HUNTING DOWN HORIZON-SCALE EFFECTS WITH MULTI-WAVELENGTH SURVEYS

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ABSTRACT

Next-generation cosmological surveys will probe ever larger volumes of the universe, including the largest scales, near and beyond the horizon. On these scales, the galaxy power spectrum carries signatures of local primordial non-Gaussianity (PNG) and horizon-scale general relativistic (GR) effects. However, cosmic variance limits the detection of horizon-scale effects. Combining different surveys via the multi-tracer method allows us to reduce the effect of cosmic variance. This method benefits from large bias differences between two tracers of the underlying dark matter distribution, which suggests a multi-wavelength combination of large volume surveys that are planned on a similar timescale. We show that the combination of two contemporaneous surveys, a large neutral hydrogen intensity mapping survey in SKA Phase 1 and a Euclid-like photometric survey, will provide unprecedented constraints on PNG as well as detection of the GR effects. We forecast that the error on local PNG will break through the cosmic variance limit on cosmic microwave background surveys and achieve $\sigma(f_{\rm NL}) \simeq 1.4 - 0.5$, depending on assumed priors, bias, and sky coverage. GR effects are more robust to changes in the assumed fiducial model, and we forecast that they can be detected with a signal-to-noise of about 14.

Key words: cosmological parameters - cosmology: miscellaneous - large-scale structure of universe

1. INTRODUCTION

Upcoming cosmological surveys will probe larger volumes of the universe, opening new windows to studying cosmological effects on horizon scales (see e.g., Yoo et al. 2012; Alonso et al. 2015a; Camera et al. 2015c; Raccanelli et al. 2015). These effects include primordial non-Gaussianity (PNG) and general relativistic (GR) horizon-scale effects in the observed power spectrum.

PNG is a key discriminator between different classes of inflation models. Local-type PNG (characterized by the parameter $f_{\rm NL}$) leaves a frozen imprint on horizon-scale power, allowing us to probe the primordial universe via the cosmic microwave background (CMB) and large-scale structure surveys. The Planck constraint (Planck Collaboration et al. 2015), $\sigma(f_{\rm NL}) \simeq 6.5$ (using the large-scale structure convention), is far stronger than those from current galaxy surveys, but is close to the maximum achievable with CMB experiments, which can only rule out inflation models with relatively large PNG.

Local PNG induces a scale-dependent correction to the bias of any dark matter tracer (Dalal et al. 2008; Matarrese & Verde 2008). This scale dependence can be probed through the two-point correlation function of the tracer on very large scales, allowing next-generation surveys to significantly improve upon the CMB constraints (see, e.g., Giannantonio et al. 2012; Camera et al. 2013; Camera et al. 2015d).

At this level of sensitivity, neglecting GR horizon-scale effects would bias results. Moreover, they might hint at something new if GR breaks down on these scales. They arise via lightcone observations of dark matter tracers such as the number counts of galaxies (Yoo 2010; Challinor & Lewis 2011; Bonvin & Durrer 2011) or maps of intensity (e.g., the integrated 21 cm signal from neutral hydrogen (H I) galaxies; Hall et al. 2013), including Doppler, Sachs–Wolfe, integrated Sachs–Wolfe, and time-delay-type terms. The lensing

contribution to the clustering power, mediated by magnification bias, can also be significant on horizon scales (Alonso et al. 2015a; Montanari & Durrer 2015).

Cosmic variance becomes a serious obstacle for horizonscale measurements where PNG and GR signals are strongest. Forecasts for next-generation surveys show that GR effects will not be detectable using a single tracer and PNG detection is limited to $\sigma(f_{\rm NL}) > 1$ (Alonso et al. 2015a; Raccanelli et al. 2015). This calls for the multi-tracer technique (MT) to reduce the effect of cosmic variance (McDonald & Seljak 2009; Seljak 2009).

MT has been used to explore improvements in the measurement of $f_{\rm NL}$ (see e.g., McDonald & Seljak 2009; Hamaus et al. 2011; Abramo & Leonard 2013; Ferramacho et al. 2014; Yamauchi et al. 2014). In these works, the lensing and GR contributions to clustering power were ignored. While this may have little effect on σ ($f_{\rm NL}$), it can significantly bias the best-fit value extracted from the data (Namikawa et al. 2011; Camera et al. 2015b). MT has also been used to forecast detectability of GR effects by Yoo et al. (2012), but neglecting the lensing contribution and the integrated GR effects. Here we include all lensing and GR effects without making any flat-sky approximation.

The MT technique opens a new observational window into probing large-scale signatures in the universe. In addition to reducing cosmic variance, it also cancels the individual systematics of the two experiments and removes foreground residuals. We show here that MT is a game-changer in the way we design surveys to probe these scales, as volume is no longer the ultimate goal and noise reduction becomes a priority again.

2. THE MULTI-TRACER TECHNIQUE

The theoretical observed fluctuations for a given dark matter tracer A can be written in Fourier space and Newtonian gauge in the form:

$$\Delta^{A} = \delta \left\{ b_{G}^{A} + \Delta b^{A} + f \frac{k_{\parallel}^{2}}{k^{2}} + E \left(\mathcal{Q}^{A} - 1 \right) \frac{k_{\perp}^{2}}{k^{2}} \right. \\ \left. + \left[F - if \left(b_{e}^{A} - 2\mathcal{Q}^{A} + \frac{2\mathcal{Q}^{A}}{\mathcal{H}\chi} \right) \frac{k_{\parallel}}{k} \right] \frac{\mathcal{H}}{k} \right. \\ \left. + \left[G + f b_{e}^{A} + I \mathcal{Q}^{A} + J \left(b_{e}^{A} - 2\mathcal{Q}^{A} + \frac{2\mathcal{Q}^{A}}{\mathcal{H}\chi} \right) \right] \frac{\mathcal{H}^{2}}{k^{2}} \right\},$$
(1)

where $\mathbf{k} = (\mathbf{k}_{\perp}, k_{\parallel})$ and \mathcal{H} is the conformal Hubble parameter. The first line contains the RSD and lensing terms, while the next two lines constitute the horizon-scale GR terms. The density contrast δ is in the comoving-synchronous gauge in order to consistently define the bias on large scales, and f is the growth rate. The correction to the Gaussian bias $b_{\rm G}$ due to local PNG is given by $\Delta b(z, k) = 3f_{\text{NL}} [b_{\text{G}}(z) - 1] \Omega_m H_0^2 \delta_c / [D(z)T(k)k^2].$ Here $\delta_c \simeq 1.69$ is the critical matter density contrast for spherical collapse, T(k) is the transfer function (normalized to 1 on large scales), and D(z) is the growth factor (normalized to 1 at z = 0). The magnification bias is $\mathcal{Q}(z, \mathcal{F}_*) \equiv$ $[-\partial \ln N_{\rm s}/\partial \ln \mathcal{F}]_{\mathcal{F}=\mathcal{F}_*}$, where $N_{\rm s}(z, \mathcal{F} > \mathcal{F}_*)$ is the background number density of sources at redshift z with flux \mathcal{F} above the detection threshold \mathcal{F}_* . The evolution bias is $b_e(z) = -\partial \ln[(1+z)^{-3}N_s]/\partial \ln(1+z)$. For details on the tracer-independent background functions E, F, G, I, and J see Challinor & Lewis (2011).

We see that PNG grows as k^{-2} , while GR terms grow as \mathcal{H}/k or \mathcal{H}^2/k^2 . The difference in the scale dependence and amplitude of the different terms allows GR corrections to be distinguished from PNG. More importantly, cosmic variance uncertainties come from δ , since it is a single realization of the underlying probability distribution. The MT technique relies on the fact that the ratio of different Δs is independent of cosmic variance, since both tracers are linear in δ .

The estimator we use is the sky map itself in the form of the $a_{\ell m}$. Assuming that the $a_{\ell m}$ are Gaussian, all the information will be encoded in the angular power spectrum, $\langle a_{\ell m} a_{\ell' m'}^* \rangle = \delta_{\ell \ell'} \delta_{m m'} C_{\ell}$. Using a Gaussian likelihood for the $a_{\ell m}$, the corresponding Fisher matrix will then be enough to account for the MT effects. Extending the single tracer case (Challinor & Lewis 2011) to multi-tracers, the angular power spectrum is given by

$$C_{\ell}^{AB}(z_i, z_j) = 4\pi \int d \ln k \, \Delta_{\ell}^{W_A}(z_i, k) \Delta_{\ell}^{W_B}(z_j, k) \mathcal{P}_{\zeta}(k).$$
(2)

Here, z_i are the redshift bin centers and \mathcal{P}_{ζ} is the dimensionless power spectrum of the primordial curvature perturbation. The measurable transfer function in the bin is

$$\Delta_{\ell}^{W_A}(z_i, k) = \int dz \ p^A(z) W(z_i, z) \Delta_{\ell}^A(z, k), \tag{3}$$

where $p^{A}(z)$ is the redshift distribution function of tracer A.

Equation (1) gives the theoretical transfer function $\Delta_{\ell}^{A}(z, k)$, while $W(z_i, z)$ is the window function centered on z_i , namely the probability distribution function of a source to be inside the *i*th bin. The product $p^{A}(z)W(z_i, z)$ is the effective tracer's redshift distribution function inside the bin, normalized so that $\int dz \ p^{A}(z)W(z_i, z) = 1$ for all z_i .

3. MULTI-WAVELENGTH SURVEYS

In order to optimally exploit the MT method, we look for two surveys with significant difference in bias. We focus on the H1 intensity mapping (IM) survey that will be performed with SKA phase 1 in "single" dish mode (Maartens et al. 2015; Santos et al. 2015) together with a Euclid-type galaxy survey (Laureijs et al. 2011; Amendola et al. 2013). We opt for the planned photometric survey because it will detect a larger number of galaxies than the spectroscopic option. We also consider variations to this survey with different noise and sky coverage, similar to a second-generation galaxy survey such as the Large Synoptic Survey Telescope (LSST; LSST Dark Energy Science Collaboration 2012; Bacon et al. 2015).

H I IM Experiment

In H_I IM, all galaxies with H_I contribute to the signal. We compute the Gaussian H_I bias, $b_{\rm G}^{\rm H_I}(z)$, weighting the halo bias with the H_I content in the dark matter halos (Santos et al. 2015). The number of observed sources is independent of the flux limit but the fractional temperature perturbation is equal to Equation (1) with $Q^{\rm H_I} = 1$ (Hall et al. 2013). The signal's redshift distribution follows the H_I temperature, $p^{\rm H_I}(z) \propto T_{\rm H_I}(z)$, which we fit using the results of Santos et al. (2015). For H_I IM, $N_{\rm s}/(1 + z)^3 = n_{\rm H_I}(z) \propto T_{\rm H_I}(z)/(1 + z)^2$ is the comoving density of H_I atoms (Hall et al. 2013), which is used to compute $b_e^{\rm H_I}$. The noise angular power spectrum in the *i*th bin of frequency width Δv_i for an experiment with N_d collecting dishes, total observation time $t_{\rm tot}$, and observed fraction of the sky $f_{\rm sky}$, is given by

$$\mathcal{N}_{\rm H\,I}^{ij} = \frac{4\pi f_{\rm sky} T_{\rm sys}^2}{2N_d t_{\rm tot} \Delta \nu_i} \delta^{ij},\tag{4}$$

where $T_{\rm sys} = 25 + 60(300 \text{ MHz}/\nu)^{2.55} \text{ K}$ is the system temperature. For SKA1, we assume $N_d t_{\rm tot} = 2 \times 10^6 \text{ hr}$ and $f_{\rm sky} = 0.72$. For H₁ IM one can neglect shot noise (Gong et al. 2011).

Photometric Galaxy Survey (PG)

The bias, magnification bias, and photometric galaxies' redshift distribution that we adopt are (Amendola et al. 2013; Raccanelli et al. 2015)

$$b_{\rm G}^{\rm PG}(z) = \sqrt{1+z},\tag{5}$$

$$Q^{\rm PG}(z) = 0.2985 + 0.5305z - 0.1678z^2 + 0.2578z^3$$
, (6)

$$p^{\rm PG}(z) \propto z^2 \exp\left[-(1.412z/0.9)^{3/2}\right],$$
 (7)

where the proportionality in the last equation is set by the total number of galaxies detected. For a galaxy survey the noise angular power spectrum is dominated by shot noise, i.e.,

$$\mathcal{N}_{\rm PG}^{ij} = \frac{\delta^{ij}}{N_{\rm PG}^i},\tag{8}$$

where N_{PG}^{i} is the number of galaxies per steradian in the *i*th bin. For the scatter between the photometric redshift estimate and the true redshift, we use $\sigma_{ph}(z) = 0.05(1 + z)$ (Ma et al. 2005).

We consider three observational scenarios: (*i*) Euclid-like survey detecting 30 galaxies per square arcminute and covering 15,000 deg², with 50% overlap with the SKA1 H_I experiment; (*ii*) the same case but with 100% overlap; (*iii*) a more futuristic

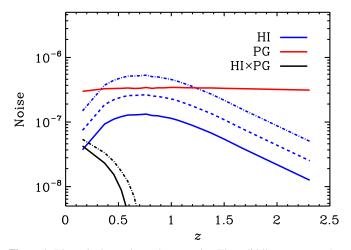


Figure 1. Dimensionless noise and cross-noise. The solid lines correspond to $f_{sky} = 0.18$, dashed to $f_{sky} = 0.36$, and dotted–dashed to 0.72, i.e., the scenarios in Equation (9). Note that (*i*) and (*ii*) have the same cross-noise since $N_{PG}^{(i)} = N_{PG}^{(ii)}$.

LSST-like survey detecting 40 galaxies per square arcminute and covering the whole SKA1 sky. Summarizing,

$$f_{\rm sky}^{(i)} = \frac{1}{2} f_{\rm sky}^{(ii)} = \frac{1}{4} f_{\rm sky}^{(iii)} = 0.18$$
(9)

$$N_{\rm PG}^{(i)} = N_{\rm PG}^{(ii)} = \frac{3}{4} N_{\rm PG}^{(iii)} = 30 \, {\rm arcmin}^{-2}.$$
 (10)

When referring to a single tracer, we use its value of f_{sky} , while in the case of MT for (*i*) and (*ii*) only the overlapping sky fraction is considered. Note that for (*i*) and (*ii*), we allow the SKA survey to be optimized for the smaller sky area, while maintaining the same total observation time, which will make the noise decrease (Equation (4)). The same cannot be done for the galaxy survey, since its sky coverage is assumed to be already fixed.

HI-PG Cross-noise

We also need to take into account the possible shot noise cross-power spectrum in each bin. This is due to an overlap in the halo mass range which the tracers probe. Even if this is small, it might be important for the MT as this is the only noise showing up in the cross-correlation between tracers. The crossnoise is given by

$$\mathcal{N}_{\mathrm{H\,I},PPG}^{ij} = \frac{\delta^{ij} T_{\mathrm{H\,I}}(z_i)}{\rho_{\mathrm{H\,I}}(z_i) N_{\mathrm{PG}}^i} \int dM \ \frac{dN}{dM} M_{\mathrm{H\,I}}(M) \Theta(M), \quad (11)$$

where dN/dM is the halo mass function, $M_{\rm H\,I}$ is the mass of H I in a halo of mass M, and $\rho_{\rm H\,I}(z_i)$ is the H I density. If the halo masses probed by the two surveys overlap, then $\Theta(M) = 1$; otherwise it is zero. For further details on the halo mass range for the SKA H I IM survey, see Santos et al. (2015). The mass range for the photo-z survey is found by matching the number of galaxies in the bin given by the halo mass function with the number given by the redshift distribution of photometric galaxies.

Figure 1 shows the (dimensionless) noise and cross-noise for the chosen binning. The cross-noise is only different from zero at low redshifts when some halos will be detected by both surveys. While H₁ IM is sensitive to the low halo mass galaxies, galaxy surveys are more sensitive to massive halos. For the bins with cross-noise, this is at least two orders of magnitude lower than the cross-angular power spectrum at large-angle multipoles. This might not be the case for more sensitive future surveys (like LSST), which may detect lower mass halos.

4. FORECASTING METHODOLOGY

We perform a Fisher analysis (Tegmark et al. 1997) for a set of 64 parameters

$$\vartheta_{\alpha} = \left\{ f_{\text{GR}}, f_{\text{NL}}, \ln A_s, \ln \Omega_{\text{DM}}, \ln b_i^A, \ln \mathcal{Q}_i \right\}, \qquad (12)$$

with $b_i^A \equiv b_G^A(z_i)$ (40 parameters) and $Q_i \equiv Q^{PG}(z_i)$ (20 parameters). We consider a bias parameter per bin, since due to the high precision measurements we are achieving with the MT technique, using a smaller number of bins with interpolation of parameters might impose a rather strong prior.

We assume a fiducial concordance cosmology with $H_0 = 67.74 \text{ km s}^{-1} \text{ Mpc}^{-1}$, cold dark matter fraction $\Omega_{\text{DM}} = 0.26$, baryon fraction $\Omega_{\text{b}} = 0.05$, amplitude of primordial scalar perturbations $A_{\text{s}} = 2.142 \times 10^{-9}$, and $f_{\text{NL}} = 0$. The fiducial values for the galaxy bias and magnification bias are set by the survey specifications in Section 3. The new parameter f_{GR} , used as a quantifier of the measurability of GR effects, is defined by

$$C_{\ell} = C_{\ell}^{\delta + \text{RSD} + \text{lens}} + f_{\text{GR}} C_{\ell}^{\text{GR}}, \qquad (13)$$

where we take $f_{\rm GR} = 1$ as the fiducial value and $C_{\ell}^{\delta+\rm RSD+\rm lens}$ accounts for auto- and cross-correlations between the density, RSD, and lensing terms. The term $C_{\ell}^{\rm GR}$ includes all correlations that contain a GR horizon-scale term. We chose this definition in order to compute $\partial C_{\ell}/\partial f_{\rm GR}$ analytically. We include all effects in Equation (1) at the same time in order not to bias the accuracy on parameter reconstruction (see, e.g., Namikawa et al. 2011; Camera et al. 2015a, 2015b).

We use 20 redshift bins in the range 0 < z < 3, with variable size such that approximately the same number of photo-*z* galaxies resides in each bin. Then, we adopt exactly the same redshift binning for SKA1 H I IM, so that there is a complete overlap between the two tracers. We can do so thanks to the high resolution of an IM experiment, which allows us to tune the frequency windows. For the MT covariance matrix, we follow Ferramacho et al. (2014). We focus on large scales, neglecting beam effects from limited angular resolution, which should be negligible for small ℓ .

The angular power spectra for MT are computed for a Gaussian window function using a modified version of the publicly available CAMB_sources code (Challinor & Lewis 2011). The code was changed to include the tracers' redshift distribution and to compute $b_e^{\rm H\,I}$.

We improve the Fisher matrix analysis to ensure numerical stability. Derivatives with respect to f_{GR} and $\ln b_i^A$ are analytical, and numerical differentiation with respect to other parameters is done using a five-point stencil method. We use the logarithm of the parameter making the Fisher matrix less prone to numerical issues.

The high dimensionality of the tomographic matrices and the large number of nuisance parameters (60) require utmost control on the matrix operations. Therefore, we perform matrix inversion via "inverse diagonalization": given a square matrix,

 Table 1

 Marginal Errors on f_{GR} and f_{NL} for a Euclid-like Photo-z Survey, H I Intensity Mapping with SKA1, and Their Combined MT Analysis

		$\sigma(f_{\rm GR})$					$\sigma(f_{\rm NL})$					
	Н		PG		MT		Н		PG		MT	
$\ell_{\rm max}$	60	300	60	300	60	300	60	300	60	300	60	300
without {ln b_i^A , ln Q_i }	1.36	1.33	1.58	1.55	0.075	0.072	4.57	4.31	5.34	5.13	1.23	1.12
with $\{\ln b_i^A, \ln Q_i\}$	1.39		1.90		0.079		5.24		6.02		1.37	
with $\{\ln b_i^A, \ln Q_i\} + 5\%$ prior	1.38		1.68		0.076		5.62		5.53		1.36	

its inverse is $A^{-1} = U\Lambda^{-1}U^{-1}$, where U and Λ are respectively the matrices of the eigenvectors and eigenvalues of A. Thus, Λ is diagonal by construction and its inversion is trivial. This also helps in removing degeneracies in the Fisher matrix. Indeed, when marginalizing over the set of nuisance parameters, if one or more eigenvalues (nearly) vanish, then this degeneracy does not propagate into the cosmological parameters of interest (Camera et al. 2012).

5. DISCUSSION OF RESULTS

Unless otherwise stated, we assume configuration (i) for the photo-z survey, which is the more near-term scenario. For the Fisher matrix, we set the maximum available angular scale, ℓ_{\min} , to 2, and consider two different minimum angular scales, $\ell_{\rm max} = 60$ and 300. Table 1 shows the 1σ marginal error on $f_{\rm GR}$ and $f_{\rm NL}$ for the different tracer configurations, $\ell_{\rm max}$ and 60 nuisance parameters. Given the high dimensionality of both the C_{ℓ} tomographic matrix and the Fisher matrix, as well as the various implementations ensuring numerical stability of the matrix operations, the computation of the Fisher matrix can become unwieldy as ℓ_{max} increases. Therefore in Table 1 we present results for $\ell_{\rm max} = 300$ for the cosmological parameter set, and show the trend due to the inclusion of the nuisance parameters for $\ell_{max} = 60$ only. The big improvement of MT over the single tracer is apparent: for $f_{\rm NL}$ we get constraints ~ 5 times tighter, and for f_{GR} the improvement is even more impressive as the bound shrinks by a factor >20. Moreover, MT is more robust when we allow for full uncertainty on the bias-related nuisance parameters, as can be seen in the impact of a 5% prior on the nuisance parameters.

The use of two different maximum angular multipoles is done not for the sake of a conservative versus optimistic comparison: both ℓ_{max} s are well within the linear régime and the inclusion, besides A_s , of Ω_{DM} and the nuisance parameters ensures that we do not overestimate the constraining power on f_{NL} or f_{GR} , even when pushing to small scales. Instead, we want to understand to what extent smaller scales contribute to the signal of PNG or GR effects, both of which are strongest on ultra-large scales. This also enables us to monitor the impact of noise. Noise usually dominates on small scales and is negligible in the cosmic variance limited régime, but Seljak (2009) and Ferramacho et al. (2014) suggested that the more MT is effective in "removing" cosmic variance, the larger the scales at which noise starts becoming relevant.

In Figure 2 we show the forecast marginal errors on f_{GR} (solid) and f_{NL} (dashed) as a function of the noise level. We multiply the noise of both IM and galaxy number counts by a fudge factor and let it vary from 0 to 1, where 0 means a noiseless experiment and 1 is the real setting. As we remove noise, single tracers soon reach the cosmic variance limited

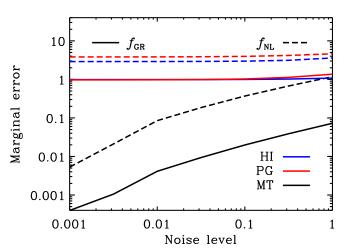


Figure 2. Marginal 1σ error on f_{GR} (solid) and f_{NL} (dashed) vs. noise level for SKA1 intensity mapping (blue), Euclid-like photo-*z* galaxies (red) and MT (magenta) for $\ell_{max} = 100$.

plateau, while MT keeps improving. So, as cosmic variance fades, the more important the signal-to-noise ratio becomes.

Seljak et al. (2009) proposed a mass-dependent weighting of the detected sources, which can considerably suppress the stochasticity between halos and dark matter, thus reducing the shot noise contribution. By doing so, they showed that it will be possible for a next-generation Euclid-like survey to reduce the Poisson noise even by 30%. From Figure 2, the resulting improvement appears very clear (see also Hamaus et al. 2011).

The impact of f_{sky} and b_G^{PG} is studied by performing the same analysis as before, but comparing the three photo-*z* scenarios (*i*), (*ii*), and (*iii*) (Section 3) and changing from the bias of Equation (5) to a higher bias $b_G^{PG} = 1 + z$. In Figure 3 we present the corresponding forecast joint 1σ marginal error contours in the (f_{GR}, f_{NL}) plane. The collapse from the outermost to the middle ellipses is simply caused by the doubling of the surveyed sky areas, but the innermost contours are also affected by a reduced photometric galaxy shot noise as the number density grows by 30%. Also, we remind the reader that the H I IM noise is linearly dependent on f_{sky} . Summarizing, in Table 2 we quote the forecast marginal errors on the measurement of GR effects and PNG for the three photo-*z* scenarios and two biases. Results are shown for $\ell_{max} = 300$ and no nuisance parameters.

To conclude, the MT technique will allow synergies between SKA and Euclid to provide game-changing measurements on horizon scales. We have shown that this can break through the PNG barrier of $\sigma(f_{\rm NL}) = 1$ and make the first ever detections of GR effects. Moreover, our analysis shows that with this new

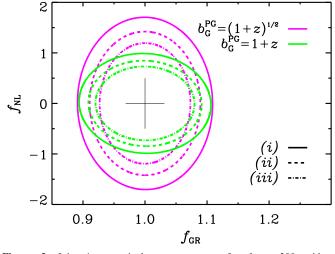


Figure 3. Joint 1σ marginal error contours for $\ell_{max} = 300$ without $\{\ln b_i^A, \ln Q_i\}$ for the two photo-z Gaussian bias models and the three survey scenarios

Table 2 Marginal Errors from the MT Analysis for the Three Photo-z Scenarios with Gaussian Bias $\sqrt{1+z}$ (or 1+z) and $\ell_{max} = 300$

	σ	$(f_{\rm GR})$	$\sigma(f_{\rm NL})$		
<i>(i)</i>	0.071	(0.070)	1.12	(0.65)	
<i>(ii)</i>	0.059	(0.060)	0.94	(0.56)	
(iii)	0.048	(0.053)	0.79	(0.48)	

method, we need to rethink the way large-scale surveys are being designed. Ultra-large volumes are no longer the ultimate goal, as we can cancel cosmic variance when probing these features. Instead, we only need to probe up to the required scale and maximization of the signal-to-noise should be the priority instead. In this context a survey of about 10,000 deg² should be enough. This at the same time will make it easier for the SKA1 cosmology survey to be commensal with other science cases. Finally, although we have not addressed specifically the issue of foreground contamination, it is expected that the MT technique will alleviate this problem even further since any possible residuals from the cleaning process (Alonso et al. 2015b) and even systematics should be uncorrelated between H I IM and the photo-z galaxy survey.

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Note added. While this paper was being completed, Alonso & Ferreira (2015) appeared, covering a similar topic.

REFERENCES

- Abramo, L. R., & Leonard, K. E. 2013, MNRAS, 432, 318
- Alonso, D., Bull, P., Ferreira, P. G., Maartens, R., & Santos, M. G. 2015a, arXiv:1505.07596
- Alonso, D., Bull, P., Ferreira, P. G., & Santos, M. G. 2015b, MNRAS, 447, 400
- Alonso, D., & Ferreira, P. G. 2015, arXiv:1507.03550
- Amendola, L., Appleby, S., Bacon, D., et al. 2013, LRR, 16, 6
- Bacon, D., Bridle, S., Abdalla, F. B., et al. 2015, PoS, AASKA1, 4, 145
- Bonvin, C., & Durrer, R. 2011, PhRvD, 84, 063505
- Camera, S., Carbone, C., Fedeli, C., & Moscardini, L. 2015a, PhRvD, 91, 043533
- Camera, S., Maartens, R., & Santos, M. G. 2015b, MNRAS, 451, L80
- Camera, S., Raccanelli, A., Bull, P., et al. 2015c, PoS AASKA1, 4, 025
- Camera, S., Santos, M. G., Bacon, D. J., et al. 2012, MNRAS, 427
- Camera, S., Santos, M. G., Ferreira, P. G., & Ferramacho, L. 2013, PhRvL, 111, 171302
- Camera, S., Santos, M. G., & Maartens, R. 2015d, MNRAS, 448, 1035
- Challinor, A., & Lewis, A. 2011, PhRvD, 84, 043516
- Dalal, N., Dore, O., Huterer, D., & Shirokov, A. 2008, PhRvD, 77, 123514
- Ferramacho, L. D., Santos, M. G., Jarvis, M. J., & Camera, S. 2014, MNRAS,
- 442, 2511 Giannantonio, T., Porciani, C., Carron, J., Amara, A., & Pillepich, A. 2012, MNRAS, 422, 2854
- Gong, Y., Chen, X., Silva, M., Cooray, A., & Santos, M. G. 2011, ApJL, 740 L20
- Hall, A., Bonvin, C., & Challinor, A. 2013, PhRvD, 87, 064026
- Hamaus, N., Seljak, U., & Desjacques, V. 2011, PhRvD, 84, 083509
- Laureijs, R., Amiaux, J., Arduini, S., et al. 2011, ESA-SRE, 12
- LSST Dark Energy Science Collaboration 2012, arXiv:1211.0310
- Ma, Z.-M., Hu, W., & Huterer, D. 2005, ApJ, 636, 21
- Maartens, R., Abdalla, F. B., Jarvis, M., & Santos, M. G. 2015, PoS, AASKA14, 016
- Matarrese, S., & Verde, L. 2008, ApJL, 677, L77
- McDonald, P., & Seljak, U. 2009, JCAP, 0910, 007
- Montanari, F., & Durrer, R. 2015, arXiv:1506.01369
- Namikawa, T., Okamura, T., & Taruya, A. 2011, PhRvD, 83, 123514
- Planck Collaboration Ade, P. A. R., Aghanim, N., Arnaud, M., et al. 2015, arXiv:1502.01592
- Raccanelli, A., Montanari, F., Bertacca, D., Doré, O., & Durrer, R. 2015, arXiv:1505.06179
- Santos, M., Bull, P., Alonso, D., et al. 2015, PoS, AASKA1, 4, 019
- Seljak, U. 2009, PhRvL, 102, 021302
- Seljak, U., Hamaus, N., & Desjacques, V. 2009, PhRvL, 103, 091303
- Tegmark, M., Taylor, A., & Heavens, A. 1997, ApJ, 480, 22
- Yamauchi, D., Takahashi, K., & Oguri, M. 2014, PhRvD, 90, 083520 Yoo, J. 2010, PhRvD, 82, 083508
- Yoo, J., Hamaus, N., Seljak, U., & Zaldarriaga, M. 2012, PhRvD, 86, 063514