

IMMEDIATE COMMUNICATION

Molecular genetic evidence for overlap between general cognitive ability and risk for schizophrenia: a report from the Cognitive Genomics consortium (COGENT)

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It has long been recognized that generalized deficits in cognitive ability represent a core component of schizophrenia (SCZ), evident before full illness onset and independent of medication. The possibility of genetic overlap between risk for SCZ and cognitive phenotypes has been suggested by the presence of cognitive deficits in first-degree relatives of patients with SCZ; however, until recently, molecular genetic approaches to test this overlap have been lacking. Within the last few years, large-scale genome-wide association studies (GWAS) of SCZ have demonstrated that a substantial proportion of the heritability of the disorder is explained by a polygenic component consisting of many common single-nucleotide polymorphisms (SNPs) of extremely small effect. Similar results have been reported in GWAS of general cognitive ability. The primary aim of the present study is to provide the first molecular genetic test of the classic endophenotype hypothesis, which states that alleles associated with reduced cognitive ability should also serve to increase risk for SCZ. We tested the endophenotype hypothesis by applying polygenic SNP scores derived from a large-scale cognitive GWAS meta-analysis (~5000 individuals from nine nonclinical cohorts comprising the Cognitive Genomics consortium (COGENT)) to four SCZ case-control cohorts. As predicted, cases had significantly lower cognitive polygenic scores compared to controls. In parallel, polygenic risk scores for SCZ were associated with lower general cognitive ability. In addition, using our large cognitive meta-analytic data set, we identified nominally significant cognitive associations for several SNPs that have previously been robustly associated with SCZ susceptibility. Results provide molecular confirmation of the genetic overlap between SCZ and general cognitive ability, and may provide additional insight into pathophysiology of the disorder.

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INTRODUCTION

Schizophrenia (SCZ) is a brain disorder characterized by, on average, a reduction in general cognitive abilities of >1 s.d. below the population mean. Given the long-standing observation of subtle cognitive abnormalities in undiagnosed and unmedicated relatives of patients with SCZ, it has been suggested that cognitive deficits may serve as an endophenotype, permitting identification of SCZ risk genes using a quantitative phenotype more closely reflecting gene function.¹ Extensive family and twin data support the role of shared additive genetic factors underpinning both SCZ and cognitive deficits.² A recent population-scale study of siblings and twins further suggested that the overlap between these phenotypes is largely genetic, but questioned the overall magnitude of the phenotypic correlation.³ However, these family-based studies have two primary limitations: (1) they lack direct molecular assays of genetic variance and (2) they have relied upon measurement of cognitive abilities in patients with SCZ, which are subject to potential confounds relating to illness process and state.

A direct test of the endophenotype hypothesis would examine molecular genetic variants associated with cognitive performance in the general (not clinically referred) population, in order to see if these variants are also associated with SCZ. However, this approach has not been adequately tested because of the lack of strongly replicated cognition-associated single-nucleotide polymorphisms (SNPs). By contrast, a number of candidate gene studies have applied a 'reverse endophenotype' approach, in which SCZ risk variants are tested for association with cognition in the general population.^{4,5}

Large-scale genome-wide association studies (GWAS) of SCZ have demonstrated that a substantial proportion of the heritability of the disorder is explained by a polygenic component consisting of thousands of common SNPs of extremely small effect.⁶ Similarly, recent GWAS of general cognitive ability have indicated that a polygenic architecture accounts for a majority of the heritability, consistent with other normally distributed traits in the general population, such as height.⁷⁻⁹ The relationship of the underlying genetic architecture between two phenotypes can be examined using polygenic score tests, in which large numbers of alleles demonstrating subtle (not statistically significant) levels of association with a given phenotype are combined to produce a genetic risk profile.¹⁰ The association of these alleles in a different cohort (measured on a different phenotype) can then be utilized to estimate the degree of molecular overlap between phenotypes.

Very recently, two reverse endophenotype studies have examined the ability of SCZ polygenic risk scores to predict cognitive abilities in independent cohorts.^{11,12} Both studies draw from the work of the Psychiatric GWAS Consortium on SCZ (PGC-SCZ), a large-scale mega-analytic GWAS of 17 Caucasian cohorts encompassing >9000 SCZ cases and >12 000 controls.¹³ In one study,¹² polygenic risk scores for SCZ were significantly associated with IQ in a cohort of patients with SCZ and controls, but these associations were not observed in the smaller ($n = 322$) healthy cohort alone. However, in a much larger study,¹¹ polygenic risk scores for SCZ were found to be significantly associated with general cognitive ability (as assessed by the first principal component of cognitive tests assessing multiple domains) in two moderately sized (total $N \sim 1500$) cohorts of psychiatrically normal older adults.

In this study, we sought to test the endophenotype hypothesis by comparing SCZ cases to controls on polygenic SNP scores derived from a large-scale meta-analysis of cognitive GWAS. We focused on general cognitive ability as the primary endophenotype for several reasons that we have discussed in detail previously.¹⁴ Although a variety of specific cognitive traits have been proposed as SCZ endophenotypes, research to date has failed to conclusively demonstrate any specific cognitive signature

of SCZ.¹⁵ Overall, deficits in general cognitive ability follow the pattern required of an endophenotype¹: they are observed in SCZ patients before illness onset, are largely independent of clinical state and medication status, and are familial in nature.^{16,17}

It has been recognized for more than 100 years that most cognitive abilities in humans are not orthogonal traits, but instead tend to covary; the underlying factor accounting for this pattern of intercorrelation is termed general cognitive ability (or Spearman's g).¹⁸ General cognitive ability is a robust phenotype, accounting for nearly half of the variance on the broad range of specific neurocognitive traits identifiable in humans.¹⁹ Moreover, general cognitive ability in the population is itself highly heritable,^{9,20,21} and heritability studies have demonstrated substantial genetic overlap across most specific cognitive domains tested.^{18,19,21} Crucially, general cognitive ability can be reliably estimated by extracting the first principal component of any appropriately diverse set of neurocognitive test scores, regardless of the specific components of the battery; it has been empirically demonstrated that estimates of g derived from varying batteries tend to be very highly correlated, with correlation coefficients often approaching 1.^{22,23}

The present study represents the first empirical report of an international collaborative effort entitled, 'The Cognitive Genomics consorTium (COGENT)'.¹⁴ COGENT aims to bring together human genetic data sets with both: (1) high-density genome-wide genotype data and (2) phenotype data on cognitive function in individuals drawn from the general population. At the time of the first data freeze, COGENT consists of nine sites across seven countries with approximately 5000 individuals with available genotype and phenotype data. Although genotyping platforms and phenotype measures vary by site, genetic imputation and factor analysis of cognitive scores were used to harmonize data across sites. Because generalized cognitive ability (g) can be robustly estimated from a variety of test batteries, we were able to perform meta-analysis of GWAS associations to g across the nine COGENT cohorts. From this meta-analysis, we derived polygenic allele scores associated with general cognitive ability. These allele scores were then applied to four SCZ case-control cohorts consisting of more than 11 000 independently ascertained subjects (>5000 cases and >5000 controls), as described in detail below. We in addition performed 'reverse endophenotype' analyses, examining the effects of SCZ risk alleles (derived from PGC-SCZ) on cognitive scores in the nine COGENT cohorts.

MATERIALS AND METHODS

Subjects–cognitive GWAS cohorts

Volunteers for cognitive studies were drawn from nine cohorts, for which study investigators agreed to share data as part of the COGENT. Details on subject recruitment procedures for each cohort are described in the Supplementary Materials; summaries of each cohort are presented in Table 1. Although screening procedures differed somewhat across cohorts, subjects were drawn from the general population, either as epidemiologically representative cohorts or as recruited control cohorts for studies of SCZ and/or other mental illnesses. All subjects were of Caucasian descent (as confirmed by principal components analysis (PCA) of genetic data). All subjects provided written, informed consent to protocols approved by their institutional ethics boards in accordance with the Helsinki declaration.

Subjects–SCZ case-control cohorts

The primary test of the endophenotype hypothesis was performed in the Molecular Genetics of Schizophrenia (MGS) European–American case-control cohort. This data set was selected for several reasons: it is large ($n > 5000$), publicly available, has been extensively studied,²⁴⁻²⁶ and contains an ethnic distribution that is comparable to nine COGENT cohorts (primarily Northern European in ancestry but with a non-negligible Southern European component as well). To replicate and extend our findings, we secondarily tested three additional SCZ case-control cohorts

Table 1. Description of COGENT cohorts

| Data set ^a | Genotyping platform | Concordance ^b | Genotypes after QC | N | % Male | Mean age (s.d.) | Lambda ^c |
|-----------------------|--------------------------|--------------------------|--------------------|------|--------|-----------------|---------------------|
| Germany | Illumina OE ^d | 99.50% | 1,078,289 | 594 | 51% | 54.0 (15.0) | 1.01 |
| LOGOS | Illumina OE | 99.40% | 835,287 | 802 | 100% | 22.3 (3.8) | 1.03 |
| IBG | Affymetrix 6.0 | 99.59% | 938,800 | 299 | 77% | 15.9 (1.5) | 1.00 |
| LBC1936 | Illumina 610 | 99.60% | 1,058,722 | 1005 | 51% | 69.5 (0.8) | 1.01 |
| TOP | Affymetrix 6.0 | 94.23% | 917,315 | 351 | 48% | 34.2 (9.8) | 1.01 |
| NCNG | Illumina 610 | 99.40% | 944,135 | 629 | 32% | 47.6 (18.3) | 1.00 |
| Manchester | Illumina 610 | 99.60% | 1,059,916 | 697 | 30% | 67.7 (2.8) | 1.01 |
| HBCS | Illumina 610 | 99.60% | 1,043,380 | 318 | 100% | 67.7 (2.3) | 1.00 |
| ZHH | Illumina OE | 99.40% | 1,043,785 | 201 | 47% | 39.1 (1.8) | 1.06 |

Abbreviations: COGENT, Cognitive Genomics consortium; GWAS, genome-wide association studies; QC, quality control. ^aDetailed descriptions of each cohort provided in Supplementary Text. ^bConcordance between imputed and genotyped SNPs. ^cLambda refers λ_{GC} , a measure of the degree of statistical inflation in GWAS. ^dOE refers to the Illumina OmniExpress genotyping bead chip.

Table 2. Description of SCZ case-control cohorts

| SCZ data set | N cases | N controls | GWAS platform |
|--|---------|------------|---------------------|
| MGS European–American, Shi et al. ²⁵ | 2681 | 2653 | Affymetrix 6.0 |
| Japan, Ikeda et al. ²⁷ | 575 | 564 | Affymetrix 5.0 |
| Ashkenazi Jewish (Israel), Guha et al. ²⁸ | 904 | 1640 | Illumina Omni1-Quad |
| MGS African–American, Shi et al. ²⁵ | 1286 | 973 | Affymetrix 6.0 |
| Total | 5446 | 5830 | |

Abbreviations: GWAS, genome-wide association studies; SCZ, schizophrenia.

of varying ethnicities: (1) a Japanese cohort with >1000 subjects;²⁷ (2) an Ashkenazi Jewish cohort with >2500 subjects;²⁸ and the African–American subcohort ($n > 2000$) of the MGS sample.²⁵ Demographic details of these cohorts are presented in Table 2. It should be noted that increasing evidence suggests substantial common architecture of complex traits (including SCZ) across populations,^{29,30} but it would still be anticipated that replication samples would demonstrate attenuated results due to residual differences in allele frequencies and effect sizes.³¹

Genotyping, quality control and imputation

As described in detail in the Supplementary Materials, all COGENT subjects were genotyped on one of three microarray platforms: Affymetrix 6.0 (Santa Clara, CA, USA; ~900 K SNPs), Illumina 610 K or Illumina OmniExpress (San Diego, CA, USA; ~770 K SNPs). A standardized quality control (QC) pipeline was applied to each COGENT GWAS data set: SNP call rate >95%; sample call rate >90%; SNP Hardy–Weinberg equilibrium $P > 10^{-6}$; and X chromosome sex match with reported gender. For any pair of subjects with cryptic relatedness (π -hat >0.125 in PLINK (ref. 32) 1.07), the sample with the lower call rate was eliminated. For each data set, a PCA was performed (in SVS 7.7.4, GoldenHelix, Bozeman, MT, USA), and samples demonstrating non-Caucasian ancestry were eliminated.

After QC, all SNPs within a given cohort were strand-aligned to HapMap3 and phased using SHAPEIT³³ before imputation with IMPUTE2.³⁴ As recently recommended to increase imputation accuracy,³⁵ a large, cosmopolitan HapMap3 reference panel ($n = 1011$ individuals from Africa, Asia and Europe, and the Americas) was utilized (except for NCNG, which was the only data set for which imputation was not performed centrally). Because our phenotype is a quantitative trait, we sought to avoid potentially spurious findings introduced by random association of rare alleles with a few extreme scores.³⁶ Therefore, imputed SNPs receiving a probability call >0.90 were retained and converted to PLINK-format genotype calls. The imputed data were then re-cleaned using the same call rate and Hardy–Weinberg equilibrium criteria described above; in addition, SNPs with minor allele frequency < 2.5% were dropped. For each cohort, ~3000 randomly selected genotyped SNPs were held out for concordance

analysis with imputation results. As shown in Table 1, concordance exceeded 99% for eight COGENT cohorts, and ~1 M SNPs were available for analysis in each cohort.

Neurocognitive assessment

Details of neurocognitive batteries for each cohort are provided in the Supplementary Materials. Although the specific instruments varied across cohorts, each cohort was required to have test scores available across at least three domains of cognitive ability for computation of Spearman's ρ ,^{18,37} an estimate of general cognitive ability derived from PCA.^{22,23} (For one cohort, a validated estimate of general cognitive ability derived from two subscales of the Wechsler Adult Intelligence Scale was utilized.) For each of the cohorts, available measures were entered into PCA and the first unrotated component was extracted. Any variable with more than 5% of missing data was dropped from the analysis. Normality is not a strict requirement of PCA implemented for the purpose of data reduction and so no variable was subject to transformation.³⁸ Moreover, inspection of box plots indicated that variables were generally normally distributed and no noticeable outliers were observed. In each cohort, as expected based on hundreds of prior studies,³⁹ the first principal component significantly loaded all measures and accounted for ~40% of the variance on average.

The dependent measure for the cognitive GWAS in each cohort was this first PC score, corrected for the following (using linear regression before GWAS): age, sex, age \times sex, age², and age² \times sex, based on consistent evidence demonstrating the presence of both linear and quadratic effects of age on general cognitive ability across the life span.⁴⁰

Statistical analysis

Genome-wide association analysis of the quantitative cognitive phenotype was performed in each COGENT cohort using linear regression (additive model) in SVS7.7.4. As shown in Table 1, lambda (genomic control) values for each cohort were at or near 1, indicating no significant effect of subtle population structure on association results. Fixed effects meta-analysis of β -weights from the linear regression analyses was performed in PLINK 1.07,³² using data from all available cohorts possessing high-QC genotyped or imputed data for each given SNP. Only SNPs with data available in three or more cohorts were retained. By convention, a positive β -weight for a given allele indicated an additive (allele-dose) relationship in the direction of higher cognitive phenotype scores.

Based on results of the meta-analysis, polygenic scores were computed in PLINK using β -weights of alleles at varying statistical thresholds (nominal $P < 0.10, 0.20, 0.30, 0.40$ and 0.50), following the procedure originally described by Purcell et al.¹⁰ For each statistical threshold, the clump procedure in PLINK was utilized to prune the set of SNPs for linkage disequilibrium (using r^2 threshold of 0.50 within a 250-kb window), so as to avoid redundancy of SNPs representing a given association signal.

For each of the five statistical thresholds, the weighted allele scores of each 'clumped' SNP were summed for each subject in each SCZ case-control cohort, thus creating a 'cognitive polygene score.' Thus, each subject in each SCZ case-control data set had a set of five cognitive polygene scores: one for each of the statistical thresholds applied to the original COGENT meta-analytic results. For each SCZ case-control data set, five logistic regression analyses were then used to compare cases and controls on cognitive polygene scores at each threshold. Nagelkerkes'

pseudo R^2 was utilized to reflect estimated percent variance in the SCZ phenotype accounted for by cognitive polygene scores at each threshold.¹⁰

RESULTS

GWAS results for general cognitive ability

Meta-analytic results of the nine cognitive GWAS cohorts are presented as a Manhattan plot in Figure 1; no SNPs reached genome-wide significance, and the overall lambda was 1.031 (see QQ plot in Supplementary Figure 1). This lambda value is higher than that for any individual cohort (Table 1), indicative of polygenic signal,⁴¹ but is lower than reported in similar recent studies of cognitive ability.^{7,9} A list of top SNPs ($P < .001$) emerging from the meta-analysis are presented in Supplementary Table 1.

Primary test of endophenotype hypothesis—cognitive polygenic score analysis

We utilized a large, publicly available, Caucasian SCZ GWAS case-control cohort to test our primary endophenotype hypothesis. Using 'clumped' SNPs at five different thresholds (nominal $P < 0.10, 0.20, 0.30, 0.40$ and 0.50), polygenic cognitive scores were computed for each of 2886 cases and 2056 controls in the MGS/GAIN European-American cohort. As predicted, cases had significantly lower cognitive polygenic scores across each of the five thresholds (P -values ranging from 6.56×10^{-6} to 3.73×10^{-7} ; see Figure 2). In other words, SCZ cases had fewer alleles associated with good cognitive performance and more alleles associated with poorer cognitive performance in the COGENT meta-analysis. The overall magnitude of the polygenic effect was small ($R^2 < 0.01$), but comparable to the total variance explained by a similar approach applied from one cognitive cohort to another⁷ and from a reverse endophenotype study.¹¹

As replication, the same cognitive SNPs were tested in three smaller SCZ case-control cohorts of differing ethnicities. Because of the apparent dip in variance explained that appears at $P < 0.2$ in the MGS EA cohort (Figure 1), we applied a $P < 0.3$ threshold in our replication data sets. As depicted in Table 3, nominally

significant results were observed in two cohorts (Japanese and Ashkenazi Jewish). Meta-analysis of these results across the four SCZ case-control cohorts yielded a strongly significant effect of cognitive polygenic scores on prediction of SCZ risk ($P = 3.6 \times 10^{-7}$). Very similar results were observed using a clump threshold of $P < 0.5$ (Supplementary Table 2; meta-analytic $P = 3.8 \times 10^{-7}$), and somewhat stronger results were observed when the IBG cohort (the only non-adult COGENT cohort) was removed before calculation of polygenic allele weights (Supplementary Table 3; meta-analytic $P = 4.0 \times 10^{-9}$).

Reverse endophenotype approach—examination of prior SCZ GWAS hits

As an additional examination of the relationship between cognitive GWAS results and the SCZ phenotype, we applied the commonly employed 'reverse endophenotype' approach described

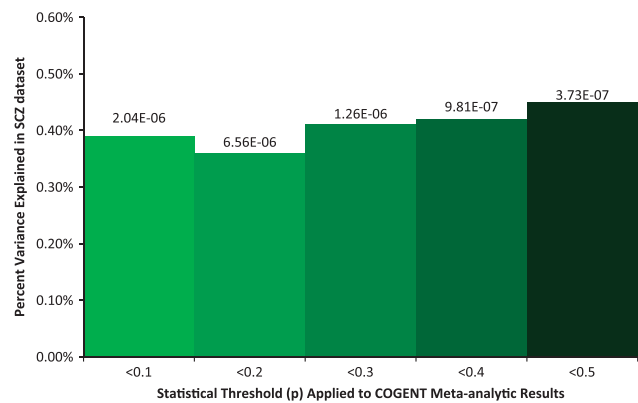


Figure 2. Polygenic overlap between cognitive allele scores (derived from the Cognitive Genomics consortium (COGENT) meta-analysis thresholded at varying P -values) and schizophrenia (SCZ) case-control status in the Molecular Genetics of Schizophrenia (MGS) European-American cohort.

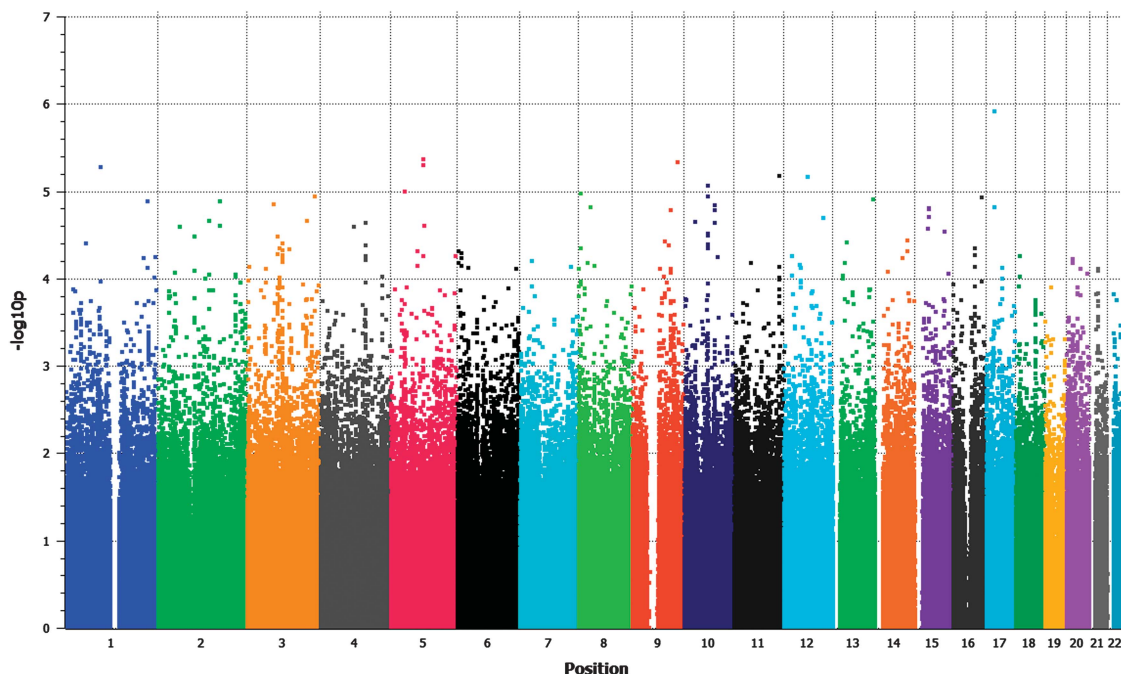


Figure 1. Manhattan plot depicting results of the Cognitive Genomics consortium (COGENT) meta-analysis.

above. First, we selected all SNPs that have demonstrated genome-wide significance ($P < 5 \times 10^{-8}$) in large-scale (total $n > 5000$) published SCZ GWAS obtained from the NHGRI GWAS catalog (<http://www.genome.gov/gwastudies/>, downloaded January 10, 2013). As shown in Table 4, four of the 13 independent SNPs (excluding the major histocompatibility complex) associated with SCZ have nominally significant ($P < 0.05$) associations with general cognitive ability in volunteers from the general population. Although effect sizes for these SNPs on cognitive phenotypes are extremely modest, the number of SNPs achieving nominal significance exceeded that expected by chance (binomial test, $P = 0.006$, two-tailed). Although extensive linkage disequilibrium in the major histocompatibility complex has prohibited clear discernment of the source of SCZ GWAS signal, Supplementary Table 4 demonstrates results for five SNPs in the region derived from published SCZ GWAS.

Reverse endophenotype approach—SCZ susceptibility polygenic score analysis

Finally, we performed a polygenic score analysis, comparable to the one presented in Table 2, but in the 'reverse' direction. Specifically, we downloaded "clumped" SNPs derived from the publicly available PGC¹³ data set (<https://pgc.unc.edu/Sharing.php#SharingOpp>). Because polygenic score approaches to SCZ have

tended to demonstrate increasing variance explained at higher P -value thresholds,¹⁰ we utilized a threshold of $P < 0.50$ on the resulting clumped SNPs. These polygenic scores were then computed for each subject in each of the nine COGENT cohorts separately; scores were then compared with cognitive ability (g) using linear regression in each cohort. As shown in Table 5, SCZ-derived polygenic scores were correlated with cognitive ability in the predicted direction (greater SCZ load associated with lower cognitive scores) in seven of the nine cohorts, with three cohorts demonstrating nominally significant (or nearly so) results. As with the primary analysis, total variance accounted for was small (R^2 range from 0 to 2%) but strongly significant ($P = 1.4 \times 10^{-4}$) in the combined analysis.

DISCUSSION

To our knowledge, this is the first study to present molecular genetic evidence supporting general cognitive ability as a true endophenotype for SCZ susceptibility. In a large Caucasian SCZ case-control cohort, we demonstrated that a set of polygenic alleles associated with lower general cognitive ability strongly predicted increased likelihood for the disorder. Consistent, though less strong, associations were observed in additional cohorts,

Table 3. Polygenic overlap between cognitive alleles (derived from COGENT meta-analysis using $P < 0.3$ threshold) and SCZ in four case-control cohorts

| SCZ data set | # overlapping SNPs | R^2 for SCZ | P-value | Direction |
|-----------------------|--------------------|---------------|----------------------|-----------|
| MGS European–American | 17,237 | 0.41% | 1.3×10^{-6} | Negative |
| Japan | 6468 | 0.38% | 0.039 | Negative |
| Ashkenazi Jewish | 15,151 | 0.16% | 0.041 | Negative |
| MGS African–American | 17,382 | 0.00% | 0.958 | Positive |
| Meta-analysis | | | 3.6×10^{-7} | Negative |

Abbreviations: COGENT, Cognitive Genomics consortium; MGS, Molecular Genetics of Schizophrenia; SCZ, schizophrenia; SNPs, single-nucleotide polymorphisms.

Table 5. Polygenic overlap between schizophrenia risk alleles (derived from PGC meta-analysis using $P < 0.5$ threshold) and general cognitive ability in nine COGENT cohorts

| COGENT data set | No. of overlapping SNPs | R^2 for g | P-value | Direction |
|-----------------|-------------------------|---------------|----------------------|-----------|
| Germany | 89,360 | 0.08% | 0.475 | Positive |
| LOGOS | 69,069 | 0.00% | 0.985 | Positive |
| IBG | 89,353 | 0.04% | 0.72 | Negative |
| LBC1936 | 96,820 | 1.17% | 0.0006 | Negative |
| TOP | 88,946 | 0.10% | 0.5481 | Negative |
| NCNG | 87,934 | 0.04% | 0.5925 | Negative |
| Manchester | 96,907 | 1.61% | 0.0007 | Negative |
| HBSC | 93,890 | 0.11% | 0.5467 | Negative |
| ZHH | 85,681 | 1.90% | 0.0532 | Negative |
| Meta-analysis | | | 1.4×10^{-4} | Negative |

Abbreviations: COGENT, Cognitive Genomics consortium; GWAS, genome-wide association studies; PGC, psychiatric GWAS Consortium.

Table 4. Examination of cognitive associations (in COGENT meta-analysis) for SNPs identified in published GWAS for schizophrenia (excluding major histocompatibility complex)

| Source | SNP | Region | Gene | Risk-allele frequency ^a | OR_Sz ^a | COGENT studies | P_COGENT meta-analysis |
|--|------------|----------|---------|------------------------------------|--------------------|----------------|------------------------|
| Ripke (PGC) Nat Gen 2011 | rs1625579 | 1p21.3 | MIR137 | 0.8 | 1.12 | 8 | 0.4789 |
| Shi (China) Nat Gen 2011 | rs10489202 | 1q24.2 | MPC2 | 0.141 | 1.23 | 9 | 0.9024 |
| Ripke (PGC) Nat Gen 2011 | rs6703335 | 1q43 | SDCCAG8 | 0.56 | 1.09 | 6 | 0.9127 |
| O'Donovan Nat Gen 2011 | rs1344706 | 2q32.1 | ZNF804A | 0.59 | 1.12 | 9 | 0.4048 |
| Ripke (PGC) Nat Gen 2011 | rs17662626 | 2q32.3 | PCGEM1 | 0.91 | 1.2 | 5 | 0.4509 |
| Bergen (Swe) Mol Psy 2012 | rs7709645 | 5q12.1 | ZSWIM6 | 0.475 | 1.11 | 9 | 0.7207 |
| Bergen (Swe) Mol Psy 2012 ^b | rs12666575 | 7p22.3 | MAD1L1 | 0.673 | 1.12 | 8 | 0.0320 |
| Ripke (PGC) Nat Gen 2011 | rs10503253 | 8p23.2 | CSMD1 | 0.19 | 1.16 | 7 | 0.5874 |
| Shi (China) Nat Gen 2011 ^c | rs16887244 | 8p11.23 | LSM1 | 0.683 | 1.19 | 9 | 0.0171 |
| Ripke (PGC) Nat Gen 2011 ^b | rs7914558 | 10q24.32 | CNNM2 | 0.59 | 1.10 | 9 | 0.0368 |
| Ripke (PGC) Nat Gen 2011 | rs11191580 | 10q24.33 | NT5C2 | 0.91 | 1.15 | 9 | 0.6575 |
| Decode Nature 2009 ^c | rs12807809 | 11q24.2 | NRGN | 0.83 | 1.15 | 8 | 0.0399 |
| Ripke (PGC) Nat Gen 2011 | rs12966547 | 18q21.2 | TCF4 | 0.58 | 1.09 | 8 | 0.7327 |

Abbreviations: COGENT, Cognitive Genomics consortium; GWAS, genome-wide association studies; PGC, psychiatric GWAS Consortium; SNPs, single-nucleotide polymorphisms. ^aBased on source publication. ^bSchizophrenia risk allele associated with lower cognitive ability. ^cSchizophrenia risk allele associated with higher cognitive ability.

despite the fact that they were drawn from populations of differing ethnicities.

Although the statistical evidence (P -values ranging from $\sim 10^{-6}$ to $\sim 10^{-9}$) for association was strong, the overall amount of variance explained, even in the Caucasian SCZ cohort ($\sim 0.5\%$) was modest, a result that must be interpreted in the context of other studies of polygenic overlap. For example, in a recent study of cognition,⁷ polygenic SNP scores accounted for only $\sim 1\%$ of the variance in a separate cognitive cohort (these cohorts are now included in COGENT). Nevertheless, these authors were able to demonstrate that nearly half of the variation in cognitive ability could ultimately be accounted for by common SNPs. Similarly, the initial study of polygenic effects in SCZ identified only $\sim 2\text{--}3\%$ overlap between any two SCZ cohorts, and $\sim 1\text{--}2\%$ overlap between SCZ and bipolar cohorts.¹⁰ Again, extensive simulations demonstrated that the polygenic SNPs tagged common genetic variation accounting for approximately one-third of the total variance in SCZ risk; this estimate has been replicated, indicating that the empirically observed cross-sample allelic overlap is a substantial under-estimate of the total polygenic effect.⁶ Moreover, observed polygenic overlap tends to rapidly increase as a function of sample size; for example, the empirically observed variance explained by polygenic effects in SCZ has increased 10-fold with larger sample sizes.⁴² Although the COGENT cohort represents the largest genetic sample of cognition to date, sample size remains small relative to polygenic studies of other quantitative traits such as height.^{43,44}

Thus, the results of the present study are consistent with a model in which a substantial fraction of the molecular basis of general cognitive ability is shared with genetic risk for SCZ. These results are consistent with a large body of evidence from family-based studies that use twin and sibling correlations to model genetic effects.² Notably, a recent population-based study using similar modeling strategies³ found only limited phenotypic overlap between intelligence and psychosis; even so, the source of this overlap was estimated to be largely ($\sim 90\%$) comprised of additive genetic variation. The present study is unique in directly testing molecular genetic variation, and utilizing non-clinical volunteers for the estimation of the cognitive genetic component. Our demonstration of parallel effects when examining genome-wide overlap in the reverse direction (SCZ risk alleles predicting cognitive scores, Table 5) add further confidence to our conclusions, and are also consistent with a recent genome-wide reverse endophenotype study, which demonstrated significant overlap between polygenic SCZ risk alleles and cognitive decline in two aging cohorts.¹¹

In addition, following the conventional 'reverse endophenotype' approach, multiple GWAS identified SCZ risk SNPs (4/13, or 31%) demonstrated nominal evidence of association to general cognitive ability in a large meta-analytic cohort of nonclinical volunteers. Three of these are intronic SNPs, in *MAD1L1*, *LSM1* and *CNNM2*, and the present study represents the first report of human neurocognitive correlates of variants at these loci. Little is known about the functions of these genes in the central nervous system, and the structural and functional properties of their associated proteins vary widely. The fourth nominally significant locus is < 5 kb, 5' to *NRGN* (and is in a linkage disequilibrium block encompassing the gene). *NRGN* encodes neurogranin, a well-characterized postsynaptic protein that binds to calmodulin and thereby modulates postsynaptic calcium signaling.⁴⁵ Although this locus did not show significant association with cognitive variables in a smaller prior study,⁴⁶ neuroimaging studies have previously associated this locus with structural and functional variation in the frontal cortex, cingulate and hippocampus.^{47–50} Although these results were significant in the aggregate ($P=0.006$ by binomial test), it should be emphasized that the effect sizes for individual SNPs were extremely small, and P -values would not survive Bonferroni correction.

Based on the present results, it is likely that larger sample sizes will afford greater ability to predict SCZ using cognitive polygenic scores. This overlap can then be utilized to enhance SCZ gene detection using recently developed 'pleiotropic enrichment' techniques.⁵¹ At the same time, the study of the genetic basis of normal variation in cognition is likely to further our understanding of the mechanisms by which SCZ risk genes affect the central nervous system.

Several caveats should be placed on the interpretation of this study. First, the present study did not directly evaluate the genetic source of cognitive deficits in patients with SCZ. It could be argued that GWAS of cognitive ability in SCZ cohorts would be required to test whether this putative endophenotype actually mediates the relationship between cognitive polygene score and SCZ risk. However, cognitive performance in patients with SCZ can be influenced by potential confounds such as effects of medication or acute symptomatology, which would tend to attenuate any genetic signal. Consequently, our approach of utilizing nonclinical samples was designed to maximize the potential power of GWAS.

Finally, it should be noted that no genome-wide significant loci for cognition were identified in the present study, despite being the largest GWAS of cognitive ability in predominately adult cohorts. This result was anticipated based on recent large-scale GWAS results for childhood intelligence,⁹ as well as early GWAS (with comparable sample size to the present study) of potentially comparable quantitative traits, height⁵² and weight.⁵³ It is also possible that power to detect genetic signals was reduced due to unavoidable heterogeneity in cognitive assessment methods across cohorts; such an interpretation is consistent with the relatively low lambda observed in our study. Future studies, ideally with prospectively collected cohorts utilizing harmonized approaches to phenotype assessment, will be required to tease out genome-wide significant loci for cognitive ability.

CONFLICT OF INTEREST

The authors declare no conflict of interest.

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