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# "KEEPING BOTH CORRUPTION AND THE SHADOW ECONOMY IN CHECK: THE ROLE OF DECENTRALIZATION"

• Roberto Dell'Anno, (U. Salerno)

• Désirée Teobaldelli, (U. Urbino)

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#### Keeping both Corruption and the Shadow Economy in Check: The Role of Decentralization

Roberto Dell'Anno<sup>†</sup> and Désirée Teobaldelli<sup>‡</sup>

#### Abstract

This paper puts forward a framework for evaluating the effects of governmental decentralization on the shadow economy and corruption. The theoretical analysis demonstrates that decentralization exerts both a direct and an indirect impact on the shadow economy and corruption. Firstly, decentralization helps to mitigate government-induced distortions, thus limiting the extent of corruption and the informal sector in a direct way. Secondly, in more decentralized systems, individuals have the option to avoid corruption by moving to other jurisdictions, rather than going underground. This limits the impact of corruption on the shadow economy and implies that decentralization is also beneficial in an indirect way. As a result, our analysis documents a positive relationship between corruption and the shadow economy; however, this link proves to be lower in decentralized countries. To test these predictions, we developed an empirical analysis based on a cross-country database of 145 countries that includes different indexes of decentralization, corruption and shadow economy. The empirical evidence is consistent with the theory. Results are robust and significant even after controlling for the endogeneity bias.

#### JEL-Codes: O17, H77, H11, D73.

Keywords: Shadow economy, federalism, decentralization, corruption.

<sup>&</sup>lt;sup>†</sup> Roberto Dell'Anno, Department of Economics and Statistics, University of Salerno, Via Ponte Don Melillo, 84084 Fisciano (Sa), Italy; e-mail: rdellanno@unisa.it.

<sup>&</sup>lt;sup>‡</sup> Désirée Teobaldelli, Department of Law, University of Urbino "Carlo Bo", Via Matteotti 1, 61029 Urbino, Italy; Phone-No: +39-0722303208; Fax-No: +39-07222955; e-mail: desiree.teobaldelli@uniurb.it. (corresponding author)

#### 1. Introduction

This paper intends to cast light on the relationship between decentralization, corruption, and the shadow economy. Corruption and shadow economies are pervasive and significant around the world and are widely believed to constitute a major obstacle to economic and social development. At the same time, they are two related phenomena that prove to affect and reinforce each other, as countries characterized by high levels of corruption also exhibit larger informal sectors. This is probably due to the fact that they share common roots: they are both illegal, are deeply rooted in cultural and social attitudes, and represent a consequence of inefficient and low-quality governments.

For these reasons, a more in-depth understanding of the causes and the channels of the relationship between corruption and shadow economies deserve particular attention, especially in relation to the debate over institutional design, such as, for example, the optimal degree of decentralization.

The existing literature on related topics follows three main directions. A number of studies concentrate on the relationship between the decentralization of government activities and corruption (e.g., Treisman, 2000; Fan et al., 2009; Enikolopov and Zhuravskaya, 2007), while other analyses focus on the relationship between corruption and the shadow economy (e.g., Friedman et al., 2000; Dreher and Schneider, 2010; Buehn and Schneider, 2012a). Moreover, some more recent works investigate the effect of decentralization on the size of the informal sector, finding a negative correlation (e.g., Torgler et al., 2010; Teobaldelli, 2011).

Our work attempts to contribute to this debate by assuming a different perspective. We aim to understand whether decentralization may help keeping both corruption and the shadow economy in check and, at the same time, we want to identify how decentralization affects the relationship between corruption and the shadow economy.

To our knowledge, the literature that explicitly addresses this issue is scant. A previous attempt in this respect is provided by Alexeev and Habodaszova (2012), who examine the implications of decentralization for the incentives of local governments to provide productivity-enhancing local public goods and extort bribes from local entrepreneurs. They show that locally raised tax revenues help limit the size of the informal sector, while corruption—measured by the size of bribes that local officials extort for issuing licenses—may increase or decrease, depending on the extent to which public goods are capable of enhancing the entrepreneur's productivity. Echazu and Bose (2008) study the impact of a centralized bureaucracy on corruption, taking into account economies with formal and informal sectors. The authors demonstrate that when corrupt officials are active in both sectors, bureaucratic centralization is advantageous only if restricted to the formal sector, since cross-sector centralization can lead to higher corruption and lower welfare. They conclude that the shadow economy may cause adverse effects on bribes and welfare, depending on the organization of bureaucracy and the productivity of the informal sector.

In line with this field of research, we try to advance and improve upon the literature in two ways. In terms of theory, we develop a model that provides an explanation of the transmission channels through which decentralization may influence, both directly and indirectly, the size of the informal sector as well as corruption. Our analysis indicates that the link between the shadow economy and corruption is higher in centralized systems than in decentralized ones. In a unified country, individuals can avoid inefficient business regulations and taxes either by exiting from the official sector and going underground, or by bribing public officials, when this is possible. This implies that the level of corruption should be closely (and positively) linked to the size of the shadow economy. In a decentralized country, the competition among jurisdictions and the mobility of the agents might generate two kinds of effects on both corruption and the shadow economy. A first effect lies in the improvement of policies that leads directly to a reduction of both corruption and the shadow economy. A second effect proves to be indirect and relates to the fact that producers may now avoid the consequences of corruption by moving to other jurisdictions, that is, they do not necessarily need to go underground. This implies that in a federal system, a higher degree of corruption exerts a lower impact on the size of the shadow economy relative to a centralized one, because some of the agents will prefer moving to other jurisdictions and remain in the formal sector rather than going underground. As a result, the impact of corruption on the shadow economy is expected to be larger in centralized political systems relative to decentralized ones.

To test these predictions, we developed an empirical analysis based on a cross-country database of 145 countries that includes different indexes of decentralization, corruption and shadow economy.

The empirical evidence is consistent with the theory. We find that decentralized countries have smaller informal sectors than centralized ones and the difference between the sizes of the unofficial economy between the two institutional settings is important (on average, this is about four percent of the official GDP, ceteris paribus). Moreover, we find a larger effect of corruption on the shadow economy in centralized states relative to decentralized ones. Results are robust and significant even after controlling for the endogeneity bias.

The paper is organized as follows. Section 2 reviews the related theoretical and empirical literature. Section 3 presents our theoretical approach and develops our hypotheses. Section 4 describes the empirical methodology, database and the estimation results. Section 5 concludes with a summary and discussion of the main results.

#### 2. Related Literature

#### 2.1 Previous studies on decentralization and corruption

A number of studies concentrated on the relationship between the decentralization of government activities and corruption, intended as the abuse of public power for private gains through rent extraction.

Theories suggested two alternative perspectives on why the structure of government institutions and of the political process may have an impact on the level of corruption. On the one hand, the classical channels through which decentralization exerts its beneficial effects, in terms of higher government accountability and efficiency, are competition among the different levels of government for mobile resources (Tiebout, 1956; Oates, 1999) and informational advantages (Hayek, 1948; Oates, 1972). On the other hand, the main drawback of decentralization is its propensity to impede coordination between the different levels of government, which lead to inefficiently high taxes and regulatory burdens.

At the empirical level, several scholars analyze the relationship between decentralization and corruption, reaching conflicting findings. Most studies find a higher degree of decentralization to be associated with lower levels of corruption (Shleifer and Visny, 1993; Fisman and Gatti, 2002a,b; Arikan, 2004; Lessmann and Markwardt, 2010), yet other studies reach the opposite conclusion (Treisman, 2000; Treisman, 2006; Fan et al. 2009).<sup>1</sup>

Treisman (2000) analyzes the determinants of corruption and finds that federal countries are associated with higher levels of corruption, controlling for the level of economic development. The author interprets this result, arguing that "in unitary states more effective hierarchies of control enable central officials to limit the extraction of sub national officials to more reasonable levels" (p. 441). He concludes that in countries characterized by low levels of development that are more exposed to corruption, the decentralization of political power may be problematic.

Fan et al. (2009) advance this analysis by using a new dataset constructed by combining a crossnational dataset on different indicators of decentralization and firm-level survey data on corruption collected for 80 countries. They show that a larger number of administrative or governmental tiers and a larger fraction of local bureaucratic personnel are related to a greater occurrence of reported corruption. However, greater sub national revenues are linked to lower corruption. The authors conclude that as government structures become more complex, the possibility of uncoordinated rent-seeking increases, whereas providing local governments with more autonomy in the administration of revenues helps to reduce the public officials' incentive to accept bribes.

Fisman and Gatti (2002a) analyze the cross-country relation between decentralization and corruption and find that fiscal decentralization in government expenditure is associated with lower corruption. Their empirical results are strong and robust, even controlling, for a wide number of control variables as well as endogeneity bias, thus providing some support for theories of decentralization that highlight its benefits.

Taking a different perspective on the same topic, Fisman and Gatti (2002b) investigate whether distinct forms of decentralization differently affect corruption. In particular, they focus on the mismatch between revenue generation and expenditure decentralization in U.S. states in order to test the theoretical hypothesis that expenditure decentralization is effective only when combined with the devolution of revenue generation to local governments. The authors find empirical evidence that large federal transfers are related to a higher rate of conviction for bureaucratic corruption.

Arikan (2004) analyzes, both theoretically and empirically, the influence of fiscal decentralization on the level of corruption and shows that a higher degree of decentralization is associated with a lower level of corruption. At the empirical level, this negative relationship is less clear and depends on the explanatory variables used.

Lederman et al. (2005) study the impact of political institutions, including decentralization, on corruption. Their empirical analysis, based on panel data, indicates that political decentralization seems to

<sup>&</sup>lt;sup>1</sup> There are some contributions that analyze the implications of decentralization from a political economy perspective, e.g. Hindriks and Lockwood (2009) study the impact of decentralization on electoral discipline, while Baskaran (2011) investigates the connection between ideology and decentralization and their joint impact on the size of government, finding that decentralization results in a smaller public sector size when the federal government is formed by a right-wing party.

increase corruption, while fiscal decentralization is likely to lower corruption. The authors however suggest that these conclusions need to be investigated more in depth.

Enikolopov and Zhuravskaya (2007) develop an empirical analysis on a sample of 75 developing and transition countries based on a cross-section and panel data for 25 years in order to study the impact of fiscal decentralization on growth, public goods provision, and corruption as a measure of government quality. Their findings suggest that the beneficial effects of decentralization on corruption and economic outcomes crucially depend on the national political party system's strength that favors local politicians' discipline.

Lessmann and Markwardt (2010) examine a cross-section of 64 countries using alternative decentralization and corruption measures in order to evaluate the positive effects of decentralization on corruption. They show that decentralization is effective in counteracting corruption if there is a supervisory body, such as a free press, able to guarantee the monitoring of bureaucrats' behaviors.

The empirical analysis of the effects of decentralization on economic outcomes clearly needs to place problems of the decentralization proxies' choice at its forefront. It has been argued that decentralization is beneficial in terms of accountability and good governance if accompanied by the devolution of decisionmaking powers to local units. Fisman and Gatti (2002a) highlight that it would be useful to have "a set of homogeneous and informative indicators of the extent of decision-making decentralization" in order to develop informative comparative analysis at cross-country level. The issue is also discussed in Rodden (2004) and Stegarescu (2006). An index sometimes used in the literature to this purpose is a dummy variable that reflects whether a country has a political federal structure, based on the Riker's (1964) definition of a federal state, where the constitution guarantees sub national governments the power to autonomously rule and legislate.<sup>2</sup> However, it has been noticed that federalism may be an imperfect measure to the intention pursued, as "there can be both centralized and decentralized federations and, similarly, centralized and decentralized unitary states" (Lijphart 1984: 176), even if federalism and decentralization tend to go hand in hand, especially in developed countries. An alternative measure often employed in the literature is the sub national share of total government expenditures (or revenues), but this measure is not immune from criticism either. Fisman and Gatti (2002a) recognize that there could be a weak correspondence between budgetary items and actual decision making. If the budgets of local governments are mandated from above, then greater decentralization does not correspond to effective expenditures and the revenue-raising power of subnational units. We try to contribute to the debate by proposing a new index of governmental decentralization that reflects the extent of the resources' devolution as well as the transfer of political responsibility to local entities. Table 1 recapitulates this strand of empirical literature.

<sup>&</sup>lt;sup>2</sup>According to Riker's definition, a federal state implies "(1) [at least] two levels of government rule the same land and people, (2) each level has at least one area of action in which it is autonomous, and (3) there is some guarantee (even though merely a statement in the constitution) of the autonomy of each government in its own sphere."

#### 2.2 Previous studies on corruption and the shadow economy

A growing body of literature analyzes the relationship between corruption and the shadow economy (and vice versa), trying to assess the complementarity and substitution effects, that is, whether the shadow economy and corruption are positively or negatively related to each other. The underlying idea is that in countries with a large informal sector, individuals bribe public officials in order to avoid taxation and regulatory burden. As a result, the shadow economy supports corruption since bureaucrats are induced to abuse their position because of the firm's propensity to pay bribes in order to hide their economic activities. At the same time, a pervasive corruption may act as an additional tax in the official sector, leading individuals to go underground in order to avoid government-induced distortions. This, in turn, may increase the size of the informal sector and trigger a detrimental vicious circle in which the shadow economy and corruption foster each other, making it difficult to assess the direction of the causal link between the two related phenomena.<sup>3</sup>

Both theoretical and empirical analyses have been put forward in trying to disentangle the interaction between the shadow economy and corruption. Johnson et al. (1998; 1999) develop a full-employment model in which individuals are employed either in the official or in the shadow economy in order to analyze the relationship between taxation and the provision of productive public goods. They derive implications about the effect of taxes and regulatory burden on the size of the informal sector and economic growth and highlight two different equilibriums that may characterize the economies. In the first case, tax distortions and the regulatory burden are low, public revenues are high, and the provision of public services is efficient, which leads to a small size of the informal sector. In the opposite scenario, the burden of taxes and regulation in the context of government-induced distortions induces a low quality provision of public services; as a consequence, people escape the official sector's inefficiencies by going underground. In this case, especially the financing of market-supporting institutions, including regulatory agencies and an honest public administration play a key role in limiting the informal sector development. Following the predictions of the theoretical model, the authors analyze the relationship between regulatory discretion and the unofficial economy by developing a cross-section analysis based on a sample of 49 countries from OECD, Latin America, and transition economies during the mid-1990s. They find evidence that countries with an intensive regulatory burden are characterized by higher shares of the shadow economy, in a context of great bureaucracy inefficiencies and discretion in the implementation of regulatory rules. Corruption—as perceived by business—and a rule of law that is weak and ineffective in protecting economic activity against the public officials' abuses of power tend to be associated with larger unofficial economies.

In line with these findings, Friedman et al. (2000) evaluate the determinants of underground activity in 169 countries using data from the 1990s and find that firms operate unofficially not to avoid taxation,

<sup>&</sup>lt;sup>3</sup> For a comprehensive analysis of the determinants of the informal sector, see Schneider 2005.

but to mitigate regulatory burden and corruption. Corruption, bureaucracy and a weak legal system are systematically associated with a higher level of the informal sector.

Choi and Thum (2005) present a model in which the entrepreneur's option to move underground constrains the corrupt bureaucrat's ability to ask for bribes. They argue that the existence of the shadow economy helps to mitigate the public institutions' failures to support an efficient official economy and reduces corruption.

Dreher et al. (2009) reach similar results. They develop a theoretical model that explains the relationship among institutional quality, corruption, and the shadow economy. They show that corruption and the shadow economy are substitutes as the presence of the informal sector limits the ability of bureaucrats to extract bribes from economic activities. Moreover, institutional quality is likely to reduce both the shadow economy and corruption. The predictions of the model are tested and confirmed by using a structural equation model that treats corruption and the shadow economy as latent variables in a sample of 18 OECD countries.

Dreher and Schneider (2010) develop an empirical analysis based on both a cross-section of 120 countries and a panel of 70 countries for the period 1994–2002. They find that corruption and the shadow economy are substitutes in high-income countries while they are complements in low income countries. In high income countries, characterized by a good rule of law, firms have the option to bring corrupted public officials to the court, when asked for bribes. In this case, corruption is more likely to take place in order to facilitate the official economic activity and obtain benefits from the public sector. In low income countries, on the contrary, the shadow economy and corruption are likely to reinforce each other, since corruption is employed by firms to keep their activity underground. Consequently, in these countries, it is natural to observe a positive relationship between the shadow economy and corruption, while in high income countries, the opposite holds true. However, the authors specify that this mixed result depends on the indicators chosen to measure corruption as well as on how regressions are specified.

Buehn and Schneider (2012b) test the relationship between the shadow economy and corruption using a structural equation model that treats the shadow economy and corruption as latent variables. They find a positive relationship between the shadow economy and corruption and show that the causal effect of the shadow economy on corruption is stronger than the effect of corruption on the shadow economy. Table 2 summarizes this empirical literature.

Table 2. Previous empirical studies on shadow economy and corruption

#### 3. The Model

#### 3.1 The Framework

We consider an economy with a unique final good that can be produced by two sectors, the formal and the informal one. Each agent *i* is a consumer–producer that can produce in the formal sector using a Cobb–Douglas production function with constant returns to scale in labor  $x_{i,f}$  and in the quantity of per-

capita public goods and services g, that is,

$$y_{i,f} = x_{i,f}^{\alpha} g^{1-\alpha} \tag{1}$$

where  $0 < \alpha < 1$ . Production in the informal sector does not require the input provided by the public sector and the production function of this sector is

$$y_{i,s} = A x_{i,s}^{\alpha} \tag{2}$$

where  $x_{i,s}$  is the amount of labor employed in shadow activities and A is a positive constant related to the efficiency of underground production.<sup>4</sup>

Each agent *i* supplies inelastically 1 unit of labor,  $x_{i,f} + x_{i,s} = 1$ , and chooses his optimal allocation between the two sectors, maximizing the income produced net of taxes. We assume that income in the formal sector is perfectly observable by the tax authorities and can be taxed at a constant rate  $t \in [0,1]$ , while production in the unofficial economy is completely unobservable and, therefore, cannot be taxed. The economy consists of a continuum of individuals of measure 1.

We also assume that each individual is controlled by the bureaucracy when it comes to the respect of regulations and that the bureaucrat can impose a sanction equal to a fraction  $a \in [0,1]$  of the net income in the formal sector even when the producer is respecting the rules, that is, the net income in the formal sector is equal to  $(1-a)(1-t)y_{i,f}$ . The inefficiency of the bureaucratic system and regulations implies that producers will have to incur the payment of the sanction imposed with a certain probability. We assume that this probability is a function of the inefficiency of the bureaucratic system, which is denoted by  $\theta \in [0,1]$ , so that with probability  $q(\theta)$ , with  $\partial q(\theta)/\partial \theta > 0$ , the producer has to pay the sanction, while with probability  $1-q(\theta)$ , the sanction is cancelled. The individual can avoid the risk of paying the sanction by bribing the bureaucrat. The bribe that has to be paid is set by the bureaucrat by making a takeit-or-leave-it offer to the individual and is denoted by *B*. This assumption aims at capturing the idea that higher levels of inefficiencies in the organization of bureaucracy and regulations favor corruption as widely documented in the literature. To simplify the analysis and without loss of generality, we assume that  $q(\theta)$  is a linear function, that is,  $q(\theta) = q\theta$ , with  $q \in [0,1]$ .

The revenues of the public sector are spent on the provision of the productive public services and the government budget constraint can be written as

$$G = \omega \int_{i=0}^{n} t y_{i,f} di$$
(3)

where G is the total amount of public goods (so that g=G/n), n is the number of individuals, and  $\omega$  is a parameter measuring the efficiency of the public sector in producing the public goods and services. We assume that higher inefficiency  $\theta$  of the bureaucracy and regulations reduces the efficiency of the public

<sup>&</sup>lt;sup>4</sup> The technology that characterized the informal production in our framework does not require public services as a productive input. Here, we want to capture the idea that because of their illegal status, informal agents do not benefit from government-provided goods and services that can facilitate the production by allowing them full, enforceable property rights over their capital and output. For example, Loayza (1996) refers to the protection of the police and the legal and judicial system from crimes committed against the property. Moreover, informal producers are unable to take full advantage of other public services such as social welfare, skill-training programs, and government-sponsored credit facilities.

sector in this production, so that  $\omega \equiv \omega(\theta)$ , with  $\omega_{\theta} \equiv \partial \omega / \partial \theta < 0$ ,  $\omega_{\theta\theta} \equiv \partial^2 \omega / \partial \theta^2 < 0$ .

The degree of inefficiency  $\theta$  and the tax rate *t* are chosen by an incumbent politician whose monetary rent *R* corresponds to a fraction  $\beta \in [0,1]$  of the bribes of bureaucrats. In choosing these two policy variables, the politician maximizes a weighted average of its own utility and the utility of the median voter (e.g., Panizza, 1999). The objective function of the politician can be written as

$$V = \lambda R + (1 - \lambda) y_i \tag{4}$$

where  $y_i$  is the net income of the median voter. The parameter  $\lambda$  can be interpreted as an inverse measure of the democratic quality of the country.

We will consider two organizational models of society—a centralized state in which policy decisions  $(t, \theta)$  are made at the centralized level and a federal state within which policy decisions are the responsibility of each jurisdiction. In order to simplify the analysis and exposition, we assume the existence of two identical jurisdictions. Each individual can move freely among the two jurisdictions by bearing a cost  $C_i$ . Individuals are heterogeneous with respect to moving costs and the probability distribution function is denoted by  $f(C_i)$ .

#### 3.2 Characterization of the equilibrium

We now solve the model by determining the equilibrium policy in a unitary state. Then, we characterize the equilibrium policy in a federal system and discuss the implications for the shadow economy and corruption.

#### 3.2.1 The equilibrium in a centralized state

The level of bribe set by the bureaucrat is such that each producer is indifferent to choosing between bribing the bureaucrat and accepting the gamble represented by the sanction, that is,

$$q\theta(1-a)(1-t)y_f + (1-q\theta)(1-t)y_f = (1-t)y_f - B$$
(5)

where the left hand side represents the expected income when the individual does not bribe the bureaucrat while the right hand side is the income when the latter is bribed. In the first case, the citizen pays a fine  $a(1-t)y_f$  with probability  $q\theta$  and nothing is extorted with the complementary probability. From (5), it follows that the level of bribe paid to the bureaucrat is

$$B = aq\theta(1-t)y_f \tag{6}$$

which implies a monetary rent  $R = \beta B$  for the politician equal to

$$R = \beta a q \theta (1-t) y_f \tag{7}$$

where we have used the fact that all agents are identical and have a mass equal to 1.

From (1) and (3) and the fact that all agents are identical, and therefore  $x_i = x$  for all *i*, it follows that the government budget constraint can be rewritten as

$$g = (\omega t)^{1/\alpha} (1-x). \tag{8}$$

Taking into account (1), (2) and (6), the disposable income of agent *i* is

$$y_i = (1 - aq\theta)(1 - t)(1 - x_i)^{\alpha} g^{1 - \alpha} + Ax_i^{\alpha}$$
(9)

where  $x_i$  denotes the amount worked by agent *i* in the informal sector.

For any given level of taxation t and inefficiency  $\theta$  of bureaucracy and regulations, the individual chooses an allocation of labor in the two sectors that maximizes  $y_i$ . The following lemma reports the solution to this problem.

### **Lemma 1.** The optimal amount of labor employed in the informal sector by each individual is $x(t,\theta) = A^{1/(1-\alpha)} (1 - aq\theta)^{-1/(1-\alpha)} (1 - t)^{-1/(1-\alpha)} (\omega t)^{-1/\alpha}.$ (10)

The amount worked in the informal sector is monotonically decreasing in  $\theta$  and is minimum at  $t=1-\alpha$ ; it is decreasing in t when  $t < 1-\alpha$ , and vice versa. x=1 for t=0 and t=1.

**Proof**: From (9), we obtain that the first order condition is

$$\partial y_i / \partial x_i = -\alpha (1 - aq\theta)(1 - t)(1 - x_i)^{\alpha - 1} g^{1 - \alpha} + \alpha x_i^{\alpha - 1} = 0$$

Substituting the government budget constraint (8) and taking into account that  $x_i = x$  for all *i*, we obtain (10).  $x(t, \theta)$  is monotonically decreasing in  $\theta$  since it is equal to

$$\frac{\partial x(t,\theta)}{\partial \theta} = -x \left[ \frac{aq}{(1-\alpha)(1-aq\theta)} - \frac{\omega_{\theta}}{\alpha \omega} \right]$$
(11)

and  $\omega_{\theta} < 0$ . By deriving  $x(t, \theta)$  with respect to t, we obtain

$$\frac{\partial x(t,\theta)}{\partial t} = -x \frac{1-\alpha-t}{\alpha t(1-\alpha)(1-t)}$$
(12)

which means that the function has a minimum at  $t = 1 - \alpha$ . It is decreasing for  $t < 1 - \alpha$  and vice versa.  $x(t, \theta) = 1$  at t = 0 and t = 1 is immediate from (9).

The optimal allocation of labor is such that the marginal revenues in the two sectors are equalized. A higher inefficiency of the bureaucracy and regulations increases the amount of labor employed in the shadow economy through two channels. On the one hand, it reduces the productivity in the formal sector by increasing corruption by bureaucrats and politicians and, on the other hand, it reduces the level of public goods and services for any given level of taxation. Instead, a higher level of taxation has nonmonotonic effects on the shadow economy. It reduces the labor in the informal sector when it is relatively low and increases it at higher levels. In fact, a higher tax rate reduces the income in the formal sector, but it also allows more provision of productive public services and, therefore, a higher marginal productivity of labor in the formal sector. At low levels of taxation, the latter effect dominates and vice versa.

We now determine the optimal policy choice  $(t, \theta)$  of the politician. As he maximizes a convex combination of its rent and the utility of the median voter, we first consider the extreme cases and focus on

the values of  $(t, \theta)$  that maximize the R and  $y_i$ .

From (9),  $y_i$  is monotonically decreasing in  $\theta$ , and therefore all citizens (and therefore the median voter) prefer  $\theta = \theta_m \equiv 0$ . Using the government budget constraint (8), we obtain that the disposable income of each citizen is

$$y_i = (1 - aq\theta)(1 - t)(1 - x_i)(\omega t)^{(1 - \alpha)/\alpha} + Ax_i^{\alpha}.$$
 (13)

Using the envelope theorem, it follows that

$$\frac{\partial y_i}{\partial t} = (1 - aq\theta)(1 - x_i)(\omega t)^{(1 - \alpha)/\alpha} \frac{1 - \alpha - t}{\alpha t} = 0$$
(14)

which implies that  $y_i$  is maximized at  $t = t_m \equiv 1 - \alpha$ .

Using the expression (1) of the production function in the formal sector and the government budget constraint (8), the monetary rent of the politician can be rewritten as

$$R = \beta a q \theta (1-t) (\omega t)^{(1-\alpha)/\alpha} (1-x).$$
(15)

It can be shown (see the proof of Lemma 2) that the values of  $(t, \theta)$  maximizing the politician's rent are  $t_p = t_m \equiv 1 - \alpha$  and  $0 < \theta_p < 1$ . The following lemma summarizes these results.

**Lemma 2.** The optimal policy for the citizens (median voter) is setting the tax rate at level  $t = t_m \equiv 1 - \alpha$ and setting the most efficient bureaucracy and regulations corresponding to  $\theta = \theta_m \equiv 0$ . The policy that maximizes the politician's rent R involves the same level of taxation but a higher level of inefficiency of bureaucracy and regulations  $0 < \theta_p \le 1$ .

**Proof**: See the Appendix.

The intuition behind the result of Lemma 2 is the following. The optimal level of taxation that maximizes the citizens' net income corresponds to setting  $t=1-\alpha$  and the efficiency of the state apparatus at maximal level ( $\theta = 0$ ). The politician also prefers a tax rate  $t=1-\alpha$ , because the monetary rent he can exploit through corruption is related to the production in the formal sector, and the level of taxation  $t=1-\alpha$  maximizes the net income in the formal sector. However, the politician has a different preference for the level of efficiency of the state apparatus, whose optimal value corresponds in this case to the level which balances the two opposite effects that  $\theta$  exerts on R. On the one hand, a higher  $\theta$  allows the politician to appropriate a higher fraction of productive public services provided. This latter effect increases the incentive to supply more labor in the shadow economy, which reduces both the production in the formal sector and the amount of politician's monetary rents deriving from corruption. The presence of these opposite effects leads the politician to prefer intermediate values of  $\theta$ .

The following proposition characterizes the optimal policy for the politician in a centralized system.

Proposition 1. The optimal policy for the politician in a centralized state is setting taxation at level

 $t_C = t_m \equiv 1 - \alpha$  and the efficiency of the state apparatus and regulations at a level  $\theta_C \equiv \theta_C(\lambda)$ , where  $\theta_C$  is increasing in  $\lambda$ ,  $\partial \theta_C / \partial \lambda \ge 0$ , that is,  $\theta_C$  is decreasing in the quality of democracy, with  $\theta_C(0) = 0$  and  $\theta_C(1) = \theta_p$ .

**Proof**: The politician's objective function V in (4) is a convex combination of R in (15) and  $y_i$  in (13). Both are maximized at  $t = 1 - \alpha$  for any  $\theta \in [0,1]$ , so that V is also maximized at  $t = 1 - \alpha$ . The optimal level  $\theta = \theta_C(\lambda)$  follows from the fact that  $y_i$  is monotonically decreasing in  $\theta$  and maximized at  $\theta = 0$ , while R is single-peaked and maximized at  $\theta = \theta_p$ .

#### 3.2.2 The equilibrium in a federal state

We now move to characterize the policy  $(t, \theta)$  in a federal system. In this case, the optimal policy of the median voter is unchanged. The rent of the politician in jurisdiction 1 is

$$R_{1} = \beta a q \theta_{1} (1 - t_{1}) (\omega_{1} t_{1})^{(1 - \alpha)/\alpha} (1 - x_{1}) n_{1}$$
(16)

where  $\omega_1 \equiv \omega(\theta_1)$ ,  $x_1 \equiv x(t_1, \theta_1)$  and  $n_1 = (1/2) + h(t_1, \theta_1, t_2, \theta_2)$  is the size of individuals in jurisdiction 1 and  $h(t_1, \theta_1, t_2, \theta_2)$  is the net flow of individuals from jurisdiction 2 to jurisdiction 1. Note that the expression for  $R_1$  has the same expression as the one at the centralized level reported in (15), with the difference that the size of the citizens is now  $n_1$  rather than 1. Similarly, the rent of the politician in jurisdiction 2 is

$$R_2 = \beta a q \theta_2 (1 - t_2) (\omega_2 t_2)^{(1 - \alpha)/\alpha} (1 - x_2) n_2$$
(17)

where  $\omega_2 \equiv \omega(\theta_2)$ ,  $x_2 \equiv x(t_2, \theta_2)$  and  $n_2 = (1/2) - h(t_1, \theta_1, t_2, \theta_2)$ .

Assume that the policy vector  $(t_1, \theta_1, t_2, \theta_2)$  is such that  $y_f(t_1, \theta_1) > y_f(t_2, \theta_2)$ , so that  $h(\cdot) > 0$  and there is a net flow of individuals from jurisdiction 2 to jurisdiction 1. All individuals of jurisdiction 2 with migration cost  $C_i$  such that  $y_f(t_1, \theta_1) - y_f(t_2, \theta_2) \ge C_i$  will move to jurisdiction 1. Let  $F(C_i)$  denotes the cumulative probability distribution function of  $C_i$ . Then, it follows that

$$\begin{aligned} h(t_1, \theta_1, t_2, \theta_2) &= F(y_f(t_1, \theta_1) - y_f(t_2, \theta_2)) > 0 \quad if \quad y_f(t_1, \theta_1) \ge y_f(t_2, \theta_2), \\ h(t_1, \theta_1, t_2, \theta_2) &= -F(y_f(t_1, \theta_1) - y_f(t_2, \theta_2)) < 0 \quad if \quad y_f(t_1, \theta_1) < y_f(t_2, \theta_2). \end{aligned}$$
(18)

From (18), it can be verified that the function  $h(t_1, \theta_1, t_2, \theta_2)$  has the following properties:  $\partial h(\cdot)/\partial \theta_1 < 0$ ;  $\partial h(\cdot)/\partial \theta_2 > 0$ ;  $\partial h(\cdot)/\partial t_1 < 0$  for  $t_1 < 1 - \alpha$  and vice versa;  $\partial h(\cdot)/\partial t_2 > 0$  for  $t_2 < 1 - \alpha$  and vice versa;  $h(\cdot) = 0$  if  $t_1 = t_2$  and  $\theta_1 = \theta_2$  (as the citizens have no incentives to move when the policy is the same).

The following lemma characterizes the Nash equilibrium policy in a federal state when politicians maximize their rents and agents can move across jurisdictions.

**Lemma 3.** The Nash equilibrium policy in a federal state when politicians maximize their monetary rent R is  $t_1 = t_2 = 1 - \alpha$  and  $\theta_1 = \theta_2 = \theta_r$  where  $0 < \theta_r < \theta_p$ .

**Proof**: See the Appendix. ■

Lemma 3 states that the competition among jurisdictions leads them to improve the efficiency of the state apparatus, that is, to choose a more efficient bureaucracy and regulations (i.e., a lower  $\theta$ ). The next proposition defines the Nash equilibrium policy in a federal state when politicians also take into account the utility of the median voter, and therefore maximize the utility reported in (4).

**Proposition 2.** The Nash equilibrium policy in a federal state is  $t_F = t_m \equiv 1 - \alpha$  and  $\theta_F \equiv \theta_F(\lambda)$ , where  $\theta_F$  is decreasing in  $\lambda$  with  $\theta_F(0) = 0$  and  $\theta_F(1) = \theta_r$ . Moreover,  $\theta_F(\lambda) < \theta_C(\lambda)$  for any  $\lambda \in [0,1]$ . **Proof:** The results are straightforward from the proof of Lemma 3 and Proposition 1.

From Proposition 2, it follows that the tax rate in a centralized and in a federal state are the same. The rationale is that the politicians choose the tax rate maximizing the income in the formal sector from which they can extract rents through bribing. However, the level of inefficiency of the state apparatus is lower in a federal state than in a centralized system. These two facts imply that both the level of corruption and of the shadow economy are lower in a federal system as shown in the following proposition.

**Proposition 3.** The level of the shadow economy and corruption are lower in a federal state than in a centralized country.

Proof: The result on the shadow economy follows from the fact that the labor employed in the informal sector is monotonically decreasing in  $\theta$  (see Lemma 1) and that  $\theta_F(\lambda) < \theta_C(\lambda)$  while  $t_F = t_C$  (see Proposition 2). The result of corruption comes from the fact that the level of corruption *B* in (6) and the monetary rent *R* of the politician are positively related. Hence, from  $R(\theta_F(\lambda)) < R(\theta_C(\lambda))$ , it follows that  $B(\theta_F(\lambda)) < B(\theta_C(\lambda))$ .

As we have clarified above, a higher level of corruption increases the production of the informal sector. However, this effect is likely to be different in a centralized state with respect to a federal system. In particular, an increase in corruption in a federal system may have a lower impact on the shadow economy as long as the individuals of some jurisdictions (those with lower moving costs) find it optimal to move to other jurisdictions. To clarify this point, let us consider an exogenous increase in the level of corruption *B* in a centralized state due to an increase in  $\theta$  or in the fraction of product *a* in the formal sector that can be extorted by a bureaucrat. This leads to an increase in the labor employed in the shadow economy equal to  $\partial x(B)/\partial B$ . Now, let us consider this increase in corruption in a federal state and assume that this is not uniformly distributed in all jurisdictions (e.g., let us consider the extreme case where it is concentrated in jurisdiction 1). As the individuals with lower migration costs find it optimal to move in

other jurisdictions, the same increase in corruption implies an increase in the labor employed in the shadow economy equal to  $\left[1 + \frac{\partial h(\cdot)}{\partial B}\right] \frac{\partial x(B)}{\partial B}$ , where  $\partial h(\cdot)/\partial B < 0$  represents the size of agents moving from jurisdiction 1 to 2. In other words, the same increase in corruption should increase the shadow economy in a centralized state by  $-(\partial h(\cdot)/\partial B)(\partial x(B)/\partial B) > 0$  relative to the increase in a federal state.

**Proposition 4.** The positive effect of corruption on the shadow economy is higher in a centralized country than in a federal state.

**Proof**: In the text. ■

This analysis has led to two main results that we will test in the next section. First, federalism or higher degrees of decentralization should help reduce both corruption and the shadow economy. Second, in federal or more decentralized countries, the impact of corruption on the shadow economy should be lower.

#### 4. Empirical analysis

In this section, we empirically test the two main policy implications of the theoretical model (Propositions 3 and 4). However, there are some practical problems associated with the empirical analysis of the relationships among measures of decentralization, the shadow economy and corruption. Section 4.1 describes these econometric hitches and how we deal with them. Section 5.1 illustrates the empirical methodology and the estimation results.

#### 4.1 Data Description

There are three major practical problems associated with the empirical analysis of the relationships among measures of decentralization, estimates of the shadow economy and indexes of (perceived) corruption. The first corresponds to the measurement errors of the existing proxies of these phenomena. Decentralization is a multidimensional concept that has political, fiscal, cultural and geographical features. Thus, defining a suitable index to compare the degree of decentralization among countries is a particularly hard task. Concerning the shadow economy, as specialists in this field know quite well, all the estimates of the shadow economy are indicative and no one can really claim to be confident of the full reliability of their estimates (Dell'Anno and Schneider, 2008). A similar concern exists for corruption. While there are practical reasons for using the index of perception of the size of corruption, several studies consider this measurement approach as biased by cultural factors (e.g., Mocan 2004; Andvig 2005; Søreide 2005). As a result, the criticism that these indexes reflect the quality of a country's institutions more than the actual size of "misuse of public power for private benefit"<sup>5</sup> is sound.

The second practical problem is the issue of endogeneity. In this study, this issue is due to the aim of explaining the interaction between two related phenomena, as the shadow economy and corruption

<sup>&</sup>lt;sup>5</sup> This is the definition of corruption followed by Dreher and Schneider (2010).

conditioned to the degree of decentralization. For instance, Buehn and Schneider (2012a) find a bidirectional causal effect between the two phenomena. They estimate that the positive effect of the shadow economy on corruption is stronger than the positive effect of corruption on the shadow economy.

A third departure from the assumptions of the classical linear model is the potential bias due to multicollinearity. While highly correlated variables in the right hand side of a regression equation does not reduce the reliability of the model as a whole, the coefficients are less significant with wider confidence intervals; therefore, the interpretation of the regression coefficients is harder. To examine the joint effect of decentralization and corruption on the shadow economy means to include among the regressors of informal economy some index of decentralization together with other control variables.<sup>6</sup> It makes this violation of one of the assumptions of the classical linear model a difficult issue to overcome.

In the following, we deal with these three concerns. On the first issue – measurement errors – several indexes of the perception of corruption, decentralization and the shadow economy are used. Although this approach does not correct the measurement errors, it is a reasonable way to check indirectly the robustness of the estimates across different "inexact" measures.

The consequences of measurement errors and the lack of reliable indexes for corruption, decentralization and the shadow economy is one of the most relevant shortcomings in this strand of literature. As a result, we describe below how we attempt to minimize these inaccuracies in the data.<sup>7</sup>

About the measures of corruption, we standardize the most-used indexes of perceived corruption proposed in literature (Transparency International –  $Corr_i^{CPI}$ ;<sup>8</sup> Worldwide Governance Indicators –  $Corr_i^{WGI}$ ; International Country Risk Guide –  $Corr_i^{ICRG}$ ; World Development Indicators based on Enterprise Surveys –  $Corr_i^{Bribe1}$ ; Enterprise Surveys from World Bank –  $Corr_i^{Bribe2}$ ). These indexes are calculated as the average of annual scores for all available countries around the world. With the exclusion of  $Corr_i^{Bribe2}$ , the averages are aimed to cover the decade from 1999–2007. This purpose, however, encounters several obstacles due to the considerable presence of missing values. Thus, our main concern is to define the time period in order to preserve the sample size. To make it easier for the interpretation of the proxies of corruption, the five indexes are rescaled to take values between 0 and 10, with high scores to mean high corruption. Standardizations of the indexes are as follows:

$$Corr_i^{CPI} = 10 - Av.CPI_i^{1999-2005}$$
(19)

$$Corr_{i}^{WGI} = 10 - \left[\frac{1}{7} \sum_{t=1998}^{2005} 2(WGI_{it} + 2.5)\right]$$
(20)

$$Corr_{i}^{ICGR} = 10 - \frac{10* \left[ Av.CorrICGR_{i}^{1995 \& 2011} - Min_{i} \left( Av.CorrICGR_{i}^{1995 \& 2011} \right) \right]}{Max_{i} \left( Av.CorrICGR_{i}^{1995 \& 2011} \right) - Min_{i} \left( Av.CorrICGR_{i}^{1995 \& 2011} \right)}$$
(21)

<sup>7</sup> Definitions and sources of data are provided in Appendix B.

<sup>&</sup>lt;sup>6</sup> Buehn and Schneider (2012a) estimate shadow economy by the MIMIC approach. Their structural equation includes the size of the government, share of direct taxation, fiscal freedom, business freedom and GDP per capita. It implies that high correlation between the estimates of the shadow economy and these variables is expected.

<sup>&</sup>lt;sup>8</sup> Paldam (2002) investigates and explains the large pattern in the corruption index from Transparency International.

$$Corr_{i}^{Bribe1} = \frac{10*\left[Bribe_{i}^{1999-2007} - Min\left(Av.Bribe_{i}^{1999-2007}\right)\right]}{Max\left(Av.Bribe_{i}^{1999-2007}\right) - Min\left(Av.Bribe_{i}^{1999-2007}\right)}$$
(22)

$$Corr_{i}^{Bribe2} = \frac{10*\left[Av.Corr_{i}^{2005-2010} - Min\left(Av.Corr_{i}^{2005-2010}\right)\right]}{Max\left(Av.Corr_{i}^{2005-2010}\right) - Min\left(Av.Corr_{i}^{2005-2010}\right)}$$
(23)

with:

$$Av.Corr_{i}^{2005-2010} = \begin{pmatrix} Av.CorrGIFT_{i}^{'05-'10} + Av.CorrTAX_{i}^{'05-'10} + Av.CorrGOV_{i}^{'05-'10} + Av.CorrLIC_{i}^{'05-'10} + Av.CorrIMP_{i}^{'05-'10} + Av.CorrCONS_{i}^{'05-'10} + Av.CorrELEC_{i}^{'05-'10} + Av.CorrWAT_{i}^{'05-'10} + Av.CorrPAST_{i}^{'05-'10} + Av.CorrREL_{i}^{'05-'10} + Av.CorrJUS_{i}^{'05-'10} \end{pmatrix} / (11)$$

where the subscript *i* indicates the countries. Sample ranges from a minimum of 88 ( $Corr_i^{Bribe2}$ ) to 145 ( $Corr_i^{CPI}$ ).

Data on shadow economy are extracted by Buehn and Schneider (2012a). They estimate the size of the shadow economy as a percentage of the official GDP by a MIMIC approach (*Shadow<sub>i</sub>*). From their panel of countries, we calculate the average of all available estimates over the period from 1999 to 2007 for the 145 countries of our sample. To check robustness, the correlations with other indexes of informality are estimated (Table 3). These alternative proxies are calculated as the averages, from 2005 to 2010, of "Percent of firms competing against unregistered or informal firms" (*Inf\_Comp*) and "Percent of firms identifying practices of competitors in the informal sector as a major constraint" (*Inf\_Relev*). The source of data for both the variables is the World Bank (Enterprise Surveys).

#### Table 3. Correlation matrix of proxies of the Shadow economy

Table 3 shows that these indicators are positively correlated as expected. Taking into account that  $Inf\_Comp$  and  $Inf\_Rel$  do not measure the same "item" estimated by Buehn and Schneider (2012a)<sup>9</sup>, we are quite confident that this output supports, at least in a rough way, the reliability of Buehn and Schneider (2012a) estimates. Considering that Buehn and Schneider (2012a) estimates have both the widest coverage of countries (145 instead of 108) and they are much closer to the concept of shadow economy relevant for this study, these estimates are thus employed in the following.

The most relevant measurement issue encountered in collecting the dataset is to look for a suitable measure of decentralization. Two main approaches exist in literature to measure the degree of country decentralization. The first kind of measures focuses on fiscal decentralization. They usually estimate a ratio between the expenditure (or revenue) of sub national government and the total government expenditure (or revenue) at the national level. Following this approach, two indexes from the IMF – Government Finance Statistics over the period 1999–2007 (*Exp\_Dec* and *Rev\_Dec*) are calculated. To

<sup>&</sup>lt;sup>9</sup> Authors estimate Shadow economy as percentage of official GDP, where shadow economy includes "All marketbased legal production of goods and services that are deliberately concealed from public authorities for any of the following reasons: (a) to avoid payment of income, value added, or other taxes; (b) to avoid payment of social security contributions; (c) to avoid having to meet certain legal labor market standards, such as minimum wages, maximum working hours, safety standards, etc.; (d) to avoid complying with certain administrative procedures, such as completing statistical questionnaires, or other administrative forms." Buehn and Schneider's (2012a: 141).

check the robustness of our own indicators, we also consider two proxies of fiscal decentralization estimated by the World Bank (Fiscal Decentralization Indicators) calculated for a longer time period (1972–2000). These are sub national expenditures as a percentage of total expenditures (*Subexp*<sub>72-00</sub>) and sub national revenues as a percentage of total revenues (*Subrev*<sub>72-00</sub>).<sup>10</sup> A second set of proxies of decentralization looks at the political and administrative dimensions of the public decision-making process. These are three dichotomous variables (where "1" denotes the federal state) provided by Treisman (2007, here labelled *D1\_Dec*), Persson and Tabellini (2003, *D2\_Dec*) and Fan et al. (2009, *D3\_Dec*).

A preliminary investigation of the reliability of these proxies to measure the same topic and of the signs of correlation among the three analyzed phenomena is possible by examining the correlation matrix (Table 4).

#### Table 4. Correlation matrix and descriptive statistics

From Table 4, two main results emerge. First, looking at the correlations among alternative measures of corruption and decentralization, we find that the indexes of corruption are positively correlated to each other. On the contrary, the proxies of decentralization show a lack of robustness. While  $D1_Dec_i$ ,  $D2\_Dec_i$ , Subexp<sub>72-00</sub> and Subrev<sub>72-00</sub> positively correlate to each other,  $D3\_Dec_i$ , Exp\_Dec<sub>i</sub> and Rev\_Dec<sub>i</sub> are not correlated to the other measures of decentralization. As there are no theoretical reasons behind these differences among indexes, we interpret this result to be a consequence of data limitation.  $D1_Dec_i$ , D2\_Dec<sub>i</sub>, Subexp<sub>72-00</sub> and Subrev<sub>72-00</sub> include between 70 and 140 countries while D3\_Dec<sub>i</sub>, Exp\_Dec<sub>i</sub> and  $Rev_Dec_i$  have a sample size that ranges from 45 to 68. Thus, we hypothesize that different signs of correlations are due to the sample selection bias. Examining the distributions of data across the countries, missing values on the second group for decentralization proxies are concentrated on the less developed countries. Seeing that these countries are frequently characterized by higher levels of corruption, there is thus the suspicion that the orthogonality between  $D3\_Dec_i$ ,  $Exp\_Dec_i$  and  $Rev\_Dec_i$  and the indexes of corruption may be caused by a selection bias. This hypothesis points out a potential drawback for crosscountry analyses that use just one proxy of decentralization. To control this shortcoming and preserve the sample size, a new dichotomous variable based on the seven indexes collected in our dataset is proposed. It takes values of 1 if the country is defined as decentralized. It is calculated as follows:

$$DT\_Dec_{i} = \begin{cases} 1 & \sum_{k=1}^{7} D_{i}^{k} > 0 \\ 0 & if & \sum_{k=1}^{7} D_{i}^{k} < 0 \\ m.v. & \sum_{k=1}^{7} D_{i}^{k} = 0 \end{cases}$$
(24)

where:

<sup>&</sup>lt;sup>10</sup> Dataset is available at: http://www1.worldbank.org/publicsector/decentralization/fiscalindicators.htm#Formulas.

$$D_{i}^{k} = \begin{cases} 1 & K_{i} = 1 \\ -1 & if \quad K_{i} = 0 \text{ for } K = D1\_Dec_{i}, D2\_Dec_{i}, D3\_Dec_{i}; \\ 0 & K_{i} = m.v. \end{cases}$$
$$D_{i}^{k} = \begin{cases} 1 & K_{i} > Median(K) \\ -1 & if \quad K_{i} < Median(K) \text{ for } k = Exp\_Dec_{i}, Rev\_Dec_{i}, Subexp_{72-00}, Subrev_{72-00}. \end{cases}$$

According to the  $DT\_Dec$  index, among the 145 countries of our sample, we count 20 decentralized states,<sup>12</sup> 120 centralized (i.e.,  $D^k$  is a negative value) and 5 countries with  $D^k$  equals to zero as a consequence of missing values for all the seven proxies (i.e., Croatia, Syria and Yemen) or due to contrasting scores among the original indexes (i.e., Colombia and Taiwan).

The second result of correlation analysis refers to the signs of the statistical relationships among the three phenomena (Table 5). In line with the main literature, estimated correlations validate a positive sign between the shadow economy and corruption (e.g., Buehn and Schneider, 2012b), a negative sign between the shadow economy and decentralization (e.g., Teobaldelli, 2011) and a negative sign between corruption indexes and the decentralization dummy (e.g., Fan et al., 2009).

#### Table 5: Correlation matrix of DT\_Dec

With reference to the second statistical difficulty of this analysis – endogeneity – we apply the instrumental variables (IV) approach to address the endogeneity of corruption, the shadow economy and other variables related to economic development. However, the estimation bias associated with ordinary least squares (OLS) estimates of endogenous variable has no severe effects on statistical outputs.

On the third concern common in this strand of literature – multicollinearity – the Coefficient Variance Decomposition (CVD) proposed by Belsley et al.  $(1980)^{13}$  is applied to investigate if and which coefficient estimates are strongly affected by high correlation among regressors. If the degree of collinearity is severe, we apply the main conventional methods proposed in literature to deal with this issue. First, we use alternative model specifications with several variables to check the robustness of statistical tests. Second, the high correlate explanatory variables are dropped. Third, to attenuate multicollinearity consequences, the models including variables with a larger sample size are preferred (e.g.,  $Corr_i^{CPI}$ , *Shadow<sub>i</sub>*,  $DT_Dec_i$ ). This is because the larger the sample size, the smaller is the standard

<sup>&</sup>lt;sup>11</sup> We also discriminate between centralized and decentralized country according to the threshold of 50% of the ratio between sub national and national expenditure and revenue. However, following this procedure, only 13 countries out of 145 can be defined as "decentralized". In the authors' view, it is an unreasonable binding criterion to identify the decentralized country.

<sup>&</sup>lt;sup>12</sup> According to D<sup>k</sup> ranking: Canada (7); Switzerland (7); Argentina(5); Brazil (3); India(3); Japan (3); Turkey (3); United Kingdom (3); Austria (1); Denmark (1); Ireland (1); Lithuania (1); Luxembourg (1); Malaysia (1); Mexico (1); Morocco (1); Pakistan (1); Russia (1); Sweden (1); United States (1).
<sup>13</sup> This approach providing information on the eigenvector decomposition of the coefficient covariance matrix. It is a

<sup>&</sup>lt;sup>15</sup> This approach providing information on the eigenvector decomposition of the coefficient covariance matrix. It is a useful tool to diagnose multicollinearity problems amongst the regressors. Belsley et al. (1980) propose that degrading collinearity exists when one observes both (a) at least a condition number of the matrix smaller than  $1/30^2$  (0.001) and (b) two or more variables with values greater than 0.5.

error. This approach leads to specify a model where the estimated coefficients of the variables of main interest are not affected by severe collinearity.

Following previous guidelines, we include several control variables into the dataset. They capture a broad range of theoretically plausible determinants of corruption and the shadow economy in the hope of reducing omitted variable bias, endogeneity and multicollinearity sources.

With reference to institutional quality, La Porta et al. (1999) state that greater protections of property against the state embodied in common law systems improve government performances, including reducing corruption. Thus, following Treisman (2000), we consider dummies for the legal origin. It is because civil law systems differ on this dimension from common law systems. Taking into account that these variables have good performances to control for endogeneity, we also include the colonial heritage of countries as instruments in several TSLS model specifications. Sources of these variables are Treisman (2000) for former British colony (*BritCol*), former French colony (*FrenCol*), former Spanish or Portuguese colony (*SpanPorCol*), former colony of state other than Britain, France, Spain, or Portugal (*OtherCol*) and never a colony (*NonCol*). For the set of dummies taking into account the "legal origin", the source is Global Development Network Growth Database. These variables identify the origin of the Company Law or Commercial Code in each country: British legal origin (*LegBrit*), French legal origin (*LegSoc*), German legal origin (*LegGerm*) and Scandinavian legal origin (*LegScand*).

A second set control variable is selected by following Ades and Di Tella's (1999) findings. According to the authors countries that are more open to foreign trade tend to be less corrupt. Consequently, two indicators extracted by World Development Indicators of World Bank are collected: Foreign direct investment as percentage of GDP (*FDI*) and an index of international trade calculated as the percentage of exports and imports of goods and services on the Gross national expenditure ( $t_open$ ).

There is abundant empirical literature that includes, among the controls of corruption, regressions proxies of economic development. One of the most used is the Real GDP per capita ( $LGDP\_cap$ ). Other controls are the level of education (*School*), the size of public sector ( $Gov\_size$ ) and some proxies of institutional environment, for example, index of political rights published by Gastil index, Freedom House, etc. About the latter, considering that the political indexes are relevant sources of multicollinearity and endogeneity for our variables of interest, we opt to drop these variables from regressions as their variability is already adequately accounted by the other explanatory variables. We also consider the urbanization rate (*urban*) in TSLS specifications to instrument the *LGDP\\_cap*.

With reference to the controls of the shadow economy, we further include self-employment (*Self*) and tax revenue as percentage of GDP ( $T_rev$ ).

#### **4.2 Empirical results**

In this section, we convert the main conclusion of the theoretical model (proposition 3 and 4) in testable regressions.

#### 4.2.1 Test on Proposition 3.

**Proposition 3.** The level of the shadow economy and corruption are lower in a federal state than in a centralized country.

The econometric translation of this proposition is as follows:

$$Corr_{i} = \alpha_{1} Dec_{i} + \alpha'_{2} X^{C} + \varepsilon_{i}$$
<sup>(25)</sup>

$$Shadow_i = \beta_1 Dec_i + \beta_2 X^S + \varepsilon_i$$
<sup>(26)</sup>

with i = 1, ..., n and test if  $\alpha_1 < 0$  and  $\beta_1 < 0$ .

In Table 6 and 7 are tested regressions (25) and (26), respectively. The regressions are run by OLS and TSLS to deal with the endogeneity issue.<sup>14</sup> Endogeneity problems in regressions are estimated by the Durbin-Wu-Hausmann (DWH) test. The null hypothesis of DWH is that endogeneity does not affect the OLS estimator. We also provide diagnostic information on the instruments used during estimation, that is, Stock and Yogo's (2005) test. Unfortunately, although we cannot reject the DWH null hypothesis that the variables in question are exogenous (i.e. OLS is consistent and efficient), Stock and Yogo's tests reveal a problem of weak instruments. As a result, DWH results should be regarded with caution.

We perform a set of tests for residual normality, homoskedasticity and serial independence. As regard the hypothesis of the normality of the residuals, we employ the Jarque-Bera test. With reference to homoskedasticity tests, we carry out Harvey (1976) tests. Combining Harvey's outcomes with serial correlation testing we choose the suitable correction of standard errors. In particular, with the exclusion of 1a, 2a and 3a, for each regression first-order serial correlation tests (Durbin-Watson test) advice not to reject the null hypothesis of no serial correlation. According to these results, White (1980) heteroskedasticity consistent covariance matrix estimator for regressions 1a, 2a and 3a, that provides consistent estimates of the coefficient covariances in the presence of conditional heteroskedasticity of unknown form, is applied. For the remaining regressions, we apply HAC consistent covariances proposed by Newey and West (1987). They are consistent in the presence of both heteroskedasticity and autocorrelation of unknown form.

#### Table 6: Dependent Variable: Indexes of Corruption

Regressions are specified in order to minimize the estimation bias for multicollinearity. The estimated eigenvalues have condition numbers smaller than 0.001, which would indicate a significant amount of collinearity. The vpd columns display the decomposition proportions associated with the smallest condition number. As a result, we find that the logarithm of real GDP per capita and the constant and/or legal origin dummies have value larger than 0.5. This indicates that there is a high level of collinearity between those variables, as expected. However, coefficient of decentralization index is not significantly biased.

<sup>&</sup>lt;sup>14</sup> TSLS estimator is not applied for models 4 and 5 because Corr<sup>Bribe1</sup> and Corr<sup>Bribe2</sup> have small sample sizes (i.e. lower than 10) thus these estimates are unreliable.

In Table 7, we find a negative correlation between decentralization and the size of the shadow economy, as expected. This result is robust to multicollinearity and endogeneity.

According to these findings, we conclude that empirical evidence supports Proposition 3: Corruption (Table 6) and Shadow Economy (Table 7) will be lower in countries with greater decentralization.

#### 4.2.2 Test on Proposition 4

**Proposition 4.** The positive effect of corruption on the shadow economy is higher in a centralized country than in a federal state.

The statistical test is performed by splitting the sample into two groups: *d* (decentralized) and *c* (centralized countries) and testing if  $\alpha_1^d \in [0, \alpha_1^c]$ . Consequently, the following regression is estimated:

Shadow<sub>i</sub> = 
$$\alpha_1 Corr + \alpha'_2 X^s + \varepsilon_i$$
 with  $i = 1, ..., d$ , (27)

where d+c=n. In particular, decentralized or federal states are the countries with  $DT\_Dec_i$  equal to one. We also include different controls. For models 7, 8 and 9, we include among the determinants of the shadow economy the index of decentralization also.

Models from 10–F to 14–F are estimated on the whole sample of countries. On the contrary, estimates for models 10–C to 14–C are based on centralized countries (i.e.  $DT_Dec_i = 0$ ), while from 10–D to 12–D, the coefficients are estimated on the sub-samples of decentralized states.

#### Table 8. Dependent Variable: Shadow Economy - All Countries

In Table 8, we report robust empirical findings supporting theoretical model. Decentralized countries have lower shadow economy than centralized states. In average, the difference between the sizes of the unofficial economy and the two institutional settings is about four percentage points, ceteris paribus.

Table 9 shows the estimates of corruption carried out on the two subsamples of centralized and decentralized states. The main limitation of this econometric exercise is the inadequate size of the subsample for decentralized countries. Furthermore, combining the small number of federal countries with the missing values in Corr<sup>Bribe1</sup> and Corr<sup>Bribe2</sup> variables makes it unfeasible to estimate regressions 13–D and 14–D.

#### Table 9. Dependent Variable: Shadow Economy – All Countries

To take into account the problem of endogeneity for the control variables, we use both the set of dummy variables of legal origin and colonial dominance. Focusing on the differences between corruption indexes in the two sub samples, we find that a larger effect of corruption is associated with a centralized state. Thus, Proposition 4 that stipulates that the impact of corruption on the shadow economy will be lower in countries with greater decentralization is empirically supported by data.

#### 5. Conclusions

In this paper, we have examined theoretically and empirically how decentralization affects the relationship between corruption and the shadow economy. The literature that specifically analyzes this relationship appears to be limited and only focuses either on the implications of decentralization for the development of both the informal sector and corruption, or on the impact of a centralized bureaucracy on corruption, taking into account the existence of underground economic activities and their adverse effects on bribes and welfare.

We try to advance the existing literature on the subject in two ways. Firstly, we theoretically analyze the transmission channels through which decentralization may influence both corruption and the development of the informal sector, by also considering the occurrence of indirect effects that could lead these two related phenomena to affect each other. Secondly, from a methodological perspective, we propose a new proxy of decentralization that reflects the different indexes of decentralization. This is aimed to overcome the difficulty related to the empirical use of decentralization measures. It is due to both the complexity of the concept – which implies a scarce adherence of the existing indexes, singularly considered, to its real extent – and the large presence of missing values in available indicators.

In particular, our theoretical model predicts that decentralization is conducive to improving the quality of government intervention, leading to both a lower size of the shadow economy and a lower level of corruption. Moreover, we have demonstrated that decentralization also works indirectly in keeping the shadow economy and corruption at bay, as in more decentralized systems, individuals have the possibility to avoid corruption by moving to other jurisdictions, rather than operating informally. This implies that a higher degree of corruption exerts a lower impact on the shadow economy in more decentralized countries.

In order to test the model's implications, we have developed an empirical analysis based on a crosscountry database of 145 countries that includes a number of different indices of corruption, decentralization and the shadow economy.

The empirical evidence supported the two main results of the theoretical model: (1) decentralization is significantly associated with lower shadow economy and corruption and (2) the impact of corruption on the shadow economy will be lower in countries with greater decentralization. While the tests performed should be treated with caution, due to the problems associated with measurement issues and endogeneity, we can still derive substantial policy implications. Decentralization proves to be an essential precondition for the efficient government provision of productive goods and services as well as the accountability of governments, which facilitates the producer's choice to operate in the official sector. This turns out to be particularly relevant in the debate over the optimal institutional design that concerns both developing and developed countries, before proceeding to welfare state reforms.

#### **APPENDIX A**

#### **Proof of Lemma 2**

From Lemma 1, it follows that politician's monetary rent *R* is zero at t=0 and t=1 since x=1 in this case. R=0 also when  $\theta=0$  as there is no corruption and rents for the politician. This implies that the policy  $(t_p, \theta_p)$  that maximizes monetary rents involves an intermediate level of the tax rate and a positive value of  $\theta$ .

The first order conditions defining  $(t_p, \theta_p)$  are the following

$$\frac{\partial R}{\partial \theta} = \beta a q (1-t) (\omega t)^{(1-\alpha)/\alpha} q \left[ (1-x) + \frac{(1-\alpha)}{\alpha \omega} \omega_{\theta} (1-x) - \frac{\partial x}{\partial \theta} \right] = 0$$
(A.1)

$$\frac{\partial R}{\partial t} = \beta a q \theta(\omega t)^{(1-\alpha)/\alpha} \left[ -(1-x) + \frac{1-\alpha}{\alpha t} (1-t)(1-x) - (1-t)\frac{\partial x}{\partial t} \right] = 0.$$
(A.2)

Substituting (11) and (12) into respectively (A.1) and (A.2), we obtain

$$\frac{\partial R}{\partial \theta} = \beta a q (1-t) (\omega t)^{(1-\alpha)/\alpha} q \left[ (1-x) + \frac{(1-\alpha)}{\alpha} \frac{\omega_{\theta}}{\omega} (1-x)\theta + \frac{\omega_{\theta}}{\omega} \frac{x\theta}{\alpha} - \frac{xaq\theta}{(1-\alpha)(1-aq\theta)} \right] = 0 \quad (A.3)$$
$$\frac{\partial R}{\partial t} = \beta a q \theta (\omega t)^{(1-\alpha)/\alpha} \left[ -(1-x) + \frac{1-\alpha}{\alpha t} (1-t)(1-x) - \frac{x}{1-\alpha} + (1-t)\frac{x}{\alpha t} \right] = 0 \quad (A.4)$$

First, note that the sign of both derivatives is determined by the components within the square brackets. Rearranging terms in the square bracket of (A.4), we obtain the tax rate solving this equation to be  $t = 1 - \alpha$ . From the square bracket of (A.3), it follows that there is a unique level of  $\theta$ , which we call  $\theta_p$ , solving the equation

$$(1-x) + \frac{(1-\alpha)}{\alpha} \frac{\omega_{\theta}}{\omega} (1-x)\theta + \frac{\omega_{\theta}}{\omega} \frac{x\theta}{\alpha} - \frac{xaq\theta}{(1-\alpha)(1-aq\theta)} = 0.$$
(A.5)

In fact, the first component is positive while the other three are negative since  $\omega_{\theta} < 0$ . Moreover, the first component is decreasing in  $\theta$  as a higher level of  $\theta$  increases the labor *x* employed in the shadow economy (see Lemma 1), while the other three components are increasing (in absolute value) in  $\theta$ . This can be shown by considering the absolute value of these three components, and we can define

$$Z = -\frac{1 - \alpha - \alpha x}{\alpha} \frac{\omega_{\theta}}{\omega} \theta + \frac{xaq\theta}{(1 - \alpha)(1 - aq\theta)}$$

Then, we have that

$$\frac{\partial Z}{\partial \theta} = -\frac{1-\alpha+\alpha x}{\alpha}\frac{\omega_{\theta}}{\omega} - \frac{\omega_{\theta}}{\omega}\theta\frac{\partial x}{\partial \theta} - \frac{1-\alpha+\alpha x}{\alpha}\theta\frac{\partial(\omega_{\theta}/\omega)}{\partial \theta} + \frac{aq\theta}{(1-\alpha)(1-aq\theta)}\frac{\partial x}{\partial \theta} + \frac{xaq}{(1-\alpha)}\frac{1}{(1-aq\theta)^{2}} > 0,$$

as  $\omega_{\theta} < 0$ ,  $\partial x / \partial \theta > 0$  and

$$\frac{\partial(\omega_{\theta} / \omega)}{\partial \theta} = \frac{\omega_{\theta\theta} \omega - (\omega_{\theta})^{2}}{\omega^{2}} < 0$$

since  $\omega_{\theta\theta} < 0$ . Finally, notice that  $\theta_p$  is not necessarily lower than 1 as it is possible that the square bracket of (A.3) is positive for all  $\theta \in [0,1]$ .

#### **Proof of Lemma 3**

From Lemma 2, it is straightforward that both politicians tax at the rate  $t_1 = t_2 = 1 - \alpha$  as this maximizes their rent *R* as well as the net income  $y_i$  of the individuals. Taking into account (11) and (12), the reaction functions of the two politicians with respect to  $\theta$  are the following:

$$\frac{\partial R_1}{\partial \theta_1} = \beta aq(1-t_1)(\omega_1 t_1)^{(1-\alpha)/\alpha} qn_1$$

$$\left[ (1-x_1) + \frac{(1-\alpha)}{\alpha} \frac{(\partial \omega_1/\partial \theta_1)}{\omega_1} (1-x_1)\theta_1 + \frac{(\partial \omega_1/\partial \theta_1)}{\omega_1} \frac{x_1\theta_1}{\alpha} - \frac{x_1aq\theta_1}{(1-\alpha)(1-aq\theta_1)} + (1-x_1)\theta_1 \frac{(\partial n_1/\partial \theta_1)}{n_1} \right] = 0$$

$$\frac{\partial R_2}{\partial R_2} = \beta aq(1-t_2)(\omega_2 t_2)^{(1-\alpha)/\alpha} qn_2$$
(A.6)

$$\frac{\partial \theta_2}{\partial \theta_2} + \frac{\partial \theta_2}{$$

where  $\partial n_1 / \partial \theta_1 = \partial h(\cdot) / \partial \theta_1 < 0$  and  $\partial n_2 / \partial \theta_2 = -\partial h(\cdot) / \partial \theta_2 < 0$ . Since the terms outside the square bracket are positive, the Nash equilibrium is the solution of the following system of equations:

$$(1-x_1) + \frac{(1-\alpha)}{\alpha} \frac{(\partial \omega_1/\partial \theta_1)}{\omega_1} (1-x_1)\theta_1 + \frac{(\partial \omega_1/\partial \theta_1)}{\omega_1} \frac{x_1\theta_1}{\alpha} - \frac{x_1aq\theta_1}{(1-\alpha)(1-aq\theta_1)} = -\frac{(1-x_1)\theta_1}{n_1} \frac{\partial n_1}{\partial \theta_1}$$
(A.8)

$$(1-x_2) + \frac{(1-\alpha)}{\alpha} \frac{(\partial \omega_2 / \partial \theta_2)}{\omega_2} (1-x_2) \theta_2 + \frac{(\partial \omega_2 / \partial \theta_2)}{\omega_2} \frac{x_2 \theta_2}{\alpha} - \frac{x_2 a q \theta_2}{(1-\alpha)(1-a q \theta_2)} = -\frac{(1-x_2) \theta_2}{n_2} \frac{\partial n_2}{\partial \theta_2}.$$
 (A.9)

Since the equilibrium is symmetric,  $\theta_1 = \theta_2 = \theta_r$  where  $\theta_r$  is the solution to the following equation:

$$(1-x) + \frac{(1-\alpha)}{\alpha} \frac{(\partial \omega/\partial \theta)}{\omega} (1-x)\theta + \frac{(\partial \omega/\partial \theta)}{\omega} \frac{x\theta}{\alpha} - \frac{xaq\theta}{(1-\alpha)(1-aq\theta)} = -\frac{(1-x)\theta}{n} \frac{\partial n}{\partial \theta}.$$
 (A.10)

From the comparison of (A.10) and (A.5) defining  $\theta_p$ , it follows that  $\theta_r < \theta_p$  since the left hand side of both expressions are the same while the right hand side of (A.10) is positive (as  $\partial n/\partial \theta = \partial h(\cdot)/\partial \theta < 0$ ).

# Appendix B: Database (1)

Symbol	Description	Source	Obs	Min	Max
Av.CPI	Transparency International's Corruption Perceptions Index (CPI) ranks countries based on how	http://archive.transparency.org/policy_research/surveys_i	145	1.2	9.8
	corrupt their public sector is perceived to be. Average (1999-2005)	ndices/cpi			
Av.WGI	Control of Corruption: Reflects perceptions of the extent to which public power is exercised for	Worldwide Governance Indicators. World Bank.	144	-1.7	1.8
	private gain. Average (1998-2005, not available data for 1999),	(www.govindicators.org)			
Av.Bribe1	Percentage of firms expected to make informal payments to public officials to "get things done"	World Development Indicators. [IC.FRM.CORR.ZS]	88	4.4	98.3
	with regard to customs, taxes, licenses, regulations, services, and the like. Average (1999-2007)	based on Enterprise Surveys- World Bank.	10-		6.0
Av.CorrICRG	Political risk points by component "F - Corruption". (Scores 2005 + Scores 2011)/2.	International Country Risk Guide. The PRS Group, Inc.	127	0.8	6.0
Av.CorrGIFT	Percent of firms expected to give gifts to public officials "to get things done". Average('05-10)		108	0.0	85.1
Av.CorrTAX	Percent of firms expected to give gifts in meetings with tax officials. Average (2005-'10)		108	0.0	66.7
Av.CorrGOV	Percent of firms expected to give gifts to secure government contract. Average (2005-2010)		107	0.0	97.7
Av.CorrLIC	Percent of firms expected to give gifts to get an operating license . Average (2005-2010)		107	0.0	16.4
Av.CorrIMP	Percent of firms expected to give gifts to get an import license. Average (2005-2010)		107	0.0	88.2
Av.CorrCONS	Percent of firms expected to give gifts to get a construction permit. Average (2005-2010)		108	0.0	61.1
Av.Corr.ELEC	Percent of firms expected to give gifts to get an electrical connection. Average (2005-2010)	Enterprise Surveys; World Bank.	107	0.0	91.6
Av.CorrWAT	Percent of firms expected to give gifts to get a water connection. Average (2005-2010)	http://www.enterprisesurveys.org	106	0.0	71.1
Av.CorrPAST	Percent of firms experiencing at least one bribe payment request. Average (2005-2010)		103	0.0	70.9
Av.CorrREL	Percent of firms identifying corruption as a major constraint. Average (2005-2010)		108	0.0	65.2
Av.CorrJUS	Percent of firms identifying the courts system as a major constraint. Average (2005-2010).		108	0.0	68.9
INF_comp	Percent of firms competing against unregistered or informal firms. Average (2005-2010)		108	0.0	89.8
INF rel	Percent of firms identifying practices of competitors in the informal sector as a major constraint.		108	0.0	72.4
	Average (2005-2010)				
Shadow	Shadow Economy as percentage of official GDP. Average (1999-2007)	Buehn and Schneider (2012a), Table 3	145	8.64	66.9
Exp_Dec	100*{1-[Total Expenditure, Noncash (Cen. Govt.)/Total Expenditure, Noncash (Gen. Govt.)]}. Average (1999-2007)	Government Finance Statistics. IMF	45	12.5	129.5
Rev_Dec	100*{1-[Revenue, Noncash (Cen. Govt.)/Revenue, Noncash (Gen. Govt.)]}. Average (1999-2007)	Government Finance Statistics. IMF	48	12.4	130.8
D1_Fed	Classified as a federation by Elazar (1995) and updated by Treisman (2007) by adding Ethiopia, Serbia-Montenegro, Bosnia-Herzegovina, which became federal after Elazar' article. (Federal =1).	Available from: www.sscnet.ucla.edu/polisci/faculty/ treisman/Papers/what_have_we_learned_data.xls, [Fedelaz]	141	0	1
D2_Fed	Persson and Tabellini (2003). (Federal =1).	Available from: http://didattica.unibocconi.it/mypage/ index.php?IdUte=48805&idr=4273,[Federal]	81	0	1
D3_Fed	Dummy variable = 1 if (a) constitution reserves decision making on at least one topic exclusively to subnational legislatures and/or (b) constitution assigns to subnational legislatures exclusive right to legislate on issues that it does not specifically assign to one level of government.	Treisman (2008) [Auton]	68	0	1
DT_Fed	Dummy variable: Decentralized =1. (See paragraph 4.1 for details)	Own estimation based on GFS (IMF) and Treisman (2007, 2008) Pearsson and Tabellini (2003)	140	0	1

# Appendix B: Database (2)

Symbol	Description	Source	Obs	Min	Max
BritCol	Former British colony		134	0	1
FrenCol	Former French colony	Treisman (2000). Available from: www.sscnet.ucla.edu/	134	0	1
SpanPorC	Former Spanish or Portuguese colony	polisci/faculty/treisman/Papers/what_have_we_learned_data.xls. We	134	0	1
OtherCol	Former colony of state other than Britain, France, Spain, or Portugal,	correct four digit errors in this file.	134	0	1
NonCol	Never a colony		134	0	1
LegBritish	legal origin: British,		126	0	1
LegFren	legal origin: French	CDN Crowth Databasa Available from www.coopet.vola.edu/rolicei	126	0	1
LegSoc	legal origin: Socialist	/faculty/traisman/Dapars/what have we learned date vis	126	0	1
LegGerm	legal origin: German,	/lacuity/iteismail/Fapers/what_have_we_learneu_data.xis	126	0	1
LegScand	legal origin: Scandinavian,		126	0	1
LGDPcap	Logarithm of GDP per capita (constant 2000 US\$). Average (1999-2007)	World Development Indicators Log[NY.GDP.PCAP.KD]	143	4.5	10.8
FDI	Average fromForeign direct investment, net inflows (% of GDP). Average (1999-2007)	World Development Indicators. [BX.KLT.DINV.WD.GD.ZS]	143	-6.5	388.4
t_open	(Exports of goods and services (current LCU)+Imports of goods and services (current LCU)/(Gross national expenditure (current LCU). Average (1999-2007)	World Development Indicators. 100*[(NE.EXP.GNFS.CN+ NE.IMP.GNFS.CN)/ NE.DAB.TOTL.CN]	136	0.6	514.8
School	School enrollment, secondary (% gross). Average (1999-2007)	World Development Indicators. [SE.SEC.ENRR]	135	8.4	152.6
Gov_size	General government final consumption expenditure (current LCU)/ Final consumption expenditure (current LCU). Average (1999-2007)	World Development Indicators 100*[NE.CON.GOVT.CN /NE.CON.TOTL.CN]	116	5.4	53.0
Self	Self-employed, total (% of total employed). Average (1999-2007)	World Development Indicators [SL.EMP.SELF.ZS]	121	0.9	94.7
T_rev	Tax revenue (% of GDP). Average (1999-2007)	World Development Indicators. [GC.TAX.TOTL.GD.ZS]	117	0.2	45.2
Urban	Urban population (% of total). Average (1999-2007)	World Development Indicators. [SP.URB.TOTL.IN.ZS]	142	9.0	100

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Authors	Data <sup>a,b,c</sup>	Methodology	Findings
Treisman (2000)	<ul> <li>(a) 69 countries</li> <li>(b) CPI, Business international and Global competitiveness</li> <li>(c) Federal dummy</li> </ul>	Cross-section WLS and OLS	The level of corruption is higher in federal countries
Fan et al. (2009)	<ul> <li>(a) 80 countries</li> <li>(b) WBES</li> <li>(c) Different indexes of decentralization</li> </ul>	Cross country cross firm PROBIT	A larger number of vertical government tiers and a larger fraction of local bureaucratic personnel are related to more corruption. However, greater subnational revenues are linked to lower corruption.
Fisman and Gatti (2002a)	<ul> <li>(a) 55 countries</li> <li>(b) CPI, ICRG, WCR, Business international and GCS</li> <li>(c) Expenditure decentralization</li> </ul>	Cross-section OLS and TSLS	Fiscal decentralization in government expenditure is associated with lower corruption.
Fisman and Gatti (2002b)	<ul> <li>(a) U.S. states</li> <li>(b) Convictions for abuse of public office</li> <li>(c) Shares of federal transfers</li> </ul>	Cross-section OLS	Large federal transfers are related to a higher rate of convictions for bureaucratic corruption.
Arikan (2004)	<ul> <li>(a) 40 countries</li> <li>(b) CPI</li> <li>(c) Expenditure decentralization, number of local jurisdictions</li> </ul>	Cross-section OLS and TSLS	A higher degree of decentralization is associated with a lower level of corruption.
Lederman et al. (2005)	<ul><li>(a) 102 countries</li><li>(b) ICRG</li><li>(c) Central transfers</li></ul>	Panel Pooled OLS PROBIT	related to higher corruption, while fiscal decentralization is associated with lower corruption.
Enikalopov and Zhuravskaya (2007)	<ul> <li>(a) 75 countries</li> <li>(b) CPI, WBC</li> <li>(c) Revenue decentralization, expenditure decentralization</li> </ul>	Cross-section, panel OLS and TSLS	The beneficial effects of decentralization on corruption depend on the national political party system's strength.
Lessmann and Markwardt (2010)	<ul> <li>(a) 64 countries</li> <li>(b) CPI, ICRG, WBC</li> <li>(c) Federal dummy, vertical government tiers, revenue decentralization, expenditure decentralization, sub-nationals government employment</li> </ul>	Cross-section OLS and TSLS	Decentralization is effective in counteracting corruption if there is a supervisory body, such as a free press, able to guarantee the monitoring of bureaucrats' behaviour.

Table 1. Previous empirical studies on decentralization and corruption

(a) Sample size; (b) Corruption indexes; (c) Decentralization measures.

Authors	Data <sup>a,b,c</sup>	Methodology	Findings
Johnson et al. (1998; 1999)	<ul> <li>(a) 49 countries</li> <li>(b) ICRG, TIC, GCSB</li> <li>(c) Own estimates based on electricity consumption; MIMIC model estimates from both Loayza (1996) and Schneider (1997)</li> <li>(a) 60 countries</li> </ul>	Cross-section OLS	Countries with an intensive regulatory burden are characterized by higher share of shadow economy, in a context of great discretion in the implementation of regulatory rules.
Friedman et al. (2000)	<ul> <li>(a) 09 countres</li> <li>(b) GCSB index; IT index; Political Risk Service index; German business people survey</li> <li>(c) MIMIC values from Schneider and Enste (2000); electricity consumption estimates</li> </ul>	Cross-section OLS and TSLS	Corruption is associated with a higher level of shadow economy.
Dreher et al. (2009)	<ul><li>(a) 18 OECD countries</li><li>(b) Own estimates</li><li>(c) Own estimates</li></ul>	SEM	Corruption and shadow economy are substitutes: the existence of the informal sector limits the ability of bureaucrats to extract bribes from economic activities.
Dreher and Schneider (2010)	<ul> <li>(a) 120 and 70 countries</li> <li>(b) ICRG</li> <li>(c) MIMIC model estimates from Schneider</li> </ul>	Cross-section, panel OLS, TSLS and SEM	Corruption and shadow economy are substitutes in high income countries, while they are complements in low income countries.
Buehn and Schneider (2012b)	<ul> <li>(a) 51 countries</li> <li>(b) Own estimates; CPI</li> <li>(c) Own estimates; MIMIC model estimates from Schneider</li> </ul>	SEM	Shadow economy is conducive to increasing the level of corruption.

Table 2.	Previous	empirical	studies	on	shadow	economy	and co	orruption

(a) Sample size; (b) Corruption indexes; (c) Shadow economy measures.

Table 3. Correlation matrix of	proxies of Shadow economy
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	Shadow <sub>i</sub>	Inf_Comp <sub>i</sub>	Inf_Rel <sub>i</sub>
Shadow <sub>i</sub>	1.000		
Inf_Comp <sub>i</sub>	0.286	1.000	
Inf_Rel <sub>i</sub>	0.257 (2.742)	0.586 (7.446)	1.000

The numbers in parenthesis are the t-ratios.

	$Corr_i^{CPI}$	$Corr_i^{WGI}$	$Corr_i^{ICGR}$	$Corr_i^{Bribe1}$	$Corr_i^{Bribe2}$	Shadow <sub>i</sub>	D1_Dec <sub>i</sub>	$D2\_Dec_i$	$D3\_Dec_i$	Exp_Dec <sub>i</sub>	Rev_Dec <sub>i</sub>	Subexp <sub>72-00</sub>	Subrev <sub>72-00</sub>
Corr <sup>CPI</sup>	1												
Coni	(145)												
Corr <sup>WGI</sup>	0.924	1											
com	(0.00;143)	(143)											
Corr. <sup>ICGR</sup>	0.866	0.837	1										
1	(0.00;127)	(0.00;125)	(127)										
$Corr_{i}^{Bribe1}$	0.523	0.569	0.455	1									
l	(0.00;88)	(0.00;86)	(0.00;76)	(88)	_								
$Corr_{i}^{Bribe2}$	0.558	0.583	0.419	0.752	1								
L	(0.00;107)	(0.00;105)		(0.00;83)	(107)	1							
Shadow <sub>i</sub>	0.700	0.035	0.094	(0.22.98)	(0.140)	1 (145)							
	(0.00;145)	(0.00;143)	(0.00;127)	(0.55;88)	(0.13;107)	(143)	1						
D1 $Dec_i$	-0.167	-0.168	-0.102	0.008	-0.010	-0.042	I						
	(0.05;141)	(0.05;139)	(0.26;123)	(0.94;86)	(0.92;105)	(0.62;141)	(141)						
D2 Dec:	-0.128	-0.094	-0.088	-0.059	0.144	-0.258	0.398	1					
	(0.26;81)	(0.40;79)	(0.44;78)	(0.70;47)	(0.28;59)	(0.02;81)	(0.00;79)	(81)					
D3 Dec;	-0.103	-0.078	-0.086	-0.130	-0.068	-0.202	0.358	0.539	1				
	(0.40;68)	(0.53;67)	(0.49;66)	(0.45;36)	(0.66;45)	(0.10;68)	(0.00;66)	(0.00;67)	(68)				
Exp Dec <sub>i</sub>	-0.214	-0.190	-0.132	0.240	-0.171	-0.207	-0.133	0.042	0.041	1			
1 = .	(0.14;48)	(0.20;48)	(0.38;46)	(0.28;22)	(0.37;29)	(0.16;48)	(0.37;47)	(0.79;43)	(0.80;40)	(48)			
$Rev_Dec_i$	-0.225	-0.207	-0.159	0.187	-0.199	-0.245	-0.153	-0.020	0.004	0.959	1		
	(0.14;45)	(0.17;45)	(0.30;44)	(0.43;20)	(0.32;27)	(0.11;45)	(0.32;44)	(0.90;40)	(0.98;37)	(0.00;45)	(45)	1	
Subexp <sub>72-00</sub>	-0.390	-0.325	-0.389	-0.181	(0.6251)		0.276	0.457	0.135	(0.54.42)	0.062	[ (71)	
		(0.01;09)	(0.00;70)	(0.26;40)	0.106		(0.02;70)	(0.00;71)	(0.31;00)	(0.54;42)	(0.70;40)	(/1)	1
Subrev <sub>72-00</sub>	-0.290	-0.252	-0.28/	-0.092	(0.46.50)	-0.232	0.245	0.438	(0.092)	(0.21.42)	(0.130)	U.948 (0.00.70)	1 (70)
Mana	(0.02;70)	(0.00;08)	(0.02;09)	(0.38,39)	(0.40,30)	(0.05;70)	(0.04;09)	(0.00;70)	(0.49,00)	(0.31,42)	(0.43,40)	21.00	(70)
Mediar	5.00	4.90	5.30	3.73	2.00	33.43	0.10	0.20	0.22	39.03	41.03	21.09	12.00
Min	0.00	0.08	0.00	0.00	2.30	864	0.00	0.00	0.00	12 27	12 16	20.30	1 10
Max	8.78	8.02	10.00	10.00	10.00	66.04	1	1	1	12.37	12.40	<u> </u>	52.18
wias	0.70	0.02	10.00	10.00	10.00	00.90	1	1	1	130.78	129.30	57.45	52.10

# Table 4. Correlation matrix and descriptive statistics

In parenthesis p-value of  $H_0$ :  $r_{xy}=0$  and number of observations. In bold are the correlations significant at the 0.05 level.

Table 5: Correlation matrix of DT_Dec										
	$Corr_i^{CPI}$	$Corr_i^{WGI}$	Shadow <sub>i</sub>	$DT\_Dec_i$						
$Corr_i^{CPI}$	1 (145)									
$Corr_i^{WGI}$	0.924 (28.65; 143)	1 (143)								
Shadow <sub>i</sub>	0.700 (11.71; 145)	0.635 (9.76; 143)	1 (145)							
DT_Dec <sub>i</sub>	-0.332 (-4.13; 140)	-0.290 (-3.533; 138)	-0.324 (-4.018; 140)	1 (140)						

The numbers in parenthesis are the t-ratios and number of observations

Table 6: Dependent Variable: Indexes of Corruption

		Corr <sup>CPI</sup>			Corr <sup>WGI</sup>		Corr	ICGR	Corr <sup>Bribe1</sup>	Corr <sup>Bribe2</sup>
	1a: OLS	vdp	1b: TSLS	2a: OLS	vdp	2b: TSLS	3a: OLS	3b: TSLS	4a: OLS	5a OLS
$DT\_Dec_i$	-0.84 <sup>**</sup> (-2.03)	0.06	-1.99 <sup>***</sup> (-2.75)	-0.49 (-1.22)	0.05	-1.62 <sup>**</sup> (-2.58)	-0.87 <sup>*</sup> (-1.69)	-1.87 <sup>**</sup> (-2.32)	-1.34 <sup>*</sup> (-1.77)	-0.55 (-0.67)
FDI	-0.004 <sup>**</sup> (-2.00)	0.23	0.002 (0.25)	-0.002 (-0.89)	0.26	0.004 (0.41)	$-0.01^{**}$ (-2.46)	0.00 (0.02)	0.04 (0.39)	0.11 (1.14)
LGDPcap	-0.94 (-6.34)	$0.54^{\circ}$	0.50 (-0.58)	-0.86 (-5.09)	$0.64^{\circ}$	0.48 (0.62)	$-0.80^{-0.71}$	0.46 (0.55)	_	- *
Gov_size	$-0.04^{**}$ (-2.11)	0.07	$-0.17^{**}$ (-2.08)	-0.03 (-1.48)	0.00	$-0.14^{*}$ (-1.84)	0.02 (0.75)	-0.05 (0.58)	-0.05 (-0.94)	$-0.11^{\circ}$ (-2.21)
$Sch_2g$	-0.00 (-0.06)	0.15	-0.05 (-1.58)	-0.00 (-0.05)	0.22	$-0.05^{*}$ (-1.67)	-0.01 (-0.78)	$-0.06^{*}$ (-1.76)	-0.02 (-1.31)	-0.02 (-1.32)
legBrit	$-1.41^{-3.33}$	0.32	-1.08 (-1.23)	$-0.88^{**}$ (-2.40)	0.19	-0.66 (-0.95)	-0.35 (-0.77)	0.20 (0.25)	$-2.11^{-1}$ (3.00)	-0.12 (-0.21)
legFren	-0.83** (-2.38)	0.44	-0.15 (-0.18)	-0.63** (-2.29)	0.26	0.14 (0.27)	0.11 (0.34)	0.73 (1.17)	$-1.61^{-1.61}$ (-2.68)	0.66 (1.25)
legGerm	-0.30 (-0.43)	0.23	-0.52 (-0.44)	-0.39 (-0.50)	0.17	-0.59 (-0.50)	0.76 (0.85)	0.50 (0.36)	-0.30 (0.65)	-0.01 (-0.02)
legSoc	-1.25 (-2.72)	$0.55^{\circ}$	-0.33 (0.76)	-0.66 (-1.57)	0.43	0.13 (0.17)	-0.17 (-0.29)	0.48 <sup>**</sup> (0.48)	-1.23 (-1.27)	0.22 (0.73)
constant	15.27 (15.79)	0.93 <sup>°</sup>	9.66 <sup>**</sup> (2.58)	13.05 <sup>(13.82)</sup>	0.94 <sup>°</sup>	7.77 <sup>**</sup> (2.33)	12.30 <sup>(9.73)</sup>	6.98 <sup>**</sup> (1.91)	7.43 <sup>(4.90)</sup>	5.32 (4.28)
Observ.	94	94	74	94	94	74	85	69	56	65
R <sup>2</sup> -adj	0.696	-	0.486	0.598	-	0.420	0.476	0.356	0.003	0.169
D- $W$ test <sup>a</sup>	1.960	-	1.751	1.998	-	1.854	1.993	2.136	0.943	1.141
Het. test $b$	0.067	—	0.719	0.264	—	0.740	0.053	0.308	0.000	0.000
Nor. test $^{c}$	0.808	—	0.702	0.361	—	0.486	0.520	0.996	0.249	0.022
DWH test <sup>d</sup>	-	_	0.172	-	_	0.141	_	0.337	_	_
Vcd: Eig.	-	1.23	-	-	1.08	-	-	-	—	-
Vcd:Cond.	-	0.00	-	-	0.00	-	-	-	—	-
C. Std. Err.	White	_	HAC	White	_	HAC	White	HAC	HAC	HAC

\*\*\*\*Denotes significant at 1% level; \*\*Denotes significant at 5% level; \*Denotes significant at 10% level. The numbers in parenthesis are the t-ratios. vdp: Variance Decomposition Proportions, ° denotes coefficients strongly affected by multicollinearity (vdp>0.50). Regression 3, has same qualitative vdp results or model 2a.

<sup>a</sup> Durbin–Watson (statistic).<sup>b</sup> Harvey (1976) Test where the null hypothesis is no heteroskedasticity. The p-values of F-statistic are reported. <sup>c</sup>Jargue-Bera Test, (p-value) A small p-value leads to the rejection of the null hypothesis of a normal distribution.

<sup>d</sup>Durbin-Wu-Hausman Test. P-values of differences in J-statistics is reported. Null Hypothesis from 1b-3b is: FDI, LGDPcap and Gov\_size are exogenous. Instruments are: t\_open; t\_rev; urban; britcol; frencol; spanporc; othercol and noncol.

	Sh	adow Econor	my	Shadow Economy			
	6a: OLS	vpd	6b: TSLS	6c: OLS	vpd	6d: TSLS	
$DT\_Dec_i$	-6.38 <sup>**</sup> (-2.02)	0.03	-7.25 (-0.44)	-7.66 <sup>***</sup> (-2.67)	0.12	-6.98 <sup>***</sup> (-2.67)	
T_rev	-0.06 (-0.54)	0.16	-0.53 (-0.90)	_	_	_	
Self	0.12 (1.51)	$0.68^{\circ}$	0.12 (0.60)	_	_	-	
Gov_size	$-0.47^{***}$ (-0.04)	0.39	-0.15 (-0.32)	$-0.50^{***}$ (-4.14)	0.16	-0.45 <sup>**</sup> (-2.29)	
Urban	-0.04 (-0.54)	0.32	-0.09 (0.55)	$-0.16^{***}$ (-3.54)	0.21	-0.15 <sup>***</sup> (-2.97)	
BritCol	2.34	0.00	4.82	6.97**	0.44	$6.27^{**}$	
FrenCol	-0.95	0.02	3.58	3.28	0.76 <sup>°</sup>	6.61*	
SpanPorC	-1.86	0.23	1.46	1.42	0.73 <sup>°</sup>	5.00	
OtherCol	-1.86	0.04	4.99	5.57	0.56	5.52	
Constant	39.54***	<i>0.96</i> °	$40.38^{*}$	48.16***	0.57 <sup>°</sup>	44.51***	
Obs.	78	78	70	101	101	74	
R <sup>2</sup> -adj	0.352	_	0.334	0.334	_	0.381	
D- $W$ test <sup>a</sup>	1.509	_	1.337	1.592	_	1.218	
Het. test <sup>b</sup>	0.904	_	0.906	0.601	_	0.371	
Nor. test <sup>c</sup>	0.000	_	0.004	0.001	_	0.000	
DWH test <sup>d</sup>	_	_	0.828	-	_	0.800	
Vcd: Eig.	_	81.93	_	-	49.48	-	
Vcd:Cond.	_	0.000	_	—	0.000	—	
C. Std. Err.	HAC	_	HAC	HAC	_	HAC	

 Table 7. Dependent Variable: Shadow Economy

Notes: see Table 6.

<sup>d</sup>Durbin-Wu-Hausman Test. P-values of differences in J-statistics is reported. Null Hypothesis of 6b is: t\_rev and self are exogenous. Null Hypothesis of 6d is gov\_size is exogenous. Instruments for 6b are: fdi; t\_open; legBrit; legFren; legGerm; legSoc and legScand. Instruments for 6d are: t\_rev; FDI; t\_open; legBrit; legFren; legGerm; legSoc; legScand.

	79	vnd	7h	89	Qa	10_F	11_F	12_F	13_F	14_F
	OLS	vpu	TSLS	OLS	OLS	OLS	OLS	OLS	OLS	OLS
	-3.34*	0.05	-1.11	-4.63**	-3.83*					
$DI\_Dec_i$	(-1.75)	0.05	(-0.44)	(-2.14)	(-1.82)	-	-	_	—	-
Corr <sup>CPI</sup>	4.03 <sup>***</sup> (10.91)	0.01	4.69 <sup>***</sup> (8.81)	-	-	4.22 <sup>***</sup> (12.30)	-	-	-	-
Corr <sup>WGI</sup>	-	-	-	3.95 <sup>***</sup> (8.03)	-	_	4.19 <sup>***</sup> (10.03)	-	_	-
Corr <sup>ICGR</sup>	-	-	-	-	4.27 <sup>***</sup> (11.44)	_	-	4.50 <sup>***</sup> (11.73)	_	-
Corr <sup>Bribe1</sup>	-	-	_	-	_	_	-	_	0.78 (1.40)	
Corr <sup>Bribe2</sup>	-	-	-	-	-	-	-	-	_	0.66 (1.06))
legBrit	$8.20^{**}$	0.73 <sup>°</sup>	4.11	12.66***	5.17	$7.97^{**}$	12.42**	5.13	30.45***	36.40***
legFren	4.96	0.89°	1.58	$10.01^{***}$	0.37	3.32	8.12**	-1.41	30.30***	34.39***
legGerm	$4.65^{*}$	0.21°	0.81	9.53***	3.36	1.56	5.51**	-0.18	$27.70^{***}$	29.47***
legSoc	3.98	$0.71^{\circ}$	0.38	$7.82^*$	-2.02	2.81	6.44	-3.29	$27.89^{***}$	35.81***
legScand	1.13	$0.45^{\circ}$	-6.60	6.68	2.47	-0.79	4.36	0.45	$25.16^{**}$	29.69***
BritCol	2.85	0.69 <sup>°</sup>	3.41	1.45	$4.88^{*}$	1.79	0.16	3.41	2.68	-
FrenCol	6.14	$0.72^{\circ}$	4.39	5.41	9.67***	6.24	5.433	9.60***	3.84	-
SpanPorC	4.38	$0.74^{\circ}$	3.14	4.54	$8.84^{**}$	4.58	4.64	$8.95^{**}$	3.16	-
OtherCol	$7.27^{**}$	0.63 <sup>°</sup>	5.34	$6.62^{*}$	12.23***	$6.65^{*}$	5.70	$11.28^{***}$	5.64	-
Obs.	119	119	105	117	103	121	119	105	75	94
R <sup>2</sup> -adj	0.446	-	0.420	0.403	0.467	0.441	0.392	0.463	-0.088	-0.02
D- $W$ test <sup>a</sup>	1.481	-	1.502	1.575	1.470	1.476	1.56	1.438	1.324	1.426
Het. test <sup>b</sup>	0.081	-	0.001	0.066	0.485	0.054	0.075	0.223	0.040	0.053
Nor. test <sup>c</sup>	0.002	—	0.034	0.000	0.074	0.001	0.000	0.070	0.085	0.322
$DWH test^d$	-	—	0.194	-	-	-	-	-	-	-
Vcd: Eig.	-	90.20	-	-	-	-	—	-	-	-
Vcd:Cond.	-	0.000	-	-	-	-	-	-	-	-
C. Std. Err.	HAC	-	HAC	HAC	HAC	HAC	HAC	HAC	HAC	HAC

Table 8. Dependent Variable: Shadow Economy – All Countries

Notes: see Table 6. <sup>d</sup>Durbin-Wu-Hausman Test. P-values of differences in J-statistics is reported. Null Hypothesis is Corr<sup>CPI</sup> is exogenous. Instruments are: FDI; t\_open; school; urban.

	10-C	11–C	12-С	13–C	14–C	10-D	11–D	12–D
	Centralized Countries					Decentralized Countries		
Corr <sup>CPI</sup>	4.25 <sup>***</sup> (8.84)	_	-	_	-	3.61 <sup>**</sup> (2.78)	_	_
Corr <sup>WGI</sup>	_	4.03 <sup>***</sup> (6.08)	-	_	_	_	3.22 <sup>**</sup> (2.49)	_
Corr <sup>ICGR</sup>	-	_	4.57 <sup>***</sup> (9.40)	_	-	-	_	3.22 <sup>*</sup> (2.27)
Corr <sup>Bribe1</sup>	_	_	-	0.51 (0.91)		_	_	_
Corr <sup>Bribe2</sup>	-	_	_	_	0.75 (1.18)	_	_	_
legBrit	6.29	$11.70^{**}$	2.00	31.18***	35.55***	$15.78^{**}$	20.94	19.19**
legFren	3.35	$9.07^*$	-2.29	31.98***	34.81***	10.15	19.93***	16.51
legGerm	2.93	$9.79^{***}$	-5.02	$27.82^{***}$	28.44***	5.34	10.16	9.39
legSoc	1.49	6.40	-5.55	$29.76^{***}$	34.15***	10.92	16.81	12.22
legScand	-1.10	6.75	-1.54	$25.46^{**}$	$28.27^{***}$	-0.32	2.05	3.40
BritCol	3.07	1.57	$5.65^{*}$	2.82	0.55	-1.67	-5.77	-4.70
FrenCol	6.50	6.12	$11.16^{***}$	3.42	1.50	-3.17	-10.44	-12.26
SpanPorC	5.35	5.45	$10.28^{***}$	3.57	-0.72	-5.30	-8.47	-7.16
OtherCol	$8.23^{*}$	$7.58^*$	14.54***	7.64	2.47	-0.33	-4.23	-2.79
Obs.	101	99	85	66	79	18	18	18
R <sup>2</sup> -adj	0.342	0.294	0.392	-0.102	-0.081	0.479	0.355	0.438
D- $W$ test <sup>a</sup>	1.223	1.239	1.158	1.052	1.211	1.692	2.694	1.028
Het. test <sup>b</sup>	0.000	0.000	0.000	_	0.000	0.000	0.000	0.000
Nor. test <sup>c</sup>	0.024	0.000	0.151	0.123	0.207	0.635	0.393	0.965
C. Std. Err.	HAC	HAC	HAC	HAC	HAC	HAC	HAC	HAC

 Table 9. Dependent Variable: Shadow Economy – All Countries

Notes: see table 6.