1 Multiple measures are needed to quantify training loads in professional rugby league.

2 Dan Weaving, Phil Marshall, B. Jones, K. Till, Grant Abt

3 Abstract

4 To investigate the effect of training mode (conditioning and skills) on multivariate training load relationships in professional rugby league via principal component analysis. Four 5 6 measures of training load (internal: heart rate exertion index, session rating of perceived exertion; external: PlayerLoad[™], individualised high-speed distance) were collected from 23 7 8 professional male rugby league players over the course of one 12-wk preseason period. 9 Training was categorised by mode (skills or conditioning) and then subjected to a principal 10 component analysis. Extraction criteria were set at an eigenvalue of greater than 1. Modes 11 that extracted more than 1 principal component were subject to a Varimax rotation. Skills 12 extracted 1 principal component, explaining 57% of the variance. Conditioning extracted 2 principal components (1st: internal; 2nd: external), explaining 85% of the variance. The 13 presence of multiple training load dimensions (principal components) during conditioning 14 15 training provides further evidence of the influence of training mode on the ability of individual measures of external or internal training load to capture training variance. 16 Consequently, a combination of internal- and external- training load measures is required 17 during certain training modes. 18

- 19
- 20
- 21 22
- 23
- 24
- 25

26 Introduction

To develop the wide range of physical qualities needed to succeed in professional rugby 27 league competition, multiple modes are prescribed such as skills and traditional conditioning 28 29 training [16, 27, 35]. Theoretically, the frequency, intensity and duration of the activities (e.g. 30 sprinting, accelerations, collisions) performed by players during these modes (i.e. the external load) induce multiple psycho-physiological and mechanical responses termed the internal 31 32 load [22, 33]. For a given external load, both the magnitude and type of internal load is likely to vary between players due to differences in individual characteristics which result in 33 34 multiple fitness and fatigue effects and ultimately varied training outcomes [14, 22]. Understanding these dose-response relationships are therefore important to balance the 35 promotion of adaptations whilst minimising negative outcomes such as injury [9]. To ensure 36 37 precision of an appropriate training prescription, it is important that practitioners use valid methods to quantify the internal and external loads placed onto players across all training 38 modes. 39

40

There are numerous measurements to quantify the internal and external training load 41 including heart rate (HR) based [2, 27, 36], perceptual based (session rating of perceived 42 exertion [sRPE]) [26, 34], global positioning systems (GPS), [9, 27, 36] and accelerometer 43 44 based methods [9, 27, 65]. Methods using HR to quantify the internal load include Banisters' 45 training impulse (TRIMP) [6] and the individualised TRIMP (iTRIMP) [6, 28, 29, 36] while those used to determine high-speed distance include both arbitrary [27, 36] and individualised 46 methods [1] derived from 5 Hz [15], 5 Hz with 15 Hz interpolation [27, 36] and 10 Hz [31] 47 48 GPS sampling frequencies. To infer validity, typical research designs involve correlating a practical training load method with a single criterion which is selected to represent the true 49 value of the measurement [3, 19, 22, 27, 32]. As this is typically conducted in ecologically 50

51 valid environments, the selection of the criterion method is constrained by its ability to be measured in this setting and therefore, it is also important to evaluate the extent to which the 52 criterion reflects the true value of the measurement [19]. Methods such as radar guns are 53 commonly adopted to assess the validity of external load methods such as GPS to measure 54 speed [32] whilst HR-based measurements are frequently adopted as a sole criterion method 55 to validate other internal load methods due to the difficulty in collecting additional 56 57 physiological markers in the field [3, 27, 22]. For example, the validity of sRPE is inferred due to the large within-individual correlations found with Edward's TRIMP which have been 58 59 found to range from r = 0.54 [95% confidence interval (CI): 0.14 to 0.86] to 0.78 [0.45 to 0.92] [22]. 60

61

62 Whilst we can be confident that a radar gun represents the true speed value, given the multifacetted nature of training load described previously [33], it is likely that HR-based criterion 63 measurements represent only an aspect of the actual internal load imposed. Therefore, the 64 validity of adopting a single training load measure remains unclear. Given these difficulties, 65 it is regularly suggested that a more robust approach to infer validity is to adopt the changes 66 in training outcomes, such as measures of fatigue [30], injury incidence [14] or physical 67 qualities [2. 28, 29], as the criterion method. As the theoretical internal load governs training-68 69 induced adaptations, the quantification of this construct is preferred for these load-outcome 70 relationships [21]. However, external load methods have also been found to possess doseresponse relationships with training outcomes. For example, total-distance (r = 0.86 [95% CI: 71 (0.70 to 0.95]) and high-speed distance (r = 0.76 [95% CI: 0.51 to 0.91]) were associated with 72 73 the changes in creatine kinase concentration 24-hours after professional rugby league match 74 play [30]. Therefore, it is likely that both external and internal training load methods can contribute information to the outcomes of training, the extent to which is likely to change 75

between modes of training [27, 36]. However, in most research investigating load-outcome
relationships, single training load variables are used and there is limited consideration of
whether a multivariate approach is needed to represent the training load and how his changes
across modes of training.

80

In our previous study [36], we examined the influence of training mode on the multivariate 81 82 relationships of external and internal training load measures in professional rugby league players across two 12-week pre-season periods. We reported that a combination of internal 83 84 load (iTRIMP, sRPE) and external load (BodyloadTM, total impacts and high-speed distance) explained a greater proportion of the variance during certain training modes (skills, speed, 85 strongman and wrestle) when compared to either internal or external load measures alone. 86 87 Moreover, the training load measures contributing to each principal component (PC) changed depending on the training mode. For example, during skills training the external load 88 measures explained 48% of the variance with internal load measures explaining a further 89 90 20%. However, during speed training it was the opposite, with internal load measures explaining 46% of the variance and external load measures explaining a further 21%. This 91 92 strongly suggests that a single external or internal load measure is unable to capture all training-related stress across all training types. Alterations in the strength of the relationships 93 94 between training load measures have also been shown in previous studies [27]. 95 Using a single method to quantifying the training load therefore is likely to be suboptimal in representing the multifaceted nature of the load imposed during certain training modes. For 96 certain training modes, the variability in external and internal load measurements might be 97 98 similar and could be used interchangeably. Equally, in other training modes a combination of load measures could be more sensitive in highlighting the training stress elicited. However, 99 100 despite previous findings [36], differences in microtechnology could confound the findings

101 including both GPS sampling frequency which influences the validity and reliability of highspeed movement quantification [24, 31] and accelerometer reliability and validity [7, 26]. In 102 addition, contextual influences such as different players, coaching philosophies and team 103 periodisation could all influence the conclusions drawn. As a result, due to the paucity of 104 105 current information available detailing the multivariate relationships between training load measures and how these changes across modes of training, plus the wide range of methods 106 107 used to quantify both theoretical constructs in practice, a replication study is warranted to increase the generalisability of the findings [5, 23]. 108

109

Therefore, the aim of the current study was to replicate our previous study [36], while using 110 different but commonly utilised methods to represent the external (PlayerLoadTM and 111 112 individualised high-speed-distance) and internal (sRPE and heart-rate-exertion-index [HREI]) training load, together with a shorter training period, and with players competing at a 113 different standard of competition. For the current study we focused on two of the most 114 frequently utilised training modes in rugby league (skills and traditional conditioning) [27, 115 36] and aimed to determine the structure of the interrelationships among measures of training 116 load to define common underlying dimensions in the variables via a principal component 117 analysis (PCA). PCA is a mathematical technique used to reduce the dimensionality of any 118 given data set that consists of a number of highly correlated variables, while still keeping as 119 120 much of the variation in the data set as possible [11, 25]. We hypothesised that the different external load structures of skills and conditioning training would influence the strength of the 121 variance explained by an individual training load measure. If multiple principal components 122 123 (PC) are extracted this would suggest an individual measure is unable to account for the variance of multiple measures. Within the PCA, by including only four training load 124

variables (two external and two internal) rather than the many more available to practitioners,

126 we were able to provide the most conservative test to this hypothesis.

- 127
- 128 Methods

129 Participants

- 130 Twenty-three professional rugby league players from the same Kingston Press Rugby League
- 131 Championship team participated in this study. The Championship is the 2nd highest level of
- rugby league competition in England. The participants had the following characteristics ($24 \pm$
- 133 3 years, 184.8 ± 6.7 cm, body mass 95.4 ± 8.6 kg). The study conforms with international
- ethical standards [18] was granted ethics approval by the Department of Sport, Health and
- 135 Exercise Science human research ethics committee at The University of Hull. Written

136 informed consent was obtained from each player before the start of the study.

137

138 Design

The study used a longitudinal observational research design in which training load data were
collected during one 12-week preseason preparatory period during the 2014-2015 Kingston

141 Press Rugby League Championship season.

142

143 Methodology

Training load was quantified via sRPE and microtechnology which incorporated heart rate,
GPS and tri-axial accelerometer during each training session. Prior to the commencement of
the study, all players were familiarised with these methods. The training program was
prescribed by the club's coaching staff during the course of the study. During the study
period, players typically participated in 3 field-based training sessions per week which
included conditioning (Monday) and skills (Tuesday and Friday) training. Other field-based

training modes (e.g. speed, small-sided-games) were prescribed sporadically (Thursday)
within the study period and so only modes identified as skills or conditioning were included

in the analysis, and were defined as:

Skills: Focus on enhancing individual rugby league skills and team technical-tacticalstrategies

Conditioning: focus on linear- and shuttle-running which aimed to improve players
capabilities to tolerate high-intensity running bouts. The distances for these running drills
were prescribed for each player based on a percentage of the velocity they achieved during
the 30-15 Intermittent Fitness Test (30-15_{IFT}).

159

sRPE was calculated for each player during the study period using the method of Foster et al.

161 [12] Exercise intensity for sRPE was determined using the Borg CR-10 scale [6]. sRPE was

then multiplied by the training-session duration to calculate the sRPE training load in

arbitrary units (AU). All players who participated in the study had been familiarised with the

164 RPE scale, including the interpretation of exertion in relation to the verbal anchors placed on

the scale. sRPE for each player were collected ~30 minutes after the completion of each

training session by the lead researcher into a custom-made spreadsheet with no third-party

167 observation present throughout the study period.

168

169 Manufacturer-derived heart rate exertion index (HREI) was used to calculate the heart rate-

170 derived internal load. This method follows the same principles as Edwards²² but utilises

arbitrary exponential weighting factors:

172 (Duration in Zone 1 x 1) + (Duration in Zone 2 x 1.20) + (Duration in Zone 3 x 1.50) +

173 (Duration in Zone $4 \ge 2.20$) + (Duration in Zone $5 \ge 4.50$)

174 Where zone 1 = 50-59% of HR_{max} , zone 2 = 60-69% HR_{max} , zone 3 = 70-79% HR_{max} , Zone 4 175 = 80-89% HR_{max} and zone 5 = 90-100% HR_{max}

176 HR was measured at 5 s intervals during each training session using Polar HR straps (T31

177 coded, Polar, Oy, Finland) that transmitted continuously to the GPS device (Optimeye X4,

178 Catapult Innovations, Scoresby, Victoria).

179

180 External training load measures of the distance run above a player's individualised high speed threshold (high-speed distance) and PlayerLoad[™] were collected concurrently during 181 182 each session using 10 Hz GPS devices with in-built 100 Hz tri-axial accelerometer (Optimeye X4, Catapult Innovations, Scoresby, Victoria). PlayerLoad[™] was chosen as an overall 183 measure of external load experienced by players that also includes accelerations and 184 185 collision-based activity [13] which are key considerations within rugby league [15]. PlayerLoadTM is a modified vector magnitude and is expressed as the square root of the sum 186 of the squared instantaneous rate of change in acceleration in each of the three axes (X, Y, 187 and Z) and divided by 100. PlayerLoad[™] data were expressed in arbitrary units (AU). 188 PlayerLoadTM has previously been shown to possess acceptable reliability [7]. High-speed-189 190 distance was chosen as an external load measure to represent the individualised "highintensity" running demands experienced during training for each player [1]. In order to 191 192 individualise each player's demarcated high-speed threshold, players completed the 30-15_{IFT}. 193 The 30-15_{IFT} consisted of 30 s shuttle runs interspersed with 15 s passive recovery periods as per previously described methods [8]. Speed was set at 8 km \cdot h⁻¹ for the initial 30 s run after 194 which speed was increased by $0.5 \text{ km} \cdot \text{h}^{-1}$ every 30 s [8]. Players were required to run back 195 and forth between two lines that were set 40 m apart at a speed governed by an audio signal. 196 The speed $(km \cdot h^{-1})$ achieved by each player during the last successfully completed stage of 197 the test was recorded as their maximal running speed during the test and subsequently used to 198

demarcate their high-speed threshold. The mean (SD) speed achieved during the $30-15_{IFT}$ was 200 $19.6 \pm 0.6 \text{ km}\cdot\text{h}^{-1}$.

201

202 Statistical Analysis

203 Prior to performing a principal component analysis (PCA), training load data were centred and scaled with the Pearson correlation matrix was visually inspected to determine the 204 factorability of the data for PCA [34]. The suitability of the data was assessed using the 205 Kaiser-Meyer-Olkin (KMO) measure of sampling adequacy and the Bartlett test of sphericity 206 [4]. KMO (~chi-square) values were 0.60 (284) and 0.59 (562) for conditioning and skills 207 training. A KMO value of 0.5 or above has been suggested as a threshold, above which the 208 209 data set is suitable for PCA [17, 25, 36]. Bartlett's test of sphericity was significant for both training modes (P < 0.001). The four training load measures (HREI, PlayerLoadTM, high-210 speed distance, sRPE) were subjected to a PCA for each training mode using a prior 211 communality estimate of less than 1. The stages involved in the PCA method are deletion of 212 the mean, calculation of the covariance matrix of the data, determination of the eigenvalues 213 214 and eigenvectors of the covariance matrix, and rotation of the original data onto a coordinate system spanned by the eigenvectors of the covariance matrix [11]. A principal-axis method 215 was used to extract the PC. As the number of PC will always equal the number the number of 216 original inputted variables, PC with an eigenvalue of less than 1 (Kaiser criterion) were not 217 retained for extraction [25]. This is due to the notion that any component displaying an 218 eigenvalue greater than 1.00 is accounting for a greater proportion of variance than that 219 contributed by any 1 variable. Varimax rotation was performed when two or more PC were 220 retained and with the goal of making the component loadings more easily interpretable. For 221 each extracted PC, only the original variables that possessed a PC loading of greater than 222

223	0.70 were retained for interpretation [17, 37]. The Statistical Package for the Social Sciences					
224	(SPSS, version 20.0 for Windows, SPSS Inc, Chicago, IL) was used to conduct the analysis.					
225						
226	Results					
227	A total of 640 individual training sessions were observed during the study with 23 players					
228	providing 28 ± 5 sessions each. Table 1 highlights the number of sessions and mean training					
229	loads for conditioning and skills training.					
230						
231	**INSERT TABLE 1 ABOUT HERE**					
232						
233	Table 2 displays the PCA, including eigenvalues for each principal component during skills					
234	and conditioning training and the total variance explained by each principal component for					
235	each training mode. There was a single principal component identified for skills training and					
236	two principal components identified for conditioning training, explaining 56.62% and					
237	85.44% of the variance respectively. Pearson correlations including 95% confidence intervals					
238	between the training load methods for the two training modes are presented in Table 3.					
239						
240	***INSERT TABLE 2 ABOUT HERE***					
241						
242	***INSERT TABLE 3 ABOUT HERE***					
243						
244	Discussion					
245	The main finding of the study is the identification of multiple dimensions (two principal					
246	components) in one of the modes of training, thereby confirming the results of our previous					
247	study [36]. In the current study, we identified one and two PC during skills and conditioning					

248 training, respectively. These findings demonstrate further evidence that a single training load measure, either external or internal, is unable to capture the variance of multiple measures 249 across different modes of training in professional rugby league players. This has important 250 251 implications for training load monitoring. Within a concurrent training programme, the load imposed during each mode contributes to the accumulation of load across acute (e.g. 7-day 252 rolling mean) and chronic (e.g. 28-day rolling mean) training periods [14]. As single training 253 254 load methods are commonly adopted to investigate load-outcome relationships such as injury [20] and changes in fitness [2, 28, 29], further research is required to determine whether a 255 256 multivariate training load model (using methods such as PCA) provide a better representation of load for such investigations. This is important as despite the current and previous findings 257 [35] only training load methods, including either singular or multiple measurements, that 258 259 show a dose-response relationship with training outcomes such as changes in fitness or performance should be used [2, 27, 28]. Ideally, this should involve a wide range of the 260 currently utilised training load methods that are available to examine the most influential 261 individual training load variables that contribute to a multivariate training load model. 262 More specifically, in our previous study [36] we identified a single PC during conditioning 263 training, suggesting that the training load measures were providing similar information. 264 However, in the current study we identified two PC during conditioning with the first PC 265 including HREI, sRPE and PlayerLoadTM. High-speed distance, individualised based on the 266 267 maximal speed achieved during the 30-15 IFT, explained additional variance during conditioning as it was the only variable to provide a meaningful component loading on the 268 second PC. In our previous study [36], an arbitrary (>15 km \cdot h⁻¹) high-speed distance method 269 270 was unable to account for additional variance, as only a single PC was identified during conditioning. As the major aim of conditioning training is to provide a high-intensity running 271 stimulus, the speed in which players reach 'high-intensity' will likely differ between players 272

273 [1]. Therefore, the use of an individualised approach would provide practitioners with additional information of the load prescribed during this mode. Additionally, differences in 274 GPS sampling rate could have also influenced the findings, as greater validity of high-speed 275 276 running quantification has been reported for the 10 Hz MinimaxX GPS devices when compared to the GPSports SPI Pro X 15 Hz devices used in our previous study [15]. 277 The presence of one PC during skills training suggests that a single training load variable 278 279 accounts for a meaningful proportion of the variance (56.6%) of four training load measures during this mode. As only HREI (0.78) and PlayerLoad[™] (0.92) demonstrated meaningful (> 280 281 0.70) component loadings with the extracted PC, the methods could be used interchangeably to represent the variance of the four training load variables during skills training. However, 282 the presence of a single PC conflicts with our previous findings [36]. Previously, we reported 283 284 that external training load measures (BodyloadTM, total impacts, high-speed distance) accounted for the greatest proportion of the total variance (48%) with internal load measures 285 (iTRIMP, sRPE) contributing an additional 21%. Differences in the methods used to quantify 286 the heart rate TRIMP could explain some of the discrepancies between the results. The use of 287 arbitrary heart rate zones and weightings within the HREI method have previously been 288 289 criticised [2] as they do not reflect the individualised response to exercise [1]. The iTRIMP method, adopted in the previous study [36], is based on each individual's relationship 290 291 between the fractional elevation in heart rate and blood lactate concentration, with each 292 individual heart rate data point recorded during each training bout weighted according to this relationship. This method has previously shown dose-response validity with changes in 293 294 fitness over a given training period in both endurance [27] and team sports players [2, 28]. It 295 is also important to consider that whilst only 1 PC was eligible for extraction during skills training in the current study, the total variance explained (56.6%) by this PC leaves 43.4% of 296 297 the total variance unexplained between the four training load measures. The Kaiser criterion

298 (eigenvalue > than 1) is considered a conservative approach to extracting meaningful PC [17] and is one of multiple criteria that can be adopted [37] which include the assessment of the 299 scree plot [37] and/or extraction of the number of PC that equal a set percentage of total 300 variance explained [17]. Therefore, it is possible that the second PC (Table 2) could explain 301 additional meaningful variance and therefore, multiple measures could actually be required 302 during skills. A limitation of the current study is that due to the variety of skill and tactical 303 304 qualities needed to succeed in rugby league competition, skills training will involve a wide range of activities that will subject players to different compositions of external load 305 306 intensities (e.g. walk, run, sprint, collisions) between sessions including collision activity. Therefore, as skills training is prescribed frequently within training periods [27, 35], future 307 research should determine the relationships between training load methods (either single or 308 309 combined) and acute training outcomes such as changes in fatigue markers [30, 35] during skills training and consider the influence of collision based activity on those relationships to 310 further elucidate their validity during this training mode. Finally, despite the discrepancies 311 between the current and previous results [36], the findings highlight the importance of 312 investigations that replicate previous research findings [23]. 313 314 **Practical Applications** 315 Questions the use of a single measure when making decisions of the load imposed 316

- 317 onto players.
- Consider the influence that the training mode has on the capability of individual
 methods used to reflect the actual load imposed during that training session.
- Consider measuring the training load using combinations of external and internal
 load. During conditioning, it appears one of either PlayerLoad[™], HREI or sRPE plus
- 322 individualised high-speed-distance should be adopted.

323 Conclusions

324	The current study has shown that the training mode (conditioning and skills) influences the
325	capability of a single training load measure to explain the variation in multiple measures of
326	the external and internal training load in professional rugby league players. This suggests
327	practitioners shouldn't rely on a single measure to inform decisions regarding the load
328	imposed onto players and should use both an internal and external training load measure to
329	monitor their prescription of training. The findings provide further evidence that a
330	multivariate training load model that combines internal and external training load measures
331	should be considered. However, further research is needed to establish how this can be
332	implemented in practice and whether this provides a better model of load-outcome
333	relationships compared to a single measure.
334	
335	
336	
337	
338	
339	
340	
341	
342	
343	
344	
345	
346	
347	

348 **References**

- 349 1. Abt G, Lovell R. The use of individualised speed and intensity thresholds for determining
- the distance run at high-intensity in professional soccer. J Sports Sci. 2009; 27: 893-898.
- 2. Akubat I, Patel E, Barrett S, Abt G. Methods of monitoring the training and match load and
- their relationship to changes in fitness in professional youth soccer players. J Sports Sci.
- 353 2012; 30: 1473–1480.
- 3. Alexiou H, Coutts AJ. A comparison of methods used for quantifying internal training load
 in women soccer players. Int J Sports Physiol Perform. 2008; 3: 320-330.
- 4. Bartlett MS. A note on the multiplying factors for various chi square approximations. J R
- 357 Stat Soc Ser C Appl Stat. 1954; 16: 296–298.
- 358 5. Bishop D. An applied research model for the sports sciences. Sports Med. 2008; 38: 253263.
- 360 6. Borg G, Ljunggren G, Ceci R. The increase of perceived exertion, aches and pain in the
- legs, heart rate and blood lactate during exercise on a bicycle ergometer. Eur J Appl
- 362 Physiol.1985; 54: 343–349.
- 363 7. Boyd LJ, Ball K, Aughey RJ. The reliability of MinimaxX accelerometers for measuring
- 364 physical activity in Australian football. Int J Sports Physiol Perform. 2011; 6:311–321.
- 365 8. Buchheit M. The 30-15 intermittent fitness test: accuracy for individualising interval
- training of young intermittent sport players. J Strength Cond Res. 2008; 22: 365-374.
- 367 9. Colby MJ, Dawson B, Heasman J, Rogalski B, Gabbett TJ. Accelerometer and GPS-
- 368 derived running loads and injury risk in elite Australian footballers. J Strength Cond Res
- 369 2014; 28: 2244-2252.
- 10. Edwards S. High performance training and racing. In: The Heart Rate Monitor Book.
- Edwards, ed. Sacramento, CA: Feet Fleet press. 1993.

- 11. Federolf P, Reid R, Gilgien M, Haugen P, Smith G. The application of principal
- 373 component analysis to quantify technique in sports. Scand J Med Sci Sports. 2014; 24: 491374 499.
- 12. Foster C, Florhaug JA, Franklin J, et al. A new approach to monitoring exercise training.
- 376 J Strength Cond Res. 2001; 15: 109–115.
- 13. Gabbett TJ. Relationship between accelerometer load, collisions, and repeated high-
- intensity-effort activity in rugby league players. J Strength Cond Res 2015; 29: 3424-3431.
- 14. Gabbett TJ. The training-injury prevention paradox: should athletes be training smarter
- and harder? Br J Sports Med 2016; 50: 273-280.
- 15. Gabbett TJ, Jenkins DG, Abernethy B. Physical demands of professional rugby league
- training and competition using microtechnology. J Sci Med Sport. 2012; 15: 80-86.
- 383 16. Gabbett TJ, Stein JG, Kemp JG, Lorenzen C. Relationship between tests of physical
- qualities and physical match performance in elite rugby league players. J Strength Cond Res
 2013; 27: 1539-1545.
- 17. Hair J, Anderson RE, Tatham RL, Black WC. Multivariate Data Analysis. 4th ed. Upper
 Saddle River, NJ: Prentice-Hall; 1995.
- 18. Harriss DJ, Atkinson G. Ethical standards in sport and exercise science research: 2016
- 389 update. Int J Sports Med 2015; 36: 1121-1124.
- 19. Hopkins WG. Measures of reliability in sports medicine and science. Sports Med 2000;
 30, 1-15.
- 20. Hulin BT, Gabbett TJ, Lawson DW, Caputi P, Sampson JA. The acute:chronic workload
- ratio predicts injury: high chronic workload may decrease injury risk in elite rugby league
- 394 players. Br J Sports Med 2016; 50: 231-236.
- 21. Impellizzeri F, Rampinini E, Marcora S. Physiological assessment of aerobic training in
- 396 soccer. J Sports Sci. 2005; 23: 583-592.

- 22. Impellizzeri FM, Rampinini E, Coutts AJ, Sassi A, Marcora SM. Use of RPE-based
- training load in soccer. Med Sci Sports Exerc. 2004; 36: 1042-1047.
- 23. Ioannidis JPA (2005) Why most published research findings are false. PLoS Med. 2005;
 2: e124
- 401 24. Johnston RJ, Watsford ML, Kelly SJ, Pine MJ, Spurrs RW. Validity and interunit
- 402 reliability of 10 Hz and 15 Hz GPS units for assessing athlete movement demands. J Strength
- 403 Cond Res. 2014; 28: 1649-1655.
- 404 25. Kaiser HF. The application of electronic computers to factor analysis. Educ Psychol
 405 Meas. 1960; 20: 141–151.
- 406 26. Kelly SJ, Murphy AJ, Watsford ML, Austin D, Rennie M. Reliability and validity of
- 407 sports accelerometers during static and dynamic testing. Int J Sports Physiol Perform. 2015;
- 408 10: 106-111.
- 409 27. Lovell TWJ, Sirotic AC, Impellizzeri FM, Coutts AJ. Factors affecting perception of
- 410 effort (session rating of perceived exertion) during rugby league training. Int J Sports Physiol
- 411 Perform 2013; 8: 62–69.
- 412 28. Manzi V, Iellamo F, Impellizzei F, D'Ottavio S, Castagna C. Relation between
- 413 individualized training impulses and performance in distance runners. Med Sci Sports Exerc.
- 414 2009; 41: 2090–2096.
- 415 29. Manzi V, Bovenzi A, Impellizzeri FM, Carminati I, Castagna C. Individual training-load
- and aerobic fitness variables in premiership soccer players during the precompetitive season.
- 417 J Strength Cond Res. 2013; 27: 631–636.
- 418 30. Oxendale CL, Twist C, Daniels M, Highton, J. The relationship between match-play
- 419 characteristics of elite rugby league and indirect markers of muscle damage. Int J Sports
- 420 Physiol Perform. 2016; 11: 515-521.

- 421 31. Rampinini E, Alberti G, Florenza M, Riggio M, Sassi R, Borges TO, Coutts AJ. Accuracy
- 422 of GPS devices for measuring high-intensity running in field-based team sports. Int J Sports
 423 Med. 2015; 36: 49-53.
- 424 32. Roe G, Darrall-Jones J, Black C, Shaw W, Till K, Jones B. Validity of 10 Hz GPS and
- 425 timing gates for assessing maximum velocity in professional rugby union players. Int J Sports
- 426 Physiol Perform 2016; 13: 1-14.
- 427 33. Soligard T, Schwellnus M, Alonso JM, Bahr R, Clarsen B, Dijkstra HP, Gabbett T,
- 428 Gleeson M, Hagglund M, Hutchinson MR, Janse van Rensburg C, Khan KM, Meeusen R,
- 429 Orchard JW, Pluim BM, Raferty M, Budgett R, Engebretsen L. How much is too much? (Part
- 430 1) International Olympic Committee concensus statement on load in sport and risk of injury.
- 431 Br J Sports Med 2016; 50: 1030-1041.
- 432 34. Tabachnick BG, Fidell LS. Using Multivariate Statistics. Boston, MA:Pearson Education;
 433 2007.
- 434 35. Twist C, Highton J. Monitoring fatigue and recovery in rugby league players. Int J Sports
 435 Physiol Perfom 2013; 8: 467-474.
- 436 36. Weaving D, Marshall P, Earle K, Nevill A, Abt G. Combining internal- and external-
- training-load measures in professional rugby league. Int J Sports Physiol Perform 2014; 9:
- 438 905-912.
- 439 37. Williams S, Trewartha G, Cross MJ, Kemp, SP, Stokes KA. Monitoring what matters: a
- 440 systematic process for selecting training load measures. Int J Sports Physiol Perform. 2016;
- 441 11: 1-20.
- 442
- 443

Table 1. Mean \pm SD training load measures and session durations during each training mode.

Training Mode	n	Duration (min)	HREI (AU)	sRPE (AU)	PlayerLoad TM (AU)	HSD (m)
Skills	448	40 ± 24	100 ± 69	309 ± 183	351 ± 150	202 ± 265
Conditioning	192	25 ± 12	59 ± 32	183 ± 345	232 ± 81	599 ± 455

sRPE: Session rating of perceived exertion; HREI: Heart rate exertion index; HSD: High-speed distance

	Principal Component			
	1	2	3	4
<u>Skills</u>				
Eigenvalue	2.27	0.80	0.72	0.22
% of Variance	56.62	20.03	17.92	5.42
Cumulative Variance %	56.62	76.66	94.58	100.00
Unrotated Component Loadings				
HREI	0.78	-	-	-
sRPE	0.65	-	-	-
Playerload	0.92	-	-	-
HSD	0.62	-	-	-
Conditioning				
Eigenvalue	2.24	1.18	0.32	0.27
% of Variance	56.01	29.42	7.90	6.66
Cumulative Variance %	56.01	85.44	93.34	100.00
Rotated Component Loadings				
HREI	0.89	-0.12	-	-
sRPE	0.90	-0.15	-	-
Playerload	0.80	0.48	-	-
HSD	-0.14	0.96	-	-

Table 2. Results of the PCA, showing the Eigenvalue, percentage (%) of variance explained and the cumulative % of variance explained by each Principal Component (PC) for skills and conditioning. Also showing the unrotated (1 PC extracted) or rotated (> 1 PC extracted) training load component loadings for each PC that were extracted. Loadings that met interpretation criteria (> 0.70) are highlighted in bold.

sRPE: Session rating of perceived exertion; HREI: Heart rate exertion index; HSD: High-speed distance

Table 3: Pearson's product-moment coefficients for each training load measure during skills and conditioning training. Includes 95% Confidence Intervals (CI) for each significant correlation. * Significant at 0.05 level ** Significant at 0.001 level *** Significant at 0.0001 level Hopkins (2002) qualitative correlation coefficient descriptors: t: trivial (0-0.09), s: small (0.1-0.29), m: moderate (0.3-0.49), 1: large (0.7-0.89), v1: very large (0.9-0.99)

	sRPE	95% CI	PlayerLoad TM	95% CI	HSD	95% CI
Skills						
HREI	0.30*** ^m	[0.23 to 0.40]	0.72^{***1}	[0.67 to 0.76]	0.22^{***s}	[0.13 to 0.31]
sRPE	1.00	-	0.47*** ^m	[0.39 to 0.54]	0.27*** ^s	[0.18 to 0.35]
PlayerLoad™	-	-	1.00	-	0.47*** ^m	[0.39 to 0.54]
Conditioning						
HREI	0.73*** ¹	[0.66 to 0.79]	0.55^{***l}	[0.44 to 0.64]	-0.19^{**s}	[-0.32 to -0.05]
sRPE	1.00	-	0.56^{***l}	[0.45 to 0.65]	-0.21^{**s}	[-0.34 to -0.07]
PlayerLoad TM	-	-	1.00	-	0.24^{***s}	[0.10 to 0.37]

sRPE: Session rating of perceived exertion; HREI: Heart rate exertion index; HSD: High-speed distance