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# Exploring the consequences of pairing algorithms for binary stars 

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#### Abstract

Knowledge of the binary population in stellar groupings provides important information about the outcome of the star forming process in different environments. Binarity is also a key ingredient in stellar population studies, and is a prerequisite for calibrating the binary evolution channels. In this paper we present an overview of several commonly used methods of pairing individual stars into binary systems, which we refer to as pairing functions. These pairing functions are frequently used by observers and computational astronomers, either for their mathematical convenience or because they roughly describe the expected outcome of the star forming process. We discuss the consequences of each pairing function for interpreting observations and numerical simulations. The binary fraction and mass ratio distribution generally depend strongly on the selection of the range in primary spectral type in a sample. The mass ratio distribution and binary fraction derived from a binarity survey among a mass-limited sample of targets is thus not representative of the population as a whole. Neither theory nor observations indicate that random pairing of binary components from the mass distribution, the simplest pairing function, is realistic. It is more likely that companion stars are formed in a disc around a star or that a pre-binary core fragments into two binary components. The results of our analysis are important for (i) the interpretation of the observed mass ratio distribution and binary fraction for a sample of stars; (ii) a range of possible initial condition algorithms for star cluster simulations; and (iii) how to distinguish between the different star formation scenarios.


Key words. stars: binaries: general - stars: formation - methods: $N$-body simulations -
Galaxy: open clusters and associations: general

## 1. Introduction

Observations and simulations suggest that most stars form in binary systems (e.g. Duquennoy \& Mayor 1991; Mason et al. 1998; Goodwin \& Kroupa 2005; Kobulnicky \& Fryer 2007; Kouwenhoven et al. 2005, 2007b; Goodwin et al. 2007), and that a substantial fraction are part of a triple or higher-order system (e.g. Tokovinin \& Smekhov 2002; Correia et al. 2006; Tokovinin et al. 2006; Hu et al. 2008). Multiplicity is thus a fundamental property of the star forming process. Detailed knowledge of a young binary population can be employed to study the outcome of star formation, consequently the star formation process itself.

Surveys indicate that the properties of the binary population are a function of the spectral type of the primary star. Practically all O-type stars (Mason et al. 1998) and B/A-type stars (Shatsky \& Tokovinin 2002; Kobulnicky \& Fryer 2007; Kouwenhoven et al. 2007b) are found in binary or multiple systems. Abt \& Levy (1976) report a multiplicity fraction of $55 \%$ among F3-G2 stars, and in their CORAVEL spectroscopic study of F7-G9 stars, Duquennoy \& Mayor (1991) find a binary fraction of $\sim 60 \%$. The binary fraction among M-type stars is $30-40 \%$ (Fischer \& Marcy 1992; Leinert et al. 1997; Reid \& Gizis 1997). For late M-type stars and brown dwarfs the binary fraction decreases to 10-30\% (e.g., Gizis et al. 2003; Close et al. 2003; Bouy et al. 2003; Burgasser et al. 2003; Siegler et al. 2005; Ahmic et al. 2007; Maxted et al. 2008; Joergens 2008).

In this paper we discuss in detail several methods of pairing individual stars into binary stars. We refer to the latter algorithms as "pairing functions". Several of these have a physical motivation, others are discussed because of their mathematical simplicity. All these pairing functions have in common that they are frequently used in the literature. The main goal of this paper is to explain the consequences of adopting a particular pairing function when doing a numerical simulation, or when interpreting observations. A good understanding of the consequences of each pairing algorithm for the binary population is important for

- The interpretation of observations. Are the measured properties (e.g. mass ratio distribution and binary fraction) for a certain sample representative of the population as a whole? What is the role of selection effects, and how can the different pairing functions be distinguished?
- Initial conditions for simulations. In $N$-body simulations, such as STARLAB (e.g., Portegies Zwart et al. 2001) and NBODY6 (e.g., Aarseth 1999), a mass ratio distribution independent of primary mass is often adopted. What are the consequences of this approach?
- Star formation. Which pairing functions are expected from the different star forming scenarios? What numerical simulations of clustered star formation (e.g., Bate et al. 2003; Bate 2008) predict? Do we expect random pairing from the initial mass function? Which observations are necessary to be able to distinguish between these scenarios?

This paper is organised as follows. In Sect. 2 we introduce the terminology used and discuss our assumptions. In Sect. 3 we discuss the mass distribution, and in Sect. 4 the mass ratio distribution. We briefly describe the origin of the pairing between binary stars in Sect. 5. The major part of this paper is Sect. 6, where we discuss the different binary populations resulting from the choice of the mass distribution and pairing function. The specific differences are discussed in detail in Sect. 7, and the dependence on the generating properties in Sect. 8. Observational complications and a strategy to recover the pairing function are described in Sect. 9. Finally, we summarise our results in Sect. 10.

## 2. Method and terminology

We study the differences in the various methods of pairing individual stars into binary systems by analysing numerically simulated binary populations. For each binary system, we refer to the most massive component $M_{1}$ as the primary star, and the least massive component $M_{2}$ as the companion star. Our adopted definition is purely based on the current mass of the components, i.e., irrespective of their relative luminosity or initial mass. We define the mass ratio as $q \equiv M_{2} / M_{1}$, so that $0<q \leq 1$. The total mass is denoted $M_{\mathrm{T}}=M_{1}+M_{2}$ for a binary system, and $M_{\mathrm{T}}=M_{\mathrm{S}}$ for a single star with mass $M_{\mathrm{S}}$. In several cases, we construct binary systems from star forming cores of mass $M$ with star forming efficiency $\epsilon$, so that the total mass of the resulting objects is $M_{\mathrm{T}}=\epsilon M_{\mathrm{C}}$.

In our simulations each star is given a mass. A subset of the stars is assigned a companion, the other stars remain single stars. The companion is given a mass according to a pairing algorithm. We refer to the algorithm that is used to combine individual stars into binary systems as the pairing function. The pairing function of a binary population may for example be random pairing of both companions from the mass distribution. In the particular case of random pairing, primary and companion are swapped, if necessary, so that the primary is the most massive star.

Depending on which pairing function is used, the mass of the primary star, and in several cases the companion star, is drawn from a mass distribution $f_{M}(M)$. We refer to $f_{M}(M)$ as the generating mass distribution. We denote the resulting mass distributions for primary stars, companion stars, systems, and single stars with $f_{M_{1}}(M), f_{M_{2}}(M), f_{M_{\mathrm{T}}}(M)$, and $f_{M_{\mathrm{S}}}(M)$, respectively. Note that the distributions over primary and companion mass in a stellar grouping are never independent, $f_{M_{1}, M_{2}}\left(M_{1}, M_{2}\right) \neq$ $f_{M_{1}}\left(M_{1}\right) f_{M_{2}}\left(M_{2}\right)$, as by definition $M_{1} \geq M_{2}$. The resulting mass distribution for all single and primary stars is denoted with $f_{M_{1, S}}(M)$, and the mass distribution for all individual objects (i.e., singles, primaries and companions), is denoted with $f_{\text {all }}(M)$. The distribution $f_{\text {all }}(M)$ that is present immediately after star formation is called the initial mass function (IMF). Note that $f_{\text {all }}(M)$ is unequal to the generating mass distribution $f_{M}(M)$, except in the random pairing case (see Sect. 6.1).

For several pairing functions we generate companions by drawing the mass ratio for a binary from a (generating) mass ratio distribution $f_{q}(q)$ Depending on the additional constraints specific for each pairing function, the resulting overall mass ratio distribution $f_{q, \text { all }}(q)$ may or may not be equal to $f_{q}(q)$. The resulting mass ratio distribution may or may not be a function of primary spectral type. We refer to the specific mass ratio distribution for all binaries with a primary of spectral type A or B, for example, with $f_{q, \mathrm{AB}}(q)$. Throughout this paper we will mostly use the expression $f_{q ; M_{1}}(q)$ for the specific mass ratio distribution for an ensemble of binaries with a limited primary mass range.

A common expression to quantify the multiplicity of a stellar population is the multiplicity fraction $\mathcal{B}$ (which is often referred to as the binary fraction), defined as
$\mathcal{B}=\frac{B+T+\ldots}{S+B+T+\ldots}$
where $S$ is the number of single stars, $B$ the number of binaries, and $T$ the number of triple systems (e.g., Reipurth \& Zinnecker 1993). Throughout this paper we consider only single and binary stars, and do not consider higher-order systems, so that $\mathcal{B}=B /(S+B)$. The number of systems is given by $N=S+B$ and the total number of (individual) stars is $S+2 B=N(\mathcal{B}+1)$.

Each pairing algorithm is provided with a (generating) binary fraction $\mathcal{B}$, which describes the fraction of stars that is assigned a preliminary companion star. For most pairing functions this preliminary companion is accepted as such, so that the overall binary fraction $\mathcal{B}_{\text {all }}$ equal to the generating value $\mathcal{B}$. For several pairing functions, however, additional constraints are set to the properties of the companion. For example, when a pairing function generates a Jupiter-mass companion around a solarmass star, the "primary" is usually considered to be a "single" star. For such pairing functions, the resulting overall binary fraction $\mathcal{B}_{\text {all }}$ is lower than $\mathcal{B}$. We denote the specific binary fraction for the set of all single stars of spectral type $A / B$ and binary systems with primary spectral type $\mathrm{A} / \mathrm{B}$ as $\mathcal{B}_{\mathrm{AB}}$.

During our analysis we make several assumptions for reasons of clarity; our models are simplifications of reality. Our results are not limited by these assumptions, and the models can easily be extended. Our main goal is to illustrate the implications of adopting a particular pairing function.

We assume that no triples or higher-order systems are present. Although observations have shown that a significant fraction ( $\gtrsim 15 \%$ ) of the stars are part of a multiple system (Tokovinin \& Smekhov 2002; Correia et al. 2006; Tokovinin et al. 2006; Eggleton \& Tokovinin 2008; Hu et al. 2008), the properties of higher-order systems are not well understood. Observational selection effects complicate the derivation of the properties of these systems significantly. Higher-order systems are often ignored in $N$-body simulations due to computational complications (e.g., van den Berk et al. 2007). A full understanding of star formation, however, ultimately requires a full knowledge of the formation and evolution of higher-order systems.

In our models the generating binary fraction for the population can be described with a single value $\mathcal{B}$. For most pairing functions this results in a specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ that is independent of primary mass $M_{1}$. However, in Sect. 6 we describe several cases where, as a result of the pairing properties, the binary fraction is a function of primary mass, even though this dependency is not included explicitly.

Selection effects play a major role in the interpretation of binary star observations. A detailed description of the selection effects, such as in Kouwenhoven et al. (2007b), is necessary to derive the pairing function, the mass ratio distribution and the binary fraction from observations. A major bias is generally introduced by studying the binary population in a certain primary mass range; we describe this effect in detail for the different pairing functions. Throughout most of this paper we ignore the other selection effects.

## 3. The mass distribution

The mass distribution $f_{M}(M)$ defines the spectrum of masses in a stellar population, and is usually expressed as a singlecomponent power law (Salpeter 1955), a multi-component


Fig. 1. The cumulative mass distributions (top) and mass distributions (bottom) derived by Kroupa (Eq. (2); solid curve) and Salpeter (Eq. (3); dashed curve). The masses at which the slope of the Kroupa mass distribution changes are indicated with the vertical dotted lines.

Table 1. The distribution of stars over main-sequence spectral type for the Kroupa and Salpeter mass distributions ( $\mathrm{BD}=$ brown dwarfs).

| SpT | Mass range | Kroupa |  | Salpeter |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | $\left(M_{\odot}\right)$ | $\mathcal{F}_{N}(\%)$ | $\mathcal{F}_{M}(\%)$ | $\mathcal{F}_{N}(\%)$ | $\mathcal{F}_{M}(\%)$ |
| B | $3.0-20$ | 1.45 | 24.76 | 0.10 | 8.66 |
| A | $1.5-3.0$ | 2.28 | 13.05 | 0.18 | 5.31 |
| F | $1.0-1.5$ | 2.72 | 9.15 | 0.22 | 3.86 |
| G | $0.8-1.0$ | 2.22 | 5.52 | 0.18 | 2.30 |
| K | $0.5-0.8$ | 7.45 | 12.92 | 0.61 | 5.45 |
| M | $0.08-0.5$ | 51.48 | 30.28 | 14.09 | 32.07 |
| BD | $0.02-0.08$ | 32.41 | 4.33 | 84.61 | 42.34 |
| All | $0.02-20$ | 100.00 | 100.00 | 100.00 | 100.00 |

power law (e.g., Kroupa 2001) or a Gaussian distribution (e.g., Chabrier 2003). In our analysis we consider two mass distributions: the Kroupa mass distribution and the Salpeter mass distribution. The main difference between these is the presence or absence of a turnover in the low-mass regime. The mass distribution derived by Kroupa (2001) is given by
$f_{\text {Kroupa }}(M) \propto\left\{\begin{array}{l}M^{-0.3} \text { for } 0.01 \leq M / M_{\odot}<0.08 \\ M^{-1.3} \text { for } 0.08 \leq M / M_{\odot}<0.5 \\ M^{-2.3} \text { for } 0.5 \leq M / M_{\odot}<\infty\end{array}\right.$.
The classical Salpeter mass distribution (Salpeter 1955) is given by
$f_{\text {Salpeter }}(M) \propto M^{-2.35} \quad M \gtrsim 1 M_{\odot}$.
The Salpeter mass distribution is derived for intermediate-mass stars in the Galactic field. Although it is known that the Salpeter mass distribution is incorrect for masses below $\sim 1 M_{\odot}$, we use this mass distribution for comparison, to illustrate the effect of the slope of the mass distribution for low-mass stars. Note that, although, Eq. (2) introduces a turnover in Fig. 1, which is logarithmically plotted, there is no real turnover in the mass distribution (see also Appendix A.1.3).

Table 1 lists the fraction of stars of a given spectral type for the Kroupa and Salpeter mass distributions for stars in the mass range $0.02 M_{\odot} \leq M \leq 20 M_{\odot}$. For both distributions we list the fraction $\mathcal{F}_{N}$ of objects and the fraction $\mathcal{F}_{M}$ of the total
mass in each mass range. The corresponding probability distributions $f_{M}(M)$ and cumulative distributions $\int_{0}^{M} f_{M}\left(M^{\prime}\right) \mathrm{d} M^{\prime}$ are shown in Fig. 1. Most young stellar populations in our Galaxy are described accurately with the Kroupa mass distribution (Bonnell et al. 2007).

In our simulations we draw objects from the mass distribution in the range $0.02 M_{\odot} \leq M \leq 20 M_{\odot}$, i.e., stellar and substellar-mass objects. Objects less massive than $M_{\min }=$ $0.02 M_{\odot}$ are considered to be planets, and form in a different way than stars or brown dwarfs (see, e.g., Pollack et al. 1996; Kouwenhoven et al. 2007a). The absolute maximum stellar mass is of order $M_{\text {max,abs }} \approx 150 M_{\odot}$ (Zinnecker \& Yorke 2007, and references therein). The most massive star in a cluster may well depend on the total mass of the star cluster (Reddish 1978; Vanbeveren 1982; Weidner \& Kroupa 2006). Stars more massive than $20 M_{\odot}$, however, are extremely rare, are very short-lived, and possibly even form by a different mechanism than most other stars (e.g., Zinnecker \& Yorke 2007).

The mass distribution of Kroupa (2001) is for all stars in a population, including single stars, primaries and companions. We note that choosing primaries from an IMF and then choosing secondaries from a mass ratio distribution (Sect. 6.3) will not recover the original IMF. Thus the primary mass distribution function cannot be exactly the same as the desired IMF (see Sect. 7.1; Mal'Kov \& Zinnecker 2001; Goodwin et al. 2008). Technically, the mass distribution should therefore not be used to generate, for example, a primary mass distribution. For simplicity, however, we adopt the Kroupa mass distribution as the generating mass distribution for each pairing function. Ideally, one should iteratively determine the generating mass distribution by comparing the outcome of the pairing process with the Kroupa mass distribution. As this is computationally a very expensive exercise, we skip the iteration, and simply adopt the Kroupa mass distribution as the generating mass distribution. For a proper analysis of real observations, one should keep this issue in mind.

## 4. The mass ratio distribution

The mass ratio distribution describes the distribution over mass ratio $q=M_{2} / M_{1}$ for a population of binary systems. The mass ratio distribution for binary systems has been studied thoroughly over the last decades (see, e.g., Zinnecker 1984; Hogeveen 1992a; Mazeh et al. 2003; Halbwachs et al. 2003). In this paper we distinguish between three different types of mass ratio distributions. The generating mass ratio distribution is an input distribution that is used by most pairing function algorithms to generate binaries, although some pairing functions (e.g., random pairing) do not require a generating mass ratio distribution. Note that the mass ratio distribution is not the same as the pairing function; the mass ratio distribution is a property of several (not all) pairing functions; see Sect. 6 for details. The overall mass ratio distribution is the mass ratio distribution resulting from the pairing mechanism, for all binaries in the population. The specific mass ratio distribution is that for a sample of stars with primaries in a given mass range. The latter is measured in observations, as a binarity survey is in most cases focused on a particular set of targets with given spectral types.

The (specific) mass ratio distribution is usually obtained by a fit to the observed mass ratio distribution of the sample. The observed distribution is often described with a simple power-law:
$f_{\gamma_{q}}(q) \propto q^{\gamma_{q}}$ for $q_{0}<q \leq 1$,
where the exponent $\gamma_{q}$ is fitted (e.g. Hogeveen 1992b; Shatsky \& Tokovinin 2002; Kobulnicky \& Fryer 2007;

Kouwenhoven et al. 2007b). Distributions with $\gamma_{q}=0$ are flat, while those with $\gamma_{q}<0$ and $\gamma_{q}>0$ are falling and rising with increasing $q$, respectively. Usually the adopted minimum value for the fit $q_{0}$ is the value of $q$ below which the observations become incomplete due to selection effects. For distributions with $\gamma_{q} \leq-1$, a minimum value $q_{0}>0$ is necessary such that $f_{q}(q)$ can be normalised. Sometimes the necessity of $q_{0}>0$ is avoided by fitting a mass ratio distribution of the form
$f_{\Gamma_{q}}(q) \propto(1+q)^{\Gamma_{q}} \quad$ for $\quad 0<q \leq 1$
to the data (e.g. Kuiper 1935; van Rensbergen et al. 2006). This distribution is more commonly used to describe the mass ratio distribution of high-mass spectroscopic binaries, while the distribution in Eq. (4) is often used for visual binaries. As Eqs. (4) and (5) show a similar behaviour (both are either falling, flat, or rising), we will only consider the distribution of Eq. (4).

Alternatively, to allow for a mass ratio distribution with a peak in the range $0<q \leq 1$, we also consider the Gaussian mass ratio distribution:
$f_{\text {Gauss }}(q) \propto \exp \left\{-\frac{\left(q-\mu_{q}\right)^{2}}{2 \sigma_{q}}\right\} \quad$ for $\quad 0<q \leq 1$,
where $\mu_{q}$ and $\sigma_{q}$ are free parameters, corresponding to the mean and standard deviation of a Gaussian distribution without the imposed limits on $q$. Models with $\mu_{q}<0$ and $\mu_{q}>1$ show a distribution $f_{\text {Gauss }}(q)$ with an exclusively negative and positive slope, respectively. In the case where $\mu_{q} \ll 0$ and $\mu_{q} \gg 1$, the distribution may be approximated by a power-law (Eq. (4)). Models with $\sigma_{q} \gg 0.5$ can be approximated with a flat mass ratio distribution. Note that the values of $\mu_{q}$ and $\sigma_{q}$ do not necessarily reflect physical properties. A value $\mu_{q} \ll 0$, for example, merely means that the mass ratio distribution in the interval $0<q \leq 1$ can be described by Eq. (6) in this interval. A Gaussian mass ratio distributions was reported by, Duquennoy \& Mayor (1991) for solar-type stars in the solar neighbourhood, who find, based on the work on the initial mass function by Kroupa et al. (1990), $\mu_{q}=0.23$ and $\sigma_{q}=0.42$.

For our default model we adopt a (generating) mass ratio distribution of the form $f_{q}(q)=1$ with $0<q \leq 1$, the flat mass ratio distribution.

## 5. The origin of the pairing function

We define the pairing function as the algorithm which is used to pair individual stars into binary systems. A well-known pairing function is random pairing from the mass distribution. Others include, for example, a fixed mass ratio distribution. The pairing of binary components in a stellar population results from the combined effect of star formation, stellar evolution, binary evolution, and dynamical interactions. By studying the pairing of binary stars, the contributions of the latter three effects can be evaluated, and an estimate for the primordial binary population can be obtained. This primordial pairing function allows us to constrain the process of star formation.

It is worth considering what pairing we might expect from the actual star formation process, as opposed to the various theoretical constructs we describe in this paper.

Often random pairing has been used to construct binary systems for various models. Random pairing has the obvious advantage that the chosen IMF is, by design, automatically recovered. However, there is no good theoretical reason to suppose that the star formation process would produce a randomly
paired distribution. Furthermore, random pairing is ruled out observationally. The observed mass ratio distribution among intermediate mass stars (Shatsky \& Tokovinin 2002; Kouwenhoven et al. 2005, 2007a,b) and brown dwarfs (Kraus et al. 2008) in the nearby OB association Scorpius-Centaurus have indicated that the binary components are not randomly paired from the mass distribution. The same result is found for the Cygnus OB2 association (Kobulnicky \& Fryer 2007), and for F5-K5 stars in the solar neighbourhood (Metchev \& Hillenbrand 2008). Random pairing is further excluded by the high prevalence of massive binaries with a mass ratio close to unity, often referred to as the "twin peak" in the mass ratio distribution (Lucy \& Ricco 1979; Tokovinin 2000; García \& Mermilliod 2001; Pinsonneault \& Stanek 2006; Lucy 2006; Söderhjelm 2007, see also Sect. 9.1).

Simulations have shown that it is impossible to form significant numbers of binary systems from an initially single star distribution (Kroupa 1995a), therefore stars in binary systems must have predominantly formed in binary systems. Observations of pre-main sequence stars also suggest that they have a higher multiplicity than field stars (at least for $>1 M_{\odot}$ ), suggesting that most stars form in binary (or higher-order) systems (e.g., Goodwin \& Kroupa 2005). It is supposed that a primordial population with a multiplicity of almost $100 \%$ evolves into a field-like binary population through (a) the decay of higher-order multiple systems (e.g., Goodwin \& Kroupa 2005), and (b) the dynamical destruction of binaries in binary-binary encounters in clusters (e.g., Kroupa 1995a,c). Thus, the currently observed binary population is a complex mixture of primordial binaries (i.e. in the same dynamical state as when they formed), and dynamically evolved binaries (which may have different characteristics, or even companions to their initial state).

Simulations of binary star formation have comprehensively failed to produce systems that match observations, even when the dynamical evolution of the initial states is accounted for (see Goodwin et al. 2007, and references therein). However, hydrodynamic simulations of star formation suggest that companions usually form by the fragmentation of massive, disc-like circumstellar accretion regions around young stars (see Goodwin et al. 2007, and references therein).

In such a scenario for companion formation it would be expected that the secondary should have a roughly similar (i.e. within a factor of three or four) mass to the primary, especially at small separations. A massive enough region to fragment is only present during the earliest (e.g. class 0/I) phases of star formation before the star(s) have accreted the majority of their natal core. Thus, the secondary will be present whilst a large reservoir of gas is also present around it. In the case of a star that will eventually grow to be (for example) $5 M_{\odot}$ the secondary will form whilst the primary is only $O\left(1 M_{\odot}\right)$ and several solar masses of gas are present in the circumstellar environment. A secondary will presumably form with an initial mass close to the opacity limit for fragmentation, $\sim 10^{-2} M_{\odot}$. However, it is difficult to imagine a scenario in which the secondary fails to accrete at least some of the circumstellar material, especially as the secondary will form with angular momentum similar to that of the accretion region, and so will be more able to accrete material (see Goodwin et al. 2007, and references therein; also Delgado Donate \& Clarke 2005). Therefore, the secondary mass is expected to be a reasonably high fraction of the primary mass. In particular, it should be difficult for a companion to a B-star to remain at brown dwarf or M -star mass due to the large amount of material available for accretion. In particular, we would expect a rough correlation between separation and mass ratio, with
closer companions being generally more massive as is observed (e.g. Mazeh et al. 1992, for field G-dwarfs).

In addition, dynamical evolution in clusters will act to destroy the most weakly bound systems (i.e. the widest and lowestmass companions), further biasing the mass ratio distributions away from small- $q$.

We would therefore argue that random pairing over the full mass range is the last type of pairing that would be expected from the star formation process (see, however, Sect. 6.1.1).

## 6. Analysis of frequently used pairing functions

There are many ways to obtain a population of binary systems from a mass distribution $f_{M}(M)$. We analyse the most frequently used algorithms in the sections below. In general, the masses of the members of a stellar grouping are drawn from the mass distribution. A fraction of the stars is assigned a companion star (either from the mass distribution, or using a mass ratio distribution), or the mass is split into a primary and companion star. Four commonly used mechanisms are the following:

- Random pairing (RP). The masses of both binary components are randomly drawn from the mass distribution $f_{M}(M)$. For each system, the most massive component is labelled "primary star", the other component "companion star" (see Sect. 6.1).
- Primary-constrained random pairing (PCRP). The primary mass $M_{1}$ is drawn from the mass distribution $f_{M}(M)$, and the companion mass $M_{2}$ is chosen from the same mass distribution, with the additional constraint that $M_{2} \leq M_{1}$ (see Sect. 6.2).
- Primary-constrained pairing (PCP). The primary mass is drawn from the mass distribution $f_{M}(M)$. The companion mass is then determined by a mass ratio that is drawn from a distribution $f_{q}(q)$, with $0<q \leq 1$ (see below).
- Split-core pairing (SCP). The total mass of the binary is drawn from the mass distribution $f_{M}(M)$. The mass ratio of the binary is then determined by a mass ratio distribution which is drawn from $f_{q}(q)$ with $0<q \leq 1$. Finally, the primary mass $M_{1}=M(1+q)^{-1}$ and companion mass $M_{2}=M\left(1+q^{-1}\right)^{-1}$ are determined (see below).

In the case of PCP and SCP there is another complication, which occurs if $M_{2}=q M_{1}<M_{2, \text { min }}$. In this case, the mass ratio distribution generates a companion mass smaller than the permitted value $M_{2, \text { min }}$, for example the deuterium-burning limit, while such objects are usually not considered as companions in a binary system. There are three straightforward choices on how to handle such companions:

- Accept all companions (PCP-I/SCP-I). All companions are accepted, regardless of their mass. The resulting mass ratio distribution obtained with this method is equal to the generating mass ratio distribution (see Sects. 6.3.1 and 6.4.1).
- Reject low-mass companions (PCP-II/SCP-II). If $M_{2, \min } / M_{1, \max }<q_{\min }$, a fraction of the companions has $M_{2}<M_{2, \min }$. All companions with $M_{2}<M_{2, \min }$ are rejected, and the corresponding "primaries" are classified as single stars (see Sects. 6.3.2, 6.4.2, and below).
- Redraw low-mass companions (PCP-III/SCP-III). For all binaries with a companion mass $M_{2}<M_{2 \text {,min }}$ the mass ratio is redrawn from $f_{q}(q)$. This procedure is repeated until $M_{2} \geq$ $M_{2, \min }$. This method is equivalent to drawing a mass ratio from the distribution $f_{q}(q)$ with limits $M_{2, \min } / M_{1} \leq q \leq 1$ (see Sects. 6.3.3, 6.4.3, and below).

It is possible in pairing function SCP-II that the resulting single star mass is smaller than $M_{\text {min }}$, and for SCP-III the splitting up is not possible if the binary system mass is smaller than $2 M_{\text {min }}$. In these cases there are three possibilities of dealing with this problem:

- Accept all singles/primaries (SCP-IIa/SCP-IIIa). If a pairing mechanism produces a single or primary star less massive than $M_{2, \min }$, it is accepted, and included in the model.
- Reject low-mass singles/primaries (SCP-IIb/SCP-IIIb). If a pairing mechanism produces a single or primary star less massive than $M_{2, \min }$, it is rejected and removed from the model.
- Do not split-up low-mass cores (SCP-IIc/SCP-IIIc). Cores with a mass $M_{\mathrm{T}}<2 M_{\text {min }}$ are not split up; these become single stars of mass $M_{\mathrm{S}}=M_{\mathrm{T}}$.

For reasons of simplicity and clarity, we use a generating mass distribution with a minimum value of $2 M_{\text {min }}$ for the SCP pairing functions, and adopt a minimum mass $M_{\min }$ for the companions. Implicitly, we therefore only consider the variants SCP-IIb and SCP-IIIb in this paper. From hereon, we use "SCP-II" and "SCPIII" to refer to SCP-IIb and SCP-IIIb, respectively.

The above-mentioned pairing functions (RP, PCRP, and the three variations of PCP and SCP) are described in detail in the subsections below, while their differences are discussed in Sect. 7. Unless stated otherwise, we have adopted a Kroupa generating mass distribution, a flat generating mass ratio distribution (where applicable), and a generating binary fraction $\mathcal{B}$ of $100 \%$ (see Sect. 8 for a discussion of these assumptions). The main differences between the (renormalised) resulting mass ratio distributions are show in Fig. 2, where the top panel represents the overall mass ratio distribution. The middle and bottom panels represent the specific mass ratio distributions $f_{q ; M_{1}}(q)$ (the subscript $M_{1}$ indicates a restricted primary mass range) for binaries with high-mass primaries and low-mass primaries, respectively. Note that the derived mass ratio distribution for a sample of stars does not only depend strongly on the pairing function, but also on the targeted sample of stars. For the same reason, the (renormalised) companion mass distribution, shown in Fig. 3, depends strongly on the primary mass range. A two-dimensional version of Fig. 2 is shown in Fig. 4. In the sections below we discuss in detail the pairing functions and the above-mentioned figures.

Note that the choices made in this paper do not imply that stellar populations indeed have these properties. It is not known how binary stars are formed, so that no robust predictions of their properties can be made. Different binary formation mechanisms may produce different mass ratio distributions, possibly varying with primary mass, period, or eccentricity (e.g. Heggie 1975; Krumholz \& Thompson 2007; Zinnecker \& Yorke 2007). In addition, dynamical evolution after the formation process may alter the binary fraction and the mass and mass ratio distributions (e.g., Hills 1975; Heggie 1975), possibly as a function of environment (see, e.g., Kroupa et al. 1999; Preibisch et al. 2003; Duchêne et al. 2004; Köhler et al. 2006; Reipurth et al. 2007; Gaburov et al. 2008, and Sect. 9.2). Other pairing functions suggested in literature include random pairing over a restricted mass range (Kroupa 1995a, c, b; Kroupa et al. 2003; Thies \& Kroupa 2007, see also Sect. 6.1.1), "special pairing" (Weidner et al. 2008), ordered pairing (Oh et al, in prep.), pairing resulting from the dissolution of small- $N$ clusters (Clarke 1996b), "two-step" pairing (Durisen et al. 2001), binary formation from ring fragmentation (Hubber \& Whitworth 2005), and numerous others.


Fig. 2. The (renormalised) mass ratio distributions resulting from the different pairing functions. From top to bottom, the panels show the overall mass ratio distribution (i.e., for all binaries), the mass ratio distribution for binaries with $1.5 M_{\odot} \leq M_{1} \leq 20 M_{\odot}$, and for binaries with $0.02 M_{\odot} \leq$ $M_{1} \leq 0.08 M_{\odot}$. The models consist of $N=S+B=5 \times 10^{5}$ particles and all have a generating binary fraction of $\mathcal{B}=100 \%$. For each model we adopt a Kroupa generating mass distribution in the mass range $0.02-20 M_{\odot}$, and, when applicable, a flat generating mass ratio distribution $f_{q}(q)=1(0<q \leq 1)$. This figure illustrates that each pairing function results in a different overall or specific mass ratio distribution (see also Figs. 8 and 7).


Fig. 3. The (renormalised) companion mass distributions resulting from the different pairing functions, for the models shown in Fig. 2. The curves indicate the distribution for all binaries (dashed curves), for binaries with $1.5 M_{\odot} \leq M_{1} \leq 20 M_{\odot}$ (dotted curves) and for binaries with $0.02 M_{\odot} \leq M_{1} \leq 0.08 M_{\odot}$ (solid curves).

### 6.1. Pairing function RP (random pairing)

In the case of random pairing (RP), the primary and companion mass are both independently drawn from $f_{M}(M)$, and swapped, if necessary, so that the most massive star is the primary. As a result of this swapping, neither the resulting primary mass distribution $f_{M_{1}}\left(M_{1}\right)$, nor the companion mass distribution $f_{M_{2}}\left(M_{2}\right)$, nor the system mass distribution $f_{M_{\mathrm{T}}}\left(M_{\mathrm{T}}\right)$ is equal to the generating mass distribution $f_{M}(M)$; see, e.g., Warner (1961); Tout (1991); Mal'Kov \& Zinnecker (2001). On the other hand, the mass distribution of all stars $f_{\text {all }}(M)$, i.e., all singles, primaries and companions, is equal to $f_{M}(M)$. Mal' Kov \& Zinnecker (2001) derived general expressions for the distribution over primary star mass $M_{1}$, companion star mass $M_{2}$, and system mass $M_{\mathrm{T}}=M_{1}+M_{2}$, respectively:
$f_{M_{1}}\left(M_{1}\right)=2 f_{M}\left(M_{1}\right) \int_{\mathrm{c}}^{M_{1}} f_{M}(M) \mathrm{d} M$
$f_{M_{2}}\left(M_{2}\right)=2 f_{M}\left(M_{2}\right) \int_{M_{2}}^{d} f_{M}(M) \mathrm{d} M$
$f_{M_{\mathrm{T}}}\left(M_{\mathrm{T}}\right)=2 \int_{\mathrm{c}}^{M_{\mathrm{T}}-c} f_{M}(M) f_{M}\left(M_{\mathrm{T}}-M\right) \mathrm{d} M$.
Below we calculate the mass ratio distributions resulting from random pairing. Let the generating mass distribution $f$ be defined on the interval $[c, d]$. For random pairing we draw two stars with masses $x$ and $y$ from $f$, and swap, if necessary, so that the primary is the most massive star. The resulting mass ratio distribution is bounded by $0<c / d \leq q \leq 1$. To derive the overall
mass ratio distribution we follow the appendix of Piskunov \& Mal'Kov (1991) and derive $f_{q}(q)$ from its cumulative distribution function $F_{q}(q)=P(y / x \leq q)$ :

$$
\begin{equation*}
F_{q}(q)=\iint_{\mathrm{S}} f_{\mathrm{rp}}(x, y) \mathrm{d} x \mathrm{~d} y=\int_{c / q}^{d} \int_{\mathrm{c}}^{q x} 2 f(x) f(y) \mathrm{d} y \mathrm{~d} x \tag{10}
\end{equation*}
$$

where the factor 2 accounts for the fact that the pairs of masses are swapped in order to ensure that $y \leq x$. The integration domain is shown in the left-hand panel of Fig. 5. The probability density $f_{q}$ for $q$ (i.e., the overall mass ratio distribution) is then given by the derivative of Eq. (10).

For observational reasons, surveys for binarity are often restricted to a certain range of primary spectral types. To derive the effects of the selection on primary mass, the derivation of $f_{q}(q)$ again proceeds via the cumulative distribution function, which is now given by:
$F_{q}(q)=P\left(y / x \leq q \mid x_{1} \leq x \leq x_{2}\right) \propto \iint_{S^{\prime}} f_{\mathrm{rp}}(x, y) \mathrm{d} x \mathrm{~d} y$,
where the primary mass range is restricted to the range $\left[x_{1}, x_{2}\right]$. The integration domain $S^{\prime}$ is as shown in Fig. 5. The integration limits now depend on whether $q$ is larger or smaller than $c / x_{1}$. For $q \leq c / x_{1}$ the integration domain $S^{\prime}$ is given by $c / q \leq x \leq x_{2} \wedge c \leq y \leq q x$ (middle panel in Fig. 5), while for $q>c / x_{1} S^{\prime}$ is defined by $x_{1} \leq x \leq x_{2} \wedge c \leq y \leq q x$ (right panel in Fig. 5). The minimum possible value of $q$ is $c / x_{2}$, so that $F(q)=0$ for $q<c / x_{2}$. The general expression for the cumulative


Fig. 4. The distribution of binary systems over primary mass $M_{1}$ and mass ratio $q$ for the different pairing functions. All populations have a Kroupa generating mass distribution, a generating mass ratio distribution of the form $f_{q}(q)=1$ (when applicable), and a generating binary fraction $\mathcal{B}$ of $100 \%$. From left to right, the panels show the results for pairing functions PCP-I, PCP-II, PCP-III and RP in the top row, and SCP-I, SCP-II, SCP-III and PCRP on the bottom row. From light to dark, the contours bound values of $0.01 \%, 0.02 \%, 0.05 \%$ (thick curve), $0.1 \%$ and $0.2 \%$ for $\left(q, \log M_{1}\right)$. Each model contains $10^{6}$ binaries. Each dot represents a bin in mass ratio and logarithmic primary mass. No selection effects have been applied.


Fig. 5. The integration domains for pairing functions RP and PCRP. Left: the integration domain in the ( $x, y$ ) plane (where $x$ and $y$ are the primary and companion mass, respectively) for the determination of the cumulative distribution of $q=y / x$. The lower and upper limits on the generating mass distribution are given by $c$ and $d$, which implies that $c / d \leq q \leq 1$. The integration domain $S$ is given by: $c / q \leq x \leq d \wedge c \leq y \leq q x$. Middle and right: same, for the case that the primary mass $x$ is restricted by $x_{1}<x<x_{2}$. If $q \leq c / x_{1}$ the integration domain $S^{\prime}$ is given by $c / q \leq x \leq x_{2} \wedge c \leq y \leq q x$ (middle panel), while for $q>c / x_{1}$, the domain $S^{\prime}$ is defined by $x_{1} \leq x \leq x_{2} \wedge c \leq y \leq q x$ (right panel).
mass ratio distribution for a sample of binaries with a restricted primary mass range is given in Appendix A.1. The dependence of the mass ratio distribution on the primary mass range for RP is shown in Fig. 6 for the Kroupa (top) and Salpeter (bottom) generating mass distributions. Note that each mass ratio distribution is renormalised such that its maximum is unity. Note the differences between the overall mass ratio distributions $f_{q, \text { all }}(q)$ in the left-hand panels of Fig. 6. The Salpeter mass distribution results in a peak in the overall mass ratio distribution at $q \approx 1$. The Kroupa mass distribution, on the other hand, is on average much shallower, and therefore produces a peak at small $q$ (cf. the solid lines in the top panels of Fig. A.1).

In the special case where the mass distribution is of the form $f_{M}(M) \propto M^{-\alpha}$ with $\alpha \neq 1$, the overall mass ratio distribution is given by
$f_{q}(q)=\frac{\gamma}{\left(1-(c / d)^{\gamma}\right)^{2}}\left(q^{\alpha-2}-(c / d)^{2 \gamma} q^{-\alpha}\right)$,
where $\gamma=\alpha-1$. Realistic mass distributions cover a broad range of masses, i.e., $c \ll d$, for which the expression simplifies to:
$f_{q}(q) \approx\left\{\begin{array}{ll}(\alpha-1) q^{\alpha-2} & \alpha>1 \\ (1-\alpha) q^{-\alpha} & \alpha<1\end{array}\right.$.
(Piskunov \& Mal'Kov 1991). For distributions with $\alpha \approx 1$, the above expression is not a good approximation, and Eq. (12) should be used. In the case of a power-law mass ratio distribution with $c \ll d$, the overall mass ratio distribution resulting from random pairing can thus be described by a simple power-law. Note that Eq. (13) does not depend on $c$ or $d$; see also the bottom-left panel in Fig. 6. For a Salpeter mass distribution ( $\alpha=2.35$ ), for example, the resulting overall mass ratio distribution is $f_{q, \text { all }}(q) \propto q^{0.35}$. The general expression for $f_{q}$ when the primary mass range is restricted to the range $\left[x_{1}, x_{2}\right]$ is given in Appendix A.1. For a sample of binaries with a


Fig. 6. How does the specific mass ratio distribution $f_{q ; M_{1}}(q)$ resulting from random pairing (RP) depend on the mass distribution and the selected primary mass range? This figure shows the (renormalised) resulting mass ratio distributions for the Kroupa mass distribution (top panels; Monte Carlo simulations) and the Salpeter mass distribution (bottom panels; analytical calculations) in the mass range $c \leq M_{1} \leq 20 M_{\odot}$. The lower limit of $f_{M}(M)$ is set to $c=0.02 M_{\odot}$ for the solid curves, and $c=0.08 M_{\odot}$ for the dashed curves; the difference between these models is thus the presence of brown dwarfs. In all panels the models with brown dwarfs have on average a smaller mass ratio. The left-hand panel shows the overall mass ratio distribution (for all binaries in the population). The other panels show the specific mass ratio distribution for the binaries with a primary of the spectral type shown in the top-right corner of each panel ( $\mathrm{BD}=$ systems with brown dwarf primaries). All distributions are normalized so that their maximum is unity. Each model contains $5 \times 10^{5}$ binaries and has a binary fraction of $100 \%$.


Fig. 7. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ as a function of primary mass for pairing functions RP, PCP-II, SCP-I, SCP-II and SCP-III. Models with a generating binary fraction $\mathcal{B}$ of (from bottom to top) $5 \%, 50 \%$ and $90 \%$ are shown. The horizontal dotted lines in each panel represent the generating binary fraction. We adopt a Kroupa generating mass distribution with $c=0.02 M_{\odot}$ and $d=20 M_{\odot}$. For pairing functions PCP-II, SCP-I, SCP-II and SCP-III we show results for models with generating mass ratio distributions $f_{q}(q) \propto q^{\gamma}$, with $\gamma=0$ (solid curves), $\gamma=-0.3$ (dash-dotted curves), and $\gamma=-0.6$ (dashed curves).
narrow primary mass range ( $x_{1} \approx x_{2}$ ), $f_{q}(q)$ is proportional to $q^{-\alpha}$ for $c / x_{1} \leq q \leq 1$ and zero otherwise:
$f_{q ; M_{1}}(q)=\left\{\begin{array}{lr}0 & q<c / x_{1} \\ \frac{\gamma}{1-\left(c / x_{1}\right)^{2}} q^{-\alpha} & c / x_{1} \leq q \leq 1\end{array}\right.$.
Random pairing from a mass distribution that is approximately a single power-law (e.g. the Kroupa IMF) thus generally results in a three-segment mass ratio distribution, which is zero for $q \lesssim$ $c / x_{2}$ and exhibits a peak at $q \approx c / x_{1}$ (Eq. (A.15)). As already noted by Zinnecker (1984), a sample of binaries with high-mass primaries thus shows a peak in $f_{q}(q)$ at small $q$, while a sample with low-mass primaries peaks at large $q$.

For pairing function RP the overall binary fraction $\mathcal{B}_{\text {all }}$ is equal to the generating binary fraction $\mathcal{B}$. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$, however, depends on the surveyed primary mass range: the larger the primary mass, the higher the specific binary fraction (unless $\mathcal{B}=100 \%$ ). This can be understood as follows. For random pairing, the mass of the primary stars is drawn from the mass distribution. A fraction $\mathcal{B}$ of the primaries is assigned a companion, and primary and companion are swapped, if necessary, so that the primary is the most massive
star. This swapping leads to an increased number of binaries with a high-mass primary, and a decreased number of binaries with a low-mass primary, and hence a mass-dependent binary fraction. The relation between specific binary fraction and primary mass for random pairing is given by
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\left(\frac{\mathcal{B}^{-1}-1}{2 F_{M}\left(M_{1}\right)}+1\right)^{-1}$,
where $F_{M}\left(M_{1}\right)$ is the cumulative mass distribution evaluated at mass $M_{1}$ (see Appendix A.1.2). Figure 7 shows the binary fraction as a function of primary mass for several binary fractions, for the Salpeter mass distribution. Clearly, we have $\mathcal{B}_{M_{1}}\left(M_{1}\right) \equiv 1$ when $\mathcal{B}=100 \%, \mathcal{B}_{M_{1}}\left(M_{1}\right) \equiv 0$ when $\mathcal{B}=0 \%$, and $\mathcal{B}_{M_{1}}\left(M_{\text {min }}\right)=$ 0 . Systems with $\left.M_{1}\right\rangle\langle M\rangle$ have $\mathcal{B}_{M_{1}}\left(M_{1}\right) \geq \mathcal{B}$, while systems with $M_{1}<\langle M\rangle$ have $\mathcal{B}_{M_{1}}\left(M_{1}\right) \leq \mathcal{B}$ (see Appendix A.1.2 for a full derivation). The magnitude of this difference depends on the shape of the mass distribution and the generating binary fraction. Note that the variation of binary fraction with primary mass is purely a result of the choice of pairing function; no explicit variation of binary fraction with primary mass is included in the simulations.

### 6.1.1. Restricted random pairing (RRP)

Restricted random pairing (RRP) is very similar to random pairing (RP) as described in Sect. 6.1, with the difference that the binary components are now drawn from a limited mass range. All properties derived in Sect. 6.1 are thus applicable to the resulting binary sub-population resulting from RRP. However, the nature of RRP implies the presence of one or more other subpopulations that have a formed via another process. The other sub-populations could have alternative pairing function, such as RRP with different lower and upper mass limits, or a completely different pairing function.

Kroupa (1995a,c,b) finds that observations of binary systems are consistent with the population being born with pairing function RRP in the stellar mass range, prior to the effects of pre-main sequence eigenevolution. Further motivated by the difference between the observed mass ratio distribution and semimajor axis distribution of binary systems with a stellar primary and those with a brown dwarf binary (e.g., Bouy et al. 2003; Burgasser et al. 2003; Martín et al. 2003; Close et al. 2003), this implies that the brown dwarf population has formed with a different process (Kroupa et al. 2003; Thies \& Kroupa 2007, 2008). Their proposed model with stellar and substellar sub-populations is further supported by the existence of the brown dwarf desert among solar-type stars (see Sect. 9.1).

### 6.2. Pairing function PCRP (primary-constrained random pairing)

For primary-constrained random pairing (PCRP), each primary mass $M_{1}$ is drawn from the mass distribution $f_{M}(M)$ with limits $c \leq M \leq d$. The companion mass $M_{2}$ is also drawn from the same mass distribution, but with the additional constraint that $M_{2} \leq M_{1}$. The limits on the resulting mass ratio distribution are equivalent to those of random pairing $0 \leq c / d \leq q \leq 1$. Writing the primary mass distribution as $f(x)$ and the re-normalised companion mass distribution as $f^{\prime}(y)$, the expression for the joint probability distribution is $f_{\text {pcrp }}(x, y)=f_{x}(x) f^{\prime}{ }_{y}(y)$, which is normalised to unity due to the re-normalisation of $f_{y}(y)$. To derive the overall mass distribution $f_{q}$ one can proceed as for the RP case (see Sect. 6.1). The integration domain $S$ is again as shown in Fig. 5, and
$F(q)=P(y / x \leq q)=\int_{c / q}^{d} \int_{c}^{q x} f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x$.
Note that the normalisation constant for $f^{\prime}(y)$ depends on $x$. The expression for the sample with a restricted primary mass range [ $x_{1}, x_{2}$ ] now becomes:
$F_{q}(q) \propto\left\{\begin{array}{lr}0 & q<c / x_{2} \\ \int_{c / q}^{x_{2}} \int_{\mathrm{c}}^{q x} f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x & c / x_{2} \leq q<c / x_{1} \\ \int_{x_{1}}^{x_{2}} \int_{\mathrm{c}}^{q x} & f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x\end{array}, \quad c / x_{1} \leq q \leq 1, ~\right.$
where the integration domain $S^{\prime}$ is as in Fig. 5. The expression for the resulting $f_{q, \text { all }}(q)$ in the case of a single power-law generating mass distribution is given in Appendix A.2. For PCRP the distribution $f_{q, \text { all }}(q)$ contains a term that diverges for $q \uparrow 1$, which can be seen in Fig. 2 (see also Eq. (A.22) in Appendix A.2).

For a sample with primary masses restricted to the range [ $x_{1}, x_{2}$ ], Eq. (17) has to be worked out. We will not explicitly show the results here. Most importantly, $f_{q}(q)$ is zero for $q<c / x_{2}$, and exhibits a peak at $q \approx c / x_{1}$. The distribution for a sample of binaries with high-mass primaries thus peaks
at small $q$, and the distribution for low-mass binaries peaks at large $q$. If the primary mass range is narrow $\left(x_{1} \approx x_{2}\right), f_{q}(q)$ can be approximated with Eq. (14). For a sample of stars with a very narrow primary mass range, the mass ratio distributions resulting from PCRP and RP thus give the same results. Differences between the two pairing functions become larger for realistic primary mass ranges.

For pairing function PCRP, the companion mass distribution $f_{M_{2}}\left(M_{2}\right)$ for a set of primaries of identical mass is equal to the generating mass distribution $f_{M}(M)$ in the mass range $M_{\text {min }} \leq M_{2} \leq M_{1}$. The companion mass distribution can thus in principle be used to derive the properties of the generating mass distribution. For example, in a set of binaries with a primary mass of $1 M_{\odot}$, those with mass ratio $q<0.08$ have brown dwarf companions. If the observations are of good enough quality to study the mass ratio distribution below $q=0.08$, and it is known a priori that the pairing function is PCRP, the mass ratio distribution can be used to constrain the slope of the mass distribution in the brown dwarf regime.

The pairing algorithms RP and PCRP appear similar, but their difference is for example seen in the primary mass distribution. For RP there is a larger number of binary systems with high-mass primaries, which can be understood as follows. Suppose a primary mass $M_{1}$ of $5 M_{\odot}$ is drawn from the mass distribution. For PCRP, the companion mass $M_{2}$ is always smaller than the primary mass, while for RP the companion mass can take any value permitted by the mass distribution (i.e. also $M_{2}>M_{1}$, after which the components are switched). Another difference is that, unlike RP, for PCRP the binary fraction is independent of primary spectral type: $\mathcal{B}_{\text {all }}=\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B}$. For realistic mass distributions (e.g., Salpeter- or Kroupa-like), the overall mass ratio distribution $f_{q, \text { all }}(q)$ of PCRP is peaked towards large values of $q$, while that of RP is peaked towards small values of $q$. The pairing functions RP and PCRP can be excluded if more than a $1-2 \%$ of the intermediate-mass stars are "twins" ( $q \geq 0.8$ ).

### 6.3. Pairing function PCP (primary-constrained pairing)

In models with primary-constrained pairing (PCP), each binary system is generated by drawing a primary mass $M_{1}$ from $f_{M}(M)$ in the range $c \leq M \leq d$, and a mass ratio $q$ from the generating distribution $h_{q}(q)$. The companion mass is then calculated from $M_{2}=q M_{1}$. Due to the nature of this pairing mechanism it is possible that the resulting companion is of very low mass, for example a planetary mass if a very small mass ratio is drawn. Below we describe three variants of pairing function PCP, each of which handles very low mass companions in a different way: accepting all companions (PCP-I), rejecting the very low-mass companions (PCP-II), and redrawing the mass ratio if the companion mass is of very low mass (PCP-III).

### 6.3.1. Pairing function PCP-I

PCP-I is the simplest variant of PCP: the primary mass is drawn from $f_{M}(M)$ and the mass ratio from $h_{q}(q)$, and no further constraints are set. As a result, the specific mass ratio distribution $f_{q ; M_{1}}(q)$ and overall mass ratio distribution $f_{q, \text { all }}(q)$ are equal to the generating mass ratio distribution $h_{q}(q)$. Additionally, the specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ and overall binary fraction $\mathcal{B}_{\text {all }}$ are equal to the generating binary fraction $\mathcal{B}$.

The companion mass for PCP-I can be arbitrarily small: $M_{2, \min }=0$. Several companions may thus have masses
significantly lower than the deuterium burning limit $\left(\sim 0.02 M_{\odot}\right)$. Even planetary companions are considered as companion "stars" for the pairing function PCP-I. However, if we do include planets, we make the implicit assumption that the star formation process is scalable down to planetary masses. This assumption is in contradiction with the theories that suggest that stars and brown dwarfs form by fragmentation (Goodwin et al. 2007; Whitworth et al. 2007), while planets form by core-accretion (see, e.g., Pollack et al. 1996). This is an important point to keep in mind when using PCP-I, i.e., when adopting a mass ratio distribution that is fully independent of primary spectral type.

### 6.3.2. Pairing function PCP-II

For PCP-II, companions with $M_{2}<c$, are rejected, and the corresponding primary stars are classified as single stars. There are two reasons why one may want to consider using PCP-II. First, one may wish to use this prescription if a minimum companion mass is expected from theory, for example the Jeans mass or the opacity limit for fragmentation (e.g., Hoyle 1953; Rees 1976; Low \& Lynden-Bell 1976; Silk 1977a,b, 1995; Tohline 1982; Larson 1969, 1992, 2005; Masunaga \& Inutsuka 2000). The second, more observational approach may be to "ignore" the low-mass companions. Although this seems somewhat artificial, this method is often used in practice. Planets are usually not considered as companions (the Sun is a "single star"), which implies a limit $c=0.01-0.02 M_{\odot}$.

Due to the rejection of low-mass companions the overall mass ratio distribution is zero for $0<q<c / d$. In the range $c / d \leq q \leq 1$, the expression for the overall mass ratio distribution is given by:
$f_{q, \mathrm{all}}(q)=k \int_{c / q}^{d} f_{\mathrm{pcp}}\left(M_{1}, q\right) \mathrm{d} M_{1}=k \int_{c / q}^{d} h_{q}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1},(18)$ where $h_{q}(q)$ is the generating mass ratio distribution and $k$ is a normalisation constant which ensures that $\int_{0}^{1} f_{q, \text { all }}(q) \mathrm{d} q=1$. The distribution $f_{q, \text { all }}(q)$ has a larger average mass ratio than the generating mass ratio distribution $h_{q}(q)$ as a result of rejecting the low-mass companions. For a a sample with a restricted primary mass range, $c \leq x_{1} \leq M_{1} \leq x_{2} \leq d$, the expression for the specific mass ratio distribution is given in Appendix A.3, by Eq. (A.29). The specific mass ratio distribution is given by a three-segment powerlaw, with slope changes at $q=c / x_{2}$ and $q=c / x_{1}$. Note that for $q>c / x_{1}$, the distribution $f_{q ; M_{1}}(q)$ is equal to the generating mass ratio distribution. For a sample of highmass primaries, where $c / x_{1} \leq 1$, we thus have $f_{q ; M_{1}}(q) \approx h_{q}(q)$. For a sample of low-mass primaries, $c / x_{2} \leq q<c / x_{1}$ for most of the mass ratio range, and the corresponding term dominates. For these binaries, $f_{q ; M_{1}}(q)$ differs significantly from $h_{q}(q)$.

The specific binary fraction for a sample of systems with primary mass $M_{1}$ is given by
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} \int_{q_{\min ( }\left(M_{1}\right)}^{1} h_{q}(q) \mathrm{d} q<\mathcal{B}$.
Note that $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is independent of $f_{M}(M)$ for PCP-II. Figure 7 shows the specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ as a function of $\mathcal{B}$ and $f_{q}(q)$. The specific binary fraction depends on the shape of the mass ratio distribution, and is independent of the mass distribution. The overall binary fraction $\mathcal{B}_{\text {all }}$ after rejection of the low-mass companion is lower than the generating binary fraction $\mathcal{B}$ (see Eq. (A.32) in Appendix A.3.2). The binaries with high-mass primaries are hardly affected by the rejection algorithm. For these binaries the specific mass ratio distribution and
specific binary fraction are practically equal to those for PCP-I: the specific binary fraction is equal to $\mathcal{B}$, and the specific mass ratio distribution is equal to $h_{q}(q)$. For the very low-mass primaries, however, a high fraction of the companions is rejected, and therefore the specific binary fraction is low. The remaining companions of these stars have a mass comparable to that of their primary, and the resulting mass ratio distribution for the lowest-mass binaries is peaked to unity (see Fig. 2).

As an example, consider a stellar population with $\mathcal{B}=$ $100 \%$ and a generating mass ratio distribution $f_{q}(q)=1$ and $M_{2, \min }=0.02 M_{\odot}$, systems with B-type primaries and systems with M-type primaries have a resulting specific binary fraction of $99 \%$ and $87 \%$, respectively. If we also do not consider brown dwarfs as companions (so $M_{2, \min }=0.08 M_{\odot}$ ), then the specific binary fractions are $96 \%$ and $51 \%$, respectively.

### 6.3.3. Pairing function PCP-III

For PCP-III the primary mass is drawn from $f_{M}(M)$ in the range $c \leq M_{1} \leq d$, and the mass ratio is drawn from $h_{q}(q)$. If the resulting companion star mass is smaller than $c$, the mass ratio is redrawn from $h_{q}(q)$ until a companion with mass $M_{2} \geq c$ is obtained. This is equivalent to renormalising $h_{q}(q)$ in the range $q_{\min }\left(M_{1}\right) \leq q \leq 1$, where $q_{\min }\left(M_{1}\right)=c / M_{1}$. This effectively results in a mass-dependent generating mass ratio distribution $h_{q}^{\prime}(q)$. The expression for the resulting overall mass ratio distribution $f_{q, \text { all }}(q)$ are then:
$f_{q, \mathrm{all}}(q)=\int_{c / q}^{d} f_{\mathrm{pcp}}\left(M_{1}, q\right) \mathrm{d} M_{1}=\int_{c / q}^{d} h_{q}^{\prime}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$.
The specific mass ratio distribution for the restricted primary mass range $\left[x_{1}, x_{2}\right]$ is given in Appendix A.3. A sample of binaries with high-mass primaries has $c / x_{1} \ll 1$. For high-mass binaries $h_{q}^{\prime}(q) \approx h_{q}(q)$, and therefore $f_{q ; M_{1}}(q) \approx h_{q}(q)$. For the very low-mass primaries, however, a high fraction of the companions is redrawn. Consequently, all binaries with a low-mass primary have a mass ratio close to unity. The resulting mass ratio distribution for the lowest-mass binaries is thus peaked to unity (see Fig. 2).

As a result of the redrawing of the companions for pairing function PCP-III, the resulting overall binary fraction $\mathcal{B}_{\text {all }}$ equals the generating binary fraction $\mathcal{B}$, and the specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ equals $\mathcal{B}$ for any primary mass range.

### 6.4. Pairing function SCP (split-core pairing)

For split-core pairing (SCP) one assumes that the system "core" mass $M_{\mathrm{C}}$ is drawn from a core mass distribution $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right)$ with $2 \epsilon c \leq M_{\mathrm{C}} \leq 2 \epsilon d$, where $\epsilon$ is the star forming efficiency. Splitcore pairing is frequently inferred from observations of dense cores in star forming regions, assuming that a fraction of the cores fragment into binaries (see, e.g., Goodwin et al. 2008; Swift \& Williams 2008). As a core collapses, it forms one or two stars with a total mass $M_{\mathrm{T}}=\epsilon M_{\mathrm{C}}$. The resulting minimum and maximum primary masses are thus $c$ and $2 d$, respectively. The star forming efficiency may be a function of various parameters, for example the mass of the core. For simplicity in our analysis, however, we keep the star forming efficiency fixed to $\epsilon=1$ for all values of $M_{\mathrm{C}}$. The total mass of each binary is thus $M_{\mathrm{T}}=M_{\mathrm{C}}$. Note that "random fragmentation" (random splitup of a clump into two stellar components) is very different from random pairing of two components from the IMF (see, e.g., Figs. 2 and 3).

For pairing function SCP, the binary total mass is thus drawn from a distribution $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right)$. Note that, although we adopt $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right)=f_{\mathrm{Kroupa}}(M)$ in this paper, there is no obvious prerequisite that $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right)$ should be a standard IMF. The binary is split up according to a mass ratio that is drawn from a generating mass ratio distribution $h_{q}(q)$. Given the core mass $M_{\mathrm{C}}=M_{1}+M_{2}$ and the mass ratio $q$, the primary and companion mass are given by
$M_{1}=\frac{M_{\mathrm{C}}}{q+1} \quad$ and $\quad M_{2}=\frac{M_{\mathrm{C}}}{q^{-1}+1}$,
respectively. As a result of this procedure, it may happen that a companion mass smaller than the minimum mass $c$ is drawn. Similar to pairing function PCP (Sect. 6.3), there are three ways to address this issue: accepting the low-mass companions (SCP-I), rejecting the low-mass companions (SCP-II), and redrawing the mass ratio if the companion mass is too low (SCPIII). We discuss these three variants of SCP in the sections below.

### 6.4.1. Pairing function SCP-I

For pairing function SCP-I, all binary components resulting from the split-up mechanism are accepted, irrespective of their mass. Stars with substellar and planetary companions are thus also considered as "binary stars". Due to the nature of this pairing process, the overall mass ratio distribution is equal to the generating mass ratio distribution. The specific mass ratio distribution, however, is a function of spectral type. A full derivation of the specific mass ratio distribution is given in Appendix A.4; see also Clarke (1996a). In most cases the primary mass range is contained within the range [ $2 c, d$ ], which corresponds to "case 7 " in Appendix A.4. Consider the special case of a single powerlaw mass distribution $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right) \propto M_{\mathrm{C}}^{-\alpha}(\alpha \neq 1)$ and a uniform mass ratio distribution $h_{q}(q)=1$. Under these assumptions, the expression for the specific mass ratio distribution is $f_{q ; M_{1}}(q) \propto(1+q)^{1-\alpha}$ for $0<q \leq 1$, if either the primary mass range is contained within the range $[2 c, d]$ (case 7), or if $x_{1}=x_{2}$. Note that this expression is identical to that in Eq. (5). The highest mass binaries thus have on average a small mass ratio, and the lowest mass binaries have on average a large mass ratio. Note that these trends are present, even though the generating mass ratio distribution produces mass ratios in the range $0<q \leq 1$, irrespective of the core mass $M_{\mathrm{T}}$.

Pairing function SCP-I naturally results in a mass-dependent binary fraction:
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\frac{\mathcal{B} f_{M_{1}}\left(M_{1}\right)}{\mathcal{B} f_{M_{1}}\left(M_{1}\right)+(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right)}$,
where $f_{M_{1}}\left(M_{1}\right)$ is the primary mass distribution, $f_{M_{\mathrm{C}}}\left(M_{1}\right)$ the generating (core) mass distribution evaluated at mass $M_{1}$, and $\mathcal{B}$ the generating binary fraction; see Appendix A.4.2 for the derivation. The mass-dependence of the binary fraction for SCP-I is illustrated in the middle panel of Fig. 7. In general, $\mathcal{B}=100 \%$ for $c \leq M_{1}<2 c$, as $f_{M_{\mathrm{C}}}=0$ in this mass range. For $2 c<M_{1}<d$ the specific binary fraction is more or less independent of $M_{1}$, while beyond $M_{1}=d$, it decreases down to zero at $M_{1}=2 d$. The latter dependence is due to the fact that the cores are more massive than the primary stars they potentially form. As a result, the high-mass targets are dominated by cores that have not split up. The overall binary fraction $\mathcal{B}_{\text {all }}$ is always equal to $\mathcal{B}$.

### 6.4.2. Pairing function SCP-II

For SCP-II the companion is rejected if $M_{2}<c$, and the primary star becomes single. The resulting primary mass range is then $c \leq M_{1} \leq 2 d-c$, and the companion mass is in the range $c \leq$ $M_{2} \leq d$. The full derivation for the overall and specific mass ratio distributions is given in Appendix A.4. In most realistic cases the primary mass range $\left[x_{1}, x_{2}\right]$ is fully enclosed in the mass range $2 c<M_{1}<d$ ("case 1"). Unlike SCP-I, the resulting overall mass ratio distribution is unequal to the generating mass ratio distribution, but contains more large- $q$ binaries instead.

The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ resulting from SCP-II varies with primary mass $M_{1}$ :
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\frac{\mathcal{B}_{100}\left(M_{1}\right) \mathcal{B} f_{M_{1}}\left(M_{1}\right)}{\mathcal{B} f_{M_{1}}\left(M_{1}\right)+(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right)}$,
where $f_{M_{1}}$ is the primary mass distribution, $f_{M_{\mathrm{C}}}$ is the generating (core) mass distribution, and $\mathcal{B}$ is the generating binary fraction. The quantity $\mathcal{B}_{100}\left(M_{1}\right)$ represents the value of $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ for a population with $\mathcal{B}=100 \%$. The full derivation of Eq. (23) is given in Appendix A.4.2.

Figure 7 shows the specific binary fraction $\mathcal{B}\left(M_{1}\right)$ as a function of $\mathcal{B}$ and $h_{q}(q)$. The specific binary fraction decreases with decreasing primary mass, as, on average, more low-mass companions are rejected among lower-mass primaries. The majority of the newly formed single stars (due to rejection of low-mass companions) is thus of most low mass. For the very lowestmass stars, however, the binary fraction increases to unity, as $\mathcal{B}_{100}\left(M_{1}\right) \approx 1$ and $f_{M_{\mathrm{C}}}=0$ for $M_{1} \approx c$. The specific binary fraction then rises again to a maximum around $M_{1}=d$, and then rapidly drops to zero at $M_{1}=2 d-c$. As a result of the rejection of low-mass companions, the overall binary fraction is always lower than $\mathcal{B}$.

### 6.4.3. Pairing function SCP-III

For SCP-III the mass ratio is redrawn when a companion with mass $M_{2}<c$ is produced by the splitting algorithm (similar to PCP-III). This effectively corresponds to a (mass-specific) renormalised mass ratio distribution $h_{q}^{\prime}(q)$ in the range $c /\left(M_{\mathrm{C}}-\right.$ $c) \leq q \leq 1$. The resulting overall mass ratio distribution for SCP-III is given by:
$f_{q, \text { all }}(q)=k \int_{c(1+1 / q)}^{2 d} h_{q}^{\prime}(q) f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right) \mathrm{d} M_{\mathrm{C}}$,
where the lower integration limit is set by the condition $M_{2} \geq c$ and $k$ is again a normalisation constant. The expressions for the specific distribution $f_{q ; M_{1}}(q)$ for a restricted primary mass range are identical to those for the SCP-II case (see Appendix A.4), except that $h^{\prime}{ }_{q}(q)$ replaces $h_{q}(q)$ everywhere. As a result of the renormalisation of the mass ratio distribution, the mass ratio distribution is a function of spectral type. The specific mass ratio distribution for the lowest-mass cores is strongly peaked to $q=1$.

The overall binary fraction $\mathcal{B}_{\text {all }}$ equals $\mathcal{B}$ for SCP-III. The specific binary fraction is given by Eq. (22). Note however, that the primary mass distribution $f_{M_{1}}\left(M_{1}\right)$ resulting from SCP-III is different from that of SCP-III; see Appendix A.4.2 for details.

## 7. Differences between pairing functions

As discussed in the previous sections, each pairing function results in different properties of the binary population. In this


Fig. 8. The mass distributions resulting from the different pairing functions, for the models shown in Fig. 2. From top to bottom, the panels show the primary mass distribution $f_{M_{1}}$, companion mass distribution $f_{M_{2}}$, the system mass distribution $f_{M_{\mathrm{T}}}$, the mass distribution of the primaries and single stars combined $f_{M_{1, S}}$ (i.e., the targets in an observational binarity study), the single star mass distribution $f_{M_{\mathrm{S}}}$, and the distribution of all stellar masses $f_{M_{\text {all }}}$. Each distribution is normalised so that its maximum is unity. The vertical axis has a linear scale (rather than a logarithmic scale; cf. Fig. 1) so that the differences are clearly shown. Each column corresponds to a different pairing function, which is indicated at the top.
section we provide an overview of the major differences and similarities. In general, each of the properties described below depends on the choice of the pairing function, a generating mass distribution $f_{M}(M)$, a generating mass ratio distribution $f_{q}(q)$, and a generating binary fraction $\mathcal{B}$. Our example models have a Kroupa generating mass distribution with $0.02 \leq M \leq 20 M_{\odot}$, a flat generating mass ratio distribution (if applicable), and a generating binary fraction of $100 \%$. The resulting mass distributions and mass ratio distributions for the different pairing functions are shown in Figs. 2 and 8, respectively.

### 7.1. The mass distributions

For pairing functions PCP-I, PCP-III and PCRP the primary mass distribution is identical to $f_{M}(M)$. For pairing function PCP-II, $f_{M_{1}}(M)$ is very similar to $f_{M}(M)$, but it contains less primaries of low mass due to rejection of very low-mass companions (which mostly occurs among low-mass primaries). The latter "primaries" are considered as single stars after removal of their companions. For pairing function RP the primary mass distribution is more massive than $f_{M}(M)$ due to swapping of
the components that are drawn from $f_{M}(M)$. Pairing functions SCP-I, SCP-II and SCP-III result in smaller average primary masses than $f_{M}(M)$ due to the core splitting.

For all pairing functions the companion mass distribution $f_{M_{2}}(M)$ is shifted to lower masses with respect to the generating mass distribution $f_{M}(M)$. Pairing functions PCP-III, PCRP and SCP-III result in a large number of low mass companions; the companion mass distribution is sharply peaked at $M_{2} \approx c$; see Figs. 3 and 8. For a Kroupa generating mass distribution, the other pairing functions show a peak in the companion mass distribution around $M_{2} \approx 0.05 M_{\odot}$. Pairing functions PCP-I and SCP-I may result in arbitrarily small companion masses, but all other pairing functions have $M_{2, \min }=c$. All pairing functions have $M_{2, \max }=d$. For each pairing function described in this paper, the companion mass distribution depends strongly on the selected primary mass range (see Fig. 3), but is independent of $\mathcal{B}$.

We define the binary system mass distribution $f_{M_{\mathrm{T}}}(M)$ as the distribution of masses $M_{\mathrm{T}}=M_{1}+M_{2}$ of all binary systems in a population. For pairing function SCP-I the system mass distribution equals the generating mass distribution: $f_{M_{\mathrm{T}}}(M)=f_{M}(M)$. For all other pairing functions the expression for the system mass
distribution is different from $f_{M}(M), f_{M_{1}}(M)$ and $f_{M_{2}}(M)$. For random pairing, for example, the system mass distribution can be described as a convolution (Eq. (9)).

Pairing functions with $\mathcal{B}<100 \%$ result in a population of single stars. For PCP-II and SCP-II, even those with $\mathcal{B}=100 \%$ result in single stars. In our analysis we distinguish between the single mass distribution $f_{M_{S}}(M)$ and the primary/single mass distribution $f_{M_{1, S}}(M)$. The latter distribution includes both single stars and primaries and is important for observers. Target lists for binarity surveys are often defined by $f_{M_{1, S}}(M)$ for a certain mass range, as in practice it is often unknown whether a surveyed star has a companion star. For most pairing functions the single mass distribution $f_{M_{\mathrm{S}}}(M)$ is equal to the generating mass distribution $f_{M}(M)$, and the fraction of single stars among the systems is given by $1-\mathcal{B}$, where $\mathcal{B}$ is the generating binary fraction. For PCP-II and SCP-II additional single stars are created due to the rejection of very low-mass companions. The generated single stars are mostly of low mass. Only for pairing functions PCP-I, PCP-II, PCP-III and PCRP, the primary/single star mass distribution $f_{M_{1, S}}(M)$ is equal to the generating mass distribution $f_{M}(M)$. The primary/single star mass distribution is biased to higher masses for RP, and to lower masses for SCP-I, SCP-II and SCP-III.

The bottom row of Fig. 8 shows the mass distribution $f_{\text {all }}(M)$ of all stars, including primaries, companions, and single stars. This mass distribution, which includes all stars in the population, is referred to as the mass function of a stellar population. For a zero-age population, $f_{\text {all }}(M)$ is an initial mass function (IMF), e.g., the Kroupa IMF (Eq. (2)). Only for pairing function RP is the individual star mass distribution $f_{\text {all }}(M)$ equal to the generating mass distribution $f_{M}(M)$; the other pairing functions result in a distribution $f_{\text {all }}(M)$ which is biased to lower values with respect to $f_{M}(M)$.

### 7.2. The mass ratio distributions

The resulting overall mass ratio distribution $f_{q, \text { all }}(q)$ depends on the pairing function. The top panels in Fig. 2 show the significant difference between the overall mass ratio distribution for the different pairing functions. The overall mass ratio distribution $f_{q \text {,all }}(q)$ is equal to the generating mass ratio distribution $f_{q}(q)$ for pairing functions PCP-I and SCP-I. For pairing functions PCP-II, PCP-III, SCP-II and SCP-III the overall mass ratio distribution is biased to larger values of $q$ with respect to $f_{q}(q)$. For these pairing functions binaries with very low companion masses, and thus often very low mass ratios, are either rejected or redrawn, resulting in systematically larger values of $q$. The overall mass ratio distribution of pairing functions RP and PCRP are purely a result of the mass distribution $f_{M}(M)$ and depend strongly on its lower and upper limits.

Only for pairing function PCP-I is the mass ratio distribution independent of spectral type, and thus equal to the overall and generating mass ratio distributions. For all other pairing functions the specific mass ratio distribution is a function of the primary mass. The middle and bottom panels in Fig. 2 show for each pairing function the specific mass ratio distribution for target samples of different spectral types. For each of these pairing functions, high-mass binaries have on average a lower mass ratio than low-mass binaries. The lowest-mass binaries have a mass ratio distribution peaked to $q=1$.

Figures 2 and 6 illustrate the strong dependence of the specific mass ratio distribution on the targeted sample in the survey. Figure 4 shows a generalised version of these figures. Each panel shows the two-dimensional distribution $f\left(q, M_{1}\right)$ for the different

Table 2. The mass ratio distribution resulting from different pairing functions, for different samples, as compared to the generating mass ratio distribution $h_{q}(q)$.

| Pairing | $f_{q, \text { all }}(q)$ | $f_{q}(q)_{M_{1} \approx M_{\text {max }}}$ | $f_{q}(q)_{M_{1} \approx M_{\text {min }}}$ |
| :--- | :---: | :---: | :---: |
| RP | - | $\downarrow$ | $\uparrow$ |
| PCRP | - | $\downarrow$ | $\uparrow$ |
| PCP-I | $=$ | $=$ | $=$ |
| PCP-II | $\downarrow$ | $\approx$ | $\uparrow$ |
| PCP-III | $\uparrow$ | $\approx$ | $\uparrow$ |
| SCP-I | $=$ | $\downarrow$ | $\uparrow$ |
| SCP-II | $\uparrow$ | $\downarrow$ | $\uparrow$ |
| SCP-III | $\uparrow$ | $\downarrow$ | $\uparrow$ |

pairing functions. Care should thus be taken when extrapolating the results to the population as a whole. The interpretation of the observations is further complicated by the instrument bias and observational errors, an effect we will discuss in Sect. 9.

An overview of the mass ratio distribution changes is presented in Table 2. The columns show the pairing function, the overall mass ratio distribution, the specific mass ratio distribution for high-mass stars, the specific mass ratio distribution for low-mass stars. The symbols the table indicate whether the mass ratio distribution is equal to $(=)$, almost equal to $(\approx)$, or biased to low mass ratios $(\downarrow)$ or large mass ratios $(\uparrow)$, with respect to $h_{q}(q)$. As the distribution $h_{q}(q)$ is undefined for RP and PCRP, the properties of $f_{q ; M_{1}}(q)$ with respect to $f_{q, \text { all }}(q)$ are indicated.

### 7.3. The binary fractions

The overall binary fraction $\mathcal{B}_{\text {all }}$ is equal to the generating binary fraction $\mathcal{B}$ for most pairing functions. Only for PCP-II and SCP-II the overall binary fraction is lower than $\mathcal{B}$ because of the rejection of low-mass companions.

Table 3 provides an overview of the changes in the binary fraction as a function of spectral type and primary mass range. The columns show the pairing function, the overall binary fraction $\mathcal{B}_{\text {all }}$, the specific binary fraction for high-mass stars $\mathcal{B}_{M_{1} \approx M_{\max }}$ and for low-mass stars $\mathcal{B}_{M_{1} \approx M_{\text {min }}}$. The last column shows whether the resulting binary fraction is equal to $\mathcal{B}$ for any primary mass range.

Most pairing functions result in a mass-dependent binary fraction. If the generating binary fraction is lower than $100 \%$, all pairing functions except PCP-I, PCP-III and PCRP result in a specific binary fraction that depends on primary mass (see, e.g., Fig. 7). On the other hand, if $\mathcal{B}=100 \%$, only pairing functions PCP-II and SCP-II result in a varying $\mathcal{B}_{M_{1}}\left(M_{1}\right)$, due to the rejection of low-mass companions.

For a sample of binaries with high-mass primaries, $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is approximately equal to $\mathcal{B}$ for pairing functions PCP-I, PCPII, PCP-III and PCRP. For RP, the specific binary fraction for high-mass binaries is higher than $\mathcal{B}$ (unless $\mathcal{B}=100 \%$ ), while for SCP-I, SCP-II and SCP-III the binary fraction for high-mass stars is lower than $\mathcal{B}$ (unless $\mathcal{B}=100$ ). For low-mass binaries, $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is equal to $\mathcal{B}$ for pairing functions PCP-I, PCPIII, and PCRP. For RP, the specific binary fraction for low-mass stars is only equal to the generating binary fraction if the latter is $\mathcal{B}=0 \%$ or $\mathcal{B}=100 \%$, and lower in the other cases. For PCP-II the specific binary fraction for low-mass stars is lower than $\mathcal{B}$. For SCP-I and SCP-III the binary fraction for low-mass stars is higher than $\mathcal{B}$. For SCP-II the specific binary fraction for

Table 3. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ as compared to the generating binary fraction $\mathcal{B}$, for different the pairing functions.

| Pairing | $\mathcal{B}_{\text {all }}$ | $\mathcal{B}_{M_{1} \approx M_{\max }}$ | $\mathcal{B}_{M_{1} \approx M_{\min }}$ | $\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} ?$ |
| :--- | :--- | :--- | :--- | :--- |
| RP | $=$ | $>$ | $<$ | no, unless $\mathcal{B}=100 \%$ |
| PCRP | $=$ | $=$ | $=$ | yes |
| PCP-I | $=$ | $=$ | $=$ | yes |
| PCP-II | $<$ | $\approx$ | $<$ | no |
| PCP-III $=$ | $=$ | $=$ | yes |  |
| SCP-I | $=$ | $<, \approx,>$ | $<, \approx,>$ | no, unless $\mathcal{B}=100 \%$ |
| SCP-II | $<$ | $<, \approx,>$ | $<, \approx,>$ | no |
| SCP-III $=$ | $<, \approx,>$ | $<, \approx,>$ | no, unless $\mathcal{B}=100 \%$ |  |

low-mass stars may be higher or lower than $\mathcal{B}$, depending on the properties of $f_{M}(M)$ and $f_{q}(q)$ and the value of $\mathcal{B}$.

## 8. Dependence on generating properties

In this section we discuss how the properties of a binary population depend on the attributes for the pairing functions: the generating mass distribution $f_{M}(M)$, the generating mass ratio distribution $f_{q}(q)$, and the generating binary fraction $\mathcal{B}$. For the most important properties of a population (with respect to binarity), we list in Table 4 whether or not they depend on $f_{M}(M)$, $f_{q}(q)$, or $\mathcal{B}$, for each of the eight pairing functions described in this paper. Note that the system mass distribution $f_{M_{\mathrm{T}}}$ includes both single stars and binary systems. The specific mass ratio distribution $f_{q ; M_{1}}(q)$ and specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ in Table 4 only refer to samples where all binaries have identical primary mass $M_{1}$. For a sample with a finite primary mass, range, the results for the specific mass ratio distribution and specific binary fraction are mostly identical to those of $f_{q \text {,all }}(q)$ and $\mathcal{B}_{\text {all }}$, respectively. Table 4 also illustrates which properties of the population can be used to recover $f_{M}(M), f_{q}(q)$, and $\mathcal{B}$. For example, for RP, the specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ provides information on the generating mass distribution (e.g., the IMF).

The generating mass distribution. For obvious reasons, all mass distributions listed in Table 4 depend on $f_{M}(M)$. The mass ratio distributions for RP and PCRP are defined by, and depend strongly on the properties of $f_{M}(M)$ and $\mathcal{B}$. For PCP-I/II/III, $f_{q ; M_{1}}(q)$ does not depend on $f_{M}(M)$, as both $M_{1}$ and $q$ are drawn independently from their generating distributions. Obtaining the overall mass ratio distribution $f_{q \text {,all }}(q)$ requires integration over the primary mass distribution; as for PCP-II and PCP-III $f_{q ; M_{1}}(q)$ varies with $M_{1}$, so does $f_{q, \text { all }}(q)$. The overall binary fraction $\mathcal{B}_{\text {all }}$ is independent of $f_{M}(M)$, except for PCP-II and SCP-II, for which low-mass companions are rejected. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$, does not depend on $f_{M}(M)$ for PCRP and PCP-I/II/III due to the independent drawing of $M_{1}$ and $M_{2}$ ( or $q$ ), while it does vary with $f_{M}(M)$ for SCP-I/II/III as $M_{1}$ and $q$ are not drawn independently, and for RP as $M_{1}$ and $M_{2}$ are not drawn independently (due to the swapping of the components; see Sect. 6.1).

The generating mass ratio distribution. The generating mass ratio distribution $f_{q}(q)$ is undefined for RP and PCRP. For the other pairing functions, the dependence of $f_{q ; M_{1}}(q)$ and $f_{q, \text { all }}(q)$ on $f_{q}(q)$ is obvious. As companion masses are derived from $q$, the distributions $f_{M_{2}}$ and $f_{M_{\text {all }}}$ depend on the choice of $f_{q}(q)$. Note that that for SCP-II, all parameters vary with $f_{q}(q)$, and for PCP-II most parameters (except the primary/single mass distribution) vary with $f_{q}(q)$. The properties of the single stars do not depend on $f_{q}(q)$, except for PCP-II and SCP-II, where
additional single stars are created due to the rejection of lowmass companions.

The generating binary fraction. The dependence of $\mathcal{B}_{\text {all }}$ and $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ on the generating binary fraction $\mathcal{B}$ is trivial. The distributions that do not involve single stars, such as $f_{M_{1}}, f_{M_{2}}$, $f_{q, \text { all }}(q)$ and $f_{q ; M_{1}}(q)$, by definition never depend on the choice of $\mathcal{B}$. The mass distribution of all stars $f_{M_{\text {all }}}$ (the "IMF") depends on the choice of $\mathcal{B}$ for all pairing functions, except for RP.

## 9. Interpretation of observations

Binarity and multiplicity provide important information about the outcome of the star forming process in different environments (Blaauw 1961, 1991). In this paper we explore this issue by making the assumption that binary stars are formed through a simple "pairing function". In reality the distribution of stars in $f_{M_{1}, M_{2}}\left(M_{1}, M_{2}\right)$ is the result of complex physics involving the collapse of a molecular cloud into stars and stellar systems with discs (which can themselves fragment), followed by dynamical evolution of the protocluster (see Sect. 5). The resulting "pairing function" may thus not be describable in terms of the simple probability distributions given in this paper (Sect. 6). However, even if we proceed from our assumption that the pairing of binary stars involves the random selection of a primary mass followed by the secondary (RP, PCRP, PCP-I/II/III) or the random splitting of cores (SCP-I/II/III), the interpretation of the observations is not trivial:

- there is a large space of possible models for the formation and evolution of a binary population. In this paper we describe eight pairing mechanisms. For each pairing function there are a large number of possibilities for the generating mass distributions $f_{M}(M)$, mass ratio distributions $f_{q}(q)$, and binary fractions $\mathcal{B}$. In addition, we have no a priori restriction on the plausible formation mechanisms (see, however, Sect. 5);
- the observations of a binary population are generally limited. The surveys are incomplete, and are affected by selection effects and observational biases, and often only a limited set of binary population parameters is measured.

When the inverse problem of obtaining the binary formation mechanism from the data is so poorly constrained, it is not possible to find the "best model" for binary formation from a fit to the data. The only thing one can realistically do is to exclude models that are not capable of reproducing the data and accept that all other models offer plausible binary formation prescriptions.

### 9.1. Constraints from observations

Below we list several properties that have been identified for various binary populations over the last decades. These provide important information on the primordial pairing function, and the formation and evolution of binary populations.

The observed mass-dependent binary fraction. The observed binary fraction is known to increase with increasing primary spectral type; see, e.g., Sterzik \& Durisen (2004); Köhler et al. (2006); Lada (2006); Bouy et al. (2006) for an overview. For early-type ( $\mathrm{O} / \mathrm{B} / \mathrm{A}$ ) stars the binary fraction approaches $100 \%$ (e.g., Abt et al. 1990; Mason et al. 1998; Shatsky \& Tokovinin 2002; Kobulnicky \& Fryer 2007; Kouwenhoven et al. 2007b). The binary fraction decreases to $50-60 \%$ for F/G-type stars (Abt \& Levy 1976; Duquennoy \& Mayor 1991). For M-type stars the binary fraction is 30-40\% (Fischer \& Marcy 1992;

Table 4．Do the properties of a population resulting from a certain pairing function depend on its generating properties？The three generating properties $f_{M}(M), f_{q}(q)$ and $\mathcal{B}$ are listed in the top row．For each of the ten quantities listed in the left－most column，we list whether or not they depend on the choice of $f_{M}(M), f_{q}(q)$ and $\mathcal{B}$ ．We indicate dependence and independence with the symbols $\sqrt{ }$ and $\times$ ，respectively．The results for $f_{q ; M_{1}}(q)$ and $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ are valid only for a population of binaries with identical primary mass $M_{1}$ ．The distribution over total system mass $f_{M_{\mathrm{T}}}$ includes both binary systems and single stars．

| Generating prop． | $f_{M}(M)$ |  |  |  |  |  |  |  | $f_{q}(q)$ |  |  |  |  |  |  |  | $\mathcal{B}$ |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pairing function | \％ | $\stackrel{\widetilde{\partial}}{\square}$ | $\stackrel{\rightharpoonup}{0}$ | $\stackrel{0}{0}$ | $\begin{aligned} & \stackrel{\rightharpoonup}{7} \\ & \text { 自 } \end{aligned}$ | $\begin{aligned} & \tilde{1} \\ & \underset{\sim}{2} \end{aligned}$ | $\begin{aligned} & \tilde{\sim} \\ & \underset{\sim}{9} \end{aligned}$ | $\begin{aligned} & \tilde{n}^{2} \\ & \text { 号 } \end{aligned}$ | て | $\stackrel{\widetilde{Z}}{\square}$ | $\stackrel{\rightharpoonup}{7}$ | $\begin{aligned} & \stackrel{\rightharpoonup}{7} \\ & \stackrel{1}{4} \end{aligned}$ | $\begin{aligned} & \text { چٌ } \\ & \text { 自 } \end{aligned}$ | $\begin{aligned} & \text { n } \\ & \square \end{aligned}$ | $\begin{aligned} & \text { حٌ } \\ & \underset{\sim}{4} \end{aligned}$ | 解 | T | $\stackrel{\rightharpoonup}{0}$ | $\stackrel{7}{7}$ | $\stackrel{\rightharpoonup}{7}$ | $\begin{aligned} & \stackrel{\rightharpoonup}{3} \\ & \text { 首 } \end{aligned}$ | $\begin{aligned} & \sqrt{n} \\ & \end{aligned}$ | $\begin{aligned} & \text { n } \\ & \end{aligned}$ | \％ |
| $f_{M_{1}}$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | － | － | $\times$ | $\checkmark$ | $\times$ | $\checkmark$ | ， | $\checkmark$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ |
| $f_{M_{2}}$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | － | － | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ |
| $f_{M_{\mathrm{T}}}$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | － | － | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\sqrt{ }$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\times$ | $\checkmark$ | $\times$ |
| $f_{M_{1, S}}$ |  | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | － | － | $\times$ | $\times$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ |
| $f_{M_{S}}$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | － | － | $\times$ | $\sqrt{ }$ | $\times$ | $\times$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\checkmark$ | $\times$ | $\times$ | $\sqrt{ }$ | $\times$ |
| $f_{M_{\text {all }}}$（＂IMF＂） | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | － | － | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ |
| $f_{q, \text { all }}(q)$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\sqrt{ }$ | $\checkmark$ | $\times$ | $\sqrt{ }$ | $\sqrt{ }$ | － | － | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ |
| $f_{q ; M_{1}}(q)$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | － | － | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ | $\times$ |
| $\mathcal{B}_{\text {all }}$ |  | $\times$ | $\times$ | $\sqrt{ }$ | $\times$ | $\times$ | $\sqrt{ }$ |  |  | － | $\times$ | $\sqrt{ }$ | $\times$ | $\times$ | $\sqrt{ }$ | $\times$ |  | $\checkmark$ | $\checkmark$ | $\checkmark$ | ， | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ |
| $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ | $\sqrt{ }$ | $\times$ | $\times$ | $\times$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\checkmark$ | － | － | $\times$ | $\sqrt{ }$ | $\times$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\sqrt{ }$ | $\sqrt{ }$ | $\sqrt{ }$ | $\checkmark$ | $\checkmark$ | $\sqrt{ }$ |

Leinert et al．1997；Reid \＆Gizis 1997），and for late M－type stars and brown dwarfs the binary fraction decreases to $10-20 \%$ （Gizis et al．2003；Close et al．2003；Bouy et al．2003；Burgasser et al．2003；Siegler et al．2005；Maxted et al．2008；Joergens 2008）．Note that this correlation between mass and binary frac－ tion is also predicted by the hydrodynamical／sink particle sim－ ulations of Bate（2008）．Assuming that this observational trend is not induced by selection effects，is inconsistent with pairing functions PCP－I，PCP－III and SCP－I，for which the binary frac－ tion is independent of primary mass．Furthermore，observations have ruled out pairing functions RP and PCRP in various stel－ lar groupings（see Sect．5）．Pairing functions PCP－II，SCP－II and SCP－III remain options to describe the binary population，as for these pairing functions the binary fraction increases with in－ creasing stellar mass，and the average mass ratio decreases with increasing mass．However，in this paper we merely describe sim－ plistic（but frequently used）pairing functions．A deeper analysis， including a study of more complicated pairing functions，com－ bined with further observations，is necessary for a full descrip－ tion of the pairing function in the different stellar populations．

Twin binaries．Observationally，there is a high prevalence of massive binaries with a mass ratio close to unity，often referred to as the＂twin peak＂in the mass ratio distribution（Lucy \＆Ricco 1979；Tokovinin 2000；Pinsonneault \＆Stanek 2006；Lucy 2006； Söderhjelm 2007）．High－mass twin binaries are extremely rare for RP and PCRP．For pairing functions PCP and SCP，high－ mass twin binaries only occur frequently when this is explicitly put into the generating mass ratio distribution．A high prevalence of low－mass twin binaries，on the other hand，naturally results from all pairing functions except for PCP－I．In general，peaks in the mass ratio distribution can occur for any mass ratio（see，e．g．， Fig．6）．The location of the peak depends on the pairing function and the primary mass range，and，if applicable，the mass ratio distribution．In general，the peaks occur at low－$q$ for a sample of high－mass binaries，and at large－$q$ for a sample of low－mass stars．Pairing functions RP and PCRP are thus excluded，while pairing functions PCP and SCP can only result in massive twin binaries if the corresponding generating mass ratio distribution is strongly peaked to $q=1$ ．

The brown dwarf desert．The brown dwarf desert is defined as a deficit（not necessarily a total absence）of brown dwarf
companions，either relative to the frequency of stellar com－ panions，or relative to the frequency of planetary companions （McCarthy \＆Zuckerman 2004；Grether \＆Lineweaver 2006）． Theories have been developed that explain the existence of the brown dwarf desert using migration（Armitage \＆Bonnell 2002）or ejection（Reipurth \＆Clarke 2001）of brown darfs． The latter scenario＂embryo ejection＂is most popular，and pre－ dicts ejectino of brown dwarfs soon after their formation．In this scenario，brown dwarfs could be considered as failed stars． Kouwenhoven et al．（2007a），however，show that the scarcity of brown dwarf companions among intermediate－mass stars can also be explained by an extrapolation of the mass ratio distri－ bution into the brown dwarf regime．Metchev \＆Hillenbrand （2008）reach a similar conclusion．PCP－I／II／III are thus not ex－ cluded by the presence of the brown dwarf desert．

The（initial）mass distribution．The initial or present－day mass distribution $f_{\text {all }}(M)$ of a stellar population sets strong con－ straints on the star formation process，and is an important feature of each pairing function．The distribution is often derived after its members are securely identified（e．g．，Kroupa 2001；Preibisch et al．2002；Harayama et al．2008；Stolte et al．2008）．The mea－ sured mass distribution is often the distribution of single／primary masses $f_{M_{1, S}}(M)$ ，as it is not known which members are sin－ gle and which are binary，which results in a measured mass distribution that is biased to higher masses with respect to the overall mass distribution $f_{\text {all }}(M)$ ，which，if measured just after star formation，is the IMF．As stellar masses are often derived from measured luminosities，the presence of unresolved binaries and crowding may further bias the measured（initial or present－ day）mass distribution to higher masses（e．g．，Vanbeveren 1982； Maíz Apellániz 2008；Weidner et al．2008）．Over the last decade， considerable effort has been put into studying possible envi－ ronmental dependences of the IMF（see，e．g．，Elmegreen 2007； Kroupa 2008，for an overview and examples）．The IMF of a pop－ ulation is presumably a result of the form of the initial core mass function，and the primordial pairing function（i．e．，how these cores fragment into multiple systems，see Goodwin et al．2008）． An environmental dependence of the primordial pairing func－ tion（e．g．，mass ratio distribution，binary fraction）implies a dif－ ferent outcome of the star formation process with environment， it also almost certainly implies an environmental dependence of
the IMF (unless the core mass function changes in such as way as to mask this change).

### 9.2. Recovering the pairing function

The pairing function, $f_{M}(M), f_{q}(q)$ and $\mathcal{B}$ can in principle be derived from observations of binary systems, provided that the observations cover a large part of the parameter space $\left\{M_{1}, q\right\}$; see Sect. 8. A significant complication, however, is introduced by selection effects, in particular by detection limits that prevent the detection of faint companion stars. As an example, a twin peak for high mass binaries would rule out all of the models presented in this paper if $f_{q}(q)$ is assumed to be flat for PCP and SCP (see Sect. 9.1). On the other hand, a twin peak at the low mass end only rules out PCP-I (if the generating $f_{q}(q)$ is flat). So how should we proceed when interpreting the observational data? If we assume for the moment that we somehow know that one of the eight pairing mechanisms discussed in this paper occurs in Nature, we can advise the following:

- make sure the survey of the binary population is complete, i.e., that all primary masses are sampled. As is clear from Fig. 2, looking only at the low mass stars does not allow differentiating PCP-II, PCP-III, RP, PCRP, and SCP-I/II/III (all show a twin peak);
- examine not only the overall distribution of a certain parameter (such as $q$ or the binary fraction) but also study how it varies with primary mass. Again, from Fig. 2 it can be seen that when considering only a single row of panels it is not possible to easily differentiate the pairing mechanisms. However, when looking at the overall and specific mass ratio distributions, the differences do become clear;
- examine the combined behaviour of each parameter (the mass ratio distribution and the binary fraction) as a function of primary mass. The combination of these parameters constrains the possible pairing mechanisms significantly further;
- list all mechanisms capable of reproducing the observations as possible solutions to the inverse problem. Do not try to give a single answer if this is not warranted by the data.
In reality the number of possible models is of course much larger, especially if we start from arbitrary probability distributions that are not constrained by an understanding of the physics of binary formation. This is illustrated by the simple example in Sect. 9.3 which shows that allowing a power-law distribution for $f_{q}(q)$ means that only RP and PCRP can be excluded based on an observed flat distribution of $q$.

The only practical way of excluding models of binary formation based on observations is to treat the inverse problem with Monte Carlo methods where the observations are predicted from the model and compared to the real observations. In this method selection and observational biases should be included (see, e.g., Kouwenhoven 2006; Kouwenhoven et al. 2007b). Starting from models based on probability distributions for a set of parameters may not be the most useful way of constraining the formation mechanism for binaries as this leaves a lot of freedom. It is more fruitful (but also more difficult) to start from actual physical models of binary formation and see if these are capable of reproducing the observations.

A further complication occurs when one wants to recover the primordial pairing function, i.e., the pairing function that is present just after star formation, as the pairing function of a stellar population evolves over time as a result of several processes. During the first stages of star formation, the newly formed proto-binaries are affected by pre-main sequence eigenevolution
(Kroupa 1995a,b,c) due to interaction with the remaining gas in the circumbinary disc. Dynamical interactions can result in ionisation of binaries, the formation of new binary systems, and exchange interactions, and thus alters the pairing function of a stellar population (see, e.g., Kroupa et al. 1999, 2001; Preibisch et al. 2003; Duchêne et al. 2004; Köhler et al. 2006; Reipurth et al. 2007). Thie process is enhanced by gravitational focusing (e.g., Gaburov et al. 2008). The pairing function also changes due to stellar evolution, which can change the mass of one or both of the components of a binary system, and in some cases in a merger (e.g., Sills et al. 2002; Gaburov et al. 2008), or in the ejection of one of the components during a supernova event (e.g., Blaauw 1961). The primordial pairing function can be constrained using the technique of inverse dynamical population synthesis (e.g., Kroupa 1995c,b), in which the outcome of N -body simulations is compared with present-day binary population in a Monte Carlo way.

Finally we stress that any interpretation of observations of a binary population in terms of the formation of binaries should start by stating the assumptions one makes in order to restrict the number of solutions to explore. That is, in the context of what class of binary formation mechanisms are the observations interpreted?

### 9.3. An example - an observed flat mass ratio distribution

Most pairing functions result in a mass ratio distribution that varies with primary spectral type. For this reason one has to be cautious when interpreting the observations of a sample of binaries. Given the observed dataset, what is the pairing function, and what is the generating mass ratio distribution $f_{q}(q)$ ? The answer partially depends on the generating mass distribution, which we assume to be the Kroupa mass distribution for now. More importantly, the answer depends on the properties of the surveyed targets. In this example we analyse three cases: a sample where all binaries are studied, a sample of spectral type A/B targets, and a sample of brown dwarf targets. Using Monte-Carlo techniques we determine which pairing function is consistent with an observed flat mass ratio distribution $f_{q, \text { obs }}(q)=1$ for each subset, and, if applicable, which generating mass ratio distribution. We assume that the generating mass ratio distribution has the form $f_{q}(q) \propto q^{\gamma_{q}}$. We further assume that the observations are complete in the range $q_{\text {min }}<q \leq 1$ with $q_{\text {min }}=0.1$, and that no binaries with $q<q_{\text {min }}$ are observed due to incompleteness.

Table 5 shows the best-fitting values of $\gamma_{q}$ for each pairing function, assuming binaries with $q<0.1$ cannot be detected. We ignore the other selection effects. The numbers in the three columns represent the most compatible values of $\gamma_{q}$ for the overall mass ratio distribution $f_{q}(q)$, and for the specific mass ratio distribution of binaries with $\mathrm{A} / \mathrm{B}$ primaries and brown dwarf primaries, respectively. For each for the three samples, pairing functions RP and PCRP are excluded with high confidence; these are unable to reproduce the observed flat mass ratio distribution. The best-fitting value for pairing function PCP-I is $\gamma_{q}=0$ for all samples. This is not surprising, as for $f_{q ; M_{1}}(q)=$ $f_{q}(q)$ for this pairing function. For PCP-II and PCP-III the derived $\gamma_{q}$ for high-mass stars equals the observed value, but the other two samples contain more binaries with large mass ratios. For pairing functions SCP, the best-fitting intrinsic value of $\gamma_{q}$ is smaller than the observed value of $\gamma_{q}$ for high-mass binaries, but larger for low-mass binaries. This example illustrates that the inferred intrinsic pairing properties may be significantly different from the observed pairing properties, depending on the pairing function and the selected sample of binaries.

Table 5. Suppose that a stellar population has an observed mass ratio distribution $f_{q \text {,obs }}(q)=1$ for $q_{\text {min }}<q \leq 1$. What is the pairing function and the generating mass distribution $f_{q}(q)$ ? This table lists the exponent $\gamma_{q}$ of the generating mass distribution $f_{q}(q) \propto q^{\gamma_{q}}$ that is most compatible with the synthetic observations.

| Pairing function | $\gamma_{q}$ (all stars) | $\gamma_{q}$ (AB stars) | $\gamma_{q}(\mathrm{BDs})$ |
| :--- | :---: | :---: | :---: |
| Observed | 0.00 | 0.00 | 0.00 |
| PCP-I | 0.00 | 0.00 | 0.00 |
| PCP-II | 0.35 | 0.00 | 1.30 |
| PCP-III | 0.55 | 0.00 | 1.80 |
| RP | excluded | excluded | excluded |
| PCRP | excluded | excluded | excluded |
| SCP-I | 0.00 | -0.40 | 0.10 |
| SCP-II | 0.30 | -0.40 | 1.50 |
| SCP-III | 0.55 | -0.40 | 1.75 |

In practice, parameter distributions are often represented with a functional form. Suppose, for example, that we assume that the generating mass ratio distribution has the form $f_{q}(q) \propto$ $q^{\gamma_{q}}$. If we use this functional form for our model population, and compare simulated observations with the true observations for different values of $\gamma_{q}$, we will likely find a best-fitting $\gamma_{q}$. This does not necessarily mean that the generating mass ratio distribution has indeed the form $f_{q}(q) \propto q^{\gamma_{q}}$. In this example we have added another assumption, i.e., that the mass ratio distribution has the form $f_{q}(q) \propto q^{\gamma_{q}}$.

## 10. Summary and discussion

We have described several methods of pairing individual stars into binary systems. We refer to these algorithms as pairing functions. Each pairing function is characterized by a generating mass distribution $f_{M}(M)$ and a generating binary fraction $\mathcal{B}$, and most additionally by a generating mass ratio distribution $f_{q}(q)$. Each pairing function results in a significantly different binary population. Depending on the pairing function and the mass range of the binaries studied, the resulting binary population may or may not have a mass ratio distribution or binary fraction that is equal to $f_{q}(q)$ or $\mathcal{B}$, respectively. The binary fraction and mass ratio distribution generally depend strongly on the number of substellar objects in the population, and on the properties of the surveyed sample.

Eight pairing mechanisms are discussed in detail. For random pairing (RP) both components are randomly drawn from the mass distribution $f_{M}(M)$. For primary-constrained random pairing (PCRP), both components are drawn from $f_{M}(M)$, with the constraint that the companion is less massive than the primary. For primary-constrained pairing (PCP-I, PCP-II, and PCP-III), the primary is drawn from $f_{M}(M)$, and the companion mass is determined using a mass ratio distribution $f_{q}(q)$. For split-core pairing (SCP-I, SCP-II, and SCP-III), the core mass is drawn from $f_{M}(M)$, and the masses of the binary components are determined by the mass ratio $f_{q}(q)$, which splits up the core into two stars. The difference between the variants of pairing functions PCP and SCP lies in the treatment of low-mass companions (see Sect. 6). Seven pairing functions naturally result in a specific mass ratio distribution that depends on primary spectral type, and five naturally result in a mass-dependent binary fraction. Seven out of eight pairing functions always produce a twin peak for low-mass binaries, while none result in a twin peak for high-mass binaries, unless the generating mass ratio distribution is strongly peaked to $q=1$.

The differences between pairing functions are important for (i) the interpretation of observations; (ii) initial conditions of numerical simulations; and (iii) understanding the outcome of star formation:
(i) The interpretation of observations. The choice of the observational sample may mislead the observer in deriving the overall properties of a stellar population, as most pairing functions have a mass-dependent binary fraction and mass ratio distribution. If the binary fraction or mass ratio distribution of two samples (e.g., systems with B-type primaries and those with M-type primaries) are different, this does not necessarily mean that the underlying pairing function is different. A significant further complication is introduced by observational selection effects, which artificially decrease the binary fraction and increase the average mass ratio. The only practical way to account for these is by using a Monte Carlo approach, and to compare simulated observations of a model population with the results of the binarity survey, taking into account all sampling and selection effects (Sect. 9).
(ii) Initial conditions for numerical simulations. A choice for the pairing function has to be made when generating initial conditions for simulations of star cluster simulations with binaries. The simplest choice is random pairing (RP), although this pairing function is excluded from observations and not expected from star formation. When modelling star clusters, one has to be aware that most pairing mechanisms result in mass-dependent properties, such as mass ratio distribution and binary fraction. The choice of the pairing function affects the outcome of the simulations, such as the dynamical evolution of star clusters, mass segregation, and the number of contact binaries and supernovae.
(iii) Star formation. Different star formation scenarios result in different mechanisms of pairing stars into binary systems. After star formation, the pairing function is altered by dynamical interactions and stellar evolution. Random pairing, however, is not predicted by star formation models, and is excluded by observations (see Sect. 5). The binary fraction and mass ratio distribution can be used to distinguish between the different pairing functions. Although the pairing functions described in this paper are common in literature, we do not suggest that one of these pairing functions indeed describes the natural outcome of the star forming process. It is, for example, possible that the pairing properties are a function of primary mass or core mass. If this is the case, it may indicate different formation processes for different masses. Due to the lack of observations, this has not been studied in detail, apart from the extreme ends of the mass distribution (very massive stars and brown dwarfs). Kroupa et al. (2003) and Thies \& Kroupa (2007), for example, find that the observed IMF and binary population can be explained by separate formation mechanisms for stars and brown dwarfs (see Sect. 6.1.1). Nevertheless, the pairing functions described in this paper are useful tools to describe the outcome of star formation simulations. In order to constrain the primordial binary population from observations, one does not only have to take into account the selection effects, but also the change in the pairing function that has occured due to the effects of stellar and dynamical evolution. Over the last decade, considerable effort has been put into studying possible environmental dependences of the IMF. The IMF of a population is presumably a result of the form of the initial core mass function, and the
primordial pairing function (i.e., how these cores fragment into multiple systems, see Goodwin et al. 2008). An environmental dependence of the primordial pairing function (e.g., mass ratio distribution and binary fraction) implies a different outcome of the star formation process and almost certainly an environmental dependence of the IMF (unless the core mass function changes in such as way as to exactly mask this change); see Sect. 8 .

Each pairing function, as well as each subset of stars, results in a different mass ratio distribution and binary fraction. It is therefore of great importance to carefully study selection effects in observations, and to clearly state the pairing mechanism used in simulations, in order to make statements about the star formation process. The pairing functions described in this paper are likely too simplistic to describe a realistic stellar population. However, they are frequently used to describe observations and simulations. The next step forward is to fully characterize the binary population of several young stellar groupings; only in this way the star formation process can be recovered.

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## Appendix A: The pairing function and mass ratio distributions

In this appendix we discuss for each of the pairing mechanisms described in the main paper how to calculate the mass ratio distribution function $f_{q}(q)$. The masses of primary $\left(M_{1}\right)$ and secondary $\left(M_{2}\right)$ are drawn from the same generating mass distribution $f_{M}(M)$, or alternatively, the primary is drawn from $f_{M}(M)$ while the secondary is drawn from a generating mass ratio distribution. The generating mass distribution is treated throughout this appendix as a probability density:
$P(M \leq t)=F_{M}(t)=\int_{-\infty}^{t} f_{M}(x) \mathrm{d} x$,
and
$\int_{-\infty}^{\infty} f_{M}(M) \mathrm{d} M=1$.
In the subsequent section we describe for each of the pairing functions first how to calculate $f_{q}(q)$ without specifying the generating mass distribution and then we work out the resulting mass-ratio distributions for the single power-law $f_{M}(M)$ which is given here as:
$f(M)=a M^{-\alpha} \quad$ with $\quad a=\frac{1-\alpha}{d^{1-\alpha}-c^{1-\alpha}}=\frac{\gamma c^{\gamma}}{1-(c / d)^{\gamma}}$,
where $\alpha \neq 1$ and $\gamma=\alpha-1$ and $c$ and $d$ are the lower and upper limits on the mass distribution.

## A.1. Random pairing

The simplest choice for a pairing function is that of "random pairing" (RP). In this case both component masses are drawn
independently from $f_{M}(M)$ and swapped, if necessary, to ensure that $M_{2} \leq M_{1}$. In this case the joint distribution function $f_{\mathrm{rp}}\left(M_{1}, M_{2}\right)$ for the component masses is given by:
$f_{\mathrm{rp}}\left(M_{1}, M_{2}\right)=2 f_{M_{1}}\left(M_{1}\right) f_{M_{2}}\left(M_{2}\right)$,
where the factor 2 accounts for the fact that the pairs of masses are swapped in order to ensure that $M_{2} \leq M_{1}$. The masses are restricted by a lower limit $c$ and and upper limit $d$ which leads to the domain of $f_{\mathrm{rp}}\left(M_{1}, M_{2}\right)$ being defined as $c \leq M_{1} \leq d$ and $c \leq M_{2} \leq M_{1}$.

## A.1.1. Mass ratio distributions for RP

First, general expressions for the mass ratio distributions are derived before working out specific examples. To derive the mass ratio distribution we follow the appendix of Piskunov \& Mal' Kov (1991) and derive $f_{q}(q)$ from its cumulative distribution function:

$$
\begin{align*}
P(y / x \leq q)=F_{q}(q) & =\iint_{\mathrm{S}} f_{\mathrm{rp}}(x, y) \mathrm{d} x \mathrm{~d} y \\
& =\int_{c / q}^{d} \int_{\mathrm{c}}^{q x} 2 f(x) f(y) \mathrm{d} y \mathrm{~d} x \tag{A.5}
\end{align*}
$$

where for ease of notation we use $M_{1}=x$ and $M_{2}=y$. The integration domain $S$ is shown in the left-hand panel of Fig. 5. The probability density $f_{q}$ is then given by:
$f_{q}(q)=\frac{\mathrm{d}}{\mathrm{d} q} F_{q}(q)$.
In observational surveys for binarity one is often restricted to a certain range of primary spectral types for observational reasons. To derive the effects of the selection on primary mass, the derivation of $f_{q}(q)$ again proceeds via the cumulative distribution function, which is now given by:
$F_{q}(q)=P\left(y / x \leq q \mid x_{1} \leq x \leq x_{2}\right) \propto \iint_{S^{\prime}} f_{\mathrm{rp}}(x, y) \mathrm{d} x \mathrm{~d} y$,
where the primary mass range is restricted to $x_{1} \leq M_{1} \leq x_{2}$ and the integration domain $S^{\prime}$ is as shown in Fig. 5. The integration domain limits now depend on whether $q$ is larger or smaller than $c / x_{1}$. For $q \leq c / x_{1}$ the integration domain $S^{\prime}$ is given by: $c / q \leq x \leq x_{2} \wedge c \leq y \leq q x$ (middle panel in Fig. 5), while for $q>c / x_{1} S^{\prime}$ is defined by $x_{1} \leq x \leq x_{2} \wedge c \leq y \leq q x$ (right-hand panel in Fig. 5). The minimum possible value of $q$ is $c / x_{2}$ in this case, implying that $F(q)=0$ for $q<c / x_{2}$. Thus the expression for the mass ratio distribution now becomes:
$F_{q}(q)=\left\{\begin{array}{lr}0 & q<\frac{c}{x_{2}} \\ k \int_{c / q}^{x_{2}} \int_{c}^{q x} 2 f(x) f(y) \mathrm{d} y \mathrm{~d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}}^{x_{2}} \int_{c}^{q x} 2 f(x) f(y) \mathrm{d} y \mathrm{~d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
and the probability density $f_{q}$ is again derived according to Eq. (A.6). The normalisation constant $k$ can be calculated from the condition $F_{q}(1)=1$.

## A.1.2. Binary fractions for RP

For RP, the overall binary fraction $\mathcal{B}_{\text {all }}$ is equal to the generating binary fraction $\mathcal{B}$. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$, however, is generally a function of primary mass. For a given
primary mass, the number of single stars $S\left(M_{1}\right)$ with mass $M_{1}$ and the number of binary stars $B\left(M_{1}\right)$ with primary mass $M_{1}$ is given by:
$S_{M_{1}}=S f_{M}\left(M_{1}\right) \mathrm{d} M_{1}$,
where $S$ is the total number of single stars in the system, and
$B_{M_{1}}=2 B f_{M}\left(M_{1}\right) \mathrm{d} M_{1} \int_{\mathrm{c}}^{M_{1}} f_{M}\left(M^{\prime}\right) \mathrm{d} M^{\prime}$,
where $B$ is the total number of binary stars in the system and $c$ is the minimum stellar mass. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)=B_{M_{1}} /\left(S_{M_{1}}+B_{M_{1}}\right)$ is then given by:
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\frac{2 B F_{M}\left(M_{1}\right)}{2 B F_{M}\left(M_{1}\right)+S}=\left(\frac{\mathcal{B}^{-1}-1}{2 F_{M}\left(M_{1}\right)}+1\right)^{-1}$,
where $F_{M}\left(M_{1}\right)$ is the cumulative mass distribution (i.e., the primitive of $f_{M}$ ) evaluated at mass $M_{1}$.

An example of the mass-dependent binary fraction resulting from RP is shown in Fig. 7. Clearly, we have $\mathcal{B}_{M_{1}}\left(M_{1}\right) \equiv 1$ when $S=0$, and $\mathcal{B}_{M_{1}}\left(M_{1}\right) \equiv 0$ when $B=0$. Also, $\mathcal{B}_{M_{1}}\left(M_{\max }\right)=$ $2 B /(2 B+S) \geq \mathcal{B}$ and $\mathcal{B}_{M_{1}}\left(M_{\min }\right)=0$. The mass at which the specific binary fraction equals the generating binary fraction can be found by solving $\mathcal{B}=\mathcal{B}_{M_{1}}\left(M_{1}\right)$, which gives $F_{M}\left(M_{1}\right)=0.5$. In other words, for pairing function RP the overall binary fraction can be found at the median stellar mass $\langle M\rangle$. For higher primary masses, $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is higher, and for lower primary masses, $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is lower.

## A.1.3. Single power-law mass distribution

From Eqs. (A.5), (A.6), and (A.8) mass ratio distributions for specific choices of the generating mass distribution can be derived. The single power-law defined in Eq. (A.3) is considered here. Without restrictions on the primary mass range, the cumulative distribution for $q$ is given by:

$$
\begin{aligned}
F_{q}(q) & =\int_{c / q}^{d} \int_{c}^{q x} 2 a^{2} x^{-\alpha} y^{-\alpha} \mathrm{d} y \mathrm{~d} x \\
& =\frac{a^{2}}{\gamma^{2} c^{2 \gamma}}\left((c / d)^{2 \gamma} q^{-\gamma}+q^{\gamma}-2(c / d)^{\gamma}\right) .
\end{aligned}
$$

The probability density for $q$ is then obtained as the derivative with respect to $q$ of $F_{q}$ which after some algebraic manipulation leads to:
$f_{q}(q)=\frac{\gamma}{\left(1-(c / d)^{\gamma}\right)^{2}}\left(q^{\alpha-2}-(c / d)^{2 \gamma} q^{-\alpha}\right)$.
For $c / d \ll 1$ one obtains:
$f_{q}(q) \approx\left\{\begin{array}{ll}(\alpha-1) q^{\alpha-2} & \alpha>1 \\ (1-\alpha) q^{-\alpha} & \alpha<1\end{array}\right.$.
These approximations are poor when $\alpha \approx 1$.
Figure A. 1 shows a number of examples of the resulting mass ratio distributions for different values of $c / d$ and the powerlaw slope $\alpha$. The values for $\alpha$ represent the Salpeter mass distribution (2.35), the slopes at the lower mass end for the Kroupa (2001) mass distribution (1.3 and 0.3), and a possible slope at the very low mass end of the mass distribution, where the number of stars increases with $m$. The latter value may occur in a multi-part power-law mass distribution with a real turnover at the low mass end. For $\alpha>1$ the mass distribution decreases
with $\log m$ and for $\alpha<1$ it increases with $\log m$. However this does not represent a real turnover in the mass distribution, the number of stars still increases as the mass goes down as long as $\alpha>0$. Note how the peak in the $q$ distribution changes as the values of $\alpha$ and $c / d$ are changed. For very small values of $c / d$ one can see from the approximation (A.13) that $f_{q}$ will be flat for $\alpha \approx 0$ and $\alpha \approx 2$ and that it will peak at small values of $q$ for $0<\alpha<2$. For values of $\alpha$ larger than 2 the number of low-mass stars is so dominant that large values of $q$ are favoured (i.e., both $M_{1}$ and $M_{2}$ are likely to be small). Conversely, for $\alpha<0$ the rise of the number of stars with $m$ again favours large values of $q$. For $0<\alpha<2$ the ratio of probabilities to obtain low- or high-mass stars is such that drawing two equal mass star is unlikely thus favouring small values of $q$.

To find the expression for $f_{q}$ when the primary mass range is restricted Eq. (A.8) has to be worked out for the single powerlaw mass distribution. For $0<q<c / x_{2} F_{q}(q)=0$, while for $c / x_{2} \leq q<c / x_{1} F_{q}(q)$ is given by:

$$
\begin{aligned}
F_{q}(q) \propto & \int_{c / q}^{x_{2}} \int_{\mathrm{c}}^{q x} x^{-\alpha} y^{-\alpha} \mathrm{d} y \mathrm{~d} x=\frac{1}{1-\alpha} \\
& \times \int_{c / q}^{x_{2}} x^{-\alpha}\left(q^{1-\alpha} x^{1-\alpha}-c^{1-\alpha}\right) \mathrm{d} x \\
\propto & \frac{1}{\gamma^{2}}\left(\frac{1}{2}\left(q^{-\gamma} x_{2}^{-2 \gamma}+c^{-2 \gamma} q^{\gamma}\right)-c^{-\gamma} x_{2}^{-\gamma}\right) .
\end{aligned}
$$

For $c / x_{1} \leq q \leq 1$ the cumulative distribution for $q$ is given by:
$F_{q}(q) \propto \int_{x_{1}}^{x_{2}} \int_{c}^{q x} x^{-\alpha} y^{-\alpha} \mathrm{d} y \mathrm{~d} x$

$$
\propto \frac{1}{\gamma^{2}}\left(\frac{1}{2} q^{-\gamma}\left(x_{2}^{-2 \gamma}-x_{1}^{-2 \gamma}\right)-c^{-\gamma}\left(x_{2}^{-\gamma}-x_{1}^{-\gamma}\right)\right) .
$$

The normalisation constant $k$ for the probability density of $q$ can now be found by substituting $q=1$ in the last expression for $F_{q}(q)$ :
$k=\gamma^{2}\left(\frac{1}{2}\left(x_{2}^{-2 \gamma}-x_{1}^{-2 \gamma}\right)-c^{-\gamma}\left(x_{2}^{-\gamma}-x_{1}^{-\gamma}\right)\right)^{-1}$.
Now the expressions for $f_{q}(q)$ can be derived by taking the derivative of the integrals above:
$f_{q}(q)= \begin{cases}0 & q<\frac{c}{x_{2}} \\ \frac{k}{2 \gamma c^{2 \gamma}}\left(q^{\alpha-2}-\left(\frac{c}{x_{2}}\right)^{2 \gamma} q^{-\alpha}\right) \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}}, \\ \frac{k}{2 \gamma x_{1}^{2 \gamma}}\left(1-\left(\frac{x_{1}}{x_{2}}\right)^{2 \gamma}\right) q^{-\alpha} & \frac{c}{x_{1}} \leq q \leq 1\end{cases}$
If $x_{1}=x_{2} f_{q}(q)$ will be proportional to $q^{-\alpha}$ for $c / x_{1} \leq q \leq 1$ and zero otherwise:
$f_{q}(q)= \begin{cases}0 & q<\frac{c}{x_{1}} \\ \frac{1-\alpha}{1-\left(c / x_{1}\right)^{1-\alpha}} q^{-\alpha} & \frac{c}{x_{1}} \leq q \leq 1\end{cases}$
Figure A. 2 shows examples of the behaviour of $f_{q}(q)$ when the primary mass range is restricted. The solid lines show the distribution of $q$ for a complete binary sample. The dotted lines show what happens if there is only a lower limit $(>c)$ on the primary mass range. For $\alpha>1$ the large mass ratios are preferentially removed because low-mass primaries are removed. Conversely


Fig. A.1. Mass ratio distributions $f_{q}(q)$ for random pairing from the single power-law $f_{M}(M)$. The curves are shown for $\alpha=2.35,1.30,0.30,-2.50$ and four ratios of the lower to the upper mass limit of $f_{M}(M): c / d=0.01,0.1,0.3,0.5$.
if there is only an upper limit on $M_{1}(<d)$ only small mass ratios are removed and the resulting $f_{q}(q)$ is given by the dashed lines. The dot-dashed lines show a generic case with $c<x_{1}<x_{2}<d$. The latter case for very narrow primary mass ranges $f_{q}(q)$ will converge to $f_{q}(q) \propto q^{-\alpha}$. For $0<\alpha<1$ the effect of mass selection is to remove the small values of $q$, thus flattening the distribution and moving the peak. For $\alpha<0$ the mass selection does not have much effect. This figure illustrates that the interpretation of mass ratio distributions in terms of random pairing is not straightforward unless the generating mass function is well known and the observations are indeed complete over a known primary mass range.

The expression for the specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ resulting from random pairing is listed in Table A.3, and, for the Salpeter generating mass distribution $(\alpha=2.35)$, shown in Fig. 7.

## A.2. Primary-constrained random pairing

In this case (PCRP) the primary and secondary are again drawn independently from the generating mass distribution, however for the secondary the condition $M_{2} \leq M_{1}$ is imposed before drawing the secondary mass. That is the probability density $f_{M_{2}}\left(M_{2}\right)$ is re-normalised to the interval $\left[c, M_{1}\right]$ (recall that $c$ is the lower mass limit on the mass distribution).

## A.2.1. Mass ratio distributions for PCRP

Writing the re-normalised secondary mass distribution as $f_{M_{2}}^{\prime}\left(M_{2}\right)=f^{\prime}(y)$, the expression for the joint probability distribution $f_{\text {pcrp }}\left(M_{1}, M_{2}\right)$ is:
$f_{\text {pcrp }}\left(M_{1}, M_{2}\right)=f_{M_{1}}\left(M_{1}\right) f^{\prime}{ }_{M_{2}}\left(M_{2}\right)$,
which is normalised to 1 due to the re-normalisation of $f^{\prime}{ }_{M_{2}}\left(M_{2}\right)$.

To derive $f_{q}$ one can proceed as for the RP case. The integration domain is as shown in the left-hand panel of Fig. 5 and the expression for $F(q)$ can be written as:
$F(q)=\int_{c / q}^{d} \int_{c}^{q x} f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x$.
Note that the normalisation constant for $f^{\prime}(y)$ depends on $x$. The expression for the restricted primary mass range becomes:
$F_{q}(q)=\left\{\begin{array}{lc}0 & q<\frac{c}{x_{2}} \\ k \int_{c / q}^{x_{2}} \int_{c}^{q x} f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}}, \\ k \int_{x_{1}}^{x_{2}} \int_{c}^{q x} f(x) f^{\prime}(y) \mathrm{d} y \mathrm{~d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$
where the integration domains for $c / x_{2} \leq q<c / x_{1}$ and $c / x_{1} \leq$ $q \leq 1$ are shown in the middle and right-hand panels of Fig. 5, respectively.


Fig. A.2. Mass ratio distributions $f_{q}(q)$ for random pairing from a single power-law $f_{M}(M)$. The solid curve shows the complete mass ratio distribution (for all binaries in the population). The other curves show what happens to the observed $f_{q}(q)$ if the primary mass $M_{1}$ is restricted to $x_{1} \leq M_{1} \leq x_{2}$. The curves are shown for $\alpha=2.35,1.30,0.30,-2.50$. The value of $c / d$ is 0.01 and $x_{1}$ and $x_{2}$ are listed in the panels.

## A.2.2. Binary fractions for PCRP

For PCRP, the overall binary fraction $\mathcal{B}_{\text {all }}$ equals $\mathcal{B}$, and specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is equal to the generating binary fraction $\mathcal{B}$ for any mass $M_{1}$.

## A.2.3. Single power-law mass distribution

For the single power-law mass distribution, with lower and upper mass limits $c$ and $d$, the joint distribution of $x=M_{1}$ and $y=$ $M_{2}$ is:
$f_{\text {pcrp }}(x, y)=a_{1} x^{-\alpha} a_{2}(x) y^{-\alpha}$,
where the normalisation constant for $f(y)$ depends on the primary mass $x$ :
$a_{1}=\frac{\gamma c^{\gamma}}{1-(c / d)^{\gamma}} \quad$ and $\quad a_{2}(x)=\frac{\gamma c^{\gamma}}{1-(c / x)^{\gamma}}$.
A distinction has to be made between the cases $\alpha<1$ and $\alpha>1$ as will become clear below.
$F_{q}(q)$ is given by:

$$
\begin{aligned}
F_{q}(q) & =\int_{c / q}^{d} \int_{\mathrm{c}}^{q x} \frac{\gamma^{2} c^{2 \gamma} x^{-\alpha}}{\left(1-(c / d)^{\gamma}\right)\left(1-(c / x)^{\gamma}\right)} y^{-\alpha} \mathrm{d} y \mathrm{~d} x \\
& =a_{1} \int_{c / q}^{d} x^{-\gamma-1}\left[\frac{q^{-\gamma}(c / x)^{\gamma}-1}{(c / x)^{\gamma}-1}\right] \mathrm{d} x .
\end{aligned}
$$

Using the substitution $z=c / x$ (which implies $\mathrm{d} x=-\left(c / z^{2}\right) \mathrm{d} z$ ) the integral can be written in a more convenient form and its solution can be written as linear combination of the terms $z^{\gamma}$ and $\ln \left(1-z^{\gamma}\right)$ for $\alpha>1$, while for $\alpha<1$ the terms $z^{\gamma}$ and $\ln \left(z^{\gamma}-1\right)$ are involved.

The resulting expression for $f_{q}(q)$ is:
$f_{q}(q)=\left\{\begin{array}{l}\frac{\gamma q^{-\gamma-1}}{1-(c / d)^{\gamma}}\left[\ln \left(\frac{(c / d)^{\gamma}-1}{q^{\gamma}-1}\right)+(c / d)^{\gamma}-q^{\gamma}\right] \alpha<1 \\ \frac{\gamma q^{-\gamma-1}}{1-(c / d)^{\gamma}}\left[\ln \left(\frac{1-(c / d)^{\gamma}}{1-q^{\gamma}}\right)+(c / d)^{\gamma}-q^{\gamma}\right] \alpha>1\end{array}\right.$.
This expression diverges as $q \rightarrow 1$ which is due to the lowmass end of the primary mass distribution. As $M_{1}$ approaches $c$ the values of $q$ will increasingly all be close to 1 . The rate at which $f_{q}$ diverges depends on the value of $\alpha$. For $\alpha>0$ the mass distribution peaks at the low-mass end thus causing a rapid divergence.

Figure A. 3 shows four examples of the mass ratio distribution for PCRP from a single power-law mass distribution. The curves are for the values of $c / d$ of $0.01,0.1,0.3$, and 0.5 .

For the restricted primary mass range Eq. (A.19) has to be worked out. For $c / x_{2} \leq q<c / x_{1}$ the expression for $F_{q}$ is:
$F_{q}(q) \propto \int_{c / q}^{x_{2}} \int_{c}^{q x} \frac{\gamma c^{\gamma} x^{-\alpha}}{1-(c / x)^{\gamma}} y^{-\alpha} \mathrm{d} y \mathrm{~d} x$,


Fig. A.3. Mass ratio distributions $f_{q}(q)$ for PCRP from a single power-law mass distribution. The curves are shown for $\alpha=2.35,1.30,0.30,-2.50$ and four ratios of the lower to the upper mass limit of the mass distribution: $c / d=0.01,0.1,0.3,0.5$.
which can be worked out to:
$F_{q}(q) \propto \int_{c / q}^{x_{2}} \frac{c^{\gamma} q^{-\gamma} x^{-2 \gamma-1}-x^{-\gamma-1}}{(c / x)^{\gamma}-1} \mathrm{~d} x$.
This integral can be worked out in the same way as for the full primary mass range. For $c / x_{1} \leq q \leq 1$ the expression for $F_{q}$ becomes:
$F_{q}(q) \propto \int_{x_{1}}^{x_{2}} \frac{c^{\gamma} q^{-\gamma} x^{-2 \gamma-1}-x^{-\gamma-1}}{(c / x)^{\gamma}-1} \mathrm{~d} x$.
The expression for $f_{q}(q)$ for $\alpha>1$ is:
$f_{q}(q)=\left\{\begin{array}{ll}0 & q<\frac{c}{x_{2}} \\ \frac{k}{q^{\gamma+1}}\left(\ln \left(\frac{1-\left(c / x_{2}\right)^{\gamma}}{1-q^{\gamma}}\right)+\left(\frac{c}{x_{2}}\right)^{\gamma}-q^{\gamma}\right) & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ \frac{k}{q^{\gamma+1}}\left(\ln \left(\frac{\left.1-(c / x)^{\gamma}\right)^{\gamma}}{1-\left(c / x_{1}\right)^{\gamma}}\right)+\left(\frac{c}{x_{2}}\right)^{\gamma}-\left(\frac{c}{x_{1}}\right)^{\gamma}\right) & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$.
For $\alpha<1$ the expression for $f_{q}$ is:
$f_{q}(q)= \begin{cases}0 & q<\frac{c}{x_{2}} \\ \frac{k}{q^{\gamma+1}}\left(\ln \left(\frac{\left(c / x_{2}\right)^{\gamma}-1}{q^{\gamma}-1}\right)+\left(\frac{c}{x_{2}}\right)^{\gamma}-q^{\gamma}\right) & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ \frac{k}{q^{\gamma+1}}\left(\ln \left(\frac{1-\left(c / x_{2}\right)^{\gamma}}{1-\left(c / x_{1}\right)^{\gamma}}\right)+\left(\frac{c}{x_{2}}\right)^{\gamma}-\left(\frac{c}{x_{1}}\right)^{\gamma}\right) \frac{c}{x_{1}} \leq q \leq 1\end{cases}$
The normalisation constant $k$ is:
$k=\gamma\left(\left(c / x_{1}\right)^{\gamma}-\left(c / x_{2}\right)^{\gamma}\right)^{-1}$.

Again, if $x_{1}=x_{2}, f_{q}(q)$ will be proportional to $q^{-\alpha}$ for $c / x_{1} \leq$ $q \leq 1$ and zero otherwise and the expression is the same as Eq. (A.16).

Examples of what happens in the case of PCRP when a selection is done on primary mass are shown in Fig. A.4. The behaviour is qualitatively the same as for the random pairing case.

## A.3. Primary-constrained pairing

For the pairing mechanism discussed now the assumption is that there is a physical process which sets the primary mass and the mass ratio of the binary, rather than setting the masses of primary and secondary. For the primary constrained pairing mechanism (PCP) the assumption is that $M_{2}$ is determined from $M_{1}$ through the mass ratio $q$. The probability densities for $M_{1}$ and $q$ are specified in this case and they are assumed to be independent. That is:
$f_{\text {pcp }}\left(M_{1}, q\right)=f_{M_{1}}\left(M_{1}\right) h_{q}(q)$,
where the generating mass ratio distribution is written as $h_{q}(q)$ in order to distinguish it from the observed mass ratio distribution $f_{q}(q)$. The latter can be obtained by integrating $f_{\mathrm{pcp}}\left(M_{1}, q\right)$ over $M_{1}$ :
$f_{q}(q)=\int f_{\mathrm{pcp}}\left(M_{1}, q\right) \mathrm{d} M_{1}$.
There are a number of choices one can make in generating a binary population from $f_{\mathrm{pcp}}$. The generating mass ratio distribution


Fig. A.4. Mass ratio distributions $f_{q}(q)$ for PCRP from a single power-law mass distribution. The solid curves show the complete mass ratio distribution (for all binaries in the population). The other curves show what happens to the observed $f_{q}(q)$ if the primary mass $M_{1}$ is restricted to $x_{1} \leq M_{1} \leq x_{2}$. The value of $c / d$ is 0.01 and $x_{1}$ and $x_{2}$ are listed in the panels.
is assumed to be specified for the interval $0<q \leq 1$ and $c \leq M_{1} \leq d$, which leads to the following three possibilities:

PCP-I All values of $q$ are allowed which means that for a given primary mass $M_{1}, 0<M_{2} \leq M_{1}$. Thus binary systems with "sub-stellar" secondary components are also allowed.
PCP-II All values of $q$ are allowed but only binary systems for which the secondary is "stellar" (i.e. $M_{2} \geq c$ ) are retained. This amounts to integrating $f_{\text {pcp }}$ over the range $c / q \leq M_{1} \leq d$ and re-normalising the resulting distribution of $q$ to 1 .
PCP-III Only values of $q$ for which $M_{2} \geq c$ are allowed, i.e. $c / M_{1} \leq q \leq 1$. This is equivalent to re-normalising the generating distribution $h_{q}(q)$ to the interval $\left[c / M_{1}, 1\right]$.

## A.3.1. Mass ratio distributions for PCP

Again we first derive the general expressions for $f_{q}(q)$ before discussing specific examples.

PCP-I In this case one always obtains $f_{q}(q)=h_{q}(q)$ because the distributions of $q$ and $M_{1}$ are independent. Also when restricting the primary mass range the observed mass ratio distribution is equal to the generating distribution.

PCP-II Here the systems with $M_{2}<c$ are discarded and then:
$f_{q}(q)=k \int_{c / q}^{d} f_{\mathrm{pcp}}\left(M_{1}, q\right) \mathrm{d} M_{1}=k \int_{c / q}^{\mathrm{d}} h_{q}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$,
where $k$ is a normalisation constant which ensures that $\int f_{q}(q) \mathrm{d} q=1$. Of course $f_{q}(q)=0$ for $0<q<c / d$. For a restricted primary mass range, $c \leq x_{1} \leq M_{1} \leq x_{2} \leq d$, the expression for $f_{q}$ is:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c / q}^{x_{2}} h_{q}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1} & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}}^{x_{2}} h_{q}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1} & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
where $k$ is again a normalisation constant. For $q<c / x_{2}$ the value of $c / q$ is larger than $x_{2}$ so the lower integration limit of the integral in Eq. (A.28) becomes $x_{2}$, i.e., all companions with $q<c / x_{2}$ are rejected, as their mass is smaller than the minimum mass $c$. For $q>c / x_{1}$ we have $c / q<x_{1}$ so the lower integration limit of the integral should be fixed at $x_{1}$.

PCP-III In this case the generating mass ratio distribution is renormalised to the interval $\left[c / M_{1}, 1\right]$, resulting in a generating distribution $h^{\prime}{ }_{q}(q)$. The expressions for $f_{q}$ are then derived as for the PCP-II case:
$f_{q}(q)=\int_{c / q}^{d} f_{\mathrm{pcp}}\left(M_{1}, q\right) \mathrm{d} M_{1}=\int_{c / q}^{d} h_{q}^{\prime}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$,
where a normalisation constant is now not needed $\left(h_{q}^{\prime}(q)\right.$ is normalised). For the restricted primary mass range the expression is:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c / q}^{x_{2}} h_{q}^{\prime}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1} & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}}^{x_{2}} h^{\prime}{ }_{q}(q) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1} & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$.

## A.3.2. Binary fractions for PCP

PCP-I The overall and specific binary fractions are always equal to the generating binary fraction $\mathcal{B}$ for PCP-I

PCP-II Due to the rejection of low-mass companions with $M_{2}<$ $c$, i.e., with $q<q_{\min }\left(M_{1}\right)=c / M_{1}$, the binary fraction varies with primary mass:
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} \int_{q_{\text {min }}\left(M_{1}\right)}^{1} h_{q}(q) \mathrm{d} q<\mathcal{B}$,
where $h_{q}(q)$ is the generating mass ratio distribution. Note that $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ is independent of the generating mass distribution. An example of the mass-dependent binary fraction resulting from PCP-II is shown in Fig. 7. For high-mass binaries with $M_{1} \approx d$, very few companions are rejected as $q_{\min }\left(M_{1}\right) \ll 1$, and hence $\mathcal{B}_{M_{1}}\left(M_{1}\right) \approx \mathcal{B}$. For the lowest mass binaries in the sample, on the other hand, $q_{\min }\left(M_{1}\right) \approx 1$, and therefore $\mathcal{B}_{M_{1}}\left(M_{1}\right) \approx 0$. The overall binary fraction can be found by integrating over primary mass
$\mathcal{B}_{\text {all }}=\int_{c}^{d} \mathcal{B}_{M_{1}}\left(M_{1}\right) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$,
and is always lower than $\mathcal{B}$.

PCP-III The overall and specific binary fractions are always equal to the generating binary fraction $\mathcal{B}$ for PCP-III

## A.3.3. Uniform mass ratio distribution and single power-law mass distribution

The following is assumed for $h_{q}$ and $f_{M_{1}}\left(M_{1}\right)=f(x)$ (using $x=$ $M_{1}$ for ease of notation):

$$
\begin{aligned}
h_{q}(q) & =1 \quad 0<q \leq 1 \\
f(x) & =a x^{-\alpha} \quad c \leq x \leq d,
\end{aligned}
$$

where $a=\gamma c^{\gamma} /\left(1-(c / d)^{\gamma}\right)$ and $\alpha \neq 1$.

The PCP-II case For the full primary mass range we have from Eq. (A.28):
$f_{q}(q)=k \int_{c / q}^{d} a x^{-\alpha} \mathrm{d} x=k \frac{q^{\gamma}-(c / d)^{\gamma}}{1-(c / d)^{\gamma}}$.
The expression for $k$ can be found by solving $\int_{c / d}^{1} f_{q}(q) \mathrm{d} q=1$ for $k$ which leads to the final expression for the observed mass ratio distribution:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{d} \\ \frac{(\gamma+1)\left(q^{\gamma}-(c / d)^{\gamma}\right)}{1-(\gamma+1)(c / d)^{\gamma}+\gamma(c / d)^{\gamma+1}} & \frac{c}{d} \leq q \leq 1\end{array}\right.$.

If the primary mass range is restricted to $x_{1} \leq q \leq x_{2}$ the expression for $f_{q}$ can be derived using Eq. (A.29):
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k\left(q^{\gamma}-\left(c / x_{2}\right)^{\gamma}\right) & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k\left(\left(c / x_{1}\right)^{\gamma}-\left(c / x_{2}\right)^{\gamma}\right) & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
where the normalisation constant is:
$k=\frac{\gamma+1}{(\gamma+1)\left[\left(c / x_{1}\right)^{\gamma}-\left(c / x_{2}\right)^{\gamma}\right]-\gamma\left[\left(c / x_{1}\right)^{\gamma+1}-\left(c / x_{2}\right)^{\gamma+1}\right]}$.
Note that if $x_{1}=x_{2}$ the distribution of $q$ will be uniform on the interval $\left[c / x_{1}, 1\right]$, i.e. $f_{q}(q)=1 /\left(1-c / x_{1}\right)$. The equations above do not hold for $\alpha=0$. This case is easily evaluated and the result for $\alpha=0$ is:
$f_{q}(q)=\frac{1}{1-c / d+(c / d) \ln (c / d)}\left(1-\frac{c / d}{q}\right)$.

The PCP-III case Now the mass ratio distribution is renormalised to the interval $[c / x, 1]$ which gives:
$h_{q}^{\prime}(q)=\frac{1}{1-c / x}$.
For the full primary mass range $f_{q}$ is obtained from Eq. (A.30):
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{d} \\ a \int_{c / q}^{d} \frac{x^{1-\alpha}}{x-c} \mathrm{~d} x & \frac{c}{d} \leq q \leq 1\end{array}\right.$.
Restricting the primary mass range leads to:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c / q}^{x_{2}} \frac{x^{1-\alpha}}{x-c} \mathrm{~d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}}^{x_{2}} \frac{x^{1-\alpha}}{x-c} \mathrm{~d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
The expressions for the PCP-III case can be written in a slightly more convenient form when using the substitution $z=c / x$. For the full primary mass range:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{d} \\ \frac{\gamma}{1-(c / d)^{\gamma}} \int_{c / d}^{q} \frac{z^{\gamma-1}}{1-z} \mathrm{~d} z & \frac{c}{d} \leq q \leq 1\end{array}\right.$,
and for the restricted range:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \frac{\gamma}{1-(c / d)^{\gamma}} \int_{c / x_{2}}^{q} \frac{z^{\gamma-1}}{1-z} \mathrm{~d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \frac{\gamma}{1-(c / d)^{\gamma}} \int_{c / x_{2}}^{c / x_{1}} \frac{z^{\gamma-1}}{1-z} \mathrm{~d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$.
The integrals involving the term $z^{\gamma-1} /(1-z)$ can be solved and the general expression involves the hypergeometric function ${ }_{2} F_{1}$. The expression is:
$\int \frac{z^{\gamma-1}}{1-z} \mathrm{~d} z=\frac{z^{\gamma}}{\gamma}+\frac{{ }_{2} F_{1}(\gamma+1,1 ; \gamma+2 ; z) z^{\gamma+1}}{\gamma+1}$.
From the properties of the hypergeometric function (see Gradshteyn \& Ryzhik 2007, Chapt. 9) it follows that because $(\gamma+1)+1-(\gamma+2)=0$ the expression above converges throughout the unit circle in the complex plane except at $|z|=1$. So
there will be a singularity at $q=1$ for the full primary mass range case.

For integer values of $\gamma$ (or $\alpha$ ) special care should be taken. From the expression above it is clear that $\gamma=0$ or $\gamma=-1$ should be treated separately. Furthermore the hypergeometric series for ${ }_{2} F_{1}(a, b ; c ; z)$ is indeterminate for $c=-n$ where $n=0,1,2, \ldots$ if neither $a$ nor $b$ is equal to $-m$ (where $m<n$ and $m$ is a natural number). Here this means that all cases $\gamma+2=-n$ should be treated separately which combined with the condition $\gamma \neq 0,-1$ implies that whenever $\gamma=-n$ (i.e. $\alpha=-n+1$ ) the expression above will not apply. The case $\alpha=1, \gamma=0$, is excluded, so the special cases are $\gamma=-n-1, \alpha=-n$, and then the expression for the integrals becomes:

$$
\begin{equation*}
\int \frac{\mathrm{d} z}{z^{n+1}(1-z)}=\sum_{k=1}^{n} \frac{-1}{(n+1-k) z^{n+1-k}}-\ln \frac{1-z}{z} \tag{A.45}
\end{equation*}
$$

where the solution can be found from formula 2.117(4) in Gradshteyn \& Ryzhik (2007). In the integrals above the integration constants were left out.

For restricted primary mass ranges the results above can be used to evaluate the integrals listed in Eq. (A.43). The normalisation constant $k$ can be obtained from numerical integration of $f_{q}(q)$. Again, if $x_{1}=x_{2}$ one obtains $f_{q}(q)=1 /\left(1-c / x_{1}\right)$ for $q$ in the interval $\left[c / x_{1}, 1\right]$. Figures A. 5 and A. 6 show examples of the resulting mass ratio distributions for $h_{q}(q)=1$ and various single power-law mass distributions for the primaries. The hypergeometric function was calculated using the routine from Press et al. (2007, Sect. 6.13).

Finally, we have listed the expressions for the binary fraction as a function of primary mass in Table A.3.

## A.4. Split-core pairing

The mechanism of split-core pairing (SCP) works on the assumption that binaries are formed by the splitting of star-forming cores into two components. The component masses are specified through their mass ratio. The distribution function for the core masses $M_{\mathrm{c}}$ is given by the core mass function $f_{\mathrm{c}}\left(M_{\mathrm{c}}\right)$ and the mass ratio distribution $f_{q}(q)$ is specified independently. The masses of the primary and secondary are then given by:
$M_{1}=\frac{1}{1+q} M_{\mathrm{c}} \quad$ and $\quad M_{2}=\frac{1}{1+q^{-1}} M_{\mathrm{c}}$,
where for simplicity the star formation efficiency is assumed to be $100 \%$ once the core mass is set. For a constant efficiency as a function of core mass this assumption has no influence on the results. Furthermore it is assumed that the minimum core mass is large enough to ensure that the primary is always of "stellar" mass. That is the minimum core mass has to be at least twice the minimum stellar mass.

The joint probability density $f_{\text {scp }}$ for $M_{\mathrm{c}}$ and $q$ is written as:
$f_{\mathrm{scp}}\left(M_{\mathrm{c}}, q\right)=f_{\mathrm{c}}\left(M_{\mathrm{c}}\right) h_{q}(q)$,
where again the generating mass ratio distribution is written as $h_{q}(q)$ in order to distinguish it from the observed mass ratio distribution $f_{q}(q)$. The latter can be obtained by integrating $f_{\text {scp }}\left(M_{\mathrm{c}}, q\right)$ over $M_{\mathrm{c}}$ :
$f_{q}(q)=\int f_{\mathrm{scp}}\left(M_{\mathrm{c}}, q\right) \mathrm{d} M_{\mathrm{c}}$.
As for PCP there are three cases: The generating mass ratio distribution is assumed to be specified for the interval $0<q \leq 1$
and we assume $2 c \leq M_{\mathrm{c}} \leq 2 d$ (with $c$ and $2 d$ being the minimum and maximum stellar mass, respectively), which leads to the following three possibilities:

SCP-I All values of $q$ are allowed, so that for a given core mass $M_{\mathrm{c}}, M_{1}=M_{\mathrm{c}} /(1+q)$ and $0<M_{2} \leq M_{1}$. Thus binary systems with "sub-stellar" companions are allowed.
SCP-II All values of $q$ are allowed, but only binaries for which the secondary is "stellar" (i.e. $M_{2} \geq c$ ) are retained. This amounts to integrating $f_{\text {scp }}$ over the range $c(1+1 / q) \leq M_{\mathrm{c}} \leq$ $2 d$ and then re-normalising the resulting distribution of $q$ to 1.

SCP-III Only values of $q$ for which $M_{2} \geq c$ are allowed, i.e. $c /\left(M_{\mathrm{c}}-c\right) \leq q \leq 1$. This is equivalent to re-normalising the generating distribution $h_{q}(q)$ to the interval $\left[c /\left(M_{\mathrm{c}}-c\right), 1\right]$.

If we would have chosen our core mass distribution in the interval $c \leq M_{\mathrm{c}} \leq 2 d$ rather than $2 c \leq M_{\mathrm{c}} \leq 2 d$, we would have encountered a further complication. For pairing function SCP-II the resulting single star mass is then occasionally smaller than $c$, and for SCP-III the splitting up is not possible if the binary system mass is smaller than $2 c$ (see Sect. 6 for details). For simplicity we avoid this issue in our analysis below, and simply draw masses from the core mass distribution in the range $2 c \leq M_{\mathrm{c}} \leq 2 d$.

## A.4.1. Mass ratio distributions for SCP

We give the general expressions for $f_{q}(q)$ before discussing specific examples. For ease of notation we use $x=M_{c}$ and $f(x)=f_{\mathrm{c}}\left(M_{\mathrm{c}}\right)$.

SCP-I If no selection on primary mass is made one always obtains $f_{q}(q)=h_{q}(q)$ because the distributions of $q$ and $M_{\mathrm{c}}$ are independent. Unlike PCP-I, restricting the primary mass range now does cause changes of the observed mass ratio distribution with respect to the generating one. Several cases have to be distinguished on the basis of the value of the primary mass selection limits $x_{1}$ and $x_{2}$ compared to the values of $c, d, 2 c$, and $2 d$. Note that $x_{1}$ and $x_{2}$ are limits on $M_{1}$, not $M_{\mathrm{c}}$.

In all cases discussed below a selection $x_{1} \leq M_{1} \leq x_{2}$ in principle translates to $x_{1}(1+q) \leq x \leq x_{2}(1+q)$ but the upper and/or lower limits on $x$ and $q$ used for integrating $f_{\text {scp }}$ and normalising $f_{q}(q)$ are different for each case. We introduce the following variables to distinguish the different cases:

$$
\begin{array}{ll}
q_{0}=2 c / x_{2}-1 & q_{2}=2 d / x_{2}-1 \\
q_{1}=2 c / x_{1}-1 & q_{3}=2 d / x_{1}-1
\end{array}
$$

The value of $x$ (core-mass) is restricted to $[2 c, 2 d]$, hence for $q<q_{0}$ or $q>q_{3}$ the probability distribution for $q$ vanishes (as these conditions imply $M_{1}>x_{2}$ or $M_{1}<x_{1}$, both of which are not allowed). For $q<q_{1}$ the lower integration limit for $x$ when determining $f_{q}$ is fixed at $2 c$ and for $q>q_{2}$ the upper limit is fixed at $2 d$.

Case 1: $x_{1} \leq c<x_{2} \leq 2 c$ In this case the value of $q_{0}$ is between 0 and 1 and all other values of $q_{i}$ are larger than 1 . This means that two cases should be distinguished for $f_{q}, q<q_{0}$ and $q>q_{0}$, and that the lower integration limit for $x$ is always fixed at $2 c$. The expression for $f_{q}$ is then given by:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<q_{0} \\ k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & q_{0} \leq q \leq 1\end{array}\right.$,


Fig. A.5. Mass ratio distributions $f_{q}(q)$ for PCP-II and III for a uniform generating mass ratio distribution and primary masses from the single power-law mass distribution for $\alpha=2.35,0.30,-2.50$. The curves are shown for three ratios of the lower to the upper mass limit of the mass distribution: $c / d=0.1,0.3,0.5$.
where $k$ is a normalisation constant which follows from:
$\int_{q_{0}}^{1} f_{q}(q) \mathrm{d} q=1$.

Case 2: $x_{1} \leq c \wedge 2 c<x_{2} \leq d \quad$ This is the same as case 1 except that now $q_{0}<0$ so that there is only one part to the expression for $f_{q}$. The lower integration limit for $x$ is still $2 c$ :
$f_{q}(q)=k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x \quad 0<q \leq 1$,


Fig. A.6. Mass ratio distributions $f_{q}(q)$ for PCP-II and III for a uniform generating mass ratio distribution and primary masses from a single powerlaw mass distribution. The solid curve shows the complete mass ratio distribution (for all binaries in the population) for $c=0.2$ and $d=2.0$. The other curves show what happens to the observed $f_{q}(q)$ if the primary mass $M_{1}$ is restricted to $x_{1} \leq M_{1} \leq x_{2}$.
where the normalisation constant $k$ is obtained by integrating $f_{q}$ over $[0,1]$.

Case 3: $x_{1} \leq c \wedge d<x_{2} \leq 2 d$ Now the value of $q_{2}$ is in the interval $[0,1]$ which means that for the case $q>q_{2}$ the upper
integration limit for $x$ is fixed at $2 d$ which leads to a plateau in the probability density $f_{q}$ :
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q \leq q_{2} \\ k \int_{2 c}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2}<q \leq 1\end{array}\right.$,
where the normalisation constant $k$ is obtained through integration over the intervals $\left[0, q_{2}\right]$ and $\left[q_{2}, 1\right]$.

Case 4: $c<x_{1}<x_{2} \leq 2 c \quad$ Now the values of $q_{0}$ and $q_{1}$ are in the interval $[0,1]$ which means that the lower integration limit for $x$ depends on $q$. The probability density for $q$ now consists of three parts:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<q_{0} \\ k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & q_{0} \leq q<q_{1} \\ k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & q_{1} \leq q \leq 1\end{array}\right.$,
where $k$ now has to be obtained from condition (A.50) by integrating over the intervals $\left[q_{0}, q_{1}\right]$ and $\left[q_{1}, 1\right]$.

Case 5: $c<x_{1} \leq 2 c \wedge 2 c<x_{2} \leq d$ The value of $q_{0}$ becomes less than zero and we have:
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q<q_{1} \\ k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & q_{1} \leq q \leq 1\end{array}\right.$,
where $k$ is now has to be obtained from condition (A.50) by integrating over the intervals $\left[0, q_{1}\right]$ and $\left[q_{1}, 1\right]$.

Case 6: $c<x_{1} \leq 2 c \wedge d<x_{2} \leq 2 d$ Now the values of $q_{1}$ and $q_{2}$ are in the interval $[0,1]$. For $q<q_{1}$ the lower integration limit is fixed to $2 c$ and for $q>q_{2}$ the upper integration limit is fixed to $2 d$. In addition it can happen that $q_{1}<q_{2}$ or $q_{1} \geq q_{2}$. The latter case will cause a plateau of constant probability density $f_{q}$ for $q_{1} \leq q \leq q_{2}$. So now there are two "sub-cases".

Case 6a: $q_{1}<q_{2}$
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{2 c}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q<q_{1} \\ k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & q_{1} \leq q<q_{2}, \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q \leq 1\end{array}\right.$,
where the normalisation constant $k$ is obtained through integration over three intervals $\left[0, q_{1}\right],\left[q_{1}, q_{2}\right]$, and $\left[q_{2}, 1\right]$.

Case 6b: $q_{1} \geq q_{2}$
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{2 c}^{x(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q<q_{2} \\ k \int_{2 c}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q<q_{1}, \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q \leq 1\end{array}\right.$,
where the normalisation constant $k$ is obtained through integration over three intervals $\left[0, q_{2}\right],\left[q_{2}, q_{1}\right]$, and $\left[q_{1}, 1\right]$.

Case 7: $2 c<x_{1}<x_{2} \leq d \quad$ Now the values of $q_{0}$ and $q_{1}$ are less than zero and both $q_{2}$ and $q_{3}$ are larger than 1 . Hence the probability density consists of one part only:
$f_{q}(q)=k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x \quad 0<q \leq 1$,
where $k$ is obtained by integrating $f_{q}$ over $[0,1]$.
Case 8: $2 c<x_{1} \leq d \wedge d<x_{2} \leq 2 d$ The value of $q_{2}$ is now in $[0,1]$ so $f_{q}$ will consist of two parts:
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q<q_{2} \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[0, q_{2}\right]$ and $\left[q_{2}, 1\right]$.

Case 9: $d<x_{1}<x_{2} \leq d$ Now the values of both $q_{2}$ and $q_{3}$ are in $[0,1]$ and the probability density will vanish if $q>q_{3}$ :
$f_{q}(q)=\left\{\begin{array}{ll}k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & 0<q<q_{2} \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q<q_{3} \\ 0 & q_{3} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[0, q_{2}\right]$ and $\left[q_{2}, q_{3}\right]$.
SCP-II The expression for $f_{q}$ is obtained by integrating the joint distribution $f_{\text {scp }}$ over $x$ over the range $c(1+1 / q) \leq x \leq 2 d$ and normalising the resulting expression for $f_{q}$ to 1 :
$f_{q}(q)=k \int_{c(1+1 / q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x$,
where $k$ is obtained from the condition:
$\int_{c /(2 d-c)}^{1} f_{q}(q) \mathrm{d} q=1$,
with $c /(2 d-c)$ being the minimum possible value for $q$.
When the mass range is restricted there are again a number of cases to consider, depending on the values of $x_{1}$ and $x_{2}$. However the situation is less complicated than for SCP-I. First of all the value of $c(1+1 / q)$ is always larger than $2 c$ for $0<q \leq 1$ which means that the value of $q_{1}$ plays no role. Secondly the minimum possible value of $q$ is $c / x_{2}$ and this quantity is always larger than $q_{0}$ for $x_{2} \geq c$ (which is mandatory) and therefore also the value of $q_{0}$ plays no role. For $q<c / x_{2} f_{q}(q)$ is always zero. The values of $q_{2}$ and $q_{3}$ do matter as discussed below.

Case 1: $x_{1}<x_{2} \leq d$ In this case the integration limits for $x, x_{1}(1+q)$ and $x_{2}(1+q)$ are guaranteed to be less than $2 d$. For $q<c / x_{1}$ the value of $x_{1}(1+q)$ is less than $c(1+1 / q)$ so the lower limit for integral over $x$ is then fixed at $c(1+1 / q)$. The expression for $f_{q}(q)$ becomes:
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c} x_{2(1+1 / q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained from the condition:
$\int_{c / x_{2}}^{1} f_{q}(q) \mathrm{d} q=1$.
Case 2: $x_{1} \leq d \wedge x_{2}>d$ Now the value of $q_{2}$ is less than 1 and for $q>q_{2}$ the upper limit of the integral over $x$ is fixed at $2 d$ (the value of $x_{2}(1+q)$ being larger than $2 d$ ). Now, $q_{2} \geq c / x_{2}$ but $q_{2}$ may be larger or smaller than $c / x_{1}$. So there are two sub-cases:

Case 2a $q_{2} \leq c / x_{1}$
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c(1+1 / q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{2}} \leq q<q_{2} \\ k \int_{c(1+1 / q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{1}} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[c / x_{2}, q_{2}\right],\left[q_{2}, c / x_{1}\right]$, and $\left[c / x_{1}, 1\right]$, and applying condition (A.63).

Case 2b $q_{2}>\frac{c}{x_{1}}$
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c(1+1 / q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{1}} \leq q<q_{2} \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[c / x_{2}, c / x_{1}\right],\left[c / x_{1}, q_{2}\right]$, and $\left[q_{2}, 1\right]$, and applying condition (A.63).

Case 3: $x_{1}>d$ This is the same as case 2 except that $q_{3}<1$ which means that $f_{q}(q)=0$ for $q>q_{3}$ (this is due to the value of $x_{1}(1+q)$ becoming larger than $2 d$ ). There are the same two sub-cases:

Case 3a: $q_{2} \leq c / x_{1}$
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c(1+1 / q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{2}} \leq q<q_{2} \\ k \int_{c(1+1 / q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{1}} \leq q<q_{3} \\ 0 & q_{3} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[c / x_{2}, q_{2}\right],\left[q_{2}, c / x_{1}\right]$, and $\left[c / x_{1}, q_{3}\right]$, and applying condition (A.63).

Case 3b $q_{2}>c / x_{1}$
$f_{q}(q)=\left\{\begin{array}{ll}0 & 0<q<\frac{c}{x_{2}} \\ k \int_{c(1+1 / q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k \int_{x_{2}(1+q)}^{x_{2}(1+q)} h_{q}(q) f(x) \mathrm{d} x & \frac{c}{x_{1}} \leq q<q_{2}, \\ k \int_{x_{1}(1+q)}^{2 d} h_{q}(q) f(x) \mathrm{d} x & q_{2} \leq q<q_{3} \\ 0 & q_{3} \leq q \leq 1\end{array}\right.$,
where $k$ is obtained by integrating $f_{q}$ over $\left[c / x_{2}, c / x_{1}\right],\left[c / x_{1}, q_{2}\right]$, and $\left[q_{2}, q_{3}\right]$, and applying condition (A.63).

SCP-III Now the generating mass ratio distribution $h_{q}$ is restricted to the range $c /\left(M_{\mathrm{c}}-c\right) \leq q \leq 1$ and re-normalised. The corresponding distribution is $h^{\prime}(q)$ and the expression for $f_{q}$ becomes:
$f_{q}(q)=\int_{c(1+1 / q)}^{2 d} h_{q}^{\prime}(q) f(x) \mathrm{d} x$,
where the lower integration limit is set by the condition $M_{2} \geq c$ and the distribution is normalised.

When the primary mass range is restricted to $x_{1} \leq x \leq x_{2}$ the cases and expressions for $f_{q}(q)$ are the same as for the SCP-II case, except that $h_{q}^{\prime}(q)$ replaces $h_{q}(q)$ everywhere.

## A.4.2. Binary fractions for SCP

In order to find the specific binary fraction as a function of primary mass, we first need to find the primary mass distribution $f_{M_{1}}\left(M_{1}\right)$, which can be calculated as:
$f_{M_{1}}\left(M_{1}\right)=\int_{q_{\min }\left(M_{1}\right)}^{q_{\max }\left(M_{1}\right)} f_{M_{1}, q}\left(M_{1}, q\right) \mathrm{d} q$,
where $q_{\min }\left(M_{1}\right)$ and $q_{\max }\left(M_{1}\right)$ are the minimum and maximum mass ratios for a given primary mass $M_{1}$, and $f_{M_{1}, q}\left(M_{1}, q\right)$ is the joint probability density function for $M_{1}$ and $q$. The latter can be derived from the generating (core) mass distribution $f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\right)$ and the generating mass distribution $h_{q}(q)$ :
$f_{M_{1}, q}\left(M_{1}, q\right)=f_{M_{\mathrm{C}}}\left(M_{\mathrm{C}}\left(M_{1}, q\right)\right) h_{q}(q)(1+q)$,
where the factor $(1+q)$ is the Jacobian of the transformation $M_{1}=M_{\mathrm{C}}(q+1)^{-1}$ and $q=q$.

SCP-I The number of binary systems with a mass $M_{1}$ is given by $N \mathcal{B} f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$, where $N$ is the total number of systems (singles plus binaries) in the population. The number of single stars with a mass $M_{1}$ is given by $N(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right) \mathrm{d} M_{1}$. The specific binary fraction is thus:
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\frac{B_{M_{1}}}{S_{M_{1}}+B_{M_{1}}}=\frac{\mathcal{B} f_{M_{1}}\left(M_{1}\right)}{\mathcal{B} f_{M_{1}}\left(M_{1}\right)+(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right)}$.
The integration limits for $f_{M_{1}}\left(M_{1}\right)$ are indicated with domain $S$ in Fig. A.8, and are given by:
$q_{\text {min }}=\left\{\begin{array}{ll}2-M_{1} / c & c \leq M_{1}<2 c \\ 0 & 2 c \leq M_{1} \leq 2 d\end{array}\right.$,
and
$q_{\max }=\left\{\begin{array}{lc}1 & c \leq M_{1}<d \\ 2-M_{1} / d & d \leq M_{1} \leq 2 d\end{array}\right.$.
An example for the specific binary fraction resulting from SCP-I is shown in Fig. 7. The specific binary fraction for $c \leq M_{1} \leq 2 c$ equals unity, as $f_{M_{\mathrm{C}}}\left(M_{1}\right)=0$ in this mass range. In the mass range $2 c<M_{1} \leq d$ the specific binary fraction is practically independent of $M_{1}$. Beyond $M_{1}=d$, the specific binary fraction rapidly drops to zero at $M_{1}=2 d$. The overall binary fraction $\mathcal{B}_{\text {all }}$ is always equal to $\mathcal{B}$ for SCP-I.

SCP-II The integration domain for SCP-II to obtain $f_{M_{1}}\left(M_{1}\right)$ is indicated with region $S^{\prime}$ in Fig. A.8, i.e., for $c \leq M_{1} \leq 2 d-c$. The integration limits $q_{\text {min }}$ and $q_{\text {max }}$ are given by:
$q_{\text {min }}=c / M_{1} \quad c \leq M_{1}<2 d-c$,
and
$q_{\text {max }}=\left\{\begin{array}{lc}1 & c \leq M_{1}<d \\ 2-M_{1} / d & d \leq M_{1} \leq 2 d-c\end{array}\right.$.
The specific binary fraction for a population with a generating binary fraction of unity, i.e., $\mathcal{B}=100 \%$ ), is given by:
$\mathcal{B}_{100}\left(M_{1}\right)=\frac{\int_{q_{\text {mina } 2}\left(M_{1}\right)}^{q_{\text {max }}}\left(M_{1}\right)}{q_{M_{1}, q}\left(M_{1}, q\right) \mathrm{d} q}$,
where the joint probability distribution is given by Eq. (A.70). For a given primary mass $M_{1}$, the value of $\mathcal{B}_{100}\left(M_{1}\right)$ given by the ratio between the (weighted) lengths of the horizontal line segments of domain $S^{\prime}$ and $S$ in Fig. A.8. The integration limits $q_{\min , 2}$ and $q_{\text {max, } 2}$ are thus given by Eqs. (A.74) and (A.75), and the limits $q_{\text {min, } 1}$ and $q_{\text {max, } 1}$ are given by Eqs. (A.72) and (A.73), respectively. For systems with an arbitrary value of $\mathcal{B}$, the number of binary systems is given by
$B_{M_{1}}=N \mathcal{B B}_{100}\left(M_{1}\right) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1}$,

Table A.1. The specific mass ratio distribution $f_{q}(q)$ for pairing function SCP-I, resulting from a power-law generating mass distribution $f_{m}(m)$ and a uniform generating mass ratio distribution $h_{q}(q)$. See Appendix A.4.3 for details.

| SCP-I | Mass ratio distribution |
| :--- | :---: |
| Case 1 |  |

$$
f_{q}(q)=\left\{\begin{array}{ll}
0 & 0<q<q_{0} \\
k\left[1-\left(2 c / x_{2}\right)^{\gamma}(1+q)^{-\gamma}\right] & q_{0} \leq q \leq 1
\end{array} \quad k=\frac{1-\gamma}{\left(1-q_{0}\right)(1-\gamma)-2^{1-\gamma}\left(2 c / x_{2}\right)^{\gamma}+2 c / x_{2}}\right.
$$

Case 2

$$
f_{q}(q)=k\left[1-\left(2 c / x_{2}\right)^{\gamma}(1+q)^{-\gamma}\right] \quad 0<q \leq 1 \quad k=\frac{1-\gamma}{(1-\gamma)-\left(2 c / x_{2}\right)^{\gamma}\left(2^{1-\gamma}-1\right)}
$$

Case 3

$$
f_{q}(q)=\left\{\begin{array}{ll}
k \frac{1-\left(2 c / x_{2}\right)^{\gamma}(1+q)^{-\gamma}}{1-(c / d)^{\gamma}} & 0<q \leq q_{2} \\
k & q_{2}<q \leq 1
\end{array} \quad k^{-1}=\frac{1}{1-(c / d)^{\gamma}}\left[q_{2}-\frac{\left(2 c / x_{2}\right)^{\gamma}\left(\left(1+q_{2}\right)^{1-\gamma}-1\right)}{1-\gamma}\right]+1-q_{2}\right.
$$

## Case 4

$$
f_{q}(q)= \begin{cases}0 & 0<q<q_{0} \\ k\left[1-\left(\frac{2 c}{x_{2}}\right)^{\gamma}(1+q)^{-\gamma}\right] & q_{0} \leq q<q_{1} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma} & q_{1} \leq q \leq 1\end{cases}
$$

$$
k=\frac{1-\gamma}{2^{1-\gamma}\left[\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}\right]-\gamma\left(2 c / x_{1}-2 c / x_{2}\right)}
$$

Case 5

$$
f_{q}(q)=\left\{\begin{array}{ll}
k\left[1-\left(\frac{2 c}{x_{2}}\right)^{\gamma}(1+q)^{-\gamma}\right] & 0<q<q_{1} \\
k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma} & q_{1} \leq q \leq 1
\end{array} \quad k=\frac{1-\gamma}{q_{1}(1-\gamma)+\left[\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}\right] 2^{1-\gamma}+\left(2 c / x_{2}\right)^{\gamma}-2 c / x_{1}}\right.
$$

Case 6a

$$
f_{q}(q)= \begin{cases}k\left[1-\left(\frac{2 c}{x_{2}}\right)^{\gamma}(1+q)^{-\gamma}\right] & 0<q<q_{1} \\ k\left[\left(\frac{2 c}{x_{1} \gamma^{\gamma}}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma}\right. & q_{1} \leq q<q_{2} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}(1+q)^{1-\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k= & \frac{1-\gamma}{q_{1}(1-\gamma)-\left(2 c / x_{2}\right)^{\gamma}\left[\left(1+q_{1}\right)^{1-\gamma}-1\right]} \\
& +\frac{1-\gamma}{\left[\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}\right]\left[\left(1+q_{2}\right)^{1-\gamma}-\left(1+q_{1}\right)^{1-\gamma}\right]} \\
& +\frac{1-\gamma}{\left(2 c / x_{1}\right)^{\gamma}\left(2^{1-\gamma}-\left(1+q_{2}\right)^{1-\gamma}\right)-(c / d)^{\gamma}\left(1-q_{2}\right)(1-\gamma)}
\end{aligned}
$$

Case 6b

$$
f_{q}(q)= \begin{cases}k\left[1-\left(\frac{2 c}{x_{2}}\right)^{\gamma}(1+q)^{-\gamma}\right] & 0<q<q_{2} \\ k\left[1-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q<q_{1} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}(1+q)^{1-\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{1} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k= & \frac{1-\gamma}{q_{2}(1-\gamma)-\left(2 c / x_{2}\right)^{\gamma}\left[\left(1+q_{2}\right)^{1-\gamma}-1\right]} \\
& +\frac{1-\gamma}{\left(1-(c / d)^{\gamma}\right)\left(q_{1}-q_{2}\right)(1-\gamma)} \\
& +\frac{1-\gamma}{\left(2 c / x_{1}\right)^{\gamma}\left(2^{1-\gamma}-\left(1+q_{1}\right)^{1-\gamma}\right)-(c / d)^{\gamma}\left(1-q_{1}\right)(1-\gamma)}
\end{aligned}
$$

## Case 7

$$
f_{q}(q)=k(1+q)^{-\gamma}
$$

$$
k=\frac{1-\gamma}{2^{1-\gamma}-1}
$$

$$
f_{q}(q)= \begin{cases}k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma} & 0<q<q_{2} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}(1+q)^{1-\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k= & \frac{1-\gamma}{\left[\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}\right]\left[\left(1+q_{2}\right)^{1-\gamma}-1\right]} \\
& +\frac{1-\gamma}{\left(2 c / x_{1}\right)^{\gamma}\left(2^{1-\gamma}-\left(1+q_{2}\right)^{1-\gamma}\right)} \\
& -\frac{1-\gamma}{(c / d)^{\gamma}\left(1-q_{2}\right)(1-\gamma)}
\end{aligned}
$$

Case 9

$$
\left.f_{q}(q)=\left\{\begin{array}{ll}
k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma} & 0<q<q_{2} \\
k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}(1+q)^{1-\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q<q_{3} \\
0 & q_{3} \leq q \leq 1
\end{array} \quad k=\frac{1-\gamma}{\left[\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}\right]\left[\left(1+q_{2}\right)^{1-\gamma}-1\right]}\right]+\frac{1-\gamma}{\left(2 c / x_{1}\right)^{\gamma}\left(\left(1+q_{3}\right)^{1-\gamma}-\left(1+q_{2}\right)^{1-\gamma}\right)}\right)
$$



Fig. A.7. Left: mass ratio distributions $f_{q}(q)$ for split-core pairing I for a uniform generating mass ratio distribution and core masses from single power-law mass distribution, where $2 c \leq M_{\mathrm{c}} \leq 2 d$. The values of $\alpha$ are 2.35, 0.30 , and -2.50 . The black solid curve shows the complete mass ratio distribution (for all binaries in the population) for $c=0.2$ and $d=2.0$. The other curves show what happens to the observed $f_{q}(q)$ if the primary mass $M_{1}$ is restricted to $x_{1} \leq M_{1} \leq x_{2}$. The ten cases for SCP-I from section A.4.1 are listed in the legend. Right: same, but with the vertical scale changed to bring out some of the details in the cases close to the line $f_{q}(q)=1$.

Table A.2. The specific mass ratio distribution $f_{q}(q)$ for pairing function SCP-II, resulting from a power-law generating mass distribution $f_{\mathrm{m}}(m)$ and a uniform generating mass ratio distribution $h_{q}(q)$. See Appendix A.4.3 for details.

| SCP-II | Mass ratio distribution | Normalization constant |
| :--- | :---: | :---: |
| Case 1 |  |  |

$$
f_{q}(q)= \begin{cases}0 & 0<q<\frac{c}{x_{2}} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{-\gamma} & \frac{c}{x_{1}} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k^{-1}= & \int_{c / x_{2}}^{c / x_{1}}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q-\frac{\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}-\left(1+\frac{c}{x_{2}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[2^{1-\gamma}-\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}\right]
\end{aligned}
$$

Case 2a

$$
f_{q}(q)= \begin{cases}0 & 0<q<\frac{c}{x_{2}} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{2}} \leq q<q_{2} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q<\frac{c}{x_{1}} \\ k\left[\left(\frac{2 c / x_{1}}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & \frac{c}{x_{1}} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k^{-1}= & \int_{c / x_{2}}^{c / x_{1}}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q \\
& -\frac{\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{2}\right)^{1-\gamma}-\left(1+\frac{c}{x_{2}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}}{1-\gamma}\left[2^{1-\gamma}-\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}\right]-\left(\frac{c}{d}\right)^{\gamma}\left(1-q_{2}\right)
\end{aligned}
$$

## Case 2b

$$
f_{q}(q)= \begin{cases}0 & 0<q<\frac{c}{x_{2}} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k\left[\left(\frac{2 c}{x_{1}}\right)^{\gamma}-\left(\frac{2 c}{x_{2}}\right)^{\gamma}\right](1+q)^{\gamma} & \frac{c}{x_{1}} \leq q<q_{2} \\ k\left[\left(\frac{2 c / x_{1}}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k^{-1}= & \int_{c / x_{2}}^{c / x_{1}}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q \\
& -\frac{\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}-\left(1+\frac{c}{x_{2}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{2}\right)^{1-\gamma}-\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}}{1-\gamma}\left[2^{1-\gamma}-\left(1+q_{2}\right)^{1-\gamma}\right]-\left(\frac{c}{d}\right)^{\gamma}\left(1-q_{2}\right)
\end{aligned}
$$

## Case 3a

$$
f_{q}(q)= \begin{cases}0 & 0<q<\frac{c}{x_{2}} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{2}} \leq q<q_{2} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q<\frac{c}{x_{1}} \\ k\left[\left(\frac{2 c / x_{1}}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & \frac{c}{x_{1}} \leq q<q_{3} \\ 0 & q_{3} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k^{-1}= & \int_{c / x_{2}}^{c / x_{1}}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q \\
& -\frac{\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{2}\right)^{1-\gamma}-\left(1+\frac{c}{x_{2}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{3}\right)^{1-\gamma}-\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}\right]-\left(\frac{c}{d}\right)^{\gamma}\left(q_{3}-q_{2}\right)
\end{aligned}
$$

Case 3b

$$
f_{q}(q)= \begin{cases}0 & 0<q<\frac{c}{x_{2}} \\ k\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{2}} \leq q<\frac{c}{x_{1}} \\ k\left[\left(\frac{2 c x_{1}}{1+q}\right)^{\gamma}-\left(\frac{2 c / x_{2}}{1+q}\right)^{\gamma}\right] & \frac{c}{x_{1}} \leq q<q_{2} \\ k\left[\left(\frac{2 c / x_{1}}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right] & q_{2} \leq q<q_{3} \\ 0 & q_{3} \leq q \leq 1\end{cases}
$$

$$
\begin{aligned}
k^{-1}= & \int_{c / x_{2}}^{c / x_{1}}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q \\
& -\frac{\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}-\left(1+\frac{c}{x_{2}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}-\left(2 c / x_{2}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{2}\right)^{1-\gamma}-\left(1+\frac{c}{x_{1}}\right)^{1-\gamma}\right] \\
& +\frac{\left(2 c / x_{1}\right)^{\gamma}}{1-\gamma}\left[\left(1+q_{3}\right)^{1-\gamma}-\left(1+q_{2}\right)^{1-\gamma}\right]-\left(\frac{c}{d}\right)^{\gamma}\left(q_{3}-q_{2}\right)
\end{aligned}
$$

Table A.3. The specific binary fraction $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ for the nine pairing functions described in this paper, resulting from a power-law mass distribution $f_{M}(M)$ and a uniform mass ratio distribution $h_{q}(q)$. See Appendices A.1-A. 4 for a detailed description, and Fig. 7 for a visualisation.

| Pairing function $\quad$ Specific binary fraction |  |
| :--- | :--- |
| RP |  |
| $\mathcal{B}_{M_{1}}\left(M_{1}\right)=\left(\frac{\mathcal{B}^{-1}-1}{2 F_{M}\left(M_{1}\right)}+1\right)^{-1} \quad c \leq M_{1} \leq d$ | $F_{M}\left(M_{1}\right)=\frac{M_{1}^{1-\alpha}-c^{1-\alpha}}{d^{1-\alpha}-c^{1-\alpha}}$ |

PCRP

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} \quad c \leq M_{1} \leq d
$$

PCP-I

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} \quad c \leq M_{1} \leq d
$$

PCP-II

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B}\left(1-\frac{c}{M_{1}}\right) \quad c \leq M_{1} \leq d
$$

## PCP-III

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)=\mathcal{B} \quad c \leq M_{1} \leq d
$$

SCP-I

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)= \begin{cases}1 & c \leq M_{1} \leq 2 c \\ \left(1+\frac{\alpha+2}{\mathcal{B} \cdot R\left(M_{1}\right)}\right)^{-1} & 2 c \leq M_{1} \leq d \\ \left(1+\frac{\alpha+2}{\mathcal{B} \cdot S\left(M_{1}\right)}\right)^{-1} & d \leq M_{1} \leq 2 d\end{cases}
$$

$$
R\left(M_{1}\right)=2^{\alpha+2}-\left(3-\frac{M_{1}}{c}\right)^{\alpha+2}
$$

$$
S\left(M_{1}\right)=\left(3-\frac{M_{1}}{d}\right)^{\alpha+2}-1
$$

## SCP-II

SCP-III

$$
\mathcal{B}_{M_{1}}\left(M_{1}\right)= \begin{cases}1 & c \leq M_{1} \leq 2 c \\ \left(1+\frac{\alpha+2}{\mathcal{B} \cdot R\left(M_{1}\right)}\right)^{-1} & 2 c \leq M_{1} \leq d \\ \left(1+\frac{\alpha+2}{\mathcal{B} \cdot S\left(M_{1}\right)}\right)^{-1} & d \leq M_{1} \leq 2 d-c \\ 0 & 2 d-c \leq M_{1} \leq 2 d\end{cases}
$$

$R\left(M_{1}\right)=\left(1-\frac{c}{M_{1}}\right)^{-1}\left[2^{\alpha+2}-\left(1+\frac{c}{M_{1}}\right)^{\alpha+2}\right]$
$S\left(M_{1}\right)=\left(1-\frac{c}{M_{1}}\right)^{-1}\left[\left(3-\frac{M_{1}}{d}\right)^{\alpha+2}-\left(1+\frac{c}{M_{1}}\right)^{\alpha+2}\right]$
and the number of single stars by

$$
\begin{align*}
S_{M_{1}}= & N(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right) \mathrm{d} M_{1} \\
& +N \mathcal{B}\left(1-\mathcal{B}_{100}\left(M_{1}\right)\right) f_{M_{1}}\left(M_{1}\right) \mathrm{d} M_{1} \tag{A.78}
\end{align*}
$$

where the first term refers to the fraction of clumps that were assigned to be single, and the second term refers to the primaries
that have become single due to the rejection of their low-mass companions. The specific binary fraction is then given by:
$\mathcal{B}_{M_{1}}\left(M_{1}\right)=\frac{B_{M_{1}}}{S_{M_{1}}+B_{M_{1}}}=\frac{\mathcal{B}_{100}\left(M_{1}\right) \mathcal{B} f_{M_{1}}\left(M_{1}\right)}{\mathcal{B} f_{M_{1}}\left(M_{1}\right)+(1-\mathcal{B}) f_{M_{\mathrm{C}}}\left(M_{1}\right)}$,

$$
\begin{aligned}
& \mathcal{B}_{M_{1}}\left(M_{1}\right)= \begin{cases}\mathcal{B}_{100, R}\left(M_{1}\right) & c \leq M_{1} \leq 2 c \\
\frac{\mathcal{B}_{100, S}\left(M_{1}\right)(\alpha+2)}{1+S\left(M_{1}\right)\left(\mathcal{B}^{-1}-1\right)} & 2 c \leq M_{1} \leq d \\
\frac{\mathcal{B}_{100, T}\left(M_{1}\right)(\alpha+2)}{1+T\left(M_{1}\right)\left(\mathcal{B}^{-1}-1\right)} & d \leq M_{1} \leq 2 d-c \\
0 & 2 d-c \leq M_{1} \leq 2 d\end{cases} \\
& \mathcal{B}_{100, R}\left(M_{1}\right)=\frac{2^{\alpha+2}-\left(1+c / M_{1}\right)^{\alpha+2}}{2^{\alpha+2}-\left(3-M_{1} / c\right)^{\alpha+2}} \\
& \mathcal{B}_{100, S}\left(M_{1}\right)=\frac{2^{\alpha+2}-\left(1+c / M_{1}\right)^{\alpha+2}}{2^{\alpha+2}-1} \\
& \mathcal{B}_{100, T}\left(M_{1}\right)=\frac{\left(3-M_{1} / d\right)^{\alpha+2}-\left(1+c / M_{1}\right)^{\alpha+2}}{\left(3-M_{1} / c\right)^{\alpha+2}-1} \\
& S\left(M_{1}\right)=2^{\alpha+2}-\left(1+\frac{c}{M_{1}}\right)^{\alpha+2} \\
& T\left(M_{1}\right)=\left(3-\frac{M_{1}}{d}\right)^{\alpha+2}-\left(1+\frac{c}{M_{1}}\right)^{\alpha+2}
\end{aligned}
$$



Fig. A.8. The integration domains for obtaining the primary mass distribution and specific binary fraction for pairing functions SCP-I, SCP-II and SCP-III. The domain $S$ for SCP-I is the parallelogram enclosed by the solid lines. The domain $S^{\prime}$ for SCP-II and SCP-III is indicated with the shaded region.

An example of the specific binary fraction resulting from SCP-II is shown in Fig. 7. For $c \leq M_{1} \leq 2 c$, the above equation reduces to $\mathcal{B}\left(M_{1}\right)=\mathcal{B}_{100}\left(M_{1}\right)$ as $f_{M_{\mathrm{C}}}\left(M_{1}\right)=0$. The binary fraction among the lowest mass stars ( $M_{1}=c$ ) equals unity, after which it drops until $M_{1}=2 c$. Beyond that minimum mass it rises again and reaches its maximum at $M_{1}=d$, after which it drops again go zero at $M_{1}=2 d$, for the same reasons as for the SCP-I case. The overall binary fraction can be calculated by integrating $\mathcal{B}\left(M_{1}\right)$ over primary mass, weighed by the primary mass distribution $f_{M_{1}}\left(M_{1}\right)$, see Eq. (A.33).

SCP-III The values for $q_{\min }$ and $q_{\max }$ for SCP-III are identical to those for SCP-II, and are given by Eqs. (A.74) and (A.75). The expression for the specific binary fraction for SCP-III is the same as in (the one for SCP-I). Note, however, that the primary mass distribution $f_{M_{1}}\left(M_{1}\right)$ is different due to the different limits $q_{\text {min }}$ and $q_{\text {max }}$. The right-hand panel in Fig. 7 shows an example of $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ resulting from SCP-III. The term $f_{M_{\mathrm{C}}}\left(M_{1}\right)$ vanishes for $c \leq M_{1} \leq 2 c$, so that $\mathcal{B}_{M_{1}}\left(M_{1}\right)=1$ in this mass range. A discontinuity appears at $M_{1}=2 c$, beyond which $\mathcal{B}_{M_{1}}\left(M_{1}\right)$ reaches its lowest point, after which it steadily rises until $M_{1}=$ $d$. Beyond $M_{1}=d$, the specific binary fraction decreases again to $\mathcal{B}_{M_{1}}\left(M_{1}\right)=0$ at $M_{1}=2 d$, for the same reasons as for SCP-I and SCP-II. The overall binary fraction $\mathcal{B}_{\text {all }}$ is always equal to $\mathcal{B}$ for SCP-III.

## A.4.3. Uniform mass ratio distribution and single power-law mass distribution

In the expressions given in this section the term $1-(c / d)^{\gamma}$ has often been absorbed in the normalisation constants (which are therefore not strictly consistent with the expressions above). The following is assumed for $h_{q}$ and $f_{\mathrm{c}}\left(M_{\mathrm{c}}\right)=f(x)$ (using $x=M_{\mathrm{c}}$ for ease of notation):

$$
\begin{aligned}
h_{q}(q) & =1 \quad 0<q \leq 1 \\
f(x) & =a x^{-\alpha} \quad 2 c \leq x \leq 2 d
\end{aligned}
$$

where $a=\gamma(2 c)^{\gamma} /\left(1-(c / d)^{\gamma}\right)$ and $\gamma=\alpha-1$, and $\alpha \neq 1$.

The SCP-I case Without restricting the primary mass range the expression for the mass ratio distribution is simply $f_{q}(q)=1$. When the primary mass range is restricted the cases listed above have to be worked out. The corresponding expressions for $f_{q}(q)$ are listed in Table A.1. Figure A. 7 shows an example of what $f_{q}(q)$ looks like for all the SCP-I cases discussed above.

The SCP-II case Without restrictions on the primary mass range the expression for the mass ratio distribution follows from Eq. (A.60) and is:

$$
\begin{equation*}
f_{q}(q)=\frac{k}{1-(c / d)^{\gamma}}\left[\left(\frac{2 q}{1+q}\right)^{\gamma}-\left(\frac{c}{d}\right)^{\gamma}\right], \tag{A.80}
\end{equation*}
$$

where the normalisation constant follows from:

$$
\begin{align*}
1 & =\int_{q_{\min }}^{1} f_{q}(q) \mathrm{d} q \\
& =\frac{k}{1-(c / d)^{\gamma}}\left(\int_{q_{\min }}^{1}\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q-\left(\frac{c}{d}\right)^{\gamma}\left(1-q_{\min }\right)\right), \tag{A.81}
\end{align*}
$$

where $q_{\text {min }}=(c / d) /(2-c / d)$. The integral in this expression will be dealt with below. With a restricted primary mass range the expressions for $f_{q}$ are obtained for the three cases discussed above. These expressions are listed in Table A.2. Again the term $1-(c / d)^{\gamma}$ is absorbed in the normalisation constants $k$. All the normalisation constants for the SCP-II case contain the following integral:
$\int\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q$.
For $\gamma \neq 0,-1,-2,-3, \ldots$ the integral evaluates to an expression involving a hypergeometric function:
$\frac{2^{\gamma} q^{1+\gamma}{ }_{2} F_{1}(\gamma+1, \gamma ; \gamma+2 ;-q)}{\gamma+1}$.
where the result was obtained by using the website integrals.wolfram.com. The value $\gamma=0$ is not allowed as $\alpha=1$ was excluded. For $\gamma=-1(\alpha=0)$ the integral is:
$\frac{1}{2}(q+\ln q)$.
From the properties of the hypergeometric function (see PCP-III case) it follows that for $\gamma+2=-n$ the expression above does not converge. This means that for $\gamma=-p$, where $p=2,3,4, \ldots$ the integral has to be evaluated separately. In this case we can write:

$$
\begin{aligned}
\int\left(\frac{2 q}{1+q}\right)^{\gamma} \mathrm{d} q & =2^{-p} \int\left(1+\frac{1}{q}\right)^{p} \mathrm{~d} q \\
& =2^{-p} \int \sum_{r=0}^{p}\binom{p}{r} q^{-r} \mathrm{~d} q
\end{aligned}
$$

which evaluates to:
$2^{-p}\left(1+p \ln q+\sum_{r=2}^{p}\binom{p}{r} \frac{1}{1-r} q^{1-r}\right)$.







Fig. A.9. Mass ratio distributions $f_{q}(q)$ for split-core pairing II and III for a uniform generating mass ratio distribution and core masses from single power-law mass distribution for $\alpha=2.35,0.30,-2.50$, where $2 c \leq M_{\mathrm{c}} \leq 2 d$. The curves are shown for $c / d=0.01,0.1,0.3,0.5$.

The SCP-III case Without restrictions on the primary mass range the expression for the mass ratio distribution is:
where the term for $h^{\prime}{ }_{q}$ follows from $h^{\prime}{ }_{q}(q)=1 /(1-c /(x-c))$. Note that the distribution is normalised. To bring out better the dependence on $c / d$ one can also write (substituting $z=x / d$ ):
$f_{q}(q)=\int_{c(1+1 / q)}^{2 d}\left(\frac{x-c}{x-2 c}\right) a x^{-\alpha} \mathrm{d} x$,

$$
\begin{equation*}
f_{q}(q)=\frac{\gamma(2 c / d)^{\gamma}}{1-(c / d)^{\gamma}} \int_{(c / d)(1+1 / q)}^{2}\left(\frac{z-(c / d)}{z-(2 c / d)}\right) z^{-\alpha} \mathrm{d} z . \tag{A.87}
\end{equation*}
$$



Fig. A.10. Mass ratio distributions $f_{q}(q)$ for split-core pairing II and III for a uniform generating mass ratio distribution and core masses from a single power-law mass distribution with $\alpha=2.35$, where $2 c \leq M_{\mathrm{c}} \leq 2 d$. The black solid curve shows the complete mass ratio distribution (for all binaries in the population) for $c=0.2$ and $d=2.0$. The other curves show what happens to the observed $f_{q}(q)$ if the primary mass $M_{1}$ is restricted to $x_{1} \leq M_{1} \leq x_{2}$. Results are shown for $\alpha=2.35$ (top), $\alpha=0.3$ (middle) and $\alpha=-2.5$ (bottom). The examples cases are listed in the legend.

The integral over $z$ :
$\int\left(\frac{z-(c / d)}{z-(2 c / d)}\right) z^{-\alpha} \mathrm{d} z$
can again be evaluated using the hypergeometric function for $\alpha \neq 0,1,2,3, \ldots$, and the result is:
$\frac{z^{1-\alpha}\left({ }_{2} F_{1}\left(1-\alpha, 1 ; 2-\alpha ; \frac{z}{2 c / d}\right)-2\right)}{2(\alpha-1)}$.

The excluded values of $\alpha$ again follow from the convergence properties of the hypergeometric function. For $\alpha=0$ the integral evaluates to:
$z-(c / d) \ln (z-2 c / d)$.
The case $\alpha=1$ is excluded and for $\alpha=p=2,3,4, \ldots$ the integral can be written as:
$\int\left(\frac{1}{z^{p-1}(z-2 c / d)}-\frac{c / d}{z^{p}(z-2 c / d}\right) \mathrm{d} z$.
The solution can be found again from formula 2.117(4) in Gradshteyn \& Ryzhik (2007):

$$
\begin{align*}
& \sum_{r=1}^{p-2} \frac{1}{(p-1-r)(2 c / d)^{r} z^{p-1-r}}+\left(\frac{2 c}{d}\right)^{1-p} \ln \left(\frac{z-2 c / d}{z}\right) \\
& -\left(\frac{c}{d}\right)\left(\sum_{r=1}^{p-1} \frac{1}{(p-r)(2 c / d)^{r} z^{p-r}}+\left(\frac{2 c}{d}\right)^{-p} \ln \left(\frac{z-2 c / d}{z}\right)\right) \tag{A.92}
\end{align*}
$$

With restrictions on the primary mass range the integral discussed above has to be evaluated for the various integration intervals corresponding to the cases discussed for SCP-II. The normalisation constants $k$ can be obtained from a numerical integration of the functions $f_{q}(q)$. For $x_{1}=x_{2}$ the mass ratio distribution becomes $f_{q}(q)=1 /\left(1-c /\left(x_{1}-c\right)\right)=\left(x_{1}-c\right) /\left(x_{1}-2 c\right)$. Figures A. 9 and A. 10 show examples of what $f_{q}(q)$ looks like for all the SCP-II and SCP-III cases discussed above. Finally, we have listed the expressions for the binary fraction as a function of primary mass in Table A.3.

## References

Aarseth, S. J. 1999, PASP, 111, 1333
Abt, H. A., \& Levy, S. G. 1976, ApJS, 30, 273
Abt, H. A., Gomez, A. E., \& Levy, S. G. 1990, ApJS, 74, 551
Ahmic, M., Jayawardhana, R., Brandeker, A., et al. 2007, ApJ, 671, 2074
Armitage, P. J., \& Bonnell, I. A. 2002, MNRAS, 330, L11
Bate, M. 2008, MNRAS, in press [arXiv:0811.0163]
Bate, M. R., Bonnell, I. A., \& Bromm, V. 2003, MNRAS, 339, 577
Blaauw, A. 1961, Bull. Astron. Inst. Netherlands, 15, 265
Blaauw, A. 1991, in NATO ASIC Proc. 342: The Physics of Star Formation and Early Stellar Evolution, 125
Bonnell, I. A., Larson, R. B., \& Zinnecker, H. 2007, in Protostars and Planets V, ed. B. Reipurth, D. Jewitt, \& K. Keil, 149
Bouy, H., Brandner, W., Martín, E. L., et al. 2003, AJ, 126, 1526
Bouy, H., Martín, E. L., Brandner, W., et al. 2006, A\&A, 451, 177
Burgasser, A. J., Kirkpatrick, J. D., Reid, I. N., et al. 2003, ApJ, 586, 512
Chabrier, G. 2003, PASP, 115, 763
Clarke, C. J. 1996a, MNRAS, 283, 353
Clarke, C. J. 1996b, in The Origins, Evolution, and Destinies of Binary Stars in Clusters, ed. E. F. Milone, \& J.-C. Mermilliod, ASP Conf. Ser., 90, 242
Close, L. M., Siegler, N., Freed, M., \& Biller, B. 2003, ApJ, 587, 407
Correia, S., Zinnecker, H., Ratzka, T., \& Sterzik, M. F. 2006, A\&A, 459, 909
Delgado Donate, E. J., \& Clarke, C. J. 2005, Astron. Nachr., 326, 940
Duchêne, G., Bouvier, J., Bontemps, S., André, P., \& Motte, F. 2004, A\&A, 427, 651
Duquennoy, A., \& Mayor, M. 1991, A\&A, 248, 485
Durisen, R. H., Sterzik, M. F., \& Pickett, B. K. 2001, A\&A, 371, 952
Eggleton, P. P., \& Tokovinin, A. A. 2008, MNRAS, 389, 869
Elmegreen, B. G. 2007, ApJ, 668, 1064
Fischer, D. A., \& Marcy, G. W. 1992, ApJ, 396, 178
Gaburov, E., Gualandris, A., \& Portegies Zwart, S. 2008, MNRAS, 384, 376
García, B., \& Mermilliod, J. C. 2001, A\&A, 368, 122
Gizis, J. E., Reid, I. N., Knapp, G. R., et al. 2003, AJ, 125, 3302
Goodwin, S. P., \& Kroupa, P. 2005, A\&A, 439, 565
Goodwin, S. P., Kroupa, P., Goodman, A., \& Burkert, A. 2007, in Protostars and Planets V, ed. B. Reipurth, D. Jewitt, \& K. Keil, 133
Goodwin, S. P., Nutter, D., Kroupa, P., Ward-Thompson, D., \& Whitworth, A. P. 2008, A\&A, 477, 823

Gradshteyn, I., \& Ryzhik, I. 2007, Table of Integrals, Series, and Products, 7th edn. (Academic Press)
Grether, D., \& Lineweaver, C. H. 2006, ApJ, 640, 1051
Halbwachs, J. L., Mayor, M., Udry, S., \& Arenou, F. 2003, A\&A, 397, 159
Harayama, Y., Eisenhauer, F., \& Martins, F. 2008, ApJ, 675, 1319
Heggie, D. C. 1975, MNRAS, 173, 729
Hills, J. G. 1975, AJ, 80, 809
Hogeveen, S. J. 1992a, Ap\&SS, 196, 299
Hogeveen, S. J. 1992b, Ap\&SS, 196, 299
Hoyle, F. 1953, ApJ, 118, 513
Hu, Y., Deng, L., de Grijs, R., Goodwin, S. P., \& Qiang, L. 2008, MNRAS, submitted [arXiv:0801.2814]
Hubber, D. A., \& Whitworth, A. P. 2005, A\&A, 437, 113
Joergens, V. 2008, A\&A, 492, 545
Kobulnicky, H. A., \& Fryer, C. L. 2007, ApJ, 670, 747
Köhler, R., Petr-Gotzens, M. G., McCaughrean, M. J., et al. 2006, A\&A, 458, 461
Kouwenhoven, M. B. N. 2006, Ph.D. Thesis, University of Amsterdam [arXiv:astro-ph/0610792]
Kouwenhoven, M. B. N., Brown, A. G. A., Zinnecker, H., Kaper, L., \& Portegies Zwart, S. F. 2005, A\&A, 430, 137
Kouwenhoven, M. B. N., Brown, A. G. A., \& Kaper, L. 2007a, A\&A, 464, 581
Kouwenhoven, M. B. N., Brown, A. G. A., Portegies Zwart, S. F., \& Kaper, L. 2007b, A\&A, 474, 77
Kraus, A. L., Ireland, M. J., Martinache, F., \& Lloyd, J. P. 2008, ApJ, 679, 762
Kroupa, P. 1995a, MNRAS, 277, 1491
Kroupa, P. 1995b, MNRAS, 277, 1522
Kroupa, P. 1995c, MNRAS, 277, 1507
Kroupa, P. 2001, MNRAS, 322, 231
Kroupa, P. 2008, in The Cambridge $N$-body Lectures, ed. S. Aarseth, C. Tout, \& R. Mardling [arXiv:0803.1833]

Kroupa, P., Tout, C. A., \& Gilmore, G. 1990, MNRAS, 244, 76
Kroupa, P., Petr, M. G., \& McCaughrean, M. J. 1999, New Astron., 4, 495
Kroupa, P., Aarseth, S., \& Hurley, J. 2001, MNRAS, 321, 699
Kroupa, P., Bouvier, J., Duchêne, G., \& Moraux, E. 2003, MNRAS, 346, 354
Krumholz, M. R., \& Thompson, T. A. 2007, ApJ, 661, 1034
Kuiper, G. P. 1935, PASP, 47, 15
Lada, C. J. 2006, ApJ, 640, L63
Larson, R. B. 1969, MNRAS, 145, 271
Larson, R. B. 1992, MNRAS, 256, 641
Larson, R. B. 2005, MNRAS, 359, 211
Leinert, C., Henry, T., Glindemann, A., \& McCarthy, Jr., D. W. 1997, A\&A, 325, 159
Low, C., \& Lynden-Bell, D. 1976, MNRAS, 176, 367
Lucy, L. B. 2006, A\&A, 457, 629
Lucy, L. B., \& Ricco, E. 1979, AJ, 84, 401
Maíz Apellániz, J. 2008, ApJ, 677, 1278
Mal'Kov, O., \& Zinnecker, H. 2001, MNRAS, 321, 149
Martín, E. L., Barrado y Navascués, D., Baraffe, I., Bouy, H., \& Dahm, S. 2003, ApJ, 594, 525
Mason, B. D., Gies, D. R., Hartkopf, W. I., et al. 1998, AJ, 115, 821
Masunaga, H., \& Inutsuka, S.-I. 2000, ApJ, 531, 350
Maxted, P. F. L., Jeffries, R. D., Oliveira, J. M., Naylor, T., \& Jackson, R. J. 2008, MNRAS, 385, 2210
Mazeh, T., Goldberg, D., Duquennoy, A., \& Mayor, M. 1992, ApJ, 401, 265
Mazeh, T., Simon, M., Prato, L., Markus, B., \& Zucker, S. 2003, ApJ, 599, 1344
McCarthy, C., \& Zuckerman, B. 2004, AJ, 127, 2871
Metchev, S., \& Hillenbrand, L. 2008 [arXiv:0808.2982]
Pinsonneault, M. H., \& Stanek, K. Z. 2006, ApJ, 639, L67
Piskunov, A. E., \& Mal'Kov, O. I. 1991, A\&A, 247, 87
Pollack, J. B., Hubickyj, O., Bodenheimer, P., et al. 1996, Icarus, 124, 62
Portegies Zwart, S. F., McMillan, S. L. W., Hut, P., \& Makino, J. 2001, MNRAS, 321, 199
Preibisch, T., Brown, A. G. A., Bridges, T., Guenther, E., \& Zinnecker, H. 2002, AJ, 124, 404
Preibisch, T., Stanke, T., \& Zinnecker, H. 2003, A\&A, 409, 147
Press, W., Teukolsky, S., Vetterling, W., \& Flannery, B. 2007, Numerical Recipes: The Art of Scientific Computing, 3rd edn. (Cambridge University Press)
Reddish, V. C. 1978, Stellar formation, International Series in Natural Philosophy (Oxford: Pergamon)
Rees, M. J. 1976, MNRAS, 176, 483
Reid, I. N., \& Gizis, J. E. 1997, AJ, 113, 2246
Reipurth, B., \& Clarke, C. 2001, AJ, 122, 432
Reipurth, B., \& Zinnecker, H. 1993, A\&A, 278, 81
Reipurth, B., Guimarães, M. M., Connelley, M. S., \& Bally, J. 2007, AJ, 134, 2272
Salpeter, E. E. 1955, ApJ, 121, 161

Shatsky, N., \& Tokovinin, A. 2002, A\&A, 382, 92
Siegler, N., Close, L. M., Cruz, K. L., Martín, E. L., \& Reid, I. N. 2005, ApJ, 621, 1023
Silk, J. 1977a, ApJ, 214, 152
Silk, J. 1977b, ApJ, 214, 718
Silk, J. 1995, ApJ, 438, L41
Sills, A., Adams, T., Davies, M. B., \& Bate, M. R. 2002, MNRAS, 332, 49
Söderhjelm, S. 2007, A\&A, 463, 683
Sterzik, M. F., \& Durisen, R. H. 2004, in Rev. Mex. Astron. Astrofis. Conf. Ser., 21, ed. C. Allen, \& C. Scarfe, 58
Stolte, A., Ghez, A. M., Morris, M., et al. 2008, ApJ, 675, 1278
Swift, J. J., \& Williams, J. P. 2008, ApJ, 679, 552
Thies, I., \& Kroupa, P. 2007, ApJ, 671, 767
Thies, I., \& Kroupa, P. 2008, MNRAS, 390, 1200
Tohline, J. E. 1982, Fundamentals of Cosmic Physics, 8, 1
Tokovinin, A., Thomas, S., Sterzik, M., \& Udry, S. 2006, A\&A, 450, 681

Tokovinin, A. A. 2000, A\&A, 360, 997
Tokovinin, A. A., \& Smekhov, M. G. 2002, A\&A, 382, 118
Tout, C. A. 1991, MNRAS, 250, 701
van den Berk, J., Portegies Zwart, S. F., \& McMillan, S. L. W. 2007, MNRAS, 379, 111
van Rensbergen, W., de Loore, C., \& Jansen, K. 2006, A\&A, 446, 1071
Vanbeveren, D. 1982, A\&A, 115, 65
Warner, B. 1961, PASP, 73, 439
Weidner, C., \& Kroupa, P. 2006, MNRAS, 365, 1333
Weidner, C., Kroupa, P., \& Maschberger, T. 2008, MNRAS, in press [arXiv:0811.3730]
Whitworth, A., Bate, M. R., Nordlund, Å., Reipurth, B., \& Zinnecker, H. 2007, in Protostars and Planets V, ed. B. Reipurth, D. Jewitt, \& K. Keil, 459
Zinnecker, H. 1984, Ap\&SS, 99, 41
Zinnecker, H., \& Yorke, H. W. 2007, ARA\&A, 45, 481

