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Genome-Wide Association Study (GWAS) and Genome-Wide Environment Interaction Study (GWEIS) of Depressive Symptoms in African American and Hispanic/Latina Women

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Abstract

Background—Genome-wide association studies (GWAS) have been unable to identify variants linked to depression. We hypothesized that examining depressive symptoms and considering gene-environment interaction (G×E) might improve efficiency for gene discovery. We therefore conducted a GWAS and genome-wide environment interaction study (GWEIS) of depressive symptoms.

Methods—Using data from the SHARe cohort of the Women’s Health Initiative, comprising African Americans (n=7179) and Hispanics/Latinas (n=3138), we examined genetic main effects and G×E with stressful life events and social support. We also conducted a heritability analysis using genome-wide complex trait analysis (GCTA). Replication was attempted in four independent cohorts.

Results—No SNPs achieved genome-wide significance for main effects in either discovery sample. The top signals in African Americans were rs73531535 (located 20kb from *GPR139*, $p=5.75\times10^{-8}$) and rs75407252 (intronic to *CACNA2D3*, $p=6.99\times10^{-7}$). In Hispanics/Latinas, the top signals were rs2532087 (located 27kb from *CD38*, $p=2.44\times10^{-7}$) and rs4542757 (intronic to *DCC*, $p=7.31\times10^{-7}$). In the GWEIS with stressful life events, one interaction signal was genome-wide significant in African Americans (rs4652467; $p=4.10\times10^{-10}$; located 14kb from *CEP350*). This interaction was not observed in a smaller replication cohort. Although heritability estimates for depressive symptoms and stressful life events were each less than 10%, they were strongly genetically correlated ($r_G=0.95$), suggesting that common variation underlying depressive symptoms and stressful life event exposure, though modest on their own, were highly overlapping in this sample.

Conclusions—Our results underscore the need for larger samples, more GWEIS, and greater investigation into genetic and environmental determinants of depressive symptoms in minorities.

Keywords

genome-wide association study; gene-environment interaction; depression; stressful life events; social support

Introduction

Although family and twin studies show that depression is driven partly by genetic variation¹, until just recently², genome-wide association studies (GWAS) have made little progress in identifying specific loci linked to depression³. Several factors could explain the lack of success, including the complex genetic architecture of depression, small samples, and heterogeneity in the “depression” phenotype^{4,5}. Moreover, with the exception of two studies^{6,7}, including a large meta-analysis⁶, most prior GWAS have examined diagnoses, rather than quantitative traits (e.g., depressive symptoms). In light of evidence suggesting the diagnostic categories have been artificially imposed on a continuity of depression risk⁸, such case-control analyses may be limited. For example, simulations studies demonstrate that for common phenotypes (i.e., with prevalence greater than 10%), the quantitative trait approach may have power advantages under certain conditions in population-based samples⁹. GWAS have also neglected the role of gene-environment interaction (G×E)¹⁰, which many believe

contributes to the etiology of depression^{11,12}. Previous G×E studies have been limited to candidate genes; these results have been highly controversial^{13–16}. Studies of G×E in the context of GWAS for psychiatric phenotypes are needed and may be informative for identifying novel genomic loci^{17,18}. Indeed, G×E studies using genome-wide data for other complex phenotypes have revealed genotype-phenotype associations not apparent in genetic main effect analyses^{19–21}.

Further, genetic studies of depression and other psychiatric phenotypes have almost exclusively comprised samples of European ancestry, leaving racial/ethnic minorities underrepresented in psychiatric genetics work. Extending genetic association studies to more diverse racial/ethnic populations – especially of women – is therefore needed. These studies are likely to be informative, as depression appears at least as heritable (around 40%) among African Americans^{22,23} and Hispanics²⁴ compared to European Americans¹. Such extensions are also important given known racial/ethnic (as well as sex) disparities. For example, epidemiological studies have observed lower lifetime prevalence estimates for major depressive disorder (MDD) among non-Whites²⁵, despite a higher burden of social-environmental adversity from stressful life events²⁶, discrimination^{27,28} and lower socioeconomic status²⁹. Epidemiological studies have also consistently showed a two-fold elevated risk of MDD in women compared to men³⁰.

Here, we aimed to address these limitations by conducting a GWAS of depressive symptoms and performing a genome-wide environment interaction study (GWEIS) using data from a large population-based epidemiological sample of African American and Hispanic/Latina women drawn from the Women's Health Initiative (WHI).

Methods and Materials

Overview

As described elsewhere^{31,32} (www.whi.org), the WHI consists of an observational study (WHI-OS) and randomized clinical trial (WHI-CT). The WHI-OS prospectively followed 93,676 postmenopausal women ages 50–79 recruited from 40 clinical centers in the United States between 1993 and 1998. The WHI-CT enrolled 68,132 postmenopausal women of the same age and between the same time period to participate in one of three prevention trials: (1) hormone therapy; (2) dietary modification; and (3) calcium/vitamin D supplementation. We analyzed data from women genotyped as part of the WHI SNP Health Association Resource (SHARe), a sub-study of self-reported minority women in WHI (n=7,480 African American and 3,352 Hispanic/Latina women). All participants consented to be included in studies for general research use. Data were downloaded from the database of Genotypes and Phenotypes (dbGaP; accession #phs000200.v9.p3).

Phenotype Definition

Depressive symptoms were assessed at enrollment using total scores from a six-item version of the Center for Epidemiological Studies of Depression Scale (CES-D)³³, a widely-used measure of depressive symptoms in epidemiological studies. The six-item CES-D captured core symptoms of depression in the past week, including anhedonia, depressed mood, and

behavioral symptoms (e.g., felt depressed; sleep was restless; enjoyed life; had crying spells; felt sad; felt people disliked you). The six item scale correlates highly with the full 20-item CES-D ($r=0.88$)³². Brief versions of the CES-D correlate highly in older adults with diagnoses of MDD obtained from structured interviews³⁴.

As CES-D scores in this population-based sample could have been influenced by antidepressant medication use, we used a nonparametric imputation algorithm developed in a previous GWAS of depressive symptoms⁶ to adjust the CES-D score of women taking antidepressants (as determined by pill bottles women brought to the baseline interview). This algorithm, which increased the CES-D score for all antidepressant users, was based on one used to adjust blood pressure for persons on antihypertensive medications³⁵ (see Supplemental Materials).

We tested for statistical G×E interaction with two environmental exposures – stressful life events and social support – both of which were shown to correlate with depressive symptoms in WHI³² and numerous other studies³⁶. These two social-environmental exposures were measured at enrollment, concurrently with depressive symptoms. Stressful life events were assessed using a scale modified from the Almeida County Study^{37,38}, which asked women to indicate whether they had experienced 11 different major losses or traumatic events in the past year (see Supplemental Materials for specific items). Items were summed to create a total count of the number of past-year stressors among those with complete data on all stressors (ranging from 0–11). Social support was assessed using nine items from the 19-item Medical Outcome Survey³⁹. We summed across these items to obtain a measure for level of perceived social support.

SNP Genotyping and Imputation

All participants were genotyped using the Affymetrix 6.0 chip designed to human genome build 36. Genotyping, on all samples plus 2% blinded duplicates, was performed at Affymetrix, Inc., Santa Clara, CA. A total of 720,101 (African Americans) and 709,042 (Hispanics/Latinas) SNPs passed pre-imputation filters.

Quality control procedures were performed at the Fred Hutchinson Cancer Research Center (FHCRC) in Seattle, WA. As described elsewhere (refer to⁴⁰ and Supplemental Materials), the WHI GARNET Coordinating Center (www.garnetstudy.org) performed the imputation using the 1000 Genomes Interim reference panel (release December 2010) and BEAGLE software version 3.3.1⁴¹.

Quality Control (QC) of SNPs and Samples

In addition to the QC standards imposed by WHI, we additionally excluded SNPs with a MAF of 2 % or imputation quality score $< r^2=0.80$. Population stratification was assessed by WHI investigators using a principal components analysis estimated by the program EIGENSTRAT⁴². A total of 61 genetic outliers were removed from the African American analysis based on their PCA scores. After QC, 10,771 women (7,419 African American and 3,352 Hispanic/Latina women) were available for analysis. Allele dosages (meaning the probability of each of the three genotypes), rather than hard-called or “best guess” genotypes were used for both the GWAS and GWEIS analyses.

Statistical Analyses

GWAS Analysis—We performed a GWAS, using PLINK version 1.07⁴³, separately for African Americans and Hispanics/Latinas. We used linear regression for all analyses, modeled each SNP additively, and used the standard 5×10^{-8} as our threshold for statistical significance. After obtaining GWAS results, SNPs were clumped according to linkage disequilibrium (LD) to identify independent loci represented by a single best SNP⁴³. This clump procedure used the following thresholds to identify independent SNPs: (1) SNPs that had LD $r^2 \leq 0.25$; and (2) SNPs that were within 250 kb. We also analyzed SNPs on the X chromosome.

Both GWAS analyses (and the GWEIS, described below) adjusted for the following covariates, measured at baseline: age, income, education, marital status, and four principal components adjusting for population structure⁴⁰. These covariates were included because each was associated with depressive symptoms in either the SHARe or larger WHI cohort³², and prior studies have suggested inclusion of covariates in GWAS of common phenotypes may increase power⁴⁴. Quantile-quantile (QQ) and Manhattan plots were generated using R⁴⁵. Regional association plots were generated using Locus Zoom⁴⁶. Inverse variance weighted fixed-effect meta-analyses were conducted using METAL (<http://www.sph.umich.edu/csg/abecasis/metal/>;⁴⁷).

GWEIS Analysis—We performed the GWEIS using probABEL⁴⁸. Both stressful life events and social support were modeled separately using a categorical variable derived by taking quartiles of the total score distribution (0=first quartile; 1=second quartile; 2=third quartile; 3=fourth quartile). The lowest quartile group (0) indicated the lowest social-environmental risk group, whereas the highest quartile group (3) indicated the highest social-environmental risk group. We used quartiles to facilitate interpretation and address the skewed distribution of these variables; categorization (into four or more categories) does not result in the loss of information (and power) that occurs when continuous variables are dichotomized⁴⁹. We tested for G×E by including dummy variables for quartile group as well as a SNP by quartile-group (treated as ordinal) interaction term in the model. We used a Bonferroni correction to establish a significance threshold accounting for multiple testing of two environmental exposures ($\alpha = 2.5 \times 10^{-8}$). To reduce the likelihood of spurious G×E findings, we used model-robust estimates of standard errors (also known as sandwich standard errors)⁵⁰ in all tests of G×E. Robust variance estimates can reduce the possibility of inflated Type I errors found for G×E effects if the environmental main effect is misspecified or if there is departure from the presumed linear model^{51–53}. P-values corresponding to the interaction term (in the multiple regression model) were calculated in R based on a Wald chi-square test.

Replication

We sought replication of top GWAS findings ($p < 1 \times 10^{-6}$) in each sample using data from four independent cohorts (see Supplemental Materials); two cohorts (HRS and HCHS/SOL) were also used to replicate the GWEIS results. For the African American replication, we analyzed data from African American women in the Health and Retirement Study (HRS; $n = 1231$; mean age 62.09)^{54,55}, where depressive symptoms were measured using a 8-item

version of the CES-D, social support was measured through 3 items asking about support received from a spouse, children, family, and friends, and stressful life events were measured through a composite measure developed to most closely approximate the discovery analysis. We also analyzed data from African Americans in the Grady Trauma Project (GTP; $n=2010$ women ages 18–65)⁵⁶, where depressive symptoms were assessed using the Beck Depression Inventory⁵⁷. For the Hispanic/Latino replication, we analyzed data from the Hispanic Community Health Study/Study of Latinos (HCHS/SOL; $n=3371$ women ages 50–76), where depressive symptoms were measured by a 10-item CES-D, social support was measured through the 12-item version of the Interpersonal Support Evaluation List⁵⁸, and stressful life events measured through a composite measure designed to match the discovery sample. We also assessed top GWAS findings for both Hispanics and African Americans in the Cohorts for Heart and Aging Research in Genomic Epidemiology (CHARGE Consortium), which performed the largest meta-analysis GWAS of depressive symptoms to date using 17 European-ancestry population-based studies ($n=51,258$ individuals) of older adults where depressive symptoms were measured through the CES-D⁶.

Secondary Analyses

We performed five secondary analyses. First, we conducted two sets of meta analyses to determine the degree to which the top GWAS SNPs ($p<10^{-5}$) obtained in African Americans also showed evidence of nominal association in Hispanics/Latinas and vice versa. Second, we reran the GWAS in each sample after additionally adjusting for both environmental exposures, as both stressful life events and social support were found to make large and unique contributions to the variance in depressive symptoms. Third, we performed an analysis using genome-wide complex trait analysis (GCTA), which uses restricted maximum likelihood (REML) to obtain an estimation of the additive effect of common variants or “SNP-chip heritability”⁵⁹. We conducted these analyses, focusing on depressive symptoms, stressful life events, and social support separately, to evaluate the unique genetic contribution to these phenotypes and the potential presence of gene-environment correlation. We also examined the genetic contribution to depressive symptoms after adjusting for each of these environmental exposures individually. These analyses were performed only in African Americans, as a power calculation indicated the Hispanic/Latino sample would be underpowered to detect SNP heritability estimates in the range reported in previous studies of European Americans (ranging from 21%⁶⁰ to 32%⁶¹ for major depressive disorder; MDD). We also performed a bivariate REML analysis to determine the genetic correlation between depressive symptoms and these two environmental exposures⁶². Finally, to evaluate the strength of our findings given the skewed distribution of our outcome, we repeated the top GWAS ($p<1\times 10^{-5}$) and GWEIS ($p<1\times 10^{-6}$) tests of association using a non-parametric bootstrap. For the top GWAS SNPs, we fit a linear regression on 1000 bootstrap samples using the boot package in R^{63,64} and compared the effective sizes (betas) from the bootstrap samples to the betas obtained in the original analysis for each top SNP. For the top GWEIS SNPs, we fit linear regressions to 5000 datasets simulated under the null hypothesis⁶⁵ and generated p-values for each top SNP. These p-values represent the number of betas that were more extreme than the beta obtained in the original analysis divided by 5000 replicates. A significant p-value therefore indicates that the G×E interaction is significant at that level.

Results

Discovery Sample: GWAS

There were 7,179 African American and 3,138 Hispanic/Latina women in the analysis. See Supplemental Table 1 (Supplemental Materials) for sample demographic characteristics. Depressive symptoms scores were slightly higher in Hispanics/Latinas (mean=3.27; sd=3.20) and skewed towards lower values (skew=1.22; kurtosis=1.24), particularly in African Americans (mean=2.52; sd=2.71; skew=1.55; kurtosis=2.97). However, as linear regression is robust to minor violations of normality⁶⁶ and tests of G×E are sensitive to changing the scale of the phenotype⁶⁷, we did not perform any transformations.

Manhattan and QQ plots are shown in Supplemental Figure 1. As shown in the QQ plot, there was no evidence of inflation in either the African American ($\lambda = 1.004$, median=0.458) or Hispanic/Latina ($\lambda = 0.998$, median=0.455) GWAS.

No SNPs achieved genome-wide significance in either sample (Table 1). The peak signal in African Americans ($p = 5.75 \times 10^{-8}$) was for an imputed SNP rs73531535 located 20kb from *GPR139* (the G protein-coupled receptor 139), although several other SNPs in the region that also showed support were genotyped (Supplemental Figure 2). The second strongest association signal in African Americans was observed at rs75407252 ($p = 6.99 \times 10^{-7}$), an intron of *CACNA2D3*, the gene that encodes a protein in the voltage-dependent calcium channel subunit (Supplemental Figure 3).

In Hispanics/Latinas, the peak signal ($p = 2.44 \times 10^{-7}$) was for an imputed SNP rs2532087 located approximately 27kb away from *CD38* (Supplemental Figure 3). The second strongest association signal was for the imputed SNP rs4542757 ($p = 7.31 \times 10^{-7}$) located in an intron of *DCC* (deleted in colorectal cancer; Supplemental Figure 5). All GWAS results at $p < 1 \times 10^{-4}$ are shown for African Americans (Supplemental Table 2) and Hispanics/Latinas (Supplemental Table 3).

No SNPs achieved genome-wide significance on the X chromosome for either sample (Supplemental Table 4).

Replication Samples: GWAS

16 SNPs from the African American analysis and 18 SNPs from the Hispanic/Latina analysis with $p < 1 \times 10^{-5}$ were evaluated in four independent samples. For the African American replication (Table 2), one SNP was nominally significant in the HRS (rs418207; $p = 0.015$), though was not statistically associated after correction for multiple testing ($\alpha = 0.003$). This SNP showed the same direction and magnitude of effect in WHI and HRS and is intronic to *SRGAP3*, the gene that encodes the enzyme SLIT-ROBO Rho GTPase-activating protein 3.

In the Hispanic/Latina replication (Table 3), the peak WHI signal (rs2532087) also had the lowest p-value of the 18 SNPs in HCHS/SOL ($p = 0.00964$), though this result was not significantly associated after multiple testing correction ($\alpha = 0.003$). However, the direction and effect size were nearly identical in both the discovery and replication samples

(WHI beta=0.54; HCHS/SOL beta=0.56). The Hispanic/Latina discovery and HCHS/SOL replication results were also highly concordant, with 72% of linear regression beta coefficients (13 out of 18 SNPs) yielding the same direction of effect (sign test $p=0.05$).

None of the top GWAS findings in African Americans or Hispanics/Latinas were significantly associated with depressive symptoms in the CHARGE consortium of European Americans (refer to Supplemental Table 4).

Discovery Sample: GWEIS

Women in each sample reported a similar number of stressful life events (African American mean=2.15, sd=1.57; Hispanic/Latina mean=2.13, sd=1.68) and levels of social support (African American mean=35.29, sd=7.63; Hispanic/Latina mean=34.27, sd=8.92). The number of stressful life events and depressive symptoms were *positively* associated in both African Americans ($r^2=0.10$; $p<0.001$) and Hispanics/Latinas ($r^2=0.10$; $p<0.001$). Social support was *negatively* associated with depressive symptoms in both African Americans ($r^2=0.09$; $p<0.001$) and Hispanics/Latinas ($r^2=0.15$; $p<0.001$).

There was no evidence of genomic inflation for the African American stressful life events ($\lambda=0.99$) and social support analyses ($\lambda=1.02$) or the Hispanic/Latina stressful life events ($\lambda=1.01$) and social support analyses ($\lambda=1.03$) (Supplemental Figure 6 and 7).

One association signal was genome-wide significant (rs4652467; $p=4.10\times 10^{-10}$) in African Americans for the stressful life events GWEIS (Table 4). This SNP, located within 20kb of *CEP350*, was imputed, as were other SNPs in the region with $p<2.4\times 10^{-8}$ (Figure 1). The second strongest signal in African Americans was rs7275997 ($p=1.22\times 10^{-7}$), a genotyped intronic SNP located in *TMPRSS15* (transmembrane protease, serine 15; Supplemental Figure 8).

The GWEIS of social support in African Americans did not yield any genome-wide significant results (Table 4). The top two loci were rs77966298 ($p=2.43\times 10^{-7}$; Supplemental Figure 9) and rs6419121 ($p=3.98\times 10^{-7}$; Supplemental Figure 10).

In Hispanics/Latinas, we did not find any genome-wide significant association signals for either GWEIS (Table 5). The top two loci in the GWEIS of stressful life events were rs58707171 ($p=3.02\times 10^{-7}$) and rs6579218 ($p=4.94\times 10^{-7}$) (Supplemental Figure 11). The top two loci in the GWEIS of social support were rs35612712 ($p=3.42\times 10^{-7}$) and rs61973969 ($p=9.41\times 10^{-7}$) (Supplemental Figure 12).

Replication Samples: GWEIS

No top variants were significant in any replication sample (Table 6).

Secondary Analyses

The top loci in African Americans did not have similarly low p-values in Hispanics/Latinas and vice versa (see Supplemental Materials). Rerunning the GWAS after including the environmental exposures did not systematically change the results (see Supplemental Materials). SNP heritability estimates for depressive symptoms and the environmental

exposures were low (less than 10%) when each examined on their own and only significant for stressful life events, after adjusting for covariates (Table 7). The numerically largest and statistically significant estimate was found for stressful life events (8%). Interestingly, a very large genetic correlation was detected in the bivariate REML for depressive symptoms and stressful life events ($r_G=0.97$; $p=0.04$) after adjusting for covariates, suggesting that the genetic influences on depressive symptoms and stressful life events are largely shared. Indeed, after adjusting for each environmental measure in the REML analysis, no significant heritable signal for depressive symptoms remained. The GWAS and GWEIS results using a non-parametric bootstrap were similar to our original findings (see Supplemental Materials), suggesting our results were not sensitive to distributional assumptions.

Discussion

This study involved two major innovations in efforts to identify the genetic basis of depression. First, to our knowledge, this was the first genome-wide G×E analysis of depression. Prior G×E studies have focused on a relatively limited set of candidate gene polymorphisms, many of which have showed mixed results^{10,68}. Second, our study was also the largest GWAS of depressive symptoms conducted specifically in African Americans and Hispanics/Latinas. To our knowledge, only one prior GWAS was conducted among these groups; this study had a much smaller sample (African Americans $n=1603$; Hispanics $n=1443$) and did not examine G×E⁶⁹.

We highlight three findings. First, although no genome-wide significant loci were detected in our GWAS, three of the strongest signals were in genes previously implicated in depression-related phenotypes. In African Americans, our top SNP was located 20kb from *GPR139*. Recent studies show that GPR139 encodes a highly conserved G-protein coupled receptor whose ligands are tryptophan and phenylalanine⁷⁰. Expression of GPR139 appears to be restricted to the central system and evidence from mouse studies suggests that it is specifically expressed in the lateral habenula and septum, two regions previously implicated in the pathophysiology of depression⁷¹. Based on these results, Bonaventure and colleagues suggested that GRP139 may mediate the well-established depressogenic effects of tryptophan depletion⁷⁰. Our second best SNP in African Americans was located in a calcium channel gene (*CACNA2D3*). Variants in calcium channel signaling genes have been associated with MDD and other psychiatric disorders in large-scale genome-wide association studies^{72,73}. However, the *CACNA2D3* variant did not show evidence of association in either the GTP or HRS replication samples. In the analysis of Hispanics/Latinas, the second strongest signal was located in *DCC* (deleted in colorectal cancer), which encodes the netrin-1 receptor⁷⁴. *DCC* regulates transmembrane signaling receptor activity and is mutated or downregulated in colorectal cancer and esophageal carcinoma. Manitt, Nessler, and colleagues recently found *DCC* signaling aids in establishing medial prefrontal cortex dopamine synaptic connectivity and that higher expression of *DCC* may be linked to suicide⁷⁵. The *DCC* variant, however, was not associated with depressive symptoms in our replication sample. However, the *DCC* variant (as well as other top loci) showed similar directions of effect across the discovery and replication results, suggesting that our study may have been underpowered. Indeed, a post-hoc power calculation suggested we had poor power among the top results ($p<1\times 10^{-5}$) to detect the effect sizes observed

given our discovery sample sizes (African Americans=7,179; Hispanics=3,138). Specifically, the average power among the top SNPs was 0.26 in the African American GWAS and 0.23 in the Hispanic discovery GWAS. Thus, it appears that even larger samples sizes are needed to detect SNPs associated with depressive symptoms.

Second, in the African American sample, we observed a genome-wide significant interaction between rs4652467, a variant 14kb away from *CEP350*, and stressful life events. This interaction suggested depressive symptoms were highest among those with more exposure to stressful life events who also had more copies of the major allele. However, this G×E was not observed in the HRS replication. Whether this lack of replication indicates a spurious G×E result or is due to the differences in WHI and HRS phenotype definitions is unclear. To that end, only three of the six depressive symptoms assessed in WHI were also assessed in HRS; the stressful life events measures also had limited overlap (see Supplemental Materials for comparisons). The failure to identify more genome-wide significant G×E loci or replicate the one genome-wide significant finding may also be due to the small discovery sample size or smaller size of the HRS sample. Our discovery GWEIS analysis could have been underpowered, especially since G×E studies are known to require even larger samples than primary genetic association studies, perhaps as much as four times the size^{76,77}. However, a post-hoc power calculation we ran suggested our discovery GWEIS had high power (<90%) to detect the effect estimates we observed. This power estimate is likely inflated due to Winner's Curse (or the phenomena by which detected effects are larger than they really are)⁷⁸ and also does not take into account measurement error. Future studies are needed to identify optimal methods to estimate Winner's curse adjusted effect sizes for G×E interaction effects that also address measurement error.

Third, we were able to estimate the SNP heritability of depressive symptoms as well as the two social-environmental exposures in African Americans. SNP heritability estimates were low (less than 10%) for all three phenotypes. The SNP heritability for depressive symptoms (5%) was numerically the lowest and about one-quarter the size of estimates that have been observed in case-control studies of MDD with European-ancestry samples^{60,61}. SNP-chip heritability estimates of psychiatric and behavioral symptoms have been shown elsewhere^{79,80} to produce similarly lower heritability estimates than those obtained from studies examining disorders. Moreover, the largest and only statistically significant estimate observed was for stressful life events (8%), suggesting there may be some degree of gene-environment correlation. Our SNP heritability estimate for stressful life events was lower than a previous study, which found that SNPs explained 29% of the variance in stressful life events⁸¹. That study, however, was of European ancestry adults and focused on 6-month, rather than past year stressors and was drawn from a case-control sample of adults with recurrent MDD. Interestingly, we also found a very large genetic correlation for depressive symptoms with stressful life events ($r_G=0.95$), suggesting that common variation underlying depressive symptoms and stressful life event exposure, though modest on their own, were highly overlapping in this sample. This finding could be an artifact of the correlated nature of these variables when assessed in cross-sectional studies. Indeed, stressful life events ($r=0.32$) and social support ($r=0.30$) were modestly correlated with depressive symptoms, and thus these GCTA results could reflect shared genetic contribution to self-reported measures. Future studies are needed to replicate these findings and determine the impact of

this degree of gene-environment correlation (as well as environment-depression correlation) for studying G×E.

Another area for future research relates to whether and how to adjust for use of antidepressant medications in studies of depressive symptoms. In the current study, we followed the precedent set by the CHARGE consortium⁶, which conducted the largest meta-analysis of depressive symptoms to date, and used an algorithm to modify our depressive symptom score to account for medication use. By harmonizing our depressive symptoms phenotype to theirs, we aimed to facilitate future replication efforts and increase interpretation of results across individual studies. However, there are certainly many alternative approaches, such as conducting the GWAS and GWEIS analyses after excluding medication users, or accounting for medication use using alternative adjustment algorithms (of note, including antidepressant medication use would not have been appropriate, for reasons outlined in the Supplemental Materials). Simulation studies are needed to fully evaluate the strengths and drawbacks of alternative approaches. Such studies could evaluate the extent to which different conditions (e.g., the percentage of the sample taking medications, the shape of the distribution of the outcome, the average effect sizes for the efficacy of medications, and differences in the distribution of outcome by medication use) produce different GWAS and GWEIS effect estimates.

It is also worth noting that in the absence of a large sample, researchers can use several alternative approaches to GWEIS, including: (1) testing for G×E with replicable variants identified from GWAS, including the two loci observed in the CONVERGE study²; (2) pursuing two-stage genome-wide G×E⁸²; and (3) conducting gene pathway-by-environment interaction analyses⁸³ or polygenic risk score-by-environment interaction analyses^{84–86}.

Several limitations should be noted. First, the outcome was based on a brief inventory of depressive symptoms during the past week, rather than levels of depressive symptoms captured over a longer period of time. Thus, it is unclear how long these symptoms lasted. However, the CES-D has demonstrated excellent psychometric properties, including in predicting DSM-IV diagnoses^{33,34}, and its widespread use in epidemiological studies enabled us to conduct discovery and replication analyses. Future studies of trait or diagnostic measures of depressive symptoms in minority populations are needed. Second, the social-environmental exposures included in our G×E analyses were based on retrospective reporting and in the case of stressful life events, only captured the prior year. Thus, our study was not designed to capture whether genetic variation interacted with stressors experienced earlier in the lifespan. Prospective studies examining G×E at different stages of the lifespan are needed. Moreover, stressful life events and social support were assessed concurrently with depressive symptoms in the discovery sample as well as both replications. This may not be ideal, especially when studying the effects of stress, as prior work suggests the odds of depression is greatest in the same month of the stressor⁸⁷. Longitudinal, prospective studies measuring social-environmental exposures antecedent to and close in time to depressive symptoms are necessary. These study designs are particularly important, as prior work suggests support for the 5-HTTLPR G×E, for example, is more consistent when structured interviews of stressful life events are used instead of self-report questionnaires^{88,89}. Finally, our replication samples were smaller and more phenotypically heterogeneous than the

discovery sample. For example, the WHI and HRS samples were of older adults, GTP comprised mostly middle-aged adults, and HCHS/SOL comprised a broader age range. The phenotypes also varied across these samples. Unfortunately, these limitations reflect the state of the field. Harmonizing data for GWAS and G×E analyses on a large scale in racial/ethnic minority populations is challenging. Whether our failure to replicate reflects Type I error in the discovery sample or Type II error in the replication is unknown. By undertaking these analyses, we hope to spark more large-scale epidemiological studies to incorporate such measures and to study the genetic determinants of depression in women, who are more burdened by the disorder than men.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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References

1. Sullivan PF, Neale MC, Kendler KS. Genetic epidemiology of major depression: Review and meta analysis. *American Journal of Psychiatry*. 2000; 157:1552–1562. [PubMed: 11007705]
2. CONVERGE Consortium. Sparse whole-genome sequencing identifies two loci for major depressive disorder. *Nature*. 2015; 523(7562):588–591. [PubMed: 26176920]
3. Ripke S, Wray NR, et al. Major Depressive Disorder Working Group of the Psychiatric GWAS Consortium. A mega-analysis of genome-wide association studies for major depressive disorder. *Molecular Psychiatry*. 2013; 18(4):497–511. [PubMed: 22472876]
4. Levinson DF, Mostafavi S, Milaneschi Y, et al. Genetic studies of major depressive disorder: why are there no genome-wide association study findings and what can we do about it? *Biological Psychiatry*. 2014; 76(7):510–512. [PubMed: 25201436]
5. Flint J, Kendler KS. The genetics of major depression. *Neuron*. 2014; 81:484–503. [PubMed: 24507187]
6. Hek K, Demirkan A, Lahti J, et al. A genome-wide association study of depressive symptoms. *Biological Psychiatry*. 2013; 73(7):667–678. [PubMed: 23290196]
7. Terracciano A, Tanaka T, Sutin AR, et al. Genome-wide association scan of trait depression. *Biol Psychiatry*. 2010; 68(9):811–817. [PubMed: 20800221]

8. Kendler KS, Gardner CO. Boundaries of major depression: An evaluation of DSM-IV criteria. *American Journal of Psychiatry*. 1998; 155:172–177. [PubMed: 9464194]
9. Yang J, Wray NR, Visscher PM. Comparing apples and oranges: equating the power of case-control and quantitative trait association studies. *Genetic Epidemiology*. 2010; 34:254–257. [PubMed: 19918758]
10. Dunn EC, Brown RC, Dai Y, et al. Genetic determinants of depression: recent findings and future directions. *Harv Rev Psychiatry*. 2015; 23(1):1–18. [PubMed: 25563565]
11. Caspi A, Sugden K, Moffitt TE, et al. Influence of life stress on depression: Moderation by a polymorphism in the 5-HTT gene. *Science*. 2003; 301:386–389. [PubMed: 12869766]
12. Dunn EC, Uddin M, Subramanian SV, Smoller JW, Galea S, Koenen KC. Gene-environment interaction (G×E) research in youth depression: A systematic review with recommendations for future research. *Journal of Child Psychology and Psychiatry*. 2011; 52(12):1223–1238. [PubMed: 21954964]
13. Risch N, Herrell R, Lehner T, et al. Interaction between the serotonin transporter gene (5-HTTLPR), stressful life events, and risk of depression: A meta analysis. *Journal of the American Medical Association*. 2009; 301(23):2462–2471. [PubMed: 19531786]
14. Munafò MR, Durrant C, Lewis G, Flint J. Gene X environment interactions at the serotonin transporter locus. *Biological Psychiatry*. 2009; 65:211–219. [PubMed: 18691701]
15. Karg K, Burmeister M, Shedden K, Sen S. The serotonin transporter promoter variant (5-HTTLPR), stress, and depression meta analysis revisited: Evidence of genetic moderation. *Archives of General Psychiatry*. 2011; 68(5):444–454. [PubMed: 21199959]
16. Duncan LE, Pollastri AR, Smoller JW. Mind the gap: Why many geneticists and psychological scientists have discrepant views about gene-environment interaction (G×E) research. *American Psychologist*. 2014; 69(3):249–268. [PubMed: 24750075]
17. van Os J, Rutten BP. Gene-environment-wide interaction studies in psychiatry. *Am J Psychiatry*. 2009; 166(9):964–966. [PubMed: 19723791]
18. Winham SJ, Biernacka JM. Gene-environment interactions in genome-wide association studies: Current approaches and new directions. *Journal of Child Psychology and Psychiatry*. 2013
19. Wu C, Kraft P, Zhai K, et al. Genome-wide association analyses of esophageal squamous cell carcinoma in Chinese identify multiple susceptibility loci and gene-environment interactions. *Nature Genetics*. 2012; 44(10):1090–1097. [PubMed: 22960999]
20. Seigert S, Hampe J, Schafmayer C, et al. Genome-wide investigation of gene-environment interactions in colorectal cancer. *Human Genetics*. 2013; 132(2):219–231. [PubMed: 23114982]
21. Manning AK, Hivert MF, Scott RA, et al. A genome-wide approach accounting for body mass index identifies genetic variants influencing fasting glycemic traits and insulin resistance. *Nature Genetics*. 2012; 44(6):659–669
22. Whitfield KE, Edwards CL, Brandon D, McDougald C. Genetic and environmental influences on depressive symptoms by age and gender in African American twins. *Aging Ment Health*. 2008; 12(2):221–227. [PubMed: 18389402]
23. Duncan AE, Munn-Chernoff MA, Hudson DL, et al. Genetic and environmental risk for major depression in African-American and European-American women. *Twin Res Hum Genet*. 2014; 17(4):244–253. [PubMed: 24910290]
24. Olvera RL, Bearden CE, Velligan DI, et al. Common genetic influences on depression, alcohol, and substance use disorders in Mexican-American families. *Am J Med Genet B Neuropsychiatr Genet*. 2011; 156B(5):561–568. [PubMed: 21557468]
25. Kessler RC, Berglund P, Demler O, Jin R, Koretz D, et al. The epidemiology of major depressive disorder: Results from the National Comorbidity Survey Replication (NCS-R). *Journal of the American Medical Association*. 2003; 289:3095–3015. [PubMed: 12813115]
26. Hatch SL, Dohrenwend BP. Distribution of traumatic and other stressful life events by race/ethnicity, gender, SES, age: A review of the research. *American Journal of Community Psychology*. 2007; 40:313–332. [PubMed: 17906927]
27. Williams DR, Mohammed SA. Discrimination and racial disparities in health: Evidence and needed research. *Journal of Behavioral Medicine*. 2009; 32:20–47. [PubMed: 19030981]

28. Mays VM, Cochran SD, Barnes NW. Race, race-based discrimination, and health outcomes among African Americans. *Annual Review of Psychology*. 2007; 58:201–225.
29. Crimmins, EM.; Hayward, MD.; Seeman, TE. Race/Ethnicity, Socioeconomic Status, and Health. Anderson, NB.; Bulatao, RA.; Cohen, B., editors. Washington, DC: The National Academies Press; 2004. Critical Perspectives on Racial and Ethnic Differences in Health in Late Life
30. Kessler RC, Berglund P, Demler O, Jin R, Merikangas KR, Walters EE. Lifetime prevalence and age-of-onset distributions of DSM-IV disorders in the National Comorbidity Survey Replication. *Archives of General Psychiatry*. 2005; 62(6):593–602. [PubMed: 15939837]
31. Design of the Women's Health Initiative clinical trial and observational study. The Women's Health Initiative Study Group. *Control and Clinical Trials*. 1998; 19:61–109.
32. Wassertheil-Smoller S, Shumaker S, Ockene J, et al. Depression and cardiovascular sequela in postmenopausal women: The Women's Health Initiative (WHI). *Archives of Internal Medicine*. 2004; 164:289–298. [PubMed: 14769624]
33. Radloff LS. The CES-D scale: A self-report depression scale for research in the general population. *Applied Psychological Measurement*. 1977; 1:385–401.
34. Irwin M, Artin KH, Oxman MN. Screening for depression in the older adult: criterion validity of the 10-item Center for Epidemiological Studies Depression Scale (CES-D). *Arch Intern Med*. 1999; 159(15):1701–1704. [PubMed: 10448771]
35. Levy D, DeStefano AL, Larson MG, et al. Evidence for a gene influencing blood pressure on chromosome 17: Genome scan linkage results for longitudinal blood pressure phenotypes in subjects from the Framingham Heart Study. *Hypertension*. 2000; 36(4):477–483. [PubMed: 11040222]
36. Wade TD, Kendler KS. The relationship between social support and major depression: cross-sectional, longitudinal, and genetic perspectives. *J Nerv Ment Dis*. 2000; 188(5):251–258. [PubMed: 10830561]
37. Berkman LF, Syme SL. Social networks, host resistance, and mortality: a nine-year follow-up study of Alameda County residents. *Am J Epidemiol*. 1979; 109(2):186–204. [PubMed: 425958]
38. Smoller JW, Pollack MH, Wassertheil-Smoller S, et al. Panic attacks, daily life ischemia, and chest pain in postmenopausal women. *Psychosom Med*. 2006; 68(6):824–832. [PubMed: 17101813]
39. Sherbourne CD, Stewart AL. The MOS social support survey. *Soc Sci Med*. 1991; 32(6):705–714. [PubMed: 2035047]
40. WOMEN'S HEALTH INITIATIVE SHARE PROJECT. Imputation Report- 1000 Genomes Project reference panel. 2011 https://www.garnetstudy.org/sites/www/content/files/dataflowcleaning/WHI_SHARe_qc_report_1000G_final.pdf
41. Browning B, Browning SA. A unified approach to genotype imputation and haplotype-phase inference for large data sets of trios and unrelated individuals. *American Journal of Human Genetics*. 2009; 84:210–223. [PubMed: 19200528]
42. Price AL, Patterson NJ, Plenge RM, Weinblatt ME, Shadick NA, Reich D. Principal components analysis corrects for stratification in genome-wide association studies. *Nature Genetics*. 2006; 38:904–909. [PubMed: 16862161]
43. Purcell S, Neale B, Todd-Brown K, et al. PLINK: a toolset for whole-genome association and population-based linkage analysis. *American Journal of Human Genetics*. 2007; 81(3):559–575. [PubMed: 17701901]
44. Pirinen M, Donnelly P, Spencer CCA. Including known covariates can reduce power to detect genetic effects in case-control studies. *Nature Genetics*. 2012; 44(8):848–851. [PubMed: 22820511]
45. R Development Core Team. R: A language and environment for statistical computing. Vienna, Austria: R Foundation for Statistical Computing; 2011.
46. Pruim RJ, Welch RP, Sanna S, et al. LocusZoom: Regional visualization of genome-wide association scan results. *Bioinformatics*. 2010; 26(18):2336–2337. [PubMed: 20634204]
47. Willer CJ, Li Y, Abecasis GR. METAL: fast and efficient meta-analysis of genomewide association scans. *Bioinformatics*. 2010; 26(17):2190–2191. [PubMed: 20616382]
48. Aulchenko YS, Struchalin MV, vanDuijn CM. probABEL package for genome-wide association analysis of imputed data. *BMC Bioinformatics*. 2010; 11:134. [PubMed: 20233392]

49. Altman DG, Royston P. The cost of dichotomising continuous variables. *BMJ*. 2006; 332(7549): 1080. [PubMed: 16675816]
50. White H. A heteroskedasticity-consistent covariance-matrix estimator and a direct test for heteroskedasticity. *Econometrica*. 1980; 48:817–838.
51. Voorman A, Lumley T, McKnight B, Rice K. Behavior of qq-plots and genomic control in studies of gene-environment interaction. *Plos One*. 2011; 65(5):e19416. [PubMed: 21589913]
52. Cornelis MC, Tchetgen EJ, Liang L, et al. Gene-environment interactions in genome-wide association studies: A comparative study of tests applied to empirical studies of type 2 diabetes. *American Journal of Epidemiology*. 2012; 175(3):191–202. [PubMed: 22199026]
53. Almli LM, Duncan R, Feng H, et al. Correcting systematic inflation in genetic association tests that consider interaction effects: application to a genome-wide association study of posttraumatic stress disorder. *JAMA Psychiatry*. 2014; 71(12):1392–1399. [PubMed: 25354142]
54. Juster FT, Suzman R. An overview of the health and retirement study. *Journal of Human Resources*. 1995; 30
55. Sonnega A, Faul JD, Ofstedal MB, Langa KM, Phillips JWR, Weir DR. Cohort profile: the health and retirement study (HRS). *International Journal of Epidemiology*. 2014; 43(2):576–585. [PubMed: 24671021]
56. Gillespie CF, Bradley B, Mercer K, et al. Trauma exposure and stress-related disorders in inner city primary care patients. *Gen Hosp Psychiatry*. 2009; 31(6):505–514. [PubMed: 19892208]
57. Beck AT, Ward CH, Mendelson M, Mock J, Erbaugh J. An inventory for measuring depression. *Archives of General Psychiatry*. 1961; 4:561–571. [PubMed: 13688369]
58. Cohen, S.; Mermelstein, R.; Kamarck, T.; Hoberman, H. Measuring the functional components of social support. In: Sarason, IG.; Sarason, B., editors. *Social support: Theory, research, and application*. Netherlands: Springer; 1985. p. 73-94.
59. Yang J, Lee SH, Goddard ME, Visscher PM. GCTA: A tool for genome-wide complex trait analysis. *American Journal of Human Genetics*. 2011; 88:76–82. [PubMed: 21167468]
60. Lee SH, Ripke S, Neale BM, et al. Genetic relationship between five psychiatric disorders estimated from genome-wide SNPs. *Nat Genet*. 2013; 45(9):984–994. [PubMed: 23933821]
61. Lubke GH, Hottenga JJ, Walters R, et al. Estimating the genetic variance of major depressive disorder due to all single nucleotide polymorphisms. *Biological Psychiatry*. 2012; 72(8):707–709. [PubMed: 22520966]
62. Lee SH, Yang J, Goddard ME, Visscher PM, Wray NR. Estimation of pleiotropy between complex diseases using single-nucleotide polymorphism-derived genomic relationships and restricted maximum likelihood. *Bioinformatics*. 2012; 28(19):2540–2542. [PubMed: 22843982]
63. Canty A, Ripley B. boot: Bootstrap R (S-Plus) Functions. R package version 1.3-17. 2015
64. Davison, AC.; Hinkley, DV. *Bootstrap Methods and Their Applications*. Cambridge: Cambridge University Press; 1997.
65. Buzkova P, Lumley T, Rice K. Permutation and parametric bootstrap tests for gene-gene and gene-environment interactions. *Ann Hum Genet*. 2011; 75(1):36–45. [PubMed: 20384625]
66. van Belle, G. *STRUTS: Statistical rules of thumb*. New York, NY: John Wiley and Sons; 2002.
67. Kraft, P.; Hunter, DJ.; Khoury, MJ.; Bedrosian, SR.; Gwinn, M.; Higgins, JPT.; Ioannidis, JPA.; Little, J., editors. *The challenge of assessing complex gene-environment and gene-gene interactions*. Second. New York, NY: Oxford University Press; 2009. *Human genome epidemiology: Building the evidence for using genetic information to improve health and prevent disease*
68. Duncan LE, Keller MC. A critical review of the first 10 years of candidate gene-by-environment interaction research in psychiatry. *American Journal of Psychiatry*. 2011; 168(10):1041–1049. [PubMed: 21890791]
69. Ware EB, Mukherjee B, Sun YV, Diez-Roux AV, Kardina SL, Smith JA. Comparative genome-wide association studies of a depressive symptom phenotype in a repeated measures setting by race/ethnicity in the Multi-Ethnic Study of Atherosclerosis. *BMC Genet*. 2015; 16:118. [PubMed: 26459564]
70. Bonaventure, P.; Liu, C.; Lee, G., et al. GPR139, an Orphan Receptor Highly Enriched in the Habenula and Septum, is Activated by the Essential Amino Acids L-Tryptophan and L-

Phenylalanine. Paper presented at: American College of Neuropsychopharmacology; Hollywood, Florida. 2015.

71. Li K, Zhou T, Liao L, et al. betaCaMKII in lateral habenula mediates core symptoms of depression. *Science*. 2013; 341(6149):1016–1020. [PubMed: 23990563]
72. Cross-Disorder Group of the Psychiatric Genomics Consortium. Identification of risk loci with shared effects on five major psychiatric disorders: A genome-wide analysis. *Lancet*. 2013; 381(9875):1371–1379. [PubMed: 23453885]
73. Schizophrenia Working Group of the Psychiatric Genomics C. Biological insights from 108 schizophrenia-associated genetic loci. *Nature*. 2014; 511(7510):421–427. [PubMed: 25056061]
74. Manitt C, Labelle-Dumais C, Eng C, et al. Peri-pubertal emergence of UNC-5 homologue expression by dopamine neurons in rodents. *PLoS One*. 2010; 5(7):e11463. [PubMed: 20628609]
75. Manitt C, Eng C, Pokinko M, et al. dcc orchestrates the development of the prefrontal cortex during adolescence and is altered in psychiatric patients. *Transational Psychiatry*. 2013; 3:e338.
76. Luan JA, Wong MY, Day NE, Wareham NJ. Sample size determination for studies of gene-environment interaction. *Int J Epidemiol*. 2001; 30(5):1035–1040. [PubMed: 11689518]
77. Thomas D. Methods for investigating gene-environment interactions in candidate pathway and genome-wide association studies. *Annual Review of Public Health*. 2010; 31:21–36.
78. Hardy J, Low NC. Genes and environment in psychiatry: Winner's curse or cure? *Archives of General Psychiatry*. 2011; 68(5):455–456. [PubMed: 21536973]
79. Trzaskowski M, Dale PS, Plomin R. No genetic influence for childhood behavior problems from DNA analysis. *J Am Acad Child Adolesc Psychiatry*. 2013; 52(10):1048–1056. e1043. [PubMed: 24074471]
80. Sieradzka D, Power RA, Freeman D, Cardno AG, Dudbridge F, Ronald A. Heritability of Individual Psychotic Experiences Captured by Common Genetic Variants in a Community Sample of Adolescents. *Behav Genet*. 2015; 45(5):493–502. [PubMed: 26049723]
81. Power RA, Wingenbach T, Cohen-Woods S, et al. Estimating the heritability of reporting stressful life events captured by common genetic variants. *Psychological Medicine*. 2013; 43(9):1965–1971. [PubMed: 23237013]
82. Gauderman WJ, Zhang P, Morrison JL, Lewinger JP. Finding novel genes by testing G×E interactions in a genome-wide association study. *Genetic Epidemiology*. 2013; 37:603–613. [PubMed: 23873611]
83. Sharafeldin N, Slaterry ML, Liu Q, et al. A Candidate-Pathway Approach to Identify Gene-Environment Interactions: Analyses of Colon Cancer Risk and Survival. *J Natl Cancer Inst*. 2015; 107(9)
84. Musliner KL, Seifuddin F, Judy JA, Pirooznia M, Goes FS, Zandi PP. Polygenic risk, stressful life events and depressive symptoms in older adults: a polygenic score analysis. *Psychol Med*. 2015; 45(8):1709–1720. [PubMed: 25488392]
85. Peyrot WJ, Milaneschi Y, Abdellaoui A, et al. Effect of polygenic risk scores on depression in childhood trauma. *The British journal of psychiatry : the journal of mental science*. 2014; 205(2): 113–119. [PubMed: 24925986]
86. Mullins N, Power RA, Fisher HL, et al. Polygenic interactions with environmental adversity in the aetiology of major depressive disorder. *Psychol Med*. 2015:1–12.
87. Kendler KS, Karkowski LM, Prescott CA. Causal relationship between stressful life events and the onset of major depression. *Am J Psychiatry*. 1999; 156(6):837–841. [PubMed: 10360120]
88. Uher R, McGuffin P. The moderation by the serotonin transporter gene of environmental adversity in the aetiology of mental illness: Review and methodological analysis. *Molecular Psychiatry*. 2008; 13:131–146. [PubMed: 17700575]
89. Caspi A, Hariri AR, Holmes A, Uher R, Moffitt TE. Genetic sensitivity to the environment: The case of the serotonin transporter gene and its implications for studying complex diseases and traits. *American Journal of Psychiatry*. 2010; 167(5):1–19. [PubMed: 20068118]

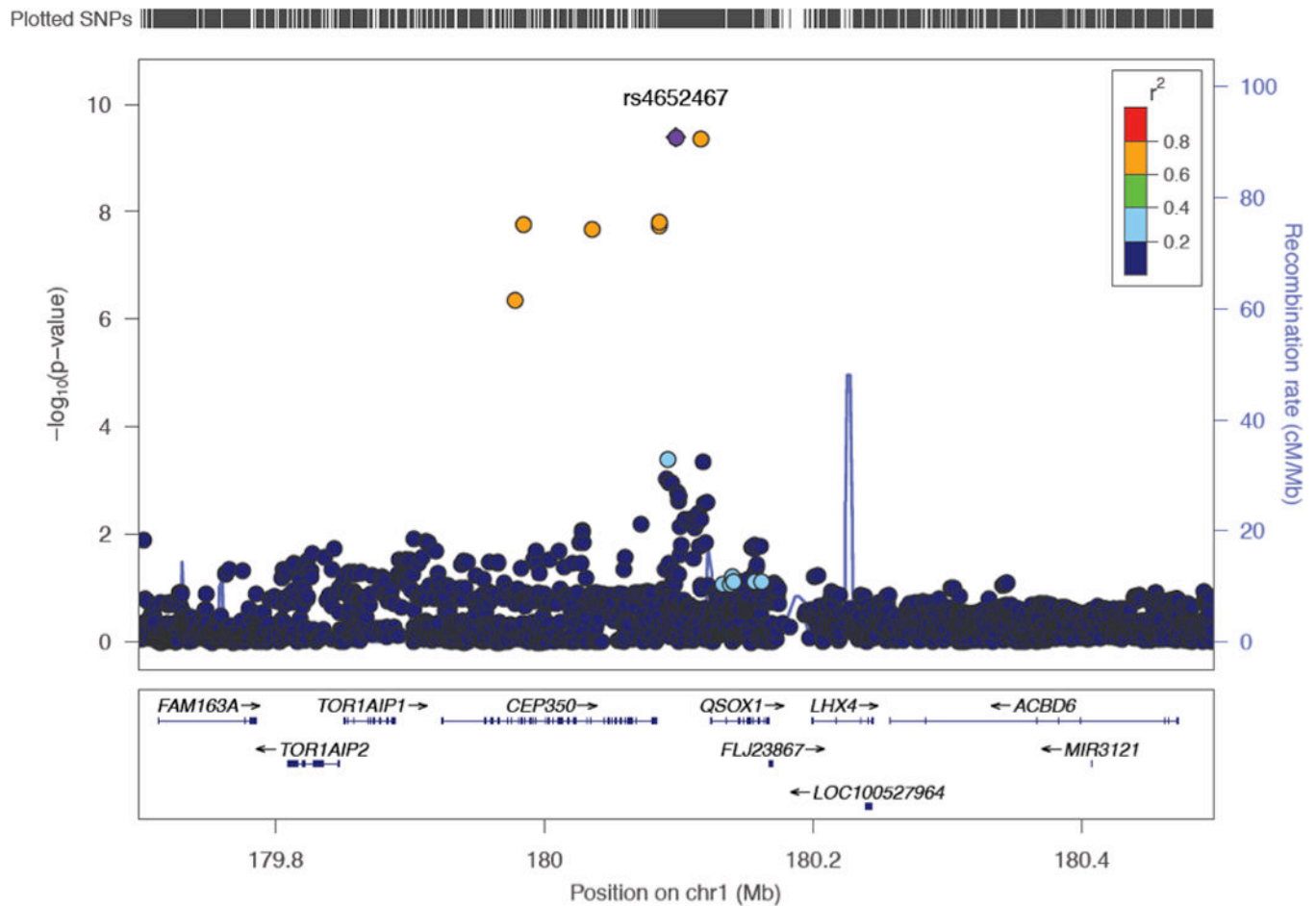


Figure 1.

Regional association plot for the top SNP (rs4652467) identified in the African American genome-wide environment interaction study (GWEIS) of stressful life events. The regional association plot was generated using LocusZoom (<http://csg.sph.umich.edu/locuszoom/>). We present results for the African American reference panel only as the SNP was monomorphic in Europeans (100% G allele). The left-side y-axis refers to the log of the p-value corresponding to the test of association between each SNP (denoted as a colored dot) and stressful life events (in the test of $G \times E$) and levels of depressive symptoms. SNPs are colored based on the level of linkage disequilibrium (LD) between each SNP and the index SNP. r^2 values are determined based on the HG19/1000 Genomes (March 2012 build) data. The index SNP (rs4652467, purple diamond) and its closest neighbors (shown in orange) are imputed.

Table 1

Genome-wide association study (GWAS) results for the top loci ($p < 1 \times 10^{-5}$) in African Americans and Hispanics/Latinos

SNP	chr	position	A1	A2	MAF	G/I	Info	Beta	SE	P-value	Location	Closest Gene (<20kb)
African Americans												
rs73531535	16	20105038	C	T	0.229	I	0.938	-0.297	0.055	5.75E-08		<i>GPR139</i>
rs75407252	3	54241886	C	T	0.053	I	0.813	-0.548	0.110	6.99E-07	intron variant	<i>CACNA2D3</i>
rs11233283	11	82415904	A	G	0.016	I	0.986	-0.298	0.062	1.41E-06		
rs34257140	17	42675053	G	T	0.149	I	0.999	-0.313	0.065	1.59E-06		
rs580112	3	177289895	A	G	0.184	I	0.945	0.279	0.060	2.84E-06	intron variant	<i>LINC00578</i>
rs1413154	13	83240729	G	T	0.205	I	0.825	-0.283	0.061	3.54E-06		
rs1893586	21	43286918	A	G	0.483	I	0.947	0.211	0.046	4.19E-06	intron variant	<i>PRDM15</i>
rs10777901	12	98492992	A	C	0.481	I	0.993	0.206	0.045	4.27E-06		
rs10125319	9	133426729	C	T	0.491	I	0.920	-0.214	0.047	4.27E-06		
rs10221121	16	56840328	A	G	0.229	G	0.986	-0.241	0.053	5.98E-06	intron variant	<i>NUP93</i>
rs210329	14	54059800	G	T	0.332	I	0.989	0.217	0.048	6.35E-06		<i>RPS3AP46</i>
rs28493952	3	95747804	C	T	0.320	I	0.850	0.234	0.052	6.41E-06		
rs7312307	12	106441333	C	G	0.087	I	0.904	0.376	0.084	6.92E-06		<i>NUAK1</i>
rs17030391	2	43353504	A	G	0.143	I	0.808	0.316	0.071	7.84E-06		
rs48666976	5	45579793	A	G	0.086	G	0.996	-0.361	0.081	8.04E-06	intron variant	<i>HCN1</i>
rs418207	3	9225376	A	G	0.477	I	0.811	-0.219	0.050	9.59E-06	intron variant	<i>SRGAP3</i>
Hispanics/Latinos												
rs2532087	4	15878327	C	G	0.231	I	0.8154	0.5379	0.104	2.44E-07		
rs4542757	18	50198724	C	T	0.418	I	0.9304	-0.4135	0.0833	7.31E-07	intron variant	<i>DCC</i>
rs10249677	7	50650831	G	T	0.042	I	0.8816	1.0497	0.2157	1.20E-06		<i>GRB10</i>
rs1129411	2	231077725	A	G	0.085	I	0.9941	0.6637	0.1417	2.94E-06	missense/intron variant	<i>SPI10</i>
rs11738766	5	8214282	A	G	0.279	G	1.0253	-0.4134	0.0885	3.11E-06		
rs34359572	1	194036781	A	G	0.072	I	0.8899	-0.7471	0.1619	4.10E-06		
rs609508	20	54167720	C	G	0.214	I	0.9752	0.4429	0.097	5.21E-06		
rs16823787	2	183692791	A	G	0.084	G	0.9823	0.6591	0.1444	5.21E-06		<i>FRZB</i>
rs17345417	4	95948486	A	G	0.111	G	0.9949	-0.574	0.1259	5.30E-06	intron variant	<i>BMPR1B</i>

SNP	chr	position	A1	A2	MAF	G/I	Info	Beta	SE	P-value	Location	Closest Gene (<20kb)
rs2822657	21	15774729	C	T	0.457	I	0.9875	-0.3657	0.0802	5.35E-06		<i>HSPA13</i>
rs13033587	2	52857818	C	T	0.475	I	0.9342	0.376	0.0828	5.85E-06		
rs9601962	13	83312889	G	T	0.185	I	0.8544	0.4945	0.1098	6.91E-06		
rs2282123	6	89907561	C	G	0.255	G	1.0034	0.4114	0.0915	7.19E-06	intron variant	<i>GABRR1</i>
rs10886733	10	122402887	C	T	0.117	I	0.9766	0.556	0.1237	7.23E-06		<i>MIR5694</i>
rs61848143	10	24746704	C	G	0.177	I	0.8403	-0.5023	0.1122	7.85E-06	intron variant	<i>KIAA1217</i>
rs10166852	2	183450923	C	G	0.474	I	0.9666	-0.3752	0.0838	7.85E-06		
rs6736484	2	45146524	G	T	0.075	G	0.8556	-0.7206	0.1621	9.05E-06		
rs2912513	8	69968166	A	T	0.033	I	0.9774	-0.996	0.2243	9.33E-06	intron variant	<i>LINC01592</i>

The table lists all LD-pruned SNPs associated with depressive symptoms at $p < 1 \times 10^{-5}$. A1 is the tested allele using an additive model, where alleles were analyzed as dosages. The closest gene within 20kb upstream/downstream of the SNP is provided. All SNPs are on the positive (5' to 3') strand.

Chr: chromosome; position: base pair position, G/I: genotyped or imputed.

Table 2

Replication of genome-wide association study (GWAS) results for the top loci ($p<1\times10^{-5}$) in African Americans

SNP	chr	position	WHI			HRS			GTP			Discovery: WHI (n=7179)			Replication: HRS (n=1231)			Replication: GTP (n=2010)		
			A1	A2	G/I	info	MAF	A1	A2	G/I	info	MAF	A1	A2	Beta	SE	p-value	Beta	SE	p-value
rs73531535	16	20105038	C	T	I	0.995	0.245	C	T	I	0.973	0.246	T	C	-0.297	0.055	5.75E-08	0.060	0.106	0.573
rs75407252	3	54241886	C	T	I	0.922	0.051	C	T	I	0.868	0.048	T	C	-0.548	0.110	6.99E-07	0.285	0.208	0.171
rs11233283	11	82415904	A	G	I	0.999	0.192	A	G	I	0.968	0.185	A	G	-0.298	0.062	1.41E-06	0.194	0.117	0.098
rs34257140	17	42675053	G	T	G	1	0.102	G	T	I	0.971	0.106	T	G	-0.313	0.065	1.59E-06	0.112	0.152	0.461
rs580112	3	177289895	A	G	I	0.999	0.158	A	G	I	0.985	0.152	A	G	0.279	0.060	2.84E-06	-0.130	0.123	0.290
rs1413154	13	83240729	G	T	G	1	0.219	G	T	G	0.987	0.223	T	G	-0.283	0.061	3.54E-06	0.141	0.105	0.181
rs1893586	21	43286918	A	G	I	0.999	0.462	A	G	I	0.993	0.478	A	G	0.211	0.046	4.19E-06	-0.046	0.094	0.626
rs10777901	12	98492992	A	C	G	1	0.498	A	C	G	1.018	0.476	A	C	0.206	0.045	4.27E-06	0.155	0.090	0.085
rs10125319	9	133426729	C	T	I	0.996	0.495	C	T	I	1.008	0.487	T	C	-0.214	0.047	4.27E-06	-0.039	0.089	0.664
rs10221121	16	56840328	A	G	G	1	0.230	A	G	G	1.001	0.210	A	G	-0.241	0.053	5.98E-06	-0.020	0.109	0.857
rs210329	14	54059800	G	T	I	0.996	0.303	G	T	I	0.970	0.283	T	G	0.217	0.048	6.35E-06	-0.018	0.101	0.856
rs28493952	3	95747804	C	T	I	0.979	0.291	C	T	I	0.936	0.295	T	C	0.234	0.052	6.41E-06	0.065	0.104	0.530
rs7312307	12	106441333	C	G	I	0.982	0.090	C	G	I	0.927	0.091	C	G	0.376	0.084	6.92E-06	0.124	0.162	0.447
rs17030391	2	43353504	A	G	I	0.998	0.122	A	G	I	0.961	0.132	A	G	0.316	0.071	7.84E-06	-0.009	0.143	0.947
rs4866976	5	45579793	A	G	I	0.989	0.076	A	G	I	0.943	0.082	A	G	-0.361	0.081	8.04E-06	0.205	0.171	0.230
rs418207	3	9225376	A	G	G	1	0.486	A	G	G	1.029	0.494	A	G	-0.219	0.050	9.59E-06	-0.228	0.093	0.015

HRS = Health and Retirement Study. These models were estimated using R 3.0.1. Covariates in HRS were age, income, education, marital status and the top 10 principal components. Imputation was conducted using IMPUTE2. GTP = Grady Trauma Project. In these analyses with dosage data, PLINK models A1 as the tested allele. SNPs were analyzed using additive coding, where alleles were analyzed as dosages. Covariates in the GTP were age, income per month, education, marital status, and 5 principal components. In the GTP, quality control and imputation were performed by the PGC Statistical Analysis Group. Methods for imputation are described in the Supplemental Materials. The Bonferroni adjusted alpha level in these analyses was 0.05/16=0.003.

Table 3

Replication of genome-wide association study (GWAS) results for the top loci ($p < 1 \times 10^{-5}$) in Hispanics

			WHI		HCHS/SOL (n=3371)				Discovery: WHI (n=3138)				Replication: HCHS/SOL			
SNP	chr	position	A1	A2	G/I	info	MAF	A1	A2	Beta	SE	p-value	Beta	SE	p-value	
rs2532087	4	15878327	C	G	I	0.926	0.210	C	G	0.538	0.104	2.44E-07	0.556	0.215	0.00964	
rs4542757	18	50198724	C	T	G	1.000	0.415	C	T	-0.413	0.083	7.31E-07	-0.072	0.174	0.68	
rs10249677	7	50650831	G	T	I	0.979	0.065	T	G	1.050	0.216	1.20E-06	-0.582	0.354	0.1	
rs1129411	2	231077725	A	G	G	1.000	0.084	A	G	0.664	0.142	2.94E-06	0.262	0.305	0.39	
rs11738766	5	8214282	A	G	G	1.000	0.286	A	G	-0.413	0.089	3.11E-06	-0.277	0.191	0.147	
rs34359572	1	194036781	A	G	I	0.999	0.083	A	G	-0.747	0.162	4.10E-06	-0.019	0.311	0.951	
rs609508	20	54167720	C	G	I	0.998	0.215	C	G	0.443	0.097	5.21E-06	0.069	0.208	0.738	
rs16823787	2	183692791	A	G	I	0.987	0.091	A	G	0.659	0.144	5.21E-06	-0.099	0.297	0.739	
rs17345417	4	95948486	A	G	I	0.996	0.102	A	G	-0.574	0.126	5.30E-06	0.194	0.284	0.495	
rs2822657	21	15774729	C	T	G	1.000	0.436	T	C	-0.366	0.080	5.35E-06	-0.098	0.171	0.569	
rs13033587	2	52857818	C	T	I	0.997	0.489	T	C	0.376	0.083	5.85E-06	0.298	0.174	0.0869	
rs9601962	13	83312889	G	T	I	0.973	0.210	T	G	0.494	0.110	6.91E-06	0.098	0.211	0.644	
rs2282123	6	89907561	C	G	I	0.995	0.238	C	G	0.411	0.092	7.19E-06	0.049	0.202	0.808	
rs10886733	10	122402887	C	T	I	0.989	0.106	T	C	0.556	0.124	7.23E-06	0.054	0.272	0.843	
rs61848143	10	24746704	C	G	I	0.985	0.175	G	C	-0.502	0.112	7.85E-06	0.410	0.223	0.0662	
rs10166852	2	183450923	C	G	I	0.989	0.471	C	G	-0.375	0.084	7.85E-06	0.141	0.176	0.423	
rs6736484	2	45146524	G	T	I	0.948	0.068	G	T	-0.721	0.162	9.05E-06	-0.285	0.338	0.399	
rs2912513	8	69968166	A	T	I	0.999	0.035	A	T	-0.996	0.224	9.33E-06	-0.119	0.472	0.8	

HCHS/SOL = Hispanic Community Health Study/Study of Latinos. These models were estimated using a linear mixed model fit by maximum likelihood with age, education, study center, 5 principal components, and covariates adjusting for the sampling design. Imputation was conducted using IMPUTE2. In HCHS/SOL, A1 was the tested allele. The Bonferroni adjusted alpha level in these analyses was $0.05/18=0.003$.

Table 4

genome-wide environment interaction study (GWEIS) top results for the top loci ($p < 1 \times 10^{-6}$) in African Americans

Stressful Life Event Results (n=6,982)												
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	SNP Main Effect			Closest Gene (<20kb)
									Beta	SE	p-value	
rs4652467	1	180097705	I	0.945	0.026	A	G	0.026	-0.662	0.167	1	<i>CEP350</i>
rs7275997	21	19663487	G	0.993	0.180	A	G	0.820	0.264	0.069	1	<i>TMPPRSS15</i>
rs28377528	7	153884444	I	0.874	0.420	A	G	0.580	-0.212	0.058	1	<i>DPP6</i>
rs2852310	18	43093004	I	0.996	0.027	A	G	0.027	-0.560	0.183	1	<i>SLC14A2</i>
rs12183135	6	151353805	G	0.996	0.024	C	G	0.024	-0.137	0.155	1	<i>MTHFD1L</i>
Social Support Results (n=6,908)												
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	SNP Main Effect			Closest Gene (<20kb)
									Beta	SE	p-value	
rs77966298	2	10984514	I	0.891	0.034	A	G	0.966	0.796	0.223	1	<i>PDIA6</i>
rs6419121	4	88490040	I	0.921	0.178	C	G	0.178	-0.375	0.096	1	
rs10836421	11	35581792	I	0.971	0.315	A	G	0.315	-0.184	0.071	1	
rs78012311	21	33634345	I	0.981	0.104	C	G	0.104	0.468	0.094	1	<i>MIS18A</i>

Robust (sandwich) standard errors are presented. In these tests of statistical interaction (on the additive scale and using allele dosages), probABEL uses A2 as the tested (non-reference) allele. The beta coefficients in these models can be interpreted as follows. The SNP main effect beta coefficient indicates the average difference in levels of depressive symptoms for women with a zero value on all covariates, who have 1 copy of the tested allele, and who are in the lowest quartile of stressful life events. The G×E interaction term indicates the average estimated difference in the effect of each tested allele on depressive symptoms associated with a one-unit different in stressful life events, adjusting for covariates. The Bonferroni adjusted alpha level in these analyses was 2.5×10^{-8} .

genome-wide environment interaction study (GWELS) top results for the top loci ($p < 1 \times 10^{-6}$) in Hispanics

Table 5

Stressful Life Event Results (n=2,989)																
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	SNP Main Effect			Closest Gene (<20kb)				
									Beta	SE	p-value					
SNP*Stressful Life Events Interaction Term																
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	Beta	SE	p-value					
rs58707171	4	36317832	I	0.921	0.037	A	C	0.963	0.649	0.254	1	-0.778	0.152	3.02E-07	intron variant	<i>DTHD1</i>
rs6579218	20	33709846	I	0.989	0.156	C	G	0.844	0.425	0.142	1	-0.505	0.100	4.94E-07	intron variant	<i>EDEM2</i>
rs10227305	7	3272267	I	0.849	0.207	A	C	0.793	0.308	0.133	1	-0.454	0.093	9.36E-07		
Social Support Results (n=3,012)																
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	SNP Main Effect			Closest Gene (<20kb)				
									Beta	SE	p-value					
SNP*Social Support Interaction Term																
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	Beta	SE	p-value					
rs35612712	4	187347203	I	0.941	0.416	C	T	0.416	-0.375	0.119	1	0.376	0.074	3.42E-07	intron variant	<i>F11-AS1</i>
rs61973969	13	96689182	I	0.985	0.032	C	T	0.032	-0.979	0.317	1	0.839	0.171	9.41E-07	intron variant	<i>UGGT2</i>

Robust (sandwich) standard errors are presented. In these tests of statistical interaction (on the additive scale and using allele dosages), probABEL uses A2 as the tested (non-reference) allele, consistent with MACH. The beta coefficients in these models can be interpreted as follows. The SNP main effect beta coefficient indicates the average difference in levels of depressive symptoms for women with a zero value on all covariates, who have 1 copy of the tested allele, and who are in the lowest quartile of stressful life events. The G×E interaction term indicates the average estimated difference in the effect of each tested allele on depressive symptoms associated with a one-unit different in social support, adjusting for covariates. The Bonferroni adjusted alpha level in these analyses was 2.5×10^{-8} .

Table 6

Replication of genome-wide environment interaction study (GWEIS) results for the top loci ($p < 1 \times 10^{-6}$) in African Americans and Hispanics

African Americans														
Stressful Life Event Results (n=952)														
						SNP Main Effect			SNP*Stressful Life Events Interaction Term					
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	Beta	SE	p-value	Beta	SE	p-value
rs4652467	1	180097705	I	0.919	0.029	A	G	0.971	0.143	0.310	0.645	-1.227	0.765	0.109
rs7275997	21	19663487	I	0.999	0.176	A	G	0.176	-0.242	0.151	0.109	0.124	0.279	0.655
rs28377528	7	153884444	I	0.960	0.439	A	G	0.439	-0.018	0.116	0.878	0.074	0.206	0.718
rs2852310	18	43093004	I	0.961	0.028	A	G	0.972	0.192	0.273	0.483	0.660	0.422	0.118
rs12183135	6	151353805	I	0.953	0.025	C	G	0.975	-0.693	0.374	0.064	-0.035	0.889	0.969
Social Support Results (n=952)														
						SNP Main Effect			SNP*Societal Support Interaction Term					
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	Beta	SE	p-value	Beta	SE	p-value
rs77966298	2	10984514	I	0.995	0.027	A	G	0.027	0.177	0.686	0.796	-0.073	0.319	0.820
rs6419121	4	88490040	I	0.994	0.177	C	G	0.823	-0.136	0.259	0.598	0.105	0.131	0.422
rs10836421	11	35581792	I	0.998	0.284	A	G	0.716	0.325	0.193	0.092	-0.166	0.100	0.098
rs78012311	21	33634345	I	0.995	0.082	C	G	0.918	0.470	0.343	0.171	-0.243	0.176	0.167
Hispanics														
Stressful Life Event Results (n=1,117)														
						SNP Main Effect			SNP*Societal Support Interaction Term					
SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	Beta	SE	p-value	Beta	SE	p-value
rs58707171	4	36317832	I	0.992	0.040	A	C	0.960	0.374	0.413	0.365	-0.771	0.788	0.328
rs6579218	20	33709846	I	0.997	0.141	G	C	0.141	0.554	0.241	0.022	0.210	0.411	0.609
rs10227305	7	3272267	G	1	0.196	A	C	0.804	0.115	0.211	0.586	0.233	0.362	0.521

Social Support Results (n=1,117)

SNP	chr	position	G/I	Info	MAF	A1	A2	Freq1	SNP Main Effect			SNP*Social Support Interaction Term		
									Beta	SE	p-value	Beta	SE	p-value
rs35612712	4	187347203	I	0.999	0.401	T	C	0.401	-0.026	0.174	0.883	-0.170	0.268	0.525
rs61973969	13	96689182	I	0.998	0.035	T	C	0.035	-0.118	0.458	0.797	0.467	0.706	0.509

HRS = Health and Retirement Study. These models were estimated using R 3.0.1. Covariates in HRS were age, income, education, marital status, and the top 10 principal components. Imputation was conducted using IMPUTE2. The Bonferroni adjusted alpha level in these analyses was 0.05/9=0.006. Robust SE were used.

Table 7

Results of genome-wide complex trait analysis based on the GREML method

	Model 1				Model 2			
	V(G)/Vp	SE	P		V(G)/Vp	SE	P	
Depressive symptoms	0.05	0.04	0.07		0.04	0.04	0.16	
Stressful life events	0.08	0.04	0.02		0.06	0.04	0.06	
Social support	0.04	0.04	0.13		0.03	0.04	0.25	
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	V(G)/Vp	SE	P		V(G)/Vp	SE	P	
Depressive symptoms, controlling for stress	0.03	0.04	0.18		0.02	0.04	0.29	
Depressive symptoms, controlling for support	0.04	0.04	0.11		0.03	0.04	0.19	
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	rG	SE	P		rG	SE	P	
Depressive symptoms and stressful life events	0.95	0.32	0.01		0.97	0.48	0.04	
Depressive symptoms and social support	-0.80	0.45	0.08		-0.79	0.76	0.21	

V(G)/Vp = SNP heritability estimate

rG = bivariate REML analysis

Model 1: Adjusted for age, principal components, and imputation group

Model 2: Adjusted for Model 1 covariates and income, education, marital status

All phenotypes were treated as continuous measures.

p-values for the bivariate REML analysis are one-sided and test whether the genetic correlation between depressive symptoms and each of the two environmental exposures is significantly different from zero.