# Limits on the Macho content of the Galactic Halo from the EROS-2 Survey of the Magellanic Clouds ${ }^{\star, \star \star}$ 

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#### Abstract

Aims. The EROS-2 project was designed to test the hypothesis that massive compact halo objects (the so-called "machos") could be a major component of the dark matter halo of the Milky Way galaxy. To this end, EROS-2 monitored over 6.7 years $33 \times 10^{6}$ stars in the Magellanic clouds for microlensing events caused by such objects. Methods. In this work, we use only a subsample of $7 \times 10^{6}$ bright stars spread over $84 \mathrm{deg}^{2}$ of the LMC and $9 \mathrm{deg}^{2}$ of the SMC. The strategy of using only bright stars helps to discriminate against background events due to variable stars and allows a simple determination of the effects of source confusion (blending). The use of a large solid angle makes the survey relatively insensitive to effects that could make the optical depth strongly direction dependent. Results. Using this sample of bright stars, only one candidate event was found, whereas $\sim 39$ events would have been expected if the Halo were entirely populated by objects of mass $M \sim 0.4 M_{\odot}$. Combined with the results of EROS-1, this implies that the optical depth toward the Large Magellanic Cloud (LMC) due to such lenses is $\tau<0.36 \times 10^{-7}$ ( $95 \% \mathrm{CL}$ ), corresponding to a fraction of the halo mass of less than $8 \%$. This optical depth is considerably less than that measured by the MACHO collaboration in the central region of the LMC. More generally, machos in the mass range $0.6 \times 10^{-7} M_{\odot}<M<15 M_{\odot}$ are ruled out as the primary occupants of the Milky Way Halo.


Key words. Galaxy: halo - cosmology: dark matter - gravitational lensing

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## 1. Introduction

Since the proposal (Paczyński 1986; Petrou 1981) that dark matter in the form of faint compact objects (machos ${ }^{1}$ ) could be found through gravitational microlensing, the EROS, MACHO, OGLE, MOA and SuperMACHO collaborations have monitored millions of stars in the Magellanic Clouds to search for microlensing events. Such events would be due to a lensing object passing near the line of sight toward a background Magellanic star, causing a transient magnification of the star's primary image as well as creating a secondary image. Neither the image separation nor the image size are normally resolvable, so the only easily observable effect during an event is an apparent transient amplification of the star's flux. The amplification is greater than a factor 1.34 if the line of sight to the star passes within the lens's Einstein ring of squared radius $r_{\mathrm{E}}^{2}=4 G M D_{\mathrm{s}} x(1-x) / c^{2}$ where $D_{\mathrm{s}}$ is the distance to the star and $x D_{\mathrm{s}}$ is the distance to

[^1]the lens of mass $M$. The optical depth for microlensing, i.e. the probability that at a given time a given star is amplified by more than a factor 1.34 , is
\[

$$
\begin{equation*}
\tau=\frac{4 \pi G D_{\mathrm{s}}^{2}}{c^{2}} \int_{0}^{1} \mathrm{~d} x \rho(x) x(1-x) \tag{1}
\end{equation*}
$$

\]

where $\rho$ is the mass density of lenses. For source stars in the Magellanic Clouds, the order of magnitude of $\tau$ is $f v_{\mathrm{rot}}^{2} / c^{2} \sim$ $f \times 10^{-6}$ where $v_{\text {rot }} \sim 220 \mathrm{~km} \mathrm{~s}^{-1}$ is the rotation velocity of the Milky Way and $f$ is the fraction of the halo mass that is comprised of lensing objects. The factor of proportionality between $\tau$ and $f v_{\mathrm{rot}}^{2} / c^{2}$ depends on the structure of the Halo. The benchmark value is often taken to be that for a spherical isothermal halo of core radius 5 kpc , the so-called " S model" used by the MACHO collaboration (Alcock et al. 2000b; Griest 1991). For the Large Magellanic Cloud (LMC) this gives
$\tau_{\text {lmc }}=4.7 f \times 10^{-7}$.
For the Small Magellanic Cloud (SMC), the S model gives $\tau_{\text {smc }} \sim 1.4 \tau_{\text {lmc }}$. For a flattened halo, one finds a smaller value, typically $\tau_{\mathrm{smc}} \sim \tau_{\mathrm{lmc}}$ (Sackett \& Gould 1993).

Magellanic stars can also be lensed by non-halo stars, either in the Magellanic Clouds or in the Milky Way disk. Lensing by disk stars is expected to have an optical depth of order $10^{-8}$ (Alcock et al. 2000b). The optical depth for lensing by "selflensing", i.e. lensing by stars in the Clouds, is expected to range from $\sim 5 \times 10^{-8}$ in the center of the LMC bar to $\sim 0.5 \times 10^{-8}$ at 3 deg from the bar (Mancini et al. 2004). For the SMC, the self-lensing optical depth is expected to be somewhat larger, $\sim 4 \times 10^{-8}$ averaged over the central $10 \operatorname{deg}^{2}$ (Graff \& Gardiner 1999).

Microlensing events are characterized by a timescale $t_{\mathrm{E}}$ giving the time for the lens to travel a distance corresponding to its Einstein radius, $t_{\mathrm{E}}=r_{\mathrm{E}} / v_{t}$ where $v_{t}$ is the lens's transverse velocity relative to the line of sight. For high amplification events, $2 t_{\mathrm{E}}$ is the time over which the amplification is $A>1.34$. Since $r_{\mathrm{E}}$ is proportional to the square root of the lens mass $M$, the mean $t_{\mathrm{E}}$ will scale like $M^{1 / 2}$. The S model has a 3 -dimensional macho velocity dispersion of $270 \mathrm{~km} \mathrm{~s}^{-1}$ and gives
$\left\langle t_{\mathrm{E}}\right\rangle \sim 70\left(\frac{M}{M_{\odot}}\right)^{1 / 2}$ days.
Much excitement was generated by the MACHO collaboration's measurement of the LMC microlensing rate which suggested that a significant amount of the Milky Way's Halo is comprised of machos. Their latest analysis (Alcock et al. 2000b) used 13/17 observed events ${ }^{2}$ to measure an optical depth of $\tau_{\text {lmc }} / 10^{-7}=$ $1.2_{-0.3}^{+0.4}$ (stat) with an additional $20 \%$ to $30 \%$ systematic error. This would correspond to a Halo fraction $0.08<f<0.50(95 \%$ $\mathrm{CL})$. The mean $t_{\mathrm{E}}$ of their events was 40 d corresponding to machos in the mass range $0.15 M_{\odot}<M<0.9 M_{\odot}$. On the other hand the EROS collaboration (Lasserre et al. 2000; Afonso et al. 2003a) has placed only an upper limit on the halo fraction, $f<$ 0.2 ( $95 \% \mathrm{CL}$ ) for objects in this mass range, ruling out a large part of the range of $f$ favored by the MACHO collaboration.

Bennett (2005) argued that the MACHO optical depth should be reduced to $\tau_{\mathrm{lmc}} / 10^{-7}=1.0 \pm 0.3$ in order to take into account

[^2]contamination by variable stars. This paper made use of the observation by the EROS collaboration (Tisserand 2004) of further variability of one of the MACHO A candidates, indicating intrinsic stellar variability. The paper also noted that the spectrum of the MACHO B candidate MACHO-LMC-22 indicated that the source is an active background galaxy, as reported in Alcock et al. (2001b) where the event was eliminated from the sample for studying high-mass lenses. Using four MACHO A candidates whose microlensing nature was confirmed by precision photometry and the one A candidate rejected as a variable star, Bennett (2005) performed a likelihood analysis to argue that $11 \pm 1$ of the 13 A candidates are likely to be microlensing events, yielding the revised optical depth.

Machos can also be searched for by monitoring M31 and looking for temporal variations of surface brightness consistent with a star in M31 being microlensed. Candidate events have been reported by the VATT (Uglesich et al. 2004), WeCAPP (Riffeser et al. 2003), POINT-AGAPE (Calchi Novati et al. 2005), MEGA (de Jong et al. 2006) and Nainital (Joshi et al. 2005) collaborations. The POINT-AGAPE and MEGA collaborations presented efficiency calculations allowing them to constrain the content of the M31 and Milky Way halos. The disagreement between these two collaborations parallels that between the MACHO and EROS collaborations with the AGAPE collaboration finding a halo fraction in the range $0.2<f<0.9$, while the MEGA collaboration finds a halo fraction $f<0.3$.

In this paper, we extend our previous analysis to find $\tau_{\text {lmc }}<$ $0.36 \times 10^{-7}(95 \% \mathrm{CL})$ for $M \sim 0.4 M_{\odot}$, corresponding to $f<0.08$. Unlike the previous EROS limit, this is significantly lower than the optical depth measured by the MACHO collaboration. Unlike all previous analyses, we use only a bright, wellmeasured subsample of the Magellanic stars, about $20 \%$ of the total. We believe that the use of this bright subsample gives more reliable limits on the optical depth than measurements using faint stars. There are two reasons for this. First, bright stars have well reconstructed light curves that permit discrimination of intrinsically variable stars. Second, the use of bright stars makes it relatively simple to estimate so-called blending effects where reconstructed fluxes can receive contributions from more than one star, complicating the interpretation of events.

EROS-2 is a second generation microlensing experiment. The first generation, EROS-1, consisted of two programs, both at the European Southern Observatory (ESO) at La Silla, Chile. The first program (Ansari et al. 1996a) used Schmidt photographic plates to monitor a $27 \mathrm{deg}^{2}$ region containing the LMC bar during the southern summer from October, 1990 through April, 1993. With a sampling frequency of up to one image per night, it was sensitive mostly to machos in the range $10^{-4} M_{\odot}<M<1 M_{\odot}$. The second program (Renault et al. 1998) used a $0.4 \mathrm{deg}^{2}$ CCD mosaic from December 1991 through March, 1995 to monitor one field in the LMC bar and another in the SMC. With up to 40 images taken per night, this program was sensitive mostly to machos in the range $10^{-7} M_{\odot}<M<$ $10^{-3} M_{\odot}$. The results of these two EROS-1 programs are summarized in Renault et al. (1997).

The second generation program described here, EROS-2, used the Marly 1 meter telescope at ESO, La Silla. The telescope was equipped with two $0.95 \mathrm{deg}^{2} \mathrm{CCD}$ mosaics to monitor $93 \mathrm{deg}^{2}$ in the Magellanic Clouds, $63 \mathrm{deg}^{2}$ in the Galactic Bulge, and $28 \mathrm{deg}^{2}$ in the spiral arms of the Milky Way. The observations were performed between July 1996 and February 2003 (JD between 2450300 and 2452700 ).

Besides the Magellanic results presented here, EROS-2 has also published measurements of the optical depth toward the

Galactic Bulge (Afonso et al. 2003b; Hamadache et al. 2006). The measured optical depth is in agreement with that measured by MACHO (Popowski et al. 2005), and OGLE-2 (Sumi et al. 2006) and with the predictions of Galactic models (Evans \& Belokurov 2002; Bissantz \& Gerhard 2002; Han \& Gould 2003; Wood \& Mao 2005).

EROS-1 and EROS-2 overlapped with the MACHO program that monitored $\sim 13.4 \mathrm{deg}^{2}$ of the LMC from July, 1992 through January, 2000. Three other wide-angle microlensing searches are now in operation: $\mathrm{MOA}^{3}$ (since August, 1998), OGLE-3 ${ }^{4}$ (since June, 2001) and SuperMACHO ${ }^{5}$ (since October, 2001).

In this article, we report on the analysis of the full EROS-2 data set (July 1996 till February 2003) toward the Magellanic Clouds. Our previous analyses reported on 5 years of data and 5 million stars toward the SMC (Afonso et al. 2003a), and 2 years of data and 17 million stars toward the LMC (Lasserre et al. 2000). An update of the LMC analysis was reported in Milsztajn et al. (2001) and Lasserre (2000), which dealt with 25 million stars and 3 years. The limits determined in the previous analyses are refined in the analysis presented here. More details about the present analysis can be found in Tisserand (2004).

The plan of the article is as follows. In Sect. 2, we recall the basics of the EROS-2 setup, give the general characteristics of the data sample and describe the data reduction steps used to produce the light curves. Section 3 presents the Bright-Star Sample of stars to be used in the measurement of the optical depth. In Sect. 4, we describe the selection criteria used to choose the microlensing candidates in the Bright-Star Sample. (Selection of events in the full sample is described in Appendix A.) Section 5 presents the final sample of selected events from the Bright-Star Sample as well as events found by relaxing the selection criteria. In Sect. 6 we discuss the status of former EROS-1 and EROS-2 microlensing candidates as well as those of the MACHO collaboration. In Sect. 7, we describe the computation of the EROS-2 detection efficiency. Section 8 presents the limit on the optical depth and on the abundance of machos in the Galactic halo by combining all EROS-1 and EROS-2 data. We conclude in Sect. 9 with a discussion of the significance of the limit, and a comparison with the results of Alcock et al. (2000b).

## 2. Experimental setup, observations and data reduction

The EROS-2 Marly telescope, camera, telescope operations and data reduction are described in Bauer et al. (1997), Palanque-Delabrouille et al. (1998), and references therein. Here we give only general information, and details or modifications that are specific to the present analysis.

### 2.1. The setup and the data

The Marly telescope is a one meter diameter Ritchey-Chrétien ( $f=5.14 \mathrm{~m}$ ), equipped with two wide angle CCD cameras. Each camera is a mosaic of 8 CCDs, 2 along right ascension and 4 along declination. Each CCD has $2048 \times 2048$ pixels of $15 \times 15 \mu \mathrm{~m}^{2}$ size, corresponding to $0.602 \times 0.602 \mathrm{arcsec}^{2}$. Images were taken simultaneously in two wide passbands, so-called $R_{\text {eros }}$ centered close to the $I_{\mathrm{C}}$ standard band, and $B_{\text {eros }}$ intermediate between the standard $V$ and $R$ bands. While no results presented here depend on the photometric calibration, almost all of

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Fig. 1. Map of the EROS-2 LMC and SMC fields in equatorial coordinates. A total of 88 LMC and 10 SMC fields were monitored. The first number in each field is the field number and the second is the number of bright stars (as defined in Sect. 3) in the field in units of $10^{4}$. The two shaded regions (the larger one centered on the LMC bar) are the $13.4 \mathrm{deg}^{2}$ used by the MACHO collaboration to measure the optical depth (Alcock et al. 2000b).
our fields could be calibrated using stars from the catalogs of the Magellanic Clouds Photometric Survey (Zaritsky et al. 2004). For $4.5 \mathrm{deg}^{2}$, the calibration was checked with the OGLE-II cata$\log$ (Udalski et al. 2000b). To a precision of $\sim 0.1 \mathrm{mag}$, the EROS magnitudes satisfy
$R_{\text {eros }}=I_{\mathrm{C}} \quad B_{\text {eros }}=V_{\mathrm{J}}-0.4\left(V_{\mathrm{J}}-I_{\mathrm{C}}\right)$.
The observed fields correspond to $0.95 \mathrm{deg}^{2}$ each and they are twice as large in declination than in right ascension ( $1.38 \times$ $0.69 \mathrm{deg}^{2}$ ). A total of 88 fields have been monitored toward the LMC and 10 toward the SMC. The positions of the fields are shown in Fig. 1. The exposure times ranged from 180 s to 900 s . The fields have been sampled differently, according to their stellar density and distance from the optical centers of the LMC and the SMC. Broadly, there are three families of LMC fields: in the first 3 years of operation, 22 outer fields were seldom imaged, 22 fields were imaged about 100 times and the remaining 44 inner fields were imaged over 200 times. Later, from July 1999 on, all 88 LMC fields were imaged with a similar sampling. The number of photometric measurements, per star and per band, ranges from 300 to 600 in the LMC fields.

The ten SMC fields were imaged with a similar sampling, except in 2001 and 2002 when the inner six fields were imaged twice as often as the four outer ones. The number of photometric measurements, per star and per band, ranges from 600 to 900 in these SMC fields.

After rejection of bad images (11\%) due to bad seeing, high sky background or instrumental problems, the average numbers of measurements used in the analysis, per star and per band, are about 430 and 780 in the LMC and SMC fields respectively.

### 2.2. Template images

The production of light curves proceeded in three steps: template image construction, star catalog production from the templates, and photometry of individual images to obtain the light curves.

The template images were obtained by co-adding 15 images of each field; they were resampled so that the templates have twice as many pixels as the original images. The 15 images were chosen among the best ones available, i.e. with low sky background, good seeing, and with a large number of stars (as estimated from a quick first look algorithm). All images susceptible to enter the template construction were checked for the absence of long tracks, caused by satellites, planes or meteors. For technical reasons linked to computing and to PSF variation within one CCD, the CCD images were divided in four quadrants, such that there were in total 6272 template images ( 98 fields, 8 CCDs, 4 quadrants, 2 passbands). In order to ensure relatively uniform zeropoints of the 6272 templates, we required that the first image used in template construction (to which the other 14 images were photometrically aligned) be registered within a short time interval with good and uniform sky conditions (23rd to 26th, November 2000). For each field, the same epochs were used in the construction of the templates in the two passbands.

The stars were identified on the template images using a pseudo-image that we call a correlation image. Each pixel of this image contains the correlation coefficient of neighboring pixels of the template itself with a two-dimensional Gaussian PSF. Each group of pixels satisfying some threshold value on the pseudo-image was retained as a star in the catalog. In previous EROS-2 analyses, the thresholds were identical for all templates; this had led to over $20 \%$ failures in this cataloging step. The present analysis has chosen to progressively relax the thresholds when such failures occur; in this way, the cataloging step failures were drastically reduced. (The number of identified stars on fields using relaxed thresholds was lower on average.) The star catalogs obtained from the template images in the two passbands were then merged. A star was retained only if it was detected in both. The overall efficiency of the template plus star catalog construction was excellent; only 24 CCD quadrants could not be processed out of a total of 3136 .

Examples of color-magnitude diagrams can be found in Figs. 7 to 10 and A. 1 to A.5. They are all characterized by a prominent group of clump giants and a main sequence whose relative strength varies from field to field. There are also stars that are redder than Magellanic red giants. Most of them are likely to be foreground stars in the disk of the Milky Way and their number is consistent with the predictions of Galactic models (Robin et al. 2003).

### 2.3. Light curves

Photometry was then performed on each image of a given quadrant in turn with software specifically designed for crowded fields, PEIDA (Photométrie et Étude d'Images Destinées à l'Astrophysique) (Ansari et al. 1996b). First, the image was geometrically aligned with the template. Then, imposing the star position determined from the template image, PSF-fitting photometry was performed for all stars by a linear least-squares method involving the star and all neighbors closer than 11 pixels, plus a sky background. An estimate of the error on this flux measurement was computed that depends on the photon statistics and on the overall quality of the photometered image. Typical photometric precisions are shown in Fig. 2.


Fig. 2. The photometric precision as a function of $R_{\text {eros }}$ for the dense field $\operatorname{lm} 009$ (top) and the sparse field $\operatorname{lm} 048$ (bottom). Each point represents the rms dispersion of the measurements for a single star after elimination of outliers. The vertical line shows the position of the bright-star magnitude cut (5) for the field, $R_{\text {eros }}=18.23$ for 1 m 009 CCDs $0-3$ and $R_{\text {eros }}=19.7$ for $\operatorname{lm} 048$ CCDs 0-3.

Before the analysis was started, we removed from the light curves measurements taken under far from normal conditions. This happened not infrequently, as the data taking policy was to work whenever possible, leaving to the analysis the task of rejecting these abnormal measurements. These were identified by extreme values of the sky background, seeing or absorption. In addition, images where the photometry failed for over $40 \%$ of cataloged stars were eliminated, as well as images for which more than $12 \%$ of the stars showed an excursion from their average flux larger than three standard deviations. Depending on the CCD, the number of rejected images varied between 7 and $18 \%$, with an average of $11 \%^{6}$.

To reduce systematic errors in the photometry, each light curve was searched for significant linear correlations between the measured flux and three observational variables, the seeing, the hour angle and the airmass. This was done independently in the two passbands. The measured fluxes were corrected linearly by requiring a vanishing correlation coefficient between the corrected fluxes and the given variable. The largest correlation was found with the seeing, in both passbands. For the bright stars considered in this work, the correction has only a small influence on the point-to-point dispersion of the lightcurves, reducing it on average by $6 \%$.

[^4] for these 5 CCDs.

A potential problem with this correction is that an artificial correlation can be induced if a real flux variation happens to occur during a period of poor seeing (for example). In this case, application of the correction would reduce the amplitude of the real flux variation but increase the point-to-point dispersion of the curve by making an incorrect flux correction. To guard against this possibility, we did not apply the correction if it increases the point-to-point dispersion of the light curves.

### 2.4. The full sample of stars

A total of 58.4 million objects were found on the template images of both passbands - 51.8 million in the LMC and 6.6 million in the SMC. We chose to ignore those stars for which the association between the objects detected separately in the two passbands was doubtful or ambiguous ( 8.4 million) and objects that are dimmer than about twice the typical size of sky background fluctuations ( 14.9 million). We rejected light curves for which more than half of the photometry points are absent ( 1.0 million). Finally, we did not consider stars whose photometry is unstable due to its environment. This includes stars close to a very bright field star ( $V_{\mathrm{J}}<10.1$, probably in the Galactic disk) and stars close to a visible diffraction feature in the PSF of bright stars; these two categories contain respectively $1.1 \%$ and $0.8 \%$ of the remaining stars.

We removed from consideration stars in field $\operatorname{lm} 003$, CCD0 which has numerous events caused by light echos from SN1987a. The echos generate arc-like images that appear to move a few arcsec per year ${ }^{7}$, causing false variations of a star's flux as the arc passes through the star's position.

After these cuts, 33.4 million objects remained, 29.2 million in the LMC and 4.2 million in the SMC. This constitutes the Full Sample of EROS-2 stars. In the next section we describe the selection of the Bright-Star Sample used for the measurement of optical depths.

## 3. The bright-star sample

We have chosen to restrict our analysis to the Bright Sample of stars defined below. Besides the obvious advantage of ensuring a good photometric resolution, we do this to simplify the evaluation of the number of expected events predicted by a model. In crowded fields, this evaluation is complicated by "blending", i.e. the fact that photometry of a given object can be influenced by more than one star. We shall see that these effects are rather small and simple to evaluate for the Bright-Star Sample (which should, strictly speaking, be called the Bright-Object Sample).

Limiting the number of studied objects limits the sensitivity of the experiment so the magnitude cut must be a compromise between quantity and quality of objects. The efficiency calculations of Sect. 7 for unblended sources indicated that keeping only the $\sim 20 \%$ brightest stars reduces by only $\sim 50 \%$ the number of simulated microlensing events that pass our selection criteria. We therefore initially decided to accept a nominal $\sim 50 \%$ loss of events by requiring $R_{\text {eros }}<R_{\text {median }}$, where $R_{\text {median }}$ is the median $R_{\text {eros }}$ for simulated unblended microlensing events passing our event selection criteria. $R_{\text {median }}$ ranges from $\sim 18.2$ in the densest fields to $\sim 20.5$ in the sparsest. However, for sparse fields far from the LMC bar, in order to have a reasonable object

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Fig. 3. The flux precision, $\Delta \phi / \phi$, for the Bright-Star Sample ( $2 \%$ of the sample are in the histograms). Each entry is the mean flux uncertainty for a star on its light curve. The top (bottom) histograms are for the $R_{\text {eros }}$ ( $B_{\text {eros }}$ ) bands.
reconstruction efficiency, a stricter cut was found to be necessary, $R_{\text {eros }}<19.7$.

The final sample of bright stars is therefore defined by

$$
\begin{equation*}
16.0<R_{\text {eros }}<R_{\max } \quad R_{\max }=\min \left(R_{\text {median }}, 19.7\right) \tag{5}
\end{equation*}
$$

The minimum magnitude $R_{\text {eros }}=16$ was chosen to avoid the large number of variable stars brighter than this.

The position of the magnitudes cut are shown in the color-magnitude diagrams of Figs. 7 to 10 and A. 1 to A. 5 . Generally speaking, the cut includes clump giants but not the numerous main sequence stars seen far below the clump. Other than the small number of bright main sequence stars, we therefore employ stars of colors and magnitudes similar to those used in Galactic Bulge measurements that use clump giants (Hamadache et al. 2006).

It turns out that the cut (5) gives a rather uniform photometric precision for the Bright-Star Sample. Figure 3 shows the distribution of precision with a mean in both bands of $\sim 7 \%$. The precision in dense fields hardly differs from that in sparse fields and the precision in the LMC hardly differs from that in the SMC.

The number of EROS-2 objects in the Bright-Star Sample defined by (5) is 6.05 million in the LMC and 0.90 million in the SMC. We must subtract from these numbers the expected number of foreground stars in the Milky Way disk. The Besançon model of the Milky Way (Robin et al. 2003) predicts that 0.56 million (LMC) and 0.04 million (SMC) are foreground Milky Way stars, consistent with the number of stars we observe that are redder than Magellanic red giants. The number of

Bright-Sample stars to be used for optical depth measurements must therefore be reduced by $9 \%$ ( $4 \%$ ) for the LMC (SMC), giving
$N_{\text {objects }}=5.49 \times 10^{6} \quad(\mathrm{LMC}) ;$
$N_{\text {objects }}=0.86 \times 10^{6} \quad(\mathrm{SMC})$.
We now show how we evaluate the effect of blending on the Bright-Star Sample. For a given lens model (defined by the spatial, velocity and mass distribution of lenses), the expected number of detected events is
$N_{\mathrm{ex}}=\frac{2}{\pi} \frac{T_{\mathrm{obs}}}{\left\langle t_{\mathrm{E}}\right\rangle} \tau \sum_{i=1}^{N_{\text {sars }}}\left\langle\epsilon_{i}\right\rangle$
where $T_{\text {obs }}$ is the observing period, $\left\langle t_{\mathrm{E}}\right\rangle$ is the mean Einstein radius crossing time given by the model, e.g. (3), and $\tau$ is the optical depth. The sum is over all observed stars and the "efficiency", $\left\langle\epsilon_{i}\right\rangle$, is the expected ratio between the number of events passing the selection cuts of Sect. 4 and the number of events with $u<1$ occurring during the observation period. The efficiency must be averaged over the $t_{\mathrm{E}}$ distribution given by the model.

EROS photometry is performed on a set of "objects" found on a reference image. In crowded fields, the correspondence between objects and stars is not straightforward. Generally speaking, most objects are dominated by one star, i.e. by the brightest star in the object's seeing disk. However, if a fainter star in the seeing disk is microlensed, EROS photometry will assign some of the extra flux to the object with the remainder assigned to sky background. It is thus convenient to rewrite (8) as
$N_{\mathrm{ex}}=\frac{2}{\pi} \frac{T_{\text {obs }}}{\left\langle t_{\mathrm{E}}\right\rangle} \tau \sum_{j=1}^{N_{\text {objects }}} \sum_{i}\left\langle\epsilon_{i j}\right\rangle$
where the second sum is over the stars $i$ near the object $j$. The efficiency, $\left\langle\epsilon_{i j}\right\rangle$, is again the expected ratio (averaged over $t_{\mathrm{E}}$ ) between the number of microlensings of star $i$ that pass selection cuts and the number of microlensings of star $i$ with $u<1$ occurring during the observation period. Since we do not know the stars in the seeing disk of each object, we must estimate the sum statistically using knowledge of the luminosity function. We will see that sum is dominated by the brightest star, $i=1$, with contributions from the second brightest, $i=2$, of order $10 \%$.

If there were no blending, the efficiency could be estimated by modifying a sample of light curves with measured fluxes $F_{0}(t)$ by
$F(t)=F_{0}(t) \frac{u^{2}+2}{u \sqrt{u^{2}+4}}, \quad u(t)^{2}=u_{0}^{2}+\left[\left(t-t_{0}\right) / t_{\mathrm{E}}\right]^{2}$,
where $u(t)$ is the distance of the lens to the star's line-of-sight in units of the Einstein radius, $u_{0}$ is this distance at the time, $t_{0}$, of maximum amplification, and $t_{\mathrm{E}}$ is the Einstein ring radius crossing time. The modified lightcurves can then be subjected to the selection criteria and the efficiency deduced by averaging over $t_{\mathrm{E}}$ and $u_{0}$.

For blended events only a fraction $\alpha$ of an object's flux is amplified. If the fraction is independent of amplification, the resulting light curve is
$F(t)=F_{0}(t)\left[(1-\alpha)+\alpha \frac{u^{2}+2}{u \sqrt{u^{2}+4}}\right]$.

The efficiency to see such events now depends on $\alpha: \epsilon_{i j}=\epsilon_{i j}\left(\alpha_{i}\right)$. Stars with small $\alpha_{i}$ require a small impact parameter $u_{0}$ to give a sufficiently large reconstructed amplification. Since our selection criteria require a reconstructed amplification greater than $1.34\left(u_{0}=1\right.$ for $\left.\alpha=1\right)$, the primary effect of blending is to reduce the efficiency by a factor $u_{i}\left(\alpha_{i}\right)$, the impact parameter needed for star $i$ to produce a reconstructed amplification of 1.34 . The efficiency is reduced by this factor because of the flat a priori $u$ distribution. Additionally, the event appears shorter by a factor $u_{i}$ further modifying the efficiency.

A simple example is two superimposed stars of the same magnitude and color. In this case $\alpha_{1}=\alpha_{2}=0.5$ and $u_{1}=$ $u_{2} \sim 0.7$. Before efficiency corrections, the event rate would be $0.7+0.7=1.4$ times the rate calculated ignoring blending. If the two stars have colors differing by $\Delta(B-R)_{\text {eros }}=0.7$, the requirement that $A_{\max }>1.34$ in both colors yields a rate that is increased by a factor 1.3 over the rate for $\alpha=1$.

To statistically evaluate the distribution of the $\alpha_{i}$ for BrightSample objects, a spectrum of 8000 artificial stars was placed at random positions on real EROS images $(\operatorname{lm} 009, \operatorname{lm} 019$, $\operatorname{lm} 034, \operatorname{lm} 048$ ) to give artificial images ( $\mathrm{xm} 009, \mathrm{xm} 019, \mathrm{xm} 034$, xm 048 ). These fields were chosen as representative of the crowding variations over the LMC fields. The stars were given fluxes according to randomly chosen microlensing events (one per artificial star). Reference xm images were created using these images and photometry performed in the same manner as for the normal (lm) images.

After photometry, the light curves of xm objects falling near artificial stars were studied to find values of the $\alpha_{i}$. In practice, one median value of $\alpha$ was calculated for each lightcurve using points for which $(A-1)_{\text {input }}>0.5$, though no nonlinearity was observed that made this a crucial point. Of the 8000 artificial stars that were added to each of the four fields, most are not usable for various (understandable) reasons: star in a masked region of the CCD, event during a period with no observations, or the reconstructed magnitude not satisfying the Bright Sample cuts. For the field xm009, a total of 1123 lightcurves were usable for determining $\alpha$. Of these lightcurves, 982 concerned objects for which the artificial star is the primary component, while in the remaining 141 objects the artificial star falls underneath a preexisting bright object on the original image. We use the first type of object to determine the distribution of $\alpha_{1}$, the value of $\alpha$ for the primary star associated with each object.

Figure 4 (top) shows the distribution of $\alpha_{1}$ for the dense field $\operatorname{lm} 009$ and the sparse field $\operatorname{lm} 048$. The other two fields give similar distributions. The distributions of $\alpha_{1}$ are characterized by peaks at $\alpha_{1} \sim 1$ due to artificial stars falling more than $\sim 2 \operatorname{arcsec}$ from any pre-existing $\operatorname{lm}$ object. This happens for about half the artificial stars in the densest field (lm009) so blending has little effect on about half the bright stars in dense fields. Stars falling on pre-existing stars yield the tail at $\alpha_{1}<1$. This can be seen in the middle plots that show, for artificial stars within 2 arcsec of pre-existing stars, the correlation between $\alpha_{1}$ and the magnitude difference between artificial and pre-existing stars. The distribution follows closely the expected form $\alpha_{1}=1 /\left(1+f_{2} / f_{1}\right)$ where $f_{1}$ is the flux of the artificial star and $f_{2}$ is the flux of the preexisting star. For example, when $\alpha_{1} \sim 0.5$ the cloud of points passes near $\Delta R \sim 0$.

The mean value of $\alpha_{1}$ is $\sim 0.86$ equivalent to a mean value of $u_{1}$ of 0.92 . This means that the (efficiency uncorrected) rate of events with amplifications of the primary component greater than 1.34 is reduced by a factor 0.92 from the rate calculated assuming no blending.


Fig. 4. The effects of blending on artificial stars added to the dense field $\operatorname{lm} 009$ (left) and the sparse field lm048 (right). The top two panels show the distribution of $\alpha_{1}$ using artificial stars that do not fall under a brighter pre-existing object. Stars further than 2 arcsec from the nearest preexisting object are concentrated in the peak at $\alpha_{1} \sim 1$ while stars nearer to pre-existing objects form the tail at $\alpha_{1}<1$. The middle two panels show, for objects within 2 arcsec of pre-existing objects, the correlation between $\alpha_{1}$ and the difference in magnitude between the artificial star and the pre-existing object. Artificial objects with $\alpha_{1}$ significantly below unity are associated with pre-existing objects of similar magnitude. The line shows the expected relation for two superimposed objects: $\alpha_{1}=1 . /\left(1 .+10^{0.4 \Delta R}\right)$. The bottom two panels show the distribution of ( $\alpha_{1}, \alpha_{2}$ ) with $\alpha_{2}$ calculated from (12) and the magnitude of the pre-existing object. The line shows the expected relation for two superimposed objects: $\alpha_{1}+\alpha_{2}=1$.

The loss of $8 \%$ of the events is compensated by extra events due to stars under the primary component. Since $\alpha_{1} \sim 1 /(1+$ $\left.f_{2} / f_{1}\right)$, we can expect that $\alpha_{2} \sim 1 /\left(1+f_{1} / f_{2}\right)$ for superimposed objects. However, the superposition is not perfect so $\alpha_{2}$ must be a decreasing function of the separation between the object and star 2. The artificial events due to artificial stars falling under brighter pre-existing objects are well described by
$\alpha_{2}(d) \sim \frac{f_{2}}{f_{1}+f_{2}} \exp \left[-\left(d / d_{0}\right)^{2}\right] \quad d_{0}=1.9 \operatorname{arcsec}$
where $f_{1}$ is the flux of the major component, $f_{2}$ is the flux of the minor (amplified) component and $d$ is the separation between the xm object and the minor component. This formula can then be applied to the objects where the artificial star is the major component since the flux and position of the minor component (on the original image) is known. In the bottom panels of Fig. 4 we show $\alpha_{l}$ vs. $\alpha_{2}$ for the fields $\operatorname{lm} 009$ and $\operatorname{lm} 048$. The anticorellation between $\alpha_{l}$ and $\alpha_{2}$ means that the loss of events due to $\alpha_{1}<1$ is partially compensated by events from the second star on an object by object basis.


Fig. 5. The distribution of $u_{1}$ (solid line) and $\left(u_{1}+u_{2}\right)$ dashed line for the sparse EROS field lm048 (top panel) the dense EROS field lm009 (middle panel) and the very dense HST field in $\operatorname{lm} 009$, ccd3 (bottom panel) as described in the text. The distributions of $u_{1}$ are all characterized by a peak near $u_{1} \sim 1$ and a tail at $u_{1}<1$. With increasing star density, the mean of $u_{1}+u_{2}$ increases because of the increasing importance of secondary stars.

The compensation is seen in Fig. 5 histogramming $u_{1}$ and $u_{1}+u_{2}$. The sum $u_{1}+u_{2}$ gives the rate for event with amplifications $>1.34$ compared to that calculated assuming $\alpha=1$. In $\operatorname{lm} 048$, events due to the second star compensate for the loss of events on the first star while in the dense field $\operatorname{lm} 009$ there is a $10 \%$ overcompensation. As we will see in Sect. 7, the efficiency to observe the blended events is degraded compared to unblended events, so the overall number of expected events is about $10 \%$ less than what one would calculate ignoring blending.

We have checked our calculations by studying public HST images of a small part of the densest EROS region, that of lm009 ccd 3. A catalog of stars for this region was produced and images with EROS seeing were fabricated. As in the previous analysis, images with some amplifications were compared with reference images. Figure 5 shows the distribution of $u_{1}$ for bright objects on the convoluted images. Because of the very high density of stars, $u_{1}$ is somewhat smaller than that calculated with the artificial images: $\left\langle u_{1}\right\rangle=0.85$ as opposed to $\left\langle u_{1}\right\rangle=0.92$. This lower value of $u_{1}$ is, as expected, compensated by higher values of $u_{2}$. Figure 5 shows the histogram of $u_{1}+u_{2}$ indicating that the event rate is raised by $15 \%$ over the unblended rate. Including the loss of efficiency modifies this so that the event rate is only 5\% higher than the unblended rate.

## 4. Candidate selection from the bright-star sample

The present analysis aims at detecting luminosity excursions, due to microlensing, on otherwise constant light curves. The smallest reachable microlensing timescale, $t_{\mathrm{E}}$, is determined by the sampling of the fields; in practice, except for a few fields, the detection efficiency is very low below 2-3 days. The largest detectable timescale, about 800 days, is determined by the reduction in the detection efficiency when the timescale becomes so large that the baseline flux is not seen during the observing period.

The analysis is guided using the simulation of microlensing described in Sect. 7, which also serves to determine the efficiency of the selection procedure.

In the present analysis, luminosity excursions are defined with respect to the baseline stellar flux; the first task is thus to determine this baseline. To that end, we order the $N$ measured fluxes $\Phi_{i}$ of the light curve by increasing values of $\Phi$. For each of the $N-1$ middle values of the flux intervals $\Phi_{i, \text { mid }}=$ $0.5\left(\Phi_{i}+\Phi_{i+1}\right)$, we count the number $N_{\mathrm{run}, i}$ of runs, i.e. groups of consecutive points on the light curve that are on the same side of $\Phi_{i, \text { mid }}{ }^{8}$. The baseline flux is defined as that value of $\Phi_{i, \text { mid }}$ which maximizes $N_{\text {run }, i}$. This way of defining the baseline has a precision similar to that of a simple average of the fluxes, but it proves much more robust to aberrant measurements, and less biased for most variable stars. We have checked on the simulated microlensing light curves that there is no visible bias for timescales $t_{\mathrm{E}}$ shorter than $\sim 200 \mathrm{~d}$. The bias is non-negligible above 600 d , but this is taken into account in the efficiency calculation.

Next, we determine the point-to-point dispersion in the light curve $\sigma_{\mathrm{ptp}}$, from the comparison of each measured flux with the linear interpolation of its two neighboring (in time) fluxes. This is done separately for the 7 seasons of data taking, as we have observed a progressive degradation of the photometric scatter in the last 3 seasons.

Using the baseline flux and the photometric scatter, we then search for luminosity excursions defined as a group of consecutive points with fluxes sufficiently far from the baseline flux. These should have at least 5 consecutive measurements more than $1.5 \sigma_{\text {ptp }}$ from the baseline. In practice, to increase the detection efficiency for short duration phenomena, we allow small "holes" within the excursion (series of points less than $1.5 \sigma_{\mathrm{ptp}}$ from the baseline) provided that each hole contains exactly one measurement. We call positive (negative) excursions those comprised of fluxes higher (lower) than the baseline.

[^6]Table 1. The number of light curves surviving the cuts $C 1$ to $C 12$ as described in the text.

| Cut | LMC | SMC |
| :--- | ---: | ---: |
| Bright-Star Sample | 6052102 | 900809 |
| Filter $(C 1, C 2, C 3)$ | 33876 | 5787 |
| $L P_{N, \text { pos,, }, 2} / L P_{N, \text { pos }, 2}>10(C 4)$ | 6778 | 847 |
| good $\chi_{\text {base }}^{2}(C 5)$ | 2636 | 266 |
| $t_{0}(C 6)$ | 2065 | 195 |
| $N_{\text {peak }}>4(C 7)$ | 1624 | 149 |
| significant $\Delta \chi^{2}(C 8)$ | 350 | 32 |
| good $\chi_{\text {peak }}^{2}(C 9)$ | 152 | 17 |
| BlueBumper $(C 10)$ | 10 | 4 |
| SN $(S<0.3)(C 11)$ | 7 | 4 |
| $u_{0}<1(C 12)$ | $1(\mathrm{SN})$ | 1 |

The relative significance of each excursion is then estimated using the probability that it corresponds to a statistical fluctuation of a normal law. We use the variable $L P_{N}$, the co-logarithm of this probability
$L P_{N}=-\sum_{i=1}^{i=N} \log \left(\frac{1}{2} \operatorname{Erfc}\left(\frac{x_{i}}{\sqrt{2}}\right)\right)$,
where $x_{i}$ is the deviation from the baseline of the point taken at time $t_{i}$, in units of its error $\sigma_{i}$, and $N$ is the number of points within the excursion. The excursions are ranked by decreasing values of $L P_{N}$ (decreasing significance).

We now describe the 12 selection criteria, $C 1-C 12$, used to select microlensing candidates. $C 1$ requires at least one significant positive excursion:

$$
C 1: \quad L P_{N, \text { pos }, 1}>20 .
$$

The largest negative excursion should be much smaller than the largest positive one:

$$
C 2: \quad L P_{N, \mathrm{pos}, 1} / L P_{N, \mathrm{neg}, 1}>10 .
$$

There should be less than 10 excursions in total:
C3: $\quad N_{\text {excursions }}<10$.
Cuts $C 1, C 2$ and $C 3$ are applied independently in the two passbands and a star is retained for further analysis only if it passes the three cuts in both bands. There are 33876 LMC light curves and 5787 SMC light curves that survive. The progression of the number of surviving light curves with the subsequent cuts is shown in Table 1.

All light curves passing $C 1-3$ are fitted with a simple microlensing curve (10), independently in the two passbands. The results of the fits, i.e. the fitted parameters $t_{0}, u_{0}$ and $t_{\mathrm{E}}$ and the values of $\chi^{2}$, are used in cuts C5-C10.

Cuts $C 4$ and $C 5$ eliminate light curves with significant variability outside the main positive excursion. We first require that the second positive excursion be much less significant than the first one:

## $C 4: \quad L P_{N, \mathrm{pos}, 1} / L P_{N, \mathrm{pos}, 2}>10$.

To discriminate against low amplitude variations (either real or due to photometric problems) we require that the normalized $\chi_{\text {base }}^{2}$ for a time independent light curve outside the main excursion (points with $u>2$ ) be sufficiently good:
$C 5: \quad \frac{\chi_{\text {base }}^{2}-N_{\text {dof,base }}}{\sqrt{2 N_{\text {dof,base }}}}<20$.


Fig. 6. The $u_{0}$ (mean of those for $R_{\text {eros }}$ and $B_{\text {eros }}$ ) distribution for LMC and SMC light curves surviving cuts $C 1-C 7, C 1-C 8$, and $C 1-C 9$. The positions of the 14 light curves surviving cuts $C 1-C 10$ are shown with the 4 supernovae as dashed lines.

This cut may seem loose as it would correspond to cuts at 20 standard deviations, for a normal law. However, the errors we use to compute $\chi_{\text {norm }}^{2}$ are only estimates of the actual errors. For isolated stars, the errors are overestimated. The values of the cuts have thus been tuned using the simulated light curves, that have the same uncertainties as the original data; the two cuts reject about $10 \%$ of the simulated microlensing light curves.

Cuts $C 6$ and $C 7$ require that the main fluctuation occur at a time when the light curve is sufficiently well sampled. In order to discriminate against stellar variations with very asymmetric peaks we thus require that the time $t_{0}$ be within the observing interval:

C6: $\quad T_{\text {beg }}+t_{\mathrm{E}}<t_{0}<T_{\text {end }}-t_{\mathrm{E}}$,
where $T_{\text {beg }}$ and $T_{\text {end }}$ are the first and last date of EROS-2 observations. We also require that there be more than 4 flux measurements within the main fluctuation $(u<2)$

C7: $\quad N_{\text {peak }}>4$.
There are 1624 LMC light curves and 149 SMC light curves that survive cuts $C 1-C 7$. Their distribution of $u_{0}$ and those of light curves satisfying subsequent cuts are shown in Fig. 6. It is seen that most light curves show only low amplitude variations with $u_{0}>0.6$.

We next require that the microlensing fit to the light curves in both passbands be significantly better than the fit of a constant flux:
$C 8: \quad \frac{\chi_{\mathrm{ct}}^{2}-\chi_{\mathrm{ml}}^{2}}{\chi_{\mathrm{ml}}^{2} / N_{\text {dof }}} \frac{1}{\sqrt{2 N_{\text {dof,peak }}}}>40$,
where $\chi_{\mathrm{ct}}^{2}$ and $\chi_{\mathrm{ml}}^{2}$ are the chi-squared values for the constant fit and the simple microlensing fits, and $N_{\text {dof }}$ and $N_{\text {dof,peak }}$ are the number of measured points in the full light curve and in the peak. The event is required to pass this cut in both passbands. This cut eliminates light curves that have low amplifications, poor sampling, and/or poor photometric errors.

We then require a reasonably good fit to the microlensing curve within the peak:
$C 9: \quad \frac{\chi_{\text {peak }}^{2}-N_{\text {dof,peak }}}{\sqrt{2 N_{\text {dof, peak }}}}<10$.


Fig. 7. The light curves of EROS-2 microlensing candidate EROS2-SMC-1 (star sm005-4m-5761). Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.

There are 152 LMC and 17 SMC light curves that pass cuts $C 1-C 9$. As can be seen in Fig. 6, almost all of these light curves have $u_{0}>0.8$. They are mostly bright main-sequence stars that display low-amplitude variations that are difficult to distinguish from microlensing, though they often show $\sim 20 \%$ more variation in the red than in the blue. They were first mentioned in Alcock et al. (1997a) and are commonly called "blue bumpers".

We identify blue bumpers first from their position in the color-magnitude diagram and their low, chromatic amplification. For stars with $(B-R)_{\text {eros }}<0.2$ and $R_{\text {eros }}<18.6$ (LMC) or $R_{\text {eros }}<18.9$ (SMC), we require
$C 10: \quad \min \left(A_{R}, A_{B}\right)>1.6 \quad$ and $\quad \frac{A_{R}-1}{A_{B}-1}<1.2$,
where $A_{R}$ and $A_{B}$ are the maximum amplifications observed in the two bands. These two amplifications are taken to be the fitted amplifications at the time of the maximum observed amplification. This value is used since the fitted amplification at $t=t_{0}$ may significantly overestimate a blue bumper amplitude if the peak region is not well sampled.

There are 10 LMC light curves and 4 SMC light curves that pass cuts $C 1-C 10$. Four of these light curves are most likely supernovæ (SN) exploding in galaxies far behind the Magellanic Clouds. In the analysis of the Full Sample of EROS-2 Magellanic light curves (Tisserand 2004), 31 such supernovae were found. All SN in our sample have fitted timescales in the range $25-50$ days, asymmetric light curves with a faster rise time, and larger variations in the bluer passband. For a fraction of them, about $20 \%$, the host galaxy is visible, which makes them indisputable SN. In order to identify the remaining SN, we have devised a fitting function with a time asymmetry parameter $S$; the function reduces to simple microlensing for $S=0$. The fitting function is obtained from
simple microlensing (10) by replacing the Einstein timescale $t_{\mathrm{E}}$ by the "supernova" varying timescale
$t_{\mathrm{S}}=t_{\mathrm{E}}\left[1+S \arctan \left(\frac{t-t_{0}}{t_{\mathrm{E}}}\right)\right]$.
For positive (negative) values of $S$, the rise time is faster (slower) than the decrease time. When fitting this function to the remaining light curves, we expect to find $S>0$ for the SN . Based on the study of the supernovae found in the Full Sample analysis and on simulated microlensing events, we choose to eliminate light curves with blue-band asymmetries by requiring

C11: $\quad S_{b}<0.30$.
This cut eliminates 3 light curves from the Bright-Star Sample, leaving 11 light curves, mostly with very small amplitudes, $u_{0}>$ 1. Since the $u_{0}$ distribution for microlensing events is flat, the events with $u_{0}>1$ in Fig. 6 must be mostly background. We therefore require
C12: $\quad u_{0}<1.0$.
This leaves two light curves. One of them is superimposed on a clear background galaxy. It is therefore most likely a supernova and we eliminate this event leaving one event to be discussed in the next section.

## 5. Candidate sample and verifications

After the cuts described in the previous two sections have been applied, only one candidate microlensing event remains in the Bright-Star Sample, EROS2-SMC-1. Its light curve is shown in Fig. 7 and its characteristics are given on the first entry in Table 2. It has been known since 1997 (Alcock et al. 1997b; Palanque-Delabrouille et al. 1998) and is identical to candidate MACHO-97-SMC-1. It has one of the two longest

Table 2. Events found in the analysis of the full sample of EROS-2 stars (Tisserand 2004). Only event EROS2-SMC-1 is in the Bright-Star Sample and passes the cuts $C 1-C 12$ and therefore is used for limits on the optical depth. All fits assume zero blending and no intrinsic variability. The two values of the fitted parameters $t_{0}, t_{\mathrm{E}}$ and $u_{0}$ are for $B_{\text {eros }}$ (first line) and $R_{\text {eros }}$ (second line). (There are no $R_{\text {eros }}$ points during the event for EROS2-LMC-12.) The time of maximum, $t_{0}$ is given as JD-2450000.

| Candidate | EROS-2 star | $\mathrm{RA}(\mathrm{deg})$ | $\mathrm{Dec}(\mathrm{deg})$ | $R_{\text {eros }}$ | $(B-R)_{\text {eros }}$ | $t_{0}$ | $t_{\mathrm{E}}$ | $u_{0}$ | $\chi^{2} /$ d.o.f. |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| EROS2-SMC-1 | sm005-4m-5761 | 15.0233 | -72.2507 | 18.14 | -0.13 | 459.0 | 105.8 | 0.53 | $972.4 / 873$ |
|  |  |  |  |  |  | 460.5 | 101.6 | 0.52 | $862.4 / 748$ |
| EROS2-SMC-5 | sm006-6k-19475 | 12.6625 | -72.6983 | 17.83 | 0.34 | 1414.8 | 29.5 | 0.90 | $485.1 / 875$ |
| EROS2-LMC-8 | $\operatorname{lm} 055-7 m-23303$ | 76.6396 | -71.6624 | 20.34 | 1.16 | 1432.0 | 15.1 | 1.12 | $560.0 / 573$ |
|  |  |  |  |  |  | 1594.1 | 15.3 | 0.01 | $344.0 / 572$ |
| EROS2-LMC-9 | $\operatorname{lm} 042-11-2622$ | 75.3717 | -65.0486 | 20.93 | 0.23 | 1594.1 | 8.0 | 0.03 | $540.9 / 510$ |
|  |  |  |  |  |  | 2204.1 | 57.8 | 0.36 | $640.8 / 369$ |
| EROS2-LMC-10 | $\operatorname{lm} 045-5 n-26323$ | 80.9925 | -65.8733 | 20.55 | 0.81 | 1600.3 | 37.1 | 0.33 | $580.0 / 505$ |
|  |  |  |  |  | 1602.2 | 35.9 | 0.51 | $649.1 / 447$ |  |
| EROS2-LMC-11 | $\operatorname{lm} 061-4 m-15782$ | 88.4392 | -71.2536 | 20.55 | 0.55 | 451.9 | 51.3 | 0.36 | $456.5 / 444$ |
|  |  |  |  |  |  | 446.6 | 46.1 | 0.37 | $458.7 / 394$ |
| EROS2-LMC-12 | $\operatorname{lm} 085-61-14234$ | 89.8750 | -74.5676 | 20.19 | 0.34 | 2664.0 | 49.6 | 0.27 | $580.4 / 391$ |




Fig. 8. The light curves of EROS-2 star $\operatorname{lm} 057-0 \mathrm{n}-29305(\mathrm{RA}=79.9488 \mathrm{deg}, \mathrm{Dec}=-70.7741 \mathrm{deg})$. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.
timescales ( 103 d ) and the highest luminosity of all published microlensing candidates reported toward the Magellanic Clouds. The star is separated by 1.6 arcsec from another star that is 1.2 mag fainter (Udalski et al. 1997) causing $\sim 30 \%$ blending in EROS-2 images. Including the blend in the light curve fit increases $t_{\mathrm{E}}$ to $125 \mathrm{~d}^{9}$. The star displays a $4-5 \%$ variability with a period $P=5.12 \mathrm{~d}$ (Palanque-Delabrouille et al. 1998). For further details on this event, see Palanque-Delabrouille et al. (1998), Afonso et al. (1999, 2003a) and Assef et al. (2006).

As a check on the cuts $C 1-12$ leading to EROS2-SMC-1 and to search for nonstandard microlensing events (e.g. those due to binary lenses), a large number of light curves were visually scanned. Among them were all events satisfying C1-C8

[^7]and $u_{0}<1$, all events satisfying $C 1-C 7$ and $u_{0}<0.5$, all events satisfying $C 1-C 10$, and all events satisfying $C 1-C 12$ but with $R_{\text {max }}$ increased by 0.2 mag. Only one interesting event was found, EROS2-LMC-15 shown in Fig. 8. The beginning of the event is very similar to a standard microlensing event but subsequent points are too high ${ }^{10}$. It was rejected by the $\chi_{\text {peak }}^{2}$ cut (C9) but it may, in fact, be a lensing event due to a binary lens. Note that its position in the color-magnitude diagram suggests that the star is either not in the LMC or behind a foreground Milky Way star.

An important verification of the analysis of the BrightStar Sample presented here comes from an analysis of the Full Sample of Magellanic light curves described in

[^8]Table 3. The 17 microlensing candidates of the MACHO collaboration (Alcock et al. 2000b). Candidates 20, 22 and 27 (marked B) are low signal-to-noise candidates not satisfying the MACHO " A " requirements. Candidate 9 is due to a binary lens and does not satisfy their A requirements. The values of $t_{\mathrm{E}}$ and maximum amplification $A_{\max }$ are from the MACHO fit assuming no blending, and $t_{0}$ is given as JD- 2450000 . Candidates 18 and 25 are each within 1 arcsec of an object of similar magnitude that is resolved by the MACHO analysis but blended in the EROS analysis. Candidates 14 and 20 occurred during EROS-2 observations, and the EROS-2 values of $t_{\mathrm{E}}$ and $A_{\text {max }}$ are shown in parentheses (averages of $R_{\text {eros }}$ and $B_{\text {eros }}$ measurements).

| MACHO candidate | EROS-2 star | $R_{\text {eros }}$ | $(B-R)_{\text {eros }}$ | $t_{0}$ | $t_{\text {E }}$ | $A_{\text {max }}$ | Comments |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| MACHO-LMC-1 | lm019-6k-14879 | 18.97 | 0.61 | -942.9 | 17.20 | 7.15 |  |
| MACHO-LMC-4 | lm057-01-27696 | 19.85 | 0.24 | -353.1 | 22.70 | 2.92 |  |
| MACHO-LMC-5 | lm057-0k-27469 | 19.97 | 1.52 | -976.1 | 37.80 | 47.28 |  |
| MACHO-LMC-6 | too faint |  |  | -802.9 | 45.80 | 2.43 |  |
| MACHO-LMC-7 | lm010-2k-21383 | 20.47 | 0.52 | -536.5 | 51.45 | 5.91 |  |
| MACHO-LMC-8 | too faint |  |  | -612.2 | 33.20 | 2.19 |  |
| MACHO-LMC-9 B-binary | lm001-01-6864 | 20.92 | 0.83 | -400.1 | 89.60 | 1.95 |  |
| MACHO-LMC-13 | $\operatorname{lm020-6m-19733~}$ | 21.03 | -0.07 | 134.0 | 50.05 | 2.36 |  |
| MACHO-LMC-14 | $\operatorname{lm} 002-7 \mathrm{n}-24616$ | 18.84 |  | 391.3 | 50.05 (38.2) | 3.37 (2.77) |  |
| MACHO-LMC-15 | too faint |  |  | 472.4 | 18.40 | 2.83 |  |
| MACHO-LMC-18 | lm060-2n-16057 | 18.35 | 0.47 | -217.2 | 37.10 | 1.54 | Bright Sample |
| MACHO-LMC-20 B | $\operatorname{lm} 012-5 n-23157$ | 20.09 | 1.03 | 774.7 | 36.35 (23.2) | 2.95 (3.1) |  |
| MACHO-LMC-21 | dead zone |  |  | -786.8 | 46.60 | 5.64 |  |
| MACHO-LMC-22 B | dead zone |  |  | -43.3 | 114.65 | 2.70 |  |
| MACHO-LMC-23 | $\operatorname{lm} 055-3 \mathrm{n}-4994$ | 19.82 | 0.37 | -237.7 | 42.60 | 2.41 | variable (Fig. 10) |
| MACHO-LMC-25 | lm017-2m-2056 | 17.95 |  | -265.7 | 42.60 | 1.50 | Bright Sample |
| MACHO-LMC-27 B | $\operatorname{lm} 010-0 \mathrm{n}-23830$ | 19.00 | 0.14 | -485.9 | 25.25 | 1.45 |  |

Tisserand (2004). Because faint stars were analyzed, the cuts, given in Appendix A, were generally slightly stricter than those described in Sect. 4. However, the analysis did not require observation of the event in both colors. It therefore provides an important check on our analysis. The candidates found in this analysis are listed in Table 2 and their light curves shown in Appendix A. Two candidates EROS2-SMC-1 and 5 are in the Bright-Star Sample. Candidate 5 does not pass the cuts presented here because it fails $C 12$. We also note that it shows non-microlensing-like variations in the light curve of the MACHO collaboration ${ }^{11}$.

Two other independent analyses were performed on the SMC data, restricted to the first five years of data. In both cases, the only event found in the Bright-Star Sample was EROS2-SMC-1. The first analysis, reported in Afonso et al. (2003a), was based on a larger set of stars ( 5.2 million); the number of analysis cuts common to it and the present analysis is small; the computer programs were written independently. The second analysis followed a complete reprocessing using a new implementation of differential photometry developed by Le Guillou (2003). The technique would allow us to find events not on cataloged stars.

We attempted to check our efficiency for finding microlensing events by considering the events published by the MACHO collaboration. The 13-17 events (Alcock et al. 2000b) used by them to measure the optical depth toward the LMC are listed in Table 3. Only 2 of the 17 stars (MACHO-LMC-18 and 25) are bright enough to be in our Bright-Star Sample. Three of the 17 events occurred after the beginning of EROS-2 operations but none of these three events occurred on stars in our Bright-Star Sample. One of the three (MACHO-LMC-15) was on a star too dim to be seen by EROS- $2^{12}$. The two other events (MACHO-LMC-14 and 20) are seen in the EROS-2 images and give microlensing parameters compatible with those measured by MACHO (see Table 3).

[^9]MACHO-LMC-14 was selected in the MACHO A analysis. In EROS-2, it was located in a defective zone of $200 \times 2000$ pixels on CCD 7 of our blue camera, near the edge of the mosaic. The corresponding star was not cataloged in this band, and the star failed the requirement to be observed in both passbands. Consequently, it does not appear in Table 2.

MACHO-LMC-20 was selected in the MACHO B analysis. In EROS-2, it is located 35 arcsec from a very bright Galactic star (about $V=10 \mathrm{vs}$. 21 for the candidate). In this analysis, stars too near bright stars were eliminated so this candidate does not appear in Table 2.

## 6. The status of published Magellanic microlensing candidates

In this section, we review and update the status of the published Magellanic microlensing candidates of the EROS and MACHO collaborations.

EROS has, in the past, used 11 candidates to place upper limits on the optical depth toward the LMC and to measure the depth toward the SMC. Six of the candidates are in the BrightStar Sample considered here. The candidates and their present status are given in Table 4.

Of the candidates in Table 4 only EROS2-SMC-1 remains as a candidate microlensing event. The others have been eliminated either because continued observations of the same stars show further variability on the light curves, or because improved photometry, in some cases complemented by spectroscopy, lead to re-interpreting the candidates as variable stars. Before the analysis presented here, the following EROS candidates were eliminated: EROS1-LMC-2, presented in Aubourg et al. (1993) and which displayed a new variation 8 years later (Lasserre et al. 2000); and candidate EROS2-LMC-4, presented in Lasserre et al. (2000) and eliminated in Milsztajn et al. (2001).

Before this analysis started, there were 5 surviving LMC microlensing candidates from EROS, one from EROS-1 (number EROS1-LMC-1) and four from EROS-2 (numbered EROS2-LMC-3, 5, 6 and 7).

Table 4. The 11 events of the EROS collaboration used in the past to set upper limits on the microlensing optical depth toward the LMC and to measure the depth toward the SMC. Candidates marked "B-S" in Col. 2 occur on stars in the EROS-2 Bright-Star Sample. All candidates except EROS2-SMC-1 have been eliminated as variable stars or as supernovae.

| Candidate | EROS-2 star | $R_{\text {eros }}$ | $(B-R)_{\text {eros }}$ | Original ref. | Status |
| :--- | :--- | :--- | :--- | :--- | :--- |
| EROS1-LMC-1 | lm058-2k-21915 | 18.75 | 0.34 | Aubourg et al. (1993) | 2nd variation (Tisserand 2004) (Figure 9) |
| EROS1-LMC-2 | $\operatorname{lm} 043-6 m-9377$ B-S | 19.32 | -0.04 | Aubourg et al. (1993) | 2nd variation (Lasserre et al. 2000) |
| EROS2-LMC-3 | $\operatorname{lm} 034-61-20493$ | 20.90 | 0.61 | Lasserre et al. (2000) | 2nd variation (Tisserand 2004) |
| EROS2-LMC-4 | $\operatorname{lm} 018-6 \mathrm{n}-23236$ | 19.10 | 1.87 | Lasserre et al. (2000) | 2nd variation (Milsztajn et al. 2001) |
| EROS2-LMC-5 | $\operatorname{lm} 015-3 \mathrm{n}-22431 \mathrm{~B}-\mathrm{S}$ | 19.17 | 0.14 | Milsztajn et al. (2001) | Supernova (Tisserand 2004) |
| EROS2-LMC-6 | $\operatorname{lm} 067-5 \mathrm{~m}-14700$ | 21.01 | 0.63 | Milsztajn et al. (2001) | Supernova (Tisserand 2004) |
| EROS2-LMC-7 | $\operatorname{lm} 070-3 \mathrm{n}-23389$ | 21.00 | 0.76 | Milsztajn et al. (2001) | Supernova (Tisserand 2004) |
| EROS2-SMC-1 | sm005-4m-5761 B-S | 18.13 | -0.13 | Palanque-D. et al. (1998) | Fig. 7 |
| EROS2-SMC-2 | sm001-6l-13221 B-S | 19.56 | 0.44 | Afonso et al. (2003a) | long period variable (Tisserand 2004) |
| EROS2-SMC-3 | sm001-6n-16904 B-S | 19.31 | 0.59 | Afonso et al. (2003a) | long period variable (Tisserand 2004) |
| EROS2-SMC-4 | sm002-7m-21331 B-S | 19.48 | 0.32 | Afonso et al. (2003a) | long period variable (Tisserand 2004) |

The EROS-1 candidate EROS1-LMC-1 displayed a new variation in the EROS-2 data (Fig. 9) in 1998, 6.3 years after the first one, of similar amplitude (a factor two) and timescale (27 days). This second variation is well fitted by a microlensing light curve. Because they are separated in time by more than 80 Einstein timescales, the probability that these two bumps correspond to the microlensing of a double source star is lower than half a percent, even in the favorable case of the two stars being of equal luminosity. This candidate was thus rejected. Let us recall that it was already known to be a Be star (Beaulieu et al. 1995) and was thus already suspected of being variable (Paczyński 1996). Candidate EROS2-LMC-3 also displayed new variations between 1999 and 2002, of a more irregular nature, and was thus rejected.

The light curves of candidates EROS2-LMC-5, 6 and 7 have been improved, due to the better template images in the present analysis. The reduced photometric scatter, compared to Milsztajn et al. (2001), made apparent an asymmetry in rise and fall times; the asymmetry test ( $C 11$ ) allowed us to identify and reject them as supernovæ. EROS2-LMC-5 is identical to MACHO-LMC-26, which had been rejected by Alcock et al. (2000b) for the same reason.

The conclusion is that none of the former EROS LMC microlensing candidates are still considered valid. Four displayed further variability, and three were identified as SN thanks to improved photometry.

There have been four EROS-2 SMC candidates, EROS2-SMC-1 discussed in Sect. 5, and EROS2-SMC-2, 3 and 4 presented in Afonso et al. (2003a). Candidates 2, 3 and 4 were described as doubtful candidates, as all three display very long timescale variations, and look more like irregular variable stars. Another analysis of the same SMC data using differential photometry (Le Guillou 2003) allowed us to reduce the photometric errors and substantiate this interpretation. The additional two years of data since Afonso et al. (2003a) have confirmed this, as further irregular variations have been observed. They are not selected in the present analysis, and are now considered as definitively rejected.

We have also searched for further variations in the 17 stars used by the MACHO collaboration to measure the optical depth toward the LMC (Table 3). One of the candidates, MACHO-LMC-23, showed a further variation in the EROS-2 data (Tisserand 2004), 6.8 years after its first variation in the MACHO data. Its EROS-2 lightcurve is shown in Fig. 10. As such, we can eliminate it as a microlensing candidate. In spite of this, both the variation shown in Fig. 10 and the original variation in the MACHO data are quite achromatic, indicating
that achromaticity is not a fool-proof criterion for selecting microlensing events. We note however that Bennett et al. (2005) argued that, even without considering its further variation seen in the EROS-2 data, the form of its light curve made MACHO-LMC-23 a weak microlensing candidate.

The MACHO collaboration has also reported candidate events found by their alert system. The most notable is MACHO99-LMC-2 that was studied by Bond et al. (2002). This impressive microlensing candidate was on a star too faint to be in the EROS Full Sample.

Besides the 17 LMC events of the MACHO collaboration, they have reported two candidates in the SMC, though they have reported no systematic search for SMC events. The first MACHO SMC candidate (Alcock et al. 1997b) is identical to EROS2-SMC-1 (Sect. 5). The second candidate, MACHO-98-SMC-1, concerned a star too faint to be included in our Full Sample of SMC stars. The event was detected by the alert system of the MACHO group in May 1998 and a probable caustic crossing due to a double lens was announced soon after. The source star was monitored by most microlensing groups, including EROS-2. The second caustic crossing was fully measured, which allowed the determination of the relative proper motion between the lens and source, allowing one to conclude that the event was due to a lens in the SMC (Afonso et al. 1998, 2000) ${ }^{13}$.

## 7. The detection efficiency

To measure the optical depth from the detected events, or limits on this quantity, we first need to evaluate the detection efficiency as a function of the time scale $t_{\mathrm{E}}$. This was determined by using Monte Carlo simulated light curves: we superimpose artificial microlensing events on a representative sample of light curves, corresponding to $2 \%$ of the Bright-Star Sample from each of the 98 monitored fields. The light curve for a simple microlensing event (i.e. point-source point-lens zero-blending) is described by three parameters of (10): date of maximum amplification $t_{0}$, impact parameter $u_{0}$ and time scale $t_{\mathrm{E}}$. Blended events have the additional parameter $\alpha$ in (11). The microlensing parameters are chosen at random: $t_{0}$ follows a flat distribution over our 2500 days observing period, JD 2450242 till 2452742 ; $u_{0}$ is

[^10]

Fig. 9. The EROS-2 light curve of EROS-1 microlensing candidate EROS1-LMC-1. The curve shows a second variation, 6.3 years after the variation observed in EROS-1. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.


Fig. 10. The EROS-2 light curve of MACHO microlensing candidate MACHO-LMC-23. The curve shows a second variation, 6.8 years after the variation seen by MACHO. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.
picked randomly between 0 and 1.2 ; and $t_{\mathrm{E}}$ is chosen at random from a distribution flat in $\ln \left(t_{\mathrm{E}}\right)$, between 1 and 1000 days. Each star in the $2 \%$ sample is actually used thrice in the simulation, once per decade in $t_{\mathrm{E}}$. The simulation takes into account the relative variation of photometric errors.

Simulated light curves were then fed into the analysis chain to find the fraction that are recovered by our detection algorithm.

The detection efficiency in a given $t_{\mathrm{E}}$ bin, $\epsilon\left(t_{\mathrm{E}}\right)$, is then given by the ratio between the number of events passing all selection criteria in this bin and the number of microlensing events generated in the same bin with $u_{0}<1$.

Figure 11 shows the LMC and SMC efficiencies for unblended events $(\alpha=1)$ as a function of $t_{\mathrm{E}}$. For the range of $t_{\mathrm{E}}$


Fig. 11. The detection efficiency for unblended $(\alpha=1)$ microlensing light curves (10) as a function of $t_{\mathrm{E}}$ for LMC (bold line) and SMC (light line) fields. The efficiency applies to a 2500 d total observing period. The efficiency used for calculation of the optical depth is the efficiency shown here, multiplied by 0.9 to take into account lensing by binary lenses.


Fig. 12. The ratio between the summed efficiency, $\epsilon_{1}\left(\alpha_{1}\right)+\epsilon_{2}\left(\alpha_{2}\right)$ and the unblended efficiency $\epsilon(\alpha=1)$ for pairs $\left(\alpha_{1}, \alpha_{2}\right)$ taken from the artificial star analysis of Sect. 3. The ratio is a function of $t_{\mathrm{E}}$. The four curves correspond to the four studied fields, $1 \mathrm{~m} 009,019,034$ and 048.
of interest for this work, the efficiency varies from $\sim 0.25$ at $t_{\mathrm{E}}=10 \mathrm{~d}$ to $\sim 0.45$ at $t_{\mathrm{E}}=200 \mathrm{~d}$.

Blended events have an efficiency that is reduced because an impact parameter smaller than unity is necessary to produce an amplification of 1.34 . The efficiency is also modified because the time scale is reduced. To sufficiently good approximation we find
$\epsilon_{i j}\left(\alpha_{i}, t_{\mathrm{E}}\right) \sim \frac{u_{i}\left(2.18, \alpha_{i}\right)}{0.5} \epsilon\left(\alpha=1, u_{i}\left(1, \alpha_{i}\right) t_{\mathrm{E}}\right)$
where $u_{i}\left(A_{\max }, \alpha_{i}\right)$ is the impact parameter necessary to produce a maximum amplification $A_{\max }$.

To evaluate the summed detection efficiency for a realistic distribution of $\alpha_{1}$ and $\alpha_{2}$ we used the pairs from the artificial images of Sect. 3, Fig. 4. Figure 12 shows the ratio between the calculated sum (for the four studied fields) and the unblended efficiency. For $t_{\mathrm{E}}=40 \mathrm{~d}$ the efficiency is reduced by a factor ranging from 0.90 in the sparse fields $\operatorname{lm} 048$ and $\operatorname{lm} 034$ to 0.92 and 0.97 in the denser fields $\operatorname{lm} 019$ and $\operatorname{lm} 009$. The brightest star, $i=1$ accounts for $95 \%$ of the rate in the sparse fields and $88 \%$ in the dense fields.

Figure 13 shows the efficiency for unblended events as a function of $u_{0}$ and of $R_{\text {eros }}$ for events with $10 \mathrm{~d}<\mathrm{t}_{\mathrm{E}}<200 \mathrm{~d}$.


Fig. 13. The detection efficiency for simple microlensing light curves (10) as a function of $u_{0}(t o p)$ and of $R_{\text {eros }}$ (bottom) for LMC and SMC events in the range $10 \mathrm{~d}<\mathrm{t}_{\mathrm{E}}<200 \mathrm{~d}$.

The efficiency for the Bright-Star Sample has a much weaker dependence on $u_{0}$ and $R_{\text {eros }}$ than that for the Full Sample. For the Full Sample the efficiency falls rapidly with increasing $u_{0}$ and $R_{\text {eros }}$ (Tisserand 2004).

The efficiencies in Figs. 11 and 13 are for the detection of microlensing events due to simple lenses. Events due to binary lenses with caustic crossings are discriminated against, mostly by C9. Of the 17 LMC events of the MACHO collaboration (Alcock et al. 2000b) only 1 event, MACHO-LMC-9, is of this type and would not pass our selection criteria. This event corresponds to $\sim 10 \%$ of their optical depth. We note also that toward the Galactic Bulge $\sim 10 \%$ of the observed microlensing events are due to binary lenses (Udalski et al. 2000a). To compensate for this loss of efficiency, we conservatively reduce the efficiency of Fig. 11 by a factor 0.9 when calculating limits on the optical depth.

## 8. Limits on the abundance of machos

The microlensing optical depth, $\tau$, is defined as the probability that any given star, at a given time, is amplified by at least 1.34, i.e. with an impact parameter $u<1$. From a set of $N_{\mathrm{ev}}$ events, $\tau$ can be estimated from
$\tau=\frac{\pi N_{\mathrm{ev}}}{2 N_{\mathrm{stars}} T_{\mathrm{obs}}}\left\langle\frac{t_{\mathrm{E}}}{\epsilon}\right\rangle$,
where $\left\langle t_{\mathrm{E}} / \epsilon\right\rangle$ is the mean $t_{\mathrm{E}}$ divided by efficiency for the observed events.

In the LMC we have found no events so we can only give an upper limit on $\tau$ by calculating the expected number of events


Fig. 14. The $t_{\mathrm{E}}$ distribution $\mathrm{d} \Gamma / \mathrm{d} t_{\mathrm{E}}$ expected for $1 M_{\odot}$ lenses in a spherically symmetric isothermal halo with core radius 5 kpc , i.e. the S model used by the MACHO collaboration (Alcock et al. 2000b).
as a function of $\tau$ as given by (9). For this analysis we use the $t_{\mathrm{E}}$ distribution of the S model (Alcock et al. 2000b) shown in Fig. 14. It relates $\left\langle t_{\mathrm{E}}\right\rangle$ to the macho mass (assumed unique): $\left\langle t_{\mathrm{E}}\right\rangle=70 \mathrm{~d} \sqrt{M / M_{\odot}}$. Limits using other halo models or macho mass distributions can be found to often good approximation by simply scaling (9) with $\epsilon\left(\left\langle t_{\mathrm{E}}\right\rangle\right) /\left\langle t_{\mathrm{E}}\right\rangle$.

The expected number of LMC events for $\tau_{\mathrm{lmc}}=4.7 \times 10^{-7}$ as a function of lens mass, $M$, is shown in Fig. 15a. For $M=$ $0.4 M_{\odot}$, we have $\left\langle t_{\mathrm{E}}\right\rangle=44 \mathrm{~d},\langle\epsilon\rangle=0.35, N_{\text {star }}=5.5 \times 10^{6}$ and $T_{\text {obs }}=2500 \mathrm{~d}$, giving 32 LMC events for EROS-2. We add 7 LMC events for EROS-1 to give a total of 39 expected events for $\tau_{\mathrm{lmc}}=4.7 \times 10^{-7}$.

For no observed events ( $N<3.0,95 \% \mathrm{CL}$ ), the $95 \%$ CL upper limit on the optical depth is
$\frac{\tau}{4.7 \times 10^{-7}}<\frac{3}{N_{\text {ex }}\left(4.7 \times 10^{-7}\right)}$.
For 39 expected events, The upper limit is then $\tau_{\mathrm{lmc}}<0.36 \times$ $10^{-7}$. The limit on $\tau_{\text {lmc }}$ as a function of $M$ is shown in Fig. 15b. In the $t_{\mathrm{E}}$ range favored by the MACHO collaboration, we find
$\tau_{\operatorname{lmc}}<0.36 \times 10^{-7} \times\left[1+\log \left(M / 0.4 M_{\odot}\right)\right] \quad 95 \% \mathrm{CL}$,
i.e.
$f<0.077 \times\left[1+\log \left(M / 0.4 M_{\odot}\right)\right] \quad 95 \% \mathrm{CL}$,
where $f \equiv \tau_{\text {lmc }} / 4.7 \times 10^{-7}$ is the halo mass fraction within the framework of the S model. This limit on the optical depth is significantly below the value for the central region of the LMC measured by the MACHO collaboration (Alcock et al. 2000b), $\tau_{\text {lmc }} / 10^{-7}=1.2_{-0.3}^{+0.4}$ (stat.) $\pm 0.36$ (sys.) and the revised value of Bennett (2005), $\tau_{\mathrm{lmc}} / 10^{-7}=1.0 \pm 0.3$. The Alcock et al. (2000b) optical depth used for the entire LMC predicts that EROS would see $\sim 9$ LMC events whereas none are seen.

For the SMC, the one observed event corresponds to an optical depth of $1.7 \times 10^{-7}\left(N_{\text {star }}=0.86 \times 10^{6}\right)$. Taking into account only Poisson statistics on one event, $0.05<N_{\text {obs }}<4.74$ ( $90 \% \mathrm{CL}$ ) this gives
$0.085 \times 10^{-7}<\tau_{\text {smc }}<8.0 \times 10^{-7} \quad 90 \% \mathrm{CL}$.
This is consistent with the expectations of lensing by objects in the SMC itself, $\tau_{\mathrm{smc}} \sim 0.4 \times 10^{-7}$ (Graff \& Gardiner 1999). The value of $t_{\mathrm{E}}=125 \mathrm{~d}$ is also consistent with expectations for selflensing $\left\langle t_{\mathrm{E}}\right\rangle \sim 100 \mathrm{~d}$ for a mean lens mass of $0.35 M_{\odot}$.


Fig. 15. The top panel shows the numbers of expected events as a function of macho mass $M$ for the S model of Alcock et al. (2000b). The expectations for EROS-2-LMC, SMC (this work) are shown along with those of EROS-1 (Renault et al. 1997) with contributions from the photographic plate program (Ansari et al. 1996a) and CCD program (Renault et al. 1998). The number of events for EROS-2-SMC supposes $\tau_{\mathrm{smc}}=1.4 \tau_{\text {lmc }}$. In the lower panel the solid line shows the EROS $95 \%$ CL upper limit on $f=\tau_{\text {lmc }} / 4.7 \times 10^{-7}$ based on no observed events in the EROS-2 LMC data and the EROS-1 data. The dashed line shows the EROS upper limit on $\tau_{\text {lmc }}$ based on one observed SMC event in all EROS-2 and EROS-1 data assuming $\tau_{\text {smc-halo }}=1.4 \tau_{\text {lmc }}$. The MACHO $95 \%$ CL. curve is taken from Fig. 12 (A, no lmc halo) of Alcock et al. (2000b).

We also note that the self-lensing interpretation is favored from the absence of an indication of parallax in the light curve (Assef et al. 2006).

We can combine the LMC data and the SMC data to give a limit on the halo contribution to the optical depth by supposing that the SMC optical depth is the sum of a halo contribution, $\tau_{\text {smc-halo }}=\alpha \tau_{\text {lmc }}(\alpha \sim 1.4)$ and a self-lensing contribution $\tau_{\mathrm{sl}}$. (We conservatively ignore contributions from LMC self-lensing and from lensing by stars in the disk of the Milky Way.) For one observed SMC event with $t_{\mathrm{E}}=125 \mathrm{~d}$ and zero observed LMC events, the likelihood function is
$L\left(\tau_{\mathrm{lmc}}, \tau_{\mathrm{sl}}\right) \propto\left[\alpha \tau_{\mathrm{lmc}} \Gamma_{\mathrm{h}}^{\prime}\left(t_{\mathrm{E}}\right)+\tau_{\mathrm{sl}} \Gamma_{\mathrm{sl}}^{\prime}\left(t_{\mathrm{E}}\right)\right] \exp \left[-N\left(\tau_{\mathrm{lmc}}, \tau_{\mathrm{sl}}\right)\right]$
where $N\left(\tau_{\mathrm{lmc}}, \tau_{\mathrm{sl}}\right)$ is the total number of expected events (LMC and SMC) as a function of the two optical depths as calculated with Eq. (8). The function $\Gamma_{\mathrm{h}}^{\prime}\left(t_{\mathrm{E}}\right)$ is the distribution (normalized to unit integral) expected for halo lenses of mass $M$ (Fig. 14) and $\Gamma_{\mathrm{sl}}^{\prime}\left(t_{\mathrm{E}}\right)$ is the expected distribution for SMC self-lensing taken from Graff \& Gardiner (1999). We assume the SMC self-lensing optical depth is that calculated by Graff \& Gardiner (1999) though the results are not sensitive to
this assumption. For macho masses less than $1 M_{\odot}$, the likelihood function is maximized for $\tau_{\text {lmc }}=0$ because there are no LMC events in spite of the greater number of LMC source stars. For $M<0.1 M_{\odot}$ the limit on the halo contribution approaches that one would calculate for no candidates in either the LMC or the SMC because the observed $t_{\mathrm{E}}$ of 125 d is too long for a halo event. The calculated upper limit is shown as the dashed line in Fig. 15b. In the mass range favored by the MACHO collaboration, the limit is slightly lower than that using only the LMC data. The combined limit would be somewhat stronger if we assumed an oblate halo $(\alpha<1.4)$ and somewhat weaker if we assumed a prolate halo ( $\alpha>1.4$ ). Constraints on the shape of the Milky Way halo were recently summarized by Fellhauer et al. (2006) who argued that the observed bifurcation of the Sagittarius Stream can be explained if the halo is close to spherical.

A possible systematic error in our result could come from our assumption that the optical depth due to binary lenses is small, $10 \%$ of the total. An alternative strategy would have been to relax the cuts so as to include the event shown in Fig. 8. We have chosen not to do this because the light curve itself is not sufficiently well sampled to establish the nature of the event (other than that it is not a simple microlensing event) and also because of its anomalous position in the color-magnitude diagram. We note also that the optical depth associated with the event, $\tau=0.7 \times 10^{-8}$, is a factor $\sim 4$ below the upper limit (17).

Another important question concerns the influence on our results of the Bright-Sample magnitude cut. Since the cut was not established before the event search, it is natural to ask if the position of the cut was chosen to give a strong limit. In fact, elimination of the cut would not change significantly the conclusions of this paper. Four additional events (EROS2-LMC-8, 9,10 and 11 from Table 2) were found with the analysis described in Appendix A. The baselines of the source stars of these events are on average 0.9 mag below the Bright Sample cut. An additional event (EROS2-LMC-12) with no $R_{\text {eros }}$ data during the variation was found. The values of $t_{\mathrm{E}}$ for these events are in the range $10-60$ days, similar to those in the MACHO sample. The optical depth associated with the four events seen in two colors is $(1.7 \pm 0.6) \times 10^{-8}$ (Tisserand 2004), not in contradiction with the limit from the Bright-Sample analysis. The $95 \%$ CL limit on the halo fraction is $12 \%$ at $M=0.4 M_{\odot}$. We prefer to use the limit from the Bright Sample for the reasons already mentioned: better understanding of blending and a superior photometry allowing better rejection of variable stars. Indeed, inspection of the light curves of EROS2-LMC-8-12 in Figs. A.1-A. 5 indicates that most are not especially convincing candidates. An exception is EROS2-LMC-8 but its position in the color-magnitude diagram and strongly chromatic magnification indicate that the event is most likely due to a lens in the Milky Way disk.

## 9. Discussion of the results

The limits shown in Fig. 15 rule out machos as the majority of Galactic dark matter over the range $0.6 \times 10^{-7} M_{\odot}<M<15 M_{\odot}$. The limits are $f<0.04$ for $10^{-3} M_{\odot}<M<10^{-1} M_{\odot}$ and $f<$ 0.1 for $10^{-6} M_{\odot}<M<1 M_{\odot}$. We note that even stronger limits in the range $10^{-7} M_{\odot}<M<10^{-3} M_{\odot}$ were found (Alcock et al. 1998) by combining the data from the EROS-1 CCD program with those of the MACHO program. These data gave $f<0.07$ for $M \sim 3 \times 10^{-6} M_{\odot}$. It should also be possible to improve limits in the range $10 M_{\odot}<M<100 M_{\odot}$ by combining the results presented here with the MACHO high mass results (Alcock et al. 2001b). This may narrow the small remaining allowed macho
mass range between the range excluded by microlensing and that excluded by the abundance of halo wide binary stars (Yoo et al. 2004).

Initially, the EROS and MACHO programs were primarily motivated by the search for halo brown dwarfs of mass $0.02-0.08 M_{\odot}$. Such objects are clearly ruled out as primary components of the Milky Way halo by Fig. 15 (implying $f<$ 0.04 ) and the data of the MACHO collaboration (Alcock et al. 2000b).

The observation of 17 events by MACHO with $\left\langle t_{\mathrm{E}}\right\rangle \sim 40 \mathrm{~d}$ suggested the existence of machos of $M \sim 0.4 M_{\odot}$. Such objects could be invisible (e.g. primordial black holes) or faint (e.g. cool white dwarfs). However, the latter are not seen at the corresponding level in multi-color surveys (Gates et al. 2004) and high proper-motion surveys (Goldman et al. 2002; Reid 2005, and references therein).

At any rate, the EROS limit (17) is significantly less than the MACHO result $\tau_{\text {lmc }}=1.2_{-0.3}^{+0.4} \times 10^{-7}$. However, there are considerable differences between the EROS and MACHO data sets that may help to resolve the conflict. Generally speaking, MACHO uses faint stars in dense fields $\left(1.1 \times 10^{7}\right.$ stars over $13.4 \mathrm{deg}^{2}$ ) while EROS-2 uses bright stars in sparse fields $(0.7 \times$ $10^{7}$ stars over $84 \mathrm{deg}^{2}$ ). Of these bright EROS-2 stars, $0.2 \times 10^{7}$ are in MACHO fields.

The use of dense fields by the MACHO group suggests that the higher MACHO optical depth may be due, in part, to self-lensing in the inner parts of the LMC. This would contradict LMC models (Mancini et al. 2004) which suggest that only 1-3 MACHO events should be expected to be due to selflensing. In fact, MACHO-LMC-14 is known to be due to selflensing (Alcock et al. 2001c) because it has a binary source and the form of the accompanying deformation of the lightcurve with respect to the simple microlensing lightcurve requires that the lens be in the LMC. A second event, MACHO-LMC-9, is due to a binary lens and the self-lensing interpretation can be avoided only by assuming that the source is also a binary system and that each of the two widely separated components happened to land on the caustic on the two successive observations made of the caustic entrance (Alcock et al. 2000a).

If it turns out that the self-lensing rate is higher than the model estimates in the MACHO fields but still negligible in the outer fields of the LMC, the disagreement between MACHO and this work is considerably reduced. Since only $1 / 3$ of our Bright Sample stars are in MACHO fields, the EROS-2 95\% CL upper limit on $\tau$ for the MACHO fields is the limit (17) multiplied by a factor 3, consistent with the Alcock et al. (2000b) result as modified by Bennett (2005).

A possible explanation for the discrepancy that is similar to self-lensing is the possibility that the events are due to halo lenses but the Halo is clumpy and that the MACHO-fields lie behind a clump of size less than that of the EROS-2 fields. The effect of a clumpy halo on the optical depth was discussed by Holopainen et al. (2006) though they did not discuss directly the possibility that it could resolve the EROS-MACHO controversy. At any rate, if this is the cause of the discrepancy, the EROS-2 result gives the more representative optical depth because it is based on a larger solid angle.

The use of faint stars by MACHO may also give an explanation of the disagreement. Only two of the 17 MACHO candidates (MACHO-LMC-18 and 25) are sufficiently bright to be in our Bright-Star Sample. Of these two, MACHO-LMC-18 is in the EROS Bright-Sample only because the EROS starfinder mixed two similar objects that are resolved by the MACHO starfinder.

The use of faint stars by MACHO suggests two possible explanations for the disagreement. The first would be contamination by variable stars that, in our Bright-Star Sample, are either not present or identified as such because of superior photometric precision. The low photometric precision for the faint LMC stars makes most of the events less convincing than the events on bright stars in the Galactic Bulge or the one EROS event in the SMC. Indeed, one of the MACHO A events, MACHO-LMC-23, has been identified as a variable star by a second variation in the EROS-2 data shown in Fig. 10 (Tisserand 2004). As already noted, Bennett et al. (2005) argued that MACHO-LMC-23 was, in any case, a weak candidate and that its variability doesn't call into question the nature of the other MACHO candidates. Indeed, some of the MACHO candidates are very convincing microlensing candidates. In particular, MACHO-LMC 1, 5, 9, 14 and 21 are strong candidates based only on MACHO photometry while Bennett et al. (2005) argued that MACHO-LMC-4, 13 and 15 are strong candidates because of high precision followup photometry. We might note, however, that of these candidates, MACHO-LMC-14 is most likely due to LMC self-lensing and MACHO-LMC- 5 is due to lensing by a normal red-dwarf star in the disk of the Milky Way (Alcock et al. 2001d; Gould 2004).

The second possible explanation related to the use of faint source stars is the large blending effects that must be understood. Alcock et al. (2000b, 2001a) suggest a $30 \%$ systematic error to reflect this uncertainty. The experience with the use of faint stars in the Galactic Bulge suggests that this uncertainty may be underestimated, though in principle the star distribution is better understood in the LMC than in the Bulge. At any rate, the extension of the limits presented here, either by EROS or by OGLE-3 or SuperMACHO, will require the use of faint stars and a good understanding of blending.

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## Online Material

## Appendix A: Candidate selection for the full sample

In this paper we have concentrated on the analysis of the BrightStar Sample of stars because this leads to the most reliable limits of the optical depth. A search for microlensing events on the full sample of stars was also performed (Tisserand 2004). The candidates found in this search are listed in Table 2 and displayed in Figs. A.1-A.5. In this section we list the selection criteria leading to this set of candidates.

The criteria are very similar to those described in Sect. 4 for the Bright-Star Sample. However, the problems encountered with analyzing low precision light curves necessarily led to more complicated criteria to avoid the many spurious events caused by photometric problems. In all, 17 criteria, $c 1-c 17$, were applied compared to the 12 criteria, $C 1-C 12$, applied to the Bright-Star Sample.

The first four criteria are identical to those applied to the Bright-Star Sample:
$c 1=C 1 \quad c 2=C 2 \quad c 3=C 3 \quad c 4=C 4$.
We then required that the total amplitude of luminosity variation along the light curve be greater than 5 times the point-to-point dispersion, $\sigma_{\text {base }}$ in the light curve, recomputed after excluding the most significant excursions:
$c 5: \quad F_{5 \max }-F_{5 \min }>5 \sigma_{\text {base }}$.
Here $F_{5 \text { max }}$ and $F_{5 \text { min }}$ are the maximum and minimum fluxes averaged over any 5 neighboring measurements.

For excursions with a regular variation, the point-to-point dispersion of the measured fluxes within the excursion, $\sigma_{\mathrm{ptp}, \mathrm{exc}}$, obtained from the comparison of each measured flux with the linear interpolation of its two neighbors (in time), is normally smaller than the global dispersion, $\sigma_{\mathrm{exc}}$, i.e. the width of the distribution of all fluxes within the excursion. To exclude irregular variations, we require that their ratio be sufficiently small;
$c 6: \quad \frac{\sigma_{\mathrm{ptp}, \mathrm{exc}}}{\sigma_{\mathrm{exc}}}<0.90$.
Cuts $c 4-c 6$ are applied independently in the two passbands. Light curves are retained if they are selected in one passband at least. This reduces the star sample to slightly less than $0.1 \%$ (28500 objects) of the full sample.

The next two criteria are similar to $C 5$ and $C 9$ applied to the Bright-Star Sample
$c 7: \quad\left[\frac{\chi^{2}-N_{\text {dof }}}{\sqrt{2 N_{\text {dof }}}}\right]_{\text {base }}<15$,
c8: $\left[\frac{\chi^{2}-N_{\text {dof }}}{\sqrt{2 N_{\text {dof }}}}\right]_{\text {peak }}<10$.
The next cut, $c 9$, deals with a background connected to the realignment of the telescope optics in May 1998, which slightly changed the PSF. Faint stars near a diffraction feature in the PSF of bright stars were affected; this was seen as light curves with two plateaus, one (higher) before and one (lower) after May 1998. A rather complicated algorithm was developed to identify such light curves. It relies mostly on the relative flux values of the two stars, their distance and the relative height of the two plateaus in the light curve. The details can be found in (Tisserand 2004) but we note that it affects our efficiency only for $t_{\mathrm{E}}>150 \mathrm{~d}$.
$c 9$ : no influence of telescope realignment.

At this point we accepted only light curves that passed criteria $c 1-c 9$ in both colors:
$c 10$ : $\quad c 1-c 9$ red and blue.
Criterion c11 is similar to $C 7$ :
$c 11: \quad N_{\text {peak }}>4 ; \geq 2$ points in rise $; \geq 2$ points in fall.
The remaining cuts follow closely the cuts for the Bright-Star Sample
$c 12=C 6$
$c 13: \quad u_{0 r}<1$ or $u_{0 b}<1$
$c 14: \quad \frac{\chi_{\mathrm{ct}}^{2}-\chi_{\mathrm{ml}}^{2}}{\chi_{\mathrm{ml}}^{2} / N_{\text {dof }}} \frac{1}{\sqrt{2 N_{\text {dof,peak }}}}>50$ in $R$ or $B$.
Criterion $c 15$ eliminates light echos from SN1987a by requiring

## $c 15$ : Distance event - SN1987a > 30 arcmin.

Finally, $c 16$ and $c 17$ are the blue bumper and supernovae criteria $C 10$ and $C 11$ supplemented by a visual rejection of events superimposed on background galaxies.

The first six events in Table 2 passed the cuts $c 1-c 17$. A monochromatic analysis was performed where $c 10$ was removed and candidates were only required to pass the other criteria in either of the two colors. About 500 events were scanned by eye. Only one event, EROS2-LMC-12, was found, the others being due to photometric problems or long period variable stars.


Fig. A.1. The light curves of EROS-2 star $\operatorname{lm} 055-7 \mathrm{~m}-23303$. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.


Fig. A.2. The light curves of EROS-2 star lm042-11-2622. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.


Fig. A.3. The light curves of EROS-2 star $\operatorname{lm} 045-5 n-26323$. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event. Note the visible asymmetry in the B light curve, possibly indicative of a supernova.


Fig. A.4. The light curves of EROS-2 star $\operatorname{lm} 061-4 m-15782$. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the excursion of the event.


Fig. A.5. The light curves of EROS-2 star lm085-6l-14234. Also shown is the color-magnitude diagram of the star's CCD-quadrant and the position of the source's baseline.


[^0]:    * Based on observations made with the Marly telescope at the European Southern Observatory, La Silla, Chile.
    ** Appendix A is only available in electronic form at http://www.aanda.org
    *ᄎぇ Now at Research School of Astronomy and Astrophysics, Australian National University, Mount Stromlo Obs., Cotter Rd., Weston, ACT 2611, Australia.
    $\dagger$ Now at LPNHE, CNRS-IN2P3 and Universités Paris 6 et Paris 7, 4 place Jussieu, 75252 Paris Cedex 05 France.
    $\ddagger$ Now at Max-Planck-Institut für Astronomie, Koenigstuhl 17, 69117 Heidelberg, Germany.
    § Also at APC, 10 rue Alice Domon et Lonie Duquet, 75205 Paris Cedex 13, France.
    II Now at Division of Medical Imaging Physics, Johns Hopkins University Baltimore, MD 21287-0859, USA.
    ॥ Now at San Pedro de Atacama Celestial Exploration, Casilla 21, San Pedro de Atacama, Chile.

[^1]:    ${ }^{1}$ For "Massive Astrophysical Compact Halo Objects" (Griest 1991), not to be confused with the "MACHO collaboration."

[^2]:    ${ }^{2} 13$ of the 17 events satisfied their so-called A criteria intended to identify high signal-to-noise events. The other 4 events, so-called $B$ events, are selected by looser cuts.

[^3]:    ${ }^{3}$ http://www.phys.canterbury.ac.nz/moa/
    ${ }^{4}$ http://bulge.astro.princeton.edu/~ogle/
    5 http://www.ctio.noao.edu/~supermacho/

[^4]:    ${ }^{6}$ The largest single cause for rejection was the malfunction that affected 5 CCDs of the $R_{\text {eros }}$ passband camera starting in January 2002. We have thus chosen to reject all measurements taken in 2002 and 2003

[^5]:    ${ }^{7}$ Using 1000 pictures of this region co-added by groups of 10 , we have produced a film that shows the motion of the light echos; it can be found at the URL http://eros.in2p3.fr/EchoesSN1987a/

[^6]:    ${ }^{8}$ Or, equivalently, $N_{\text {run }, i}$ is the number of times the time-ordered light curve crosses the constant flux line at $\Phi_{i, \text { mid }}$, plus one.

[^7]:    ${ }^{9}$ When blending is taken into account, this event has the longest $t_{\mathrm{E}}$ of any published Magellanic event.

[^8]:    ${ }^{10}$ OGLE-3 data confirm the high flux of the post-peak points (A. Udalski, private communication).

[^9]:    ${ }^{11}$ http://wwwmacho.mcmaster.ca/Data/MachoData.html. There is no information at this site on the EROS-2 LMC candidates of Table 2 since none are in MACHO fields.
    ${ }^{12}$ The 1.27 m telescope of the MACHO collaboration allowed them to use fainter stars than EROS-2 with its 1 m telescope.

[^10]:    ${ }^{13}$ Note that there exists a third, unpublished, microlensing candidate toward the SMC, OGLE-2005-SMC-001 (http:// bulge.astro.princeton.edu/~ogle/). Its long timescale ( $>150 \mathrm{~d}$ ), large amplification ( $>12$ ) and bright source $(I=18.2$ ) offer the prospect of completely resolving the microlensing parameters degeneracy through measurements of lightcurve deformations, e.g. of that due to parallax.

