Hypoexponential Distribution with Different Parameters

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ABSTRACT

The Hypoexponential distribution is the distribution of the sum of $n \geq 2$ independent Exponential random variables. This distribution is used in moduling multiple exponential stages in series. This distribution can be used in many domains of application. In this paper we consider the case of n exponential Random Variable having distinct parameters. Using convolution, some properties of Laplace transform and the moment generating function, we analyse this case and give new properties and identities. Moreover, we shall study particular cases when α_i are arithmetic and geometric.

1. Introduction

The Random Variable (RV) plays an important role in modeling many events [1,2]. In particular the sum of exponential random has important applications in the modeling in many domains such as communications and computer science [3,4], Markov process [5,6], insurance [7,8] and reliability and performance evaluation [4,5,9,10]. Nadarajah [11], presented a review of some results on the sum of random variables.

Many processes in nature can be divided into sequential phases. If the time the process spends in each phase is independent and exponentially distributed, then the overall time is hypoexponentially distributed. The service times for input-output operations in a computer system often possess this distribution. The probability density function (pdf) and cummulative distribution function (cdf) of the hypoexponential with distinct parameters were presented by many authors [5,12,13]. Moreover, in the domain of reliability and performance evaluation of systems and software many authors used the geometric and arithmetic parameters such as [10,14,15].

In this paper we study the hypoexponential distribution in the case of n independent exponential R. V. with distinct parameters $\alpha_i \neq \alpha_j$ for $i \neq j$, written as $\text{hypoexp}(\alpha_1, \alpha_2, \cdots, \alpha_n)$. We use in our work the properties of convolution, Laplace transform and moment generating function in finding the kth derivative of the pdf of this sum and the moment of this distribution of order k. In addition, we deduce some new equalities related to these parameters. Also we shall study the case when the parameters form an arithmetic and geometric sequence considered by [10,14,15] and find some new results.

2. Definitions and Notations

Let $X_1, X_2, \cdots, X_n$ be independent exponential random variables with different respective parameters α_i, $i = 1, 2, \cdots, n$, written as $X_i \sim \text{Exp}(\alpha_i)$. We define the random variable

$$S_n = \sum_{i=1}^{n} X_i$$

to be the Hypoexponential random variable with parameters α_i, $i = 1, 2, \cdots, n$, written as

$$S_n \sim \text{hypoexp}(\alpha_1, \alpha_2, \cdots, \alpha_n)$$

Some notations used throughout the paper.

X_i : $\text{Exp}(\alpha_i)$.
S_n : $\text{hypoexp}(\alpha_1, \alpha_2, \cdots, \alpha_n)$.
f_X : The pdf of the random variable X.
F_X : The cdf of the random variable X.
$f_X^{(k)}$: The kth derivative of the pdf f_X.

$\mathcal{L}\{\cdot\}$: Laplace-Stieltjes Transform.

$\mathcal{L}^{-1}\{\cdot\}$: Laplace Inverse.

$\Phi_X(t)$: The moment generating function of X.

$E[X^k]$: The moment of order k of the RV X.

α : $\prod_{i=1}^n \alpha_i$ product of all parameters.

P_i : $\prod_{j=1, j\neq i}^n \left(1 - \dfrac{\alpha_i}{\alpha_j}\right)$.

γ_i : $\prod_{j=1, j\neq i}^n (\alpha_j - \alpha_i)$.

E_k : $\left\{ (l_1, \cdots, l_n) / 0 \leq l_i \leq k; \sum_{i=1}^n l_i = k; 1 \leq i \leq n \right\}, E_0 = 0$.

3. Applications on pdf and cdf Using Laplace Transform

The pdf and cdf of the hypoexponential with distinct parameters were presented by many authors [2,7,11-13]. We shall state in thoerem 1 and propostion 1 these results and provide another proof using Laplace transform. Next, we give some new properties of its pdf, where new identities are obtained.

Theorem 1. Let $n \geq 2$ and $t > 0$. Then

$$f_{S_n}(t) = \sum_{i=1}^n \frac{f_{X_i}(t)}{P_i}$$

and

$$F_{S_n}(x) = 1 - \sum_{i=1}^n \frac{e^{-\alpha_i x}}{P_i} I_{(0,\infty)}(x).$$

Proof. We have

$$\mathcal{L}\{f_{X_i}(x)\} = \frac{\alpha_i}{s + \alpha_i},$$

where $s > \max\{-\alpha_i\}$ for $i = 1, 2, \cdots, n$. Since X_i are independent then $f_{S_n}(t)$ is the convolutions of f_{X_i}, $i = 1, 2, \cdots, n$ written as

$$f_{S_n}(t) = \left(f_{X_1} * f_{X_2} * \cdots * f_{X_n} \right)(t)$$

and the Laplace transform of convolution of functions is the product of their Laplace transform, thus

$$\mathcal{L}\{f_{S_n}(t)\} = \prod_{i=1}^n \mathcal{L}\{f_{X_i}(t)\}$$

$$= \prod_{i=1}^n \frac{\alpha_i}{s + \alpha_i} = \alpha \prod_{i=1}^n \frac{1}{s + \alpha_i} \quad (1)$$

where $s > \max\{-\alpha_i\}$. However, by Heaviside Expansion Theorem [16], for distinct poles gives that

$$\mathcal{L}\{f_{S_n}(t)\} = \alpha \sum_{i=1}^n \frac{A_i}{s + \alpha_i},$$

where

$$A_i = \frac{1}{\prod_{j=1, j\neq i}^n (\alpha_j - \alpha_i)}.$$

Therefore,

$$f_{S_n}(t) = \mathcal{L}^{-1}\left\{ \alpha \sum_{i=1}^n \frac{A_i}{s + \alpha_i} \right\}$$

$$= \alpha \sum_{i=1}^n A_i e^{-\alpha_i t} I_{(0,\infty)}(t).$$

But $\alpha A_i = \dfrac{\alpha_i}{P_i}$. Thus

$$f_{S_n}(t) = \sum_{i=1}^n \frac{f_{X_i}(t)}{P_i}.$$

On the other hand we have

$$F_{S_n}(x) = \int_0^x \sum_{i=1}^n \frac{f_{X_i}(t)\,\mathrm{d}t}{P_i}$$

$$= \sum_{i=1}^n \frac{1}{P_i} F_{X_i}(x) = \sum_{i=1}^n \frac{1 - e^{-\alpha_i x}}{P_i}.$$

But $\lim_{x \to \infty} F_{S_n}(x) = 1$, then $\sum_{i=1}^n \dfrac{1}{P_i} = 1$ and we conclude that

$$F_{S_n}(x) = 1 - \sum_{i=1}^n \frac{e^{-\alpha_i x}}{P_i} I_{(0,\infty)}(x). \qquad \square$$

Next we shall discuss the kth derivative of $f_{S_n}(t)$ and many equalities are obtained concerning P_i form and some similar forms.

We start by noting from the previous proof that $\sum_{i=1}^n \dfrac{1}{P_i} = 1$. Here, we shall state another simple proof using Laplace transform.

Proposition 1. Let $n \geq 2$. Then

$$\sum_{i=1}^n \frac{1}{P_i} = 1.$$

Proof. We have from Equation (1),

$$\mathcal{L}\{f_{S_n}(t)\} = \prod_{i=1}^n \left(\frac{\alpha_i}{s + \alpha_i} \right)$$

where $s > \max\{-\alpha_i\}, i = 1, 2, \cdots, n$. But from Theorem 1,

$$f_{S_n}(t) = \sum_{i=1}^n \frac{f_{X_i}(t)}{P_i}$$

and

$$\mathcal{L}\{f_{S_n}(t)\} = \sum_{i=1}^n \frac{\mathcal{L}\{f_{X_i}(t)\}}{P_i}$$

$$= \sum_{i=1}^n \frac{\alpha_i}{P_i(s + \alpha_i)} I_{(0,\infty)}(t).$$

Hence, $\prod_{i=1}^n \left(\dfrac{\alpha_i}{s + \alpha_i} \right) = \sum_{i=1}^n \dfrac{\alpha_i}{P_i(s + \alpha_i)}$. For $s = 0$,

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$$\prod_{i=1}^{n}\left(\frac{\alpha_i}{\alpha_i}\right) = \sum_{i=1}^{n}\frac{\alpha_i}{P_i(\alpha_i)}. \quad \text{Therefore,} \quad \sum_{i=1}^{n}\frac{1}{P_i} = 1. \quad \square$$

Lemma 1. Let $n \geq 2$. Then

$$\mathcal{L}\left\{f_{S_n}^{(k)}(t)\right\} = s^k \prod_{i=1}^{n}\left(\frac{\alpha_i}{s+\alpha_i}\right)$$

for $0 \leq k \leq n-1$.

Proof. The proof is done by induction. For $k = 0$, we have from Equation (1)

$$\mathcal{L}\left\{f_{S_n}(t)\right\} = \prod_{i=1}^{n}\left(\frac{\alpha_i}{s+\alpha_i}\right).$$

However, by Initial Value Theorem, we have

$$\lim_{t \to 0} f_{S_n}(t) = \lim_{s \to +\infty} s\mathcal{L}\left\{f_{S_n}(t)\right\}$$
$$= \lim_{s \to +\infty} s\prod_{i=1}^{n}\left(\frac{\alpha_i}{s+\alpha_i}\right) = 0$$

and for $k = 1$ we have

$$\mathcal{L}\left\{f_{S_n}^{(1)}(t)\right\} = s\mathcal{L}\left\{f_{S_n}(t)\right\} - f_{S_n}(0^+)$$
$$= s\mathcal{L}\left\{f_{S_n}(t)\right\} = s\prod_{i=1}^{n}\left(\frac{\alpha_i}{s+\alpha_k}\right).$$

Moreover

$$\mathcal{L}\left\{f_{S_n}^{(k+1)}(t)\right\} = s\mathcal{L}\left\{f_{S_n}^{(k)}(t)\right\} - f_{S_n}^{(k)}(0^+)$$

Continuing in the same manner till the $(n-1)^{th}$ derivative, we obtain the result. $\square$

In the following propostion we shall prove that the first $(n-2)^{th}$ derivative of the pdf of S_n are zeros, which verifies the fact that the coefficient of variation of the hypoexponential distribution is less than one unlike the hyperexponential distribution that have the coefficient of variation greater than 1.

Proposition 2. Let $n \geq 2$. Then

$$\lim_{t \to 0} f_{S_n}^{(k)}(t) = \begin{cases} 0, & \text{if } 0 \leq k \leq n-2 \\ \alpha, & \text{if } k = n-1 \end{cases}$$

Proof. Let $n \geq 2$, we have from Lemma 1,

$$\mathcal{L}\left\{f_{S_n}^{(k)}(t)\right\} = s^k \prod_{i=1}^{n}\left(\frac{\alpha_i}{s+\alpha_i}\right)$$

for $0 \leq k \leq n-1$ and from Initial Value Theorem, we have

$$\lim_{t \to 0} f_{S_n}^{(k)}(t) = \lim_{s \to +\infty} s\mathcal{L}\left\{f_{S_n}^{(k)}(t)\right\} = \lim_{s \to +\infty} \frac{s^{k+1}}{s^n}\alpha$$
$$= \lim_{s \to +\infty} \frac{1}{s^{n-k-1}}\alpha = \begin{cases} 0, \text{if } n-k-1 \geq 1 \\ \alpha, \text{if } n-k-1 = 0 \end{cases}$$
$$\square$$

Corollary 1. Let $n \geq 2$. Then

$$\sum_{i=1}^{n}\frac{\alpha_i^k}{P_i} = \begin{cases} 0, & \text{if } 1 \leq k \leq n-1 \\ (-1)^{n-1}\alpha, & \text{if } k = n \end{cases}$$

Proof. We have $f_{X_i}(t) = \alpha_i e^{-\alpha_i t} I_{(0,\infty)}(t)$. Then the rth derivative of f_{X_i} is

$$f_{X_i}^{(r)}(t) = (-1)^r \alpha_i^{r+1} e^{-\alpha_i t} I_{(0,\infty)}(t).$$

However, from Theorem 1,

$$f_{S_n}(t) = \sum_{i=1}^{n}\frac{f_{X_i}(t)}{P_i},$$

then

$$f_{S_n}^{(r)}(t) = \sum_{i=1}^{n}\frac{f_{X_i}^{(r)}(t)}{P_i} = \sum_{i=1}^{n}(-1)^r \frac{\alpha_i^{r+1} e^{-\alpha_i t}}{P_i} I_{(0,\infty)}$$

and

$$\lim_{t \to 0} f_{S_n}^{(r)}(t) = (-1)^r \sum_{i=1}^{n}\frac{\alpha_i^{r+1}}{P_i}. \tag{2}$$

By Proposition 2, we obtain that

$$\sum_{i=1}^{n}\frac{\alpha_i^{r+1}}{P_i} = \begin{cases} 0, & \text{if } 0 \leq r \leq n-2 \\ (-1)^r \alpha, & \text{if } r = n-1 \end{cases}$$

By replacing $r+1$ with k we obtain the result. $\square$

4. Applications on pdf and cdf Using Moment Generating Function

In the previous section we saw the use of Laplace properties in the proofs of the theorems and propositions. In a similar manner, in this section we use the moment generating function to obtain more new related results. A new form of the moment generating function of S_n and the moment of S_n of order k is given. Moreover, we deduce more new related equalities concerning P_i and higher order derivatives of pdf of S_n.

Proposition 3. Let $n \geq 2$. Then

$$\Phi_{S_n}(t) = \sum_{i=1}^{n}\frac{\Phi_{X_i}(t)}{P_i}.$$

Proof. We have

$$\Phi_{S_n}(t) = E\left[e^{tS_n}\right] = \int_{-\infty}^{+\infty} e^{tx} f_{S_n}(x)\,\mathrm{d}x$$

and from Theorem 1,

$$f_{S_n}(t) = \sum_{i=1}^{n}\frac{f_{X_i}(t)}{P_i},$$

then

$$\Phi_{S_n}(t) = \sum_{i=1}^{n}\frac{1}{P_i}\int_{-\infty}^{+\infty} e^{tx} f_{X_i}(x)\,\mathrm{d}x = \sum_{i=1}^{n}\frac{\Phi_{X_i}(t)}{P_i}. \quad \square$$

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Proposition 4. Let $n \geq 2$ and $k \geq 0$. Then

$$E\left[S_n^k\right] = \sum_{i=1}^n \frac{k!}{P_i \alpha_i^k}$$

Proof. We have from Proposition 3,

$$\Phi_{S_n}(t) = \sum_{i=1}^n \frac{\Phi_{X_i}(t)}{P_i}.$$

Then

$$\frac{d^k \Phi_{S_n}(t)}{dt^k} = \sum_{i=1}^n \frac{1}{P_i} \frac{d^k \Phi_{X_i}(t)}{dt^k}$$

and

$$\left.\frac{d^k \Phi_{S_n}(t)}{dt^k}\right|_{t=0} = \sum_{i=1}^n \frac{1}{P_i} \left.\frac{d^k \Phi_{X_i}(t)}{dt^k}\right|_{t=0}$$

which gives $E\left[S_n^k\right] = \sum_{i=1}^n \frac{E\left[X_i^k\right]}{P_i}$. But $E\left[X_i^k\right] = \frac{k!}{\alpha_i^k}$.

Thus we obtain the result. □

Next, we shall use the Proposition 3 and 4 to find other identities on P_i and higher orders for $f_{S_n}^{(k)}(t)$. We start by noting that $\Phi_{S_n}(0) = 1$ and by taking $t = 0$ in Proposition 3, we again obtain the result in Proposition 1, that is $\sum_{i=1}^n \frac{1}{P_i} = 1$.

Proposition 5. Let $n \geq 2$ and $k \geq 0$. Then

$$\sum_{i=1}^n \frac{1}{P_i \alpha_i^k} = \sum_{E_k} \frac{1}{\alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}}.$$

where

$$E_k = \left\{ (l_1, \cdots, l_n) / 0 \leq l_i \leq k; \sum_{i=1}^n l_i = k; 1 \leq i \leq n \right\}.$$

Note that we may write

$$\sum_{E_k} \frac{1}{\alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}} = \sum_{I_k} \frac{1}{\alpha_{i_1} \alpha_{i_2} \alpha_{i_3} \cdots \alpha_{i_k}}, \quad (3)$$

where

$$I_k = \left\{ (i_1, \cdots, i_k) / 1 \leq i_1 \leq i_2 \leq \cdots \leq i_k \leq n \right\}.$$

However E_k and I_k are equivalent representing a set of combination with repetition having $\binom{n+k-1}{k}$ possibilities and $E_0 = I_0 = 0$, thus the above summation (3) shall be 1.

Proof. Let $k \geq 0$ and $n \geq 2$. We have

$$E\left[S_n^k\right] = E\left[(X_1 + X_2 + \cdots + X_n)^k\right]$$

and using multinomial expansion formula, we obtain

$$E\left[S_n^k\right] = E\left[\sum_{E_k} \frac{k!}{l_1! l_2! \cdots l_n!} \left(X_1^{l_1} X_2^{l_2} \cdots X_n^{l_n}\right)\right].$$

Knowing that expectation is linear and X_i, $i = 1, 2, \cdots, n$ are independent with

$$E\left[X_i^{l_i}\right] = \frac{l_i!}{\alpha_i^{l_i}},$$

then

$$E\left[S_n^k\right] = \sum_{E_k} \frac{k!}{\alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}}. \quad (4)$$

Since from Proposition 4,

$$E\left[S_n^k\right] = \sum_{i=1}^n \frac{k!}{P_i \alpha_i^k}.$$

Therefore,

$$\sum_{i=1}^n \frac{1}{P_i \alpha_i^k} = \sum_{E_k} \frac{1}{\alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}}. \quad \square$$

The following corollary is direct consequence of Proposition 5 and Equation (4), taking $k = 0, 1$ and 2 respectively.

Corollary 2. Let $n \geq 2$. Then

1) $\sum_{i=1}^n \frac{1}{P_i} = 1$.

2) $\sum_{i=1}^n \frac{1}{P_i \alpha_i} = \sum_{i=1}^n \frac{1}{\alpha_i}$ and $E\left[S_n\right] = \sum_{i=1}^n \frac{1}{\alpha_i}$.

3) $\sum_{i=1}^n \frac{1}{P_i \alpha_i^2} = \sum_{1 \leq i \leq j \leq n} \frac{1}{\alpha_i \alpha_j}$ and $E\left[S_n^2\right] = \sum_{1 \leq i \leq j \leq n} \frac{2!}{\alpha_i \alpha_j}$.

In Proposition 2, we found the first $(n-1)^{th}$ derivative of f_{S_n} at 0, However to find higher order derivaties we recall Equation (2), that shows a direct relation between the kth derivative f_{S_n} and $\sum_{i=1}^n \frac{\alpha_i^k}{P_i}$. Hence, in the next propostion we shall use Propostion 5, to find an equation for $\sum_{i=1}^n \frac{\alpha_i^k}{P_i}$ by finding a relation between hypoexp$(\alpha_1, \alpha_2, \cdots, \alpha_n)$ and hypoexp$\left(\frac{1}{\alpha_1}, \frac{1}{\alpha_2}, \cdots, \frac{1}{\alpha_n}\right)$.

Proposition 6. Let $n \geq 2$ and $k \geq n$. Then

$$\sum_{i=1}^n \frac{\alpha_i^k}{P_i} = (-1)^{n-1} \alpha \sum_{E_{k-n}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}.$$

Proof. Let $n \geq 2, \beta_i = \frac{1}{\alpha_i}, i = 1, 2, \cdots, n$ and

$C_n \sim$ hypoexp$(\beta_1, \beta_2, \cdots, \beta_n)$. Then by Theorem 1, the pdf of C_n is

$$f_{C_n}(t) = \sum_{i=1}^n \frac{f_{Y_i}(t)}{B_i}$$

where $Y_i \sim \text{Exp}(\beta_i)$ and $B_i = \prod_{j=1, j \neq i}^{n}\left(1 - \frac{\beta_i}{\beta_j}\right)$.

Next, we shall find P_i in terms of B_i. We have

$$P_i = \prod_{j=1, j \neq i}^{n}\left(1 - \frac{\alpha_i}{\alpha_j}\right) = \prod_{j=1, j \neq i}^{n}\left(1 - \frac{\beta_j}{\beta_i}\right)$$

$$= \frac{(-1)^{n-1} \prod_{j=1, j \neq i}^{n}(\beta_j - \beta_i)}{\beta_i^{n-1}},$$

multiplying in the numerator and denominator by $\prod_{j=1, j \neq i}^{n}(\beta_j)$, we obtain $P_i = (-1)^{n-1} \frac{\beta}{\beta_i^{n}} B_i$ where $\beta = \prod_{j=1}^{n}(\beta_j)$. Hence, we may write

$$\sum_{i=1}^{n} \frac{\alpha_i^{k}}{P_i} = \frac{(-1)^{n-1}}{\beta} \sum_{i=1}^{n} \frac{1}{\beta_i^{k-n} B_i}.$$

But, for $k \geq n$ Proposition 5 gives that

$$\sum_{i=1}^{n} \frac{1}{\beta_i^{k-n} B_i} = \sum_{E_{k-n}} \frac{1}{\beta_1^{l_1} \beta_2^{l_2} \cdots \beta_n^{l_n}}.$$

Therefore,

$$\sum_{i=1}^{n} \frac{\alpha_i^{k}}{P_i} = \frac{(-1)^{n-1}}{\beta} \sum_{E_{k-n}} \frac{1}{\beta_1^{l_1} \beta_2^{l_2} \cdots \beta_n^{l_n}}$$

$$= (-1)^{n-1} \alpha \sum_{E_{k-n}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}.$$

□

Proposition 7. Let $n \geq 2$ and $k \geq n-1$. Then

$$\lim_{t \to 0} f_{S_n}^{(k)}(t) = (-1)^{k+n-1} \alpha \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}.$$

Proof. We have from Equation (2),

$$\lim_{t \to 0} f_{S_n}^{(k)}(t) = (-1)^{k} \sum_{i=1}^{n} \frac{\alpha_i^{k+1}}{P_i}.$$

and from Proposition 6,

$$\sum_{i=1}^{n} \frac{\alpha_i^{k+1}}{P_i} = (-1)^{n-1} \alpha \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}.$$

for $k \geq n-1$. Then,

$$\lim_{t \to 0} f_{S_n}^{(k)}(t) = (-1)^{k+n-1} \alpha \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}. \quad □$$

Many authors used the identity

$$\sum_{i=1}^{n}\left[\frac{1}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}\right] = 0$$

and proved it in many long and complicated methods. Here we shall submit a more simple prove. In addition,

we shall find more related identities using the above results.

Proposition 8. Let $n \geq 2$. Then

$$\sum_{i=1}^{n}\left[\frac{1}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}\right] = 0.$$

Proof. Let $n \geq 2$. By Corollary 1, taking $i = 1$ we have $i \leq n-1$, then

$$\sum_{i=1}^{n} \frac{\alpha_i}{P_i} = 0.$$

However,

$$\sum_{i=1}^{n} \frac{\alpha_i}{P_i} = \sum_{i=1}^{n} \frac{\alpha_i \prod_{j=1, j \neq i}^{n}(\alpha_j)}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}$$

$$= \alpha \sum_{i=1}^{n} \frac{1}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)} = 0$$

Therefore,

$$\sum_{i=1}^{n}\left[\frac{1}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}\right] = 0. \quad □$$

Next we shall find a more general equality using our previous results.

Proposition 9. Let $n \geq 2$. Then

$$\sum_{i=1}^{n}\left[\frac{\alpha_i^{k}}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}\right]$$

$$= \begin{cases} 0, & \text{if } 0 \leq k \leq n-2 \\ (-1)^{n-1} \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}, & \text{if } k \geq n-1 \end{cases}$$

Proof. Let $n \geq 2$. Then,

$$\sum_{i=1}^{n} \frac{\alpha_i^{k}}{P_i} = \sum_{i=1}^{n} \frac{\alpha_i^{k} \prod_{j=1, j \neq i}^{n}(\alpha_j)}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}$$

$$= \alpha \sum_{i=1}^{n} \frac{\alpha_i^{k-1}}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}$$

(5)

Suppose that $1 \leq k \leq n$. We have from Corollary 1,

$$\sum_{i=1}^{n} \frac{\alpha_i^{k}}{P_i} = \begin{cases} 0, & \text{if } 1 \leq k \leq n-1 \\ (-1)^{n-1} \alpha, & \text{if } k = n \end{cases}$$

and Equation (5) gives that

$$\sum_{i=1}^{n}\left[\frac{\alpha_i^{k-1}}{\prod_{j=1, j \neq i}^{n}(\alpha_j - \alpha_i)}\right] = \begin{cases} 0, & \text{if } 1 \leq k \leq n-1 \\ (-1)^{n-1}, & \text{if } k = n \end{cases}$$

Replace $k-1$ with k, we obtain the first case and

the case when $k = n-1$, where $\sum_{E_0} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n} = 1$.

Now, suppose $k \geq n$. By Proposition 6,

$$\sum_{i=1}^n \frac{\alpha_i^k}{P_i} = (-1)^{n-1} \alpha \sum_{E_{k-n}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}$$

and the Equation (5) gives that

$$\sum_{i=1}^n \left[\frac{\alpha_i^{k-1}}{\prod_{j=1, j \neq i}^n (\alpha_j - \alpha_i)} \right] = (-1)^{n-1} \sum_{E_{k-n}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}.$$

Also, replace $k-1$ by k we obtain the last case when $n \geq k-1$. □

5. The Main Results

We summarize Proposition 2 and 7 in the following theorem.

Theorem 2. Let $n \geq 2$. Then

$$\lim_{t \to 0} f_{S_n}^{(k)}(t)$$
$$= \begin{cases} 0, & \text{if } 0 \leq k \leq n-2 \\ (-1)^{k+n-1} \alpha \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}, & \text{if } k \geq n-1 \end{cases}$$

Also Corollary 1 and Proposition 5 and 6 can be summarized in the following theorem.

Theorem 3. Let $n \geq 2$ and $k \geq 0$. Then

1) $\sum_{i=1}^n \frac{\alpha_i^k}{P_i} = \begin{cases} 0, & \text{if } 0 \leq k \leq n-1 \\ (-1)^{n-1} \alpha \sum_{E_{k-n}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}, & \text{if } k \geq n \end{cases}$

and

2) $\sum_{i=1}^n \frac{1}{P_i \alpha_i^k} = \sum_{E_k} \frac{1}{\alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}}.$

We recall Propostion 9 in the following corollary of Theorem 3.

Corollary 3. Let $n \geq 2$. Then

$$\sum_{i=1}^n \left[\frac{\alpha_i^k}{\prod_{j=1, j \neq i}^n (\alpha_j - \alpha_i)} \right]$$
$$= \begin{cases} 0, & \text{if } 0 \leq k \leq n-2 \\ (-1)^{n-1} \sum_{E_{k-n+1}} \alpha_1^{l_1} \alpha_2^{l_2} \cdots \alpha_n^{l_n}, & \text{if } k \geq n-1 \end{cases}$$

6. Case of Arithmetic and Geometric Parameters

The study of reliability and performance evaluation of systems and softwares use in general sum of independent exponential R.V. with distinct parameters. The model of Jelinski and Moranda [14], considered that the parameters changes in an arithmetic sequence $\alpha_i = \alpha_{i-1} + d$. Moreover, Moranda [15], considered the model when

α_i changes in an geometric sequence $\alpha_i = \alpha_{i-1} r$. In this section, we study the hypoexponential in these two cases when the parameters are arithmetic and geometric, and we present their pdf.

6.1. Case of Arithmetic Parameters

We first consider the case when $\alpha_i, i = 1, 2, \cdots, n$ form an arithmetic sequence of common difference d.

Lemma 2. For all $1 \leq i \leq n$.

$$\gamma_i = (-1)^{i-1} \frac{(n-1)!}{\binom{n-1}{i-1}} d^{n-1}.$$

Proof. Suppose that α_i form an arithmetic sequence of common difference d. Then $\alpha_j - \alpha_i = (j-i)d$. We have

$$\gamma_i = \prod_{j=1, j \neq i}^n (\alpha_j - \alpha_i).$$

Hence,

$$\gamma_i = (\alpha_1 - \alpha_i)(\alpha_2 - \alpha_i) \cdots (\alpha_{i-1} - \alpha_i)(\alpha_{i+1} - \alpha_i) \cdots (\alpha_n - \alpha_i)$$
$$= (-(i-1)d) \cdots (-2d)(d)(2d) \cdots ((n-i)d)$$
$$= (-1)^{i-1}(i-1)!(n-i)! d^{n-1}$$

However,

$$(i-1)!(n-i)! = \frac{n!}{i\binom{n}{i}} = \frac{(n-1)!}{\binom{n-1}{i-1}}.$$

Then

$$\gamma_i = (-1)^{i-1} \frac{(n-1)!}{\binom{n-1}{i-1}} d^{n-1}, \qquad \square$$

Lemma 3. *For all* $1 \leq i \leq n$.

$$\gamma_i = (-1)^{n-1} \gamma_{n-(i-1)}.$$

Proof. We have from Lemma 2,

$$\gamma_i = (-1)^{i-1} \frac{(n-1)!}{\binom{n-1}{i-1}} d^{n-1}$$

for all $1 \leq i \leq n$. Replace i by $n-(i-1)$, we obtain

$$\gamma_{n-(i-1)} = (-1)^{n-i} \frac{(n-1)!}{\binom{n-1}{n-(i-1)}} d^{n-1}$$
$$= (-1)^{n-i} \frac{(n-1)!}{\binom{n-1}{i-1}} d^{n-1} = (-1)^{n-1} \gamma_i.$$

Thus we obtain the result.   □

Proposition 10. Let $n \geq 2$. Then

$$f_{S_n}(t) = \alpha \sum_{i=1}^{n} \frac{e^{-\alpha_i t}}{\gamma_i} I_{(0,\infty)}(t),$$

where

$$\gamma_i = (-1)^{i-1} \frac{(n-1)!}{\binom{n-1}{i-1}} d^{n-1} = (-1)^{n-1} \gamma_{n-(i-1)}$$

for all $1 \leq i \leq n$.

Proof. We have from Theorem 1

$$f_{S_n}(t) = \sum_{i=1}^{n} \frac{\alpha_i e^{-\alpha_i t}}{\prod_{j=1, j\neq i}^{n} \left(1 - \frac{\alpha_i}{\alpha_j}\right)} I_{(0,\infty)}(t)$$

$$= \sum_{i=1}^{n} \frac{\alpha_i e^{-\alpha_i t} \prod_{j=1, j\neq i}^{n}(\alpha_j)}{\prod_{j=1, j\neq i}^{n}(\alpha_j - \alpha_i)} I_{(0,\infty)}(t),$$

that can be written as

$$f_{S_n}(t) = \alpha \sum_{i=1}^{n} \frac{e^{-\alpha_i t}}{\gamma_i} I_{(0,\infty)}(t),$$

where $\gamma_i = \prod_{j=1, j\neq i}^{n}(\alpha_j - \alpha_i)$ and by the Lemmas 2 and 3 we obtain the result.   □

6.2. Case of Arithmetic Parameters

Next, we consider the case when $\alpha_i, i = 1, 2, \cdots, n$ form a geometric sequence of common ratio r.

Proposition 11. Let $n \geq 2$. Then

$$f_{S_n}(t) = \sum_{i=1}^{n} \frac{\alpha_i e^{-\alpha_i t}}{\prod_{j=1, j\neq i}^{n}(1 - r^{i-j})} I_{(0,\infty)}(t).$$

Proof. We have from Theorem 1,

$$f_{S_n}(t) = \sum_{i=1}^{n} \frac{\alpha_i e^{-\alpha_i t}}{\prod_{j=1, j\neq i}^{n}\left(1 - \frac{\alpha_i}{\alpha_j}\right)} I_{(0,\infty)}(t).$$

Suppose now the parameter α_i form geometric sequence of common ratio r. Then $\alpha_i = \alpha_j r^{i-j}$ and

$$P_i = \prod_{j=1, j\neq i}^{n}\left(1 - \frac{\alpha_i}{\alpha_j}\right) = \prod_{j=1, j\neq i}^{n}(1 - r^{i-j}).   □$$

We may also note that the equalities obtained for P_i represent here a special case and worth mentioning such as

$$\sum_{i=1}^{n} \frac{1}{\prod_{j=1, j\neq i}^{n}(1 - r^{i-j})} = 1.$$

7. Conclusion

The pdf and cdf and some related properties of the hypoexponential distribution with distinct parameters were established. The proofs have been done by using Laplace transform and moment generating function technique. Also with the help of some known computational theorems as Heaviside expansion theorem and multinomial expansion formula the kth order derivative of f_{S_n} and the moment of this distribution of order k were established, in addition for some new related equalities. Eventually, the pdf for models when the parameters α_i are arithmetic and geometric were presented. However the other two cases for hypoexponential distribution when the parameters are equal or not all equal can be studied and observed for future studies. It may be checked if they have the same properties as in this paper.

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