# OPTIMUM BANDWIDTHS AND KERNELS FOR ESTIMATING CERTAIN DISCONTINUOUS DENSITIES

## B. K. GHOSH AND WEI-MIN HUANG

Department of Mathematics, Lehigh University, Bethlehem, PA 18015, U.S.A.

(Received August 22, 1990; revised June 21, 1991)

**Abstract.** Rosenblatt and Parzen proposed a well-known estimator  $f_n$  for an unknown density function f, and later Schuster suggested a modification  $\hat{f}_n$  to rectify certain drawbacks of  $f_n$ . This paper gives the asymptotically optimum bandwidth and kernel for  $\hat{f}_n$  under the standard measure of IMSE when f is discontinuous at one or both endpoints of its support. We also consider an alternative definition of the IMSE under which the optimum bandwidths and kernels for  $f_n$  are derived. The latter supplement van Eeden's results.

*Key words and phrases:* Estimation of discontinuous densities, alternative notions of IMSE, modified kernel density estimates, optimal bandwidths, optimal kernels.

### 1. Introduction

Let  $X_1, \ldots, X_n$  be independent and identically distributed random variables having a common density function  $f \in \mathcal{F}$  with known support  $[\alpha, \beta]$ ,  $[\alpha, \infty)$  or  $(-\infty, \beta]$ . Consider the problem of estimating f(x) at a given point x using the sample  $(X_1, \ldots, X_n)$ . A common estimator, proposed by Rosenblatt (1956) and Parzen (1962), is given by

(1.1) 
$$f_n(x) = (nb_n)^{-1} \sum_{i=1}^n w(b_n^{-1}[x - X_i]).$$

In (1.1),  $b_n$  is a predetermined bandwidth satisfying

(1.2)  $b_n \to 0 \text{ and } nb_n \to \infty \text{ as } n \to \infty$ 

and  $w \in \mathcal{W}$  is a suitably chosen *kernel*, where the class  $\mathcal{W}$  is defined by (all integrals without limits are over  $\mathbb{R}$ )

(1.3) 
$$w \ge 0, \qquad \int w(z)dz = 1, \qquad \int z^2 w(z)dz < \infty,$$
$$\int w^2(z)dz < \infty, \qquad \int |z|w^2(z)dz < \infty.$$

The class  $\mathcal{W}$  defined by (1.3) is slightly more general than what one usually encounters in the literature; w need not be bounded or symmetric (e.g.,  $w(z) = (16)^{-1}z^{-1/4}$  for  $0 \le z \le 16$  and  $w(z) = 4^{-1}|z|^{-1/4}$  for  $-1 \le z \le 0$  is permissible in the developments below). The varying nature of the family  $\mathcal{F}$  will be described in the third paragraph. The properties of  $f_n$  are generally quite difficult to investigate for arbitrary n. Consequently, the decision on the best choice of the pair  $(b_n, w)$  is usually made from the asymptotic behavior of  $f_n$ . We stipulate that the best choice of  $(b_n, w)$  should, for large n, minimize the *integrated mean square error* (IMSE), which is defined by

(1.4) 
$$R_n(f, b_n, w) = \int E[f_n(x) - f(x)]^2 dx$$
$$= \int [Ef_n(x) - f(x)]^2 dx + \int [\operatorname{Var} f_n(x)] dx.$$

Suppose one can show that, for each  $f \in \mathcal{F}$  and large  $n, R_n$  in (1.4) reduces to

(1.5) 
$$R_{n}(f, b_{n}, w) = \begin{cases} (nb_{n})^{-1}A(f, w) + b_{n}^{\delta}B(f, w) \\ +o(n^{-1}b_{n}^{-1}) + o(b_{n}^{\delta}) & \text{for } w \in \mathcal{W}_{o} \\ (nb_{n})^{-1}A'(f, w) + b_{n}^{\delta'}B'(f, w) \\ +o(n^{-1}b_{n}^{-1}) + o(b_{n}^{\delta'}) & \text{for } w \in \mathcal{W} - \mathcal{W}_{o}, \end{cases}$$

where A > 0, B > 0, A' > 0, B' > 0,  $\delta > \delta' > 0$ , and  $\mathcal{W}_o$  is nonempty. Then, for each fixed  $w \in \mathcal{W}_o$  and  $n \ge 1$ , the choice (see Parzen (1962), Lemma 4a)

(1.6) 
$$b_n(f,w) = [A(f,w)/\delta B(f,w)]^{1/(\delta+1)_n - 1/(\delta+1)_n}$$

minimizes the dominant part of (1.5) with respect to  $b_n$  under  $w \in \mathcal{W}_o$ . In fact, for large n,

$$\min_{b_n} R_n(f, b_n, w) = R_n(f, b_n(f, w), w) + o(n^{-\delta/(\delta+1)}),$$

where

(1.7) 
$$R_n(f, b_n(f, w), w) = (\delta + 1) [\delta^{-\delta} A(f, w)^{\delta} B(f, w)]^{1/(\delta + 1)} n^{-\delta/(\delta + 1)} + o(n^{-\delta/(\delta + 1)}).$$

It is easily verified from (1.5) and (1.7) that, for sufficiently large n, (1.7) is less than  $R_n(f, b_n, w')$  for each  $w \in \mathcal{W}_o$  and  $w' \in \mathcal{W} - \mathcal{W}_o$  so that the class  $\mathcal{W} - \mathcal{W}_o$  can be ignored for asymptotic purposes. We call  $b_n(f, w)$  in (1.6) the (asymptotically) optimum bandwidth for  $f_n$  under  $w \in \mathcal{W}_o$  (note that (1.6) need not be the optimum bandwidth under  $w \in \mathcal{W} - \mathcal{W}_o$ ). Moreover, if there exists a  $w^* \in \mathcal{W}_o$  which minimizes the leading term in  $R_n(f, b_n(f, w), w)$  or, equivalently, the functional  $A(f, w)^{\delta}B(f, w)$  then we call  $w^*$  and  $b_n(f, w^*)$  the (asymptotically) optimum kernel and bandwidth for  $f_n$  under  $\mathcal{W}$  ( $w^*$  may not minimize  $A(f, w)^{\delta}B(f, w)$  in  $\mathcal{W}$  but clearly  $R_n(f, b_n, w')/R_n(f, b_n(f, w^*), w^*) \to \infty$  for any  $w' \in \mathcal{W} - \mathcal{W}_o$ ).

Consider now three families of densities,

 $\mathcal{F}(-\infty,\infty)$ : f is continuous and bounded in  $\mathbb{R}$ ,  $f \ge 0$ ,  $\int f(x)dx = 1$ , f'' is continuous and square-integrable;

 $\mathcal{F}(\alpha,\infty)$ : f is continuous and bounded on  $[\alpha,\infty)$ ,  $f \ge 0$  on  $(\alpha,\infty)$ , f = 0 on  $(-\infty,\alpha)$ ,  $f(\alpha) > 0$ ,  $\int f(x)dx = 1$ , f' is continuous and square-integrable on  $[\alpha,\infty)$ ;

 $\mathcal{F}(\alpha,\beta): f \text{ is continuous on } [\alpha,\beta], f(\alpha) > 0, f(\beta) > 0, f \ge 0 \text{ on } (\alpha,\beta) \text{ and}$ zero elsewhere,  $\int f(x)dx = 1, f'$  is bounded on  $[\alpha,\beta]$ .

Observe that we are demanding from  $\mathcal{F}(\alpha, \infty)$  that f is right-continuous at  $\alpha$  and the right-derivative  $f'(\alpha)$  exists. It is, of course, understood that  $\mathcal{F}(-\infty, \infty)$  and  $\mathcal{F}(\alpha, \infty)$  can contain densities with finite support  $[\alpha, \beta]$ , which implies  $f(\alpha) = 0 =$  $f(\beta)$  in the former and  $f(\alpha) > 0 = f(\beta)$  in the latter. The family  $\mathcal{F}(\alpha, \infty)$  has a dual  $\mathcal{F}(-\infty, \beta)$  and the results for the latter follow from those of  $\mathcal{F}(\alpha, \infty)$  by taking  $X'_i = -X_i$ . The bulk of the literature on kernel density estimation involves  $\mathcal{F}(-\infty, \infty)$ . On the other hand, the families  $\mathcal{F}(\alpha, \infty)$ ,  $\mathcal{F}(-\infty, \beta)$  and  $\mathcal{F}(\alpha, \beta)$  arise in practice when the observed data follow the uniform, U-shaped, J-shaped and certain Pearsonian distributions as well as singly and doubly truncated versions of all continuous densities (see Johnson and Kotz (1970)). In fact, one would seldom expect an f with a finite or semi-infinite support to belong in  $\mathcal{F}(-\infty, \infty)$ .

If  $f \in \mathcal{F}(-\infty, \infty)$ , considerations of the second paragraph lead to the following conclusions under assumptions (1.2) and (1.3) (see Epanechnikov (1969), Rosenblatt (1971) and Silverman (1986)). Let

(1.8) 
$$\xi_w = \int zw(z)dz, \quad \tau_w^2 = \int (z-\xi_w)^2 w(z)dz$$

and define  $\mathcal{W}_o$  as the subclass of  $\mathcal{W}$  having  $\xi_w = 0$ . Then the optimum bandwidth for  $f_n$  under  $w \in \mathcal{W}_o$  is

(1.9) 
$$b_n(f,w) = \left[\tau_w^{-4} \int w^2(z) dz\right]^{1/5} \left[\int f''(x)^2 dx\right]^{-1/5} n^{-1/5}$$

(1.10) 
$$R_n(f, b_n(f, w), w) = \frac{5}{4} \left[ \tau_w \int w^2(z) dz \right]^{4/5} \left[ \int f''(z)^2 dz \right]^{1/5} n^{-4/5} + o(n^{-4/5}),$$

and the kernel minimizing the leading term in (1.10), subject to  $\tau_w = \tau$ , is

(1.11) 
$$w_o^*(z) = \begin{cases} 3(4\tau\sqrt{5})^{-1}(1-z^2/5\tau^2) & \text{for } |z| \le \tau\sqrt{5} \\ 0 & \text{for } |z| > \tau\sqrt{5}. \end{cases}$$

Thus, if our criterion is to minimize  $R_n$  for large n,  $w_o^*$  and  $b_n(f, w_o^*)$  are the optimum kernel and bandwidth for  $f_n$  when  $f \in \mathcal{F}(-\infty, \infty)$ . Similarly, if  $f \in \mathcal{F}(\alpha, \beta)$  or  $\mathcal{F}(\alpha, \infty)$ , then van Eeden (1985) shows that the optimum kernel and

bandwidth for  $f_n$  are, respectively,

(1.12) 
$$w_1^*(z) = (\tau \sqrt{2})^{-1} \exp[-\sqrt{2}|z|/\tau] \quad \text{for} \quad -\infty < z < \infty,$$

(1.13) 
$$b_n(f, w_1^*) = \sqrt{2} [\tau \{f^2(\alpha) + f^2(\beta)\}^{1/2}]^{-1} n^{-1/2},$$

and

(1.14) 
$$R_n(f, b_n(f, w_1^*), w_1^*) = \frac{1}{2} [f^2(\alpha) + f^2(\beta)]^{1/2} n^{-1/2} + o(n^{-1/2}),$$

where  $\tau$  is an arbitrary positive number (one takes  $f(\beta) = 0$  above when  $f \in \mathcal{F}(\alpha, \infty)$ ). van Eeden actually assumes symmetry and boundedness for w but her proof can easily be extended to  $\mathcal{W}_o$  (see Cline and Hart (1990)). It is customary to take  $\tau = 1$  in  $w_o^*$  and  $w_1^*$  because any scale change in w can be absorbed in  $b_n$ .

The purpose of the present paper is to explore two new aspects of estimation for the families  $\mathcal{F}(\alpha, \infty)$  and  $\mathcal{F}(\alpha, \beta)$ . First, it is easily shown by Parzen's (1962) method that, if  $x_o$  is an arbitrary point of discontinuity in f, then (1.2) implies that with probability one we have

$$(1.15) \quad f_n(x_o) \to f(x_o) + [f(x_o^-) - f(x_o)][1 - W(0)] + [f(x_o^+) - f(x_o)]W(0),$$

where  $W(z) = \int_{-\infty}^{z} w(x) dx$ . It follows from (1.15) that  $f_n$  under any  $w \in W$ cannot be uniformly consistent for  $f \in \mathcal{F}(\alpha, \beta)$  while the only uniformly consistent  $f_n$  for  $f \in \mathcal{F}(\alpha, \infty)$  are those which use a w with support in  $\mathbb{R}^-$ . In particular, the lack of uniform consistency renders (1.12) and (1.13) unsatisfactory. In Sections 2 and 3, we investigate Schuster's (1985) estimator  $\hat{f}_n$  to remedy the situation. Our main result (Theorem 2.1) obtains the optimum kernel and bandwidth for  $\hat{f}_n$ . Apart from being uniformly consistent, the IMSE of  $\hat{f}_n$  is of order  $n^{-3/4}$ , which tends to zero faster than (1.14). Second, in Section 4, we introduce an alternative definition  $\tilde{R}_n$  for the IMSE of  $f_n$ , which is sometimes more meaningful than  $R_n$ . It is shown (Theorems 4.1 and 4.2) that  $\tilde{R}_n$  leads to optimum kernel and bandwidth for  $f_n$  that are different from (1.12)-(1.13) for families  $\mathcal{F}(\alpha, \infty)$  and  $\mathcal{F}(\alpha, \beta)$  but the same as (1.9)-(1.11) for the family  $\mathcal{F}(-\infty, \infty)$ .

### 2. A modified estimator and the main result

Schuster (1985) proposed a simple modification  $\hat{f}_n$  of  $f_n$ . For  $f \in \mathcal{F}(\alpha, \infty)$ , define

(2.1) 
$$\hat{f}_n(x) = \begin{cases} f_n(x) + f_n(2\alpha - x), & x \ge \alpha \\ 0, & \text{elsewhere} \end{cases}$$

For  $f \in \mathcal{F}(\alpha, \beta)$ , define

(2.2) 
$$\hat{f}_n(x) = \begin{cases} f_n(x) + f_n(2\alpha - x) + f_n(2\beta - x), & \alpha \le x \le \beta \\ 0, & \text{elsewhere.} \end{cases}$$

The idea of the reflection principle (i.e., folding  $f_n$  at  $\alpha$  in (2.1) or at both  $\alpha$  and  $\beta$  in (2.2)) involved in  $\hat{f}_n$  seems to have been well known to time series analysts,

566

but Schuster (1985) and Silverman ((1986), Section 2.10) provide formal, albeit intuitive, justifications for using  $\hat{f}_n$  in the present context. Schuster proves the uniform consistency of  $\hat{f}_n$  when  $f \in \mathcal{F}(\alpha, \infty)$  or  $\mathcal{F}(\alpha, \beta)$  and w is symmetric; one can show more generally that the uniform consistency holds if and only if wsatisfies  $\int_0^\infty w(z)dz = 1/2$  (w need not be symmetric or nonnegative).

Define the IMSE  $\hat{R}_n(f, b_n, w)$  of  $\hat{f}_n$  as the right-hand side of (1.4) with  $f_n$  replaced by  $\hat{f}_n$ . Our main result is

THEOREM 2.1. Assume that (1.2) and (1.3) hold. If  $f \in \mathcal{F}(\alpha, \beta)$  and f' is continuous on  $[\alpha, \beta]$  with either  $f'(\alpha) \neq 0$  or  $f'(\beta) \neq 0$ , then the optimum kernel and bandwidth for  $\hat{f}_n$  in (2.2) are, respectively,

$$(2.3) \quad w_{2}^{*}(z) = \begin{cases} \frac{1}{2}m[(C-1)e^{m|z|}\{\cos(mz) - \sin(m|z|)\} \\ +Ce^{-m|z|}\{\cos(mz) + \sin(m|z|)\}] & \text{for } |z| \le \pi/(2m) \\ 0 & \text{for } |z| > \pi/(2m), \end{cases}$$

$$(2.4) \quad b_{n}(f, w_{2}^{*}) = 2^{1/4}m[f'(\alpha)^{2} + f'(\beta)^{2}]^{-1/4}n^{-1/4},$$

where

$$C = (1 - e^{-\pi})^{-1}, \quad m = 2^{1/2} \tau^{-1} (e^{\pi/2} - e^{-\pi/2})^{-1/2}$$

and  $\tau$  is an arbitrary positive number. If  $f \in \mathcal{F}(\alpha, \infty)$  and  $f'(\alpha) \neq 0$ , then (2.3) and (2.4) are the optimum kernel and bandwidth for  $\hat{f}_n$  in (2.1) with  $f'(\beta)$  in (2.4) replaced by zero. Moreover, in either case  $\hat{R}_n(f, b_n(f, w_2^*), w_2^*)$  is given by the right-hand side of (2.27) with L(w) given by (2.28).

We make a few remarks before proving the theorem. One consequence of Theorem 2.1 is that  $\hat{R}_n(f, b_n(f, w_2^*), w_2^*) = O(n^{-3/4})$ , which formally shows why  $\hat{f}_n$  should be a better estimate than  $f_n$ . In Section 3, we investigate more precisely the efficiency of  $\hat{f}_n$  based on (2.3) and (2.4). For the case of  $f \in \mathcal{F}(\alpha, \infty)$ , Cline and Hart (1990) derive expressions for  $\hat{R}_n$  (take their  $x_o$  as our  $\alpha$ ), but not for the optimum w, under various conditions on f and w. However, their expression for  $\hat{R}_n$  is incorrect (e.g., their (4.6)), they use stronger conditions (e.g.,  $\int |z|^3 w(z) dz < \infty$ ), and there are certain gaps in their proof (e.g., they treated  $\hat{p}$  as a constant). Cline and Hart (1990) do not attempt to derive  $\hat{R}_n$  when  $f \in \mathcal{F}(\alpha, \beta)$ .

The kernel  $w_2^*$  appears somewhat unusual but the uniform and Epanechnikov kernels can, in fact, be thought of as first and second order approximations to  $w_2^*$ . This follows from the power series expansion of (2.3),

$$w_2^*(z) = m\left(C - \frac{1}{2}\right)(1 - m^2 z^2) + O(|z|^3) \quad \text{for} \quad |z| \le \pi/(2m).$$

Finally, note that (2.1) and (2.2) require the actual values of  $\alpha$  and  $\beta$ , which must come either from a knowledge of the population that is being sampled or from some adaptive estimation method (e.g., start with  $\hat{\alpha} = \min[X_1, \ldots, X_n]$  and  $\hat{\beta} = \max[X_1, \ldots, X_n]$ ; in the adaptive case the resulting IMSE must be

reanalyzed, which is not done in this paper. On the other hand, the occurrence of  $f'(\alpha)$  and  $f'(\beta)$  in  $b_n(f, w_2^*)$  seems to be less restrictive than the occurrence of  $\int f''(x)^2 dx$  in  $b_n(f, w_o^*)$ .

The proof of Theorem 2.1 requires two lemmas.

LEMMA 2.1. If  $f \in \mathcal{F}(\alpha, \beta)$  or  $\mathcal{F}(\alpha, \infty)$ , (1.2) and (1.3) hold, f' is continuous in the support  $[\alpha, \beta]$  of f and w is symmetric about zero, then (defining W as in (1.15))

$$(2.5) \qquad \hat{R}_n(f, b_n, w) = (nb_n)^{-1} \int w^2(z) dz + 4b_n^3 \{f'(\alpha)^2 + f'(\beta)^2\} \int_0^\infty \left\{ \int_y^\infty [1 - W(z)] dz \right\}^2 dy + O(n^{-1}b_n^{-1/2}) + o(b_n^3).$$

PROOF. We will drop the subscript from  $b_n$  and prove the result (2.5) for the case  $\beta = \infty$ . The proof for the case  $\beta < \infty$  is analogous but more tedious. By definition

(2.6) 
$$\hat{R}_n(f,b_n,w) = \int_\alpha^\infty [E\hat{f}_n(x) - f(x)]^2 dx + \int_\alpha^\infty [\operatorname{Var} \hat{f}_n(x)] dx.$$

For the first integral, note that the Taylor expansion of (2.1) yields, for all  $x \ge \alpha$ ,

$$(2.7) \quad E\hat{f}_n(x) = f(x) \left[ W\left(\frac{x-\alpha}{b}\right) + W\left(-\frac{x-\alpha}{b}\right) \right] \\ - b \int_{-\infty}^{(x-\alpha)/b} zw(z)f'(x-\theta_1 bz)dz \\ + \int_{-\infty}^{-(x-\alpha)/b} (2\alpha - 2x - bz)w(z)f'(x+\theta_2\{2\alpha - 2x - bz\})dz,$$

where  $0 < \theta_1, \theta_2 < 1$ . Using the symmetry of w twice one gets

$$(2.8) \quad \int_{\alpha}^{\infty} [E\hat{f}_{n}(x) - f(x)]^{2} dx$$

$$= \int_{\alpha}^{\infty} \left[ b \int_{-\infty}^{(x-\alpha)/b} zw(z)f'(x-\theta_{1}bz)dz + \int_{-\infty}^{-(x-\alpha)/b} (bz+2x-2\alpha)w(z)f'(x+\theta_{2}\{2\alpha-2x-bz\})dz \right]^{2} dx$$

$$= b^{3} \int_{0}^{\infty} \left[ \int_{y}^{\infty} zw(z)f'(\alpha+by+\theta_{1}bz)dz + \int_{y}^{\infty} (z-2y)w(z)f'(\alpha+by-2\theta_{2}by+\theta_{2}bz)dz \right]^{2} dy.$$

By the right-continuity of f' at  $\alpha$  (see Rosenblatt (1971), p. 1819 or Apostol (1957), p. 441) one finally gets

(2.9) 
$$\int_{\alpha}^{\infty} [E\hat{f}_{n}(x) - f(x)]^{2} dx$$
$$= 4b^{3}f'(\alpha)^{2} \int_{0}^{\infty} \left[ \int_{y}^{\infty} (z - y)w(z)dz \right]^{2} dy + o(b^{3})$$

Consider next the second integral in (2.6). It follows from (2.1) that

(2.10) 
$$\int_{\alpha}^{\infty} [\operatorname{Var} \hat{f}_n(x)] dx = \int_{\alpha}^{\infty} [\operatorname{Var} f_n(x)] dx + \int_{\alpha}^{\infty} [\operatorname{Var} f_n(2\alpha - x)] dx + 2 \int_{\alpha}^{\infty} [\operatorname{Cov}(f_n(x), f_n(2\alpha - x))] dx.$$

The first term on the right-hand side of (2.10) is given by (see Rosenblatt (1971))

(2.11) 
$$\int_{\alpha}^{\infty} [\operatorname{Var} f_n(x)] dx = (nb)^{-1} \int w^2(z) dz + O(n^{-1}),$$

and the second term is

(2.12) 
$$\int_{\alpha}^{\infty} [\operatorname{Var} f_n(2\alpha - x)] dx$$
$$= (nb)^{-1} \int_{\alpha}^{\infty} \int_{-\infty}^{-(x-\alpha)/b} w^2(z) f(2\alpha - x - bz) dz dx$$
$$- n^{-1} \int_{\alpha}^{\infty} [Ef_n(2\alpha - x)]^2 dx.$$

It is easily shown that the first term in (2.12) is  $O(n^{-1})$  and the second term is  $O(bn^{-1})$ , so that (2.12) is  $O(n^{-1})$ . Since

$$|\operatorname{Cov}(f_n(x), f_n(2\alpha - x))| \le [\operatorname{Var} f_n(x)]^{1/2} [\operatorname{Var} f_n(2\alpha - x)]^{1/2}$$

for all  $x \ge \alpha$ , it also follows that the third term in (2.10) is  $O(n^{-1}b^{-1/2})$ . Thus

(2.13) 
$$\int_{\alpha}^{\infty} [\operatorname{Var} \hat{f}_n(x)] dx = (nb)^{-1} \int w^2(z) dz + O(n^{-1}b^{-1/2}).$$

Relations (2.6), (2.9) and (2.13) lead to (2.5).  $\Box$ 

If w is nonsymmetric, the first term on the right-hand side of (2.7) does not simplify to f(x) for all x. One can easily show that this, in turn, causes (2.8) to be of order O(b). Note that the symmetry of w has not been used in (2.13). If w is bounded then one can replace  $O(n^{-1}b^{-1/2})$  in (2.13) and (2.5) by  $O(n^{-1})$ . The reason is that the third term in (2.10) can then be written as  $2n^{-1}(A_n - B_n)$ , where

$$egin{aligned} A_n &= b^{-1} \int_lpha^\infty \int_{-\infty}^{(x-lpha)/b} w(z) w(z-2(x-lpha)b^{-1}) f(x-bz) dz dx \ &\leq 4 [\sup w(z)] [\sup f(x)] \int |z| w(z) dz \end{aligned}$$

and

$$B_n = \int_{\alpha}^{\infty} [Ef_n(x)] [Ef_n(2\alpha - x)] dx \le b [\sup f(x)]^2 \int |z| w(z) dz.$$

LEMMA 2.2. Let  $W_o$  be the class of all functions w on  $\mathbb{R}$  satisfying

(2.14) 
$$w(z) = w(-z) \ge 0, \quad \int w(z)dz = 2\gamma, \quad \int z^2 w(z)dz = 2\delta,$$

where  $\gamma$  and  $\delta$  are given positive numbers. Then the functional

(2.15) 
$$L(w) = \left[\int_0^\infty w^2(z)dz\right]^3 \left[\int_0^\infty \left\{\int_y^\infty [1-W(z)]dz\right\}^2 dy\right].$$

is minimized on  $\mathcal{W}_o$  uniquely by

$$(2.16) \quad w_2^*(z) = \begin{cases} m\gamma(C-1)e^{m|z|}\{\cos(mz) - \sin(m|z|)\} \\ +m\gamma Ce^{-m|z|}\{\cos(mz) + \sin(m|z|)\} \\ 0 & \text{for } |z| \le \pi/(2m), \end{cases}$$

where

(2.17) 
$$C = (1 - e^{-\pi})^{-1}, \quad m = (2\gamma/\delta)^{1/2} (e^{\pi/2} - e^{-\pi/2})^{-1/2}.$$

PROOF. For each  $w \in \mathcal{W}_o$ , define a function g on  $\mathbb{R}^+$  by

(2.18) 
$$g(z) = \int_{z}^{\infty} w(x) dx \quad \text{for} \quad z \ge 0,$$

and express (2.15) under (2.14) as

(2.19) 
$$L(g) = \left[\int_0^\infty g'(z)^2 dz\right]^3 \left[\int_0^\infty \left\{\int_y^\infty g(z) dz\right\}^2 dy\right],$$

(2.20) 
$$g'(z) \le 0 \text{ for } z \ge 0, \quad g(0) = \gamma, \quad \int_0^\infty z g(z) dz = \delta/2,$$

where g' denotes the first derivative of g. Denote by  $\mathcal{G}$  the class of all g on  $\mathbb{R}^+$  satisfying (2.20). Since there is a one-to-one correspondence between  $w \in \mathcal{W}_o$  and  $g \in \mathcal{G}$ , it suffices to show that the g corresponding to  $w_2^*$ , i.e.,

(2.21) 
$$g_*(z) = \int_z^{\pi/(2m)} w_2^*(x) dx$$
$$= \begin{cases} \gamma [Ce^{-mz} - (C-1)e^{mz}] \cos(mz) & \text{for } 0 \le z \le \pi/(2m) \\ 0 & \text{for } z > \pi/(2m) \end{cases}$$

minimizes L of (2.19) uniquely on  $\mathcal{G}$ . The calculus of variations involved in proving this is quite similar to that in the proof of Theorem 1 in Ghosh and Huang (1991) and we mention here the essential steps. It is easily shown that L is strictly convex on  $\mathcal{G}$ . Simple computations show that  $g_*$  indeed satisfies the constraints in (2.20) and  $g'_*(\pi/2m) = 0$ . Let  $h \in (0, 1)$  be a real number and define  $R(z) = g(z) - g_*(z)$ for an arbitrary  $g \in \mathcal{G}$ . One gets from (2.19), by integrating by parts, the Gateaux differential as

(2.22) 
$$\partial L(g_* + hR)/\partial h \mid_{h=0}$$
$$= \int_0^\infty \left[ -g_*''(z) + (A/3B) \int_0^z \int_y^\infty g_*(x) dx dy \right] R(z) dz,$$

where

(2.23) 
$$A = \int_0^\infty g'_*(x)^2 dx, \quad B = \int_0^\infty \left\{ \int_y^\infty g_*(x) dx \right\}^2 dy.$$

It is easily verified that

$$A/(3B) = 4m^4, \qquad \int_0^{\pi/(2m)} \int_y^\infty g_*(x) dx dy = 1/4,$$

where m is defined in (2.17). Consequently,

(2.24) 
$$-g_*''(z) + (A/3B) \int_0^z \int_y^\infty g_*(x) dx dy = \begin{cases} 0 & \text{for } 0 \le z \le \pi/(2m) \\ m^4 & \text{for } z > \pi/(2m). \end{cases}$$

It follows from (2.22) and (2.24) that

$$(2.25) \qquad \partial L(g_* + hR)/\partial h \mid_{h=0} \\ = \int_{\pi/(2m)}^{\infty} \left[ -g_*''(z) + (A/3B) \int_0^z \int_y^\infty g_*(x) dx dy \right] R(z) dz \\ = m^4 \int_{\pi/(2m)}^\infty g(z) dz \ge 0 \quad \text{for every } g \in G.$$

The first equation in (2.24) and the convexity of L show that  $g_*$  minimizes L uniquely among all  $g \in \mathcal{G}$  with support in  $[0, \pi/(2m)]$ . The result in (2.25) globalizes the conclusion to  $\mathcal{G}$ .  $\Box$ 

PROOF OF THEOREM 2.1. It follows from (1.6) and (2.5) that the bandwidth which asymptotically minimizes  $\hat{R}_n$  is given by

(2.26) 
$$b_n(f,w) = 6^{-1/4} [\{f'(\alpha)\}^2 + \{f'(\beta)\}^2]^{-1/4} \\ \cdot \left[\int_0^\infty w^2(z) dz\right] [L(w)]^{-1/4} n^{-1/4},$$

where L(w) is defined in (2.15). Substituting (2.26) in (2.5) one gets

(2.27) 
$$\min_{b_n} \hat{R}_n(f, b_n, w) = 2^{13/4} 3^{-3/4} [f'(\alpha)^2 + f'(\beta)^2]^{1/4} [L(w)]^{1/4} n^{-3/4} + o(n^{-3/4})$$

It follows from the remark after Lemma 2.1 that, for any nonsymmetric w,  $\hat{R}_n(f, b_n, w)$  is  $an^{-1/2} + o(n^{-1/2})$  for some a > 0, which is larger than (2.27) for sufficiently large n. Consequently, it suffices to restrict attention to symmetric w for the purpose of determining the optimal kernel. The latter problem is equivalent to minimizing L(w) of (2.15) with respect to symmetric  $w \in \mathcal{W}$ . It follows from Lemma 2.2 with  $\gamma = 1/2$  and  $\delta = \tau^2/2$  that L(w) is minimized, subject to  $\tau_w = \tau$ , uniquely by  $w_2^*$  of (2.3). Finally, substituting  $w_2^*$  in (2.26) one gets (2.4).  $\Box$ 

It is easily verified that (2.15) is invariant in  $\tau_w$  for any symmetric w. Substitution of (2.16) in (2.15) yields

(2.28) 
$$L(w_2^*) = (27/262144)(e^{\pi} + 1)^4(e^{\pi} - 1)^{-4}$$

# 3. Relative efficiency of $\hat{f}_n$ and $f_n$

Suppose  $f \in \mathcal{F}(\alpha, \infty)$  or  $\mathcal{F}(\alpha, \beta)$ . Consider first the efficiency of  $\hat{f}_n$  based on an arbitrary w relative to  $\hat{f}_n$  based on  $w_2^*$ . An asymptotic measure of this relative efficiency is biven by

(3.1) 
$$\lim_{n \to \infty} [\hat{R}_n(f, b_n(f, w_2^*), w_2^*) / \hat{R}_n(f, b_n(f, w), w)] = \begin{cases} [L(w_2^*) / L(w)]^{1/4} & \text{if } w \text{ is symmetric} \\ 0 & \text{if } w \text{ is nonsymmetric}, \end{cases}$$

where L(w) is defined in (2.15) and  $L(w_2^*)$  is given by (2.28). Theorem 2.1 shows that  $L(w_2^*) < L(w)$  for any symmetric  $w \neq w_2^*$ . Table 1 shows values of L(w)and (3.1) for some popular symmetric kernels (Epanechnikov (1969), Rosenblatt (1971)). It can be seen that, in practical terms, there is very little difference between  $w_2^*$  and the quadratic kernel in the table. This quadratic kernel happens to be the best choice  $w_o^*$  for  $f_n$  when  $f \in \mathcal{F}(-\infty, \infty)$ . Conversely, it is interesting to note here that there is very little difference between  $R_n$  under the kernels  $w_o^*$ and  $w_2^*$  when  $f \in \mathcal{F}(-\infty, \infty)$  (e.g., in Epanechnikov's (1969) Table 1, r = 1.004 under  $w_2^*$ ). Consider next the efficiency of  $f_n$  based on its optimum kernel  $w_1^*$  relative to  $\hat{f}_n$  based on any symmetric w. It follows from (1.14) and (2.27) that the relative efficiency of  $f_n$  is asymptotically zero, in addition to the fact that  $\hat{f}_n$  is globally consistent but  $f_n$  is not.

Kernel	L(w)	$[L(w_2^*)/L(w)]^{1/4}$
$w_2^*(z)$ of (2.3)	see (2.28)	
$w(z) = (\sqrt{6} -  z )/6  ext{ for }  z  \le \sqrt{6}^{+}$	1/6804	.9976
$w(z) = 3(4\sqrt{5})^{-1}(1-z^2/5)$ for $ z  \le \sqrt{5}^{\dagger}$	33/224000	.9970
$w(z) = (2\pi)^{-1/2} \exp(-z^2/2)$	$(\sqrt{2}-1)/192\sqrt{2}\pi^2$	.9851
$w(z) = (2\sqrt{3})^{-1}$ for $ z  \le \sqrt{3}^{\dagger}$	1/5120	.9291
$w(z) = (\sqrt{2})^{-1} \exp(-\sqrt{2} z )$	1/4096	.8787

Table 1. Relative efficiency of some symmetric kernels for  $\hat{f}_n$ .

 $^{\dagger}w(z) = 0$  otherwise.

In principle, one may think of using  $\hat{f}_n$  of (2.2) to estimate an f with compact support  $[\alpha, \beta]$  even when the latter is continuous at  $\alpha$  and  $\beta$ . However, folding seems to be undesirable at a continuity point for two reasons. Suppose first that  $f \in \mathcal{F}(-\infty, \infty)$ , w is symmetric, and both  $f_n$  and  $\hat{f}_n$  use the same w. The conditions underlying  $\mathcal{F}(-\infty, \infty)$  imply  $f(\alpha) = f'(\alpha) = f''(\alpha) = 0$  and  $f(\beta) =$  $f'(\beta) = f''(\beta) = 0$ , using which it is easy to show that  $|R_n - \hat{R}_n|$  is  $o(b_n^4) + O(n^{-1})$ . This means that folding will not affect the asymptotic adequacy of  $f_n$  for the family  $\mathcal{F}(-\infty, \infty)$ . Suppose next that  $f(\alpha) = 0 = f(\beta)$  but  $f'(\alpha) \neq 0$  or  $f'(\beta) \neq 0$ ; such an f does not belong to  $\mathcal{F}(-\infty, \infty)$ ,  $\mathcal{F}(\alpha, \infty)$  or  $\mathcal{F}(\alpha, \beta)$ . One can then show that  $\hat{R}_n$  is as in (2.5) and  $R_n$  is given by (2.5) with the factor 4 replaced by 2. This implies

$$\min_{b_n} \hat{R}_n(f, b_n, w) = 2^{1/4} \min_{b_n} R_n(f, b_n, w) + o(n^{-3/4}),$$

and therefore  $f_n$  is an asymptotically better estimate than  $f_n$ . It follows from the results for  $\hat{f}_n$  that, under the present conditions on f, the optimal kernel for  $f_n$  is  $w_2^*$  (and not  $w_a^*$  nor  $w_1^*$ ).

### 4. An alternative definition of IMSE of $f_n$

Suppose that the support of f is  $[\alpha, \beta], -\infty \leq \alpha < \beta \leq \infty$ , and consider

(4.1) 
$$\tilde{R}_n(f, b_n, w) = \int_{\alpha}^{\beta} E[f_n(x) - f(x)]^2 dx$$
$$= \int_{\alpha}^{\beta} [Ef_n(x) - f(x)]^2 dx + \int_{\alpha}^{\beta} [\operatorname{Var} f_n(x)] dx$$

Clearly,  $R_n \geq \hat{R}_n$  for each *n* because  $f_n$  need not be zero outside  $[\alpha, \beta]$ . In some situations,  $\tilde{R}_n$  may be a more meaningful indicator of the IMSE of  $f_n$  than  $R_n$ .

One reason is that, if  $\alpha$  and  $\beta$  are known, one would not really use  $f_n$  to estimate f = 0 outside  $[\alpha, \beta]$  and, therefore, no loss is incurred by  $f_n(x)$  for  $x \notin [\alpha, \beta]$  (see Lehmann (1983), pp. 55–56, for a discussion). A second reason is that, for estimating  $f \in \mathcal{F}(\alpha, \infty)$ ,  $f_n$  is uniformly consistent for f if and only if the support of w is in  $\mathbb{R}^-$ . This feature is not revealed by the minimization process of  $R_n$  (see (1.12) and (1.13)) whereas  $\tilde{R}_n$  does exhibit this feature as noted below. Finally, in the context of using  $f_n$  for hypothesis testing, Bickel and Rosenblatt (1973) also ignored the loss in  $f_n(x)$  for  $x \notin [\alpha, \beta]$ .

It is of some interest to know the asymptotically optimum kernel and bandwidth for  $f_n$  that minimize  $\tilde{R}_n$ . The relevant results follow from those under  $R_n$ by noting that

$$R_n(f, b_n, w) - \tilde{R}_n(f, b_n, w) = \int_{-\infty}^{\alpha} [Ef_n(x)]^2 dx + \int_{\beta}^{\infty} [Ef_n(x)]^2 dx + O(n^{-1})$$

and then analyzing the two integrals above under the appropriate assumptions of f. We summarize them without proofs which are similar to those in Rosenblatt (1971) and van Eeden (1985). First, if  $f \in \mathcal{F}(-\infty,\infty)$ , then  $w_o^*$  and  $b_n(f, w_o^*)$  are also the optimum kernel and bandwidth under  $\tilde{R}_n$ . Second, if  $f \in \mathcal{F}(\alpha,\infty)$  or  $\mathcal{F}(\alpha,\beta)$ , then we have

THEOREM 4.1. Let  $f \in \mathcal{F}(\alpha, \infty)$  and assume that (1.2) and (1.3) hold. Then, under the IMSE  $\tilde{R}_n$ , the optimum kernel and bandwidth for  $f_n$  are, repectively,

(4.2) 
$$w_3^*(z) = \begin{cases} (z+3\tau\sqrt{2})/9\tau^2 & \text{for } -3\tau\sqrt{2} \le z \le 0\\ 0 & \text{elsewhere,} \end{cases}$$

(4.3) 
$$b_n(f, w_3^*) = \tau^{-1} \left[ 9\sqrt{2} \int f'(x)^2 dx \right]^{-1/3} n^{-1/3},$$

where  $\tau$  is an arbitrary positive number. Moreover,

(4.4) 
$$ilde{R}_n(f, b_n(f, w_3^*), w_3^*) = \left[\frac{4}{3}\int f'(x)^2 dx\right]^{1/3} n^{-2/3} + o(n^{-2/3}).$$

THEOREM 4.2. Let  $f \in \mathcal{F}(\alpha, \beta)$  and assume that (1.2) and (1.3) hold. Then, under the IMSE  $\tilde{R}_n$ , the optimum kernel and bandwidth for  $f_n$  are, respectively,

(4.5) 
$$w_4^*(z) = \begin{cases} [\tau(1+\rho)]^{-1}(1+\rho^2)^{1/2}\exp[-(1+\rho^2)^{1/2}z/\tau] & \text{for } z \ge 0\\ [\tau(1+\rho)]^{-1}(1+\rho^2)^{1/2}\exp[(1+\rho^2)^{1/2}z/\rho\tau] & \text{for } z \le 0, \end{cases}$$
  
(4.6)  $b_n(f, w_4^*) = (1+\rho^2)^{1/2}[\tau f(\alpha)]^{-1}n^{-1/2},$ 

where  $\rho = f(\alpha)/f(\beta)$  and  $\tau$  is an arbitrary positive number. Moreover,

(4.7) 
$$\tilde{R}_n(f, b_n(f, w_4^*), w_4^*) = [f^{-1}(\alpha) + f^{-1}(\beta)]^{-1} n^{-1/2} + O(n^{-1}).$$



Fig. 1. Three estimates of (5.1) using n = 100 observations.

If  $f \in \mathcal{F}(\alpha, \infty)$ , then (1.15) shows that  $f_n$  under every w with support in  $\mathbb{R}^-$  is uniformly consistent for f; (4.2) effectively gives the best kernel among these w's under the criterion  $\tilde{R}_n$ . On the other hand, (1.12) is not related to the consistency property because it is constrained by the very choice of  $R_n$  as the IMSE. The best kernel among uniformly consistent  $f_n$  under the criterion  $R_n$  is, in fact,  $w(z) = \tau^{-1} \exp(z/\tau)$  for  $z \leq 0$ .

In practical terms, Theorems 4.1 and 4.2 are most because, when  $f \in \mathcal{F}(\alpha, \infty)$  or  $\mathcal{F}(\alpha, \beta)$ ,  $\hat{f}_n$  is a superior estimator also under (4.1). If  $f_n(x)$  in (4.1) is replaced by  $\hat{f}_n(x)$  of (2.1) or (2.2), it is easily shown that Theorem 2.1 remains valid so that the right-hand side of (4.1) is of order  $n^{-3/4}$ .

### 5. Numerical results

The efficacy of  $f_n$  and  $\hat{f}_n$  as global estimators of a discontinuous f can be best judged by estimating a known f over a range of x-values using actual samples from f. We consider the exponential density

(5.1) 
$$f(x) = \begin{cases} \exp(-x) & \text{for } x \ge 0\\ 0 & \text{elsewhere} \end{cases}$$

as an example of  $\mathcal{F}(\alpha, \infty)$  to illustrate the essential features. A pseudo-random sample of size n = 100 was obtained by computer from the exponential distribution. Figure 1 shows the target curve (5.1) and three estimates: (i) the  $f_n$ -curve based on (4.2)–(4.3), (ii) the  $f_n$ -curve based on (1.9) and (1.11) and (iii) the  $\hat{f}_n$ curve based on (2.3)–(2.4). Also shown are the observed integrated squared errors (ISE)  $\int [f_n(x) - f(x)]^2 dx$  and  $\int [\hat{f}_n(x) - f(x)]^2 dx$ . In Fig. 1(i), the "local kinks" are caused by the discontinuity of  $w_3^*$  at 0 while the "global waves" are due to the one-sided nature of  $w_3^*$  (i.e., only observations in  $[x, x + b_n]$  are contributing to  $f_n(x)$ ). At the expense of inconsistency of  $f_n(0)$ , the local kinks are corrected in Fig. 1(ii) by the continuity of  $w_o^*$ . Figure 1(iii) shows, as a confirmation of the formal results of Section 2, that the folded estimator rectifies the drawbacks of  $f_n$ . Actual computations show that the  $f_n$ -curve based on (1.12) and (1.13) is indistinguishable from Fig. 1(ii), and the  $\hat{f}_n$ -curve based on (1.11) and (2.26) is indistinguishable from Fig. 1(iii).

#### Acknowledgements

The authors are grateful to the referees for several comments which helped improve the paper.

### References

Apostol, T. M. (1957). Mathematical Analysis, Addison-Wesley, Reading, Massachusetts.

- Bickel, P. J. and Rosenblatt, M. (1973). On some global measures of the deviations of density function estimates, Ann. Statist., 1, 1071-1095 (correction: ibid, (1975) 3, 1370).
- Cline, D. B. H. and Hart, J. D. (1990). Kernel estimates of densities with discontinuities or discontinuous derivatives, *Statistics*, 22, 69–84.

- Epanechnikov, V. A. (1969). Nonparametric estimates of a multivariate probability density, Theory Probab. Appl., 14, 153-158.
- Ghosh, B. K. and Huang, W.-M. (1991). The power and optimal kernel of the Bickel-Rosenblatt test for goodness of fit, Ann. Statist., **19**, 999–1009.
- Johnson, N. L. and Kotz, S. (1970). Distributions in Statistics: Continuous Univariate Distributions, Houghton-Mifflin, Boston.

Lehmann, E. L. (1983). Theory of Point Estimation, Wiley, New York.

- Parzen, E. (1962). On estimation of a probability density function and mode, Ann. Math. Statist., 33, 1065-1076.
- Rosenblatt, M. (1956). Remarks on some nonparametric estimates of a density function, Ann. Math. Statist., 27, 832-837.
- Rosenblatt, M. (1971). Curve estimates, Ann. Math. Statist., 42, 1815-1842.
- Schuster, E. F. (1985). Incorporating support constraints into nonparametric estimates of densities, Comm. Statist. Theory Methods, 14, 1123–1136.
- Silverman, B. W. (1986). Density Estimation for Statistics and Data Analysis, Chapman, New York.
- van Eeden, C. (1985). Mean integrated squared error of kernel estimators when the density and its derivative are not necessarily continuous, Ann. Inst. Statist. Math., 37, 461–472.