



Merging of Opinions with Increasing Information

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MERGING OF OPINIONS WITH INCREASING INFORMATION¹

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1. History. One of us [1] has shown that if Z_n , $n = 1, 2, \dots$ is a stochastic process with D states, $0, 1, \dots, D - 1$ such that $X = \sum_{n=1}^{\infty} Z_n/D^n$ has an absolutely continuous distribution with respect to Lebesgue measure, then the conditional distribution of $R_k = \sum_{n=1}^{\infty} Z_{k+n}/D^n$ given Z_1, \dots, Z_k converges with probability one as $k \rightarrow \infty$ to the uniform distribution on the unit interval, in the sense that for each λ , $0 < \lambda \leq 1$, $P(R_k < \lambda \mid Z_1, \dots, Z_k) \rightarrow \lambda$ with probability 1 as $k \rightarrow \infty$. It follows that the unconditional distribution of R_k converges to the uniform distribution as $k \rightarrow \infty$. If $\{Z_n\}$ is stationary, the distribution of R_k is independent of k , and hence uniform, a result obtained earlier by Harris [3]. Earlier work relevant to convergence of opinion can be found in [4, Chap. 3, Sect. 6].

Here we generalize these results and also show that the conditional distribution of R_k given Z_1, \dots, Z_k converges in a much stronger sense. All probabilities in this paper are countably additive.

2. Statement of the theorem. Let \mathfrak{B}_i be a σ -field of subsets of a set X_i , $i = 1, 2, \dots$; and let $(X, \mathfrak{B}) = (X_1 \times X_2 \times \dots, \mathfrak{B}_1 \times \mathfrak{B}_2 \times \dots)$. Suppose (X, \mathfrak{B}, P) is a probability space and let P_n be the marginal distribution of $(X_1 \times \dots \times X_n, \mathfrak{B}_1 \times \dots \times \mathfrak{B}_n)$; that is, $P_n(A) = P(A \times X_{n+1} \times \dots)$ for all $A \in \mathfrak{B}_1 \times \dots \times \mathfrak{B}_n$. The probability P is *predictive* if for every $n \geq 1$, there exists a *conditional distribution* P^n for the *future* $X_{n+1} \times \dots$ given the *past* X_1, \dots, X_n ; that is, if there exists a function $P^n(x_1, \dots, x_n)(C)$ where (x_1, \dots, x_n) ranges over $X_1 \times \dots \times X_n$ and C ranges over $\mathfrak{B}_{n+1} \times \dots$ with the usual three properties: $P^n(x_1, \dots, x_n)(C)$ is $\mathfrak{B}_1 \times \dots \times \mathfrak{B}_n$ -measurable for fixed C ; a probability distribution on $(X_{n+1} \times \dots; \mathfrak{B}_{n+1} \times \dots)$ for fixed (x_1, \dots, x_n) ; and for bounded \mathfrak{B} -measurable ϕ

$$(1) \quad \int \phi dP = \int [(\phi(x_1, \dots, x_n, x_{n+1}, \dots)) dP^n(x_{n+1}, \dots \mid x_1, \dots, x_n)] \\ \cdot dP_n(x_1, \dots, x_n)$$

holds.

The assumption that P is predictive is mild and applies to all natural probabilities known to us. It is easy to verify that any probability which is absolutely continuous with respect to a predictive probability is also predictive.

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For any two probabilities μ_1 and μ_2 on the same σ -field \mathfrak{F} , the well known distance $\rho(\mu_1, \mu_2)$ between μ_1 and μ_2 is the least upper bound over $D \in \mathfrak{F}$ of $|\mu_1(D) - \mu_2(D)|$. Of course μ_i is absolutely continuous with respect to $(\mu_1 + \mu_2)/2 = m$ and has a density ϕ_i , so that $\rho(\mu_1, \mu_2) = \int_A (\phi_1 - \phi_2) dm = (1/2) \int |\phi_1 - \phi_2| dm$ where A is the set where $\phi_1 - \phi_2 > 0$.

MAIN THEOREM. *Suppose that P is a predictive probability on (X, \mathfrak{B}) and that Q is absolutely continuous with respect to P . Then for each conditional distribution P^n of the future given the past with respect to P , there exists a conditional distribution Q^n of the future given the past with respect to Q such that, with the exception of a set of histories $(x_1, \dots, x_n, x_{n+1}, \dots)$ of Q -probability 0, the distance between $P^n(x_1, \dots, x_n)$ and $Q^n(x_1, \dots, x_n)$ converges to 0 as n converges to ∞ .*

3. Martingale preliminaries. The proof of the theorem requires a slightly generalized martingale convergence theorem. Say that a sequence $\{y_n\}$ of random variables is *dominated in the sense of Lebesgue* if $\sup_n |y_n|$ has a finite expectation.

THEOREM 2. *Suppose that $\{y_n\}$, $n = 1, 2, \dots$, a sequence of random variables dominated in the sense of Lebesgue, converges almost everywhere to a random variable y . Then for every monotone increasing or monotone decreasing sequence of σ -fields \mathfrak{U}_j , $j = 1, 2, \dots$ converging to a σ -field \mathfrak{U} ,*

$$(2) \quad \lim_{\substack{j \rightarrow \infty \\ n \rightarrow \infty}} E[y_n | \mathfrak{U}_j] = E[y | \mathfrak{U}],$$

almost everywhere and in L_1 .

In this note we are primarily interested in the weaker conclusion that $\lim_{n \rightarrow \infty} E[y_n | \mathfrak{U}_n] = E[y | \mathfrak{U}]$. The two important special cases in which either y_n or \mathfrak{U}_n is independent of n are in [2].

PROOF OF THEOREM 2. Let $g_k = \sup y_n$ for $n \geq k$. Equalities and inequalities below are asserted to hold with probability 1. Fix k for a moment and let $n \geq k$. Then $y_n \leq g_k$ and

$$(3) \quad E[y_n | \mathfrak{U}_i] \leq E[g_k | \mathfrak{U}_i].$$

Letting

$$(4) \quad \begin{aligned} z &= \limsup_j E[y_n | \mathfrak{U}_i], \\ x &= \liminf_j E[y_n | \mathfrak{U}_i], \end{aligned}$$

you conclude from (3) and a usual form of martingale convergence theorem [For example, see 2, Theorem 4.3, Chap. VII] that

$$(5) \quad z \leq \lim_j \sup_{i \geq j} E[g_k | \mathfrak{U}_i] = \lim_i E[g_k | \mathfrak{U}_i] = E[g_k | \mathfrak{U}].$$

Therefore $z \leq \lim E[g_k | \mathfrak{U}] = E[y | \mathfrak{U}]$ by Lebesgue's theorem suitably generalized so as to apply to conditional expectations. [See, for example, 2, CE₆ Section 8, Chap. 1]. Similarly, $x \geq E[y | \mathfrak{U}]$, and the proof of almost everywhere convergence is complete. The proof of L_1 convergence is routine and omitted.

COROLLARY 1. Suppose that with probability 1, only a finite number of the events E_1, E_2, \dots occur. Then for any monotone sequence of σ -fields $\mathfrak{U}_1, \mathfrak{U}_2, \dots$

$$(6) \quad P\left[\bigcup_{k \geq n} E_k \mid \mathfrak{U}_j\right] \text{ and } P[E_n \mid \mathfrak{U}_j] \rightarrow 0.$$

almost surely as n and $j \rightarrow \infty$.

COROLLARY 2. If f_n is any sequence of random variables that converges almost everywhere to 0 and \mathfrak{U}_j is a monotone sequence of σ -fields, then with probability 1, for all $\epsilon > 0$,

$$(7) \quad P[\sup_{k \geq n} |f_k| > \epsilon \mid \mathfrak{U}_j], \text{ and } P[|f_n| > \epsilon \mid \mathfrak{U}_j]$$

converge to 0 as n and j converge to ∞ .

COROLLARY 3. Let $q \geq 0$ be a density function for which $Q(B) = \int_B q dP$ for all $B \in \mathfrak{G}$; let

$$(8) \quad q_n(x_1, \dots, x_n) = \int q(x_1, \dots, x_n, x_{n+1}, \dots) dP^n(x_{n+1}, \dots \mid x_1, \dots, x_n);$$

and let

$$(9) \quad d_n(x_1, \dots, x_n, x_{n+1}, \dots) = q(x_1, \dots, x_n, x_{n+1}, \dots) / q_n(x_1, \dots, x_n) \text{ or } 1,$$

according as $q_n(x_1, \dots, x_n) \neq 0$ or not. Then, with P -probability 1, for all $\epsilon > 0$,

$$(10) \quad P[d_n - 1 > \epsilon \mid x_1, \dots, x_n] \rightarrow 0 \text{ as } n \rightarrow \infty,$$

and with Q -probability 1, for all $\epsilon > 0$,

$$(11) \quad Q[|d_n - 1| > \epsilon \mid x_1, \dots, x_n] \rightarrow 0 \text{ as } n \rightarrow \infty.$$

PROOF OF COROLLARY 3. With respect to P measure,

$$(12) \quad E[q \mid x_1, \dots, x_n] = q_n(x_1, \dots, x_n),$$

so that according to Doob's martingale convergence theorem, $q_n(x_1, \dots, x_n)$ converges to $q(x_1, \dots, x_n, x_{n+1}, \dots)$ almost surely with respect to P . Consequently, $\limsup d_n \leq 1$ a.s. P and $d_n \rightarrow 1$ a.s. Q since $q > 0$ a.s. Q . An application of Corollary 2 completes the proof.

4. Proof of main theorem. Define

$$(13) \quad \begin{aligned} & Q^n(x_1, \dots, x_n)(C) \\ &= \int_C d_n(x_1, \dots, x_n, x_{n+1}, \dots) dP^n(x_{n+1}, \dots \mid x_1, \dots, x_n), \end{aligned}$$

for all $C \in \mathfrak{G}_{n+1} \times \dots$.

It is routine to verify that Q^n is a conditional distribution for the future given the past. Let $u = (x_1, \dots, x_n)$ and $v = (x_{n+1}, \dots)$, and compute thus:

$$\begin{aligned}
 & \rho(P^n(x_1, \dots, x_n), Q^n(x_1, \dots, x_n)) \\
 &= \rho(P^n(u), Q^n(u)) \\
 &= \int (d_n(u, v) - 1) dP^n(v | u) \text{ over } v: d_n(u, v) - 1 > 0 \\
 (14) \quad & \leq \epsilon + \int d_n(u, v) dP^n(v | u) \text{ over } v: d_n(u, v) - 1 > \epsilon \\
 &= \epsilon + Q^n(u)(v: d_n(u, v) - 1 > \epsilon) \\
 &= \epsilon + Q[d_n - 1 > \epsilon | x_1, \dots, x_n] \\
 &= \epsilon + \epsilon
 \end{aligned}$$

for all but a finite number of n with Q -probability 1, according to (11). This completes the proof.

5. Interpretation. Usually, there is essentially only one conditional distribution Q^n of the future given the past. Therefore, our theorem may be interpreted to imply that if the opinions of two individuals, as summarized by P and Q , agree only in that $P(D) > 0 \leftrightarrow Q(D) > 0$, then they are certain that after a sufficiently large finite number of observations x_1, \dots, x_n , their opinions will become and remain close to each other, where close means that for every event E the probability that one man assigns to E differs by at most ϵ from the probability that the other man assigns to it, where ϵ does not depend on E . Leonard J. Savage observed that our theorem applies to the particularly interesting case in which P and Q are symmetric (or exchangeable). That is, if the measures P and Q on the sequences x_i are those that arise when the x_i are, for a fixed parameter value, independent and identically distributed observations, with prior distributions p and q on the parameter, then the relations of absolute continuity between P and Q are precisely those between p and q .

6. Caution. Though the conditional distributions of the future P^n and Q^n merge as n becomes large, this need not happen to the unconditional distributions of the future. That is, let $P(n)(D) = P(X_1 \times \dots \times X_n \times D)$ for all $D \in \mathfrak{B}_{n+1} \times \dots$, and let $Q(n)$ be similarly defined. The following is a simple example of two probabilities P and Q absolutely continuous with respect to each other for which $P(n)$ and $Q(n)$ do not merge with increasing n . Let R be the probability on infinite sequences x_1, x_2, \dots of 0's and 1's determined by independent tosses of a coin which has probability r of success, and let S be the probability determined if the coin has probability s for success, with $0 \leq r \leq 1$, $0 \leq s \leq 1$, and $r \neq s$. Now let $0 < p < q < 1$ and let P and Q be mixtures of R and S : $P = pR + (1 - p)S$, $Q = qR + (1 - q)S$. Since $P(n) = P$ and $Q(n) = Q$ for all n , there is no tendency for $P(n)$ and $Q(n)$ to merge.

7. An application. By viewing the unit interval as a product of two point spaces, the interested reader will see that the main theorem yields information about the local behavior of positive integrable functions $q(x)$ defined for $0 \leq x \leq 1$.

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