such $\theta \in (0, 1)$ that $\theta \varepsilon_0 < \gamma_0$ we obtain

$$r^{(-1)} \left(\frac{\hat{I}_r(\min(\theta \varepsilon_0, \gamma_0))}{\theta \varepsilon_0} \right) \leq \exp \left\{ \ln \left(\frac{\varkappa^2}{4} \right) + \frac{2\alpha}{\alpha-1} \left(\frac{2C_Z C_Y}{\theta \varepsilon_0} \right)^{\frac{1}{\alpha}} \right\}$$

and, finally, for all $u > 0$ the following inequality emerges

$$\begin{aligned} P \left\{ \sup_{\substack{a \leq t \leq b \\ c \leq x \leq d}} |U(t, x)| > u \right\} &\leq 2\hat{A}(\theta, u) \\ &\leq 2 \exp \left\{ - \left(\frac{u(1-\theta)}{\varepsilon_0} + 1 \right) \ln \left(\frac{u(1-\theta)}{\varepsilon_0} + 1 \right) \right. \\ &\quad \left. + \frac{u(1-\theta)}{\varepsilon_0} + \ln \left(\frac{\varkappa^2}{4} \right) + \frac{2\alpha}{\alpha-1} \left(\frac{2C_Z C_Y}{\theta \varepsilon_0} \right)^{\frac{1}{\alpha}} \right\}. \end{aligned}$$

Appendix A. Proof of Theorem 3.2

The derivation of the bound (3.6) for the process $U(t, x)$ is based on the particular structure of this process and on the use of the property (2.4) of admissible function Z .

Firstly note that the process $U(t, x)$ is separable since $U(t, x)$ is continuous with probability one. The process $U(t, x)$ is strictly φ -sub-Gaussian with the determining constant C_y , and, therefore, we can write:

$$\sup_{\substack{|t-t_1| \leq h \\ |x-x_1| \leq h}} \tau_\varphi(U(t, x) - U(t_1, x_1)) \leq C_y \sup_{\substack{|t-t_1| \leq h \\ |x-x_1| \leq h}} \left(\mathbf{E}(U(t, x) - U(t_1, x_1))^2 \right)^{1/2}.$$

We can also estimate $\varepsilon_0 = \sup_{(t,x) \in D} \tau_\varphi(U(t, x))$, where $D = \{a \leq t \leq b, c \leq x \leq d\}$, as follows:

$$\begin{aligned} \varepsilon_0 &\leq C_y \sup_{(t,x) \in D} \left(\mathbf{E}|U(t, x)|^2 \right)^{\frac{1}{2}} \\ &\leq C_y \sup_{(t,x) \in D} \left(\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |I(t, x, \lambda)I(t, x, \mu)| d|\Gamma_y(\lambda, \mu)| \right)^{\frac{1}{2}} \\ &\leq C_y \left(\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} d|\Gamma_y(\lambda, \mu)| \right)^{\frac{1}{2}} =: \Gamma. \end{aligned} \quad (\text{A.1})$$

It is enough to consider one of the representations for initial condition, $\eta(x) = \int_{-\infty}^{\infty} \sin xu dy(u)$ or $\eta(x) = \int_{-\infty}^{\infty} \cos xu dy(u)$, since the proofs are similar for both cases.

Let us consider $\mathbf{E}(U(t, x) - U(t_1, x_1))^2$ for $\kappa(u) = \cos(u)$:

$$\begin{aligned} \mathbf{E}(U(t, x) - U(t_1, x_1))^2 &= \left(\int_{-\infty}^{\infty} (I(t, x, \lambda) - I(t_1, x_1, \lambda)) dy(\lambda) \right)^2 \\ &\leq \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |I(t, x, \lambda) - I(t_1, x_1, \lambda)| |I(t, x, \mu) - I(t_1, x_1, \mu)| d|\Gamma_y(\lambda, \mu)|. \end{aligned}$$

We can write $|I(t, x, \lambda) - I(t_1, x_1, \lambda)| = |\cos A - \cos B|$, where

$$A = x\lambda + t \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k, \quad B = x_1\lambda + t_1 \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k.$$

Thus

$$\begin{aligned} |I(t, x, \lambda) - I(t_1, x_1, \lambda)| &= 2 \left| \sin \frac{A+B}{2} \sin \frac{B-A}{2} \right| \\ &\leq 2 \left| \sin \frac{B-A}{2} \right| = 2 \left| \sin(C+D) \right|, \end{aligned}$$

where

$$C = \frac{\lambda(x_1 - x)}{2}, \quad D = \frac{t_1 - t}{2} \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k.$$

Therefore,

$$\begin{aligned} 2|\sin(C + D)| &= 2|\sin C \cos D + \cos C \sin D| \leq 2(|\sin C| + |\sin D|) \\ &= 2\left(\left|\sin \frac{\lambda(x_1 - x)}{2}\right| + \left|\sin \frac{(t_1 - t)}{2} \sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k\right|\right). \end{aligned}$$

Choose some admissible function $\{Z(x), x \geq 0\}$ for the space $\text{Sub}_\varphi(\Omega)$. In view of Lemma 2.1, we can write

$$\begin{aligned} |I(t, x, \lambda) - I(t_1, x_1, \lambda)| &\leq 2Z^{-1}\left(\frac{1}{|x - x_1|} + u_0\right)Z\left(\frac{|\lambda|}{2} + u_0\right) \\ &\quad + 2Z^{-1}\left(\frac{1}{|t - t_1|} + u_0\right)Z\left(\frac{1}{2}\left|\sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k\right| + u_0\right). \end{aligned}$$

Thus, we obtain:

$$\begin{aligned} \sup_{\substack{|t-t_1| \leq h \\ |x-x_1| \leq h}} \tau_\varphi(U(t, x) - U(t_1, x_1)) &\leq C_y \sup_{\substack{|t-t_1| \leq h \\ |x-x_1| \leq h}} \left(\mathbf{E}(U(t, x) - U(t_1, x_1))^2\right)^{1/2} \\ &\leq C_y \frac{2}{Z\left(\frac{1}{h} + u_0\right)} \left[\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(Z\left(\frac{|\lambda|}{2} + u_0\right) + Z\left(\frac{1}{2}\left|\sum_{k=1}^N a_k \lambda^{2k+1} (-1)^k\right| + u_0\right) \right) \right. \\ &\quad \times \left. \left(Z\left(\frac{|\mu|}{2} + u_0\right) + Z\left(\frac{1}{2}\left|\sum_{k=1}^N a_k \mu^{2k+1} (-1)^k\right| + u_0\right) \right) d|\Gamma_y(\lambda, \mu)| \right]^{1/2} \\ &= C_y C_Z \frac{2}{Z\left(\frac{1}{h} + u_0\right)}. \end{aligned}$$

For $\kappa(u) = \sin u$ we have the same inequality. So, in the notations of Theorem 3.1

$$\sigma(h) = 2C_y C_Z \left(Z\left(\frac{1}{h} + u_0\right) \right)^{-1}, \quad 0 < h < \max(b - a, d - c).$$

Acknowledgement

We are grateful to editors and reviewers for valuable remarks and suggestions, which helped us to improve the paper significantly.

References

- [1] Beghin, L., Knopova, V.P., Leonenko, N.N., Orsingher, E.: Gaussian limiting behavior of the rescaled solution to the linear Korteweg-de-Vries equation with random initial conditions. *J. Stat. Phys.* **99**(3/4), 769–781 (2000). [MR1766908. https://doi.org/10.1023/A:1018687327580](https://doi.org/10.1023/A:1018687327580)

- [2] Beghin, L., Kozachenko, Yu., Orsingher, E., Sakhno, L.: On the Solutions of Linear Odd-Order Heat-Type Equations with Random Initial Conditions. *J. Stat. Phys.* **127**(4), 721–739 (2007). [MR2319850](#). <https://doi.org/10.1007/s10955-007-9309-x>
- [3] Buldygin, V.V., Kozachenko, Yu.V.: Metric characterization of random variables and random processes. *Translations of Mathematical Monographs*, vol. 188. AMS, American Mathematical Society, Providence, RI (2000). 257 pp. [MR1743716](#)
- [4] Giuliano Antonini, R., Kozachenko, Yu.V., Nikitina, T.: Spaces of φ -subgaussian random variables. *Rend. Accad. Naz. Sci. XL Mem. Mat. Appl.* **121** **XXVII**, 95–124 (2003). [MR2056414](#)
- [5] Kozachenko, Yu.V., Koval'chuk, Yu.A.: Boundary value problems with random initial conditions and series of functions of $Sub_\varphi(\Omega)$. *Ukr. Math. J.* **50**(4), 572–585 (1998). [MR1698149](#). <https://doi.org/10.1007/BF02487389>
- [6] Kozachenko, Yu., Olenko, A.: Whitaker-Kotelnikov-Shanon approximation of φ -sub-Gaussian random processes. *J. Math. Anal. Appl.* **442**(2), 924–946 (2016). [MR3514327](#). <https://doi.org/10.1016/j.jmaa.2016.05.052>
- [7] Kozachenko, Yu., Orsingher, E., Sakhno, L., Vasylyk, O.: Estimates for Functionals of Solutions to Higher-Order Heat-Type Equations with Random Initial Conditions. *J. Stat. Phys.* **172**(6), 1641–1662 (2018). [MR3856958](#). <https://doi.org/10.1007/s10955-018-2111-0>
- [8] Kozachenko, Yu.V., Ostrovskij, E.I.: Banach spaces of random variables of sub-Gaussian type. *Theory Probab. Math. Stat.* **32**, 45–56 (1986). [MR0882158](#)
- [9] Kozachenko, Yu.V., Rozora, I.V.: Simulation of Gaussian stochastic processes. *Random Oper. Stoch. Equ.* **11**(3), 275–296 (2003). [MR2009187](#). <https://doi.org/10.1163/156939703771378626>
- [10] Kozachenko, Yu.V., Rozora, I., Pogorilyak, O.O., Tegza, A.M.: *Simulation of Stochastic Processes with Given Accuracy and Reliability*. ISTE Press Ltd, London and Elsevier Ltd, Oxford (2017). 346 p. [MR3644192](#)
- [11] Krasnosel'skii, M.A., Rutickii, Ya.B.: *Convex Functions and Orlicz Spaces*. Noordhoff, Groningen (1961). 249 p. [MR0126722](#)
- [12] Loève, M.: *Probability Theory. Foundations. Random sequences*. D. van Nostrand Co., Inc. XV, New York (1955). 515 p. [MR0066573](#)
- [13] Orsingher, E., D'Ovidio, M.: Probabilistic representation of fundamental solutions to $\frac{\partial u}{\partial t} = \kappa_m \frac{\partial^m u}{\partial x^m}$. *Electron. Commun. Probab.* **17**, 1–12 (2012). [MR2965747](#). <https://doi.org/10.1214/ECP.v17-1885>
- [14] Rao, M.M.: Harmonizable, Cramér, and Karhunen classes of processes. In: Hannan, D.E.J., et al. (eds.) *Handbook of Statistics: Time Series in Time*, vol. 5, pp. 279–310. Elsevier, Amsterdam (1985). [MR0831752](#). [https://doi.org/10.1016/S0169-7161\(85\)05012-X](https://doi.org/10.1016/S0169-7161(85)05012-X)
- [15] Swift, R.J.: Covariance analysis and associated spectra for classes of nonstationary processes. *J. Stat. Plan. Inference* **100**(2), 145–157 (2002). [MR1877184](#). [https://doi.org/10.1016/S0378-3758\(01\)00129-X](https://doi.org/10.1016/S0378-3758(01)00129-X)
- [16] Tao, T.: *Nonlinear Dispersive Equations: Local and Global Analysis*. CBMS Regional Conference Series in Mathematics, vol. 106, American Mathematical Soc. (2006). 373 pp. [MR2233925](#). <https://doi.org/10.1090/cbms/106>
- [17] Vasylyk, O.I., Kozachenko, Yu.V., Yamnenko, R.: Upper estimate of overrunning by $Sub_\varphi(\Omega)$ random process the level specified by continuous function. *Random Oper. Stoch. Equ.* **13**(2), 111–128 (2005). [MR2152102](#). <https://doi.org/10.1163/156939705323383832>