Impact of Volatility and Equity Market Uncertainty on Herd Behaviour: Evidence from UK REITs

Abstract

Purpose

This paper examines herding behaviour among investors in UK-listed Real Estate Investment Trusts (REITs) within three market regimes (low, high and extreme volatility periods) from the period June 2004 to April 2016.

Design/methodology/approach

Observations of investors in 36 REITs that trade on the London Stock Exchange as at April 2016 were used to analyse herding behaviour among investors of shares of UK REITs, employing a Markov regime-switching model.

Findings

Although a static herding model rejects the existence of herding in REITs markets, estimates of the regime-switching model reveal substantial evidence of herding behaviour within the low volatility regime. Most interestingly, we observed a shift from anti-herding behaviour within the high volatility regime to herding behaviour within the low volatility regime, with this having been caused by the FTSE 100 Volatility Index (UK VIX).

Originality/value

The results have various implications for decisions concerning asset allocation, diversification and value management within UK REITs. Market participants and analysts may take into account that collective movements and market sentiment/psychology are determinative factors of risk-return of UK REITs. In addition, general uncertainty in the equity market, proxied by the impact of the UK VIX, may also provide a signal for increasing herding-related risks within UK REITs.

JEL Classification Code: C32, G11, G15

Keywords: Herd behaviour; UK REITs; Markov-switching; Time-varying probabilities.

Decision-making in finance may inevitably involve shortfalls arising from reasons such as information asymmetries or market-/person-specific investment psychology. The literature reveals that irrational features and behavioural aspects of the investment process, ranging from non-financial assets to financial assets are among the reasons for anomalies and boom-bust cycles. In this respect, despite well-established theoretical frameworks and empirical evidence on the efficient market hypothesis, financial markets involve widespread irrationalities. Global-/market-specific uncertainties/volatilities have also exacerbated irrationalities in the leading/developing financial markets. It is not wrong to expect in light of past experiences that market and specific investment psychologies may go hand in hand, which may result in a market-wide roller-coaster effect during up- and down-market conditions in stock markets. As the signal of a crowd movement reflecting mass consensus, herd behaviour in direct/indirect property investments maybe also be analyzed in the context of this irrational/inefficient roller-coaster effect.

Herding behaviour in property stocks has received increasing attention in the property literature. However, studies that have attempted to capture herding behaviour under different market regimes in Real Estate Investment Trusts (REITs) are still at an embryonic stage in the empirical literature. By employing static and dynamic herding models, this study aims to explore herding behaviour in UK-listed REITs over the period June 2004 to April 2016. Based on static modelling, we followed Chang et al's. (2000) methodology, involving cross-sectional absolute standard deviations (CSAD) among individual firm returns, to define non-linear relations between equity return dispersions and market returns. In addition, we allowed the regime transition probabilities to be time varying, by using the time-varying transition probability Markov-Switching model (TVTP-MS). Because the observation period involves nine years with, during and after the Global Financial Crisis period, the evidence may imply that the UK REITs market inherently involves long-run and persistent herding patterns. Due to the highly sophisticated nature of the London Stock Exchange (LSE), this behavioural phenomenon provides interesting evidence of herd behaviour from a developed-market perspective.

The UK REIT market was introduced in January 2007. Its introduction resulted in a number of prior UK-listed property firms converting to REITs, as well as in the initial public offerings of newly formed REITs being listed on the LSE. As at September 2017, the UK REIT market capitalization was valued at around £56 billion¹. The top five UK-listed REITs in market capitalization as at September 2017 included the Land Securities Group plc (£7.21 billion), British Land (£6.14 billion), SEGRO plc (£5.36 billion), Hammerson plc (£4.26 billion) and Intu properties plc (£3.12 billion)². The UK market is a significant global REIT market; for instance, according to the FTSE EPRA/NAREIT global REITs index as at 29 September 2017, UK REITs rank at number four in terms of weightage, with 5.34% having a value of US\$63.28 billion, behind Japan (6.48%, US\$76.75 billion), Australia (6.90%, US\$81.71 billion) and the United States, which ranks number one, with a weight of 64.63% and a value of US\$765.51 billion).

The contribution of this study is twofold. Firstly, to the best of our knowledge, this paper is the

¹ The value of UK REITs was calculated by adding the market capitalization of UK REITs as seen on <u>http://www.bpf.org.uk/reits-and-property-companies</u> as at 30 September 2017.

² The market capitalisation of the UK REIT market as well as the top five UK REITs were sourced from the LSE. http://www.londonstockexchange.com/statistics/companies-and-issuers/companies-and-issuers.htm

³ The information was extracted from the FTSE Russell factsheet (the FTSE EPRA/NAREIT global REITs index) as at 29 September 2017.

first study to examine herding behaviour among investors in UK REITs. The intuition behind this investigation was to determine behavioural aspects in decision-making among investors in UK REITs, connected to the relations among uncertainty, volatility, and herding behaviour. Secondly, the study provides a critical observation for the role of the different regimes on herding behaviour in the UK REIT market. Based on the selected analysis period, we provide comparative knowledge on herding in low, high and extreme market regimes. Third, using timevarying transition probabilities for herding behaviour, we also provide significant knowledge on the shifts between positive and negative herding behaviour during different volatility periods. In doing so, we employed a new framework for analysing the destabilizing effects of herding in the UK REIT market. This analysis provides evidence from a developed country's REIT market, such as that of the UK, which may be found to be interesting as the prior literature reveals that herding is more likely to take place in emerging markets (Zhou and Anderson, 2013), and emerging markets have been found to accommodate higher herding levels compared with their developed market counterparts (Andronikidi and Kallinterakis, 2010). Overall, by focusing on UK REIT stocks during and after the Global Financial Crisis, the study opens a debate on whether UK REIT stocks show irrationalities from a herding perspective, as well as on whether the existing strategies of global portfolio managers and policy makers are compatible with a herding-based market structure.

The paper has four further sections. Section 2 reviews prior studies. Section 3 introduces the data and the testing methodology. The results based on the analysis of cross-sectional absolute standard deviations (CSAD) and time-varying transition probabilities are presented in section 4. Finally, the last section concludes.

2. Literature review

As indicated in Keynes' beauty-contest analogy, stock market investments would be driven by the expectations of other investors, rather than by rational decisions based on analysis of the fundamentals of the asset. This 'animal sprit' may be typically apparent during bubble (Kindleberger and Aliber, 2005; Akerlof and Shiller, 2009) or herding periods connected with mass psychology and irrational price movements in stock markets. The main consensus among theoretical herding studies is that herding can be construed as being either a rational or irrational form of investment behaviour (Zhou and Anderson, 2013).

Herding is broadly perceived as an exuberant and irrational synchronized movement of asset prices that is not justified by their fundamental values (Babalos et al., 2015). Bikhchandani and Sharma (2001) suggest that herding results from an obvious intent by investors to copy the behaviour of other investors and that imperfect information, concern for reputation, and compensation structures are the potential reasons for rational herd behaviour in financial markets. Devenow and Welch (1996) postulate that despite the difficulty of precisely defining herding, it could be defined as behaviour patterns that are correlated across individuals, and it is closely linked to such distinct phenomena as imperfect expectations, fickle changes without much new information, bubbles, fads, frenzies, and sunspot equilibria.

Empirical studies on herding focus on either the behaviour of specific groups (i.e. mutual/pension fund managers, financial analysts) or on the overall market. For example, by examining the quarterly holdings of 155 mutual funds over the period 1975-1984, Grinblatt et al. (1995) found relatively weak evidence for mutual funds tending to buy and sell the same stocks at the same time. The studies selected – Chevalier and Ellison (1999), Graham (1999), Wermers (1999), Welch (2000), Hong et al. (2000), Gleason and Lee (2003), and Clement and Tse (2005) – also provided evidence for group-wide herding. On the other hand, by analyzing 769 funds and

the behaviour of pension managers, Lakonishok et al. (1992) found no market-wide herding, but weak evidence of herding among smaller stocks, and relatively little of either herding or positive-feedback trading among the largest stocks.

Herding analyses under different market regimes provide interesting country-level outcomes. Hwang and Salmon (2004) analyzed herding in the US and South Korean stock markets and found evidence of herding towards market portfolios in both bull and bear markets. The authors further show that, contrary to common belief, the Asian Crisis and particularly the Russian crisis, involved limited herding. Andronikidi and Kallinterakis (2010) found in the case of Israel that the presence of thin trading tends to conceal the actual magnitude of herding. Analyzing Taiwanese open-end equity mutual fund herding behaviour over the period of 1996-2008, Hou et al. (2014) found evidence of both directional and directionless herding, and the abolition of qualified foreign institutional investors has reduced directionless and sell-side herding behaviour in asymmetric (bearish versus bullish contexts) and extreme market conditions, through daily data from the Shanghai and Shenzhen stock exchange markets, and found that a bullish context generates herding behaviour among investors of B-shares, while a bearish situation rather favours crowd movement among A-shares.

Herding in REIT stocks, the particular interest of this study, is a newly developing research area in the literature. As the first attempt to test herding in the REIT market, Zhou and Anderson (2011) investigated market-wide herding behaviour in the US equity REIT market by utilizing the quantile regression method, and found that herding is more likely to be present in the high quantiles of the REIT return dispersion. Authors further indicate that REIT investors tend to herd under turbulent market conditions, and that herding is more likely to occur, and becomes stronger in declining markets, rather than in rising markets, implying asymmetry in herding behaviour. Moreover, the findings also show that during the Global Financial Crisis, REIT investors may not have started to herd until the market became extremely turbulent. By examining the existence of herding effects in the US REIT market during the period of January 2004 to December 2011, Philippas et al. (2013) found that a deterioration of investor sentiment and adverse macro-shocks to REIT funding conditions were significantly related to the emergence of herding behaviour, contrary to the common belief that the recent Financial Crisis did not seem to contribute to this phenomenon. The authors also documented asymmetric herding effects during the days of negative market returns. Babalos et al. (2015) explored herding under low, high and extreme market volatility regimes among US-listed REIT investors during January 2004 and June 2013, and using a regime-switching model, reveal substantial evidence of herding behaviour for the crash regime for almost all sectors, despite the static herding model's rejection of the existence of herding. Moreover, the study suggests a shift from negative herding behaviour during low- and high-volatility regimes to positive herding behaviour under crash regimes for almost all REITs sectors. Using a Markov-switching time-varying parameter (MS-TVP) herding model for South African REITs, Akinsomi et al. (2017a) found that higher levels of gold market speculation considerably contribute to herding behaviour in the South African REIT market, and argue that herding and market volatility creates a vicious cycle in which market volatility contributes to the formation of herding, and herding drives up market volatility, making it especially challenging for policy makers. By utilizing Chang et al's. (2000) methodology over the period of July 2007 to May 2016 for Turkish REITs, Akinsomi et al. (2017b) found herding behaviour, the presence of directional asymmetry and linear relations between volatility and herding. The authors argue that herding is a persistent phenomenon and increases during periods of market stress in Borsa, Istanbul.

Maitland-Smith and Brooks (1999) applied and compared the properties of two regime-switching models for the value indices of commercial real estate in the US and the UK, and found that the Markov-switching model is better able to capture the non-stationary features of the data than the threshold autoregressive model. By employing a Markov-switching model, Krystalogianni and Tsolacos (2004) examined the structure of yields among broad asset classes (real estate, equities and government bonds) and its implications for portfolio allocation decisions and real estate investment. Liow and Zhu (2007) employed a Markov-switching model to characterize real estate security markets' risk-return and detected strong evidence of regimes in the six real estate security markets. In a different research field, Corradin and Fontana (2013) examined the house price dynamics of thirteen European countries, using a Markov-switching error correction model. Lee et al. (2013) applied the bivariate Markov-switching autoregressive model (MSVAR) to identify the turning points of real estate cycles in Taiwan.

Despite a lack of studies on herding, the literature reveals interesting market characteristics of UK REITs. For example, Barkham and Ward (1999) provide evidence regarding the relationship between the NAV discounts of UK property companies and their market capitalizations, based on various hypotheses. Analyzing long memory in the returns and volatility of REITs markets in the USA, UK, Hong Kong, Australia, and Japan, Assaf (2015) confirms that long memory in volatility is real, and is not caused by shifts in variance for all markets. Lee (2013) found a high correlation between the various property types and regions in the UK and raises the question of how well diversified current institutional portfolios are in the UK. Galariotis et al. (2015) found that there have been herding spillover effects from the US to the UK during earlier financial crises and suggest that the drivers of herding behaviour are period- and country-specific.

Analyzing UK REITs seems interesting for a number of reasons. The first point is that the UK real estate industry plays an important role in the national and international economy. As a reflection of this economic role, UK REITs are receiving increasing attention from investors. As indicated by Newell et al. (2016), UK REITs are an important property investment vehicle, being the fourth-largest REIT market globally and having delivered strong risk-adjusted returns since the post-global financial crisis. Second, although our study is the first in the literature, the literature reveals that UK REITs show some anomalies and inefficiencies, implying that the industry may have further implicit irrationalities. For example, Jadevicius Lee (2017) provide evidence that return anomalies exist among UK REITs, and investors can buy and sell them more effectively by recognising the day-of-the-week effect. Morri and Baccarin (2016) investigated the NAV discount puzzle for REITs listed in France, the Netherlands and the United Kingdom between 2003 and 2014. The study suggests that in the UK and France, REITs with more debt are traded at higher discounts, and larger REITs trade at higher discounts in France and the UK, even though the relationship is not significant in all cases.

3. Data and testing methodology

3.1. Cross-sectional absolute standard deviations (CSAD)

Following Chang et al. (2000), this study uses cross-sectional absolute standard deviations (CSAD) among individual firm returns within REITs to define the non-linear relation between the level of equity return dispersions and the overall market return.

The CSAD statistic, used as a measure of return dispersion, is formulated as follows:

$$CSAD_t = \sqrt{\frac{\sum_{i=1}^{N} (r_{i,t} - r_{m,t})^2}{N-1}}$$
(1)

where $R_{i,t}$ and $R_{m,t}$ is the return on stock *i* and the value of an equally weighted average of all REITs returns for period *t*, respectively, and *n* is the number of stocks in the portfolio. Herd behaviour assumes that individual investors make investment decisions following the collective actions of the market, and that these actions will lead security returns to converge with the overall market return. Therefore, herd behaviour implies that security dispersions (i.e. CSAD_t) will decrease with the absolute value of the market return, since each asset becomes similar with regard to sensitivity to the market return.

Chang et al. (2000) suggest that during periods of market stress, one would expect return dispersion (i.e. $CSAD_t$) and market return (i.e. $r_{m,t}$) to have a nonlinear relationship. Christie and Huang (1995) suggest that the probability of herd behaviour increases during periods of market stress and large price movements; therefore, we have a benchmark model based on the following quadratic model of return dispersion and market return:

$$CSAD_{t} = \alpha_{0} + \alpha_{1} |R_{mt}| + \alpha_{2} R_{mt}^{2} + \varepsilon_{t}$$
⁽²⁾

The presence of herding is tested through the following hypotheses: H₀: In the absence of herding effects, we expect in Eq (2) that $\alpha_1 > 0$ and $\alpha_2 = 0$.

 H_{11} . If herding behaviour exists, we expect $\alpha_2 < 0$.

 H_{12} . If anti-herding behaviour exists, we expect $\alpha_2 > 0$.

Because it is suggested in the herding literature that investor herding would be more likely to present itself within sufficiently homogeneous groups of market participants (e.g. Christie and Huang, 1995; Bikhchandani and Sharma, 2001), we focused on securities that are classified as real estate investment trusts. As mentioned earlier, the choice of REITs was largely motivated by the fact that, to the best of our knowledge, no studies exist on herding involving UK REITs. In addition, securitized real estate markets, i.e., REITs, have experienced tremendous growth in the global economy. According to the National Association of Real Estate Investment Trusts (NAREIT), global real estate markets represented more than \$1.22 trillion of equity capitalization in July of 2016. In addition, with REITs being exchange-traded funds that earn most of their income from investments in real estate, REITs have been at the epicentre of research interest because their returns do not suffer from measurement error and high transaction costs compared with other real estate investments. As indicated by Akinsomi et al. (2016), REITs constitute a very good proxy for the real estate market, simultaneously providing high frequency observable data, since REITs shares trade as common stocks. Because REITs are accessible to all investors, irrespective of their portfolio size, this asset class has been particularly successful in attracting investment capital.

For our analysis, we used daily data comprising 36 primary REITs⁴ on the LSE for the period June 2004 to April 2016,⁵ with a total of 3070 observations. The source for the closing

⁴ As at April 2016, there were 36 listed UK REITs. Our sample of firms involves listed REITs on the LSE as at April 2016; for instance, our sample includes REITs such as Land Securities group plc, with a market cap of 7.21 billion pounds as at September 2017, and Redefine International, an offshore REIT with a market cap of 686.64 million pounds that trades on the LSE. Our sample size begins with 16 REITs in June 2004 and ends with 36 REITs on 5 April 2016, highlighting the dynamic nature of our data.

⁵ In this paper, we recognize that the REITs regime begins in 2007. However, we expanded our data to 2004 to extend our timeline: the period from 2004 to 2007 tracks all REITs that converted in 2007; the year 2004 presented a significant count of 16 individual firms for estimating the CSAD. Our robustness tests also examine the REITs regimes solely between 2007 to 2016, and our results remain consistent with earlier results when we employ the

prices of the various REITs is Datastream of Thomson Reuters. In addition, we considered the FTSE 100 VIX (VIX) in estimating the regime transition probabilities of the Markov-Switching model. The VIX data was derived from the same source, with the aim of capturing aggregate equity market uncertainty in the UK.

3.2. The TVTP-MS model with VIX

It is argued that the static model in Equation (2) leads to a misleading conclusion regarding herd behaviour, as parameters are assumed to be constant over time (Balcilar et al, 2013a, b; Ngene, et al., 2017). To distinguish and examine whether herding behaviour is contingent for different market phases, we estimated the following three-state Markov switching model of the cross-sectional returns dispersions:

$$CSAD_{t} = {}_{0,S_{t}} + {}_{1,S_{t}} \left| R_{m,t} \right| + {}_{2,S_{t}} R_{m,t}^{2} + {}_{t}$$
(3)

where $\varepsilon_t \sim iid(0, \sigma_{S_t}^2)$ and S_t is a discrete regime variable taking values in $\{0,1,2\}$ and following a three-state Markov process⁶. The volatility term in Equation (3), ε_t is modelled to be heteroscedastic with three states such that

$$\sigma_t^2 = \sigma_1^2 S_{1t} + \sigma_2^2 S_{2t} + \sigma_3^2 S_{3t}$$
(4)

where $S_{kt} = 1$ if $S_t = k$ and zero otherwise (k = 1, 2, 3). The specification of allowing the volatility term in to be heteroscedastic differentiates market regimes in terms of the level of volatility in each regime, i.e. ${}_{t}^{2} = {}_{k}^{2}$ for regimes k = 1, 2, 3, and allows the variance of cross-sectional dispersions to switch across different regimes. In addition, we allowed the regime transition probabilities to be time varying by using the time-varying transition probability Markov-Switching model (TVTP-MS) to assess the role of uncertainty in the overall UK equity market in herding regimes in the British REIT market. The main advantage of the TVTP-MS model against the constant transition probability specification is that it allows the duration of herd behaviour to vary across different regimes of market volatility and the gauging of fear and market sentiment, as measured by the FTSE 100 VIX index (VIX_{UK}). Hence after modelling the role of VIX_{UK} shock, we could define the transition probabilities of the Markov chain in Equation (3) as:

$$p_{ij,t} = P(S_t = i | S_{t-1} = j, \mathbf{Z}_{t-1})$$
(5)

where \mathbf{Z}_t is a vector of exogenous VIX variables.⁷ We could also define _{*ij*} as the vector of parameters of exogenous variables associated with the transition probability of switching from

period prior to REIT conversion, similar to authors such as Akinsomi et al (2017a), who extended and justified the extension of REIT timelines in the case of South African (SA) REITs, which was not investigated due to data constraints. In addition, we ideally needed data at higher frequencies, to pick up herding and to estimate the Markov-switching model precisely, which tends to have a lot of parameters, especially in our case as we allowed for time-varying transition probabilities. In essence, the starting point of 2004 was driven by the need to use high-frequency firm-level data for the REITs sector of the UK, which begins only in 2007.

⁶ Previous studies found that a three-state Markov process fits the stock return model well (see, for example, Guidolin and Timmermann, 2006; Maheu et al., 2009; and Charfeddine and Ajmi, 2013).

⁷ The variables in \mathbf{Z}_{t} impact the transition probabilities with one lag, since the transition probabilities governing the regime switches that occur from *t*-1 to *t* must be determined at time *t*-1.

state *j* at time *t*-1 to state *i* at time *t*. The time-varying transition probabilities can be written as $p_{ij,t} = \Phi(Z_{ij,t-1}\theta_{ij}), \quad i = 0,1 \text{ and } j = 1,2,3$ (6)

where (.) is the normal cumulative distribution function (CDF), and the transition probabilities satisfy $\sum_{i=0}^{2} p_{ij,t} = 1$ for t=1, 2, ..., T.

Therefore, we include in the TVTP model the vector $\mathbf{Z} = [z_i]$ (*i*=0,1,...,2) in Equation (6) is defined as $\mathbf{Z} = (1, \text{VIX}_{\text{UK}})$, with the UK VIX variable being measured in returns.

4. Empirical results

This section presents the findings for the TVTP-MS model described in Equations (3) through (6). The findings for the static model in Equation (2) are reported in Table 1. First, we find that coefficient α_1 in Equation (2) is positive and statistically significant, as predicted by the equilibrium model of CAPM, and the cross-sectional absolute deviation of REITs returns with respect to the market return is increasing with the absolute magnitude of market returns. Second, we find anti-herding behaviour, as illustrated by the statistically significant coefficient α_2 , even though the magnitude is small.

Table 2 presents our findings for the TVTP-MS model specified in Equations (4) through (6). As is evident, the TVTP-MS model is clearly a better fit to the data than is the static model, as the former has a much lower AIC.⁸ This result is not surprising given that we obtained strong evidence (the highest possible level of significance at all possible dimensions involved in the test) of nonlinearity when we applied the Brock et al. (1996, BDS test) to the residuals of the static model (Equation 2). The results are reported in Table A1, in the Appendix of the paper. In addition, we detected as many as four breaks (3 May 2006; 27 May 2008; 2 March 2010; and 16 February 2012) in this equation when we implemented the test of multiple structural breaks, based on the global information criteria as developed by Bai and Perron (2003).

The regime-specific volatility estimates $\binom{2}{k}$, k=1, 2, 3) are reported in Table 2. Market regimes are clearly identified in terms of low (i.e. regime 2), high (i.e. regime 1) and extreme volatility (i.e. regime 3) in terms of the level of return, with the low volatility regime being primarily concentrated post the financial crisis, especially for 2011. Our main finding is that there is significant evidence of herding in the UK REITs market during the low volatility regime, which is opposite to what was detected for the US REITs market by Babalos et al. (2015), who found strong evidence of herding in the crash-regime. Our results suggest that in the UK, herding occurs when uncertainty, i.e., volatility is low, with anti-herding being observed at high and crash-regimes of volatility. However, most important is the observation that, unlike the linear static model, the TVTP-MS model detects evidence of herding in a specific regime, which happens to be the low-volatility regime. Thus our result similar to that of Babalos et al. (2015).

In January 2007 the legislation set out the rules for REITs in the United Kingdom. As a result, a number of listed property groups converted to this regime; this change provided opportunities for the growth of the property investment sector, because property companies could now get

⁸ The TVTP-MS model's AIC was also lower than that of the MS model, with constant probabilities of transition, with the latter having an AIC of 0.4965. Complete details of the results from the MS model's constant transition probabilities, which were qualitatively similar to those of the TVTP-MS model, are available upon request from the authors. We chose to work with the TVTP-MS model due its better fit, and to use the VIX in explaining the movements of the transition probabilities.

access to capital markets, and investors would have wider investment opportunities than with alternative asset classes, due to the underlying property assets, without significant tax leakage.

Therefore, we carried out a robustness check for the period 2007-2016 to see whether the results are robust before and after 2007; in particular, we found that the results for 2007-2016 are similar compared with the results in Table 1. For the TVTP-MS model, the results for 2007-2016 are similar to those for 2004-2016 (i.e. our main finding is that there is significant evidence of herding in the UK REITs market during the low volatility regime)⁹.

The first break, which occurred on 3 May 2006, is associated with the implementation of the 2007 rules for REITs in the United Kingdom, when a number of listed property groups converted to REITs. This regime change enabled UK REITs to undertake activities other than running a property rental business (for example, to be involved in property trading or services where a minimum of 75% of the business entails running a property rental business). The UK REIT regime is set out in Part 4 of the Finance Act 2006, and the date that Royal Assent was received was 19 July 2006, which is one month after the date of the first break, on 3 May 2016. The date (i.e. 27 May 2008) of the second break is associated with the global financial crisis, including the nationalization and splitting of Bradford & Bingley and the part-nationalization of RBS and Llovds TSB. The date of the third break was 2 March 2010, which is associated with the Corporation Tax Act 2010, Part 12. This Act received Royal Assent on 3 March 2010. The act specified that "no one property or leasehold interest can account for more than 40% of the fair value of the gross assets of the property rental business" and that "In the accounting period, at least 75% of a REIT's total income-profits (before tax) must arise from its tax-exempt property rental business". The date of the last break (i.e. 16 February 2012) is also associated with a significant REIT regime change, namely the draft legislation that was published on 29 March 2012. The change reduced the entry barriers and increased the incentives for investors to invest in REITs, which included the abolition of the 2% entry charge for companies entering the REITs.

[Insert Tables 1 and 2 Here]

4.1. Persistence of market regimes

The estimated regime durations in Table 2 indicate that the low volatility regime is the most persistent, with the longest average duration across market regimes. We observed that the longest average duration of the low volatility regime is 174 days for the All Equity REIT sector. This suggests that the low volatility regime is the most persistent, while the average duration for the extreme volatility regime is 14.9 days, as it has the most frequent regime switches. Our findings are consistent with the current literature on herding, using MS models (see, for example, Balcilar and Demirer, 2013)

The transition probability estimate p_{ij} and its relevant smoothed probabilities, plotted in Figure 1, provide a visual examination of the dynamic nature of regime transitions and herd behaviour in the UK REITs market. The smoothed regime probabilities for the three-regime nonlinear TVTP-MS model is plotted in Figures (a)-(c).

The smoothed probability plots generally suggest a low-high-extreme (LHE) volatility

⁹ Detailed results are available upon request.

transition order, in which the high-volatility regime (i.e. regime 1) follows the low volatility regime (i.e. regime 2), and the extreme volatility/crash regime (i.e. regime 3) follows the high volatility regime (i.e. regime 1). This finding is consistent with the evidence for advanced markets and provides market regulators with a warning signal before the extreme volatility regime (see, for example, Babalos et al., 2015). It is evident that the crash regime is followed by the high volatility one¹⁰. Another interesting pattern is exhibited by the fact that from late 2010 the market regime had entered a period of low volatility (regime 2).

4.2. VIX and time-varying transition probabilities

As explained earlier, the parameters, θ_{ij} , i = 0,1 and j = 1, 2, 3. in Equation (6) capture the dynamic effects of the UK VIX return on transition probabilities across regimes. Significant parameter estimates imply that the VIX plays a role in leading the UK REITs market from one regime to another, possibly driving herding regimes. As described earlier, the l^{th} element of the vector \hat{i}_{ij} , that is, $\hat{i}_{ij,l}$ for i = 0,1 and j = 1, 2, 3, is defined as $\{l = 0 \text{ (constant)}, 1 \text{ (VIX}_{\text{UK}} \text{ return})\}$ with two parameter estimates for the variable.

We find that the VIX_{UK} is significant in driving regime transitions in the UK REITs market, as indicated by the significant $\theta_{21,1}$ estimate. Our attention was drawn to the significant transition probability estimates for switching from the high volatility regime (i.e. regime 1) to the low volatility regime (i.e. regime 2), where the herding takes place. We therefore conclude that the UK VIX does play a role in driving regime transition from high to low volatility.

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[Insert Figure1 Here]
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5. Conclusion

Due to globally increasing investment volumes in property, real estate has become an important asset class since the 1990s. The Global Financial Crisis and the recent Brexit shock have also showed that direct and indirect real estate investments in the UK are also highly sensitive to uncertainties. This picture makes it of paramount importance to understand the risk-return characteristics of UK REIT stocks for asset/portfolio managers and policy makers. The market facts also confirm this approach. According to the LSE data, the market value of REITs – involving diversified, specialty, retail, industrial and offices, residential, and diversified REITs – is worth £ 43,544 million, and the overall market value of real estate holding and development, real estate investment trusts, and real estate service sub-sectors is worth £ 82,386 million as at 31 November 2016.¹¹ Moreover, the British Property Federation and Toscafund Asset Management (2016) estimated that the market value of commercial real estate was £1,662 billion, just over 20% of net wealth, and contributed £94bn to GDP in 2014 in the UK.

As a first in the literature, the study employs static and dynamic models to explore herding in UK REITs over the period June 2004 to April 2016. The study provides various elements of evidence and interesting implications of herding behaviour in UK REITs.

From a methodology perspective, the study first suggests, in parallel with Babalos et al. (2015), that a TVTP-MS model is a better fit for the data than is the static model. In this respect, the

¹⁰ For example, we observed that UK REITs were in the low volatility regime until the beginning of 2007, and then from the beginning 2007 to the middle of 2008, the market return was driven mainly by the global financial crisis. The market was dominated by extreme volatility between the end of 2008 and early 2009.

¹¹ Available at: <u>http://www.londonstockexchange.com/statistics/companies-and-issuers/companies-and-issuers.htm.</u>

study defines the importance of modelling nonlinearity in a herding analysis. Second, given the importance of nonlinearity, the static model, which suggests anti-herding behaviour, is econometrically misspecified. Third, although the static herding model rejects the existence of herding, the Markov regime-switching model defines three market regimes, namely the low, high, and crash volatility regimes, and provides evidence of herding behaviour under the low volatility regime but anti-herding behaviour in the high and crash regimes of volatility. In the presence of nonlinearity, the Markov-switching model is clearly the correctly specified econometric framework, and should be relied upon to draw inferences. Fourth, the evidence further suggests that the low volatility regime is the most persistent market regime, with the longest average regime duration involving 174 days, primarily in the post-2011 period (and to some extent before the global financial crisis). Interestingly, this low-volatility herding period essentially coincides with the bull market conditions of the LSE. Fifth, the model outcomes also suggest a low-high-extreme (LHE) volatility transition order, and that the UK VIX does play a role in driving regime transition from high to low volatility. In this respect, the high volatility regime follows the low volatility regime, and the extreme volatility/crash regime follows the high volatility regime. This herding cycle may translate as a shift from anti-herding behaviour during high volatility regimes to herding behaviour under low volatility regimes.

These results have implications for decisions concerning asset allocation and portfolio diversification in the UK REITs market. First, regarding the order of regime transitions, moving from low to high and to extreme volatility suggests that the market follows a consistent pattern, which warns investors and regulators of potential or imminent extreme volatility in the UK REIT market. This behavioural pattern may provide significant foresight for market participants about changes in REITs returns, depending on the consistent chain pattern in the REITs market. This behaviour in the UK REIT market is similar to those of general stocks in developed markets, where volatility transmits from low to high to crash (Balcilar et al., 2013). This result, however, contradicts the findings of studies on developing markets, such as those of the Gulf States, which shows that the market moves from low volatility to extreme volatility to high volatility (Balcilar et al., 2013). Second, defining the low volatility regime of UK REITs as the most persistent herding market regime in the rising period of the LSE suggests another interesting market insight. This evidence implies that when the general stock market is doing well, which in turn corresponds to low volatility, and hence, lower risks, agents in the REITs sector tend to herd, i.e. to behave similarly. However, when the markets are highly volatile and risky, economic agents operating in the REITs sector tend to behave differently from one another, in an attempt to maximize their profits. Third, the source of the fluctuations in risk in the REITs sector originates from the VIX, i.e. aggregate equity market risks, which spill over into the REITs sector as well, implying that there are no diversification opportunities between conventional equities and the REITs sector. Therefore, market participants in LSE and UK REITs may perceive that the risk profile in the overall stock market and the property sector may be interconnected.

Overall, the aspects of evidence in this study collectively imply that UK REITs market participants may improve their decision-making by utilizing the herding characteristics and cycles of the REITs market during different volatility regimes, in addition to signals of overall market behaviours through the LSE market index and the UK VIX.

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	Tab	le 1. Estimates	of the static 1	nodel		
α ₀	α_1	α_2	RSS	logL	AIC	adj.R ²
0.6975***	0.7320***	0.03784***	1149.17	-2847.37	1.857	0.8314

Note: The table reports the estimates for CSAD in Equation (2). All estimations were done using the ordinary least squares (OLS) with the Newey-West heteroskedasticity and autocorrelation consistent (HAC) standard errors. RSS denotes residual sum of squares, log L denotes the log likelihood of the OLS model, AIC denotes the Akaike information criterion, and adj. R^2 denotes the adjusted coefficient of determination. *** represent significance at the 1% level. A significant and positive α_2 estimate implies anti-herding behaviour.

Parameter	All Equity REITs
$lpha_{0,1}$	1.0574
$lpha_{0,2}$	0.5746
α _{0,3}	2.1972
$lpha_{\mathrm{l,l}}$	0.2666***
$lpha_{1,2}$	0.6346***
$\alpha_{1,3}$	0.0721
	Herding coefficients
$\alpha_{2,1}$	0.0473***
$lpha_{2,2}$	-0.0734***
$lpha_{2,3}$	0.1002***
	Regime volatilities
$\sigma_{\rm l}$	0.0700^{***}
σ_2	0.0245^{***}
σ_3	0.8391**
	Time-varying transition probabilities
$ heta_{11,0}$	3.8434***
$ heta_{11,1}$	0.0813
$\theta_{21,0}$	0.8527***
$ heta_{21,1}$	0.1858*
$\theta_{12,0}$	-0.2056***
$ heta_{12,1}$	-0.0379
$\theta_{22,0}$ -	6.0929***
$ heta_{22,1}$	0.0326
$-\theta_{13,0}$	-2.8175***
$ heta_{13,1}$	-0.0597
$\theta_{23,0}$	-4.6643***
$\theta_{23,1}$	0.0223
	Regime durations
$ au_{\mathrm{l}}$	37.5
$ au_2$	174
$ au_3$	14.9
AIC	0.431
$\log L$	-643.521

Table 2. Estimates for the regime-based herding model with the UK VIX

Notes: This table presents the estimates of the three regime TVTP-MSH model given in Equations (3) through (6). The asterisks ***, ** and * represent significance at the 1%, 5%, and 10% levels, respectively.

	Table A1. Es	timates of the st	tatic model (2007-2016)	I	
α ₀	α_1	α2	RSS	logL	AIC	adj.R ²
0.76431***	0.72106***	0.03813***	1052.35	-2424.7	2.0088	0.8396

 Table A1. Estimates of the static model (2007-2016)

Note: The table reports the estimates for CSAD in Equation (2). All estimations were done using the ordinary least squares (OLS) with the Newey-West heteroskedasticity and autocorrelation consistent (HAC) standard errors. RSS denotes residual sum of squares, log L denotes the log likelihood of the OLS model, AIC denotes the Akaike information criterion, and adj. R^2 denotes the adjusted coefficient of determination. *** represents significance at the 1% level. A significant and positive α_2 estimate implies anti-herding behaviour.

Parameter	All Equity REITs
$lpha_{0,1}$	0.87334***
$lpha_{0,2}$	1.85006****
$lpha_{0,3}$	0.66114***
$\alpha_{1,1}$	0.73757***
$lpha_{1,2}$	0.5531***
$\alpha_{1,3}$	0.4751***
	Herding coefficients
$\alpha_{2,1}$	-0.1064***
$lpha_{2,2}$	0.0426***
$lpha_{2,3}$	-0.0351***
	Regime volatilities
$\sigma_{\rm l}$	0.0999***
σ_2	1.8448^{***}
σ_3	0.01414**
	Time-varying transition probabilities
$ heta_{11,0}$	2.7332***
$ heta_{11,1}$	0.0562
$ heta_{21,0}$	2.9816**
$\theta_{21,1}$	-0.2267**
$ heta_{12,0}$	-1.399***
$\theta_{12,1}$	0.00412
$ heta_{22,0}$	6.3322***
$ heta_{22,1}$	-0.1168
$ heta_{13,0}$	-2.9491***
$ heta_{13,1}$	0.0536**
$ heta_{23,0}$	-22.693
$ heta_{23,1}$	0.0033
	Regime durations
$ au_1$	14.4
$ au_2$	29.6
τ ₃	21.4
AIC	0.647
$\log L$	-757.449

Table A2.	Estimates	for the re	gime-base	ed herding	model	with l	UK V	/IX ((2007-2)	016)

Notes: This table presents the estimates of the three regime TVTP-MSH model given in Equations (3) through (6). The asterisks ***, ** and * represent significance at the 1%, 5%, and 10% levels, respectively.

Figure 1. Return and smoothed probability of three-regime nonlinear TVTP-MS model for UK REITs



a) Smoothed Probability: Regime 1

b) Smoothed Probability: Regime 2







d) Time-Varying Markov Transition Probabilities

Table A1. BDS test on residual o	f
Equation 2 (Static Model)	

	BDS				
Dimension	Statistic	Std. Error	z-Statistic	Prob.	
2	0.0652	0.0020	32.7078	0.0000	
3	0.1203	0.0032	37.9167	0.0000	
4	0.1571	0.0038	41.5370	0.0000	
5	0.1794	0.0039	45.4485	0.0000	
6	0.1907	0.0038	50.0301	0.0000	
		0.00.50			
Raw ej	psilon	0.0058			
Pairs within epsilon		6628888	V-Statistic	0.7033	
Triples within epsilon		1.59E+10	V-Statistic	0.5499	
			C(1.n-(m-	c(1.n-(m-	c(1.n-(n
Dimension	C(m,n)	c(m,n)	1))	1))	1))^k
2	2634800	0.5597	3310342	0.7032	0.4944
3	2200863	0.4678	3307738	0.7031	0.3475
4	1886874	0.4013	3305260	0.7030	0.2442
5	1649088	0.3510	3302663	0.7029	0.1716
6	1461537	0 3113	3300162	0 7028	0 1205

Note: *m* stands for the number of (embedded) dimensions that embed the time series into *m*-dimensional vectors, by taking each *m* successive point in the series. The BDS *z*-statistic tests for the null of *i.i.d.* residuals.