## RESEARCH REPORT

# Marital status and longevity in the United States population 

Robert M Kaplan, Richard G Kronick

Purpose: To investigate the relation between marital status and survival.
Data sources: The US 1989 national health interview survey (NHIS) merged with the 1997 US national death index.

See end of article for authors' affiliations

Correspondence to: Dr R M Kaplan, University of California, Los Angeles, PO Box 951 1772, Room 31-293-C, CHS UCLA Los Angeles, 90025, USA; rmkaplan@ucla.edu

Accepted for publication 13 February 2006

Results: Among 1989 NHIS respondents, 5876 (8.77\%) died before 1997 and 61123 (91.23\%) were known to be alive. Controlling for demographic and socioeconomic characteristics, the death rate for people who were unmarried was significantly higher than it was for those who were married and living with their spouses. Although the effect was significant for all categories of unmarried, it was strongest for those who had never married. The never married effect was seen for both sexes, and was significantly stronger for men than for women. For the youngest age group (19-44), the predominant causes of early death among adults who had never married were infectious disease (presumably HIV) and external causes. In the middle aged and older men and women, the predominant causes were cardiovascular and other chronic diseases.
Conclusion: Current marriage is associated with longer survival. Among the not married categories, having never been married was the strongest predictor of premature mortality. It is difficult to assess the causal effect of marital status from these observational data.

More than 25 years ago cardiologist James Lynch published The Broken Heart: The Medical Consequences of Loneliness. ${ }^{1}$ Lynch argued that premature death from heart disease was more common among people who lived alone or were never married. A variety of studies have suggested that supportive social networks promote longer life expectancy. ${ }^{2-13}$ The relation between social connectedness and survival has been shown in a variety of different populations including US residents of Alameda County, California, ${ }^{13}$ Tecumseh Mighigan, ${ }^{8}$ and Evans County Georgia. ${ }^{14}$ International studies include Japanese-American men, ${ }^{9}$ Japanese men living in Hawaii, ${ }^{15}$ Israeli civil service, ${ }^{10}$ and adults living alone in Gothenberg, Sweden. ${ }^{16}$ In this paper we concentrate on categories of being unmarried. In particular, we focus on those who have never married. Adults with previous marriages are likely to have offspring that stimulate continuing family contact. In contrast, those who have never married might be more socially isolated than those who have been widowed or divorced. Thus, we predict that the health consequences of being unmarried will be strongest for the never married adults.
The studies reported in the literature have several limitations. In most studies, the number of deaths is too few to evaluate the relation between social connectedness and specific causes of death. Furthermore, most studies are too small to evaluate the relation between social connectedness and mortality at different points along the age range. The purpose of this study is to evaluate the relation between marital status and longevity using data from the US national health interview survey and the national death index.

## METHOD

## Data sources

## National health interview survey (NHIS)

Data were from the 1989 NIHS. The survey uses a probability sample of American households to estimate health status for the US civilian non-institutionalised population. The 1989 survey included over 480000 households and 115159 valid person estimates. The questionnaire also asks about acute and chronic conditions that were responsible for disabilities,
limitations, or health care visits. The NHIS includes a measure of self perceived health status using the categories, excellent, very good, good, fair, and poor. The full questionnaire, sampling weights, and summary statistics can be obtained at http://www.cdc.gov/nchs/data/series/sr_10/ srl0_176.pdf. ${ }^{17} 18$

## National death index (NDI)

The NDI captures death records on the basis of state vital statistics reports. Our analysis uses a data file created by NCHS in which NCHS linked the 1989 NHIS to the 1997 NDI using social security numbers. Although the 1989 NHIS included over 115000 respondents, social security numbers were not collected for respondents 18 years old and younger, and mortality information is not available for respondents 18 and younger, as well as for small numbers of respondents 19 and over with missing social security numbers.

## Statistical analysis

Weighted multiple logistic regressions were used to estimate the association between marital status and mortality. We estimated a total of 21 regressions: three variations for each of seven outcome variables, where the outcome variables were all cause mortality as well as mortality from each of the six causes-infectious disease ( $2.8 \%$ of all deaths), cancer ( $27.2 \%$ ), cardiovascular disease ( $41.4 \%$ ), pulmonary disease ( $7.8 \%$ ), external causes (accidents, homicides, and suicides, $5.6 \%)$, and all "other" causes ( $15.0 \%$ ). In the first regression of each set, we included an indicator variable for being "never married", to permit straightforward comparison of the effect of being ever married with the effect of being married, divorced/separated, or widowed. In the second regression we included interactions of never married with age and sex, allowing the effect of being never married to vary by age and sex. In the third regression of each set we included interactions of never married with self reported health status. All regressions also include variables

Abbreviations: NHIS, national health interview survey; NDI, national death index

Table 1 Likelihood of death by 1997 based on NHIS responses in 1989 using logit analysis*

| Parameter | p Value | Odds ratio estimate | 95\% CI |
| :---: | :---: | :---: | :---: |
| Intercept | <0.0001 |  |  |
| Age (19-24) |  |  |  |
| 25-34 | 0.1020 | 1.244 | 0.958 to 1.615 |
| 35-44 | $<0.0001$ | 3.730 | 2.727 to 5.100 |
| 45-54 | <0.0001 | 7.569 | 5.553 to 10.318 |
| 55-64 | <0.0001 | 15.552 | 11.457 to 21.109 |
| 65-74 | $<0.0001$ | 31.360 | 23.086 to 42.599 |
| 75-84 | <0.0001 | 76.171 | 55.795 to 103.988 |
| 85+ | $<0.0001$ | 196.778 | 139.258 to 278.057 |
| Sex $\times$ age (female) |  |  |  |
| Male less that 35 years | <0.0001 | 3.013 | 2.308 to 3.932 |
| Male 35 years or older | <0.0001 | 1.709 | 1.556 to 1.878 |
| Ethnicity (white-non-Hispanic) |  |  |  |
| Black | 0.0009 | 1.169 | 1.066 to 1.281 |
| Latino | 0.1761 | 0.904 | 0.782 to 1.046 |
| Marital status (married) |  |  |  |
| Widowed | $<0.0001$ | 1.386 | 1.264 to 1.520 |
| Divorced-separated | <0.0001 | 1.267 | 1.133 to 1.416 |
| Never married | <0.0001 | 1.575 | 1.394 to 1.781 |
| Education (college graduate) |  |  |  |
| Lt high school | 0.0001 | 1.251 | 1.115 to 1.404 |
| High school graduate | 0.0023 | 1.188 | 1.063 to 1.327 |
| Some college | <0.0001 | 1.300 | 1.148 to 1.473 |
| Income ( $>$ \$50K) |  |  |  |
| <\$20K | <0.0001 | 1.367 | 1.211 to 1.543 |
| \$20-35K | 0.0001 | 1.267 | 1.124 to 1.428 |
| \$35-50K | 0.9513 | 0.996 | 0.868 to 1.142 |
| Perceived health (excellent) |  |  |  |
| Very good | 0.0057 | 1.155 | 1.043 to 1.279 |
| Good | <0.0001 | 1.618 | 1.469 to 1.783 |
| Fair | <0.0001 | 2.346 | 2.103 to 2.618 |
| Poor | <0.0001 | 4.126 | 3.611 to 4.715 |
| Miscellaneous |  |  |  |
| Personal care | $<0.0001$ | 1.795 | 1.550 to 2.078 |
| Veteran | 0.0157 | 1.120 | 1.022 to 1.229 |
| Keep house | 0.0004 | 1.207 | 1.087 to 1.339 |
| Major activity limitation | <0.0001 | 1.550 | 1.411 to 1.704 |

Omitted categories are noted within parentheses in "Parameter" column. Source: authors' analysis of 1989 NHIS linked to 1997 NDI.
measuring age, sex, education, race/ethnicity, income, employment status, veteran status, receipt of personal care services (as a proxy for functional status), and self perceived health status. The analyses were completed using the SAS and SPSS programs.

## RESULTS

Among the 35141 cases in the 1989 NHIS that were not useable, $99 \%$ were excluded because the respondent's age was less than 19 years, leaving 80018 cases for analysis. Forty eight per cent of the participants were male and 52\% were female. With regard to education at baseline, $6.2 \%$ had less than an 8 th grade education, $38 \%$ were high school graduates, $31.3 \%$ had some college, $15.6 \%$ had a bachelor's degree, $7.9 \%$ had a college graduate degree, and $1 \%$ refused. About $34.3 \%$ of the participants had family income less than $\$ 20000$ per year. The most common race or ethnic identity was white, non-Hispanic ( $73.6 \%$ ), followed by black/African American (14.6\%), Hispanic (8.5\%), Asian-Pacific Islander ( $2.2 \%$ ), Native American ( $0.7 \%$ ), multiple ( $0.3 \%$ ), and "other" $(0.1 \%)$. Thirty nine per cent of the participants rated their health as excellent, $27.8 \%$ rated their health as very good, $22.7 \%$ as good, $7.2 \%$ as fair, and $2.7 \%$ as poor.
The main focus of the study is on marital status. Among the participants, $47.7 \%$ were married in 1989, $9.9 \%$ widowed, $12.5 \%$ divorced, $3.5 \%$ separated, $5 \%$ living with partner, $0.4 \%$ unknown marital status, and $21.0 \%$ had never been married. Our analyses exclude the small number of unmarried respondents who were living with a partner or had unknown
marital status. The divorced and separated groups were combined into a single category. Age was initially aggregated into 10 year intervals, with the exception of a five year block for those 19-24 years. Because of the small numbers in some cells, age was later aggregated into three larger categories: young (19-44 years), middle adult (45-64 years), and older adult ( 65 years or older).

We investigate the association between marital status and all cause mortality, as well as the association with cause specific mortality. Early in adulthood most of the population was in the never married category (fig 1). By age 35-44, less than $10 \%$ of the population remained never married. The separated or divorced category peaked during midlife (ages 35-54) and declined later in life. Widowers increased rapidly with age, beginning about age 45 .

In a multivariate logistic regression predicting all cause mortality, age was by far the strongest predictor (table 1). As expected, the risk of dying from all causes increases sharply as the population ages. Using married as the reference group, those who were widowed had a $39 \%$ greater risk of mortality and those who were divorced or separated had $27 \%$ greater chance of mortality. The strongest risk was seen for the never married people who were 1.58 (CI 1.39 to 1.78 ) more likely to have died between 1989 and 1997 than were those who were married in 1989. The point estimate for the relation between being never married and mortality was greater than the point estimate for being widowed or divorced/separated, although the $95 \%$ confidence intervals overlap. We define the loss of life expectancy associated with having never been married as the "never married penalty".

Table 2 Logistic regression comparing likelihood of death for never married with reference to married for subgroups of age and sex (upper section) and with health status (lower section)

| Effect | Odds ratio estimate | Lower CI limit | Upper CI limit |
| :--- | :--- | :--- | :--- |
| Never married/married for |  |  |  |
| Male 19-44 | 2.121 | 1.684 | 2.672 |
| Male 45-64 | 1.618 | 1.195 | 2.192 |
| Male 65+ | 1.097 | 0.704 | 1.353 |
| Female 19-44 | 1.674 | 1.204 | 2.328 |
| Female 45-64 | 1.495 | 1.064 | 2.102 |
| Female 65+ | 1.478 | 1.153 | 1.895 |
| Never married/married for perceived |  |  |  |
| health | 1.884 | 1.532 | 2.318 |
| Excellent | 1.733 | 1.415 | 2.123 |
| Very Good | 1.356 | 1.119 | 1.642 |
| Good | 1.355 | 1.058 | 1.736 |
| Fair |  |  |  |

Logistic regression controls for all variables shown in the logistic regression in table 1. For this analysis $n$ for dead $=7418, n$ for alive $=72295$. Covariates include ethnicity, education, personal care source, veteran status, head of household, and major activity limitation.

Other variables included in the model were related to mortality in expected directions. In comparison with those who reported themselves to be in excellent health, there was a systematic increase in the probability of death for those with lower self reported health, with an odds ratio over 4.0 for those reporting poor health. Death rates were higher among men than among women ( $\mathrm{OR}=3.0$ for men less than 35 , and 1.7 for men greater than 35 ), among those with less than a college degree (there is no evidence of a dose-response relation at lower levels of education), among those with lower incomes ( $\mathrm{OR}=1.4$ for family income less than $\$ 20000$ per year compared with those with incomes above $\$ 50000$ per year), and higher among those receiving personal care services, among veterans, among those keeping house (compared with working), and among people whose health prevented them from working. In comparison with white respondents, African-Americans were at a significantly higher risk for early mortality ( $\mathrm{OR}=1.17, \mathrm{CI} 1.06$ to 1.28 ) while Hispanic respondents were not at higher risk ( $\mathrm{OR}=0.90$, CI 0.78 to 1.05 ). If anything, there is a slight protective effect for being Hispanic. The logistic regression that includes interactions of never married with age and sex suggests that being never married is associated with higher death rates for both men and women (table 2, upper section). Never married men age 19 to 44 were 2.12 (CI 1.68 to 2.67) times more likely to die than married men in this age group. The never married penalty declined with age among men,


Figure 1 Marital status by age.
reaching an odds ratio of 1.62 among men 45-64; men in the $65+$ age category were not at significantly greater risk of early death if never married. Among women, there was no evidence of an age gradient in the never married penalty, with odds ratios of $1.64,1.5$, and 1.48 for women in the 19-$44,45-64$, and 65 plus age groups, respectively. The never married penalty was greater among men than among women in the youngest age group, about equal in the middle age group, and larger among women than among men among the elderly group.

One explanation for the never married penalty is that sick people require greater support. Thus we would expect that the never married penalty to be stronger for those with lower health status. We saw a clear monotonic relation between self rated health and probability of death between 1989 and 1996. We examined this issue in more detail by considering the odds ratio of death for the never married compared with currently married status separately for those who rated their own health as excellent, very good, good, or fair. The never married penalty (in odds ratio terms) was larger for those in excellent or very good health and smaller for those in good, fair, or poor health (table 2, lower section).
To investigate the association between being never married and specific causes of death, we estimated logistic regressions with specific causes of death as the dependent variable. Figure 2 summarises the risk of being never married by cause of death. Using married as the reference group, the figure compares odds ratios for the divorced/separated, widowed, and never married for primary cause of death. The figure replicates previous work showing that being divorced or widowed increases the risks of death attributable to cardiovascular disease and cancer, as well as all cause mortality. Although most authors cite the risk of being divorced/separated or widowed, the point estimates suggest that never having been married was a greater risk factor than being divorced or separated for death from all causes, external causes, other causes, and infectious diseases. Furthermore, having never been married was equivalent to being widowed as a risk for death from cardiovascular diseases.
There were only two causes (cancer and pulmonary disease) for which being never married was not a significant risk factor relative to being married.
Table 3 shows odds ratios by cause of death for all never married people, and never married men in the 19-44 and 4564 age groupings. The table also lists confidence intervals. Embolded cells show that never married people are significantly different from those who were married. For young

Table 3 Odds ratios comparing likelihood of death for never married with married by cause of death and confidence intervals for all never married compared with married (both sexes, all ages), never married men age 19-44, and never married men age 45-64

| Cause of death | Both sexes, all ages | Male $19-\mathbf{4 4}$ | Male 45-64 |
| :--- | :--- | :--- | :--- |
| Cardiovascular $(\mathrm{n}=2438)$ | $1.38(1.14$ to 1.67$)$ | $1.14(0.61$ to 2.10$)$ | $2.35(1.57$ to 3.53$)$ |
| Cancer $(\mathrm{n}=1603)$ | $0.97(0.77$ to 1.23$)$ | $0.66(0.31$ to 1.43$)$ | $0.77(0.42$ to 1.44$)$ |
| Pulmonary $(\mathrm{n}=463)$ | $1.09(0.71$ to 1.67$)$ | $2.38(1.66$ to 8.49$)$ | $0.98(0.30$ to 3.16$)$ |
| Infectious $(\mathrm{n}=167)$ | $4.99(3.32$ to 7.52$)$ | $9.08(4.86$ to 17.34$)$ | $6.26(2.32$ to 16.86$)$ |
| Accidents, homicides, suicide $(\mathrm{n}=332)$ | $2.03(1.47$ to 2.80$)$ | $2.35(1.55$ to 3.56$)$ | $1.08(0.26$ to 4.46$)$ |
| Other $(\mathrm{n}=883)$ | $1.75(1.37$ to 2.25$)$ | $2.67(1.54$ to 4.61$)$ | $1.71(0.88$ to 3.32$)$ |
| All causes $(\mathrm{n}=5876)$ | $1.58(1.39$ to 1.78$)$ | $2.12(1.68$ to 2.67$)$ | $1.61(1.19$ to 2.19$)$ |

Logistic regression controls for all variables shown in the logistic regression in table 1. Other includes all deaths not in any other category. Embolded cells show statistically significant effects. N reports on the number of deaths among both sexes, all ages. Covariates include ethnicity, education, personal care source, veteran status, head of household, and major activity limitation.
men aged 19 to 44 , the association between being never married and increased mortality is particularly notable for infectious disease (presumably a result of increased risk of death from AIDS), external causes, and "other" causes. For middle aged men, increased risk associated with being never married was greatest for cardiovascular and infectious diseases. For infectious disease and all cause mortality, there was no differentiation by age for men. For cardiovascular disease, the relative risk of being never married was greater for older men than younger men. For pulmonary disease, accidents, homicides and suicides, and other causes, the point estimates suggest that being never married was a greater relative risk factor for younger than for older men.

## DISCUSSION

A variety of studies have shown that unmarried adults have a higher probability of early death than those that are married. ${ }^{19}$ 20 However, most of these studies do not differentiate those who are separated and divorced from those that were never married. A recent detailed review of the literature ${ }^{21}$ does not treat never married as a separate category. Our results suggest that those who are never married are at greater risk than those who are separated and divorced. To observe the impact of never married status upon
different causes of death and at different ages, a large sample is required. We are not aware that the NHIS with linkage to the NDI has previously been used for this purpose. Among the published studies, only Johnson and colleagues have reported on a large cohort. Using a cohort of 281460 participants in the national longitudinal mortality study (NLMS), Johnson et al reported findings similar to ours for cardiovascular disease mortality and for cancer. However, they did not include the younger age groups, so were unable to identify the increased risk of early infectious disease deaths in younger never married men. ${ }^{22}$
We have shown, as have several other authors, ${ }^{35}{ }^{22-28}$ that there is an association between being never married and increased risk of death. Hu and Goldman have shown that the effects of being married upon mortality occurs in a wide variety of cultures. ${ }^{25}$ Reports of similar effects have emerged from the USA, ${ }^{22}$ Great Britain, ${ }^{23}$ Sweden, ${ }^{27}$ Denmark, ${ }^{26}$ and the Netherlands. ${ }^{28}$ None of the studies has established a causal relation.
Accumulated evidence suggests that social isolation increases the risk of premature death. Marriage is a rough proxy for social connectedness. Among categories on being unmarried, we suggest that having never married may be associated with more severe isolation because it is associated


Figure 2 Odds ratios by marital status and cause of death. *Shows confidence limit excludes 1.0.
with greater isolation from children and other family. The data seem to support the hypothesis that the greater level of social isolation associated with having never married is associated with larger health consequences. We consider several rival explanations for the association between social relationships and mortality.

## Illness causes disruption in social relationships

One explanation for the never married penalty is that those who are seriously ill are less suitable as marriage partners. However, half of those never married in the 25-34 age group do marry by age 35 (fig 1) and we saw the consequences of being never married in all age groups. Some of the literature on social support argues that disruption in social ties is stressful and leads to poor health outcomes. ${ }^{1515}$ However, our data suggest that the effect of never having been married is stronger than being divorced or becoming widowed. It might also be argued that the duration of social isolation has a cumulative effect. For example, Lund and colleagues followed up 10891 younger men in Denmark. They found a cumulative effect of marriage upon life expectancy. The effects of divorce may be attenuated by remarriage while duration without a female spouse may predict early mortality for younger men. ${ }^{26}$ Data from our analyses argue the opposite. With increasing age, the marriage penalty among men decreases rather than increases. The analysis adjusting for self rated health status did not confirm the hypotheses that the marriage penalty results from low health status in the never married. ${ }^{29}$ In fact, the marriage penalty is strongest among those who report themselves to be in excellent health (table 2). We acknowledge that this result is difficult to interpret because the absolute risk is higher for participants with the lowest self rated health while the relative risk is higher for the well portion of the population. In addition to a true causal relation, there are several other explanations. The greater relative risk among the well portion of the population might be explained by a selection effect (people in poor health are less attractive as marriage partners), or by an omitted variable.

## High risk behaviours

Another explanation for the marriage penalty is that those who have never been married engage in higher risk behaviours than those who are married. ${ }^{630}$ Our findings show that the never married penalty is greatest for younger adults and that the relation is strongest for infectious disease-presumably deaths related to HIV infection. It is probable that the increase in infectious disease deaths in younger men may be attributable to the prevalence of HIV in

## What the paper adds

Our study raises a series of new questions. Firstly, we found that having never been married is a better predictor of poor health outcomes than either divorce or widowhood. Secondly, the impact of social isolation is not constrained to the elderly. In fact, it is comparatively stronger early in life. This phenomenon may have been overlooked in previous studies because early death is uncommon. The early effects of having never been married are only detectable in studies with very large sample sizes. Finally, our findings challenge the belief that the effect of social isolation is upon diseases of the heart. Our findings underscore the impact on all cause mortality. Furthermore, the predominant cause of death associated with social isolation differs at different stages of the life cycle.

## Policy implications

The paper suggests that having never been married may be a risk factor for mortality in all age groups.
gay unmarried men. NHIS does not ask about sexual preference, so we cannot formally evaluate this hypothesis.

The decrease in the never married penalty with age may suggest that as the population ages, risky health behaviours decline. Despite the attractiveness of the risky behaviour hypothesis, there is some evidence to contradict it. It is unlikely that the never married penalty is associated with poor health habits. For example, we analysed data from the US NHIS and found that those who were never married were only slightly more likely than those currently married to be current every day smokers ( $18.6 \%$ compared with $15.7 \%$ ). Furthermore, those never married were less likely to be smokers than those who were divorced ( $27.9 \%$ ) or separated ( $31.4 \%$ ), or living with a partner ( $35.1 \%$ ). Never married adults were more likely to have been lifetime abstainers from alcohol and were slightly less likely than currently married people to be regular alcohol consumers. The never married group exercise slightly more than those that were married and were less likely to be overweight. Overall, those who have never married have slightly better health habits and risk factors than married peers and never married adults have notably better health habits than those who are divorced or separated.

## Usual source of care

This study was conducted in the USA where a significant proportion of the population does not have a usual source of health care. One potential explanation for our results is that never married people are less likely to have a usual source of medical care. Although this factor did not clearly emerge in our analysis of the NHIS data, evidence reported in Vital and Health Statistics suggested that the never married were more likely to lack a usual source of care ( $17 \%$ ) in comparison with those who are married ( $10.5 \%$ ), widowed ( $11 \%$ ) or divorced or separated ( $16 \%$ ). Only those "living with a partner" are less likely to have a usual source of medical care (20.7\%). The most common reason for not having a usual source of care is being uninsured. As many people gain insurance through a spouse or a former spouse, never married status may be associated with a greater degree of separation from usual health care.

Unfortunately, the NHIS data do not permit much exploration of sub-groupings within the never married category. For example, lesbian/gay status cannot be determined from the information available to us. We encourage evaluation of these factors using other datasets.

## Practical implications

Our study may have some practical implications. In particular, it suggests that those who have never married may be in need of particular attention. The risks of being never married, in terms of odds ratios, rival the risks of having increased blood pressure or high cholesterol. There are also some important research implications. In many epidemiological studies, the never married category is merged with widowed and divorced/separated. Our data suggest that being never married deserves a separate category and more research attention.

## Limitations

Our study has significant limitations. The data were limited to baseline information obtained from a wide ranging
interview and mortality follow up. As a result we are limited to associations between broadly defined variables. Because we had only baseline data, we are unable to assess the impact of changes in health status or changes in marital status over the course of time. Comparison of age groups in a cross sectional study cannot substitute for longitudinal analysis. Secular changes and cohort effects could serve as alternative explanations for some of our findings. On the other hand, we are encouraged by the concordance of our findings and those reported by other authors who used very different populations and different methodologies.

## CONCLUSIONS

Lynch argued that social isolation is a significant factor in death from heart disease. A variety of studies in the literature underscore the impact of social isolation upon diseases of the heart. ${ }^{1-5} 911131531$ Furthermore, a growing literature shows that widowhood may also be an important predictor of mortality in the elderly population. ${ }^{251331}$
Our study raises a series of new questions. Firstly, we found that having never been married is a better predictor of poor health outcomes than either divorce or widowhood. Secondly, the impact of social isolation is not constrained to the elderly. In fact, it is comparatively stronger early in life. This phenomenon may have been overlooked in previous studies because early death is uncommon. The early effects of having never been married are only detectable in studies with very large sample sizes. Finally, our findings challenge the belief that the effect of social isolation is upon diseases of the heart. Our findings underscore the impact on all cause mortality. Furthermore, the predominant cause of death associated with social isolation differs at different stages of the life cycle.

## Authors' affiliations

R M Kaplan, University of California, Los Angeles, USA
R G Kronick, University of California, San Diego, USA
Funding: none.
Conflicts of interest: none

## REFERENCES

1 Lynch JJ. The broken heart: the medical consequences of loneliness in America. New York: Basic Books, 1977.
2 The golden wedding anniversary: an increased likelihood. Stat Bull Metropol Life Insur Co 1966;47:1-2.
3 Adams O. Life expectancy in Canada-an overview. Health Rep 1990;2:361-76.
4 Berkman LF, Syme SL. Social networks, host resistance, and mortality: a nineyear follow-up study of Alameda County residents. Am J Epidemiol 1979;109:186-204.

5 Blazer DG. Social support and mortality in an elderly community population. Am J Epidemiol 1982;115:684-94.
6 Fletcher BC. Marital relationships as a cause of death: an analysis of occupational mortality and the hidden consequences of marriage-some UK data. Hum Relat 1983;36:123-33.
7 Goldman N. Changes in widowhood and divorce and expected durations of marriage. Demography 1984;21:297-307.
8 House JS, Robbins C, Metzner HL. The association of social relationships and activities with mortality: prospective evidence from the Tecumseh Community health study. Am J Epidemiol 1982;116:123-40.
9 Marmot MG, Syme SL, Kagan A, et al. Epidemiologic studies of coronary heart disease and stroke in Japanese men living in Japan, Hawaii and California: prevalence of coronary and hypertensive heart disease and associated risk factors. Am J Epidemiol 1975;102:514-25.
10 Medalie JH, Goldbourt U. Angina pectoris among 10,000 men. II. Psychosocial and other risk factors as evidenced by a multivariate analysis of a five year incidence study. Am J Med 1976;60:910-21.
11 Megson B. Life expectations of the widows and orphans of freemen in London 1375-1399. Local Popul Stud 1996;57:18-29.
12 Oechsli FW. A population model based on a life table that includes marriage and parity. Theor Popul Biol 1975;7:229-45.
13 Seeman TE, Kaplan GA, Knudsen L, et al. Social network ties and mortality among the elderly in the Alameda County study. Am J Epidemiol 1987;126:714-23.
14 Schoenbach VJ, Kaplan BH, Fredman L, et al. Social ties and mortality in Evans County, Georgia. Am J Epidemiol 1986;123:577-91.
15 Reed D, McGee D, Yano K, et al. Social networks and coronary heart disease among Japanese men in Hawaii. Am J Epidemiol 1983;117:384-96.
16 Welin L, Larsson B, Svardsudd K, et al. Social network and activities in relation to mortality from cardiovascular diseases, cancer and other causes: a 12 year follow up of the study of men born in 1913 and 1923. J Epidemiol Community Health 1992;46:127-32.
17 Kovar MG. Data systems of the National Center for Health Statistics. Vital Health Stat 1989;23:1-21.
18 Adams PF, Benson V. Current estimates from the National Health interview survey, 1989. Vital Health Stat 1990;176:1-221.
19 Verbrugge LM, Wingard DL. Sex differentials in health and mortality. Women Health 1987;12:103-45.
20 Lahorgue Z. Morbidity and marital status. J Chronic Dis 1960;12:476-98
21 Uchino BN. Social support and physical health: understanding the health consequences of relationships. New Haven, CT: Yale University Press, 2004.
22 Johnson NJ, Backlund E, Sorlie PD, et al. Marital status and mortality: the national longitudinal mortality study. Ann Epidemiol 2000;10:224-38.
23 Ebrahim S, Wannamethee G, McCallum A, et al. Marital status, change in marital status, and mortality in middle-aged British men. Am J Epidemiol 1995;142:834-42.
24 Hu Y. Patterns of mortality differentials by marital status in low mortality countries. Ingu munje noniip 1987;10:97-1 28.
25 Hu YR, Goldman N. Mortality differentials by marital status: an international comparison. Demography 1990;27:233-50.
26 Lund R, Holstein BE, Osler M: Marital history from age 15 to 40 years and subsequent 10-year mortality: a longitudinal study of Danish males born in 1953. Int J Epidemiol 2004;33:389-97.

27 Modin B. Born out of wedlock and never married-it breaks a man's heart. Soc Sci Med 2003;57:487-501.
28 van Poppel F, Joung I. Long-term trends in marital status mortality differences in the Netherlands 1850-1970. J Biosoc Sci 2001;33:279-303.
29 Tun W, Gange SJ, Vlahov D, et al. Increase in sexual risk behavior associated with immunologic response to highly active antiretroviral therapy among HIVinfected injection drug users. Clin Infect Dis 2004;38:1167-74.
30 Strathdee SA, Galai N, Safaiean M, et al. Sex differences in risk factors for hiv seroconversion among injection drug users: a 10-year perspective. Arch Intern Med 2001;161:1281-8.
31 Goldman N, Lord G. Sex differences in life cycle measures of widowhood. Demography 1983;20:177-95.

