Mendelian Randomization Methodology

# Mendelian randomization studies: a review of the approaches used and the quality of reporting 

Anna G C Boef, ${ }^{1 *}$ Olaf M Dekkers ${ }^{1,2}$ and Saskia le Cessie ${ }^{1,3}$<br>${ }^{1}$ Department of Clinical Epidemiology, ${ }^{2}$ Department of Endocrinology and Metabolic Diseases, and<br>${ }^{3}$ Department of Medical Statistics and Bioinformatics, Leiden University Medical Centre, Leiden, The Netherlands.<br>*Corresponding author. Department of Clinical Epidemiology, Leiden University Medical Centre, PO Box 9600, 2300 RC Leiden, The Netherlands. E-mail: a.g.c.boef@lumc.nl

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#### Abstract

Background: Mendelian randomization (MR) studies investigate the effect of genetic variation in levels of an exposure on an outcome, thereby using genetic variation as an instrumental variable (IV). We provide a meta-epidemiological overview of the methodological approaches used in MR studies, and evaluate the discussion of MR assumptions and reporting of statistical methods. Methods: We searched PubMed, Medline, Embase and Web of Science for MR studies up to December 2013. We assessed (i) the MR approach used; (ii) whether the plausibility of MR assumptions was discussed; and (iii) whether the statistical methods used were reported adequately. Results: Of 99 studies using data from one study population, 32 used genetic information as a proxy for the exposure without further estimation, 44 performed a formal IV analysis, 7 compared the observed with the expected genotype-outcome association, and 1 used both the latter two approaches. The 80 studies using data from multiple study populations used many different approaches to combine the data; 52 of these studies used some form of IV analysis; $44 \%$ of studies discussed the plausibility of all three MR assumptions in their study. Statistical methods used for IV analysis were insufficiently described in $14 \%$ of studies. Conclusions: Most MR studies either use the genotype as a proxy for exposure without further estimation or perform an IV analysis. The discussion of underlying assumptions and reporting of statistical methods for IV analysis are frequently insufficient. Studies using data from multiple study populations are further complicated by the combination of data or estimates. We provide a checklist for the reporting of MR studies.


Key words: Mendelian randomization, instrumental variable, aetiology

## Key Messages

- The specific methods used in Mendelian randomization studies vary widely.
- These methods broadly fall into three categories: (i) using genetic information as a proxy for the exposure without further estimation; (ii) performing an instrumental variable analysis; (iii) comparing the observed with the expected genotype-outcome association.
- Mendelian randomization studies frequently insufficiently discuss underlying assumptions or report statistical methods for IV analysis.
- A checklist for the reporting of Mendelian randomization studies is provided.


## Introduction

Observational studies are limited in their ability to identify whether exposures are causally related to disease occurrence or other outcomes. Adjustment for confounding is only possible for those factors which are identified and measured and will inevitably be incomplete: some degree of residual confounding will always remain. Reverse causation, an effect of the outcome on the studied exposure, may also explain associations found in an observational study. ${ }^{1,2}$ An approach which can circumvent both reverse causation (as first proposed in 1986$)^{3}$ and residual confounding in order to establish the causal effect of the exposure on the outcome is to investigate the effect of genetic variation in levels of the exposure on the outcome. This approach has come to be known as Mendelian randomization over the past decade. ${ }^{2}$ The random allocation of genetic variants from parents to offspring means these variants will generally be unrelated to other factors which affect the outcome. ${ }^{1}$ Furthermore, associations between the genotype and the outcome will not be affected by reverse causation because disease does not affect genotype. ${ }^{1}$

Mendelian randomization studies use genetic variation as an instrumental variable (IV) and must fulfil instrumental variable assumptions. Applied to Mendelian randomization, these assumptions are that (i) the genotype is associated with the exposure; (ii) the genotype is associated with the outcome through the studied exposure only (exclusion restriction assumption); and (iii) the genotype is independent of other factors which affect the outcome (independence assumption). ${ }^{4}$ Potential threats to the validity of these assumptions, such as population stratification, linkage disequilibrium and pleiotropic effects, are discussed in detail elsewhere. ${ }^{1,5}$

These general principles of Mendelian randomization are increasingly used in aetiological research, but the specific methods used in these studies can vary widely. In this study we review the methodology used in studies from the past 10 years which were described by the authors as Mendelian randomization studies. We provide an overview of the use of the different approaches to Mendelian
randomization and, where applicable, the specific statistical methods used for estimation. We evaluate whether the plausibility of the Mendelian randomization assumptions is discussed. Further, we evaluate whether the statistical methods used are sufficiently described (including how the confidence interval was obtained) for those studies which perform an instrumental variable analysis or compare the observed and expected genotype-outcome association.

## Methods

## Search strategy and inclusion criteria

We searched PubMed, Medline, Embase and Web of Science for studies containing the term 'Mendelian randomization' or 'genetic instrumental variable' or a related term (e.g. 'genetic instrument')from 1 January 2003 to 31 December 2013. The full search strategies for each of the databases are included in Supplementary Methods, available as Supplementary data at $I J E$ online. We excluded publications that; (i) were conference abstracts, letters, commentaries, editorials, reviews, study proposals or theoretical papers; (ii) did not use Mendelian randomization (i.e. did not state Mendelian randomization or a genetic instrumental variable was used in the text, abstract or title and did not include 'Mendelian randomization' or 'genetic instrumental variable' or a related term as a keyword); (iii) identified potential genetic instruments for future Mendelian randomization studies; (iv) were primarily methodological, using an application of Mendelian randomization as an example; or (v) were published in a health economics journal (rather than a biomedical journal).

## Classification of Mendelian randomization approach used

First we classified publications into studies which used data from a single study population and studies which used data from multiple study populations. We then classified
included studies according to their general Mendelian randomization approach: i.e. how they utilized the genetically determined variation in exposure.
A. For studies performed in a single study population we identified the following three main approaches:
i. Use of genetic variation as a proxy for the exposure, without further estimation.
These studies investigate the association between a genotype (which affects the exposure) and the outcome. No comparison is made with the expected association between this genotype and the outcome, and no IV estimate of the effect of the exposure on the outcome is obtained.
ii. Comparison of the observed and expected geno-type-outcome associations.
These studies compare the observed association between the genotype and the outcome with the association which would be expected if the observed exposure-outcome association were causal. This expected association is calculated by multiplying the observed genotype-exposure association with the observed exposure-outcome association (sometimes termed the 'triangulation' approach, although this is not a specific term); see Figure 1 for an illustration and further explanation. The confidence interval of the expected genotype-outcome association can be estimated analytically or using bootstrap techniques.
iii. Formal instrumental variable analysis using genetic variation as the instrument.
These studies perform a formal IV analysis to obtain a causal estimate of the effect of genetically determined variation in the exposure on the outcome. Different statistical techniques can be used for this purpose, as we will further explain below.
B. We classified the studies including more than one study population into the following pre-specified main categories:
i. Pooling of the data, followed by any of the approaches Ai-iii listed above.
ii. IV analysis in each of the study populations, followed by a meta-analysis.
iii. Meta-analysis using the genotype as a proxy for the exposure, without further estimation.
iv. Meta-analyses of the genotype-exposure, exposureoutcome and genotype-outcome associations, followed by comparison of observed and expected genotype-outcome associations (as in approach Aii).
v. Meta-analyses of the genotype-exposure and genotype-outcome associations, followed by a

Wald-type/ratio estimate (see Didelez et al. for a description of Wald-type estimators). ${ }^{6}$
vi. Data analysed separately for more than one population, followed by any of the approaches Ai-iii.

Further categories were added for those studies which did not fall into any of the above categories.

## Assessment of discussion of Mendelian randomization assumptions

Regardless of the approach used, Mendelian randomization studies rely on three main assumptions, as briefly mentioned in the introduction.
i. The genotype is associated with the exposure.

This assumption can and should be verified in the data. Reporting guidelines for IV analyses recommend the use of the partial F-statistic as a measure of the strength of the association between the IV and the exposure. ${ }^{7,8}$ It encompasses information on the strength of the instrument and on the number of observations in the analysis. ${ }^{9}$ We assessed whether studies reported the strength of the genotype-exposure association in the data using a partial F-statistic or using another measure (e.g. mean difference in exposure by genotype). If not, we assessed whether they reported the strength of this association from literature.
ii. The genotype is associated with the outcome through the studied exposure only (exclusion restriction assumption).
This assumption is violated if the genotype has multiple (pleiotropic) effects, if a nearby variant with which it is in linkage disequilibrium affects the outcome in other ways than through the exposure of interest, or if developmental canalization occurs. ${ }^{1}$ For all studies we evaluated whether the plausibility of this assumption was discussed. Mentioning the assumption in general terms was not deemed sufficient: a specific discussion of its plausibility in the particular study was required.
iii. The genotype is independent of other factors which affect the outcome (independence assumption).
This assumption is violated if subgroups in the study population have both different genotype frequencies and different distributions of the outcome (population stratification). ${ }^{1}$ It is also violated if there is an association between the genotype used as an instrument and confounders. For all studies we assessed whether the association between the genetic instrument and measured confounders was reported, as recommended in IV reporting guidelines. ${ }^{7}$ Furthermore, we assessed whether potential associations of the genotype with


Figure 1. Diagram of the approach used by Mendelian randomization studies which compare the observed genotype-outcome association with the expected genotype-outcome association. $\beta_{1}$, regression coefficient of the genetic variant-exposure association. $\beta_{2}$, regression coefficient of the exposure-outcome association. $\beta_{30 B 5}$, observed regression coefficient of the genetic variant-outcome association. $\beta_{3 \text { EXP }}$, expected regression coefficient of the genetic variant-outcome association. The point estimate of $\beta_{3 \text { EXP }}$ is calculated as follows: $\beta_{3 \text { EXP }}=\beta_{1} \cdot \beta_{2}$. The confidence interval of the expected genotype-outcome association can be estimated analytically or using bootstrap techniques.
unmeasured confounders were discussed and/or population stratification was discussed. Again, a specific discussion of the plausibility of the assumption in the particular study was required.

## Assessment of reporting of statistical analysis

This section only applies to the studies which used the IV approach or the observed-expected approach, because using genetic variation as a proxy for the exposure without further estimation does not involve any special statistical methods.
i. For studies which obtained an IV estimate of the effect of the exposure on the outcome, we determined which statistical method was used and assessed whether it was described sufficiently and whether a confidence interval was reported. A frequently used IV method is two-stage least squares analysis. This involves two stages of linear regression. The first stage is a linear regression with the exposure as the dependent variable and the instrument (genotype) as the independent variable, which is then used to obtain predicted exposure levels based on the instrument. The second stage is a regression with the outcome as the dependent variable and these genetically predicted exposure levels as the independent variable. Software for two-stage least squares regression takes into account the errors in both stages of the analysis to give a correct confidence interval. Additionally, we determined the type of outcome investigated (continuous, binary, time-to-event) and for binary outcomes what kind of target parameter was estimated (risk difference, odds ratio, relative risk, probit coefficient).

We also determined whether a statistical test was used to compare the IV estimate with the 'conventional' estimate of the effect of the exposure on the outcome, what type of genetic instrument was used [one single nucleotide polymorphism (SNP) or allele, multiple SNPs in separate analyses, multiple SNPs in a single analysis, combinations of SNPs e.g. haplotypes or a genetic risk score] and for those studies which used multiple SNPs in a single analysis, whether weak instrument bias was discussed. In the IV studies within one study population, we also determined whether the genetic variant used as an instrument was identified or selected in the same population or if the weights for a weighted genetic risk score were derived in the same population.
ii. For studies comparing the observed and expected geno-type-outcome association, we assessed whether the method used to obtain a point estimate of the expected genotype-outcome association was described. If the description was such that calculation of this point estimate should be possible using the data provided, we assessed whether the point estimate corresponded to our calculations (only in those studies within one population). Further, we assessed whether a confidence interval for the expected genotype-outcome association was reported, whether the method used to obtain this confidence interval was described, and whether the confidence interval incorporated the variance of both the genotype-exposure association and the exposureoutcome association.

## Results

Our search returned 1911 hits, of which 594 hits remained after exclusion of conference abstracts and duplications. After reviewing the title and abstract and if necessary the full-text article, a further 415 records were excluded for reasons listed in the flowchart in Figure 2, resulting in 179 eligible Mendelian randomization studies. Of these 179 studies, 99 studies used data from a single study population for their main analyses, ${ }^{10-108}$ and 80 studies used data from more than one study population (Table 1). ${ }^{109-188}$ The included studies were published between May 2005 and December 2013. An overview of the exposures studied and the genetic instruments used is presented in Supplementary Table 1 (available as Supplementary data at $I J E$ online). The most frequently studied exposures were C-reactive protein (29 studies) and adiposity measures such as body mass index, fat mass and percentage body fat ( 25 studies).

Of the 99 studies which used data from a single study population, 38 studies ( $38 \%$ ) used the genetic information


Figure 2. Summary of literature search.
as a proxy for the exposure by investigating the genotypeoutcome association without further estimation of either the causal effect of the exposure on the outcome or of the expected genotype-outcome association (Table 1); 48 studies ( $48 \%$ ) used IV analysis to estimate the effect of genetically determined variation in exposure levels on the outcome. Ten studies compared the observed association between the genotype and the outcome with the expected association based on the genotype-exposure association and the exposure-outcome association. One study used both these latter two approaches. For two studies we could not categorize the methods used into any of the aforementioned approaches.

Of the 80 studies which used data from multiple study populations, $26(33 \%)$ studies pooled the data from the different studies and subsequently analysed the pooled data (Table 1). Ten studies performed an IV analysis in the different studies followed by a meta-analysis; 41 studies ( $51 \%$ ) first performed a meta-analysis of one or more of the genotype-exposure, exposure-outcome and genotypeoutcome associations, 26 of which subsequently used these meta-analysed associations for further estimation of either the causal effect of the exposure on the outcome or of the expected genotype-outcome association. In total, 52 studies ( $65 \%$ ) used some form of IV analysis to obtain a causal effect of the exposure on the outcome. A further 23 studies

Table 1. Approaches used in Mendelian randomization studies

| Data from 1 study population | $(n=99)$ | Refs |
| :--- | :---: | :---: |
| 1. Genotype used as a proxy for exposure, without further estimation |  |  |
| a | $10-47$ |  |
| 2. Comparison of observed and expected genotype-outcome association | 38 |  |
| 3. IV analysis ${ }^{\text {a }}$ | 10 | $48-57$ |
| 4. Comparison of observed and expected genotype-outcome association and IV analysis | 48 | $58-105$ |
| 5. Unclear | 106 |  |
| Data from more than 1 study population | $(n=80)$ |  |


| 1. Data pooled, then analysed |  |  |
| :---: | :---: | :---: |
| a. Genotype used as a proxy for exposure, without further estimation | 2 | 109,110 |
| b. Comparison of observed and expected genotype-outcome association ${ }^{\text {a }}$ | 3 | 111-113 |
| c. IV analysis ${ }^{\text {b }}$ | 14 | 114-127 |
| d. Comparison of observed and expected genotype-outcome association and IV analysis | 7 | 128-134 |
| 2. IV analyses, then meta-analysis | 10 | 135-144 |
| 3. Meta-analysis using genotype as a proxy for exposure, without further estimation | 15 | 134-136,145-157 |
| 4. Meta-analyses* followed by comparison of observed and expected genotype-outcome association ${ }^{\text {c }}$ | 13 | 139,144,157-167 |
| 5. Meta-analyses* followed by a Wald-type/ratio estimate | 9 | 168-176 |
| 6. Data analysed and reported separately for more than 1 population |  |  |
| a. Genotype used as a proxy for exposure, without further estimation ${ }^{\text {c }}$ | 3 | 177,179,180 |
| b. IV analysis | 3 | 178,181,182 |
| 7. Multivariate meta-analysis | 2 | 155,183 |
| 8. Bayesian meta-analysis | 1 | 184 |
| 9. Separate study IV-analysis | 1 | 185 |
| 10. Meta-analysis of gene-exposure association, then ratio estimate, then meta-analysis | 1 | 186 |
| 11. Other/unclear** | 2 | 187,188 |

Some studies used multiple approaches in non-identical sets of study populations.
${ }^{\mathrm{a}}$ Two studies also performed an IV analysis for which it was unclear how the data were combined. ${ }^{112,113}$
${ }^{\mathrm{b}}$ One study performed some of the analyses in a single study population. ${ }^{118}$
${ }^{\text {c }}$ Two studies also performed an IV analysis in a single study population. ${ }^{158,177}$
*Meta-analyses of genotype-exposure, exposure-outcome and/or genotype-outcome associations.
**One study first investigated the genotype-outcome association and then performed further analyses for which the approach was unclear. ${ }^{187}$ One study used a 'likelihood-based method for combining summarised genetic association estimates'. ${ }^{188}$
compared the observed and expected genotype-outcome associations.

Table 2 summarizes the reporting of the Mendelian randomization assumptions. Reporting of assumptions was assessed in 178 studies, because the design of one study was so different from the general Mendelian randomization design that the assumptions could not be assessed. A total of 37 out of 98 studies ( $38 \%$ ) which used a single study population and 42 out of 80 studies ( $53 \%$ ) which used multiple study populations explicitly discussed the plausibility of all three Mendelian randomization assumptions in the context of their study.

Among the studies which performed an IV analysis, those using a single study population most frequently studied a continuous outcome, whereas those using multiple study populations most frequently studied a binary outcome and estimated an odds ratio (Table 3). The statistical methods used in these formal IV studies are shown in Table 4. Two-stage least squares (2-SLS) regression was
the most common method used in studies within one study population $(n=26,53 \%)$. Ten studies within multiple study populations also used this method. One study used 2-SLS with a binary outcome, but it did not mention whether heteroskedasticity robust standard errors were used. ${ }^{69}$ Among the studies which used multiple study populations, a Wald-type or ratio estimator was most frequently used ( $n=16,31 \%$ ). The method used to obtain the confidence interval for the ratio estimate was a Taylor series expansion (termed the delta method ${ }^{138,141,168,169,186}$ or Taylor expansion ${ }^{170}$ ), Fieller method, ${ }^{120,176-178}$ or was not described. Three studies in a single study population also used a Wald-type/ratio estimator, but two of these studies did not report a confidence interval. Other methods used were control functions ( $n=8$ in total), IV probit regression $(n=4)$, generalized method of moments $(n=8)$, generalized least squares regression $(n=5)$, quasilikelihood and variance function $(n=4)$ and a two-stage approach with a linear first stage and a logistic second

Table 2. Reporting of Mendelian randomization assumptions

| Criteria | 1 Study population $(n=98)$ * | $>1$ Study population $(n=80)$ |
| :---: | :---: | :---: |
| Strength of genetic instrument-exposure association (assumption 1) |  |  |
| Verified in data using F-statistic | 33 | 26 |
| Otherwise verified in data (e.g. using risk difference or odds ratio) | 53 | 45 |
| Reported from literature | 4 | 4 |
| Not reported | 8 | $5^{\ddagger}$ |
| Plausibility of exclusion restriction assumption discussed (assumption 2) | 56 | 55 |
| Independence assumption (assumption 3) |  |  |
| Instrument-confounder associations shown \& assumption further discussed theoretically ${ }^{\text {a }}$ | 20 | 16 |
| Instrument-confounder associations shown, assumption not further discussed | 30 | 21 |
| Investigation of instrument-confounder associations mentioned, not shown $\&$ assumption further discussed theoretically | 4 | 0 |
| Investigation of instrument-confounder associations mentioned, not shown $\&$ assumption not further discussed | 8 | 0 |
| Plausibility of assumption theoretically discussed only | 7 | 18 |
| Plausibility of assumption not discussed | 29 | 25 |

*Reporting of assumptions was not assessed in one study, because its design was vastly different from the general design of a Mendelian randomization study. The total number of studies within 1 study population is therefore 98 .
${ }^{\text {a P Potential association with unmeasured confounders discussed and/or population stratification discussed. }}$
${ }^{\ddagger}$ Two studies reported a $P$-value only.

Table 3. Types of outcome and parameters estimated in IV Mendelian randomization studies

| Type of outcome | 1 Study <br> population <br> $(n=49)$ | Refs | $>1$ Study <br> population <br> $(n=52)$ | Refs |
| :--- | :--- | :--- | :--- | :--- |
| Continuous |  |  | 14 | $114-116,118,126,135-139,168,170,182,185$ |
| Binary | 37 | $58-68,70-89,94,98,103-106$ |  |  |
| Risk difference | 3 | $69,93,99$ | 0 | - |
| Relative risk | 2 | 81,102 | 2 | 118,124 |
| Odds ratio | 7 | $67,82,88,90-92,100$ | 37 | $112-114,117,119-121,123,125,128-135,139-144,155,158,168,169$, |
|  |  | 83 | 1 | $12717-176,178,181,183,184,186$ |
| Probit coefficient | 1 | $95-97,101$ | 5 | $124,138,168,177,178$ |
| Time-to-event | 4 | - | 1 | 127 |

The total number of types of outcome and parameters estimated exceeds the total number of studies because some studies included multiple types of outcomes.
stage $(n=5)$. Four of the studies which used this last approach did not report how the correct confidence interval was obtained, ${ }^{125,133,135,158}$ and the fifth used a sandwich estimator. ${ }^{114}$ The IV method was insufficiently described in 14 studies. In six of these studies there was a discrepancy between the statistical method reportedly used (2-SLS) and the effect estimate reported (OR). ${ }^{100,128,129,130,131,134}$ Another study seemingly did not take into account the variance of the genotype-exposure association in the variance of the IV estimate, which would result in too narrow a confidence interval. ${ }^{101}$

Of the101 studies which used one of the approaches which yields an IV estimate, 48 reported tests of the
difference between the IV estimate and the conventional estimate: the most commonly used were (a variant of) the Durbin-Wu-Hausman test (29 studies), ${ }^{58-60,62,64-66,68,}$ 70-72,74,75-77,79-87,94,98,105,116,139 and (a variant of) the Bland-Altman test(10 studies). ${ }^{112,113,117,119,128-132,134}$ The types of genetic instrument used (e.g. a single SNP or a genetic risk score) in the IV analysis studies are listed in Supplementary Table 2 (available as Supplementary data at $I J E$ online). Of the 25 studies which used multiple SNPs in a single analysis, 13 mentioned weak instrument bias, with two studies very specifically discussing it in relation to using multiple instruments. ${ }^{117,185}$ Of the 49 studies which used IV methods and were performed in one study

Table 4. Statistical methods used in the instrumental variable studies

| Method | 1 Study population ( $n=49$ ) | Refs | $>1$ Study population* ( $n=52$ ) | Refs |
| :---: | :---: | :---: | :---: | :---: |
| Two-stage least squares | 26 | 58-83 | 10 | 114-116,118,135-139,182 |
| Instrumental variable regression in Stata, not further specified (2-SLS, GMM or LIML) | 5 | 84-88 | 0 | - |
| Control function | 6 | 81,82,89-92 | 2 | 139,143 |
| Instrumental variable probit regression | 3 | 67,83,93 | 1 | 122 |
| GMM | 2 | 94,98 | 0 | - |
| Multiplicative GMM | 0 | - | 6 | 117-121,124 |
| Generalized least squares regression | 1 | 95 | 4 | 112,113,123,132 |
| Two-stage: linear first stage, logistic second stage | 0 | - | 5 | 114,125,133,135,158 |
| Quasi-likelihood and variance function | 1 | 88 | 3 | 140,144,181 |
| Ratio/Wald-type estimator | 1 | 99 | 17 | 120,124,138,141,142,168-178,186 |
| Ratio/Wald-type estimator without confidence interval | 2 | 96,97 | 0 | - |
| Insufficiently described/unclear | 7 | 100-106 | 7 | 126-131,134 |
| Other ${ }^{\text {b }}$ | 0 | - | 4 | 155,183-185 |

The total number of statistical methods exceeds the total number of studies because some studies investigated multiple statistical methods.
2-SLS, two-stage least squares; GMM, generalized method of moments; LIML, limited-information maximum likelihood.
*Including the two studies which used multiple study populations, but performed the IV analysis in a single study population.
${ }^{\mathrm{b}}$ See Table 1.
population, 14 evidently identified or selected the genetic variant used as an instrument in the same population or derived weights for a weighted genetic risk score in the same population. ${ }^{61,62,65,66,70,71,82,85,87,91,93,102,104,105}$

In 3 of the 11 studies comparing the observed geneoutcome association with the expected gene-outcome association in one study population we could not reconstruct the point estimate of the expected association from the data. ${ }^{48,50,57}$ Four studies did not report a confidence interval for the expected genotype-outcome association. ${ }^{52,54,56,106}$ In a further five studies, the methods used to calculate this confidence interval were unclear, ${ }^{48-50,53,57}$ and in one study only the error in the exposure-outcome association seemed to have been taken into account in the calculation of this confidence interval. ${ }^{55}$ Only one study adequately described the methods used to obtain the point estimate and confidence interval (bootstrapping) of the expected genotype-outcome association. ${ }^{51}$ In the 23 studies which employed this approach using more than one study population, three only took into account the error in the exposure-outcome association and not the error in the genotype-exposure association, ${ }^{159,164,165}$ and 16 studies did not describe how the confidence interval was obtained.

## Discussion

Most Mendelian randomization studies either performed some form of IV analysis ( $49 \%$ of studies within one study population and $65 \%$ of studies within multiple study
populations) or used the genotype as a proxy for the exposure without further estimation. A third approach used less frequently was to compare the observed genotype-outcome association with the expected genotype-outcome association. Although validity of the three main Mendelian randomization assumptions is required regardless of the approach used, only $44 \%$ of studies adequately discussed the plausibility of these assumptions. The methods used to obtain an IV estimate were not always adequately described. For those studies which were performed using multiple study populations, the range of approaches used was very broad, because of further differentiation according to the way the data from the different studies were combined. Here we will discuss our findings and propose recommendations for the reporting of Mendelian randomization studies.

To our knowledge there is one paper which previously reviewed MR studies, which included a much smaller number of studies. Its main focus was on whether the Mendelian randomization studies reported results that were compatible with a causal association, which was the case for over half of their reviewed studies. ${ }^{189}$ In contrast, our review focused on the approach used and on the discussion of the assumptions and the reporting of the statistical methods used. The previous review also noted that many studies applied IV analysis to a binary outcome, using methods which had not quite been validated, ${ }^{189}$ which is an issue which we will also discuss later.

Our meta-epidemiological study has several limitations. With respect to study selection, we investigated what
methods were used in studies stating that they used Mendelian randomization or that they used a genetic IV. Importantly, we were unable to include studies which apply the same principles without using the term Mendelian randomization or genetic IV because these could not feasibly be found using a systematic search strategy. We do not know to what extent our results apply to these studies, but suspect the discussion of Mendelian randomization assumptions in particular is likely to be insufficient in many of these studies. Importantly, the focus of our review was on the quality of reporting of methods used in Mendelian randomization studies. We did not assess whether the statistical method used to obtain an IV estimate was actually appropriate. We investigated whether the statistical method used was adequately described, whether it was consistent with the estimates reported and if any evident mistakes were made. Similarly, we focused on whether plausibility of MR assumptions was discussed, not on whether we considered them likely to hold.

With regard to the Mendelian randomization approach used, we found that a majority of studies performed some form of IV analysis, but a substantial proportion of studies used the genotype as a proxy for the exposure without performing a formal IV analysis. This raises the question whether either of these approaches, or the third option of comparing the observed and expected genotype-outcome association, should be preferred. This depends on the aim of the study: for a test of causality, testing the presence of a genotype-outcome association is sufficient. ${ }^{1,190}$ Often the aim will be a quantification of the causal effect of the exposure on the outcome. We note that IV analysis is more suited to this aim than a comparison of the observed and the expected effect of the genotype on the outcome, although some may find the latter approach more intuitive. Showing the association between the genotype (or genetic score) and the outcome is always advisable as it increases the transparency of the study by showing the data as they are. Further analyses can subsequently be undertaken. ${ }^{191}$ When considering whether a formal IV analysis is appropriate, further aspects of the underlying biology of the genotype-phenotype association need to be taken into account to avoid misleading inferences. ${ }^{192}$ A recent paper discusses a number of situations in which a formal IV analysis may give biased results, but a Mendelian randomization approach looking only at the genotype-outcome association can validly be used as a test of causality. ${ }^{190}$ Another recent paper specifically discusses smoking as an example of an exposure for which the measurement does not fully capture the underlying exposure, which gives a biased estimate of the effect of the measured exposure on the outcome if an IV analysis is performed in a Mendelian randomization study. ${ }^{4}$

With regard to the discussion of the Mendelian randomization assumptions, we found that fewer than half of studies adequately discussed all three assumptions. Some studies did mention what the assumptions are and how they can be violated in general terms, but did not discuss how plausible the assumptions were for the specific setting of their study. An aspect of the assumptions which can be evaluated using the data is whether there is an association between the genetic instrument and measured confounders. This may be more difficult for studies which use multiple study populations, but an effort to obtain this information from those studies in which it is available is warranted. Among the studies which performed an IV analysis in a single study population, we identified 14 studies in which SNPs were detected or selected, or genetic risk score weights were derived in that same study population. This can bias Mendelian randomization estimates. ${ }^{116,193}$ The number of studies in which we found this to have occurred may be an underestimation, because some study populations are used for multiple Mendelian randomization studies and the later studies may not report the detection of SNPs in a previous study in the same population.

With regard to the IV methods used, we found that two-stage least squares regression and a Wald-type/ratio estimator were the most commonly used methods. We also found that a considerable number of the Mendelian randomization studies which used IV methods estimated an odds ratio or risk ratio, especially in those studies which used data from multiple study populations. However, which methods are appropriate for IV estimation of causal odds ratios or risk ratios is a methodological challenge of IV analysis that has not yet been fully resolved. Several methodological studies have investigated this issue in recent years. ${ }^{194,194-198}$ One of the reviewed MR studies mentioned that the Wald-type estimator used to estimate an odds ratio was an approximate method. ${ }^{169}$ The properties and limitations of these IV methods used to estimate a causal odds ratio deserve more attention in the Mendelian randomization studies in which they are used.

Overall, we conclude from our review the standard of reporting of Mendelian randomization studies should be improved. Existing guidelines and recommendations for the reporting of IV analyses largely apply to Mendelian randomization studies (the extent depending on the Mendelian randomization approach used). ${ }^{7,8}$ In addition to these recommendations, we have formulated a checklist of Mendelian randomization-specific reporting recommendations in Box 1.

In conclusion, studies stating that they perform a Mendelian randomization study within one study population broadly fall into three categories: studies using a

## Box 1. Proposed checklist for reporting Mendelian randomization studies

## Methods

- If an expected genotype-outcome association is calculated, report how this was calculated and how the confidence interval was obtained. Take into account the variance of both the genotype-exposure and the exposure-outcome association.
- If an instrumental variable analysis is performed, report in detail which method was used and how the confidence interval was obtained. For non-standard instrumental variable methods (e.g. methods used to estimate an odds ratio), discuss the properties of these methods.
- If data from multiple populations are used, clearly explain how and at what stage the data/estimates were combined.

Results

- Report the strength of the association between the genetic instrument and the exposure, using a partial F-statistic if possible.
- Show the association between the genetic instrument and measured confounders. If multiple study populations are used, show this for those populations for which this information is available.
- Report the association of the genotype and the outcome.


## Discussion

- Discuss the plausibility of the second and third instrumental variable assumptions in the specific setting of the study: could pleiotropy, linkage disequilibrium, canalization, population stratification or unmeasured confounding of the genotype-outcome relation affect results in this study?
genotype as a proxy for exposure without further estimation; studies performing IV analysis using a genotype as an instrument; and studies comparing observed and expected genotype-outcome associations. Plausibility of underlying Mendelian randomization assumptions is not always discussed, but as these assumptions are crucial for validity of MR studies, they should always be discussed in the specific context of the study. If IV methods are used to estimate a causal effect of the exposure, the statistical methods used should be clearly explained. Studies using data from multiple populations should also clearly report how data or estimates are combined.


## Supplementary Data

Supplementary data are available at $I J E$ online.

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## References

1. Lawlor DA, Harbord RM, Sterne JA, Timpson N, Davey Smith G. Mendelian randomization: using genes as instruments for making causal inferences in epidemiology. Stat Med 2008;27: 1133-63.
2. Davey Smith G, Ebrahim S. 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease? Int J Epidemiol 2003;32:1-22.
3. Katan MB. Apolipoprotein $E$ isoforms, serum cholesterol, and cancer. Lancet 1986;1:507-08.
4. Taylor AE, Davies NM, Ware JJ, VanderWeele T, Davey Smith G, Munafo MR. Mendelian randomization in health research: using appropriate genetic variants and avoiding biased estimates. Econ Hum Biol 2014;13:99-106.
5. Sheehan NA, Meng S, Didelez V. Mendelian randomization: a tool for assessing causality in observational epidemiology. Methods Mol Biol 2011;713:153-66.
6. Didelez V, Meng S, Sheehan NA. Assumptions of IV methods for observational epidemiology. Stat Sci 2010;25:22-40.
7. Davies NM, Davey Smith G, Windmeijer F, Martin RM. Issues in the reporting and conduct of instrumental variable studies: a systematic review. Epidemiology 2013;24:363-69.
8. Swanson SA, Hernán MA. Commentary: how to report instrumental variable analyses (suggestions welcome). Epidemiology 2013;24:370-74.
9. Martens EP, Pestman WR, de Boer A, Belitser SV, Klungel OH. Instrumental variables: application and limitations. Epidemiology 2006;17:260-67.
10. Sharma NK, Gupta A, Prabhakar S et al. Association between CFH Y402H polymorphism and age related macular degeneration in North Indian cohort. PLoS One 2013;8:e70193.
11. Shaheen SO, Rutterford C, Zuccolo L et al. Prenatal alcohol exposure and childhood atopic disease: a Mendelian randomization approach. J Allergy Clin Immunol 2014;133:225-32.
12. Humphriss R, Hall A, May M, Zuccolo L, Macleod J. Prenatal alcohol exposure and childhood balance ability: findings from a UK birth cohort study. BMJ Open 2013;3(6). doi: 10.1136/ bmjopen-2013-002716.
13. Bonilla C, Lawlor DA, Taylor AE et al. Vitamin B-12 status during pregnancy and child's IQ at age 8: a Mendelian randomization study in the Avon longitudinal study of parents and children. PLoS One 2012;7:e51084.
14. Bonilla C, Lawlor DA, Ben-Shlomo Y et al. Maternal and offspring fasting glucose and type 2 diabetes-associated genetic variants and cognitive function at age 8: a Mendelian randomization study in the Avon Longitudinal Study of Parents and Children. BMC Med Genet 2012;13:90.
15. Bonilla C, Gilbert R, Kemp JP et al. Using genetic proxies for lifecourse sun exposure to assess the causal relationship of sun exposure with circulating vitamin D and prostate cancer risk. Cancer Epidemiol Biomarkers Prev 2013;22:597-606.
16. Alegret JM, Aragonès G, Elosua R et al. The relevance of the association between inflammation and atrial fibrillation. Eur J Clin Invest 2013;43:324-31.
17. Almon R, Álvarez-León EE, Serra-Majem L. Association of the European lactase persistence variant (LCT-13910 C $>$ T polymorphism) with obesity in the Canary Islands. PLoS One 2012; 7:e43978.
18. Bjørngaard JH, Gunnell D, Elvestad MB et al. The causal role of smoking in anxiety and depression: a Mendelian randomization analysis of the HUNT study. Psychol Med 2013;43: 711-19.
19. Attermann J, Obel C, Bilenberg N, Nordenbæk CM, Skytthe A, Olsen J. Traits of ADHD and autism in girls with a twin brother: a Mendelian randomization study. Eur Child Adolesc Psychiatry 2012;21:503-09.
20. Yang Q, Bailey L, Clarke Ret al. Prospective study of methylenetetrahydrofolate reductase (MTHFR) variant C677T and risk of all-cause and cardiovascular disease mortality among 6000 US adults. Am J Clin Nutr 2012;95:1245-53.
21. van Durme YM, Lahousse L, Verhamme KM et al. Mendelian randomization study of interleukin-6 in chronic obstructive pulmonary disease. Respiration 2011;82:530-38.
22. Scott JA, Berkley JA, Mwangi I et al. Relation between falciparum malaria and bacteraemia in Kenyan children: a popula-tion-based, case-control study and a longitudinal study. Lancet 2011;378:1316-23.
23. Chmielewski M, Verduijn M, Drechsler C et al. Low cholesterol in dialysis patients - causal factor for mortality or an effect of confounding? Nephrol Dial Transplant 2011;26: 3325-31.
24. Bolton CE, Schumacher W, Cockcroft JR et al. The CRP genotype, serum levels and lung function in men: the Caerphilly Prospective Study. Clin Sci (Lond) 2011;120:347-55.
25. Kröger J, Zietemann V, Enzenbach C et al. Erythrocyte membrane phospholipid fatty acids, desaturase activity, and dietary
fatty acids in relation to risk of type 2 diabetes in the European Prospective Investigation into Cancer and Nutrition (EPIC)Potsdam Study. Am J Clin Nutr 2011;93:127-42.
26. Welsh P, Polisecki E, Robertson M et al. Unraveling the directional link between adiposity and inflammation: a bidirectional Mendelian randomization approach. J Clin Endocrinol Metab 2010;95:93-99.
27. Trompet S, Jukema JW, Katan MB et al. Apolipoprotein e genotype, plasma cholesterol, and cancer: a Mendelian randomization study. Am J Epidemiol 2009;170:1415-21.
28. Almon R, Álvarez-León EE, Engfeldt P, Serra-Majem L, Magnuson A, Nilsson TK. Associations between lactase persistence and the metabolic syndrome in a cross-sectional study in the Canary Islands. Eur J Nutr 2010;49:141-46.
29. Brennan P, McKay J, Moore L et al. Obesity and cancer: Mendelian randomization approach utilizing the FTO genotype. Int J Epidemiol 2009;38:971-75.
30. Drenos F, Talmud PJ, Casas JP et al. Integrated associations of genotypes with multiple blood biomarkers linked to coronary heart disease risk. Hum Mol Genet 2009;18:2305-16.
31. Lim LS, Tai ES, Aung T et al. Relation of age-related cataract with obesity and obesity genes in an Asian population. Am J Epidemiol 2009;169:1267-74.
32. Giltay EJ, van Reedt Dortland AK, Nissinen A et al. Serum cholesterol, apolipoprotein E genotype and depressive symptoms in elderly European men: the FINE study. J Affect Disord 2009;115:471-77.
33. Irons DE, McGue M, Iacono WG, Oetting WS. Mendelian randomization: a novel test of the gateway hypothesis and models of gene-environment interplay. Dev Psychopathol 2007;19: 1181-95.
34. Herder C, Klopp N, Baumert J et al. Effect of macrophage migration inhibitory factor (MIF) gene variants and MIF serum concentrations on the risk of type 2 diabetes: results from the MONICA/KORA Augsburg Case-Cohort Study, 1984-2002. Diabetologia 2008;51:276-84.
35. Keavney B, Danesh J, Parish S et al. Fibrinogen and coronary heart disease: test of causality by 'Mendelian randomization'. Int J Epidemiol 2006;35:935-43.
36. Bech BH, Autrup H, Nohr EA, Henriksen TB, Olsen J. Stillbirth and slow metabolizers of caffeine: comparison by genotypes. Int J Epidemiol 2006;35:948-53.
37. Zuccolo L, Lewis SJ, Davey Smith G et al. Prenatal alcohol exposure and offspring cognition and school performance. A 'Mendelian randomization' natural experiment. Int $J$ Epidemiol 2013;42:1358-70.
38. Aramini B, Kim C, Diangelo S et al. Donor surfactant protein D (SP-D) polymorphisms are associated with lung transplant outcome. Am J Transplant 2013;13:2130-36.
39. Yarwood A, Martin P, Bowes J et al. Enrichment of vitamin D response elements in RA-associated loci supports a role for vitamin D in the pathogenesis of RA. Genes Immun 2013;14:325-29.
40. Travis RC, Appleby PN, Siddiq A et al. Genetic variation in the lactase gene, dairy product intake and risk for prostate cancer in the European prospective investigation into cancer and nutrition. Int J Cancer2013;132:1901-10.
41. Veen G, Giltay EJ, Van Vliet IM et al. C-reactive protein polymorphisms are associated with the cortisol awakening
response in basal conditions in human subjects. Stress 2011;14: 128-35.
42. Koshy B, Miyashita A, St Jean P et al. Genetic deficiency of plasma lipoprotein-associated phospholipase A 2 (PLA2G7 V297F null mutation) and risk of Alzheimer's disease in Japan. J Alzheimer's Dis 2010;21:775-80.
43. Pierce BL, Ahsan H. Genetic susceptibility to type 2 diabetes is associated with reduced prostate cancer risk. Hum Hered 2010; 69:193-201.
44. Thuesen BH, Husemoen LLN, Fenger M, Linneberg A. Lack of association between the MTHFR (C677T) polymorphism and atopic disease. Clin Resp J 2009;3:102-08.
45. Granell R, Heron J, Lewis S, Davey Smith G, Sterne JAC, Henderson J. The association between mother and child MTHFR C677T polymorphisms, dietary folate intake and childhood atopy in a population-based, longitudinal birth cohort. Clin Exp Allergy 2008;38:320-28.
46. Heidrich J, Wellmann J, Doring A, Illig T, Keil U. Alcohol consumption, alcohol dehydrogenase and risk of coronary heart disease in the MONICA/KORA-Augsburg cohort 1994/1995-2002. Eur J Cardiovasc Prev Rehabil 2007;14: 769-74.
47. Almeida OP, Hankey GJ, Yeap BB, Golledge J, Flicker L. The triangular association of ADH1B genetic polymorphism, alcohol consumption and the risk of depression in older men. Mol Psychiatry 2014;19:995-1000.
48. Dai X, Yuan J, Yao P et al. Association between serum uric acid and the metabolic syndrome among a middle- and old-age Chinese population. Eur J Epidemiol 2013;28:66976.
49. Yao WM, Zhang HF, Zhu ZY et al. Genetically elevated levels of circulating triglycerides and brachial-ankle pulse wave velocity in a Chinese population. J Hum Hypertens 2013;27:265-70.
50. Gan W, Guan Y, Wu Q et al. Association of TMPRSS6 polymorphisms with ferritin, hemoglobin, and type 2 diabetes risk in a Chinese Han population. Am J Clin Nutr 2012;95:626-32.
51. Breitling LP, Koenig W, Fischer M et al. Type II secretory phospholipase A2 and prognosis in patients with stable coronary heart disease: mendelian randomization study. PLoS One 2011;6:e22318.
52. Rasmussen-Torvik LJ, Li M, Kao WH et al. Association of a fasting glucose genetic risk score with subclinical atherosclerosis: The Atherosclerosis Risk in Communities (ARIC) study. Diabetes 2011;60:331-35.
53. Wu Y, Li H, Loos RJ et al. RBP4 variants are significantly associated with plasma RBP4 levels and hypertriglyceridemia risk in Chinese Hans. J Lipid Res 2009;50:1479-86.
54. Di Paola R, Marucci A, Fontana A et al. Role of obesity on allcause mortality in whites with type 2 diabetes from Italy. Acta Diabetol 2013;50:971-76.
55. Menzaghi C, De Cosmo S, Copetti M et al. Relationship between ADIPOQ gene, circulating high molecular weight adiponectin and albuminuria in individuals with normal kidney function: Evidence from a family-based study. Diabetologia 2011;54:812-18.
56. Oei L, Campos-Obando N, Dehghan A et al. Dissecting the relationship between high-sensitivity serum C-reactive protein and increased fracture risk: the Rotterdam Study. Osteoporos Int 2014;25:1247-54.
57. Tian Q, Jia J, Ling S, Liu Y, Yang S, Shao Z. A causal role for circulating miR-34b in osteosarcoma. Eur J Surg Oncol 2014; 40:67-72.
58. Bouthoorn SH, van Lenthe FJ, Kiefte-de Jong JC et al. Genetic taste blindness to bitter and body composition in childhood: a Mendelian randomization design. Int J Obes (Lond) 2014;38: 1005-10.
59. Lee HA, Park EA, Cho SJ et al. Mendelian randomization analysis of the effect of maternal homocysteine during pregnancy, as represented by maternal MTHFR C677T genotype, on birth weight. J Epidemiol 2013;23:371-75.
60. Warodomwichit D, Sritara C, Thakkinstian A et al. Causal inference of the effect of adiposity on bone mineral density in adults. Clin Endocrinol (Oxf) 2013;78:694-99.
61. Jensen MK, Bartz TM, Djousse L et al. Genetically elevated fetuin-A levels, fasting glucose levels, and risk of type 2 diabetes: the cardiovascular health study. Diabetes Care 2013;36: 3121-27.
62. Gao H, Fall T, van Dam RM et al. Evidence of a causal relationship between adiponectin levels and insulin sensitivity: a Mendelian randomization study. Diabetes 2013; 62:1338-44.
63. Alwan NA, Lawlor DA, McArdle HJ, Greenwood DC, Cade JE. Exploring the relationship between maternal iron status and offspring's blood pressure and adiposity: a Mendelian randomization study. Clin Epidemiol 2012;4:193-200.
64. Oikonen M, Wendelin-Saarenhovi M, Lyytikäinen LP et al. Associations between serum uric acid and markers of subclinical atherosclerosis in young adults. The cardiovascular risk in Young Finns study. Atherosclerosis 2012;223:497-503.
65. Lyngdoh T, Vuistiner P, Marques-Vidal P et al. Serum uric acid and adiposity: deciphering causality using a bidirectional Mendelian randomization approach. PLoS One 2012;7: e39321.
66. Guessous I, Dobrinas M, Kutalik Z et al. Caffeine intake and CYP1A2 variants associated with high caffeine intake protect non-smokers from hypertension. Hum Mol Genet 2012;21: 3283-92.
67. Au Yeung SL, Jiang C, Cheng KK et al. Moderate alcohol use and cardiovascular disease from Mendelian randomization. PLoS One 2013;8:e68054.
68. Au Yeung SL, Jiang CQ, Cheng KK et al. Evaluation of moderate alcohol use and cognitive function among men using a Mendelian randomization design in the Guangzhou biobank cohort study. Am J Epidemiol 2012;175:1021-28.
69. Lewis SJ, Araya R, Davey Smith G et al. Smoking is associated with, but does not cause, depressed mood in pregnancy - a mendelian randomization study. PLoS One 2011;6:e21689.
70. Conen D, Vollenweider P, Rousson Vet al. Use of a Mendelian randomization approach to assess the causal relation of gamma-Glutamyltransferase with blood pressure and serum insulin levels. Am J Epidemiol 2010;172:1431-41.
71. Bochud M, Marquant F, Marques-Vidal PM et al. Association between C-reactive protein and adiposity in women. J Clin Endocrinol Metab 2009;94:3969-77.
72. Timpson NJ, Sayers A, Davey Smith G, Tobias JH. How does body fat influence bone mass in childhood? A Mendelian randomization approach. J Bone Miner Res 2009;24:522-33.
73. Kivimäki M, Lawlor DA, Davey Smith G et al. Does high Creactive protein concentration increase atherosclerosis? The Whitehall II Study. PLoS One 2008;3:e3013.
74. Brunner EJ, Kivimäki M, Witte DR et al. Inflammation, insulin resistance, and diabetes - Mendelian randomization using CRP haplotypes points upstream. PLoS Med 2008;5:e155.
75. Lawlor DA, Timpson NJ, Harbord RM et al. Exploring the developmental overnutrition hypothesis using parental-offspring associations and FTO as an instrumental variable. PLoS Med 2008;5:e33.
76. Viikari LA, Huupponen RK, Viikari JS et al. Relationship between leptin and C-reactive protein in young Finnish adults. J Clin Endocrinol Metab 2007;92:4753-58.
77. Timpson NJ, Lawlor DA, Harbord RM et al. C-reactive protein and its role in metabolic syndrome: mendelian randomization study. Lancet 2005;366:1954-59.
78. Binder AM, Michels KB. The causal effect of red blood cell folate on genome-wide methylation in cord blood: a Mendelian randomization approach. BMC Bioinformatics 2013;14:353.
79. Debette S, Wolf C, Lambert JC et al. Abdominal obesity and lower gray matter volume: a Mendelian randomization study. Neurobiol Aging 2014;35:378-86.
80. Thakkinstian A, Chailurkit L, Warodomwichit D et al. Causal relationship between body mass index and fetuin-A level in the asian population: a bidirectional mendelian randomization study. Clin Endocrinol (Oxf) 2014;81:197-203.
81. Haring R, Teumer A, Völker U et al. Mendelian randomization suggests non-causal associations of testosterone with cardiometabolic risk factors and mortality. Andrology 2013;1:17-23.
82. Islam M, Jafar TH, Wood AR et al. Multiple genetic variants explain measurable variance in type 2 diabetes-related traits in Pakistanis. Diabetologia 2012;55:2193-204.
83. Kivimäki M, Jokela M, Hamer M et al. Examining overweight and obesity as risk factors for common mental disorders using fat mass and obesity-associated (FTO) genotype-instrumented analysis: The Whitehall II Study, 1985-2004. Am J Epidemiol 2011;173:421-29.
84. Jokela M, Elovainio M, Keltikangas-Järvinen L et al. Body mass index and depressive symptoms: instrumental-variables regression with genetic risk score. Genes Brain Behav 2012; Sep 7. doi: 10.1111/j.1601-183X.2012.00846.x. [Epub ahead of print.]
85. Kivimäki M, Magnussen CG, Juonala M et al. Conventional and Mendelian randomization analyses suggest no association between lipoprotein(a) and early atherosclerosis: the Young Finns Study. Int J Epidemiol 2011;40:470-78.
86. Kivimäki M, Davey Smith G, Timpson NJ et al. Lifetime body mass index and later atherosclerosis risk in young adults: examining causal links using Mendelian randomization in the Cardiovascular Risk in Young Finns study. Eur Heart J 2008; 29:2552-60.
87. Frayling TM, Rafiq S, Murray A et al. An interleukin-18 polymorphism is associated with reduced serum concentrations and better physical functioning in older people. J Gerontol A Biol Sci Med Sci 2007;62:73-78.
88. Davey Smith G, Lawlor DA, Harbord R et al. Association of C-reactive protein with blood pressure and hypertension: life
course confounding and mendelian randomization tests of causality. Arterioscler Thromb Vasc Biol 2005;25:1051-56.
89. Lawlor DA, Nordestgaard BG, Benn M, Zuccolo L, TybjærgHansen A, Davey Smith G. Exploring causal associations between alcohol and coronary heart disease risk factors: findings from a Mendelian randomization study in the Copenhagen General Population Study. Eur Heart J 2013;34: 2519-28.
90. Theodoratou E, Palmer T, Zgaga L et al. Instrumental variable estimation of the causal effect of plasma 25-hydroxy-vitamin D on colorectal cancer risk: a mendelian randomization analysis. PLoS One 2012;7:e37662.
91. Collin SM, Metcalfe C, Palmer TM et al. The causal roles of vitamin $\mathrm{B}(12)$ and transcobalamin in prostate cancer: can Mendelian randomization analysis provide definitive answers? Int J Mol Epidemiol Genet 2011;2:316-27.
92. Lawlor DA, Harbord RM, Tybjærg-Hansen A et al. Using genetic loci to understand the relationship between adiposity and psychological distress: a Mendelian Randomization study in the Copenhagen General Population Study of 53221 adults. J Intern Med 2011;269:525-37.
93. Mumby HS, Elks CE, Li S et al. Mendelian randomization study of childhood BMI and early menarche. J Obes 2011; 2011:180729.
94. Timpson NJ, Nordestgaard BG, Harbord RM et al. C-reactive protein levels and body mass index: elucidating direction of causation through reciprocal Mendelian randomization. Int J Obes (Lond) 2011;35:300-08.
95. Zacho J, Tybjærg-Hansen A, Nordestgaard BG. C-reactive protein and all-cause mortality - the Copenhagen City Heart Study. Eur Heart J 2010;31:1624-32.
96. Verduijn M, Prein RA, Stenvinkel P et al. Is fetuin-A a mortality risk factor in dialysis patients or a mere risk marker? A Mendelian randomization approach. Nephrol Dial Transplant 2011;26:239-45.
97. Fisher E, Stefan N, Saar K et al. Association of AHSG gene polymorphisms with fetuin-A plasma levels and cardiovascular diseases in the EPIC-Potsdam study. Circ Cardiovasc Genet 2009;2:607-13.
98. Timpson NJ, Harbord R, Davey Smith G, Zacho J, TybjærgHansen A, Nordestgaard BG. Does greater adiposity increase blood pressure and hypertension risk?: Mendelian randomization using the FTO/MC4R genotype. Hypertension 2009;54: 84-90.
99. Kang H, Kreuels B, Adjei O, Krumkamp R, May J, Small DS. The causal effect of malaria on stunting: a Mendelian randomization and matching approach. Int J Epidemiol 2013;42: 1390-98.
100. Qin XY, Tian J, Fang K et al. [Mendelian randomization study of the relationship between high-density lipoprotein cholesterol and age-related macular degeneration]. Beijing Da Xue Xue Bao 2012;44:407-11.
101. Trummer O, Pilz S, Hoffmann MM et al. Vitamin D and mortality: a Mendelian randomization study. Clin Chem 2013;59: 793-97.
102. You NC, Chen BH, Song Y et al. A prospective study of leukocyte telomere length and risk of type 2 diabetes in postmenopausal women. Diabetes 2012;61:2998-3004.
103. McArdle PF, Whitcomb BW, Tanner K, Mitchell BD, Shuldiner AR, Parsa A. Association between bilirubin and cardiovascular disease risk factors: using Mendelian randomization to assess causal inference. BMC Cardiovasc Disord 2012;12:16.
104. Parsa A, Brown E, Weir MR et al. Genotype-based changes in serum uric acid affect blood pressure. Kidney Int 2012;81: 502-07.
105. Sunyer J, Pistelli R, Plana E et al. Systemic inflammation, genetic susceptibility and lung function. Eur Respir J 2008;32: 92-97.
106. Mente A, Meyre D, Lanktree MB et al. Causal relationship between adiponectin and metabolic traits: a Mendelian randomization study in a multiethnic population. PLoS One 2013;8: e66808.
107. Love-Gregory L, Sherva R, Schappe T et al. Common CD36 SNPs reduce protein expression and may contribute to a protective atherogenic profile. Hum Mol Genet 2011;20:193-201.
108. Nagele P, Zeugswetter B, Wiener C et al. Influence of methylenetetrahydrofolate reductase gene polymorphisms on homocysteine concentrations after nitrous oxide anesthesia. Anesthesiology 2008;109:36-43.
109. Klovaite J, Nordestgaard BG, Tybjærg-Hansen A, Benn M. Elevated fibrinogen levels are associated with risk of pulmonary embolism, but not with deep venous thrombosis. Am J Respir Crit Care Med 2013;187:286-93.
110. Rius-Ottenheim N, de Craen AJ, Geleijnse JM et al. C-reactive protein haplotypes and dispositional optimism in obese and nonobese elderly subjects. Inflamm Res 2012;61:43-51.
111. Stender S, Frikke-Schmidt R, Nordestgaard BG, TybjærgHansen A. Extreme bilirubin levels as a causal risk factor for symptomatic gallstone disease. JAMA Intern Med 2013;173: 1222-28.
112. Dahl M, Vestbo J, Zacho J, Lange P, Tybjærg-Hansen A, Nordestgaard BG. C reactive protein and chronic obstructive pulmonary disease: a Mendelian randomization approach. Thorax 2011;66:197-204.
113. Marott SC, Nordestgaard BG, Zacho J et al. Does elevated C-reactive protein increase atrial fibrillation risk? A Mendelian randomization of 47000 individuals from the general population. J Am Coll Cardiol 2010;56:789-95.
114. Skaaby T, Husemoen LL, Martinussen T et al. Vitamin D status, filaggrin genotype, and cardiovascular risk factors: a Mendelian randomization approach. PLoS One 2013;8: e57647.
115. Cruchaga C, Kauwe JS, Nowotny P et al. Cerebrospinal fluid APOE levels: an endophenotype for genetic studies for Alzheimer's disease. Hum Mol Genet 2012;21:4558-71.
116. Hughes K, Flynn T, de Zoysa J, Dalbeth N, Merriman TR. Mendelian randomization analysis associates increased serum urate, due to genetic variation in uric acid transporters, with improved renal function. Kidney Int 2014;85:344-51.
117. Benn M, Tybjærg-Hansen A, McCarthy MI, Jensen GB, Grande P, Nordestgaard BG. Nonfasting glucose, ischemic heart disease, and myocardial infarction: a Mendelian randomization study. J Am Coll Cardiol 2012;59:2356-65.
118. Varbo A, Benn M, Tybjærg-Hansen A, Nordestgaard BG. Elevated remnant cholesterol causes both low-grade inflammation and ischemic heart disease, whereas elevated low-density
lipoprotein cholesterol causes ischemic heart disease without inflammation. Circulation 2013;128:1298-309.
119. Stender S, Nordestgaard BG, Tybjærg-Hansen A. Elevated body mass index as a causal risk factor for symptomatic gallstone disease: a Mendelian randomization study. Hepatology 2013;58:2133-41.
120. Kamstrup PR, Nordestgaard BG. Lipoprotein(a) concentrations, isoform size, and risk of type 2 diabetes: a Mendelian randomization study. Lancet Diabet Endocrinol 2013;1:220-27.
121. Wium-Andersen MK, Orsted DD, Nordestgaard BG. Elevated C-reactive protein associated with late- and very-late-onset schizophrenia in the general population: a prospective study. Schizophr Bull 2014;40:1117-27.
122. Heikkilä K, Silander K, Salomaa V et al. C-reactive proteinassociated genetic variants and cancer risk: findings from FINRISK 1992, FINRISK 1997 and Health 2000 studies. Eur J Cancer 2011;47:404-12.
123. Allin KH, Nordestgaard BG, Zacho J, Tybjærg-Hansen A, Bojesen SE. C-reactive protein and the risk of cancer: a mendelian randomization study. J Natl Cancer Inst 2010;102:202-06.
124. Kamstrup PR, Tybjærg-Hansen A, Nordestgaard BG. Elevated lipoprotein(a) and risk of aortic valve stenosis in the general population. J Am Coll Cardiol 2014;63:470-77.
125. Pierce BL, Tong L, Argos M et al. Arsenic metabolism efficiency has a causal role in arsenic toxicity: Mendelian randomization and gene-environment interaction. Int J Epidemiol 2013;42: 1862-71.
126. Oelsner EC, Pottinger TD, Burkart KM et al. Adhesion molecules, endothelin-1 and lung function in seven populationbased cohorts. Biomarkers 2013;18:196-203.
127. Trombetta M, Bonetti S, Boselli ML et al. PPARG2 Pro12Ala and ADAMTS9 rs4607103 as 'insulin resistance loci' and 'insulin secretion loci' in Italian individuals. The GENFIEV study and the Verona Newly Diagnosed Type 2 Diabetes Study (VNDS) 4. Acta Diabetol 2013;50:401-08.
128. Varbo A, Benn M, Tybjærg-Hansen A, Jørgensen AB, FrikkeSchmidt R, Nordestgaard BG. Remnant cholesterol as a causal risk factor for ischemic heart disease. J Am Coll Cardiol 2013; 61:427-36.
129. Jørgensen AB, Frikke-Schmidt R, West AS, Grande P, Nordestgaard BG, Tybjærg-Hansen A. Genetically elevated non-fasting triglycerides and calculated remnant cholesterol as causal risk factors for myocardial infarction. Eur Heart J 2013; 34:1826-33.
130. Haase CL, Tybjærg-Hansen A, Qayyum AA, Schou J, Nordestgaard BG, Frikke-Schmidt R. LCAT, HDL cholesterol and ischemic cardiovascular disease: a Mendelian randomization study of HDL cholesterol in 54500 individuals. J Clin Endocrinol Metab 2012;97:E248-E256.
131. Benn M, Tybjærg-Hansen A, Stender S, Frikke-Schmidt R, Nordestgaard BG. Low-density lipoprotein cholesterol and the risk of cancer: a mendelian randomization study. J Natl Cancer Inst 2011;103:508-19.
132. Zacho J, Tybjærg-Hansen A, Nordestgaard BG. C-reactive protein and risk of venous thromboembolism in the general population. Arterioscler Thromb Vasc Biol 2010;30:1672-78.
133. Wium-Andersen MK, Orsted DD, Nordestgaard BG. Elevated C-reactive protein, depression, somatic diseases, and all-cause
mortality: a mendelian randomization study. Biol Psychiatry 2014;76:249-57.
134. Stender S, Frikke-Schmidt R, Benn M, Nordestgaard BG, Tybjærg-Hansen A. Low-density lipoprotein cholesterol and risk of gallstone disease: a Mendelian randomization study and meta-analyses. J Hepatol 2013;58:126-33.
135. Yaghootkar H, Lamina C, Scott RA et al. Mendelian randomization studies do not support a causal role for reduced circulating adiponectin levels in insulin resistance and type 2 diabetes. Diabetes 2013;62:3589-98.
136. Chen L, Davey Smith G, Harbord RM, Lewis SJ. Alcohol intake and blood pressure: a systematic review implementing a Mendelian randomization approach. PLoS Med 2008;5:e52.
137. Shah S, Casas JP, Drenos F et al. Causal relevance of blood lipid fractions in the development of carotid atherosclerosis: Mendelian randomization analysis. Circ Cardiovasc Genet 2013;6:63-72.
138. Palmer TM, Nordestgaard BG, Benn M et al. Association of plasma uric acid with ischaemic heart disease and blood pressure: mendelian randomization analysis of two large cohorts. BMJ 2013;347:f4262.
139. De Silva NM, Freathy RM, Palmer TM et al. Mendelian randomization studies do not support a role for raised circulating triglyceride levels influencing type 2 diabetes, glucose levels, or insulin resistance. Diabetes 2011;60:1008-18.
140. Thanassoulis G, Campbell CY, Owens DS et al. Genetic associations with valvular calcification and aortic stenosis. N Engl J Med 2013;368:503-12.
141. Nordestgaard BG, Palmer TM, Benn M et al. The effect of elevated body mass index on ischemic heart disease risk: causal estimates from a Mendelian randomization approach. PLoS Med 2012;9:e1001212.
142. Elliott P, Chambers JC, Zhang Wet al. Genetic loci associated with C-reactive protein levels and risk of coronary heart disease. JAMA 2009;302:37-48.
143. Ye Z, Haycock PC, Gurdasani D et al. The association between circulating lipoprotein(a) and type 2 diabetes: is it causal? Diabetes 2014;63:332-42.
144. Voight BF, Peloso GM, Orho-Melander M et al. Plasma HDL cholesterol and risk of myocardial infarction: a mendelian randomization study. Lancet 2012;380:572-80.
145. Mamasoula C, Prentice RR, Pierscionek, et al. Association between C677T polymorphism of methylene tetrahydrofolate reductase and congenital heart disease: meta-analysis of 7697 cases and 13125 controls. Circ Cardiovasc Genet 2013;6: 347-53.
146. Harrison SC, Smith AJ, Jones GT et al. Interleukin-6 receptor pathways in abdominal aortic aneurysm. Eur Heart J 2013;34: 3707-16.
147. Stender S, Frikke-Schmidt R, Nordestgaard BG, Grande P, Tybjærg-Hansen A. Genetically elevated bilirubin and risk of ischaemic heart disease: three Mendelian randomization studies and a meta-analysis. J Intern Med 2013;273:59-68.
148. Hingorani AD, Casas JP. The interleukin-6 receptor as a target for prevention of coronary heart disease: a mendelian randomization analysis. Lancet 2012;379:1214-24.
149. Clarke R, Bennett DA, Parish S et al. Homocysteine and coronary heart disease: meta-analysis of MTHFR case-control
studies, avoiding publication bias. PLoS Med 2012;9: e1001177.
150. Wang J, Wang H, Chen Y, Hao P, Zhang Y. Alcohol ingestion and colorectal neoplasia: a meta-analysis based on a Mendelian randomization approach. Colorectal Dis 2011;13:e71-e78.
151. Casas JP, Ninio E, Panayiotou A et al. PLA2G7 genotype, lipo-protein-associated phospholipase A2 activity, and coronary heart disease risk in 10494 cases and 15624 controls of European ancestry. Circulation 2010;121:2284-93.
152. Lewis SJ, Baker I, Davey Smith G. Meta-analysis of vitamin D receptor polymorphisms and pulmonary tuberculosis risk. Int $J$ Tuberc Lung Dis 2005;9:1174-77.
153. Davies JR, Field S, Randerson-Moor J et al. An inherited variant in the gene coding for vitamin D-binding protein and survival from cutaneous melanoma: a BioGenoMEL study. Pigment Cell Melanoma Res 2014;27:234-43.
154. Panoutsopoulou K, Metrustry S, Doherty SA et al. The effect of FTO variation on increased osteoarthritis risk is mediated through body mass index: a mendelian randomization study. Ann Rheum Dis 2013;73:2082-86.
155. Ioannidis A, Ikonomi E, Dimou NL, Douma L, Bagos PG. Polymorphisms of the insulin receptor and the insulin receptor substrates genes in polycystic ovary syndrome: a Mendelian randomization meta-analysis. Mol Genet Metab 2010;99: 174-83.
156. Rice NE, Bandinelli S, Corsi AM et al. The paraoxonase (PON1) Q192R polymorphism is not associated with poor health status or depression in the ELSA or INCHIANTI studies. Int J Epidemiol 2009;38:1374-79.
157. Rafiq S, Melzer D, Weedon MN et al. Gene variants influencing measures of inflammation or predisposing to autoimmune and inflammatory diseases are not associated with the risk of type 2 diabetes. Diabetologia 2008;51:2205-13.
158. Pfister R, Sharp S, Luben R et al. Mendelian randomization study of B-type natriuretic peptide and type 2 diabetes: evidence of causal association from population studies. PLoS Med 2011; 8:e1001112.
159. Pfister R, Barnes D, Luben Ret al. No evidence for a causal link between uric acid and type 2 diabetes: a Mendelian randomization approach. Diabetologia 2011;54:2561-69.
160. Marjot T, Yadav S, Hasan N, Bentley P, Sharma P. Genes associated with adult cerebral venous thrombosis. Stroke 2011;42: 913-18.
161. Sarwar N, Sandhu MS, Ricketts SL et al. Triglyceride-mediated pathways and coronary disease: collaborative analysis of 101 studies. Lancet 2010;375:1634-39.
162. Bentley P, Peck G, Smeeth L, Whittaker J, Sharma P. Causal relationship of susceptibility genes to ischemic stroke: comparison to ischemic heart disease and biochemical determinants. PLoS One 2010;5:e9136.
163. Perry JR, Ferrucci L, Bandinelli S et al. Circulating beta-carotene levels and type 2 diabetes - cause or effect? Diabetologia 2009;52:2117-21.
164. Perry JR, Weedon MN, Langenberg C et al. Genetic evidence that raised sex hormone binding globulin (SHBG) levels reduce the risk of type 2 diabetes. Hum Mol Genet 2010;19:535-44.
165. Boccia S, Hashibe M, Galli P et al. Aldehyde dehydrogenase 2 and head and neck cancer: a meta-analysis implementing a

Mendelian randomization approach. Cancer Epidemiol Biomarkers Prev 2009;18:248-54.
166. Casas JP, Bautista LE, Smeeth L, Sharma P, Hingorani AD. Homocysteine and stroke: evidence on a causal link from mendelian randomization. Lancet 2005;365:224-32.
167. Casas JP, Shah T, Cooper J et al. Insight into the nature of the CRP-coronary event association using Mendelian randomization. Int J Epidemiol 2006;35:922-31.
168. Fall T, Hagg S, Magi Ret al. The role of adiposity in cardiometabolic traits: a Mendelian randomization analysis. PLoS Med 2013;10:e1001474.
169. Pichler I, Del Greco MF, Gögele Met al. Serum iron levels and the risk of Parkinson disease: a Mendelian randomization study. PLoS Med 2013;10:e1001462.
170. Vimaleswaran KS, Berry DJ, Lu C et al. Causal relationship between obesity and vitamin D status: bi-directional Mendelian randomization analysis of multiple cohorts. PLoS Med 2013; 10:e1001383.
171. Ference BA, Yoo W, Alesh I et al. Effect of long-term exposure to lower low-density lipoprotein cholesterol beginning early in life on the risk of coronary heart disease: a Mendelian randomization analysis. J Am Coll Cardiol 2012;60:2631-39.
172. Niu W, Liu Y, Qi Y, Wu Z, Zhu D, Jin W. Association of inter-leukin-6 circulating levels with coronary artery disease: a metaanalysis implementing mendelian randomization approach. Int J Cardiol 2012;157:243-52.
173. Niu W, Zhang X, Qi Y. Association of an apolipoprotein E polymorphism with circulating cholesterols and hypertension: a meta-based Mendelian randomization analysis. Hypertens Res 2012;35:434-40.
174. Huang T, Ren J, Huang J, Li D. Association of homocysteine with type 2 diabetes: a meta-analysis implementing Mendelian randomization approach. BMC Genomics 2013;14:867.
175. Davey Smith G, Harbord R, Milton J, Ebrahim S, Sterne JAC. Does elevated plasma fibrinogen increase the risk of coronary heart disease? Evidence from a meta-analysis of genetic association studies. Arterioscler Thromb Vasc Biol 2005;25:2228-33.
176. Lawlor DA, Harbord RM, Timpson NJ et al. The association of C-reactive protein and CRP genotype with coronary heart disease: findings from five studies with 4610 cases amongst 18637 participants. PLoS One 2008;3:e3011.
177. Kamstrup PR, Tybjærg-Hansen A, Steffensen R, Nordestgaard BG. Genetically elevated lipoprotein $(\mathrm{a})$ and increased risk of myocardial infarction. JAMA 2009;301:2331-39.
178. Kamstrup PR, Tybjærg-Hansen A, Nordestgaard BG. Genetic evidence that lipoprotein(a) associates with atherosclerotic stenosis rather than venous thrombosis. Arterioscler Thromb Vasc Biol 2012;32:1732-41.
179. Stegeman BH, Helmerhorst FM, Vos HL, Rosendaal FR, Van Hylckama Vlieg A. Sex hormone-binding globulin levels are not causally related to venous thrombosis risk in women not using hormonal contraceptives. J Thromb Haemostat 2012;10:2061-67.
180. Adamsson Eryd S, Sjögren M, Smith JG et al. Ceruloplasmin and atrial fibrillation: evidence of causality from a population-based Mendelian randomization study. J Intern Med 2014;275:164-71.
181. Ding EL, Song Y, Manson JE et al. Sex hormone-binding globulin and risk of type 2 diabetes in women and men. N Engl J Med 2009;361:1152-63.
182. Husemoen LL, Skaaby T, Martinussen T et al. Investigating the causal effect of vitamin D on serum adiponectin using a mendelian randomization approach. Eur J Clin Nutr 2014;68: 189-95.
183. Song Y, Yeung E, Liu A et al. Pancreatic beta-cell function and type 2 diabetes risk: quantify the causal effect using a Mendelian randomization approach based on meta-analyses. Hum Mol Genet 2012;21:5010-18.
184. Wensley F, Gao P, Burgess Set al. Association between C reactive protein and coronary heart disease: mendelian randomization analysis based on individual participant data. BMJ 2011; 342:d548.
185. Zhao J, Jiang C, Lam TH et al. Genetically predicted testosterone and cardiovascular risk factors in men: a Mendelian randomization analysis in the Guangzhou Biobank Cohort Study. Int J Epidemiol 2014;43:140-48.
186. Holmes MV, Simon T, Exeter HJ et al. Secretory phospholipase $\mathrm{A}(2)$-IIA and cardiovascular disease: a mendelian randomization study. J Am Coll Cardiol 2013;62:1966-76.
187. Linsel-Nitschke P, Götz A, Erdmann J et al. Lifelong reduction of LDL-cholesterol related to a common variant in the LDLreceptor gene decreases the risk of coronary artery disease - a Mendelian Randomization study. PLoS One 2008;3:e2986.
188. Kunutsor SK, Burgess S, Munroe PB, Khan H. Vitamin D and high blood pressure: causal association or epiphenomenon? Eur J Epidemiol 2014;29:1-14.
189. Bochud M, Rousson V. Usefulness of Mendelian randomization in observational epidemiology. Int J Environ Res Public Health 2010;7:711-28.
190. VanderWeele TJ, Tchetgen Tchetgen EJ, Cornelis M, Kraft P. Methodological challenges in mendelian randomization. Epidemiology 2014;25:427-35.
191. Boef AG, Dekkers OM, le Cessie S, Vandenbroucke JP. Reporting instrumental variable analyses. Epidemiology 2013; 24:937-38.
192. Davey Smith G, Hemani G. Mendelian randomization: genetic anchors for causal inference in epidemiological studies. Hum Mol Genet 2014;23:R89-R98.
193. Burgess S, Thompson SG. Use of allele scores as instrumental variables for Mendelian randomization. Int J Epidemiol 2013; 42:1134-44.
194. Palmer TM, Sterne JA, Harbord RM et al. Instrumental variable estimation of causal risk ratios and causal odds ratios in Mendelian randomization analyses. Am J Epidemiol 2011;173: 1392-403.
195. Rassen JA, Schneeweiss S, Glynn RJ, Mittleman MA, Brookhart MA. Instrumental variable analysis for estimation of treatment effects with dichotomous outcomes. Am J Epidemiol 2009;169:273-84.
196. Vuistiner P, Bochud M, Rousson V. A comparison of three methods of Mendelian randomization when the genetic instrument, the risk factor and the outcome are all binary. PLoS One 2012;7:e35951.
197. Burgess S. Identifying the odds ratio estimated by a two-stage instrumental variable analysis with a logistic regression model. Stat Med 2013;32:4726-47.
198. Clarke PS, Windmeijer F. Instrumental variable estimators for binary outcomes. J Am Stat Assoc 2012;107:1638-52.

