1 Opposite-sex associations are linked with annual fitness, but sociality is stable over

- 2 lifetime.
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17 Abstract

18 Animal sociality, an individual's propensity to associate with others, has fitness 19 consequences through mate choice, for example, directly, by increasing the pool of 20 prospective partners, and indirectly through increased survival, and individuals benefit 21 from both. Annually, fitness consequences are realised through increased mating success 22 and subsequent fecundity. However, it remains unknown whether these consequences 23 translate to life-time fitness. Here, we quantified social associations and their link to fitness 24 annually and over lifetime, using a multi-generational, genetic pedigree. We used social network analysis to calculate variables representing different aspects of an individual's 25 26 sociality. Sociality showed high within-individual repeatability. We found that birds with 27 more opposite-sex associates had higher annual fitness than those with fewer, but this did 28 not translate to lifetime fitness. Instead, for lifetime fitness, we found evidence for 29 stabilizing selection on opposite-sex sociality, and sociality in general, suggesting that 30 reported benefits are only short-lived in a wild population, and that selection favours an 31 average sociality.

- 32 Keywords: sociality, social network analysis, selection, social behaviour, genetic pedigree,
- 33 long-term fitness, de-lifing, lifetime reproductive success
- 34

36 Background

37

38 Some individuals are consistently more sociable than others, demonstrated by within-39 individual repeatability of social traits across vertebrate groups (Aplin et al. 2015; Thys et 40 al. 2017; Dimitriadou, Croft and Darden 2019; Plaza et al. 2019; Beck, Valcu, and 41 Kempenaers 2020; Proops et al. 2021; Strickland et al. 2021). This variation in individual 42 sociality is positively linked with fitness in some taxa (Silk 2007; Silk et al. 2009) and is therefore expected to be subject to selection (Krause and Ruxton 2002). Fitness is a relative 43 44 measure of an individual's genetic contribution to the population in the next generation, 45 and thus, can only be represented comprehensively and precisely by measures of traits spanning an organism's lifetime (Endler 1986; Blankenhorn 2010; Reid et al. 2019; Moiron, 46 47 Charmantier and Bouwhuis 2022). A comprehensive definition of fitness is fundamental to 48 understand the evolutionary pressures that shape variation in sociality. In practice, 49 however, many studies must rely on fitness correlates instead (e.g., number of broods, or 50 survival, instead of genetic contribution). For example, in mammal societies, both variation 51 in within- and between-sex affiliations (Archie et al. 2014) have been linked to lifetime 52 fitness correlates, survival and longevity (Cameron, Setsaas and Linklater 2009; Silk, 53 Alberts and Altman 2003: Silk et al. 2010: Stanton and Mann 2012, but also see Thompson 54 and Cords 2018). Whereas, in birds, the subject of this study, the use of fitness correlates 55 (eggs laid, chicks fledged, within-year survival etc.) are frequent over more precise fitness 56 measures (Moiron, Charmantier and Bouwhuis 2022), that require intensive field work 57 over a long period of time.

Although some benefits are linked with sociality during the breeding period most of these
tend to be short-term and contextual (Bebbington et al. 2017; Riehl and Strong 2018).
Instead, benefits associated with reproduction are often linked with non-breeding sociality
(Firth and Sheldon 2016; Kohn 2017; Maldonado-Chaparro et al. 2018; Beck, Farine, and
Kempenaers 2020), when group cohesion is stronger (Silk et al. 2014; Kurvers 2020).
Sociality may influence fitness in different ways, through benefits to reproductive success
or increased survival, and so, the mechanism of selection acting on social traits may also

65 differ. For example, communal foraging between socially associated individuals during the non-breeding period facilitates resource information transfer (Aplin et al. 2012; Firth, 66 Sheldon and Farine 2016; Hillemann et al. 2020) and reduces predation risk (Cresswell 67 68 1994; Cresswell and Quinn 2011; Sorato et al. 2012), increasing survival. However, these benefits may also incur costs associated with competition for resources and mate choice 69 70 (Birkhead and Biggins 1987; Le Galliard et al. 2005; Forstmeier et al. 2011; Mayer and 71 Pasinelli 2013; Grant and Grant 2019; but also see Lea et al. 2010). Sociality may also 72 benefit individuals who hold more central social network positions or have access to 73 opposite-sex associates, through enhansed mate choice (McDonald 2007; Oh and Badyaev 74 2010; Firth et al. 2018; Beck, Farine, and Kempenaers 2021). Although the association 75 between annual fitness correlates and non-breeding sociality has been well described, 76 testing how selection acts on social traits requires lifetime fitness measures, and remains

77 unresolved.

78 With the recent development of tools to construct and analyse social networks (Wey et al. 79 2008; Farine and Whitehead 2015) the study of sociality has become popular among 80 behavioural ecologists. Yet, to describe the association between sociality and fitness any 81 potential study must first overcome two problems: (1) A social association must be clearly 82 defined relative to the behavior of the study system (Figure 1; Psorakis et al. 2012, 2015), 83 and (2) to study the evolution of social behaviour, precise measures of individual fitness 84 must be quantified to test for correlation with a social trait – selection. Although annual fitness correlates are widely used, lifetime fitness is more precise and without the 85 86 stochasticity of annual measures (Dobson, Murie and Viblanc 2020; Alif et al. 2022) and 87 thus, can better describe selection pressure acting on a trait (Endler 1986; Blankenhorn 88 2010: Reid et al. 2019). However, lifetime fitness requires wild animals to be monitored 89 throughout their whole lives, and all breeding attempts, and the fates of their offspring, 90 must be recorded to determine recruitment. All of these require a multi-generational, 91 genetic pedigree (Kruuk 2004; Korsten et al. 2013).

92 Our study system, an island population of house sparrows *Passer domesticus* (hereafter,
93 sparrow/s) where we monitor all individuals from birth to death, without capture bias

94 (Simons et al. 2015), overcomes both problems. (1) We have sociality data from birds that

95 are electronically registered visiting a feeder. Social centrality measures are repeatable

96 across different timescales and contexts in this and other populations (Plaza et al. 2019).

97 (2) We have lifetime recruitment data available, from a multi-generational genetic pedigree

98 that, because our population is closed, meaningly there is no movement of sparrows to or

99 from the island, and that our study covers all sparrows on the island, we can use to

100 compute precise annual and lifetime fitness estimates (Schroeder et al. 2015; Alif et al.

101 2022).

102 We tested predictions based on arguments presented above to understand the potential for

selection on sociality: (1) We confirmed that the social traits we measured were

104 meaningful by testing for individual repeatability of sociality; (2) We tested the prediction

105 that non-breeding sociality has fitness benefits, either driven by reproductive success

106 through opposite-sex association or through increased survival through network centrality

107 measures and is subject to selection.

108

109

110 Methods

111

112 Study system

113 We used data from the Lundy sparrow system, a long-term study based on the island of 114 Lundy (51.11N, 4.40W), ≈19 km off North Devon, UK, in the Bristol Channel. The sparrows 115 on Lundy breed in nest boxes, sited in groups around the only village on the island. The 116 island is rodent-free and therefore the sparrows have no predators but for the occasional 117 vagrant raptor. House sparrows are a model organism in behavioral ecology and evolution, 118 and much is known about their biology, physiology and life-history (Andersson 1994; 119 Sánchez-Tójar et al. 2018). House sparrow are socially monogamous, but 25% of broods 120 show they can be genetically promiscuous (Schroeder et al. 2016). On Lundy, they have on 121 average 2-3 broods of 4-5 eggs per breeding season (Westneat et al. 2014). The sex ratio is 122 stable, and the mean lifespan of recruits is three years (Alif et al. 2022). Although sparrows 123 are territorial during the breeding season, during the non-breeding period they form

gregarious groups that forage communally for seed and at supplementary bird feeders
(Summer-Smith 1963), both of which are available year-round on Lundy.

126 Most sparrows were first captured, and tissue sampled in nest boxes at their natal site 127 during the breeding season (April to August) or using mist nets during the post-fledging 128 period (Schroeder et al. 2011; Girndt et al. 2019). Tissue samples were either blood or 129 mouth swabs and were stored in ethanol and refrigerated at 3°C prior to analysis. We 130 genotyped sparrow DNA at <22 microsatellite loci suitable for parentage assignment in 131 sparrows (Dawson et al. 2012). Using the genetic data, we assembled a near-complete 132 genetic pedigree (Schroeder et al. 2015, 2016), which at the time of writing spans twenty 133 years, 2000–2019, and 8,379 individuals. We fitted all sparrows with a unique 134 combination of a coded metal ring and three coloured leg rings. We also provided each 135 sparrow with a subcutaneous Passive Identification Transponder (PIT tag; TROVANID100: 136 11.5×2.1 mm and 0.1 g), under the skin of the breast, which we have previously shown 137 have no detrimental effect on subsequent fitness (for details see Schroeder et al. 2011). 138 These tagged individuals were then recorded when they visited a custom made 19.8cm x 139 19.8cm Radio Frequency Identification antenna (RFID; DorsetID) mounted on a seed 140 reservoir (for photo see Sánchez-Tójar et al. 2017; Brandl et al. 2019), positioned centrally 141 within our study site. The feeder was open access, and explicitly not limited to a single bird 142 feeding at one time, as is the case at hanging bird feeders (Youngblood 2019; Beck, Farine 143 and Kempenaers 2020). Our feeder recorded visiting birds every day that the island 144 generators were running (6am – midnight, seven days a week).

145

146 Social centrality measures

To quantify the sociality of individual sparrows we calculated measures of social centrality
(hereafter centrality measures) using presence data from the RFID antenna, collected
during the non-breeding periods, November–January, of 2015/16 (15 weeks) and 2016/17
(13 weeks, hereafter referred to events). An association between birds observed from
these data can reflect either, individuals that choose to maintain some social cohesion,
which we consider to be a non-random association, or individuals without pre-established
social cohesion who coincide in time and space, which we consider random mixing. These

data record the presence of birds without distinguishing between these two types of
association. Further, dyadic interactions can also be aggressive interactions, and our data
cannot exclude these cases. Although in sparrows', dominance hierarchies are typically
linear, there is no evidence for a correlation with reproductive fitness (Sánchez-Tójar
2018).

159

160 The 'gambit of the group' is a common approach used to identify discrete groups among all 161 associating individuals (Whitehead and Dufault 1999; Figure 1A). However, given the 162 gregarious nature of sparrows and the high activity at our feeder, at which non-discrete 163 groups of sparrows accumulated, the gambit of the group approach overestimated 164 associations between individuals (Figure 1A; also see Ferreira et al. 2020). One solution to 165 this would be to define a non-random association where two individuals overlapped by a 166 defined period (Δ_t) at our bird feeder. However, in our system, again due to the near 167 constant visitations, this resulted in linear network structures, e.g., linking the first bird to 168 the second, then the second to the third, and so on (Figure 1B). To account for the social 169 behaviour of sparrows, we derived a method to infer non-random associations that 170 assumed that non-random social associations are established before they attend the feeder 171 (suggested by Summer-Smith 1963). We therefore defined an association as two 172 individuals that arrived to feed within 150 seconds (Δ_t) of each other. Here an arrival is 173 defined as the (re)appearance of the individual at the feeder after being absent for a period 174 of minimally >300 seconds (Δ_l). We defined that $\Delta_t = 150$ seconds was sufficient to detect 175 and link individuals who arrive together in a group (see Figure 1C), and the resulting data 176 better sampled non-random associations between individuals in our system, from 177 watching sparrows in the field and on pre-recorded footage, arriving at our feeder (Plaza et 178 al., 2019).

From the resulting association matrices of the two events, 2015/16 and 2016/17, we built a series of weighted, non-directional, social networks (hereafter, network/s), where the vertices represent individual sparrows and interconnecting edges their associations. First, we built individual networks for each of the 15 weeks in 2015/16, and 13 weeks in 2016/17, to estimate within-individual repeatability in centrality network metrics. This
was to validate these centrality measures against the individual repeatability already
demonstrated in this system using both RFID and video data (Plaza et al. 2019). Then, we
built a network for each event to extract non-breeding sociality. Finally, we also built two
bipartite networks from each event (sub-graphs), which only considered association
strength between opposite-sex individuals.

189 From the first non-breeding networks we extracted three centrality measures representing 190 different aspects of sociality for each individual using the 'iGraph' R package (Csardi and 191 Nepusz 2006): We selected centrality measures to reflect aspects of an individual's social 192 preference, following similar studies on sociality (McDonald 2007; Farine and Whitehead 193 2015; Beck, Farine and Kempenaers 2021). Degree represents the number of associates, 194 and may impact fitness through enhanced mate choice, where individuals position 195 themselves alongside others of lesser quality (Oh and Badyaev 2010;), or where same-sex 196 associates benefit reproduction in cooperative breeding species (Bebbington et al. 2017; 197 Riehl and Strong 2018). Strength represents the quality of those relationships and may 198 influence the structure and behaviour of reproductive communities (Firth and Sheldon 199 2016; Culina, Firth and Hinde 2020). We calculated strength following (Farine 2013), using 200 the sum of dyadic Simple Ratio Indices (the association probability between a dyad, from 0, 201 never associated, to 1, always associated), which we transformed to give a measure of net 202 association quality (Eq. 1, where S denotes strength, $d_{(i)}$ and sum SRI_(i) the degree and SRI 203 of a given individual respectively and N(V) = the number of vertices in a network)

$$(Eq. 1) \quad S = d_{(i)} \frac{SRI_{(i)}}{N(V)}$$

We used eigenvector centrality (following McDonald 2007, hereafter centrality) to quantify
the influence of an individual to all others within the network (Newman 2004; Oh and
Badyaev 2010); Finally, we extracted opposite-sex degree from the two bipartite subgraphs, the number of opposite-sex associations, which we used to represent an
individual's pool of potential reproductive partners, as a fourth measure of sociality
(following Beck, Farine and Kempenaers 2021). Given the high density of sparrows visiting

- 210 the feeder, and the frequency at which those birds were detected we did not threshold our
- 211 networks (using only a sample of the birds attending an antenna, following Farine and
- 212 Whitehead 2015) to maintain network structure, although only birds who arrived in a dyad
- 213 (where degree > 0) were included in our networks.
- 214
- 215 Fitness measures
- 216 For each of the sparrows that survived to the following breeding period we used the
- 217 genetic pedigree to calculate two fitness measures, and for each of these, we calculated one
- annual measure, and one across the lifetime of the sparrow. Both fitness measures are
- 219 based on the number of recruits, and we have shown that they correlate well with
- 220 reproductive value, thus represent fitness in an evolutionarily meaningful way (Alif et al.
- 221 2022). We defined recruits as offspring that survived and produced genetic offspring
- themselves.
- (1) *Recruits*. For the number of annual recruits, we summed individual recruits within the
 breeding year following the social events. We then again summed individual recruits across
 a lifetime, or up to 2020 as a measure of lifetime recruits. Note that the latter category only
 contained five sparrows that were still alive at the point of census, and as such, our
 recruitment data can be considered near complete. We excluded birds which did not
 survive to breed, and yearlings which had zero recruits.
- (2) *De-lifed fitness*. As a second fitness measure, we used de-lifed fitness (p_{ti} ; Eq. 2), which estimates an individual's genetic contribution to the population (Coulson et al. 2006). Delifed fitness is a retrospective measure of realized fitness, relative to the population each year, calculated by removing (de-lifing) an individual and its offspring from the pedigree and recalculating the resulting change in population growth.

$$(Eq.2) \quad p_{ti} = \frac{\xi_{t(i)} - w_{t(i)}}{N_t - 1}$$

Here, p_{ti} is the contribution of individual (*i*) to population growth during a specific period *t*. Further, $\xi_{t(i)}$ is a measure of individual performance, here the number of surviving offspring of individual *i* at the end of the breeding period *t*. We added a value of one if the

individual *i* itself survived to the next breeding period *t* + 1. The population size at time *t* is

238 N_t at the beginning of each breeding cycle (here April). To estimate the individual's

239 contribution to population growth, we use w_t , which represents the ratio of the population

size at t + 1 to the population size at t. This de-lifed fitness p_{ti} is an annual value per

individual, and we calculated it for all birds which produced at least one recruit. We then

also summed p_{ti} , within individuals as a lifetime de-lifed fitness measure, p_i .

243

244

245 Individual repeatability in sociality

246 First, we validated our centrality measures by confirming that they were repeatable within-

247 individuals between weeks. We modelled degree, strength, centrality and opposite-sex

248 degree respectively, as response variables each in four Gaussian linear mixed models

249 against the intercept, and with bird identity modelled as a random effect to compensate for

250 repeat identities between years. We then divided the variance explained by bird identity by

251 the total phenotypic variance of the trait to quantify repeatability (see Nakagawa and

252 Schielzeth 2010). We ran repeatability models using package MCMCglmm default

253 parameters and priors – the models converged robustly and reliably.

254

255 Selection on sociality – annual and lifetime fitness

256 We quantified the association between centrality measures and fitness. As all four

257 centrality measures are inherently correlated, we modeled each separately to avoid

collinearity (Webster, Schneider and Vander Wal 2020). For each centrality measure, we

ran two models, with annual recruits and annual de-lifed fitness as the response variables.

260 In the models with annual numbers of recruits as response, we assumed a Poisson error

261 distribution with a log link function, and in the models explaining de-lifed fitness we

assumed a Gaussian error distribution. We mean-centered all centrality measures within

263 each year, eliminating between-year differences, and modelled them as fixed covariates.

264 We also added each sociality variable as a quadratic effect to test for stabilizing or

disruptive selection where averages are favored over the extremes (Wolf et al. 2007). Bird

266 identity was modelled as a random effect on the intercept to account for pseudo-

267 replication, and cohort to account for environmental stochasticity. We modelled fixed

- 268 effects for sex (male, 1 or female, 0) and age (in calendar years) and age as a quadratic
- 269 effect, to account for variation in fitness as explained by demography (Schroeder et al.
- 270 2012). We added sex as an interaction term with age to account for the extra-pair
- 271 behaviour of older males (Girndt et al. 2018).

272 We modelled lifetime recruits and lifetime de-lifed fitness in the same way as the annual

273 ones, but instead of age we used lifespan, or maximum age at year 2020. Because each bird

was only represented once in this dataset, we only modelled cohort as a random effect.

275 We used Bayesian Markov Monte-Carlo simulations, using MCMCglmm (Hadfield 2010), to

276 run all models. For all models we used inverse Wishart priors for random effects, and ran

each over 343,000 iterations, with a burn-in of 3,000 and a thinning interval of 200. We

visually checked the posterior trace plots for all model outputs and ensured that

autocorrelation was below 0.1 and effective sample sizes between 1,000 and 2,000. The

280 fixed effects were considered statistically significant when the 95% credible interval (CI) of

281 its posterior distribution did not span zero.

282

283 Null models and dominance interactions

284 We ran a node-permutation null-model by shuffling the identities of birds visiting the 285 feeder between existing arrival times in our association matrices, thereby breaking any link 286 between sociality and fitness, over 1000 randomized permutations (following Farine, 287 2017). We used these randomized association matrices to construct 1000 new networks 288 and extracted the mode for our four centrality measures. We used these randomized 289 centrality measures to re-run all fitness models. Finally, to exclude the possibility that 290 dominance was interacting with our observed centrality measures, we tested for 291 correlations between the centrality measures and dominance from videos recorded during 292 the same period of our social network events. We represented individual dominance by 293 calculating ELO ratings, based on antagonistic interactions at the bird feeder (for further 294 details see Sánchez-Tójar 2018). We did not include the randomised centrality measures 295 from our null models in these correlations.

297 Results

298 The data consisted of 150 individual birds making 410,114 visits to the RFID feeder within 299 our study period (mean = 2,734 visits per bird, SD = 8,116), across both events. Excluding 300 birds that died prior to the start of our study or those that were ringed after, 160 tagged 301 birds survived in our system in November 2015, an additional 90 birds were tagged prior 302 to the 2017 breeding period, although not all survived to sampling. After constructing the 303 arrival networks, we identified 3,783 associations between 118 PIT tagged individuals 304 during 2015/2016, and 874 associations between 69 individuals in 2016/2017. These 305 networks contained 66.3% of 122 and 26.3% of 205 breeding birds in 2016 and 2017 306 respectively. Combined, we had 130 records for annual and lifetime fitness from 102 307 individuals, 33 were recorded in both years (for summary statistics see Table 1). Degree 308 and opposite-sex degree are closely correlated, implying that those with more opposite-sex 309 associates also tend to have more associates of either sex (supplementary material table 1)

310

- 311 Individual repeatability in sociality
- 312 We confirmed individual repeatability by week in all four centrality measures between 15
- 313 weeks in 2015/16: Degree, R=0.29 (0.15 0.39), Strength, R=0.22 (0.10 0.32), Centrality,
- 314 015 (0.03 0.27) and, Opposite sex degree, 0.27 (0.13 0.4); and, 13 weeks in 2016/17:
- 315 Degree, R=0.29 (0.15 0.39), Strength, R=0.22 (0.10 0.32), Centrality, 015 (0.03 0.27)
- 316 and, Opposite sex degree, 0.27 (0.13 0.4) (Table 1).

318 Selection on sociality – annual and lifetime fitness

- 319 Opposite-sex degree had a statistically significant linear relationship with the number of
- 320 annual recruits and annual de-lifed fitness. Strength and centrality had a negative quadratic
- 321 association with annual recruitment (Figure 2; Table 2; Table 3). Age and sex both also
- 322 predicted annual recruits, with younger individuals and females recruiting more offspring.
- 323 Age also positively predicted annual de-lifed fitness (Table2; Table3).
- 324
- 325 None of our centrality measures statistically significantly predicted lifetime recruitment
- 326 (Supplementary material table 4), however, all four had a statistically significant negative
- 327 quadratic relationship with lifetime de-lifed fitness (Figure 2).
- 328
- 329
- 330 Null models and dominance interactions
- 331 We found no link between fitness and sociality, nor any evidence of selection from our null
- 332 models (supplementary material figure 3). Likewise, dominance was not strongly
- 333 correlated with any centrality measure, implying that our method of assigning associations
- based on arrival time, rather than shared space at a bird feeder, is unlikely to be influenced
- 335 by dominance (supplementary material table 1).
- 336

337 Table 1. Summary statistics for recruitment and centrality measures for individual house sparrows on Lundy Island during

two non-breeding events (November–January 2015/2016 and 2016/2017). Each measure is given as mean, standard

| 339 | deviation and sample size | , and including repeatability (R) and | 95% CI for centrality measures. |
|-----|---------------------------|---------------------------------------|---------------------------------|
|-----|---------------------------|---------------------------------------|---------------------------------|

| Summary statistics of recruitment | | | | | | |
|---|----------|------------------|------------------------------|--------------|------------------------------|--|
| Fitness Variable | Scale | 201. Mean, SD | 2015/16 Mean_SD (n = 118) | | 2016/17 Mean, SD (n = 69) | |
| | Annual | 2.65, | 2.65, 1.98 | | 0.78, 1.07 | |
| Recruits | Lifetime | 5.34, 3.76 | | 4.15, 3.86 | | |
| Summary statistics and repeatability of centrality measures | | | | | | |
| Centrality measure | | Mean, SD | R, 95CI | Mean, SD | R, 95CI | |
| Degree | | 64.11, 32.03 | 0.29, 0.15-0.39 | 25.33, 17.37 | 0.35, 0.16-0.49 | |
| Strength | | 0.06, 0.05 | 0.22, 0.10-0.32 | 0.09, 0.13 | 0.33, 0.16-0.48 | |
| Centrality | | 0.35, 0.30 | 0.15, 0.03-0.27 | 0.25, 0.31 | 0.23, 0.10-0.36 | |
| Opposite-sex degree | | 32.00, 15.86 | 0.27, 0.13-0.4 | 12.18, 9.19 | 0.45, 0.24-0.61 | |

340

Table 2. Annual recruitment model outputs from GLMMs for each of our four centrality measures (Degree, Strength, Centrality

343 and Opposite-sex degree), derived of 410,114 visits to an RFID baited feeder by 150 individuals. Centrality measure of house

sparrows on Lundy Island, modeled against annual recruits. We inferred significance where the 95% CI did not span zero,

345 positive effects on the response variable are highlighted in red, and negative in blue.

| Annual recruits | Posterior mode 95% credible intervals | | | |
|---------------------|--|-------------|-------------|---------------------|
| Variable | Degree | Strength | Centrality | Opposite-sex degree |
| Fixed terms | | | | |
| (Intercent) | 3.49 | 3.69 | 4.19 | 3.78 |
| (intercept) | 1.46—6.11 | 1.58—6.46 | 1.60—6.31 | 1.58—6.50 |
| Controlity moscuro | 0.17 | 0.22 | 0.26 | 0.25 |
| Centrality measure | -0.01-0.34 | -0.03-0.44 | -0.01—0.45 | 0.05-0.41 |
| Contrality massuro? | -0.1 | -0.12 | -0.16 | -0.04 |
| Centrality measure- | -0.26—0.06 | -0.32—0.07 | -0.34-0.08 | -0.16-0.10 |
| Sov (male) | -0.67 | -0.7 | -0.74 | -0.57 |
| Sex (male) | -1.41—0.05 | -1.46—-0 | -1.48—-0 | -1.47—0 |
| ٨٥٥ | -1.14 | -1.15 | -1.27 | -1.07 |
| Age | -2.03—-0.30 | -2.04—-0.28 | -1.99—-0.26 | -2.00—-0.26 |
| A go? | 0.05 | 0.04 | 0.06 | 0.03 |
| Age- | -0.07—0.17 | -0.08—0.17 | -0.08—0.16 | -0.07—0.17 |
| Ago*Sov (malo) | 0.21 | 0.22 | 0.28 | 0.21 |
| Age Sex (male) | -0.01-0.48 | -0.02—0.46 | -0.02-0.49 | 0—0.49 |
| Random terms | | | | |
| *Bird ID | 0 | 0 | 0 | 0 |
| DITUTD | 0—0.13 | 0—0.12 | 0—0.12 | 0—0.13 |
| *Cohort | 1.88 | 1.5 | 2.09 | 2.29 |
| COHOIC | 0.25—15.69 | 0.19—15.57 | 0—15.89 | 0.11—16.90 |
| Posiduals | 0 | 0 | 0 | 0 |
| ICSIUUAIS | 0-0.27 | 0—0.28 | 0—0.29 | 0-0.24 |

347 Table 3. Lifetime recruitment model outputs from GLMMs for each of our four centrality measures (Degree, Strength,

348 Centrality and Opposite-sex degree), derived of 410,114 visits to an RFID baited feeder by 150 individuals. Centrality measure

of house sparrows on Lundy Island, modelled against lifetime recruits. We inferred significance where the 95% CI do not span

350 zero, positive effects on the response variable are highlighted in red, and negative in blue. († Age in lifetime models denotes

351 _____either lifespan, or age in 2020, whichever is greatest)

| Lifetime recruits | | Posterior mode, 95% credible intervals (lower – upper) | | |
|---------------------------------|------------|--|------------|---------------------|
| Variable | Degree | Strength | Centrality | Opposite-sex degree |
| Fixed term | | | | |
| (Intercent) | 0.85 | 0.85 | 0.99 | 0.8 |
| (intercept) | 0—1.93 | 0.02—1.95 | -0.06—1.87 | -0.03—1.83 |
| Centrality measure | 0.1 | 0.1 | 0.12 | 0.14 |
| Generality measure | -0—0.25 | -0.09—0.3 | -0.07—0.31 | -0.05—0.25 |
| Centrality measure ² | -0.05 | -0.01 | -0.07 | -0.03 |
| dente anty meabare | -0.19—0.11 | -0.24—0.11 | -0.24—0.11 | -0.16-0.12 |
| Sex (male) | 0.33 | 0.28 | 0.28 | 0.48 |
| ben (male) | -0.35—0.99 | -0.36—0.91 | -0.33—0.98 | -0.24—1.03 |
| +Aσe | 0.1 | 0.11 | 0.09 | 0.12 |
| 11.60 | 0—0.23 | 0—0.24 | 0.01—0.25 | 0—0.24 |
| +Age*Sex (male) | -0.06 | -0.07 | -0.07 | -0.06 |
| Tige sen (mare) | -0.20—0.07 | -0.19—0.07 | -0.2—0.06 | -0.21-0.04 |
| Random effects | | | | |
| *Cohort | NA | 0.47 | 0.45 | 0.39 |
| Conort | 0.07—2.78 | 0.07—2.76 | 0.06—3 | 0.10-2.77 |
| Residuals | NA | 0.12 | 0.11 | 0.14 |
| ite situais | 0—0.27 | 0.01—0.29 | 0.02-0.28 | 0—0.26 |

352

353 Discussion

354 We found evidence for annual fitness benefits of sociality, where individuals with more

- 355 opposite-sex associates had higher fitness in the breeding period, than those with fewer,
- 356 but that this did not translate to lifetime fitness. For lifetime fitness, we found evidence for
- 357 stabilising selection on sociality, including opposite-sex degree, suggesting that such
- 358 benefits are only short-lived, or contextual, in a wild population.
- 359 We constructed our networks by linking dyads of birds that arrived together to a bird
- 360 feeder, but ignored the time that they spent there, to eliminate most random associations.
- 361 Other studies have adapted similar approaches in high-density and open feeder systems, or
- 362 have considered the same implicit problems (Gomes, Boogert and Cardoso, 2021). Ferreira
- 363 et al. (2020), for example, identified flocks arriving, but then defined associations by spatial
- 364 proximity at a series of feeder boxes. Further research could optimize our approach for
- 365 other systems, either by refining the time after which an individual is determined to have
- 366 left the feeder (Δ_i), or similarly, the time it takes for all members of a group to interact with
- 367 the feeder upon arrival (Δ_t). Further work may also consider defining associations only
- 368 where a dyad visit together more often than would be expected by chance but doing so
- 369 must also consider some method of retaining peripheral associations. Although RFID
- 370 systems sample sociality well at a feeder, we cannot be sure that sociality traits are
- 371 maintained in other contexts future works might consider tracking social behaviour
- across time and space (For example see Ripperger and Carter 2021).

373 Where previous studies on wild birds have suggested links between aspects of sociality and 374 annual reproductive success (for examples see Firth and Sheldon 2016; Kohn 2017; Beck, 375 Farine, and Kempenaers 2020), we were also able to use lifetime measures, which better 376 reflect the genetic contribution of the individual to population growth, and thus, fitness. 377 This also allowed us to also describe how selection acts upon sociality across the 378 population. We found that sociality had little influence on fitness at the annual scale, apart 379 from for opposite-sex association, which was linked to increased recruitment and de-lifed 380 fitness. Our study corroborates that annual fitness benefits described elsewhere,

381 particularly regarding mate choice (Oh and Badyaev 2010; Beck, Farine, and Kempenaers 382 2021, Beck, Farine, and Kempenaers 2020) directly translate into increased annual fitness. 383 At the lifetime scale, our study also provides some insight into the evolution of social 384 behaviors, which we found to be maintained at the population average through stabilizing selection. We are therefore, to the best of our knowledge, the first study to link sociality 385 386 with lifetime fitness benefits in a wild bird (but see Formica et al. 2021; Philson and 387 Blumstein 2022). Our results may also suggest a mechanism for selection on sociality 388 through enhanced mate choice, but the impact on survival was difficult to determine in this 389 study. Sociality is predicted to increase survival through reduced predation risk or 390 information transfer (Sorato et al. 2012; Hillemann et al 2020), but we found no evidence 391 to suggest that either was selected for, through higher centrality, in our analyses.

Stabilizing selection in this case may be driven by factors such as high mate fidelity or
changing sociality with age (Oh and Badyaev 2010; Albery et al. 2022), removing the need
to constantly maintain opposite-sex associations over lifetime while maintaining individual
fitness. However, those opposite-sex associations may also be beneficial in an extra-pair
context from the male perspective (Beck, Farine, and Kempenaers 2020) and requires
further research.

398 Our centrality measures were associated with lifetime, but not with annual de-lifed fitness, 399 and only opposite-sex degree was associated with recruitment at the annual scale. We 400 found no relationship between social centrality and dominance in our study using arrival 401 time to define sociality, but aggressive interactions are probably also reduced over the 402 non-breeding period (Summer-Smith 1963). None of our centrality metrics were linked 403 with recruitment at the lifetime scales. Overall, de-lifed fitness better represents fitness as 404 it is a relative measure of the contribution to population growth (Alif et al. 2022). The 405 number of recruits, while intuitively appealing, is not relative, and in good years, more 406 birds may have a higher number of recruits, while in poor years, having one recruit may be 407 an achievement. As such, this measure is not always comparable between years and may 408 explain our results. Further, recruitment is also dependent on parental effects and 409 relationships within the breeding season, which were not quantified here, although they

410 have been suggested elsewhere (Bebbington et al. 2017; Riehl and Strong 2018), whereas

411 de-lifed fitness also captures long-term survival. We found that older males recruited more

412 offspring, likely by virtue of older males siring more extra-pair offspring (Girndt et al.

413 2018). Likewise, younger birds had lower annual de-lifed fitness, because younger birds

414 had not recruited any offspring in previous years that would contribute to their current de-

415 lifed fitness.

416 In conclusion, we suggest a link between opposite-sex association and reproductive success

417 at the annual scale, suggesting a mechanism for selection to shape social behaviour. At the

418 lifetime scale we suggest that selection on sociality is stable, suggesting greater fitness for

- 419 those at the population mean, in a wild population of passerine birds.
- 420
- 421
- 422

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431

432 Data availability

433 Analyses reported in this article can be reproduced using the data provided by Dunning,

434 Jamie et al. (2022), Opposite-sex associations are linked with annual fitness, but sociality is

435 stable over lifetime, Dryad, Dataset, https://doi.org/10.5061/dryad.z08kprrhd

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bioRxiv preprint doi: https://doi.org/10.1101/2022.01.04.474937; this version posted December 8, 2022. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license. Figure 1. Three versions of a simulated event (A, B and C) show the interval over which five individuals (1–5, black/grey bars) spent at a resource over time (t), and the derived social networks from each approach: A = gambit of the group, which links all individuals in a discrete group equally; B = time-window overlap (by Δ_t), which links individuals who overlap at a resource; and C = arrival time (developed for this study), which links individuals who arrive together to a resource. Red boxes denote the time period during which individuals are considered to be associated, and overlaps represent an association: A, all individuals within a group; B, where they are physically present at the same time (red box), or shortly after they depart to account for birds which were present, but not currently being recorded by the antenna, in that case, over-lapping by Δ_t (red over-hanging box, typically a few seconds); or, C, where they arrive within Δ_t of each other, but the subsequent time spent at a feeder is irrelevant. However, note that the function of Δ_t differs between B and C; Where in B, Δ_t functions to detect when birds are in the same place but where one (or more) are not currently being detected by the antenna, in C the function is to link all individuals which arrive together, while ignoring those already present at the resource, which has the potential to link two separate groups in A and B. In the case of C, an additional interval (Δ_i) is required to define when birds have left the resource, after which they can be recorded as arriving again.

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Figure 2. De-lifed fitness as response variables against Centrality measures from 8 linear mixed models, at two scales, from the 737 738 Lundy Island house sparrows, derived of 410,114 visits to an RFID baited feeder by 150 individuals: Explanatory variables for Annual de-lifed fitness (A); and Lifetime de-lifed fitness (B), where ² denotes a guadratic function, also shown in the four 739 740 adjacent panels for A and B, and their 95% credible intervals. Credible intervals are given as solid bars for each explanatory variable, where a solid point denotes the posterior mode. Black bars denote no effect on the response variable; red denote a 741 742 positive and blue, a negative, relationship with the response. In adjacent panels, quadratic functions of each response variable 743 presented in A and B (on the Y axis: A Centrality, Degree, Opp. Degree, Strength, and B Centrality, Degree, Opp. Degree, 744 Strength). Blue curves represent a negative interaction with fitness measures (on the X axis). Measures with no effect are not 745 shown in figure. We found no link to sex, and age was also subject to stabilising selection, (given in supplementary material 746 table 4).

