# Obesity and cancer: Mendelian randomization approach utilizing the *FTO* genotype

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| Background | Obesity is a risk factor for several cancers although appears to have an inverse association with cancers strongly related to tobacco. Studying obesity is difficult due to numerous biases and confounding.   |
| Methods    | To avoid these biases we used a Mendelian randomization approach incorporating an analysis of variants in the <i>FTO</i> gene that are strongly associated with BMI levels among 7000 subjects from a study of lung, kidney and upper-aerodigestive cancer.  |
| Results    | The <i>FTO</i> A allele which is linked with increased BMI was associated with a decreased risk of lung cancer (allelic odds ratio (OR) = 0.92, 95% confidence interval (CI) 0.84–1.00). It was also associated with a weak increased risk of kidney cancer, which was more apparent before the age of 50 (OR = 1.44, CI 1.09–1.90). |
| Conclusion | Our results highlight the potential for genetic variation to act as an<br>unconfounded marker of environmentally modifiable factors, and<br>offer the potential to obtain estimates of the causal effect of obesity.<br>However, far larger sample sizes than studied here will be required<br>to undertake this with precision.     |
| Keywords   | Obesity, cancer, Mendelian randomization   |

## Introduction

Obesity is implicated in the aetiology of several cancers including kidney, breast, colon and gallbladder,

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and may contribute to an important and increasing number of incident cancer cases in Western societies.<sup>1,2</sup> Conversely, a consistent protective effect of higher body mass index (BMI) has been observed

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for some tobacco-related cancers, in particular for lung cancer.<sup>1</sup> Evaluating a causal role of obesity and specific cancers is difficult for a number of reasons. Obesity is likely to be correlated with numerous other lifestyle factors, some of which may be unknown or unmeasured, and removing all potential confounding is not possible. Epidemiological studies of obesity and cancer also usually include measures of weight and height obtained at one point in time, and some level of measurement error is inevitable. Further, a one-off measure may not be representative of lifetime patterns of body composition, including during childhood and early adulthood. Measures of BMI also confuse weight change due to muscle formation as opposed to increased fat. Finally, early stages of cancer may result in weight loss, which may introduce bias due to reverse causation into studies of obesity and cancer.

A Mendelian randomization approach can help overcome these problems by using genetic variation as an instrument or proxy for BMI and obesity.<sup>3</sup> Genetic variants that result in increased weight would have several advantages over traditional measures of obesity in that they will not be related to potential confounding factors, can be measured with little or no error, are not affected by disease or recall bias, and they may act as a lifelong marker of increased body weight. The principal restriction of this approach is to identify genetic markers that have a sufficiently important effect on weight for this to translate into a detectable increase (or decrease) in cancer risk.

Recent results provide compelling evidence that variants in the *FTO* gene can have an important effect on BMI and risk of obesity.<sup>4,5</sup> One variant located in the first intron of FTO (rs9939609) resulted in an average increase of  $\sim$ 3 kg in weight, or one BMI unit, between the two homozygous genotype groups A/A and T/T and an odds ratio (OR) of obesity of 1.67 for a BMI >30. This association was detectable among children as young as 7, implying a long-term effect starting in childhood. Furthermore, the A/A homozygote genetic variant that is associated with this increase is common, with approximately one in six individuals of European origin being homozygous for the variant allele. We have therefore assessed the potential for this gene to act as an unconfounded marker of BMI and obesity in a large study of 4000 cases of lung, aero-digestive and kidney cancer, and 3000 controls.

## Materials and methods

Study details have been described in detail elsewhere.<sup>6</sup> Briefly, 15 centres in six countries of central and eastern Europe (Czech Republic, Hungary, Poland, Romania, Russia and Slovakia) followed an identical protocol and sought to recruit a consecutive group of newly diagnosed cases of primary lung cancer, and, in some centres, upper aero-digestive cancer and/or kidney cancer, as well as a comparable group of population or hospital controls. All participants were recruited between 1998 and 2003. The total number of potential cases for this study was 2250 lung, 811 aero-digestive (168 mouth, 113 pharynx, 326 larynx and 176 oesophagus, with 28 having overlapping sites) and 954 kidney cancer cases. All cases had been histologically confirmed locally, using standard WHO criteria. Controls in all centres except Warsaw were in-patients or outpatients with non-tobacco-related conditions at the same hospital as the cases, and were frequency matched with the cases by sex, age (within 3 years), centre and area of residence. No diagnostic category made up >20% of the overall control group in each centre. In Warsaw, population controls were recruited instead of hospital controls. In total there were 3052 participating controls. The same interviewer-administered questionnaire was used for both cases and controls that included extensive information on reported height and weight 2 years prior to interview.

The mean BMI among all controls was 26.89, significantly greater than among both lung cancer cases [mean BMI difference adjusted for country and sex (diff<sub>adj</sub>) = 1.45, P < 0.0001] and aero-digestive cancer cases (diff<sub>adj</sub> = 1.84, P < 0.0001), but less than that observed among kidney cancer cases (diff<sub>adj</sub> = -0.65, P = 0.0003). When we compared cases and controls with a BMI of  $\leq 25$ , with 5 BMI unit categories above this limit, a strong increased risk was observed for kidney cancer and increasing obesity levels ( $P = 1 \times 10^{-4}$ ), with even stronger inverse risks for lung cancer ( $P = 5 \times 10^{-18}$ ) and head and neck cancers ( $P = 6 \times 10^{-21}$ ) (Table 1).

After DNA extraction from lymphocyte samples, genotyping was performed by Taqman for three gene variants of the *FTO* gene that were reported to be strongly associated with BMI.<sup>4</sup> All three genotypes were in very strong linkage disequilibrium, providing almost identical data, and results are presented for rs9939609 only.

## Results

Similar to previous reports, controls with the A/A genotype had a higher BMI than controls with the T/T genotype (diff<sub>adj</sub> = 1.14 kg/m<sup>2</sup>, P < 0.00001) (Table 2). No particular heterogeneity was seen between the six different countries (P > 0.1 for all three cancer sites). A greater effect of the *FTO* gene on BMI was seen for those aged <50 [diff<sub>adj</sub> = 2.26, 95% confidence interval (CI) 1.14–3.38] compared with those aged  $\geq 50$ (diff<sub>adj</sub> = 0.94, 95% CI 0.42–1.46) ( $P_{\text{heterogeneity}} =$ 0.04) No effect was observed between the *FTO* gene and a variety of potential confounders including tobacco and alcohol consumption (data not shown).

When controls were compared with the three different case groups, the A/A genotype was associated

|               | BMI         | Cases   | Controls | OR <sup>a</sup> | 95% CI              | P-value             |
|---------------|-------------|---------|----------|-----------------|---------------------|---------------------|
| Lung          | ≤25         | 1180    | 1074     | 1.00            | ref                 |                     |
|               | 26-30       | 719     | 1192     | 0.60            | (0.52-0.69)         | $4 \times 10^{-13}$ |
|               | 31–35       | 231     | 449      | 0.47            | (0.39-0.58)         | $3 \times 10^{-13}$ |
|               | 36-40       | 50      | 89       | 0.54            | (0.36-0.81)         | 0.003               |
|               | 41 +        | 11      | 29       | 0.30            | (0.14-0.66)         | 0.003               |
|               | P-value for | r trend |          |                 | $5 \times 10^{-18}$ |                     |
| Head and Neck | ≤25         | 492     | 986      | 1.00            | ref                 |                     |
|               | 26-30       | 256     | 1118     | 0.52            | (0.43-0.63)         | $1 \times 10^{-11}$ |
|               | 31–35       | 47      | 407      | 0.27            | (0.19-0.37)         | $6 \times 10^{-14}$ |
|               | 36-40       | 11      | 77       | 0.35            | (0.18-0.71)         | 0.003               |
|               | 41 +        | 0       | 25       |                 | _                   | -                   |
|               | P-value for | r trend |          |                 | $6 \times 10^{-21}$ |                     |
| Kidney        | ≤25         | 282     | 917      | 1.00            | ref                 |                     |
|               | 26-30       | 412     | 1018     | 1.22            | (1.02 - 1.47)       | 0.033               |
|               | 31–35       | 206     | 368      | 1.59            | (1.27-2.01)         | $7 \times 10^{-5}$  |
|               | 36-40       | 43      | 69       | 1.59            | (1.04-2.43)         | 0.034               |
|               | 41 +        | 11      | 21       | 1.34            | (0.62-2.90)         | 0.458               |
|               | P-value for | r trend |          |                 | $1 \times 10^{-4}$  |                     |

Table 1 Association between increasing BMI and cancers of the lung, head and neck and kidney, in the IARC Central Europe study

<sup>a</sup>OR: Adjusted for age, sex, cumulative tobacco consumption, years of drinking consumption and country.

| Table 2 | Mean | BMI | $(kg/m^2)$ | in | controls | by | genotype | of th | e rs99396 | )9 1 | variant i | n tl | ne FTO | gene, | overall | and | by | country, |
|---------|------|-----|------------|----|----------|----|----------|-------|-----------|------|-----------|------|--------|-------|---------|-----|----|----------|
| age and | sex  |     |            |    |          |    |          |       |           |      |           |      |        |       |         |     |    |          |

|                     | Mean BMI <sup>a</sup> |       |       |                         |                     |       |  |  |  |  |  |
|---------------------|-----------------------|-------|-------|-------------------------|---------------------|-------|--|--|--|--|--|
|                     | TT                    | TA    | AA    | Difference <sup>b</sup> | 95% CI <sup>b</sup> | Pb    |  |  |  |  |  |
| Overall             | 26.65                 | 27.05 | 27.79 | 1.14                    | (0.66–1.61)         |       |  |  |  |  |  |
| By country          |                       |       |       |                         |                     |       |  |  |  |  |  |
| Romania             | 26.17                 | 27.72 | 27.15 | 0.98                    | (-0.89 to 2.86)     | 0.905 |  |  |  |  |  |
| Hungary             | 25.77                 | 27.40 | 27.86 | 2.09                    | (0.15-4.03)         |       |  |  |  |  |  |
| Poland              | 26.11                 | 26.61 | 27.21 | 1.1                     | (0.27-1.93)         |       |  |  |  |  |  |
| Russia              | 27.06                 | 26.95 | 28.02 | 0.97                    | (0.08-1.85)         |       |  |  |  |  |  |
| Slovakia            | 27.91                 | 28.05 | 28.66 | 0.75                    | (-0.96 to 2.46)     |       |  |  |  |  |  |
| Czech rep           | 26.80                 | 27.26 | 28.22 | 1.42                    | (0.42-2.42)         |       |  |  |  |  |  |
| By sex              |                       |       |       |                         |                     |       |  |  |  |  |  |
| Men                 | 26.24                 | 26.52 | 27.10 | 0.86                    | (0.34–1.38)         | 0.11  |  |  |  |  |  |
| Women               | 27.12                 | 27.90 | 28.96 | 1.84                    | (0.83-2.84)         |       |  |  |  |  |  |
| By age              |                       |       |       |                         |                     |       |  |  |  |  |  |
| Young ( $\leq 50$ ) | 25.21                 | 25.54 | 27.47 | 2.26                    | (1.14–3.38)         | 0.038 |  |  |  |  |  |
| Old (>50)           | 26.99                 | 27.41 | 27.93 | 0.94                    | (0.42 - 1.46)       |       |  |  |  |  |  |

<sup>a</sup>Means BMI adjusted for age, sex, country and tobacco consumption (pack-years).

<sup>b</sup>Difference and 95% CI compare data for rs9939609-AA to rs9939609-TT genotypes. P is for test of heterogeneity between differences.

0.83, 95% CI 0.69–0.99) with a more pronounced risk was observed for kidney cancer (OR = 1.13, effect for squamous cell tumours (OR = 0.69, 95% CI 0.53-0.89) and among never-smokers (OR = 0.53, 95%)

with a significantly reduced risk of lung cancer (OR = CI 0.30–0.93) (Figure 1). A marginal increase in 95% CI 0.90–1.43), which was restricted to those aged <50 (OR = 2.08, 95% CI 1.19–2.54)

ORs are derived from the per-allele model, adjusted for age sex and, when relevant, smoking and country. *P*-values are from heterogeneity tests.



Figure 1 Odds ratio (OR) and 95% confidence interval (CI) for lung, upper-aerodigestive (UADT) and kidney cancer by *FTO* genotype

 $(P_{heterogeneity} = 0.02)$ . No effect was observed for upper aero-digestive cancers.

#### Discussion

These results provide tentative support for an association between the FTO gene and both lung and kidney cancer and, if replicated in further large studies, would appear to confirm an independent and unconfounded role of obesity for these two cancers (one positive and one negative association). The observed effect size for the FTO gene and kidney cancer is potentially greater than that expected based on previous epidemiological findings, a feature expected to some degree given that the FTO gene is likely to correlate with obesity over the lifecourse. Our results on lung cancer would appear to support the hypothesis that weight change is inversely associated with risk of this cancer, independent of smoking and weight loss due to early disease. Further, the mechanism for any effect, and whether it differs for lung and kidney cancer, is not clear.

This current analysis was based on a large sample size comprising over 7000 subjects, although still only provided evidence that was borderline significant for an association with lung cancer (P=0.05) and a non-significant association with kidney cancer. Assuming that the difference in BMI with each A allele is ~0.6 BMI units (as in Table 2), and that the effect of increasing 1 BMI unit is to decrease risk of lung cancer by ~10%, and increase kidney cancer by 5% (as in Table 1) then one would expect an allele OR associated with *FTO* of 0.94 for lung cancer and 1.03 for kidney cancer (compared with the observed)

OR of 0.92 and 1.06, respectively). Based on these expected values, we had only a 34% power to detect a significant (P = 0.05) effect for lung cancer and a 10% power for kidney cancer. A larger sample size of 10 000 case–control pairs will result in a power of close to 90% for lung cancer, although still only a moderate power for kidney cancer (34%).

In conclusion, genetic markers of obesity (such as the *FTO* gene, but increasingly others), allow for an unbiased assessment of the role of obesity in cancer. However, effects are likely to be moderate and very large-scale coordinated initiatives to determine the true causal effect of obesity-related genes on cancer risk are required.

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## Commentary: A new dawn for genetic epidemiology?

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Genome-wide association studies (GWAS) have resulted in an unprecedented leap in our understanding of the common genetic variation associated with common diseases and traits. Work performed predominantly in the last 2 years means there are now more than 200 common DNA variants associated with human traits, at levels of statistical significance, which means less than 1 in 20 will be false positives. This contrasts with the situation before 2006, where, in total, less than 20 common variant-trait associations was considered robust.

Despite this progress, GWAS studies have been recently criticized as not providing the biological or clinical insight predicted.<sup>1,2</sup> Some commentators have suggested that the associations between a single base-pair change and a human trait or disease have not told us a great deal. This is a legitimate concern, given that often, associated DNA variants occur in the middle of several genes and it is unclear which gene is important. Those involved have pointed out that the concern is unfounded; that often variantdisease associations are merely the first step towards a better understanding of biological processes and that translation will take >2 years.<sup>3</sup> Even so, GWAS studies are already providing real biological insights. Type 2 diabetes provides a good example of where GWAS studies are starting to take the study of the actiology of a 'typical' complex disease. Many further studies are needed to build up a body of evidence to support some early indications, but initial fascinating clues include the primary role of beta-cell dysfunction ahead of insulin resistance,<sup>4,5</sup> the role of zinc ion transporters in beta-cell function,<sup>6,7</sup> evidence that disruption to circadian rhythm may be causal to disease rather than an epidemiological artefact,<sup>8–11</sup> evidence that type 2 diabetes is associated genetically with birth weight<sup>12</sup> and prostate cancer,<sup>13</sup> a clear distinction between normal glucose physiology and beta-cell dysfunction<sup>10</sup> and the start of the dissection of how the *FTO* gene could be altering adiposity through metabolism.<sup>14</sup>

In this issue of *IJE*, Brennan et al.<sup>15</sup> provide a further example of how results from GWAS studies can be used to increase our understanding of disease. This is through using common DNA variants known to associate with a trait as 'tools' or 'instruments' to assess the likely causal nature or otherwise of observational epidemiological associations. This approach, termed Mendelian Randomization, has been used before and is based around the principle that casecontrol-based associations between gene variants and traits are much less likely to be biased or confounded than traditional epidemiological studies.<sup>16,17</sup> Probably the best proof of principle comes from the fact that rare and common variants in genes that alter low-density lipoprotein cholesterol alter the risk of coronary artery disease<sup>18,19</sup>—something we know from randomized controlled trials of statins, but genetic studies provide the same answer. Brennan et al. have now taken the Mendelian Randomization approach a step further by using it to test an association that is much more poorly understood—adiposity and cancer. They tested the association between the body mass index (BMI) and obesity-associated variant in the FTO gene in a series of cancer case-control studies. They reasoned that this would provide an

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