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# Inhomogeneous random graphs, isolated vertices, and Poisson approximation

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## Abstract

Consider a graph on randomly scattered points in an arbitrary space, with any two points  $x, y$  connected with probability  $\phi(x, y)$ . Suppose the number of points is large but the mean number of isolated points is  $O(1)$ . We give general criteria for the latter to be approximately Poisson distributed. More generally, we consider the number of vertices of fixed degree, the number of components of fixed order, and the number of edges. We use a general result on Poisson approximation by Stein's method for a set of points selected from a Poisson point process. This method also gives a good Poisson approximation for U-statistics of a Poisson process.

## 1 Introduction

In the *inhomogeneous random graph* (IRG), each vertex has one of several possible types or *states*, where the space of possible states may be infinite. Given the states of the vertices, each possible edge is present with a probability that depends on the states of the two endpoints, independently of the other edges. This provides a very flexible class of random graph models; for example, the state of a vertex could represent spatial location, or it could represent the time at which a vertex is born, for a randomly evolving graph.

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Such models are popular in statistical network modelling, where they go under names such as *stochastic block model* [24] (in which case the state space is usually taken to be finite) and *latent space model* [11] or *latent variable model* [23]. The IRG terminology dates from [25], and is prevalent in the more probabilistic literature, where the model has been studied in depth, for example in [4] and [2]. Much of this literature is concerned with the birth of the giant component, but full connectivity has also been studied in [6].

For our purposes, the IRG is defined as follows. Let  $(\mathbb{X}, \mathcal{F}, \mu)$  be a probability space (the state space). For  $s > 0$  let  $\mathcal{P}_s$  be the (random) set of points of a Poisson point process on  $\mathbb{X}$  with mean measure (i.e., intensity measure)  $s\mu$ . Also, for  $n \in \mathbb{N}$ , let  $\mathcal{X}_n$  be the binomial point process consisting of  $n$  independent random elements of  $\mathbb{X}$  which common distribution  $\mu$ .

Suppose  $\phi : \mathbb{X} \times \mathbb{X} \rightarrow [0, 1]$  is a measurable symmetric function; we call such a function a *connection function*. Given finite  $\mathcal{V} \subset \mathbb{X}$  (possibly with multiplicity), let  $G(\mathcal{V}, \phi)$  be the random (undirected) graph with vertex set  $\mathcal{V}$ , with each pair  $\{x, y\}$  of points of  $\mathcal{V}$  connected by an edge with probability  $\phi(x, y)$ , independently of all other pairs. We are particularly interested in  $G(\mathcal{P}_s, \phi)$  and  $G(\mathcal{X}_n, \phi)$ ; we define these graphs more formally in Section 4.

In the special case where  $\mathbb{X}$  is a region of Euclidean space and  $\phi(x, y)$  is determined by the displacement  $x - y$  (typically via  $\|x - y\|$ , where  $\|\cdot\|$  denotes the Euclidean norm), the IRG is also known as the *soft random geometric graph* [22] or *random connection model* (RCM) [18]. If, in the Euclidean setting, we have  $\phi(x, y) = \mathbf{1}_{\{\|x-y\| \leq r\}}$  the IRG is known as the *random geometric graph* (RGG) [19] or *Gilbert graph*. These models are important in applications to wireless communications; see for example [7, 8, 10, 17, 26].

For any graph  $G$  and any  $j \in \mathbb{N}_0 := \mathbb{N} \cup \{0\}$ , let  $D_j(G)$  denote the number of vertices in  $G$  of degree  $j$ ; also set  $D_{\leq j}(G) := \sum_{i=0}^j D_i(G)$ . In particular,  $D_0(G)$  is the number of isolated vertices. Of interest is the question of whether  $D_{\leq k}(G(\mathcal{P}_s, \phi))$  and  $D_{\leq k}(G(\mathcal{X}_{\lfloor s \rfloor}, \phi))$  are approximately Poisson distributed for  $s$  large, with  $k$  fixed, e.g.  $k = 0$ . One reason for interest is that  $D_0(G) = 0$  (respectively  $D_{\leq k}(G) = 0$ ) is clearly a necessary condition for  $G$  to be connected (resp.  $(k + 1)$ -connected, assuming  $G$  has at least  $k + 2$  vertices), and for many choices of  $\mathbb{X}$  and  $\phi_n$  this condition is asymptotically sufficient (in probability) when  $G = G(\mathcal{X}_n, \phi_n)$  with  $n$  large; see [9, 12, 22, 21, 20]. In such cases,  $\mathbb{P}[D_{\leq k}(G(\mathcal{X}_n, \phi_n)) = 0]$  is a good approximation for the probability that  $G(\mathcal{X}_n, \phi_n)$  is  $(k + 1)$ -connected, so if we can estimate the former via Poisson approximation, we may have a useful estimate for the probability of  $(k + 1)$ -connectivity.

The references just mentioned show that in many cases where  $s$  is large and  $\phi$  is chosen so that  $\mathbb{E} D_j(G(\mathcal{P}_s, \phi)) = O(1)$ , the distributions of  $D_j(G(\mathcal{P}_s, \phi))$  and of  $D_j(G(\mathcal{X}_{\lfloor s \rfloor}, \phi))$  are approximately Poisson. In this paper we give a general criterion

for this to be the case, without making any geometrical or topological assumptions on the space  $\mathbb{X}$  (Theorem 2.1). We give similar results for the number of components of order  $k$  with  $k$  fixed (Theorems 2.2 and 2.3), and for the number of edges (Theorem 2.4). These theorems also incorporate asymptotic normality, when the mean of the variable in question grows slowly as a function of  $s$ .

We prove Theorems 2.1, 2.2, 2.3 and 2.4 in Sections 5, 6 and 7. The proofs use Theorem 3.1, a general result on Poisson approximation for functionals of Poisson processes, which is of independent interest. This theorem also gives us further results on the Poisson approximation for the number of edges under different assumptions from those of Theorem 2.4. See Theorem 7.1 and the subsequent discussion. It also provides an alternative way to derive (and slightly improve) a result of [5] on Poisson approximation of  $U$ -statistics of a Poisson process; see Section 8.

## 2 Results on inhomogeneous random graphs

Let  $\Phi$  be the class of all measurable symmetric functions from  $\mathbb{X} \times \mathbb{X}$  to  $[0, 1]$ , and for  $\varepsilon > 0$  let  $\Phi_\varepsilon$  be the class of all  $\phi \in \Phi$  that satisfy

$$\inf_{x \in \mathbb{X}} \left( \int_{\mathbb{X}} \phi(x, y) \mu(dy) \right) \geq \varepsilon \sup_{x \in \mathbb{X}} \left( \int_{\mathbb{X}} \phi(x, y) \mu(dy) \right).$$

If  $\phi \in \Phi_\varepsilon$  we say the connection function  $\phi$  is  $\varepsilon$ -homogeneous. Note that 1-homogeneity, according to our definition, is the same as homogeneity as defined in [6].

Several of our results require  $\varepsilon$ -homogeneity. There are many interesting classes of connection function which satisfy this condition. For example, suppose  $\mathbb{X}$  is a bounded convex region in Euclidean space  $\mathbb{R}^d$ , and  $\mu$  has a density (with respect to Lebesgue measure on that region) that is bounded away from 0 and infinity. Then there exists  $\varepsilon > 0$  such that all connection functions of the form  $\phi(x, y) = \psi(|x - y|)$  with  $\psi$  nonincreasing lie in  $\Phi_\varepsilon$ . In particular, we do not require any exponential decay condition on  $\psi$ , such as that imposed in [22].

For  $\phi \in \Phi$ , set  $\bar{\phi} = \sup_{x, y} \phi(x, y)$ . As well as  $\varepsilon$ -homogeneity, some of our results also require that  $\bar{\phi}$  be bounded away from 1. This condition is annoying since it excludes from consideration the standard RGG, and also some cases of the IRG which have been considered elsewhere in the literature [4, 6]. On the other hand, ‘soft’ versions of these models, which do satisfy this condition, are easily defined and arguably will often be reasonable from the point of view of applications (we shall discuss this further below). Without this condition, it may be unrealistic to expect to find a simple argument for Poisson approximation of  $D_0$  covering a large class of RGG densities without getting involved in geometrical details of any particular probability density or even assuming any Euclidean structure at all, which is what our condition on  $\bar{\phi}$  allows us to do.

Suppose we have connection functions  $\phi_s$  defined for each  $s > 0$  with  $\mathbb{E} D_{\leq j}(G(\mathcal{P}_s, \phi_s))$  bounded as  $s \rightarrow \infty$ , for some fixed  $j \in \mathbb{N}_0$ . See (5.1) below for a formula for  $\mathbb{E} D_j(G(\mathcal{P}_s, \phi_s))$ . Our first result shows that  $D_j(G(\mathcal{P}_s, \phi_s))$  and  $D_j(G(\mathcal{X}_{[s]}, \phi_{[s]}))$  are approximately Poisson and  $D_k(G(\mathcal{P}_s, \phi_s))$  is approximately normal for  $k > j$ , for large  $s$  under the condition that  $\phi_s \in \Phi_\varepsilon$  for all  $s$  and  $\bar{\phi}_s$  is bounded away from 1.

For  $\alpha \in (0, \infty)$ , let  $Z_\alpha$  denote a random variable having the Poisson distribution with parameter  $\alpha$ . Let  $\mathcal{N}$  denote a random variable having the standard normal distribution in  $\mathbb{R}$ , i.e. with probability density function  $(2\pi)^{-1/2} \exp(-x^2/2)$ ,  $x \in \mathbb{R}$ . For any graph  $G$  we define  $D_{\leq -1}(G) := 0$ . For any two  $(0, \infty)$ -valued functions  $u(s)$  and  $v(s)$ , we say  $u(s) = O(v(s))$  if  $\limsup_{s \rightarrow \infty} u(s)/v(s) < \infty$ , and  $u(s) = o(v(s))$  if  $\limsup_{s \rightarrow \infty} u(s)/v(s) = 0$ , and  $u(s) = \Theta(v(s))$  if  $u(s) = O(v(s))$  and  $v(s) = O(u(s))$ .

**Theorem 2.1.** *Let  $j, k \in \mathbb{N}_0$ ,  $\varepsilon > 0$  and  $\phi_s \in \Phi_\varepsilon$  with  $\bar{\phi}_s \leq 1 - \varepsilon$  for  $s \in (0, \infty)$ , satisfying  $\lim_{s \rightarrow \infty} \mathbb{E} D_{\leq j}(G(\mathcal{P}_s, \phi_s)) = \alpha \in (0, \infty)$ . Then as  $s \rightarrow \infty$ ,*

$$D_j(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha; \quad D_j(G(\mathcal{X}_{[s]}, \phi_{[s]})) \xrightarrow{\mathcal{D}} Z_\alpha, \quad (2.1)$$

and also

$$D_{\leq j-1}(G(\mathcal{P}_s, \phi_s)) \xrightarrow{P} 0; \quad D_{\leq j-1}(G(\mathcal{X}_{[s]}, \phi_{[s]})) \xrightarrow{P} 0, \quad (2.2)$$

and moreover

$$\mathbb{E} [D_k(G(\mathcal{P}_s, \phi_s))] = \Theta((\log s)^{k-j}) \quad (2.3)$$

and if also  $k > j$  then

$$\frac{D_k(G(\mathcal{P}_s, \phi_s)) - \mathbb{E} [D_k(G(\mathcal{P}_s, \phi_s))]}{\sqrt{\mathbb{E} [D_k(G(\mathcal{P}_s, \phi_s))]} } \xrightarrow{\mathcal{D}} \mathcal{N}. \quad (2.4)$$

It is interesting to compare the conclusions (2.1) and (2.2) of this result with the example on page 55 of [19]. In that case, for a certain sequence of RGGs one can arrange for the number of vertices of degree 2 to be asymptotically *compound* Poisson, whereas here it is asymptotically Poisson.

For  $k \in \mathbb{N}$ , and for any graph  $G$ , we refer to the components of  $G$  of order  $k$  (i.e., with  $k$  vertices) as the  $k$ -*components* of  $G$ . Let  $N_k(G)$  denote the number of  $k$ -components in  $G$ . In particular,  $N_1(G) = D_0(G)$ .

Suppose we have connection functions  $\phi_s$  defined for each  $s > 0$  with  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s))$  bounded as  $s \rightarrow \infty$ , for some fixed  $k \in \mathbb{N}$ . Theorem 2.3 below shows that  $N_k(G(\mathcal{P}_s, \phi_s))$  and  $N_k(G(\mathcal{X}_{[s]}, \phi_{[s]}))$  are approximately Poisson under the condition that  $\phi_s \in \Phi_\varepsilon$

and  $\bar{\phi}_s \leq 1 - \varepsilon$  for all  $s$  and some fixed  $\varepsilon > 0$ ; even without this condition, Theorem 2.2 shows that the Poisson approximation holds provided that  $\bar{\phi}_s = o(1/\log s)$  (in the case of  $G(\mathcal{P}_s, \phi_s)$ ), or provided that  $\bar{\phi}_{\lfloor s \rfloor} = o(s^{-1/2})$  (in the case of  $G(\mathcal{X}_{\lfloor s \rfloor}, \phi_{\lfloor s \rfloor})$ ). These rates of decay imposed on  $\bar{\phi}_s$  are significantly milder than the condition  $p_n = \Theta((\log n)/n)$  for the Erdős-Rényi random graph  $G(n, p_n)$  to be at the threshold for having no isolated vertices.

We also give a result on asymptotic normality, when  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s))$  grows slowly as  $s \rightarrow \infty$ . All asymptotics in the next two theorems are as  $s \rightarrow \infty$ .

**Theorem 2.2.** *Let  $k \in \mathbb{N}$  and  $\phi_s \in \Phi$  for  $s > 0$ , with  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha \in (0, \infty)$ .*

- (a) *If  $\bar{\phi}_s = o(1/\log s)$ , then  $N_k(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha$ .*
- (b) *If  $\bar{\phi}_s = o(s^{-1/2})$ , then  $N_k(G(\mathcal{X}_{\lfloor s \rfloor}, \phi_{\lfloor s \rfloor})) \xrightarrow{\mathcal{D}} Z_\alpha$ .*

**Theorem 2.3.** *Let  $k \in \mathbb{N}$ ,  $\varepsilon > 0$  and  $\phi_s \in \Phi_\varepsilon$  for  $s > 0$ , with  $\bar{\phi}_s \leq 1 - \varepsilon$  for all  $s$ .*

- (a) *If  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha \in (0, \infty)$ , then*

$$N_k(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha; \quad N_k(G(\mathcal{X}_{\lfloor s \rfloor}, \phi_{\lfloor s \rfloor})) \xrightarrow{\mathcal{D}} Z_\alpha. \quad (2.5)$$

- (b) *If  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \infty$ , but  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) = o(s)$ , then setting  $\alpha_s = \mathbb{E} N_k(G(\mathcal{P}_s, \phi_s))$ , we have that  $(N_k(G(\mathcal{P}_s, \phi_s)) - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ .*

Among other things, the case  $k = 1$  of Theorem 2.2 (a) generalizes Lemma 3.2 of [22] to a more general class of  $(\mathbb{X}, \phi_s)$  than is considered in [22].

Given a connection function  $\phi \in \Phi$ , let us define

$$\kappa(\phi) := \sup_{x \in \mathbb{X}} \int_{\mathbb{X}} \phi(x, y) \mu(dy). \quad (2.6)$$

If  $\phi$  is  $\varepsilon$ -homogeneous for fixed  $\varepsilon > 0$ , then for any  $x \in \mathbb{X}$  the expected degree of a vertex of  $G(\mathcal{P}_s, \phi)$  located at  $x$  is of the order of  $s\kappa(\phi)$ . Our proof of Theorems 2.1 and 2.3 actually shows that under conditions of  $\varepsilon$ -homogeneity and  $\bar{\phi}_s \leq 1 - \varepsilon$ , if  $s\kappa(\phi_s) \rightarrow \infty$  (so the ‘typical degrees’ become large) then for any fixed  $j \geq 0$  and  $k \geq 2$ , both  $D_j$  and  $N_k$  are approximately Poisson distributed (so if the mean of this Poisson distribution goes to infinity then they are approximately normally distributed). Moreover, under the conditions of Theorem 2.1 or Theorem 2.3 (a), we shall show that  $s\kappa(\phi_s) = \Theta(\log s)$  (see (5.2) and (6.12)).

In Theorem 2.1 we do not address asymptotic normality  $D_j(G(\mathcal{P}_s, \phi_s))$  when its expected value grows to infinity more slowly than  $s$ , except for the case  $j = 0$  which is covered by the case  $k = 1$  of Theorem 2.3 (b). In attempting to adapt the proof of the latter to the case  $j > 0$ , the difficulty is that in general the expression (5.8) does not vanish in the case where  $s\kappa(\phi_s)$  tends to zero. We conjecture that asymptotic

normality of  $D_j(G(\mathcal{P}_s, \phi_s))$  can be proved by other means but this is beyond the scope of this paper.

Our next result is concerned with Poisson or normal approximation for (a generalization of) the number of edges of  $G(\mathcal{P}_s, \phi_s)$ . This is of interest in itself, and will also be of use in the proof of Theorem 2.3. Given  $k \in \mathbb{N}$  with  $k \geq 2$ , for any graph  $G$  let  $H_k(G)$  denote the number of connected induced subgraphs of  $G$  of order  $k$  (so in particular,  $H_2(G)$  is the number of edges.) Again, asymptotics are as  $s \rightarrow \infty$ .

**Theorem 2.4.** *Let  $\varepsilon > 0$ ,  $\phi_s \in \Phi_\varepsilon$  for all  $s > 0$ , and  $k \in \mathbb{N}$  with  $k \geq 2$ .*

(a) *If  $\mathbb{E} H_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha \in (0, \infty)$ , then  $H_k(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha$ .*

(b) *If  $\mathbb{E} H_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \infty$  but  $\mathbb{E} H_k(G(\mathcal{P}_s, \phi_s)) = o(s)$ , then setting  $\alpha_s := \mathbb{E} H_k(G(\mathcal{P}_s, \phi_s))$ , we have  $(H_k(G(\mathcal{P}_s, \phi_s)) - \alpha_s) / \sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ .*

The proof of Theorems 2.1, 2.3 and 2.4 provides information about the rates of convergence in these results. Theorem 2.2 will be proved by the method of moments, which does not provide any information about rates. The reason part (b) of that result requires a stronger condition on  $\bar{\phi}_s$  than part (a) does, is because when bounding these moments from below, for part (a) we have a sequence of the form  $\exp(-nx_n)$  (with  $x_n$  bounded by a constant times  $\bar{\phi}_n$ ), while for (b) we have a sequence of the form  $(1 - x_n)^n$  which is asymptotic to  $\exp(-nx_n)$  only when  $x_n = o(n^{-1/2})$ .

We now discuss some of the literature related to these results. Much of this concerns the RGG. Suppose  $(\mathbb{X}, \mathcal{F}, \mu)$  is the  $d$ -dimensional Euclidean space with the Borel  $\sigma$ -algebra and with  $\mu$  having a density  $f$  with respect to Lebesgue measure, while the connection function is given by  $\phi_s(x, y) = \mathbf{1}_{\{|x-y| \leq r_s\}}$ , with  $r_s$  chosen in such a way that  $\mathbb{E} D_0(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha \in (0, \infty)$ . It has been shown that  $D_0 \xrightarrow{\mathcal{D}} Z_\alpha$  when  $\mu$  is uniform on the unit cube [21], and for certain special types of density with unbounded support [9, 12, 20]. However,  $D_0$  is not always asymptotically Poisson; see the last paragraph of [12, Section 2], where it is suggested that the Poisson limit is ‘the exception rather than the rule’ in dimension  $d = 1$ . We are a long way from having any complete characterisation of distributions for which the number of isolated vertices in the RGG is asymptotically Poisson.

Suppose  $(\mathbb{X}, \mathcal{F}, \mu)$  is the  $d$ -dimensional unit cube equipped with Lebesgue measure. If  $\phi_s$  is of the form  $\phi_s(x, y) = p_s \mathbf{1}_{\{|x-y| \leq r_s\}}$ , with  $p_s$  and  $r_s$  chosen so that  $\mathbb{E} D_0(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha$ , then  $D_0 \xrightarrow{\mathcal{D}} Z_\alpha$ . More generally, if  $\phi_s$  is of the form  $\phi_s(x, y) = \psi_s(|x - y|)$  with  $\psi_s$  a decreasing function satisfying an exponential decay condition, then the Poisson convergence of  $D_0$  is known to hold; see [22]. Our results enable us to relax the exponential decay condition and allow for other distributions, for example with density bounded away from zero and infinity on a convex compact region in  $\mathbb{R}^d$ .

Here is a simple example of a sequence of connection functions  $\phi_s$  where  $D_0(G(\mathcal{P}_s, \phi_s))$  has bounded mean, but is *not* asymptotically Poisson. Let  $(\mathbb{X}, \mathcal{F}, \mu)$  be the unit interval equipped with Lebesgue measure, and set  $\phi_s(x, y) = 1$  if  $\max(x, y) \leq s^{-1}$  or  $\min(x, y) > s^{-1}$ , otherwise setting  $\phi_s(x, y) = 0$ . Then in the large- $s$  limit the random variable  $D_0(G(\mathcal{P}_s, \phi_s))$  is Bernoulli distributed with parameter  $e^{-1}$ , not Poisson distributed. The condition of  $\varepsilon$ -homogeneity, which appears in many of our results, rules out this sort of example.

Devroye and Fraiman [6] consider  $D_0(G(\mathcal{X}_n, \phi_n))$  on a general space  $\mathbb{X}$  in the case where  $\phi_n = (a(n)\phi_0) \wedge 1$  for a fixed function  $\phi_0 : \mathbb{X} \times \mathbb{X} \rightarrow \mathbb{R}_+$  (such a function  $\phi_0$  is called a *kernel*), and some sequence  $a(n)$  (where  $\wedge$  denotes minimum). This is a common assumption for the IRG; see for example [4]. Their results may be interpreted as saying that (under certain conditions) the threshold value of  $a$  above which  $G(\mathcal{X}_n, a\phi_0 \wedge 1)$  is free of isolated vertices (which may be viewed as a random variable) satisfies a weak law of large numbers; it is asymptotic to a constant times  $(\log n)/n$ , in probability. They also derive a similar law of large numbers for the threshold for the graph to be connected.

When  $\phi_0$  is bounded, our Theorem 2.2 (b) gives a possibility of deriving convergence in distribution for this threshold (suitably transformed). For example, suppose  $\mathbb{X} = [0, 1)$  and  $\mu$  has a density  $f$  with respect to Lebesgue measure, with  $f$  bounded away from zero. Suppose also that  $\phi_0(x, y)$  is determined via the absolute value of  $x - y \pmod{1}$  (so we are in the one-dimensional torus). If  $f \equiv 1$ , or if  $f$  is smooth with  $g(x) := \int \phi_0(x, y)f(y)(dy)$  having a unique minimum, then it should be possible to derive a distributional limit law for this threshold, since in these cases it should be possible, for any  $\alpha \in (0, \infty)$ , to determine a sequence  $a_s$  such that  $\mathbb{E} D_0(G(\mathcal{P}_s, a_s\phi_0)) \rightarrow \alpha$ , and then apply Theorem 2.2 (b).

If  $\phi_0$  is unbounded but shift-invariant on the torus, for example if  $\phi_0(x, y) = ((x - y) \pmod{1})^{-\gamma}$  for some fixed positive  $\gamma$  and  $f \equiv 1$ , then for any  $\varepsilon \in (0, 1)$ , it may be possible to find a limiting distribution for a suitable transformation of the threshold value of  $a$  above which  $G(\mathcal{X}_n, a\phi_0 \wedge (1 - \varepsilon))$  is free of isolated vertices, now using Theorem 2.3 (a). Indeed, in this case the connection functions are all 1-homogeneous, and by using connection functions  $\phi_n = a_n\phi_0 \wedge (1 - \varepsilon)$  rather than the more standard  $\phi_n = a_n\phi_0 \wedge 1$ , we ensure that the extra condition  $\phi_n \leq 1 - \varepsilon$  is also satisfied.

It would be interesting to fully work out and extend these examples; to get a similar results for the connectivity threshold; and to improve the weak law of [6] to a strong law.

In the present work we consider only undirected graphs. Analogous directed graph models can be defined similarly; these have been considered in the random geometric graphs literature [13] and in the statistical literature [11]. In this case, the connection function  $\phi(x, y)$  represents the probability that there is a directed



edge from a vertex at  $x$  to a vertex at  $y$ . It is not required to be symmetric, and can be adapted so that all vertices have the same expected outdegree, as in [13]. It would be interesting to try to derive a similar result to Theorem 2.1 for the number of vertices with outdegree  $j$  in such a model.

### 3 A general result on Poisson approximation

Let  $(\mathbb{M}, \mathcal{M}, \mathbf{m})$  be a probability space (known as a *mark space*). Assume that the probability measure  $\mathbf{m}$  on  $\mathbb{M}$  is *diffuse*, by which we mean that there is a product measurable set  $A \subset \mathbb{M} \times \mathbb{M}$  with  $(\mathbf{m} \otimes \mathbf{m})(A) = 0$ , such that the diagonal  $\{(t, t) : t \in \mathbb{M}\}$  is contained in  $A$ . For example, if  $\{t\} \in \mathcal{M}$  and  $\mathbf{m}(\{t\}) = 0$  for all  $t \in \mathbb{M}$ , then  $\mathbf{m}$  is diffuse.

Suppose on a suitable probability space that we have a sequence  $((X_i, T_i), i = 1, 2, 3, \dots)$  of independent identically distributed random elements of  $\mathbb{X} \times \mathbb{M}$  with common distribution  $\mu \otimes \mathbf{m}$ , and an independent unit rate Poisson counting process  $(Z_s, s > 0)$ , so that the random variable  $Z_s$  has Poisson( $s$ ) distribution for each  $s$ , and also a further independent sequence  $(\tau, \tau_1, \tau_2, \tau_3, \dots)$  of independent random elements of  $\mathbb{M}$  with common distribution  $\mathbf{m}$ . By our assumption that  $\mathbf{m}$  is diffuse, the values of  $\tau, T_1, \tau_1, T_2, \tau_2, T_3, \tau_3, \dots$  are almost surely distinct.

A *finite point process* in  $\mathbb{X}$  is defined as a random element of the space  $\mathbf{S}(\mathbb{X})$  of all finite subsets of  $\mathbb{X}$ , where  $\mathbf{S}(\mathbb{X})$  is equipped with the smallest  $\sigma$ -algebra  $\mathcal{S}(\mathbb{X})$  containing the sets  $\{\xi \in \mathbf{S}(\mathbb{X}) : \xi(B) = k\}$  for all  $B \in \mathcal{F}$  and all  $k \in \mathbb{N}_0$ , where  $\xi(B) := |\xi \cap B|$  and  $|\xi|$  denotes the number of elements of  $\xi$ .

A finite point process in  $\mathbb{X} \times \mathbb{M}$  is defined similarly as a random element of the space  $\mathbf{S} := \mathbf{S}(\mathbb{X} \times \mathbb{M})$ , where  $\mathbb{X} \times \mathbb{M}$  is equipped with the product  $\sigma$ -algebra  $\mathcal{F} \otimes \mathcal{M}$ . For  $k \in \mathbb{N}$ , let  $\mathbf{S}_k := \{\xi \in \mathbf{S} : |\xi| = k\}$ .

Given  $s > 0$ ,  $n \in \mathbb{N}$ , define the following point processes in  $\mathbb{X} \times \mathbb{M}$ :

$$\eta_s := \cup_{i=1}^{Z_s} \{(X_i, T_i)\}; \quad \xi_n := \cup_{i=1}^n \{(X_i, T_i)\}. \quad (3.1)$$

Then  $\eta_s$  is a Poisson point process in  $\mathbb{X} \times \mathbb{M}$  with mean measure  $s\mu \times \mathbf{m}$ . Similarly,  $\xi_n$  is a binomial point process in  $\mathbb{X} \times \mathbb{M}$ .

Let  $d_{TV}$  and  $d_W$  denote total variation distance and Wasserstein distance, respectively, between probability measures on the nonnegative integers. That is, for  $\mathbb{N}_0$ -valued random variables  $X, Y$  with distribution  $\mathcal{L}(X), \mathcal{L}(Y)$  respectively, we set

$$d_{TV}(\mathcal{L}(X), \mathcal{L}(Y)) = \sup_{A \subset \mathbb{N}_0} (\mathbb{P}[X \in A] - \mathbb{P}[Y \in A]);$$

$$d_W(\mathcal{L}(X), \mathcal{L}(Y)) = \sup\{|\mathbb{E} h(X) - \mathbb{E} h(Y)| : \|\Delta h\|_\infty \leq 1\},$$

where for  $h : \mathbb{N}_0 \rightarrow \mathbb{R}$  we set  $\Delta h(i) := h(i+1) - h(i)$  for  $i \in \mathbb{N}_0$ , and  $\|h\|_\infty := \sup_{i \in \mathbb{N}_0} |h(i)|$

The following theorem is related to a well-known result on the Poisson approximation of a sum of Bernoulli random variables by Stein's method via coupling (Theorem II.24.3 of [16], or Theorem 1.B of [1]). Here the terms in the sum are themselves indexed by  $k$ -subsets of the set of points of a (marked) Poisson point process.

Let  $k \in \mathbb{N}$  and let  $f : \mathbf{S}_k \times \mathbf{S} \rightarrow \{0, 1\}$  be a measurable function. For  $\xi \in \mathbf{S}$ , set

$$F(\xi) := \sum_{\{\psi \subset \xi : |\psi| = k\}} f(\psi, \xi \setminus \psi). \quad (3.2)$$

We can think of  $f$  as a mechanism for selecting some of the  $k$ -subsets of  $\xi$ , and  $F(\xi)$  as the total number of  $k$ -subsets selected.

**Theorem 3.1.** *Let  $s > 0$ . Let  $W := F(\eta_s)$  with  $\eta_s$  and  $F$  as described above. For  $x_1, \dots, x_k \in \mathbb{X}$  set  $p(x_1, \dots, x_k) := \mathbb{E} f(\{(x_1, \tau_1), \dots, (x_k, \tau_k)\}, \eta_s)$ , and set  $\lambda := s\mu$ .*

*Suppose that  $w : \mathbb{X}^k \rightarrow [0, \infty)$  is a measurable function, and that for  $\lambda^k$ -almost every  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$  with  $p(x_1, \dots, x_k) > 0$  we can find coupled random variables  $U_{\mathbf{x}}, V_{\mathbf{x}}$  such that*

- $\mathcal{L}(U_{\mathbf{x}}) = \mathcal{L}(W)$ ;
- $\mathcal{L}(1 + V_{\mathbf{x}}) = \mathcal{L}(F(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s) | f(\cup_{i=1}^k \{(x_i, \tau_i)\}, \eta_s) = 1)$ .
- $\mathbb{E} [|U_{\mathbf{x}} - V_{\mathbf{x}}|] \leq w(\mathbf{x})$ .

Set  $\alpha = \mathbb{E} W = (1/k!) \int p(\mathbf{x}) \lambda^k(d\mathbf{x})$ . Then

$$d_{TV}(\mathcal{L}(W), \mathcal{L}(Z_\alpha)) \leq \frac{(1 \wedge \alpha^{-1})}{k!} \int w(\mathbf{x}) p(\mathbf{x}) \lambda^k(d\mathbf{x}), \quad (3.3)$$

and

$$d_W(\mathcal{L}(W), \mathcal{L}(Z_\alpha)) \leq \frac{3(1 \wedge \alpha^{-1/2})}{k!} \int w(\mathbf{x}) p(\mathbf{x}) \lambda^k(d\mathbf{x}). \quad (3.4)$$

The proof uses the (multivariate) *Mecke formula* (see e.g. [15], or [19, Theorem 1.6]), which says that if  $g : \mathbf{S}_k \times \mathbf{S} \rightarrow \mathbb{R}$  is a bounded measurable function, then

$$\mathbb{E} \sum_{\{\psi \subset \eta_s : |\psi| = k\}} g(\psi, \eta_s \setminus \psi) = \frac{1}{k!} \int \mathbb{E} [g(\{(x_1, \tau_1), \dots, (x_k, \tau_k)\}, \eta_s)] \lambda^k(d\mathbf{x}). \quad (3.5)$$

This fact gives us the assertion in the statement of the theorem that  $\mathbb{E} W = (1/k!) \int p(\mathbf{x}) \lambda^k(d\mathbf{x})$ .

Theorem 3.1 still holds in the case where the measure  $\lambda$  is taken to be  $\sigma$ -finite but infinite, and  $\eta_s$  is replaced by a Poisson point process  $\eta$  with mean measure  $\lambda \otimes \mathbf{m}$ . The proof is essentially unchanged.

*Proof of Theorem 3.1.* Let  $h : \mathbb{N}_0 \rightarrow \mathbb{R}$  be bounded. Then

$$\begin{aligned} \mathbb{E}[Wh(W)] &= \mathbb{E} \sum_{\{\psi \subset \eta_s : |\psi|=k\}} f(\psi, \eta_s \setminus \psi) h(F(\eta_s)) \\ &= \frac{1}{k!} \int \mathbb{E} [f(\{(x_1, \tau_1), \dots, (x_k, \tau_k)\}, \eta_s) h(F(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s))] \lambda^k(d\mathbf{x}) \\ &= \frac{1}{k!} \int \mathbb{E} [h(F(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s)) | f(\cup_{i=1}^k \{(x_i, \tau_i)\}, \eta_s) = 1] p(\mathbf{x}) \lambda^k(d\mathbf{x}). \end{aligned}$$

Also,  $\mathbb{E}[\alpha h(W+1)] = (1/k!) \int \mathbb{E} h(W+1) p(\mathbf{x}) \lambda^k(d\mathbf{x})$ , and therefore

$$\begin{aligned} |\mathbb{E}[\alpha h(W+1) - Wh(W)]| &\leq \frac{1}{k!} \int p(\mathbf{x}) \lambda^k(d\mathbf{x}) \\ &\quad \times |\mathbb{E} h(W+1) - \mathbb{E} [h(F(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s)) | f(\cup_{i=1}^k \{(x_i, \tau_i)\}, \eta_s) = 1]|. \end{aligned}$$

Since  $|h(i) - h(j)| \leq \|\Delta h\|_\infty \cdot |i - j|$  for  $i, j \in \mathbb{N}_0$ , we obtain for each  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$  that

$$\begin{aligned} |\mathbb{E} h(W+1) - \mathbb{E} [h(F(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s)) | f(\cup_{i=1}^k \{(x_i, \tau_i)\}, \eta_s) = 1]| \\ \leq |\mathbb{E} h(U_{\mathbf{x}}+1) - \mathbb{E} h(V_{\mathbf{x}}+1)| \\ \leq \|\Delta h\|_\infty \mathbb{E} |U_{\mathbf{x}} - V_{\mathbf{x}}| \leq \|\Delta h\|_\infty w(\mathbf{x}), \end{aligned}$$

and therefore

$$|\mathbb{E}[\alpha h(W+1) - Wh(W)]| \leq \frac{\|\Delta h\|_\infty}{k!} \int w(\mathbf{x}) p(\mathbf{x}) \lambda^k(d\mathbf{x}).$$

Given  $A \subset \mathbb{N}_0$ , set  $g = \mathbf{1}_A$  and choose  $h : \mathbb{N}_0 \rightarrow \mathbb{R}$  so that  $h(0) = 0$  and

$$\alpha h(i+1) - ih(i) = g(i) - \mathbb{E}[G(Z_\alpha)], i \in \mathbb{N}_0. \quad (3.6)$$

Then (see Lemma 1.1.1 of [1])  $h$  is bounded and  $\|\Delta h\|_\infty \leq 1 \wedge \alpha^{-1}$ , and hence

$$|\mathbb{P}[W \in A] - \mathbb{P}[Z_\alpha \in A]| \leq \frac{(1 \wedge \alpha^{-1})}{k!} \int w(\mathbf{x}) p(\mathbf{x}) \lambda^k(d\mathbf{x}).$$

The result (3.3) follows.

One obtains (3.4) similarly by choosing, for any given  $g$  with  $\|\Delta g\|_\infty \leq 1$ , a solution  $h$  to (3.6) with  $h(0) = 0$ , and using Lemma 1.1.5 of [1].  $\square$

We now give an overview of how we shall use Theorem 3.1 to prove Theorems 2.1 and 2.3. As explained in the next section, we may view  $G(\mathcal{P}_s, \phi_s)$  as being determined by a marked Poisson point process in  $\mathbb{X}$ , i.e. a Poisson point process in a product space  $\mathbf{X} \times \mathbb{M}^*$  (where  $\mathbb{M}^*$  is a certain mark space).

Then the functional  $D_k(G(\mathcal{P}_s, \phi_s))$  may be viewed as a sum of the form (3.2) for a suitable  $f$  which selects those points with degree  $k$ . For each  $x \in \mathbb{X}$  we need to find coupled variables  $U_x$  and  $V_x$  such that  $U_x$  has the distribution of  $D_k(G(\mathcal{P}_s, \phi_s))$  and  $1 + V_x$  has the conditional distribution of  $D_k(G(\mathcal{P}_s \cup \{x\}, \phi_s))$  given that  $x$  has degree  $k$ , and such that  $|U_x - V_x|$  is small (in probability). To do this we note that by the thinning theorem (see for example [15]), the point process of points of  $\mathcal{P}_s$  connected to  $x$ , and the point process of points of  $\mathcal{P}_s$  not connected to  $x$ , are independent. To generate  $V_x$  we need to condition the first point process to have exactly  $k$  points, which we can do by adding or removing points from it, while leaving the second point process unchanged. Since we modify only the first point process (i.e., the points connected to  $x$ ), we may hope that the score  $V_x$  obtained from the modified (overall) point process is similar to the score  $U_x$  obtained from the original (overall) point process. This is how we shall prove Theorem 2.1.

For Theorem 2.3, we view  $N_k(G(\mathcal{P}_s, \phi_s))$  as a sum of the form (3.2) for a suitable  $f$  which selects those  $k$ -tuples of points forming a component. For each  $\mathbf{x} \subset \mathbb{X}$  with  $k$  elements, we need to find coupled variables  $U_{\mathbf{x}}$  and  $V_{\mathbf{x}}$  such that  $U_{\mathbf{x}}$  has the distribution of  $N_k(G(\mathcal{P}_s, \phi_s))$  and  $1 + V_{\mathbf{x}}$  has the conditional distribution of  $N_k(G(\mathcal{P}_s \cup \mathbf{x}, \phi_s))$  given that the points of  $\mathbf{x}$  form a component. Again by the thinning theorem, the point process of points of  $\mathcal{P}_s$  connected to  $\mathbf{x}$  is independent of the rest of  $\mathcal{P}_s$ , so to get  $V_{\mathbf{x}}$  we condition on this point process having no elements, simply by removing those points. Again this is a small change, so the difference between  $U_{\mathbf{x}}$  and  $V_{\mathbf{x}}$  again is small (in probability).

## 4 Formal constructions of the IRG

Let  $s \in (0, \infty)$  and  $n \in \mathbb{N}$ . Let  $\phi \in \Phi$ . We now give a more formal definition of the graphs  $G(\mathcal{P}_s, \phi)$  and  $G(\mathcal{X}_n, \phi)$ ,

We make the following particular choice of mark space  $(\mathbb{M}^*, \mathcal{M}^*, \mathbf{m}^*)$ . Let  $\text{Leb}$  denote Lebesgue measure on  $[0, 1)$ . Let  $\mathbb{M}^* = [0, 1)^{\mathbb{N}_0}$  with  $\mathcal{M}^*$  the product Borel  $\sigma$ -algebra, and with  $\mathbf{m}^* := \otimes_{n=0}^{\infty} \text{Leb}$ , so that a random element of  $\mathbb{M}^*$  with distribution  $\mathbf{m}^*$  is a sequence of independent uniform  $(0, 1)$  random variables indexed by  $\mathbb{N}_0$ .

Now (and for the rest of this paper) taking  $(\mathbb{M}, \mathcal{M}, \mathbf{m}) = (\mathbb{M}^*, \mathcal{M}^*, \mathbf{m}^*)$ , let  $(X_i, T_i)_{i \geq 1}$ ,  $(Z_s)_{s > 0}$  and  $(\tau, \tau_1, \tau_2, \dots)$  be as in the preceding section. Given  $n \in \mathbb{N}$  and  $s > 0$ , let the point processes  $\eta_s$  and  $\xi_n$  be as given by (3.1). Thus  $\eta_s$  is a Poisson point process in  $\mathbb{X}$  with mean measure  $s\mu$  and with each point marked

with a sequence of independent uniform $[0, 1)$  variables indexed by the nonnegative integers. Similarly,  $\xi_n$  is a binomial point process in  $\mathbb{X} \times \mathbb{M}^*$ . We write  $\mathbf{S}^*$  for  $\mathbf{S}(\mathbb{X} \times \mathbb{M}^*)$ .

To ease notation, we shall also assume from now on that the probability measure  $\mu$  on  $\mathbb{X}$  is *diffuse*. This ensures that the values of  $X_1, X_2, \dots$  are almost surely distinct. However, this assumption is for notational convenience only; even without it, one can make sense of our results either by allowing the set  $\{X_1, \dots, X_n\}$  to have multiplicities, or by using the attached marks  $T_i$  (which are almost surely distinct) to distinguish between different points  $X_i$ .

Set  $\mathcal{X}_n := \{X_1, \dots, X_n\}$  and  $\mathcal{P}_s := \{X_1, \dots, X_{Z_s}\}$ , the canonical projections of  $\xi_n$  and  $\eta_s$  respectively onto  $\mathbb{X}$ . Define the graph  $G(\mathcal{X}_n, \phi)$  to have vertex set  $\mathcal{X}_n$  and to have an edge between vertices  $X_i$  and  $X_j$ , for  $i, j \in [n] := \{1, 2, \dots, n\}$  with  $i < j$ , if and only if  $T_{i,j} \leq \phi(X_i, X_j)$ , where  $T_i = (T_{i,0}, T_{i,1}, T_{i,2}, \dots)$ . Let  $G(\mathcal{P}_s, \phi)$  be the graph  $G(\mathcal{X}_{Z_s}, \phi)$ .

This is one way to formally define the random graphs with the properties described more informally in the Introduction. It has the advantage that  $G(\mathcal{X}_n, \phi)$  is a subgraph of  $G(\mathcal{P}_s, \phi)$  whenever  $n \leq Z_s$  and  $G(\mathcal{P}_s, \phi)$  is a subgraph of  $G(\mathcal{X}_n, \phi)$  whenever  $Z_s \leq n$ , which is useful for coupling arguments. However, it has the disadvantage that the edge-set of the graph  $G(\mathcal{P}_s, \phi)$  is not invariant under permutation of the order in which the marked points  $(X_1, T_1), \dots, (X_{Z_s}, T_{Z_s})$  are listed. Therefore we define a further graph which has the same distribution but also satisfies this permutation-invariance. This will be useful in applying Theorem 3.1 in the proof of Theorems 2.1, 2.3 and 2.4.

Given  $\xi \in \mathbf{S}^*$  and  $\phi \in \Phi$ , define the graph  $G_\phi(\xi)$  as follows. If there exist distinct  $(x, t_0, t_1, \dots) \in \xi$  and  $(y, u_0, u_1, \dots) \in \xi$  with  $t_0 = u_0$ , then set  $G_\phi(\xi)$  to be the empty graph. Otherwise, we can write  $\xi$  uniquely as

$$\xi = \cup_{i=1}^{|\xi|} \{(x_i, t_{i,0}, t_{i,1}, t_{i,2}, \dots)\}$$

with each  $x_i \in \mathbb{X}$  and with  $t_{1,0} < t_{2,0} < \dots < t_{|\xi|,0}$ . Let  $G_\phi(\xi)$  have vertex set  $\{x_1, \dots, x_{|\xi|}\}$ , and for each  $i < j \leq |\xi|$  let  $G_\phi(\xi)$  have an edge connecting  $x_i$  to  $x_j$  if and only if  $t_{i,j} \leq \phi(x_i, x_j)$ . In other words, we use the first coordinate of the marks to determine the order in which we enumerate the points of  $\xi$ ; having done so, for  $i < j$  we use the  $(j + 1)$ -st component of the mark attached to the  $i$ -th point to decide whether to connect it to the  $j$ -th point.

Let us say that two random graphs  $G$  and  $G'$  have the same distribution if any graph invariant of  $G$  has the same distribution as the same graph invariant evaluated on  $G'$ . The following is immediate from the independence of the components of the marks  $\tau_i$ .

**Proposition 4.1.** *For any finite  $\mathcal{X} = \{x_1, \dots, x_m\} \subset \mathbb{X}$  the distribution of the random graph  $G_\phi(\{(x_1, \tau_1), \dots, (x_m, \tau_m)\})$  is the same as that of  $G(\mathcal{X}, \phi)$ .*

In particular, the the distribution of  $G_\phi(\eta_s)$  is the same as that of  $G(\mathcal{P}_s, \phi)$ , although they are not the same graph because the set of edges is defined differently for the two graphs. Likewise  $G_\phi(\xi_n)$  has the same distribution as  $G(\mathcal{X}_n, \phi)$ .

In the following lemmas we check measurability of functions which will feature in the proof of Theorems 2.1 and 2.3 respectively. We use the following notation. For  $\ell \in \mathbb{N}$  and  $i \in [\ell]$ , let  $I_{\ell,i}$  denote the set of  $(x, t_0, t_1, t_2, \dots) \in \mathbb{X} \times [0, 1]^{\mathbb{N}_0}$  such that  $(i-1)/\ell \leq t_0 < i/\ell$ .

**Lemma 4.1.** *Let  $j \in \mathbb{N}_0$ . Then the function  $f : \mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^* \rightarrow \{0, 1\}$  given by*

$$f(x, \mathbf{t}, \xi) := \mathbf{1}\{x \text{ has degree } j \text{ in } G_\phi(\xi \cup \{(x, \mathbf{t})\})\}, \quad (x, \mathbf{t}, \xi) \in \mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^*$$

*is measurable.*

*Proof.* For  $\ell, n, m, i_1, \dots, i_n \in \mathbb{N}$  with  $m \leq n \leq \ell$ , and  $i_1 < i_2 < \dots < i_n \leq \ell$ , let  $A_{\ell, n, m, i_1, \dots, i_n}$  denote the set of  $(x, \mathbf{t}, \xi) \in \mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^*$  such that:

1.  $(\xi \cup \{(x, \mathbf{t})\})(I_{\ell, i_k}) = 1$ , for  $1 \leq k \leq n$ ;
2.  $(\xi \cup \{(x, \mathbf{t})\})(I_{\ell, i}) = 0$ , for  $i \in [\ell] \setminus \{i_1, \dots, i_n\}$ ;
3.  $(x, \mathbf{t}) \in I_{\ell, i_m}$ ;
4.  $(\sum_{k=1}^{m-1} \xi(\{(y, \mathbf{u}) \in I_{\ell, i_k} : u_m \leq \phi(x, y)\})) + \sum_{k=m+1}^n \xi(\{(y, \mathbf{u}) \in I_{\ell, i_k} : t_k \leq \phi(x, y)\}) = j$ , where  $\mathbf{u} = (u_0, u_1, \dots)$  and  $\mathbf{t} = (t_0, t_1, \dots)$ .

Then each  $A_{\ell, n, m, i_1, \dots, i_n}$  is measurable in  $\mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^*$ , and

$$f(x, \mathbf{t}, \xi) = \mathbf{1}\{(x, \mathbf{t}, \xi) \in \cup_{\ell, n, m, i_1, \dots, i_n \in \mathbb{N}: m \leq n \leq \ell, 1 \leq i_1 < i_2 < \dots < i_n \leq \ell} A_{\ell, n, m, i_1, \dots, i_n}\}$$

which is a measurable function. □

**Lemma 4.2.** *Let  $k \in \mathbb{N}$ . Suppose the function  $\tilde{f} : (\mathbb{X} \times \mathbb{M}^*)^k \times \mathbf{S}^* \rightarrow \{0, 1\}$  is given, for  $(x_1, \mathbf{t}_1, \dots, x_k, \mathbf{t}_k) \in (\mathbb{X} \times \mathbb{M}^*)^k$  and  $\xi \in \mathbf{S}^*$ , by*

$$\begin{aligned} & \tilde{f}(x_1, \mathbf{t}_1, \dots, x_k, \mathbf{t}_k, \xi) \\ & := \mathbf{1}\{\{x_1, \dots, x_k\} \text{ induces a component of } G_\phi(\cup_{i=1}^k \{(x_i, \mathbf{t}_i)\} \cup \xi)\}. \end{aligned}$$

*Then  $\tilde{f}$  is measurable.*

*Proof.* For  $\ell, n, m_1, \dots, m_k, i_1, \dots, i_n \in \mathbb{N}$  with  $\max_{1 \leq i \leq k} m_i \leq n \leq \ell$ , and  $m_1, \dots, m_k$  distinct, and  $1 \leq i_1 < i_2 < \dots < i_n \leq \ell$ , let  $A_{\ell, n, m_1, \dots, m_k, i_1, \dots, i_n}$  denote the set of  $(x_1, \mathbf{t}_1, \dots, x_k, \mathbf{t}_k, \xi) \in (\mathbb{X} \times \mathbb{M}^*)^k \times \mathbf{S}^*$  such that for some connected graph  $\Gamma$  on vertex set  $[k]$ :

1.  $(\cup_{h=1}^k \{(x_h, \mathbf{t}_h)\} \cup \xi)(I_{\ell, i_j}) = 1$ , for  $1 \leq j \leq n$ ;
2.  $(\cup_{h=1}^k \{(x_h, \mathbf{t}_h)\} \cup \xi)(I_{\ell, i}) = 0$ , for  $i \in [\ell] \setminus \{i_1, \dots, i_n\}$ ;
3.  $(x_h, \mathbf{t}_h) \in I_{\ell, i_{m_h}}$  for  $1 \leq h \leq k$ ;
4. for all  $h \in [k]$  and  $1 \leq j < m_h$  with  $j \notin \{m_1, \dots, m_k\}$  we have  $\xi(\{(y, \mathbf{u}) \in I_{i_j} : u_{m_h} > \phi(x_h, y)\}) = 1$ , where  $\mathbf{u} = (u_0, u_1, \dots)$ ;
5. for all  $h \in [k]$  and  $m_h < j \leq n$  with  $j \notin \{m_1, \dots, m_k\}$  we have  $\xi(\{(y, \mathbf{u}) \in I_{i_j} : t_{h,j} > \phi(x_h, y)\}) = 1$ , where  $\mathbf{t}_h = (t_{h,0}, t_{h,1}, \dots)$ , and
6. for all  $h, h' \in [k]$  with  $m_h < m_{h'}$  and  $\{h, h'\}$  an edge of  $\Gamma$ , we have  $t_{h, m_h} \leq \phi(x_h, x_{h'})$ .

Then each  $A_{\ell, n, k, m_1, \dots, m_k, i_1, \dots, i_n}$  is measurable in  $(\mathbb{X} \times \mathbb{M})^k \times \mathbf{S}^*$ , and setting  $[n]_{\neq}^k$  to be the set of  $(m_1, \dots, m_k) \in [n]^k$  such that  $m_1, \dots, m_k$  are distinct, we have that  $\tilde{f}$  is the indicator of the set

$$\cup_{\ell, n, i_1, \dots, i_n \in \mathbb{N}, (m_1, \dots, m_k) \in [n]_{\neq}^k : n \leq \ell, 1 \leq i_1 < i_2 < \dots < i_n \leq \ell} A_{\ell, n, m_1, \dots, m_k, i_1, \dots, i_n}$$

which is measurable. □

## 5 Proof of Theorem 2.1

For  $s > 0$ , let  $\eta_s, \mathcal{P}_s, Z_s, \tau, \tau_1, \tau_2, \dots$  be as in the preceding section. For  $\phi \in \Phi$  and  $i \in \mathbb{N}_0$ , we can obtain from the Mecke equation (3.5) that

$$\mathbb{E} D_i(G_\phi(\eta_s)) = s \int_{\mathbb{X}} \frac{(s \int \phi(x, y) \mu(dy))^i}{i!} \exp\left(-s \int \phi(x, y) \mu(dy)\right) \mu(dx). \quad (5.1)$$

Here is a brief explanation of (5.1). In (3.5), we take  $k = 1$  and use the space  $\mathbf{S}^*$  rather than  $\mathbf{S}$ . For  $(x, \mathbf{t}, \xi) \in \mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^*$ , our function  $g((x, \mathbf{t}), \xi)$  takes the value 1 if  $x$  has degree  $i$  in the graph  $G_\phi((x, \mathbf{t}) \cup \xi)$ , and otherwise takes the value zero. Given  $Z_s = n$ , the point process  $\mathcal{P}_s$  has  $n$  points, each of which is independently connected to  $x$  with probability  $\int \phi(x, y) \mu(y) dy$ . Hence by the thinning property of the Poisson distribution [15, Proposition 1.3], the number of points of  $\mathcal{P}_s$  connected to  $x$  in  $G_\phi((x, \tau) \cup \eta_s)$  is Poisson with mean  $s \int \phi(x, y) \mu(dy)$ .

In the sequel, other formulae for expectations of numbers of vertices, or  $k$ -tuples of vertices, having certain properties in terms of the graph  $G_\phi(\eta_s)$  (or equivalently, the graph  $G(\mathcal{P}_s, \phi)$ ), will also be justified by the Mecke formula. These arguments also justify (3.5) of [22].

Recall from (2.6) that for  $\phi \in \Phi$  we set  $\kappa(\phi) := \sup_{x \in \mathbb{X}} \int \phi(x, y) \mu(dy)$ .

**Lemma 5.1.** *Suppose the assumptions of Theorem 2.1 hold. Then (2.3) holds. Also*

$$s\kappa(\phi_s) = \Theta(\log s), \quad (5.2)$$

and

$$\mathbb{E} D_j(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha. \quad (5.3)$$

*Proof.* For all  $s > 0$ , since we assume  $\phi_s \in \Phi_\varepsilon$ , by (5.1) we have for  $i \in \mathbb{N}$  that

$$\frac{\varepsilon s\kappa(\phi_s)}{i} \leq \frac{\mathbb{E} D_i(G(\mathcal{P}_s, \phi_s))}{\mathbb{E} D_{i-1}(G(\mathcal{P}_s, \phi_s))} \leq \frac{s\kappa(\phi_s)}{i}. \quad (5.4)$$

We are assuming for some fixed  $j \in \mathbb{N}_0$  that as  $s \rightarrow \infty$  we have

$$\mathbb{E} D_{\leq j}(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha \in (0, \infty). \quad (5.5)$$

Hence  $\mathbb{E} D_0(G(\mathcal{P}_s, \phi_s))$  is bounded, but also  $\mathbb{E} D_0(G(\mathcal{P}_s, \phi_s)) \geq s e^{-s\kappa(\phi_s)}$  by (5.1), so  $s\kappa(\phi_s) \rightarrow \infty$  as  $s \rightarrow \infty$ .

By (5.5),  $\mathbb{E} D_j(G(\mathcal{P}_s, \phi_s))$  remains bounded, and since also  $s\kappa(\phi_s) \rightarrow \infty$ , if  $j \geq 1$  then using (5.4) we have  $\mathbb{E} D_{j-1}(G(\mathcal{P}_s, \phi_s)) \rightarrow 0$ , and repeating the argument we also have  $\mathbb{E} D_i(G(\mathcal{P}_s, \phi_s)) \rightarrow 0$ , for  $i = 0, 1, \dots, j-1$ . Hence by (5.5) we have (5.3).

Using (5.3) and (5.1), the assumed  $\varepsilon$ -homogeneity of the  $\phi_s$ , and the fact that  $s\kappa(\phi_s) \rightarrow \infty$ , it is straightforward to show that (5.2) holds. Then, using (5.4) repeatedly, and (5.3), we obtain (2.3).  $\square$

*Proof of Theorem 2.1.* Assume the assumptions of Theorem 2.1 apply. Then (2.3) follows from Lemma 5.1, and (2.3) gives us the first part of (2.2).

Next we show that  $D_j(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha$ . To carry out the strategy outlined at the end of Section 3, we shall apply Theorem 3.1 to the case  $\phi = \phi_s$  of the function  $f$  considered in Lemma 4.1. That is, for  $(x, \mathbf{t}, \xi) \in \mathbb{X} \times \mathbb{M}^* \times \mathbf{S}^*$  we set  $f_s(x, \mathbf{t}, \xi)$  to be the indicator of the statement that  $x$  has degree  $j$  in  $G_{\phi_s}(\xi \cup \{(x, \mathbf{t})\})$ . Then  $D_j(G_{\phi_s}(\eta_s)) = F_s(\eta_s)$ , where  $F_s$  is the function  $F$  obtained by using  $f \equiv f_s$  in (3.2) (with  $k = 1$ ).

Let  $s > 0$  and  $x \in \mathbb{X}$ . If  $j \geq 1$  suppose we also have an extra sequence  $(Y, Y_1, \dots, Y_j)$  of independent identically distributed random elements of  $\mathbb{X}$  with  $\mathbb{P}[Y \in dy] = \phi_s(x, y)\mu(dy) / \int \phi_s(x, z)\mu(dz)$ , independent of  $((X_i, T_i))_{i \geq 1}$  and  $(Z_s)_{s > 0}$  and  $(\tau, \tau_1, \tau_2, \dots)$ . Let  $\mathcal{G}_s$  be the graph  $G_{\phi_s}(\eta_s \cup \{(x, \tau), (Y_1, \tau_1), \dots, (Y_j, \tau_j)\})$ , with added edges from  $x$  to each of  $Y_1, \dots, Y_j$  (if not already included). Let  $\mathcal{P}_{s,x}$  be the set of points of  $\mathcal{P}_s$  that are connected to  $x$  in this graph, and set  $\mathcal{P}_s^x := \mathcal{P}_s \setminus \mathcal{P}_{s,x}$ .

Let  $U_x$  denote the number of vertices of degree  $j$  in the subgraph of  $\mathcal{G}_s$  induced by vertex set  $\mathcal{P}_s^x$ . By Proposition 4.1, this graph has the distribution of  $G(\mathcal{P}_s, \phi_s)$ , so  $U_x$  has the distribution of  $F_s(\eta_s)$ .



Now consider the subgraph of  $\mathcal{G}_s$  induced by  $\{x\} \cup \mathcal{P}_s$ . This has the distribution of  $G(\mathcal{P}_s \cup \{x\}, \phi_s)$ , and  $\mathcal{P}_{s,x}$  is the set of vertices in this graph lying adjacent to  $x$ . Conditioning on  $x$  having degree  $j$  amounts to conditioning on  $|\mathcal{P}_{s,x}| = j$ . We define a coupled point process  $\mathcal{P}^*$  (a subset of  $\mathcal{P}_s \cup \{Y_1, \dots, Y_j\}$ ) with the distribution of  $\mathcal{P}_s$  conditioned on  $x$  having degree  $j$ , as follows.

If  $|\mathcal{P}_{s,x}| > j$  then we select  $|\mathcal{P}_{s,x}| - j$  elements of  $\mathcal{P}_{s,x}$  uniformly at random and discard them from  $\mathcal{P}_s$  to get a point process  $\mathcal{P}^*$ . If  $|\mathcal{P}_{s,x}| < j$  we set  $\mathcal{P}^* := \mathcal{P}_s \cup \mathcal{Y}_x$ , where we set

$$\mathcal{Y}_x := \{Y_1, \dots, Y_{j-|\mathcal{P}_{s,x}|}\}.$$

Let  $\mathcal{G}_s^*$  denote the subgraph of  $\mathcal{G}_s$  induced by  $\mathcal{P}^* \cup \{x\}$ . Then  $x$  has degree  $j$  in  $\mathcal{G}_s^*$ . Let  $V_x$  be the number of vertices in  $\mathcal{G}_s^*$  having degree  $j$ , other than  $x$ . This has the conditional distribution of  $F_s(\{(x, \tau)\} \cup \eta_s) - 1$  given that  $f_s(\{(x, \tau)\}, \eta_s) = 1$ . This is because  $\mathcal{P}_{s,x}$  and  $\mathcal{P}_s^x$  are independent Poisson processes, and conditioning on  $f_s(\{(x, \tau)\}, \eta_s) = 1$  amounts to conditioning on the first of these two Poisson processes having  $j$  points.

If  $|\mathcal{P}_{s,x}| > j$  then  $|U_x - V_x| \leq U'_x + V'_x$ , where we set  $U'_x$  to be the number of  $y \in \mathcal{P}_{s,x}$  such that  $y$  has  $j$  neighbours in  $\mathcal{P}_s$ , and  $V'_x$  to be the number of pairs  $(y, z)$  with  $y \in \mathcal{P}_{s,x}$ ,  $z \in \mathcal{P}_s$ , such that  $z \neq y$ ,  $z$  is connected to  $y$  and  $z$  has at most  $j$  neighbours in  $\mathcal{P}_s^x$ . By the Mecke formula, and the assumption that  $\phi_s \in \Phi_\varepsilon$  for all  $s$ , writing  $\kappa_s := \kappa(\phi_s)$  and recalling that  $s\kappa_s \rightarrow \infty$  by (5.2), we have

$$\begin{aligned} \mathbb{E} U'_x &= \int s\phi_s(x, y) \left( \frac{(s \int \phi_s(y, z)\mu(dz))^j}{j!} \right) \exp\left(-\int s\phi_s(y, z)\mu(dz)\right) \mu(dy) \\ &= O((s\kappa_s)^{j+1}) \exp(-\Theta(s\kappa_s)) = o(1), \end{aligned} \quad (5.6)$$

uniformly over  $x \in \mathbb{X}$ . Also, using that  $\bar{\phi}_s \leq 1 - \varepsilon$  for all  $s$ , we have that

$$\begin{aligned} \mathbb{E} V'_x &= \sum_{i=0}^j \int s\phi_s(x, y) \int s\phi_s(y, z) \left( \frac{(s \int \phi_s(z, w)(1 - \phi_s(x, w))\mu(dw))^i}{i!} \right) \\ &\quad \times \exp\left(-\int s\phi_s(z, w)(1 - \phi_s(x, w))\mu(dw)\right) \mu(dz)\mu(dy) \\ &= O((s\kappa_s)^{j+2}) \times \exp(-\Theta(s\kappa_s)) \rightarrow 0, \end{aligned} \quad (5.7)$$

uniformly over  $x \in \mathbb{X}$ .

If  $|\mathcal{P}_{s,x}| < j$  then  $|U_x - V_x| \leq \tilde{U}_x + \tilde{V}_x$ , where we set  $\tilde{U}_x$  to be the number of  $y \in \mathcal{Y}_x$  having at most  $j$  neighbours in  $\mathcal{P}_s$ , and  $\tilde{V}_x$  is the number of pairs  $(y, z)$  with  $y \in \mathcal{Y}_x$ ,  $z \in \mathcal{P}_s$  connected to  $y$ , and  $z$  having at most  $j$  neighbours in  $\mathcal{P}_s$ . Then

$$\mathbb{E} \tilde{U}_x \leq j \sum_{i=0}^j \int (s\kappa_s)^i \exp(-\Theta(s\kappa_s)) \left( \frac{\phi_s(x, y)}{\int \phi_s(x, w)\mu(dw)} \right) \mu(dy) \rightarrow 0, \quad (5.8)$$

uniformly over  $x \in \mathbb{X}$ . Also

$$\begin{aligned} \mathbb{E} \tilde{V}_x &\leq \sum_{i=0}^j \int s \left( s \int \phi_s(z, w) \mu(dw) \right)^i \exp \left( - \int s \phi_s(z, w) \mu(dw) \right) \\ &\quad \times j \int \left( \frac{\phi_s(x, y) \phi_s(z, y)}{\int \phi_s(x, w) \mu(dw)} \right) \mu(dy) \mu(dz) \\ &= O((s\kappa_s)^{j+1}) \times \exp(-\Theta(s\kappa_s)) \rightarrow 0, \end{aligned} \tag{5.9}$$

uniformly over  $x \in \mathbb{X}$ .

Combining the estimates (5.6), (5.7), (5.8) and (5.9), and using Theorem 3.1, gives us the first part of (2.1), namely  $D_j(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_\alpha$ .

Now suppose  $k > j$ . Set  $\beta_{k,s} := \mathbb{E}[D_k(G(\mathcal{P}_s, \phi_s))]$ . The argument just given, with  $j$  replaced by  $k$ , shows that  $d_{TV}(D_k(G(\mathcal{P}_s, \phi_s)), Z_{\beta_{k,s}}) \rightarrow 0$  as  $s \rightarrow \infty$ . Also  $\beta_{k,s} \rightarrow \infty$  by (2.3), so that  $(Z_{\beta_{k,s}} - \beta_{k,s})/\sqrt{\beta_{k,s}} \xrightarrow{\mathcal{D}} \mathcal{N}$ . Hence  $(D_k(G(\mathcal{P}_s, \phi_s)) - \beta_{k,s})/\sqrt{\beta_{k,s}} \xrightarrow{\mathcal{D}} \mathcal{N}$ , which is (2.4).

It remains to prove the second parts of (2.1) and (2.2). For  $n \in \mathbb{N}$ , let  $s(n) = n - n^{3/4}$  and  $t(n) := n + n^{3/4}$ . By Chebyshev's inequality, with high probability (i.e. with probability tending to 1) we have  $Z_{s(n)} \leq n \leq Z_{t(n)}$  so that  $\mathcal{P}_{s(n)} \subset \mathcal{X}_n \subset \mathcal{P}_{t(n)}$ . Moreover, when this happens,  $G(\mathcal{X}_n, \phi)$  is the subgraph of  $G(\mathcal{P}_{t(n)}, \phi)$  induced by  $\mathcal{X}_n$ , and  $G(\mathcal{P}_{s(n)}, \phi)$  is the subgraph of  $G(\mathcal{X}_n, \phi)$  induced by  $\mathcal{P}_{s(n)}$ .

By (5.1), for  $i \in \mathbb{N}_0$ ,

$$\left( \frac{s(n)}{n} \right)^{i+1} \leq \frac{ED_i(G(\mathcal{P}_{s(n)}, \phi_n))}{ED_i(G(\mathcal{P}_n, \phi_n))} \leq \sup_{x \in \mathbb{X}} \left( \exp \left( n^{3/4} \int \phi_n(x, y) \mu(dy) \right) \right)$$

and by (5.2), both the upper and the lower bound tend to 1. Therefore by (2.3) and (5.3) we have as  $n \rightarrow \infty$  that

$$\mathbb{E} D_i(G(\mathcal{P}_{s(n)}, \phi_n)) \rightarrow 0, \quad i = 0, 1, \dots, j-1; \tag{5.10}$$

$$\mathbb{E} D_j(G(\mathcal{P}_{s(n)}, \phi_n)) \rightarrow \alpha. \tag{5.11}$$

Let  $k \in \mathbb{N}_0$  with  $k \leq j$ . If  $\mathcal{P}_{s(n)} \subset \mathcal{X}_n$ , then

$$D_{\leq k}(G(\mathcal{X}_n, \phi_n)) - D_{\leq k}(G(\mathcal{P}_{s(n)}, \phi_n)) = S_n - R_n,$$

where  $S_n$  denotes the number of points of  $\mathcal{X}_n \setminus \mathcal{P}_{s(n)}$  with degree at most  $k$  in  $G(\mathcal{X}_n, \phi_n)$ , and  $R_n$  is the number of points of  $\mathcal{P}_{s(n)}$  with degree at most  $k$  in  $G(\mathcal{P}_{s(n)}, \phi_n)$  but with degree at least  $k+1$  in  $G(\mathcal{X}_n, \phi_n)$ .

Let  $S'_n$  denote the number of points of  $\mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$  that are connected to at most  $k$  points of  $\mathcal{P}_{s(n)}$ , and let  $R'_n$  be the number of points of  $\mathcal{P}_{s(n)}$  with degree at most  $k$

in  $G(\mathcal{P}_{s(n)}, \phi_n)$  but with degree at least  $k + 1$  in  $G(\mathcal{P}_{t(n)}, \phi_n)$ . If  $\mathcal{P}_{s(n)} \subset \mathcal{X}_n \subset \mathcal{P}_{t(n)}$ , then  $S_n \leq S'_n$  and  $R_n \leq R'_n$ . Hence by Markov's inequality,

$$\mathbb{P}[D_{\leq k}(G(\mathcal{X}_n, \phi_n)) \neq D_{\leq k}(G(\mathcal{P}_{s(n)}, \phi_n))] \leq \mathbb{P}[\{Z_{s(n)} \leq n \leq Z_{t(n)}\}^c] + \mathbb{E} S'_n + \mathbb{E} R'_n.$$

Now

$$\begin{aligned} \mathbb{E} S'_n &= 2n^{3/4} \sum_{i=0}^k \int_{\mathbb{X}} \frac{(s(n) \int \phi_n(x, y) \mu(dy))^i}{i!} \exp\left(-s(n) \int \phi_n(x, y) dy\right) \mu(dx) \\ &= \left(\frac{2n^{3/4}}{s(n)}\right) \mathbb{E} D_{\leq k}(G(\mathcal{P}_{s(n)}, \phi_n)), \end{aligned}$$

which tends to zero by (5.10) and (5.11). Also

$$\begin{aligned} \mathbb{E} R'_n &\leq s(n) \sum_{i=0}^k \int \frac{(s(n) \int \phi_n(x, y) \mu(dy))^i}{i!} \\ &\quad \times \exp\left(-s(n) \int \phi_n(x, y) \mu(dy)\right) \left\{2n^{3/4} \int \phi_n(x, y) \mu(dy)\right\} \mu(dx) \\ &\leq 2n^{3/4} a_n \mathbb{E} D_{\leq k}(\mathcal{P}_{s(n)}, \phi_n), \end{aligned}$$

which tends to zero by (5.2), (5.10) and (5.11).

Therefore with high probability we have  $D_{\leq k}(G(\mathcal{X}_n, \phi_n)) = D_{\leq k}(G(\mathcal{P}_{s(n)}, \phi_n))$ . This holds both for  $k = j$ , and for  $k = j - 1$ . Hence using the first part of (2.1) and the first part of (2.2) we obtain the second part of (2.1) and of (2.2).  $\square$

## 6 Proof of Theorems 2.2 and 2.3

Given  $\phi \in \Phi$ , given  $k, \ell \in \mathbb{N}$ , and given  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$  and  $\mathbf{y} = (y_1, \dots, y_\ell) \in \mathbb{X}^\ell$ , set

$$\phi(\mathbf{x}, \mathbf{y}) := 1 - \prod_{i=1}^k \prod_{j=1}^{\ell} (1 - \phi(x_i, y_j)). \quad (6.1)$$

We also write  $\phi(\{x_1, \dots, x_k\}, \{y_1, \dots, y_\ell\})$  for  $\phi(\mathbf{x}, \mathbf{y})$  (allowing multiplicities in the sets  $\{x_1, \dots, x_k\}$  and  $\{y_1, \dots, y_\ell\}$ ); it is the probability that there is at least one edge in the random graph  $G(\{x_1, \dots, x_k, y_1, \dots, y_\ell\}, \phi)$  connecting one of the vertices  $x_i$  to one of the vertices  $y_j$ . If  $k = 1$  we write  $\phi(x_1, \{y_1, \dots, y_\ell\})$  for  $\phi(\{x_1\}, \{y_1, \dots, y_\ell\})$ . Also, let  $h_\phi(\mathbf{x})$  or  $h_\phi(x_1, \dots, x_k)$  denote the probability that  $G(\{x_1, \dots, x_k\}; \phi)$  is connected; more precisely, let  $h_\phi(\mathbf{x}) := 1$  if  $k = 1$  and otherwise let

$$h_\phi(x_1, \dots, x_k) := \sum_{\Gamma} \prod_{\{\{i,j\}:\{i,j\} \in E(\Gamma)\}} \phi(x_i, x_j) \prod_{\{\{i,j\}:i,j \notin E(\Gamma)\}} (1 - \phi(x_i, x_j)), \quad (6.2)$$

where the sum is over all connected graphs  $\Gamma$  on vertex set  $\{1, \dots, k\}$ , and  $E(G)$  denotes the set of edges of a graph  $G$ . By the Mecke formula (3.5), and the equality in distribution of  $G(\mathcal{P}_s, \phi)$  and  $G_\phi(\eta_s)$  as discussed in Section 4,

$$\begin{aligned} \mathbb{E} N_k(G(\mathcal{P}_s, \phi)) &= \mathbb{E} N_k(G_\phi(\eta_s)) \\ &= \frac{s^k}{k!} \int_{\mathbb{X}^k} h_\phi(x_1, \dots, x_k) \exp\left(-s \int \phi(z; \{x_1, \dots, x_k\}) \mu(dz)\right) \mu^k(d(x_1, \dots, x_k)). \end{aligned} \quad (6.3)$$

Now fix  $k \in \mathbb{N}$  and  $\alpha \in (0, \infty)$ . Assume throughout this section that  $\phi_s \in \Phi$  for  $s > 0$ , and (unless explicitly stated otherwise) that

$$\lim_{s \rightarrow \infty} \mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) = \alpha. \quad (6.4)$$

*Proof of Theorem 2.2 (a).* Assume that  $\bar{\phi}_s = o(1/\log s)$ . We shall use the method of moments.

For  $n, \ell \in \mathbb{N}$  we write  $(n)_\ell$  for the descending factorial  $n(n-1)\cdots(n-\ell+1)$ . Then  $(N_k(G(\mathcal{P}_s, \phi_s)))_\ell$  is the number of ordered  $\ell$ -tuples of distinct  $k$ -components of  $G(\mathcal{P}_s, \phi_s)$ . This equals the sum over all ordered  $k\ell$ -tuples  $x_{1,1}, \dots, x_{1,k}, \dots, x_{\ell,1}, \dots, x_{\ell,k}$  of distinct points of  $\mathcal{P}_s$ , of the indicator of the event that for each  $i \leq \ell$  the subgraph of  $G(\mathcal{P}_s, \phi_s)$  induced by  $x_{i,1}, \dots, x_{i,k}$  is connected and these vertices are not connected to any other vertices of  $G(\mathcal{P}_s, \phi_s)$ , divided by  $(k!)^\ell$ . Hence by the Mecke formula,

$$\begin{aligned} \mathbb{E} [(N_k(G(\mathcal{P}_s, \phi_s)))_\ell] &= \frac{s^{k\ell}}{(k!)^\ell} \int_{\mathbb{X}^{k\ell}} \left( \prod_{i=1}^{\ell} h_{\phi_s}(x_{i,1}, \dots, x_{i,k}) \right) u_s(x_{1,1}, \dots, x_{\ell,k}) \\ &\quad \times \exp\left(-s \int \phi_s(z, \{x_{1,1}, \dots, x_{\ell,k}\}) \mu(dz)\right) \mu^{k\ell}(d(x_{1,1}, \dots, x_{\ell,k})), \end{aligned} \quad (6.5)$$

where we set  $u_s(\{x_{1,1}, \dots, x_{\ell,k}\})$  to be the probability that the graph  $G(\{x_{1,1}, \dots, x_{\ell,k}\}, \phi_s)$  has no edge between any  $x_{i,j}$  and  $x_{i',j'}$  such that  $i \neq i'$ , that is,

$$u_s(x_{1,1}, \dots, x_{\ell,k}) := \prod_{(i_1, j_1), (i_2, j_2) \in [\ell] \times [k]: i_1 < i_2} (1 - \phi_s(x_{i_1, j_1}, x_{i_2, j_2})). \quad (6.6)$$

By our condition on  $\bar{\phi}_s$  the value of  $u_s(x_{1,1}, \dots, x_{\ell,k})$  tends to 1, uniformly over  $(x_{1,1}, \dots, x_{\ell,k})$ . Also, by the union bound  $\phi_s(z, \{x_{1,1}, \dots, x_{\ell,k}\})$  is bounded by  $\sum_{i=1}^{\ell} \phi_s(z, \{x_{i,1}, \dots, x_{i,k}\})$ . Therefore by (6.5), writing just  $N_k$  for  $N_k(G(\mathcal{P}_s, \phi_s))$ , we have

$$\begin{aligned} \mathbb{E} [(N_k)_\ell] &\geq (1 + o(1)) \left( \frac{s^k}{k!} \int_{\mathbb{X}^k} h_{\phi_s}(x_1, \dots, x_k) \right. \\ &\quad \left. \times \exp\left(-s \int \phi_s(z, \{x_1, \dots, x_k\}) \mu(dz)\right) \mu^k(d(x_1, \dots, x_k)) \right)^\ell. \end{aligned}$$

By (6.4) and (6.3), this lower bound for  $\mathbb{E}[(N_k)_\ell]$  tends to  $\alpha^\ell$  as  $s \rightarrow \infty$ .

By the Bonferroni bound and the union bound we have for  $(z, x_{1,1}, \dots, x_{\ell,k}) \in \mathbb{X}^{1+\ell k}$  that

$$\begin{aligned} \phi_s(z, \{x_{1,1}, \dots, x_{\ell,k}\}) &\geq \sum_{i=1}^{\ell} \phi_s(z, \{x_{i,1}, \dots, x_{i,k}\}) \left(1 - \sum_{j=i+1}^{\ell} \phi_s(z, \{x_{j,1}, \dots, x_{j,k}\})\right) \\ &\geq \sum_{i=1}^{\ell} \phi_s(z, \{x_{i,1}, \dots, x_{i,k}\}) (1 - k\ell\bar{\phi}_s). \end{aligned} \quad (6.7)$$

Therefore by (6.5),

$$\begin{aligned} \mathbb{E}[(N_k)_\ell] &\leq \frac{s^{k\ell}}{(k!)^\ell} \int_{\mathbb{X}^{k\ell}} \left( \prod_{i=1}^{\ell} h_{\phi_s}(x_{i,1}, \dots, x_{i,k}) \right) \\ &\quad \times \exp \left( -s \sum_{i=1}^{\ell} \int \phi_s(z, \{x_{i,1}, \dots, x_{i,k}\}) (1 - k\ell\bar{\phi}_s) \mu(dz) \right) \mu^{k\ell}(d(x_{1,1}, \dots, x_{\ell,k})) \\ &= \left( \frac{s^k}{k!} \int_{\mathbb{X}^k} h_{\phi_s}(x_1, \dots, x_k) \right. \\ &\quad \left. \times \exp \left( -s \int \phi_s(z, \{x_1, \dots, x_k\}) (1 - k\ell\bar{\phi}_s) \mu(dz) \right) \mu^k(d(x_1, \dots, x_k)) \right)^\ell. \end{aligned}$$

Since  $\mu$  is a probability measure and  $x^{1-k\ell\bar{\phi}_s}$  is a concave function on  $x \geq 0$ , we obtain by Jensen's inequality and the fact that  $h_{\phi_s}(\cdot)^{1/(1-k\ell\bar{\phi}_s)} \leq h_{\phi_s}(\cdot)$  that

$$\begin{aligned} \mathbb{E}[(N_k)_\ell] &\leq \frac{s^{k^2\ell^2\bar{\phi}_s}}{k!^{\ell^2 k\bar{\phi}_s}} \left( \frac{s^k}{k!} \int h_{\phi_s}(x_1, \dots, x_k) \right. \\ &\quad \left. \times \exp \left( -s \int \phi_s(z, \{x_1, \dots, x_k\}) \mu(dz) \right) \mu^k(d(x_1, \dots, x_k)) \right)^{\ell(1-k\ell\bar{\phi}_s)} \end{aligned}$$

which tends to  $\alpha^\ell$  by (6.3), (6.4) and our assumption on  $\bar{\phi}_s$ . Thus  $\mathbb{E}[(N_k)_\ell] \rightarrow \alpha^\ell$ , so by the method of moments (see e.g. Theorem 1.22 of [3]) the result (i) follows.  $\square$

*Proof of Theorem 2.2 (b).* Assume now that  $\bar{\phi}_s = o(s^{-1/2})$ . Again we use the method of moments. Write  $n$  for  $\lfloor s \rfloor$ , and set  $N'_k = N_k(G(\mathcal{X}_n, \phi_n))$ . Then by (6.6)

and the union bound,

$$\begin{aligned}
\mathbb{E}[(N'_k)_\ell] &= \frac{(n)_{k\ell}}{k!^\ell} \int \cdots \int \left( \prod_{i=1}^{\ell} h_{\phi_n}(x_{i,1}, \dots, x_{i,k}) \right) u_n(\{x_{i,1}, \dots, x_{\ell,k}\}) \\
&\quad \times \left( 1 - \int \phi_n(z; \{x_{1,1}, \dots, x_{\ell,k}\}) \mu(dz) \right)^{n-k\ell} \mu(dx_{1,1}) \cdots \mu(dx_{\ell,k}) \\
&\geq (1 + o(1)) \left( \frac{n^k}{k!} \right)^\ell \int \cdots \int \left( \prod_{i=1}^{\ell} h_{\phi_n}(x_{i,1}, \dots, x_{i,k}) \right) \\
&\quad \times \left( 1 - \sum_{i=1}^{\ell} \int \phi_n(z, \{x_{i,1}, \dots, x_{i,k}\}) \mu(dz) \right)^n \mu(dx_{1,1}) \cdots \mu(dx_{\ell,k}).
\end{aligned}$$

Using the bound  $1 - x \geq \exp(-x - x^2)$  for small positive  $x$ , we have for large  $n$  that

$$\begin{aligned}
\mathbb{E}[(N'_k)_\ell] &\geq (1 + o(1)) \left( \frac{n^k}{k!} \right)^\ell \int \cdots \int \left( \prod_{i=1}^{\ell} h_{\phi_n}(x_{i,1}, \dots, x_{i,k}) \right) \\
&\quad \times \exp \left( -n \left( \sum_{i=1}^{\ell} \int \phi_n(z, \{x_{i,1}, \dots, x_{i,k}\}) \mu(dz) \right) \right. \\
&\quad \left. - n \left( \sum_{i=1}^{\ell} \sum_{j=1}^k \int \phi_n(z, x_{i,j}) \mu(dz) \right)^2 \right) \mu(dx_{1,1}) \cdots \mu(dx_{\ell,k}),
\end{aligned}$$

so that

$$\begin{aligned}
\mathbb{E}[(N'_k)_\ell] &\geq (1 + o(1)) \left( \frac{n^k}{k!} \int h_{\phi_n}(x_1, \dots, x_k) \right. \\
&\quad \left. \times \exp \left( -n \int \phi_n(z, \{x_1, \dots, x_k\}) \mu(dz) \right) \mu^k(d(x_1, \dots, x_k)) \right)^\ell \times \exp(-nk^2 \ell^2 \bar{\phi}_n^2)
\end{aligned}$$

which tends to  $\alpha^\ell$  by (6.3), (6.4), and the assumption that  $\bar{\phi}_n = o(n^{-1/2})$ .

Conversely, by the bound  $e^x \geq 1 + x$  for all  $x \in \mathbb{R}$ , we also have

$$\begin{aligned}
\mathbb{E}[(N'_k)_\ell] &\leq \left( \frac{n^k}{k!} \right)^\ell \int \cdots \int \left( \prod_{i=1}^{\ell} h_{\phi_n}(x_{i,1}, \dots, x_{i,k}) \right) \\
&\quad \times \exp \left( -(n - \ell k) \int \phi_n(z; \{x_{1,1}, \dots, x_{\ell,k}\}) \mu(dz) \right) \mu(dx_{1,1}) \cdots \mu(dx_{\ell,k})
\end{aligned}$$

and using the Bonferroni bound as in (6.7), we obtain that

$$\begin{aligned} \mathbb{E}[(N'_k)_\ell] &\leq e^{k^2 \ell^2 \bar{\phi}_n} \left(\frac{n^k}{k!}\right)^\ell \int \cdots \int \left(\prod_{i=1}^{\ell} h_{\phi_n}(x_{i,1}, \dots, x_{i,k})\right) \\ &\quad \times \exp\left(-n \sum_{i=1}^{\ell} \int \phi_n(z; \{x_{i,1}, \dots, x_{i,k}\})(1 - k\ell \bar{\phi}_n) \mu(dz)\right) \mu(dx_{1,1}) \cdots \mu(dx_{\ell,k}). \end{aligned}$$

By Jensen's inequality, since  $x^{1-k\ell\bar{\phi}_n}$  is a concave function on  $x \geq 0$  and  $h_{\phi_n}(\cdot) \leq 1$ , we have

$$\begin{aligned} \mathbb{E}[(N'_k)_\ell] &\leq e^{k^2 \ell^2 \bar{\phi}_n} n^{k^2 \ell^2 \bar{\phi}_n} \left(\frac{1}{k!}\right)^\ell \left(n^k \int \cdots \int h_{\phi_n}(x_1, \dots, x_k)\right. \\ &\quad \left. \times \exp\left(-n \int \phi_n(z; \{x_1, \dots, x_k\}) \mu(dz)\right) \mu(dx_1) \cdots \mu(dx_k)\right)^{\ell(1-k\ell\bar{\phi}_n)}. \end{aligned}$$

Hence, since our assumption  $\bar{\phi}_n = o(n^{-1/2})$  implies that  $\bar{\phi}_n = o(1/(\log n))$ , we have  $\limsup \mathbb{E}[(N'_k)_\ell] \leq \alpha^\ell$ . Therefore  $\mathbb{E}[(N'_k)_\ell] \rightarrow \alpha^\ell$ , and the method of moments gives us part (ii).  $\square$

**Lemma 6.1.** *Let  $k \in \mathbb{N}$ ,  $\varepsilon > 0$  and  $\phi_s \in \Phi_\varepsilon$  for all  $s > 0$ . Set  $\kappa_s := \kappa(\phi_s) = \sup_{x \in \mathbb{X}} \int \phi_s(x, y) \mu(dy)$ . Then*

$$\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) = \Theta(s^k \kappa_s^{k-1}) \exp(-\Theta(s \kappa_s)). \quad (6.8)$$

*Proof.* Our starting point is (6.3). We first bound the exponent in the exponential factor. Let  $z, x_1, \dots, x_k \in \mathbb{X}$  and  $s > 0$ . Then by the union bound,

$$\phi_s(z, x_1) \leq \phi_s(z, \{x_1, \dots, x_k\}) \leq \sum_{i=1}^k \phi_s(z, x_i),$$

so integrating over  $z$  and using the assumed  $\varepsilon$ -homogeneity, we have

$$\varepsilon \kappa_s \leq \int_{\mathbb{X}} \phi_s(z, \{x_1, \dots, x_k\}) \mu(dz) \leq k \kappa_s. \quad (6.9)$$

Using (6.3), this already proves the result for  $k = 1$ , so from now on assume  $k \geq 2$ .

If  $x_{i-1}$  is connected to  $x_i$  for  $2 \leq i \leq k$ , then  $G(\{x_1, \dots, x_k\}; \phi_s)$  is connected; hence  $h_{\phi_s}(x_1, \dots, x_k) \geq \prod_{i=2}^k \phi_s(x_{i-1}, x_i)$ , so using (6.3) and (6.9) we obtain that

$$\begin{aligned} \mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) &\geq \frac{s^k}{k!} \exp(-k s \kappa_s) \int_{\mathbb{X}} \cdots \int_{\mathbb{X}} \prod_{i=2}^k \phi_s(x_{i-1}, x_i) \mu(dx_k) \cdots \mu(dx_1) \\ &\geq \frac{\varepsilon^{k-1} \kappa_s^{k-1} s^k}{k!} \exp(-k s \kappa_s). \end{aligned} \quad (6.10)$$

For an upper bound, observe that if  $G(\{x_1, \dots, x_k\}; \phi_s)$  is connected, then there exists a permutation  $\sigma$  of  $\{1, \dots, k\}$  with  $\sigma(1) = 1$  such that for  $2 \leq i \leq k$  this graph has an edge from  $x_{\sigma(i)}$  to  $\{x_{\sigma(1)}, \dots, x_{\sigma(i-1)}\}$ . Therefore, setting

$$h_{\phi_s}^*(x_1, \dots, x_k) := \prod_{i=2}^k \phi_s(x_i, \{x_1, \dots, x_{i-1}\}) \leq \prod_{i=2}^k \sum_{j=1}^{i-1} \phi_s(x_i, x_j), \quad (6.11)$$

we have by the union bound that  $h_{\phi_s}(x_1, \dots, x_k) \leq \sum_{\sigma} h_{\phi_s}^*(x_{\sigma(1)}, \dots, x_{\sigma(k)})$  where the sum is over all such permutations. Hence using (6.3) and (6.9), setting  $y_j = x_{\sigma(j)}$  and taking the integrals in order  $y_k, \dots, y_1$  for each permutation  $\sigma$ , we have

$$\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \leq \frac{s^k (k-1)!}{k!} \exp(-\varepsilon s \kappa_s) \int_{\mathbb{X}} \cdots \int_{\mathbb{X}} h_{\phi_s}^*(y_1, \dots, y_k) \mu(dy_k) \cdots \mu(dy_1),$$

and then using the inequality in (6.11) we have

$$\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \leq s^k (k-1)! \exp(-\varepsilon s \kappa_s) \kappa_s^{k-1}.$$

Combined with (6.10) this gives us (6.8).  $\square$

*Proof of Theorem 2.3.* Assume there exists  $\varepsilon > 0$  such that  $\phi_s \in \Phi_{\varepsilon}$  and  $\bar{\phi}_s \leq 1 - \varepsilon$  for all  $s$ . Set  $\kappa_s := \kappa(\phi_s)$ . If  $s\kappa_s$  remains bounded away from zero and infinity, then by (6.8) we have that  $\mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \infty$ , contradicting (6.4). Hence for any sequence of values of  $s$  tending to infinity, there is a subsequence such that either  $s\kappa_s \rightarrow 0$  or  $s\kappa_s \rightarrow \infty$  as  $s \rightarrow \infty$  along the subsequence.

Consider first the case with  $s\kappa_s \rightarrow 0$ . In this case, by (6.8) and (6.4) we have  $k \geq 2$  and  $s^k \kappa_s^{k-1} = \Theta(1)$ . Recalling that  $H_k(G)$  denotes the number of connected induced subgraphs of a graph  $G$  of order  $k$ , we have

$$\mathbb{E} H_{k+1}(G(\mathcal{P}_s, \phi_s)) = O(s^{k+1} \kappa_s^k) = o(1).$$

Since  $0 \leq H_k(G(\mathcal{P}_s, \phi_s)) - N_k(G(\mathcal{P}_s, \phi_s)) \leq (k+1)H_{k+1}(G(\mathcal{P}_s, \phi_s))$ , we have

$$\mathbb{E} [H_k(G(\mathcal{P}_s, \phi_s)) - N_k(G(\mathcal{P}_s, \phi_s))] \rightarrow 0,$$

so by (6.4) we have that  $\mathbb{E} H_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \alpha$ . Hence by Theorem 2.4 (a) we have  $H_k(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_{\alpha}$ . By Markov's inequality  $\mathbb{P}[H_k(G(\mathcal{P}_s, \phi_s)) - N_k(G(\mathcal{P}_s, \phi_s)) \geq 1] \rightarrow 0$ , so we also have  $N_k(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_{\alpha}$ , which is the first part of (2.5).

Suppose now that  $s\kappa_s \rightarrow \infty$ . Then by (6.8) and (6.4), one may readily deduce that

$$s\kappa_s = \Theta(\log s). \quad (6.12)$$



We seek to apply Theorem 3.1. For  $s > 0$ , let  $\eta_s, \mathcal{P}_s, Z_s, \tau, \tau_1, \tau_2, \dots$  be as in Section 4. Let  $\tilde{f}_s$  be the function  $\tilde{f}$  considered in Lemma 4.2, using the connection function  $\phi \equiv \phi_s$ . That is, let  $\tilde{f}_s(x_1, \mathbf{t}_1, \dots, x_k, \mathbf{t}_k, \xi)$  be the indicator of the statement that  $\{x_1, \dots, x_k\}$  induces a component of  $G_{\phi_s}(\cup_{i=1}^k \{(x_i, \mathbf{t}_i)\} \cup \xi)$ . Then with  $\tilde{F}_s$  denoting the function  $F$  obtained by taking  $f \equiv \tilde{f}_s$  in the definition (3.2), we have that  $\tilde{F}_s(\eta_s) = N_k(G_{\phi_s}(\eta_s))$  which has the same distribution as  $N_k(G(\mathcal{P}_s, \phi_s))$ .

For  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$ , set  $\mathcal{X} = \{x_1, \dots, x_k\}$  (allowing multiplicities). Define the graph

$$\tilde{\mathcal{G}}_s := G_{\phi_s}(\eta_s \cup \{(x_1, \tau_1), \dots, (x_k, \tau_k)\}). \quad (6.13)$$

Let  $\mathcal{P}_{s, \mathbf{x}}$  be the set of points of  $\mathcal{P}_s$  connected to at least one point of  $\mathcal{X}$  in  $\tilde{\mathcal{G}}_s$  and let  $\mathcal{P}_s^{\mathbf{x}} = \mathcal{P}_s \setminus \mathcal{P}_{s, \mathbf{x}}$ .

The subgraph of  $\tilde{\mathcal{G}}_s$  induced by vertex set  $\mathcal{P}_s$  has the same distribution as  $G(\mathcal{P}_s, \phi_s)$ , and we shall refer to this subgraph as  $G'(\mathcal{P}_s, \phi_s)$ . Likewise, we refer to the subgraph of  $\tilde{\mathcal{G}}_s$  induced by vertex set  $\mathcal{P}_s^{\mathbf{x}}$  as  $G'(\mathcal{P}_s^{\mathbf{x}}, \phi_s)$ , and we refer to the subgraph of  $\tilde{\mathcal{G}}_s$  induced by vertex set  $\mathcal{X}$  as  $G'(\mathcal{X}, \phi_s)$ .

Let  $U_{\mathbf{x}} = N_k(G'(\mathcal{P}_s, \phi_s))$ . This has the same distribution as  $\tilde{F}_s(\eta_s)$ .

Let  $V_{\mathbf{x}} = N_k(G'(\mathcal{P}_s^{\mathbf{x}}, \phi_s))$ . We claim that this has the same distribution as conditional distribution of  $\tilde{F}_s(\cup_{i=1}^k \{(x_i, \tau_i)\} \cup \eta_s) - 1$  given that  $\tilde{f}_s(\cup_{i=1}^k \{(x_i, \tau_i)\}, \eta_s) = 1$ . This is because by the Marking Theorem for Poisson processes (see e.g. [14]), the point processes  $\mathcal{P}_{s, \mathbf{x}}$  and  $\mathcal{P}_s^{\mathbf{x}}$  are independent and the statement that  $\tilde{\mathcal{G}}_s$  has  $\mathcal{X}$  as the vertex set of a component is equivalent to the statement that (i)  $\mathcal{P}_{s, \mathbf{x}}$  has no points and (ii) the graph  $G'(\mathcal{X}, \phi_s)$  is connected, which is independent of the outcome of  $G'(\mathcal{P}_s^{\mathbf{x}}, \phi_s)$ .

Also,  $U_{\mathbf{x}} - V_{\mathbf{x}} = U'_{\mathbf{x}} - V'_{\mathbf{x}}$ , where  $U'_{\mathbf{x}}$  denotes the number of  $k$ -components of  $G'(\mathcal{P}_s, \phi_s)$  with at least one vertex in  $\mathcal{P}_{s, \mathbf{x}}$ , and  $V'_{\mathbf{x}}$  is the number of  $k$ -components of  $G'(\mathcal{P}_s^{\mathbf{x}}, \phi_s)$  with at least one neighbour in  $\mathcal{P}_{s, \mathbf{x}}$ .

By the Mecke formula

$$\mathbb{E} U'_{\mathbf{x}} = \frac{s^k}{k!} \int_{\mathbb{X}^k} \phi_s(\mathbf{x}, \mathbf{y}) h_{\phi_s}(\mathbf{y}) \exp\left(-s \int \phi_s(z, \mathbf{y}) \mu(dz)\right) \mu^k(d\mathbf{y}),$$

where  $\phi_s(\mathbf{x}, \mathbf{y})$  is given by (6.1). Since we can choose the elements  $y_1, \dots, y_k$  of  $\mathbf{y}$  in an order such that  $y_1$  is connected to  $\mathcal{X}$  and for each  $j \geq 2$ ,  $y_j$  is connected to  $\{y_1, \dots, y_{j-1}\}$ , by a similar argument to the proof of Lemma 6.1 we have

$$\mathbb{E} U'_{\mathbf{x}} = O(s^k \kappa_s^k) \times \exp(-\Theta(s \kappa_s)), \quad (6.14)$$

which tends to zero (uniformly over  $\mathbf{x}$ ).

Now  $V'_{\mathbf{x}}$  is bounded by the number of pairs  $(y, \mathbf{z})$  with  $y \in \mathcal{P}_{s, \mathbf{x}}$  and  $\mathbf{z} = (z_1, \dots, z_k)$  with  $\{z_1, \dots, z_k\}$  inducing a  $k$ -component of  $G'(\mathcal{P}_s^{\mathbf{x}}, \phi_s)$  and  $\mathbf{z}$  connected

to  $y$ . Hence by the Mecke equation,

$$\begin{aligned} \mathbb{E} V_{\mathbf{x}}' &\leq s^{k+1} \int_{\mathbb{X}} \int_{\mathbb{X}^k} \phi_s(y, \mathbf{x}) \phi_s(y, \mathbf{z}) (1 - \phi_s(\mathbf{x}, \mathbf{z})) h_{\phi_s}(\mathbf{z}) \\ &\quad \times \exp\left(-s \int \phi_s(w, \mathbf{z}) (1 - \phi_s(w, \mathbf{x})) \mu(dw)\right) \mu^k(d\mathbf{z}) \mu(dy) \\ &= O((s\kappa_s)^{k+1}) \exp(-\Theta(s\kappa_s)), \end{aligned} \quad (6.15)$$

which tends to zero, uniformly over  $\mathbf{x}$ ; here we have used the assumption that  $\bar{\phi}_s \leq 1 - \varepsilon$ , so that  $1 - \phi_s(w, \mathbf{x}) \geq \varepsilon^k$ , for all  $s, w, \mathbf{x}$ . Therefore  $\mathbb{E}[|U_{\mathbf{x}} - V_{\mathbf{x}}|] = \mathbb{E}[|U_{\mathbf{x}}' - V_{\mathbf{x}}'|] \rightarrow 0$ , uniformly over  $\mathbf{x} \in \mathbb{X}^k$ . Then we can use Theorem 3.1 to get the first part of (2.5).

Before completing the proof of part (a) of Theorem 2.3, we prove part (b), so now instead of (6.4) we assume  $\alpha_s := \mathbb{E} N_k(G(\mathcal{P}_s, \phi_s)) \rightarrow \infty$ , but  $\alpha_s = o(s)$ . Then by (6.8), for every sequence of values of  $s$  tending to infinity, there is a subsequence such that either  $s\kappa_s \rightarrow \infty$  or  $k \geq 2$  and  $s\kappa_s \rightarrow 0$  as  $s \rightarrow \infty$  along the subsequence.

In both cases, the estimates (6.14) and (6.15) hold so by Theorem 3.1 we have  $d_{TV}(\mathcal{L}(N_k(G(\mathcal{P}_s, \phi_s))), \mathcal{L}(Z_{\alpha_s})) \rightarrow 0$ . Since also  $(Z_{\alpha_s} - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ , it follows that  $(N_k(G(\mathcal{P}_s, \phi_s)) - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ . That is, for every sequence of values of  $s$  tending to infinity, there exists a subsequence such that  $(N_k(G(\mathcal{P}_s, \phi_s)) - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$  as  $s \rightarrow \infty$  along the subsequence. Hence  $(N_k(G(\mathcal{P}_s, \phi_s)) - \alpha_s)/\sqrt{\alpha_s}$  converges in distribution to  $\mathcal{N}$  as  $s \rightarrow \infty$ . This completes the proof of part (b).

Now we return to part (a), so we go back to assuming (6.4). As in the corresponding part of the proof of Theorem 2.1, for  $n \in \mathbb{N}$  set  $s(n) = n - n^{3/4}$  and  $t(n) = n + n^{3/4}$ . Then  $\mathcal{P}_{s(n)} \subset \mathcal{X}_n \subset \mathcal{P}_{t(n)}$  with high probability, and also the point process  $\mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$  is a Poisson point process with mean measure  $2n^{3/4}\mu(\cdot)$ , independent of  $\mathcal{P}_{s(n)}$ . By (6.12),  $n^{3/4} \int \phi_n(y, x) \mu(dy) = o(1)$ , uniformly over  $x \in \mathbb{X}$ , and therefore by (6.4) the sequence  $(\phi_n)_{n \in \mathbb{N}}$  satisfies

$$\frac{s(n)^k}{k!} \int_{\mathbb{X}^k} \exp\left(-s(n) \int_{\mathbb{X}} \phi_n(z, \mathbf{x}) \mu(dz)\right) h_{\phi_n}(\mathbf{x}) \mu^k(d\mathbf{x}) \rightarrow \alpha. \quad (6.16)$$

For  $n, \ell \in \mathbb{N}$  with  $1 \leq \ell \leq k$ , let  $A_{n,\ell}$  be the event that at least one collection of  $\ell$  of the added vertices of  $\mathcal{X}_n \setminus \mathcal{P}_{s(n)}$  lies in a  $k$ -component of  $G(\mathcal{X}_n, \phi_n)$ . Let  $B_n$  be the event that at least one of the added vertices of  $\mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$  is connected to one of the  $k$ -components of  $G(\mathcal{P}_{s(n)}, \phi_n)$ .

If  $A_{n,\ell}$  occurs and  $\mathcal{P}_{s(n)} \subset \mathcal{X}_n \subset \mathcal{P}_{t(n)}$ , then there is at least one pair  $(\mathcal{X}, \mathcal{Y})$ , such that  $\mathcal{X} \subset \mathcal{P}_{s(n)}$  has  $k - \ell$  elements, and  $\mathcal{Y} \subset \mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$  has  $\ell$  elements, and  $\mathcal{X} \cup \mathcal{Y}$  induces a connected subgraph of  $G(\mathcal{P}_{t(n)}, \phi_n)$ , and there is no connection between any vertex of  $\mathcal{X} \cup \mathcal{Y}$  and any vertex of  $\mathcal{P}_{s(n)} \setminus \mathcal{X}$  (however, we do allow other

connections between vertices of  $\mathcal{X} \cup \mathcal{Y}$  and other vertices of  $\mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$ . By the Mecke equation, the expected number of such pairs equals

$$\begin{aligned} & \frac{(2n^{3/4})^\ell}{\ell!} \int \frac{(s(n))^{k-\ell}}{(k-\ell)!} \int h_{\phi_n}(x_1, \dots, x_{k-\ell}, y_1, \dots, y_\ell) \\ & \quad \times \exp \left( -s(n) \int \phi_n(z, \{x_1, \dots, x_{k-\ell}, y_1, \dots, y_\ell\}) \mu(dz) \right) \\ & \quad \mu^{k-\ell}(d(x_1, \dots, x_{k-\ell})) \mu^\ell(d(y_1, \dots, y_\ell)), \end{aligned}$$

which tends to zero by (6.16), so  $\mathbb{P}[A_{n,\ell}] \rightarrow 0$ . Also, the expected number of  $k$ -components in  $G(\mathcal{P}_{s(n)}, \phi_n)$  which are connected to at least one vertex of  $\mathcal{P}_{t(n)} \setminus \mathcal{P}_{s(n)}$  is at most

$$\frac{(s(n))^k}{k!} \int \exp \left( -s(n) \int \phi_n(z, \mathbf{x}) \mu(dz) \right) h_{\phi_n}(\mathbf{x}) 2n^{3/4} k a_n \mu^k(d\mathbf{x}),$$

and by (6.16) and (6.12) this tends to zero. Hence  $\mathbb{P}[B_n] \rightarrow 0$ . By the first part of (2.5) we have for  $k \in \mathbb{N}$  that  $N_k(G(\mathcal{P}_{s(n)}, \phi_n)) \xrightarrow{\mathcal{D}} Z_\alpha$ . Also

$$\mathbb{P}[N_k(G(\mathcal{X}_n, \phi_n)) \neq N_k(G(\mathcal{P}_{s(n)}, \phi_n))] \leq \mathbb{P}[\cup_{\ell=1}^k A_{n,\ell}] + \mathbb{P}[B_n] + \mathbb{P}[\{Z_{s(n)} \leq n \leq Z_{t(n)}\}^c],$$

which tends to 0, and the second part of (2.5) follows.  $\square$

## 7 Number of edges

For any graph  $G$ , according to our earlier notation  $H_2(G)$  denotes the number of edges of  $G$ . Let  $\phi \in \Phi$ . Then by the Mecke formula (3.5),

$$\mathbb{E} H_2(G(\mathcal{P}_s, \phi)) = \mathbb{E} H_2(G_\phi(\eta_s)) = \frac{1}{2} \int_{\mathbb{X}} \int_{\mathbb{X}} \phi(x, y) s^2 \mu(dx) \mu(dy). \quad (7.1)$$

**Theorem 7.1.** *Suppose  $\phi \in \Phi$ . Set  $\alpha := \mathbb{E} H_2(G(\mathcal{P}_s, \phi))$ . Then*

$$d_{TV}(H_2(G(\mathcal{P}_s, \phi)), Z_\alpha) \leq (1 \wedge \alpha^{-1}) \int \left( \int \phi(x, y) s \mu(dy) \right)^2 s \mu(dx). \quad (7.2)$$

*Proof.* We shall use Theorem 3.1. Given  $x, y \in \mathbb{X}$  with  $\phi(x, y) > 0$ , set  $\tilde{\mathcal{G}}_s = G_\phi(\eta_s \cup \{(x, \tau_1), (y, \tau_2)\})$ . Let  $G'(\mathcal{P}_s, \phi)$  denote the subgraph of  $\tilde{\mathcal{G}}_s$  induced by the vertex set  $\mathcal{P}_s$ , which has the same distribution as  $G(\mathcal{P}_s, \phi)$ .

Set  $U_{x,y} = H_2(G'(\mathcal{P}_s, \phi))$  and  $V_{x,y} = H_2(\tilde{\mathcal{G}}_s) - \mathbf{1}\{\{x, y\} \in E(\tilde{\mathcal{G}}_s)\}$ , where  $E(G)$  denotes the set of edges of a graph  $G$ . Then  $V_{x,y}$  has the conditional distribution of  $H_2(\tilde{\mathcal{G}}_s) - 1$  given that  $\{x, y\} \in E(\tilde{\mathcal{G}}_s)$ . Also  $V_{x,y} \geq U_{x,y}$  and

$$\mathbb{E}[V_{x,y} - U_{x,y}] = \int_{\mathbb{X}} (\phi(x, z) + \phi(y, z)) s \mu(dz) =: w(x, y).$$

Hence by Theorem 3.1 the left hand side of (7.2) is bounded by the expression

$$\begin{aligned} & (1/2)(1 \wedge \alpha^{-1}) \int \int w(x, y) \phi(x, y) s^2 \mu(dx) \mu(dy) \\ &= (1 \wedge \alpha^{-1}) \int \left( \int \phi(x, y) s \mu(dy) \right)^2 s \mu(dx), \end{aligned}$$

as required.  $\square$

For example, consider the geometric setting with  $\mu$  having bounded, almost everywhere continuous density  $f$  with respect to Lebesgue measure on  $\mathbb{R}^d$ . Suppose  $\phi_s(x, y) = \phi(r_s^{-1}(x - y))$  for some fixed integrable, symmetric and almost everywhere continuous  $\phi$ , and  $r_s > 0$  satisfying  $s^2 r_s^d \rightarrow \beta$  for some  $\beta > 0$ . Set  $\alpha_s := \mathbb{E} H_2(G(\mathcal{P}_s, \phi_s))$ . Then by (7.1),

$$\alpha_s \rightarrow \frac{\beta}{2} \left( \int \phi(z) dz \right) \int f(x)^2 dx =: \alpha.$$

By (7.2) we have  $d_{TV}(\mathcal{L}(H_2(G(\mathcal{P}_s, \phi_s))), \mathcal{L}(Z_{\alpha_s})) = O(s^3 r_s^{2d}) = O(s^{-1})$ .

In particular  $H_2(G(\mathcal{P}_s, \phi_s)) \xrightarrow{\mathcal{D}} Z_{\alpha}$ . This could possibly also be proved by deriving a Poisson limit for  $H_2(G(\mathcal{X}_n, \phi_n))$  by adapting the argument in [19, Theorem 3.4] to the RCM, and Poissonizing. However, the Poissonization would seem to introduce an error of at least  $s^{-1/2}$  in the total variation distance, so the rate of convergence would not be as good.

*Proof of Theorem 2.4.* Let  $k \in \mathbb{N}$  with  $k \geq 2$ . Let  $\varepsilon > 0$  and assume  $\phi_s \in \Phi_{\varepsilon}$  for all  $s > 0$ . Set  $G_s := G(\mathcal{P}_s, \phi_s)$ . Set  $\kappa_s = \sup_{x \in \mathbb{X}} \int \phi_s(x, y) \mu(dy)$ . Then with  $h_{\phi}(\cdot)$  defined at (6.2), by (3.5) we have

$$\mathbb{E} H_k(G_s) = \frac{1}{k!} \int s^k h_{\phi_s}(x_1, \dots, x_k) \mu^k(d(x_1, \dots, x_k)) = \Theta(s^k \kappa_s^{k-1}), \quad (7.3)$$

where the second relation is obtained similarly to the proof of Lemma 6.1, using the fact that for each connected graph  $\Gamma$  on  $\{1, \dots, k\}$  we can integrate the variables  $x_1, \dots, x_k$  in an order  $(x_{\sigma(1)}, \dots, x_{\sigma(k)})$  such that for  $2 \leq i \leq k$  each successive  $\sigma(i)$  is connected in  $\Gamma$  to one or more of  $\sigma(1), \dots, \sigma(i-1)$ . Then each successive integral gives another factor of  $\Theta(\kappa_s)$ .

Let  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$ , with  $x_1, \dots, x_k$  distinct and with  $h_{\phi_s}(x_1, \dots, x_k) > 0$ . Let  $\tilde{\mathcal{G}}_s$  be the graph defined by (6.13), but now conditioned on the subgraph induced by  $\{x_1, \dots, x_k\}$  being connected. Denote by  $G'(\mathcal{P}_s, \phi_s)$  the subgraph of  $\tilde{\mathcal{G}}_s$  induced by  $\mathcal{P}_s$ . Set  $U_{\mathbf{x}} = H_k(G'(\mathcal{P}_s, \phi_s))$  and  $V_{\mathbf{x}} = H_k(\tilde{\mathcal{G}}_s) - 1$ . Then  $U_{\mathbf{x}}$  has the distribution

of  $H_k(G_s)$  and  $1 + V_{\mathbf{x}}$  has the conditional distribution of  $H_k(G(\mathcal{P}_s \cup \{x_1, \dots, x_k\}, \phi_s))$  given that  $\{x_1, \dots, x_k\}$  induces a connected subgraph of this graph.

Now  $V_{\mathbf{x}} \geq U_{\mathbf{x}}$  and we assert that

$$\mathbb{E}[V_{\mathbf{x}} - U_{\mathbf{x}}] = O\left(\sum_{j=1}^{k-1} (s^j \kappa_s^j)\right), \quad (7.4)$$

uniformly over  $\mathbf{x} \in \mathbb{X}^k$ . To see this, observe that  $V_{\mathbf{x}} - U_{\mathbf{x}}$  is the number of pairs  $(\mathcal{X}, \mathcal{Y})$  with  $\mathcal{X}$  a non-empty subset of  $\{x_1, \dots, x_k\}$  and  $\mathcal{Y}$  a non-empty subset of  $\mathcal{P}_s$ , such that the subgraph of  $\tilde{\mathcal{G}}_s$  induced by vertex set  $\mathcal{X} \cup \mathcal{Y}$  is a connected graph of order  $k$  (to ease notation we ignore the issue of multiplicities in this notation). But then, similarly to the proof of Lemma 6.1, we can take the successive elements  $y_i$  of  $\mathcal{Y}$  in an order such that each of them is connected to at least one existing vertex from  $\mathcal{X} \cup \{y_1, \dots, y_{i-1}\}$ . Then each successive integral gives another factor of  $O(s\kappa_s)$ .

For part (a), assume  $\mathbb{E} H_k(G_s) \rightarrow \alpha \in (0, \infty)$ . Then by (7.3) and (7.4) we have  $\mathbb{E}|V_{\mathbf{x}} - U_{\mathbf{x}}| = o(1)$ , uniformly over  $\mathbf{x} \in \mathbb{X}^k$ . Hence by Theorem 3.1,  $H_k(G_s) \xrightarrow{\mathcal{D}} Z_{\alpha}$ , which is part (i).

For part (b), set  $\alpha_s = \mathbb{E} H_k(G_s)$ , and assume  $\alpha_s \rightarrow \infty$  but  $\alpha_s = o(s)$ . Then by (7.3), the assumption  $\alpha_s = o(s)$  implies that (recalling  $k \geq 2$ ) we have  $s\kappa_s \rightarrow 0$ . Then by (7.4) and Theorem 3.1, we have that  $d_{TV}(\mathcal{L}(H_k(G_s)), \mathcal{L}(Z_{\alpha_s})) \rightarrow 0$ . Since also  $(Z_{\alpha_s} - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ , we therefore have  $(H_k(G_s) - \alpha_s)/\sqrt{\alpha_s} \xrightarrow{\mathcal{D}} \mathcal{N}$ .  $\square$

## 8 U-statistics of a Poisson process

Let  $k \in \mathbb{N}$ , and let  $\mathbf{S}_k(\mathbb{X}) := \{\xi \in \mathbf{S}(\mathbb{X}) : |\xi| = k\}$ . Let  $h : \mathbf{S}_k(\mathbb{X}) \rightarrow \{0, 1\}$  be measurable. For  $\xi \in \mathbf{S}(\mathbb{X})$  set

$$F(\xi) = \sum_{\psi \subset \xi : |\psi| = k} h(\psi).$$

Let  $s > 0$ , and let  $\eta$  be a Poisson process on  $\mathbb{X}$  with mean measure  $\lambda := s\mu$ . We seek to apply Theorem 3.1 to  $W := F(\eta)$  for this class of choices of  $F$ , called  $U$ -statistics of the Poisson process  $\eta$ . Assume  $\mu$  is diffuse; once again, this assumption is for notational convenience only.

For  $\mathbf{x} = (x_1, \dots, x_k) \in \mathbb{X}^k$  with  $h(\{x_1, \dots, x_k\}) = 1$ , set  $U_{\mathbf{x}} = F(\eta)$ , and

$$V_{\mathbf{x}} = F\left(\bigcup_{i=1}^k \{x_i\} \cup \eta\right) - 1.$$

Clearly  $V_{\mathbf{x}}$  and  $U_{\mathbf{x}}$  have the required distributional properties in the statement of Theorem 3.1. Also  $V_{\mathbf{x}} \geq U_{\mathbf{x}}$ , and

$$|V_{\mathbf{x}} - U_{\mathbf{x}}| = \sum_{J \subset [k] : 1 \leq |J| < k} \sum_{\psi \subset \eta : |\psi| = k - |J|} h(\psi \cup \{x_i : i \in J\}),$$

so using the Mecke formula and integrating over  $\mathbf{x}$  we obtain that

$$\begin{aligned} \int (\mathbb{E} |V_{\mathbf{x}} - U_{\mathbf{x}}|) h(\mathbf{x}) \lambda^k(d\mathbf{x}) &= \sum_{\ell=1}^{k-1} \frac{1}{(k-\ell)!} \binom{k}{\ell} \int \int h(\{x_1, \dots, x_\ell, y_1, \dots, y_{k-\ell}\}) \\ &\quad \times \lambda^{k-\ell}(d(y_1, \dots, y_{k-\ell})) h(\{x_1, \dots, x_k\}) \lambda^k(d\mathbf{x}) \\ &= \sum_{\ell=1}^{k-1} \frac{k!}{\ell!(k-\ell)!^2} \int \left( \int h(\{x_1, \dots, x_\ell, y_1, \dots, y_{k-\ell}\}) \lambda^{k-\ell}(d(y_1, \dots, y_{k-\ell})) \right)^2 \\ &\quad \times \lambda^\ell(d(x_1, \dots, x_\ell)) \\ &=: \gamma(h, \lambda). \end{aligned}$$

By Theorem 3.1, if  $\mathbb{E} F(\eta) = \alpha$ , then  $d_{TV}(\mathcal{L}(F(\eta)), \mathcal{L}(Z_\alpha)) \leq (1 \wedge \alpha^{-1})\gamma(h, \lambda)/k!$ . Also  $d_W(\mathcal{L}(F(\eta)), \mathcal{L}(Z_\alpha)) \leq 3(1 \wedge \alpha^{-1/2})\gamma(h, \lambda)/k!$ . This bound is comparable to the one obtained in Theorem 7.1 of [5]. Our bound has an extra factor  $1 \wedge \alpha^{-1/2}$  in front, which may make it better when  $\alpha$  is large. Also, unlike [5] we do not make any topological assumptions on the measurable space  $\mathbb{X}$ . As remarked just after the statement of Theorem 3.1, it is possible to extend that result to the case where the measure  $\lambda$  is  $\sigma$ -finite, and hence to extend the above argument likewise, but we do not go into details here.

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## References

- [1] Barbour, A., Holst, L. and Janson, S. (1992). *Poisson Approximation*. Oxford Studies in Probability **2**, Oxford University Press, Oxford.
- [2] Bhamidi, S., van der Hofstad, R., and van Leeuwaarden, J. S. H. (2012). Novel scaling limits for critical inhomogeneous random graphs. *Ann. Probab.* **40**, 2299–2361.
- [3] Bollobás, B. (2001). *Random Graphs*. Second edition. *Cambridge Studies in Advanced Mathematics*, **73**. Cambridge University Press, Cambridge, 2001.
- [4] Bollobás, B., Janson, S. and Riordan, O. (2007). The phase transition in inhomogeneous random graphs. *Random Structures Algorithms* **31**, 3–122.

- [5] Decreusefond, L., Schulte, M. and Thäle C. (2016). Functional Poisson approximation in Kantorovich-Rubinstein distance with applications to U-statistics and stochastic geometry. *Ann. Probab.* **44**, 2147–2197.
- [6] Devroye, L. and Fraiman, N. (2014). Connectivity of inhomogeneous random graphs. *Random Structures Algorithms* **45**, 408–420.
- [7] Díaz, J., Petit, J., and Serna, M. (2000). Faulty random geometric networks. *Parallel Process. Lett.* **10**, 343–357.
- [8] Franceschetti, M. and Meester, R. (2007). *Random Networks for Communication*. Cambridge University Press, Cambridge.
- [9] Gupta, B. and Iyer, S. K. (2010). Criticality of the exponential rate of decay for the largest nearest-neighbor link in random geometric graphs. *Adv. in Appl. Probab.* **42**, 631–658.
- [10] Gupta, P. and Kumar, P. R. (1999). Critical power for asymptotic connectivity in wireless networks. *Stochastic Analysis, Control, Optimization and Applications*, 547–566, *Systems Control Found. Appl.*, Birkhäuser Boston, Boston.
- [11] Hoff, P. D., Raftery, A. E. and Handcock, M.S. (2002). Latent space approaches to social network analysis. *J. Amer. Statist. Assoc.* **97**, 1090–1098.
- [12] Hsing, T. and Rootzén, H. (2005). Extremes on trees. *Ann. Probab.* **33**, 413–444.
- [13] Iyer, S. K. and Thacker, D. (2012). Nonuniform random geometric graphs with location-dependent radii. *Ann. Appl. Probab.* **22**, 2048–2066.
- [14] Kingman, J. F. C. (1993). *Poisson Processes*. Oxford Studies in Probability **3**. Oxford University Press, Oxford.
- [15] Last, G. and Penrose, M. (2017). *Lectures on the Poisson Process*. Cambridge University Press, Cambridge.
- [16] Lindvall, T. (1992). *Lectures on the Coupling Method*. Wiley, New York.
- [17] Mao, G. and Anderson, B. D. O. (2012). Towards a better understanding of large-scale network models. *IEEE/ACM Transactions on Networking* **20**, 408–421.
- [18] Meester, R. and Roy, R. (1996). Continuum percolation. *Cambridge Tracts in Mathematics*, **119**. Cambridge University Press, Cambridge,
- [19] Penrose, M. (2003). *Random Geometric Graphs*. Oxford Studies in Probability **5**, Oxford University Press, Oxford.

- [20] Penrose, M. D. (1998). Extremes for the minimal spanning tree on normally distributed points. *Adv. in Appl. Probab.* **30**, 628–639.
- [21] Penrose, M. D. (1999). On  $k$ -connectivity for a geometric random graph. *Random Structures Algorithms* **15**, 145–164.
- [22] Penrose, M. D. (2016). Connectivity of soft random geometric graphs. *Ann. Appl. Probab.* **26**, 986–1028.
- [23] Rastelli, R., Friel, N. and Raftery, A. E. (2016). Properties of Latent Variable Network Models. *Network Science* **4**, 407–432.
- [24] Snijders, T.A.B. and Nowicki, K. (1997). Estimation and prediction for stochastic blockmodels for graphs with latent block structure. *J. Classification* **14**, 75–100.
- [25] Söderberg, B. (2002). General formalism for inhomogeneous random graphs. *Phys Rev. E* **66**, 066121
- [26] Yi, C.-W., Wan, P.-J., Lin, K.-W., and Huang, C.-H. (2006). Asymptotic distribution of the number of isolated nodes in wireless ad hoc networks with unreliable nodes and links. *Global Telecommunications Conference, 2006. GLOBECOM '06. IEEE.*