



Article Integrating Replenishment Policy and Maintenance Services in a Stochastic Inventory System with Bilateral Movements

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Abstract: We study an inventory control problem with two storage facilities: a primary warehouse (PW) of limited capacity M, and a subsidiary one (SW) of sufficiently large capacity. Two types of customers are considered: individual customers arriving at (positive and negative) linear rates governed by a Markov chain, and retailers arriving according to a Markov arrival process and bringing a (positive and negative) random number of items. The PW is managed according to a triple-parameter band policy (M, S, s), $0 \le s < S \le M$, under a lost sales assumption. Under this policy, as soon as the stock level at the PW falls below s, a refilling to S is performed by a distributor after a random lead-time. However, if the stock exceeds level S when the distributor arrives, no refilling is carried out, and only maintenance services are performed. Items that exceed level M are transferred to the SW at a negligible amount of time for those used in related products. Our cost structure includes a fixed order cost, a variable cost for each item supplied by the distributor, a cost for the additional maintenance, a salvage payment for each transferred item from the PW to the SW, and a loss cost for each unsatisfied item due to demands. We seek to determine the optimal thresholds that minimize the expected overall cost under the discounted criterion. Applying first-passage time results, we present a simple set of equations that provide managers with a useful and an efficient tool to derive the optimal thresholds. Sensitivity analysis and fruitful conclusions along with future scope of research directions are provided.

Keywords: inventory; band policy; MAP; Markov chain; first-passage times

MSC: 90B05; 60J28; 60G51

1. Introduction

An inventory model is characterized by an inconsistency and volatility in the flow of items in and out of the supply network. Inventory management has been studied for decades, and volatile market conditions have increased the complexity of modeling and analyzing supply chains. Components such as demands, returns, delivery times, and collaborative initiatives among partners have become more variable. This increasing uncertainty has a direct impact on the inventory level. Although most inventory models assume that the manager owns a storage facility of unlimited capacity, it has been observed that, in real life, this assumption may be unrealistic. When inventory capacity is limited, the on-hand inventory may exceed the warehouse's capacity and need to be transferred to an external warehouse of insufficient capacity. Managing integrated storage facilities effectively and efficiently is an increasingly important task for companies in order to gain competitive advantages [1].

In the present paper, we introduce an integrated inventory management problem with two types of storage facilities: a primary warehouse (PW) of limited capacity *M*, and a subsidiary warehouse (SW) of sufficiently large capacity. The main storage facility, the PW, is used for ongoing demands and returns. We consider two types of customers arriving to the PW, individual customers and retailers; both arrivals are characterized by a continuous-time Markov chain (CTMC). Each individual customer demands or returns



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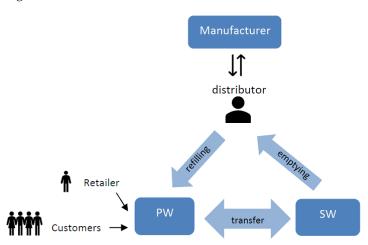
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Copyright: © 2023 by the author. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). a unit, and the inter-arrival times between successive arrivals are negligible. Thus, these arrivals can be naturally approximated by continuous linear rates, where negative rates represent demands, and positive rates represent returns. In addition, we consider a stream of retailers, each bringing a positive or negative number of items. Here, the stock level jumps down (for a demand) or up (for a return) at the arrival instances, and the batch sizes are independent and identically distributed (i.i.d.) random variables (r.v.s), having phase-type (PH) distributions (see [2] for more details on the phase-distributions).

We assume that the management of the PW is assisted by the services of an outsourcing partner (for our needs, called the*distributor*), and directly affects the inventory level of the SW. Specifically:

- (a) Management of PW. (i) The PW is managed according to a triple-parameter band policy (M, S, s) $(0 \le s < S \le M)$, i.e., when the on-hand stock drops to some level *s*, an order is placed to purchase more stock from the manufacturer in order to raise the current level up to level *S*. As in practice, the order arrives by the distributor after some random time (called the lead time); nevertheless, it is assumed that, during that lead time, new orders are not allowed. Upon arrival, the PW is refilled up to level *S*; in addition, maintenance activities are provided by the distributor. It may happen that, when the distributor arrives, the stock exceeds *S* due to returns; in such a case, no refilling is performed at the PW (however, the distributor is still employed in these maintenance activities; see Point (c) below). Furthermore, (*ii*) any demand or part of the demand that is not satisfied is lost. (*iii*) The PW has a limited storage capacity of *M* items; each time the on-hand stock exceeds level *M*, the excess amount is transferred to the SW in a negligible amount of time.
- (b) *Storage at the SW*. The SW has unlimited capacity. We assume that each unit of material that is transferred from the PW is processed as one unit of material to be used in the SW; the transferred material is accumulated and used to satisfied related products. Thus, the SW is not accessible to customers.
- (c) Inventory management. Each time the distributor arrives, he provides some basic maintenance services, including cleaning, organizing, and emptying the SW, as needed, and the SW stores with zero items. We assume that these maintenance services are an integral part of the distributor's duties; therefore, no additional payment is due for this work. However, if the PW is not refilled, the distributor is rewarded (a kind of monetary compensation) for his services.



A schematic diagram of the relationship between the existing entities is presented in Figure 1.

Figure 1. The relationship between the manufacturer, distributor, customers, and warehouses.

Our model is motivated by various practical settings. The first example is taken from the collaboration between firms and farms—specifically, the vegetable marketing firm

Chasalat Alei Katif (https://www.aleikatif.org) (27 December 2022), which specializes in growing insect-free, fresh cabbage without using pesticides. The firm markets products from several farms, each with several greenhouses for growing organic cabbage. The ripe cabbage is stored in the PW, which serves both individual customers and retailers. When the stock in the PW falls below a certain level, the farms supply the firm with additional stock, which is time-consuming. However, according to a special agreement, if the supplied cabbages cannot be stored in the PW (due to capacity constraints), they are sliced, and accumulated until they are sold privately by the farms.

The second example is taken from the collaboration between medical centers and university scientists [3]. Here, specialists are responsible for the operation and management of the blood bank. When a lack of blood donations is observed, the associated medical faculty is enlisted to help carry out a rapid and immediate blood donation campaign on campus. Since the capacity of the storage facility is limited, the surplus blood donations are transferred to the university faculty for academic use, after which it cannot be used by patients.

In this paper, we assume the following costs and rewards: (i) a fixed cost for each order; (ii) a cost for each item supplied by the distributor (this cost includes the payment to the distributor for his maintenance services). (iii) If the PW is not refilled due to a high inventory level, the distributor is rewarded for their maintenance services; (iv) a salvage payment for each item transferred from the PW to the SW; and (iv) a loss cost is charged for each unsatisfied item at the PW.

We seek to determine the optimal levels of *s* and *S* that minimize the overall discounted cost of managing the warehouses. We further assume a fixed cost to operate the PW and, thus, each item has a negligible storage cost. This assumption has many applications in reality, e.g., in blood inventory management and in organic food storage, where the cost of storing an item is negligible compared to the cost of cooling the storage facility. Moreover, we assume that the distributor arrives after an exponentially distributed lead time. This assumption is practical when the lead time depends on different logistics factors—for example, for a distributor that serves several independent companies that line up as an M/M/1 queue. The time it takes to prepare and deliver the items can be interpreted as a sojourn time that is exponentially distributed. Due to variable delivery times, the company may ask the distributor for a flexible contract that allows for delays or cancellations of outstanding orders.

The policy addressed here expands the band (S, s)-type policy. As far as we know, most of the existing studies under the (S, s) control policy consider continuous movements and allows at most unilateral jumps; there are no studies on inventory management that consider a Markov-modulated process with bilateral continuous and jump movements. However, when considering the significant increase in returns and the uncertainty in lead times, allowing double-side changes is more appropriate for modeling inventory management. Fluctuations in inventory levels, where the inventory level not only decreases due to demands but also increases due to returns, make the analysis much more challenging. From a managerial perspective, this difficulty becomes the main driving force behind for an integrated inventory management.

The main contributions of the paper are fourfold: (1) We differ from existing works in the inventory management literature by considering continuous and batch-type bilateral changes, which are both governed by the underlying Markov chain. We further assume phase-distributed batch amounts, given that any nonnegative continuous distribution can be approximated by a phase distribution. By that, we capture the uncertainty of customers' consumption habits impacted by a fluctuating environment. As a result, our framework can be adapted to a wide range of applications. (2) We extend classic inventory models to include a limited storage capacity and, thus, introduce a simple and practical (M, S, s)-type band policy. As in practice, we further assume that the fixed cost is paid at the time that the order is placed, and that the variable cost is paid at the time the PW is refilled. This difference in timing is significantly important and may lead to cost saving. (3) We also con-

tribute to the inventory management literature by implementing the new business concept integrating partnerships for managing the warehouses. (4) Most existing papers employ analytic approaches, such as a dynamic programming approach [4,5] difference-differential equations approach [6] or scale function [7,8]; by contrast, we use a more probabilistic and intuitive approach. We combine the first-step conditioning technique for the expectation, renewal theory and the strong Markov property (which says that, for a Markovian process X(t), conditioning on state at a given stopping time T, the probabilistic behavior of the process depends only on its value at time T and discards its past behavior; see [2]). By doing so, we build a relatively simple set of equations that provide managers with a useful and efficient tool to derive optimal thresholds. To the best of our knowledge, this is the first time that first-passage time results have been carried out in bilateral movements in the inventory management literature. We further note that many studies deal with the long-run average criterion, without taking timing into consideration [5,9]. Our results show that timing significantly impacts the system performance and is a crucial factor; therefore, it should be taken into account.

From our conclusions, we numerically glean some important insights. Firstly, the total cost seems to be convex in S and s (for a fixed M; see Figure 4. Although we cannot prove this result, it enables us to numerically obtain the optimal thresholds S^* and s^* using a linear search over the range $0 \le s < S < M$. Secondly, we show the impact of the limited capacity of the PW on the optimal thresholds, especially when the transfer cost is high (see Figure 4a, Table 2, and the discussions thereafter). Thirdly, as demands become more frequent, it becomes worthwhile to order more frequently and for smaller quantities. In such a case, the impact on S^* is negligible and, surprisingly, the weight of the cost's components is relatively fixed (here, the transfer cost is negligible). By contrast, more returns prompt the manager to reduce the frequency of orders and enlarge their quantities. Here, the weight of the cost of the distributor's services increases relative to the weight of the loss cost (see Figures 5–8). Fourthly, investigating the impact of the timing, we observe that, when the discount factor is high, postponing the call for the distributor becomes economically profitable (see Figure 9). The outline of the paper is as follows: In Section 2, we review the related literature. Section 3 presents the mathematical structure of the stock levels at the PW and SW; the cost components are also detailed. The core preliminaries are given in Section 4. Section 5 derives the expected discounted cost components; a summary of the costs' derivation is given in Appendix A.2. Numerical examples, observations, and insights are provided in Section 6. Finally, Section 7 concludes. Technical parts of the proofs are relegated to Appendices B.1–B.3.

2. Literature Review

It is well established in the literature that customer flows can exhibit high variability and unpredictably [10,11], state dependency [12], and batch patterns [13,14]. Thus, modeling the behavior of customers has become a key challenge in inventory management. These modeling challenges are aggravated by the rapid spread of the home shopping phenomenon, which force retailers to incorporate the occurrence of returns on a daily basis [4]. Real-word examples of such variability include the introduction of new products, where the flow rates are obviously non-stationary and evolve throughout the stages of the product's life cycle [15,16], fluctuating environments, where economic conditions, extreme weather, technological advancements, competition, and dynamic events can perturb customer flows [17–19], home shopping and internet retailing, where a high return rate, particularly for electronic and computer devices, significantly changes customers' behavior [20,21]. Any of these sources of variability on its own impacts the on-hand stock level and, thus, complicates the analysis of the inventory model. Therefore, a considerable portion of the related literature usually uses Markov processes due to their versatility in matching key statistical properties of the customers' consumption needs [10,13,22,23].

In this study, the inventory level changes are the sum of continuous linear rates forming a likewise fluid process, and instantaneous big inflows and outflows governed by a Poisson process. For example, a company that sells its product through two channels, usual daily sell contracts and one-time opportunities [24], or a manufacturer who produces a component needed for few products, as well as a replacement [25]. Consequently, an excellent example of a return in batches policy is out-of-fashion products, and parts delivered to maintenance service engineers, particularly in isolated areas [4].

We assume that the batch sizes come from the family of phase distributions. The advantage of a PH distribution is that its Markovian nature allows for an exact analysis and performance evaluation; thus, they are often used [12,13,26,27]. When general distributions are appropriate, phase distributions can be taken into account in a natural way, as any nonnegative continuous distribution of the probability can be approximated with a phase-type distribution [2].

We note that incorporating continuous rates and instantaneous jumps is well studied also in economics and cash management (e.g., loads and withdrawals), in healthcare management (e.g., daily patients and unexpected disasters), in reliability models (e.g., parts delivered to maintenance service engineers in geographically spread out areas), and in chemical production systems and gas stations. More examples are available in [28–30].

In the literature, the (S, s) control policy is one of the most widely used [24]. The optimality of (S, s)-type policies has been investigated for various inventory models, including those with discrete and continuous time reviews, different time horizons, discounted or average cost criteria, and backlogging or lost-sales. Under continuous time, the underlying stochastic stock level governing the state variables is typically linked to either varying rates [4,31], fluid processes [32], Wiener processes [9,33–36], renewal processes [5,12,37,38], superposition of deterministic rates combined with a renewal process [25], and a one-sided Lèvy process [7,39,40]. Using batch-size patterns, Bensoussan et al. [41] consider two models: one model is a mixture of a diffusion process and a compound Poisson process with exponentially distributed jump sizes, and the second model, as in Presman and Sethi [25], is a mixture of constant demand governed by a compound Poisson process. Their work was generalized by Benkherouf and Bensoussan [6] to a general compound Poisson process. Yamazaki [8] tackled the inventory problem when the demand follows a Lèvy process with jumps having infinite activity/variation. Chakravarthy and Rao [42] assume that the customers arrive according to a Markovian arrival process (MAP) with demand of varying sizes. Along with this line, Barron [13,18] studies a mixture of a fluid process and a MAP demand process. For a comprehensive survey, see [24].

All papers cited above consider a one-dimensional inventory process. For the multidimensional process, several models used in the literature focus on the derivation of the Laplace transform of the first-passage time in a Markov-modulated process with bilateral random jumps, mostly for studying various financial problems (e.g., [26,39,43–45]).

Furthermore, the increasing environment uncertainty has been mitigated over the last few decades, the phenomenon of a gradual transition to collaborative and integrative approaches to achieving optimal supply chain performance. In this paper, we assume that the distributor also assists with the maintenance and upkeep of the storage. Particularly, it has become customary for managers to motivate high-quality relationships with their distributor. Research on marketing shows that most distributors also provide a range of services such as technical support, warranty, and other complex services [46–48].

To outline our position, an overview of the most relevant literature studies concerning the continuous-review base stock policy is given in Table A1 in Appendix A.1. To the best of our knowledge, the combination of random demands and returns (bilateral continuous and jump type) in supply chain collaboration under the (S, s) policy with cancellation has not been explored in literature of inventory management; hence, the model developed here significantly contributes to the literature.

3. The Stock Level Processes at PW and SW

In this section, we use similar notations to those of Breuer [27] (see also [18,26]).

Consider two storage facilities, the PW and the SW, and two types of customers, namely, individual customers and retailers. The demands (returns) of the individual customers form a likewise fluid process with linear rates, governed by a continuous-time Markov chain (CTMC) that is used as a background environment as follows. Let $\tilde{\mathcal{J}} = \{\tilde{J}(t) : t \ge 0\}$ be a CTMC with state space \tilde{E} , initial probability vector $\eta = [\eta_1, \eta_2, \dots, \eta_{|\tilde{E}|}]$, and infinitesimal generator $\tilde{\mathbb{Q}} = (\tilde{q}_{ij})_{i,j\in\tilde{E}}$. Let π be the stationary probability vector; π is the unique solution of the equation $\pi \tilde{\mathbb{Q}} = 0$ such that $\pi e = 1$. When the state equals $i \in \tilde{E}$, products are returned at rate r_i , and demand is observed at rate d_i by the individual customers. The growth rate is denoted by $c_i = r_i - d_i$, which can be positive or negative. Thus, the state space \tilde{E} is composed of two subsets $\tilde{E} = E_p \cup E_n$, where E_p includes the increasing states, $E_p = \{i : c_i > 0\}$, and E_n includes the decreasing states, $E_n = \{i : c_i < 0\}$. Accordingly, we use the terms ascending (descending) environment when $i \in E_p (E_n)$.

Next, we describe the retailers' arrivals. The retailers arrive at random times and form a Poisson process of instantaneous big inflows (returns) or outflows (demands). Specifically, we assume that, when an ascending environment is observed, i.e., when there are more returns than demands, the manager offers retailers the opportunity to demand products, in order to balance the stock level. On the other hand, when a descending environment is observed, i.e., when there are more demands than returns, the manager offers retailers the ability to return products.

For the mathematical description, we define the process $\widehat{\mathcal{I}} = \{\widehat{I}(t) : t \ge 0\}$ as a special case of a Lèvy process $\widehat{\mathcal{I}}^{(i)}$ with a drift c_i (to describe the customers) and a Lèvy measure \widehat{v}_i (to describe the retailers) during intervals when the phase equals $i \in \widetilde{E}$. We assume that \widehat{v}_i takes the form

$$\widehat{v}_{i}(dx) = \lambda_{i}^{(+)} \mathbf{1}_{\{i \in E_{n}\}} \boldsymbol{\alpha}^{(i+)} \exp(\mathbb{T}^{(i+)}x) \boldsymbol{\eta}^{(i+)} dx + \lambda_{i}^{(-)} \mathbf{1}_{\{i \in E_{p}\}} \boldsymbol{\alpha}^{(i-)} \exp(-\mathbb{T}^{(i-)}x) \boldsymbol{\eta}^{(i-)} dx, \ x > 0$$
(1)

for all $i \in \tilde{E}$. Here, $\lambda_i^{(\pm)} \geq 0$ is the retailers' arrival rates, and the retailers' batch amounts $U^{(i\pm)}$ are PH distributed r.v.s with initial probability vectors $\boldsymbol{\alpha}^{(i\pm)}$ and transition rate matrices $\mathbb{T}^{(i\pm)}$; i.e., the arrival process is a compound Poisson process with jump sizes $U^{(i+)} \sim PH(\boldsymbol{\alpha}^{(i+)}, \mathbb{T}^{(i+)})$ of order m_i^+ when $i \in E_n$, and $U^{(i-)} \sim PH(\boldsymbol{\alpha}^{(i-)}, \mathbb{T}^{(i-)})$ of order m_i^- when $i \in E_p$. Let $\boldsymbol{\eta}^{(i\pm)} = -\mathbb{T}^{(i\pm)}\boldsymbol{e}$ be the exit vectors. Now, the process $\hat{\mathcal{I}}$ is a MAP-modulated process with upward jumps and downward jumps. Let $\boldsymbol{\lambda}^{(+)}$ be a $(1 \times |E_n|)$ vector whose i component is $\lambda_i^{(-)}$, $i \in E_p$. Finally, define the $(1 \times \sum_{i=1}^{|E_n|} m_i^+)$ vector $\boldsymbol{\alpha}^{(+)} = (\boldsymbol{\alpha}^{(1+)}, \dots, \boldsymbol{\alpha}^{(|E_n|+)})$, and the $(1 \times \sum_{i=1}^{|E_p|} m_i^-)$ vector $\boldsymbol{\alpha}^{(-)} = (\boldsymbol{\alpha}^{(1-)}, \dots, \boldsymbol{\alpha}^{(|E_p|-)})$.

The PW is controlled by a triple-parameter band policy (M, S, s), $0 \le s < S \le M$, under a lost sales policy. Let $\tilde{\mathcal{I}}_o = \{\tilde{I}_o(t) : t \ge 0\}$ be the on-hand stock level at the PW. We assume that, as soon as $\tilde{\mathcal{I}}_o$ drops to or below level *s*, a delivery is ordered and arrives by the distributor after an exponential lead time; during that lead time, new orders are not allowed. When the distributor arrives, if the on-hand stock increases to or above level *S*, no items are refilled; otherwise, the stock is refilled up to level *S*. Thus, the amount refilled at arrival time is random and is equal to $\max(S - \tilde{I}_o(t), 0) \equiv \left[S - \tilde{I}_o(t)\right]^+$ items. We further assume that the PW has a capacity of M > S units. Every time the inventory exceeds level *M*, the excess items are transferred from PW to SW in a negligible amount of time and for some fee. Hence, $\tilde{\mathcal{I}}_o$ lies in the range [0, M].

Next, we describe the management of the SW. The SW has unlimited capacity. Each time the distributor arrives, he provides some basic maintenance activities and empties the SW; no additional payment is given to him for this work. However, if the PW is not refilled, the distributor is rewarded for his service. Let $\tilde{I}_R = {\tilde{I}_R(t) : t \ge 0}$ be the on-hand stock

level at the SW; Note that, in contrast to $\hat{\mathcal{I}}_o$, the process $\hat{\mathcal{I}}_R$ is not bounded from above. For simplicity, we further assume that $\tilde{\mathcal{I}}_o(0) = S$ and $\tilde{\mathcal{I}}_R(0) = 0$.

Let \mathcal{L}_n , $n \ge 1$ be a sequence of i.i.d. $exp(\mu)$ -distributed r.v.s representing the lead times independent of $(\tilde{\mathcal{I}}_o, \tilde{\mathcal{J}})$. Let $\mathcal{Z}_1 = \inf\{t : t > 0, \tilde{I}_o(t) \le s\}$ be the first time that the stock level drops to or below level *s*, and thus a delivery is ordered. Let $K_1 = \mathcal{Z}_1 + \mathcal{L}_1$ be the first arrival time of the distributor. We define, recursively, the following stopping times:

$$\mathcal{Z}_n = \inf\left\{t: t > \sum_{i=1}^{n-1} K_i, \widetilde{I}_o(t) \le s\right\}, \quad n = 2, 3, \dots$$

$$K_n = \mathcal{Z}_n + \mathcal{L}_n, \qquad n = 2, 3, \dots$$
(2)

The r.v. Z_n , n = 1, 2, ..., represents the *n*-thorderingtime; the r.v. K_n , n = 1, 2, ... represents the *n*-th arrival time of the distributor. Recall that $\tilde{I}_o(Z_n) = s$ occurs only due to a continuous hitting linear rate, i.e., by a customer's demand, while $\tilde{I}_o(Z_n) < s$ occurs due to a downward jump, i.e., by a retailer's demand. At arrival times, $\left[S - \tilde{I}_o(K_n)\right]^+$ items are refilled, such that, if $\tilde{I}_o(K_n) < S$, then \tilde{I}_o is refilled with $S - \tilde{I}_o(K_n)$ items and starts from level *S*; otherwise, no refilling is carried out. In addition, at times K_n , n = 1, 2, ..., the SW is emptied. Thus, we have $\tilde{I}_o(K_n) \ge S$, $\tilde{I}_R(K_n) = 0$, n = 1, 2, ... Finally, we define J_i as

$$j_i = \min\{K_n > j_{i-1} : \widetilde{I}_o(K_n) < S\}, n = 1, 2, \dots\}, i = 1, 2, \dots, \text{with } j_0 = 0,$$
(3)

to represent the *i*-th (actual) *refilling* time of the PW (and emptying time of the SW). Clearly, we obtain $\tilde{I}_o(j_i) = S$ and $\tilde{I}_R(j_i) = 0$ for i = 1, 2, ... It is easy to verify that the processes $\tilde{I}_o(t)$ and $\tilde{I}_R(t)$, $t \ge 0$ are semi-regenerative processes with regenerative points j_i , i = 1, 2, ... Let $C_i = j_i - j_{i-1}$, i = 1, 2, ... be the time between two consecutive (actual) refillings. A typical sample path of $\tilde{I}_o(t)$ and $\tilde{I}_R(t)$ is depicted in Figure 2.

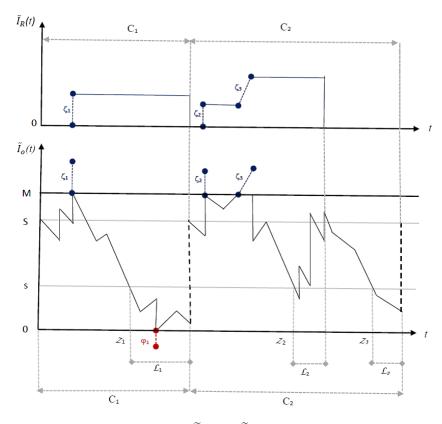


Figure 2. A typical sample path of $\tilde{I}_o(t)$ and $\tilde{I}_R(t)$.

The stock levels of PW and SW, $\tilde{I}_o(t)$ and $\tilde{I}_R(t)$, are depicted in the bottom and top parts of Figure 2, respectively. We start with the first cycle. At time Z_1 , we have

 $I_0(\mathcal{Z}_1) < s$, and thus a delivery is ordered. After lead time \mathcal{L}_1 , when the distributor arrives, $I_0(K_1 = Z_1 + L_1) < S$, and thus the refilling is carried out, the stock is raised up to level *S*, and the cycle ends. At that time, the SW is emptied by the distributor. Here, $J_1 = K_1$. By contrast, during the second cycle, a delivery is ordered twice. At the first time, Z_2 , the distributor arrives after lead time \mathcal{L}_2 when $I_o(K_2 = \mathcal{Z}_2 + \mathcal{L}_2) > S$, and thus no items are supplied and only maintenance service is performed, leading to $I_R(K_2) = 0$. The second order is placed at time \mathcal{Z}_3 , when $I_o(\mathcal{Z}_3) < s$. Here, we see that $I_o(K_3 = \mathcal{Z}_3 + \mathcal{L}_3) < S$, and thus the PW is refilled, the SW is emptied, and the cycle ends. Here, $I_2 = K_3$ and we obtain $I_0(j_2) = S$ and $I_R(j_2) = 0$. Figure 2 also shows that level *M* is exceeded three times: once during cycle 1 with ζ_1 items (due to a return by a retailer), and twice during cycle 2 with ζ_2 (due to return by a retailer) and with ζ_3 items (due to a return by a customer); these overstocking events are represented by the blue segments. The excess items are accumulated in the SW and emptied at times K_1 and K_2 (we note that maintenance services are also provided at time $K_3 = Z_3 + L_3$; however, since $I_o(t) < M$ and $I_R(t) = 0$ for $t \in (K_2, K_3]$, no change in I_R is depicted in the figure). In addition, we see one shortage event, with amounts of φ_1 (due to a demand by a retailer), represented by the red segment.

Our aim is to find the optimal parameters in order to optimize the expected discounted total cost. To that end, we let $\mathbb{E}_x(e^{-\beta C})$ be the $\left(\left|\widetilde{E}\right| \times \left|\widetilde{E}\right|\right)$ conditional Laplace transform matrix (LST) of the cycle length whose (i, j)-th component is:

$$\left(\mathbb{E}_{x}(e^{-\beta C})\right)_{i,j} = E(e^{-\beta C}\mathbf{1}_{\{\widetilde{J}(C)=j\}} \mid \widetilde{I}_{o}(0) = x, \widetilde{J}(0) = i), i, j \in \widetilde{E}.$$

Applying renewal theory, we can express all the cost functionals by using the first cycle.

(a) *Order cost* $OC(\beta)$. At ordering time, a fixed cost Y^+ is paid. This cost is non-refundable (even when no refilling is performed due to returns). Note that there may be multiple orders per cycle, each costing Y^+ . Thus, we obtain

$$OC(\beta) = \Upsilon^{+} \sum_{n} E_{S}(e^{-\beta Z_{n}}) = \Upsilon^{+} \eta \Big(\mathbb{I} - \mathbb{E}_{S}(e^{-\beta C}) \Big)^{-1} OC_{S}(\beta),$$
(4)

where $OC_S(\beta)$ is the $(\left|\widetilde{E}\right| \times 1)$ vector, representing the total expected discounted ordering times per cycle, given $\widetilde{I}_o(0) = S, \widetilde{J}(t) \in \widetilde{E}$.

(b) *Distributor's service cost*. At times j_n , n = 1, 2, ..., the PW is refilled with $S - I_o(j_n -)$ items and the cycle ends; this amount is random depending on the stock level upon arrival. Let γ be the cost of each refilled item (and is charged when the distributor arrives). When no refilling is performed, the distributor is rewarded with Y^- for his maintenance services. Summarizing, we obtain

$$DC(\beta) = \gamma \sum_{n} E_{S} \left(e^{-\beta I_{n}} (S - \widetilde{I}_{o}(j_{n})) \right) + Y^{-} \sum_{n} E_{S} \left(e^{-\beta K_{n}} \mathbf{1}_{\{\widetilde{I}_{o}(K_{n}) > S\}} \right)$$
$$= \gamma \cdot S \sum_{n} E_{S} \left(e^{-\beta C_{n}} \right) + \sum_{n} E_{S} (e^{-\beta K_{n}}) \left(Y^{-} \mathbf{1}_{\{\widetilde{I}_{o}(K_{n}) > S\}} - \gamma \widetilde{I}_{o}(K_{n}) \mathbf{1}_{\{\widetilde{I}_{o}(K_{n}) < S\}} \right)$$
$$= \eta \left[\mathbb{I} - \mathbb{E}_{S} (e^{-\beta C}) \right]^{-1} DC_{S}(\beta), \tag{5}$$

where $DC_S(\beta)$ is the expected discounted cost per cycle incurred by the distributor. Note that, while Υ^- may be charged several times per cycle (due to maintenance services), γ is charged only once, at the end of the cycle, when the PW is refilled.

(c) *Transfer cost* $SC(\beta)$. When the on-hand stock in the PW exceeds level M, the excess amount is transferred to the SW. Let v be the cost of transferring each item. Here, level M is exceeded either continuously by an individual customer at a linear rate $c_{\tilde{J}(t)}, \tilde{J}(t) \in E_p$, or by an upward jump due to a demand by a retailer. In the latter case, denote by Ξ_n the *n*-th time that level M is exceeded by an upward jump, and let ζ_n be the amount exceeding

level *M* and transferred to SW at time Ξ_n . The overall discounted transfer cost is given by (see also Remark 1, Point 1):

$$SC(\beta) = v \left(\sum_{n=1}^{\infty} E(e^{-\beta \Xi_n} \zeta_n) + \int_0^{\infty} E(e^{-\beta t} c_{\widetilde{J}(t)} \mathbf{1}_{\{\widetilde{I}_o(t) = M, \widetilde{J}(t) \in E_p\}}) dt \right)$$
$$= \eta v \left(\mathbb{I} - \mathbb{E}(e^{-\beta C}) \right)^{-1} SC_S(\beta), \tag{6}$$

where the $(|\tilde{E}| \times 1)$ vector $SC_S(\beta)$ is the expected discounted amount exceeding level M per cycle. Note that the first (second) term of (6) refers to a jump (continuous linear rate).

(d) Loss cost $LC(\beta)$. Any demand or part of a demand that is not satisfied is lost. Similarly, we distinguish between customers and retailers. Level 0 is hit either continuously by an individual customer at a linear rate $-c_{\tilde{J}(t)}, \tilde{J}(t) \in E_n$, or by a retailer by a downward jump. Let Φ_n be the *n*-th time that level 0 is hit by a downward jump, and let φ_n be the number of unsatisfied demands at time Φ_n . Let ϕ be the cost of each unsatisfied demand. The expected discounted loss cost is given by (see also Remark 1, Points 1 and 2):

$$LC(\beta) = \phi \left(\sum_{n=1}^{\infty} E(e^{-\beta\Phi_n} \varphi_n) - \int_0^{\infty} E(e^{-\beta t} c_{\widetilde{J}(t)} \mathbf{1}_{\{\widetilde{I}_O(t)=0,\widetilde{J}(t)\in E_n\}}) dt \right)$$
$$= \eta \phi \left(\mathbb{I} - \mathbb{E}_S(e^{-\beta C}) \right)^{-1} LC_S(\beta), \tag{7}$$

where the $(|\tilde{E}| \times 1)$ vector $LC_S(\beta)$ is the expected discounted unsatisfied demands per cycle.

Remark 1.

- 1. Regarding the individual customers. We emphasize that, when the on-hand stock level M is exceeding in an ascending state $i \in E_p$, demands are continuously satisfied by some of the returns (and the remaining returns are transferred to the SW at rate $c_i = r_i d_i > 0$). Similarly, when the on-hand stock level 0 is hit in a descending state $i \in E_n$, some of the demands are continuously satisfied by returns (and the remaining unsatisfied demands are lost at rate $-c_i > 0$).
- 2. As in practice, here we assume that new arrivals (both customers and retailers) are willing to accept that only some of their demands are satisfied as long as they receive some compensation for the unsatisfied demands. Nowadays, such a policy is highly valued and and even considered by the customer as a high level of service.

The expected total cost of integrated inventory management is therefore

$$TC(\beta) = OC(\beta) + DC(\beta) + SC(\beta) + LC(\beta).$$

In what follows, we denote by $(\mathbb{A})_{ij}$ the elements of a matrix \mathbb{A} , by $\mathbb{A}_{i,j}$ the submatrix of \mathbb{A} with row indices in E_i and column indices in E_j ; in particular, the matrix $\mathbf{0}_{i,j}$ represents the $(|E_i| \times |E_j|)$ zero matrix. Vectors are denoted by bold letters (e.g., E, P), and matrices by blackboard letters (e.g., \mathbb{E}, \mathbb{P}). The $(|E| \times |E|)$ conditional matrix expectation, given an initial state $(x, i \in E)$, is denoted by $\mathbb{E}_x(.) = \mathbb{E}(. | I(0) = x)$ (and, similarly, the $(|E| \times 1)$ conditional vector expectation $E_x(.)$). Specifically, we use the notations $\mathbb{E}_x(.)$ and $\mathbb{E}_x(.)^T$ to denote the submatrices of $\mathbb{E}_x(.)$, of order $(|E_p| \times |E|)$ and $(|E_n| \times |E|)$, respectively, associated with the ascending and the descending states of $\mathbb{E}_x(.)$, i.e., $\mathbb{E}_x(.) = (\mathbb{E}_x(.), \mathbb{E}_x(.))^T$. Following convention, we let $1_{\{A\}}$ be the indicator of an event $A, e = (1, \ldots, 1)^T$ be the unit column vector, and (A B) be the matrix obtained by stringing the matrix B after the matrix A, all of the appropriate size.

4. Preliminaries

4.1. Markov-Modulated Fluid Flow Process for MAP

Following Breuer [27], let $\tilde{\mathcal{J}} = \{\tilde{J}(t), t \geq 0\}$ be a CTMC with state space \tilde{E} and infinitesimal generator $\tilde{\mathbb{Q}} = (\tilde{q}_{ij})_{i,j\in\tilde{E}}$. Let $\hat{\mathcal{I}} = \{\hat{I}(t), t \geq 0\}$ be a real-valued process evolving like a Lévy process $\hat{\mathcal{I}}^{(i)}$ with a drift c_i (negative or positive) and a Lévy measure \hat{v}_i (given by (1)) when $i \in \tilde{E}$; here, we assume a compound Poisson jump. The jump sizes $U^{(i\pm)}$ are PH-distributed r.v.s $PH(\boldsymbol{\alpha}^{(i\pm)}, \mathbb{T}^{(i\pm)})$ of order m_i^{\mp} and depend on the phases $i \in \tilde{E}$ prior to the jumps. The two-dimensional process $(\hat{\mathcal{I}}, \tilde{\mathcal{J}})$ is called a MAP. The main advantage of PH-distributed jumps is the possibility of lengthening the jumps into a succession of linear periods of exponential duration (having slopes 1 and -1 for the positive and negative jumps, respectively) and retrieving the original process via a simple time change. The transformation is conducted as follows. First, we partition the phase space \tilde{E} (without the jumps) into subspaces E_p (for positive drifts, i.e., $c_i > 0$) and E_n (for negative drifts, i.e., $c_i < 0$), $\tilde{E} = E_p \cup E_n$. Then, we introduce two new phase spaces:

$$E_{\pm} := \{ (i, k, \pm) : i \in \widetilde{E}, 1 \le k \le m_i^{\pm} \},$$
(8)

to model the (positive and negative) jumps. Next, we define the enlarged state space $E \equiv \tilde{E} \cup E_- \cup E_+$. The revised phase process \mathcal{J} with state space E is determined by the generator matrix $\mathbb{Q} = (q_{ij})_{i,j \in E}$

$$q_{ij} = \begin{cases} \widetilde{q}_{ii} - \lambda_i, & j = i \in \widetilde{E}, \\ \widetilde{q}_{ij}, & j \neq i, h \in \widetilde{E}, \\ \lambda_i \alpha_k^{(i\pm)}, & j = (i, k, \pm), \end{cases}$$
(9)

for $i \in \widetilde{E}$ as well as

$$q_{(i,k,\pm),(i,l,\pm)} = T_{kl}^{(i\pm)} \text{ and } q_{(i,k,\pm),i} = \eta_k^{(i\pm)}$$
 (10)

for $i \in \tilde{E}$ and $1 \le k, l \le m_i^{\pm}$. By that, \mathcal{J} tracks the states (phases) of the PH-type jumps in addition to the states in $\tilde{\mathcal{J}}$. The revised process constructed by such a technique is denoted by $\{\mathcal{I}, \mathcal{J}\}$ and becomes a Markov-modulated fluid flow process (MMFF) without jumps. Accordingly, the infinitesimal generator \mathbb{Q} can be written in block form with respect to the subspaces $E_1 = E_p \cup E_+$ (ascending states) and $E_2 = E_n \cup E_-$ (descending states):

$$\mathbb{Q} = \left(\begin{array}{cc} \mathbb{Q}_{11} & \mathbb{Q}_{12} \\ \mathbb{Q}_{21} & \mathbb{Q}_{22} \end{array}\right)$$

Specifically, we modify the MAP $\{\tilde{\mathcal{I}}_{o}, \tilde{\mathcal{J}}\}\)$, with $\tilde{\mathcal{I}}_{o}$ being the stock level at the PW, to the revised MMFF process $\{\mathcal{I}_{o}, \mathcal{J}\}\)$. Figure 3 illustrates the first cycle of $\tilde{\mathcal{I}}_{o}$, presented in Figure 2, and its corresponding sample path \mathcal{I}_{o} ; the jumps that are lengthened to linear segments are marked by the dotted black segments. Note that downward jumps below level 0 and upward jumps above level *M* are preserved in their original form and are not lengthened to linear segments (as we will show, these quantities require a special treatment). Next, we present the main tools used for the analysis of the MMFF process $\{\mathcal{I}_{o}, \mathcal{J}\}\)$.

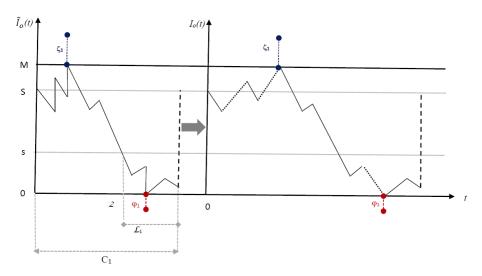


Figure 3. The MMFF process \mathcal{I}_o for $\widetilde{\mathcal{I}}_o$.

4.2. First-Passage Times

In our analysis, the Laplace transform matrices of the first passage times play an essential component. Our first step is to obtain the LST matrix $\mathbb{A}(\beta)$ of the return time for the MAP $\{\tilde{\mathcal{I}}_o, \tilde{\mathcal{J}}\}$ process. Using $\mathbb{A}(\beta)$ as a key matrix, we then obtain the conditional LST matrices of the first passage times for the MMFF process.

4.2.1. Return Times for MAPs

Let $\{\tilde{I}(t), \tilde{J}(t), t \ge 0\}$ be a MAP process with phase-type jumps. Define $\tilde{\tau}(x) = \inf\{t \ge 0, \tilde{I}(t) > x\}$ for $x \ge 0$ to be the LST of the first passage time above x. Assume that $\tilde{I}(0) = 0$ and let

$$E_{ij}(e^{-\tilde{\tau}(x)}) = E(e^{-\tilde{\tau}(x)}, J(\tau(x)) = j \mid \tilde{I}(0) = 0, \tilde{J}(0) = i), \ i, j \in E.$$
(11)

Breuer [27] shows that the matrix $\mathbb{E}(e^{-\tilde{\tau}(x)})$ can be written in a block form according to E_1 and E_2 as

$$\mathbb{E}(e^{-\widetilde{\tau}(x)}) = \begin{pmatrix} e^{\mathbb{U}(\beta)x} & \mathbf{0} \\ \mathbb{A}(\beta)e^{\mathbb{U}(\beta)x} & \mathbf{0} \end{pmatrix},$$
(12)

for some matrices $\mathbb{U}(\beta)$ and $\mathbb{A}(\beta)$ of dimensions $|E_1| \times |E_1|$ and $|E_2| \times |E_1|$, respectively (note that, since a first passage to a level above cannot occur at a descending state, we have the zero matrices). Furthermore, the matrices $\mathbb{U}(\beta)$ and $\mathbb{A}(\beta)$ can be determined by successive approximation to the limits of the sequence $((A_n, U_n) : n \ge 0)$ with specific initial values (see Section 2.2 of Breuer [27] for the detailed algorithm). Setting x = 0in (11) and (12), i.e., the return time to level 0 from below given $\tilde{I}(0) = 0$, yields that $\mathbb{E}_{11}(e^{-\tilde{\tau}(0)}) = \mathbb{I}$ and $\mathbb{E}_{21}(e^{-\tilde{\tau}(0)}) = \mathbb{A}(\beta)$.

An important variant of the MAP process \hat{I} , called the *image-reversed* process, is particularly useful in the analysis of our process. We define the*image-reversed MAP process* \tilde{I}^r by reversing the roles of the up and down states; let $\tilde{\mathcal{J}}^r(t)$ be the modulated state process for \tilde{I}^r . Then, $\mathbb{A}^r(\beta)$ is the matrix of order $(|E_1| \times |E_2|)$, whose (i, j)-th component is the LST of $\tilde{\tau}^r(0)$, i.e., the time until a downward jump to level 0 for the process \tilde{I}^r at state $j \in E_2$, given that $\tilde{I}^r(0) = 0$ and $\tilde{\mathcal{J}}^r(0) = i \in E_1$. As we show, these two LSTs, $\mathbb{A}(\beta)$ and $\mathbb{A}^r(\beta)$, are the key matrices used in our analysis; we further emphasize that these matrices are the return time for the original MAP process, meaning that the time spent in jump states E_{\pm} does not account.

Next, to extend our LSTs and include upper and lower borders, we employ results from MMFF processes.

4.2.2. First Passage Times for MMFF Processes

Let ${}_{u}^{v}\tau(x,y)$ (for $x \ge 0, y \ge 0$) be the first passage time of \mathcal{I} from level x to level y, with avoiding a visit to the levels in $[0, u] \cup [v, \infty)$ enroute (in the case of an unlimited visit, u and v are omitted). The notation ${}_{u}^{v}\widehat{f}(x,y,)$ denotes the LST matrix of the joint distribution of the first passage time ${}_{u}^{v}\tau(x,y)$ and the state of the phase process at that time, ${}_{u}^{v}\widehat{f}^{r}(x,y,)$ denotes the LST of the hitting time for \mathcal{I}^{r} . Ramaswami [49] shows that, once $\mathbb{A}(\beta)$ and $\mathbb{A}^{r}(\beta)$ are computed, the derivation of other LST matrices is straightforward; thus, all quantities of interest in this paper are characterized explicitly in terms of $\mathbb{A}(\beta)$ and $\mathbb{A}^{r}(\beta)$; the explicit expressions of the LST matrices and their probabilistic interpretations are summarized in [49]. To account for the jump-like nature of the additional movements, we substitute $\beta = 0$ in all entries of the form (i,h), (h,i), or (h,h') for $i \in \widetilde{E}$ and $h, h' \in E_+ \cup E_-$. This means that the time is not discounted for all states associated with the jumps (i.e., at states E_+ and E_-). Henceforth, we assume that all LST matrices for hitting times ($\mathbb{A}(\beta), \mathbb{A}^{r}(\beta)$, ${}_{u}^{v}\widehat{f}(x,y,)$) are modified for the original process { $\widetilde{\mathcal{I}_{o}}, \widetilde{\mathcal{J}}$.

4.3. MMFF at an Exponential Time

Corollary 1. Let $\{\mathcal{I}, \mathcal{J}\}$ be an MMFF with state space E and \mathbb{Q} , and let \mathcal{L} be an exponential *r.v.* with rate $\mu > 0$ independently of $\{\mathcal{I}, \mathcal{J}\}$.

(1) The LST matrix at time \mathcal{L} with the (ij)-th component

$$\left(\mathbb{E}(e^{-\beta\mathcal{L}})\right)_{ij} = E\left(e^{-\mathcal{L}}, J(\mathcal{L}) = j \mid J(0) = i\right), \ i \in E$$

is given by

$$\mathbb{E}(e^{-\beta\mathcal{L}}) = \mu((\beta + \mu)\mathbb{I} - \mathbb{Q})^{-1}.$$
(13)

In our analysis, the matrix form of (13) is required in order to track the initial and final states of the background CTMC. Here, we use the fact that, for a CTMC \mathcal{J} , the matrix \mathbb{Q} satisfies $(e^{\mathbb{Q}x})_{ij} = P(J(x) = j | J(0) = i)$ (see [2]). (Recall that the original state space is \tilde{E} ; thus, only rows in \tilde{E} are considered while other rows in (13) are set to zero.)

(2) Applying the complete probability equation, it is easy to verify that, for a nonnegative generally distributed r.v. X,

(i)
$$\mathbb{E}(e^{-\beta X} \mathbf{1}_{\{X < \mathcal{L}\}}) = \mathbb{E}(e^{-(\beta + \mu)X}).$$
 (14)
(ii) $\mathbb{E}(e^{\alpha I(\mathcal{L}) - \beta L} \mathbf{1}_{\{\mathcal{L} < X\}}) = \mu \mathbb{E}\left(\int_{t=0}^{X} e^{\alpha I(t) - (\beta + \mu)t} dt\right).$

Here, Equations (13) and (14) are used with respect to the lead time \mathcal{L} .

4.4. The Multidimensional Martingale

Let $\{\mathcal{X}(t), t \ge 0\}$ be a right-continuous Markov-modulated Lévy process with a modulating process $\{\mathcal{J}(t), t \ge 0\}$ that is an irreducible right-continuous Markov chain with a finite state space *E*. Let $\{Y(t), t \ge 0\}$ be an adapted continuous process with a finite expected variation on finite intervals and let $\mathcal{Z}(t) = \mathcal{X}(t) + Y(t)$. Assussen and Kella [50] have shown that the matrix with elements $E_i\left[e^{\alpha \mathcal{X}(t)}; J(t) = j\right]$ has the form of $e^{t\mathbb{K}(\alpha)}$ for some matrix $\mathbb{K}(\alpha)$. Theorem 2.1 of Assussen and Kella [50] yields that, under certain mild conditions on $\{\mathcal{Z}(t), t \ge 0\}$, the multidimensional process

$$\boldsymbol{M}(\alpha,t) = \int_{u=0}^{t} e^{\alpha \mathcal{Z}(u)} \mathbf{1}_{J(u)} du \mathbb{K}(\alpha) + e^{\alpha \mathcal{Z}(0)} \mathbf{1}_{J(0)} - e^{\alpha \mathcal{Z}(t)} \mathbf{1}_{J(t)} + \alpha \int_{u=0}^{t} e^{\alpha \mathcal{Z}(u)} \mathbf{1}_{J(u)} dY(u)$$
(15)

is a (row) vector-valued zero mean martingale. Here, the indicator $\mathbf{1}_{J(t)}$ is an |E|-row vector mean $\{\mathbf{1}_{J(t)=i}, i \in E\}$. In our model, the process $I_o(t)$ has piecewise linear sample paths (at rate c_i for $i \in \tilde{E}$, rate 1 for $i \in E_+$, and rate -1 for $i \in E_-$). Some of the cost functionals used

are obtained by applying the optional stopping theorem (OST) to the multidimensional martingale process (15).

5. Derivation of the Expected Discounted Cost Functionals

We need to derive $E_S(e^{-\beta C})$, $OC_S(\beta)$, $DC_S(\beta)$, $SC_S(\beta)$, and $LC_S(\beta)$. The derivation of $DC_S(\beta)$ is more challenging and will be carried out at the end. To this end, we first derive two unique LSTs associated with hitting the boundaries, i.e., hitting level *M* from below, and dropping to level 0 from above.

Corollary 2.

(i) Assume that level M is hit from below at ascending state $J(t) \in E_1$. It is easy to verify that the $(|E_1| \times |E_2|)$ matrix $(\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12}$ is the expected discounting time until exiting level M at a descending state (either continuously at a state in E_n or by a downward jump at a state in E_-). As we set $\beta = 0$ in all entries in $E_+ \in E_1$ due to the jumps, the matrix $(\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12}$ has the block form

$$(\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12} = \begin{pmatrix} (\beta \mathbb{I} - \widetilde{\mathbb{Q}}_{11})^{-1} \widetilde{\mathbb{Q}}_{12} & (\beta \mathbb{I} - \widetilde{\mathbb{Q}}_{11}) \boldsymbol{\lambda}^{(-)} \boldsymbol{\alpha}^{(-)} \\ \mathbb{I} \otimes \boldsymbol{e} & \boldsymbol{0} \end{pmatrix}$$

with respect to the subspaces $(E_p E_+) \times (E_n E_-)$. The matrix $\mathbb{I} \otimes e$ arises due to the zero time spent in E_+ , and thereafter $I_o(t)$ immediately returns to its original state in E_n (recall that an upward jump is allowed only when $J(t) \in E_n$).

(ii) Similarly, assume that level 0 is hit from above at a descending state $J(t) \in E_2$. It is easy to verify that the $(|E_2| \times |E_1|)$ matrix $(\beta \mathbb{I} - \mathbb{Q}_{22})^{-1} \mathbb{Q}_{21}$ is the expected discounting time until exiting level 0 at ascending state E_n (either continuously at a state in E_p or by an upward jump at a state in E_+) Here, too, by setting $\beta = 0$ in all entries in $E_- \in E_2$, the matrix $(\beta \mathbb{I} - \mathbb{Q}_{22})^{-1} \mathbb{Q}_{21}$ has the block form

$$(\beta \mathbb{I} - \mathbb{Q}_{22})^{-1} \mathbb{Q}_{21} = \begin{pmatrix} (\beta \mathbb{I} - \widetilde{\mathbb{Q}}_{22})^{-1} \widetilde{\mathbb{Q}}_{21} & (\beta \mathbb{I} - \widetilde{\mathbb{Q}}_{22}) \lambda^{(+)} \boldsymbol{\alpha}^{(+)} \\ \mathbb{I} \otimes \boldsymbol{e} & \boldsymbol{0} \end{pmatrix}$$

with respect to the subspaces $(E_n E_-) \times (E_p E_+)$.

5.1. Cycle Length and Order Cost

The cycle length is composed of two sequential periods: the time elapsed until level *s* is hit when a refilling is ordered, and the remaining cycle time. The next proposition summarizes the derivation of $E_S(e^{-\beta C})$ (recall that the superscript *o* indicates the stock-level evolution during lead time).

Proposition 1. The LST $(|E| \times 1)$ vector $\mathbf{E}_{S}(e^{-\beta C}) = (\overline{\mathbf{E}}_{S}(e^{-\beta C}) \underline{\mathbf{E}}_{S}(e^{-\beta C}))^{T}$ satisfies the following set of equations:

(remaining cycle length with no upcoming refill)

(i)
$$\overline{\mathbf{E}}_{S}(e^{-\beta C}) = {}^{M-S} \mathbb{A}^{r}(\beta) \mathbb{E}_{S}(e^{-\beta C}) + {}_{S} \hat{f}_{11}(S, M, \beta) \overline{\mathbf{E}}_{M}(e^{-\beta C}),$$

(ii) $\underline{\mathbf{E}}_{S}(e^{-\beta C}) = {}^{S} \hat{f}_{22}(S, s, \beta) \underline{\mathbf{E}}_{s}^{o}(e^{-\beta C}) + {}^{S-s} \mathbb{A}(\beta) \overline{\mathbf{E}}_{S}(e^{-\beta C}),$
(iii) $\overline{\mathbf{E}}_{M}(e^{-\beta C}) = (\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12} \underline{\mathbf{E}}_{M}(e^{-\beta C}),$
(iv) $\underline{\mathbf{E}}_{M}(e^{-\beta C}) = {}^{M} \hat{f}_{22}(M, S, \beta) \underline{\mathbf{E}}_{S}(e^{-\beta C}) + {}^{M-S} \mathbb{A}(\beta) \overline{\mathbf{E}}_{M}(e^{-\beta C}),$ (16)

(remaining cycle length during lead time)

$$\begin{aligned} (v) \quad E_{s}^{0}(e^{-\beta C}) &= 0 \hat{f}_{21}(s, S, \beta + \mu) \overline{E}_{S}^{0}(e^{-\beta C}) + S \hat{f}_{22}(s, 0, \beta + \mu) E_{0}^{0}(e^{-\beta C}) \\ &+ \frac{\mu}{\beta + \mu} \Big[e - \Big(0 \hat{f}_{21}(s, S, \beta + \mu) e^{+S} \hat{f}_{22}(s, 0, \beta + \mu) e \Big) \Big], \\ (vi) \quad \overline{E}_{S}^{0}(e^{-\beta C}) &= {}^{M-S} \mathbb{A}^{r}(\beta + \mu) \underline{E}_{S}^{0}(e^{-\beta C}) + S \hat{f}_{11}(S, M, \beta + \mu) \overline{E}_{M}^{0}(e^{-\beta C}) \\ &+ \Big({}^{M-S} \mathbb{A}^{r}(\beta) - {}^{M-S} \mathbb{A}^{r}(\beta + \mu) \Big) \overline{E}_{S}(e^{-\beta C}) \\ &+ \Big(s \hat{f}_{11}(S, M, \beta) - s \hat{f}_{11}(S, M, \beta + \mu) \Big) \overline{E}_{M}(e^{-\beta C}), \\ (vii) \quad \underline{E}_{S}^{0}(e^{-\beta C}) &= {}^{S} \hat{f}_{22}(S, s, \beta + \mu) \underline{E}_{S}^{0}(e^{-\beta C}) + {}^{S-s} \mathbb{A}(\beta + \mu) \overline{E}_{S}^{0}(e^{-\beta C}) \\ &+ \frac{\mu}{\beta + \mu} \Big[e - \Big({}^{S-s} \mathbb{A}(\beta + \mu) e^{+S} \hat{f}_{22}(S, s, \beta + \mu) e \Big) \Big], \\ (viii) \quad \underline{E}_{0}^{0}(e^{-\beta C}) &= ((\beta + \not \gg) \mathbb{I} - \mathbb{Q}_{22})^{-1} \mathbb{Q}_{21} \overline{E}_{0}^{0}(e^{-\beta C}) \\ &+ \frac{\mu}{\beta + \mu} \Big[e - ((\beta + \mu) \mathbb{I} - \mathbb{Q}_{22})^{-1} \mathbb{Q}_{21} e \Big], \\ (ix) \quad \overline{E}_{0}^{0}(e^{-\beta C}) &= 0 \hat{f}_{11}(0, S, \beta + \mu) \overline{E}_{S}^{0}(e^{-\beta C}) \\ &+ \frac{\mu}{\beta + \mu} \Big[e - (0 \hat{f}_{11}(0, S, \beta + \mu) e + \Big({}^{S} \mathbb{A}^{r}(\beta + \mu) E_{0}^{0}(e^{-\beta C}) \\ &+ \frac{\mu}{\beta + \mu} \Big[e - \Big((\beta + \mu) \mathbb{I} - \mathbb{Q}_{11} \Big)^{-1} \mathbb{Q}_{12} \underline{E}_{M}^{0}(e^{-\beta C}) \\ &+ \Big((\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12} - ((\beta + \mu) \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12} \Big) \underline{E}_{M}(e^{-\beta C}), \\ (xi) \quad \overline{E}_{M}^{0}(e^{-\beta C}) &= {}^{M-S} \mathbb{A}(\beta + \mu) \overline{E}_{M}^{0}(e^{-\beta C}) \\ &+ \Big({}^{M-S} \mathbb{A}(\beta) - {}^{M-S} \mathbb{A}(\beta + \mu) \Big) \overline{E}_{M}(e^{-\beta C}) \\ &+ \Big({}^{M-S} \mathbb{A}(\beta) - {}^{M-S} \mathbb{A}(\beta + \mu) \Big) \overline{E}_{M}(e^{-\beta C}) \\ &+ \Big({}^{M-S} \mathbb{A}(\beta) - {}^{M-S} \mathbb{A}(\beta + \mu) \Big) \overline{E}_{M}(e^{-\beta C}) \\ &+ \Big({}^{M-S} \mathbb{A}(\beta) - {}^{M-S} \mathbb{A}(\beta + \mu) \Big) \overline{E}_{N}(e^{-\beta C}). \end{aligned}$$

$$(17)$$

Proof. The proof of Proposition 1 (and all subsequent proofs) is based on the first-step conditioning method and is obtained in Appendix B.1. \Box

Corollary 3. The total number of expected discounted ordering times per cycle $OC_S(\beta) = (\overline{OC}_S(\beta) \underline{OC}_S(\beta))^T$ satisfy the set of Equations (16) and (17) with only three differences:

- (a) The vector **OC** replaces the vector $\mathbf{E}(e^{-\beta C})$ (all of the appropriate sizes and with the same superscripts and subscripts).
- (b) The first term of Case (ii) becomes ${}^{S}\hat{f}_{22}(S,s,\beta)(e+\underline{OC}_{s}^{o})$ (instead of ${}^{S}\hat{f}_{22}(S,s,\beta) \underline{E}_{s}^{o}(e^{-\beta C})$).
- (c) The last terms in Cases (v), (vii), (viii), and (ix) are deleted.

We emphasize that calling the distributor during lead time is not allowed. Thus, the order cost is charged only when dropping to *s* with no upcoming order explains Corollary 3(b). Accordingly, when $\mathcal{L} < T_S$, the distributor refills the PW at no additional order cost, and the cycle ends, which explains Corollary 3(c).

5.2. Transfer Cost

The derivation of $SC_S(\beta)$ includes two steps. First, Claim 1 below derives the expected discounted number of transferred items from hitting level M until exiting it for the first time. From that time, Proposition 2 derives the remaining number of transferred items until the end of the cycle. To see this, let $\zeta = \{\zeta_l, l \in E_1\}$ be the $(|E_1| \times 1)$ vector whose *l*-th component represents the expected discounted number of transferred items from hitting level M at state $l \in E_1$ until exiting M at a descending state.

Claim 1. The vector $\boldsymbol{\zeta} = (\zeta_l, l \in E_1)$ satisfies the following system of linear equations:

$$\zeta_{l} = \begin{cases} -\boldsymbol{e}_{r}^{(i+)}(\mathbb{T}^{(i+)})^{-1}\boldsymbol{e} & l = (i,r,+) \in E_{+}, \\ \frac{c_{l}}{\beta - q_{ll}} + \sum_{k \neq l} \frac{q_{lk}}{\beta - q_{ll}} \zeta_{k} & l,k \in E_{p}. \end{cases}$$
(18)

Proof. Clearly, level *M* can be exceeded only at an ascending state in E_1 . We distinguish between two situations. (i) Exceeding level *M* by an upward jump at state l = (i, r, +), $i \in E_n, 1 \le r \le m_i^+$, i.e., by a batch amount $U^{(i+)}$ at phase r (here, the timing is omitted due to the jump pattern). Thus, the exceeded amount is $U_r^{(i+)} \sim PH(e_r^{(i+)}, \mathbb{T}^{(i+)})$ of order m_i^+ , where $e_r^{(i+)}$ is a $(1 \times m_i^+)$ vector equal to 1 at the *r*-th entry and to 0 otherwise. From that point, level *M* is left at a descending state *i*. (ii) Hitting level *M* continuously at a linear rate $c_l, l \in E_p$. First, the items are transferred to the SW at rate c_l at a random time $\delta_l \sim \exp(-q_{ll})$ and, thus, the expected discounted amount is $E\left(\int_{t=0}^{\delta_l} c_l e^{-\beta t} dt\right)$. Then, with

probability $\frac{q_{lk}}{-q_{ll}}$, $k \neq l \in E_p$, the state changes to k, and the stock level \mathcal{I}_R continues to accrue at rate c_k ; otherwise (i.e., with probability $1 - \sum_{k \neq l \in E_p} \frac{q_{lk}}{-q_{ll}}$), level *M* is left at a descending state. Summarizing, we can express the vector ζ as

$$\zeta_{l} = \begin{cases} E(U_{r}^{(i+)}) & l = (i,r,+) \in E_{+}, \\ E\left(\int_{t=0}^{\delta_{l}} c_{l}e^{-\beta t}dt\right) + E_{l}(e^{-\beta\delta_{l}}) \sum_{k \neq l} \frac{q_{lk}}{-q_{ll}} \zeta_{k} & l,k \in E_{p}. \end{cases}$$
(19)

Substituting $E(U_r^{(i+)}) = -e_r^{(i+)}(\mathbb{T}^{(i+)})^{-1}e$, $E\left(\int_{t=0}^{\delta_l} c_l e^{-\beta t} dt\right) = \frac{c_l}{\beta} \left(1 + \frac{q_{ll}}{\beta - q_{ll}}\right) = \frac{c_l}{\beta - q_{ll}}$, and $E_l(e^{-\beta\delta_l}) = \frac{-q_{ll}}{\beta - q_{ll}}$ into (19) completes the proof. \Box

Proposition 2. The expected discounted number of transferred items to the SW during a cycle $SC_{S}(\beta) = (\overline{SC}_{S}(\beta) SC_{S}(\beta))^{T}$ satisfies the following set of equations: (transferred items with no upcoming order)

(i)
$$\overline{SC}_{S}(\beta) = {}_{S}\widehat{f}_{11}(S, M, \beta)\overline{SC}_{M}(\beta) + {}^{M-S}\mathbb{A}^{r}(\beta) \underline{SC}_{S}(\beta),$$

(ii) $\underline{SC}_{M}(\beta) = {}^{M-S}\mathbb{A}(\beta)\overline{SC}_{M}(\beta) + {}^{M}\widehat{f}_{22}(M, S, \beta) \underline{SC}_{S}(\beta),$
(iii) $\overline{SC}_{M}(\beta) = \zeta + (\beta \mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12} \underline{SC}_{M}(\beta),$
(iv) $SC_{S}(\beta) = {}^{S}\widehat{f}_{22}(S, s, \beta) SC_{s}^{o}(\beta) + {}^{S-s}\mathbb{A}(\beta) \overline{SC}_{S}(\beta).$
(20)

(transferred items during lead time)

$$(v) \ \underline{SC}_{s}^{o}(\beta) = {}^{s}\mathbb{A}(\beta+\mu) \ \overline{SC}_{s}^{o}(\beta) + {}^{s}\widehat{f}_{22}(s,0,\beta+\mu) \ \underline{SC}_{0}^{o}(\beta),$$

$$(vi) \ \overline{SC}_{s}^{o}(\beta) = {}_{s}\widehat{f}_{11}(s,S,\beta+\mu)\overline{SC}_{s}^{o}(\beta) + {}^{S-s}\mathbb{A}^{r}(\beta+\mu) \ \underline{SC}_{s}^{o}(\beta),$$

$$(vii) \ \underline{SC}_{0}^{o}(\beta) = ((\beta+\mu)\mathbb{I} - \mathbb{Q}_{22})^{-1}\mathbb{Q}_{21}\overline{SC}_{0}^{o}(\beta),$$

$$(viii) \ \overline{SC}_{0}^{o}(\beta) = {}^{s}\mathbb{A}^{r}(\beta+\mu) \ \underline{SC}_{0}^{o}(\beta) + {}_{0}\widehat{f}_{11}(0,s,\beta+\mu)\overline{SC}_{s}^{o}(\beta),$$

$$(ix) \ \underline{SC}_{s}^{o}(\beta) = {}^{s}\widehat{f}_{22}(S,s,\beta+\mu)\underline{SC}_{s}^{o}(\beta) + {}^{S-s}\mathbb{A}(\beta+\mu) \ \overline{SC}_{s}^{o}(\beta),$$

$$(x) \ \overline{SC}_{s}^{o}(\beta) = {}^{s}\widehat{f}_{11}(S,M,\beta+\mu)\overline{SC}_{M}^{o}(\beta) + {}^{M-S}\mathbb{A}^{r}(\beta+\mu) \ \underline{SC}_{s}^{o}(\beta)$$

$$+ \left({}^{s}\widehat{f}_{11}(S,M,\beta) - {}^{s}\widehat{f}_{11}(S,M,\beta+\mu)\right)\overline{SC}_{M}(\beta)$$

$$+ \left({}^{M-S}\mathbb{A}^{r}(\beta) - {}^{M-S}\mathbb{A}^{r}(\beta+\mu)\right)\underline{SC}_{s}(\beta),$$

$$(xi) \ \underline{SC}_{M}^{o}(\beta) = {}^{M-S}\mathbb{A}(\beta+\mu)\overline{SC}_{M}^{o}(\beta) + {}^{M}\widehat{f}_{22}(M,S,\beta+\mu) \ \underline{SC}_{s}^{o}(\beta)$$

$$+ \left({}^{M-S}\mathbb{A}(\beta) - {}^{M-S}\mathbb{A}(\beta+\mu)\right)\overline{SC}_{M}(\beta)$$

$$+ \left({}^{M}\widehat{f}_{22}(M,S,\beta) - {}^{M}\widehat{f}_{22}(M,S,\beta+\mu)\right) \ \underline{SC}_{s}(\beta),$$

$$(xii) \ \overline{SC}_{M}^{o}(\beta) = \zeta + \left[(\beta+\mu)\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12}\right] \ \underline{SC}_{M}(\beta)$$

$$+ \left[(\beta\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12} - ((\beta+\mu)\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12}\right] \ \underline{SC}_{M}(\beta).$$

$$(21)$$

Proof. The Proof is given in Appendix B.2. \Box

5.3. Loss Cost

Let $\Theta = (\Theta_l, l \in E_2)$ be the $(|E_2| \times 1)$ vector whose *l*-th component represents the expected discounted number of unsatisfied items when hitting level 0 at state $l \in E_2$ until exiting 0 for the first time.

Claim 2. The vector $\boldsymbol{\Theta} = (\Theta_l, l \in E_2)$ satisfies the following system of linear equations:

$$\Theta_{l} = \begin{cases} -e_{r}^{(i-)} (\mathbb{T}^{(i-)})^{-1} e & l = (i, r, -) \in E_{-} \\ \frac{-c_{l}}{\beta + \mu - q_{ll}} + \sum_{k \neq l} \frac{q_{lk}}{\beta + \mu - q_{ll}} \Theta_{k} & l, k \in E_{n}. \end{cases}$$
(22)

Proof. Similar to Claim 1 with one exception, here, demand may be lost only during lead time. Thus, if level 0 is hit at a state $l \in E_n$, lost demand is accumulated at rate $-c_l$, as long as the distributor has not arrived or until leaving state l, i.e., the random time min (δ_l, L) with exp $(\not = q_{ll})$ distribution. A similar technique that used for Claim 1 completes the proof. \Box

Proposition 3. The expected discounted number of unsatisfied items during a cycle $LC_S(\beta) = (\overline{LC}_S(\beta) \ \underline{LC}_S(\beta))^T$ satisfies the set of Equations (20) and (21) with the exception of three modifications:

- (*a*) The vector *LC* replaces the vector *SC* (all of the appropriate sizes and with the same superscripts and subscripts).
- (b) The first term of Cases (iii) and (xii), ζ , is deleted.
- (c) The expected discounted loss Θ is added to Case (vii).

5.4. Distributor's Service Cost

Recall that K_n , n = 1, 2... is the *n*-th arrival time of the distributor. Clearly, when $I_0(K_n) < S$, the stock is refilled at a cost γ for each refilled item. In addition, maintenance activities are conducted at no additional charge, and the cycle ends; in that case, $C = K_n$.

Otherwise, only maintenance activities are conducted at cost Y^- . Let N_C be the number of times that the distributor arrives during a cycle. Thus,

$$N_C = \min\{n : I_0(K_n) < S\}.$$

The distributor's service cost is therefore given by

$$DC_{S}(\beta) = \gamma E_{S} \left(e^{-\beta C} (S - I_{o}(K_{N_{C}})) + Y^{-} \sum_{n=1}^{N_{C}-1} E_{S}(e^{-\beta K_{n}}) \right)$$
$$= \gamma \left(S E_{S} \left(e^{-\beta C} \right) - E_{S}(e^{-\beta C} I_{o}(C)) + Y^{-} \sum_{n=1}^{N_{C}-1} E_{S}(e^{-\beta K_{n}}). \right)$$
(23)

The first term is the charge for to refilling the PW; the last term is the charge for maintaining the SW when no refilling is made. Note that the term $E_S(e^{-\beta C})$ is given by Equations (16) and (17). Let

$$\Delta_{x}(\beta) = \Upsilon^{-} \sum_{n=1}^{N_{C}-1} E_{x}(e^{-\beta K_{n}}) - \gamma E_{x}(e^{-\beta C}I_{o}(C)),$$

and let T_0^- be the time to exit level 0 (recall that T_0 is the time to enter level 0). The next proposition derives $\Delta_S(\beta)$ and the sequence of vectors to be used.

Proposition 4. The cost functional $\Delta_S(\beta) = (\overline{\Delta}_S(\beta) \underline{\Delta}_S(\beta))$ satisfies the following set of equations: (cost with no upcoming order)

$$\begin{array}{ll} (i) & \overline{\Delta}_{S}(\beta) = {}^{M-S} \mathbb{A}^{r}(\beta) \underline{\Delta}_{S}(\beta) + {}_{S}\widehat{f}_{11}(S,M,\beta)\overline{\Delta}_{M}(\beta), \\ (ii) & \underline{\Delta}_{S}(\beta) = {}^{M}\widehat{f}_{22}(S,s,\beta) \underline{\Delta}_{s}^{o}(\beta) + {}^{S-s} \mathbb{A}(\beta)\overline{\Delta}_{S}(\beta), \\ (iii) & \overline{\Delta}_{M}(\beta) = (\beta \mathbb{I} - \mathbb{Q}_{11})^{-1} \mathbb{Q}_{12} \underline{\Delta}_{M}(\beta), \\ (iv) & \underline{\Delta}_{M}(\beta) = {}^{M-S} \mathbb{A}(\beta)\overline{\Delta}_{M}(\beta) + {}^{M}\widehat{f}_{22}(M,S,\beta) \underline{\Delta}_{S}(\beta). \end{array}$$

(cost during lead time)

$$\begin{aligned} (v) \quad \underline{\Delta}_{s}^{o}(\beta) &= \widehat{0f_{21}}(s, S, \beta + \mu)\overline{\Delta}_{s}^{o}(\beta) + {}^{S}\widehat{f_{22}}(s, 0, \beta + \mu)\underline{\Delta}_{0}^{o}(\beta) \\ &- \gamma \underline{E}_{s}(e^{-\beta \mathcal{L}}I_{o}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}), \\ (vi) \quad \overline{\Delta}_{S}^{0}(\beta) &= {}^{M-S}\mathbb{A}^{r}(\beta + \mu)\underline{\Delta}_{S}^{o}(\beta) + s\,\widehat{f}_{11}(S, M, \beta + \mu)\Big)\Big(Y^{-}e + \overline{\Delta}_{M}(\beta)\Big) \\ &+ \Big(s\widehat{f}_{11}(S, M, \beta) - s\,\widehat{f}_{11}(S, M, \beta + \mu)\Big)\Big(Y^{-}e + \underline{\Delta}_{S}(\beta)\Big), \\ (vii) \quad \overline{\Delta}_{M}^{o}(\beta) &= \Big[(\beta + \mu)\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12}\Big]\underline{\Delta}_{M}^{o}(\beta) \\ &+ \Big[(\beta\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12} - ((\beta + \mu)\mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12}\Big]\Big(Y^{-}e + \underline{\Delta}_{M}(\beta)\Big), \\ (viii) \quad \underline{\Delta}_{M}^{o}(\beta) &= {}^{M-S}\mathbb{A}(\beta + \mu)\overline{\Delta}_{M}^{o}(\beta) + {}^{M}\widehat{f}_{22}(M, S, \beta + \mu)\underline{\Delta}_{S}^{o}(\beta) \\ &+ \Big(M^{-S}\mathbb{A}(\beta) - {}^{M-S}\mathbb{A}(\beta + \mu)\Big)\Big(Y^{-}e + \overline{\Delta}_{M}(\beta)\Big) \\ &+ \Big[{}^{M}\widehat{f}_{22}(M, S, \beta) - {}^{M}\widehat{f}_{22}(M, S, \beta + \mu)\Big]\Big(Y^{-}e + \underline{\Delta}_{S}(\beta)\Big), \\ (ix) \quad \underline{\Delta}_{S}^{o}(\beta) &= {}^{S}\mathbb{A}(\beta + \mu)\overline{\Delta}_{S}^{o}(\beta) + {}^{S}\widehat{f}_{22}(S, 0, \beta + \mu)\underline{\Delta}_{0}^{o}(\beta) \\ &- \gamma \underline{E}_{S}(e^{-\beta \mathcal{L}}I_{0}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}), \\ (xi) \quad \overline{\Delta}_{0}^{o}(\beta) &= ((\beta + \mu)\mathbb{I} - \mathbb{Q}_{22})^{-1}\mathbb{Q}_{21}\overline{\Delta}_{0}^{o}(\beta)), \\ (xi) \quad \overline{\Delta}_{0}^{o}(\beta) &= 0\widehat{f}_{11}(0, S, \beta + \mu)\overline{\Delta}_{S}^{o}(\beta) + {}^{S}\mathbb{A}^{r}(\beta + \mu)\underline{\Delta}_{0}^{o}(\beta) \\ &- \gamma \overline{E}_{0}(e^{-\beta \mathcal{L}}I_{0}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}). \end{aligned}$$

(We emphasize that $I_o(\mathcal{L})$ is the on-hand stock level at the PW just before the distributor arrives.)

Proof. We introduce only the key steps of the proof. Cases (i)-(iv) are straightforward. Cases (v), (ix), and (xi) are based on the decomposition according to $\min(T_S, T_0, \mathcal{L})$. Similarly, Cases (vi)-(viii) use similar techniques regarding the cost Y⁻. Regarding Case (*x*), note that $I_0(\mathcal{L})\mathbf{1}_{\{\mathcal{L} < T_0^-\}} = 0$ and, thus, only the event $\{\mathcal{L} > T_0^-\}$ (with LST $((\beta + \mu)\mathbb{I} - \mathbb{Q}_{22})^{-1}\mathbb{Q}_{21})$ should be considered. In order to complete the derivation, we need to obtain the last terms of Cases (v), (ix), and (xi). We concentrate on Case (v); the derivations of Cases (ix) and (xi) are similar, and are given in Corollary 4(a) and 4(b), respectively. \Box

The $(|E_2| \times 1)$ vector $\underline{E}_s(e^{-\beta \mathcal{L}} I_o(\mathcal{L}) \mathbf{1}_{\{\mathcal{L} < \min(T_S, T_0)\}})$. We shift the original time so that $I_o(0) = s$, $J(0) \in E_2$. Recall that $\varsigma = \min(T_S, T_0)$; we start by introducing the vector $\underline{E}_s(e^{\alpha I_o(\mathcal{L}) - \beta \mathcal{L}} \mathbf{1}_{\{\mathcal{L} < \varsigma\}})$. It is easy to see that

$$\underline{\mathbf{E}}_{s}(e^{-\beta\mathcal{L}}I_{o}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\varsigma\}} = \frac{d}{d\alpha} \left(\underline{\mathbb{E}}_{s}(e^{\alpha I_{o}(\mathcal{L})-\beta\mathcal{L}}\mathbf{1}_{\{\mathcal{L}<\varsigma\}})e\right)\Big|_{\alpha=0}.$$
(25)

Applying (14)(ii) yields

$$\underline{\mathbb{E}}_{s}(e^{\alpha I_{o}(\mathcal{L})-\beta\mathcal{L}}\mathbf{1}_{\{\mathcal{L}<\varsigma\}}) = \mu \underline{\mathbb{E}}_{s}\left(\int_{0}^{\varsigma} e^{\alpha I_{o}(t)-(\beta+\mu)t}\mathbf{1}_{J(t)}dt\right).$$
(26)

The basic tool to compute $\underline{\mathbb{E}}_{s}\left(\int_{0}^{\varsigma} e^{\alpha I_{o}(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt\right)$ is the OST to the Asmussen–Kella multidimensional martingale defined in Section 4.4. Consider the process $(\mathcal{X}(t))_{t \geq 0}$:

$$\mathcal{X}(t) = \mathcal{X}(0) + \int_{u=0}^{t} c_{J(u)} du \quad t \ge 0,$$
 (27)

with $\mathcal{X}(0) = s, J(0) \in E_2$. It is not difficult to see that, by conditioning on $I_0(0) = s$, $J(0) \in E_2$, the process up to time ς , i.e., $(\mathcal{X}(t))_{0 \le t < \varsigma}$, has the same distribution as $(I_0(t))_{0 \le t < \varsigma}$. Let \mathbb{C} be the rate matrix $\mathbb{C} = diag(\mathbb{C}_1, \mathbb{C}_2)$ with $\mathbb{C}_1 = diag(c_i : i \in E_1)$ and $\mathbb{C}_2 = diag(c_i : i \in E_2)$. It can be concluded from Chapter XI, p.311 of Asmussen [2] that $\mathbb{E}(e^{\alpha \mathcal{X}(t)} \mathbf{1}_{\{J(t)\}}) = e^{t\mathbb{K}(\alpha)}$, where $\mathbb{K}(\alpha) = \mathbb{Q} + \alpha \mathbb{C}$.

1 -

Claim 3. The
$$(|E_2| \times |E|)$$
 matrix $\underline{\mathbb{E}}_s \left(\int_0^{\varsigma} e^{\alpha I_o(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt \right)$ is given by

$$\underline{\mathbb{E}}_s \left(\int_0^{\varsigma} e^{\alpha I_o(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt \right) = \left[e^{\alpha S} \left(_0 \widehat{f}_{21}(s, S, \beta + \mu) \mathbf{0} \right) + \left(\mathbf{0}^{-S} \widehat{f}_{22}(s, 0, \beta + \mu) \right) - e^{\alpha S} (\mathbf{0} \ \mathbb{I}) \right] (\mathbb{K}(\alpha) - (\beta + \mu)\mathbb{I})^{-1}.$$
(28)

Proof. The proof is given in Appendix B.3. \Box

Corollary 4.

(a) The $(|E_2| \times 1)$ vector $\underline{E}_S(e^{-\beta \mathcal{L}}I(\mathcal{L})1_{\{\mathcal{L} < \varsigma\}})$ (Case (ix)) is given by

$$\underline{E}_{S}(e^{-\beta\mathcal{L}}I_{o}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\varsigma\}} = \frac{d}{d\alpha} \left[\mu \underline{\mathbb{E}}_{S} \left(\int_{0}^{\varsigma} e^{\alpha I_{o}(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt \right) e \right] \Big|_{\alpha=0},$$

$$\underline{\mathbb{E}}_{S} \left(\int_{0}^{\varsigma} e^{\alpha I(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt \right) = \left[e^{\alpha S} \left({}^{S} \mathbb{A}(\beta + \mu) \mathbf{0} \right) + \left(\mathbf{0}^{S} \widehat{f}_{22}(S, 0, \beta + \mu) \right) - e^{\alpha S}(\mathbf{0} \mathbb{I}) \right] \times (\mathbb{K}(\alpha) - (\beta + \mu)\mathbb{I})^{-1}.$$
(29)

(Applying (28) with S replacing s; note that $_{0}\widehat{f}_{21}(S, S, \beta + \mu) = {}^{S}\mathbb{A}(\beta + \mu)$.) (b) The $(|E_{1}| \times 1)$ vector $\overline{E}_{0}(e^{-\beta \mathcal{L}}I(\mathcal{L})1_{\{\mathcal{L} \leq \varsigma\}})$ (Case (xi)) is given by

$$\overline{E}_{0}(e^{-\beta\mathcal{L}}I_{o}(\mathcal{L})\mathbf{1}_{\{\mathcal{L}<\varsigma\}} = \frac{d}{d\alpha} \left[\mu \overline{\mathbb{E}}_{0} \left(\int_{0}^{\varsigma} e^{\alpha I_{o}(t) - (\beta+\mu)t} \mathbf{1}_{J(t)} dt \right) e \right] \Big|_{\alpha=0},$$

$$\overline{\mathbb{E}}_{0} \left(\int_{0}^{\varsigma} e^{\alpha I(t) - (\beta+\mu)t} \mathbf{1}_{J(t)} dt \right) = \left[e^{\alpha S} \left(_{0}\widehat{f}_{11}(0, S, \beta+\mu) \mathbf{0} \right) + \left(\mathbf{0} \ ^{S} \mathbb{A}^{r}(\beta+\mu) \right) - (\mathbb{I} \mathbf{0}) \right] \times (\mathbb{K}(\alpha) - (\beta+\mu)\mathbb{I})^{-1}.$$
(30)

(Here, we apply (28) with $(0, E_1)$ replacing (s, E_2) .)

A summary of the costs' derivation is given in Appendix A.2.

6. Numerical Examples

In this section, we illustrate the impact of the thresholds, the parameters, and the uncertainty of demands and returns on the system's performance. Clearly, our aim is to find the optimal (s^*, S^*, M^*) for different values of the parameters. The sets of equations in Propositions 1–4 show that it is difficult to provide an explicit expression for the discounted expected cost functions and, thus, it is difficult to obtain explicit terms for the optimal controllers. Hence, a numerical investigation is applied. Our base case assumes a CTMC with n = 2 states: $\tilde{E}_p = \{1\}, \tilde{E}_n = \{2\}, \tilde{E} = \tilde{E}_p \cup \tilde{E}_n = \{1, 2\}$, i.e., one state with positive net rate and one state with negative net rate. The infinitesimal generator and the stationary probability vector are given by

$$\widetilde{\mathbb{Q}} = \begin{pmatrix} -0.03 & 0.03\\ 0.05 & -0.05 \end{pmatrix}, \ \pi = (0.625, 0.325), \tag{31}$$

where, for simplicity, we let the initial probability vector $\eta = \pi$. In states 1 and 2, demand occurs at rate $d_1 = 1$ and $d_2 = 2$, respectively, and return occurs at rate $r_1 = 1.5$ and $r_2 = 0.5$, respectively. The net growth is thus $c_1 = 0.5$ and $c_2 = -1.5$. Furthermore, in state 1, a downward jump may occur at rate λ_1 with a demand size $U^{(1-)} \sim PH(\alpha^{(1-)}, \mathbb{T}^{(1-)})$ of order $m_1^- = 2$; in state 2, an upward jump may occur at rate λ_2 with a return size $U^{(2+)} \sim PH(\alpha^{(2+)}, \mathbb{T}^{(2+)})$ of order $m_2^+ = 1$. We denote by ξ^+ (ξ^-) the average growth rate of states in E_1 (E_2); thus, $E(\xi^+) = \pi_1 c_1 + \pi_2 \lambda_2 E(U^{(2+)})$, and $E(\xi^-) = \pi_2 |c_2| + \pi_1 \lambda_1 E(U^{(1-)})$. Finally, let $\xi = \xi^+ - \xi^-$ be the net average rate (which can be positive or negative). We start with

$$\lambda_{1} = 0.2, \qquad \boldsymbol{\alpha}^{(1-)} = (0.3, 0.7), \qquad \mathbb{T}^{(1-)} = \begin{pmatrix} -0.25 & 0\\ 0 & -0.5 \end{pmatrix}, \qquad (32)$$

$$\lambda_{2} = 0.161, \qquad \boldsymbol{\alpha}^{(2+)} = (1), \qquad \mathbb{T}^{(2+)} = (-0.2).$$

(i.e., $U^{(2+)}$ is an exponentially r.v.). Here, $E(\xi^+) = E(\xi^-) = 0.614$ and $E(\xi) = 0$ (i.e., the system is considered to be *balanced*). Assume a discounted factor $\beta = 0.075$, an exponential lead time $\mathcal{L} \sim \exp(\mu = 0.1)$, and cost values

$$Y^+ = 50, Y^- = 50, \gamma = 10, \phi = 50, \nu = 0.5.$$

We start by studying the impact of the parameters *S* and *s*; then, we focus on the optimal policy and cost.

Example 1. The impact of the parameters *S* and *s* on the system's performance. It is practically a commonplace that warehouses have limited and known capacity; thus, and without loss of generality, we assumed that M = 35 and study the changes of *S* and *s*. Table 1 summarizes the effect of changing *S* and *s* on the expected discounted cycle length and cost components; each line presents the effect of increasing only one parameter while keeping the other fixed. We use " \uparrow " and " \downarrow " to express increasing and decreasing functions of the parameter and " \cup " to express a convex shape. The star " \bigstar " stands for a special behavior to be discussed further on.

Table 1. The impact of the parameters on the system's performance.

	Expected Discounted							
Increasing the Parameter	Cycle Length EC	Order Cost OC	Transfer Cost SC	Loss Cost LC	Distributor's Service Cost DC			
$S\uparrow$	U	\downarrow	\uparrow	\downarrow	U			
$s\uparrow$	\uparrow	¢	*	\downarrow	1			

- *The impact of S*. As expected, increasing *S* decreases the number of orders and shortages, and increases the number of transferred items; thus, *OC* and *LC* decrease and *SC* increases. However, the changes in *EC* and *DC* are more challenging and are explained as follows. At first, increasing *S* increases the cycle length (i.e., *EC* decreases), minimizes value for money (due to β), and thus decreases *DC*. However, a highly increasing *S* increases the number of refillings, shortens the cycle length, intensifies the impact of β, and thus increases *DC*.
- The impact of *s*. Clearly, increasing *s* shortens the cycle length-resulting in early ordering times, and intensifies the impact of β ; thus, the changes in *EC*, *OC*, *LC*, and *DC* are as expected. By contrast, the change in *SC* is not straightforward. Figure 4a curves *SC* as a function of *s*, for $S = \{20, 25, 30, 34\}$ (recall that s < S and M = 35). We see that *SC* increases and then decreases, and these changes become more dramatic with *S*. At first, increasing *s* increases the stock, level *M* is hit more frequently and, thus, *SC* increases. However, as *s* approaches *S*, the ordering time and the cycle length are significantly shortened and thus *SC* decreases.

To summarize, while the impact of *S* and *s* on *OC* and *LC* is as expected, their impact on *SC* and *DC* is not straightforward and difficult to predict. To illustrate the overall impact, Figure 4b displays the expected discounted total cost per time unit TC(M = 35, S, s) as a function of *S* and *s*. We see that TC(M = 35, S, s) is jointly convex in *S* and *s*; we note that this appears in all our examples (other values of *M*, *S*, and *s*,, which are not reported here, also exhibit a similar behavior). Thus, we assume that the cost is convex in *S*, *s* for a fixed *M*. Assuming convexity enables us to numerically obtain the optimal thresholds *S*^{*} and *s*^{*} quicker by using a line search.

It should be noted that convexity of the total cost for base-stock policy is a common supposition, and it has been proven by several researchers for both periodic and continuous review contexts, however, for simpler models and under restricted assumption [51]. Although we cannot give a formal proof of this, we give the following intuitive explanation. Table 1 shows that the cost components are influenced in a different fashion with *S* and *s*. For example, increasing *S* reduces the number of orders and decreases *OC*; it also reduces the penalty cost, *LC*. However, at the same time, increasing *S* increases *SC* and the cycle length, and minimizes the impact of β . This trade-off between potential savings is due to fewer orders and timing on the one side, and more transferred items from the other side yields convexity.

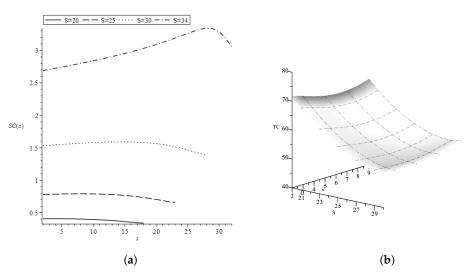


Figure 4. (a) SC(s) for $S \in \{20, 25, 30, 34\}$, M = 35. (b) TC(M = 35, S, s) as a function of *S* and *s*.

Example 2. Sensitivity analysis of the optimal thresholds and cost. Here, we investigate the impact of the different parameters on the optimal thresholds and cost; we start with a balanced system. Following Example 1, we let $E(\xi) = 0$ (the specific values are given in (31) and (32)), $\beta = 0.075$, $\mu = 0.1$, $Y^+ = 50$, and M = 35. We vary γ in $\{5, 10\}$, $Y^- \in \{50, 150, \}$, $\phi \in \{5, 25, 50, 100\}$, and $\nu \in \{0.5, 5, 15, 25, 100\}$. Table 2 summarizes the optimal (S^*, s^*) and the corresponding total cost $TC^* = TC(M = 35, S^*, s^*)$.

Table 2 shows that increasing the lost $cost \phi$ increases S^* and s^* (to avoid shortage); increasing γ decreases s^* , probably to delaying the order time and, consequently, the refilling time. Additionally, S^* is increasing in Y⁻ when ν and ϕ are low in order to reduce the number of cancellations. We further see that, usually, increasing ν increases s^* and thus shortens the until ordering. However, Table 2 shows few exceptions particularly when Y⁻ is high (Y⁻ = 150); here, the distributor's service cost cost is very high and, thus, both S^* and s^* are kept low to reduce the number of cancellations. Additional numerical results (that are not reported here) further show that increasing Y⁺ increases S^* and decreases s^* ; thereby, $S^* - s^*$ increases and the number of orders decreases. Table 3 summarizes the changes in S^* and s^* when the costs are increasing; here, too, the star " \bigstar " stands for non-monotonic behavior.

S*, s* TC*	$oldsymbol{\phi}\setminus u$	0.5	5	15	25	100
$j^{-} = 50$ $\gamma = 5$	5	20, 2 27.133	24, 2 29.457	19, 2 38.748	17, 2 44.792	11, 2 69.150
	25	30, 2 30.153	26, 2 40.540	21, 4 53.402	19, 4 61.845	13, 4 96.876
	50	30, 7 37.363	26, 8 48.063	23, 8 62.386	21, 8 72.243	15, 8 114.098
	100	30, 11 44.696	27, 11 55.907	23, 8 62.386	22, 11 83.298	17, 11 133.019
$i^{-} = 50$ $\gamma = 10$	5	20, 2 38.039	24, 2 39.811	19, 2 49.818	17, 2 56.186	11, 2 81.156
	25	30, 2 40.785	26, 2 50.760	22, 2 64.027	20, 2 72.961	13, 2 108.709
	50	30, 4 51.595	27, 5 62.358	23, 6 77.256	21, 6 87.626	15, 7 130.449
	100	30, 9 62.397	27,9 73.001	24, 10 89.145	22, 10 100.755	17, 10 151.083
; ⁻ = 150	5	29, 2 21.328	24, 2 29.542	19, 2 39.043	17, 2 45.280	11, 2 71.492
	25	31, 2 30.170	26, 2 40.592	21, 4 53.706	19, 4 62.347	14, 2 98.084
$\gamma = 5$	50	30, 7 37.430	26, 8 48.315	23, 8 62.921	21, 8 73.131	16, 7 117.311
-	100	30, 10 44.878	27, 11 56.341	24, 11 72.911	22, 11 84.850	17, 10 137.497
	5	28, 2 1.728	24, 2 39.896	20, 2 50.062	17, 2 56.918	11, 2 83.498
; ⁻ = 150	25	30, 2 40.812	26, 2 50.811	22, 2 64.167	20, 2 73.190	13, 5 115.468
$\gamma = 10$	50	30, 4 51.640	27, 5 62.446	23, 6 77.569	21, 6 88.403	15, 6 133.090
-	100	30, 8 62.871	27, 9 73.255	24, 10 89.856	22, 10 101.937	17, 9 155.096

Table 2. $S^*, s^*, TC^*(M = 35, S^*, s^*)$ for $Y^- \in \{50, 150\}$, $\gamma \in \{5, 10\}$, $\phi \in \{5, 25, 50, 100\}$, and $\nu \in \{0.5, 5, 15, 25, 100\}$.

Table 3. The impact of increasing costs on S^* and s^* .

The Cost \uparrow							
The Effect on	Service Cost Y ⁻	Unit Cost γ	Loss Cost ϕ	Transfer Cost ν	Ordering Cost Y ⁺		
S^*	\uparrow	*	\uparrow	\downarrow	\uparrow		
<i>s</i> *	*	\downarrow	\uparrow	*	\downarrow		

Our next goal is to study the effect of inflows and outflows (average demand and return) on the optimal thresholds S^* and s^* . To do so, we vary $E(\xi)$ in $\{-0.45, -0.41, -0.308, -0.24, -0.12, 0, 0.22, 0.43, 0.625\}$ by changing c_1 and $T^{(2+)}$ and fixing $c_2 = -1.5, \lambda_1 = 0.2$ and $\lambda_2 = 0.161$ (the specific pairs $(c_1, E(U^{(2+)}))$ corresponding to these values are $(0.025, 2.5), (0.03, 3), (0.5, 1), (0.2, 3), (0.5, 3), (0.5, 5), (0.85, 5), (1, 7), and (1.5, 5), respectively). Recall that, when <math>E(\xi) < 0$, demands arrive more frequently on average, and when $E(\xi) > 0$, returns arrive more frequently (note that the results for $E(\xi) = 0$ are given in

Table 2). We set $Y^+ = 50$, $Y^- = 150$, and $\nu = 5$, and focus on two scenarios: (1) $\gamma = 10$ and $\phi = 5$, and (2) $\gamma = 5$ and $\phi = 50$. Scenario 1 highlights the high cost of refilling; Scenario 2 highlights the high penalty cost of a shortage. Regarding Scenario 1, Figure 5 plots the pairs (S^*, s^*) (black and blue points, respectively) corresponding to $E(\xi)$, and Figure 6 shows the percentage of the cost components % OC, % LC, % SC, and % DC of TC^* . The blue-black scale depicts for cases where $E(\xi) \ge 0$, and the red-orange scale depicts cases where $E(\xi) < 0$; when $E(\xi) \le 0$, the transfer cost percentage % SC becomes negligible and thus is omitted. Similarly, Figures 7 and 8 plot (S^*, s^*) and the cost percentages associated with Scenario 2, respectively. (Here, although % SC is low, it is reported).

Figures 5–8 show that:

The levels s^* and S^* . When the loss cost is low (i.e., Scenario 1), the level s^* is kept low; when the loss cost is high (i.e., Scenario 2), level s^* is significantly higher and is decreasing in $E(\xi)$. Consequently, it interesting to see that, despite the high loss cost, when $E(\xi) > 0$, s^* is set low, probably due to the more frequent returns that reduce the number of shortages (see Figure 7). When $E(\xi) > 0$, Figures 5 and 7 show similar values of S^* ; we further see that S^* is increasing in $E(\xi)$. By contrast, when $E(\xi) < 0$, we see significantly higher values of S^* under Scenario 2. Here, when demands become more frequent, the impact of the high penalty cost dominates in increasing S^* . Nevertheless, the difference (S^*-s^*) is increasing in $E(\xi)$ under both scenarios, meaning that as demands become more frequent, the optimal policy is to call the distributor more frequently and for smaller quantities.

The cost components. Our results show that, as expected, when demands become more frequent ($E(\xi) < 0$) and S^* is low, the cost percentage of *SC* is negligible; this holds even when S^* is set higher due to a high penalty cost (Figure 8 showing that SC becomes a more significant but still meager component of the cost, less than 2%). Surprisingly, when $E(\xi) < 0$, the cost percentages remain generally similar in $E(\xi)$ under both scenarios. This can be explained by the relatively small changes in S^* . By contrast, when returns arrive more frequently, $E(\xi) > 0$, we see significant changes in the cost percentages as a function of $E(\xi)$. Figures 6 and 8 highlight the interplay between %DC and %LC. The percentage of the distributor's service cost %DC is significantly higher under Scenario 1 (which is directly impacted by γ); by contrast, under Scenario 2, the loss cost %*LC* forms the main component of the cost. Under both scenarios, when more returns arrive, SC (OC) becomes the more (less) dominant component of the cost. Moreover, comparing Figures 6 and 8 implies that, when $E(\xi) \ge 0$, we have $\% SC_{\{Scenario 1\}} \approx \% SC_{\{Scenario 2\}}$ and $OC_{\{Scenario 1\}} \approx OC_{\{Scenario 2\}}$. For example, we have $SC_{\{Scenario 1\}} = \{0.17, 0.30, 0.30, 0.30\}$ 0.64, 0.73}, and $SC_{Scenario 2} = \{0.18, 0.33, 0.65, 0.71\}$ corresponding to $\dot{E}(\xi) = \{0, 0.22, 0$ 0.43, 0.625}. When $E(\xi) < 0$, we have $DC_{\{Scenario 1\}} \approx DC_{\{Scenario 2\}}$; here, we obtain $DC_{\{Scenario 1\}} = \{0.37, 0.40, 0.41, 0.42\}, \text{ and } DC_{\{Scenario 2\}} = \{0.40, 0.40, 0.40, 0.38\}$ corresponding to $E(\xi) = \{-0.45, -0.41, -0.30, -0.12\}$. These observations emphasize the company's policy of balancing the cost components despite the different cost values. In general, when more returns arrive, and thus more items are transferred, the optimal policy aims to keep %SC relatively constant, whereas, when more demands arrive, and thus more orders are placed, the focus is to keep %DC relatively constant.

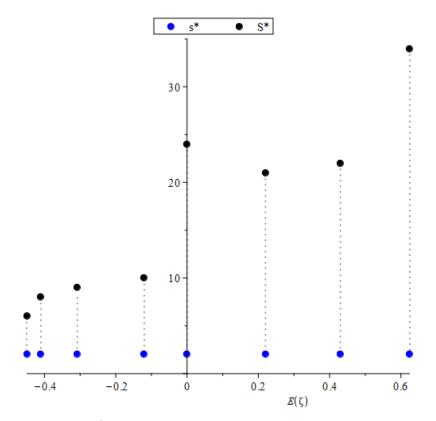


Figure 5. (S^*, s^*) corresponding to $E(\xi)$, Scenario 1.

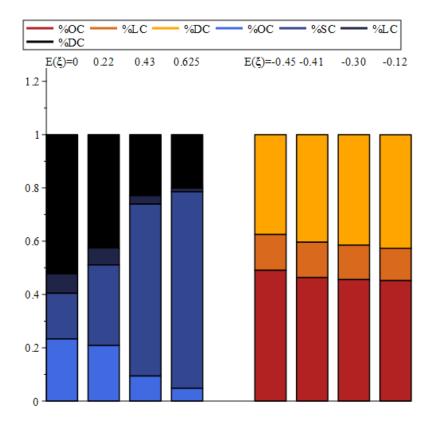


Figure 6. The %cost components, Scenario 1.

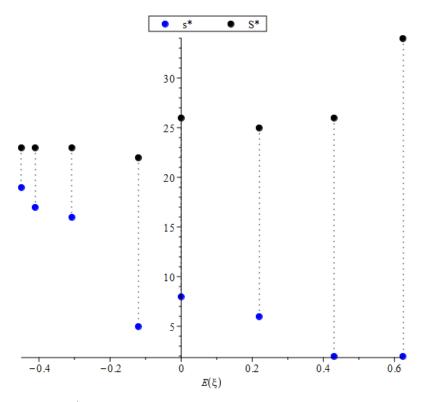


Figure 7. (s^*, S^*) corresponding to $E(\xi)$, Scenario 2.

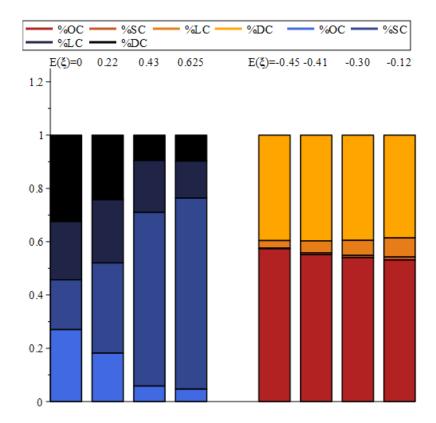


Figure 8. The %cost components, Scenario 2.

Example 3. The impact of β . We next concentrate on the impact of the discounted factor β on the optimal thresholds and cost. To do so, we let $Y^- = 50$, $\nu = 25$, $\phi = 25$, $\gamma = 5$, $\mu = 0.1$, $Y^+ \in \{15, 25, 50, 100\}$, and vary β in $\{0.025, 0.05, 0.075, 0.1\}$. Table 4 presents the optimal policy and cost; it includes five subtables that differ in their $E(\xi)$. The top subtable corresponds to

 $E(\xi) = 0$, the two left-hand subtables correspond to the negative $E(\xi) = \{-0.308, -0.41\}$, and the two right-hand subtables correspond to the positive $E(\xi) = \{0.22, 0.43\}$. Each subtable presents the levels (S^*, s^*) and the cost $TC(S^*, s^*, M = 35)$ for different values of β ; the entry $E(\xi) = 0$, $\beta = 0.075$, $Y^+ = 50$ is also given in Table 2. Figure 9a curves S^* and s^* as a function of β , for $E(\xi) = \{0, -0.308, -0.41\}$ (the black, blue, and gray curves, respectively); similarly, Figure 9b curves S^* and s^* for $E(\xi) = \{0, 0.22, 0.43\}$ (the black, blue, and gray curves, respectively). (Here, we take $Y^+ = 50$.) Note that both figures contain the case $E(\xi) = 0$, indicated by the black curves.

Table 4. (S^*, s^*) , $TC(M = 35, S^*, s^*)$ as a function of β , $E(\xi)$, and Y^+ .

	, s* C*			в							
Ε (ξ)=0	0.025	0.05	0.075	0.1						
	15	13, 8 241.121	16, 9 92.529	17, 8 48.974	19, 5 29.221						
<i>i</i> +	25	12,7 263.498	16, 8 102.091	18, 6 53.278	20, 4 31.282						
	50	12, 5 313.620	17, 7 122.909	19, 4 61.845	20, 2 34.708						
	100	13, 5 405.342	18, 2 155.221	21, 2 73.906	21, 1 40.932						
$E\left(\xi\right) =$	-0.308	0.025	0.05	0.075	0.1	$E\left(\xi\right)$ =	= 0.22	0.025	0.05	0.075	0.1
	15	14, 11 189.954	18, 16 98.580	20, 16 56.465	21, 14 35.679		15	15, 2 1193	13, 2 233.797	14, 5 83.106	15, 4 40.079
i ⁺	25	13, 10 217.648	18, 15 117.955	21, 18 70.431	21, 13 42.855	.+	25	21, 2 1145	14, 2 240.001	14, 2 86.883	16, 1 41.923
1	50	12, 9 284.880	18, 15 165.584	20,14 97.471	22, 11 59.143	1 -	i ⁺ 50	21, 2 1188	15, 2 254.774	15, 1 94.634	17, 1 46.180
	100	12, 9 418.205	16,13 259.948	19, 12 153.258	23, 8 87.600	-	100	21, 2 1274	17, 2 281.976	17, 1 108.364	18, 1 53.676
$E(\xi) =$	-0.41	0.025	0.05	0.075	0.1	Ε (ξ) =	= 0.43	0.025	0.05	0.075	0.1
	15	16, 15 190.886	18, 16 104.637	20, 16 60.542	20, 14 38.376		15	34, 2 959.741	25, 2 417.673	15, 1 172.134	13, 1 76.027
i ⁺	25	16, 15 223.120	18, 16 125.17	19, 15 73.465	21, 14 46.47	;+ -	25	34, 2 963.969	25, 2 419.423	15, 1 174.239	13, 1 77.896
1	50	14, 13 302.48	17, 15 176.372	19, 14 105.381	21, 11 64.951	1 -	50	34, 2 974.540	25, 1 423.796	16, 1 179.223	14, 1 82.263
	100	12, 11 454.858	16, 14 277.973	19, 13 167.416	23, 9 97.662	-	100	34, 2 995.681	26, 1 432.543	17, 1 197.933	15, 1 89.658

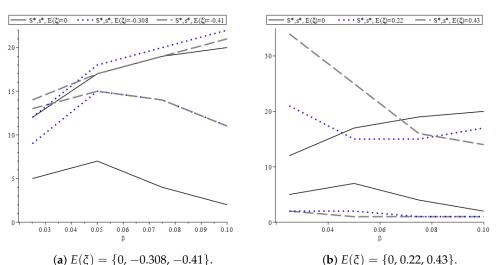


Figure 9. The levels S^* and s^* as a function of β and $E(\xi)$.

From Table 4 and Figure 9, we derive the following conclusions:

- As expected, the cost TC^* is decreasing in β and is increasing in Y^+ . We further see that, while increasing Y^+ decreases s^* (in accordance with Table 3), the changes of S^* in Y^+ are not monotone and are impacted also by β and $E(\xi)$ (in fact, S^* varies slightly in Y^+).
- When $E(\xi) \leq 0$, the level S^* is increasing in β , while s^* seems to be concave. Hence, increasing β increases the difference S^* - s^* (noted that a balanced system with $E(\xi) = 0$ is also included here; see Table 3). When the value of β is high, then, despite the fact that demands become more frequent, it is economically profitable for the company to postpone ordering and thereby increase the quantity of refilling.
- By contrast, when $E(\xi) > 0$, we see that S^* is convex in β , and s^* is decreasing. Here, the difference S^*-s^* is convex in β . In general, as more returns arrive, and for low values of β , the company's main motivation is to reduce the transfer cost by decreasing S^* , s^* and reducing S^*-s^* . However, as β increases, we see a slight increase in S^* (see Figure 9b, particularly the dotted blue curve when $\beta \ge 0.08$). The main implication is to postpone the distributor in order to achieve a full utilization of the discount factor.

Summarizing, we see that the discount factor, and the interplay between returns and demands, have a crucial impact on determining the optimal policy and obtaining a decisive economic advantage.

7. Conclusions

This paper studies an inventory control problem with two types of storage facilities that differ in purpose and capacity. We consider continuous and batch-type bilateral changes in the inventory level, where both types of changes are governed by the underlying Markov chain. Applying first passage time results, we build a relatively easy set of equations to derive the cost components under the discounted criterion and lost sales assumption. In doing so, we provide managers with a useful and efficient tool to derive, numerically, the optimal parameters. We show that the limited capacity yields lower thresholds, and as demands become more frequent, it is worthwhile to order more frequently and in smaller quantities. We further show that the timing has a significant impact on the optimal policy; for high values of β , it is worth considering postponing the distributor even at the risk of causing more shortage events.

For future research, it will be interesting to consider the case where each refilling order proceeds sequentially for a few stages, each of which has an independent exponential distributed time. Thus, the total time it would take to handle the order could be expressed by the phase-type distribution rather than by an exponential one. In particular, nowadays,

when collaborative inventories and outsourcing services are widespread, it is practical to assume that each order has to go through several procedures, e.g., due to miscellaneous certificates, multiple managers, or different hierarchical ranks. Our approach in this study appears to be simple and easy to implement; thus, we believe that it can be applied as a powerful tool to address other inventory policies. For example, upon adding the option of backordering up to some fixed level, thereafter the unsatisfied demand will be lost. Another well-known policy is the (Q, r) policy, where a fixed order of size Q is placed whenever the on-hand stock level drops to or below level r. Finally, a combination of the above-mentioned policies can also be considered. Each of the discussed policy is worth studying and hence is left for future investigation.

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Appendix A

Appendix A.1. Selected Studies on Continuous Review Base Stock Inventory Policy

Table A1 presents the most relevant studies on continuous review base stock inventory policy and random demands and returns.

Continuous-Review Base-Stock Inventory Models		Model Features						
Authors	Arrival Pattern of Demand/Return	Movement Patterns (Conti., Jumps)	Lead Time	Base-Stock Policy	Additional Features			
Benkherouf, Bensoussan (2009) [6]	Com. Poiss. + diffusion/no returns	One-sided jumps	zero	(q_i, u_i) - band policy	• Disc. criteria	 A quasi-variational inequalities approach 		
Yamazaki (2016, 2017) [8,29]	Spectrally positive Lévy processes	One-sided jumps	zero	(<i>S</i> , <i>s</i>)	• Disc. criteria	fluctuation theory of Lévy processesThe scale function		
Perera, Janakiraman, Niu (2018) [5]	Renewal proc./no returns	One-sided unit jumps	zero/ constant	(S, s) (B)	• Long run analysis	 Renewal theory Deterministic dynamic programming 		
Barron (2018) [18]	Fluid flow /MAP	One-sided jumps	zero	(Q, 0) (LS)	 Disc. criteria Obsolescence + emergency 	MartingalesNumerical investigation		
Azcue, Muler (2019) [43]	Comp. Poiss./Comp. Poiss.	Two sided jumps	zero	Multi-band structure	• Disc. criteria	 Solution of HJB equation 		
Pérez, Yamazaki, Bensoussan (2020) [7]	Spectrally positive Lévy processes	One-sided jumps	zero	Periodical replenishments	• Disc. criteria	The scale functionGuess and verify procedure		
Barron, Dreyfuss (2021) [4]	Comp. Poiss./Comp. Poiss.	Two-sided jumps. No conti. movements	exp	(S, s) (LS)	Steady-state analysis	Markov chainNumerical investigation		
Chakravarthy, Rao (2021) [42]	MAP/no returns	One-sided jumps	zero	Random replenishments (LS, B)	Steady-state analysis	 Matrix-analytic methods 		
Barron (2022) [13]	Fluid flow or jumps for demands/no returns	One-sided movements	exp	(S, s) + emergencies	Disc. CriteriaImmediate emergency	LSTsNumerical investigation		
This study	MAP/MAP	Bilateral linear rates. Two-sided-PH-Jumps	exp	(<i>M</i> , <i>S</i> , <i>s</i>) (LS)	 Disc. Criteria Joint management policy, two storages 	 Renewal set of equations LSTs Multidimensional Martingales Numerical investigation 		

Table A1. Selected studies on continuous review base stock inventory policy.	Table A1. Selected studies o	n continuous	review bas	e stock invent	ory policy.
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It can be observed that, although numerous studies deal with the (S, s)-type inventory models with uncertainties, no study discusses cancellations, double-sided barriers, bilateral movements and integrated control policy. As we pointed out in Section 2, several models used in the literature focus on the derivation of the LSTs of the first-passage time in a Markov-modulated process with bilateral random jumps. However, they do not apply any replenishment policy or cost components; thus, they are not included in Table A1.

Appendix A.2. A Summary of the Cost' Derivation

Here, we summarize the algorithm for the derivation of the cost components. The algorithm is composed of three main parts. The first part includes the derivation of the matrices $\mathbb{A}(\beta)$ and $\mathbb{A}^{R}(\beta)$. In the second part, we derive the cost components. The third part is devoted to the optimal policy. The algorithm is demonstrated by applying the values of Scenario 1 and $E(\xi) = 0$, i.e., $\widetilde{\mathbb{Q}} = \begin{pmatrix} -0.03 & 0.03 \\ 0.05 & -0.05 \end{pmatrix}$, $\eta = (0.625, 0.325)$, $c_1 = 0.5, T^{(1-)} = \begin{pmatrix} -0.25 & 0 \\ 0 & -0.5 \end{pmatrix}$, $c_2 = -1.5, T^{(2+)} = -1/5$, $\lambda_1 = 0.2$, $\lambda_2 = 0.161$, $\beta = 0.075$, $\mu = 0.1$, M = 35, and the costs Y⁺ = 50, Y⁻ = 150, $\nu = 5$, $\gamma = 10$ and $\phi = 5$.

Appendix A.2.1. Step 1 (Build the Modified Process)

- Define the enlarged state space $E \equiv \tilde{E} \cup E_- \cup E_+$ where $E_\pm := \{(i, j, k, \pm) : i, j \in \tilde{E}, 1 \le k \le m_{ij}^{\pm}$. Here, we obtain: $\tilde{E} = \{1, 2\}, E_+ = \{(2, 1, +)\}, E_- = \{(1, 1, -), (1, 2, -)\}$. Thus, the enlarged state space *E* includes five states.
- Built the generator matrix Q of the modified MMFF process; use Equations (9) and (10).
 Here, we obtain:

$$Q = \left(\begin{array}{cccccc} -0.23 & 0 & 0.03 & 0.06 & 0.14 \\ 0 & -1/3 & 1/3 & 0 & 0 \\ 0.05 & 0.1 & -0.15 & 0 & 0 \\ 0.25 & 0 & 0 & -0.25 & 0 \\ 0.5 & 0 & 0 & 0 & -0.5 \end{array} \right).$$

• Applying the algorithm of Section 2.2 of Breuer [27] for $((\mathbb{A}(\beta), U(\beta)) : n \ge 0)$ to obtain $\mathbb{A}(\beta)$. Built the process \mathcal{X}^R by reversing the roles of the up and down movements; similarly, use the algorithm to obtain $\mathbb{A}^R(\beta)$. For our values, we obtain:

$$A(\beta) = \begin{pmatrix} 0.05582 & 0.5831 \\ 0.3808 & 0.1048 \\ 0.5516 & 0.07592 \end{pmatrix}, \ A^{R}(\beta) = \begin{pmatrix} 0.08207 & 0.1981 & 0.4621 \\ 0.6229 & 0.03089 & 0.07207 \end{pmatrix}$$

• Calculating the special LSTs given in Corollary 2(*i*) and (*ii*),

$$(\beta I - Q_{11})^{-1}Q_{12} = \begin{pmatrix} 0.175 & 0.563 \\ 1 & 0 \\ 1 & 0 \end{pmatrix}, \ (\beta I - Q_{22})^{-1}Q_{21} = \begin{pmatrix} 0.098 & 0.197 & 0.459 \\ 1 & 0 & 0 \end{pmatrix}.$$

Appendix A.2.2. Step 2 (Derive the Cost Components

For *S* and $0 \le s \le S - 1$, do the following:

- Apply the algorithm of [49] to derive the LSTs ${}^v_u \hat{f}^r(x, y,)$ and ${}^b_a \hat{f}^r(x, y,)$ to be used.
- Derive the expected discounted cycle-length matrix $\mathbb{E}_{S}(e^{-C})$; use Proposition 1.
- Apply Corollary 3 to obtain the expected discounted ordering times vector $OC_S(\beta)$.
- Use Proposition 2 and Claim 2 to derive the vectors $SC_S(\beta)$ and $LC_S(\beta)$, respectively.
- Derive the vectors $\underline{E}_{s}(e^{-\beta \mathcal{L}}I_{o}(\mathcal{L})1_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}), \underline{E}_{S}(e^{-\beta \mathcal{L}}I(\mathcal{L})1_{\{\mathcal{L}<\varsigma\}})$ and $\overline{E}_{0}(e^{-\beta \mathcal{L}}I(\mathcal{L})1_{\{\mathcal{L}<\varsigma\}})$; use Claim 3 and Corollary 4. Then, apply Proposition 4 to obtain $\Delta_{S}(\beta)$. Finally, substitute into Equation (23) to obtain the vector $DC_{S}(\beta)$.

• Use Equations (4)–(7) to calculate the expected discounted costs $OC(\beta)$, $SC(\beta)$, $LC(\beta)$, $DC(\beta)$ and the total cost $TC(S, s) = OC(\beta) + SC(\beta) + LC(\beta) + DC(\beta)$.

Appendix A.2.3. Step 3 (the Optimal Policy)

Choose the optimal value $TC^* = \min_{S,s} \{TC(S,s)\}$. For the values mentioned above, we obtain optimal thresholds $S^* = 24$, $s^* = 2$, $TC^* = 39.896$ (see Table 2), and costs

$$OC(\beta) = 9.295$$
, $SC(\beta) = 6.855$, $LC(\beta) = 2.910$, $DC = 20.836$.

Appendix B. Proofs

Appendix B.1. Proof of Proposition 1

There is a significant difference between cases (i)-(iv), where a refilling has not been ordered yet, and the more challenging cases (v)-(xi) that occur during lead time (Recall that T_x denotes the time to hit level x.)

Cases (i)–(iv). We start with case (i); here, we assume an initial state (S, E_1) , i.e., $I_o(0) = S$ at an ascending state. Applying decomposition according to T_S and T_M , we obtain

$$\overline{E}_{S}(e^{-\beta C}) = \overline{E}_{S}(e^{-\beta C}\mathbf{1}_{\{T_{S} < T_{M}\}}) + \overline{E}_{S}(e^{-\beta C}\mathbf{1}_{\{T_{S} > T_{M}\}}).$$
(A1)

The first term of (A1) is the LST to return to level *S* while avoiding level $M(^{M-S}\mathbb{A}^r(\beta))$; from that point, the expected discounted cycle length is $\underline{E}_S(e^{-\beta C})$. The second term of (A1) is the LST to hit level *M* while avoiding $S(_S\hat{f}_{11}(S, M, \beta))$, and thereafter the expected cycle length is $\overline{E}_M(e^{-\beta C})$. The proofs of cases (ii)-(iv) are similar and thus are omitted. Note that, by Corollary 2(i), the term $(\beta \mathbb{I} - \mathbb{Q}_{11})^{-1}\mathbb{Q}_{12}$ in case (iii) is the LST to exit level *M* at a descending state.

Cases (v)-(xi). Here, we need to track the stock level during lead time. We focus on the proofs of Cases (v) and (vi); other proofs use similar techniques and thus are omitted. Case (v) considers the situation of hitting level *s* and ordering a refill. We distinguish between three events that affect the continuity of the process: either the distributor arrives, or level *S* is hit, or level 0 is hit. Thus,

$$\underline{E}_{s}^{o}(e^{-\beta C}) = \underline{E}_{s}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{s} < \min(\mathcal{L}, T_{0})\}}) + \underline{E}_{s}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{0} < \min(T_{s}, \mathcal{L})\}}) + \underline{E}_{s}^{o}(e^{-\beta C}\mathbf{1}_{\{\mathcal{L} < \min(T_{s}, T_{0})\}}).$$
(A2)

Applying the first passage times for MMFF, we obtain

$$\underline{E}_{s}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{s}<\min(\mathcal{L},T_{0})\}}) = \hat{f}_{21}(s,S,\beta+\mu)\overline{E}_{s}^{o}(e^{-\beta C}),$$

$$\underline{E}_{s}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{0}<\min(T_{s},\mathcal{L})\}}) = {}^{S}\hat{f}_{22}(s,0,\beta+\mu)\underline{E}_{0}^{o}(e^{-\beta C}).$$
(A3)

In the case of $\mathcal{L} < \min(T_S, T_0)$, the distributor arrives when $0 < I_o(\mathcal{L}) < S$, the refilling is made, the cycle ends and, thus, the remaining cycle length is \mathcal{L} . Formally,

$$\underline{E}_{s}^{o}(e^{-\beta \mathcal{C}}\mathbf{1}_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}) = \underline{E}_{s}^{o}(e^{-\beta \mathcal{L}}\mathbf{1}_{\{\mathcal{L}<\min\{T_{S},T_{0}\}}) = \underline{E}_{s}^{o}(e^{-\beta \mathcal{L}}) - \underline{E}_{s}^{o}(e^{-\beta \mathcal{L}}\mathbf{1}_{\{\mathcal{L}>\min\{T_{S},T_{0}\}}).$$
(A4)

The first term of (A4) is given by $\underline{E}_{s}^{o}(e^{-\beta \mathcal{L}}) = \frac{\neq}{\beta + \neq} e$. For the second term, we use the memoryless property of the exponential distribution. Define the stopping time $\zeta = \min(T_{S}, T_{0})$. When $\mathcal{L} > \zeta$, we can plug in $\mathcal{L} = \widetilde{\mathcal{L}} + \zeta$, where $\widetilde{\mathcal{L}}$ is an independent exponential r.v. with rate \neq . Altogether, we obtain

$$\underline{E}_{s}^{o}(e^{-\beta\mathcal{L}}\mathbf{1}_{\{\mathcal{L}>\varsigma\}}) = \underline{E}_{s}^{o}(e^{-\beta(\widetilde{\mathcal{L}}+\varsigma)}\mathbf{1}_{\{\mathcal{L}>\varsigma\}}) = \underline{E}_{s}^{o}(e^{-\beta\varsigma}\mathbf{1}_{\{\mathcal{L}>\varsigma\}})E(e^{-\beta\widetilde{\mathcal{L}}})$$

$$= \left({}_{0}\hat{f}_{21}(s,S,\beta+\mu)e^{+S}\hat{f}_{22}(s,0,\beta+\mu)e\right)\frac{\mu}{\beta+\mu}.$$
(A5)

The second line of (A5) arises via the decomposition according to $\zeta = T_S$ and $\zeta = T_0$. Specifically, when $\zeta = T_S$, we obtain $T_S < \mathcal{L}$, $T_S < T_0$ with LST $_0\hat{f}_{21}(s, S, \beta + \mu)$. When Next, we prove Case (vi). Here, \mathcal{I}_o hits level *S* from below during lead time (in contrast to Case (i), where no refilling is ordered). A similar decomposition yields

$$\overline{E}_{S}^{o}(e^{-\beta C}) = \overline{E}_{S}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{S} < \min(\mathcal{L}, T_{M})\}}) + \overline{E}_{S}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{M} < \min(\mathcal{L}, T_{S})\}}) + \overline{E}_{S}^{o}(e^{-\beta C}\mathbf{1}_{\{\mathcal{L} < \min(T_{S}, T_{M})\}}).$$
(A6)

It is easy to verify that

$$\overline{E}_{S}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{S}<\min(L,T_{M})\}}) = {}^{M-S}\mathbb{A}^{r}(\beta+\mu) \underline{E}_{S}^{o}(e^{-\beta C}),$$

$$\overline{E}_{S}^{o}(e^{-\beta C}\mathbf{1}_{\{T_{M}<\min(L,T_{S})\}}) = {}_{S}\widehat{f}_{11}(S,M,\beta+\mu) \overline{E}_{M}^{o}(e^{-\beta C}).$$
(A7)

The last term of (A6) describes the event where the distributor arrives when $I_o > S$ and, thus, no refilling is carried out, and the cycle continues. Hence,

$$\overline{E}_{\mathcal{S}}^{o}(e^{-\beta C}\mathbf{1}_{\{\mathcal{L}<\min\{T_{\mathcal{S}},T_{M}\}}) = \overline{E}_{\mathcal{S}}^{o}(e^{-\beta T_{\mathcal{S}}}\mathbf{1}_{\{\mathcal{L}< T_{\mathcal{S}}< T_{M}\}})\underline{E}_{\mathcal{S}}(e^{-\beta C}) + \overline{E}_{\mathcal{S}}^{o}(e^{-\beta T_{M}}\mathbf{1}_{\{\mathcal{L}< T_{M}< T_{\mathcal{S}}\}})\overline{E}_{M}(e^{-\beta C}).$$

Applying the first passage times to MMFF yields

$$\overline{E}_{S}^{o}(e^{-\beta T_{S}}\mathbf{1}_{\{\mathcal{L}< T_{S}< T_{M}\}}) = \overline{E}_{S}^{o}(e^{-\beta T_{S}}\mathbf{1}_{\{T_{S}< T_{M}\}}) - \overline{E}_{S}^{o}(e^{-\beta T_{S}}\mathbf{1}_{\{T_{S}< T_{M},\mathcal{L}>T_{S}\}})$$

$$= \binom{M-S}{R}^{r}(\beta) - \binom{M-S}{R}^{r}(\beta+\mu) \underline{E}_{S}(e^{-\beta C}). \quad (A8)$$

Similarly,

$$\overline{E}_{S}^{o}(e^{-\beta T_{M}}\mathbf{1}_{\{\mathcal{L}< T_{M}, T_{M} < T_{S}\}}) = \overline{E}_{S}^{o}(e^{-\beta T_{M}}\mathbf{1}_{\{T_{M} < T_{S}\}}) - \overline{E}_{S}^{o}(e^{-\beta T_{M}}\mathbf{1}_{\{T_{M} < T_{S}, \mathcal{L} > T_{M}\}})$$

$$= \left(s\hat{f}_{11}(S, M, \beta) - s\hat{f}_{11}(S, M, \beta + \not\preceq)\right). \quad (A9)$$

Substituting (A7)–(A9) into (A6) completes the derivation of Case (*vi*). We further note that, by Corollary 2(ii), the term $((\beta + \mu)\mathbb{I} - \mathbb{Q}_{22})^{-1}\mathbb{Q}_{21}$ in Case (*viii*) is the LST to exit level 0 at an ascending state. In addition, only in Cases (*v*), (*vii*), (*viii*), and (*ix*) does the cycle end when the distributor arrives (represented by the last term in each case).

Appendix B.2. Proof of Proposition 2

We present only the key steps of the proofs of Cases (i)-(iv). It is easy to verify that Cases (i) and (ii) follow from the decomposition according to min (T_M, T_S) ; the only difference is due to the initial state (S, E_1) (in Case (i)) and (M, E_2) (in Case (ii)). Similarly, Case (iv) is derived with regard to *S* and *s*. Applying Claim 1 and Corollary 2(i) immediately yields Case (iii).

Next, we will address the impact of the lead time. Cases (v)-(ix) are straightforward since they include the event $I_o(\mathcal{L}) < S$, and thereafter the PW is refilled and the cycle ends. Cases (x)-(xii) are more complicated; here, the event $I_o(\mathcal{L}) > S$ may happen. We focus on case (x) and distinguish between four events. (*a*) Hitting level *M* during lead time and continue with $\overline{SC}_M^o(\beta)$. (*b*) The distributor arrives when $S < \mathcal{I}_o < M$ (with no refilling); thereafter, level *M* is hit. Using the decomposition yields

$$\overline{\mathbf{SC}}^{o}_{S}(\beta)\mathbf{1}_{\{\mathcal{L} < T_{M} < T_{S}\}} = \overline{\mathbf{SC}}^{o}_{S}(\beta)\mathbf{1}_{\{T_{M} < T_{S}\}} - \overline{\mathbf{SC}}^{o}_{S}(\beta)\mathbf{1}_{\{T_{M} < T_{S}, \mathcal{L} > T_{M}\}},$$

where $\overline{SC}_{S}^{o}(\beta)\mathbf{1}_{\{T_{M} < T_{S}\}} = {}_{S}\widehat{f}_{11}(S, M, \beta)$ and $\overline{SC}_{S}^{o}(\beta)\mathbf{1}_{\{T_{M} < T_{S}, \mathcal{L} > T_{M}\}} = {}_{S}\widehat{f}_{11}(S, M, \beta + \mu).$ (c) Hitting level *S* during lead time and continues with $\underline{SC}_{S}^{o}(\beta)$. Finally, (*d*) the distributor arrives when $S < \mathcal{I}_{o} < M$ (with there is no refilling); thereafter, level *S* is hit; a similar technique to (*b*) is used for ${}^{M-S}\mathbb{A}^{r}(\beta) - {}^{M-S}\mathbb{A}^{r}(\beta + \mu)$. The proofs of Cases (*xi*)–(*xii*) are similar and thus are omitted. Let $Y(t) = -(\beta + \mu)t/\alpha$ and Z(t) = X(t) + Y(t). Applying Theorem 2.1 of Asmussen and Kella [50] yields that the process

$$\begin{aligned} \boldsymbol{M}(\boldsymbol{\alpha},t) &= \left(M^{1}(\boldsymbol{\alpha},t), M^{2}(\boldsymbol{\alpha},t), \dots, M^{|E_{2}|}(\boldsymbol{\alpha},t) \right)^{\mathsf{T}} \\ &= \int_{0}^{t} e^{\boldsymbol{\alpha}\mathcal{Z}(\boldsymbol{u})} \mathbf{1}_{J(\boldsymbol{u})} d\boldsymbol{u} \mathbb{K}(\boldsymbol{\alpha}) + e^{\boldsymbol{\alpha}\mathcal{Z}(\boldsymbol{0})} \mathbf{1}_{J(\boldsymbol{0})} - e^{\boldsymbol{\alpha}\mathcal{Z}(t)} \mathbf{1}_{J(t)} + \boldsymbol{\alpha} \int_{0}^{t} e^{\boldsymbol{\alpha}\mathcal{Z}(\boldsymbol{u})} \mathbf{1}_{J(\boldsymbol{u})} d\boldsymbol{Y}(\boldsymbol{u}) \\ &= \int_{0}^{t} e^{\boldsymbol{\alpha}\mathcal{X}(\boldsymbol{u}) - (\boldsymbol{\beta} + \boldsymbol{\mu})\boldsymbol{u}} \mathbf{1}_{J(\boldsymbol{u})} d\boldsymbol{u} (\mathbb{K}(\boldsymbol{\alpha}) - (\boldsymbol{\beta} + \boldsymbol{\mu})\mathbb{I}) + e^{\boldsymbol{\alpha}\mathcal{X}(\boldsymbol{0})} \mathbf{1}_{J(\boldsymbol{0})} - e^{\boldsymbol{\alpha}\mathcal{X}(t) - (\boldsymbol{\beta} + \boldsymbol{\mu})t} \mathbf{1}_{J(t)} \quad (A10) \end{aligned}$$

is an $|E_2|$ -row vector-valued zero mean martingale. The OST yields $\mathbb{E}(M(\alpha, \varsigma)) = \mathbb{E}(M(\alpha, 0)) = \mathbf{0}$, i.e.,

$$\underline{\mathbb{E}}_{s}\left(\int_{0}^{\varsigma} e^{\alpha I_{o}(t) - (\beta + \mu)t} \mathbf{1}_{J(t)} dt\right) = \left[\underline{\mathbb{E}}_{s}(e^{\alpha \mathcal{X}(\varsigma) - (\beta + \mu)\varsigma} \mathbf{1}_{J(\varsigma)}) - \underline{\mathbb{E}}_{s}(e^{\alpha \mathcal{X}(0)} \mathbf{1}_{J(0)})\right] (\mathbb{K}(\alpha) - (\beta + \mu)\mathbb{I})^{-1}.$$
(A11)

Since $\mathcal{X}(0) = s$ and $J(0) = i \in E_2$, we have that $\left[\mathbb{E}_s\left(e^{\alpha \mathcal{X}(0)}\mathbf{1}_{J(0)}\right)\right]_{i,i} = e^{\alpha s}, i \in E_2$; using an $(|E_2| \times |E|)$ matrix form yields

$$\underline{\mathbb{E}}_{s}\left(e^{\alpha\mathcal{X}(0)}\mathbf{1}_{J(0)}\right) = e^{\alpha s}(\mathbf{0} \ \mathbb{I}).$$
(A12)

To derive the $(|E_2| \times |E|)$ matrix $\underline{\mathbb{E}}_s(e^{\alpha \mathcal{X}(\varsigma) - (\beta + \mu)\varsigma} \mathbf{1}_{J(\varsigma)})$, we use the following decomposition (note that $\mathcal{X}(T_S) = S$ and $\mathcal{X}(T_0) = 0$):

$$\underline{\mathbb{E}}_{s}(e^{\alpha\mathcal{X}(\varsigma)-(\beta+\mu)\varsigma}\mathbf{1}_{J(\varsigma)}) = \underline{\mathbb{E}}_{s}(e^{\alpha\mathcal{X}(T_{S})-(\beta+\mu)T_{S}}\mathbf{1}_{\{\varsigma=T_{S},J(\varsigma)\}}) + \underline{\mathbb{E}}_{s}(e^{\alpha\mathcal{X}(T_{0})-(\beta+\mu)T_{0}}\mathbf{1}_{\{\varsigma=T_{0},J(\varsigma)\}})
= e^{\alpha S}\underline{\mathbb{E}}_{s}(e^{-(\beta+\mu)T_{S}}\mathbf{1}_{\{\varsigma=T_{S},J(\varsigma)\}}) + \underline{\mathbb{E}}_{s}(e^{-(\beta+\mu)T_{0}}\mathbf{1}_{\{\varsigma=T_{0},J(\varsigma)\}})
= e^{\alpha S}\left(_{0}\widehat{f}_{21}(s,S,\beta+\mu) \mathbf{0}\right) + \left(\mathbf{0}^{-S}\widehat{f}_{22}(s,0,\beta+\mu)\right).$$
(A13)

Plugging (A13) and (A12) into (A11) yields (28).

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