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Genetic risk and atrial fibrillation in patients with heart failure

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Aims

To study the association between an atrial fibrillation (AF) genetic risk score with prevalent AF and all-cause mortality in patients with heart failure.

Methods and results

An AF genetic risk score was calculated in 3759 European ancestry individuals (1783 with sinus rhythm, 1976 with AF) from the BIOlogy Study to Tailored Treatment in Chronic Heart Failure (BIOSTAT-CHF) by summing 97 single nucleotide polymorphism (SNP) alleles (ranging from 0–2) weighted by the natural logarithm of the relative SNP risk from the latest AF genome-wide association study. Further, we assessed AF risk variance explained by additive SNP variation, and performance of clinical or genetic risk factors, and the combination in classifying AF prevalence. AF was classified as AF or atrial flutter (AFL) at baseline electrocardiogram and/or a history of AF or AFL. The genetic risk score was associated with AF after multivariable adjustment. Odds ratio for AF prevalence per 1-unit increase genetic risk score was 2.12 (95% confidence interval 1.84–2.45, $P = 2.15 \times 10^{-24}$) in the total cohort, 2.08 (1.72–2.50, $P = 1.30 \times 10^{-14}$) in heart failure with reduced ejection fraction (HFrEF) and 2.02 (1.37–2.99, $P = 4.37 \times 10^{-4}$) in heart failure with preserved ejection fraction (HFpEF). AF-associated loci explained 22.9% of overall AF SNP heritability. Addition of the genetic risk score to clinical risk factors increased the C-index by 2.2% to 0.721.

Conclusions

The AF genetic risk score was associated with increased AF prevalence in HFrEF and HFpEF. Genetic variation accounted for 22.9% of overall AF SNP heritability. Addition of genetic risk to clinical risk improved model performance in classifying AF prevalence.

Keywords

Atrial fibrillation • Heart failure • Genetic association studies • Single nucleotide polymorphism • Risk factors

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Introduction

Atrial fibrillation (AF) is the most common cardiac arrhythmia and is highly prevalent in patients with heart failure.^{1–3} The co-existence of these conditions can be expected by virtue of their prevalence alone: the lifetime risk of developing AF is about one in three in individuals of European ancestry and one in five in individuals of African ancestry,^{4–6} and after age 45 the lifetime risk of heart failure ranges between 20–45%.⁷

Furthermore, both conditions have many shared risk factors which makes their co-existence more likely.^{8,9} Additionally, a reciprocal relation between both conditions seems to exist, but regardless of which condition occurs first, the concomitant presence of both AF and heart failure is associated with substantially increased risks of mortality.^{2,3}

Atrial fibrillation is common in heart failure and prevalence of the arrhythmia increases with heart failure severity, but little is known about the mechanisms that underlie AF onset in heart failure patients.^{10,11} Genetic factors could theoretically explain, at least partly, the increased risk of AF in patients with heart failure.¹² But heritability of AF is complex; in a recent study, 97 genome-wide susceptibility loci for AF were identified, and the proportion of heritability explained by the loci in individuals of European ancestry was 42%.¹³ Prevalence estimates of heart failure in population-based biobanks and case-referent studies used for AF genome-wide association studies (GWAS) is limited, and it remains unclear whether individuals with AF in the context of heart failure share a similar genetic susceptibility to the arrhythmia.

We aimed to study the association between a genetic risk score based on 97 lead single nucleotide polymorphisms (SNPs)¹³ with prevalent AF and all-cause mortality in a large sample of patients with heart failure included in The BIOlogy Study to Tailored Treatment in Chronic Heart Failure (BIOSTAT-CHF). Further, we assessed the variance in AF prevalence explained by additive SNP variation (SNP heritability), and determined the discriminatory accuracy of clinical risk factors, genetic risk factors, and the combination in classifying AF prevalence.

Methods

Study population

The prospective, observational, international BIOSTAT-CHF study included 2516 patients with heart failure from 11 European countries between December 2010 and December 2012. Another 1738 patients from Scotland were included in a validation cohort between October 2010 and April 2014. The rationale, design, and primary results have been previously published.¹⁴ Briefly, the majority of patients were hospitalized for acute heart failure, and the remainder presented with worsening signs and/or symptoms of heart failure at outpatient clinics. Patients had to have objective evidence of cardiac dysfunction documented either by left ventricular ejection fraction (LVEF) of $\leq 40\%$, previous heart failure hospitalization, or plasma concentrations of B-type natriuretic peptide (BNP) and/or N-terminal pro-B-type natriuretic peptide (NT-proBNP) >400 pg/mL

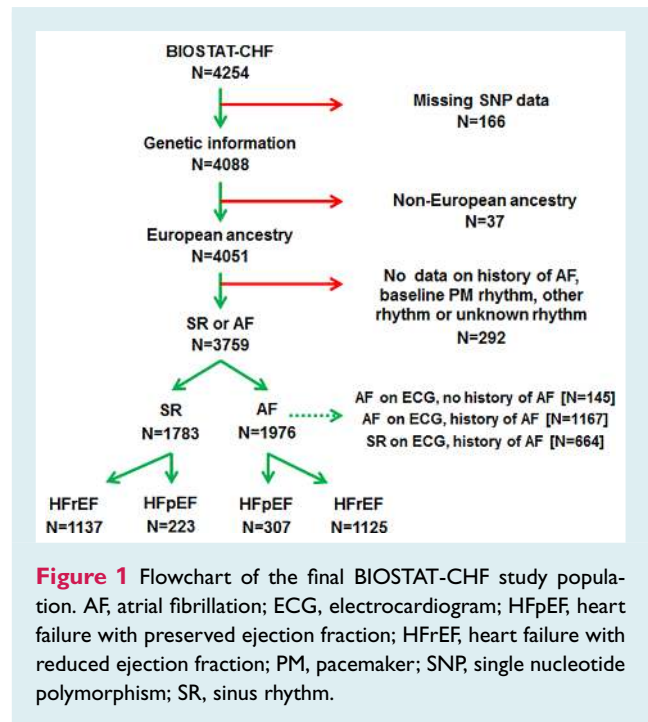


Figure 1 Flowchart of the final BIOSTAT-CHF study population. AF, atrial fibrillation; ECG, electrocardiogram; HFpEF, heart failure with preserved ejection fraction; HFpEF, heart failure with reduced ejection fraction; PM, pacemaker; SNP, single nucleotide polymorphism; SR, sinus rhythm.

or >2000 pg/mL, respectively. According to study design, all patients used diuretics but were not on optimal, evidence-based medical therapy of angiotensin-converting enzyme inhibitors/angiotensin receptor blockers and, or beta-blockers. After inclusion patients were extensively phenotyped and genotyped, underwent physical examination and quality of life measurements, and plasma, serum, and urine samples were collected for analysis. During the first 3 months of follow-up, medication was optimized. The study complies with the Declaration of Helsinki, medical ethics committee of participating centres approved the study, and all patients provided written informed consent before inclusion.

Patient selection

For the current analysis the BIOSTAT-CHF index cohort ($n = 2516$) and validation cohort ($n = 1738$) were combined to achieve a larger set of patients ($n = 4254$). Patients with no blood samples available for genotyping ($n = 166$), self-reported non-European ancestry ($n = 37$), and pacemaker rhythm or missing variables that prohibited rhythm classification ($n = 292$) were excluded (Figure 1).

Atrial fibrillation prevalence and all-cause mortality

Atrial fibrillation prevalence was defined as clinical history of AF or atrial flutter (AFL) and/or AF(L) on baseline electrocardiogram (ECG). Patients were regarded as having sinus rhythm if they had no history of AF and sinus rhythm on baseline ECG. Incident AF was not captured during follow-up.

After the optimization (3 months) and maintenance phase (6 months),¹⁴ patients were followed by standard clinical follow-up or telephone contact with 6-month intervals. Follow-up ended on April 1st 2015. Median follow-up duration was approximately

21 months. During follow-up all deaths and hospitalizations were recorded. For the current analysis, all-cause mortality was assessed.

Genotyping in BIOSTAT-CHF

The two cohorts were processed, genotyped, QC'd and imputed independently, using the same protocols. Genotyping of all patients from both BIOSTAT-CHF cohorts was performed using the Affymetrix Axiom Genome-Wide UKB WCGS genotyping array. Sample level QC was performed for X chromosome homozygosity (sex mismatch) and identity by descent (IBD) estimates (relatedness and duplicates). Prior to imputation, variants were removed if their call rate was <95% for variants with minor allele frequency (MAF) $\geq 5\%$, or <99% for variants with MAF <5%, or had a Hardy–Weinberg equilibrium $P < 1 \times 10^{-6}$. Imputation was performed using SHAPEIT¹⁵ and IMPUTE2¹⁶ with the phase 3 release 1000G reference panel.¹⁷

Genetic analysis

Atrial fibrillation genetic risk score

Genotypes of 97 SNPs associated with AF risk in the latest published GWAS¹³ with significance thresholds of $P < 1 \times 10^{-8}$ were used to calculate an individual patient AF genetic risk score by summing the dosage of each AF risk allele in BIOSTAT-CHF (ranging from 0–2) weighted by the natural logarithm of the relative risk for each SNP. Weights were determined by the latest AF GWAS¹³ (online supplementary Table S1). The SNP rs465276 was not available in BIOSTAT-CHF and was substituted with a proxy (rs361834, $r^2 = 0.91$, based on pairwise linkage disequilibrium from European ancestry samples in the Broad AF study¹³). All SNPs had an INFO score > 0.4 and a Hardy–Weinberg equilibrium $P > 1 \times 10^{-6}$. AF genetic risk scores were calculated using PLINK v2.00.¹⁸

Proportion of heritability explained

We assessed the proportion of AF phenotypic variance explained by additive genetic variation, otherwise referred to as SNP heritability (h^2_g). h^2_g was calculated with the software BOLT-LMM v2.3.2.¹⁹ The AF loci were defined as a region of 1 Mb (± 500 kb) around each of the 97 reported sentinel variants from the latest AF GWAS analysis.¹³ We used the imputed genotype data, filtered the variants for imputation quality > 0.8 , as calculated by QCTOOL v2,²⁰ hard-called the genotypes with a genotype probability threshold > 0.9 with PLINK v2.00,¹⁸ and combined the overlapping variants that remained from the index and validation cohort of BIOSTAT-CHF. Additional filtering removed variants with MAF $< 1\%$ and variant call rate missingness $> 0.5\%$. We then applied one round of pruning with the settings – indep-pairwise 50 5 0.9 in PLINK. The heritability calculation was performed on the remaining 806130 variants. We included age, sex, and the first five principal components as covariates. The observed heritability estimates were converted to the liability scale following equation 17 from Lee *et al.*²¹ and using the AF prevalence in the BIOSTAT-CHF cohorts (AF prevalence of 53%) as disease prevalence in a heart failure population.

Statistical analyses

Normally distributed variables are depicted as means \pm standard deviation and non-normally distributed variables as median with the first and third quartile (Q1, Q3). Categorical variables are presented

as numbers with percentages. Multivariable logistic regression models were used to examine whether a genetic risk score build of 97 AF genetic loci was associated with AF prevalence. Model 1 was adjusted for age, sex, and the first 10 principal components of ancestry. Model 2 was adjusted for clinical AF risk factors from the CHARGE-AF risk model,²² a model aimed to predict future risk of AF. Variables include: age, height, weight, systolic and diastolic blood pressure, current smoking, hypertension as a proxy for antihypertensive treatment, diabetes, myocardial infarction, and the first 10 principal components of ancestry. The CHARGE-AF risk model variables heart failure and race were not included since our population consists of European ancestry patients with heart failure. A total of 96 patients had missing values and were excluded. We calculated the area under the receiver operating curve (AUC) in logistic regression models for AF prevalence. All calculations included the first 10 principal components and were performed in R using the package *pROC*²³ to calculate the AUC and the 95% confidence intervals (CI) with the DeLong method. Cox proportional hazard analysis was performed to determine hazard ratios (HR) with 95% CI for the genetic risk score and all-cause mortality. All HR were adjusted for covariates of the CHARGE-AF risk model. The Cox proportional hazards assumption was assessed by visually inspecting plots of Schoenfeld residuals against time, which showed no proportionality violation (i.e. the plots showed random patterns of residuals against time). Interaction testing was performed to determine whether the effect of the genetic risk score differed between the heart failure phenotypes, with regard to AF prevalence and all-cause mortality risk. Secondary analyses were performed in subgroups based on LVEF: LVEF $< 40\%$ and LVEF $\geq 50\%$, respectively, heart failure with reduced ejection fraction (HFrEF) and heart failure with preserved ejection fraction (HFpEF). Patients with a mid-range ejection fraction (LVEF 40–49%) or missing LVEF data were not assessed separately. Analyses were performed using IBM SPSS Statistics version 23. The a priori significance threshold for all analyses was $P < 0.05$ using 2-sided tests.

Results

Patient characteristics

An overview of the cohort is shown in Table 1. A total of 3759 European ancestry individuals from BIOSTAT-CHF were included, of whom 1976 (53%) had prevalent AF. Mean age was 72.8 ± 11.5 , 30% were women. These patients were further stratified in 2262 HFrEF patients, of whom 1137 (50.3%) were in sinus rhythm and 1125 (49.7%) had AF; and 530 HFpEF patients, of whom 223 (42%) were in sinus rhythm and 307 (58%) had AF (Figure 1). Overall, patients with AF were older (75.0 ± 10.2 vs. 70.3 ± 12.3 years), more often men (73% vs. 67%), and had a higher body mass index (28.7 ± 5.9 vs. 28.0 ± 5.9 kg/m²). AF patients more often had renal disease (38% vs. 29%), but less often had coronary artery disease (43% vs. 54%) (all $P < 0.001$).

Genetic risk score and atrial fibrillation prevalence

In the total cohort, the AF genetic risk score ranged between 4.62 to 8.29 with a median of 6.37. After multivariable adjustment, the odds ratio for AF presence was 2.12 per 1-unit

Table 1 Baseline characteristics

	Overall (n = 3759)	AF (n = 1976, 53%)	SR (n = 1783, 47%)	P-value
Demographics				
Age, years	72.8 ± 11.5	75.0 ± 10.2	70.3 ± 12.3	<0.001
Women, n (%)	1128 (30)	537 (27)	591 (33)	<0.001
NYHA class I/II/III/IV, %	6/43/36/7	5/46/41/8	8/47/37/8	0.001
Clinical variables				
BMI, kg/m ²	28.3 ± 5.9	28.7 ± 5.9	28.0 ± 5.9	<0.001
Blood pressure, mmHg				
Systolic	125 ± 22	124 ± 21	127 ± 23	0.002
Diastolic	73 ± 14	73 ± 14	72 ± 13	0.01
Heart rate, bpm	78 ± 19	80 ± 21	75 ± 16	<0.001
Medical history, n (%)				
Coronary artery disease ^a	1814 (48)	856 (43)	958 (54)	<0.001
Hypertension	2295 (61)	1221 (62)	1074 (60)	0.32
Diabetes mellitus	1218 (32)	657 (33)	561 (31)	0.25
Renal disease ^a	1276 (34)	757 (38)	519 (29)	<0.001
Echocardiographic data				
LVEF, %	35 ± 13	36 ± 13	34 ± 13	<0.001
HFrEF ^b , n (%)	2262 (60)	1125 (57)	1137 (64)	<0.001
HFpEF ^c , n (%)	530 (14)	307 (16)	223 (13)	<0.001
Laboratory data				
NT-proBNP, ng/L, median (IQR)	2096 (825–4861)	2537 (1128–5122)	1588 (515–4510)	<0.001
Medications, n (%)				
ACEi/ARB	2681 (71)	1370 (69)	1311 (74)	0.005
Beta-blocker	2410 (64)	1307 (66)	1103 (62)	0.18
MRA	1670 (44)	872 (44)	798 (45)	0.37
Diuretics	3735 (99)	1960 (99)	1775 (99)	0.01

ACEi, angiotensin-converting enzyme inhibitor; AF, atrial fibrillation; ARB, angiotensin receptor blocker; BMI, body mass index; HFpEF, heart failure with preserved ejection fraction; HFrEF, heart failure with reduced ejection fraction; IQR, interquartile range; LVEF, left ventricular ejection fraction; MRA, mineralocorticoid receptor antagonist; NT-proBNP, N-terminal pro-B-type natriuretic peptide; NYHA, New York Heart Association; SR, sinus rhythm.

^aCoronary artery disease defined as: previous myocardial infarction, percutaneous coronary intervention and/or coronary artery bypass graft. Renal disease defined as estimated glomerular filtration rate < 60 mL/min/1.73 m².

^bHFrEF defined as LVEF < 40%.

^cHFpEF defined as LVEF ≥ 50%.

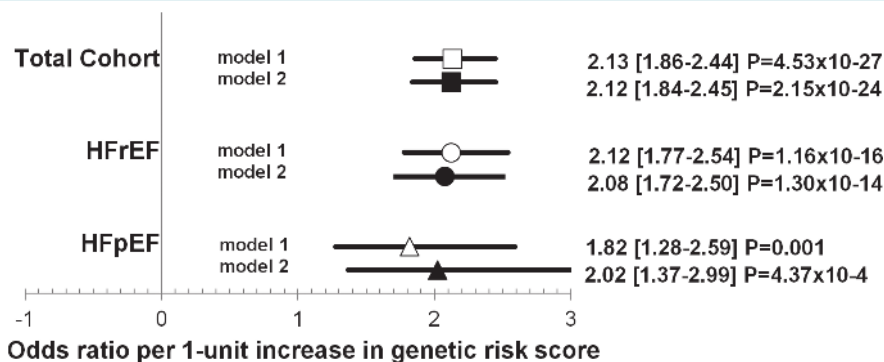


Figure 2 Genetic risk score and risk of atrial fibrillation prevalence. The bars signify the 95% confidence interval, the clear symbols represent results of model 1 and the solid symbols results of model 2. Squares indicate the total cohort, circles patients with heart failure with reduced ejection fraction (HFrEF), and triangles patients with heart failure with preserved ejection fraction (HFpEF). Model 1: adjusted for age, sex, and first 10 principal components of ancestry. Model 2: adjusted for age, height, weight, systolic and diastolic blood pressure, current smoking, hypertension, diabetes, myocardial infarction, and first 10 principal components of ancestry.

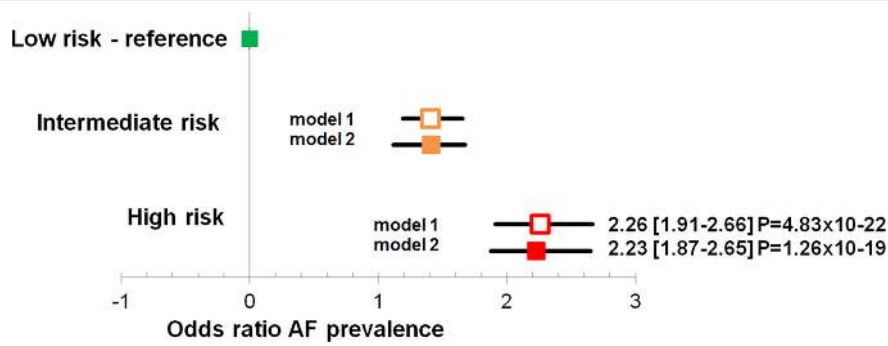


Figure 3 Increasing atrial fibrillation (AF) risk according to genetic risk score tertiles in the total cohort. The bars signify the 95% confidence interval, the clear symbols represent results of model 1 and the solid symbols results of model 2. Squares indicate the total cohort. Model 1: adjusted for age, sex, and first 10 principal components of ancestry. Model 2: adjusted for age, height, weight, systolic and diastolic blood pressure, current smoking, hypertension, diabetes, myocardial infarction, and first 10 principal components of ancestry.

Table 2 Proportion of heritability explained by atrial fibrillation loci

Study	AF-loci h^2_g observed (SE)	AF-loci h^2_g liability scale (SE)	Remaining genome h^2_g observed (SE)	Remaining genome h^2_g liability scale (SE)	Overall h^2_g liability scale	Proportion explained (%)
97 AF loci	0.0557 (0.0297)	0.0876 (0.0468)	0.1873 (0.1135)	0.2947 (0.1786)	0.3823	22.92

AF, atrial fibrillation; h^2_g , single nucleotide polymorphism heritability; SE, standard error.

Proportion of AF single nucleotide polymorphism heritability explained by AF loci, defined as a 1 Mb region around sentinel variants.

increase in genetic risk score (95% CI 1.84–2.45, $P = 2.15 \times 10^{-24}$) in the total BIOSTAT-CHF cohort (Figure 2, Model 2). The odds ratio were 2.08 per 1-unit increase in genetic risk score (95% CI 1.72–2.50, $P = 1.30 \times 10^{-14}$) in HF_rEF and 2.02 per 1-unit increase (95% CI 1.37–2.99, $P = 4.37 \times 10^{-4}$) in HF_pEF, respectively.

There was no interaction between genetic risk score and heart failure type on AF prevalence ($P = 0.99$). We estimated odds ratios comparing individuals in genetic risk score tertiles (Figure 3). The odds ratio for AF prevalence increased with higher genetic risk score categories. For the total BIOSTAT-CHF population, those in the highest tertile had 2.23 fold increased risk of AF compared to those in the lowest tertile (95% CI 1.87–2.65, $P = 1.26 \times 10^{-19}$).

Heritability and atrial fibrillation prevalence classification models

Atrial fibrillation-associated loci explain 22.9% of the overall AF SNP heritability (h^2_g) in our heart failure sample (Table 2).

The CHARGE-AF risk model had an AUC of 0.699 (95% CI 0.682–0.716) for accurately classifying AF prevalence, and was better than the genetic risk score alone (AUC 0.606; 95% CI 0.588–0.624). Combining the AF genetic risk score with the CHARGE-AF risk variables led to a model with an AUC of 0.721 (95% CI 0.704–0.737), a 2.2% increase over the CHARGE-AF risk model alone (Table 3).

Table 3 Area under the receiver operating curves for atrial fibrillation risk models

Risk model	AUC (95% CI)	P-value
CHARGE-AF clinical risk score	0.699 (0.682–0.716)	<0.001
AF genetic risk score	0.606 (0.588–0.624)	<0.001
CHARGE-AF clinical risk score + AF genetic risk score	0.721 (0.704–0.737)	<0.001

AF, atrial fibrillation; AUC, area under the receiver operating curve; CI, confidence interval.

Genetic risk score and all-cause mortality

During follow-up, with a median of 656 days (interquartile range 448–872 days), 1062 patients died (28%). In the total cohort, the genetic risk score was not associated with an increased risk for all-cause mortality after multivariable adjustment (HR 0.93, 95% CI 0.82–1.05, $P = 0.22$). Similar results were observed for the HF_rEF (HR 0.92, 95% CI 0.78–1.08, $P = 0.31$) and HF_pEF (HR 1.12, 95% CI 0.85–1.48, $P = 0.44$) subgroups. There was no interaction between heart failure subgroup and the genetic risk score on outcome ($P = 0.63$).

Discussion

In 3759 heart failure patients of European ancestry, an AF genetic risk score, based on lead SNPs at 97 AF loci, was associated with a higher prevalence of AF after adjustment for clinical AF risk variables from the CHARGE-AF risk model. We observed that 22.9% of variance in AF risk was attributable to additive genetic variation. Furthermore, addition of the AF genetic risk score to clinical risk factors improved risk model performance in classifying AF prevalence. The AF genetic risk score was not associated with all-cause mortality. Our findings support and extend the prior observation that there is, at least, a partial genetic basis for AF in patients with HFrEF and HFpEF.¹²

Genetic basis for atrial fibrillation in heart failure patients

Atrial fibrillation and heart failure frequently co-exist, but direct causality has not been unequivocally proven. Additionally, the underlying mechanisms that lead to the development of AF in HFrEF and HFpEF and vice versa remain complex and not completely understood. Previously the *ZFX3* gene was found to be associated with AF presence in a heart failure population.¹² Our comprehensive AF genetic risk score of 97 SNPs, together with the estimation that 22.9% of the phenotypic variance is explained by additive genetic variation, provide evidence of a substantial contribution of genome-wide variation to AF susceptibility in heart failure patients.

The genetic contribution to AF in our heart failure sample is less than what was previously observed in population based- and case-referent AF-GWAS studies, which also included a proportion of patients with heart failure (approximately 23% vs. 42%). Part of this portion of missing heritability may be caused by unidentified common genetic variants. Gene–environment interactions may also play a role, as genetic variants can also have associations with risk factors (pleiotropic effects). Heart failure patients have many risk factors including age, hypertension, diabetes, obesity, as well as valvular, ischaemic and non-ischaemic structural heart disease.^{10,11} On the other hand, increased cardiac filling pressures and consequently atrial stretch, cardiac fibrosis, dysregulation of intracellular calcium, and autonomic and neuroendocrine dysfunction in the setting of heart failure may evoke AF. It is possible that in the context of heart failure, with several concomitant risk factors, genetics may play a smaller role than in the general population.

It is hypothesized that AF in the presence of HFrEF is a marker of more advanced cardiac disease, with ventricular function deterioration and increased neurohormonal activation, while patients with AF and HFpEF share a more underlying substrate, albeit heterogeneous, with many shared risk factors.^{10,11} A difference in the genetic contribution to AF in HFrEF or HFpEF is not evident from current results, as no interaction between genetic risk score and heart failure type was observed.

Atrial fibrillation genetic risk score and all-cause mortality

Previous analyses in BIOSTAT-CHF have shown that worse cardiovascular outcomes were seen in heart failure patients with AF compared to sinus rhythm.²⁴ Nevertheless, after multivariable adjustment, the AF genetic risk score was not associated with all-cause mortality. However, a genetic risk score alone does not capture the clinical significance of AF presence in patients with an extensive cardiac substrate and other underlying risk factors. Additionally, current observations may be affected by survival bias.

Implications

The clinical risk factor model alone outperformed the genetic risk score, this is to be expected since compared to clinical risk factors the effect size of genetic variants is small, even when combined in a polygenic risk score. Although the genetic risk score had moderate discriminatory accuracy, we demonstrated that a combined risk model, consisting of the AF genetic risk score with clinical AF risk factors as present in the CHARGE-AF risk model, performed better than either risk model alone. But statistical significance does not automatically translate into clinical significance, and currently translation of genetics into clinical practice remains unresolved.

In the future, genetic profiling may provide insight into the mechanisms that underlie why some patients develop AF and others do not. The individual SNPs implicate genes that may reveal some of the mechanisms underlying AF (online supplementary Figure S1).¹³ Currently, most genes represent gene candidates at the loci, while the causal gene remains unknown. Experimental observations illustrate the pleiotropic nature of genes that are associated with this challenging arrhythmia and underscore the complexity of AF: so does *PITX2* encodes a transcription factor that plays a role in the formation of the pulmonary vein myocardium,²⁵ does *TBX5* encodes transcription factors that are required for patterning and maturing of the cardiac conduction system in mice²⁶ and have *KCNN3* and *SCN5A*, which both encode subunits of the potassium channel complex, been previously been linked to AF through candidate gene analyses and family-based studies.²⁷ More insights into the functional consequences of SNPs and genes is critical to identify potential therapeutic targets for this major health burden.²⁸ However, whether the genetic proportion to AF risk has a meaningful contribution to clinical risk assessment warrants further investigation.

Limitations

Current results, based on genetic data of 97 SNPs in 3759 patients from a well-defined heart failure cohort, point towards a genetic basis for AF in the context of heart failure. Analyses were limited to European ancestry individuals, and the current heart failure sample had a higher percentage of men with only 30% of women, and a higher percentage of HFrEF than is typical in the community; the findings may not be completely generalizable to individuals of different ancestral backgrounds, regions, or the general heart failure population. Additionally, women and men generally have a different

risk factor burden, which next to genetics and the underlying heart failure substrate, may be of different importance in the presence of concomitant heart failure and prevalent AF. Second, the genetic risk models were linear in nature with a single predictor variable and did not account for potential non-additive genetic effects, interactions between genetic variants, or interactions between genetic variants and environmental factors. Therefore, all observations are vulnerable to the risk of residual confounding that may bias mentioned estimates. Thirdly, AF ascertainment was partially based on physician-reported AF. This means that the percentage of AF is likely an underestimation since subclinical AF may have gone undetected. Fourthly, whether heart failure developed before the onset of AF, or AF before the onset of heart failure may be associated with a different genetic risk. Also the sequence in which the diseases develop can impact outcome. Unfortunately, we did not have information on the onset of AF and heart failure; therefore a temporal sequence of diagnoses was unknown, prohibiting time-dependent analyses. AF occurrence during follow-up was not systematically collected and therefore current analyses focus on baseline AF prevalence. Additionally, there was a lack of data on type and duration of AF, as well as applied therapies for AF. Fifthly, electro- and echocardiographic variables such as left atrial volume were omitted from the models since they were not available in a large proportion of patients. Additionally, these biomarkers will be influenced by both the underlying heart failure substrate as well as AF presence, duration, and severity. Covariates including LVEF, New York Heart Association class and NT-proBNP will be confounded by AF itself as it inhibits adequate echocardiographic determination of ejection fraction, is associated with symptoms of dyspnoea, and will lead to an increase in NT-proBNP levels. In line with the previous limitation, we did not adjust for heart failure severity in the multivariable models. We acknowledge that the CHARGE-AF model application in heart failure was not ideal, albeit the best validated AF risk score. Sixthly, in determining SNP heritability we assessed variants with MAF $\geq 1\%$, and, therefore, the contribution of rare or loss-of-function variants to total AF variance was not assessed. Furthermore, the estimates for SNP heritability have large standard errors bringing a level of uncertainty to these estimates. Seventhly, we cannot attribute the AF risk variance to functional categories. It remains challenging to identify the causal gene at each locus since the AF-associated SNPs predominantly fall within non-coding portions of the genome. Additionally, the association of genes to functional groups is based on their affiliation to enriched gene sets that were identified in an *in silico* analysis. Lastly, establishing a heart failure cohort of sufficient size is complex, and the current study is underpowered to study individual SNPs or perform extensive subgroup analyses. Larger studies, powered for outcomes, are warranted to investigate the genetic contribution to incident AF in heart failure populations, both HF_rEF and HF_pEF. Further efforts are needed to uncover the functional consequence of SNPs and genes at each locus on AF risk in patients with incident heart failure.

Conclusion

The AF genetic risk score was associated with increased AF prevalence in heart failure patients with reduced and preserved

ejection fraction. Genetic variation accounted for 22.9% overall AF SNP heritability. Addition of the AF genetic risk score to clinical risk factors improved risk model performance in classifying AF prevalence. Efforts are warranted to consider the role and mechanisms of genetic susceptibility of AF risk in heart failure patients.

Supplementary Information

Additional supporting information may be found online in the Supporting Information section at the end of the article.

Table S1. SNPs and weights used in the AF genetic risk score.

Figure S1. Venn diagram.

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