Mapping routine measles vaccination in low-and middle-income countries

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Local Burden of Disease Vaccine Coverage Collaborators*

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The safe, highly effective measles vaccine has been recommended globally since 1974, yet in 2017 there were more than 17 million cases of measles and 83,400 deaths in children under 5 years old, and more than 99% of both occurred in low- and middle-income countries (LMICs)¹⁻⁴. Globally comparable, annual, local estimates of routine first-dose measles-containing vaccine (MCV1) coverage are critical for understanding geographically precise immunity patterns, progress towards the targets of the Global Vaccine Action Plan (GVAP), and high-risk areas amid disruptions to vaccination programmes caused by coronavirus disease 2019 (COVID-19)⁵⁻⁸. Here we generated annual estimates of routine childhood MCV1 coverage at 5 × 5-km² pixel and second administrative levels from 2000 to 2019 in 101 LMICs, quantified geographical inequality and assessed vaccination status by geographical remoteness. After widespread MCV1 gains from 2000 to 2010, coverage regressed in more than half of the districts between 2010 and 2019, leaving many LMICs far from the GVAP goal of 80% coverage in all districts by 2019. MCV1 coverage was lower in rural than in urban locations, although a larger proportion of unvaccinated children overall lived in urban locations; strategies to provide essential vaccination services should address both geographical contexts. These results provide a tool for decision-makers to strengthen routine MCV1 immunization programmes and provide equitable disease protection for all children.

The safe, highly effective vaccine against measles has been recommended since 1974 by the Expanded Programme on Immunization of the WHO (World Health Organization)¹⁻³. A single valid dose of any MCV is approximately 93% effective in providing individuals with lifelong protection from measles¹. Still, in 2017, there were an estimated 17,767,037 new global cases and 83,439 deaths attributable to measles in children under 5 years old⁴. Although high-income regions, such as the USA and Europe, have recently started experiencing extended measles outbreaks due to a lapse in vaccination coverage, more than 99% of cases and deaths still occur in LMICs4,9.

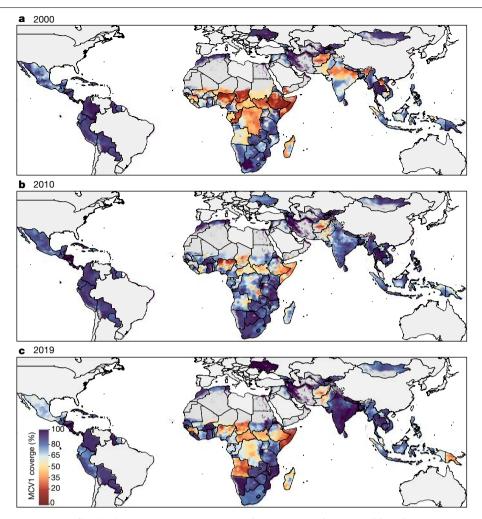
Low vaccination rates and increasing vaccine hesitancy 10-12 contribute to the persistence of measles as a major cause of childhood morbidity and mortality. National-level MCV1 estimates from the Global Burden of Diseases, Injuries and Risk Factors Study (GBD) 2019 identified only 72 out of 204 countries in which routine coverage reached approximate herd immunity targets (≥95%) in 2019, and global MCV1 coverage^{4,13} was 84.2%. Even in countries with high national coverage, these estimates mask important subnational heterogeneities that may sustain ongoing disease transmission and increase the risk of outbreaks¹⁴⁻¹⁷, especially in light of the current service disruptions associated with the COVID-19 pandemic^{7,8}. Global vaccination initiatives, such as the GVAP and Immunization Agenda 2030, recognize the importance of eliminating subnational coverage disparities, aiming for at least 90% of the target population in every country and 80% in every district to be covered^{5,6}.

Subnational routine MCV1 coverage

Since 2016, the WHO and UNICEF have collected subnational coverage data through their annual loint Reporting process, although poor data quality and biases currently limit the use of administrative data to track progress towards GVAP targets^{18–20}. For the 101 countries included in this study, 91 reported subnational data in 2018 in a total of 11,311 subnational geographical units. Of these countries, 71 reported MCV1 coverage greater than 100% in at least one unit and 55 reported such coverage in at least a quarter of units. Although researchers have estimated subnational MCV1 coverage in select countries or years for which there have been reliable surveys, to date, no comprehensive analysis of all available vaccine coverage data to produce subnational estimates of MCV1 coverage annually in all LMICs has been undertaken²¹⁻²⁴.

Building from our previous work mapping diphtheria-tetanuspertussis vaccine coverage in Africa¹⁴, here we present mapped high-spatial-resolution estimates of routine MCV1 coverage across 101 LMICs from 2000 to 2019, aggregated to policy-relevant second-level administrative units (hereafter districts). Using geolocated data on MCV1 coverage from 354 household-based surveys representing approximately 1.70 million children and a suite of environmental, sociodemographic and health-related geospatial and national-level covariates, we extended model-based geostatistical methods that have been used to map child mortality and its main components and risk factors²⁵⁻²⁸ to predict MCV1 coverage and uncertainty

^{*}A list of participants and their affiliations appears in the online version of the paper.



 $Fig.\,1|Estimated\,MCV1\,coverage\,among\,districts\,in\,101\,LMICs, 2000-2019.$ a-c, MCV1 coverage among target population in districts in 2000 (a), 2010 (b) and 2019 (c). Countries excluded from the analysis and pixels classified as

 $'barren\,or\,sparsely\,vegetated'\,based\,on\,European\,Space\,Agency\,Climate$ Change Initiative (ESA-CCI) satellite data or with fewer than 10 people per 1×1 -km² pixel based on WorldPop estimates are masked in grey^{30,50}.

(Extended Data Figs. 1, 2), while calibrating estimates to results from GBD 2019. Using these estimates, we assessed trends in geographical inequality, progress towards global targets and differential vaccination status by geographical remoteness.

Tracking uneven progress

Despite marked global progress between 2000 and 2019, considerable inequalities in routine MCV1 coverage persist, both between and within countries (Fig. 1, Extended Data Figs. 3-7, see also our visualization tool (https://vizhub.healthdata.org/lbd/mcv)). MCV1 coverage among children living in the 101 countries included in this study was 65.6% (95% uncertainty interval, 64.2-67.1%) in 2000 and 81.0% (95% uncertainty interval, 79.2-82.7%) in 2019. Coverage increased at the national level in 69.9% (95% uncertainty interval, 64.4-75.2%) of countries between 2000 and 2019 and in 57.4% (95% uncertainty interval, 50.4-64.6%) of districts (n = 20,795 districts).

The three districts with the lowest MCV1 coverage in 2000 were located in Hari Rasu, Ethiopia (4.0% (95% uncertainty interval, 1.1-9.7%)), Gabi Rasu, Ethiopia (4.8% (95% uncertainty interval, 1.4–11.4%)), and Isa, Nigeria (4.9% (95% uncertainty interval, 1.5–10.8%)). In 2000, 60 districts had a coverage below 10%; there was one such district in 2019. The three lowest-coverage districts in 2019 were all located in Afghanistan: Poruns (9.2% (95% uncertainty interval, 2.0-25.5%)), Wama (12.1% (95% uncertainty interval, 2.8–32.6%)) and Waygal (12.7% (95% uncertainty interval, 3.0-34.2%)).

In the period from 2000 to 2010, there were substantial increases in coverage and progress towards reducing subnational heterogeneity. The period from 2010 to 2019, however, showed slowing progress and, in some cases, regression of coverage compared to the 2000–2010 period (Fig. 2). Between 2000 and 2010, 70.5% (95% uncertainty interval, 66.0-75.4%) of districts increased coverage, but between 2010 and 2019, coverage increased in only 40.1% (95% uncertainty interval, 34.2-46.9%) of districts (Extended Data Fig. 8). This finding persists even when accounting for the starting level of coverage (Supplementary Information section 2.3).

Although district-level MCV1 coverage generally increased between 2000 and 2019, further gains are required to reach both 80% and 95% key coverage targets (Extended Data Fig. 9). In 2000, 38.4% of districts had a high probability (>95% posterior probability) of reaching the GVAP target of 80% district-level MCV1 coverage, which remained stagnant at 33.2% of districts in 2019. Only 15 countries had a high probability of reaching the GVAP target of greater than 80% district-level coverage in all districts.

Quantifying geographical inequalities

To further assess the effect of geographical heterogeneity in MCV1 coverage, we computed the absolute geographical inequality,

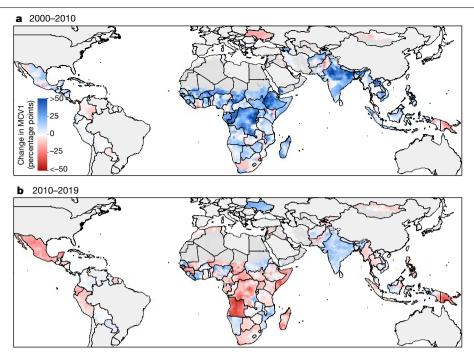


Fig. 2 | Estimated absolute changes in MCV1 coverage in the early (2000-2010) and late (2010-2019) study periods. a, b, Mean district-level absolute differences in MCV1 coverage from 2000 to 2010 (a) and from 2010 to 2019 (b).

Countries excluded from the analysis and pixels classified as 'barren or sparsely vegetated' based on ESA-CCI satellite data or with fewer than 10 $people\,per\,1\times1\text{-km}^2\,pixel\,based\,on\,WorldPop\,estimates\,are\,masked\,in\,grey^{30,50}.$

a Gini-coefficient-derived metric that ranges between zero (perfectly homogenous coverage) and one (perfectly heterogeneous coverage), at the 5 × 5-km² level. In 2000, nine countries (Cameroon, Democratic Republic of the Congo, Guinea, India, Laos, Madagascar, Mali, Nigeria and Yemen) had high absolute geographical inequality (above 0.15). In 2019, only five countries had high absolute geographical inequality (Angola, Ethiopia, Madagascar, Nigeria and Pakistan). Nigeria had absolute geographical inequality above or equal to 0.2 in both 2000 and 2019, and 25 countries had increased absolute geographical inequality in 2019 compared with 2000. Notably, absolute geographical inequality decreased considerably in India, from 0.23 in 2000 to 0.07 in 2019.

In general, and as expected, improvements in national-level coverage over time were accompanied by reductions in subnational absolute geographical inequality (Fig. 3). Changes in coverage were negatively correlated ($\rho = -0.47$, Pearson's product moment test statistic = -5.36, P < 0.001) with changes in absolute inequality. India is a true exemplar in this trend, with sizeable reductions in inequality occurring as coverage increased. This improvement was not the only pathway for a country, however; some countries with increasing coverage also experienced increasing inequality, such as Chad and Ethiopia. Other countries experienced decreasing coverage alongside increasing inequality, such as Angola.

Coverage in urban and rural areas

In a post hoc analysis, we compared vaccination status in urban and remote rural settings, using proxies of travel time of ≤ 30 min and ≥ 3 h, respectively, to the nearest major city or settlement²⁹ and the number of children under 5 years old 30 from gridded estimates. In 2019, MCV1 coverage was relatively lower in remote rural areas: in remote rural locations, 33.3% of children were MCV-unvaccinated, compared with 15.2% of children living in urban areas. Owing to the concentration of populations in urban centres, however, more unvaccinated children lived in urban locations (47.9% of all unvaccinated children) than remote rural areas (16.0% of all unvaccinated children) in 2019, although this pattern varied across countries and regions (Fig. 4). For example, in Chad, 19.3% of unvaccinated children lived in urban locations and 44.4% lived in remote rural locations in 2019. In Mexico in 2019, 72.3% of unvaccinated children lived in urban locations and 3.4% lived in remote rural locations (Extended Data Fig. 10).

Our results show the variability of urban and rural contributions to unvaccinated populations in each country and region. In Latin America and the Caribbean, for example, MCV1 coverage is generally similar between urban and rural settings (Fig. 4); the urban-remote rural distribution of unvaccinated children therefore largely reflects the underlying population distribution. In other regions, the interaction between population and coverage is more complex. In South Asia, for example, 21 times more unvaccinated children live in urban locations compared with remote rural locations. Strategies focused solely on reaching the most unvaccinated children possible in that region, therefore, might reasonably prioritize urban areas. Overall, however, MCV1 coverage in urban areas of South Asia averages 90.7%, compared to only 77.4% in remote rural areas. Approaches that focus primarily on reaching urban areas, therefore, would probably exacerbate existing urban-rural coverage inequalities.

Discussion

Our MCV1 coverage estimates show overall progress from 2000 to 2019, corresponding to the creation of benchmark targets from the Measles and Rubella Initiative and GVAP, as well as substantial financial support for comprehensive vaccination programming generated by the introduction of Gavi, the Vaccine Alliance^{5,6,31-34}. Moreover, 62 out of 101 countries increased national-level MCV1 coverage while reducing subnational geographical inequalities over time, a noteworthy achievement.

This remarkable global progress should be celebrated, but this trend was not universal. Our results show a decline and stagnation in routine MCV1 coverage in certain locations, particularly since 2010, that may be related to conflict, vaccine scepticism, available funding support and supply disruptions³⁵. Among countries with stagnant or declining coverage rates, the Central African Republic and Nigeria are experiencing

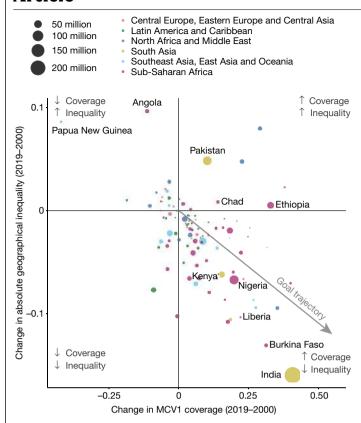


Fig. 3 | Absolute geographical inequality of MCV1 coverage in 2000 and 2019. We compared the change in geographical absolute inequality to change in national-level coverage from 2000 to 2019. Points are sized by under-5 population size.

ongoing political instability and conflict, which serve as major barriers to the success of vaccination programmes $^{36-38}$. Supply disruptions also present a major barrier to achieving and sustaining high levels of MCV coverage. For example, in 2018, the Philippines reported a national-level MCV stockout³⁹. The stockout, alongside waning public confidence in vaccination programmes stemming from misinformation related to risks of the Dengvaxia dengue vaccine, contributed to a national-level drop in coverage from 80% to 69% between 2017 and 2018⁴⁰. In Angola, economic turmoil led to a 28% decrease in governmental health spending per capita between 2010 and 2018, which may have also contributed to declines in vaccination coverage⁴¹. While global immunization initiatives have often focused on low-income countries, districts in middle-income countries also experienced recent declines, emphasizing the need for reliable immunization programmes and monitoring in these nations⁴². In Indonesia, for instance, 3 districts had coverage that reached 95% in 2000, increasing to 13 in 2010, but decreasing back to zero in 2019. In addition, countries with higher than average vaccine scepticism, such as Peru and Moldova, also experienced decreasing coverage rates and increasing within-country inequality⁴³.

Overlaid on these persistent challenges, the ongoing COVID-19 pandemic has caused the cancellation of supplementary measles immunization campaigns and puts the delivery of critical routine immunization services at risk^{7,8}. Baseline subnational estimates of routine MCV1 immunization in LMICs can help to define the geographical areas of highest vulnerability to pandemic-associated disruptions. To mitigate the risk of measles outbreaks in the setting of the COVID-19 pandemic, the maintenance of routine immunization services is crucial⁴⁴ – particularly in areas with pre-existing routine immunization weaknesses.

Even before the current pandemic, few countries reached the GVAP target of 80% coverage in all districts by 2019. Stagnant progress

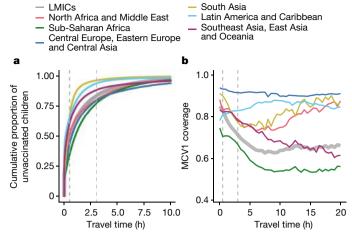


Fig. 4 | Vaccination status in 2019 and geographical remoteness. Cumulative proportion of unvaccinated children living within the spectrum of the travel time (in hours) to a major city or settlement per region (left) and coverage among children living within the spectrum of travel time to a major city or settlement per region (right). Vertical dashed grey line shows thresholds for 'urban' and 'remote rural', living within 30 min and at least 3 h from a major city or settlement, respectively.

between 2010 and 2019 further suggests that new approaches are needed to reach unvaccinated children and resolve geographical inequalities. As the era of GVAP draws to a close and the Immunization Agenda 2030 begins, these results provide a platform to identify successes and inform strategies for the next decade. India, for instance, saw exemplary improvement in both national-level coverage and geographical equality over time. This may be attributable to specific inter $ventions\,such\,as\,Mission\,Indradhanush, launched\,in\,2014\,with\,the\,goal\,and\,anush$ of targeting underserved unvaccinated populations⁴⁵. In addition, India introduced a second dose of MCV (MCV2) in select subnational geographies with MCV1 coverage below 80% starting in 2008, and $expanded\,MCV2\,to\,cover\,all\,districts\,in\,2010\,through\,the\,strengthening$ of both routine and supplementary immunization programmes 46,47. The introduction of MCV2 into the national schedule may provide a second opportunity for first-dose vaccination among children who missed the scheduled MCV1 dose. Understanding the specific drivers of simultaneous coverage and equality gains may provide critical insights for the immunization agenda in countries and regions that have fallen behind.

The Equity Reference Group for Immunization highlights the need for increased attention on vaccinating vulnerable children who live in remote rural, urban poor and conflict settings, as well as for equality in coverage by gender⁴⁸. These recommendations suggest that the agenda to leave no child unvaccinated, set by global partners and the Sustainable Development Goals, should transcend geography types and aim to eliminate coverage gaps among children who live in both urban and remote rural areas⁴⁹. These geographically resolved MCV1 estimates provide a tool for decision-makers to plan supplementary immunization activities and routine immunization strengthening programmes, to reach both the urban and remote rural communities where unvaccinated children live.

Despite large improvements made in MCV1 coverage from routine immunization programmes between 2000 and 2019, stalling progress and substantial subnational variation remain in many LMICs, leaving children at risk of preventable death. Policymakers should note where progress is most critically needed to successfully meet global immunization targets and protect the most-vulnerable children against measles. Our subnational estimates of routine MCV1 coverage at policy-relevant scales provide a tool for decision-makers to use in advocating for strong, sustainable immunization programmes that provide equitable protection for all children.

Online content

Any methods, additional references. Nature Research reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information; details of author contributions and competing interests; and statements of data and code availability are available at https://doi.org/10.1038/s41586-020-03043-4.

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Local Burden of Disease Vaccine Coverage Collaborators

Alyssa N. Sbarra¹, Sam Rolfe¹, Jason Q. Nguyen¹, Lucas Earl¹, Natalie C. Galles¹, Ashley Marks¹, Kaja M. Abbas², Mohsen Abbasi-Kangeyari³, Hedayat Abbastabar⁴, Foad Abd-Allah⁵, Ahmed Abdelalim⁵, Mohammad Abdollahi^{6,7}, Kedir Hussein Abegaz^{8,9} Hailemariam Abiy Alemu Abiy^{10,11}, Hassan Abolhassani^{12,13}, Lucas Guimarães Abreu¹⁴, Michael R. M. Abrigo¹⁵, Abdelrahman I. Abushouk^{16,17}, Manfred Mario Kokou Accrombess Maryam Adabi²⁰, Oladimeji M. Adebayo²¹, Victor Adekanmbi²², Olatunji O. Adetokunboh^{23,24}, Davoud Adham²⁵, Mohsen Afarideh^{26,27}, Mohammad Aghaali²⁸, Tauseef Ahmad²⁹, Raman Ahmadi^{30,31}, Keivan Ahmadi³², Muktar Beshir Ahmed^{33,24}, Fahad Mashhour Alanezi³⁵, Turki M. Alanzi³⁶, Jacqueline Elizabeth Alcalde-Rabanal³⁷, Birhan Tamene Alemnew³⁸, Beriwan Abdulqadir Ali^{39,40}, Muhammad Ali⁴¹, Mehran Alijanzadeh⁴², Cyrus Alinia⁴³ Reza Alipoor⁴⁴, Vahid Alipour^{45,46}, Hesam Alizade⁴⁷, Syed Mohamed Aljunid^{48,49} Ali Almasi⁵⁰, Amir Almasi-Hashiani⁵¹, Hesham M. Al-Mekhlafi^{52,53}, Khalid A. Altirkawi⁵⁴, Bekalu Amare⁵⁵, Saeed Amini⁵⁶, Mostafa Amini-Rarani⁵⁷, Fatemeh Amiri⁵⁸, Arianna Maever L. Amit^{59,6} Dickson A. Amugsi⁶¹, Robert Ancuceanu⁶², Catalina Liliana Andrei⁶³, Mina Anjomshoa⁶⁴ Fereshteh Ansari^{65,66}, Alireza Ansari-Moghaddam⁶⁷, Mustafa Geleto Ansha⁶⁸, Carl Abelardo T. Antonio 69,70, Ernoiz Antriyandarti 71, Davood Anvari 72,73, Jalal Arabloo 45, Morteza Arab-Zozani⁷⁴, Olatunde Aremu⁷⁵, Bahram Armoon^{76,77}, Krishna K. Aryal⁷⁸, Afsaneh Arzani^{79,80}, Mehran Asadi-Aliabadi⁸¹, Samaneh Asgari⁸², Zahra Atafar⁸³ Marcel Ausloos^{84,85}, Nefsu Awoke⁸⁶, Beatriz Paulina Ayala Quintanilla⁸⁷, Martin Amogre Ayanore⁸⁸, Yared Asmare Aynalem⁸⁹, Abbas Azadmehr⁹⁰, Samad Azari⁴⁵, Ebrahim Babaee⁸¹, Alaa Badawi^{91,92}, Ashish D. 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¹Institute for Health Metrics and Evaluation, University of Washington, Seattle, WA, USA ²Department of Infectious Disease Epidemiology, London School of Hygiene & Tropical Medicine, London, UK, 3Social Determinants of Health Research Center, Shahid Beheshti University of Medical Sciences, Tehran, Iran, ⁴Advanced Diagnostic and Interventional Radiology Research Center, Tehran University of Medical Sciences, Tehran, Iran, 5 Department of Neurology, Cairo University, Cairo, Egypt. ⁶The Institute of Pharmaceutical Sciences (TIPS), Tehran University of Medical Sciences, Tehran, Iran. 7School of Pharmacy, Tehran University of Medical Sciences, Tehran, Iran. 8Department of Biostatistics, Near East University, Nicosia Cyprus. 9Department of Biostatistics and Health Informatics, Madda Walabu University, Bale Robe, Ethiopia. 10 Department of Public Health, Debre Markos University, Debre Markos, Ethiopia. 11 School of Public Health, Bahir Dar University, Bahir Dar, Ethiopia. 12 Department of Laboratory Medicine, Karolinska University Hospital, Huddinge, Sweden. ¹³Research Center for Immunodeficiencies, Tehran University of Medical Sciences, Tehran, Iran. 14Department of Pediatric Dentistry, Federal University of Minas Gerais, Belo Horizonte, Brazil. 15 Department of Research, Philippine Institute for Development Studies, Quezon City, The Philippines. ¹⁶Department of Cardiovascular Medicine, Cleveland Clinic, Cleveland, OH, USA. ¹⁷Department of Medicine, Ain Shams University, Cairo, Egypt. ¹⁸Department of Disease Control, London School of Hygiene & Tropical Medicine, London, UK. 19 Clinical Research and Operations, Foundation for Scientific Research (FORS), Cotonou, Benin. 20 Hamadan University of Medical Sciences, Hamadan, Iran. 21 College of Medicine, University College Hospital, Ibadan, Ibadan, Nigeria. ²²Population Health Sciences, King's College London, London, UK. ²³Centre of Excellence for Epidemiological Modelling and Analysis, Stellenbosch University, Stellenbosch, South Africa. 24Department of Global Health, Stellenbosch University, Cape Town, South Africa, 25School of Health, Ardabil University of Medical Science, Ardabil, Iran, ²⁶Department of Dermatology, Mayo Clinic, Rochester, MN, USA. ²⁷Endocrinology and Metabolism Research Center, Tehran University of Medical Sciences, Tehran, Iran. ²⁸Department of Epidemiology and Biostatistics, Qom University of Medical Sciences, Qom, Iran. 29 Department of Epidemiology and Health Statistics, Southeast University, Nanjing, China. 30 Drug Applied Research Center, Tabriz University of Medical Sciences, Tabriz, Iran. ³¹Department of Food Science and Technology, University of Tabriz, Tabriz, Iran. ³²Lincoln Medical School, Universities of Nottingham & Lincoln, Lincoln, UK. 33 Department of Epidemiology, , Jimma University, Jimma, Ethiopia. 34Australian Center for Precision Health, , University of South Australia, Adelaide, South Australia, Australia. 35 Imam Abdulrahman Bin Faisal University, Dammam, Saudi Arabia. 36 Health Information Management and Technology Department, Imam Abdulrahman Bin Faisal University, Dammam, Saudi Arabia. 37 Center for Health System Research, National Institute of Public Health, Cuernavaca, Mexico. 38 Department of Medical Laboratory Science, Woldia University, Woldia, Ethiopia. 39 Erbil Technical Health College, Erbil Polytechnic University, Erbil, Iraq. 40 School of Pharmacy, Tishk International University, Erbil, Iraq. 41Department of Biotechnology, Quaid-i-Azam University, Islamabad, Pakistan. 42Social Determinants of Health Research Center, Qazvin University of Medical Sciences, Qazvin, Iran. 43Department of Health Care Management and Economics, Urmia University of Medical Science, Urmia, Iran. 44Student Research Committee, Hormozgan University of Medical Sciences, Bandar Abbas, Iran. ⁴⁵Health Management and Economics Research Center, Iran University of Medical Sciences, Tehran, Iran. 46Health Economics Department, Iran University of Medical Sciences, Tehran, Iran. 47 Infectious and Tropical Disease Research Center, Hormozgan University of Medical Sciences, Bandar Abbas, Iran, ¹⁸Department of Health Policy and Management, Kuwait University, Safat, Kuwait, ⁴⁹International Centre for Casemix and Clinical Coding, National University of Malaysia, Bandar Tun, Razak, Malaysia. 50 Department of Environmental Health Engineering, Kermanshah University of Medical Sciences, Kermanshah, Iran, 51Department of Epidemiology, Arak University of Medical Sciences, Arak, Iran. 52 Medical Research Center, Jazan University, Jazan, Saudi Arabia. 53 Department of Parasitology, Sana'a University, Sana'a, Yemen. 54 Pediatric Intensive Care Unit, King Saud University, Riyadh, Saudi Arabia. ⁵⁵Department of Pharmacology, Mekelle University, Mekelle, Ethiopia. ⁵⁶Health Services Management Department, Arak University of Medical Sciences, Arak, Iran. 57 Health Management and Economics Research Center, Isfahan University of Medical Sciences, Isfahan, Iran. 58 Department of Radiology and Nuclear Medicine, Kermanshah University of Medical Sciences, Kermanshah, Iran. 59 Department of Epidemiology and Biostatistics University of the Philippines Manila, Manila, The Philippines. 60 School of Public Health, Johns Hopkins University, Baltimore, MD, USA. 61 Maternal and Child Wellbeing, African Population and Health Research Center, Nairobi, Kenya. 62 Pharmacy Department, Carol Davila University of Medicine and Pharmacy, Bucharest, Romania. 63 Cardiology Department, Carol Davila University of Medicine and Pharmacy, Bucharest, Romania. ⁶⁴Social Determinants of Health Research Center, Rafsanjan University of Medical Sciences, Rafsanjan, Iran. ⁶⁵Research Center for Evidence Based Medicine, Tabriz University of Medical Sciences, Tabriz, Iran. 66Razi Vaccine and Serum Research Institute, Agricultural Research, Education, and Extension Organization (AREEO), Tehran, Iran. ⁶⁷Department of Epidemiology and Biostatistics, Zahedan University of Medical Sciences, Zahedan, Iran. ⁶⁸Department of Public Health, Debre Berhan University, Debre Berhan, Ethiopia. 69 Department of Health Policy and Administration, University of the Philippines Manila, Manila, The Philippines. 70 Department of Applied Social Sciences, Hong Kong Polytechnic University, Hong Kong, China. ⁷¹Agribusiness Study Program, Sebelas Maret University, Surakarta, Indonesia. ⁷²Department of Parasitology, Mazandaran University of Medical Sciences, Sari, Iran. 73 Department of Parasitology, Iranshahr University of Medical Sciences, Iranshahr, Iran. 74 Social Determinants of Health Research Center, Birjand University of Medical Sciences, Birjand, Iran. 75 Department of Public Health, Birmingham City University, Birmingham, UK. ⁷⁶Social Determinants of Health Research Center, Saveh University of Medical Sciences, Saveh, Iran. "Social Determinants of Health Research Center, Yasuj University of Medical Sciences, Yasuj, Iran. ⁷⁸Monitoring Evaluation and Operational Research Project, Abt Associates Nepal, Lalitpur, Nepal. 79 School of Nursing and Midwifery, Babol University of Medical Sciences, Babol, Iran. 80 Babol University of Medical Sciences, Babol, Iran. 81 Preventive Medicine and Public Health Research Center, Iran University of Medical Sciences, Tehran, Iran. 82 Prevention of Metabolic Disorders Research

Center, Shahid Beheshti University of Medical Sciences, Tehran, Iran, 83 Social Development and Health Promotion Research Center, Kermanshah University of Medical Sciences, Kermanshah, Iran. 84School of Business, University of Leicester, Leicester, UK. 85Department of Statistics and Econometrics, Bucharest University of Economic Studies, Bucharest, Romania. ⁸⁶Department of Nursing, Wolaita Sodo University, Wolaita Sodo, Ethiopia. ⁸⁷The Judith Lumley Centre, La Trobe University, Melbourne, Victoria, Australia. 88 Department of Health Policy Planning and Management, University of Health and Allied Sciences, Ho, Ghana ⁸⁹Department of Nursing, Debre Berhan University, Debre Berhan, Ethiopia. ⁹⁰Cellular and Molecular Biology Research Center, Babol University of Medical Sciences, Babol, Iran. ⁹¹Public Health Risk Sciences Division, Public Health Agency of Canada, Toronto, Ontario, Canada. 92 Department of Nutritional Sciences, University of Toronto, Toronto, Ontario, Canada. 93 Department of Forensic Science, Government Institute of Forensic Science, Nagpur, India. 94Department of Healthcare Management and Education, Shiraz University of Medical Sciences, Shiraz, Iran. 95 Unit of Biochemistry, Sultan Zainal Abidin University (Universiti Sultan Zainal Abidin), Kuala Terengganu, Malaysia. 96 Department of Health Policy, Management, and Economics, Tehran University of Medical Sciences, Tehran, Iran. ⁹⁷Department of Medical Microbiology, Haramaya University, Harar, Ethiopia, ⁹⁸Department of Hypertension, Medical University of Lodz, Lodz, Poland, 99Polish Mothers' Memorial Hospital Research Institute, Lodz, Poland. 100 Department of Non-communicable Diseases, Bangladesh University of Health Sciences, Dhaka, Bangladesh. ¹⁰¹Clinic for Infectious and Tropical Diseases, Clinical Center of Serbia, Belgrade, Serbia. 102 Faculty of Medicine, University of Belgrade, Belgrade, Serbia. 103 Razi Vaccine and Serum Research Institute, Agricultural Research, Education and Extension Organization (AREEO), Karaj, Iran. 104 Department of Epidemiology and Biostatistics, University of Gondar, Gondar, Ethiopia. 105 Center for Primary Care, Harvard University, Boston, MA, USA. 106School of Public Health, Imperial College London, London, UK. 107 Health Human Resources Research Center, Shiraz University of Medical Sciences, Shiraz, Iran. 108 Monitoring, Evaluation and Research Department, JSI Research & Training Institute, Addis Ababa, Ethiopia. 109 Department of Community Medicine, Gandhi Medical College Bhopal, Bhopal, India. 110 Jazan University, Jazan, Saudi Arabia ¹¹¹Social Determinants of Health Research Center, Lorestan University of Medical Sciences, Khorramabad, Iran. 112 School of the Environment, Yale University, New Haven, CT, USA. ¹¹³Department of Public Health, Arba Minch University, Arba Minch, Ethiopia. ¹¹⁴Department of Nutrition and Dietetics, Mekelle University, Mekelle, Ethiopia. 115 Department of Global Health, Global Institute for Interdisciplinary Studies, Kathmandu, Nepal. 116 Centre for Global Child Health, University of Toronto, Toronto, Ontario, Canada. 117 Centre of Excellence in Women & Child Health, Aga Khan University, Karachi, Pakistan. 118 Social Determinants of Health Research Center, Babol University of Medical Sciences, Babol, Iran. 119 Department of Pediatrics and Child Health Nursing, Bahir Dar University, Bahir Dar, Ethiopia. 120 Global Health Division, Research Triangle Institute International, Research Triangle Park, NC, USA. 121 School of Medicine, University of Nottingham, Nottingham, UK. 122 Department of Neurology, Institute of Post-Graduate Medical Education and Research and Seth Sukhlal Karnani Memorial Hospital, Kolkata, India, 123 Department of Veterinary Medicine, Islamic Azad University, Kermanshah, Iran. ¹²⁴Department of Biomedical Sciences, Nazarbayev University, Nur-Sultan City, Kazakhstan. ¹²⁵Department of Endocrinology, Hamadan University of Medical Sciences, Hamadan, Iran. ¹²⁶University of Genoa, Genoa, Italy. ¹²⁷Department of Biomedical Technologies, Bauman Moscow State Technical University, Moscow, Russia. 128 Department of Epidemiology and Evidence Based Medicine, I. M. Sechenov First Moscow State Medical University, Moscow, Russia. 129 Department of Community Medicine, Employee State Insurance Post Graduate Institute of Medical Sciences and Research, Bangalore, India ¹³⁰School of Public Health and Health Systems, University of Waterloo, Waterloo, Ontario, Canada. 131 Al Shifa School of Public Health, Al Shifa Trust Eye Hospital, Rawalpindi, Pakistan. ¹³²Internal Medicine Department, Hospital Italiano de Buenos Aires, Buenos Aires, Argentina. $^{\rm 133}\textsc{Board}$ of Directors, Argentine Society of Medicine, Buenos Aires, Argentina. $^{\rm 134}\textsc{Health}$ and Nutrition Research Center, National Institute of Public Health, Cuernavaca, Mexico. 135 Centre for Population Health Sciences, Nanyang Technological University, Singapore, Singapore. ¹³⁶Department of Primary Care and Public Health, Imperial College London, London, UK. ¹³⁷Department of Health Care, Metropolitan Autonomous University, Mexico City, Mexico. ¹³⁸Research Unit on Applied Molecular Biosciences (UCIBIO), University of Porto, Porto, Portugal. 139 Department of Psychiatry, University of São Paulo, São Paulo, Brazil. 140 Gorgas Memorial Institute for Health Studies, Panama City, Panama. 141 Department of Medicine, University of Toronto, Toronto, Ontario, Canada. 142 Department of Biostatistics and Epidemiology, Babol University of Medical Sciences, Babol, Iran. 143 Epidemiology Research Center, Royan Institute, Tehran, Iran. 144 Department of Epidemiology and Preventive Medicine, Monash University, Melbourne, Victoria, Australia. 145 Melbourne Medical School, University of Melbourne, Parkville, Victoria, Australia. 146 Faculty of Biology, Hanoi National University of Education, Hanoi, Vietnam. 147 Environmental Health Department, Imam Abdulrahman Bin Faisal University, Dammam, Saudi Arabia. 148 Clinical Dermatology, IRCCS Istituto Ortopedico Galeazzi, University of Milan, Milan, Italy, ¹⁴⁰Pepartment of Dermatology, Case Western Reserve University, Cleveland, OH, USA. ¹⁵⁰Public Health Foundation of India, Gurugram, India. 151 Department of Health Metrics Sciences, School of Medicine, University of Washington, Seattle, WA, USA. 152 Indian Council of Medical Research, New Delhi, India. ¹⁵³Immunization, Vaccines and Biologicals (IVB), World Health Organization (WHO), Geneva, Switzerland. 154Global Delivery Programs, Bill & Melinda Gates Foundation, Seattle, WA, USA. ¹⁵⁵Department of Information Technology, University of Human Development, Sulaymaniyah, Irac. ¹⁵⁶Department of Pediatrics, Tanta University, Tanta, Egypt. ¹⁵⁷Division of Women and Child Health, Aga Khan University, Karachi, Pakistan. ¹⁵⁸Heidelberg Institute of Global Health (HIGH), Heidelberg University, Heidelberg, Germany. ¹⁵⁰Department of Medical Laboratory Sciences, Bahir Dar University, Bahir Dar, Ethiopia. ¹⁵⁰Department of Nursing, Woldia University, Woldia, Ethiopia. 161 School of Nursing, Jimma University, Jimma, Ethiopia. 162 Center for Nutrition and Health Research, National Institute of Public Health, Cuernavaca, Mexico. 163 School of Nursing and Midwifery, Haramaya University, Harar, Ethiopia. 164 School of Pharmacy, Mekelle University, Mekelle, Ethiopia. 165 Department of Community Medicine, University of Peradeniya, Peradeniya, Sri Lanka. 166 Department of Microbiology, Far Western University, Mahendranagar, Nepal. ¹⁶⁷Department of Epidemiology and Biostatistics, Shahroud University of Medical Sciences, Shahroud, Iran. 168 Department of Epidemiology, Shiraz University of Medical Sciences, Shiraz, Iran. 169 Center of Complexity Sciences, National Autonomous University of Mexico, Mexico City, Mexico. 170 Faculty of Veterinary Medicine and

Zootechnics, Autonomous University of Sinaloa, Culiacán Rosales, Mexico, 171 Department of $Health\ Promotion\ and\ Education,\ University\ of\ Ibadan,\ Ibadan,\ Nigeria.\ ^{172}Institute\ of\ Health$ Economics and Technology, Hanoi, Vietnam. 173 Department of Medical Laboratory Sciences, Iran University of Medical Sciences, Tehran, Iran. 174 Department of Health Policy and Management, Tabriz University of Medical Sciences, Tabriz, Iran. ¹⁷⁵School of Medicine, Western Sydney University, Sydney, New South Wales, Australia. ¹⁷⁶Health Sciences, Royal Melbourne Institute of Technology University, Melbourne, Victoria, Australia. 177 Centre for Tropical Medicine and Global Health, University of Oxford, Oxford, UK. 178 Mahidol-Oxford Tropical Medicine Research Unit, Bangkok, Thailand. 179School of Medicine, Federal University of Bahia, Salvador, Brazil. 180 Department of Internal Medicine, Escola Bahiana de Medicina e Saúde Pública, Salvador, Brazil. 181 Department of Epidemiology, Florida International University, Miami, FL, USA. 182 Department of Bacteriology and Virology, Tabriz University of Medical Sciences, Tabriz, Iran. 183 School of Health Sciences, Universiti Sains Malaysia, Kubang Kerian, Malaysia. 184 Centre Clinical Epidemiology and Biostatistics, University of Newcastle, Newcastle, New South Wales, Australia. 185 Department of Pharmacology and Toxicology, Maragheh University of Medical Sciences, Maragheh, Iran. 186 Department of Pharmacology and Toxicology. The John Paul II Catholic University of Lublin, Lublin, Poland, 187 Biomedical Informatics and Medical Statistics Department, Alexandria University, Alexandria, Egypt, ¹⁸⁸Reference Laboratory of Egyptian Universities Hospitals, Ministry of Higher Education and Research, Cairo, Egypt. 189 Department of Food Science and Nutrition, Arsi University, Asella, Ethiopia. ¹⁹⁰Neurophysiology Department, Cairo University, Cairo, Egypt. ¹⁹¹Endemic Medicine and Hepatogastroentrology Department, Cairo University, Cairo, Egypt. ¹⁹²Ophthalmic Epidemiology Research Center, Shahroud University of Medical Sciences, Shahroud, Iran. ¹⁹³Department of Microbiology and Immunology, Suez Canal University, Ismailia, Egypt. ¹⁹⁴Department of Medicinal Chemistry, Kerman University of Medical Sciences, Kerman, Iran. ¹⁹⁵Pharmaceutics Research Center, Kerman University of Medical Sciences, Kerman, Iran. ¹⁹⁶Multiple Sclerosis Research Center, Tehran University of Medical Sciences, Tehran, Iran. ¹⁹⁷Department of Physiology, Tarbiat Modares University, Tehran, Iran. ¹⁹⁸Tehran Medical Sciences Branch, Islamic Azad University, Tehran, Iran. 199 Department of Public Health, Maragheh University of Medical Sciences, Maragheh, Iran. 200 Unit of Medical Physiology, Hawassa University, Hawassa, Ethiopia. 201 School of Public Health, Tehran University of Medical Sciences, Tehran, Iran. 202 College of Medicine, Imam Mohammad Ibn Saud Islamic University, Riyadh, Saudi Arabia. ²⁰³Department of Medical Parasitology, Mazandaran University of Medical Sciences, Sari, Iran. 204 Department of Medical and Surgical Sciences, University of Bologna, Bologna, Italy. 205 Non-communicable Diseases Research Center, Tehran University of Medical Sciences, Tehran, Iran. ²⁰⁶Research Center for Environmental Determinants of Health, Kermanshah University of Medical Sciences, Kermanshah, Iran. Department of Environmental Health Engineering, Ardabil University of Medical Science, Ardabil, Iran. 208 Department of Environmental Health Engineering, Tehran University of Medical Sciences, Tehran, Iran. 209 Department of Neurobiology, Karolinska Institute, Stockholm, Sweden. 210 Division of Neurology, University of Ottawa, Ottawa, Ontario, Canada. Associated Laboratory for Green Chemistry (LAQV), University of Porto, Porto, Portugal. ²¹²Psychiatry Department, Kaiser Permanente, Fontana, CA, USA. ²¹³School of Health Sciences, A. T. Still University, Mesa, AZ, USA. ²¹⁴Institute of Gerontological Health Services and Nursing Research, Ravensburg-Weingarten University of Applied Sciences, Weingarten, Germany ²¹⁵Institute of Gerontology, National Academy of Medical Sciences of Ukraine, Kyiv, Ukraine. ²¹⁶Department of Child Dental Health, Obafemi Awolowo University, Ile-Ife, Nigeria. ²¹⁷Department of Medical Parasitology, Abadan Faculty of Medical Sciences, Abadan, Iran. ²¹⁸Department of Dermatology, Kobe University, Kobe, Japan. ²¹⁹Gillings School of Global Public Health, University of North Carolina Chapel Hill, Chapel Hill, NC, USA. 220 Department of Public Health, Madda Walabu University, Bale Robe, Ethiopia. 221 Menelik-II College of Medical and Health Sciences, Kotebe Metropolitan University, Addis Ababa, Ethiopia. ²²²School of Public Health, Mekelle University, Mekelle, Ethiopia. ²²³Department of Nursing and Midwifery, Addis Ababa University, Addis Ababa, Ethiopia. 224 Department of Nursing, Aksum University, Aksum, Ethiopia. 225 Department of Nursing, Mekelle University, Mekelle, Ethiopia. ²²⁶Department of Reproductive Health, Mekelle University, Mekelle, Ethiopia. ²²⁷Department of Midwifery, Woldia University, Woldia, Ethiopia. 228 Department of Biostatistics, Mekelle University, Mekelle, Ethiopia. 229 Infectious Disease Research Center, Kermanshah University of Medical Sciences, Kermanshah, Iran. 230 Pediatric Department, Kermanshah University of Medical Sciences, Kermanshah, Iran. ²³¹Medical Education Research Center, Tabriz University of Medical Sciences, Tabiz, Iran. 232 Department of Parasitology and Entomology, Tarbiat Modares University, Tehran, Iran. ²³³Department of Neurology, Tehran University of Medical Sciences, Tehran, Iran. 234 Faculty of Nursing and Midwifery, Kurdistan University of Medical Sciences, Sanandaj, Iran. 235 Student Research Committee, Iran University of Medical Sciences, Tehran, Iran. 236 Institute of Health Research, University of Exeter, Exeter, UK. ²³⁷Young Researchers and Elite Club, Islamic Azad University, Rasht, Iran. ²³⁸Department of Biology, Islamic Azad University, Tehran, Iran. 239 Faculty of Allied Health Sciences, The University of Lahore, Lahore, Pakistan. ²⁴⁰Afro-Asian Institute, Lahore, Pakistan. ²⁴¹Medical School, University of Warwick, Coventry, UK. 242 Department of Chemistry, University of Porto, Porto, Portugal. ²⁴³Hudson College of Public Health, University of Oklahoma Health Sciences Center, Oklahoma City, OK, USA. ²⁴⁴Department of Health and Social Affairs, Government of the Federated States of Micronesia, Palikir, Federated States of Micronesia. ²⁴⁵Postgraduate Program in Epidemiology, Federal University of Rio Grande do Sul, Porto Alegre, Brazil. ²⁴⁶Department of Dermatology, Boston University, Boston, MA, USA. ²⁴⁷Institute of Tropical Pathology and Public Health (IPTSP), Federal University of Goias, Goiânia, Brazil. ²⁴⁸Department of Epidemiology, Binzhou Medical University, Yantai City, China. ²⁴⁹Medical Resources, March of Dimes, Arlington, VA, USA. 250 Health Policy, Management and Leadership, West Virginia University School of Public Health, Morgantown, WV, USA ²⁵¹Department of Radiology and Radiological Sciences, Johns Hopkins University, Baltimore, MD, USA. 252 School of Medicine, Tehran University of Medical Sciences, Tehran, Iran. ²⁵³Department of Pharmacology, Tehran University of Medical Sciences, Tehran, Iran. ²⁵⁴Obesity Research Center, Shahid Beheshti University of Medical Sciences, Tehran, Iran. ²⁵⁵Department of Public Health, Wachemo University, Hossana, Ethiopia. ²⁵⁶University Institute of Public Health, The University of Lahore, Lahore, Pakistan. ²⁵⁷Tabriz University of Medical Sciences, Tabriz, Iran. 258 Department of Zoology and Entomology, Al Azhar University, Cairo, Egypt. 259 Institute for Social Science Research, The University of Queensland, Indooroopilly, Queensland, Australia. 260 ARC Centre of Excellence for Children and Families over the Life

Course, The University of Queensland, Indooroopilly, Queensland, Australia. 261 Department of Healthcare Management, Maragheh University of Medical Sciences, Maragheh, Iran ²⁶²Department of Microbiology, Maragheh University of Medical Sciences, Maragheh, Iran. ²⁶³Department of Microbiology, Tehran University of Medical Sciences, Tehran, Iran. ²⁶⁴Gastrointestinal and Liver Diseases Research Center, Guilan University of Medical Sciences, Rasht, Iran. ²⁶⁵Caspian Digestive Disease Research Center, Guilan University of Medical Sciences, Rasht, Iran. ²⁶⁸School of Nursing and Midwifery, Tabriz University of Medical Sciences, Tabriz, Iran. ²⁶⁸Independent Consultant, Tabriz, Iran. ²⁶⁸School of Nursing and Midwifery, Tehran University of Medical Sciences, Tehran, Iran. ²⁶⁹Big Data Institute, University of Oxford, Oxford, UK. 270 School of Business, London South Bank University, London, UK. ²⁷¹Medical Biology Research Center, Kermanshah University of Medical Sciences, Kermanshah, Iran. 272 Guilan Road Trauma Research Center, Guilan University of Medical Sciences, Rasht, Iran. 273 Centre for Bio Cultural Studies (CBiCS), Manipal Academy of Higher Education, Manipal, India. 274 Department of Pharmacology, Bangladesh Industrial Gases Limited, Tangail, Bangladesh. ²⁷⁵Department of Epidemiology and Biostatistics, Tehran University of Medical Sciences, Tehran, Iran. ²⁷⁶Pediatric Chronic Kidney Disease Research Center, Tehran University of Medical Sciences, Tehran, Iran. 277 Institute of Research and Development, Duy Tan University, Da Nang, Vietnam. ²⁷⁸Department of Computer Science, University of Human Development, Sulaymaniyah, Iraq. ²⁷⁸College of Science and Engineering, Hamad Bin Khalifa University, Doha, Qatar. ²⁸⁰Department of Epidemiology and Health Statistics, Central South University, Changsha, China. 281 School of Public Health, University of Sydney, Sydney, New South Wales, Australia. 282 Maternal and Child Health Division, International Centre for Diarrhoeal Disease Research, Bangladesh, Dhaka Bangladesh. ²⁸³Department of Public Health and Community Medicine, Shaikh Khalifa Bin Zayed Al-Nahyan Medical College, Lahore, Pakistan. ²⁸⁴Department of Community Medicine, University of Ibadan, Ibadan, Nigeria. ²⁸⁵Department of Community Medicine, University College Hospital, Ibadan, Ibadan, Nigeria. 286 Department of Epidemiology, University of Kragujevac, Kragujevac, Serbia. 287 Department of Public Health, Lorestan University of Medical Sciences, Khorramabad, Iran. 288 Division of Community Health and Family Medicine, Bangalore Baptist Hospital, Bangalore, India. 289 College of Public Health, Taipei Medical University, Taipei, Taiwan. 290 Research Institute for Endocrine Sciences, Shahid Beheshti University of Medical Sciences, Tehran, Iran. 291 Institute for Physical Activity and Nutrition, Deakin University, Burwood, Victoria, Australia. 292 Sydney Medical School, University of Sydney, Sydney, New South Wales, Australia. 293 School of Psychology and Public Health, La Trobe University, Melbourne, Victoria, Australia. 294 South African Medical Research Council, Cape Town, South Africa. 295 School of Health Systems and Public Health, University of Pretoria, Pretoria, South Africa. ²⁹⁶Department of Immunology, Tabriz University of Medical Sciences, Tabriz, Iran. 297 Department of Immunology, Isfahan University of Medical Sciences, Isfahan, Iran. 298 School of Management and Medical Education, Shahid Beheshti University of Medical Sciences, Tehran, Iran. ²⁹⁹Safety Promotion and Injury Prevention Research Center, Shahid Beheshti University of Medical Sciences, Tehran, Iran. 300 N. A. Semashko Department of Public Health and Healthcare, I. M. Sechenov First Moscow State Medical University, Moscow, Russia. 301 Department of Global Health, Economics and Policy, University of Kragujevac, Kragujevac, Serbia. 302 Health Institute, Kermanshah University of Medical Sciences, Kermanshah, Iran. 303 Substance Abuse Prevention Research Center, Kermanshah University of Medical Sciences, Kermanshah, Iran. ³⁰⁷Department of Medical Mycology, Mazandaran University of Medical Sciences, Sari, Iran. ³⁰⁵Autism Spectrum Disorders Research Center, Hamadan University of Medical Sciences, Hamadan, Iran. 306The George Institute for Global Health, New Delhi, India. 307 Manipal Academy of Higher Education, Manipal, India. 308 Environmental Research Center, Duke Kunshan University, Kunshan, China. 309 Nicholas School of the Environment, Duke University, Durham, NC, USA. 310 Renal and Cardiovascular Division, The George Institute for Global Health, New Delhi, India ³¹¹Department of Medicine, University of New South Wales, Sydney, New South Wales, Australia. 312 Department of Family Medicine and Public Health, University of Opole, Opole, Poland. 313 School of Public Health, University College Cork, Cork, Ireland. 314 Minimally Invasive Surgery Research Center, Iran University of Medical Sciences, Tehran, Iran. 315 Infectious Diseases Research Center, Golestan University of Medical Sciences, Gorgan, Iran. 316 School of Management and Medical Informatics, Tabriz University of Medical Sciences, Tabriz, Iran. 317 Institute for Prevention of Non-communicable Diseases, Qazvin University of Medical Sciences, Qazvin, Iran. 318 Health Services Management Department, Qazvin University of Medical Sciences, Qazvin, Iran. 319 Department of Pharmacy, Shaheed Benazir Bhutto University, Upper Dir, Pakistan. 320 School of Pharmacy, Shanghai Jiao Tong University, Shanghai, China. 321 Department of Forensic Medicine and Toxicology, All India Institute of Medical Sciences, Jodhpur, India. 322 Department of Epidemiology, Hamadan University of Medical Sciences, Hamadan, Iran. 323 Institute for Epidemiology and Social Medicine, University of Münster, Münster, Germany. 324 Social Determinants of Health Research Center, Tabriz University of Medical Sciences, Tabriz, Iran, 325 International Research $Center of Excellence, Institute of Human Virology Nigeria, Abuja, Nigeria. \\ ^{326} Julius Centre for Center of Excellence, Institute of Human Virology Nigeria, Abuja, Nigeria, Niger$ Health Sciences and Primary Care, Utrecht University, Utrecht, The Netherlands. 327Open, Distance and eLearning Campus, University of Nairobi, Nairobi, Kenya. 328 Department of Public Health, Jordan University of Science and Technology, Irbid, Jordan. 329 Social Determinants of Health Research Center, Ahvaz Jundishapur University of Medical Sciences, Ahvaz, Iran. 330 Health Promotion Research Center, Zahedan University of Medical Sciences, Zahedan, Iran. 331 Department of Internal Medicine, John H. Stroger, Jr Hospital of Cook County, Chicago, IL, USA. 332 Department of Internal Medicine, Dow University of Health Sciences, Karachi, Pakistan. 333 Department of Epidemiology and Biostatistics, Health Services Academy, Islamabad, Pakistan. ³³⁴Department of Population Studies, International Institute for Population Sciences, Mumbai, India. ³³⁵Department of Population Science, Jatiya Kabi Kazi Nazrul Islam University, Mymensingh, Bangladesh. 336 Faculty of Health and Wellbeing, Sheffield Hallam University, Sheffield, UK. 337 College of Arts and Sciences, Ohio University, Zanesville, OH, USA. 338 Department of Medical Parasitology, Cairo University, Cairo, Egypt. 339Global Evidence Synthesis Initiative, Datta Meghe Institute of Medical Sciences, Wardha, India. ³⁴⁰Shahid Beheshti University of Medical Sciences, Tehran, Iran. ³⁴¹The Iranian Academy of Medical Sciences, Tehran, Iran. 342 Department of Neurology, Hamadan University of Medical Sciences, Hamadan, Iran. 343Deputy for Public Health, Ministry of Health and Medical Education, Tehran, Iran. 344 Health Equity Research Center, Tehran University of Medical Sciences, Tehran, Iran. 345 Department of Public Health, New Mexico State University, Las

Cruces, NM, USA. 346 Department of Public Health, Kermanshah University of Medical Sciences, Kermanshah, Iran. ³⁴⁷School of Traditional Chinese Medicine, Xiamen University Malaysia, Sepang, Malaysia. ³⁴⁸Department of Nutrition, Simmons University, Boston, MA, USA. 349 School of Health Sciences, Kristiania University College, Oslo, Norway. 350 Global Community Health and Behavioral Sciences, Tulane University, New Orleans, LA, USA. ³⁵¹Department of Nursing and Health Promotion, Oslo Metropolitan University, Oslo, Norway. ³⁵²Department of Public Health, Ambo University, Ambo, Ethiopia. ³⁵³Neurophysiology Research Center, Hamadan University of Medical Sciences, Hamadan, Iran. 354 Brain Engineering Research Center, Institute for Research in Fundamental Sciences, Tehran, Iran. 355 Independent Consultant, Jakarta, Indonesia. 356 Department of Internal and Pulmonary Medicine, Sheri Kashmir Institute of Medical Sciences, Srinagar, India. 357 CIBERSAM, San Juan de Dios Sanitary Park, Sant Boi de Llobregat, Spain. 358 Catalan Institution for Research and Advanced Studies (ICREA), Barcelona, Spain. 359 Department of Zoology, University of Oxford, Oxford, UK. 360 Harvard Medical School, Harvard University, Boston, MA, USA. 361 Department of Anthropology, Panjab University, Chandigarh, India. 362 Department of Demography, University of Montreal, Montreal, Quebec, Canada. 363 Department of Social and Preventive Medicine, University of Montreal, Montreal, Quebec, Canada. 364 Department of Psychiatry, University of Nairobi, Nairobi, Kenya. 365 Division of Psychology and Language Sciences, University College London, London, UK. 366 International Institute for Population Sciences, Mumbai, India. ³⁶⁷Imperial College Business School, Imperial College London, London, UK. ⁶⁸Faculty of Public Health, University of Indonesia, Depok, Indonesia. ³⁶⁹Department of Clinical Sciences and Community Health, University of Milan, Milan, Italy. 370 Nuffield Department of Population Health, University of Oxford, Oxford, UK, 371 National Institute of Health Research (NIHR), Tehran University of Medical Sciences, Tehran, Iran. 372 Department of Pediatrics, Post Graduate Institute of Medical Education and Research, Chandigarh, India. ³⁷³Department of Essential Medicines and Health Products, Clinton Health Access Initiative, Boston, MA, USA. ³⁷⁴Department of Community and Family Medicine, University of Baghdad, Baghdad, Iraq. ³⁷⁵HelpMeSee, New York, NY, USA. ³⁷⁶Mexican Institute of Ophthalmology, Queretaro, Mexico. 377 Department of Otorhinolaryngology, Father Muller Medical College, Mangalore, India. 378 School of Pharmacy, Monash University, Bandar Sunway, Malaysia. ³⁷⁹School of Pharmacy, Taylor's University Lakeside Campus, Subang Jaya, Malaysia. ³⁸⁰School of Nursing, Hong Kong Polytechnic University, Hong Kong, China. 381 School of Public Health, Wolaita Sodo University, Wolaita Sodo, Ethiopia. 382 School of Public Health and Preventive Medicine, Monash University, Melbourne, Victoria, Australia. 383 Chinese Center for Disease Control and Prevention, Beijing, China. 384 David Geffen School of Medicine, University of California Los Angeles, Los Angeles, CA, USA. 385 Center for Integration of Data and Health Knowledge, Oswald Cruz Foundation (FIOCRUZ), Salvador, Brazil. 386 Centre for Global Mental Health (CGMH), London School of Hygiene & Tropical Medicine, London, UK. 387 Grants, Innovation and Product Development Unit, South African Medical Research Council, Cape Town, South Africa. 388 Department of Public Health, Urmia University of Medical Science, Urmia, Iran. 389 Department of Clinical Physiology, Tribhuvan University, Kathmandu, Nepal. Department of Forensic Medicine, Rajiy Gandhi University of Health Sciences, Dharwad. India. 391 Department of Forensic Medicine, Shri Dharmasthala Manjunatheshwara University, Dharwad, India 392 Clinical Research Development Center, Kermanshah University of Medical Sciences, Kermanshah, Iran. 393 Department of Humanities and Social Sciences, Indian Institute of Technology, Roorkee, Roorkee, India. 394 Department of Development Studies, International Institute for Population Sciences, Mumbai, India. ³⁹⁵Department of Maternal and Child Nursing and Public Health, Federal University of Minas Gerais, Belo Horizonte, Brazil. ³⁹⁶Department of Health Education and Promotion, Iran University of Medical Sciences, Tehran, Iran. ³⁹⁷Campus Caucaia, Federal Institute of Education, Science and Technology of Ceará, Caucaia, Brazil. ³⁹⁸Faculty of Health and Education, Botho University-Botswana, Gaborone, Botswana. 399ICF International, DHS Program, Rockville, MD, USA. 400Neurology Department, Janakpuri Super Specialty Hospital Society, New Delhi, India. 401 Department of Neurology, Govind Ballabh Institute of Medical Education and Research, New Delhi, India. ⁴⁰²Nutrition Health Research Center, Iran University of Medical Sciences, Hamadan, Iran. ⁴⁰³Department of Epidemiology and Biostatistics, University of California San Francisco, San Francisco, CA, USA. 404 Institute of Human Virology, University of Maryland, Baltimore, MD, USA. 405 Peru Country Office, United Nations Population Fund (UNFPA), Lima, Peru. 406 Forensic Medicine Division, Imam Abdulrahman Bin Faisal University, Dammam, Saudi Arabia. ⁴⁰⁷Department of Midwifery, Adigrat University, Adigrat, Ethiopia. ⁴⁰⁸Department of Reproductive Health and Population Studies, Bahir Dar University, Bahir Dar, Ethiopia. ⁴⁰⁹Clinical Microbiology and Parasitology Unit, Dr Zora Profozic Polyclinic, Zagreb, Croatia. ⁴¹⁰University Centre Varazdin, University North, Varazdin, Croatia. ⁴¹¹Department of Epidemiology and Biostatistics, Bahir Dar University, Bahir Dar, Ethiopia. 412 Department of Health Research Methods, Evidence and Impact, McMaster University, Hamilton, Ontario, Canada. 413 Department of Computer Science and Software Engineering, University of Western Australia, Perth, Western Australia, Australia, 414 Fatemen Zahra Infertility and Reproductive Health Center, Babol University of Medical Sciences, Babol, Iran, 415 Internal Medicine Programme, Kyrgyz State Medical Academy, Bishkek, Kyrgyzstan. 416 Department of Atherosclerosis and Coronary Heart Disease, National Center of Cardiology and Internal Disease, Bishkek, Kyrgyzstan. 417 Comprehensive Research Laboratory, Iran University of Medical Sciences, Tehran, Iran. 418Water Quality Research Center, Tehran University of Medical Sciences, Tehran, Iran. ⁴¹⁹Department of Rehabilitation and Sports Medicine, Kermanshah University of Medical Sciences, Kermanshah, Iran. ⁴²⁰Department of Medical Immunology, Tehran University of Medical Sciences, Tehran, Iran. 421Research Center for Biochemistry and Nutrition in Metabolic Diseases, Kashan University of Medical Sciences, Kashan, Iran. ⁴²²Institute of Addiction Research (ISFF), Frankfurt University of Applied Sciences, Frankfurt, Germany. 423 Biotechnology Research Center, Tabriz University of Medical Sciences, Tabriz, Iran. 424 Molecular Medicine Research Center, Tabriz University of Medical Sciences, Tabriz, $Iran.\ ^{425} Department\ of\ Forestry,\ Salahaddin\ University-Erbil,\ Erbil,\ Iraq.\ ^{426} Department\ of\ Salahaddin\ University-Erbil,\ Erbil,\ Erbil,\$ Medicine-Huddinge, Karolinska Institute, Stockholm, Sweden. 427 Internal Medicine Department, King Saud University, Riyadh, Saudi Arabia. 428 Department of Biology, Salahaddin University-Erbil, Erbil, Iraq. ⁴²⁹Department of Biostatistics, Hamadan University of Medical Sciences, Hamadan, Iran. ⁴³⁰Department of Epidemiology and Biostatistics, Shahrekord University of Medical Sciences, Shahrekord, Iran. 431Department of Nursing, Mashhad University of Medical Sciences, Mashhad, Iran. 432 Health Systems and Policy Research Unit, Ahmadu Bello University, Zaria, Nigeria. 433 School of Pharmacy, Haramaya University, Harar,

Ethiopia. 434 Department of Public Health, Dire Dawa University, Dire Dawa, Ethiopia. 435 Clinical Epidemiology and Public Health Research Unit, Burlo Garofolo Institute for Maternal and Child Health, Trieste, Italy. ⁴³⁶Department of Molecular Medicine, National Institute of Genetic Engineering and Biotechnology, Tehran, Iran. ⁴³⁷Health Sciences Research Center, Mazandaran University of Medical Sciences, Sari, Iran. ⁴³⁸Social Determinants of Health Research Center, Kurdistan University of Medical Sciences, Sanandaj, Iran. 439 Department of Epidemiology and Biostatistics, Kurdistan University of Medical Sciences, Sanandaj, Iran. ⁴⁴⁰National Center for Health Insurance Research, Iran Health Insurance Organization, Tehran, Iran. 441 Computer, Electrical, and Mathematical Sciences and Engineering Division, King Abdullah University of Science and Technology, Thuwal, Saudi Arabia. 442Department of Clinical Biochemistry, Babol University of Medical Sciences, Babol, Iran. 443 Department of Clinical Biochemistry, Tarbiat Modares University, Tehran, Iran. 444 Department of Food Science, University of Campinas (Unicamp), Campinas, Brazil. 445 Federal Institute for Population Research, Wiesbaden, Germany. 446 Center for Population and Health, Wiesbaden, Germany. 447 Indian Institute of Public Health, Public Health Foundation of India, Hyderabad, India. 448 Department of Microbiology and Immunology, Mekelle University, Mekelle, Ethiopia. 449Research and Analytics Department, Initiative for Financing Health and Human Development, Chennai, India. 450 Department of Research and Analytics, Bioinsilico Technologies, Chennai, India. ⁴⁵¹Suraj Eye Institute, Nagpur, India. ⁴⁵²Department for the Control of Disease, Epidemics, and Pandemics, Ministry of Public Health, Yaoundé, Cameroon. ⁴⁵³Department of Public Heath, University of Yaoundé I, Yaoundé, Cameroon. ⁴⁵⁴Department of Forensic Medicine and Toxicology, Manipal Academy of Higher Education, Manipal, India. 455 Department of Pediatrics, Arak University of Medical Sciences, Arak, Iran. 456Cochrane South Africa, South African Medical Research Council, Cape Town, South Africa. ⁴⁵⁷Department of General Surgery, Carol Davila University of Medicine and Pharmacy, Bucharest, Romania. ⁴⁵⁸Department of General Surgery, Emergency Hospital of Bucharest, Bucharest, Romania. 459Department of Biological Sciences, University of Embu, Embu, Kenya. ⁴⁶⁰Institute for Global Health Innovations, Duy Tan University, Da Nang, Vietnam. ⁴⁶¹Center of Excellence in Behavioral Medicine, Nguyen Tat Thanh University, Ho Chi Minh City, Vietnam. ⁴⁶²Institute for Mental Health and Policy, Centre for Addiction and Mental Health, Toronto, Ontario, Canada. 463 Department of Clinical Epidemiology, Institute for Clinical Evaluative Sciences, Ottawa, Ontario, Canada. 464 Department of Pharmacoeconomics and Pharmaceutical Administration, Tehran University of Medical Sciences, Tehran, Iran. 5 Hormozgan University of Medical Sciences, Bandar Abbas, Iran. 466 Public Health Department, Universitas Negeri Semarang, Kota Semarang, Indonesia. 467 Graduate Institute of Biomedical Informatics, Taipei Medical University, Taipei, Taiwan. 468 School of Public Health and Family Medicine, University of Cape Town, Cape Town, South Africa. 469 Department of Preventive Medicine, Kyung Hee University, Dongdaemun-gu, South Korea. 470 Gorgan Congenital Malformations Research Center, Golestan University of Medical Sciences, Gorgan, Iran. 471Department of Psychiatry and Behavioural Neurosciences, McMaster University, Hamilton, Ontario, Canada. 472 Department of Psychiatry, University of Lagos, Lagos, Nigeria. 473 Centre for Healthy Start Initiative, Lagos, Nigeria. 474 Diplomacy and Public Relations Department, University of Human Development, Sulaimaniyah, Iraq. ⁴⁷⁵Department of Public Health, Jigjiga University, Jijiga, Ethiopia. 476 Department of Pharmacology and Therapeutics, University of Nigeria Nsukka, Enugu, Nigeria. 477 Department of Medicine, University of Ibadan, Ibadan, Nigeria. ⁴⁷⁸Department of Medicine, University College Hospital, Ibadan, Ibadan, Nigeria. ⁴⁷⁹Department of Respiratory Medicine, Jagadguru Sri Shivarathreeswara Academy of Health Education and Research, Mysore, India. $^{\rm 480}{\rm Department}$ of Forensic Medicine, Manipal Academy of Higher Education, Mangalore, India. 481 Department of Parasitology and Mycology, Shiraz University of Medical Sciences, Shiraz, Iran. 482 Department of Health Metrics, Center for Health Outcomes & Evaluation, Bucharest Romania. 483 Department of Research, Public Health Foundation of India, Gurugram, India. ⁴⁸⁴Infectious Disease Research Center, National Institute of Public Health, Cuernavaca, Mexico. 485 Environmental Health Research Center, Kurdistan University of Medical Sciences, Sanandaj, Iran. 486 Division of General Internal Medicine, University of Pittsburgh Medical Center, Pittsburgh, PA, USA. 487School of Medicine, University of Sinu, Cartagena, Colombia. ⁸Department of Pediatrics, University of Melbourne, Melbourne, Victoria, Australia ⁴⁸⁹Population Health Theme, Murdoch Childrens Research Institute, Melbourne, Victoria, Australia. 490 Department of Physiology, Iran University of Medical Sciences, Tehran, Iran. ⁴⁹¹Physiology Research Center, Iran University of Medical Sciences, Tehran, Iran. ⁴⁹²Center for Research and Innovation, Ateneo De Manila University, Pasig City, The Philippines. ⁴⁹³Department of Cardiology, University of Bern, Bern, Switzerland. ⁴⁹⁴Institute of Microbiology and Immunology, University of Ljubljana, Ljubljana, Slovenia. 495 University Medical Center Groningen, University of Groningen, Groningen, The Netherlands. 496School of Economics and Business, University of Groningen, Groningen, The Netherlands ⁴⁹⁷Department of Nutrition and Food Sciences, Maragheh University of Medical Sciences, $Maragheh, Iran. \ ^{498} Dietary \ Supplements \ and \ Probiotic \ Research \ Center, \ Alborz \ University \ of \ Supplements \ and \ Probiotic \ Research \ Center, \ Alborz \ University \ of \ Supplements \ Only \ Only$ Medical Sciences, Karaj, Iran. ⁴⁹⁹School of Population and Public Health, University of British Columbia, Vancouver, British Columbia, Canada. 500 Department of Emergency Medicine, Kermanshah University of Medical Sciences, Kermanshah, Iran. 501 Clinical Research Center, Valle del Lili Foundation (Centro de Investigaciones Clinicas, Fundación Valle del Lili), Cali, Colombia. 502 PROESA, ICESI University (Centro PROESA, Universidad ICESI), Cali, Colombia. ⁵⁰³Department of Neurology, Smt. B.K.S. Medical Institute and Research Center, Vadodara, India. 504 Department of Community Medicine, Datta Meghe Institute of Medical Sciences, Wardha, India. 505 Department of Chemistry, Sharif University of Technology, Tehran, Iran. ⁵⁰⁶Biomedical Engineering Department, Amirkabir University of Technology, Tehran, Iran. 507 College of Medicine, University of Central Florida, Orlando, FL, USA. 508 Department of Immunology, Mazandaran University of Medical Sciences, Sari, Iran. 509 Molecular and Cell Biology Research Center, Mazandaran University of Medical Sciences, Sari, Iran. ⁵¹⁰Thalassemia and Hemoglobinopathy Research Center, Ahvaz Jundishapur University of Medical Sciences, Ahvaz, Iran. 511 Metabolomics and Genomics Research Center, Tehran University of Medical Sciences, Tehran, Iran. 512 College of Medicine & Health Sciences University of Gondar, Gondar, Ethiopia. 513 Department of Pharmacology, Shahid Beheshti University of Medical Sciences, Tehran, Iran. 514 Research Department, Policy Research Institute, Kathmandu, Nepal. 515 Health and Public Policy Department, Global Center for Research and Development, Kathmandu, Nepal. 516 Department of Oral Pathology, Srinivas Institute of Dental Sciences, Mangalore, India. 517 Institute of Collective Health, Federal

University of Bahia, Salvador, Brazil. 518 Department of Forensic Medicine and Toxicology, Manipal Academy of Higher Education, Mangalore, India. 519 Kasturba Medical College, Mangalore, Manipal Academy of Higher Education, Manipal, India. 520 Academic Public Health England, Public Health England, London, UK. SZIWHO Collaborating Centre for Public Health Education and Training, Imperial College London, UK. SZZUniversity College London Hospitals, London, UK. 523 School of Health, Medical and Applied Sciences, CQ University, Sydney, New South Wales, Australia. 524 Department of Computer Science, Boston University, Boston, MA, USA. 525 Department of Mathematical Demography & Statistics, International Institute for Population Sciences, Mumbai, India. 526 School of Nursing and Midwifery, Royal College of Surgeons in Ireland - Bahrain, Muharraq Governorate, Bahrain. 527 School of Social Sciences and Psychology, Western Sydney University, Penrith, New South Wales, Australia. 528 Translational Health Research Institute, Western Sydney University, Penrith, New South Wales, Australia. 529 Department of Health Information Management, Manipal Academy of Higher Education, Manipal, India. 530 Department of Medical Microbiology, University of Pretoria, Pretoria, South Africa. 531 Network of Immunity in Infection, Malignancy and Autoimmunity (NIIMA), Universal Scientific Education and Research Network (USAERN), Tehran, Iran, 532 Pediatric Infectious Diseases Research Center, Mazandaran University of Medical Sciences. Sari, Iran. 533 Cardiovascular Diseases Research Center, Biriand University of Medical Sciences, Birjand, Iran. 534 Epidemiology Research Unit Institute of Public Health (EPIUnit-ISPUP), University of Porto, Porto, Portugal. ⁵³⁵Department of Surgery, University of Minnesota, Minneapolis, MN, USA. ⁵³⁶Department of Surgery, University Teaching Hospital of Kigali, Kigali, Rwanda. ⁵³⁷Research Department, Faculty of Medical Sciences, National University of Caaguazu, Coronel Oviedo, Paraguay. ⁵³⁸Department of Research and Publications, National Institute of Health, Asunción, Paraguay. ⁵³⁹Department of Clinical Publications, National Institute of Health, Asuncion, Paraguay. —Department of Clinical Research, Federal University of Uberlândia, Uberlândia, Brazil. 540School of Medicine, Gonabad University of Medical Sciences, Gonabad, Iran. 541Pepartment of Biomedical Sciences, University of Sassari, Sassari, Italy. 542Pepartment of Internal Medicine, University of Botswana, Gaborone, Botswana. 543Heart and Vascular Institute, Cleveland Clinic, Cleveland, OH, USA. 544 Department of Cardiovascular Medicine, Mayo Clinic, Rochester, MN, USA ⁵⁴⁵Department of Epidemiology, Shahid Beheshti University of Medical Sciences, Tehran, Iran. ⁵⁴⁶Department of Psychiatry, All India Institute of Medical Sciences, New Delhi, India. ⁵⁴⁷Halal Research Center of IRI, Food and Drug Administration of the Islamic Republic of Iran, Tehran, Iran. 548 Neurogenic Inflammation Research Center, Mashhad University of Medical Sciences, Mashhad, Iran. 549 Department of Phytochemistry, Soran University, Soran, Iraq. 550 Department of Nutrition, Cihan University-Erbil, Erbil, Iraq. 551 Department of Anatomical Sciences, Kermanshah University of Medical Sciences, Kermanshah, Iran. 552 Department of Microbiology, Central University of Punjab, Bathinda, India. 553 Urology Department, Cairo University, Cairo, Egypt. ⁵⁵⁵Public Health and Community Medicine Department, Cairo University, Giza, Egypt. ⁵⁵⁵Center for Health Policy & Center for Primary Care and Outcomes Research, Stanford University, Stanford, CA, USA. 556 Department of Entomology, Ain Shams University, Cairo, Egypt. 557 Department of Community Medicine, PSG Institute of Medical Sciences and Research, Coimbatore, India. 558 PSG-FAIMER South Asia Regional Institute, Coimbatore, India. 559 Department of Health and Society, Faculty of Medicine, University of Applied and Environmental Sciences, Bogota, Colombia. ⁵⁶⁰National School of Public Health, Carlos III Health Institute, Madrid, Spain. ⁵⁶¹Department of Community Medicine, Mahatma Gandhi Memorial Medical College, Indore, India. 562 Faculty of Infectious and Tropical Diseases, London School of Hygiene & Tropical Medicine, London, UK. 563 Colorectal Research Center, Iran University of Medical Sciences, Tehran, Iran. 564 Department of Geriatrics and Long Term Care, Hamad Medical Corporation, Doha, Qatar. 565 Faculty of Health & Social Sciences, Bournemouth University, Bournemouth, UK. 566 Population Health Research Institute, McMaster University, Hamilton, Ontario, Canada. 567 Department of Psychology, University of Alabama at Birmingham, Birmingham, AL, USA. ⁵⁶⁸Emergency Department, Manian Medical Centre, Erode, India. ⁵⁶⁹Population Health Sciences Institute, Newcastle University, Newcastle Upon Tyne, UK. 570 Department of Health Services Management, Iran University of Medical Sciences, Tehran, Iran. ⁵⁷¹Health Policy Research Center, Shiraz University of Medical Sciences, Shiraz, Iran. ⁵⁷²Public Health Division, An-Najah National University, Nablus, Palestine. 573 Independent Consultant, Karachi, Pakistan. 574 Neurology Department, Ain Shams University, Cairo, Egypt. 575 School of Medicine, Alborz University of Medical Sciences, Karaj, Iran. 576 Department of Sports Medicine and Rehabilitation, Kermanshah University of Medical Sciences, Kermanshah, Iran. 577 Faculty of Caring Science, Work Life and Social Welfare, University of Borås, Borås, Sweden. 578 HIV/STI Surveillance Research Center and WHO Collaborating Center for HIV Surveillance, Kerman University of Medical Sciences, Kerman, Iran. 579 Centre for Medical Informatics, University of Edinburgh, Edinburgh, UK. 580 Division of General Internal Medicine, Harvard University, Boston, MA, USA. ¹Health Information Management, Iran University of Medical Sciences, Tehran, Iran 582Department of Community Medicine, Manipal Academy of Higher Education, Manipal, India. 583 National Institute of Infectious Diseases, Tokyo, Japan. 584 College of Medicine, Yonsei University, Seoul, South Korea. 585 Cancer Research Institute, Tehran University of Medical Sciences, Tehran, Iran. 586 Cancer Biology Research Center, Tehran University of Medical Sciences, Tehran, Iran. ⁵⁸⁷Department of Health Education and Health Promotion, Kermanshah University of Medical Sciences, Kermanshah, Iran. 588 School of Health, University of Technology Sydney, Sydney, New South Wales, Australia. Sep Department of Medicine, Dow University of Health Sciences, Karachi, Pakistan. Sep Department of Dermatology, George Washington University, Washington, DC, USA. 591 Department of Law, Economics, Management and Quantitative Methods, University of Sannio, Benevento, Italy. 592WSB University in Gdańsk, Gdansk, Poland. 593 School of Medicine, University of Alabama at Birmingham, Birmingham, AL, USA. 594 Medicine Service, USA Department of Veterans Affairs (VA), Birmingham, AL, USA. ⁵⁹⁵Department of Epidemiology, School of Preventive Oncology, Patna, India. ⁵⁹⁶Department of Epidemiology, Healis Sekhsaria Institute for Public Health, Mumbai, India. 597 Program Services Unit, Pathfinder International, Addis Ababa, Ethiopia. ⁵⁹⁹Nursing Care Research Center, Semnan University of Medical Sciences, Semnan, Iran. ⁵⁹⁹Department of Infectious

Diseases, Kharkiv National Medical University, Kharkiv, Ukraine. 600 Hull York Medical School, University of Hull, Hull, UK. 601 Department of Parasitology and Mycology, Tabriz University of Medical Sciences, Tabriz, Iran. 602 Division of Community Medicine, International Medical University, Kuala Lumpur, Malaysia. 603 Nursing, Muhammadiyah University of Surakarta, Surakarta, Indonesia. 604 Department of Community Medicine, Ahmadu Bello University, Zaria, Nigeria. 605 Department of Agriculture and Food Systems, University of Melbourne, Melbourne, Victoria, Australia. 606 Department of Statistics, Manonmaniam Sundaranar University, Abishekapatti, India. 607 National Institute of Epidemiology, Indian Council of Medical Research, Chennai, India. 608 Research Center for Molecular Medicine, Hamadan University of Medical Sciences, Hamadan, Iran. 609 Non-communicable Diseases Research Center, Hamadan University of Medical Sciences, Hamadan, Iran. 610 University Institute 'Egas Moniz', Monte da Caparica, Portugal. 611 Research Institute for Medicines, University of Lisbon, Lisbon, Portugal. 612 Department of Public Health, Kurdistan University of Medical Sciences, Sanandaj, Iran. 613 School of Public Health, University of Adelaide, Adelaide, South Australia, Australia. ⁵¹⁴Department of Environmental Health, Wollo University, Dessie, Ethiopia. ⁶¹⁵Department of Community and Family Medicine, Iran University of Medical Sciences, Tehran, Iran. ⁶¹⁶Department of Public Health, Adigrat University, Adigrat, Ethiopia. ⁶¹⁷Department of Pharmacognosy, Mekelle University, Mekelle, Ethiopia, 618 Department of Medical Microbiology, University of Gondar, Gondar, Ethiopia. ⁶¹⁹Department of Public Health and Community Medicine, Central University of Kerala, Kasaragod, India. 620 Institute of Public Health, Jagiellonian University Medical College, Kraków, Poland. 621 Agency for Health Technology Assessment and Tariff System, Warsaw, Poland. 622 Department of Pathology and Legal Medicine, University of São Paulo, Ribeirão Preto, Brazil. 623 Modestum, London, UK ⁶²⁴Department of Health Economics, Hanoi Medical University, Hanoi, Vietnam. ⁶²⁵Institute for Physical Activity and Nutrition, Deakin University, Melbourne, Queensland, Australia. 626School of Health and Rehabilitation Sciences, The University of Queensland, Brisbane, Queensland, Australia. 627 Department of Allied Health Sciences, Iqra National University, Peshawar, Pakistan. 628 Department of Community Medicine, Alex Ekwueme Federal University Teaching Hospital Abakaliki, Abakaliki, Nigeria. 629 Kasturba Medical College, Manipal Academy of Higher Education, Mangalore, India. 630 Amity Institute of Biotechnology, Amity University Rajasthan, Jaipur, India. 631 Alzahra Teaching Hospital, Tabriz University of Medical Sciences, Tabriz, Iran. 632 Women's Reproductive Health Research Center, Tabriz University of Medical Sciences, Tabriz, Iran. 633 Clinical Cancer Research Center, Milad General Hospital, Tehran, Iran. ⁶³⁴Department of Microbiology, Islamic Azad University, Tehran, Iran. ⁶³⁵Argentine Society of Medicine, Buenos Aires, Argentina. 636 Velez Sarsfield Hospital, Buenos Aires, Argentina. 637 Psychosocial Injuries Research Center, Ilam University of Medical Sciences, llam, Iran. ⁶³⁸Occupational Health Unit, Sant'Orsola Malpighi Hospital, Bologna, Italy. ³⁹Department of Economics, University of Göttingen, Göttingen, Germany, ⁶⁴⁰Foundation University Medical College, Foundation University Islamabad, Islamabad, Pakistan. ⁶⁴¹Department of Statistics, University of Washington, Seattle, WA, USA. ⁶⁴²Department of Biostatistics, University of Washington, Seattle, WA, USA. ⁶⁴³Department of Epidemiology and Biostatistics, Wuhan University, Wuhan, China. 644 Demographic Change and Aging Research Area, Federal Institute for Population Research, Wiesbaden, Germany. ⁶⁴⁵Competence Center of Mortality-Follow-Up of the German National Cohort, Federal Institute for Population Research, Wiesbaden, Germany. ⁶⁴⁶Department of Physical Therapy, Naresuan University, Phitsanulok, Thailand. ⁶⁴⁷School of Pharmacy, Aksum University, Aksum, Ethiopia. Department of Pharmacology and Toxicology, Mekelle University, Mekelle, Ethiopia. ⁶⁴⁹Department of Pharmacology, Addis Ababa University, Addis Ababa, Ethiopia. ⁶⁵⁰Department of Orthopaedics, Wenzhou Medical University, Wenzhou, China. ⁶⁵¹Psychology Department, University of Sheffield, Sheffield, UK, 652 Department of Diabetes and Metabolic Diseases, University of Tokyo, Tokyo, Japan. 653 School of International Development and Global Studies, University of Ottawa, Ottawa, Ontario, Canada. 654 The George Institute for Global Health, University of Oxford, Oxford, UK. 655 Health Services Management Research Center, Kerman University of Medical Sciences, Kerman, Iran. 656 Department of Health Management, Policy, and Economics, Kerman University of Medical Sciences, Kerman, Iran. ⁶⁵⁷School of Nursing, Hawassa University, Hawassa, Ethiopia. ⁶⁵⁸Pediatrics Department, University of Jos, Jos, Nigeria. 659 Department of Pediatrics, Jos University Teaching Hospital, Jos, Nigeria. 660 Centre for Suicide Research and Prevention, University of Hong Kong, Hong Kong, China. 661 Department of Social Work and Social Administration, University of Hong Kong, Hong Kong, China. 662 Department of Neuropsychopharmacology, National Center of Neurology and Psychiatry, Kodaira, Japan. 663 Department of Public Health, Juntendo University, Tokyo, Japan. ⁶⁶⁴Department of Health Policy and Management, Jackson State University, Jackson, MS, USA. 665 School of Medicine, Tsinghua University, Beijing, China. Department of Environmental Health, Mazandaran University of Medical Sciences, Sari, Iran. 667 Injury Prevention and Safety Promotion Research Center, Shahid Beheshti University of Medical Sciences, Tehran, Iran. ⁶⁶⁸Duke Global Health Institute, Duke University, Durham, NC, USA. 669 Social Determinants of Health Research Center, Ardabil University of Medical Science, Ardabil, Iran. ⁶⁷⁰The School of Clinical Sciences at Monash Health, Monash University, Melbourne, Victoria, Australia. 671 Student Research Committee, Babol University of Medical Sciences, Babol, Iran. $^{672} \mbox{Department}$ of Community Medicine, Ardabil University of Medical Science, Ardabil, Iran. ⁶⁷³Department of Health Education, Tarbiat Modares University, Tehran, Iran. ⁶⁷⁴College of Medicine and Health Sciences, Dilla University, Dilla, Ethiopia. ⁶⁷⁵Public Health Department, University of Edinburgh, Edinburgh, UK. ⁶⁷⁶School of Public Health, Wuhan University of Science and Technology, Wuhan, China. 677 Hubei Province Key Laboratory of Occupational Hazard Identification and Control, Wuhan University of Science and Technology, Wuhan, China. ⁶⁷⁸School of Medicine, Wuhan University, Wuhan, China. ⁶⁷⁹School of Biology and Pharmaceutical Engineering, Wuhan Polytechnic University, Wuhan, China. ⁶⁸⁰School of Health Sciences, Wuhan University, Wuhan, China. ⁶⁸¹National Center for Chronic and Noncommunicable Disease Control and Prevention, Chinese Center for Disease Control and Prevention, Beijing, China. 682 These authors jointly supervised this work: Simon I. Hay, Stephen S. Lim, Jonathan F. Mosser. [⊠]e-mail: sihay@uw.edu; stevelim@uw.edu; jmosser@uw.edu

Methods

Data reporting

As this is a modelling study, no statistical methods were used to predetermine sample size, the experiments were not randomized and the investigators were not blinded to allocation during experiments and outcome assessment.

Overview

Building from our previous study of diphtheria–tetanus–pertussis vaccination coverage in Africa 14 , we fitted a geostatistical model with correlated errors across space and time to predict $5 \times 5 \cdot \text{km}^2$ level estimates of MCV1 coverage from 2000 to 2019 using a suite of geospatial and national-level covariates for $101 \, \text{LMICs}$. This overall process has been summarized in Extended Data Fig. 1. We spatially aggregated estimates using population-weighted averages to second administrative units from a modified version of the Database of Global Administrative Units (GADM), referred to as districts, and performed post hoc analyses to assess geographical inequality to examine progress towards GVAP targets, absolute geographical inequality and vaccination status as a function of geographical remoteness 5 . This study is compliant with the Guidelines for Accurate and Transparent Health Estimates Reporting (GATHER) recommendations 51 (Supplementary Table 1).

We defined routine MCV1 coverage as evidence of receipt of at least one dose of a MCV from either a home-based record (HBR) or parental recall among the target population in concordance with country-specific vaccination schedules in 2019^{52} . Despite our best efforts to remove doses delivered through supplemental immunizion activities (SIAs) (Supplementary Information section 1.3.4), there is likely to be residual misclassification of some SIA doses due to the limitations of the available data, and these estimates of routine coverage should be viewed in the context of this limitation.

Countries were selected for this analysis if they were a LMIC or were a 'Decade of Vaccine' priority country with available subnational survey data on MCVI coverage between 2000 and 2019⁵³. We defined LMICs based on the socio-demographic index, a metric combining education, fertility and income to summarize development, as determined by GBD 2019⁵⁴. For 13 countries (Bhutan, Brazil, China, Dominica, Georgia, Grenada, Libya, Oman, Palestine, Saint Lucia, Saint Vincent and the Grenadines, Seychelles and Venezuela), no available subnational vaccine coverage data met the inclusion and exclusion criteria; these countries were therefore excluded from this analysis. A full list of included countries is provided in Supplementary Table 3. Countries were assigned to one of 13 continuous geographical modelling regions. These regions were adapted from regions defined by GBD 2019, which are constructed to group countries together by epidemiological similarity and geographical proximity (Extended Data Fig. 2).

Data

Using the Global Health Data Exchange (GHDx)⁵⁵, we identified and compiled a total of 354 population-based household surveys from 101 LMICs from 2000 to 2019 containing individual MCV1 vaccination status and subnational geolocation information. Surveys were included if they contained MCV1 coverage information and subnational geolocation, and excluded if they contained areal data and were missing key survey design variables (strata, primary sampling units and design weights), did not include children aged 12–59 months, contained no subnational individual-level geographical information or if coverage estimates were implausible (Supplementary Tables 4, 5).

Coverage was computed at the cluster level when global positioning system (GPS) data were available. If GPS information was not collected or was not available, we calculated mean coverage at the most-granular geographical area possible while accounting for sampling weights and survey design. These aggregated coverage estimates were then included in the geospatial modelling process using a

previously described method^{14,56} that leverages population weights and a *k*-means clustering algorithm to propose a set of GPS coordinates as a proxy for locations where survey data collection could have occurred (Supplementary Fig. 3). These coordinates were then used to represent the areal data in the geospatial model. The following data were extracted from each survey source: vaccine card or HBR doses, parental recall vaccine doses, age (in months), survey weight and design variables, and GPS cluster or areal location. Individuals with evidence of vaccination either from HBR or recall were considered to have been vaccinated. Individuals were excluded from the analysis if they were missing age, spatial or survey design information or were outside of the study age or year range. The study included all years between 2000 and 2019. A comprehensive overview of data from all study geographies included can be found in Supplementary Figs. 1, 2.

Individual age, in months, at the time of survey collection was used to assign each child to a birth cohort (12-23 months, 24-35 months, 36-47 months and 48-59 months). Data corresponding to each birth cohort were included in the modelling process in the year in which that birth cohort was aged 0-12 months old. For countries recommending MCV1 within the first year of life, we included data from children aged 12-47 months. If the first dose was not recommended until the second year of life, we included data from children aged 24-59 months. A full list of schedules by country can be found in Supplementary Table 2. This yielded a dataset of 1,697,570 total children. This method allows the inclusion of additional individuals, which increases overall geographical representation but requires assumptions such as negligible catch-up vaccination and no differential mortality or migration. However, the overall influence of including older cohorts in our model on the key findings appeared to be minor (Supplementary Information section 2.2). Additional information on the benefits and limitations of this approach can be found in the Supplementary Information sections 1.3.4, 2.2, Supplementary Figs. 18-20 and Supplementary Tables 14, 15.

We included 26 geospatial covariates as possible predictors of MCV1 coverage in the modelling process, including maternal education, access to major cities or settlements, a binary urban or rural indicator, total population and a suite of 22 environmental covariates (Supplementary Fig. 4 and Supplementary Table 6). Four national-level covariates were also included: lag-distributed income, prevalence of the completion of the fourth antenatal care visit among pregnant women, mortality due to war and terror, and bias-adjusted national-level administrative data on MCV1 coverage reported through the WHO/UNICEF Joint Reporting Form (Supplementary Information section 1.5.1). For each region, an optimized set of geospatial covariates was selected from these 26 possible covariates, using a variance inflation factor (VIF) algorithm⁵⁷ in which covariates were selected with a VIF < 3. This method was used to ensure non-collinearity between covariates within each region to facilitate model convergence. Selected covariates varied by region (Supplementary Table 7).

Other spatial data used in our analyses included gridded population estimates, administrative boundaries and gridded estimates of travel time to major cities or settlements. These sources are described in detail in Supplementary Information section 1.3.

Geostatistical model

First, stacked generalization was used to capture potential nonlinear and complex relationships between covariates and vaccination coverage. This approach has previously been shown to increase the predictive accuracy of geospatial models⁵⁸. Using the optimized set of covariates selected for each region by the VIF algorithm, three different child models—generalized additive models, lasso regression and boosted regression trees—were fit, with each model predicting MCVI coverage as the outcome of interest. When fitting boosted regression trees, country-level fixed effects were included to allow relationships between coverage and covariates to differ by country. In this initial modelling step, there were no explicitly defined temporal or spatial

effects included in the models beyond those inherently present in the covariate patterns and correlations between covariates.

Each child model was fit using fivefold cross-validation to avoid overfitting. This generated out-of-sample predictions of coverage for each location and year per region. Each model in each region was also fit using the full set of vaccine coverage outcome data, which yielded a corresponding set of in-sample predictions. The predictions of MCV1 coverage obtained from each child model were in turn used as predictors in the second-step geostatistical model described below. Out-of-sample predictions from each child model were used as explanatory covariates when fitting the geostatistical model. In-sample predictions from each model and region were used when generating predictions from the fitted geostatistical model.

After the first step (stacked generalization), a second-step Bayesian geostatistical modelling framework was used to model vaccination coverage as counts in a binomial space with a logit link through a generalized linear regression with explicit spatial and temporal terms. This second step leverages the covariate relationships estimated through stacked generalization while also accounting for additional correlation in coverage across space and time.

A separate model of MCV1 coverage was fit for each of the 13 regions as defined below:

$$C_d|p_{i(d),t(d)}, N_d \sim \text{Binomial}(p_{i(d),t(d)}, N_d) \forall \text{ observed clusters } d$$

$$logit(p_{i,t}) = \beta_0 + X_{i,t}\beta + Z_{i,t} + \epsilon_{country_{(i)}} + \epsilon_{i,t} \forall i \in spatial domain \forall$$

$$t \in time domain$$

$$\sum_{h=1}^{3} \beta_h = 1$$

$$\epsilon_{i,t} \sim N(0, \sigma_{\mathrm{nugget}}^2)$$

$$\epsilon_{\text{country}(i)} \sim N(0, \sigma_{\text{country}(i)}^2)$$

$$Z \sim \mathsf{GP}(0, \Sigma^{\mathsf{space}} \otimes \Sigma^{\mathsf{time}})$$

$$\Sigma^{space} = \frac{2^{1-\nu}\sigma_{space}^2}{\Gamma(\nu)} \times \left(\frac{\sqrt{8}}{\delta}D\right)^{\nu} \times K_{\nu}\left(\frac{\sqrt{8}}{\delta}D\right)$$

$$\Sigma_{j,k}^{\text{time}} = \rho^{|k-j|}$$

This model, adopted from widely used Bayesian hierarchical models 89,60 , has been described in detail in other work 14,25,56,61,62 . In brief, this method estimates the number of children, C, in cluster d at location i and time t with sample size N that have been vaccinated with a specific antigen-dose combination. $p_{i(d),t(d)}$ is the proportion of children vaccinated with MCV1 among the target age population in cluster d. Each child model generates a prediction $X_{i,t}$ for each child model h. Residual terms ϵ_* are unique to each particular location in space and time across all modelled geographies and years.

In this generalized linear regression framework, the proportion of children vaccinated $p_{i,t}$ is modelled using the out-of-sample predictions of vaccine coverage $X_{i,t}$ from each of three stacked generalization child models (h) as explanatory variables. The β_h coefficients are constrained to sum to 1, via the 'extraconstr' R-INLA parameter ⁶³, to improve computational tractability ⁵⁸ and are representative of the predictive weighting used in the stacking process.

 $\epsilon_{\mathrm{country}(i)}$ represents a country-level random effect. $\epsilon_{i,t}$ represents an independent nugget effect for irreducible error for a given observation, which accounts for true variation that is unable to be captured by the

model and variation from measurement error. $Z_{i,t}$ represents a correlated spatiotemporal error term, for any residual autocorrelation across space and time that remains after accounting for the predictive capacity of the stacked-modelled covariates, country-specific variation in vaccine coverage and observation-specific irreducible error.

These additional spatiotemporal residuals $Z_{i,t}$ were modelled as a three-dimensional spatiotemporal Gaussian process with a mean of zero and a covariance matrix formed from the Kronecker product of spatial and temporal covariance kernels. The temporal covariance Σ^{time} was modelled via an annual autoregressive order 1 function from all study years from 2000 to 2019, where ρ is the autocorrelation function and k and j are points in the annual time series. The spatial covariance Σ^{space} was assumed to be an isotropic, stationary Matérn function, where Γ is the gamma function, K_{ν} is the modified Bessel function of the second kind of order v > 0, σ_{space}^2 is the marginal variance, v is a scaling constant, δ is a range parameter with a penalized complexity prior, and D is a spatial distance matrix⁶⁴, measured along the great circle in kilometres. The generalized linear model was fitted using an integrated nested Laplace approximation in R-INLA with a stochastic partial differential equation (SPDE) solver in package SPDE⁶⁵. Additional detailed information on priors, spatial mesh construction and model fitting is provided in Supplementary Information sections 1.4.2-1.4.6 and Supplementary Fig. 5. This process produces a set of 1,000 posterior draws, each representing an estimate of vaccine coverage for each location and year - in other words, a set of 1000 candidate maps of coverage from 2000 to 2019.

Post-estimation

To leverage data from additional national-level sources, including administrative data, and maintain internal consistency, the set of candidate maps was calibrated to MCV1 coverage estimates produced for GBD 2019. This post hoc calibration preserves the overall spatial variation of estimates, while ensuring that the population-weighted averages of the geospatial estimates are equivalent to those produced by GBD⁴. This step allows for the calibrated estimates to reflect information from data sources that are only available at the national level, such as surveys for which no subnational data are available, which are included in the GBD estimates but excluded from the geospatial model described in the 'Geostatistical model' section. A description of the estimation of MCV1 coverage for GBD 2019 can be found in Supplementary Information section 1.5.1.

In this calibration process, each 5×5 -km² pixel in each modelled region was first assigned to a second-level administrative unit. In locations in which boundary definitions transect a given pixel, the fraction of area of that pixel belonging to each overlapping second-level administrative unit was calculated. Because of the nested hierarchy of administrative units, this additionally allowed for the assignment of pixels and partial pixels to first administrative units and countries. Assuming that the population density within each pixel was uniform, WorldPop population values of children under 5 years old were divided for each whole or partial pixel proportional to fractional area. After pixel and partial pixel populations were assigned, population-level estimates were calibrated to GBD population estimates for each country and year.

Calibration methods similar to those used in this study have been described previously¹⁴. To ensure vaccination coverage estimates post-calibration remained between 0 and 100%, calibration was performed in logit space such that for each country c and year t, national-level estimates of coverage from GBD ($V_{\text{GBD},c,t}$) and population-weighted national averages of coverage from the model-based geostatistical (MBG) model ($V_{\text{MBG},c,t}$) can be related via a country-year-specific calibration factor ($k_{c,t}$) in the following equation:

$$logit(V_{GBD,c,t}) = logit(V_{MBG,c,t}) + k_{c,t}$$

Calibration factors were applied to each 5×5 -km² pixel and partial pixel per draw per country-year. Pixels that were fractionally assigned to multiple countries were combined using a weighted average proportional to the fraction of each area. This process resulted in a set of calibrated draw-level estimates of vaccination coverage, which were used for all subsequent analyses.

Population-weighted averages of coverage for each pixel or partial pixel within a first or second administrative unit were then calculated. Fractional pixel membership was determined as described above. This process was repeated for each of the 1,000 posterior pixel-level draws, which yielded 1,000 posterior draws of MCV1 coverage per administrative unit per year. Estimates for first and second administrative units with uncertainty were derived from mean, 2.5th and 97.5th percentiles.

Model validation

We assessed the predictive performance of the models using fivefold out-of-sample cross-validation. We stratified data by first and second administrative units and ran models leaving out one-fifth of the spatially stratified data at a time. Predicted estimates of MCV1 coverage were then compared to the withheld observed data by calculating the mean error, root mean square error, correlation and other predictive validity metrics for all years for which survey data were available (2000–2018). Fitted model parameters can be found in Supplementary Tables 9–12 and Supplementary Figs. 6–13, respectively. Additional information regarding uncertainty of estimates can be found in Supplementary Figs. 14–17.

Post hoc geospatial inequality analyses

Lorenz curves were generated using the relationship between the number of children and the number of vaccinated children for each pixel. Pixel-level Gini coefficients were calculated for 2000 and 2019 from corresponding Lorenz curves^{66,67} (Supplementary Table 13). Absolute geographical inequality per country was calculated from the national-level Gini coefficients and national MCV1 coverage using the following formula:

Absolute geographical inequality = 2 × coverage × Gini

We chose to use the absolute geographical inequality metric to represent inequality over the Gini coefficient alone. As the mean is related to Gini, we wanted to account for this relationship. Estimates are scaled by 2 as this puts the absolute geographical inequality coefficient back to the same scale as the mean 68 .

Additionally, we assessed vaccination status as a function of geographical remoteness. Using a gridded surface of travel time to major cities or settlements, we classified each 5×5 -km² pixel as remote rural, urban or neither²9. Pixels with travel times of less than 30 min were classified as urban, and pixels with travel times greater than 3 h were classified as remote rural. Overlaid with a gridded population surface from WorldPop³0, the number of unvaccinated children per pixel was also calculated.

We constructed concentration curves of the cumulative proportion of unvaccinated children as well as plots of MCV1 coverage by travel time to assess patterns across countries and regions. Country-specific concentration curves of the cumulative proportion of unvaccinated children as a function of travel time for select countries are shown in Extended Data Fig. 10. Summary metrics, such as the proportion of unvaccinated individuals living in each urban and remote rural location, were computed.

Limitations

This work is subject to several limitations. First, the primary data used in this analysis came from child-level survey data with varying degrees of representativeness, consistency, accuracy and comparability, from

both HBR and parental recall^{69,70}. The magnitude and direction of recall bias varies, and we therefore were unable to correct for it⁷¹. We estimate coverage using data from children aged 12–59 months, and while we accounted for target age at vaccination, this does not fully account for differential mortality due to vaccine status or catch-up vaccination. We aim to estimate routine coverage and have excluded doses delivered via SIAs from the analysed survey data wherever possible (Supplementary Information section 1.3.4), but misclassification of SIA doses is still likely, particularly in cases of parental recall—especially for older children—and in cases in which survey methodology does not distinguish clearly between SIA and routine doses.

In data-sparse areas for which covariate relationships may not fully capture coverage patterns, results may be biased. Additionally, data representativeness among vulnerable populations, such as those living in urban slums or migrant populations, might vary due to data collection in survey design. We include as much data on MCV1 coverage as possible, including data that are only geo-resolved to areal locations. The methodology that we used to assign areal data to specific locations for modelling could lead to oversmoothing in final estimates, obscure relationships between coverage and covariates, and underestimate uncertainty, but this method has been shown to have a higher predictive validity compared with the exclusion of the data⁷². Limitations due to data availability should not be taken lightly and should reinforce to stakeholders and policymakers the need for additional resources to collect high-quality data that are representative of all populations, especially the most vulnerable for being unvaccinated, and to increase the quality of routinely collected subnational administrative data.

Because the estimates that we used to assess geographical remoteness in post hoc analyses were also used as spatial covariates in the geospatial model, these results are limited by the possibility of circularity and subsequent confounding. In addition, we used a stacked generalization method to allow for complex and nonlinear relationships between covariates and vaccination coverage. These methods are optimized for prediction, not causal inference. For that reason, these results cannot be used to identify the specific effect of any particular covariate on MCV1 coverage. In addition, owing to limitations in the underlying data and computational feasibility, we were unable to incorporate several potentially important sources of uncertainty into this analysis, including from covariates, population estimates, the incorporation of areal data and the stacked generalization process.

We fitted our geostatistical models using R-INLA, as opposed to a full Markov chain Monte Carlo sampler. Although using a more traditional Bayesian model fitting approach that takes true samples from the posterior typically results in increased parameter identifiability. the Laplace approximation approach used by R-INLA is more computationally feasible given our current modelling scale. Our model is separable, yet symmetric, across time and space. This approach assumes that, for each region, the covariance has the same functional form between years and locations regardless of the locations themselves; the use of a non-separable covariance function could relax these assumptions^{73,74}. However, owing to the additional computational challenges associated with fitting a non-separable model, as well as data sparsity in several regions throughout space and time, we determined that fitting a non-separable model would be challenging and complex, and would probably yield little benefit compared to our current modelling approach.

In some settings with high levels of natural immunity (derived from previous infection), greater than 95% vaccination coverage may not be required to prevent disease transmission⁷⁵. These estimates only focus on the first routine dose of MCV, and immunity can also be obtained through later vaccination via SIA or natural infection. In an ideal long-term measles elimination scenario, all immunity would be vaccine-derived, and no natural infections would occur. A 95% coverage target for routine immunization, therefore, still has practical programmatic relevance.

Finally, our study describes spatial heterogeneity in coverage, but not pockets of low coverage within social or age groupings that can facilitate ongoing disease transmission, particularly in densely populated areas, despite nominally high average vaccine coverage 76 . Although these results provide a powerful tool for policymakers to identify weaknesses in routine immunization systems and plan for SIA, they should be used in conjunction with other data sources that can be used to make decisions about vaccine policy, including analyses of cost effectiveness, determinants of high or low coverage, and specific coverage initiatives to reduce disease burden.

Reporting summary

Further information on research design is available in the Nature Research Reporting Summary linked to this paper.

Data availability

The findings of this study are supported by data available in public online repositories and data publicly available upon request from the data provider. A detailed table of data sources and availability can be found in Supplementary Table 4 and at http://ghdx.healthdata.org/lbd-publication-data-input-sources. Administrative boundaries were modified from the Database for Global Administrative Areas (GADM) dataset⁷⁷. Populations were retrieved from WorldPop³⁰, and gridded estimates of travel time to the nearest city or settlement were available online from a previously published study²⁹. This study complies with the GATHER recommendations⁵¹.

Code availability

This study is compliant with the GATHER recommendations 51 ; as such, all computer code is available from GitHub (https://github.com/ihmeuw/lbd/tree/mcv1-lmic-2020). All maps and figures presented in this study are generated by the authors using RStudio (R version 3.6.1), ArcGIS Desktop 10.6 and Python 2.7.

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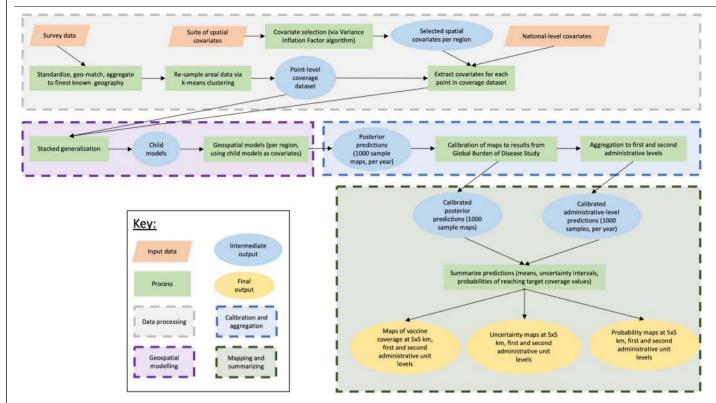
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Additional information

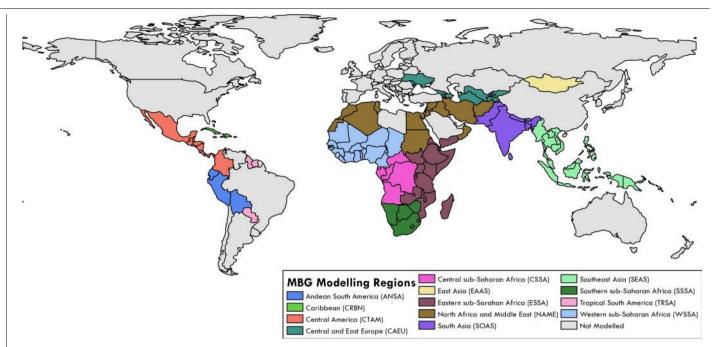
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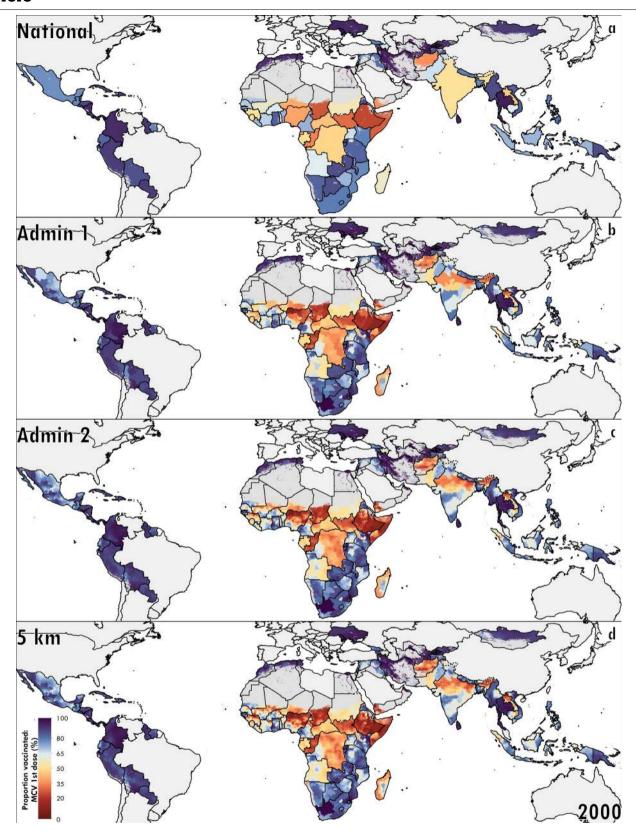
Extended Data Fig. 1 | **Data processing and geospatial modelling flowchart.** Survey data and the suite of covariates used in modelling are first compiled and processed (orange and grey). The modelling process (purple) consists of data being used in a stacked generalization ensemble modelling process via boosted regression tree, lasso and generalized additive models, fitting the second-stage spatiotemporal model using integrated nested Laplace

approximation, and finally calibration to GBD estimates (blue). Steps in dark green and outputs in yellow indicate the post-estimation process when the full posterior distribution of predications is transformed to both 5×5 -km² and first and second administrative-unit-level maps and their various final results. Intermediate outputs throughout the process are shown in blue and overall processes are shown in light green.



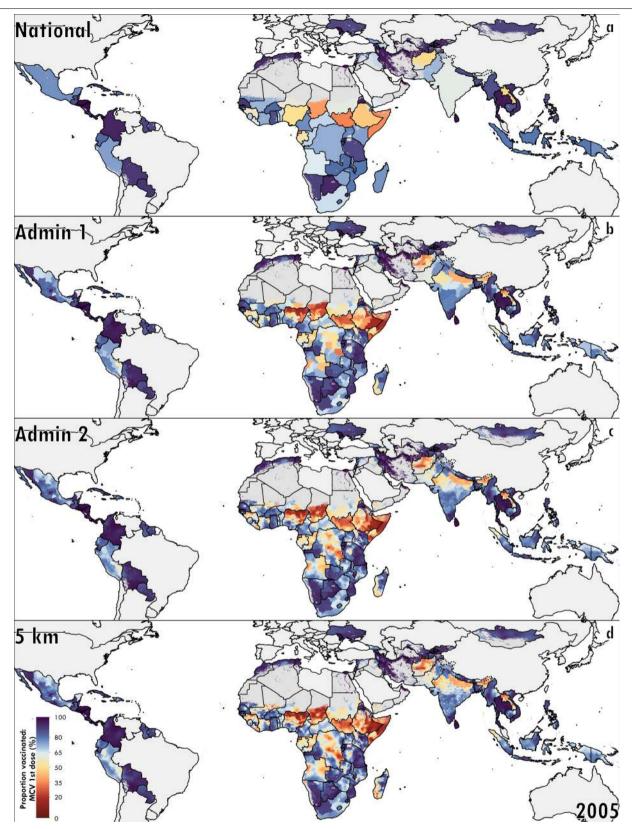
Extended Data Fig. 2 | **Regions of countries used in modelling.** Analyses were divided into 13 regions based on the GBD super-regions to allow for locations similar in data availability and patterns of vaccine coverage to be

analysed using similar covariate and modelling relationships. Each colour represents a different region, as described in the legend.



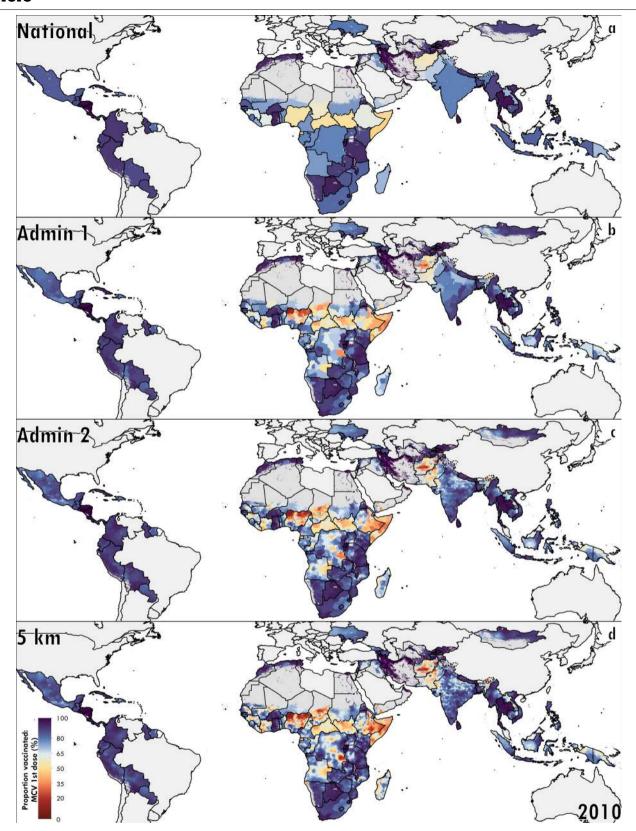
Extended Data Fig. 3 | National, first- and second-administrative-unit level, and pixel-level MCV1 coverage, 2000. \mathbf{a} – \mathbf{d} , Posterior means are represented at the national (\mathbf{a}), first-administrative-unit (\mathbf{b}), second-administrative-unit (\mathbf{c})

 $and \ 5\times 5\cdot km^2 \ pixel \ (\textbf{d}) \ levels. \ Pixels \ that \ are \ grey \ in \ colour \ are \ either \ not \ included \ in \ the \ analysis, \ or \ have \ been \ classified \ as \ being \ 'barren \ or \ sparsely \ vegetated' \ or \ had \ fewer \ than \ 10 \ people \ per \ 1\times 1\cdot km^2 \ pixel^{30,50}.$



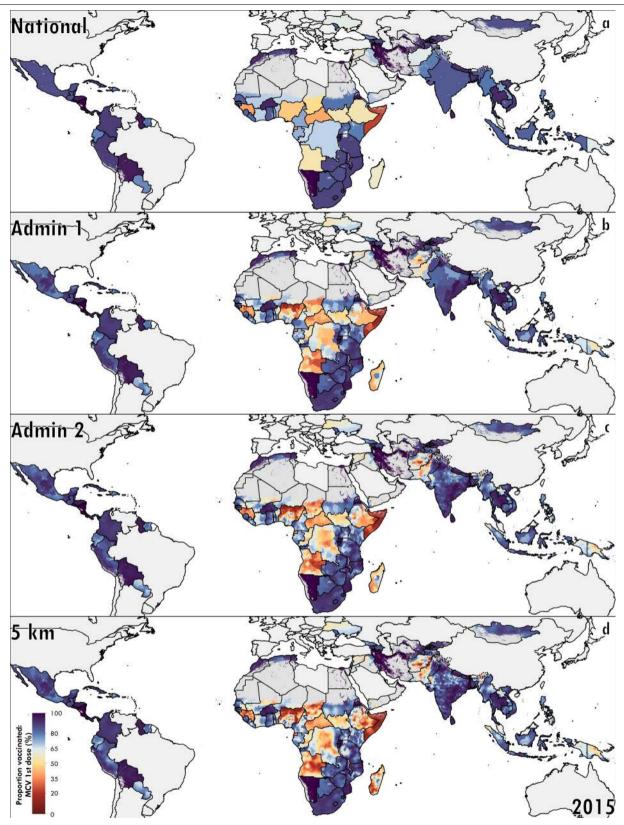
Extended Data Fig. 4 | National, first- and second-administrative level, and pixel-level MCV1 coverage, 2005. a-d, Posterior means are represented at the national (a), first-administrative-unit (b), second-administrative-unit (c) and

 $5\times5\cdot km^2pixel~\textbf{(d)}~levels.~Pixels~that~are~grey~in~colour~are~either~not~included~in~the~analysis,~or~have~been~classified~as~being~barren~or~sparsely~vegetated~or~had~fewer~than~10~people~per~1\times1\cdot km^2~pixel~^{30,50}.$



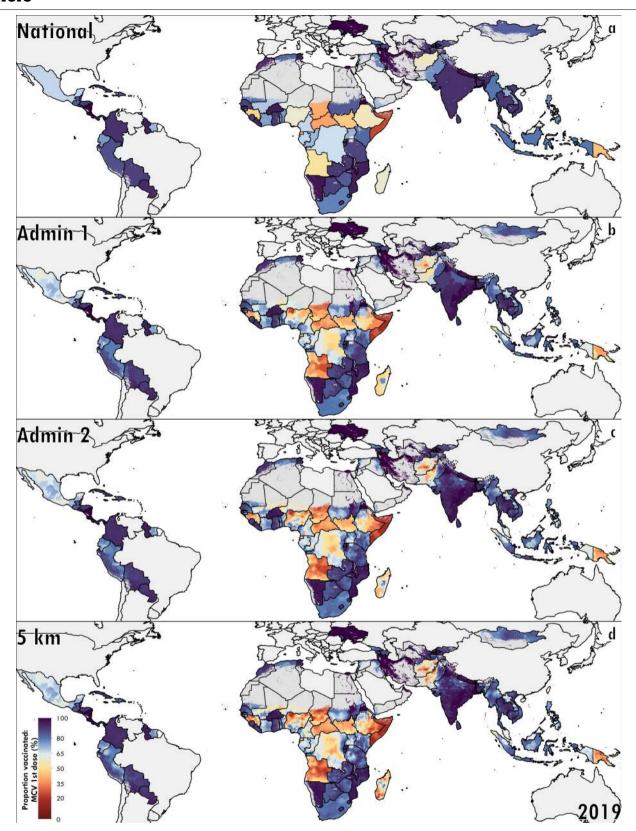
Extended Data Fig. 5 | National, first- and second-administrative-unit level, and pixel-level MCV1 coverage, 2010. a-d, Posterior means are represented at the national (a), first-administrative-unit (b), second-administrative-unit (c)

 $and \ 5\times 5\cdot km^2 \ pixel \ (\textbf{d}) \ levels. \ Pixels \ that \ are \ grey \ in \ colour \ are \ either \ not \ included \ in \ the \ analysis, \ or \ have \ been \ classified \ as \ being \ 'barren \ or \ sparsely \ vegetated' \ or \ had \ fewer \ than \ 10 \ people \ per \ 1\times 1\cdot km^2 \ pixel^{30,50}.$



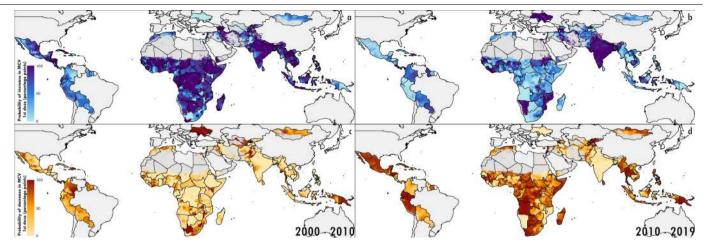
Extended Data Fig. 6 | National, first- and second-administrative-unit level, and pixel-level MCV1 coverage, 2015. a-d, Posterior means are represented at the national (a), first-administrative-unit (b), second-administrative-unit (c)

 $and \ 5\times 5\cdot km^2 \ pixel \ (\textbf{d}) \ levels. \ Pixels \ that \ are \ grey \ in \ colour \ are \ either \ not \ included \ in \ the \ analysis, \ or \ have \ been \ classified \ as \ being \ 'barren \ or \ sparsely \ vegetated' \ or \ had \ fewer \ than \ 10 \ people \ per \ 1\times 1\cdot km^2 \ pixel^{30,50}.$

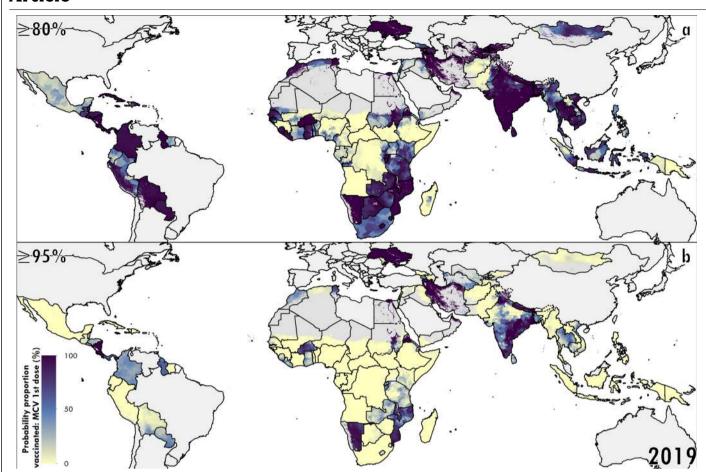


Extended Data Fig. 7 | National, first- and second-administrative-unit level, and pixel-level MCV1 coverage, 2019. a-d, Posterior means are represented at the national (a), first-administrative-unit (b), second-administrative-unit (c)

 $and \ 5\times 5\cdot km^2 \ pixel \ (\textbf{d}) \ levels. \ Pixels \ that \ are \ grey \ in \ colour \ are \ either \ not \ included \ in \ the \ analysis, \ or \ have \ been \ classified \ as \ being \ 'barren \ or \ sparsely \ vegetated' \ or \ had \ fewer \ than 10 \ people \ per \ 1\times 1\cdot km^2 \ pixel^{30,50}.$

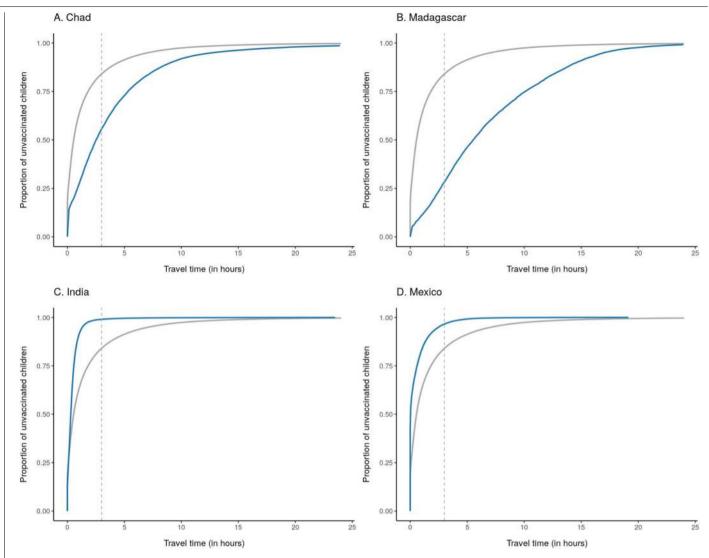


Extended Data Fig. 8 | Probability of increased or decreased coverage from 2000 to 2010 and 2010 to 2019. a-d, Probability of an increase in coverage in each district (a, b) and probability of decrease in coverage in each district (c, d) from 2000 to 2010 (a, c) and 2010 to 2019 (b, d).



Extended Data Fig. 9 | Estimated district-level probabilities of reaching MCV1 coverage targets in 2019. a, b, Probability of districts having achieved 80% GVAP and Measles Rubella Initiative targets (a) and 95% critical proportion to immunize coverage targets to reach herd immunity (b). Countries excluded

from the analysis and pixels classified as 'barren or sparsely vegetated' based on ESA-CCI satellite data or with fewer than 10 people per 1×1 -km² pixel based on WorldPop estimates are masked in grey 30,50 .



 $\textbf{Extended Data Fig. 10} \ | \ \textbf{Country examples of concentration curves.} \ Concentration \ curves of the cumulative proportion of unvaccinated children as a function of travel time (in hours) in Chad (a), Madagascar (b), India (c) and Mexico (d). Curves for both the indicated countries (blue) and all LMICs (grey) are shown.$



Corresponding author(s):	Jonathan F. Mosser
Last updated by author(s):	Dec 3, 2020

Reporting Summary

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Statistics			
For all statistical analyses, confirm that the following items are present in the figure legend, table legend, main text, or Methods section.			
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\square The exact sample size (n) for each experimental group/condition, given as a discrete number and unit of measurement			
A statement on whether measurements were taken from distinct samples or whether the same sample was measured repeatedly			
The statistical test(s) used AND whether they are one- or two-sided Only common tests should be described solely by name; describe more complex techniques in the Methods section.			
A description of all covariates tested			
A description of any assumptions or corrections, such as tests of normality and adjustment for multiple comparisons			
A full description of the statistical parameters including central tendency (e.g. means) or other basic estimates (e.g. regression coefficient) AND variation (e.g. standard deviation) or associated estimates of uncertainty (e.g. confidence intervals)			
For null hypothesis testing, the test statistic (e.g. <i>F</i> , <i>t</i> , <i>r</i>) with confidence intervals, effect sizes, degrees of freedom and <i>P</i> value noted Give <i>P</i> values as exact values whenever suitable.			
For Bayesian analysis, information on the choice of priors and Markov chain Monte Carlo settings			
For hierarchical and complex designs, identification of the appropriate level for tests and full reporting of outcomes			
\square Estimates of effect sizes (e.g. Cohen's d , Pearson's r), indicating how they were calculated			
Our web collection on <u>statistics for biologists</u> contains articles on many of the points above.			
Software and code			
Policy information about <u>availability of computer code</u>			

Data collection No primary data collection was carried out for this analysis.

Data analysis

This analysis was carried out using R version 3.6.1 and using R-INLA v.20.01.29.9000. Maps were produced using ArcGIS Desktop 10.6 and Python 2.7. All code used for these analyses will be made publicly available upon publication.

For manuscripts utilizing custom algorithms or software that are central to the research but not yet described in published literature, software must be made available to editors/reviewers. We strongly encourage code deposition in a community repository (e.g. GitHub). See the Nature Research guidelines for submitting code & software for further information.

Data

Policy information about availability of data

All manuscripts must include a data availability statement. This statement should provide the following information, where applicable:

- Accession codes, unique identifiers, or web links for publicly available datasets
- A list of figures that have associated raw data
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The findings of this study are supported by data available in public online repositories and data publicly available upon request of the data provider. A detailed table of data sources and availability can be found in Supplementary Table 4 and http://ghdx.healthdata.org/lbd-publication-data-input-sources. Administrative boundaries were modified from the Database for Global Administrative Areas (GADM) dataset. Populations were retrieved from WorldPop and gridded estimates of travel time to nearest city or settlement were available online from work by Weiss, et al 2018. This study complies with the Guidelines for Accurate and Transparent Health Estimates Reporting (GATHER) recommendations. All maps and figures presented in this study are generated by the authors; no permissions are required for publication.

Field-specific reporting					
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Life sciences study design					
All studies must dis	sclose on these points even when t	he disclosure is negative.			
Sample size	This observational study incorporated all available survey data sources that met the inclusion criteria as described in detail in the manuscript and supplementary information. The combined dataset from 354 household based surveys contained information on vaccination status from 1.70 million individual children.				
Data exclusions	Surveys were excluded due to unrealistic national or geographic trends compared to other surveys in nearby country-years, inability to match the microdata to geographic locations, or non-standard methodology. These criteria were pre-established prior to reviewing the data. A full list of excluded surveys is included in Supplementary Table 5.				
Replication	All code and data are available publicly for reproducibility.				
Randomization	As this work is an observational mapping study, there were no experimental groups.				
Blinding	As this work is an observational mapping study, there was no need for blinding.				
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Antibodies		ChIP-seq			
Eukaryotic cell lines Palaeontology		Flow cytometry MRI-based neuroimaging			
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Human resea

Animals and other organisms

Human research participants