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International Society for Bipolar Disorders Task Force on Suicide: meta-analyses and meta-regression of correlates of suicide attempts and suicide deaths in bipolar disorder

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Abstract

Objectives: Bipolar disorder is associated with a high risk of suicide attempts and suicide death. The main objective of this paper was to identify and quantify the demographic and clinical correlates of attempted and completed suicide in people with bipolar disorder.

Methods: Within the framework of the *International Society for Bipolar Disorders Task Force on Suicide*, a systematic review of articles published since 1980 characterized by both key terms *bipolar disorder* and *'suicide attempts* or *suicide'* was conducted, and data extracted for analysis from all eligible articles. Demographic and clinical variables for which \geq 3 studies with usable data were available were meta-analyzed using fixed or random-effects models for association with suicide attempts and suicide deaths. There was considerable heterogeneity in the methods employed by included studies.

Results: Variables significantly associated with suicide attempts were: female sex, younger age of illness onset, depressive polarity of first illness episode, depressive polarity of current or most recent episode, comorbid anxiety disorder, any comorbid substance use disorder, alcohol use disorder, any illicit substance use, comorbid cluster B/borderline personality disorder, and first-degree family history of suicide. Suicide deaths were significantly associated with male sex and first-degree family history of suicide.

Conclusions: This paper reports on the presence and magnitude of the correlates of suicide attempts and suicide deaths in bipolar disorder. These findings do not address causation, and the heterogeneity of data sources should limit the direct clinical ranking of correlates. Our results nonetheless support the notion of incorporating diagnosis-specific data in the development of models of understanding suicide in bipolar disorder.

Keywords

bipolar disorder; meta-analysis; suicide

Mental illness is present in nearly all people who attempt or die by suicide, and among psychiatric diagnoses, bipolar disorder (BD) may be associated with the highest suicide risk (1–14). Among people with BD, the estimated rate of death by suicide is 0.2–0.4 per 100 person-years (5, 15–17), however these rates may reflect data from higher risk periods of time in the life of someone with BD, and simple extrapolation to estimates of lifetime risk are inherently unreliable. The absolute risk of suicide among patients with a diagnosis of BD at first hospital contact has been found to be around 8% for men and 5% for women over a median of 18 years follow-up (12), and the standardized mortality ratio of BD suicide deaths compared to the general population has been reported to be 10–30-fold (18–21). Rates of suicide attempts are also very high, with an estimated annual risk of 0.9% per year, and a lifetime risk of up to one-half of sufferers with BD (22–25).

There are a number of well-described correlates of suicide attempts and suicide deaths identified in the broader mental illness or public health literature, including sex, depression, anxiety, substance use, family history of suicide, and others (26–28). Suicide is a behavioral endpoint that results from a multitude of factors, and confirming whether these broadly identified correlates of risk of suicide attempts and suicide are also present in BD populations is imperative. Analogous investigation of BD-specific clinical factors such as illness subtype and polarity of first and most recent episode is also warranted. These type of data would inform the emerging effort to move towards more diagnosis-specific approaches to understanding suicide risk, risk assessments, and prevention (29–33).

Recently published reviews on suicide attempts and suicide in BD shed light on the scope of the public health, clinical assessment and management challenges, and also illuminate glaring gaps in the available data (18, 29, 30, 34–37). One such gap is the absence of up to date, quantitative, meta-analytic models of the specific demographic and clinical factors putatively associated with suicide attempts and suicide in BD. While individual studies can identify correlates within a specific study population, these findings require both replication and an estimate of the magnitude of the association in order to inform risk assessments and suggest avenues for prevention. Prior meta-analyses by Hawton et al. (38) and Novick et al. (39) laid the foundation for this approach in BD, however, a major expansion of data has since occurred in the past few years, which is not captured in these earlier publications. This is especially relevant for suicide deaths, which have now been more thoroughly examined in several large epidemiological BD samples (4, 8, 12, 15, 40).

The *International Society for Bipolar Disorders* (ISBD) is a leading organization devoted to promoting international collaboration in the study of BD. Under the auspices of the ISBD, a Task Force on Suicide comprised of 20 international experts was launched, with the objectives of completing a comprehensive review of the available literature and conducting meta-analyses on putative correlates of suicide attempts and suicide in BD. This publication is the product of the work done by this task force to achieve the goal of identifying and quantifying the degree of associations between demographic and clinical variables with the risk of suicide attempts and suicide death in people with BD.

Methods

Study search and selection

We conducted a systematic review of English-language articles using keywords *bipolar disorder* and *'suicide attempts'* or *'suicide'* published between January 1, 1980 and June 30, 2013. The search was then expanded via the ancestry approach by manual examination of reference lists of included articles and recent published reviews on suicide or suicide attempts in BD (18, 29, 30, 34–37). Included articles had to have met all the following criteria: (i) subjects with BD comprised all or a large majority (> 80%) of the study population; (ii) study population was exclusively > 13-years-old; (iii) a binary measure of suicide attempts or suicide deaths was reported; (iv) a non-suicide attempt or non-suicide group was included; and (v) a binary measure of a demographic or clinical variable of interest was reported or could be calculated from the published data. Age of onset was included as a continuous variable, since there was no uniform definition of early or later age

of onset used across studies. Both prospective and retrospective studies were included, as were studies from either clinical or epidemiological samples. There is no uniform definition of a suicide attempt in relation to intent or lethality, however studies were excluded if they only reported on non-suicidal self-injury, suicidal ideation or if suicidality ratings were reported as a continuous measure.

Data extraction

Using the PRISMA framework for systematic reviews (41), the initial search yielded 1,700 abstracts. These were screened for eligibility criteria and duplication of samples, resulting in 74 full-text articles being assessed by two trained investigators (AS and CR) who reviewed each article for inclusion and exclusion criteria. Articles with any uncertainty about eligibility were discussed and a decision made through consensus. Studies were excluded at this stage if they contained subject overlap with another article or did not report sufficient details for data analysis (n = 33 studies excluded). This resulted in an initial group of 41 eligible studies, including one in which corresponding author was contacted for clarification (one study). Whenever possible, task force members, as a geographically diverse group of experts who were authors on many of the included studies, provided additional details of data when insufficient information was available in the published article (additional three studies). This resulted in a total of 44 eligible studies.

In total, 34 papers reported on suicide attempts (total N = 50,004 subjects with BD) across one or more variables, with 31/34 using clinical samples, 29/34 using non-representative samples, and 30/34 reporting suicide attempts in a retrospective manner. There were 12 papers that reported on suicide deaths (total N = 75,137 subjects with BD), with 8/12 using clinical samples, 4/12 using non-representative samples, and 8/12 identifying suicide in a retrospective sample. Only two studies reported on both suicide attempts and deaths (15, 40).

For each article, the following data were extracted and coded: (i) author name(s); (ii) year of publication; (iii) number of subjects with or without a suicide attempt for each demographic and clinical variable of interest; and (iv) number of suicide deaths or non-deaths for each demographic and clinical variable of interest. No study included all variables of interest, so data were only extracted for those variables examined in the publication. Variables of interest were chosen based on the general and BD-specific literature on suicide attempts and deaths, and included: (i) sex; (ii) age of onset of BD; (iii) subtype of BD [bipolar I disorder (BD-I) or bipolar II disorder (BD-II)]; (iv) polarity of first mood episode; (v) polarity of current or most recent mood episode (mixed episodes were generally reported as a subset of mania, and there was insufficient data to specifically analyze mixed episodes); (vi) lifetime history of past suicide attempts (for analysis of suicide deaths only); (vii) current or lifetime anxiety disorder; (viii) presence of lifetime history of psychotic symptoms; (ix) current or lifetime substance use disorder (most studies reported as either abuse, dependence, or both), as well as the following three subcategories: (x_A) current or lifetime alcohol use disorder; (x_B) current or lifetime cannabis use; and (x_C) any current or lifetime illicit substance use disorder; (xi) current or lifetime personality disorder; and (xii) first-degree family history of death by suicide.

Data analysis

Cochrane Information Management System–Review Manager (RevMan) version 5.2 (November 2012) was utilized to conduct the meta-analyses. For each variable of interest, a meta-analysis was only conducted if there was a minimum of three studies with usable data. A total of 13 variables had sufficient data for analysis on suicide attempts (all except past suicide attempts) and four variables had sufficient data for analysis on suicide. The same study could yield data for multiple analyses, as many studies reported on more than one variable of interest. Odds ratios (OR) with 95% confidence intervals (CI) were calculated for all binary measures and weighted mean difference was calculated for age of illness onset, the only continuous measure. Random-effects models were conducted when significant heterogeneity was present (Cochrane test p < 0.1); otherwise fixed-effects models were conducted. As a sensitivity analysis on the effect of very large studies, we re-ran metaanalyses removing any single study that had a > 40% weighting.

A meta-regression was conducted (STATA software) with suicide attempts in males versus females as the meta-analytic outcome variable, and covariates included based on a sufficient number of observations within studies that reported on sex-differences. Covariates included polarity of first mood episode, mean age of illness onset, BD subtype, psychosis, any substance use disorder, alcohol or illicit substance use disorder, and family history of suicide. Tests of meta-bias (effect of small studies) and meta-influence (influence of a single study) were also conducted. Other meta-regressions could not be completed due to an insufficient number of observations.

Results

Findings related to suicide attempts

Thirty-four papers reported suicide attempts across one or more variables in a manner suitable for meta-analysis, with a total of 50,004 subjects with BD included in the non-overlapping samples. (15, 24, 40, 42–72). We examined 13 variables, including sex, age of illness onset, BD subtype (BD-I or BD-II), polarity of first mood episode, polarity of current or most recent mood episode, lifetime comorbid anxiety disorder, lifetime history of psychotic symptoms, any substance use disorder, alcohol use disorder, cannabis use, any illicit substance use disorder, comorbid cluster B/borderline personality disorder, and first-degree family history of suicide. Each of Figures 1–9 and Supplementary Figures S1–S3 shows the results of the meta-analysis for the specific variable(s) of interest.

Numerous variables were found to be significantly correlated with presence of suicide attempts. Women were significantly more likely to attempt suicide (OR = 1.54, 95% CI: 1.44-1.66, p < 0.00001), with 8/20 studies reporting a significant effect in this direction, and the remaining 12/20 studies reporting no significant sex-based difference (Fig. 1). A meta-regression conducted with sex-differences in suicide attempts as the meta-analytic outcome found no significant independent association for any other tested covariate. There were also no significant small study or single study effects.

Age of illness onset was 2.99 years younger (95% CI: 2.20–3.78 years, p < 0.00001) among those with a history of suicide attempt (Fig. 2), with a standardized mean difference of -0.29 (95% CI: -0.36 to -0.21, p < 0.0001).

BD subtypes, BD-I or BD-II, were examined in 14 relatively evenly weighted studies, with two studies finding higher rates of suicide attempt in BD-I, two studies finding a higher rate in BD-II, and the remainder finding no difference, resulting in no overall effect of subtype being identified (OR = 1.07, 95% CI: 0.79-1.45, p = 0.68) (Supplementary Fig. S1).

Subjects with a depressive polarity of first mood episode were nearly twice as likely to attempt suicide (OR = 1.92, 95% CI: 1.39-2.65, p < 0.0001) (Fig. 3), with all seven studies reporting a similar direction of effect, and most reaching statistical significance.

Depressive polarity of the current or most recent mood episode had the strongest association with a suicide attempt (OR = 5.99, 95% CI = 1.75-20.5, p = 0.004), but only three studies reported on this variable, resulting in wide confidence intervals (Fig. 4).

The presence of a lifetime comorbid anxiety disorder was significantly associated with suicide attempts in 8/13 studies, with an OR of 1.81 (95% CI: 1.66–1.97, p < 0.0001) (Fig. 5). Although one study had a very large weighting in this analysis, there was a consistent direction of effect with the other studies.

Of the seven studies that examined history of psychosis, one reported a higher rate of suicide attempt among subjects with a history of psychosis, one reported a higher rate among subjects without psychosis, and the remainder found no significant difference, resulting in no significant association being identified (OR = 0.91, 95% CI: 0.64-1.30, p = 0.61) (Supplementary Fig. S2).

The presence of a current or lifetime comorbid substance use disorder was separated into 4 non-mutually exclusive groups, including (i) any substance use disorder (OR = 1.81, 95% CI: 1.31-2.50, p < 0.0001) (Fig. 6); (ii) alcohol use disorder (OR = 1.60, 95% CI: 1.31-1.97, p < 0.0001) (Fig. 7A); (iii) any cannabis use (OR = 1.29, 95% CI: 0.85-1.94, p = 0.23) (Supplementary Fig. S3); and (iv) any illicit substance use disorder (OR = 1.72, 95% CI: 1.23-2.39, p = 0.001) (Fig. 7B). Each of these substance use variables, except cannabis use, was significantly associated with suicide attempts.

Comorbid cluster B/borderline personality disorder was strongly associated with suicide attempts in all 5/5 studies, resulting in an OR of 2.51 (95% CI: 1.91-3.31, p < 0.0001) (Fig. 8). Data on other personality disorders or traits were not available in a sufficient number of studies to permit meta-analysis.

Finally, a first-degree family history of death by suicide was found to be associated with suicide attempts in all 11 studies, among a total of 7,452 subjects with BD (OR = 1.69, 95% CI: 1.25-2.27, p = 0.0006) (Fig. 9).

Findings related to suicide deaths

Twelve studies on completed suicide were available (2, 4, 8, 12, 15, 21, 40, 73–77). We could examine four variables: sex, lifetime history of psychotic symptoms, any substance use disorder, and first-degree family history of suicide. Despite the smaller number of studies, a total of 75,137 subjects with BD were included in these analyses, mostly comprised of several large US and European cohort studies (8, 12, 15, 40).

The sex-based analysis from 11 studies included a large sample size of 75,055 subjects with BD and a total of 1,149 suicide deaths. Suicide deaths were significantly associated with male sex (OR = 1.83,95% CI: 1.41-2.39, p < 0.0001) (Fig. 10), with each study reporting an effect in the same direction, most at a significant level.

Similar to the data for suicide attempts, a history of psychosis had no significant association with suicide deaths (OR = 0.93, 95% CI: 0.50–1.74, p = 0.82) (Supplementary Fig. S4). However, in contrast to the finding for suicide attempts, the presence of any substance use disorder was not significantly associated with suicide deaths (OR = 1.20, 95% CI: 0.93–1.56, p = 0.17) (Supplementary Fig. S5), with 3/4 individual studies finding no significant association.

Only four studies reported on first-degree family history of suicide, but nonetheless a significant association was found (OR = 2.91, 95% CI: 1.54–5.48, p = 0.001) (Fig. 11).

Figure 12 provides a list of variables that were or were not associated with suicide attempts or suicide in people with BD.

Sensitivity analyses for effect of very large studies found no switch from variables being significant to non-significant, or vice versa. Furthermore, using random effect analyses even for variables with non-significant heterogeneity resulted in only very minor changes to ORs and modest widening of CIs, with the exception being the association between presence of any substance use disorder and suicide deaths, which increased to OR = 1.49 (95% CI = 0.88–2.55).

Discussion

This paper reports on a comprehensive set of meta-analyses conducted to identify and quantify the correlates of suicide attempts and suicide deaths in BD populations. It was undertaken as part of the work of the ISBD Task Force on Suicide—an international collaborative effort to study suicide in BD. There have been prior meta-analyses published on BD subtype and suicide attempts (39), and on a broader examination of demographic and clinical correlates of suicide attempts and suicide in BD, but this latter work by Hawton et al. (38) was published in 2005, and therefore did not include the large number of studies published in the last decade. Our results are noteworthy in reporting on a number of general and BD-specific variables that are relevant for understanding risk of suicide attempts and suicide specific to a population with BD. Quantifying the associations permits ranking of correlates and informs risk estimates in a more meaningful way.

Of the variables for which sufficient data were available to conduct meta-analyses, 10 out of 13 were significantly associated with suicide attempts, and two out of four were significantly associated with suicide deaths. Factors significantly associated with suicide attempts were (ranked from highest to lowest ORs): depressive polarity of current or recent episode (OR = 5.99), comorbid cluster B/borderline personality disorder (OR = 2.51), depressive polarity of first illness episode (OR = 1.92), comorbid anxiety disorder (OR = 1.81), any substance use disorder (OR = 1.81), any illicit substance use (OR = 1.72), first-degree family history of suicide (OR = 1.69), alcohol use disorder (OR = 1.60), and female sex (OR = 1.54). Earlier age of illness onset was also significantly associated with suicide attempts (mean difference 2.99 years, OR = 1.69).

The evidence for both depressive polarity of current/most recent episode and first mood episode each being correlated with suicide attempts was generated from clinical samples. Mixed symptoms have previously been associated with elevated risk for suicide attempts (29, 64), but rates of mixed episodes were low in the studies we analyzed, and were most often classified together with manic episodes. Recent evidence suggests that broadly defined mixed states may be associated with the highest risk of suicide attempts per time period spent in a specific phase of illness (78), and with the new broader definition of mixed states in DSM-5, future studies should be able to examine more accurately the impact of mixed symptoms on suicide risk, whether during a manic or depressive phase. Nonetheless, our analyses identified current or most recent depressive episode as the strongest correlate of suicide attempts, likely as a result of a combination of elevated risk per time period, as well as the predominance of the depressive phase of illness in the natural course of BD.

The comorbidity between cluster B/borderline personality disorder and BD has long been a source of diagnostic complexity, requiring comprehensive etiological and management considerations (79). Our data identified that elevated risk of suicide attempt (OR = 2.51) is another important factor to consider, with all 5 studies demonstrating a strong association. It is possible that having recurrent suicidal behavior or self-harm without intent to die as part of the diagnostic criteria for borderline personality disorder may have resulted in an inflation of the association, and we could not address this issue in the available data, however other cluster B personality disorders do not include this criterion, and as such it is unlikely to fully account for the identified association.

The finding of comorbid anxiety being significantly associated with suicide attempts was highly consistent across studies. A majority of people with BD will experience a comorbid anxiety disorder at some point in their lives (24, 80), thus elevated rates of suicide attempts are highly clinically relevant. There is some evidence that the association between anxiety and suicide attempts in BD may be partially mediated through increased rumination (24), as well as by comorbid cluster B personality disorders (47). Unfortunately, most studies do not report on the sequencing of the comorbid symptoms or disorders and the suicide attempt(s), which may or may not be confluent. It is therefore difficult to develop a clear attribution model without sufficient prospective data to elucidate the onset and timing of the comorbid anxiety in relation to the suicide attempt.

The association between suicide attempt and substance use disorders was of a similar magnitude to anxiety (both OR = 1.81), but the results for substance use were more variables across studies, with several showing a trend in the opposite direction. This led us to break down the substance use category into alcohol use disorder, any illicit drug use, and specific cannabis use. Significant associations were identified among studies of alcohol use disorder (OR = 1.60) and illicit drug use (OR = 1.72), but not for cannabis use (OR = 1.29). While the largest study of cannabis use did report a significant association with suicide attempts, the other three studies found no significant effect. In a recent review by Watkins and Meyer (81), preliminary evidence identified impulsivity as mediating the association between alcohol use and suicide attempts in BD, but whether this is true across substance use disorders is not known.

Female sex was associated with suicide attempts (OR = 1.54), but this was a weaker association than has most commonly been reported in non-BD samples (27, 82, 83), and only 8/20 studies in our analysis reported any significant sex-based association. Results of the meta-regression did not identify any significant covariates to this association, but there may be other diagnosis-specific factors such as a greater female preponderance for a more depression-prone course of bipolar illness (84, 85) that are additionally relevant. It is also worth noting that men with BD were underrepresented in the available trials, accounting for only 36.7% of all subjects in the studies that reported on suicide attempts. Given the lack of sex-based differences in prevalence of BD, this suggests an ascertainment or sampling bias may also be relevant when interpreting these results. Overall, the findings highlight the importance of a diagnostic-specific examination of correlates of suicide attempts such as sex, since extrapolation from a broader literature on sex-based differences may be inaccurate.

Earlier age of onset of BD was also significantly associated with a history of suicide attempts, with a mean difference of 2.99 years. Early onset of illness has consistently been shown to have negative prognostic implications (86, 87), and our data support this observation. It is worth noting that we only extracted data from reports of study populations exclusively > 13 years-of-age, and 27/32 sample groups for this analysis had a mean age of illness onset of \geq 19 years, therefore our findings primarily relate to adult-onset BD.

There were three variables tested that were not significantly associated with suicide attempts. In contrast to the limited data on cannabis use, there were more studies on BD subtype (14 studies) and history of psychosis (seven studies). Considerable variability was present in the BD subtype studies, with ORs for BD-I varying from 2.38 to 0.27. There were over twice the number of subjects with BD-I compared to subjects with BD-II, which is not in line with the relative equal lifetime prevalence of the two subtypes (88), and again suggests that there may be a sample bias in the available literature. Nonetheless, analysis of current data suggests that subjects with BD-II are just as likely to attempt suicide as those with BD-I.

History of psychosis was examined in relation to both suicide attempts and suicide deaths, and neither analysis identified a significant association, although there was a wide confidence interval (0.50-1.74) in the suicide death analysis. It is important to note that

studies identified a history of psychotic symptoms at any point in the course of illness, and did not focus on specific phases of illness. Given the strong correlation between current or recent depressive episode and suicide attempts, it is possible that a current or recent psychotic depression may be associated with suicide attempts or even suicide deaths, but this has not clearly been shown in other studies (89), and data was not available to test this hypothesis in BD.

Another challenge is that there are far fewer studies of correlates of suicide deaths in BD, as compared to suicide attempts. Suicide deaths are rarer and inherently more difficult to study, as evidenced by only four variables being sufficiently examined to allow for meta-analysis. In addition to psychosis, other variables tested included sex, any substance use disorder and first-degree family history of suicide. Men with BD were nearly twice as likely to die by suicide as compared to women with BD (OR = 1.83). This is in keeping with the consistent evidence on sex-differences in suicide rates, but as with suicide attempts in BD, the size of the difference is smaller than in general suicide samples (28), which show up to a four-fold difference, and is more in line with an approximate 2:1 ratio identified in a review of suicide in depression (26). Sex-based differences in methods of suicide have been reported in the broad literature, and may also be relevant to understanding sex-based differences in BD, but there was insufficient data to examine this possibility in detail.

A noteworthy finding is that a lifetime history of any substance use disorder was not associated with higher rates of suicide deaths in BD, however the OR of 1.20 (95% CI: 0.93–1.56) was in the direction of a positive correlation. Nonetheless, this contrasts with the elevated risk of suicide attempts in our analyses and the literature on higher rates of suicide deaths in general samples (90) as well as specific to depression (26) and schizophrenia (91). One possible explanation relates to sample characteristics, as the two most heavily weighted studies in our meta-analysis (Supplementary Fig. S4) were large population-based epidemiological samples that did not find a significant association, in contrast to the one pure clinical sample from the Systematic Treatment Enhancement Program for Bipolar Disorder (STEP-BD) that reported a strong association (OR = 8.05). This suggests the possible impact of Berkson's bias (92), with a more severe sample of comorbid patients entering into a clinical protocol. This hypothesis, however, is speculative at best, and reinforces the importance of having more data on the specific connection between substance use and suicide deaths in BD, including mediating factors and greater details of the substance use in relation to phases of illness. Until these can be examined, the methodological limitations of our analysis and the divergence from the bulk of the broader literature on substance use and suicide suggest that our result must be interpreted with appropriate caution. It would also be important to further understand the association as it pertains to timing of substance use, as evidenced by recent data on self-poisoning deaths among people with BD which identified 41% of cases having alcohol in the system at the time of death (93).

Finally, first-degree family history of suicide was found to have the strongest association with suicide deaths in BD (OR = 2.91). This supports prior literature on the significance of family history of suicide on elevating suicide risk from general psychiatric samples (94–96),

and reinforces the importance of integrating genetic, epigenetic, and social learning effects when building models of psychobiological contributors to suicidal behavior in BD (36, 97).

There are a number of important limitations that must be considered when interpreting these data. First, considerable heterogeneity exists in the available literature, but since the number of studies is not very large, we included all studies with either epidemiological or clinical samples. The largest studies included were usually prospective epidemiological samples, and the predominant concordance of results with smaller clinical samples should serve to strengthen the generalizability of our results. However it is possible that the magnitude or even presence of specific associations may be different within certain BD subpopulations. This degree of granularity could not be determined based on the available literature. Second, as noted earlier, the relationship in time between the variables of interest and the outcome of a suicide attempt or suicide death was not always known. For instance, higher lifetime rates of suicide attempts in subjects with a lifetime comorbid anxiety disorder does not provide information on whether these were contemporaneous and therefore cannot be used to determine the direction of the effect or other interactions. This was not the case with all variables, since a number are not time-based, but nonetheless the usual caution around association not equaling causation is relevant here. In a related limitation, we included studies with retrospective or prospective designs; therefore it is important to consider the significant findings as being evidence of correlation rather than for any attempt to assign levels of risk of suicide attempts or suicide for a particular patient. Retrospective studies also carry the limitation of recall bias, therefore while the reported data can serve to inform future studies aimed at determining this type of risk, prospective stratified designs are required. In a similar vein, the findings do not address causation, and the heterogeneity of data sources should limit the direct clinical ranking of correlates. The results of one variable should therefore not be compared to another, since different studies were used in the analyses. An additional limitation is that data on correlates of suicide attempts should not be assumed to be relevant for understanding risk of suicide deaths, as evidenced by the well described gender paradox (98) also seen here, as well as the discordant results on substance use disorders. Finally, a number of putatively important variables such as past suicide attempts, hopelessness, recent hospitalization, childhood abuse and smoking could not be meta-analyzed due to lack of sufficient number of studies with the required data reporting. This precludes considering the correlates reported in this paper as an exhaustive list of variables associated with suicide attempts or suicide in BD, and reinforces the importance of including all potential variables of interest in future studies of suicidal behavior in BD.

The principal objective of this report was to identify correlates of suicide attempts and suicide deaths in studies specific to BD. The meta-analytic results identified ten significant correlates of suicide attempts and two significant correlates of suicide deaths, some of which are only relevant to BD, and others that have been studied more broadly. As has been discussed earlier, there are a few noteworthy differences between the results we obtained and the broader literature on suicidal behavior and deaths among general samples (28), or even those focused on unipolar depressive disorder (26) or schizophrenia (91). While appropriate caution is required when interpreting the findings, our results support the importance of incorporating diagnosis-specific data on clinical variables relevant to BD, as well on the diagnosis-specific results from a broader set of correlates into the development of models for

understanding suicide risk in BD (29, 36, 99). This conclusion has implications for future work on developing specific preventative programs and therapeutic interventions for people with BD.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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	Male	es	Fema	les		Odds Ratio (Non-event)	Odds Ratio (Non-event)
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% CI
Azorin et al. 2009(67)	133	458	249	632	7.7%	1.59 [1.23, 2.05]	-
Bellivier et al.2011(68)	213	932	450	1226	12.5%	1.96 [1.62, 2.37]	-
Dalton et al. 2003(69)	32	129	54	207	2.4%	1.07 [0.65, 1.77]	
Dennehy et al. 2011(40)	56	1881	118	2479	5.0%	1.63 [1.18, 2.25]	
Engstrom et al. 2004(70)	4	40	19	60	0.3%	4.17 [1.30, 13.40]	
Finseth et al. 2012(71)	37	93	56	113	1.7%	1.49 [0.85, 2.59]	+
Galfalvy et al 2006(66)	18	32	19	32	0.6%	1.14 [0.42, 3.07]	
Gazalle et al. 2007(65)	20	47	62	122	1.2%	1.40 [0.71, 2.75]	
Goldberg et al. 1998(64)	30	85	33	99	1.8%	0.92 [0.50, 1.69]	
Grunebaum et al. 2006(63)	23	33	41	63	0.9%	0.81 [0.33, 2.00]	
Leverich et al. 2003(72)	83	274	136	374	5.0%	1.31 [0.94, 1.83]	-
Moor et al. 2012(62)	6	24	41	76	0.3%	3.51 [1.26, 9.83]	
Parmentier et al.2012(61)	104	274	176	378	5.3%	1.42 [1.04, 1.96]	
Perroud et al. 2007(60)	67	179	97	227	3.5%	1.25 [0.84, 1.86]	
Ryu et al. 2010(59)	30	262	46	317	2.3%	1.31 [0.80, 2.15]	
Sanchez-Gistau et al 2009(58)) 43	177	69	213	2.6%	1.49 [0.95, 2.34]	
Simon GE et al. 2007(15)	450	11160	1269	21200	45.4%	1.52 [1.36, 1.69]	
Simon NM et al. 2007(24)	15	49	31	71	0.8%	1.76 [0.82, 3.78]	
Swann et al. 2005(57)	8	17	16	31	0.4%	1.20 [0.37, 3.92]	·
Valtonen et al. 2006(56)	7	85	20	91	0.5%	3.14 [1.25, 7.87]	· · · · · ·
Total (95% CI)		16231		28011	100.0%	1.54 [1.44, 1.66]	•
Total events	1379		3002				· · · · · · · · · · · · · · · · · · ·
Heterogeneity: Chi ² = 23.86, d	f = 19 (P	= 0.20);	l ² = 20%				
Heterogeneity: $Chi^2 = 23.86$, di Test for overall effect: Z = 11.6	Sector Constant Sector		l ² = 20%				0.02 0.1 1 Favours [Males] Favour

Fig. 1.

Meta-analysis of suicide attempts among males and females with bipolar disorder. CI = confidence interval.

	Age of On	set (Attemp	oters)	Age of Ons	set (NonAtte	empt)		Mean Difference	Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	I IV, Random, 95% CI	
Azorin et al. 2009(67)	26.73	9.44	382	29.32	11.23	708	10.2%	-2.59 [-3.85, -1.33]	-	
Bellivier et al.2011(68)	27.4	9.8	663	31	11.5	1556	11.5%	-3.60 [-4.54, -2.66]	i -	
Dalton et al. 2003(69)	19.16	8.36	86	20.45	7.27	250	7.4%	-1.29 [-3.27, 0.69]	i	
Dennehy et al. 2011(40)	14.76	7.05	174	17.21	8.74	4186	10.9%	-2.45 [-3.53, -1.37]	i —	
Engstrom et al. 2004(70)	25.1	8.4	23	31.9	10.2	77	2.9%	-6.80 [-10.92, -2.68]	i	
Finseth et al. 2012(71)	19	12	93	19.6	12	113	4.1%	-0.60 [-3.89, 2.69]	j <u> </u>	
Gazalle et al. 2007(65)	23.68	12.05	82	26.7	11.7	87	3.6%	-3.02 [-6.60, 0.56]	1	
Goldberg et al. 1998(64)	26.2	11.9	63	27.4	12.4	121	3.5%	-1.20 [-4.88, 2.48]	i	
Grunebaum et al. 2006(63)	19.8	10.9	64	28.7	10.6	32	2.5%	-8.90 [-13.44, -4.36]	n	
Leverich et al. 2003(72)	17	10	219	21	11	429	8.5%	-4.00 [-5.68, -2.32]	i	
Parmentier et al.2012(61)	23.11	9.13	280	27.03	11.11	372	9.0%	-3.92 [-5.48, -2.36]	i —	
Perroud et al. 2007(60)	23	9.74	164	27.4	11.4	242	7.1%	-4.40 [-6.47, -2.33]	i —	
Ryu et al. 2010(59)	26.5	8.4	76	30.5	12.7	503	6.7%	-4.00 [-6.19, -1.81]	j	
Sanchez-Gistau et al 2009(58)	25.7	11.2	112	27.3	11.3	278	5.9%	-1.60 [-4.06, 0.86]	i	
Suominen et al.2009(100)	21.7	9.5	35	24.1	9.7	141	3.7%	-2.40 [-5.93, 1.13]	1	
Swann et al. 2005(57)	17.8	8.29	24	14.3	6.9	24	2.7%	3.50 [-0.82, 7.82]	1	
Total (95% CI)			2540			9119	100.0%	-2.99 [-3.78, -2.20]	1	
Heterogeneity: Tau ² = 1.19; Ch	ni² = 33.79, c	if = 15 (P =	0.004); 12	= 56%						
Test for overall effect: Z = 7.41	(P < 0.0000	01)	sana an an Albairtí						-20 -10 0 1 Younger Age of Onset Older Age	0 20 e of Onset

Fig. 2.

Meta-analysis of suicide attempts based on age of onset of bipolar disorder. SD = standard deviation; CI = confidence interval.

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	1st Episode Dep	ression	1st Episode	Mania	Odds Ratio		Odd	s Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Ran	dom, 95% CI
Azorin et al. 2009(67)	173	377	208	713	18.8%	2.06 [1.59, 2.67]		+
Bellivier et al.2011(68)	342	1095	321	1121	20.0%	1.13 [0.94, 1.36]		-
Chaudhury et al. 2007(52)	61	83	14	30	8.4%	3.17 [1.33, 7.54]		3
Finseth et al. 2012(71)	60	109	32	96	12.8%	2.45 [1.39, 4.32]		
Neves et al. 2009(51)	52	102	25	66	11.7%	1.71 [0.91, 3.21]		—
Ryu et al. 2010(59)	40	200	26	304	13.5%	2.67 [1.57, 4.54]		
Sanchez-Gistau et al 2009(58)) 78	234	34	156	14.7%	1.79 [1.12, 2.86]		
Total (95% CI)		2200		2486	100.0%	1.92 [1.39, 2.65]	6	•
Total events	806		660					
Heterogeneity: Tau ² = 0.13; Ch	ni ² = 25.70, df = 6	(P = 0.000)	3); l ² = 77%					
Test for overall effect: Z = 3.98	8 (P < 0.0001)	A. 1940393					0.02 0.1 Favours [Mania 1st	1 10 50] Favours [Depression 1st

Fig. 3.

Meta-analysis of suicide attempts based on polarity of first episode of bipolar disorder. BD-I = bipolar I disorder; BD-II = bipolar II disorder; CI = confidence interval.

	Current/Recent De	Current/Recent Depressed		ania/Mix		Odds Ratio	Odds Ratio			
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI		M-H, Ran	dom, 95% Cl	
Liu et al. 1996(50)	2	26	1	681	17.8%	56.67 [4.97, 646.72]			1	<u> </u>
Tondo et al. 1999(53)	8	32	33	346	45.5%	3.16 [1.32, 7.60]				
Valtonen et al. 2007(49)	24	105	3	48	36.6%	4.44 [1.27, 15.58]				
Total (95% CI)		163		1075	100.0%	5.99 [1.75, 20.50]				
Total events	34		37							
Heterogeneity: Tau ² = 0.6	6; Chi ² = 4.80, df = 2 (l	= 0.09); I	² = 58%				+		1 10	
Test for overall effect: Z =	2.85 (P = 0.004)						0.002 Favours	0.1 [Mania/Mixed]] Favours [De	500 [pressed]

Fig. 4.

Meta-analysis of suicide attempts based on polarity of current or most recent episode of bipolar disorder. CI = confidence interval.

	Comorbid /	Anxiety	NoAnx	ciety		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI
Azorin et al. 2009(67)	88	214	294	876	9.2%	1.38 [1.02, 1.88]	
Dalton et al. 2003(69)	27	88	59	248	2.9%	1.42 [0.83, 2.43]	
Dennehy et al. 2011(40)	114	2011	60	2349	7.1%	2.29 [1.67, 3.15]	
Gazalle et al. 2007(65)	10	17	72	152	0.8%	1.59 [0.57, 4.39]	
_everich et al. 2003(72)	64	164	58	216	4.1%	1.74 [1.13, 2.69]	
Malloy-Diniz et al. 2009(48)	14	22	4	17	0.2%	5.69 [1.38, 23.48]	
Nakagawa et al. 2008(47)	35	60	33	56	1.9%	0.98 [0.47, 2.04]	
Parmentier et al.2012(61)	23	47	257	605	2.6%	1.30 [0.72, 2.35]	
Perroud et al. 2007(60)	71	124	93	282	3.3%	2.72 [1.76, 4.20]	
Sanchez-Gistau et al 2009(58) 21	42	91	348	1.3%	2.82 [1.47, 5.41]	
Simon GE et al. 2007(15)	576	7389	1143	25345	64.5%	1.79 [1.61, 1.99]	
Simon NM et al. 2007(24)	35	75	11	45	1.0%	2.70 [1.19, 6.12]	
Valtonen et al. 2006(56)	18	95	9	81	1.1%	1.87 [0.79, 4.43]	
Total (95% CI)		10348		30620	100.0%	1.81 [1.66, 1.97]	•
Total events	1096		2184				
Heterogeneity: Chi ² = 18.53, d	f = 12 (P = 0	.10); 12 = 3	35%				
Test for overall effect: Z = 13.7	2 (P < 0.000	001)					0.02 0.1 1 10 Favours [No Anxiety] Favours [Anxiety

Fig. 5.

Meta-analysis of suicide attempts in bipolar disorder based on the presence of comorbid anxiety disorder. CI = confidence interval.

	Any Substan	ce Use	No Substan	ce Use	Odds Ratio		Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% Cl	
Azorin et al. 2009(67)	260	785	122	305	7.1%	0.74 [0.57, 0.98]		
Dalton et al. 2003(69)	42	136	44	200	6.4%	1.58 [0.97, 2.60]		
Finseth et al. 2012(71)	35	56	58	150	5.8%	2.64 [1.40, 4.98]	200	
Gazalle et al. 2007(65)	5	8	77	161	3.0%	1.82 [0.42, 7.86]		
Goldberg et al. 1998(64)	26	51	37	133	5.7%	2.70 [1.38, 5.26]		
Grunebaum et al. 2006(63)	35	45	29	51	4.8%	2.66 [1.08, 6.50]		
everich et al. 2003(72)	96	251	123	397	6.9%	1.38 [0.99, 1.92]		
Malloy-Diniz et al. 2009(48)	12	18	6	21	3.2%	5.00 [1.28, 19.53]	· · · · · · · · · · · · · · · · · · ·	
Oquendo et al. 2010(46)	222	881	113	762	7.2%	1.93 [1.51, 2.49]	-	
Perroud et al. 2007(60)	29	58	135	348	6.1%	1.58 [0.90, 2.76]	⊢	
Potash et al. 2000(45)	38	99	33	152	6.1%	2.25 [1.28, 3.93]		
Sanchez-Gistau et al 2009(58)	12	57	100	333	5.6%	0.62 [0.32, 1.22]		
Simon GE et al. 2007(15)	603	4935	1116	27799	7.4%	3.33 [3.00, 3.69]		
Simon NM et al. 2007(24)	24	51	22	69	5.4%	1.90 [0.90, 4.01]	<u> </u>	
Sublette et al. 2009(44)	38	45	38	57	4.5%	2.71 [1.02, 7.21]		
Swann et al. 2005(57)	16	24	8	21	3.6%	3.25 [0.96, 11.04]		
Tondo et al. 1999(53)	30	142	35	362	6.2%	2.50 [1.47, 4.26]		
Valtonen et al. 2006(56)	9	84	18	92	4.9%	0.49 [0.21, 1.17]		
Fotal (95% CI)		7726		31413	100.0%	1.81 [1.31, 2.50]	•	
Total events	1532		2114				545.0	
Heterogeneity: Tau ² = 0.37; Ch	i ² = 159.90, df	= 17 (P <	0.00001); l ²	= 89%		t	02 0.1 1 10 50	
est for overall effect: Z = 3.57	(P = 0.0004)	100	390			0.0	Favours [No Use] Favours [Substance]	

Fig. 6.

Meta-analysis of suicide attempts in bipolar disorder based on the presence of any substance use disorder. CI = confidence interval.

	Alcohol	Use	No Alcoho	olUse		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Azorin et al. 2009(67)	194	572	188	518	10.4%	0.90 [0.70, 1.16]	
Bellivier et al.2011(68)	229	553	434	1636	11.0%	1.96 [1.60, 2.39]	-
Cassidy 2011(43)	23	57	33	100	5.3%	1.37 [0.70, 2.69]	
Dalton et al. 2003(69)	27	98	59	238	6.7%	1.15 [0.68, 1.96]	
Dennehy et al. 2011(40)	88	1719	86	2641	9.7%	1.60 [1.18, 2.17]	
Gazalle et al. 2007(65)	1	3	81	166	0.7%	0.52 [0.05, 5.90]	
Leverich et al. 2003(72)	59	155	63	225		1.58 [1.02, 2.44]	
Malloy-Diniz et al. 2009(48)	9	14	9	25	1.9%	3.20 [0.82, 12.52]	
Oquendo et al. 2010(46)	222	881	113	762	10.4%	1.93 [1.51, 2.49]	-
Parmentier et al.2012(61)	75	127	205	525	8.5%	2.25 [1.52, 3.34]	
Perroud et al. 2007(60)	20	37	143	369		1.86 [0.94, 3.67]	
Potash et al. 2000(45)	38	99	33	152	6.4%	2.25 [1.28, 3.93]	
Sublette et al. 2009(44)	57	69	48	69	4.2%	2.08 [0.93, 4.66]	
Swann et al. 2005(57)	16	24	8	21	2.3%	3.25 [0.96, 11.04]	· · · · ·
Tondo et al. 1999(53)	14	74	51	430	5.5%	1.73 [0.90, 3.33]	
Valtonen et al. 2006(56)	9	84	18	92	3.9%	0.49 [0.21, 1.17]	
Total (95% CI)		4566		7969	100.0%	1.60 [1.31, 1.97]	*
Total events	1081		1572				
Heterogeneity: Tau ² = 0.09;	Chi² = 43.	31, df =	15 (P = 0.0	001); l ² =	65%		
Test for overall effect: Z = 4.			67	88 -			0.02 0.1 1 10 5 Favours [No Alcohol] Favours [Alcohol Us

	Illicit Substan	ce Use	No Illicit Substan	ice Use		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Azorin et al. 2009(67)	66	213	316	877	13.1%	0.80 [0.58, 1.10]	
Bellivier et al.2011(68)	191	431	472	2057	14.1%	2.67 [2.15, 3.32]	-
Dalton et al. 2003(69)	15	38	71	298	8.9%	2.09 [1.03, 4.21]	
Dennehy et al. 2011(40)	63	1188	111	3172	13.1%	1.54 [1.13, 2.12]	
Gazalle et al. 2007(65)	4	5	78	164	1.9%	4.41 [0.48, 40.31]	
Leverich et al. 2003(72)	37	96	85	295	11.3%	1.55 [0.96, 2.51]	
Malloy-Diniz et al. 2009(48)	3	4	15	35	1.7%	4.00 [0.38, 42.37]	· · · · · · · · · · · · · · · · · · ·
Parmentier et al.2012(61)	41	79	239	573	11.5%	1.51 [0.94, 2.42]	
Perroud et al. 2007(60)	9	21	155	385	7.2%	1.11 [0.46, 2.70]	
Sublette et al. 2009(44)	50	59	55	79	7.5%	2.42 [1.03, 5.71]	
Tondo et al. 1999(53)	16	68	49	436	9.6%	2.43 [1.29, 4.58]	
Total (95% CI)		2202		8371	100.0%	1.72 [1.23, 2.39]	•
Total events	495		1646				S
Heterogeneity: Tau ² = 0.19;	Chi ² = 43.01, df	= 10 (P <	0.00001); l ² = 77%				
Test for overall effect: Z = 3.	20 (P = 0.001)		00				0.02 0.1 1 10 50 Favours [No Use] Favours [Illicit Use

Fig. 7.

Meta-analyses of suicide attempts in bipolar disorder based on the presence of alcohol use disorder or any illicit substance use disorder. CI = confidence interval.

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	Cluste	r B	No Clus	ter B		Odds Ratio	Odds Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixe	ed, 95% CI	
Grunebaum et al. 2006(63)	28	35	36	61	8.2%	2.78 [1.05, 7.35]			
Leverich et al. 2003(72)	118	288	79	304	70.7%	1.98 [1.40, 2.80]			
Malloy-Diniz et al. 2009(48)	7	8	11	31	0.9%	12.73 [1.38, 117.27]			
Sanchez-Gistau et al 2009(58)) 22	40	90	350	12.9%	3.53 [1.81, 6.88]			
Valtonen et al. 2006(56)	20	79	7	97	7.3%	4.36 [1.73, 10.95]			
Total (95% CI)		450		843	100.0%	2.51 [1.91, 3.31]	t i i i i i i i i i i i i i i i i i i i	•	
Total events	195		223					(539)	
Heterogeneity: Chi2 = 6.29, df	= 4 (P =)	0.18); l ^a	² = 36%						
Test for overall effect: Z = 6.55	6 (P < 0.0	0001)					0.01 0.1 Favours [No Cluster B]	1 10 100 Favours [Cluster B]	

Fig. 8.

Meta-analysis of suicide attempts in bipolar disorder based on the presence of a comorbid cluster B/borderline personality disorder. CI = confidence interval.

	Family History of	Suicide	NoFamilyH	listory	Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Cassidy 2011(43)	11	21	45	136	6.6%	2.22 [0.88, 5.63]	
Dalton et al. 2003(69)	33	108	53	228	11.9%	1.45 [0.87, 2.42]	
Dennehy et al. 2011(40)	16	284	158	4076	11.7%	1.48 [0.87, 2.51]	
Engstrom et al. 2004(70)	7	21	16	79	5.5%	1.97 [0.68, 5.68]	
Finseth et al. 2012(71)	5	12	88	194	4.7%	0.86 [0.26, 2.81]	
Galfalvy et al 2006(66)	7	11	29	52	3.9%	1.39 [0.36, 5.33]	
Leverich et al. 2003(72)	96	232	123	416	14.9%	1.68 [1.20, 2.35]	
Parmentier et al.2012(61)	68	162	212	490	14.5%	0.95 [0.66, 1.36]	-
Romero et al. 2007(42)	24	46	81	318	10.1%	3.19 [1.70, 6.00]	
Sanchez-Gistau et al 2009(58)) 29	52	83	338	10.5%	3.87 [2.12, 7.06]	1.
Valtonen et al. 2006(56)	5	30	22	146	5.5%	1.13 [0.39, 3.26]	
Total (95% CI)		979		6473	100.0%	1.69 [1.25, 2.27]	•
Total events	301		910				
Heterogeneity: Tau ² = 0.13; Ch	ni ² = 23.50, df = 10	(P = 0.009)); l ² = 57%				
Test for overall effect: Z = 3.44	(P = 0.0006)						0.02 0.1 1 10 50 Favours [No History] Favours [Family Histor

Fig. 9.

Meta-analysis of suicide attempts in bipolar disorder based on the presence of a family history of suicide. CI = confidence interval.

	Male	s	Fema	les		Odds Ratio	Od	ds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Ra	ndom, 95% Cl
Black et al. 1988(76)	6	266	1	320	1.5%	7.36 [0.88, 61.53]		
Dennehy et al. 2011(40)	6	1819	2	2359	2.5%	3.90 [0.79, 19.35]		
Dutta et al. 2007(77)	7	102	1	133	1.5%	9.73 [1.18, 80.37]		
Kessing et al. 2005(21)	39	4845	63	8341	15.6%	1.07 [0.71, 1.59]		+
Nordentoft et al.2011(12)	97	2571	78	3356	18.4%	1.65 [1.22, 2.23]		-
Osby et al. 2001(8)	345	6578	327	8808	22.4%	1.44 [1.23, 1.68]		
Sani et al. 2011(4)	31	733	26	1032	12.3%	1.71 [1.01, 2.90]		
Sharma & Markar 1994(75)	3	37	5	77	2.8%	1.27 [0.29, 5.63]		
Simon GE et al. 2007(15)	41	11160	32	21200	13.9%	2.44 [1.54, 3.88]		
Tondo et al. 2007(2)	15	398	7	503	6.3%	2.78 [1.12, 6.87]		
Weeke & Vaeth 1986(74)	15	185	2	232	2.8%	10.15 [2.29, 44.96]		
Total (95% CI)		28694		46361	100.0%	1.83 [1.41, 2.39]		•
Total events	605		544					
Heterogeneity: Tau ² = 0.08; C	Chi ² = 22.3	5, df = 1	0 (P = 0.0)	01); l ² =	55%			
Test for overall effect: Z = 4.4	8 (P < 0.0	0001)	76 1	07.3			0.01 0.1 Favours [Female	1 10 100 s] Favours [Males]

Fig. 10.

Meta-analysis of suicide deaths among males and females with bipolar disorder. CI = confidence interval.

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	Family History of Suicide		NoFamilyHistory		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	M-H, Fixed, 95% Cl
Dennehy et al. 2011(40)	0	284	8	4076	11.3%	0.84 [0.05, 14.61	
Sani et al. 2011(4)	9	19	87	269	61.5%	1.88 [0.74, 4.80	j +∎
Tondo et al. 2007(2)	5	111	20	2182	18.8%	5.10 [1.88, 13.85	1
Tsai et al. 2002(73)	7	8	34	74	8.4%	8.24 [0.96, 70.32	i
Total (95% CI)		422		6601	100.0%	2.91 [1.54, 5.48	1 •
Total events	21		149				5. Sec. Sec. Sec. Sec. Sec. Sec. Sec. Sec
Heterogeneity: Chi ² = 3.67	, df = 3 (P = 0.30); l ²	= 18%					
Test for overall effect: Z =							0.01 0.1 1 10 100 Favours [No Fam. History] Favours [Fam. History]

Fig. 11.

Meta-analysis of suicide deaths in bipolar disorder based on the presence of a family history of suicide. BD = bipolar disorder; CI = confidence interval.

Variables associated with suicide attempts

- Female sex
- Younger age of illness onset
- Depressive polarity of first illness episode
- Depressive polarity of current or most recent episode
- Any comorbid anxiety disorder
- Any comorbid substance use disorder
- Comorbid alcohol use disorder
- Any illicit substance use disorder
- Comorbid cluster B / borderline personality disorder
- First-degree family history of suicide

Variables NOT associated with suicide attempts

- Specific bipolar disorder subtype (I or II)
- History of psychosis
- Any cannabis use

Variables associated with suicide deaths

- Male sex
- First-degree family history of suicide

Variables NOT associated with suicide deaths

- History of psychosis
- Any comorbid substance use disorder

Fig. 12.

List of variables that were or were not associated with suicide attempts or suicide deaths in people with bipolar disorder based on meta-analytic results. *Could not be ranked because a continuous measure of age of illness onset was used.